- An investigation of factors affecting inferences from and
- reliability of state-space age-structured assessment
- 3 models
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## Abstract

State-space models are increasingly used for stock assessment, and evaluations of their statistical reliability and best practices for selecting among process error configurations are 26 needed. We simulated 72 operating models that varied fishing pressure and observation 27 error across process errors in recruitment, survival, selectivity, catchability, and/or natural mortality. We fit estimating models with different assumptions on the process error source and whether median natural mortality or a stock-recruit relationship were estimated. Estimating models without a stock-recruit relationship that assumed the correct process error 31 source and median natural mortality had high convergence rates and low bias. Bias was also 32 low under many incorrect process error assumptions when there was contrast in fishing pres-33 sure and low observation error. Marginal AIC most accurately distinguished process errors on recruitment, survival, and selectivity, and other process error sources when variability 35 was greater. Retrospective patterns were generally small but were sizable for recruitment when observation error was high. These results help establish the statistical reliability of state space assessment models and pave the way for the next-generation of fisheries stock assessment.

## 40 Introduction

panded dramatically within International Council for the Exploration of the Sea (ICES), Canada, and the Northeast US (Nielsen and Berg 2014; Cadigan 2016; Pedersen and Berg 2017; Stock and Miller 2021). State-space models latent population characteristics as statistical time series with periodic observations that also may have error due to sampling or other sources of measurement error. Traditional assessment models may use state-space approaches to account for temporal variability in population characteristics (Legault and Restrepo 1999; Methot and Wetzel 2013), but these models treat the annual parameters as penalized fixed effects parameters where the variance parameters controlling the penalties are assumed known (Thorson and Minto 2015). Modern state-space models can estimate the annually varying parameters as random effects with variance parameters estimated using maximum marginal likelihood or corresponding Bayesian approaches. These latter approaches are considered best practice and a recommended for the next generation of stock assessment models (Hoyle et al. 2022; Punt 2023). State-space stock assessment models, with nonlinear functions of latent parameters and multiple types of observations with varying distributional assumptions, are one of the most complex examples of this analytical approach. Statistical aspects of state-space models and their application within fisheries have been studied extensively, but previous work has focused primarily on linear and Gaussian state-space models (Aeberhard et al. 2018; Auger-Méthé et al. 2021). Therefore, current understanding of the reliability of state-space models does not extend to usage for stock assessment. 61 As state-space models provide greater flexibility by allowing multiple processes to vary as 62 random effects (Nielsen and Berg 2014; Aeberhard et al. 2018; Stock et al. 2021), one of the most immediate questions regards the implications of mis-specification among alternative sources of process error. Incorrect treatment of population attributes as temporally varying

Application of state-space models in fisheries stock assessment and management has ex-

(Trijoulet et al. 2020; Liljestrand et al. 2024) could lead to misidentification of stock status and biased population estimates, ultimately impacting fisheries management decisions (Legault and Palmer 2016; Szuwalski et al. 2018; Cronin-Fine and Punt 2021). Furthermore, biological, fishery, and observational processes are often confounded in catch-at-age data, which may adversely affect ability to distinguish between true process variability and observational error (Li et al. In review; Punt et al. 2014; Stewart and Monnahan 2017; Cronin-Fine and Punt 2021; Fisch et al. 2023).

Li et al. (2024) conducted a full-factorial simulation-estimation study to assess model reliability when confounding random-effects processes (numbers-at-age, fishery selectivity, and
natural mortality) were included. Their results suggest that while state-space models can
generally identify sources of process error, overly complex models, even when misspecified
(i.e., incorporating process error that did not exist in reality), often performed similarly to
correctly specified models, with little to no bias in key management quantities. Similarly,
Liljestrand et al. (2024) found little downside in assuming process error in recruitment or
selectivity, even when it was absent.

Despite mounting efforts, several limitations remain. First, confounding processes that can be treated as random effects in the model were not thoroughly examined or tested within a simulation-estimation framework. Second, previous studies relied on operating models conditioned on specific fisheries, limiting their generalizability (Li et al. In review; Liljestrand et al. 2024). In particular, the effects of observation error and underlying fishing history have not been fully isolated in simulation study designs, making it challenging to disentangle the interplay between process and observation error magnitudes, as demonstrated in Fisch et al. (2023). Third, explicitly modeling stock-recruit relationships (SRRs) as mechanistic drivers of population dynamics is promising (Fleischman et al. 2013; Pontavice et al. 2022), but reliability of inferences within integrated state-space age-structured models has not been evaluated. Evidence from other studies suggests that when both process and observation errors are unknown, estimating density dependence parameters becomes highly

that stronger density dependence becomes increasingly difficult to estimate in the presence of observation error. Therefore, it is crucial to assess whether density dependence mechanisms can be estimated with sufficient precision for use in fisheries management (Auger-Méthé et al. 2016). Finally, although the importance of autocorrelation in process errors is recognized, 97 investigations of the ability to distinguish state-space assessment models with and without autocorrelation and whether such misspecification is detrimental to estimation of important population metrics are lacking (Johnson et al. 2016; Xu et al. 2019). 100 In the present study, we conduct a simulation study with operating models (OMs) varying by 101 degree of observation error, source and variability of process error, and fishing history. The 102 simulations from these OMs are fitted with estimation models (EMs) that make alternative 103 assumptions for sources of process error, whether a SRR was estimated, and whether natural mortality is estimated. Given the confounding nature of process errors, developing diagnostic 105 tools to detect model misspecification is of great scientific interest and could aid the next 106 generation of stock assessments (Auger-Méthé et al. 2021). We evaluate whether convergence 107 and Akaike Information Criterion (AIC) can correctly determine the source of process error 108 and the existence of a SRR. We also evaluate when retrospective patterns occur and the 109 degree of bias in the outputs of the assessment model that are important for management. 110

uncertain (Knape 2008; Polansky et al. 2009). In particular, Knape (2008) demonstrated

## III Methods

We used the Woods Hole Assessment Model (WHAM) to configure OMs and EMs in our simulation study (Miller and Stock 2020; Stock and Miller 2021). WHAM is an R package freely available via a github repository and is built on the Template Model Builder package (Kristensen et al. 2016). For this study we used version 1.0.6.9000, commit 77bbd94. WHAM has also been used to configure OMs and EMs for closed loop simulations evaluating index-based assessment methods (Legault et al. 2023) and is currently used or accepted for

use in management of numerous NEUS fish stocks (e.g., NEFSC 2022a, 2022b; NEFSC 2024).

We completed a simulation study with a number of OMs that can be categorized based 120 on where process error random effects were assumed: recruitment (R, assumed present in 121 all models), apparent survival (denoted R+S), natural mortality (R+M), fleet selectivity 122 (R+Se), or index catchability (R+q). We refer to the (R+S) OMs as modeling apparent 123 survival because on logscale the random effects  $(\epsilon_{a,y})$  are additive to the total mortality 124 (F+M) between numbers at age, thus they modify the survival term. However, as Stock and 125 Miller (2021) note, these random effects can be due to events other than mortality, such as 126 immigration, emigration, missreported catch, and other sources of misspecification. For each 127 OM, assumptions about the magnitude of the variance of process errors and observations 128 are required and the values we used were based on a review of the range of estimates from Northeast Unite States (NEUS) assessments using WHAM. 130

In total, we configured 72 OMs with alternative assumptions about the source and magnitude of process errors, magnitude of observation error in indices and age composition data, and contrast in fishing pressure over time. We fitted 20 EMs to observations generated from each of 100 simulations where process errors were also simulated. Each EM differed in assumptions about the source of process errors, whether natural mortality (or the median for models with process error in natural mortality) was estimated, and whether a Beverton-Holt SRR was estimated within the EM. Details of each of the OMs and EMs are described below.

We did not use the log-normal bias-correction feature for process errors or observations described by (Stock and Miller 2021) for OMs and EMs to simplify interpretation of the study results (Li et al. In review). All code we used to perform the simulation study and summarize results can be found at https://github.com/timjmiller/SSRTWG/tree/main/Project\_0/code.

## 143 Operating models

### 144 Population

The population consists of 10 age classes, ages 1 to 10+, with the last being a plus group that accumulates ages 10 and older. We assume spawning occurs annually 1/4 of the way through the year. The maturity at age was a logistic curve with  $a_{50} = 2.89$  and slope = 0.88 (Figure S1, top left).

Weight at age was generated with a von Bertalanffy growth function

$$L_a = L_\infty \left( 1 - e^{-k(a - t_0)} \right)$$

where  $t_0 = 0$ ,  $L_{\infty} = 85$ , and k = 0.3, and a L-W relationship such that

$$W_a = \theta_1 L_a^{\theta_2}$$

where  $\theta_1 = e^{-12.1}$  and  $\theta_2 = 3.2$  (Figure S1, top right).

We assumed a Beverton-Holt SRR with constant pre-recruit mortality parameters for all OMs. All biological inputs to calculations of spawning biomass per recruit (i.e., weight, maturity, and natural mortality at age) are constant in the apparent survival (R+S) selectivity (R+Sel), and survey catchability (R+q) process error OMs. Therefore, steepness and unfished recruitment are also constant over the time period for those OMs (Miller and Brooks 2021). We specified unfished recruitment equal to  $e^{10}$  and  $F_{\rm MSY} = F_{40\%} = 0.348$ , which equates to a steepness of 0.69 and a = 0.60 and  $b = 2.4 \times 10^{-5}$  for the Beverton-Holt parameterization

$$N_{1,y} = \frac{aSSB_{y-1}}{1 + bSSB_{y-1}}$$

(Figure S1, bottom right). We assumed a value of 0.2 for the natural mortality rate in OMs without process errors on natural mortality and for the median rate for OMs with process

errors on natural mortality.

We used two fishing scenarios for OMs. In the first scenario, the stock experiences overfishing at  $2.5F_{\rm MSY}$  for the first 20 years followed by fishing at  $F_{\rm MSY}$  for the last 20 years (denoted  $2.5F_{\rm MSY} \to F_{\rm MSY}$ ). In the second scenario, the stock is fished at  $F_{\rm MSY}$  for the entire time period (40 years). The magnitude of the overfishing assumptions is based on average estimates of overfishing for NEUS groundfish stocks from Wiedenmann et al. (2019) and similar to the approach in Legault et al. (2023).

We specified initial population abundance at age at the equilibrium distribution that corresponds to fishing at either  $F=2.5\times F_{\rm MSY}$  or  $F=F_{\rm MSY}$ . This implies that, for a deterministic model, the abundance at age would not change from the first year to the next.

For OMs with time-varying random effects for M, steepness is not constant. However, we used the same a and b parameters as other OMs, which equates to a steepness and R0 at the median of the time series process for M. For OMs with time-varying random effects for fishery selectivity,  $F_{\rm MSY}$  is also not constant, but since we use the same F history as other OMs, this corresponds to  $F_{\rm MSY}$  at the mean selectivity parameters.

#### 177 Fleets

We assumed a single fleet operating year round for catch observations with logistic selectivity for the fleet ( $a_{50} = 5$  and slope = 1; Figure S1, bottom left). This selectivity was used to define  $F_{\rm MSY}$  for the Beverton-Holt SRR parameters above. We assumed a logistic-normal distribution with no correlation on the multivariate normal scale for the age-composition observations for the fleet.

#### 183 Indices

Two time series of fishery-independent surveys in numbers are generated for the entire 40 year period with one occurring in the spring (0.25 of each year) and one in the fall (0.75 of

each year). Catchability of both surveys are assumed to be 0.1. Like the fishing fleet, we assumed logistic selectivity for both indices ( $a_{50} = 5$  and slope = 1) and a logistic-normal distribution with no correlation on the multivariate normal scale for the age-composition observations.

### Observation Uncertainty

The standard deviation for log-aggregate catch was 0.1. Two levels of observation error 191 variance (high and low) were specified for indices and all age composition observations (both 192 indices and catch). The low uncertainty specification assumed a standard deviation of 0.1 for 193 both series of log-aggregate index observations, and the standard deviation of the logistic-194 normal for age composition observations was 0.3. In the high uncertainty specification, 195 the standard deviation for log-aggregate indices was 0.4 and that for the age composition 196 observations was 1.5. For all EMs, the standard deviation for log-aggregate observations 197 was assumed known whereas that for the logistic-normal age composition observations was 198 estimated. 199

#### 200 Operating models with random effects on numbers at age

For operating models with random effects on recruitment and(or) apparent survival (R, R+S), we assumed marginal standard deviations for recruitment of  $\sigma_R \in \{0.5, 1.5\}$  and marginal standard deviations for older age classes of  $\sigma_{2+} \in \{0, 0.25, 0.5\}$ . The full factorial combination of these process error assumptions (2x3 levels) and scenarios for fishing history (2 levels) and observation error (2 levels) scenarios described above results in 24 different R  $\sigma_{2+} = 0$  and R+S operating models (Table S3).

### Operating models with random effects on natural mortality

All R+M OMs treat natural mortality as constant across age, but with annually varying random effects. WHAM treats natural mortality as a log-transformed parameter

$$\log M_{u,a} = \mu_M + \epsilon_{M,u}$$

that is a linear combination of a mean log-natural mortality parameter that is constant across ages ( $\mu_M = \log(0.2)$ ) and any annual random effects are marginally distributed as  $\epsilon_{M,y} \sim N(0, \sigma_M^2)$ . The marginal standard deviations we assumed for log natural mortality random effects were  $\sigma_M \in \{0.1, 0.5\}$  and the random effects were either uncorrelated or first-order autoregressive (AR1,  $\rho_M \in \{0, 0.9\}$ ). Uncorrelated random effects were also included on recruitment with  $\sigma_R = 0.5$  (hence, we denote these OMs as R+M). The full factorial combination of these process error assumptions and fishing history (2 levels) and observation error (2 levels) scenarios described above results in 16 different R+M OMs (Table S4).

#### Operating models with random effects on fleet selectivity

WHAM treats selectivity parameter s as a logit-transformed parameter

$$\log\left(\frac{p_{s,y} - l_s}{u_s - p_{s,y}}\right) = \mu_s + \epsilon_{s,y}$$

tributed as  $\epsilon_{s,y} \sim N(0, \sigma_s^2)$ , where the lower and upper bounds of the parameter ( $l_s$  and  $u_s$ ) can be specified by the user. All selectivity parameters ( $a_{50}$  and slope parameters) were bounded by 0 and 10 for all OMs and EMs. The marginal standard deviations we assumed for logit scale random effects were  $\sigma_s \in \{0.1, 0.5\}$  and AR1 autocorrelation parameters of  $\rho_s \in \{0, 0.9\}$ . Like R+M OMs, the full factorial combination of these process error assumptions (2x2 levels) and scenarios described above for fishing history (2 levels) and observation

error (2 levels) results in 16 different R+Sel OMs (Table S5).

## Operating models with random effects on index catchability

Like selectivity parameters, WHAM treats catchability for an index i as a logit-transformed parameter

$$\log\left(\frac{q_{i,y} - l_i}{u_i - q_{i,y}}\right) = \mu_i + \epsilon_{i,y}$$

that is a linear combination of a mean  $\mu_i$  and any annual random effects marginally distributed as  $\epsilon_{i,y} \sim N\left(0, \sigma_i^2\right)$  where the lower and upper bounds of the catchability ( $l_i$  and  $u_i$ ) can be specified by the user. We assumed bounds of 0 and 1000 for all OMs and EMs. For all OMs and EMs with process errors on catchability, the temporal variation only applies to the first index, which could be interpreted as capturing some unmeasured seasonal process that affects availability to the survey. The marginal standard deviations we assumed for logit scale random effects were  $\sigma_i \in \{0.1, 0.5\}$  and AR1 autocorrelation parameters of  $\rho_i \in \{0, 0.9\}$ . Like R+M and R+Sel OMs, the full factorial combination of these process error assumptions and fishing history (2 levels) and observation error (2 levels) scenarios described above results in 16 different R+q OMs (Table S6).

### Estimation models

For each of the data sets simulated from an OM, 20 EMs were fit. A total of 32 different EMs were fit across OMs where the subset of 20 depended on the source of process error in the OM (Table S7). The EMs have different assumptions about the source of process error (R+S, R+M, R+Sel, R+q) and whether or not 1) there is temporal autocorrelation, 246 2) a Beverton-Holt SRR is estimated, and 3) the natural mortality rate ( $\mu_M$ , the constant or mean on log scale for R+M EMs) is estimated. For simplicity we refer to the derived estimate  $e^{\mu_M}$  as the median natural mortality rate regardless of whether natural mortality random effects are estimated in the EM.

Subsets of 20 EMs in Table S7 were fit to simulate data sets from each of the OM process error categories. For R and R+S OMs, fitted EMs had matching process error assumptions 251 as well as R+Sel, R+M, and R+q assumptions without autocorrelation. Similarly, For 252 other OM process error categories, we fit EMs with matching process error assumptions as 253 well as other process error types without autocorrelation. The maturity at age, weight at 254 age for catch and SSB, and observation error variance of aggregate catch and indices were 255 all assumed known at the true values. However, the variance parameters for the logistic-256 normal distributions for age composition observations were estimated in the EMs. As such, 257 EMs would either be configured completely correctly for the OM, or there could be mis-258 specification in assumptions of process error autocorrelation, the type of process error, or 250 the SRR (Beverton-Holt or none). 260

## 261 Measures of reliability

## 262 Convergence

The first measure of reliability we investigated was frequency of convergence when fitting each EM to the simulated data sets. There are various ways to assess convergence of the fit (e.g., Carvalho et al. 2021; Kapur et al. 2025), but given the importance of estimates of un-265 certainty when using assessment models in management, we estimated probability probability 266 of convergence as measured by occurrence of a positive-definite hessian matrix at the opti-267 mized negative log-likelihood that could be inverted (i.e., providing hessian-based standard 268 error estimates). We also provide results in the Supplementary Materials for the maximum 269 convergence defined by the maximum absolute gradient  $< 1^{-6}$  and the maximum of the 270 absolute values among all gradients gradient values for all fits of a given EM to all simulated 271 data sets from a given OM that produced hessian-based standard errors for all estimated 272 fixed effects. This provides an indication of how poor the calculated gradients can be, but still presumably converged adequately enough for parameter inferences. We used the <sup>275</sup> Clopper-Pearson exact method for constructing 95% confidence intervals of the probabilities <sup>276</sup> of convergence (Clopper and Pearson 1934; Thulin 2014).

#### AIC for model selection

We estimated the probability of selection also investigated the reliability of AIC-based model selection for two purposes. First, we analyzed selection of each process error model structure (R, R+S, R+M, R+Sel, R+q) using marginal AIC. For a given operating modelOM simulated data set, we compared AIC for EMs that were all configured the same for with 281 different process error assumption conditional on whether median natural mortality (known 282 or estimated) and the SRR (rate and the Beverton-Holt or none). 283 We also estimated the probability of correctly selecting stock-recruit relationship were 284 estimated. We tabulated the models providing the lowest AIC across simulated data sets. 285 Second, we analyzed AIC-based selection between EMs with and without the Beverton-Holt 286 SRR assumed and models without the SRR (null model). We made these comparisons 287 between models that otherwise assumed the same process error structure as the operating 288 model and both of the compared models either estimate median natural mortality or 289 assume it is knownstock-recruit relationship assumed. Contrast in fishing pressure and time series with recruitment at low stock size has been shown to improve estimation of SRR 291 parameters (Magnusson and Hilborn 2007; Conn et al. 2010). Our preliminary inspections 292 of the proportions of simulations where the correct recruitment model was chosen indicated generally poor performance of AIC in determining the Beverton-Holt model for a given set 294 of OM factors (including contrast in fishing pressure) indicated generally poor performance 295 of AIC. Therefore, we fit logistic regression models to the indicator of Beverton-Holt 296 models having lower AIC as a function also considered the effect of of the log-standard 297 deviation of the true log(SSB) (similar to the log of the coefficient of variation for SSB) on 298 model selection since simulations with realized SSB producing low and high recruitments recruitment would have larger variation in realized SSB.

