

¹ An investigation of factors affecting inferences from and
² reliability of state-space age-structured assessment
³ models

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23 patterns

24 **Abstract**

25 State-space models ~~are increasingly used for stock assessment~~, ~~have been promoted as the~~
26 ~~next-generation of fisheries stock assessment~~ and evaluations of their ~~statistical reliability~~
27 ~~and best practices for selecting among process error configurations are~~ ~~reliability is~~ needed.
28 We simulated ~~72~~ operating models that varied fishing pressure~~and observation error~~ across
29 ~~process errors in recruitment, survival, selectivity, catchability, and /or natural mortality.~~ We
30 ~~fit~~, magnitude of observation error, and sources of process error. For each operating model,
31 ~~we fit a range~~ estimating models with ~~different assumptions on the process error source and~~
32 ~~whether median natural mortality or a stock-recruit relationship were estimated.~~ Estimating
33 ~~models without a stock-recruit relationship that assumed the correct process error source and~~
34 ~~median natural mortality had high convergence rates and low bias.~~ Bias was also low under
35 ~~many incorrect process error assumptions when there was a range of correct and incorrect~~
36 ~~assumptions.~~ We measured reliability of estimating models by convergence rate, accuracy
37 ~~of marginal AIC, estimation bias, and magnitude of retrospective patterns.~~ All reliability
38 ~~measures were generally better with lower observation error,~~ contrast in fishing pressure
39 ~~and low observation error.~~ Marginal AIC most accurately distinguished process errors on
40 ~~recruitment, survival, and selectivity, and other process error sources when variability was~~
41 ~~greater over time, and when median natural mortality rate is known.~~ Magnitude of the
42 ~~log-likelihood gradients was not a reliable indicator of convergence.~~ AIC can generally
43 ~~distinguish process error type with lower observation error and higher true process error~~
44 ~~variability.~~ Distinguishing the stock recruit relationship with AIC required large contrast
45 ~~in spawning biomass and low recruitment variation, but bias in stock-recruit parameter~~
46 ~~estimation was prevalent.~~ Retrospective patterns were ~~generally small but were sizable for~~

47 recruitment when observation error was high. These results help establish the statistical
48 reliability of state space assessment models and pave the way for the next generation of
49 fisheries stock assessmentnot large for mis-specified models. These findings improve our
50 understanding of when results from state space models will be reliable.

51 **Introduction**

52 Application of state-space models in fisheries stock assessment and management has ex-
53 panded dramatically within International Council for the Exploration of the Sea (ICES),
54 Canada, and the Northeast US (Nielsen and Berg 2014; Cadigan 2016; Pedersen and Berg
55 2017; Stock and Miller 2021). State-space models treat latent population characteristics as
56 statistical time series with periodic observations that also may have error due to sampling
57 or other ~~sources of measurement error~~other measurement properties. Traditional assessment
58 models may use state-space approaches to account for temporal variability in population
59 characteristics (Legault and Restrepo 1999; Methot and Wetzel 2013), but these models
60 treat the annual parameters as penalized fixed ~~effects~~effect parameters where the variance
61 parameters controlling the penalties are assumed known (Thorson and Minto 2015). Modern
62 state-space models can estimate the annually varying parameters as random effects with vari-
63 ance parameters estimated using maximum marginal likelihood or corresponding Bayesian
64 approaches. These latter approaches are considered best practice and ~~a~~are recommended
65 for the next generation of stock assessment models (Hoyle et al. 2022; Punt 2023).

66 State-space stock assessment models, with nonlinear functions of latent parameters and
67 multiple types of observations with varying distributional assumptions, are one of the most
68 complex examples of this analytical approach. Statistical aspects of state-space models and
69 their application within fisheries have been studied extensively, but previous work has focused
70 primarily on linear and Gaussian state-space models (Aeberhard et al. 2018; Auger-Méthé
71 et al. 2021). Therefore, current understanding of the reliability of state-space models does
72 not extend to usage for stock assessment.

73 As state-space models provide greater flexibility by allowing multiple processes to vary as
74 random effects (Nielsen and Berg 2014; Aeberhard et al. 2018; Stock et al. 2021), one of the
75 most immediate questions regards the implications of mis-specification among alternative
76 sources of process error. Incorrect treatment of population attributes as temporally varying

77 (Trijoulet et al. 2020; Liljestrand et al. 2024) could lead to misidentification of stock
78 status and biased population estimates, ultimately impacting fisheries management decisions
79 (Legault and Palmer 2016; Szuwalski et al. 2018; Cronin-Fine and Punt 2021). Furthermore,
80 biological, fishery, and observational processes are often confounded in catch-at-age data,
81 which may adversely affect the ability to distinguish between true process variability and
82 observational error (Li et al. In reviewpress; Punt et al. 2014; Stewart and Monnahan 2017;
83 Cronin-Fine and Punt 2021; Fisch et al. 2023).

84 Li et al. (2024) conducted a full-factorial simulation-estimation study to assess model reli-
85 ability when confounding random-effects processes (numbers-at-age, fishery selectivity, and
86 natural mortality) were included. Their results suggest that while state-space models can
87 generally identify sources of process error, overly complex models, even when misspecified
88 (i.e., incorporating process error that did not exist in reality), often performed similarly to
89 correctly specified models, with little to no bias in key management quantities. Similarly,
90 Liljestrand et al. (2024) found little downside in assuming process error in recruitment or
91 selectivity, even when it was absent.

92 Despite mounting efforts, several limitations remain. First, confounding processes that can
93 be treated as random effects in the model were not have not been thoroughly examined or
94 tested within a simulation-estimation framework. Second, previous studies relied on oper-
95 ating models conditioned on specific fisheries, limiting their generalizability (Li et al. In
96 reviewpress; Liljestrand et al. 2024). In particular, the effects of observation error and un-
97 derlying fishing history have not been fully isolated in simulation study designs, making it
98 challenging to disentangle the interplay between process and observation error magnitudes,
99 as demonstrated in Fisch et al. (2023). Third, explicitly modeling stock-recruit relation-
100 ships (SRRs) as mechanistic drivers of population dynamics is promising (Fleischman et al.
101 2013; Pontavice et al. 2022), but reliability of inferences within integrated state-space age-
102 structured models has not been evaluated. Evidence from other studies suggests that when
103 both process and observation errors are unknown, estimating density dependence parameters

¹⁰⁴ becomes highly uncertain (Knape 2008; Polansky et al. 2009). In particular, Knape (2008)
¹⁰⁵ demonstrated that stronger density dependence becomes increasingly difficult to estimate in
¹⁰⁶ the presence of observation error. Therefore, it is crucial to assess whether density depen-
¹⁰⁷ dence mechanisms can be estimated with sufficient precision for use in fisheries management
¹⁰⁸ (Auger-Méthé et al. 2016). Finally, although the importance of autocorrelation in pro-
¹⁰⁹ cess errors is recognized, investigations of the ability to distinguish state-space assessment
¹¹⁰ models with and without autocorrelation and whether such misspecification is detrimental
¹¹¹ to estimation of important population metrics are lacking (Johnson et al. 2016; Xu et al.
¹¹² 2019).

¹¹³ In the present study, we conduct a simulation study with operating models (OMs) varying by
¹¹⁴ degree of observation error, source and variability of process error, and fishing history. The
¹¹⁵ simulations from these OMs are fitted with estimation models (EMs) that make alternative
¹¹⁶ assumptions for sources of process error, whether a SRR was estimated, and whether natural
¹¹⁷ mortality is estimated. Given the confounding nature of process errors, developing diagnostic
¹¹⁸ tools to detect model misspecification is of great scientific interest and could aid the next
¹¹⁹ generation of stock assessments (Auger-Méthé et al. 2021). We evaluate whether ~~convergencee~~
¹²⁰ ~~and~~-OM and EM attributes affect rates of convergence and the ability of Akaike Information
¹²¹ Criterion (AIC) ~~ean~~-to correctly determine the source of process error ~~and~~-or the existence
¹²² of a SRR. We also evaluate when retrospective patterns occur and the degree of bias in ~~the~~
¹²³ outputs of the assessment model that are important for management.

¹²⁴ Methods

¹²⁵ We used the Woods Hole Assessment Model (WHAM) to configure OMs and EMs in our
¹²⁶ simulation study (~~Miller and Stock 2020; Stock~~ Stock and Miller 2021; Miller et al. 2025).
¹²⁷ WHAM is an R package freely available via a ~~github~~-Github repository and is built on the
¹²⁸ Template Model Builder package (Kristensen et al. 2016). For this study we used version

¹²⁹ 1.0.6.9000, commit 77bbd94. WHAM has also been used to configure OMs and EMs for
¹³⁰ closed loop simulations evaluating index-based assessment methods (Legault et al. 2023)
¹³¹ and is currently used or accepted for use in management of numerous NEUS fish stocks
¹³² (e.g., NEFSC 2022a, 2022b; NEFSC 2024).

¹³³ We completed a simulation study with a number of OMs that can be categorized based
¹³⁴ on where process error random effects were assumed: recruitment (R, assumed present in
¹³⁵ all models), apparent survival (denoted R+S), natural mortality (R+M), fleet selectivity
¹³⁶ (R+Sel), or index catchability (R+q). We refer to the (R+S) OMs as modeling appar-
¹³⁷ ent survival because on ~~logscale~~ ~~log-scale~~ the random effects ($\epsilon_{a,y}$) are additive to the total
¹³⁸ mortality (F+M) between numbers at age, thus they modify the survival term. However, as
¹³⁹ Stock and Miller (2021) note, these random effects can be due to events other than mortality,
¹⁴⁰ such as immigration, emigration, missreported catch, and other sources of misspecification.
¹⁴¹ For each OM, assumptions about the magnitude of the variance of process errors and obser-
¹⁴² vations are required and the values we used were based on a review of the range of estimates
¹⁴³ from Northeast United States (NEUS) assessments using WHAM.

¹⁴⁴ In total, we configured 72 OMs with alternative assumptions about the source and magnitude
¹⁴⁵ of process errors, magnitude of observation error in indices and age composition data, and
¹⁴⁶ contrast in fishing pressure over time. ~~We fitted 20 EMs to observations generated from each~~
¹⁴⁷ ~~of~~ For each OM, we simulated 100 ~~simulations where process errors were also simulated.~~ Each
¹⁴⁸ ~~EM~~ population time series with process errors and, for each time series, simulated observation
¹⁴⁹ data sets. For each data set, we fitted a number of EMs that differed in assumptions about
¹⁵⁰ the source of process errors, whether natural mortality (or the median for models with process
¹⁵¹ error in natural mortality) was estimated, and whether a Beverton-Holt SRR was estimated
¹⁵² within the EM. Details of each of the OMs and EMs are described below.

¹⁵³ We did not use the log-normal bias-correction feature for process errors or observations
¹⁵⁴ described by (Stock and Miller 2021) for OMs and EMs to simplify interpretation of the

₁₅₅ study results (Li et al. [In-review2025](#)). All code we used to perform the simulation study
₁₅₆ and summarize results can be found at <https://github.com/timjmiller/SSRTWG/tree/main/>
₁₅₇ Project_0/code.

₁₅₈ **Operating models**

₁₅₉ **Population**

₁₆₀ We intended the population demographics and observation types to represent a general
₁₆₁ NEUS groundfish stock. The population consists of 10 age classes, ages 1 to 10+, with the
₁₆₂ last being a plus group that accumulates ages 10 and older. We assume spawning occurs
₁₆₃ annually 1/4 of the way through the year. The maturity at age was a logistic curve with a_{50}
₁₆₄ = 2.89 and slope = 0.88 (Figure S1, top left).

₁₆₅ Weight at age was generated with a von Bertalanffy growth function

$$L_a = L_\infty \left(1 - e^{-k(a-t_0)}\right)$$

₁₆₆ where $t_0 = 0$, $L_\infty = 85$, and $k = 0.3$, and a L-W relationship such that

$$W_a = \theta_1 L_a^{\theta_2}$$

₁₆₇ where $\theta_1 = e^{-12.1}$ and $\theta_2 = 3.2$ (Figure S1, top right).

₁₆₈ We assumed a Beverton-Holt SRR with constant pre-recruit mortality parameters for all
₁₆₉ OMs. All biological inputs to calculations of spawning biomass per recruit (i.e., weight,
₁₇₀ maturity, and natural mortality at age) are constant in the apparent survival (R+S) se-
₁₇₁ lectivity (R+Sel), and survey catchability (R+q) process error OMs. Therefore, steepness
₁₇₂ and unfished recruitment are also constant over the time period for those OMs (Miller and
₁₇₃ Brooks 2021). We assumed a value of 0.2 for the natural mortality rate in OMs without

174 process errors on natural mortality. We specified unfished recruitment equal to e^{10} and
175 $F_{\text{MSY}} = F_{40\%} = 0.348$, which equates to a steepness of 0.69 and $a = 0.60$ and $b = 2.4 \times 10^{-5}$
176 for the Beverton-Holt parameterization

$$N_{1,y} = \frac{a\text{SSB}_{y-1}}{1 + b\text{SSB}_{y-1}}$$

177 (Figure S1, bottom right). We assumed a value of 0.2 for the natural mortality rate in
178 OMs without process errors on natural mortality and for the median rate for OMs with
179 process errors on For OMs with time-varying random effects for natural mortality, steepness
180 is not constant. However, we used the same a and b parameters as other OMs, which equates
181 to a steepness and R_0 at the median of the time series process for natural mortality. Similarly,
182 For OMs with time-varying random effects for fishery selectivity, F_{MSY} also varies temporally,
183 so equilibrium conditions for these OMs are defined for mean selectivity parameters.

184 We used two fishing scenarios for OMs. In the first scenario, the stock experiences overfishing
185 at $2.5F_{\text{MSY}}$ for the first 20 years followed by fishing at F_{MSY} for the last 20 years (denoted
186 $2.5F_{\text{MSY}} \rightarrow F_{\text{MSY}}$). In the second scenario, the stock is fished at F_{MSY} for the entire time
187 period (40 years). The magnitude of the overfishing assumptions is based on average esti-
188 mates of overfishing for NEUS groundfish stocks from Wiedenmann et al. (2019) and similar
189 to the approach in Legault et al. (2023).

190 The second scenario represents the ideal situation where the stock is fished at an optimal
191 level, but provides less contrast in stock sizes over time. We specified initial population
192 abundance at age at the equilibrium distribution that corresponds to fishing at either
193 $F = 2.5 \times F_{\text{MSY}}$ or $F = F_{\text{MSY}}$. This implies that, for a deterministic model,
194 the abundance at age would not change from the first year to the next.

195 For OMs with time-varying random effects for M , steepness is not constant. However, we
196 used the same a and b parameters as other OMs, which equates to a steepness and R_0 at
197 the median of the time series process for M . For OMs with time-varying random effects for

198 ~~fishery selectivity, is also not constant, but since we use the same F history as other OMs,~~
199 ~~this corresponds to at the mean selectivity parameters.~~

200 **Fleets**

201 We assumed a single fleet operating year round for catch observations with logistic selectivity
202 ~~for the fleet~~ ($a_{50} = 5$ and slope = 1; Figure S1, bottom left). This selectivity was used to
203 define F_{MSY} for the Beverton-Holt SRR parameters above. We assumed a logistic-normal
204 distribution with no correlation on the multivariate normal scale for the corresponding annual
205 age-composition observations~~for the fleet~~.

206 **Indices**

207 Two time series of fishery-independent surveys measured in numbers are generated for the
208 entire 40 year period with one occurring in the spring (0.25 of each year) and one in the
209 fall (0.75 of each year), representing current bottom trawl surveys conducted in the NEUS.
210 Catchability of both surveys are assumed to be 0.1. Like the fishing fleet, we assumed logistic
211 selectivity for both indices ($a_{50} = 5$ and slope = 1) and a logistic-normal distribution with
212 no correlation on the multivariate normal scale for the annual age-composition observations.

213 **Observation Uncertainty**

214 The standard deviation for log-aggregate catch was 0.1 for all OMs, a common assumption for
215 NEUS stock assessments. Two levels of observation error variance (high and low) were speci-
216 fied for indices and all age composition observations (both indices and catch). The low uncer-
217 tainty specification assumed a standard deviation of 0.1 for both series of log-aggregate index
218 observations, and the standard deviation of the logistic-normal for age composition observa-
219 tions was 0.3. In the high uncertainty specification, the standard deviation for log-aggregate
220 indices was 0.4 and that for the age composition observations was 1.5. The low standard

221 deviation for index observations is typical for fish stocks that are consistently sampled across
222 survey stations whereas the high value is typical for more sporadically sampled stocks. The
223 standard deviations for the age composition observations were determined from the range of
224 values estimated from WHAM fits to NEUS stocks that assumed the logistic-normal model.
225 For all EMs, the standard deviation for log-aggregate observations was assumed known
226 whereas that for the logistic-normal age composition observations was estimated.

227 Operating models with random effects on numbers at age

228 For operating models with random effects on recruitment and(or) apparent survival (R,
229 R+S), we assumed marginal standard deviations for recruitment of $\sigma_R \in \{0.5, 1.5\}$ and. The
230 marginal standard deviations for apparent survival random effects at older age classes of
231 were $\sigma_{2+} \in \{0, 0.25, 0.5\}$. The full factorial combination of these process error assumptions
232 (~~2x3~~² x 3 levels) and scenarios for fishing history (2 levels) and observation error (2 levels)
233 scenarios described above results in 24 different R ($\sigma_{2+} = 0$) and R+S operating models
234 (Table S1).

235 Operating models with random effects on natural mortality

236 All R+M OMs treat natural mortality as constant across age, but with annually varying
237 random effects. WHAM treats natural mortality as a log-transformed parameter

$$\log M_{y,a} = \mu_M + \epsilon_{M,y}$$

238 that is a linear combination of a mean log-natural mortality parameter that is constant
239 across ages ($\mu_M = \log(0.2)$) and any annual random effects are marginally distributed as
240 $\epsilon_{M,y} \sim N(0, \sigma_M^2)$. The marginal standard deviations we assumed for log natural mortality
241 random effects were $\sigma_M \in \{0.1, 0.5\}$ and the random effects were either uncorrelated or first-
242 order autoregressive (AR1, $\rho_M \in \{0, 0.9\}$). Uncorrelated random effects were also included

243 on recruitment with $\sigma_R = 0.5$ (hence, we denote these OMs as R+M). The full factorial
 244 combination of these process error assumptions and fishing history (2 levels) and observation
 245 error (2 levels) scenarios described above results in 16 different R+M OMs (Table S2).

246 Operating models with random effects on fleet selectivity

247 WHAM treats selectivity parameter s as a logit-transformed parameter

$$\log \left(\frac{p_{s,y} - l_s}{u_s - p_{s,y}} \right) = \mu_s + \epsilon_{s,y}$$

248 that is a linear combination of a mean μ_s and any annual random effects marginally dis-
 249 tributed as $\epsilon_{s,y} \sim N(0, \sigma_s^2)$, where the lower and upper bounds of the parameter (l_s and
 250 u_s) can be specified by the user. All selectivity parameters (a_{50} and slope parameters) were
 251 bounded by 0 and 10 for all OMs and EMs. The marginal standard deviations we assumed
 252 for logit scale random effects were $\sigma_s \in \{0.1, 0.5\}$ and AR1 autocorrelation parameters of
 253 $\rho_s \in \{0, 0.9\}$. Like R+M OMs, the full factorial combination of these process error assump-
 254 tions (2x2 levels) and scenarios described above for fishing history (2 levels) and observation
 255 error (2 levels) results in 16 different R+Sel OMs (Table S3).

256 Operating models with random effects on index catchability

257 Like selectivity parameters, WHAM treats catchability for an index i as a logit-transformed
 258 parameter

$$\log \left(\frac{q_{i,y} - l_i}{u_i - q_{i,y}} \right) = \mu_i + \epsilon_{i,y}$$

259 that is a linear combination of a mean μ_i and any annual random effects marginally dis-
 260 tributed as $\epsilon_{i,y} \sim N(0, \sigma_i^2)$ where the lower and upper bounds of the catchability (l_i and u_i)
 261 can be specified by the user. We assumed bounds of 0 and 1000 for all OMs and EMs. For
 262 all OMs and EMs with process errors on catchability, the temporal variation only applies

263 to the first index, which could be interpreted as capturing some unmeasured seasonal pro-
264 cess that affects availability to the survey. The marginal standard deviations we assumed
265 for logit scale random effects were $\sigma_i \in \{0.1, 0.5\}$ and AR1 autocorrelation parameters of
266 $\rho_i \in \{0, 0.9\}$. Like R+M and R+Sel OMs, the full factorial combination of these process
267 error assumptions and fishing history (2 levels) and observation error (2 levels) scenarios
268 described above results in 16 different R+q OMs (Table S4).

269 **Estimation models**

270 For each of the data sets simulated from an OM, 20 EMs were fit. A total of 32 different
271 EMs were fit across OMs where the subset of 20 depended on the source of process error
272 in the OM (Table S5). The EMs have different assumptions about the source of process
273 error (R+S, R+M, R+Sel, R+q) and whether or not 1) there is temporal autocorrelation,
274 2) a Beverton-Holt SRR is estimated, and 3) the natural mortality rate (μ_M , the constant
275 or mean on log scale for R+M EMs) is estimated. For simplicity we refer to the derived
276 estimate e^{μ_M} as the median natural mortality rate regardless of whether natural mortality
277 random effects are estimated in the EM.

278 Subsets of 20 EMs in Table S5 were fit to simulate data sets from each of the OM process error
279 categories. For R and R+S OMs, fitted EMs had matching process error assumptions as well
280 as R+Sel, R+M, and R+q assumptions without autocorrelation. ~~Similarly,~~ For other OM
281 process error categories, we fit EMs with ~~matching correct~~ process error assumptions ~~as well~~
282 ~~as other process error~~, ~~the correct process error type but incorrect correlation assumption,~~
283 ~~and the incorrect process error~~ types without autocorrelation. ~~As such, EMs were configured~~
284 ~~completely correctly for the OM, or they had mis-specification in assumptions of process error~~
285 ~~autocorrelation, the type of process error, and(or) the SRR (Beverton-Holt or none).~~

286 The maturity at age, weight at age for catch and ~~SSB~~spawning stock biomass (SSB), and
287 observation error ~~variance of~~standard deviations for aggregate catch and indices were all

288 assumed known at the true values. However, the variance parameters for the logistic-
289 normal distributions for age composition observations were estimated in the EMs. ~~As~~
290 ~~such, EMs would either be configured completely correctly for the OM, or there could be~~
291 ~~mis-specification in assumptions of process error autocorrelation, the type of process error,~~
292 ~~or the SRR (Beverton-Holt or none).~~

293 Measures of reliability

294 Convergence

295 The first measure of reliability we investigated was frequency of convergence when fitting
296 each EM to the simulated data sets. There are various ways to assess convergence of the fit
297 (e.g., Carvalho et al. 2021; Kapur et al. 2025), but given the importance of estimates of un-
298 certainty when using assessment models in management, we estimated ~~probability~~probability
299 of convergence as measured by occurrence of a positive-definite ~~hessian~~Hessian matrix at
300 the optimized negative log-likelihood that could be inverted ~~(i.e., providing Hessian-based~~
301 ~~standard error estimates)~~. We also provide results in the Supplementary Materials for ~~the~~
302 ~~maximum convergence defined by the maximum absolute gradient < 1^-6 and the maximum~~
303 of the absolute ~~values among all gradients~~gradient values for all fits of a given EM to all
304 simulated data sets from a given OM that produced ~~hessian-based~~Hessian-based standard
305 errors for all estimated fixed effects. This provides an indication of how poor the calculated
306 gradients can be, but still presumably converged adequately enough for parameter inferences.
307 ~~We used the Clopper-Pearson exact method for constructing 95% confidence intervals of the~~
308 ~~probabilities of convergence (Clopper and Pearson 1934; Thulin 2014).~~

309 AIC for model selection

310 We ~~estimated the probability of selection~~investigated the reliability of AIC-based model
311 ~~selection for two purposes. First, we analyzed selection~~ of each process error model struc-

ture (R, R+S, R+M, R+Sel, R+q) using marginal AIC. For a given operating model OM simulated data set, we compared AIC for EMs that were all configured the same for with different process error assumption conditional on whether median natural mortality (known or estimated) and the SRR (rate and the Beverton-Holt or none).

We also estimated the probability of correctly selecting SRR were estimated. Second, we analyzed AIC-based selection between EMs with Beverton-Holt SRR assumed and models without the SRR (null model). We made these comparisons between models that otherwise assumed the same process error structure as the operating model and both of the compared models either estimate median natural mortality or assume it is known and without the Beverton-Holt SRR assumed. Contrast in fishing pressure and time series with recruitment at low stock size has been shown to improve estimation of SRR parameters (Magnusson and Hilborn 2007; Conn et al. 2010). Our preliminary inspections of the proportions of simulations where the correct recruitment model was chosen indicated generally poor performance of AIC in determining the Beverton-Holt model for a given set of OM factors (including contrast in fishing pressure) indicated generally poor performance of AIC, even when the EM was configured with the correct process error type. Therefore, we fit logistic regression models to the indicator of Beverton-Holt models having lower AIC as a function conditioned on the EMs having the correct process error assumption and also considered the effect of the log-standard deviation of the true log(SSB) ($\log SD_{SSB}$; similar to the log of the coefficient of variation for SSB) on model selection since simulations with realized SSB producing low and high recruitment would have larger variation in realized SSB.

All model selection results condition on whether all of the compared estimating models completed the only on completion of the optimization process without failure for all of the compared EMs. We did not condition on convergence as defined by a gradient threshold or invertibility of the hessian above because optimization could correctly determine an inappropriate process error assumption by estimating variance parameters at the lower bound

339 of zero. Such an optimization could indicate poor convergence but the likelihood would be
340 equivalent to that without the mis-specified random effects and the AIC would be appro-
341 priately higher because more (variance) parameters were estimated. All other measures of
342 reliability described below (bias and Mohn's ρ) use these same criteria for inclusion of EM
343 fits in the summarized results.

