

¹ Factors affecting inferences on natural mortality and
² associated environmental effects in state-space
³ age-structured assessment models

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19

20 **Abstract**

21 Treatment of natural mortality is a major consideration in assessment models and state-space
22 approaches allow estimation of temporal variation in this mortality rate as well as effects of
23 specific covariates. However, there has been no investigation of the reliability of inferences
24 made regarding natural mortality, associated covariate effects and important assessment
25 output from state-space assessment models exists. We conducted a large-scale simulation
26 study that considers models fit to data simulated from operating models with alternative
27 assumptions defined by several factors, but we focus on scenarios where there is temporal
28 contrast in fishing pressure and lower uncertainty in population observations (age composi-
29 tion and indices of abundance). The estimating models we fit to simulated observations had
30 alternative assumptions on whether to include the environmental effect, whether the median
31 natural mortality rate was estimated or known, and the source of temporal variability in
32 the population demography. Lower uncertainty in covariate observations, higher covariate
33 temporal variability and the source of demographic temporal variability were important for
34 better inferences regarding the covariate effect on natural mortality.

35 **keywords:** state-space assessment models, time-varying natural mortality, bias, AIC

³⁶ Introduction

³⁷ State-space population models are now used widely for fisheries stock assessment in Europe,
³⁸ the United States, and Canada (Nielsen and Berg, 2014; Cadigan, 2016; Pedersen and Berg,
³⁹ 2017; Stock and Miller, 2021). Because application of these methods are considered best
⁴⁰ practice and recommended for the next generation of stock assessment models (Hoyle et al.,
⁴¹ 2022; Punt, 2023), it is expected their use will only grow globally. An appeal of state-space
⁴² models lies in their formulation treating latent population characteristics as statistical time
⁴³ series with periodic observations that also may have error due to sampling or other sources
⁴⁴ of measurement error and therefore separating these sources of biological and measurement
⁴⁵ variability. Through advances in computational capacity, we can use sophisticated numer-
⁴⁶ ical approaches to estimate model parameters as mixed effects (Thorson and Minto, 2015;
⁴⁷ Kristensen et al., 2016).

⁴⁸ State-space stock assessment models, with non-linear functions of latent processes and nu-
⁴⁹ merous observation types with different probability distribution assumptions represent one
⁵⁰ of most complex classes of state-space models. The literature on the effects of various fac-
⁵¹ tors on reliability of inferences from state-space assessment models is growing (Li et al., 2024;
⁵² Miller et al., In reviewa). The importance of contrast in population size and fishing mor-
⁵³ tality and quality of data used to fit assessment models including the state-space variety is
⁵⁴ known (Magnusson and Hilborn, 2007; Miller et al., In reviewa). Furthermore, estimation
⁵⁵ of natural mortality, and even temporal variability is possible in many scenarios (Lee et al.,
⁵⁶ 2011; Cadigan, 2016; Miller and Hyun, 2018; Miller et al., In reviewa).

⁵⁷ The effects of temporal variation in recruitment via undefined or explicit environmental
⁵⁸ factors have been extensively investigated in both traditional assessment models and state
⁵⁹ space models (Myers, 1998; Haltuch and Punt, 2011; Johnson et al., 2016; Miller et al., 2016).
⁶⁰ Reliability of estimating environmental and spawning biomass effects on recruitment in state-
⁶¹ space assessment models requires a combination of strong effects, good age composition data

62 quality, contrast in the environmental covariate and lower recruitment variability (Britten
63 et al., In review; Miller et al., In reviewa).

64 A critical aspect of fisheries assessment models and their use in management is short-term
65 projections that are used to determine catch advice. While understanding drivers of re-
66 cruitment is important particularly for subsequent effects on reference points, recruitment
67 in short-term projections typically has little impact on the exploitable biomass in the first
68 few projection years. However, assumptions for natural mortality (M) have immediate and
69 larger effects on projected biomass because they affect the abundances at older age classes
70 at the end of the data time series that constitute spawning biomass and catch (Brodziak
71 et al., 2008; Stock et al., 2021).

72 Because of the effects of M on both biological reference points and short term projections,
73 better understanding sources of variation in M would provide more accurate estimation of
74 abundance and productivity and therefore improved management. Temporal variation in M
75 is less studied than recruitment, but its importance for explaining variability in observations
76 has been demonstrated in state-space assessment models for Atlantic cod and yellowtail
77 flounder (Cadigan, 2016; Stock et al., 2021). Deriso et al. (2008) also demonstrated the
78 importance of several factors affecting Pacific herring natural mortality.

79 Assessment models could include temporal variation in many aspects of population dynamics
80 or how observations are related to the population. For example the Woods Hole Assessment
81 Model (WHAM) can include process errors treated as random effects for transition in cohorts
82 over time, catchability for indices of abundance, selectivity of fishing fleets or indices, move-
83 ment between regions, or in M (Stock and Miller, 2021; Miller et al., In reviewb). However,
84 misspecified temporal population process errors could lead to biased population and stock
85 status estimation, and, therefore, poor fisheries management decisions (Legault and Palmer,
86 2016; Szwalski et al., 2018). Studies of the reliability of inferences we might make regarding
87 the presence of temporal variability in M are limited. Miller et al. (In reviewa) found AIC

88 could accurately distinguish process errors in the changes in abundance for cohorts, but not
89 for those specifically due to M except when uncertainty in population observations (indices,
90 catch and age composition) was low and there was greater temporal variation in natural mor-
91 tality. In their simulation studies looking a models with multiple sources of process error,
92 Li et al. (2024) found including more sources of process error than existed in the operating
93 model was a better model-building approach than excluding them a priori.

94 Here we conduct a simulation study with operating models (OMs) varying by degree of
95 observation error uncertainty, sources of process error, fishing history, temporal variation in
96 environmental covariates, and magnitude of the effect of the covariate on natural mortality.
97 The simulated observations from these OMs are fitted with estimating models that make
98 alternative assumptions for sources of process error, and whether median M and covariate
99 effects are estimated. We evaluate the effects of these factors on convergence of fitted models,
100 whether Akaike's information criterion (AIC) can correctly determine the correct source of
101 process error and correct assumption about covariate effects on natural mortality, and the
102 degree of bias in inferences for relevant parameters and outputs of the assessment model.