All model selection results condition on whether all of the compared estimating models 301 completed the only on completion of the optimization process without failure for all of the 302 compared EMs. We did not condition on convergence as defined by a gradient threshold 303 or invertibity of the hessian above because optimization could correctly determine an inap-304 propriate process error assumption by estimating variance parameters at the lower bound 305 of zero. Such an optimization could indicate poor convergence but the likelihood would be 306 equivalent to that without the mis-specified random effects and the AIC would be appro-307 priately higher because more (variance) parameters were estimated. All other measures of 308 reliability described below (bias and Mohn's  $\rho$ ) use these same criteria for inclusion of EM 309 fits in the summarized results.

#### 311 Bias

314

We also investigated bias in estimation of various model attributes as a measure of reliability.

For a given model attribute we calculated the relative error

$$RE(\theta_i) = \frac{\widehat{\theta}_i - \theta_i}{\theta_i}$$

 $\operatorname{RE}\left(\theta_{i}\right) = \frac{\widehat{\theta}_{i} - \theta_{i}}{\theta_{i}} \tag{1}$  from fitting a given estimating model to simulated data set i configured for a given OM

where  $\hat{\theta}_i$  and  $\theta_i$  are the estimated and true values for simulation i. We estimated bias as the median of the relative errors across all simulations for a given OM and EM combination.

We constructed 95% confidence intervals for the median relative bias using the binomial distribution approach as in Miller and Hyun (2018) and Stock and Miller (2021). We present results for bias in analyzed simulation results for estimates of terminal year estimates of SSB and recruitment, Beverton-Holt stock recruit parameters (a and b), and median natural

mortality rate. Results for terminal year fishing mortality were strongly negatively correlated
with those for SSB and are provided in the Supplementary Materials.

### Mohn's $\rho$

Finally, we investigated presence of retrospective patterns in fitted models as a measure of reliability. We calculated Mohn's  $\rho$  for SSB, fully-selected fishing mortality fishing mortality (averaged over all age classes), and recruitment for each EM fit to each OM simulated data set (Mohn 1999). We fit 7 peels for each EM and calculated median 95% confidence intervals for P = 7 peels to each simulated data set and calculated Mohn's  $\rho$  using the same methods as that for relative biasfor a given attribute  $\theta$  as

$$\rho(\theta) = \frac{1}{P} \sum_{p=1}^{P} \frac{\widehat{\theta}_{Y-p,Y-p} - \widehat{\theta}_{Y-p,Y}}{\widehat{\theta}_{Y-p,Y}}$$
(2)

where  $\hat{\theta}_{i,j}$  is the estimate for attribute  $\theta$  in year i from a model fit using data up to year j.

Results

### Summarizing results across OM and EM attributes

For many R and R+S OMs, convergence rate declined when either the median natural 334 mortality rate or the Beverton-Holt SRR was estimated even when the process error 335 assumptions of the EMs and OMs matched (Figure S14, A). When there was high 336 observation error and constant fishing pressure ( $F = F_{MSY}$  for all 40 years), convergence 337 was poor for of all The measures of central tendency and variability of observed values for 338 specific OM and EM attributes (e.g., low or high observation error) that we described above 339 can indicate scenarios that provide better or poorer reliability. However, the OM and EM 340 attributes that we investigated are numerous, so we used two methods to summarize the 341 most important factors for differences in results. The first method was fitting regression 342

models with the response being each of the measures of reliability described above and predictor variables were defined based on OM and EM characteristics (e.g., MacKinnon et 344 al. 1995; Wang et al. 2017; Harwell et al. 2018). For the binary indicators of convergence 345 and AIC-based selection of a stock-recruit relationship, we performed logistic regressions. 346 For indicators of AIC-based selection of EM process error configurations other than R EMs 347 when fitted to R OMs ( $\sigma_{2+} = 0$ ) regardless of whether median natural mortality and SRRs 348 were estimated. Convergence of R EMs was high for all R and R+S OMs except when there 349 was high observation error and constant fishing pressure, type (multiple categories) we 350 performed multinomial regressions. For other measures of reliability we fit linear regression 351 models to transformed responses. Because relative errors (Eq. 1) and when median natural 352 mortality and SRRs were estimated. R+S EMs fit to R OMs exhibited poor convergence 353 regardless of whether natural mortality or a SRR was estimated. R+S EMs fit to R+S 354 OMs had highest convergence rates when there was contrast in fishing pressure and low 355 observation error. Convergence rates were high for all EMs when fit to data from R+S 356 OMs with lower observation error except those where median natural mortality and/or 357 SRRs were estimated. Mohn's  $\rho$  for the various parameters are bounded below at -1, we 358 used a transformation of these values

$$y_i = \log\left[f\left(\widehat{\theta}_i, \theta_i\right) + 1\right] \tag{3}$$

where f is either the relative error (Eq. 1) or Mohn's  $\rho$  (Eq. 2) for simulation i, so that values are unbounded. For relative errors,  $y_i$  is the log-scale error. We omitted simulation estimates equal to zero (RE = -1). For all regressions we fit separate models with individual factors included, with all factors combined, with including all second order interactions, and including all third order interactions. For the multinomial regression, we used the vglm function from the VGAM package (Yee 2008; Yee 2015). We tabulated percent reduction in residual deviance for each of regression fits. We did not perform formal statistical

analyses of effects of OM and EM attributes on results (e.g., ANOVA) because of the lack of independence of the "observations" that results from fitting multiple EMs to each simulated 368 data set. 369 Convergence of all EMs fitted to R+M OMs was highest when the OMs had higher natural 370 mortality process error variability, low observation error, and contrast in fishing pressure 371 (Figure S14, B) The second method involved fitting classification and regression trees 372 (Breiman et al. 1984) to show how the OM and EM attributes, and their interactions, split 373 the values for each measure of reliability (e.g., Gonzalez et al. 2018; Collier et al. 2022). 374 We used classification trees for categorical measures (convergence and AIC) and regression 375 trees for the other measures with continuous scales (relative error and Mohn's  $\rho$ ). The 376 response variables were the same as the regressions for the deviance reduction analyses. We 377 used the rpart function in the rpart package (https://cran.r-project.org/package=rpart) 378 to fit trees. Full trees were determined using default settings except that we increased the 379 number of cross-validations to 100. For clarity, we pruned the full trees to show just the 380 primary branches. We also provide detailed results for all measures of reliability at each combination of OM 382 and EM attributes in the Supplementary Materials. For confidence intervals of probability 383 of convergence, we used the Clopper-Pearson exact method (Clopper and Pearson 1934; 384 Thulin 2014). For AIC selection of process error configuration we provide estimates of the 385 proportions of simulations where each EM type was selected. For AIC selection of the 386 stock-recruit relationship, we provide predicted probabilities from logistic regressions as a 387 function of SSB variability for each OM and EM type. We estimated bias as the median of the 388 relative errors across all simulations for a given OM and EM combination. We constructed 389 95% confidence intervals for the median relative bias, and Mohn's  $\rho$  using the binomial 390 distribution approach as in Miller and Hyun (2018) and Stock and Miller (2021). For each 391 EM we calculated median and 95% confidence intervals using the same methods as that for relative bias.

## Results

## Convergence performance

For probability of convergence, EM process error assumption was the single attribute that 396 resulted in the largest percent reduction in deviance for all OM process error types other 397 than R+M EMs that estimated autocorrelation of process errors had poor convergence for 398 R+M OMs when there was low natural mortality process error variability regardless of 399 autocorrelation of the simulated process errors. R+S EMs fitted to data generated from 400 R+M OMs always converged poorly whether or not median natural mortality and the 401 Beverton-Holt SRR were estimated. 402 The R+S EMs, in particular, had poor convergence when fit to data generated from R+Sel 403 OMs with lower selectivity process error variability or higher observation error (Figure S14, C). OMs where the EM M assumption explained the most residual deviance (Table 1). 405 However, including interactions of OM and EM factors also provided large reductions in residual deviance, suggesting successful convergence depended on a combination OM and 407 EM attributes. Classification trees for each OM process error type, all had the primary 408 branch defined using the same attribute that provided the largest reduction in deviance 409 (Figure 1). EMs that assumed R+Sel EMs generally converged better than other EMs for S 410 process errors converged poorly for all OMs that were simulated with the alternative process 411 error assumptions (R, R+Sel OMs with higher process error variability, lower observation 412 error, and contrast in fishing pressure regardless of whether median natural mortality or a 413 SRR was estimated. 414 For M, R+q OMs, convergence of Sel, an R+q EMs was generally better than that of 415 other EMs when there was contrast OMs). For all trees, branches based on the OM fishing 416 mortality history showed better convergence when the OM included a change in fishing pressure (Figure (S14, D). Convergence of R+S EMs was generally worse than that of all other EMs across all OMs whether or not. Branches based on whether the Beverton-Holt
stock-recruit model was assumed or not, showed better convergence when it was not
estimated and branches based on whether the EM estimated the median natural mortality
or a SRR was estimated. Again, convergence probability generally declined for all EMs
when median natural mortality or a SRR was estimated at showed better convergence
when it was treated as known. For certain R+M and R+Sel OMs, better convergence was
also observed when there was lower observation uncertainty.

When convergence is defined by a gradient threshold, the primary factor explaining deviance reduction is the same for all OM process error types, but there are some differences in deviance reduction for secondary factors (Table S8), and probability of convergence, overall, was lower (Figure S2). We found a wide range of maximum absolute values of gradients for models that converged had invertible hessians (Figure S3). The largest value observed for a given EM and OM combination was typically < 10<sup>-3</sup>, but many converged models had values greater than 1. For many OMs, EMs that assumed the correct process error type and did not estimate median natural mortality or the Beverton-Holt SRR produced the lowest gradient values.

## 435 AIC performance

436 Marginal AIC accurately determined

#### 437 Process error structure

For AIC selection of the correct process error assumptions in EMs when data were generated from R and configuration, the magnitude of observation and process error were the attributes that resulted in the largest percent reductions in deviance across OM process error types other than R OMs (Table 2). Both variance of apparent survival random effects ( $\sigma_{2+}$ ) and degree of observation error explained the largest reductions for R+S OMs and R+Sel OMs,

whereas variance of process errors provided the largest reductions in R+M and R+q OMs. Comparatively, none of the OM or EM attributes explained particularly large reductions 444 in deviance for R OMs, but fishing history, whether a stock-recruit model was estimated, 445 regardless of whether median natural mortality or a SRR was estimated (Figure S16, A). 446 Attempting to estimate and whether median natural mortality or a SRR separately had a 447 negligible effect on the accuracy of determining the correct process error assumption. When 448 both were estimated, there was a noticeable reduction in accuracy when OMs had a constant 449 fishing pressure, low observation error, and larger variability in recruitment process errors. 450 was known or estimated provided similar and the largest reductions. Inclusion of second 451 and third order interactions, did not provide large reductions in deviance for any of the OM 452 process error types. 453 For R+M OMs, marginal AIC only accurately determined the correct process error model 454 and correlation structure when observation error was low and variability in natural mortality 455 process errors was high (Figures S16, B). Of these OMs, estimating the median natural 456 mortality rate only reduced the accuracy of AIC when natural mortality process errors were 457 independent and fishing pressure was constant. For OMswith poor model selection accuracy, 458 AIC most frequently selected EMs with process errors in catchability (R+q)or selectivity 459 (R+Sel). Selection of R+S EMs was generally unlikely. 460 Marginal AIC most accurately determined the correct source of process error and correlation 461 structure for R +Sel OMs with low-all OM process error types other than R OMs, the 462 attributes defining the primary branches of classification trees matched those that provided 463 the largest reductions in deviance (Figure 2). Across all OMs, AIC was more accurate for 464 the process error type when process error variability was greater and when observation error 465 was lower. No branches were estimated for classification trees fit to the R OMs, likely because accuracy was high across all simulations (0.94), although inspection of the fine-scale results shows there is some degradation in AIC selection when a stock recruit relationship and median natural mortality rate are estimated for R OMs with constant fishing pressure

and high observation error (Figures S16, C). When there was low variability in selectivity process errors and high observation error, Figure S16, top left). 471

#### Stock-recruit relationship 472

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Logistic regressions for AIC selection of the Beverton-Holt stock recruit relationship, showed OM fishing history and variation in SSB (log SD<sub>SSB</sub>) provided substantial reductions in deviance for R+q or M, R+S EMs were more likely to have the best AIC. Whether median 475 natural morality or SRRs were estimated appeared to have little effect on the performance 476 of AIC. 477 Marginal AIC most accurately determined the correct source of process error and correlation 478 structure for Sel, and R+q OMs with high variability in catchability process errors (Figures 479 S16,D). The R (Table 3). For R OMs, fishing history provided the largest reduction in 480 deviance whereas none of the attributes individually provided large reductions in deviance 481 for R+q OMswith low variability in catchability process errors and high observation error 482 had the least model selection accuracy. However, for these OMs, the marginal AIC accurately 483 determined the correct source of process error (but not correlation structure) except when 484 OMs assumed a constant fishing pressure and EMs estimated both median natural morality 485 and the SRR. Our comparisons of model performance conditioned on assuming the true process error 487 configuration is known (EM and OM process error types match) and we focus on results where 488 the EMs assume median natural mortality is known because there was little difference in 489 results when the EMs estimated this parameter. Broadly, we found generally poor accuracy 490 of AIC in selecting models assuming a Beverton-Holt SRR over the null model without an 491 SRR for all S OMs. However, we also found increased accuracy of AIC in determining the 492 Beverton-Holt SRR when the simulated population exhibited greater variation in spawning 493 biomass for nearly every OM (Figure S17).