344 **Bias**

345 We also investigated bias in estimation of various model attributes as a measure of reliability.

346 For a given model attribute we calculated the relative error

$$\text{RE}(\theta_i) = \frac{\hat{\theta}_i - \theta_i}{\theta_i}$$

347

$$\text{RE}(\theta_j) = \frac{\hat{\theta}_j - \theta_j}{\theta_j} \quad (1)$$

348 from fitting a given estimating model to simulated data set i, j configured for a given OM
349 where $\hat{\theta}_i$ and θ_i , $\hat{\theta}_j$ and θ_j are the estimated and true values for simulation i . We estimated
350 bias as the median of the relative errors across all simulations for a given OM and EM
351 combination. We constructed 95% confidence intervals for the median relative bias using the
352 binomial distribution approach as in Miller and Hyun (2018) and Stock and Miller (2021).
353 We present results for bias in j . We analyzed simulation results for estimates of terminal year
354 estimates of SSB and recruitment, Beverton-Holt stock-recruit SRR parameters (a and b),
355 and median natural mortality rate. Results for terminal year fishing mortality were strongly
356 negatively correlated with those for SSB and are provided in the Supplementary Materials.

357 **Mohn's ρ**

358 Finally, we investigated presence of retrospective patterns in fitted models as a measure of
359 reliability. We calculated Mohn's ρ for SSB, fully-selected fishing mortality fishing mortality

360 (averaged over all age classes), and recruitment for each EM fit to each OM simulated data
361 set (Mohn 1999). We fit 7 peels for each EM and calculated median 95% confidence intervals
362 for $P = 7$ peels to each simulated data set and calculated Mohn's ρ using the same methods
363 as that for relative bias for a given attribute θ as

$$\rho(\theta) = \frac{1}{P} \sum_{p=1}^P \frac{\hat{\theta}_{Y-p,Y-p} - \hat{\theta}_{Y-p,Y}}{\hat{\theta}_{Y-p,Y}} \quad (2)$$

364 where Y is last year of the full set of observations and $\hat{\theta}_{y,y'}$ is the estimate for attribute θ in
365 year y from a model fit using data up to year $y' \geq y$.

366 **Results**

367 **Summarizing results across OM and EM attributes**

368 For many R and R+S OMs, convergence rate declined when either the median natural
369 mortality rate or the Beverton-Holt SRR was estimated even when the process error
370 assumptions of the EMs and OMs matched (Figure S17, A). When there was high
371 observation error and constant fishing pressure ($F = F_{MSY}$ for all 40 years), convergence was
372 poor for of all. Because the OM and EM attributes that we investigated are numerous, we
373 used two methods to summarize the most important factors for differences in results within
374 a given OM process error type. The first method was fitting regression models with the
375 response being each of the measures of reliability described above and predictor variables
376 were defined based on OM and EM characteristics (e.g., MacKinnon et al. 1995; Wang et al.
377 2017; Harwell et al. 2018). For the binary indicators of convergence and AIC-based selection
378 of a SRR, we performed logistic regressions. For indicators of AIC-based selection of EM
379 process error configurations other than R EMs when fitted to R OMs ($\sigma_{2+} = 0$) regardless
380 of whether median natural mortality and SRRs were estimated. Convergence of R EMs was
381 high for all R and R+S OMs except when there was high observation error and constant

382 fishing pressure, type (multiple categories) we performed multinomial regressions. For other
383 measures of reliability we fit linear regression models to transformed responses. Because
384 relative errors (Eq. 1) and when median natural mortality and SRRs were estimated. R+S
385 EMs fit to R OM~~s~~s exhibited poor convergence regardless of whether natural mortality or
386 a SRR was estimated. R+S EMs fit to R+S OM~~s~~s had highest convergence rates when
387 there was contrast in fishing pressure and low observation error Mohn's ρ for the various
388 parameters are bounded below at -1, we used a transformation of these values

$$y_j = \log [f(\hat{\theta}_j, \theta_j) + 1] \quad (3)$$

389 where f is either the relative error (Eq. 1) or Mohn's ρ (Eq. 2) for simulation j , so that
390 values are unbounded. For relative errors, y_j is the log-scale error. We omitted simulations
391 where estimated attributes equal to zero (RE = -1). For all regressions we fit separate models
392 with just individual OM and EM factors included, with all factors included, with all second
393 order interactions, and with all third order interactions. For the multinomial regression,
394 we used the `vglm` function from the VGAM package (Yee 2008; Yee 2015). We tabulated
395 percent reduction in residual deviance for each of regression fits. We did not perform formal
396 statistical analyses of effects of OM and EM attributes on results (e.g., ANOVA) because
397 of the lack of independence of the "observations" that results from fitting multiple EMs to
398 each simulated data set.

399 The second method involved fitting classification and regression trees (Breiman et al. 1984)
400 to show how the OM and EM attributes, and their interactions, split the values for each
401 measure of reliability (e.g., Gonzalez et al. 2018; Collier et al. 2022). We used classification
402 trees for categorical measures (convergence and AIC) and regression trees for the other
403 measures with continuous scales (relative error and Mohn's ρ). The response variables were
404 the same as the regressions for the deviance reduction analyses. We used the `rpart` function
405 in the `rpart` package (<https://cran.r-project.org/package=rpart>) to fit trees. Full trees were

406 determined using default settings except that we increased the number of cross-validations
407 to 100. For clarity, we pruned the full trees to show just the primary branches. Convergence
408 rates were high for all EMs when fit to data from R+S OM_s with lower observation error
409 except those where median natural mortality and/or SRRs were estimated.

410 Convergence of all EMs fitted to R+M OM_s was highest when the OM_s had higher natural
411 mortality process error variability, low observation error, and contrast in fishing pressure
412 (Figure S17, B). We also provide detailed results for all measures of reliability at each
413 combination of OM and EM attributes in the Supplementary Materials. For confidence
414 intervals of probability of convergence, we used the Clopper-Pearson exact method (Clopper
415 and Pearson 1934; Thulin 2014). For AIC selection of process error configuration we provide
416 estimates of the proportions of simulations where each EM type was selected. For AIC
417 selection of the SRR (a binary indicator for each simulated data set), we fit logistic regressions
418 and present resulting predicted probabilities of correctly selecting the SRR as a function
419 of SSB variability (log SD_{SSB} described above). We estimated bias as the median of the
420 relative errors across all simulations for a given OM and EM combination. We constructed
421 95% confidence intervals for the median relative bias, and Mohn's ρ using the binomial
422 distribution approach (Thompson 1936) as in Miller and Hyun (2018) and Stock and Miller
423 (2021).

424 Results

425 Convergence performance

426 For probability of convergence, EM process error assumption was the single attribute that
427 resulted in the largest percent reduction in deviance (14-28%) for all OM process error types
428 other than R+M EMs that estimated autocorrelation of process errors had poor convergence
429 for R+M OM_s when there was low natural mortality process error variability regardless of

430 autocorrelation of the simulated process errors. R+S EMs fitted to data generated from R+M
431 OMs always converged poorly whether or not OMs where the EM median natural mortal-
432 ity and the Beverton-Holt SRR were estimated. rate assumption (estimated or known)
433 explained the most residual deviance (>11%; Table 1). However, including interactions of
434 OM and EM factors also provided large reductions in residual deviance (35-47%), suggesting
435 successful convergence depended on a combination OM and EM attributes.

436 The RClassification trees for each OM process error type, all had the primary branch defined
437 using the same attribute that provided the largest reduction in deviance (Figure 1). EMs
438 that assumed R+S EMs, in particular, had poor convergence when fit to data generated
439 from R+Sel OMs with lower selectivity process error variability or higher observation error
440 (Figure S17, C). S process errors converged poorly for all OMs that were simulated with the
441 alternative process error assumptions (R, R+Sel EMs generally converged better than other
442 EMs for M, R+Sel OMs with higher process error variability, lower observation error, and
443 contrast Sel, an R+q OMs). For all trees, branches based on the OM fishing mortality history
444 showed better convergence when the OM included a change in fishing pressure regardless of
445 whether. Branches based on whether the Beverton-Holt SRR was assumed or not, showed
446 better convergence when it was not estimated and branches based on the median natural
447 mortality or a SRR was estimated.

448 For rate assumption showed better convergence when it was treated as known. For some
449 R+q OMs, convergence of RM and R+q EMs was generally better than that of other EMs
450 Sel OMs, better convergence was also observed when there was contrast in fishing pressure
451 (Figure (S17, D). Convergence of R+S EMs was generally worse than that of all other EMs
452 across all OMs whether or not median natural mortality or a SRR was estimated. Again,
453 convergence probability generally declined for all EMs when median natural mortality or a
454 SRR was estimated. lower observation uncertainty.

455 When convergence is defined by a gradient threshold, the primary factor explaining deviance

456 reduction is the same that for Hessian-based convergence for all OM process error types,
457 but there are some differences in deviance reduction for secondary factors (Table S6), and
458 probability of convergence, overall, was lower (Figure S2). We found a wide range of maxi-
459 mum absolute values of gradients for models that converged had invertible Hessians (Figure
460 S3). The largest value observed for a given EM and OM combination was typically $< 10^{-3}$,
461 but many converged models had values greater than 1. For many OMs, EMs that assumed
462 the correct process error type and did not estimate median natural mortality or the Beverton-
463 Holt SRR produced the lowest gradient values.

464 **AIC performance**

465 Marginal AIC accurately determined

466 **Process error structure**

467 For AIC selection of the correct process error assumptions in EMs when data were generated
468 from R and configuration, the magnitude of observation and process error variation were
469 the attributes that resulted in the largest percent reductions in deviance across OM process
470 error types other than R OMs (Table 2). Both sources of variation explained large reductions
471 in deviance for R+S OMs, regardless of whether median natural mortality or (17-22%) and
472 R+Sel (8-26%) OMs, whereas variance of process errors provided the major reductions for
473 R+M ($>9\%$) and R+q ($>13\%$) OMs. Comparatively, none of the OM or EM attributes
474 explained particularly large reductions in deviance for R OMs, but fishing history, whether
475 a SRR was estimated (Figure S4, A). Attempting to estimate , and whether median natu-
476 ral mortality or a SRR separately had a negligible effect on the accuracy of determining
477 the correct process error assumption. When both were estimated, there was a noticeable
478 reduction in accuracy when OM had a constant fishing pressure, low observation error, and
479 larger variability in recruitment process errors. was known or estimated provided similar and

480 the largest reductions (approximately 5-6%). Inclusion of second and third order interactions,
481 did not provide large reductions in deviance for any of the OM process error types.

482 For all OM process error types other than R OM_s, the attributes defining the primary
483 branches of classification trees matched those that provided the largest reductions in deviance
484 (Figure 2). Across all OM_s, AIC was more accurate for the process error type when process
485 error variability was greater and when observation error was lower. For R+M OM_s, marginal
486 AIC only accurately determined the correct process error model and correlation structure S₂.
487 OM_s, there was a tendency to select R OM_s when observation error was low and variability
488 in natural mortality process errors was high (Figures S4, B). Of these OM_s, estimating
489 the median natural mortality rate only reduced the accuracy of AIC when natural mortality
490 process errors were independent and fishing pressure was constant. For OM_s with poor model
491 selection accuracy, AIC most frequently selected EM_s with process errors in catchability
492 (higher and apparent survival variation was lower ($\sigma_{2+} = 0.25$), but accuracy for the process
493 error was otherwise highly accurate. Similarly, low accuracy for process error type occurred
494 for R+q or selectivity (R+Sel). Selection of Sel and R+S EM_s was generally unlikely.
495 Marginal AIC most accurately determined the correct source of process error and correlation
496 structure q OM_s only with high observation error and lower process error variability, but
497 high accuracy for R+Sel OM_s with M OM_s required both low observation error (Figures S4,
498 C). When there was low variability in selectivity process errors and high observation error,
499 R+q or R+S EM_s were more likely to have the best AIC. Whether median natural morality
500 or SRRs were estimated appeared to have little effect on the performance of AIC.

501 Marginal AIC most accurately determined the correct source of process error and high
502 process error variance. This was also required for accurate identification of the correct cor-
503 relation structure for R+q OM_s with high variability in catchability process errors (Figures
504 S4,D). The R-M, R+q, and R+q OM_s with low variability in catchability process errors and
505 high observation error had the least model selection accuracy. However, for these OM_s, the

506 marginal AIC accurately determined the correct source of process error (but not correlation
507 structure) except when OMs assumed a M OMs (Figure S4). No branches were estimated for
508 classification trees fit to the R OMs, likely because accuracy was high across all simulations
509 (0.94), although inspection of the fine-scale results shows there is some degradation in AIC
510 selection when a SRR and median natural mortality rate are estimated for R OMs with con-
511 stant fishing pressure and EMs estimated both median natural morality and the SRRhigh
512 observation error (Figure S4, top left).

513 AIC performance for the stock-recruit

514 Stock-recruit relationship

515 Our comparisons of model performance conditioned on assuming the true process error
516 configuration is known (EM and OM process error types match) and we focus on results where
517 the EMs assume median natural mortality is known because there was little difference in
518 results when the EMs estimated this parameter. Broadly, we found generally poor accuracy
519 of AIC in selecting models assuming a Logistic regressions for AIC selection of the Beverton-
520 Holt SRR over the null model without an SRR for all OMs SRR, showed OM fishing history
521 and log SDSSB provided substantial reductions in deviance for R+M (>13%), R+Sel (>26%),
522 and R+q (>24%) OMs (Table 3). For R OMs, fishing history provided the largest reduction
523 in deviance (>9%), whereas none of the attributes individually provided large reductions
524 in deviance for R+S OMs (all <5%). However, we also found increased accuracy of AIC
525 in determining the Beverton-Holt SRR when the simulated population exhibited greater
526 variation in spawning biomass for nearly every OM (Figure S19).

527 With inclusion of all attributes provided larger reductions in deviance than the sum of
528 individual contributions for both R (>30%) and R+S (~19%) OMs. Further fits for R and
529 R+S process error assumptions, probability of lowest AIC for the B-H SRR as a function
530 of SSB variability were greatest for OMs with contrast in fishing pressure and lower process

531 variability in recruitment (Figure S19, A). The largest variation in SSB occurred in OMs
532 with larger recruitment variability ($\sigma_R = 1.5$; Figure S19, A, right column group), but the
533 same high AIC accuracy was achieved for OMs with lower recruitment variability at lower
534 levels of SSB variation. The level of observation error had little effect on AIC accuracy OMs
535 including all combinations two factors showed those with log SD_{SSB} and recruitment variation
536 provided essentially the same reduction in deviance as the models with all factors. For all
537 OM process error types, inclusion of interaction terms provided relatively little reduction in
538 residual deviance.

539 For R+M OMs, probability of lowest AIC for the Beverton-Holt SRR increased steeply
540 with variation in SSB whether it was induced by Attributes defining the primary branches
541 of classification trees for AIC selection of the SRR assumption were the same as those
542 explaining the largest reductions in deviance for the logistic regression models (Figure 3).
543 All branches based on log SD_{SSB} showed better accuracy with larger variability in SSB and
544 all branches based on fishing history showed better accuracy when there was contrast in
545 fishing or variation in natural mortality process error. (Figure S19, B). There was little
546 difference in AIC accuracy whether the natural mortality process errors were correlated and
547 similar to pressure. Branched based on OM observation error or recruitment variability (R
548 and R+S OMs, there was also little effect due to level of observation error.

549 For R+Sel OMs,) showed better accuracy when they were lower. For R OMs, a combination
550 of lower recruitment variability contrast in fishing pressure over time was the primary source
551 of variation in SSB and these are the OMs where AIC accuracy for the Beverton-Holt SRR
552 was greatest (Figure S19, C). There was little effect of variability or correlation of selectivity
553 process errors or the level of observation error on AIC accuracy .

554 Like the , and higher SSB variability produced AIC accuracy over 0.8. For R+S OMs,
555 lower recruitment variability and observation error and higher SSB variability produced
556 AIC accuracy of 0.79. For R+Sel OMs, the greatest accuracy for AIC in selecting the

557 Beverton-Holt SRR occurred for M, R+Sel, and R+q OMs where there was contrast in
558 fishing pressure over time which is also where there was the greatest variation in SSB (Figure
559 S19,D). There was also little effect of variability or correlation of catchability process errors
560 or the level of observation error on AIC accuracy. OM's accuracy of 0.87 to 0.94 was observed
561 with just increased SSB variability.

562 Bias

563 Terminal year spawning stock biomass, fishing mortality, and recruitment

564 For R OMs ($\sigma_{2+} = 0$), there was no indication of bias (95% confidence intervals included
565 0) in terminal year SSB. Regression models for log-scale errors in SSB that included the
566 various OM and EM factors showed little reduction in deviance (<5%) for any of the
567 estimating models regardless of process error assumptions, except when no SR assumption
568 was made, recruitment variability was low, and there was contrast in fishing mortality and
569 high observation error (Figure S10, A). However, errors in terminal SSB estimates were
570 highly variable when factors across all OM process error types (Table 4). The attributes
571 producing the largest reductions were the EM assumption for median natural mortality was
572 estimated and there was constant fishing pressure and high observation error (Figure S10,
573 A, second row).

574 For (known or estimated) for R, R+S OMs, the EMs with matching process error
575 assumptions generally produced unbiased estimation of terminal SSB except when median
576 natural mortality was estimated and there was high observation error. In M, R+S OMs
577 with low observation error, EMs with incorrect process error assumptions typically provided
578 biased estimation of terminal year SSB. Estimating the Beverton-Holt SRR had little
579 discernible effect on bias of terminal year SSB estimation whereas estimating median M
580 tended to produce more variability in errors in terminal SSB estimation similar to ROMs.
581 For RSel, and R+M OMs with low variability in natural mortality process errors, low

582 observation error and contrast in fishing mortality over time all EMs produced low variability
583 in SSB estimation error that indicated unbiasedness (Figure S10, B, third row). However,
584 larger variability in natural mortality process errors increased bias of EMs without the
585 correct process error type. Estimating median natural mortality increased variability of
586 SSB estimation error particularly for OMs with high observation error and constant fishing
587 pressure over time. It also increased bias in SSB estimation for many R+q OMs (1-3%), EM
588 process error type for R+M OMs. Like Rand RS OMs (4%) and fishing history for all OM
589 types (1-5%). Including second order interactions provided the largest reductions in residual
590 deviance (10- 26%). Including third order interactions also provided further reductions for
591 R, R+SOMs, estimating a SRR had little discernible effect on SSB bias.

592 For , and R+Sel OMs , there was no evidence of bias for any EMs when variability in
593 selectivity process error and observation error was low, and q OMs between 5 and 11%.

594 In all regression trees, branches based on fishing history and level of observation error
595 generally showed less bias in SSB with contrast in fishing mortality (Figure S10, C). The
596 largest bias occurred for any EMs that estimated median natural mortality when the OMs
597 had high observation error , constant fishing pressure, and greater variability in selectivity
598 process errors ($\sigma_{Sel} = 0.5$) or low selectivity process errors ($\sigma_{Sel} = 0.1$) and low observation
599 error. However, there was no evidence of SSB bias for correctly specified R+Sel EMs when
600 observation error was low and variation in selectivity process errors was larger, whether
601 median natural mortality was estimated or not (Figure S10, C, third row). We only observed
602 an effect of estimating the Beverton-Holt SRR for R+Sel OMs that had high observation
603 error and contrast in fishing pressure where estimating the SRR produced less biased SSB
604 estimation for many EMs (Figure S10, C, top row).

605 All EMs fit to data from R+q OMs with low observation error and contrast in fishing
606 pressure exhibited little evidence of bias in terminal SSB estimation except for R+M EMs
607 when there was no AR1 correlation in catchability process errors (Figure S10, D). Many EMs

608 also performed well in and lower observation error (Figure 4). For scenarios where there was
609 bias, it was generally positive (over-estimation). For branches based on treatment of median
610 natural mortality rate, bias was generally less when it was known rather than estimated. For
611 some R+q OM_s with low observation error, but no contrast in fishing pressure. For Sel and
612 R+q OM_s with high observation error and contrast in fishing pressure, EMs that estimated
613 the Beverton-Holt SRR exhibited less SSB bias than those that did not. Estimating median
614 natural mortality in the EMs only resulted in much more variable SSB estimation errors
615 when there was no contrast in fishing pressure (Figure S10, D, first and third rows), less bias
616 in SSB was shown when the EM process error configuration was correct.

617 For all OM process error types, relative errors in terminal year recruitment were generally
618 more variable than SSB, but effects of R and R+S Results for bias in fishing mortality
619 and recruitment generally matched those for SSB, except that directions of bias for fishing
620 mortality were opposite to those for SSB and recruitment. Effects of individual OM and
621 EM attributes on bias (i.e., negative or positive or none) were similar (Figure S12, A).
622 Furthermore, for EM configurations where bias in terminal SSB was evident, median relative
623 errors in recruitment often indicated stronger bias in recruitment of the same sign. factors on
624 regression models were similarly small as measured by reduction in deviance (Tables S7 and
625 S8). Factors defining the primary branches of regression trees were in most cases identical
626 to those for SSB (Figures S5 and S6).

627 Stock-recruit parameters

628 Across all OM_s, there was generally less bias and (or)lower variability in estimation of the
629 Beverton-Holt *a* parameter than the Regression models for log-scale errors of estimates of
630 both the Beverton-Holt *a* and *b* parameter. In Rand Rparameters showed none of the factors
631 explained large percent reductions in deviance (Table 5). The OM fishing history provided
632 the largest deviance reduction for most OM process error types for both parameters, but

633 reductions were less than 5.6% except for R+S OM_s, EM_s with the correct assumptions
634 about process errors provided the least biased estimation of Beverton-Holt SRR parameters
635 when there was a change in fishing pressure over time and lower variability of recruitment
636 process errors, but there was little effect of estimating median natural mortality and a small
637 increase in bias for those OM_s that had high observation error (Figure S7, A). For other R
638 and R Sel OM_s where the reductions were 11.37% and 7.97% for the *a* and *b* parameters,
639 respectively and for just the *b* parameter for R+S OM_s, estimating natural mortality often
640 resulted in less biased estimation of SRR parameters. There was generally large variability
641 in relative errors of the SRR q OM_s (10%). The EM process error assumption provided
642 similar reductions in deviance for both parameters for R OM_s. Including interactions also
643 did not produce important reductions in deviance.

644 For regression trees of log-scale errors in Beverton-Holt *a* and *b* parameter estimates, but
645 the lowest variability occurred with low variability in recruitment and little or no variability
646 in survival process errors ($\sigma_{2+} \in \{0, 0.25\}$), and contrast in fishing pressure.

647 In R+M OM_s, the most accurate estimation of SRR parameters for all EM process error
648 assumptions occurred when there was a change in fishing pressure, greater variability in
649 natural mortality process errors, and lower observation error (Figure S7, B). Relative to the
650 R, less bias was indicated with contrast in OM fishing pressure for all branches in trees
651 for each OM process error type (Figures 5 and 6). For all branches based on recruitment
652 variability in trees for R and R+S OM_s, there was even less effect of estimating median
653 natural mortality on estimation bias for the SRR parameters.

654 Bias for SRR parameters was large and variability in relative errors was greatest for most
655 EM_s fit to R+Sel OM_s with constant fishing pressure (Figure S7, C). Less bias in parameter
656 estimation occurred for OM_s with a change less bias in both *a* and *b* was observed with less
657 recruitment variability. For R OM_s with contrast in fishing pressure and the best accuracy
658 occurred for those OM_s that had low observation error and more variable and uncorrelated

659 ~~selectivity process errors~~, and when the EMs had with the correct process error assumption.

660 There was little effect of estimating natural mortality on relative errors for SRR parameters.

661

662 Like ~~R~~ greater recruitment variability EMs that assumed the incorrect ~~R+Sel~~ OM~~s~~, relative
663 errors in SRR parameters for ~~R+q~~ OM~~s~~ were more accurate for most EM process error
664 types when OM~~s~~ had contrast in fishing pressure and lower observation error (Figure S7, D).
665 However, the best accuracy occurred for those OM~~s~~ that had ~~M~~ process errors produced less
666 bias in both *a* and *b* than other process error assumptions. Across all combinations of OM
667 and EM attributes, some bias was observed for both parameters, but there was generally
668 less bias and (or) lower variability in ~~catchability~~ process errors. The worst accuracy of SRR
669 parameter estimation regardless of EM type occurred when ~~R+q~~ OM~~s~~ had low observation
670 error and constant fishing pressure estimation of the *a* parameter than the *b* parameter
671 (Figure S7, D, fourth row).

672 Median natural mortality rate

673 Across all OM~~s~~ and EM~~s~~ there was little effect of estimating SRRs on the bias in estimation
674 of Fitted regression models for log-scale errors in median natural mortality (Figure S13).
675 Median natural mortality rate was estimated accurately by all rate showed largest percent
676 reductions in residual deviance for R+S and R+M models (Table 6). The largest reductions
677 for a single attribute was the EM process error types for all R OM~~s~~ except those with
678 high observation error and constant fishing pressure, in which case relative errors were high
679 (Figure S13, A, $\sigma_{2+} = 0$) . For assumption ($>20\%$) and fishing history ($>15\%$) for R+S
680 OM~~s~~ estimation of median natural mortality rate was most accurate when observation error
681 was low and there was contrast in fishing pressure and the EM process error type was correct.

682

683 For ~~R~~. Fishing history also provided $>10\%$ reduction for R+M OM~~s~~, median natural mortality

684 was estimated most accurately, regardless of EM process error type, when OMs had a change
685 in fishing pressure and low observation error (Figure S13, B). However, those but reductions
686 for all factors in R, R+M OMs that also had greatest variability in AR1 correlated natural
687 mortality process errors only had unbiased estimation when the EM process error type was
688 correct Sel, and R+q OMs were relatively low (<6%). Interactions of OM and EM factors
689 also provided substantial further reductions for R+S and R+M OMs (between 8 and 15%
690 for second order interactions).