103 Methods

104 All of our analyses used the Woods Hole Assessment Model (WHAM) to construct both
105 OMs and EMs (Miller and Stock, 2020; Stock and Miller, 2021; Miller et al., In reviewb).
106 The WHAM package has been used extensively to configure OMs and EMs for several other
107 simulation studies (Stock et al., 2021; Legault et al., 2023; Li et al., 2024; Britten et al.,
108 In review; Li et al., In reviewa) and is used to assess many commercially important stocks
109 in the Northeast U.S. (e.g., NEFSC, 2022a,b, 2024). We used version 1.0.6.9000, commit
110 77bbd94 for to generate all results.

111 We completed a simulation study with a 288 operating models. The factors defining the
112 configuration of each operating model which are described in detail in subsequent sections

113 include source of population process error (3 levels), index and catch observation uncertainty
114 (2 levels), environmental covariate uncertainty (2 levels), latent environmental covariate
115 process stochasticity (4 levels), and fishing history (2 levels). We simulated 100 data sets for
116 each operating model that included simulations of process errors.

117 For each simulated data set we fit a set of 12 estimating models (EMs). The factors that
118 distinguish the estimating models which are also described in detail below include source
119 of population process error type (3 levels) whether (median) M was estimated or assumed
120 known (2 levels), and whether the effect of the environmental covariate on M was estimated
121 or not (2 levels).

122 The sources of population process error that were used in the OMs or assumed in the EMs
123 were on recruitment only (R), recruitment and changes in cohort abundance over time (R+S),
124 or recruitment and M (R+M). We did not use the log-normal bias-correction feature for pro-
125 cess errors or observations described by Stock and Miller (2021) for operating and EMs (Li
126 et al., In reviewb). Simulations were all carried out on the University of Massachusetts Green
127 High-Performance Computing Cluster. Code for completing the simulations and summariz-
128 ing results can be found at https://github.com/timjmiller/SSRTWG/ecov_study/mortality.

129 Operating models

130 Environmental covariate

131 In the WHAM model, environmental covariates are assumed to be described as state-space
132 processes with annual observations of the true latent covariate (Miller et al., 2016; Stock
133 and Miller, 2021). In our simulations, the latent covariate is assumed to be a stationary first
134 order autoregressive (AR1) process

$$X_y | X_{y-1} \sim N \left(\mu_E (1 - \rho_E) + \rho_E X_{y-1}, (1 - \rho_E^2) \sigma_E^2 \right)$$

¹³⁵ with marginal mean $\mu_E = 0$ and variance σ_E^2 . The four configuration of the latent envi-
¹³⁶ ronmental covariate in the operating models assume the one of two values for the marginal
¹³⁷ standard deviation $\sigma_E \in \{0.1, 0.5\}$ and for the autocorrelation parameter $\rho_E \in \{0, 0.5\}$.

¹³⁸ The observations of the latent environmental covariate are assumed to be unbiased and
¹³⁹ Gaussian

$$x_y | X_y \sim N(X_y, \sigma_e^2)$$

¹⁴⁰ The standard deviation of the environmental observations in the operating models is one of
¹⁴¹ two values $\sigma_e \in \{0.1, 0.5\}$. Figure S2 provides example simulations of the latent process and
¹⁴² observations under the alternative configurations.

¹⁴³ Population

¹⁴⁴ Many of the characteristics of the population biology and structure including the age classes
¹⁴⁵ (10 age classes (ages 1 to 10+)), time span (40 years), maturity (Figure S1, top left), growth
¹⁴⁶ (Figure S1, top right), time of spawning (1/4 of the year), and recruitment (Figure S1,
¹⁴⁷ bottom right) are identical to Miller et al. (In reviewa). The maturity at age is a logistic
¹⁴⁸ function of age with age at 50% maturity ($a_{50} = 2.89$) and slope = 0.88 and weight at age
¹⁴⁹ is derived from a von Beralanffy growth function where $t_0 = 0$, $L_\infty = 85$, and $k = 0.3$, and
¹⁵⁰ a length-weight relationship

$$W_a = \theta_1 L_a^{\theta_2}$$

¹⁵¹ where $\theta_1 = e^{-12.1}$ and $\theta_2 = 3.2$.

¹⁵² The general model for M in year y is a log-linear function of both covariate effects X_y and
¹⁵³ process errors $\varepsilon_{M,y}$ and a parameter β_M that defines median M

$$\log M_y = \beta_M + \beta_E X_y + \varepsilon_{M,y}$$

¹⁵⁴ where the process errors are modeled as random effects that may, in general, be autocorre-

155 lated normal random variables

$$\varepsilon_{M,y} | \varepsilon_{M,y-1} \sim N(\varepsilon_{M,y-1}, (1 - \rho_M^2) \sigma_M^2)$$

156 (Stock and Miller, 2021), but we assume $\rho_M = 0$ in our R+M OMs. We assume the median
157 M rate $e^{\beta_M} = 0.2$ is constant across ages. For R and R+S OMs and EMs, $\varepsilon_{M,y} = 0$. For
158 all R+M OMs, we assume the same standard deviation $\sigma_M = 0.3$ and is estimated in the
159 R+M EMs. The covariate effect is one of 3 alternative values in the operating models,
160 $\beta_E \in \{0, 0.25, 0.5\}$. The parameters defining the simulated covariate time series, size of the
161 covariate effect, and any M random effects result in a range of different levels of variation
162 in annual values (Figure S3).

163 We assumed expected recruitment each year is from a Beverton-Holt stock-recruit relation-
164 ship (SRR)

$$R_y = \frac{aSSB_{y-1}}{1 + bSSB_{y-1}}.$$

165 All biological inputs to calculations of spawning biomass per recruit (i.e., weight, matu-
166 rity, and M at age) are constant in the R and R+S OMs without covariate effects on M .
167 Therefore, steepness and equilibrium unfished recruitment are also constant over the time
168 period for those OMs (Miller and Brooks, 2021). As in Miller et al. (In reviewa), our as-
169 sumed biological inputs and selectivity (defined below) with constant M result in equilibrium
170 fishing mortality that reduces spawning biomass per recruit to 40% of the unfished level is
171 $F_{40\%} = 0.348$. With an assumed unfished recruitment of $R_0 = e^{10}$, setting $F_{MSY} = F_{40\%}$
172 results in a steepness of 0.69 and $a = 0.60$ and $b = 2.4 \times 10^{-5}$. For R+M OMs and all
173 OMs with covariate effects on M , steepness is not constant, but we used the same a and b
174 parameters as other operating models which equates to a steepness and R_0 at the median of
175 the time series processes for M and the covariate.