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With inclusion of all attributes provided larger reductions in deviance than the sum of
    individual contributions for both R and R+S process error assumptions, probability of lowest
496
    AIC for the B-H SRR as a function of SSB variability were greatest for OMs with contrast
497
    in fishing pressure and lower process variability in recruitment (Figure S17, A). The largest
498
    OMs. For all OM process error types, inclusion of interaction terms provided relatively little
499
    reduction in residual deviance.
500
    Attributes defining the primary branches of classification trees for AIC selection of the
501
    stock-recruit assumption were the same as those explaining the largest reductions in deviance
502
    for the logistic regression models (Figure 3). All branches based on variation in SSB occurred
503
    in OMs with larger recruitment variability (\sigma_R = 1.5; Figure S17, A, right column group),
504
    but the same high AIC accuracy was achieved for OMs with lower recruitment variability at
505
    lower levels of SSB variation. The level of observation error had little effect on AIC accuracy
507
    For R +M OMs, probability of lowest AIC for the Beverton-Holt SRR increased steeply with
508
    variation in SSB whether it was induced by showed better accuracy with larger variability
509
    in SSB and all branches based on fishing history showed better accuracy when there was
510
    contrast in fishing pressure. Branched based on OM observation error or recruitment
511
    variability showed better accuracy when they were lower. For R OMs, a combination of
512
    lower recruitment variability, contrast in fishing or variation in natural mortality process
513
    error. (Figure S17, B). There was little difference in AIC accuracy whether the natural
514
    mortality process errors were correleted and, similar to R+S OMs, there was also little
515
    effect due to level of observation error.
516
    pressure, and higher SSB variability produced AIC accuracy over 0.8. For R+Sel OMs,
517
    contrast in fishing pressure over time was the primary source of variation in SSB and these
518
    are the OMs where AIC accuracy for the Beverton-Holt SRR was greatest (Figure S17, C).
519
    There was little effect of variability or correlation of selectivity process errors or the level of
```

observation error on AIC accuracy.

Like the S OMs, lower recruitment variability and observation error and higher SSB variability produced AIC accuracy of 0.79. For R+Sel OMs, the greatest accuracy for AIC in selecting the Beverton-Holt SRR occurred for M, R+Sel, and R+q OMs where there was contrast in fishing pressure over time which is also where there was the greatest variation in SSB (Figure S17,D). There was also little effect of variability or correlation of catchability process errors or the level of observation error on AIC accuracy OM,s accuracy of 0.87 to 0.94 was observed just with higher SSB variability.

## Bias Bias

Spawning stock biomass For R OMs ( $\sigma_{2+}=0$ ), there was no indication of bias 530 (95%confidence intervals included 0) in terminal year SSB Regression models for log-scale 531 errors in SSB that included the various OM and EM factors showed little reduction in 532 deviance (< 4.3%) for any of the estimating models regardless of process error assumptions, 533 except when no SR assumption was made, recruitment variability was low, and there was 534 contrast in fishing mortality and high observation error (Figure S6, A). However, errors 535 in terminal SSB estimates were highly variable when factors across all OM process error 536 types (Table 6). The attributes producing the largest reductions were the treatment of 537 median natural mortality was estimated and there was constant fishing pressure and high 538 observation error (Figure S6, A, second row). 539 For (for R, R+S OMs, the EMs with matching process error assumptions generally produced 540 unbiased estimation of terminal SSB except when median natural mortality was estimated 541 and there was high observation error. In M. R+S OMswith low observation error, EMs with 542 incorrect process error assumptions typically provided biased estimation of terminal year 543 SSB. Estimating the Beverton-Holt SRR had little discernible effect on bias of terminal year SSB estimation whereas estimating median M tended to produce more variability in errors 46 in terminal SSB estimation similar to ROMs.

For RSel, and R+M OMswith low variability in natural mortality process errors, low 547 observation error and contrast in fishing motality over time all EMs produced low variability 548 in SSB estimation error that indicated unbiasedness (Figure S6, B, third row). However, 549 larger variability in natural mortality process errors increased bias of EMs without the 550 correct process error type. Estimating median natural mortality increased variability of 551 SSB estimation error particularly for OMs with high observation error and constant fishing 552 pressure over time. It also increased bias in SSB estimation for many Rq OMs), EM process 553 error type (R+M OMs. Like RS OMs) and fishing history (all OM types). Including second 554 interactions provided larger further reductions in residual deviance, between 6 and 15%. 555 Including third order interactions also provided large further reductions for R, R+S, and R+S OMs, estimating a SRR had little discernible effect on SSB biasq OMs (between 5 and 11%). 558 For R+Sel OMs, there was no evidence of bias for any EMs when variability in selectivity 550 process error and observation error was low, and generally all regression trees, all branches 560 based on fishing history and level of observation error showed lower bias in SSB with 561 contrast in fishing mortality (Figure S6, C). The largest bias occurred for any EMs that 562 estimated median natural mortality when the OMs had high observation error, constant 563 fishing pressure, and greater variability in selectivity process errors ( $\sigma_{\text{Sel}} = 0.5$ ) or low 564 selectivity process errors ( $\sigma_{Sel} = 0.1$ ) and low and lower observation error. However, there 565 was no evidence of SSB bias for correctly specified For branches based on treatment of 566 median natural mortality rate, bias was generally lower when it was known rather than 567 estimated. For some R+Sel EMs when observation error was low and variation in selectivity 568 process errors was larger, whether median natural mortality was estimated or not (Figure S6, C, third row). We only observed an effect of estimating the and R+q OMs, less bias in SSB was shown when the EM process error configuration was correct.

### Stock-recruit parameters

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Regression models for transformed relative errors of estimates of both the Beverton-Holt
573
   SRR for a and b parameters showed none of the factors explained large percent reductions
574
   in deviance (Table 4). The OM fishing history provided the largest reduction for most OM
575
   process error types for both parameters, but percent reductions were less than 5.6% except for
576
   R+Sel OMs that had high observation error and contrast in fishing pressure where estimating
577
   the SRR produced less biased SSB estimation for many EMs (Figure S6, C, top row).
578
   All EMs fit to data from where the percent reductions were 11.37\% and 7.97\% for the a and b
579
   parameters, respectively and for just the b parameter for R+q OMs with low observation error
580
   and contrast in fishing pressure exhibited little evidence of bias in terminal SSB estimation
   except for R+M EMs when there was no AR1 correlation in catchability process errors
582
   (Figure S6, D). Many EMs also performed well in R+q OMs with low observation error, but
583
   no contrast in fishing pressure. (10%). The EM process error assumption provided similar
584
   reductions in deviance for both parameters for R OMs. Including interactions also did not
585
   produce important reductions in deviance (increase of .
586
   For regression trees of log-scale error in Beverton-Holt a and b parameters, lower bias was
587
   indicated with contrast in OM fishing pressure for all branches in trees for each OM process
588
   error type (Figures 4 and 5). For all branches based on recruitment variability in trees for
589
   R +q OMs with high observation error and contrast in fishing pressure, EMs that estimated
590
   the Beverton-Holt SRR exhibited less SSB bias than those that did not. Estimating median
591
   natural mortality in the EMs only resulted in much more variable SSB estimation errors
592
   when there was no contrast in fishing pressure (Figure S6, D, first and third rows).
593
   For all OM process error types, relative errors in terminal year recruitment were generally
594
   more variable than SSB, but effects of R and R+S OM and EM attributes on bias (i.e,
595
   negative or positive or none) were similar (Figure S8, A). Furthermore, for EM configurations
596
   where bias in terminal SSB was evident, median relative errors in recruitment often indicated
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stronger bias in recruitment of the same sign.

OMs, lower bias in both a and b was observed with less recruitment variability. For R 599 OMs with contrast in fishing pressure and greater recruitment variability EMs that assumed 600 the incorrect R+M process errors produced lower bias in both a and b than other process 601 error assumptions. Across all OMs, there was generally less bias and(or) lower variability in 602 estimation of the Beverton Holt a parameter than the b parameter. In R and R+S OMs, 603 EMs with the correct assumptions about process errors provided the least biased estimation 604 of Beverton-Holt SRR parameters when there was a change in fishing pressure over time 605 and lower variability of recruitment process errors, but there was little effect of estimating 606 (Figure S9). 607

### 608 Median natural mortality rate

Fitted regression models for log-scale errors in median natural mortality and a small increase 609 in bias for those OMs that had high observation error (Figure S9, A). For other Rand Rrate 610 showed largest percent reductions in residual deviance for R+S OMs, estimating natural 611 mortality often resulted in less biased estimation of SRR parameters. There was generally 612 large variability in relative errors of the SRR parameter estimates, but the lowest variability 613 occurred with low variability in recruitment and little or no variability in survival process 614 errors  $(\sigma_{2+} \in \{0, 0.25\})$ , and contrast in fishing pressure. 615 In R+M OMs, the most accurate estimation of SRR parameters for all and R\_M 616 models (Table 5). The largest reductions for a single attribute was the EM process error 617 assumptions occurred when there was a change in fishing pressure, greater variability in 618 natural mortality process errors, and lower observation error (Figure S9, B). Relative to the 619 R, and Rassumption (>20%) and fishing history (>15%) for R+S OMs, there was even less 620 effect of estimating median natural mortality on estimation bias for the SRR parameters. 621 Bias for SRR parameters was large and variability in relative errors was greatest for most

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EMs fit to. Fishing history also provided >10% reduction for R+Sel OMswith constant
   fishing pressure (Figure S9, C). Less bias in parameter estimation occurred for OMs with
624
   a change in fishing pressure and the best accuracy occurred for those OMs that had low
625
   observation error and more variable and uncorrelated selectivity process errors, and when
626
   the EMs had with the correct process error assumption. There was little effect of estimating
627
   natural mortality on relative errors for SRR parameters.
628
   Like—M OMs, but reductions for all factors in R, R+SelOMs, relative errors in SRR
629
   parameters for and R+q OMs were more accurate for most EM process error types
630
   when OMs had contrast in fishing pressure and lower observation error (Figure S9, D).
631
   However, the best accuracy occurred for those OMs that had lower variability in catchability
632
   process errors. The worst accuracy of SRR parameter estimation regardless of EM type
633
   occurred when relatively low (<6%). Interactions of OM and EM factors also provided
   substantial further reductions for R+q OMs had low observation error and constant fishing
635
   pressure (Figure S9, D, fourth rowS and R+M OMs (between 8 and 15 % for second order
636
   interactions).
637
   Across all OMs and EMs there was little effect of estimating SRRs on the bias in estimation of
638
   Regression trees with branches based on fishing history showed lower bias in median natural
639
   mortality (Figure S10). Median natural mortality rate was estimated accurately by all EM
640
   process error types for all R OMs except those with high observation error and constant
641
   fishing pressure, in which case relative errors were high (Figure S10, A, \sigma_{2+} = 0). For R+S
642
   OMs estimation of median natural mortality rate was most accurate when observation error
643
   was low and there was rate with contrast in fishing pressure and the EM process error type
   was correct.
645
   branches based on level of observation error showed lower bias with more precise observations
646
   (Figure 6). For R +M OMs, median natural mortality was estimated most accurately,
647
   regardless of EM process error type, when OMs had a change in fishing pressure and
648
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low observation error (Figure S10, B). However, those R +M OMsthat also had greatest variability in AR1 correlated natural mortality process errors only had unbiased estimation 650 when the OMs, branches based on EM process error type was correct. 651 All EM process error types accurately estimated median natural mortality rate for R 652 assumption showed lower bias with EMs assuming the correct R and the incorrect R+Sel 653 OMs that had contrast in fishing pressure, low observation error, and low selectivity process 654 error variability (Figure S10, C). When selectivity process error variability increased, the 655 incorrect EM process errors produce more biased estimation of median natural mortality 656 rate. The least accurate estimation occurred for all EM process error types when observation 657 error was high and fishing pressure was constant. 658 Like S assumption. For R+S and R+Sel OMs, all M OMs, branches based on EM pro-659 cess error types produced accurate estimation of median natural mortality rate when fit 660 to R+q OMs with contrast in fishing pressure, low observation error and low catchability 661 process error variability (Figure S10, D). Most showed only the correct EM process error 662 types produced biased estimation of median natural mortality when R+q OMs had high 663 observation error and constant fishing pressure assumption with less bias. 664