691 All EM process error types accurately estimated Regression trees with branches based on
692 fishing history showed less bias in median natural mortality rate for R+Sel OMs that had
693 with contrast in fishing pressure, low observation error, and low selectivity process error
694 variability (Figure S13, C). When selectivity process error variability increased, the incorrect
695 EM process errors produce more biased estimation of median natural mortality rate. The
696 least accurate estimation occurred for all EM process error types when observation error was
697 high and fishing pressure was constant.

698 Like and branches based on level of observation error showed less bias with more precise
699 observations (Figure 7). For R OMs, branches based on EM process error assumption showed
700 less bias with EMs assuming the correct R and the incorrect R+S assumption. For R+Sel
701 OMs, all EM process error types produced accurate estimation of median natural mortality
702 rate when fit to S and R+q OMs with contrast in fishing pressure, low observation error and
703 low catchability process error variability (Figure S13, D). Most M OMs, branches based on
704 EM process error showed only the correct EM process error types produced biased estimation
705 of median natural mortality when R+q OMs had high observation error and constant fishing
706 pressure assumption with less bias.

707 Mohn's ρ

708 Regression models for Mohn's ρ for SSB was small in absolute value for all R and R+S OMs,

709 regardless of EM process error types, and whether median natural mortality rate or SRRs
710 were estimated (Figure S14, A) of SSB showed little reduction in deviance for any of the OM
711 an EM attributes (<2%; Table 7). The strongest retrospective patterns (highest absolute
712 lack of explanatory power is also reflected in the regression trees where median Mohn's ρ
713 values) occurred in OM with the largest apparent survival process error variability are near
714 zero unless a large combinations of OM and EM conditions occur (Figure 8). For example,
715 In R+S OMs, with constant fishing pressure, high observation error, and contrast in fishing
716 pressure, but only for EMs with the incorrect process error type and where median natural
717 mortality rate was assumed known (median ρ was approximately -0.15). For higher apparent
718 survival process error, EMs that assume R+M, R+Sel, and R+q OMs, process errors have
719 a median Mohn's $\rho = -0.068$.

720 Similarly, poor explanatory power of the OM and EM attributes occurred when we fit
721 regression models for Mohn's ρ was also small in absolute value, but median values were
722 all closer to 0 than the largest values in the R and R+S OMs (Figure S14, B-D). For these
723 OMs, there was no noticeable effect of estimation of median natural mortality rate or SRRs
724 on of fishing mortality and recruitment (Tables S9 and S10). Regression trees for Mohn's ρ
725 for any EM process error types.

726 of fishing mortality were similar to those for SSB in that median values of Mohn's ρ for
727 recruitment was small in absolute value for all R OMs with low variability in recruitment
728 process errors, regardless of EM process error type, and whether median natural mortality
729 rate or SRRs were estimated (Figure S16, A) were close to zero for most combinations of OM
730 and EM attributes (Figure S8). However, R and R+S OMs with greater recruitment process
731 variability and higher observation error had we observed median Mohn's ρ for recruitment
732 greater than zero for most EMs even when the EM process error type was correct. In R+S
733 OMs with lower 0.1 at branches much closer to the base of the trees with fewer interactions of
734 the OM and EM attributes (Figure S9). These branches with consistently large retrospective
735 patterns were typically defined by larger OM observation error, EMs with the correct process

736 ~~error type exhibited better median Mohn's ρ close to 0 than EMs with the incorrect process~~
737 ~~error type. For R+M, R+Sel, and R+q OM results for OM constant fishing pressure,~~
738 ~~or incorrect EM process error configuration. Comparing regression model and regression~~
739 ~~tree fits, attributes defining the primary branches for all regression trees of all Mohn's ρ for~~
740 ~~recruitment are similar to those for SSB, but the range in median values and variation in~~
741 ~~Mohn's ρ values for a given OM are generally larger for recruitment (Figure S16, B-D) values~~
742 ~~(SSB, fishing mortality, and recruitment) generally matched those that explained the largest~~
743 ~~reductions in deviance.~~

744 Discussion

745 Assessing convergence

746 Analyses of model convergence across simulations can be useful for understanding the util-
747 ity of alternative convergence criteria used in applications to real data for directing the
748 practitioner to more appropriate random effects configurations. It is common during the
749 assessment model fitting process to check that the maximum absolute gradient component
750 is less than some threshold prior to inspecting the Hessian of the optimized likelihood for
751 invertibility (Carvalho et al. 2021). However, there is no accepted standard for the gradient
752 threshold (e.g., Lee et al. 2011; Hurtado-Ferro et al. 2014; Rudd and Thorson 2018) and
753 some thresholds would exclude models that in fact have an invertible Hessian. We found the
754 Hessian at the optimized log-likelihood can often be invertible when the maximum absolute
755 gradient was much larger than what ~~would~~ might be perceived to be a sensible threshold.

756 Li et al. (2024) found that convergence rate could be a useful diagnostic especially for sepa-
757 rating the correct model from overly complex models. However, the criteria for convergence
758 used in their study may also lead to limited ability to distinguish the correct model from
759 overly simplistic models, a pattern that was also noted by Liljestrand et al. (2024) in which

760 one process error may absorb all sources of process error when the magnitude of other process
761 errors are low.

762 Often poor convergence ~~result occurs~~ when parameter estimates are at their bounds (Car-
763 valho et al. 2021), and this also applies to variance parameters for random effects with
764 state-space assessment models. Even when the Hessian is invertible, parameters that are
765 poorly informed will have extremely large variance estimates. This further inspection can
766 lead to a more appropriate and often more parsimonious model configuration where the
767 problematic parameters are not estimated. For example, process error variance parameters
768 that are estimated close to 0 indicates that the random effects are estimated to have little or
769 no variability and removing these process errors is warranted. Generally, our results suggest
770 we can expect lower probability of convergence of state-space assessment models when esti-
771 mating natural mortality or SRRs because of the difficulty distinguishing these parameters
772 from others being estimated in assessment model ~~with data that are typically available in~~
773 common scenarios where data quality is less than ideal. Our experiments did not aim to
774 emulate the practitioner decision process in developing model configurations (e.g. removing a
775 source of process error and refitting the model when process error variance parameters were
776 estimated close to 0). Evaluating the efficacy of such a decision process when applying EMs
777 might be important in closed loop simulations (e.g. MSE) aimed at quantifying management
778 performance.

779 A factor affecting the convergence criteria, particularly for maximum likelihood estimation
780 of models with random effects, is numerical accuracy. All optimizations performed in these
781 simulations are of the Laplace approximation of the marginal likelihood and, therefore, gra-
782 dients and Hessians are also with respect to this approximation (see TMB::sdreport in the
783 Template Model Builder package). Functionality within the Template Model Builder pack-
784 age exists (i.e., TMB::checkConsistency) to check the validity of the Laplace approximation
785 and the utility of this as a diagnostic for state-space assessment models should be explored
786 further. Furthermore, numerical methods are used to calculate and invert the Hessian for

787 variance estimation for models with random effects. ~~Along with our results, Our results,~~
788 ~~along with~~ the potential lack of accuracy imposed by these approximations, ~~suggests suggest~~
789 at least investigating whether the Hessian is positive definite when the calculated absolute
790 gradients are not terribly large (e.g, < 1).

791 Configuring process error

792 ~~Of the OM process error configurations we considered, we found AIC to be accurate for~~
793 ~~selecting models with process errors on recruitment and apparent survival (Rand We found~~
794 ~~accuracy of marginal AIC for process error type required only low observation error for~~
795 ~~R, R+S). Fitting models to other OMs rarely preferred R+S EMsSel, and Rand R+S~~
796 ~~EMs were nearly always selected for the matching OMs; a similar result was reported by~~
797 ~~Liljestrand et al. (2024). For other sources of process error , accuracy of AIC was improved~~
798 ~~when there was larger variability in the process errors and/or lower q OMs. R+M OMs~~
799 ~~further required higher process error variability, but this also improved accuracy for the~~
800 ~~other OM process errors types when there was higher observation error.~~

801 ~~Across all OM process error configurations, AIC performed poorly in identifying that the~~
802 ~~presence of the~~

803 Stock recruitment relationships

804 Variation in SSB was the most important factor for using marginal AIC to distinguish the
805 ~~the~~ Beverton-Holt SRR ~~in the OM unless there was contrast in fishing pressure possibly in~~
806 ~~combination with other factors such as lower variability in recruitment process errors (in R,~~
807 ~~For R+M, R+Sel, and R+S models) or greater variation in natural mortality process errors~~
808 ~~(for R + M OMs , Fig. S19) . As such, properly accounting for process error in natural~~
809 ~~mortality could be important (Li et al. 2024) when evaluating SRRs in state-space models.~~
810 ~~Curiously, we did not find a marked effect of the level of observation error on ability q OMs,~~

811 the SRR was accurately detected when the CV of SSB over the time series was at least 40 to
812 detect the SRR, but it is possible that AIC would perform better if observations have even
813 lower uncertainty than we considered 50% ($\log SD_{SSB} = -0.9$ to -0.7) regardless of any other
814 OM or EM attributes. Detection of the SRR for R and R+S OMs required lower recruitment
815 variability, but this lower level ($\sigma_R = 0.5$) was assumed for all of the other OM process error
816 types. Our results assumed that the EM process error configuration was correct, but this
817 may be a strong limitation given the ability of AIC to distinguish the process error type in
818 many scenarios.

819 Although we did not compare models with alternative SRRs (e.g., Ricker and Beverton-
820 Holt), we do not expect AIC to perform any better distinguishing between relationships and
821 may be more difficult than distinguishing from the null model even with larger variability
822 in SSB. Our finding that AIC tended to choose simpler recruitment models in ~~most~~many
823 cases contrasts with the noted bias in AIC for more complex models (Shibata 1976; Katz
824 1981; Kass and Raftery 1995), but, whereas those findings apply to the much more common
825 comparison of models that are fit to raw and independent observations, here we are comparing
826 state-space models which account for observation error and estimate process errors in latent
827 variables.

828 Our results comport with those of de Valpine and Hastings (2002) who found AIC could not
829 distinguish among state-space SRRs that were fit just to SSB and recruitment observations
830 (i.e., not an assessment model). Similarly, Britten et al. (In review) found AIC could not
831 reliably distinguish alternative environmental effects on SRR parameters. However, Miller
832 et al. (2016) did find AIC to prefer a SRR with environmental effects when applied to data
833 for the SNEMA yellowtail flounder stock and AIC also selected an environmental covariate
834 on a SRR for the most recent stock assessment of Georges Bank yellowtail flounder (NEFSC
835 2025). Both of these yellowtail flounder stocks have large changes in stock size and the
836 values of environmental covariates over time. Additionally, this species is well-observed by
837 the bottom trawl survey that is used for an index in assessment models.

838 As expected, bias in all parameters and assessment output was generally improved with lower
839 observation error. Estimation However, estimation of SRR parameters was only moderately
840 reliable in ideal scenarios of low observation error and contrast in fishing for some R+Sel and
841 R+M OMs, but generally with large temporal variability in process errors. Otherwise, SRR
842 parameter estimation was biased and(or) highly variable. We found substantial bias in esti-
843 mated SRR parameters in R and R+S OMs particularly with high variability in recruitment
844 and apparent survival process errors, suggesting that practitioners should be cautious SRR
845 inferences when fitted assessment models have these properties. We only evaluated effects
846 of SSB variability on accuracy of AIC in identifying the SRR, but those results suggests
847 we might find less bias for the SRR parameters in such cases as well. Similarly, restricting
848 results to fits that converged may also yield better accuracy of SRR parameter estimation.

849 On the other hand, estimation

850 Estimating assessment model quantities

851 As expected, bias in parameters, SSB, and other assessment output was generally improved
852 with lower observation error. Estimation of median natural mortality was reliable in many
853 OM scenarios with contrast in fishing pressure, consistent with Hoenig et al. (2025). In
854 some OMs, when EMs estimated the SRR parameters and median natural mortality, bias
855 for those parameters was improved. Conversely, for some R+Sel and R+q OMs where there
856 was bias in natural mortality due to high observation error, estimating the SRR reduced the
857 bias in median natural mortality rate. However, estimating median natural mortality did
858 cause However, we found poor accuracy in SSB estimation terminal SSB estimation when
859 estimating median natural mortality in many OMs without when there was no contrast in
860 fishing pressure over time and with higher observation error. Thus Therefore, estimating me-
861 dian natural mortality should be approached with caution in state-space assessment models,
862 particularly given its significant impact on determination of reference point and stock status

863 (Li et al. 2024).

864 **Negligible retrospective patterns**

865 Incorrect EM process error assumptions did not produce strong retrospective patterns for
866 SSB for any OMs regardless of whether median natural mortality or a SRR was estimated
867 ~~, but some weak retrospective patterns occur although some weak patterns occurred~~ when
868 observation error was high and there was contrast in fishing pressure. However, retrospective
869 patterns tended to be more variable for recruitment and were sometimes large even when
870 the EM was correct. Therefore, we recommend ~~emphasis de-emphasis~~ on inspection of
871 ~~retrospective patterns primarily for SSB and F patterns for recruitment~~, but further research
872 on retrospective patterns in other assessment model parameters, management quantities such
873 as biological reference points, and projections may be beneficial (Brooks and Legault 2016).

874 The general lack of retrospective patterns with mis-specified process errors is perhaps to be
875 expected. Retrospective patterns are often induced in simulation studies by rapid changes
876 in a quantity such as index catchability, natural mortality, or perceived catch during years
877 toward the end of the time series (Legault 2009; Miller and Legault 2017; Huynh et al. 2022;
878 Breivik et al. 2023). In our simulations, the process errors changing over time may have
879 trends in particular simulations, particularly when strong autocorrelation is imposed, but
880 the random effects have no trend on average across simulations. Szuwalski et al. (2018) and
881 ~~Li et al. 2024~~ also found relatively small retrospective patterns when the source of
882 mis-specification was temporal variation in demography attributes. Indeed, it is common for
883 the flexibility provided by temporal random effects to reduce retrospective patterns (Miller
884 et al. 2018; Stock et al. 2021; Stock and Miller 2021), though it does not necessarily
885 indicate a more accurate assessment model (Perretti et al. 2020; Li et al. 2024; Liljestrand
886 et al. 2024). Our results together with the existing literature seem to suggest that when
887 a strong retrospective pattern is observed in an assessment it is more likely to be due to a

888 mis-specification of a rapid shift in some model attribute rather than whether a particular
889 process is assumed to be randomly varying temporally.

890 **Summarization approach**

891 Our simulation study examined the importance of several factors for reliable inferences from
892 state-space age-structured assessment models. Contrast in fishing pressure was consistently
893 an important factor across all. We found the use of regression models and classification
894 and regression trees extremely useful in understanding the most important OM and EM
895 attributes explaining variation in the measures of reliability we examined. AIC accurately
896 distinguished models with process errors on recruitment only (R) or on recruitment and
897 apparent survival (R+S). Accuracy for other process error types required a strong signal
898 (high process variability) with low noise (low observation uncertainty). Therefore, we expect
899 practitioners will find R+S configurations to provide satisfactory diagnostics across a range
900 of life history and data quality scenarios. AIC generally performed poorly for selecting the
901 SRR, but performance was improved with across all simulations. The classification and
902 regression trees are generally a good tool for determining the values of the OM and EM
903 attributes that produce better or worse measures of reliability. However, determining the
904 combination of attributes that produce the best or worst measures of reliability can be
905 challenging using the trees alone. For example, in the regression tree for median natural
906 mortality rate estimates in R OMs (Figure 7), both of the first branches imply bias is
907 low regardless of OM fishing history, but when OM fishing pressure is constant, results
908 are much better when OM observation error is low (median RE about -6%) than when
909 OM observation error is high (median RE about 40%). The default pruning of the trees
910 can exclude these lower branches. However, inspection of deviance explained by various
911 regression models shows the ~9% reduction in residual deviance by including second order
912 interaction of all OM and EM factors (Table 6), indicating that the interaction of factors may

913 be important, thereby complimenting the regression tree analysis. Higher order interactions
914 of some factors could also provide reductions in deviance and, therefore, inspection of results
915 for each combinations of OM and EM factors, as provided in the Supplementary Materials,
916 can also be important.

917 **Recommendations and conclusions**

918 Our findings regarding model convergence suggests practitioners using state-space models
919 and maximum marginal likelihood for estimation should not heavily weight the magnitude
920 of the gradient values in determining convergence as long as the maximum absolute values
921 is around 1 or lower. Instead, positive-definiteness of the Hessian of the minimized negative
922 log-likelihood should be evaluated.

923 Unfortunately, whether the practitioner includes a Beverton-Holt SRR will often depend
924 on biological plausibility of this particular SRR because using AIC to determine its validity
925 required a combination of low recruitment variability and contrast in fishing pressure. Some
926 , large variation in SSB over time, and lower observation error, which applies to a limited
927 number of managed stocks. Furthermore, some bias in estimation in at least one of the SRR
928 parameters existed in nearly all OM-EM combinations should be expected, which presumably
929 also applies to MSY-based reference points. Because bias in terminal SSB and retrospec-
930 tive patterns were indifferent to whether or not the SRR was estimated, and convergence
931 was slightly better the prevalence of bias in SRR parameter estimation, and often better
932 convergence without the SRR, we recommend a sensible default would be to fit models
933 without an assumed SRR is to exclude a SRR when fitting assessment models, as also
934 suggested by Brooks (2024).

935 We found marginal AIC can, in many cases, accurately distinguished models with process
936 errors. We saw the best accuracy for models with process errors on recruitment only
937 (R), recruitment and apparent survival (R+S), and recruitment and selectivity (R+Sel),

938 especially with lower observation error. However, AIC could also distinguish R+M and R+q
939 process errors when variability of those processes was greater. The R+S assumption for
940 process errors is common in applications of WHAM in the NEUS and the SAM assessment
941 framework (Nielsen and Berg 2014) in ICES, and we can have some confidence that
942 practitioners are correctly arriving at this assumption over other sources of process error
943 using marginal AIC.

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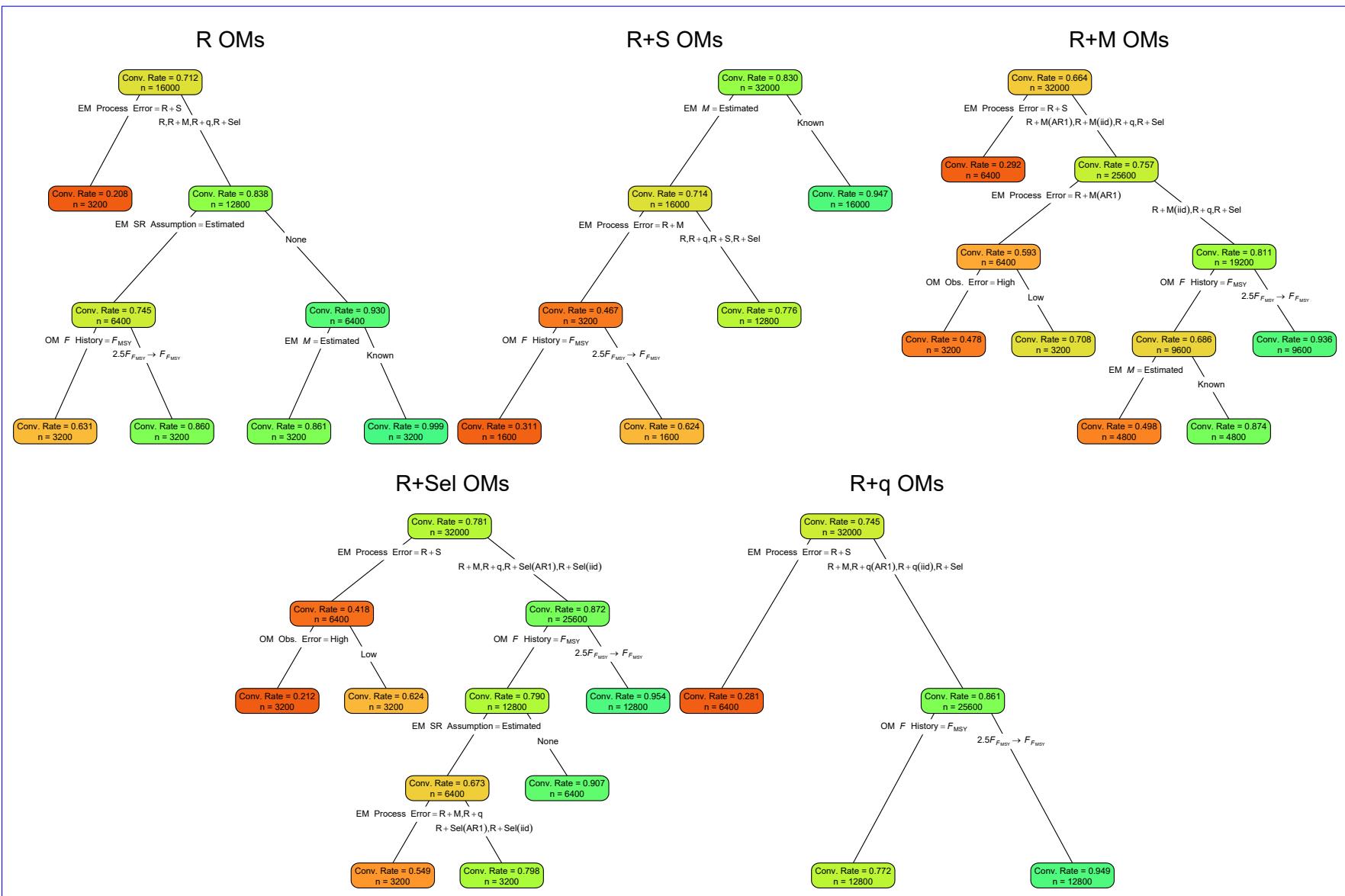


Fig. 1. Estimated probability of fits Classification trees indicating primary factors determining convergence as defined by providing hessian-based standard errors for EMs assuming alternative process error (colored points and lines) R, and median natural mortality (estimated or known) and Beverton-Holt stock-recruit relationships (estimated or not; along x-axis) when fitted to operating models that have Rand R+S(A), R+Sel(B), R+M(C), or Sel and R+q(D) process error structures(OMs). Circled values indicate results where the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals. Lower or higher convergence rates are indicated by more red or green polygons.

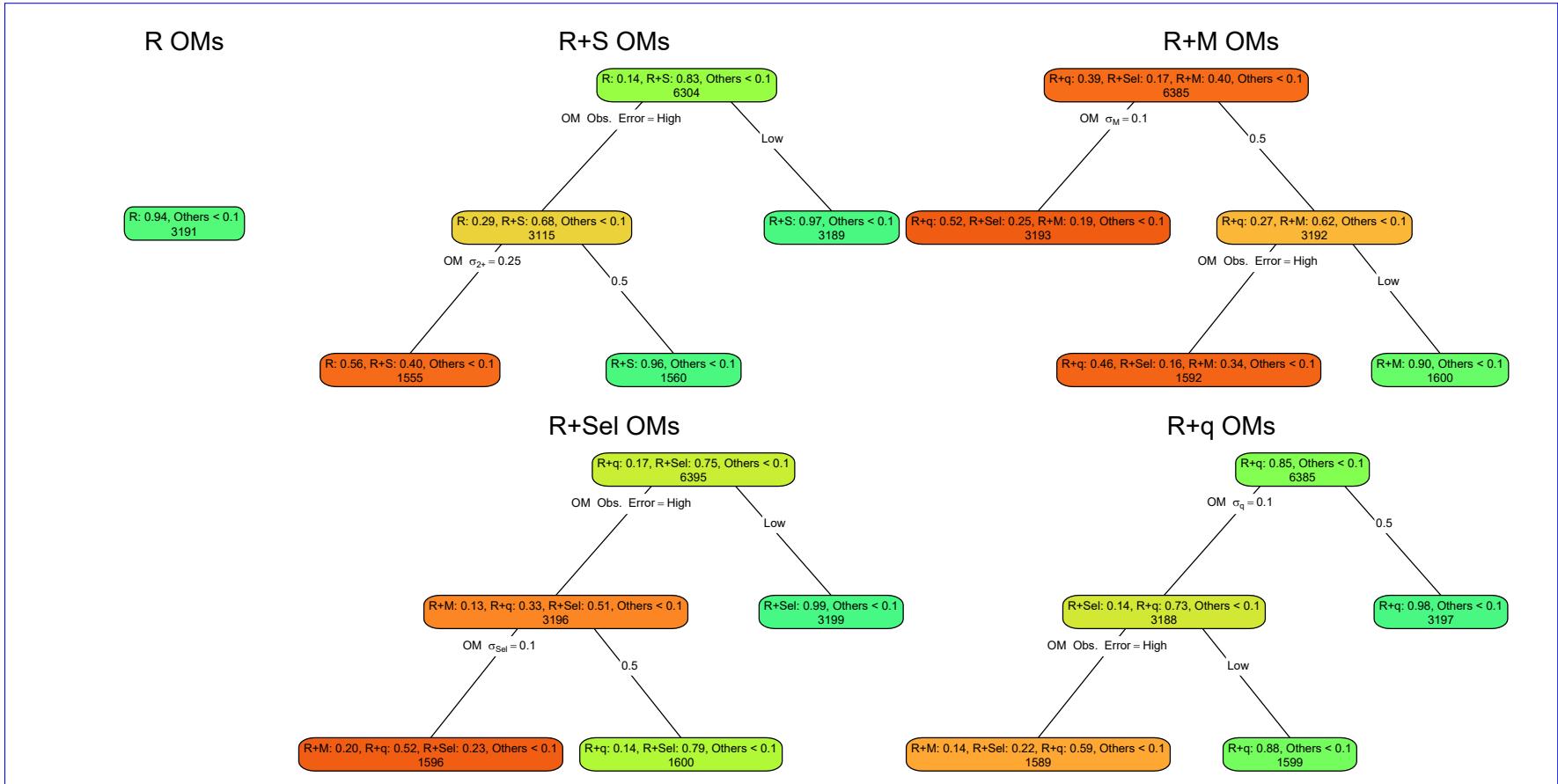


Fig. 2. Classification trees indicating primary factors determining which EM process error assumption provides the lowest AIC for R+S, R+M, R+Sel and R+q OM. Each node shows the proportion of EM process error models with lowest AIC (top) and number of observations (bottom) for the corresponding subset. Lower or higher accuracy of the process error assumption are indicated by more red or green polygons, respectively.