176 We also used the same two fishing scenarios as Miller et al. (In reviewa) for OMs. In the
177 first scenario, the stock experiences overfishing at $2.5F_{MSY}$ for the first 20 years followed by

178 fishing at F_{MSY} for the last 20 years (denoted $2.5F_{MSY} \rightarrow F_{MSY}$). In the second scenario, the
179 stock is fished at F_{MSY} for the entire time period (40 years).

180 We configured all R, R+S, and R+M OMs with uncorrelated random effects on recruitment
181 with standard deviation on log(recruitment) $\sigma_R = 0.5$. This same assumption was used by
182 Miller et al. (In reviewa) for R+M OMs and other OMs with fishery selectivity and index
183 catchability process errors. For R+S OMs, cohort temporal transition process errors were
184 uncorrelated with $\sigma_{2+} = 0.3$

185 Catch and index observations

186 We define the generation of observations of total catch, aggregate indices, and corresponding
187 age composition identical to Miller et al. (In reviewa). There is a single fleet operating
188 year round for catch observations with logistic selectivity for the fleet with $a_{50} = 5$ and
189 slope = 1 (Figure S1, bottom left). Observations are generated for all 40 years of the
190 model. There are two index time series intended to represent fishery-independent surveys
191 occurring in the spring (0.25 way through the year) and the fall (0.75 way through the
192 year). Catchability of both surveys are assumed to be 0.1. We assumed catch and index
193 age composition observations aer generated from a logistic-normal distribution where errors
194 on the multivariate normal scale are independent. The standard deviation parameter is also
195 constant across ages.

196 Standard deviation for log-aggregate catch was 0.1. There were two levels of observation error
197 variance for indices and age composition for both indices and fleet catch. A low uncertainty
198 specification assumed standard deviation of both series of log-aggregate index observations
199 was 0.1 and the standard deviation of the logistic-normal for age composition observations
200 was 0.3. In the high uncertainty specification the standard deviation for log-aggregate indices
201 was 0.4 and that for the age composition observations was 1.5. For all estimating models,
202 standard deviation for log-aggregate observations was assumed known whereas that for the

203 logistic-normal age composition observations was estimated.

204 **Estimating models**

205 For each data set simulated from an operating model 12 estimating models were fit. There
206 were three factors defining the configuration of each estimating model: 1) whether β_M was
207 estimated or assumed known, 2) whether an environmental effect β_E was estimated or not,
208 and 3) whether the process errors were assumed on recruitment only (R), recruitment and
209 survival (R+S), or recruitment and M (R+M).

210 The configuration of the process errors in the estimating models generally matched the
211 corresponding options in the operating models. For example, uncorrelated R+S was assumed
212 for both the estimating and operating model. However, R+M EMs did not assume M random
213 effects were uncorrelated (ρ_M was estimated). The environmental covariate observations
214 were included in all estimation models to ensure comparability of AIC. All fixed effects
215 parameters for selectivity, catchability, fully-selected fishing mortality, mean recruitment,
216 initial abundance at age, and variances for logistic-normal age composition distributions
217 were estimated. Any process error variance parameters for recruitment, survival, and M
218 were also estimated. The observation error variance of the environmental observations and
219 aggregate catch and indices were all assumed known at the true values.

220 **Performance measures**

221 **EM convergence**

222 We measured the frequency of convergence when fitting each EM to the simulated data
223 sets for each OM. There are various ways to assess convergence of the fit (e.g., Carvalho
224 et al., 2021), but we defined successful convergence as the Hessian of the marginal log-
225 likelihood being invertible and providing variance estimates for the fixed effects parameters

226 as recommended by Miller et al. (In reviewa). We calculated 95% confidence intervals for
227 probability of convergence using the Clopper-Pearson exact method (Clopper and Pearson,
228 1934; Thulin, 2014).

229 AIC for model selection

230 We measured the frequency of correct model selection using marginal AIC. For a given
231 operating model the set of models that were considered all made the same assumptions on
232 whether or not to estimate β_M or it is assumed at the true value. For model m , the marginal
233 AIC is a function of the marginal log-likelihood maximized with respect to the fixed effects
234 in the model $\boldsymbol{\theta}$ and the number of fixed effects $n(\boldsymbol{\theta})$ estimated,

$$\text{AIC}_m = -2 [\text{argmax}_{\boldsymbol{\theta}} \log L_m(\boldsymbol{\theta}) - n(\boldsymbol{\theta})].$$

235 All model fits that successfully completed the optimization were used in these results. We
236 used all of these fits because some lack of convergence would be expected for the correct
237 behavior of more complicated models that include process errors that did not exist in the
238 operating model. For example R+M EMs fit to R OMs would be expected to estimate no
239 variance in the M random effects and the estimated variance parameter going to zero would
240 cause poor convergence.

241 Parameter estimation

242 All results here use OM simulations with fits that satisfied the convergence criterion described
243 above. We used this conditioning, to reflect how practitioners would proceed in analyses of
244 model fits with real assessment data. That is, practitioners would ensure models converged
245 such that Hessian-based standard errors were available for all model parameter estimates.

246 We focused on statistical behavior of estimators of the covariate effect on M ($\hat{\beta}_E$), the esti-

247 mator of the median M parameter ($\hat{\beta}_M$), and estimators of terminal year natural mortality
 248 \widehat{M} , spawning stock biomass $\widehat{\text{SSB}}$, and fully-selected fishing mortality rate \widehat{F} . In prelimi-
 249 nary analyses we results of the estimators of all the annual values for M , SSB and fishing
 250 mortality over the whole time series, but we found no appreciable differences in patterns
 251 across the various factors defining the OMs and EMs. Furthermore, results for terminal year
 252 fishing mortality were generally inversely related to those for spawning stock biomass, and
 253 are provided in the Supplementary Materials and not discussed further.