## Mohn's ho

## Deviance reduction and regression trees

Fits of glms showed little explanatory power of any of the OM an EM attributes for Mohn's  $\rho$  for SSBwas small in absolute value for all R and R+S OMs, regardless of EM process error types, and whether median natural mortality rate or SRRs were estimated (Figure S11, A). The strongest retrospective patterns (highest absolute Mohn's  $\rho$  values) occurred in OMs with the largest apparent survival process error variability, high observation error, and contrast in fishing pressure, but only for EMs with the incorrect process error type and where median natural mortality rate was assumed known (median  $\rho$  was approximately

-0.15). For R+M, R+Sel, and R+q OMs, rho for SSB, F, and R. Note that the factors that showed the largest reductions in deviance (which were very small) often do not comport 675 with factors that show differences in median Mohn's  $\rho$  was also small in absolute value, but 676 median values were all closer to 0 than the largest values in the R and R+S OMs (Figure 677 S11,B-D). For these OMs, there was no noticeable effect of estimation of median natural 678 mortality rate or SRRs on rho results. This is because the glms and regression trees are 679 reducing residual variance and there could be large variability in values, but the medians are 680 less sensitive to extreme values. For example we can see that there is large Mohn's  $\rho$  for any 681 EM process error types. 682 Mohn's  $\rho$  for recruitment was small in absolute value for all rho for R in R OMs with low 683 variability in recruitment process errors, regardless of EM process error type, and whether 684 median natural mortality rate or SRRs were estimated (Figure S13, A). However, R and 685 R+S OMs with greater recruitment process variability and higher observation error had 686 median Mohn's  $\rho$  for recruitment greater than zero for most EMs even when the EM process 687 error type was correct. In R+S OMs with lower observation error, EMs with the correct 688 process error type exhibited better median Mohn's  $\rho$  close to 0 than EMs with the incorrect 689 process error type. For R+M, R+Sel, and R+q OMs, results for Mohn's  $\rho$  for recruitment 690 are similar to those for SSB, but the range in median values and variation in Mohn's  $\rho$ 691 values for a given OM are generally larger for recruitment (Figure S13, B-Dlarger variability) 692 assumed in recruitment and higher observation error (Figure X), but these factors suggest 693 little reduction in residual deviance (Table X). 694

## Discussion

## Convergence

Analyses of model convergence across simulations can be useful for understanding the util-697 ity of alternative convergence criteria used in applications to real data for directing the 698 practitioner to more appropriate random effects configurations. It is common during the 699 assessment model fitting process to check that the maximum absolute gradient component 700 is less than some threshold prior to inspecting the Hessian of the optimized likelihood for 701 invertibility (Carvalho et al. 2021). However, there is no accepted standard for the gradient 702 threshold (e.g., Lee et al. 2011; Hurtado-Ferro et al. 2014; Rudd and Thorson 2018) and 703 some thresholds would exclude models that in fact have an invertible Hessian. We found the 704 Hessian at the optimized log-likelihood can often be invertible when the maximum absolute 705 gradient was much larger than what would perceived to be a sensible threshold. 706 Li et al. (2024) found that convergence rate could be a useful diagnostic especially for separating the correct model from overly complex models. However, the criteria for convergence 708 used in their study may also lead to limited ability to distinguish the correct model from 709 overly simplistic models, a pattern that was also noted by Liljestrand et al. (2024) in which 710 one process error may absorb all sources of process error when the magnitude of other process 711 errors are low. 712 Often poor convergence result when parameter estimates are at their bounds (Carvalho et al. 713 2021), and this also applies to variance parameters for random effects with state-space assess-714 ment models. Even when the Hessian is invertible, parameters that are poorly informed will 715 have extremely large variance estimates. This further inspection can lead to a more appropri-716 ate and often more parsimonious model configuration where the problematic parameters are 717 not estimated. For example, process error variance parameters that are estimated close to 0 indicates that the random effects are estimated to have little or no variability and removing

these process errors is warranted. Generally, our results suggest we can expect lower probability of convergence of state-space assessment models when estimating natural mortality or SRRs because of the difficulty distinguishing these parameters from others being estimated 722 in assessment model with data that are typically available. Our experiments did not aim to 723 emulate the practitioner decision process in developing model configurations (e.g. removing a 724 source of process error and refitting the model when process error variance parameters were 725 estimated close to 0). Evaluating the efficacy of such a decision process when applying EMs 726 might be important in closed loop simulations (e.g. MSE) aimed at quantifying management 727 performance. 728

A factor affecting the convergence criteria, particularly for maximum likelihood estimation 729 of models with random effects, is numerical accuracy. All optimizations performed in these 730 simulations are of the Laplace approximation of the marginal likelihood and, therefore, gradients and Hessians are also with respect to this approximation (see TMB::sdreport in the Template Model Builder package). Functionality within the Template Model Builder pack-733 age exists (i.e., TMB::checkConsistency) to check the validity of the Laplace approximation and the utility of this as a diagnostic for state-space assessment models should be explored 735 further. Furthermore, numerical methods are used to calculate and invert the Hessian for 736 variance estimation for models with random effects. Along with our results, the potential 737 lack of accuracy imposed by these approximations, suggests at least investigating whether 738 the Hessian is positive definite when the calculated absolute gradients are not terribly large 739 (e.g. < 1).740

## AIC

Of the OM process error configurations we considered, we found AIC to be accurate for selecting models with process errors on recruitment and apparent survival (R and R+S). Fitting models to other OMs rarely preferred R+S EMs, and R and R+S EMs were nearly

always selected for the matching OMs; a similar result was reported by Liljestrand et al. (2024). For other sources of process error, accuracy of AIC was improved when there was larger variability in the process errors and/or lower observation error.

Across all OM process error configurations, AIC performed poorly in identifying that the 748 presence of the Beverton-Holt SRR in the OM unless there was contrast in fishing pressure 740 possibly in combination with other factors such as lower variability in recruitment process 750 errors (in R and R+S models) or greater variation in natural mortality process errors (for 751 R+M OMs, Fig. S17). As such, properly accounting for process error in natural mortality 752 could be important (Li et al. 2024) when evaluating SRRs in state-space models. Curiously, 753 we did not find a marked effect of the level of observation error on ability to detect the SRR, 754 but it is possible that AIC would perform better if observations have even lower uncertainty 755 than we considered.

Although we did not compare models with alternative SRRs (e.g., Ricker and Beverton-757 Holt), we do not expect AIC to perform any better distinguishing between relationships. 758 Our finding that AIC tended to choose simpler recruitment models in most cases contrasts 759 with the noted bias in AIC for more complex models (Shibata 1976; Katz 1981; Kass and 760 Raftery 1995), but, whereas those findings apply to the much more common comparison of 761 models that are fit to raw and independent observations, here we are comparing state-space 762 models which account for observation error and estimate process errors in latent variables. 763 Our results comport with those of de Valpine and Hastings (2002) who found AIC could not 764 distinguish among state-space SRRs that were fit just to SSB and recruitment observations (i.e., not an assessment model). Similarly, Britten et al. (In review) found AIC could not reliably distinguish alternative environmental effects on SRR parameters. However, Miller

et al. (2016) did find AIC to prefer a SRR with environmental effects when applied to data

for the SNEMA yellowtail flounder stock and AIC also selected an environmental covariate

on a SRR for the most recent stock assessment of Georges Bank yellowtail flounder (NEFSC

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769

2025). Both of these yellowtail flounder stocks have large changes in stock size and the values of environmental covariates over time. Additionally, this species is well-observed by the bottom trawl survey that is used for an index in assessment models.

#### **Bias**

As expected, bias in all parameters and assessment output was generally improved with lower observation error. Estimation of SRR parameters was reliable in ideal scenarios of 776 low observation error and contrast in fishing for some R+Sel and R+M OMs, but generally 777 estimation was biased and(or) highly variable. We found substantial bias in estimated SRR 778 parameters in R and R+S OMs particularly with high variability in recruitment and ap-779 parent survival process errors, suggesting that practitioners should be cautious when fitted assessment models have these properties. 781 On the other hand, estimation of median natural mortality was reliable in many OM scenarios with contrast in fishing pressure, consistent with Hoenig et al. (2025). In some OMs, 783 when EMs estimated the SRR parameters and median natural mortality, bias for those 784 parameters was improved. Conversely, for some R+Sel and R+q OMs where there was 785 bias in natural mortality due to high observation error, estimating the SRR reduced the 786 bias in median natural mortality rate. However, estimating median natural mortality did 787 cause poor accuracy in SSB estimation in many OMs without contrast in fishing pressure 788 over time and with higher observation error. Thus, estimating median natural mortality 789 should be approached with caution in state-space assessment models, particularly given its 790 significant impact on determination of reference point and stock status (Li et al. 2024).

## Retrospective patterns

791

Incorrect EM process error assumptions did not produce strong retrospective patterns for SSB for any OMs regardless of whether median natural mortality or a SRR was estimated, but some weak retrospective patterns occur when observation error was high and there was
contrast in fishing pressure. However, retrospective patterns tended to be more variable
for recruitment and were sometimes large even when the EM was correct. Therefore, we
recommend emphasis on inspection of retrospective patterns primarily for SSB and F, but
further research on retrospective patterns in other assessment model parameters, management quantities such as biological reference points, and projections may be beneficial (Brooks
and Legault 2016).

The general lack of retrospective patterns with mis-specified process errors is perhaps to be 802 expected. Retrospective patterns are often induced in simulation studies by rapid changes 803 in a quantity such as index catchability, natural mortality, or perceived catch during years 804 toward the end of the time series (Legault 2009; Miller and Legault 2017; Huynh et al. 2022; 805 Breivik et al. 2023). In our simulations, the process errors changing over time may have trends in particular simulations, particularly when strong autocorrelation is imposed, but the random effects have no trend on average across simulations. Szuwalski et al. (2018) and 808 lietal24 Li et al. (2024) also found relatively small retrospective patterns when the source of mis-specification was temporal variation in demography attributes. Indeed, it is common for 810 the flexibility provided by temporal random effects to reduce retrospective patterns (Miller 811 et al. 2018; Stock et al. 2021; Stock and Miller 2021), though it does not necessarily 812 indicate a more accurate assessment model (Perretti et al. 2020; Li et al. 2024; Liljestrand 813 et al. 2024). Our results together with the existing literature seem to suggest that when 814 a strong retrospective pattern is observed in an assessment it is more likely to be due to a 815 mis-specification of a rapid shift in some model attribute rather than whether a particular 816 process is assumed to be randomly varying temporally. 817

## 818 Conclusions

Our simulation study examined the importance of several factors for reliable inferences from 819 state-space age-structured assessment models. Contrast in fishing pressure was consistently 820 an important factor across and lower observation error consistently improved all measures 821 of reliability we examined. AIC accurately distinguished models with process errors on 822 recruitment only (R) or on recruitment and apparent survival (R+S). Accuracy for other 823 process error types required a strong signal (high process variability) with low noise (low 824 observation uncertainty). Therefore, we expect practitioners will find R+S configurations 825 to provide satisfactory diagnostics across a range of life history and data quality scenarios. 826 AIC generally performed poorly for selecting the SRR, but performance was improved with 827 Accurate AIC selection of the Beverton-Holt stock-recruit relationship generally required 828 a combination of low recruitment variability and, contrast in fishing pressure. Some and large variation in SSB over time. However, when the Beverton-Holt was correctly assumed, 830 some bias in estimation in at least one of the SRR parameters existed in nearly all OM-EM 831 combinations parameters existed regardless of any of the other OM and EM configurations. 832 Because bias in terminal SSB and retrospective patterns were indifferent to whether or not 833 the SRR was estimated, and convergence was slightly better without the SRR, a sensible default would be to fit models without an assumed SRR. 835