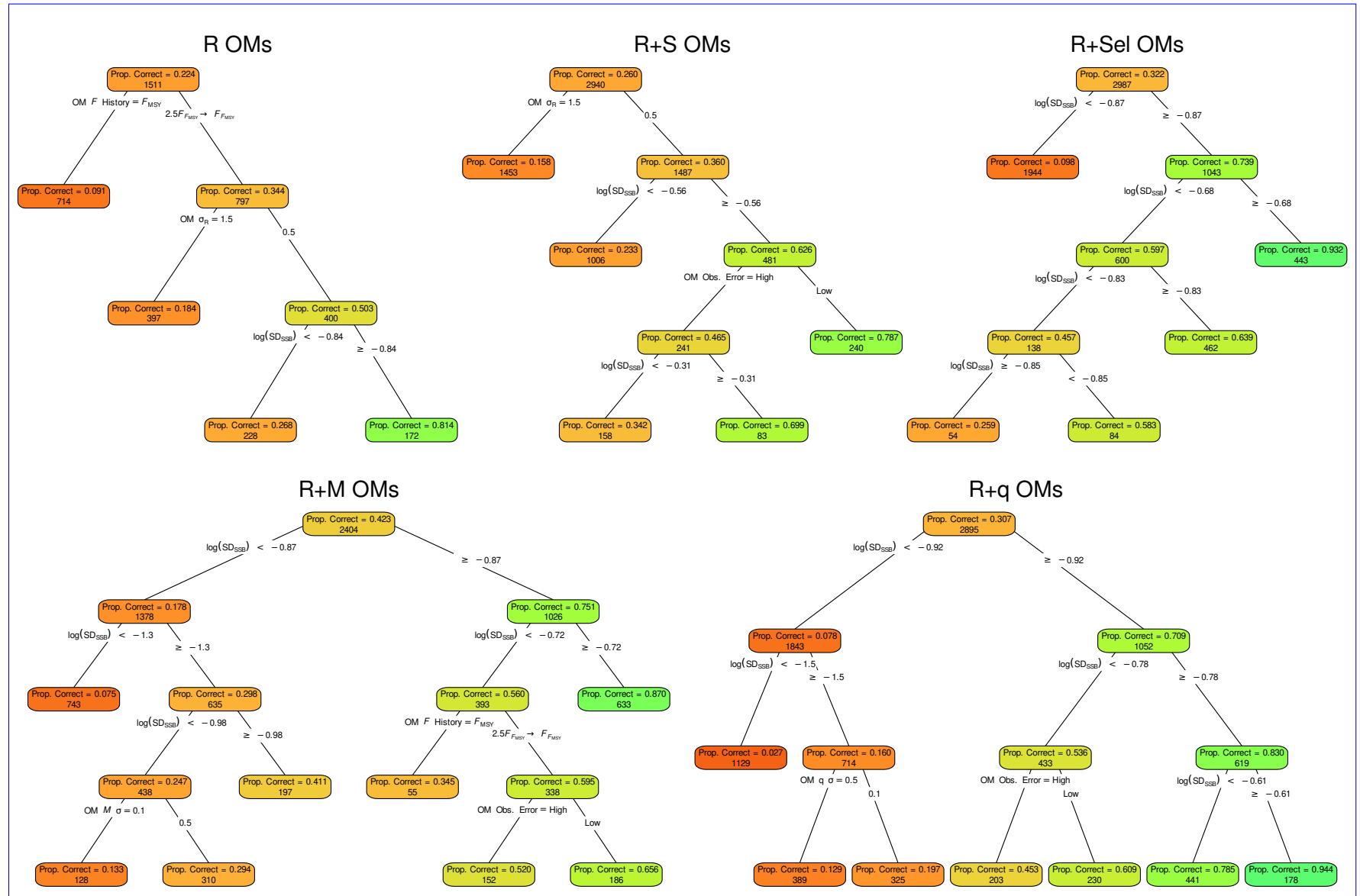


Fig. 3. Estimated probability of lowest AIC for EMs assuming alternative process error structures. Classification trees indicating primary factors determining which EM SRR assumption (colored bars) conditional on alternative assumptions for median natural mortality (estimated none or known) and Beverton-Holt stock-recruit relationships (estimated or not; along x-axis) when fitted to operating models that have provides the lowest AIC for Rand, R+S(A), R+Sel (B)M, R+M (C), or Sel and R+q OM. Each node shows the proportion of EMs that assume the SRR with lowest AIC (Dtop) process-error structures and number of observations (bottom) for the corresponding subset. Striped bars indicate results where Lower or higher accuracy of

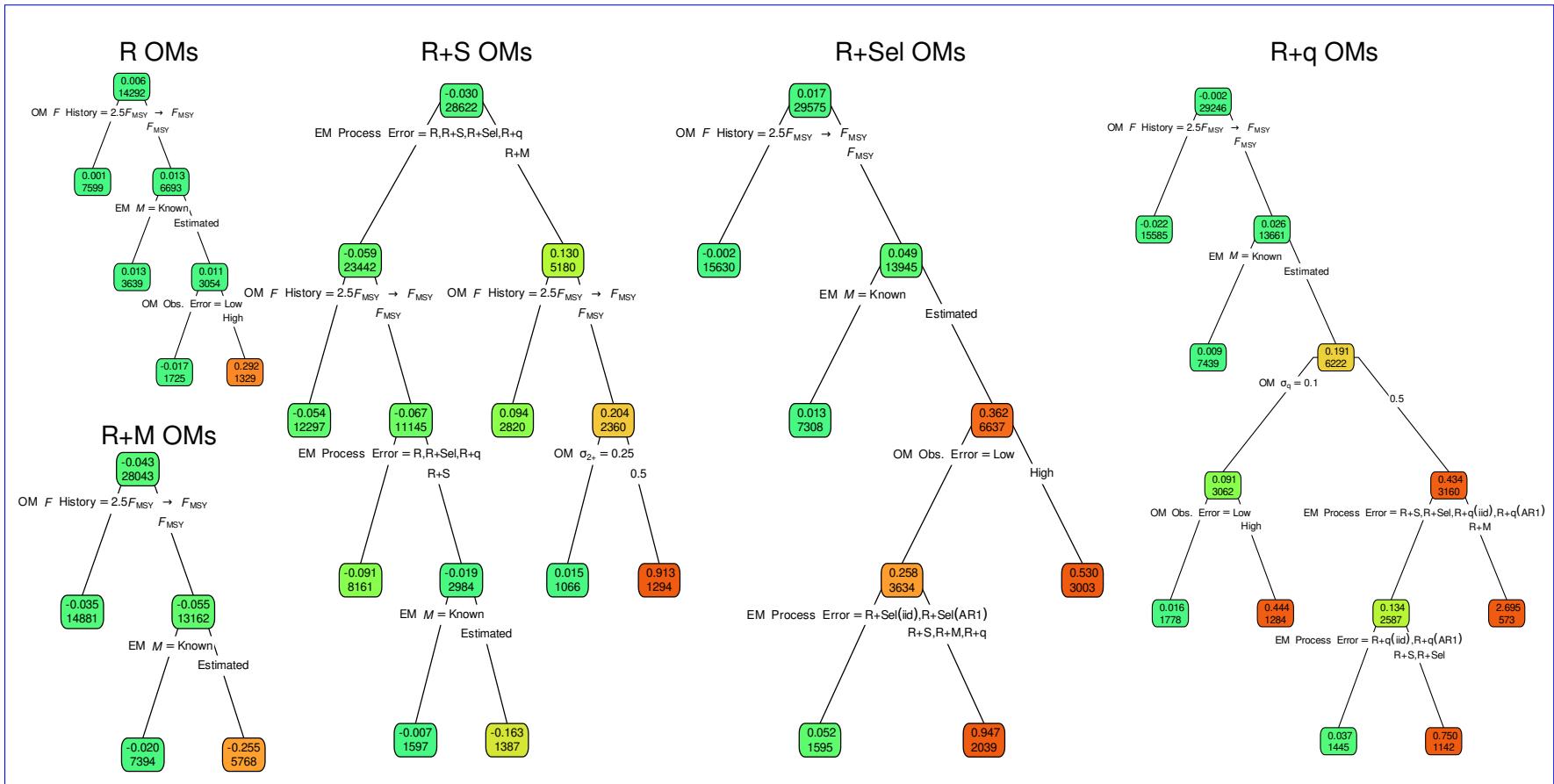


Fig. 4. Regression trees indicating primary factors determining reductions in sums of squares of errors in estimation measured by Eq. 3 for terminal year SSB for R+S, R+M, R+Sel and R+q OMs. Each node shows the median error (top) and number of observations (bottom) for the corresponding subset. Median errors closer to or further from zero are indicated by more green or red polygons, respectively.

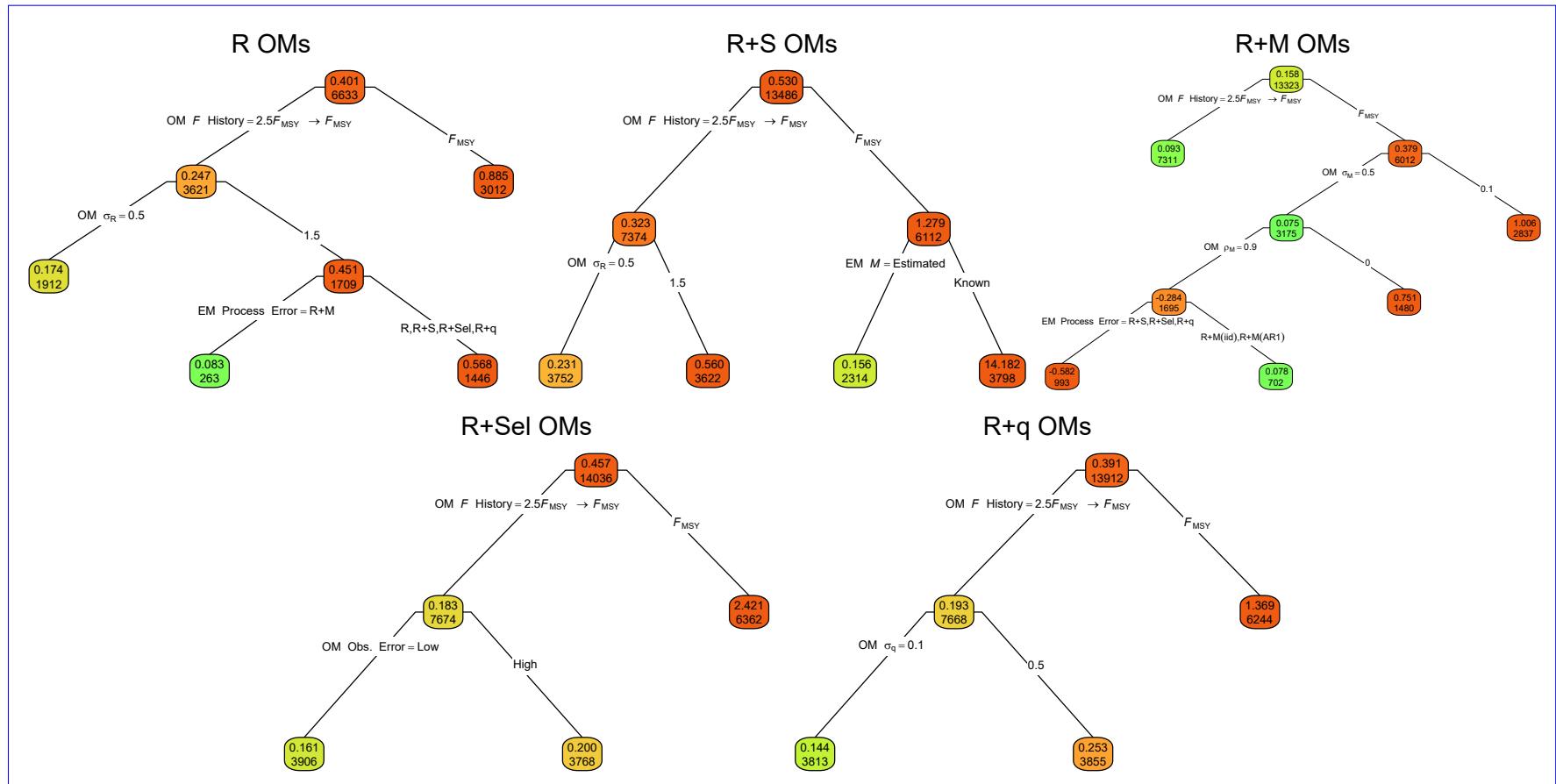


Fig. 5. Estimated probability Regression trees indicating primary factors determining reductions in sums of lowest AIC from logistic regression on the log-standard deviation squares of the true log(SSB) errors in each simulation estimation measured by Eq. 3 for estimating model with Beverton-Holt stock-recruit relationships, rather than the otherwise equivalent EM without the stock-recruit relationship. Results are conditional on median M is known in the EM and alternative assumptions EMs having the correct process error structure: Beverton-Holt SRR parameter a for R and R+S(A), R+Sel(B), R+M(C), or Sel and R+q OMs. Each node shows the median error (D_{top}), and median M is assumed known in the EM. Solid and dashed lines are for OMs with and without temporal contrast in fishing pressure, respectively, and polygons represent 95% confidence intervals. Range number of results indicates the range of log-standard deviation of logobservations (SSBbottom) for simulations the corresponding subset. Lower or higher median absolute errors of the particular OM process error assumption are indicated by more green or red polygons, respectively.

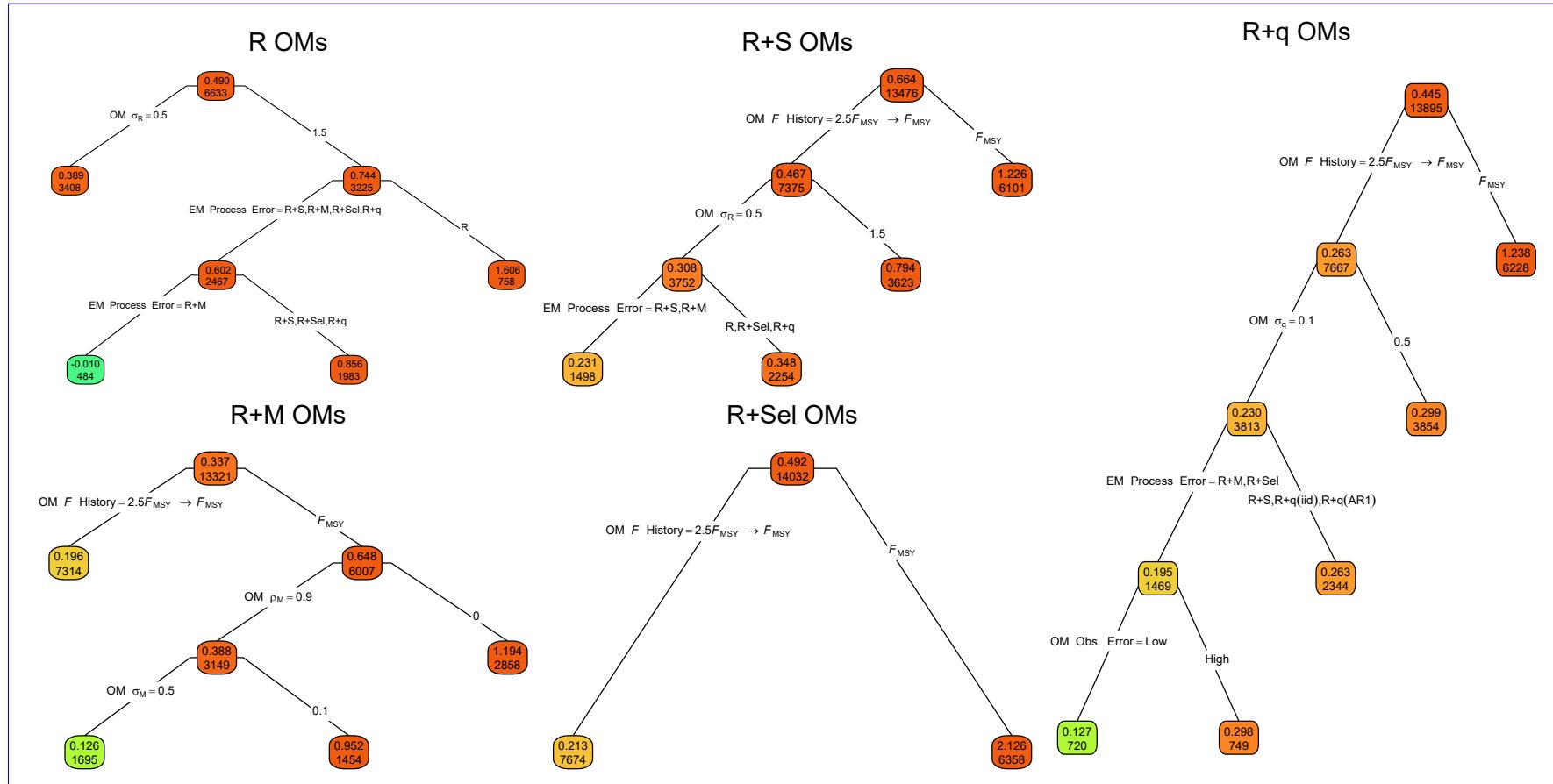


Fig. 6. **Median relative error** Regression trees indicating primary factors determining reductions in sums of terminal year SSB squares of errors in estimation measured by Eq. 3 for estimating models fitted to data sets simulated with alternative process error structures: the Beverton-Holt SRR parameter b for R and R+S(A), R+Sel(B), M, R+M(C), or Sel and R+q OMs. Each node shows the median error (D_{top}) and number of observations (bottom) for the corresponding subset. Circled values indicate results where Lower or higher median absolute errors of the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals assumption are indicated by more green or red polygons, respectively.

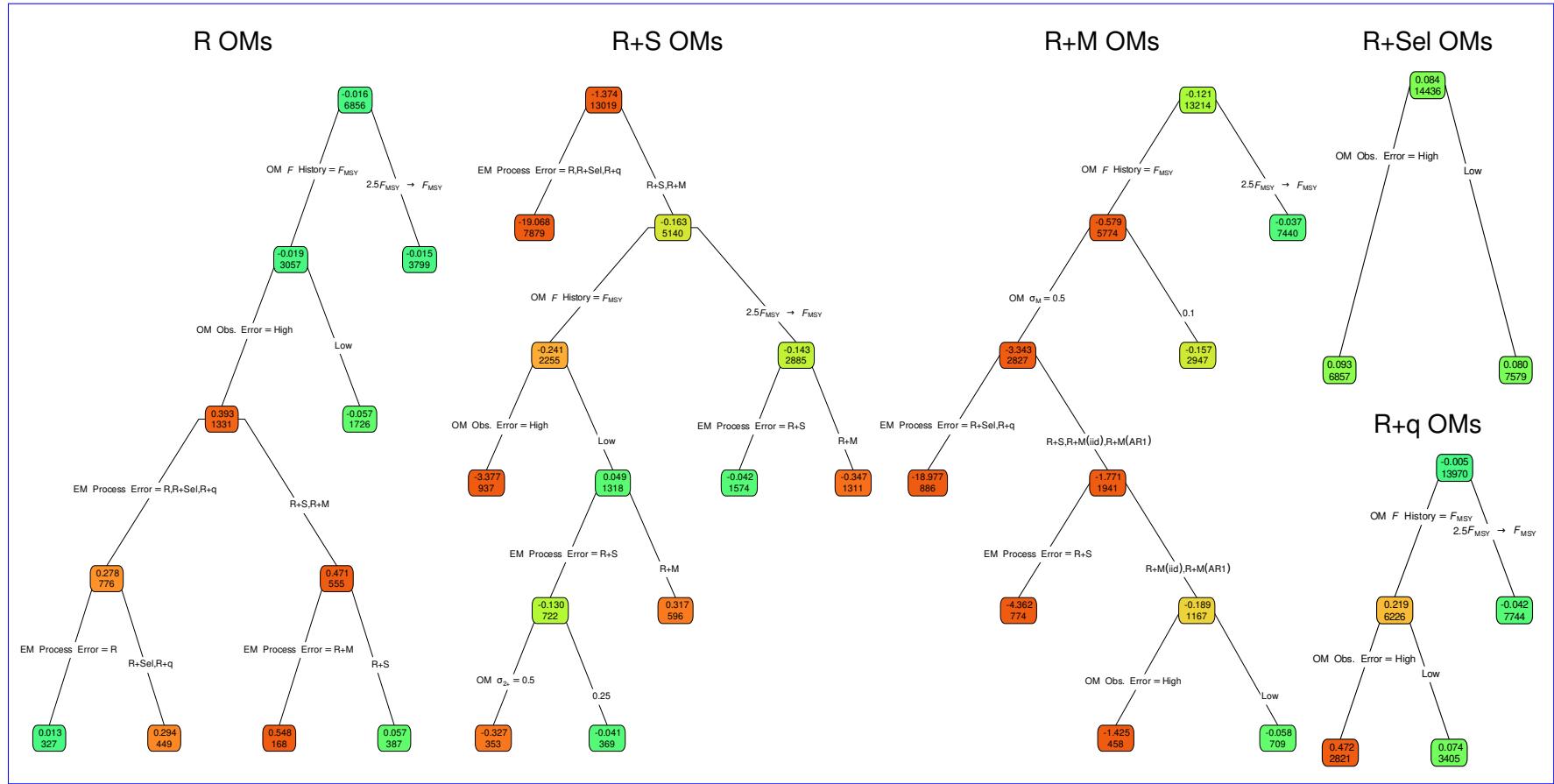


Fig. 7. ~~Median Mohn's rho~~ Regression trees indicating primary factors determining reductions in sums of squares of errors in estimation measured by Eq. 3 for ~~SSB~~ the median natural mortality rate for ~~estimating models fitted to data sets simulated with alternative process error structures: Rand R+S(A), R+Sel (B)M, R+M (C), or Sel and R+q OMs. Each node shows the median error (D_{top}) and number of observations (bottom) for the corresponding subset. Circled values indicate results where Lower or higher median absolute errors of the EM process error structure matches that of the operating model and vertical lines represent 95% confidence intervals assumption are indicated by more green or red polygons, respectively.~~

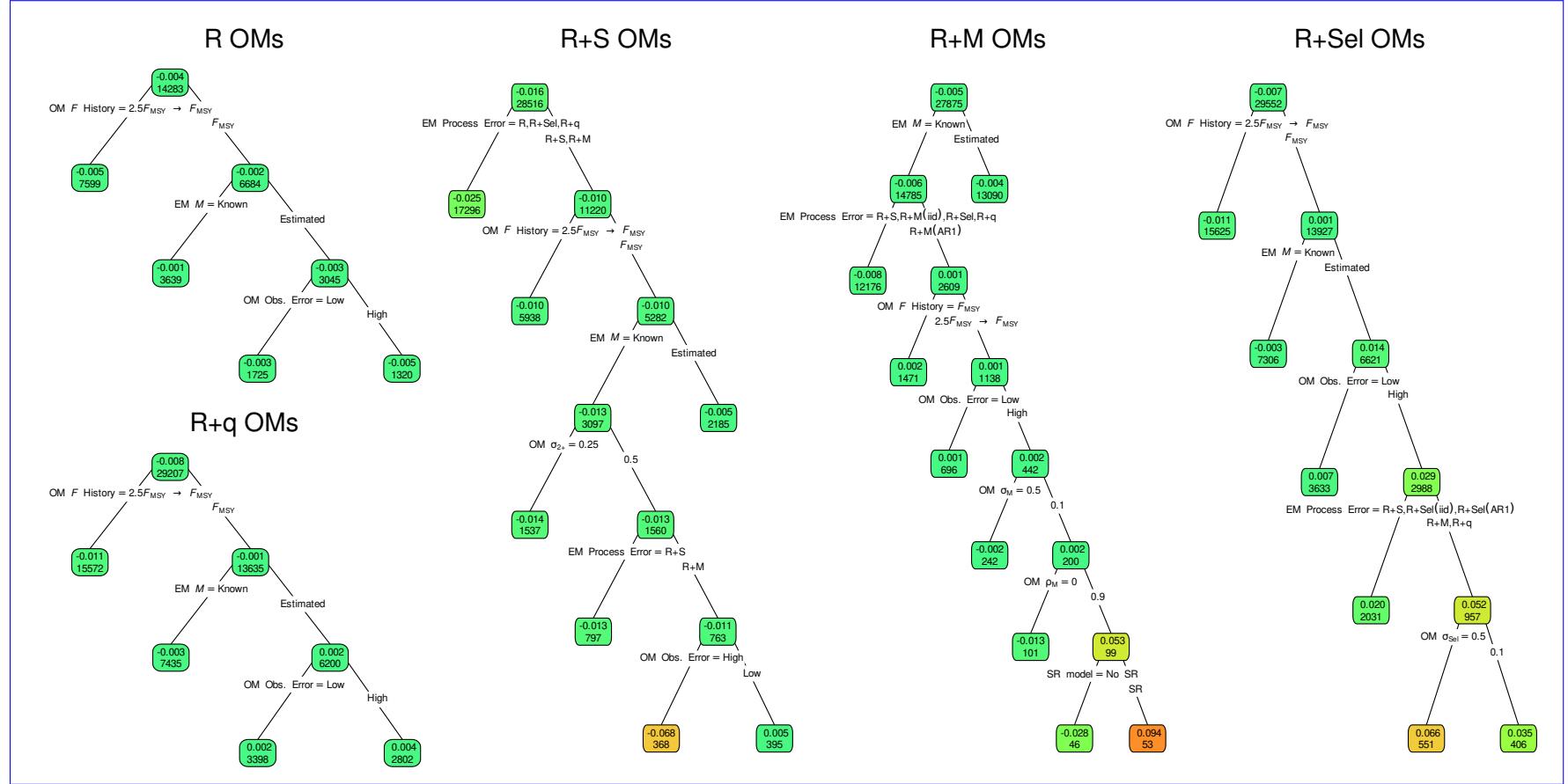


Fig. 8. Regression trees indicating primary factors determining reductions in sums of squares of errors in transformed Mohn's ρ (Eq. 3) for SSB for R+S, R+M, R+Sel and R+q OMs. Each node shows the median Mohn's ρ (top) and number of observations (bottom) for the corresponding subset. Median Mohn's ρ closer to or further from zero are indicated by more green or red polygons, respectively.