254 We calculated median errors (ME) of $\hat{\beta}_E$, $\hat{\beta}_M$, and the Hessian-based standard error esti-
 255 mators of these parameters ($\widehat{SE}(\hat{\beta}_E)$ and $\widehat{SE}(\hat{\beta}_M)$). We calculated the median relative
 256 errors (MRE) of terminal year \widehat{M} , $\widehat{\text{SSB}}$, and \widehat{F} . We also calculated the root mean square
 257 error (RMSE) and estimated probability of coverage of constructed 95% confidence intervals
 258 for $\hat{\beta}_E$ and $\hat{\beta}_M$ for EMs that estimated these parameters. We constructed the confidence
 259 intervals for probabilities of CI coverage using the same methods as those for probabilities
 260 of convergence.

261 The true values for terminal year SSB and fishing mortality rate vary among simulations.
 262 For the i th simulated data the relative error for terminal year value θ_i provided from the
 263 fitted estimation model is

$$\text{RE}_i(\theta) = \frac{\hat{\theta}_i - \theta_i}{\theta_i}$$

264 We calculated RMSE as

$$\text{RMSE}(\hat{\theta}) = \sqrt{\frac{1}{n} \sum_{i=1}^n (\hat{\theta}_i - \theta_i)^2}$$

265 For ME and RME of estimators, we constructed 95% confidence intervals using the binomial
 266 distribution approach as in Stock and Miller (2021) and Miller et al. (In reviewa). Because
 267 of the inverse relationship of the the confidence interval and the null hypothesis significance
 268 test, when the confidence interval contains zero we do not have evidence of bias from the
 269 simulation study.

270 **Results**

271 Miller et al. (In reviewa) found inferences are most reliable for distinguishing process error
272 sources and stock-recruit relationships in scenarios with lower observation error for indices
273 and age composition data and with temporal contrast in fishing pressure. Our expectation is
274 that inferences for covariate effects on M would require data to be at least that informative
275 and inspection of the comprehensive set of results across all OMs generally confirms this (See
276 supplemental Materials). Therefore, we restrict our attention to these OM scenarios with
277 more informative data in the main paper (temporal contrast in fishing and lower population
278 observation error).

279 **EM Convergence**

280 In OMs with contrast in fishing pressure and lower uncertainty in population observations
281 (catch, indices and age composition), R EMs generally converged with high frequency except
282 when EMs estimated covariate effects on M and the covariate uncertainty (σ_e) was high
283 (Figure 1). R+S EMs generally converged with low frequency when the OM process error
284 did not match the EM configuration. R+M EMs generally did not converge with high
285 frequency except for R+S OMs with low covariate uncertainty and where the EMs assumed
286 median M as known. There was no strong effect of the treatment of median M (known or
287 estimated) on convergence for any of the other EMs. The largest effect of the size of the
288 true covariate effect (β_E) on convergence was observed for R and R+M EMs that estimated
289 covariate effects in R and R+M OMs with high uncertainty in covariate observations (σ_e)
290 and high variability in covariate process errors (σ_E). There was an increase in convergence
291 frequency with increased variability in covariate process errors for many EMs in OMs with
292 large covariate uncertainty.

²⁹³ **AIC performance**

²⁹⁴ We only present results for EMs where the median M rate parameter was estimated because
²⁹⁵ results differed little whether this parameter was estimated or assumed known (see Figures
²⁹⁶ S7 to S9). When there is temporal contrast in fishing pressure and lower uncertainty in
²⁹⁷ population observations, the only OM scenarios where AIC was consistently accurate for
²⁹⁸ both covariate effect and process error configurations were R OM with high variability in
²⁹⁹ the covariate and low uncertainty in observations of the covariate (Figure 2). R+M OM
³⁰⁰ with those same covariate attributes were only accurate for the covariate effect (low Type I
³⁰¹ and II errors) and the incorrect process error almost always chosen was R rather than R+S.
³⁰² We also observed high accuracy for both covariate effect and process error configurations for
³⁰³ R+S OM where either no effect of the covariate or the strongest effect of the covariate was
³⁰⁴ simulated. All R and R+S OM were accurate for process error configuration and across
³⁰⁵ all OM, AIC accurately selects models without covariate effects when there are none (low
³⁰⁶ Type I error). See Figures S7 to S9 for full results.

³⁰⁷ **Covariate effect estimation**

³⁰⁸ When there is temporal contrast in fishing pressure, lower uncertainty in population obser-
³⁰⁹ vations, low uncertainty in environmental observations, and larger temporal contrast in the
³¹⁰ simulated true environmental covariate, we observed generally little or no bias in estimation
³¹¹ of covariate effects ($\hat{\beta}_E$) across all EM and OM process error assumptions, both EM config-
³¹² urations of median M , and all true covariate effect sizes (Figure 3, rows 3 and 4). The only
³¹³ exception to this was when R EMs with median M estimated were fit to R+S OM with
³¹⁴ those configurations. Decreasing trends in bias of $\hat{\beta}_E$ for EMs fit to many configurations of
³¹⁵ R and R+M OM indicate that $\hat{\beta}_E$ values were closer to zero on average even when the true
³¹⁶ effect increased. For R+S OM, EMs with the matching process error configuration showed
³¹⁷ little evidence of bias in estimating β_E across a range of covariate and covariate observation

318 configurations.

319 Bias of Hessian-based standard error estimation for $\hat{\beta}_E$ was also close to zero in the same OM
320 scenarios where bias in estimation of β_E itself was close to zero (Figure S13, rows 3 and 4).

321 However, in these high information OMs bias of standard error estimation was observed in
322 for several EMs that did not have the matching R+S and R+M process error configuration.

323 Estimation of standard errors was essentially unreliable for most EMs when OMs had higher
324 uncertainty in covariate observations (Figure S13, rows 5 to 8).

325 Despite the reliable estimation of covariate effects and standard errors with low uncertainty
326 in environmental observations, and larger temporal contrast in the simulated true environ-

327 mental covariate, we found poor confidence interval coverage for many EMs fitted to R+S
328 and R+M OMs (Figure 4, rows 3 and 4). For example, R+M EMs fit to R+M OMs showed

329 little bias in estimation of the covariate effect and standard errors, but confidence interval
330 coverage was negatively biased (rows 3 and 4 in right two columns of Figures 3, S13, and

331 4. However, there appears to be a positive correlation of the covariate and standard error
332 estimates such that estimates less than the true value had negatively biased standard error

333 estimates which would make confidence intervals too small (Figure S24). Confidence inter-
334 val coverage was most reliable for R+S OMs when the EMs had the matching process error

335 configuration (Figure 4, middle 2 columns).