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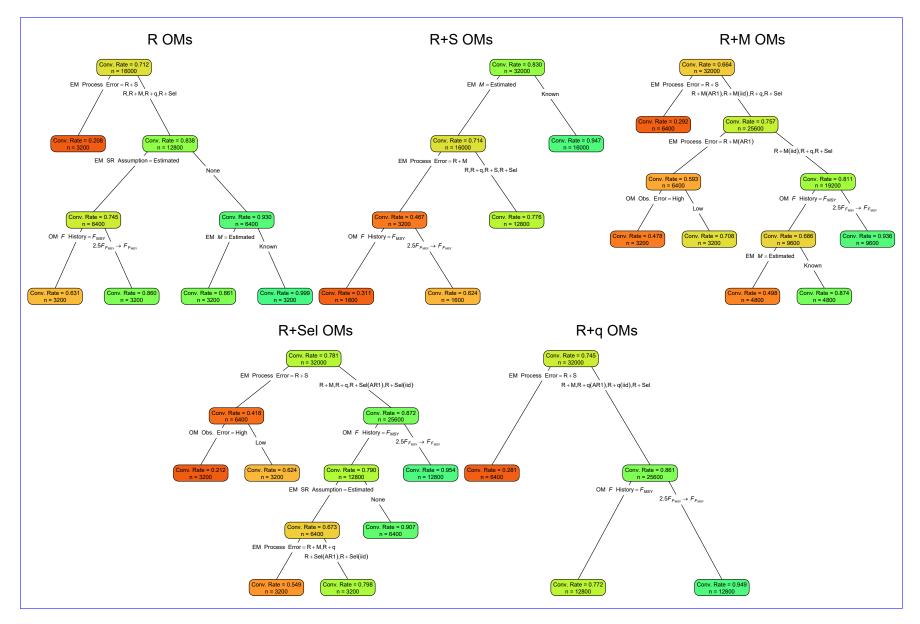


Fig. 1. Estimated probability of fits Classification trees indicating primary factors determining convergence as defined by providing hessian-based standard errors for EMs assuming alternative process error (colored points and lines)R, and median natural mortality (estimated or known) and Beverton-Holt stock-recruit relationships (estimated or not; along x-axis) when fitted to operating models that have Rand R+S(A), R+Sel (B)M, R+M (C), or Sel and R+q (D) process error structuresQMs. Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals. Lower or higher convergence rates are indicated by more red or green polygons, respectively

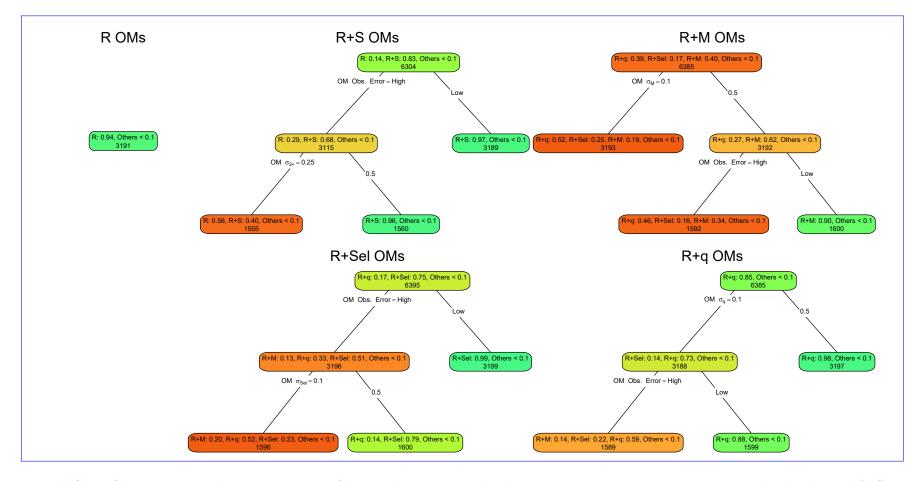


Fig. 2. Classification trees indicating primary factors determining which EM process error assumption provides the lowest AIC for R+S, R+M, R+Sel and R+q OMs. Each node shows the proportion of EM process error models with lowest AIC (top) and number of observations (bottom) for the corresponding subset. Lower or higher accuracy of the process error assumption are indicated by more red or green polygons, respectively.

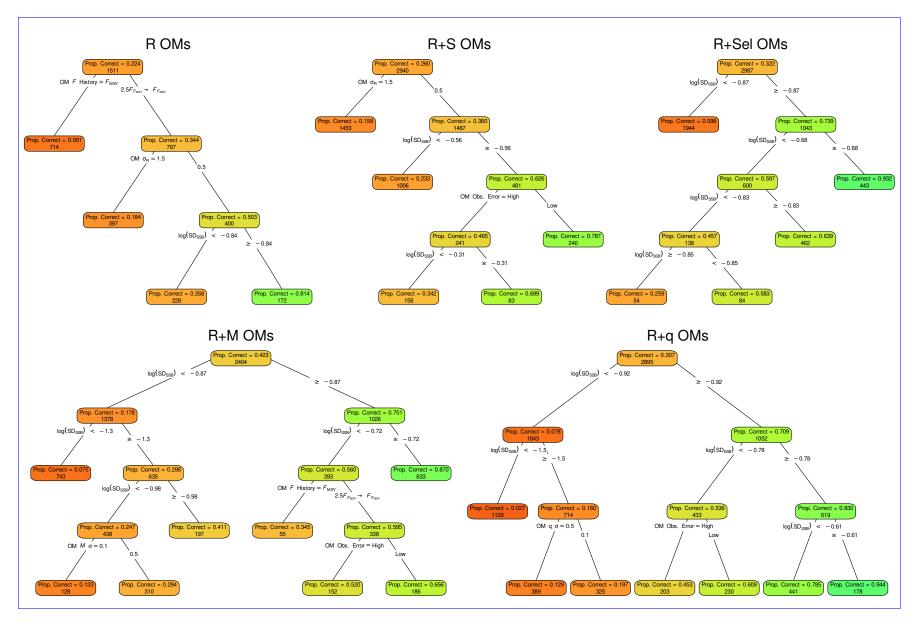


Fig. 3. Estimated probability of lowest AIC for EMs assuming alternative process error structures Classification trees indicating primary factors determining which EM stock recruitment assumption (colored bars) conditional on alternative assumptions for median natural mortality (estimated none or known) and Beverton-Holtstock-recruit relationships (estimated or not; along x-axis) when fitted to operating models that have provides the lowest AIC for Rand, R+S(A), R+Sel (B)M, R+M (C), or Sel and R+q OMs. Each node shows the proportion of EMs that assume the stock-recruit relationship with lowest AIC (Dtop) process error structures and number of observations (bottom) for the corresponding subset. Striped bars indicate results where

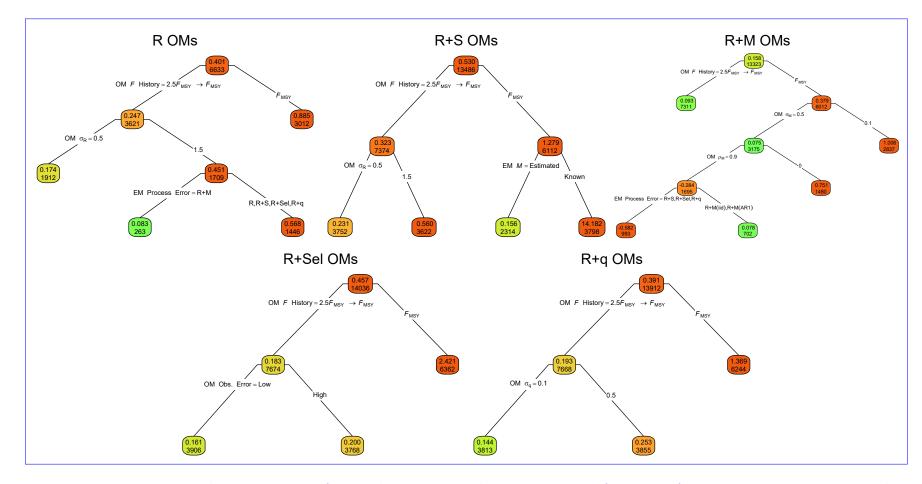


Fig. 4. Regression trees indicating primary factors determining reductions in sums of squares of errors in estimation measured by Eq. 3 for the Beverton-Holt stock-recruit parameter a for R+S, R+M, R+Sel and R+q OMs. Each node shows the median error (top) and number of observations (bottom) for the corresponding subset. Lower or higher median absolute errors of the process error assumption are indicated by more green or red polygons, respectively.

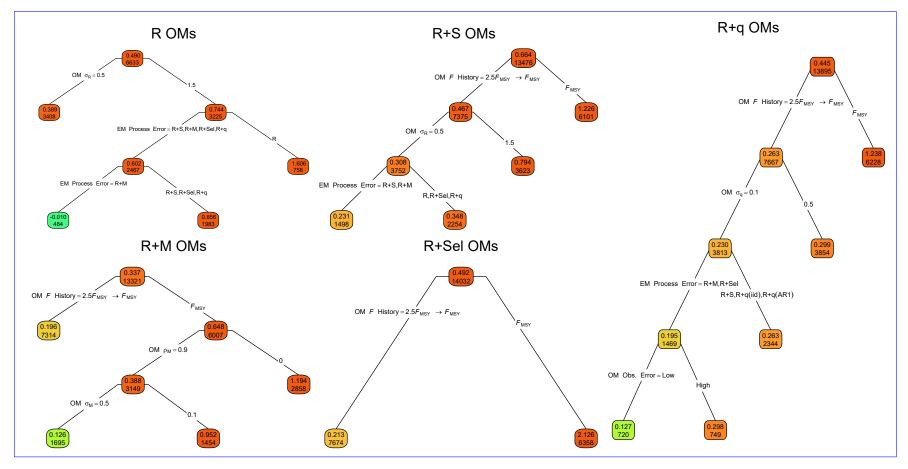


Fig. 5. Estimated probability Regression trees indicating primary factors determining reductions in sums of lowest AIC from logistic regression on the log-standard deviation squares of the true log(SSB) errors in each simulation estimation measured by Eq. 3 for estimating model with the Beverton-Holt stock-recruit relationships, rather than the otherwise equivalent EM without the stock-recruit relationship. Results are conditional on median M is known in the EM and alternative assumptions EMs having the correct process error structure: parameter b for Rand R+S(A), R+Sel (B)M, R+M (C), or Sel and R+q OMs. Each node shows the median error (Dtop), and median M is assumed known in the EM. Solid and dashed lines are for OMs with and without temporal contrast in fishing pressure, respectively, and polygons represent 95% confidence intervals. Range number of results indicates the range of log standard deviation of logobservations (SSBbottom) for simulations the corresponding subset. Lower or higher median absolute errors of the particular OMprocess error assumption are indicated by more green or red polygons, respectively.

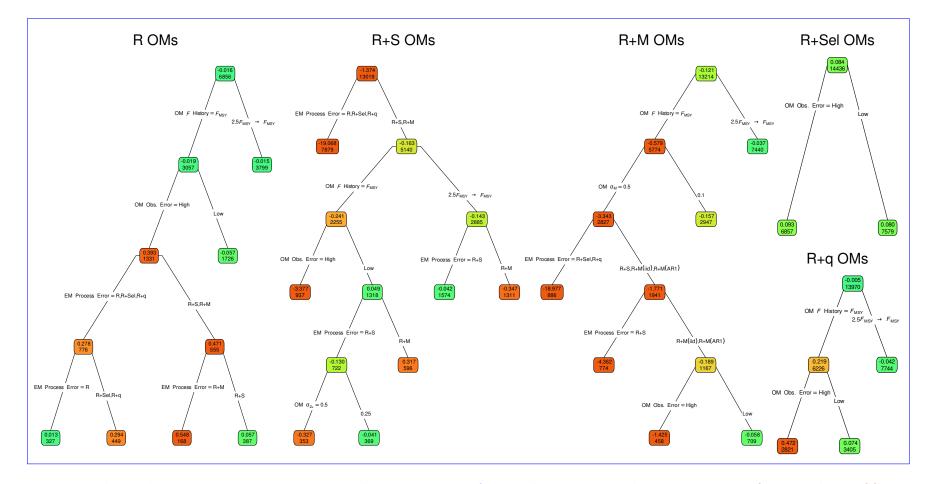


Fig. 6. Median relative error Regression trees indicating primary factors determining reductions in sums of terminal year SSB squares of errors in estimation measured by Eq. 3 for estimating models fitted to data sets simulated with alternative process error structures: the median natural mortality rate for Rand R+S(A), R+Sel (B)M, R+M (C), or Sel and R+q OMs. Each node shows the median error (Dtop) and number of observations (bottom) for the corresponding subset. Circled values indicate results where Lower or higher median absolute errors of the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals assumption are indicated by more green or red polygons, respectively.

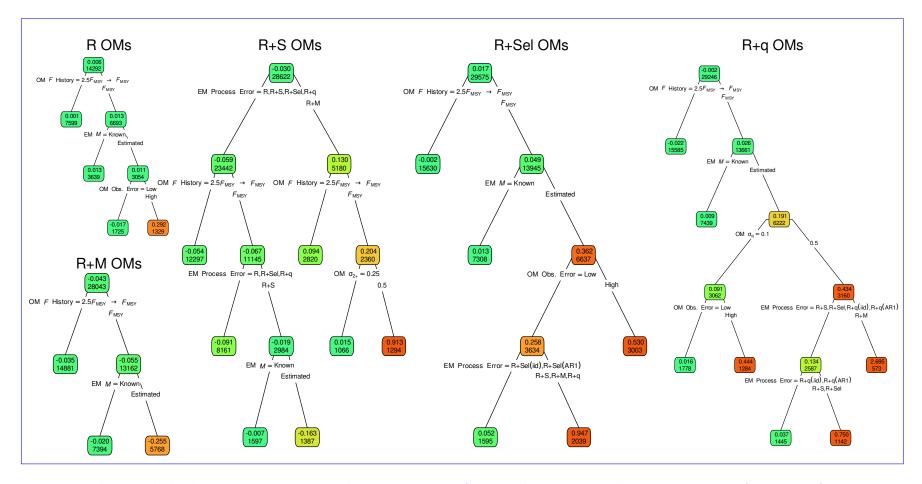


Fig. 7. Median Mohn's rho Regression trees indicating primary factors determining reductions in sums of squares of errors in estimation measured by Eq. 3 for SSB-terminal year spawning stock biomass for estimating models fitted to data sets simulated with alternative process error structures: Rand R+S(A), R+Sel (B)M, R+M (C), or Sel and R+q (D)OMs. Circled values indicate results where Each node shows the EM process median error structure matches that (top) and number of observations (bottom) for the operating model and vertical lines represent 95% confidence intervals corresponding subset. Median errors closer to or further from zero are indicated by more green or red polygons, respectively.

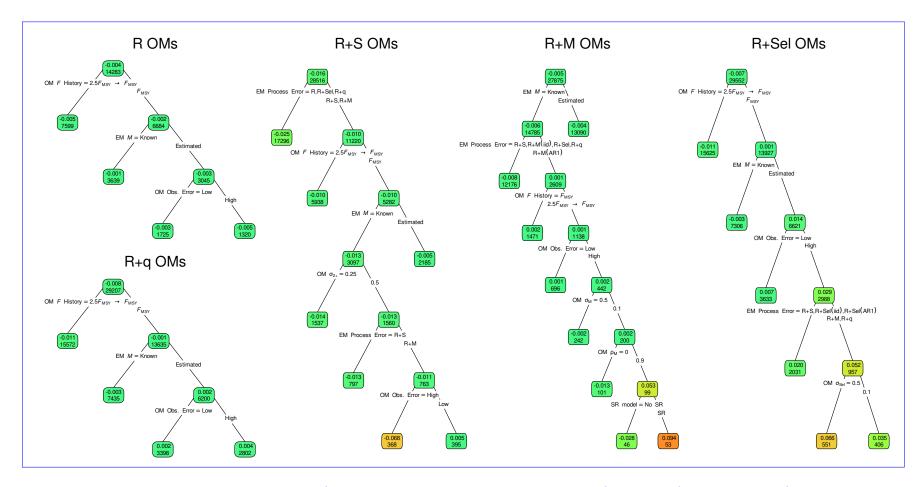


Fig. 8. Regression trees indicating primary factors determining reductions in sums of squares of errors in transformed Mohn's  $\rho$  (Eq. 3) for spawning stock biomass for R+S, R+M, R+Sel and R+q OMs. Each node shows the median Mohn's  $\rho$  (top) and number of observations (bottom) for the corresponding subset. Median Mohn's  $\rho$  closer to or further from zero are indicated by more green or red polygons, respectively.

Table 1. For each OM process error type (columns), percent reduction in deviance for logistic regression models fit to indicators of convergence (providing hessian-based standard errors) with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	<u>R</u>	<u>R</u> +S	<u>R+M</u>	<u>R+Sel</u>	R±q
EM Process Error	27.95	4.58	14.68	17.24	24.66
EM M Assumption	1.07	11.43	2.45	0.56	1.46
EM SR Assumption	2.88	3.30	1.24	2.47	1.59
OM Obs. Error	0.75	4.64	2.06	4.54	1.60
OM F History	2.32	3.37	1.63	3.30	2.59
$OM \sigma_R$	0.10	0.02	~	~	<del>_</del> ~
$OM \sigma_{2+}$	<del>_</del>	0.40	~	~	<del>_</del> ~
$ \underbrace{\mathrm{OM}}_{\sigma_{\mathcal{M}}} \sigma_{\mathcal{M}} $	<del>_</del>	<del>_</del>	0.22	~	_ ~
$OM \rho_R$	~	<del>~</del>	0.17	~	~
$\underbrace{\mathrm{OM}}_{\sigma_{Sel}}$	~	<del>~</del>	~	1.81	~
$\underbrace{\mathrm{OM}}_{\rho Sel}$	~	<del>~</del>	~	0.02	~
$\underbrace{\mathrm{OM}}_{\sigma_q} \sigma_q$	<del>_</del>	~	~	~	0.34
$\underbrace{\mathrm{OM}}_{} \rho_q$	<del>_</del>	~	~	~	≤0.01
<u>All factors</u>	39.54	31.46	24.85	34.83	36.31
+ All Two Way	45.03	39.89	35.20	42.81	43.70
+ All Three Way	47.02	44.57	37.88	45.51	46.87

Table 2. For each OM process error type (columns), percent reduction in deviance for multinomial logistic regression models fit to indicators of EM process error assumption with lowest AIC with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	$\underline{\underline{R}}$	<u>R+S</u>	<u>R+M</u>	<u>R+Sel</u>	R±g
EM M Assumption	5.52	1.05	0.52	0.61	1.32
EM SR Assumption	5.60	0.75	1.13	0.93	1.95
OM Obs. Error	2.96	22.46	3.42	25.67	5.03
OM F History	5.77	0.62	0.94	0.91	2.05
$OM \sigma_R$	0.10	0.66	~	~	<del>~</del>
$OM \sigma_{2+}$	<del>~</del>	16.86	~	~	<del>~</del>
$\underbrace{\mathrm{OM}}_{\sigma_{\mathcal{M}}} \sigma_{\mathcal{M}}$	<del>~</del>	<del>_</del>	9.06	~	~~
$\underbrace{\mathrm{OM}}_{\mathcal{P}_{\mathcal{R}}} \mathcal{P}_{\mathcal{R}}$	<del>~</del>	<del>_</del>	0.38	~	~~
$\underbrace{\mathrm{OM}}_{\sigma_{Sel}}$	<del>~</del>	<del>~</del>	~	7.59	<del>~</del>
$\underbrace{\mathrm{OM}}_{\rho_{Sel}}$	<del>~</del>	<del>_</del>	~	0.60	<del>~</del>
$\underbrace{\mathrm{OM}}_{\sigma_q} \sigma_q$	<del>~</del>	<del>~</del>	~	~	<u>13.50</u>
$\underbrace{\mathrm{OM}}_{} \rho_q$	<del>~</del>	<del>_</del>	~	~	0.75
<u>All factors</u>	20.98	46.12	16.58	40.83	25.99
+ All Two Way	22.02	48.94	21.63	44.08	30.17
+ All Three Way	22.05	49.98	22.36	44.54	31.38

Table 3. For each OM process error type (columns), percent reduction in deviance for logistic regression models fit to indicators of EM stock-recruit assumption (none or Beverton-Holt) with lowest AIC with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	<u>R</u> +S	<u>R+M</u>	<u>R+Sel</u>	<u>R</u> +g
EM M Assumption	0.04	0.21	0.18	0.02	0.01
OM Obs. Error	<u>&lt;0.01</u>	0.65	0.14	0.04	0.02
OM F History	$\underbrace{9.17}$	3.79	13.08	26.56	24.60
$OM \sigma_R$	3.54	4.74	~	~	~~
$OM \sigma_{2+}$	~	0.14	~	~	~~
$OM \sigma_M$	~	<del>~</del>	1.14	~	~~
$OM \rho_R$	~	<del>~</del>	0.05	~	~~
$\underbrace{\mathrm{OM}}_{\sigma_{Sel}} \sigma_{Sel}$	~	<del>~</del>	~	$\underbrace{0.02}_{00000000000000000000000000000000000$	~~
$ \underbrace{\text{OM}}_{\rho_{Sel}} $	~	<del>~</del>	~	0.17	~~
$\underbrace{\mathrm{OM}}_{\sigma_q} \sigma_q$	~	<del>~</del>	~	~	0.36
$\underbrace{\mathrm{OM}}_{} \rho_q$	~	<del>~</del>	~	~	0.02
$\log(\mathrm{SD}_{\mathrm{SSB}})$	4.11	1.59	33.39	41.36	39.23
All factors	31.52	18.99	34.23	43.77	42.31
+ All Two Way	34.79	22.24	35.99	45.84	44.04
+ All Three Way	35.41	23.09	37.57	46.39	44.63

Table 4. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the Beverton-Holt stock recruit relation ship parameters with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor		Beverton-Holt a						Bev	erton-I	$\frac{\mathbf{Holt}}{\mathbf{b}}$	
	$ \underbrace{\mathbb{R}} $	$\mathbb{R}+\mathbb{S}$	$\underbrace{\mathbb{R}+\mathbb{M}}$	<u>R+Sel</u>	$\underline{R}+\underline{q}$	$\widetilde{\mathbb{R}}$	ا پا	$\underbrace{R+S}$	<u>R+M</u>	<u>R+Sel</u>	$\mathbb{R}+\mathbf{q}$
EM M Assumption	0.02	1.05	0.02	0.11	0.02	$\widetilde{0}$ .	<u>05</u>	1.06	0.03	0.01	0.40
EM Process Error	2.74	0.18	0.20	1.25	1.90	2.	<u>29</u>	1.21	0.12	1.40	3.06
OM Obs. Error	0.16	<u>&lt;0.01</u>	0.01	0.04	≤0.01	$\leq 0$ .	<u>01</u>	0.01	0.05	0.01	0.01
OM F History	3.15	3.34	5.60	11.37	10.00	1.	<u>16</u>	1.17	2.01	7.97	3.87
$\underbrace{\mathrm{OM}}_{\sigma_R} \sigma_R$	2.31	0.74	~	~	~	<u>1</u> .	<u>67</u>	0.52	~	<del>_</del> ~	~
$\underbrace{\mathrm{OM}}_{\sigma_{2+}}$	<del>_</del>	0.29	~	~	~		~	0.01	~	<del>_</del> ~	~
$OM\sigma_{M}$	<del>_</del>	<del>-</del> ~	0.30	~	~		~	<del>-</del> ~	0.13	~	<del>_</del>
$OM \rho_R$	<del>_</del>	<del>-</del> ~	0.51	~	~		~	<del>-</del> ~	0.22	~	<del>_</del>
$\widecheck{\mathrm{OM}}$ $\sigma_{Sel}$	<del>_</del>	<del>-</del> ~	~	0.13	~		~	<del>-</del> ~	<del>_</del> ~	0.05	<del>_</del>
$\widecheck{\mathrm{OM}}$ $ ho_{Sel}$	<del>_</del>	<del>-</del> ~	~	0.07	~		~	<del>_</del> ~	~	0.04	~
$\underbrace{\mathrm{OM}}_{\sigma_q} \sigma_q$	<del>~</del>	$_{\sim}^{-}$	~	~	0.04		~	$\overline{\sim}$	~	~	0.10
$ \underbrace{\mathrm{OM}}_{} \rho_q $	<del>~</del>	$_{\sim}^{-}$	~	~	≤0.01		~	$\overline{\sim}$	~	~	≤0.01
All factors	8.07	5.15	6.73	12.64	11.79	4.	91	3.75	2.55	$\underbrace{9.12}_{}$	7.22
+ All Two Way	9.96	7.37	9.76	13.59	13.65	<u>7</u> .	<u>55</u>	7.15	4.32	10.08	12.16
+ All Three Way	11.22	8.15	11.13	14.48	14.87	9.	<u>78</u>	9.02	5.26	11.08	14.73

Table 5. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the median natural mortality rate parameter with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	<u>R</u>	<u>R</u> +S	<u>R+M</u>	<u>R+Sel</u>	R±g
EM SR assumption	0.21	0.38	0.11	0.26	0.43
EM Process Error	1.98	20.36	3.16	0.94	1.31
OM Obs. Error	4.74	0.79	0.40	2.23	1.88
OM F History	5.07	<u>15.11</u>	10.65	0.24	2.38
$OM \sigma_R$	<u>&lt;0.01</u>	0.01	~	~	<del>~</del>
$OM\sigma_{2+}$	~	5.04	~	~	<del>~</del>
$\underbrace{\mathrm{OM}}_{\sigma_{M}} \sigma_{M}$	~	~	5.32	~	~
$OM \rho_R$	~	~	0.85	~	~
$\underbrace{\mathrm{OM}\;\sigma_{Sel}}$	~	<del>~</del>	~	1.30	<del>~</del>
$OM$ $ ho_{Sel}$	~	<del>~</del>	~	0.37	<del>~</del>
$\underbrace{\mathrm{OM}}_{\sigma_q} \sigma_q$	~	~	~	~	0.46
$\underbrace{\mathrm{OM}}_{} \rho_q$	~	~	~	~	0.06
All factors	12.64	40.10	21.29	5.54	6.52
+ All Two Way	21.17	48.12	36.19	9.87	11.71
+ All Three Way	23.03	50.38	42.82	11.58	14.64

Table 6. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the terminal year spawning biomass with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	<u>R</u>	<u>R+S</u>	<u>R+M</u>	<u>R+Sel</u>	R±q
EM M Assumption	2.28	1.15	1.04	2.92	3.26
EM SR assumption	0.10	0.06	0.08	0.06	0.08
EM Process Error	0.43	4.28	0.40	0.11	$\underbrace{1.05}_{}$
OM Obs. Error	1.63	0.07	0.78	0.32	<u>&lt;0.01</u>
OM F History	2.62	3.15	1.28	3.22	4.72
$OM \sigma_R$	0.03	0.01	~	~	~
$OM \sigma_{2\pm}$	<del>~</del>	0.93	~	~	~
$\underbrace{\mathrm{OM}}_{\sigma_{M}} \sigma_{M}$	<del>_</del>	<del>_</del>	0.18		<del>_</del> ~
$OM \rho_R$	<del>_</del>	<del>~</del>	0.01	~	<del>_</del> ~
$OM \sigma_{Sel}$	<del>_</del>	<del>~</del>	<del>_</del> ~	0.16	<del>_</del> ~
OM PSel	<del>~</del>	<del>_</del>	~	0.04	~
$\underbrace{\mathrm{OM}}_{} \underbrace{\sigma_q}$	<del>_</del>	<del>_</del>	<del>_</del> ~		1.02
$\underbrace{\mathrm{OM}}_{} \rho_q$	<del>_</del>	<del>_</del>	<del>_</del> ~		0.06
All factors	7.59	9.86	3.93	7.04	10.64
+ All Two Way	<u>17.99</u>	25.56	10.06	13.44	22.43
+ All Three Way	23.39	36.74	13.76	16.55	31.11

Table 7. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to transformed Mohn's  $\rho$  values for each simulation (Eq. 3) for spawning biomass with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	<u>R</u>	$\mathbb{R}+\mathbb{S}$	<u>R+M</u>	<u>R+Sel</u>	R±q
EM M Assumption	0.79	0.18	0.15	0.95	1.24
EM SR assumption	≤0.01	0.01	<u>&lt;0.01</u>	≤0.01	<u>&lt;0.01</u>
EM Process Error	<0.01	0.22	0.14	0.08	0.04
OM Obs. Error	0.12	0.03	0.05	0.18	0.21
OM F History	0.84	0.14	0.07	1.08	1.56
$OM \sigma_R$	0.01	0.01	<del>-</del> ~	~	~
$OM \sigma_{2\pm}$	<del>-</del> ~	0.02	<del>-</del> ~	_~	~
$\underbrace{\mathrm{OM}}_{\sigma_{M}} \sigma_{M}$	_ ~	~	0.01	_ ~	<del>-</del> ~
$OM \rho_R$	~	$\overline{\sim}$	≤0.01	~	$\stackrel{-}{\sim}$