Table 1. For each OM process error type (columns), percent reduction in deviance for logistic regression models fit to indicators of convergence (providing Hessian-based standard errors) with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM Process Error	27.95	4.58	14.68	17.24	24.66
EM M Assumption	1.07	11.43	2.45	0.56	1.46
EM SR Assumption	2.88	3.30	1.24	2.47	1.59
OM Obs. Error	0.75	4.64	2.06	4.54	1.60
OM F History	2.32	3.37	1.63	3.30	2.59
OM σ_R	0.10	0.02	—	—	—
OM σ_{2+}	—	0.40	—	—	—
OM σ_M	—	—	0.22	—	—
OM ρ_M	—	—	0.17	—	—
OM σ_{Sel}	—	—	—	1.81	—
OM ρ_{Sel}	—	—	—	0.02	—
OM σ_q	—	—	—	—	0.34
OM ρ_q	—	—	—	—	<0.01
All factors	39.54	31.46	24.85	34.83	36.31
± All Two Way	45.03	39.89	35.20	42.81	43.70
± All Three Way	47.02	44.57	37.88	45.51	46.87

Table 2. For each OM process error type (columns), percent reduction in deviance for multinomial logistic regression models fit to indicators of EM process error assumption with lowest AIC with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM M Assumption	5.52	1.05	0.52	0.61	1.32
EM SR Assumption	5.60	0.75	1.13	0.93	1.95
OM Obs. Error	2.96	22.46	3.42	25.67	5.03
OM F History	5.77	0.62	0.94	0.91	2.05
OM σ_R	0.10	0.66	—	—	—
OM σ_{2+}	—	16.86	—	—	—
OM σ_M	—	—	9.06	—	—
OM ρ_M	—	—	0.38	—	—
OM σ_{Sel}	—	—	—	7.59	—
OM ρ_{Sel}	—	—	—	0.60	—
OM σ_q	—	—	—	—	13.50
OM ρ_q	—	—	—	—	0.75
All factors	20.98	46.12	16.58	40.83	25.99
± All Two Way	22.02	48.94	21.63	44.08	30.17
± All Three Way	22.05	49.98	22.36	44.54	31.38

Table 3. For each OM process error type (columns), percent reduction in deviance for logistic regression models fit to indicators of EM SRR assumption (none or Beverton-Holt) with lowest AIC with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM M Assumption	0.04	0.21	0.18	0.02	0.01
OM Obs. Error	<0.01	0.65	0.14	0.04	0.02
OM F History	9.17	3.79	13.08	26.56	24.60
OM σ_R	3.54	4.74	—	—	—
OM σ_{2+}	—	0.14	—	—	—
OM σ_M	—	—	1.14	—	—
OM ρ_M	—	—	0.05	—	—
OM σ_{Sel}	—	—	—	0.02	—
OM ρ_{Sel}	—	—	—	0.17	—
OM σ_q	—	—	—	—	0.36
OM ρ_q	—	—	—	—	0.02
log(SD _{SSB})	4.11	1.59	33.39	41.36	39.23
All factors	31.52	18.99	34.23	43.77	42.31
± All Two Way	34.79	22.24	35.99	45.84	44.04
± All Three Way	35.41	23.09	37.57	46.39	44.63

Table 4. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the terminal year SSB with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM M Assumption	2.28	1.15	1.04	2.92	3.26
EM SR assumption	0.10	0.06	0.08	0.06	0.08
EM Process Error	0.43	4.28	0.40	0.11	1.05
OM Obs. Error	1.63	0.07	0.78	0.32	<0.01
OM F History	2.62	3.15	1.28	3.22	4.72
OM σ_R	0.03	0.01	—	—	—
OM σ_{2+}	—	0.93	—	—	—
OM σ_M	—	—	0.18	—	—
OM ρ_M	—	—	0.01	—	—
OM σ_{Sel}	—	—	—	0.16	—
OM ρ_{Sel}	—	—	—	0.04	—
OM σ_q	—	—	—	—	1.02
OM ρ_q	—	—	—	—	0.06
All factors	7.59	9.86	3.93	7.04	10.64
± All Two Way	17.99	25.56	10.06	13.44	22.43
± All Three Way	23.39	36.74	13.76	16.55	31.11

Table 5. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the Beverton-Holt SRR parameters with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	Beverton-Holt <i>a</i>					Beverton-Holt <i>b</i>				
	R	R+S	R+M	R+Sel	R+q	R	R+S	R+M	R+Sel	R+q
EM M Assumption	0.02	1.05	0.02	0.11	0.02	0.05	1.06	0.03	0.01	0.40
EM Process Error	2.74	0.18	0.20	1.25	1.90	2.29	1.21	0.12	1.40	3.06
OM Obs. Error	0.16	<0.01	0.01	0.04	<0.01	<0.01	0.01	0.05	0.01	0.01
OM <i>F</i> History	3.15	3.34	5.60	11.37	10.00	1.16	1.17	2.01	7.97	3.87
OM σ_R	2.31	0.74	—	—	—	1.67	0.52	—	—	—
OM σ_{2+}	—	0.29	—	—	—	—	0.01	—	—	—
OM σ_M	—	—	0.30	—	—	—	—	0.13	—	—
OM ρ_M	—	—	0.51	—	—	—	—	0.22	—	—
OM σ_{Sel}	—	—	—	0.13	—	—	—	—	0.05	—
OM ρ_{Sel}	—	—	—	0.07	—	—	—	—	0.04	—
OM σ_q	—	—	—	—	0.04	—	—	—	—	0.10
OM ρ_q	—	—	—	—	<0.01	—	—	—	—	<0.01
All factors	8.07	5.15	6.73	12.64	11.79	4.91	3.75	2.55	9.12	7.22
± All Two Way	9.96	7.37	9.76	13.59	13.65	7.55	7.15	4.32	10.08	12.16
± All Three Way	11.22	8.15	11.13	14.48	14.87	9.78	9.02	5.26	11.08	14.73

Table 6. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the median natural mortality rate parameter with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM SR assumption	0.21	0.38	0.11	0.26	0.43
EM Process Error	1.98	20.36	3.16	0.94	1.31
OM Obs. Error	4.74	0.79	0.40	2.23	1.88
OM F History	5.07	15.11	10.65	0.24	2.38
OM σ_R	≤ 0.01	0.01	—	—	—
OM σ_{2+}	—	5.04	—	—	—
OM σ_M	—	—	5.32	—	—
OM ρ_M	—	—	0.85	—	—
OM σ_{Sel}	—	—	—	1.30	—
OM ρ_{Sel}	—	—	—	0.37	—
OM σ_q	—	—	—	—	0.46
OM ρ_q	—	—	—	—	0.06
All factors	12.64	40.10	21.29	5.54	6.52
± All Two Way	21.17	48.12	36.19	9.87	11.71
± All Three Way	23.03	50.38	42.82	11.58	14.64

Table 7. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to transformed Mohn's ρ values for each simulation (Eq. 3) for SSB with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM M Assumption	0.79	0.18	0.15	0.95	1.24
EM SR assumption	<0.01	0.01	<0.01	<0.01	<0.01
EM Process Error	<0.01	0.22	0.14	0.08	0.04
OM Obs. Error	0.12	0.03	0.05	0.18	0.21
OM F History	0.84	0.14	0.07	1.08	1.56
OM σ_R	0.01	0.01	—	—	—
OM σ_{2+}	—	0.02	—	—	—
OM σ_M	—	—	0.01	—	—
OM ρ_M	—	—	<0.01	—	—
OM σ_{Sel}	—	—	—	0.01	—
OM ρ_{Sel}	—	—	—	0.02	—
OM σ_q	—	—	—	—	0.01
OM ρ_q	—	—	—	—	0.01
All factors	1.89	0.63	0.43	2.43	3.29
± All Two Way	3.63	1.10	0.91	4.75	6.22
± All Three Way	4.27	1.65	1.50	5.73	7.53

1170 **Supplementary Materials**

1171 Referenced Figures

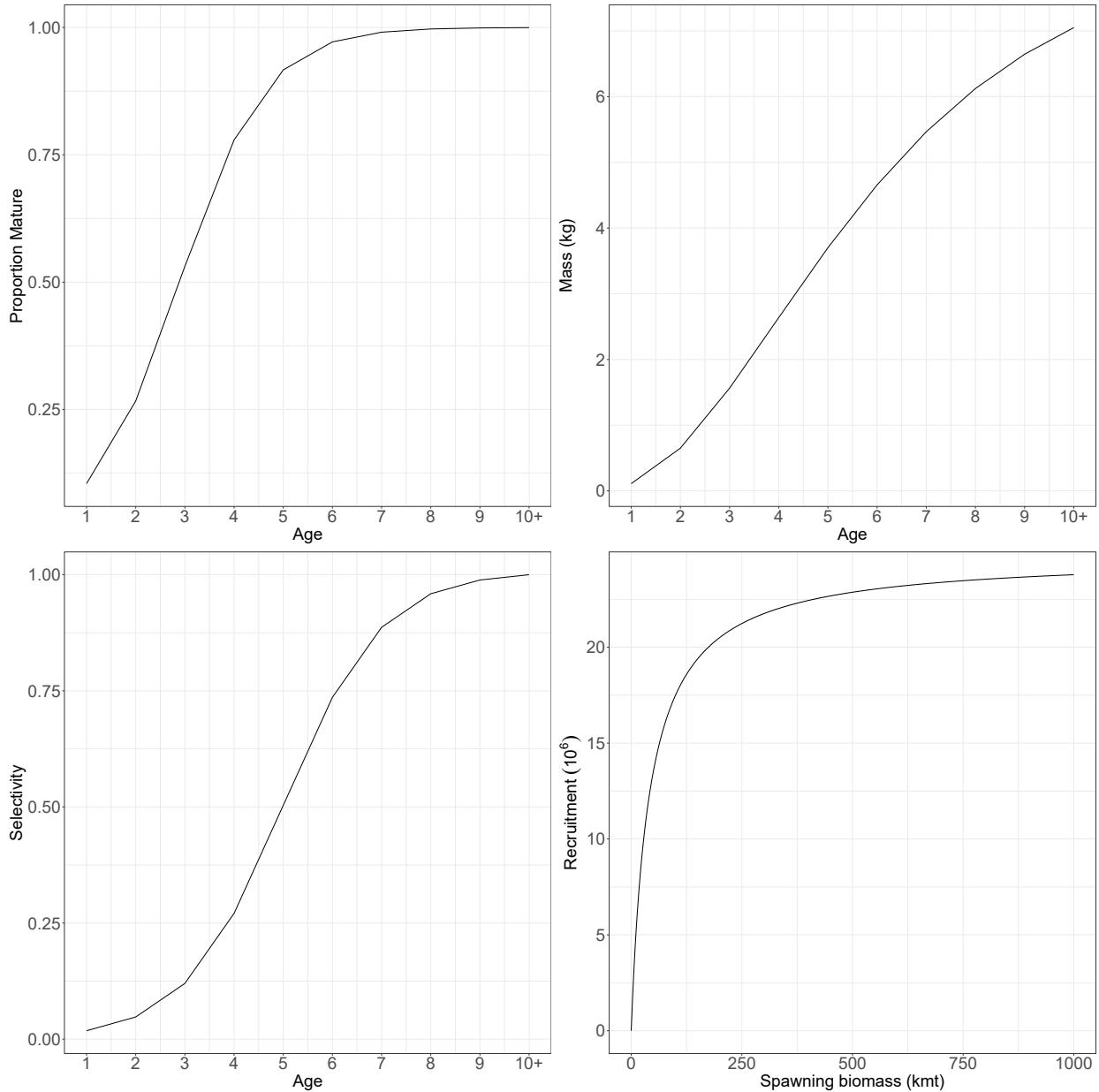


Fig. S1. The proportion mature at age, weight at age, fleet and index selectivity at age, and Beverton-Holt stock-recruit relationship-SRR assumed for the population in all operating modelsOMs. For operating models-OMs with random effects on fleet selectivity, this represents the selectivity at the mean of the random effects.

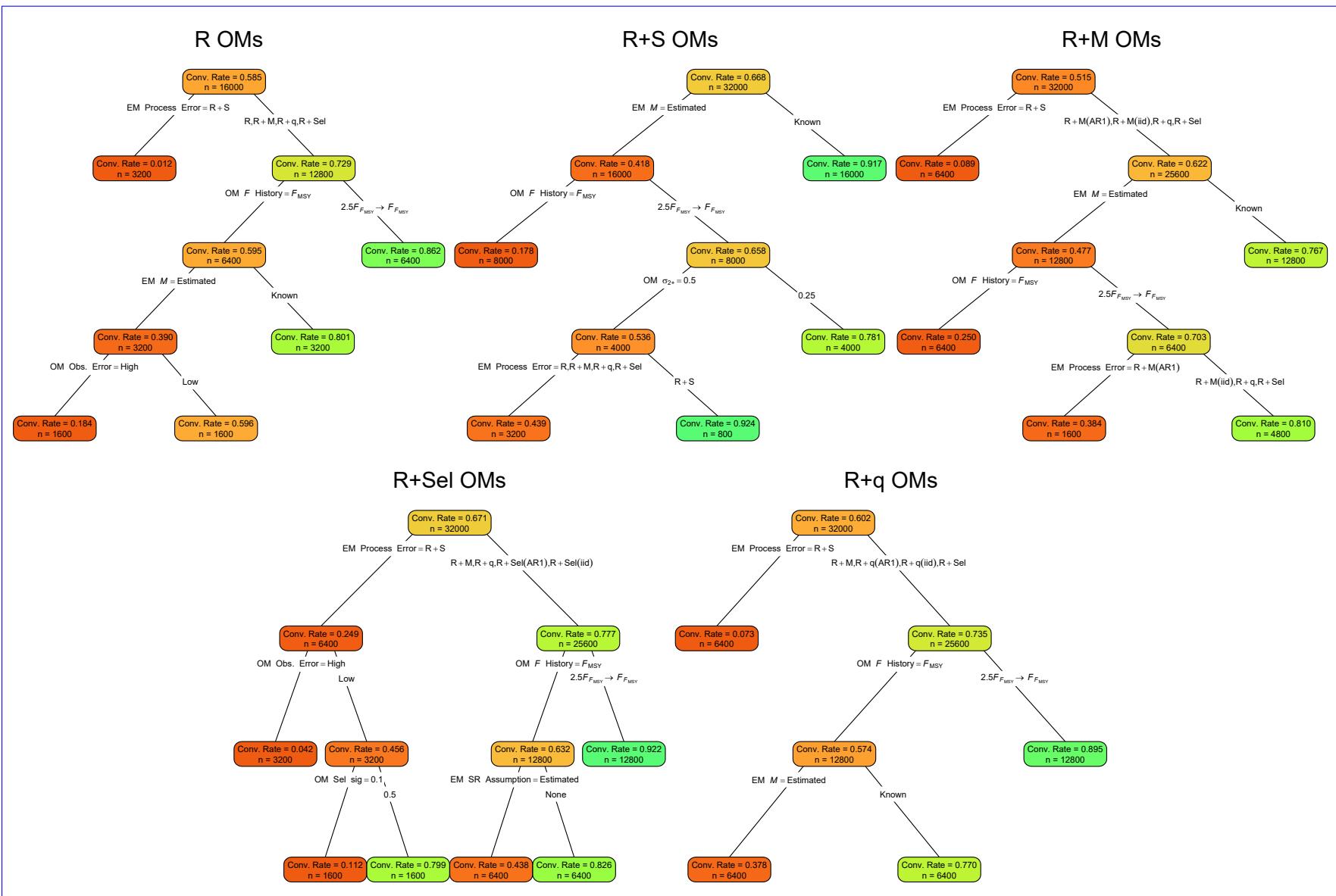


Fig. S2. Classification trees indicating primary factors determining convergence as defined by a maximum absolute gradient $< 10^{-6}$ for R, R+S, R+M, R+Sel and R+q OMs. Lower or higher convergence rates are indicated by more red or green polygons, respectively

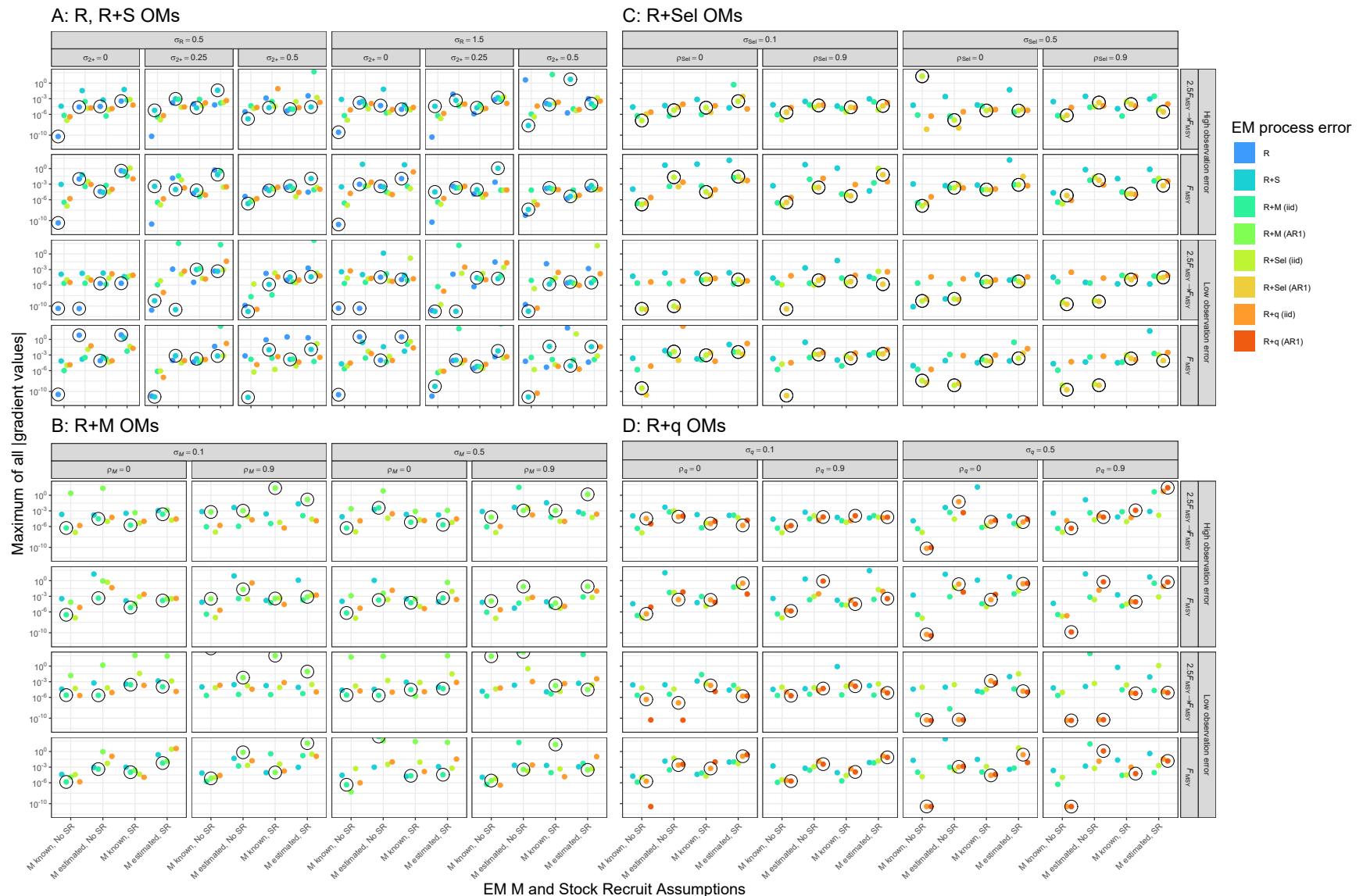


Fig. S3. The maximum of the absolute values of all gradient values for all fits that provided Hessian-based standard errors across all simulated data sets of a given OM configuration (A: R and R+S, B: R+M, C: R+Sel, or D: R+q). Results are conditional on EM fits with alternative process error type (colored points and lines), median natural mortality (estimated or known) and recruitment assumptions (Beverton-Holt SRR or not). Circled values indicate results where the EM process error structure matches that of the OM and vertical lines represent 95% confidence intervals.

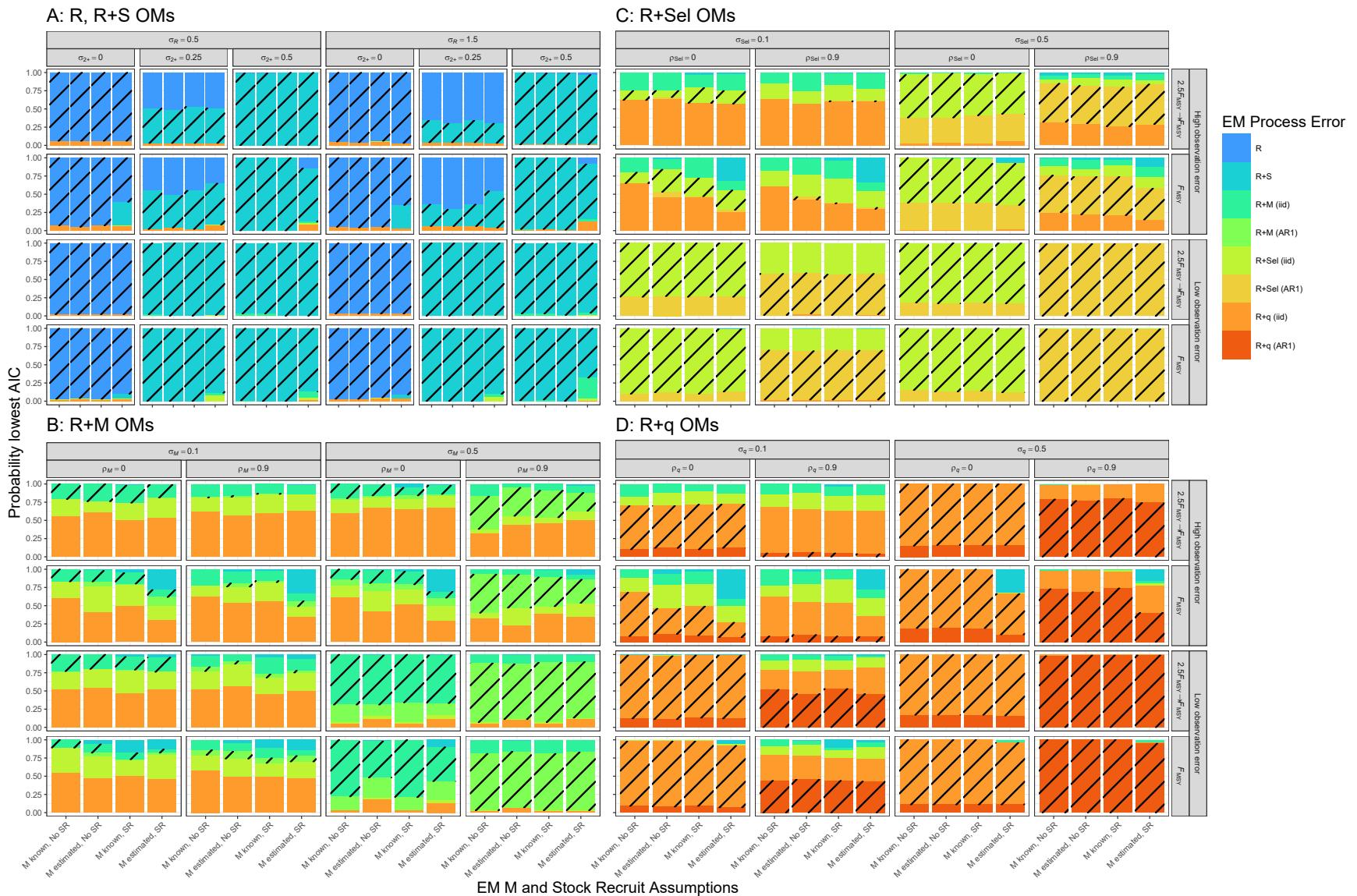


Fig. S4. Estimated probability of lowest AIC for EMs assuming alternative process error structures (colored bars) conditional on alternative assumptions for median natural mortality (estimated or known) and Beverton-Holt SRR (estimated or not; along x-axis) when fitted to OMs that have R and R+S (A), R+Sel (B), R+M (C), or R+q (D) process error structures. Striped bars indicate results where the EM process error structure matches that of the OM.