336 Median natural mortality rate estimation

337 We observed negligible bias in estimating the median M parameter (β_M) over a much wider
338 set of operating models than that for estimating the covariate effect on M . When there is

339 temporal contrast in fishing pressure, lower uncertainty in population observations, all EMs
340 fit to R and R+M OMs showed little evidence of bias for $\hat{\beta}_M$ (Figure 5, left and right sets

341 of columns). For R+S OMs with those conditions, EMs with the matching process error
342 assumption showed little evidence of bias.

343 We also found little or no bias of Hessian-based standard error estimation for β_M in the many
344 of the same OM scenarios where bias of $\hat{\beta}_M$ was close to zero except for some EMs with the
345 process error source mis-specified or where the OM simulated the strongest covariate effect
346 (Figure S28). Like $\hat{\beta}_M$, reliability of standard error estimation in R+S OMs required the
347 EMs to have the matching process error assumption.

348 However, we found confidence interval coverage to be generally unreliable even for OM and
349 EM combinations where bias in $\hat{\beta}_M$ and its standard error estimators was negligible, much
350 like that for $\hat{\beta}_E$ (Figure 6). In the same EM OM combination we investigated above for $\hat{\beta}_E$,
351 we observed an opposite negative correlation of $\hat{\beta}_M$ and its standard error estimates (Figure
352 S35), which would result in confidence intervals being too narrow when $\hat{\beta}_M$ values are larger
353 than average.

354 Terminal year natural morality rate

355 The EMs with R and R+S process error configurations with median M assumed known and
356 covariate effects not estimated will have terminal year M correctly specified when fitted to
357 R and R+S OMs with no covariate effect because M is constant in both the EM and OM
358 at the same value. We exclude those OM-EM combinations from our results here.

359 In many of the OMs with temporal contrast in fishing pressure and lower uncertainty in
360 population observations, there was little difference in bias for terminal year M among EMs
361 whether the median M parameter (β_M) was estimated (Figure 7). However, when there were
362 differences, bias in terminal year M was closer to zero when the EM assumed β_M known,
363 As would be expected. For R OMs, all estimating models exhibited little evidence of bias
364 except when OMs simulated the largest covariate effect ($\beta_E = 0.5$). However, bias estimates
365 were at worst around 10% (Figure 7, left columns). For R+S OMs, the R+S EMs generally
366 exhibited the least evidence for bias and R+M EMs generally estimated terminal M lower
367 than the true value, particularly when observation uncertainty in the covariate was lower.

368 For R+M OMs, many of the EM configurations provided little bias in terminal \widehat{M} , but
369 the biases were generally further from zero than the results for the R OMs (Figure 7, right
370 columns). Accuracy of terminal year \widehat{M} was similar for all of the EMs when fitted to R+M
371 OMs (Figure S43). For R+S OMs, accuracy was a bit better for R+S EM than others. As
372 would be expected accuracy of estimate where median M was assumed known was better
373 than when it was estimated, when there were differences.

374 Terminal year spawning stock biomass

375 In OMs with temporal contrast in fishing pressure and lower uncertainty in population
376 observations, we found little bias in estimates of terminal year SSB for all EM configurations
377 when fitted to R or R+M OMs (Figure 8). We also found little bias for R+S OMs when the
378 EMs assumed the correct process error configuration. The worst bias in terminal year SSB
379 estimation we found was in R+S OMs for R EMs with β_M estimated. Like the bias results,
380 accuracy as measured by RMSE was similar for many of EMs when fit to R and R+M OMs.
381 Accuracy was generally worse when EMs were fit to R+S OMs (Figure S50).

382 Discussion

383 Our simulation study demonstrated that estimation of environmental effects on M is possible
384 and reliable in certain scenarios even when the process error was misspecified (e.g., R and
385 R+S EMs fit to R+M OM). In many of these same OMs frequency of convergence of model
386 fits did not appear to suffer when covariate effects on M were estimated even when there
387 was no effect simulated. However, these scenarios are information rich in that there was
388 contrast in population size (due to changes in fishing pressure) and the covariate affecting
389 the population, and low uncertainty in population and covariate observations. Previous
390 research has shown that estimation of a constant M parameter requires contrast in time

391 series and informative data (Lee et al., 2011), so it is no surprise that estimation of these
392 effects also requires relatively good information via more precise observations and higher
393 contrast in the covariate time series.

394 On the other hand, AIC could only reliably detect covariate effects for R and R+M OMs
395 with contrast in covariates and low covariate uncertainty and confidence interval coverage
396 for the covariate effect was often biased even when bias in the estimators of the effect and the
397 standard error were negligible. Cadigan et al. (2024) found confidence interval coverage to
398 be biased for SSB and fishing mortality estimation in a state-space model in some scenarios,
399 but they attributed the poor coverage to bias in Hessian-based standard error estimation
400 and their simulations held any process error random effects constant. The coverage bias
401 we observed, at least for some OM-EM combinations, may be a related to correlation of
402 the estimators of the covariate effect and the corresponding standard error and therefore
403 consideration of other methods of calculating confidence intervals may be warranted (e.g.,
404 those based on profile likelihood and/or MCMC sampling of log-likelihood surface).

405 Miller et al. (In reviewa) investigated R+M OMs with two levels of M process error variability
406 ($\sigma_M \in 0.1, 0.5$ and only found AIC able to accurately distinguish R+M process errors with
407 the higher level of process error variability ($\sigma_M = 0.5$). We assumed $\sigma_M = 0.3$, intermediate
408 to the values investigated by Miller et al. (In reviewa), and found process error inferences
409 unreliable for the source of process error, indicating that the level of variability required for
410 detecting M process errors must be greater than $\sigma_M = 0.3$, but may still be less than 0.5. In
411 our results and those by Miller et al. (In reviewa), AIC typically chose R EMs which would
412 indicate the fitted R+M EMs would estimate no variability in the M process errors. Future
413 studies like ours where R+M OMs are simulated with greater variability in M process errors
414 would better inform reliability of covariate effect inferences under such scenarios.