$OM \sigma_{Sel}$	~	$\overline{\sim}$	$\stackrel{-}{\sim}$	0.01	$\stackrel{-}{\sim}$
OM Psel	<del>-</del> ~	<del>~</del>	<del>-</del> ~	0.02	~
$\underbrace{\mathrm{OM}}_{\sigma_q} \sigma_q$	_ ~	~	<del>_</del> ~	~	0.01
$\underbrace{\mathrm{OM}}_{} \varrho_q$	_ ~	~	<del>-</del> ~	_ ~	0.01
All factors	1.89	0.63	0.43	2.43	3.29
+ All Two Way	3.63	1.10	0.91	4.75	<u>6.22</u>
+ All Three Way	4.27	1.65	1.50	5.73	7.53

## Supplementary Materials

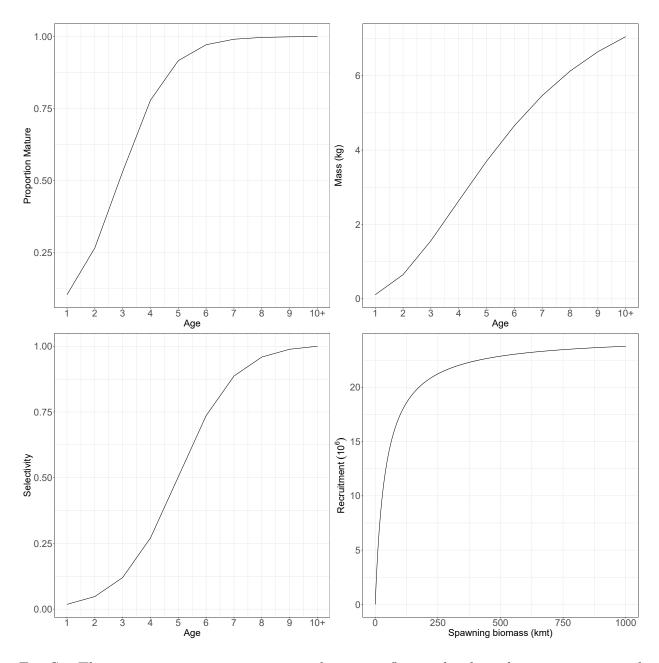


Fig. S1. The proportion mature at age, weight at age, fleet and index selectivity at age, and Beverton-Holt stock-recruit relationship assumed for the population in all operating models. For operating models with random effects on fleet selectivity, this represents the selectivity at the mean of the random effects.

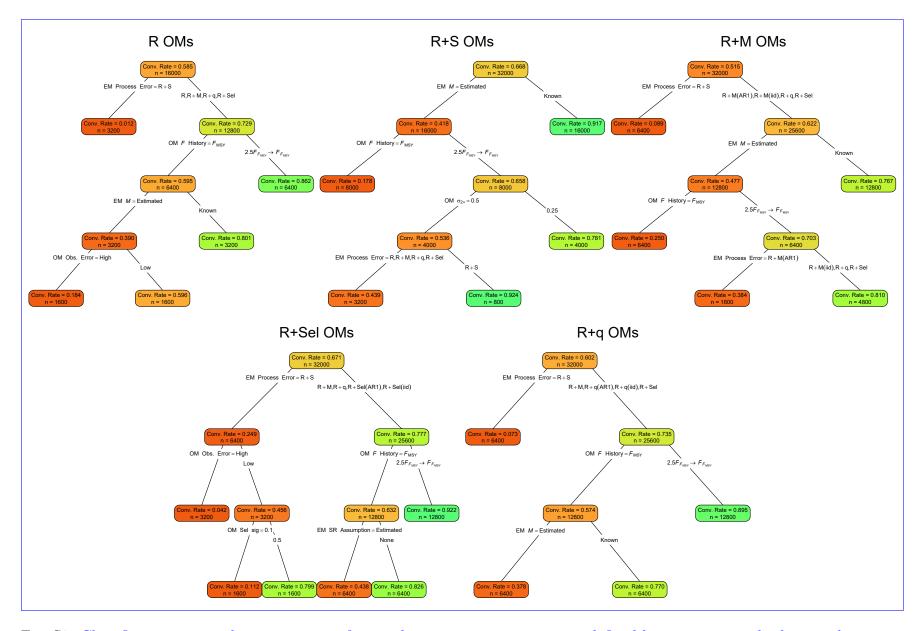


Fig. S2. Classification trees indicating primary factors determining convergence as defined by a maximum absolute gradient  $< 10^{-6}$  for R, R+S, R+M, R+Sel and R+q OMs. Lower or higher convergence rates are indicated by more red or green polygons, respectively

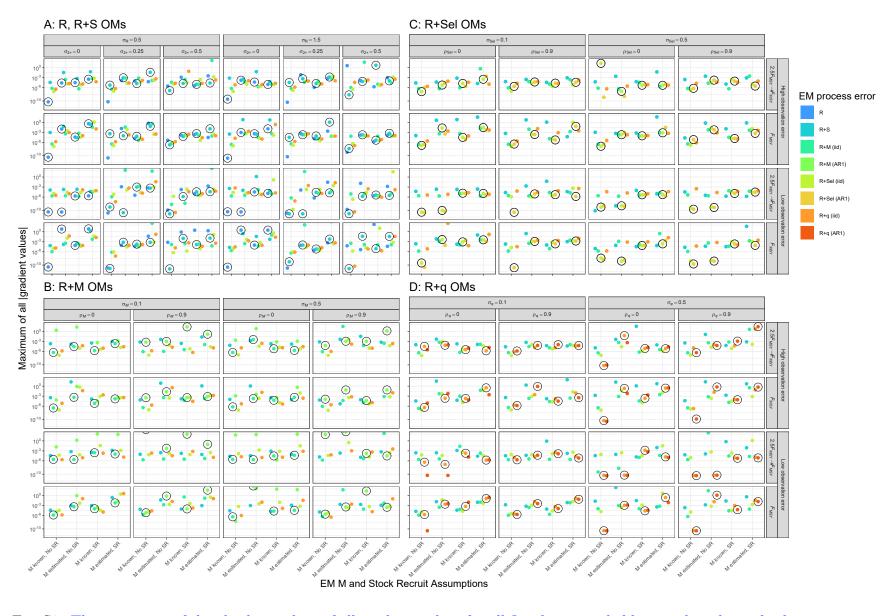


Fig. S3. The maximum of the absolute values of all gradient values for all fits that provided hessian-based standard errors across all simuated data sets of a given OM configuration (A: R and R+S, B: R+M, C: R+Sel, or D: R+q). Results are conditional on EM fits with alternative process error type (colored points and lines), median natural mortality (estimated or known) and recruitment assumptions (Beverton-Holt stock-recruit relationship or not). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

Table S1. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the terminal year fully-selected fishing mortality with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	<u>R</u>	<u>R</u> +S	<u>R+M</u>	<u>R+Sel</u>	R±q
EM M Assumption	2.26	1.33	1.26	2.93	3.26
EM SR assumption	0.11	0.07	0.08	0.07	0.09
EM Process Error	0.46	4.18	0.38	0.13	1.02
OM Obs. Error	1.61	0.06	0.86	0.41	<u>&lt;0.01</u>
OM F History	2.49	3.23	1.42	3.22	4.55
$OM \sigma_R$	0.02	0.02	<del>_</del> ~	~	~
$OM \sigma_{2\pm}$	<del>_</del>	0.87	<del>_</del> ~	~	~
$\underbrace{\mathrm{OM}}_{\sigma_{M}} \underbrace{\sigma_{M}}$	<del>_</del>	<del>_</del>	0.16	~	<del>-</del> ~
$OM \rho_R$	<del>~</del>	<del>~</del>	0.01	~	$\stackrel{-}{\sim}$
$OM \sigma_{Sel}$	<del>~</del>	<del>~</del>	~	0.24	$\stackrel{-}{\sim}$
$OM$ $ ho_{Sel}$	<del>_</del>	<del>~</del>	<del>_</del> ~	0.05	~
$\underbrace{\mathrm{OM}}_{} \underbrace{\sigma_q}$	<del>_</del>	<del>_</del>	_ ~	~	1.03
$\underbrace{\mathrm{OM}}_{} \varrho_q$	<del>_</del>	<del>_</del>	_ ~	~	0.05
All factors	7.42	9.96	4.37	7.26	10.43
+ All Two Way	<u>17.63</u>	25.76	10.94	13.88	22.07
+ All Three Way	22.97	<u>37.03</u>	14.74	17.32	30.74

Table S2. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the terminal year recruitment with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	<u>R</u>	<u>R</u> +S	<u>R+M</u>	<u>R+Sel</u>	R±q
EM M Assumption	1.96	0.40	0.69	3.52	3.03
EM SR assumption	0.06	0.02	0.05	0.02	0.05
EM Process Error	0.39	4.74	0.41	0.12	1.16
OM Obs. Error	1.47	0.08	0.64	0.18	<0.01
OM F History	2.54	2.66	1.11	4.18	5.06
$OM \sigma_R$	0.03	0.01	~	~	~
$OM \sigma_{2\pm}$	<del>~</del>	1.05	~	~	~
$\underbrace{\mathrm{OM}}_{\sigma_{M}} \sigma_{M}$	<del>_</del>	<del>_</del>	0.36		<del>_</del> ~
$OM \rho_R$	<del>_</del>	<del>~</del>	0.02	~	<del>-</del>
$OM \sigma_{Sel}$	<del>_</del>	<del>~</del>	<del>_</del> ~	0.23	<del>-</del>
$OM$ $ ho_{Sel}$	<del>_</del>	<del>~</del>	<del>_</del> ~	0.06	<del>-</del>
$\underbrace{\mathrm{OM}}_{\sigma_q} \sigma_q$	<del>_</del>	<del>_</del>	_ ~	~	1.09
$\underbrace{\mathrm{OM}}_{} \varrho_q$	<del>_</del>	<del>_</del>	_ ~	~	0.06
All factors	<u>6.90</u>	9.01	3.43	8.58	10.90
+ All Two Way	16.48	24.64	9.73	15.76	22.75
+ All Three Way	21.46	35.60	13.56	19.07	31.15

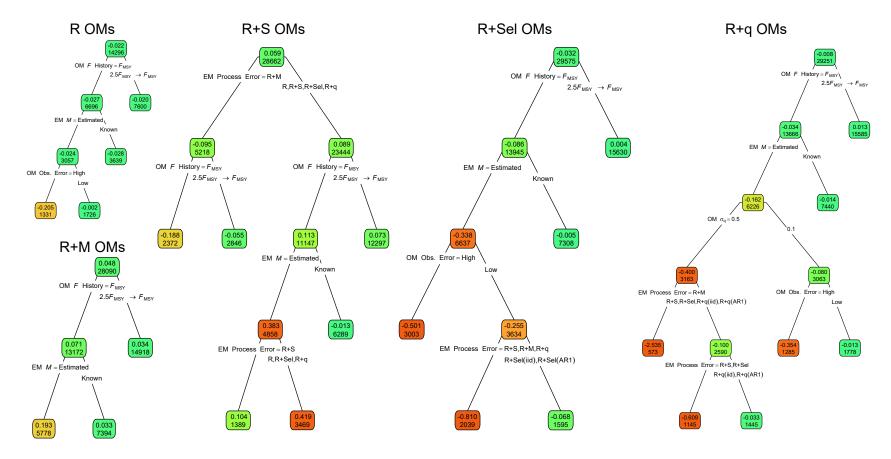


Fig. S4. Regression trees indicating primary factors determining reductions in sums of squares of errors in estimation measured by Eq. 3 for terminal year fully-selected fishing mortality for R+S, R+M, R+Sel and R+q OMs. Each node shows the median error (top) and number of observations (bottom) for the corresponding subset. Median errors closer to or further from zero are indicated by more green or red polygons, respectively.

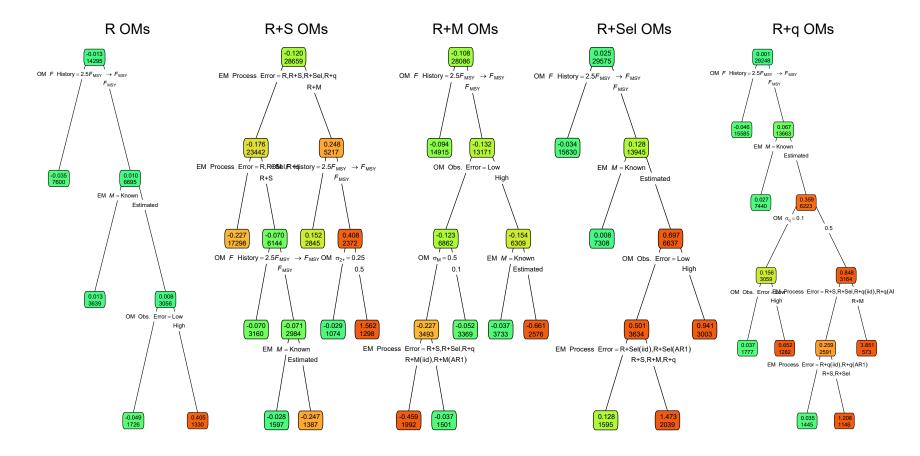


Fig. S5. Regression trees indicating primary factors determining reductions in sums of squares of errors in estimation measured by Eq. 3 for terminal year recruitment for R+S, R+M, R+Sel and R+q OMs. Each node shows the median error (top) and number of observations (bottom) for the corresponding subset. Median errors closer to or further from zero are indicated by more green or red polygons, respectively.

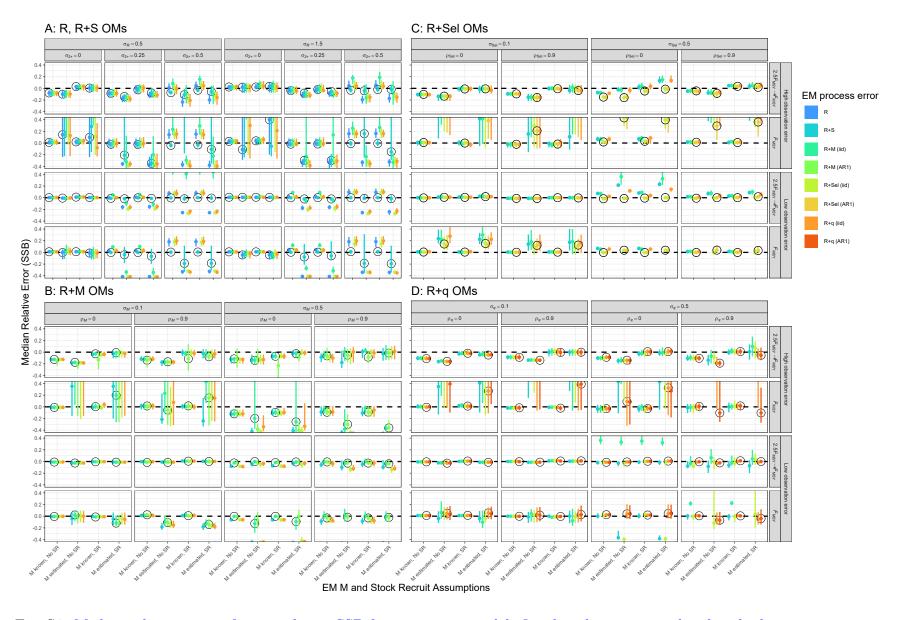


Fig. S6. Median relative error of terminal year SSB for estimating models fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

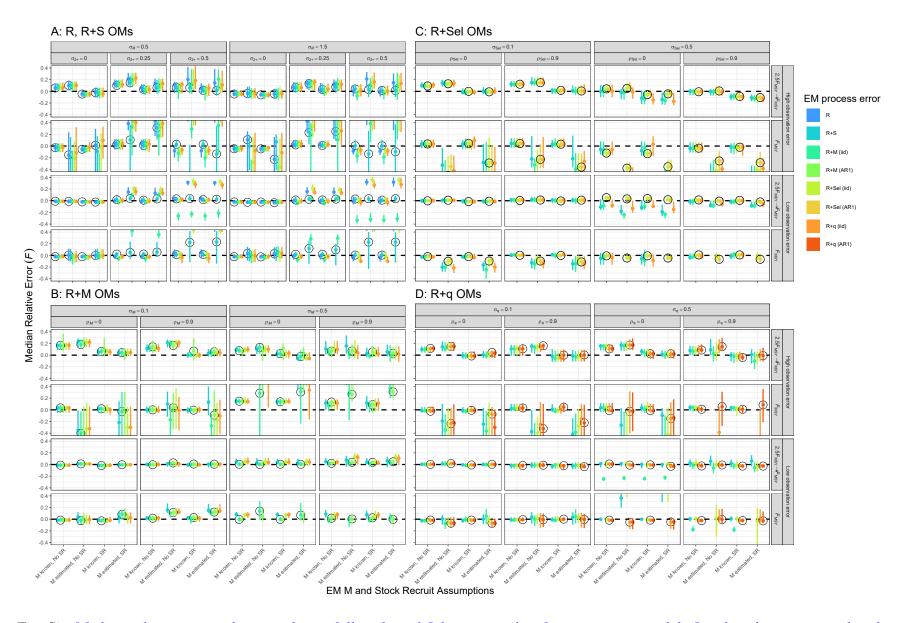


Fig. S7. Median relative error of terminal year fully-selected fishing mortality for estimating models fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

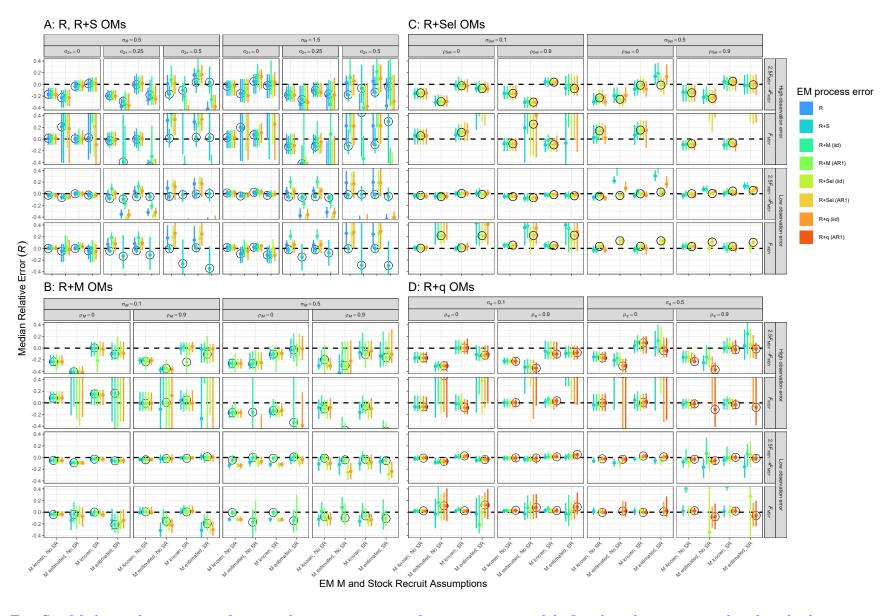


Fig. S8. Median relative error of terminal year recruitment for estimating models fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

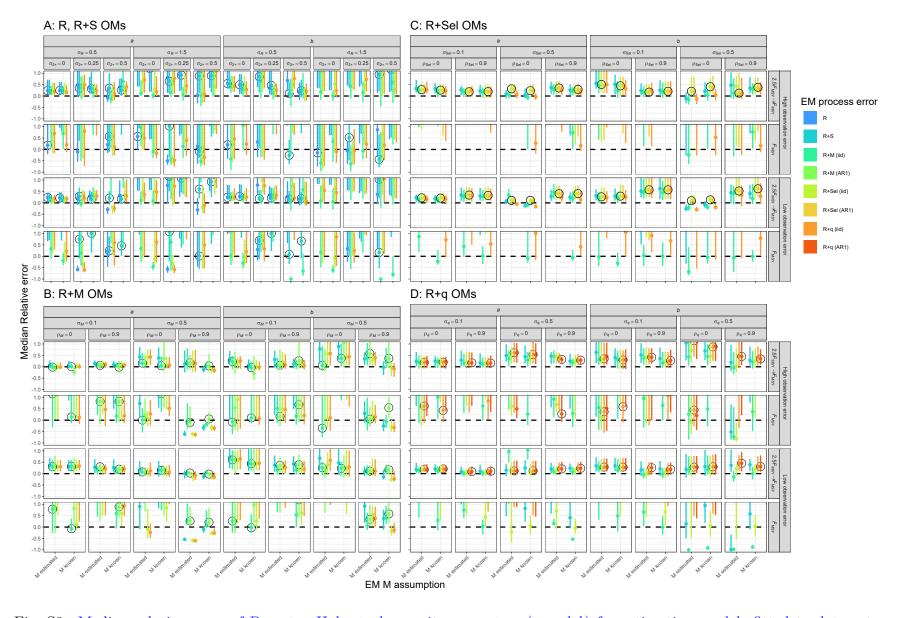


Fig. S9. Median relative error of Beverton-Holt stock-recruit parameters (a and b) for estimating models fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

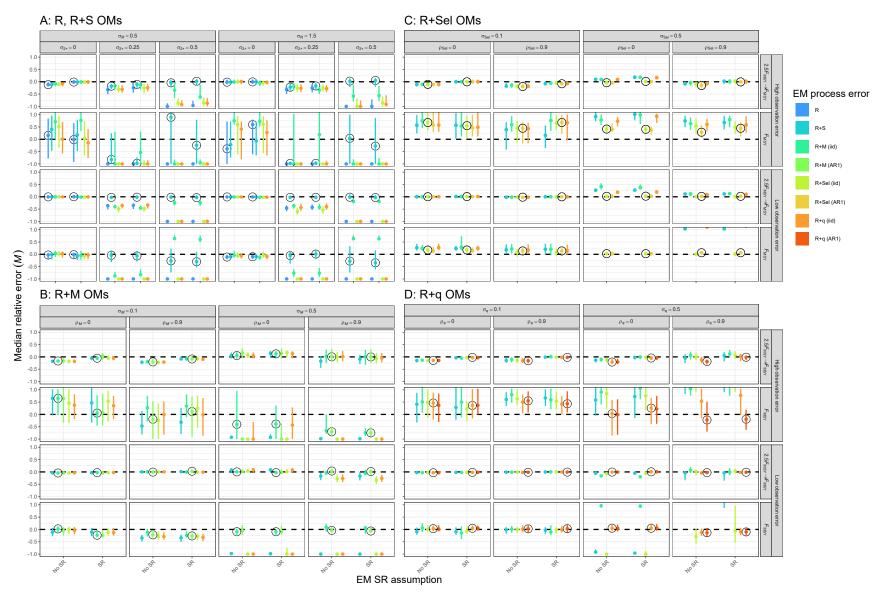


Fig. S10. Median relative error of median natural mortality for estimating models fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

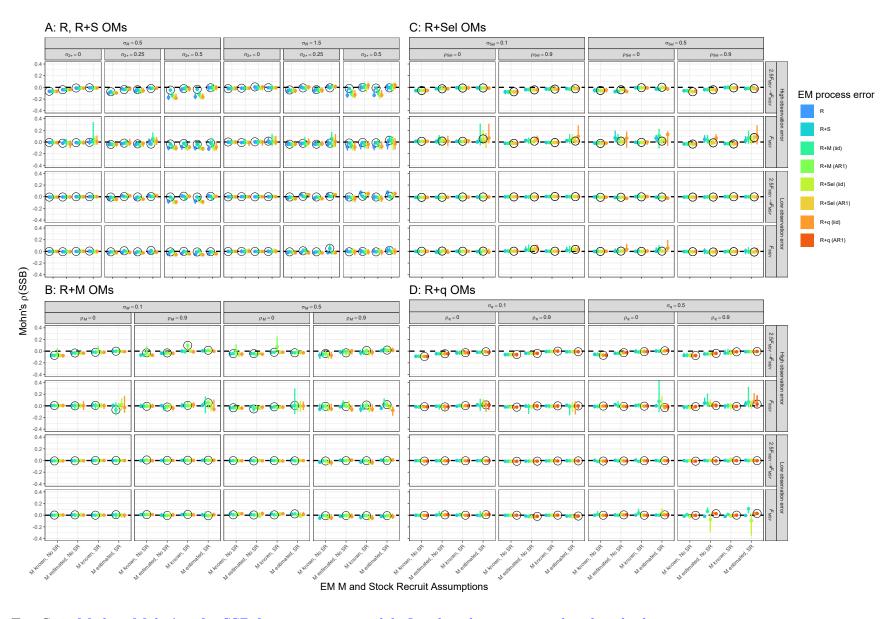


Fig. S11. Median Mohn's  $\rho$  for SSB for estimating models fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

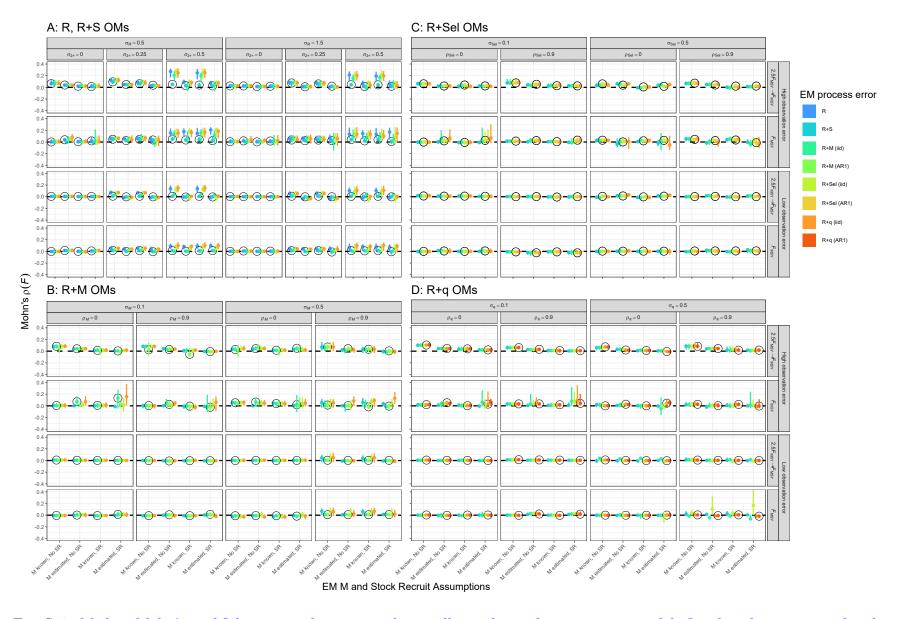


Fig. S12. Median Mohn's  $\rho$  of fishing mortality averaged over all age classes for estimating models fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

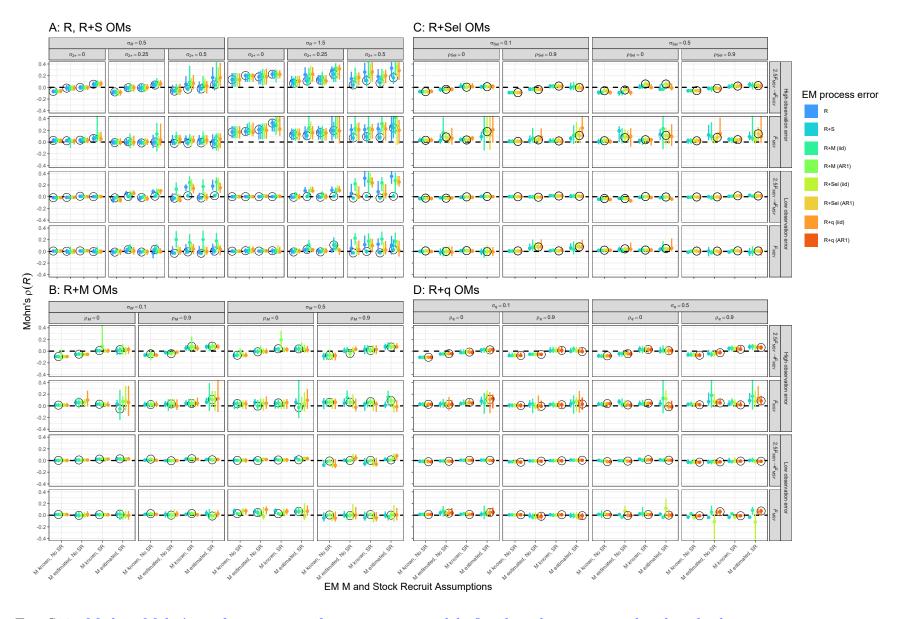


Fig. S13. Median Mohn's  $\rho$  of recruitment for estimating models fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

Model	$\sigma_R$	$\sigma_{2+}$	Fishing History	Observation Uncertainty
$NAA_1$	0.5		$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_2$	1.5		$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_3$	0.5	0.25	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_4$	1.5	0.25	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_5$	0.5	0.50	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_6$	1.5	0.50	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_7$	0.5		$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_8$	1.5		$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_9$	0.5	0.25	$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_{10}$	1.5	0.25	$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_{11}$	0.5	0.50	$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_{12}$	1.5	0.50	$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$NAA_{13}$	0.5		$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{14}$	1.5		$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{15}$	0.5	0.25	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{16}$	1.5	0.25	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{17}$	0.5	0.50	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{18}$	1.5	0.50	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{19}$	0.5		$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{20}$	1.5		$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{21}$	0.5	0.25	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{22}$	1.5	0.25	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{23}$	0.5	0.50	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$NAA_{24}$	1.5	0.50	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$

Table S4. Distinguishing characteristics of the operating models with random effects on recruitment and natural mortality (R+M). Standard deviations (SD) are for log-normal distributed indices and logistic normal distributed age composition observations (fleet and indices). Fishing mortality changes after year 20 (of 40) for fishing histories where fishing mortality is not constant. For AR1 process errors,  $\sigma$  is defined for the marginal distribution of the processes.

Model	$\sigma_R$	$\sigma_{M}$	$ ho_M$	Fishing History	Observation Uncertainty
$M_1$	0.5	0.1	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.1$ , Age composition SD = $0.3$
$M_2$	0.5	0.5	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.1$ , Age composition SD = $0.3$
$M_3$	0.5	0.1	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.1$ , Age composition SD = $0.3$
$M_4$	0.5	0.5	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.1$ , Age composition SD = $0.3$
$M_5$	0.5	0.1	0.0	$F_{ m MSY}$	Index SD = $0.1$ , Age composition SD = $0.3$
$M_6$	0.5	0.5	0.0	$F_{ m MSY}$	Index SD = $0.1$ , Age composition SD = $0.3$
$M_7$	0.5	0.1	0.9	$F_{ m MSY}$	Index SD = $0.1$ , Age composition SD = $0.3$
$M_8$	0.5	0.5	0.9	$F_{ m MSY}$	Index SD = $0.1$ , Age composition SD = $0.3$
$M_9$	0.5	0.1	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$M_{10}$	0.5	0.5	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$M_{11}$	0.5	0.1	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$M_{12}$	0.5	0.5	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$M_{13}$	0.5	0.1	0.0	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$M_{14}$	0.5	0.5	0.0	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$M_{15}$	0.5	0.1	0.9	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$M_{16}$	0.5	0.5	0.9	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$

Model	$\sigma_R$	$\sigma_{ m Sel}$	$ ho_{ m Sel}$	Fishing History	Observation Uncertainty
$\mathrm{Sel}_1$	0.5	0.1	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.1$ , Age composition SD = $0.3$
$\mathrm{Sel}_2$	0.5	0.5	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.1$ , Age composition SD = $0.3$
$\mathrm{Sel}_3$	0.5	0.1	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.1$ , Age composition SD = $0.3$
$\mathrm{Sel}_4$	0.5	0.5	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.1$ , Age composition SD = $0.3$
$\mathrm{Sel}_5$	0.5	0.1	0.0	$F_{ m MSY}$	Index SD = $0.1$ , Age composition SD = $0.3$
$\mathrm{Sel}_6$	0.5	0.5	0.0	$F_{ m MSY}$	Index SD = $0.1$ , Age composition SD = $0.3$
$\mathrm{Sel}_7$	0.5	0.1	0.9	$F_{ m MSY}$	Index SD = $0.1$ , Age composition SD = $0.3$
$\mathrm{Sel}_8$	0.5	0.5	0.9	$F_{ m MSY}$	Index SD = $0.1$ , Age composition SD = $0.3$
$\mathrm{Sel}_9$	0.5	0.1	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$\mathrm{Sel}_{10}$	0.5	0.5	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$\mathrm{Sel}_{11}$	0.5	0.1	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$\mathrm{Sel}_{12}$	0.5	0.5	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$\mathrm{Sel}_{13}$	0.5	0.1	0.0	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$\mathrm{Sel}_{14}$	0.5	0.5	0.0	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$\mathrm{Sel}_{15}$	0.5	0.1	0.9	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$Sel_{16}$	0.5	0.5	0.9	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$

Table S6. Distinguishing characteristics of the operating models with random effects on recruitment and catchability (R+q). Standard deviations (SD) are for log-normal distributed indices and logistic normal distributed age composition observations (fleet and indices). Fishing mortality changes after year 20 (of 40) for fishing histories where fishing mortality is not constant. For AR1 process errors,  $\sigma$  is defined for the marginal distribution of the processes.

Model	$\sigma_R$	$\sigma_q$	$ ho_q$	Fishing History	Observation Uncertainty
$q_1$	0.5	0.1	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.1$ , Age composition SD = $0.3$
$q_2$	0.5	0.5	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = 0.1, Age composition SD = $0.3$
$q_3$	0.5	0.1	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = 0.1, Age composition SD = $0.3$
$q_4$	0.5	0.5	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = 0.1, Age composition SD = $0.3$
$q_5$	0.5	0.1	0.0	$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$q_6$	0.5	0.5	0.0	$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$q_7$	0.5	0.1	0.9	$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$q_8$	0.5	0.5	0.9	$F_{ m MSY}$	Index SD = 0.1, Age composition SD = $0.3$
$q_9$	0.5	0.1	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$q_{10}$	0.5	0.5	0.0	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$q_{11}$	0.5	0.1	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$q_{12}$	0.5	0.5	0.9	$2.5F_{\mathrm{MSY}} \to F_{\mathrm{MSY}}$	Index SD = $0.4$ , Age composition SD = $1.5$
$q_{13}$	0.5	0.1	0.0	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$q_{14}$	0.5	0.5	0.0	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$q_{15}$	0.5	0.1	0.9	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$
$q_{16}$	0.5	0.5	0.9	$F_{ m MSY}$	Index SD = $0.4$ , Age composition SD = $1.5$

## 1056 Convergence full results

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Median relative error of terminal year fully-selected fishing mortality for estimating models 1057 fitted to data sets simulated with alternative process error structures: R and R+S (A), 1058 R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error 1059 structure matches that of the operating model and vertical lines represent 95\% confidence 1060 intervals. For many R and R+S OMs, convergence rate declined when either the median 1061 natural mortality rate or the Beverton-Holt SRR was estimated even when the process error 1062 assumptions of the EMs and OMs matched (Figure S14, A). When there was high observation 1063 error and constant fishing pressure ( $F = F_{MSY}$  for all 40 years), convergence was poor for 1064 of all EM process error configurations other than R EMs when fitted to R OMs ( $\sigma_{2+}=0$ ) 1065 regardless of whether median natural mortality and SRRs were estimated. Convergence of 1066 R EMs was high for all R and R+S OMs except when there was high observation error and 1067 constant fishing pressure, and when median natural mortality and SRRs were estimated. 1068 R+S EMs fit to R OMs exhibited poor convergence regardless of whether natural mortality 1069 or a SRR was estimated. R+S EMs fit to R+S OMs had highest convergence rates when 1070 there was contrast in fishing pressure and low observation error. Convergence rates were 1071 high for all EMs when fit to data from R+S OMs with lower observation error except those 1072 where median natural mortality and/or SRRs were estimated. 1073 Convergence of all EMs fitted to R+M OMs was highest when the OMs had higher natural 1074 mortality process error variability, low observation error, and contrast in fishing pressure 1075 (Figure S14, B). R+M EMs that estimated autocorrelation of process errors had poor 1076 convergence for R+M OMs when there was low natural mortality process error variability 1077 regardless of autocorrelation of the simulated process errors. R+S EMs fitted to data 1078 generated from R+M OMs always converged poorly whether or not median natural mortality 1079 and the Beverton-Holt SRR were estimated. 1080