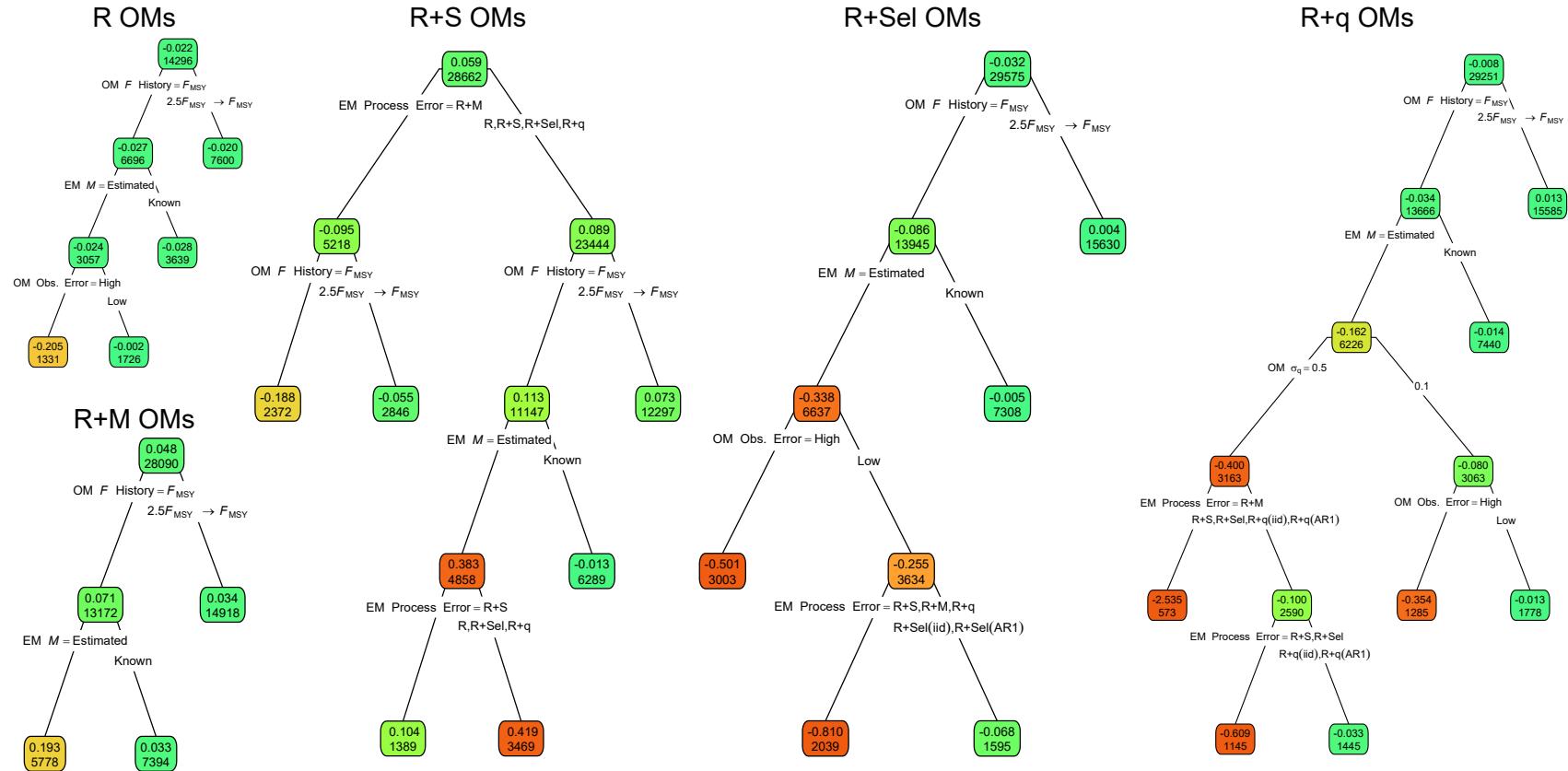


Fig. S5. Regression trees indicating primary factors determining reductions in sums of squares of errors in estimation measured by Eq. 3 for terminal year fully-selected fishing mortality for R+S, R+M, R+Sel and R+q OMs. Each node shows the median error (top) and number of observations (bottom) for the corresponding subset. Median errors closer to or further from zero are indicated by more green or red polygons, respectively.

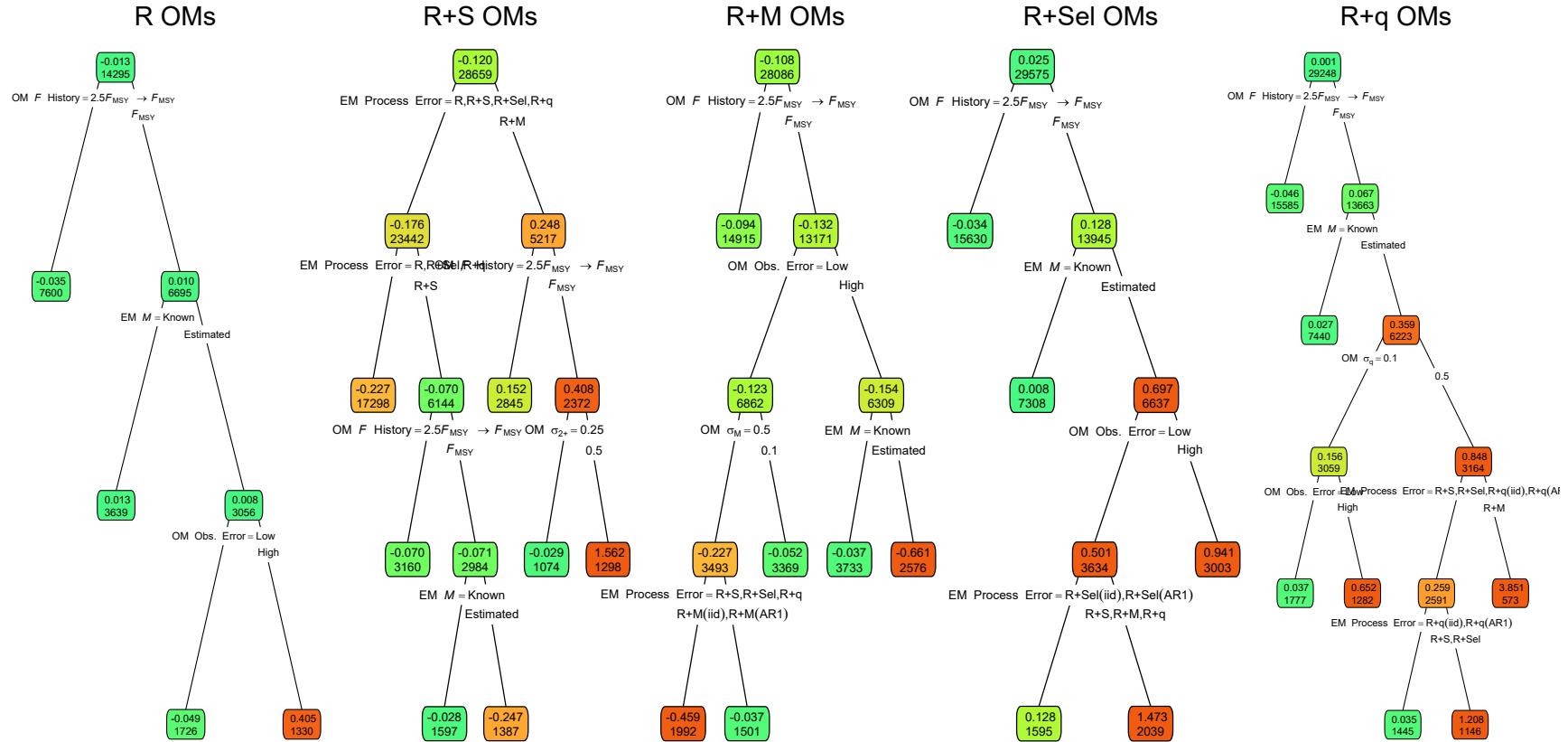


Fig. S6. Regression trees indicating primary factors determining reductions in sums of squares of errors in estimation measured by Eq. 3 for terminal year recruitment for R+S, R+M, R+Sel and R+q OM scenarios. Each node shows the median error (top) and number of observations (bottom) for the corresponding subset. Median errors closer to or further from zero are indicated by more green or red polygons, respectively.

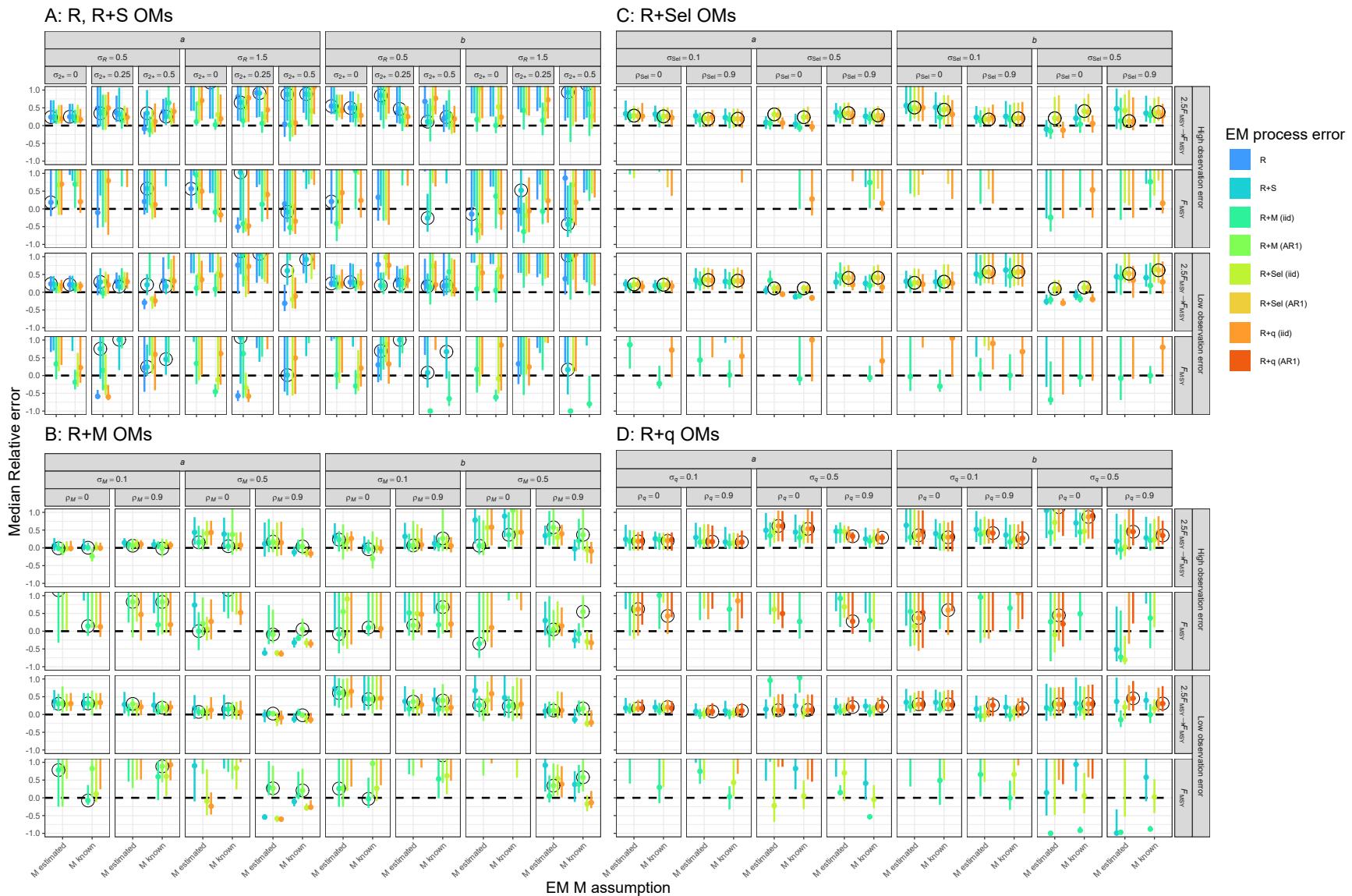


Fig. S7. Median relative error of Beverton-Holt SRR parameters (*a* and *b*) for EMs fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the OM and vertical lines represent 95% confidence intervals.

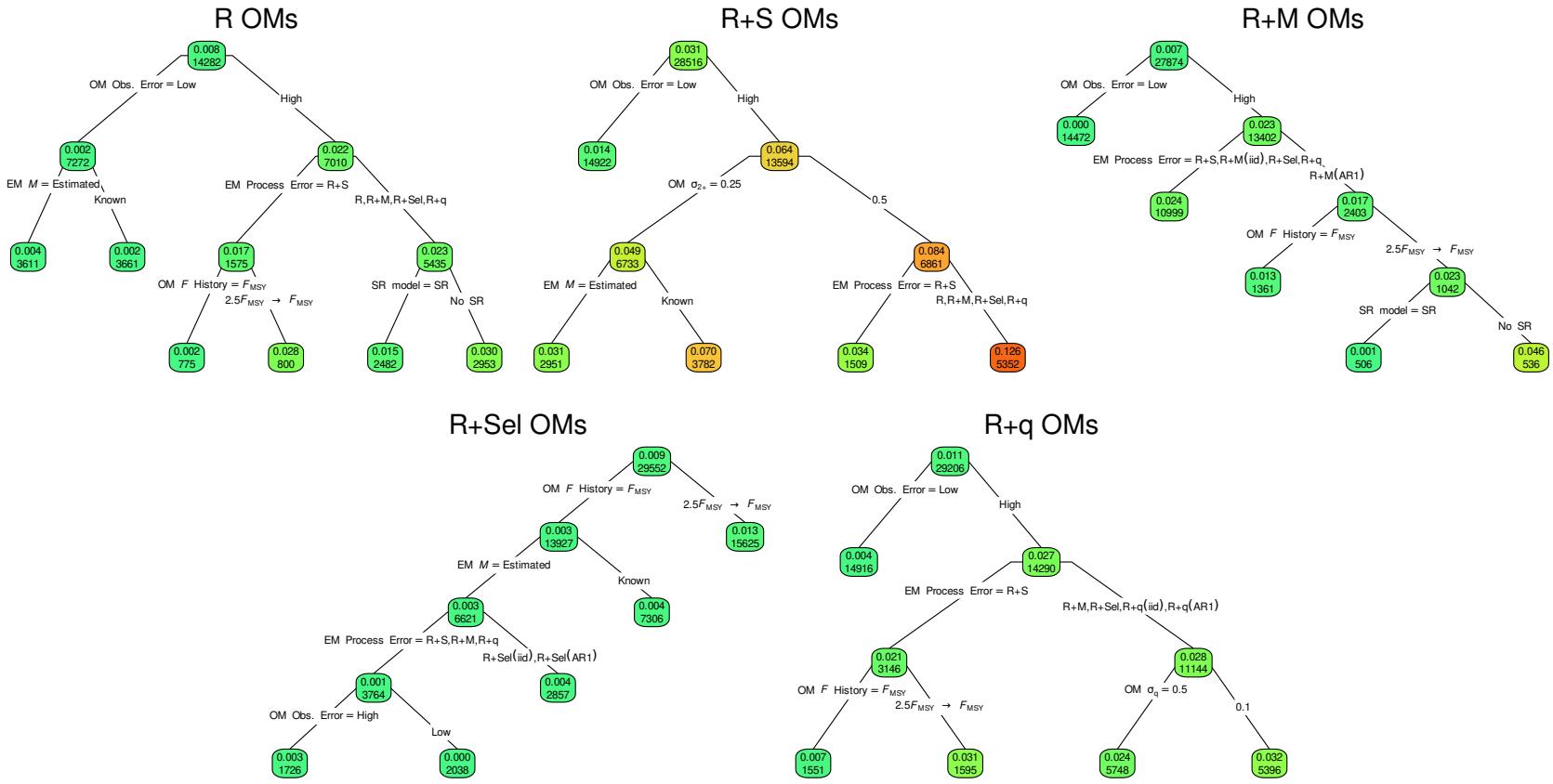


Fig. S8. Regression trees indicating primary factors determining reductions in sums of squares of errors in transformed Mohn's ρ (Eq. 3) for fishing mortality averaged over all age classes for R+S, R+M, R+Sel and R+q OM models. Each node shows the median Mohn's ρ (top) and number of observations (bottom) for the corresponding subset. Median Mohn's ρ closer to or further from zero are indicated by more green or red polygons, respectively.

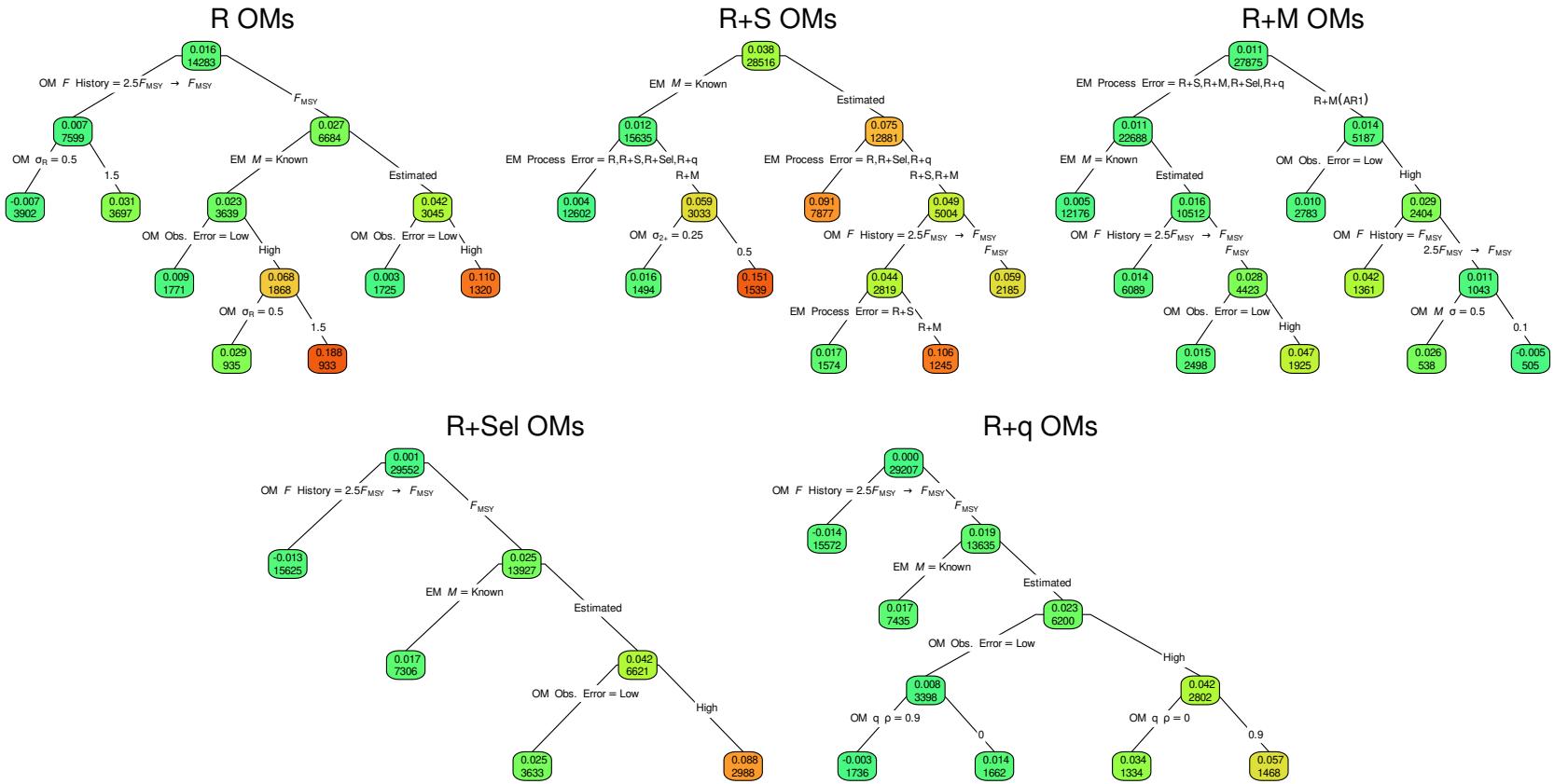


Fig. S9. Regression trees indicating primary factors determining reductions in sums of squares of errors in transformed Mohn's ρ (Eq. 3) for recruitment for R+S, R+M, R+Sel and R+q OM models. Each node shows the median Mohn's ρ (top) and number of observations (bottom) for the corresponding subset. Median Mohn's ρ closer to or further from zero are indicated by more green or red polygons, respectively.

1172 Referenced Tables

Table S1. Distinguishing characteristics of the **operating models** QMs with random effects on recruitment and apparent survival (R , $R+S$). Standard When observation uncertainty is low, standard deviations (SD) are for log-normal distributed indices and logistic normal distributed age composition observations (fleet are 0.1 and indices) 0.3, respectively, and when it is high, standard deviations are 0.4 and 1.5, respectively. Fishing mortality either changes from $2.5F_{MSY}$ to F_{MSY} after year 20 (of 40) for fishing histories where fishing mortality or is not constant at F_{MSY} over all years.

Model	σ_R	σ_{2+}	Fishing History	Observation Uncertainty
1	0.5		$2.5F_{MSY} \rightarrow F_{MSY}$	Low
2	1.5		$2.5F_{MSY} \rightarrow F_{MSY}$	Low
3	0.5	0.25	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
4	1.5	0.25	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
5	0.5	0.50	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
6	1.5	0.50	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
7	0.5		F_{MSY}	Low
8	1.5		F_{MSY}	Low
9	0.5	0.25	F_{MSY}	Low
10	1.5	0.25	F_{MSY}	Low
11	0.5	0.50	F_{MSY}	Low
12	1.5	0.50	F_{MSY}	Low
13	0.5		$2.5F_{MSY} \rightarrow F_{MSY}$	High
14	1.5		$2.5F_{MSY} \rightarrow F_{MSY}$	High
15	0.5	0.25	$2.5F_{MSY} \rightarrow F_{MSY}$	High
16	1.5	0.25	$2.5F_{MSY} \rightarrow F_{MSY}$	High
17	0.5	0.50	$2.5F_{MSY} \rightarrow F_{MSY}$	High
18	1.5	0.50	$2.5F_{MSY} \rightarrow F_{MSY}$	High
19	0.5		F_{MSY}	High
20	1.5		F_{MSY}	High
21	0.5	0.25	F_{MSY}	High
22	1.5	0.25	F_{MSY}	High
23	0.5	0.50	F_{MSY}	High
24	1.5	0.50	F_{MSY}	High

Table S2. Distinguishing characteristics of the ~~operating models~~ QMs with random effects on recruitment and natural mortality (R+M). Standard When observation uncertainty is low, standard deviations (SD) are for log-normal distributed indices and logistic normal distributed age composition observations (fleet are 0.1 and indices) 0.3, respectively, and when it is high, standard deviations are 0.4 and 1.5, respectively. Fishing mortality either changes from $2.5F_{MSY}$ to F_{MSY} after year 20 (of 40) for fishing histories where fishing mortality or is not constant at F_{MSY} over all years. For AR1 process errors, σ_{σ_M} is defined for the marginal distribution of the processes.

Model	σ_R	σ_M	ρ_M	Fishing History	Observation Uncertainty
1	0.5	0.1	0.0	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
2	0.5	0.5	0.0	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
3	0.5	0.1	0.9	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
4	0.5	0.5	0.9	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
5	0.5	0.1	0.0	F_{MSY}	Low
6	0.5	0.5	0.0	F_{MSY}	Low
7	0.5	0.1	0.9	F_{MSY}	Low
8	0.5	0.5	0.9	F_{MSY}	Low
9	0.5	0.1	0.0	$2.5F_{MSY} \rightarrow F_{MSY}$	High
10	0.5	0.5	0.0	$2.5F_{MSY} \rightarrow F_{MSY}$	High
11	0.5	0.1	0.9	$2.5F_{MSY} \rightarrow F_{MSY}$	High
12	0.5	0.5	0.9	$2.5F_{MSY} \rightarrow F_{MSY}$	High
13	0.5	0.1	0.0	F_{MSY}	High
14	0.5	0.5	0.0	F_{MSY}	High
15	0.5	0.1	0.9	F_{MSY}	High
16	0.5	0.5	0.9	F_{MSY}	High

Table S3. Distinguishing characteristics of the ~~operating models~~ QMs with random effects on recruitment and selectivity (R+Sel). Standard When observation uncertainty is low, standard deviations (SD) are for log-normal distributed indices and logistic normal distributed age composition observations (fleet are 0.1 and indices) 0.3, respectively, and when it is high, standard deviations are 0.4 and 1.5, respectively. Fishing mortality either changes from $2.5F_{MSY}$ to F_{MSY} after year 20 (of 40) for fishing histories where fishing mortality or is not constant at F_{MSY} over all years. For AR1 process errors, σ_{Sel} is defined for the marginal distribution of the processes.

Model	σ_R	σ_{Sel}	ρ_{Sel}	Fishing History	Observation Uncertainty
1	0.5	0.1	0.0	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
2	0.5	0.5	0.0	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
3	0.5	0.1	0.9	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
4	0.5	0.5	0.9	$2.5F_{MSY} \rightarrow F_{MSY}$	Low
5	0.5	0.1	0.0	F_{MSY}	Low
6	0.5	0.5	0.0	F_{MSY}	Low
7	0.5	0.1	0.9	F_{MSY}	Low
8	0.5	0.5	0.9	F_{MSY}	Low
9	0.5	0.1	0.0	$2.5F_{MSY} \rightarrow F_{MSY}$	High
10	0.5	0.5	0.0	$2.5F_{MSY} \rightarrow F_{MSY}$	High
11	0.5	0.1	0.9	$2.5F_{MSY} \rightarrow F_{MSY}$	High
12	0.5	0.5	0.9	$2.5F_{MSY} \rightarrow F_{MSY}$	High
13	0.5	0.1	0.0	F_{MSY}	High
14	0.5	0.5	0.0	F_{MSY}	High
15	0.5	0.1	0.9	F_{MSY}	High
16	0.5	0.5	0.9	F_{MSY}	High

Table S4. Distinguishing characteristics of the ~~operating models~~ QMs with random effects on recruitment and catchability ($R+q$). ~~Standard~~ When observation uncertainty is low, standard deviations (SD) are for log-normal distributed indices and logistic normal distributed age composition observations (fleet are 0.1 and indices) 0.3, respectively, and when it is high, standard deviations are 0.4 and 1.5, respectively. Fishing mortality either changes from $2.5F_{\text{MSY}}$ to F_{MSY} after year 20 (of 40) ~~for fishing histories where fishing mortality or~~ is not constant at F_{MSY} over all years. For AR1 process errors, σ_q is defined for the marginal distribution of the processes.