415 Deriso et al. (2008) attempted to estimate process errors as well as covariate effects with
416 M for Pacific herring, but similarly found no variability in M , suggesting a there was too

417 much uncertainty in the available observations relative to the true temporal variability in
418 M . Given that we found covariate effect inferences using R EMs was reliable in R+M OMs
419 with apparently little variability in M process errors, the findings of Deriso et al. (2008) on
420 covariate effects for Pacific herring would presumably also be robust to true low variability in
421 M . However, they did not account for uncertainty in covariate observations, some of which
422 would probably have substantial uncertainty (e.g., competition and predation covariates).
423 We found higher covariate observation uncertainty to cause true covariate effects to be less
424 detectable using AIC, but we did not investigate the implications for incorrectly assuming
425 no covariate uncertainty for covariate inferences.

426 Any bias or poor accuracy for annual SSB estimation was primarily a function of whether or
427 not the median M parameter was estimated or known and the types of process errors than
428 the treatment of the covariate effects on M . For example, we found estimation of SSB was
429 better when the EM had the process error correctly specified for R+S OMs. Fortunately,
430 our results and those by Miller et al. (In reviewa) demonstrate that marginal AIC seems to
431 be a good tool for determining whether this source of process error should be included in the
432 model. However, the reliability of the estimation of SSB does break down in the less ideal
433 scenarios when there is higher population observation error, and lack of contrast in fishing
434 pressure (e.g., Figures S47 to S49).

435 The ability to accurately infer covariate effects on M in some realistic situations indicates
436 that such investigations may be fruitful. Ability to make inferences could improve further
437 when WHAM is extended to incorporate tagging data as expected (Miller et al., In reviewb).
438 Tagging data can greatly inform natural mortality estimation (Pollock et al., 1991; Hampton,
439 2000), and this impact on M estimation should also apply to estimation of covariate effects
440 or unexplained temporal variation in M . Given our findings coupled with future WHAM
441 development, we expect investigations of and accounting for covariate effects on M to become
442 more common within the fisheries stock assessment process.

⁴⁴³ **Acknowledgements**

⁴⁴⁴ This work was funded by NOAA Fisheries Northeast Fisheries Science Center.

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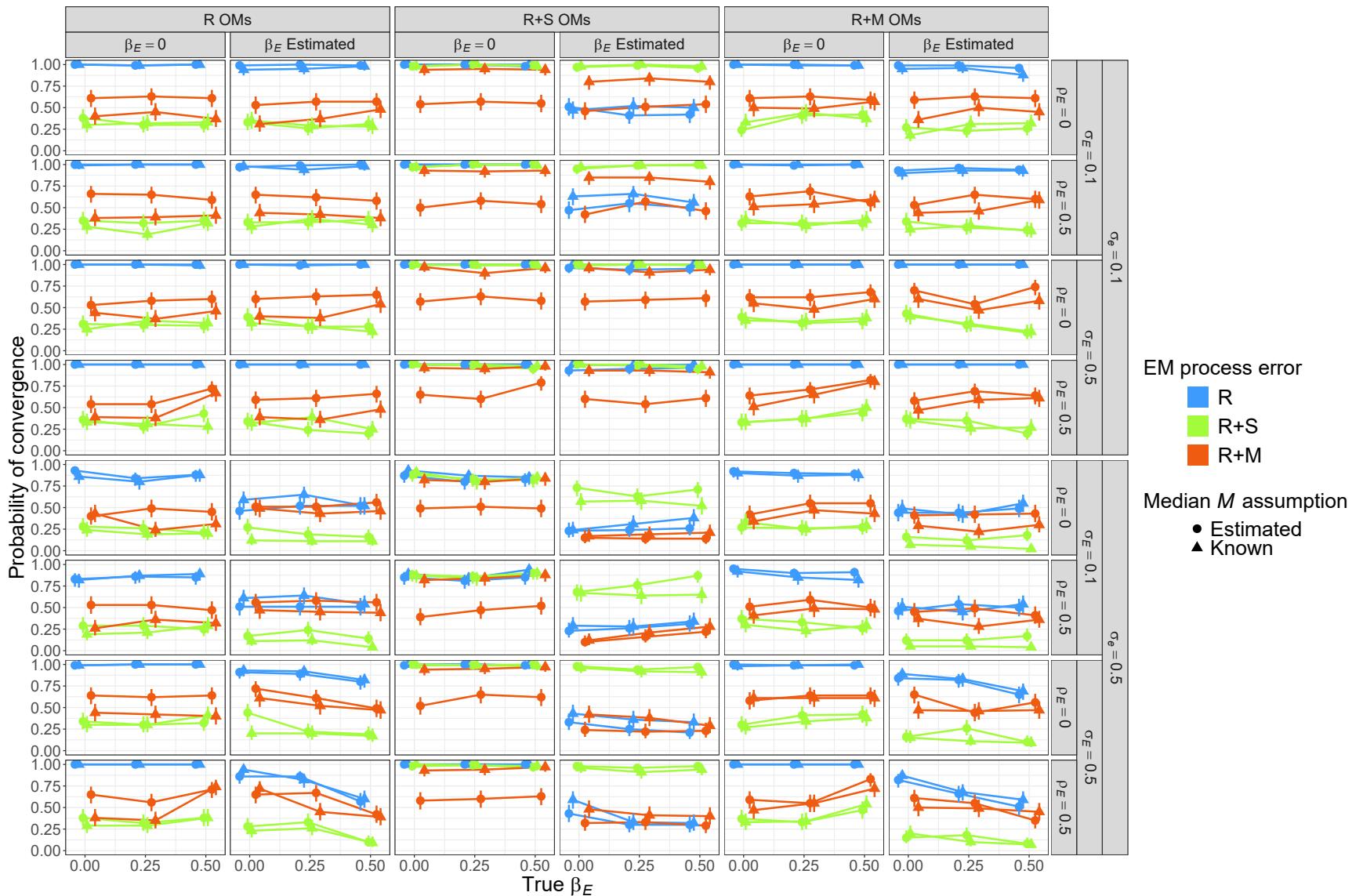


Fig. 1. Estimated probability of fits providing Hessian-based standard errors for EMs with alternative process error assumptions, treatment of median natural mortality (e_M^β known or estimated), and treatment of covariate effect ($\beta_E = 0$ or estimated). The OMs have R (left) and R+S (middle), or R+M (right) process error structures, alternative configurations of covariate time series structure and levels of observation uncertainty (rows), and three levels of true covariate effect on median natural mortality (x axis). All OMs had low observation error for fish population observations and temporal contrast in fishing pressure. Vertical lines represent 95% confidence intervals.