The R+S EMs, in particular, had poor convergence when fit to data generated from R+Sel

Table S7. Distinguishing characteristics of the estimating models and operating model process error categories (R, R+S, R+M, R+Sel, R+q) where used.

Model	Recruitment model	Median $M$	Process error	R,R+S OMs	R+M OMs	R+Sel OMs	R+q OMs
$\mathrm{EM}_1$	Mean recruitment	0.2	$R (\sigma_{2+} = 0)$	+	_	_	_
$\mathrm{EM}_2$	Beverton-Holt	0.2	$R (\sigma_{2+} = 0)$	+	_	_	_
$EM_3$	Mean recruitment	Estimated	$R (\sigma_{2+} = 0)$	+	_	_	_
$\mathrm{EM}_4$	Beverton-Holt	Estimated	$R (\sigma_{2+} = 0)$	+	_	_	_
$\mathrm{EM}_5$	Mean recruitment	0.2	R+S ( $\sigma_{2+}$ estimated)	+	+	+	+
$\mathrm{EM}_6$	Beverton-Holt	0.2	R+S ( $\sigma_{2+}$ estimated)	+	+	+	+
$\mathrm{EM}_7$	Mean recruitment	Estimated	R+S ( $\sigma_{2+}$ estimated)	+	+	+	+
$\mathrm{EM}_8$	Beverton-Holt	Estimated	R+S ( $\sigma_{2+}$ estimated)	+	+	+	+
$EM_9$	Mean recruitment	0.2	R+M ( $\rho_M = 0$ )	+	+	+	+
$\mathrm{EM}_{10}$	Beverton-Holt	0.2	R+M ( $\rho_M = 0$ )	+	+	+	+
$\mathrm{EM}_{11}$	Mean recruitment	Estimated	R+M $(\rho_M = 0)$	+	+	+	+
$\mathrm{EM}_{12}$	Beverton-Holt	Estimated	R+M $(\rho_M = 0)$	+	+	+	+
$\mathrm{EM}_{13}$	Mean recruitment	0.2	R+Sel $(\rho_{Sel} = 0)$	+	+	+	+
$\mathrm{EM}_{14}$	Beverton-Holt	0.2	R+Sel $(\rho_{Sel} = 0)$	+	+	+	+
$\mathrm{EM}_{15}$	Mean recruitment	Estimated	R+Sel $(\rho_{Sel} = 0)$	+	+	+	+
$\mathrm{EM}_{16}$	Beverton-Holt	Estimated	R+Sel $(\rho_{Sel} = 0)$	+	+	+	+
$\mathrm{EM}_{17}$	Mean recruitment	0.2	R+q $(\rho_q = 0)$	+	+	+	+
$\mathrm{EM}_{18}$	Beverton-Holt	0.2	R+q $(\rho_q = 0)$	+	+	+	+
$EM_{19}$	Mean recruitment	Estimated	R+q $(\rho_q = 0)$	+	+	+	+
$\mathrm{EM}_{20}$	Beverton-Holt	Estimated	R+q $(\rho_q = 0)$	+	+	+	+
$\mathrm{EM}_{21}$	Mean recruitment	0.2	R+M ( $\rho_M$ estimated)	_	+	_	_
$\mathrm{EM}_{22}$	Beverton-Holt	0.2	R+M ( $\rho_M$ estimated)	_	+	_	_
$\mathrm{EM}_{23}$	Mean recruitment	Estimated	R+M ( $\rho_M$ estimated)	_	+	_	_
$\mathrm{EM}_{24}$	Beverton-Holt	Estimated	R+M ( $\rho_M$ estimated)	_	+	_	_
$\mathrm{EM}_{25}$	Mean recruitment	0.2	R+Sel ( $\rho_{\rm Sel}$ estimated)	_	_	+	_
$\mathrm{EM}_{26}$	Beverton-Holt	0.2	R+Sel ( $\rho_{\rm Sel}$ estimated)	_	_	+	_
$\mathrm{EM}_{27}$	Mean recruitment	Estimated	R+Sel ( $\rho_{\rm Sel}$ estimated)	_	_	+	_
$\mathrm{EM}_{28}$	Beverton-Holt	Estimated	R+Sel ( $\rho_{\rm Sel}$ estimated)	_	_	+	_
$EM_{29}$	Mean recruitment	0.2	R+q ( $\rho_q$ estimated)	_	_	_	+
$EM_{30}$	Beverton-Holt	0.2	R+q ( $\rho_q$ estimated)	_	_	_	+
$EM_{31}$	Mean recruitment	Estimated	R+q ( $\rho_q$ estimated)	_	_	_	+
$EM_{32}$	Beverton-Holt	Estimated	R+q ( $\rho_q$ estimated)	_	_	_	+

Table S8. For each OM process error type (columns), percent reduction in deviance for logistic regression models fit to indicators of convergence (maximum absolute gradient  $< 10^{-6}$ ) with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	<u>R</u>	R+S	<u>R+M</u>	<u>R+Sel</u>	<u>R+q</u>
EM Process Error	30.40	0.45	17.57	16.04	24.03
EM M Assumption	2.38	24.11	4.42	1.02	2.66
EM SR Assumption	1.80	0.32	0.96	3.38	2.13
OM Obs. Error	0.12	0.77	0.33	1.76	0.28
OM F History	3.51	6.33	2.36	5.86	5.30
$OM\sigma_R$	<u>&lt;0.01</u>	<u>≤0.01</u>	~	~	~
$OM \sigma_{2\pm}$	~	<u>&lt;0.01</u>	~	~	$\overline{\sim}$
$\widecheck{\mathrm{OM}}_{}\sigma_{\mathcal{M}}$	<del>-</del> ~	<del>-</del>	0.39	_~	<del>_</del> ~
$OM_{\rho_R}$	<del>-</del> ~	<del>-</del>	0.09	~	<del>-</del>
$OM \sigma_{Sel}$	<del>-</del> ~	<del>-</del>	<del>_</del> ~	1.08	<del>-</del>
$\widetilde{\mathrm{OM}}$ $ ho_{Sel}$	<del>-</del> ~	<del>-</del>	<del>_</del> ~	0.01	~
$\widecheck{\mathrm{OM}}_{\sigma_q}$	_ ~	<del>-</del> ~	_ ~	_ ~	0.06
$\underbrace{\mathrm{OM}}_{} \varrho_q$	_ ~	<del>-</del> ~	_ ~	_ ~	≤0.01
All factors	43.69	35.72	29.33	34.57	40.69
+ All Two Way	50.53	42.99	43.91	45.93	48.62
+ All Three Way	52.30	48.41	46.81	47.71	50.40

The maximum of the absolute values of all gradient values for all fits that provided hessian-based standard errors across all simuated data sets of a given OM configuration (A: R and R+S, B: R+M, C: R+Sel, or D: R+q). Results are conditional on EM fits with alternative process error type (colored points and lines), median natural mortality (estimated or known) and recruitment assumptions (Beverton-Holt stock-recruit relationship or not). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

Table S9. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to transformed Mohn's  $\rho$  values for each simulation (Eq. 3) for fishing mortality averaged over all age classes with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	<u>R+S</u>	<u>R+M</u>	<u>R+Sel</u>	$\underbrace{R+q}$
EM M Assumption	0.06	0.09	0.01	0.12	0.01
EM SR assumption	0.01	<u>&lt;0.01</u>	0.01	0.02	0.01
EM Process Error	0.03	0.07	0.02	0.06	0.03
OM Obs. Error	0.16	0.10	0.05	0.02	0.07
OM F History	0.07	0.02	0.03	0.24	0.03
$OM \sigma_R$	<u>&lt;0.01</u>	0.01	~	~	$\bar{\sim}$
$OM\sigma_{2+}$	~	0.09	~	~	$\bar{\sim}$
$OM \sigma_M$	~	~	<u>&lt;0.01</u>	~	$\bar{\sim}$
$OM \rho_R$	_ ~	_ ~	<0.01	_~	<del>_</del> ~
$OM \sigma_{Sel}$	~	~	~	0.01	$_{\sim}^{-}$
$\widetilde{\mathrm{OM}}$ $\rho_{Sel}$	<del>_</del> ~	<del>-</del> ~	<del>-</del> ~	≤0.01	<del>-</del>
$\widecheck{\mathrm{OM}}_{\sigma_q}$	<del>_</del> ~	<del>-</del> ~	<del>-</del> ~	~	≤0.01
$ \underbrace{\mathrm{OM}}_{} \rho_q $	_ ~	_ ~	_ ~	_ ~	0.01
All factors	0.32	0.38	0.12	0.48	0.15
+ All Two Way	0.65	0.67	0.30	0.95	0.43
+ All Three Way	1.18	1.11	0.63	1.34	0.90

Probability of estimating models providing maximum absolute values of gradients less than  $10^{-6}$  assuming alternative process error (colored points and lines), and median natural mortality (estimated or known) and Beverton-Holt stock-recruit relationships (estimated or not; along x-axis) when fitted to operating models that have R and R+S (A), R+Sel (B), R+M (C), or R+q (D) process error structures. Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

Table S10. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to transformed Mohn's  $\rho$  values for each simulation (Eq. 3) for recruitment with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	<u>R</u>	<u>R+S</u>	<u>R+M</u>	<u>R+Sel</u>	<u>R+q</u>
EM M Assumption	0.86	0.56	0.16	1.00	1.27
EM SR assumption	<u>≤0.01</u>	0.02	0.01	0.01	0.01
EM Process Error	0.01	0.59	0.18	0.07	0.04
OM Obs. Error	0.34	0.01	$\underbrace{0.08}_{}$	0.24	0.27
OM F History	0.91	0.22	0.06	1.20	<u>1.67</u>
$OM \sigma_R$	<u>≤0.01</u>	0.14	~	~	~
$OM\sigma_{2\pm}$	~	0.11	~	~	=
$\underbrace{\mathrm{OM}}_{\sigma_{M}} \sigma_{M}$	_ ~	<del>_</del> ~	0.01	_~	<del>_</del> ~
$OM \rho_R$	~	$\overline{\sim}$	≤0.01	~	$\overline{\sim}$
$OM \sigma_{Sel}$	~	$\overline{\sim}$	$\stackrel{-}{\sim}$	0.01	$\overline{\sim}$
$OM \rho_{Sel}$	~	$\overline{\sim}$	$\stackrel{-}{\sim}$	0.01	$\overline{\sim}$
$\underbrace{\mathrm{OM}}_{\sigma_q} \sigma_q$	~	~	~	~	0.01
$\underbrace{\mathrm{OM}}_{} \varrho_q$	_ ~	~	<del>-</del> ~	_ ~	0.01
All factors	2.28	1.74	0.51	2.66	3.51
+ All Two Way	4.20	2.74	1.08	5.08	6.51
+ All Three Way	4.83	3.79	1.79	6.03	7.82

Estimated probability of lowest AIC from logistic regression on the log-standard deviation of the true  $\log(SSB)$  in each simulation for estimating model with Beverton-Holt stock-recruit relationships, rather than the otherwise equivalent EM without the stock-recruit relationship. Results are conditional on alternative assumptions for median natural mortality (estimated or known) and on EMs having the correct process error structure: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Rug along x-axis denotes  $SD(\log(SSB))$  values for each simulation and polygons represent 95% confidence intervals.

- OMs with lower selectivity process error variability or higher observation error (Figure S14,
- 1083 C). R+Sel EMs generally converged better than other EMs for R+Sel OMs with higher
- process error variability, lower observation error, and contrast in fishing pressure regardless
- of whether median natural mortality or a SRR was estimated.
- For R+q OMs, convergence of R+q EMs was generally better than that of other EMs
- when there was contrast in fishing pressure (Figure (S14, D). Convergence of R+S EMs was
- generally worse than that of all other EMs across all OMs whether or not median natural
- mortality or a SRR was estimated. Again, convergence probability generally declined for all
- EMs when median natural mortality or a SRR was estimated.
- We found a wide range of maximum absolute values of gradients for models that converged
- (Figure S3). The largest value observed for a given EM and OM combination was typically
- 1003 <  $10^{-3}$ , but many converged models had values greater than 1. For many OMs, EMs that
- assumed the correct process error type and did not estimate median natural mortality or the
- Beverton-Holt SRR produced the lowest gradient values.

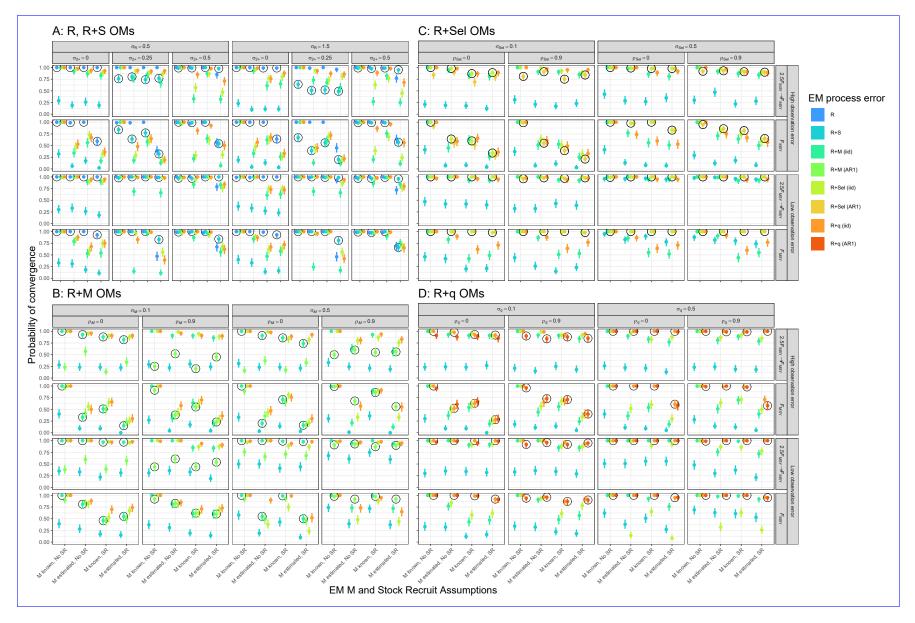


Fig. S14. Median relative error Estimated probability of terminal year recruitment fits providing hessian-based standard errors for estimating models fitted to data sets simulated with EMs assuming alternative process error structures: (colored points and lines), and median natural mortality (estimated or known) and Beverton-Holt stock-recruit relationships (estimated or not; along x-axis) when fitted to operating models that have R and R+S (A), R+Sel (B), R+M (C), or R+q (D) process error structures. Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

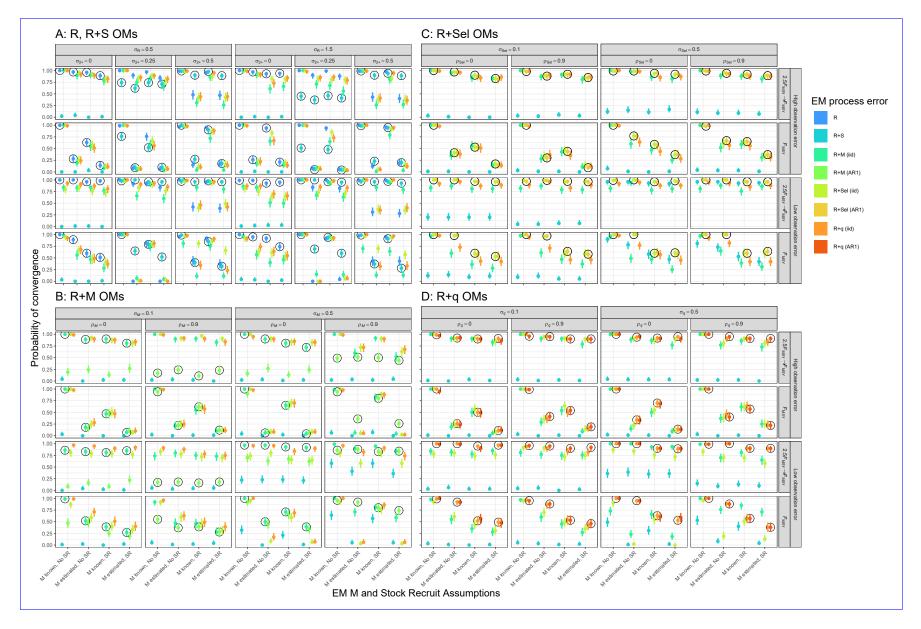


Fig. S15. Median relative error Probability of estimating models providing maximum absolute values of gradients less than  $10^{-6}$  assuming alternative process error (colored points and lines), and median natural mortality (estimated or known) and Beverton-Holt stock-recruit parameters relationships (a and bestimated or not; along x-axis) for estimating models when fitted to data sets simulated with alternative process error structures: operating models that have R and R+S (A), R+Sel (B), R+M (C), or R+q (D) process error structures. Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals.

## 1096 AIC for process error full results

Marginal AIC accurately determined the correct process error assumptions in EMs when 1097 data were generated from R and R+S OMs, regardless of whether median natural mortality 1098 or a SRR was estimated (Figure S16, A). Attempting to estimate median natural mortality 1099 or a SRR separately had a negligible effect on the accuracy of determining the correct process 1100 error assumption. When both were estimated, there was a noticeable reduction in accuracy 1101 when OMs had a constant fishing pressure, low observation error, and larger variability in 1102 recruitment process errors. 1103 For R+M OMs, marginal AIC only accurately determined the correct process error model 1104 and correlation structure when observation error was low and variability in natural mortality 1105 process errors was high (Figures S16, B). Of these OMs, estimating the median natural 1106 mortality rate only reduced the accuracy of AIC when natural mortality process errors were 1107 independent and fishing pressure was constant. For OMs with poor model selection accuracy, 1108 AIC most frequently selected EMs with process errors in catchability (R+q) or selectivity 1109 (R+Sel). Selection of R+S EMs was generally unlikely. 1110 Marginal AIC most accurately determined the correct source of process error and correlation 1111 structure for R+Sel OMs with low observation error (Figures S16, C). When there was low 1112 variability in selectivity process errors and high observation error, R+q or R+S EMs were 1113 more likely to have the best AIC. Whether median natural morality or SRRs were estimated 1114 appeared to have little effect on the performance of AIC. 1115 Marginal AIC most accurately determined the correct source of process error and correlation 1116 structure for R+q OMs with high variability in catchability process errors (Figures S16,D). 1117 The R+q OMs with low variability in catchability process errors and high observation error 1118 had the least model selection accuracy. However, for these OMs, the marginal AIC accurately 1110 determined the correct source of process error (but not correlation structure) except when 1120 OMs assumed a constant fishing pressure and EMs estimated both median natural morality 1121

1122 and the SRR.

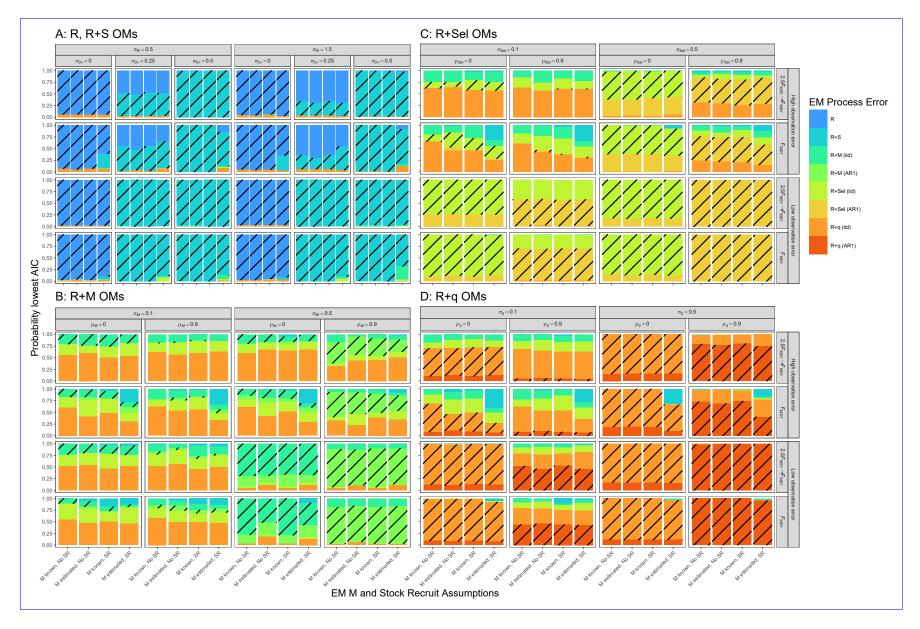


Fig. S16. Median relative error Estimated probability of median natural mortality lowest AIC for estimating models fitted to data sets simulated with EMs assuming alternative process error structures :—(colored bars) conditional on alternative assumptions for median natural mortality (estimated or known) and Beverton-Holt stock-recruit relationships (estimated or not; along x-axis) when fitted to operating models that have R and R+S (A), R+Sel (B), R+M (C), or R+q (D) process error structures. Circled values—Striped bars indicate results where the EM process error structure matches that of the operating modeland vertical lines represent 95% confidence intervals.