Model	σ_R	σ_q	ρ_q	Fishing History	Observation Uncertainty
1	0.5	0.1	0.0	$2.5F_{\text{MSY}} \rightarrow F_{\text{MSY}}$	Low
2	0.5	0.5	0.0	$2.5F_{\text{MSY}} \rightarrow F_{\text{MSY}}$	Low
3	0.5	0.1	0.9	$2.5F_{\text{MSY}} \rightarrow F_{\text{MSY}}$	Low
4	0.5	0.5	0.9	$2.5F_{\text{MSY}} \rightarrow F_{\text{MSY}}$	Low
5	0.5	0.1	0.0	F_{MSY}	Low
6	0.5	0.5	0.0	F_{MSY}	Low
7	0.5	0.1	0.9	F_{MSY}	Low
8	0.5	0.5	0.9	F_{MSY}	Low
9	0.5	0.1	0.0	$2.5F_{\text{MSY}} \rightarrow F_{\text{MSY}}$	High
10	0.5	0.5	0.0	$2.5F_{\text{MSY}} \rightarrow F_{\text{MSY}}$	High
11	0.5	0.1	0.9	$2.5F_{\text{MSY}} \rightarrow F_{\text{MSY}}$	High
12	0.5	0.5	0.9	$2.5F_{\text{MSY}} \rightarrow F_{\text{MSY}}$	High
13	0.5	0.1	0.0	F_{MSY}	High
14	0.5	0.5	0.0	F_{MSY}	High
15	0.5	0.1	0.9	F_{MSY}	High
16	0.5	0.5	0.9	F_{MSY}	High

Table S5. Distinguishing characteristics of the ~~estimating models~~ EMs and ~~operating model indication (+) of which OM~~ process error categories (R, R+S, R+M, R+Sel, R+q) ~~where used each EM configuration was fit.~~

Model	Recruitment model	Median M	Process error	R,R+S OMs	R+M OMs	R+Sel OMs	R+q OMs
1	Mean recruitment	0.2	R ($\sigma_{2+} = 0$)	+			
2	Beverton-Holt	0.2	R ($\sigma_{2+} = 0$)	+			
3	Mean recruitment	Estimated	R ($\sigma_{2+} = 0$)	+			
4	Beverton-Holt	Estimated	R ($\sigma_{2+} = 0$)	+			
5	Mean recruitment	0.2	R+S (σ_{2+} estimated)	+	+	+	+
6	Beverton-Holt	0.2	R+S (σ_{2+} estimated)	+	+	+	+
7	Mean recruitment	Estimated	R+S (σ_{2+} estimated)	+	+	+	+
8	Beverton-Holt	Estimated	R+S (σ_{2+} estimated)	+	+	+	+
9	Mean recruitment	0.2	R+M ($\rho_M = 0$)	+	+	+	+
10	Beverton-Holt	0.2	R+M ($\rho_M = 0$)	+	+	+	+
11	Mean recruitment	Estimated	R+M ($\rho_M = 0$)	+	+	+	+
12	Beverton-Holt	Estimated	R+M ($\rho_M = 0$)	+	+	+	+
13	Mean recruitment	0.2	R+Sel ($\rho_{Sel} = 0$)	+	+	+	+
14	Beverton-Holt	0.2	R+Sel ($\rho_{Sel} = 0$)	+	+	+	+
15	Mean recruitment	Estimated	R+Sel ($\rho_{Sel} = 0$)	+	+	+	+
16	Beverton-Holt	Estimated	R+Sel ($\rho_{Sel} = 0$)	+	+	+	+
17	Mean recruitment	0.2	R+q ($\rho_q = 0$)	+	+	+	+
18	Beverton-Holt	0.2	R+q ($\rho_q = 0$)	+	+	+	+
19	Mean recruitment	Estimated	R+q ($\rho_q = 0$)	+	+	+	+
20	Beverton-Holt	Estimated	R+q ($\rho_q = 0$)	+	+	+	+
21	Mean recruitment	0.2	R+M (ρ_M estimated)		+		
22	Beverton-Holt	0.2	R+M (ρ_M estimated)		+		
23	Mean recruitment	Estimated	R+M (ρ_M estimated)		+		
24	Beverton-Holt	Estimated	R+M (ρ_M estimated)		+		
25	Mean recruitment	0.2	R+Sel (ρ_{Sel} estimated)			+	
26	Beverton-Holt	0.2	R+Sel (ρ_{Sel} estimated)			+	
27	Mean recruitment	Estimated	R+Sel (ρ_{Sel} estimated)			+	
28	Beverton-Holt	Estimated	R+Sel (ρ_{Sel} estimated)			+	
29	Mean recruitment	0.2	R+q (ρ_q estimated)				+
30	Beverton-Holt	0.2	R+q (ρ_q estimated)				+
31	Mean recruitment	Estimated	R+q (ρ_q estimated)				+
32	Beverton-Holt	Estimated	R+q (ρ_q estimated)				+

Table S6. For each OM process error type (columns), percent reduction in deviance for logistic regression models fit to indicators of convergence (maximum absolute gradient $< 10^{-6}$) with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM Process Error	30.40	0.45	17.57	16.04	24.03
EM M Assumption	2.38	24.11	4.42	1.02	2.66
EM SR Assumption	1.80	0.32	0.96	3.38	2.13
OM Obs. Error	0.12	0.77	0.33	1.76	0.28
OM F History	3.51	6.33	2.36	5.86	5.30
OM σ_R	≤ 0.01	≤ 0.01	—	—	—
OM σ_{2+}	—	≤ 0.01	—	—	—
OM σ_M	—	—	0.39	—	—
OM ρ_M	—	—	0.09	—	—
OM σ_{Sel}	—	—	—	1.08	—
OM ρ_{Sel}	—	—	—	0.01	—
OM σ_q	—	—	—	—	0.06
OM ρ_q	—	—	—	—	≤ 0.01
All factors	43.69	35.72	29.33	34.57	40.69
± All Two Way	50.53	42.99	43.91	45.93	48.62
± All Three Way	52.30	48.41	46.81	47.71	50.40

Table S7. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the terminal year fully-selected fishing mortality with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM M Assumption	2.26	1.33	1.26	2.93	3.26
EM SR assumption	0.11	0.07	0.08	0.07	0.09
EM Process Error	0.46	4.18	0.38	0.13	1.02
OM Obs. Error	1.61	0.06	0.86	0.41	<0.01
OM F History	2.49	3.23	1.42	3.22	4.55
OM σ_R	0.02	0.02	—	—	—
OM σ_{2+}	—	0.87	—	—	—
OM σ_M	—	—	0.16	—	—
OM ρ_M	—	—	0.01	—	—
OM σ_{Sel}	—	—	—	0.24	—
OM ρ_{Sel}	—	—	—	0.05	—
OM σ_q	—	—	—	—	1.03
OM ρ_q	—	—	—	—	0.05
All factors	7.42	9.96	4.37	7.26	10.43
± All Two Way	17.63	25.76	10.94	13.88	22.07
± All Three Way	22.97	37.03	14.74	17.32	30.74

Table S8. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to errors in estimation measured by Eq. 3 for the terminal year recruitment with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM M Assumption	1.96	0.40	0.69	3.52	3.03
EM SR assumption	0.06	0.02	0.05	0.02	0.05
EM Process Error	0.39	4.74	0.41	0.12	1.16
OM Obs. Error	1.47	0.08	0.64	0.18	<0.01
OM F History	2.54	2.66	1.11	4.18	5.06
OM σ_R	0.03	0.01	—	—	—
OM σ_{2+}	—	1.05	—	—	—
OM σ_M	—	—	0.36	—	—
OM ρ_M	—	—	0.02	—	—
OM σ_{Sel}	—	—	—	0.23	—
OM ρ_{Sel}	—	—	—	0.06	—
OM σ_q	—	—	—	—	1.09
OM ρ_q	—	—	—	—	0.06
All factors	6.90	9.01	3.43	8.58	10.90
± All Two Way	16.48	24.64	9.73	15.76	22.75
± All Three Way	21.46	35.60	13.56	19.07	31.15

1173 **Further Detailed Results**

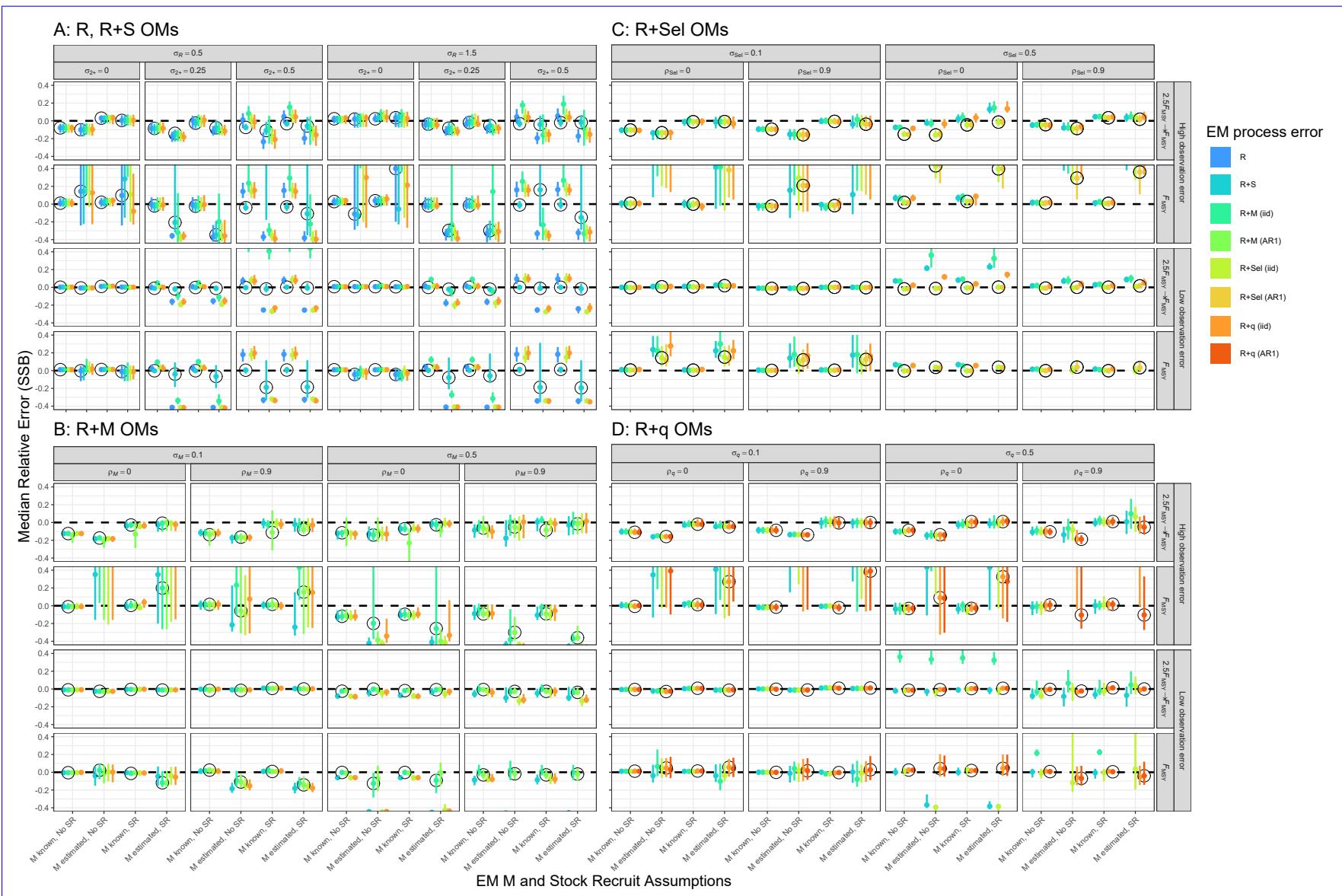


Fig. S10. The maximum Median relative error of the absolute values of all gradient values terminal year SSB for all fits that provided hessian-based standard errors across all simulated EMs fitted to data sets of a given OM configuration (A simulated with alternative process error structures: R and R+S (A), B: R+MSel (B), C: R+SelM (C), or D: R+q). Results are conditional on EM fits with alternative process error type (colored points and linesD), median natural mortality (estimated or known) and recruitment assumptions (Beverton-Holt stock-recruit relationship or not). Circled values indicate results where the EM process error structure matches that of the operating model OM and vertical lines represent 95% confidence intervals.

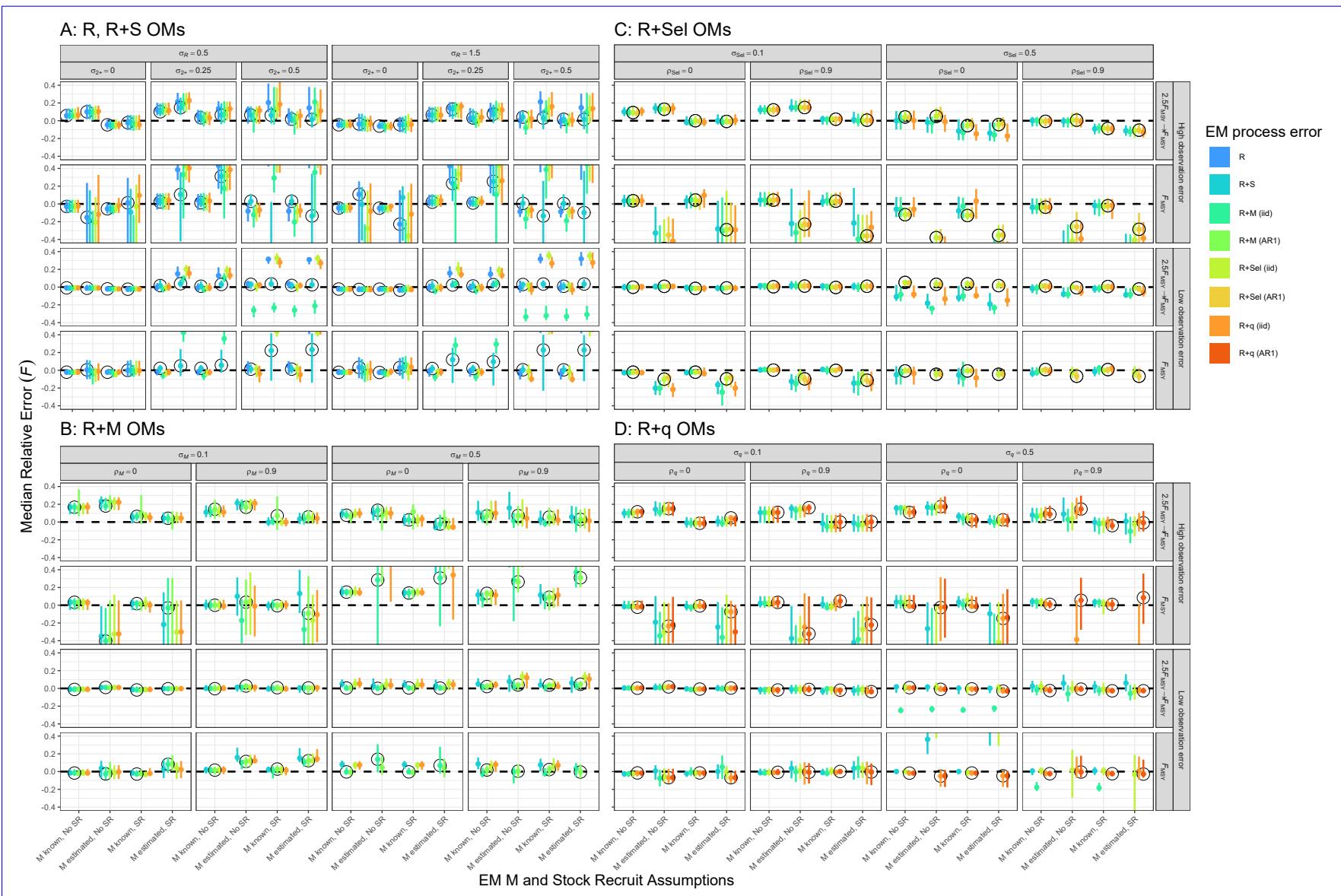


Fig. S11. Probability of estimating models providing maximum absolute values of gradients less than 10^{-6} assuming alternative process Median relative error (colored points and lines), and median natural of terminal year fully-selected fishing mortality (estimated or known) and Beverton-Holt stock-recruit relationships (estimated or not; along x-axis) when for EMs fitted to operating models that have data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D) process error structures. Circled values indicate results where the EM process error structure matches that of the operating model OM and vertical lines represent 95% confidence intervals.

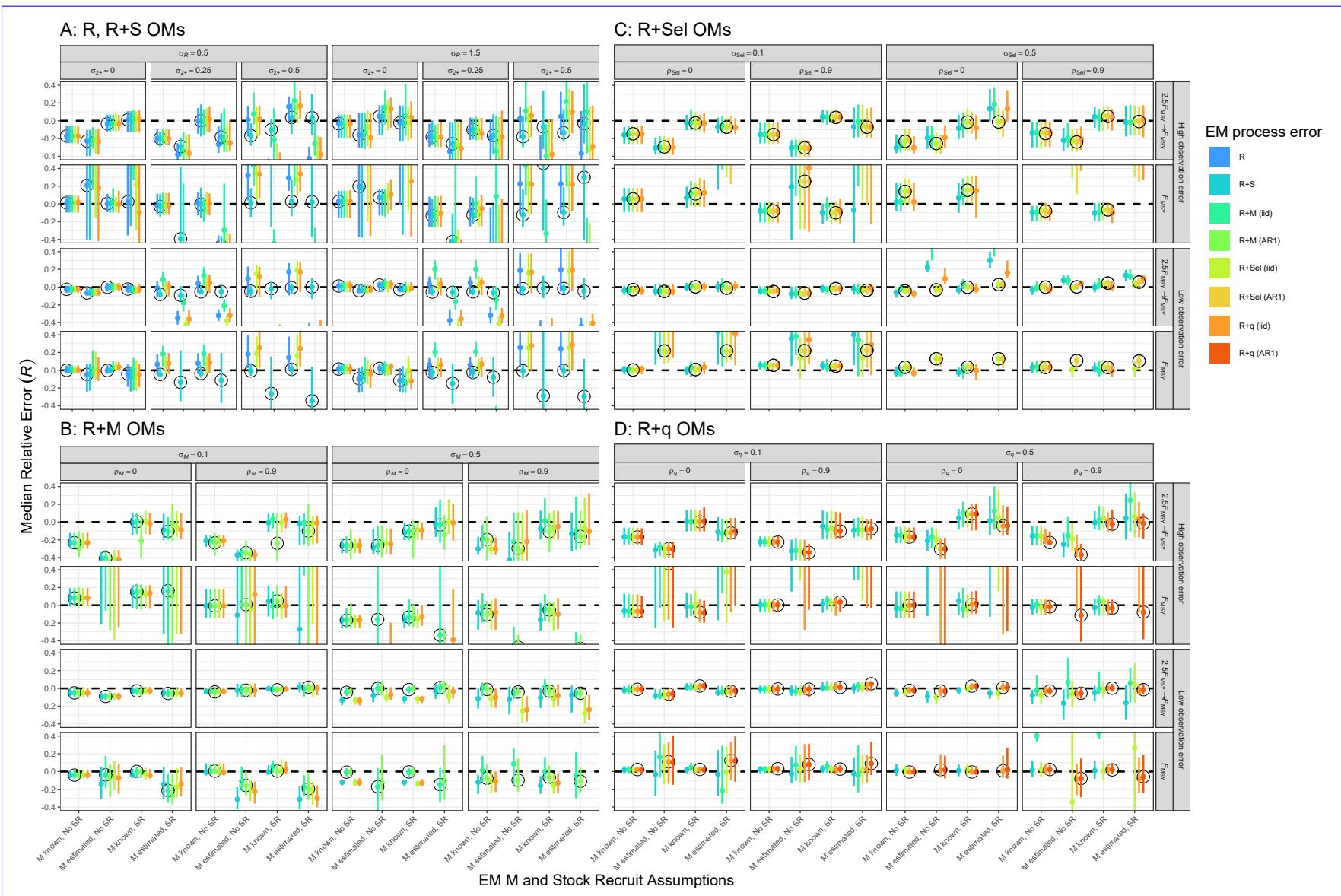


Fig. S12. Estimated probability Median relative error of lowest AIC from logistic regression on the log-standard deviation of the true log(SSB) in each simulation terminal year recruitment for estimating model EMs fitted to data sets simulated with Beverton-Holt stock-recruit relationships, rather than the otherwise equivalent EM without the stock-recruit relationship. Results are conditional on alternative assumptions for median natural mortality (estimated or known) and on EMs having the correct process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Rug along x-axis denotes $SD(\log(SSB))$. Circled values for each simulation indicate results where the EM process error structure matches that of the

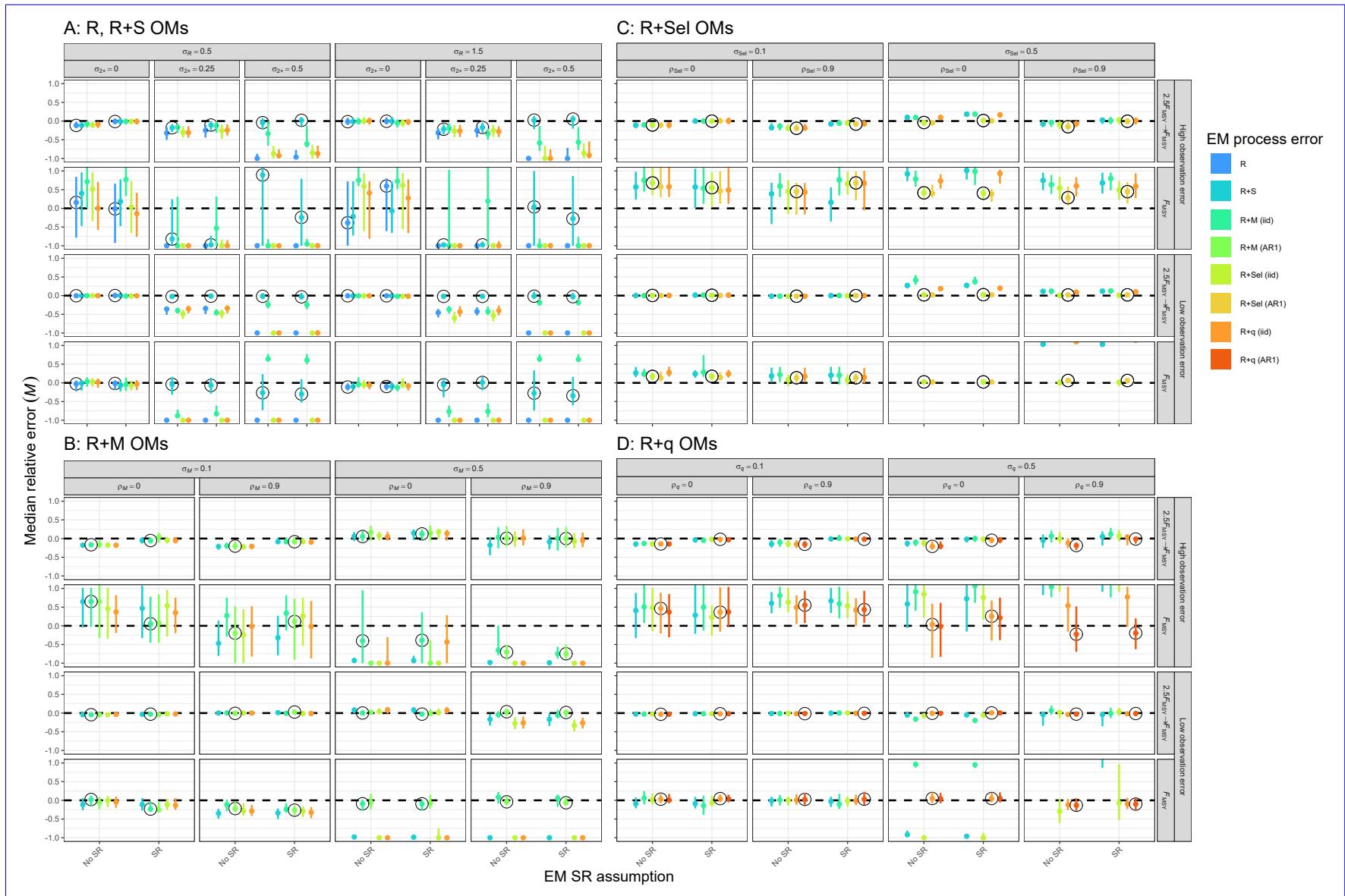


Fig. S13. Median relative error of median natural mortality for EMs fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the OM and vertical lines represent 95% confidence intervals.