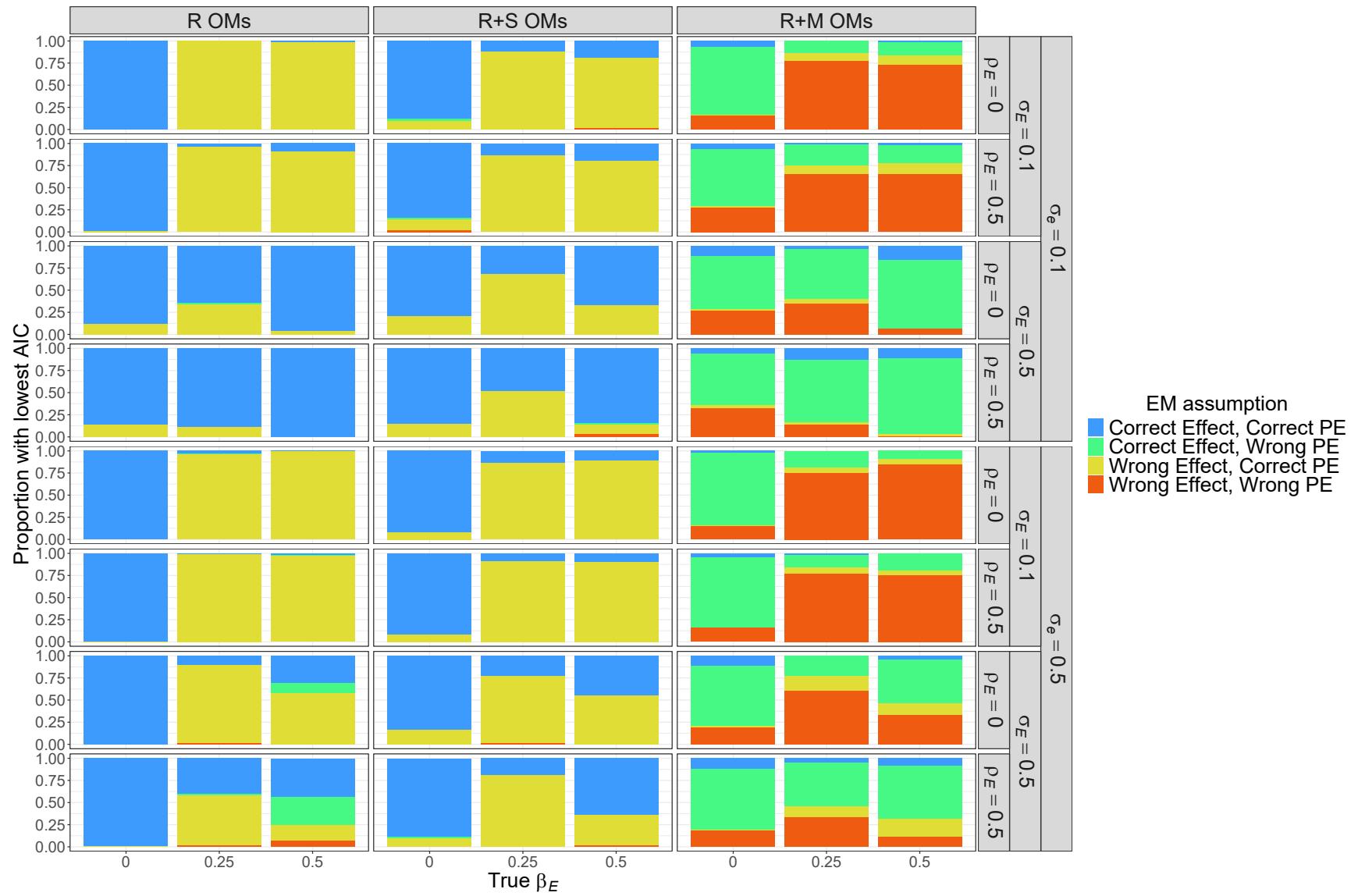


Fig. 2. For each OM, the proportion of simulated data sets where the EM type (treatment of environmental covariate effect and assumed process error type) had the lowest AIC. All OMs had low observation error for fish population observations and temporal contrast in fishing pressure. All EMs estimated median natural mortality rate.