# AIC for stock-recruit relationship full results

For each EM and OM combination, we fit logistic regression models to the indicator of 1124 Beyerton-Holt models having lower AIC as a function of the log-standard deviation of the 1125 true log(SSB) (similar to the log of the coefficient of variation for SSB) since simulations 1126 with realized SSB producing low and high recruitment would have larger variation in realized 1127 SSB. 1128 Our comparisons of model performance conditioned on assuming the true process error 1129 configuration is known (EM and OM process error types match) and we focus on results where 1130 the EMs assume median natural mortality is known because there was little difference in 1131 results when the EMs estimated this parameter. Broadly, we found generally poor accuracy 1132 of AIC in selecting models assuming a Beverton-Holt SRR over the null model without an 1133 SRR for all OMs. However, we also found increased accuracy of AIC in determining the 1134 Beverton-Holt SRR when the simulated population exhibited greater variation in spawning 1135 biomass for nearly every OM (Figure S17). 1136 With R and R+S process error assumptions, probability of lowest AIC for the B-H SRR as a 1137 function of SSB variability were greatest for OMs with contrast in fishing pressure and lower 1138 process variability in recruitment (Figure S17, A). The largest variation in SSB occurred 1139 in OMs with larger recruitment variability ( $\sigma_R = 1.5$ ; Figure S17, A, right column group), 1140 but the same high AIC accuracy was achieved for OMs with lower recruitment variability at 1141 lower levels of SSB variation. The level of observation error had little effect on AIC accuracy. 1142 1143 For R+M OMs, probability of lowest AIC for the Beverton-Holt SRR increased steeply 1144 with variation in SSB whether it was induced by contrast in fishing or variation in natural 1145 mortality process error. (Figure S17, B). There was little difference in AIC accuracy whether 1146 the natural mortality process errors were correleted and, similar to R+S OMs, there was also 1147 little effect due to level of observation error. 1148

For R+Sel OMs, contrast in fishing pressure over time was the primary source of variation in SSB and these are the OMs where AIC accuracy for the Beverton-Holt SRR was greatest (Figure S17, C). There was little effect of variability or correlation of selectivity process errors or the level of observation error on AIC accuracy.

Like the R+Sel OMs, the greatest accuracy for AIC in selecting the Beverton-Holt SRR occurred for R+q OMs where there was contrast in fishing pressure over time which is also where there was the greatest variation in SSB (Figure S17, D). There was also little effect of variability or correlation of catchability process errors or the level of observation error on AIC accuracy.

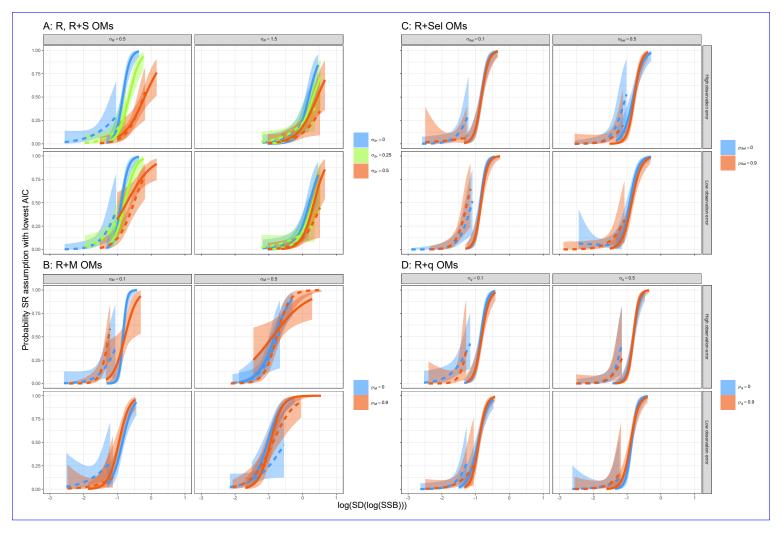


Fig. S17. Median Mohn's  $\rho$  Estimated probability of fishing mortality averaged over all age classes lowest AIC from logistic regression on the log-standard deviation of the true log(SSB) in each simulation for estimating models fitted to data sets simulated model with Beverton-Holt stock-recruit relationships, rather than the otherwise equivalent EM without the stock-recruit relationship. Results are conditional on median M is known in the EM and alternative assumptions EMs having the correct process error structures R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where, and median M is assumed known in the EMprocess error structure matches that of the operating model. Solid and vertical dashed lines are for OMs with and without temporal contrast in fishing pressure, respectively, and polygons represent 95% confidence intervals. Range of results indicates the range of log-standard deviation of log(SSB) for simulations of the particular OM.

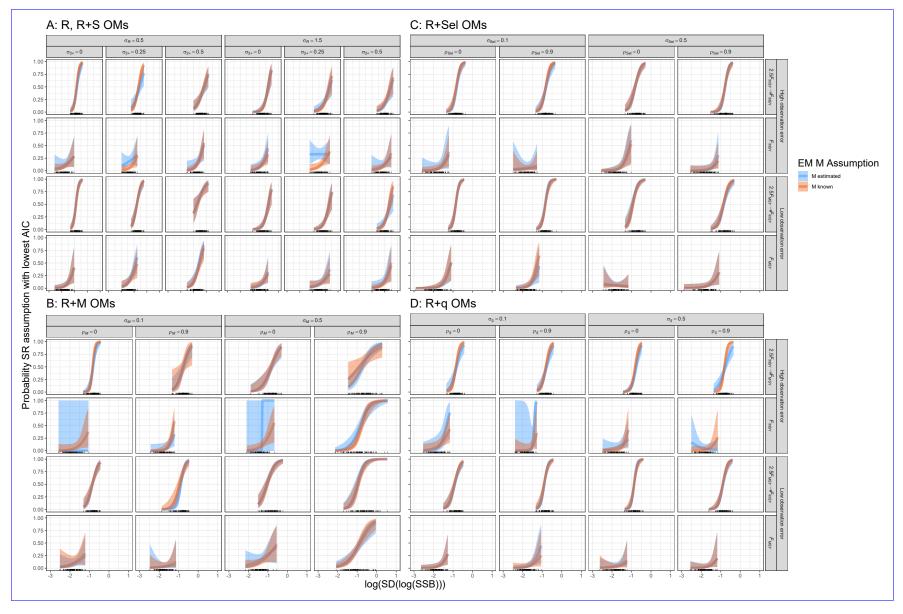


Fig. S18. Median Mohn's  $\rho$  Estimated probability of recruitment lowest AIC from logistic regression on the log-standard deviation of the true log(SSB) in each simulation for estimating models fitted to data sets simulated model with Beverton-Holt stock-recruit relationships, rather than the otherwise equivalent EM without the stock-recruit relationship. Results are conditional on alternative assumptions for median natural mortality (estimated or known) and on EMs having the correct process error structures structure: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled Rug along x-axis denotes  $SD(\log(SSB))$  values indicate results where the EM process error structure matches that of the operating model for each

### Beverton-Holt parameter bias full results

Across all OMs, there was generally less bias and(or) lower variability in estimation of the 1159 Beyerton Holt a parameter than the b parameter. In R and R+S OMs, EMs with the correct 1160 assumptions about process errors provided the least biased estimation of Beverton-Holt SRR 1161 parameters when there was a change in fishing pressure over time and lower variability of 1162 recruitment process errors, but there was little effect of estimating median natural mortality 1163 and a small increase in bias for those OMs that had high observation error (Figure S9, A). For 1164 other R and R+S OMs, estimating natural mortality often resulted in less biased estimation 1165 of SRR parameters. There was generally large variability in relative errors of the SRR 1166 parameter estimates, but the lowest variability occurred with low variability in recruitment 1167 and little or no variability in survival process errors ( $\sigma_{2+} \in \{0, 0.25\}$ ), and contrast in fishing 1168 pressure. 1169 In R+M OMs, the most accurate estimation of SRR parameters for all EM process error 1170 assumptions occurred when there was a change in fishing pressure, greater variability in 1171 natural mortality process errors, and lower observation error (Figure S9, B). Relative to 1172 the R, and R+S OMs, there was even less effect of estimating median natural mortality on 1173 estimation bias for the SRR parameters. 1174 Bias for SRR parameters was large and variability in relative errors was greatest for most 1175 EMs fit to R+Sel OMs with constant fishing pressure (Figure S9, C). Less bias in parameter 1176 estimation occurred for OMs with a change in fishing pressure and the best accuracy occurred 1177 for those OMs that had low observation error and more variable and uncorrelated selectivity 1178 process errors, and when the EMs had with the correct process error assumption. There was 1179 little effect of estimating natural mortality on relative errors for SRR parameters. 1180 Like R+Sel OMs, relative errors in SRR parameters for R+q OMs were more accurate 1181 for most EM process error types when OMs had contrast in fishing pressure and lower 1182 observation error (Figure S9, D). However, the best accuracy occurred for those OMs that 1183

had lower variability in catchability process errors. The worst accuracy of SRR parameter estimation regardless of EM type occurred when R+q OMs had low observation error and constant fishing pressure (Figure S9, D, fourth row).

#### 1187 Median natural mortality rate bias full results

Across all OMs and EMs there was little effect of estimating SRRs on the bias in estimation of median natural mortality (Figure S10). Median natural mortality rate was estimated accurately by all EM process error types for all R OMs except those with high observation error and constant fishing pressure, in which case relative errors were high (Figure S10, A,  $\sigma_{2+} = 0$ ). For R+S OMs estimation of median natural mortality rate was most accurate when observation error was low and there was contrast in fishing pressure and the EM process error type was correct.

For R+M OMs, median natural mortality was estimated most accurately, regardless of EM process error type, when OMs had a change in fishing pressure and low observation error (Figure S10, B). However, those R+M OMs that also had greatest variability in AR1 correlated natural mortality process errors only had unbiased estimation when the EM process error type was correct.

All EM process error types accurately estimated median natural mortality rate for R+Sel
OMs that had contrast in fishing pressure, low observation error, and low selectivity process
error variability (Figure S10, C). When selectivity process error variability increased, the
incorrect EM process errors produce more biased estimation of median natural mortality
rate. The least accurate estimation occurred for all EM process error types when observation
error was high and fishing pressure was constant.

Like R+Sel OMs, all EM process error types produced accurate estimation of median natural mortality rate when fit to R+q OMs with contrast in fishing pressure, low observation error and low catchability process error variability (Figure S10, D). Most EM process error

types produced biased estimation of median natural mortality when R+q OMs had high observation error and constant fishing pressure.

### 1211 Spawning stock biomass bias full results

For R OMs ( $\sigma_{2+}=0$ ), there was no indication of bias (95% confidence intervals included 0) in 1212 terminal year SSB for any of the estimating models regardless of process error assumptions, 1213 except when no SR assumption was made, recruitment variability was low, and there was 1214 contrast in fishing mortality and high observation error (Figure S6, A). However, errors in 1215 terminal SSB estimates were highly variable when median natural mortality was estimated 1216 and there was constant fishing pressure and high observation error (Figure S6, A, second 1217 row). 1218 For R+S OMs, the EMs with matching process error assumptions generally produced 1210 unbiased estimation of terminal SSB except when median natural mortality was estimated 1220 and there was high observation error. In R+S OMs with low observation error, EMs with 1221 incorrect process error assumptions typically provided biased estimation of terminal year 1222 SSB. Estimating the Beverton-Holt SRR had little discernible effect on bias of terminal 1223 year SSB estimation whereas estimating median M tended to produce more variability in 1224 errors in terminal SSB estimation similar to R OMs. 1225 For R+ M OMs with low variability in natural mortality process errors, low observation error 1226 and contrast in fishing motality over time all EMs produced low variability in SSB estimation 1227 error that indicated unbiasedness (Figure S6, B, third row). However, larger variability in 1228 natural mortality process errors increased bias of EMs without the correct process error 1229 type. Estimating median natural mortality increased variability of SSB estimation error 1230 particularly for OMs with high observation error and constant fishing pressure over time. It 1231 also increased bias in SSB estimation for many R+M OMs. Like R and R+S OMs, estimating 1232

a SRR had little discernible effect on SSB bias.

1233

For R+Sel OMs, there was no evidence of bias for any EMs when variability in selectivity 1234 process error and observation error was low, and with contrast in fishing mortality (Figure 1235 S6, C). The largest bias occurred for any EMs that estimated median natural mortality 1236 when the OMs had high observation error, constant fishing pressure, and greater variability 1237 in selectivity process errors ( $\sigma_{\rm Sel} = 0.5$ ) or low selectivity process errors ( $\sigma_{\rm Sel} = 0.1$ ) and low 1238 observation error. However, there was no evidence of SSB bias for correctly specified R+Sel 1230 EMs when observation error was low and variation in selectivity process errors was larger, 1240 whether median natural mortality was estimated or not (Figure S6, C, third row). We 1241 only observed an effect of estimating the Beverton-Holt SRR for R+Sel OMs that had high 1242 observation error and contrast in fishing pressure where estimating the SRR produced less 1243 biased SSB estimation for many EMs (Figure S6, C, top row). 1244 All EMs fit to data from R+q OMs with low observation error and contrast in fishing 1245 pressure exhibited little evidence of bias in terminal SSB estimation except for R+M EMs 1246 when there was no AR1 correlation in catchability process errors (Figure S6, D). Many 1247 EMs also performed well in R+q OMs with low observation error, but no contrast in fishing 1248 pressure. For R+q OMs with high observation error and contrast in fishing pressure, EMs 1249 that estimated the Beverton-Holt SRR exhibited less SSB bias than those that did not. 1250 Estimating median natural mortality in the EMs only resulted in much more variable SSB 1251 estimation errors when there was no contrast in fishing pressure (Figure S6, D, first and 1252 third rows). 1253 For all OM process error types, relative errors in terminal year recruitment were generally 1254 more variable than SSB, but effects of R and R+S OM and EM attributes on bias (i.e, 1255 negative or positive or none) were similar (Figure S8, A). Furthermore, for EM configurations 1256 where bias in terminal SSB was evident, median relative errors in recruitment often indicated 1257 stronger bias in recruitment of the same sign. 1258

### Mohn's $\rho$ Full results

Mohn's  $\rho$  for SSB was small in absolute value for all R and R+S OMs, regardless of EM 1260 process error types, and whether median natural mortality rate or SRRs were estimated 1261 (Figure S11, A). The strongest retrospective patterns (highest absolute Mohn's  $\rho$  values) 1262 occurred in OMs with the largest apparent survival process error variability, high observation 1263 error, and contrast in fishing pressure, but only for EMs with the incorrect process error type 1264 and where median natural mortality rate was assumed known (median  $\rho$  was approximately 1265 -0.15). For R+M, R+Sel, and R+q OMs, Mohn's  $\rho$  was also small in absolute value, but 1266 median values were all closer to 0 than the largest values in the R and R+S OMs (Figure 1267 S11,B-D). For these OMs, there was no noticeable effect of estimation of median natural 1268 mortality rate or SRRs on Mohn's  $\rho$  for any EM process error types. 1269 Mohn's  $\rho$  for recruitment was small in absolute value for all R OMs with low variability in 1270 recruitment process errors, regardless of EM process error type, and whether median natural 1271 mortality rate or SRRs were estimated (Figure S13, A). However, R and R+S OMs with 1272 greater recruitment process variability and higher observation error had median Mohn's  $\rho$ 1273 for recruitment greater than zero for most EMs even when the EM process error type was 1274 correct. In R+S OMs with lower observation error, EMs with the correct process error type 1275 exhibited better median Mohn's  $\rho$  close to 0 than EMs with the incorrect process error type. 1276 For R+M, R+Sel, and R+q OMs, results for Mohn's  $\rho$  for recruitment are similar to those 1277 for SSB, but the range in median values and variation in Mohn's  $\rho$  values for a given OM 1278 are generally larger for recruitment (Figure S13, B-D). 1279