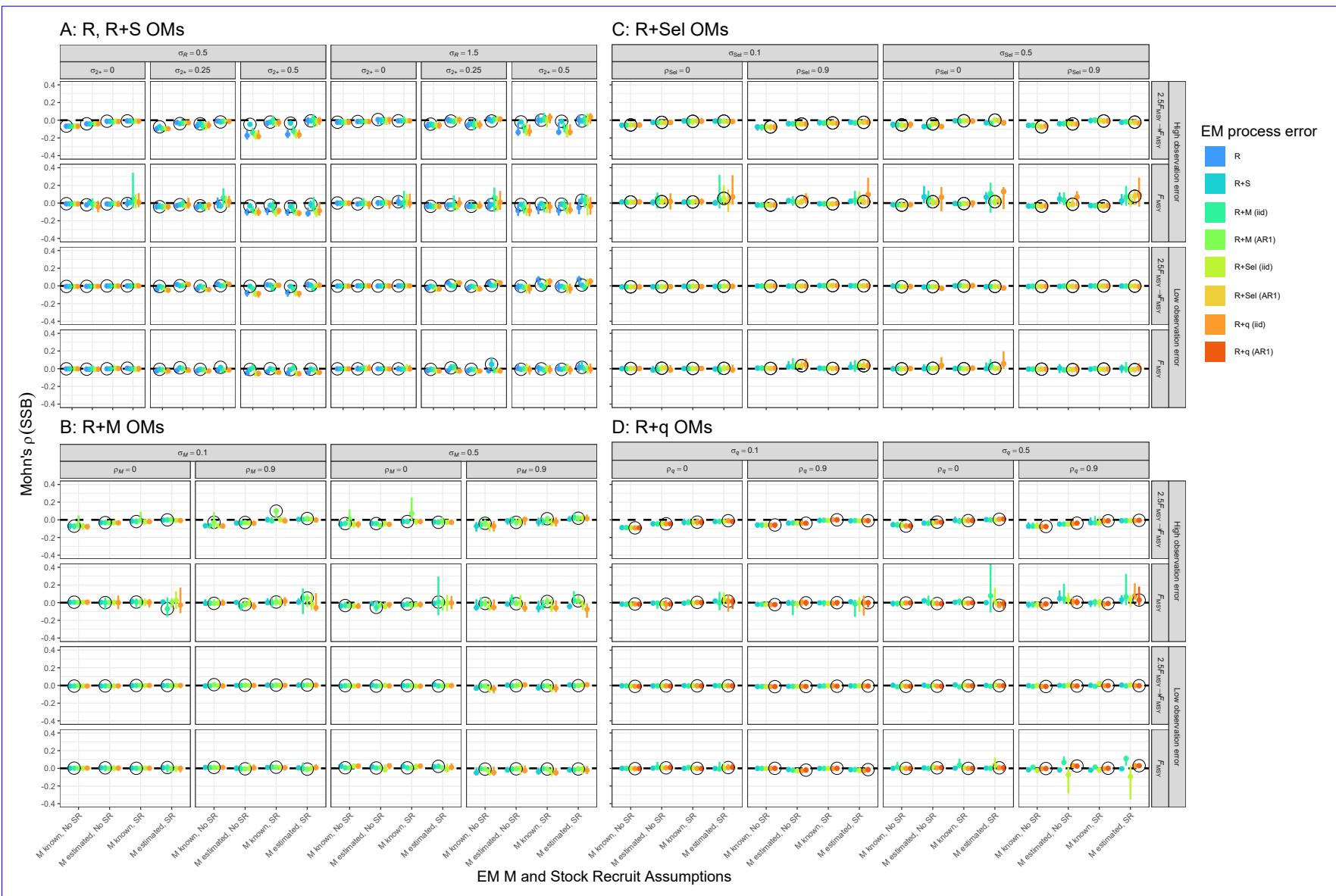


Fig. S14. Median relative error of terminal year fully selected fishing mortality Mohn's ρ for estimating models SSB for EMs fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model OM and vertical lines represent 95% confidence intervals.

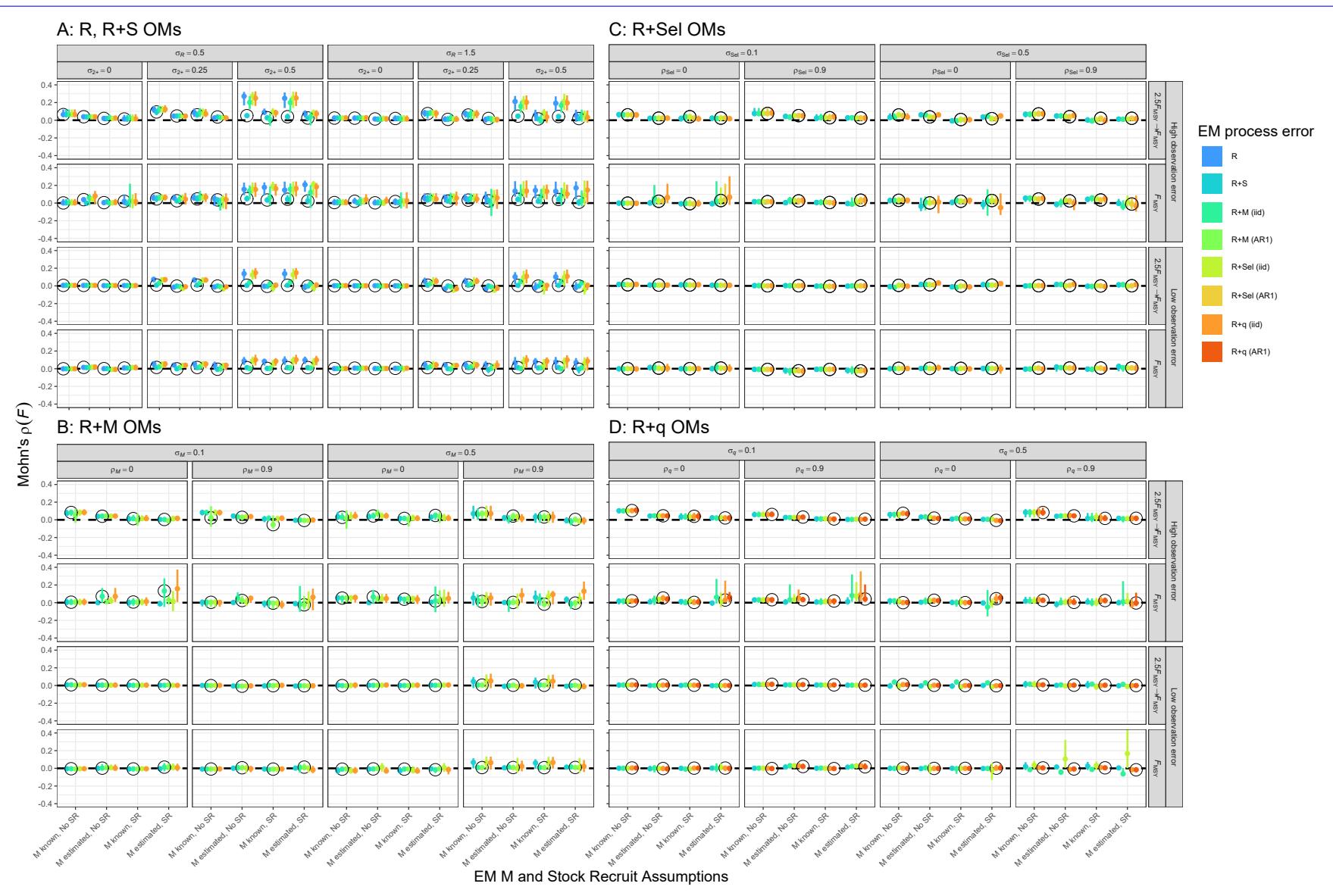


Fig. S15. Median relative error—Mohn's ρ of terminal year recruitment-fishing mortality averaged over all age classes for estimating models—EMs fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model—OM and vertical lines represent 95% confidence intervals.

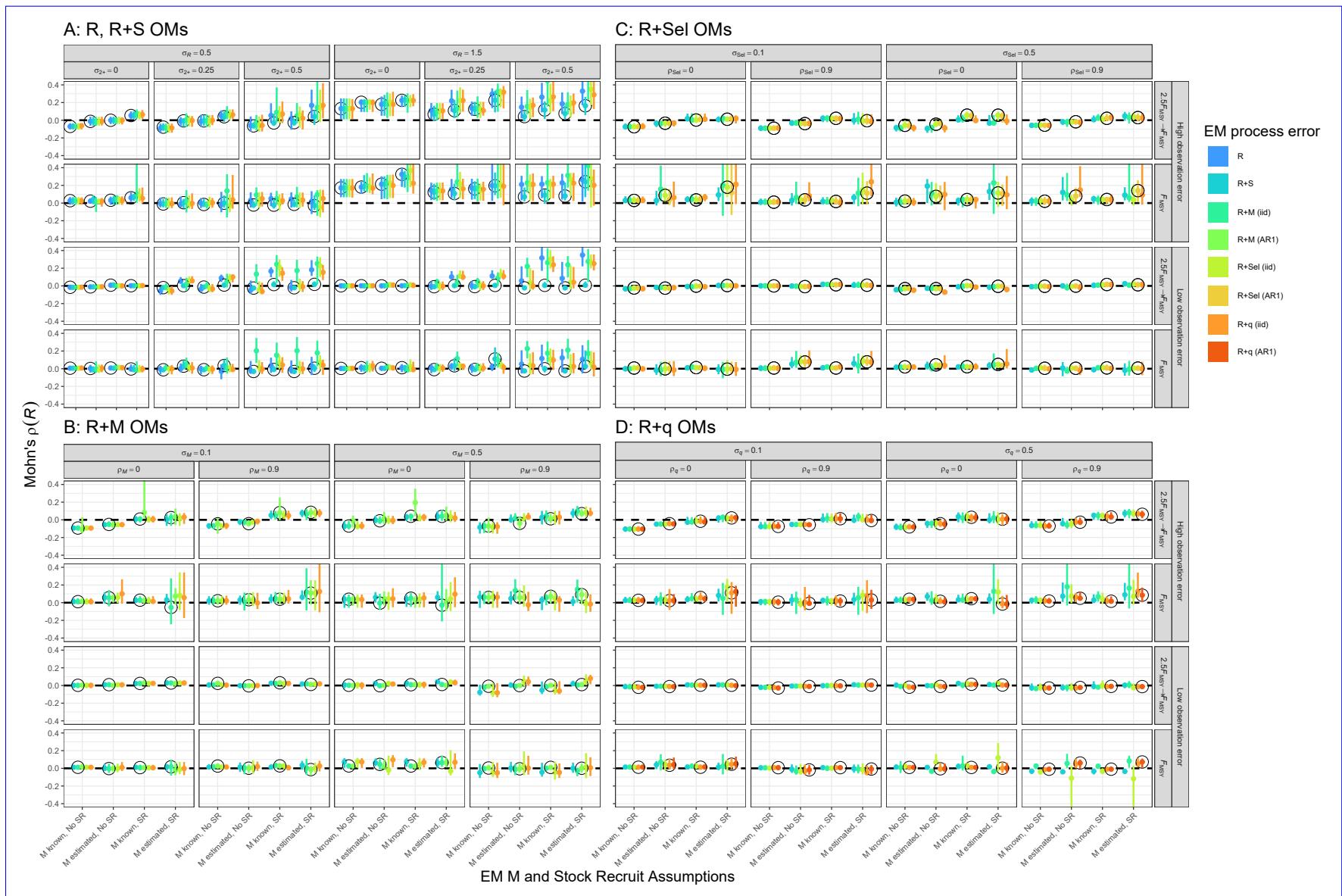


Fig. S16. Median relative error Mohn's ρ of Beverton-Holt stock-recruit parameters (*a* and *b*) recruitment for estimating models EMs fitted to data sets simulated with alternative process error structures: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where the EM process error structure matches that of the operating model OM and vertical lines represent 95% confidence intervals.

Table S9. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to transformed Mohn's ρ values for each simulation (Eq. 3) for fishing mortality averaged over all age classes with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM M Assumption	0.06	0.09	0.01	0.12	0.01
EM SR assumption	0.01	<0.01	0.01	0.02	0.01
EM Process Error	0.03	0.07	0.02	0.06	0.03
OM Obs. Error	0.16	0.10	0.05	0.02	0.07
OM F History	0.07	0.02	0.03	0.24	0.03
OM σ_R	<0.01	0.01	~	~	~
OM σ_{2+}	~	0.09	~	~	~
OM σ_M	~	~	<0.01	~	~
OM ρ_M	~	~	<0.01	~	~
OM σ_{Sel}	~	~	~	0.01	~
OM ρ_{Sel}	~	~	~	<0.01	~
OM σ_q	~	~	~	~	<0.01
OM ρ_q	~	~	~	~	0.01
All factors	0.32	0.38	0.12	0.48	0.15
+ All Two Way	0.65	0.67	0.30	0.95	0.43
+ All Three Way	1.18	1.11	0.63	1.34	0.90

Table S10. For each OM process error type (columns), percent reduction in deviance for linear regression models fit to transformed Mohn's ρ values for each simulation (Eq. 3) for recruitment with each OM and EM factor (rows) included individually, combined, and with second and third order interactions.

Factor	R	R+S	R+M	R+Sel	R+q
EM M Assumption	0.86	0.56	0.16	1.00	1.27
EM SR assumption	<0.01	0.02	0.01	0.01	0.01
EM Process Error	0.01	0.59	0.18	0.07	0.04
OM Obs. Error	0.34	0.01	0.08	0.24	0.27
OM F History	0.91	0.22	0.06	1.20	1.67
OM σ_R	<0.01	0.14	—	—	—
OM σ_{2+}	—	0.11	—	—	—
OM σ_M	—	—	0.01	—	—
OM ρ_M	—	—	<0.01	—	—
OM σ_{Sel}	—	—	—	0.01	—
OM ρ_{Sel}	—	—	—	0.01	—
OM σ_q	—	—	—	—	0.01
OM ρ_q	—	—	—	—	0.01
All factors	2.28	1.74	0.51	2.66	3.51
+ All Two Way	4.20	2.74	1.08	5.08	6.51
+ All Three Way	4.83	3.79	1.79	6.03	7.82



Fig. S17. Median relative error Probability of EMs providing Hessian-based standard errors with alternative process error (colored points and lines), and median natural mortality for estimating models (estimated or known) and Beverton-Holt SRR (estimated or not; along x-axis) assumptions when fitted to data sets simulated with alternative process error structures: OMs that have R and R+S (A), R+Sel (B), R+M (C), or R+q (D) process error structures. Circled values indicate results where the EM process error structure matches that of the operating model OM and vertical lines represent 95% confidence intervals.

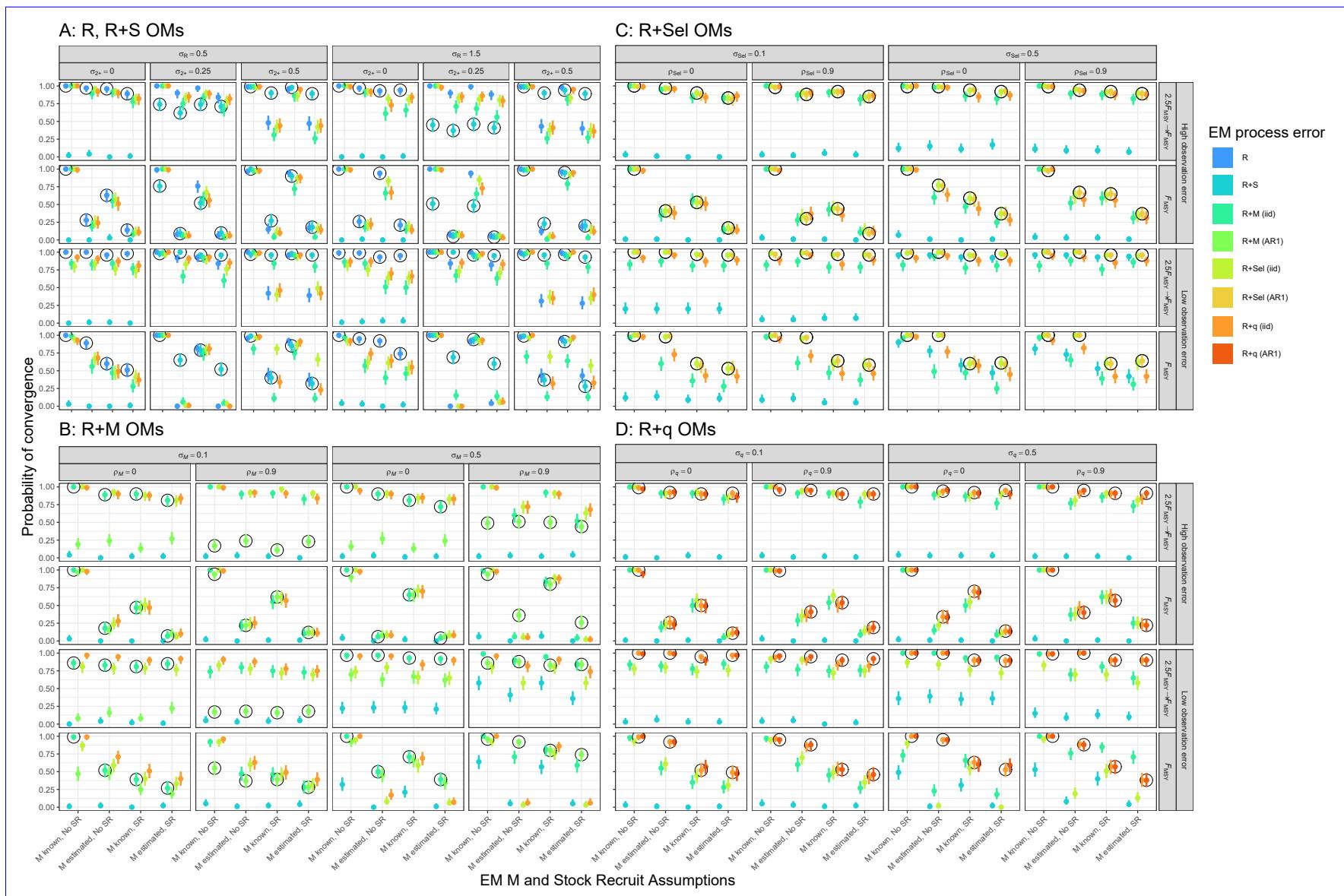


Fig. S18. Median Mohn's ρ -Probability of fishing mortality averaged over all age classes for estimating models fitted to data sets simulated EMs providing maximum absolute values of gradients less than 10^{-6} with alternative process error structures (colored points and lines), and median natural mortality (estimated or known) and Beverton-Holt SRR (estimated or not; along x-axis) assumptions when fitted to OMs that have R and R+S (A), R+Sel (B), R+M (C), or R+q (D) process error structures. Circled values indicate results where the EM process error structure matches that of the operating model OM and vertical lines represent 95% confidence intervals.

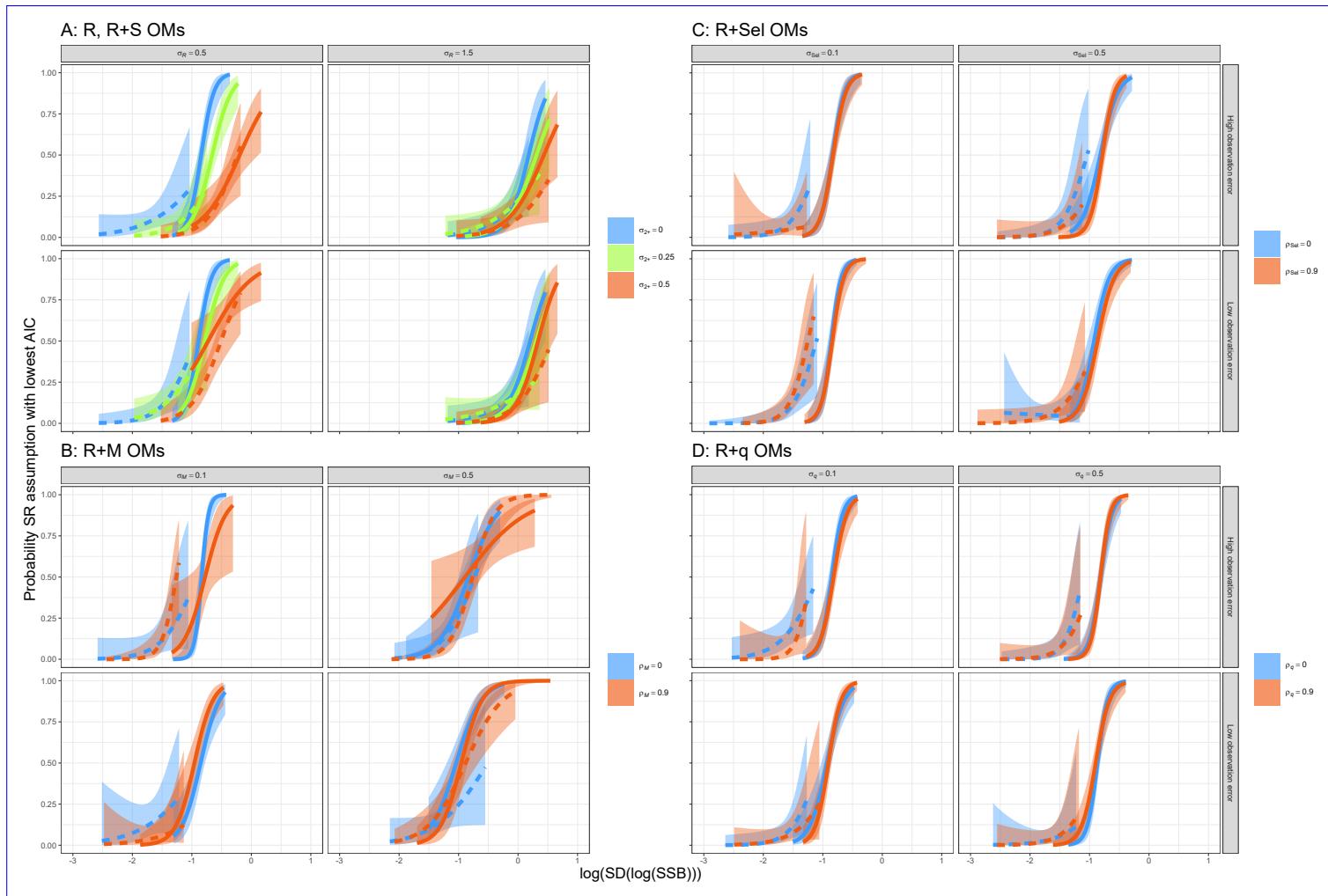


Fig. S19. Median Mohn's ρ -Probability of recruitment lowest AIC from logistic regression on the log-standard deviation of the true $\log(SSB)$ in each simulation for estimating models fitted to data sets simulated EM with Beverton-Holt SRRs, rather than the otherwise equivalent EM without the SRR. Results are conditional on median M is known in the EM and alternative assumptions EMs having the correct process error structure: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Circled values indicate results where median M is assumed known in the EM process error structure matches that of the operating model. Solid and vertical dashed lines are for OMs with and without temporal contrast in fishing pressure, respectively, and polygons represent 95% confidence intervals. Range of results indicates the range of log-standard deviation of $\log(SSB)$ for simulations of the particular OM.

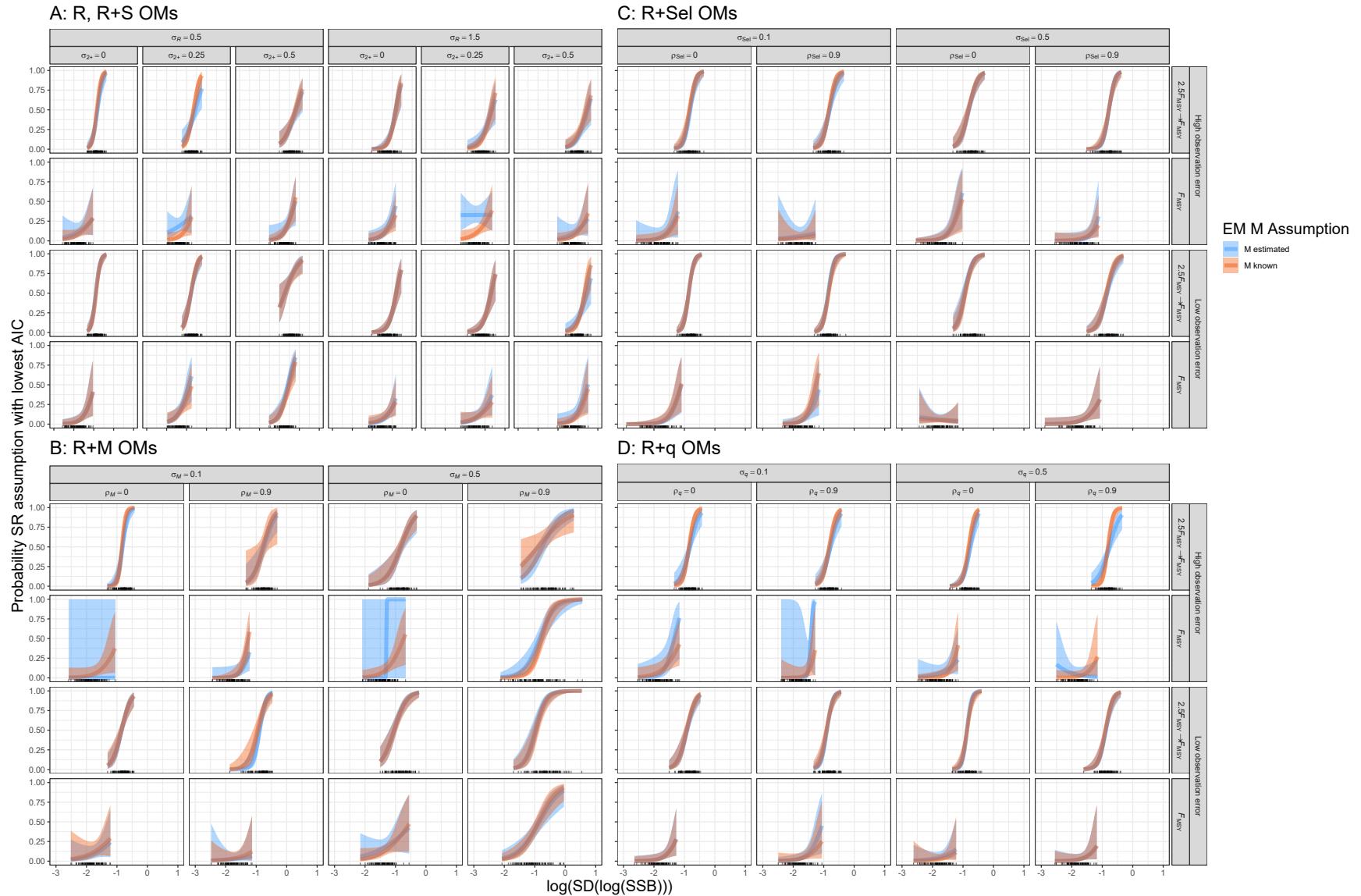


Fig. S20. Estimated probability of lowest AIC from logistic regression on the log-standard deviation of the true log(SSB) in each simulation for EM with Beverton-Holt SRRs, rather than the otherwise equivalent EM without the SRR. Results are conditional on alternative assumptions for median natural mortality (estimated or known) and on EMs having the correct process error structure: R and R+S (A), R+Sel (B), R+M (C), or R+q (D). Rug along x-axis denotes $SD(\log(SSB))$ values for each simulation and polygons represent 95% confidence intervals.