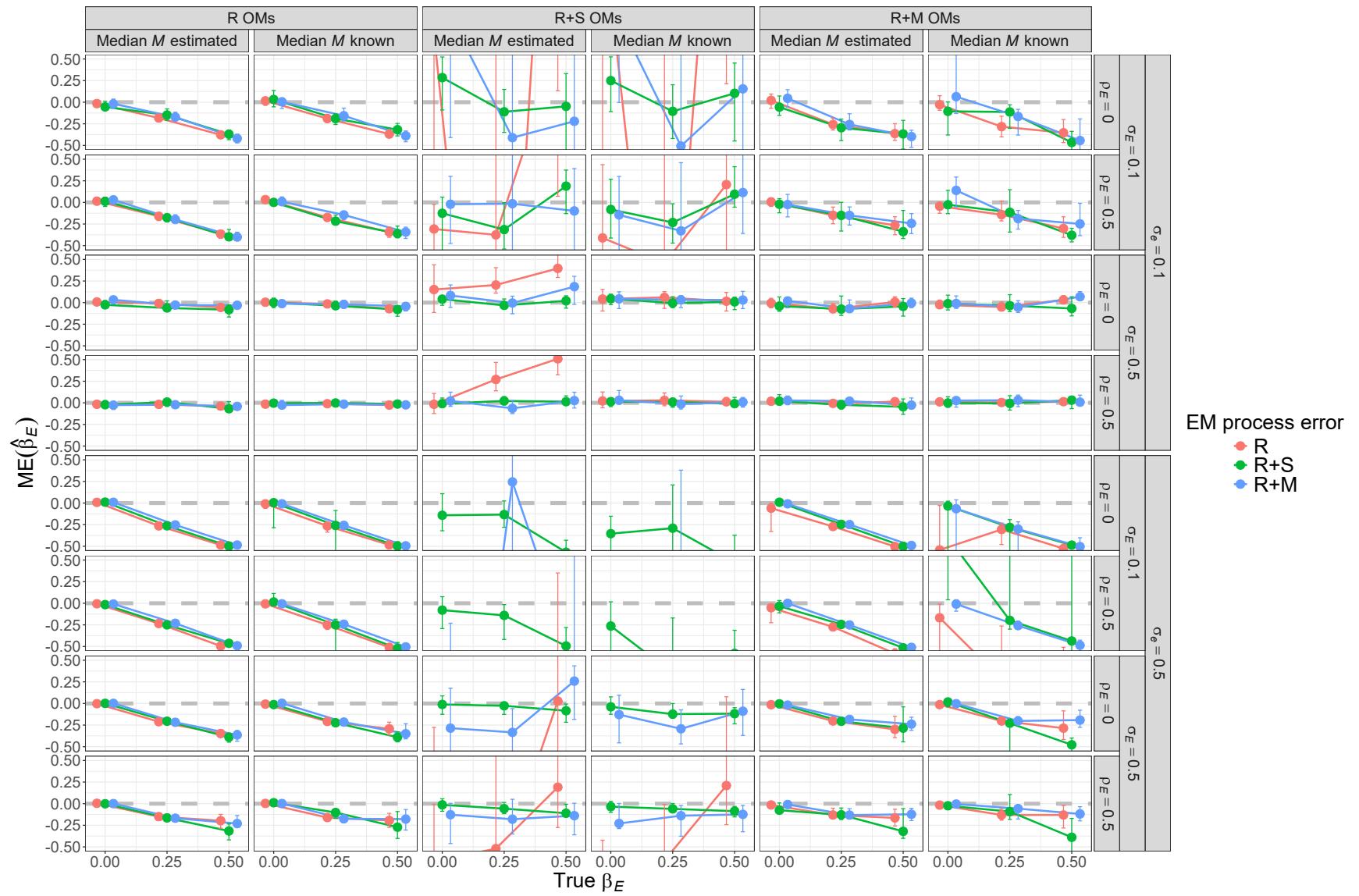


Fig. 3. Median error (ME) of estimates of environmental effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). All OMs had low observation error and contrast in fishing mortality. Vertical lines represent 95% confidence intervals.

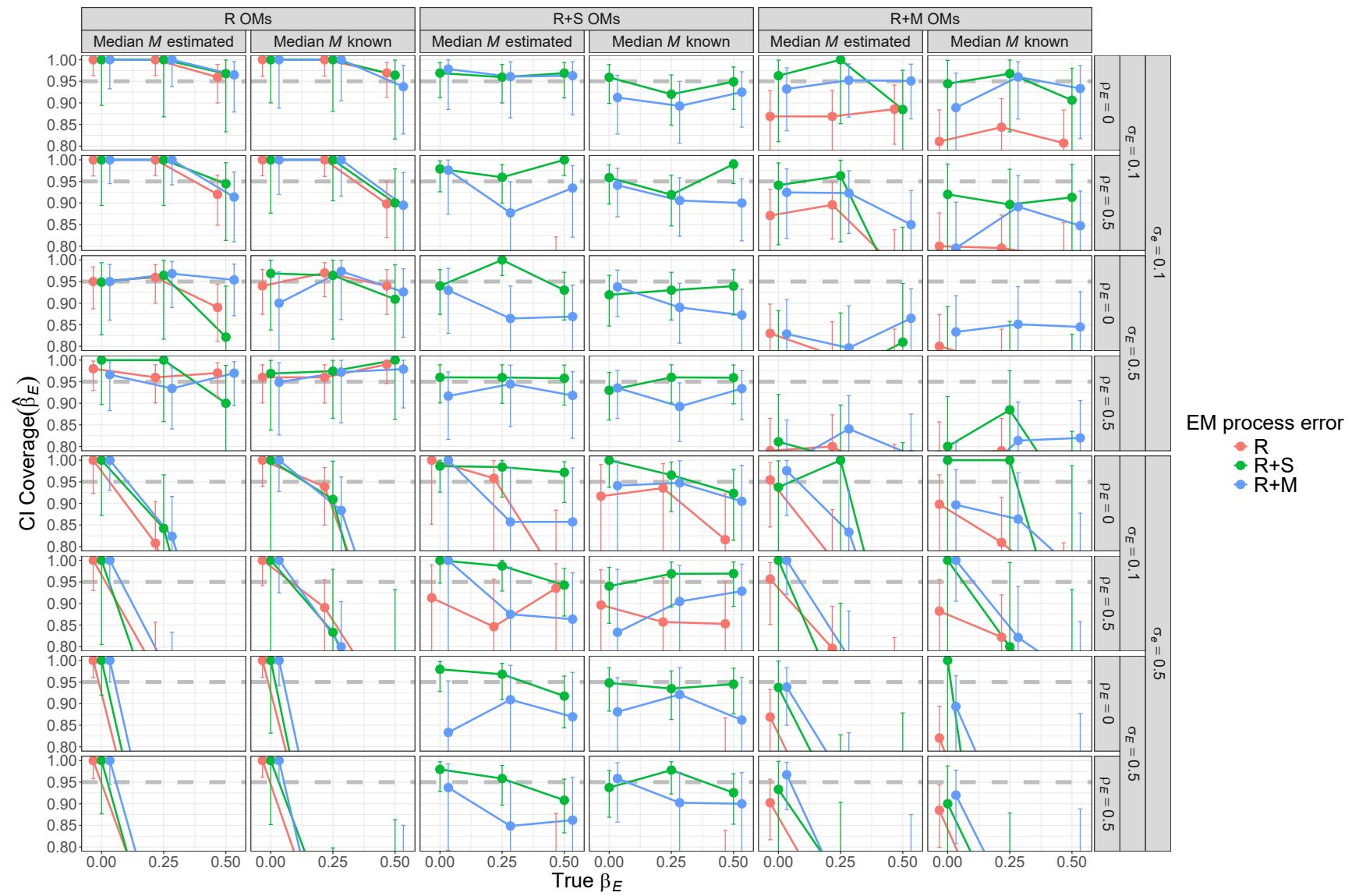


Fig. 4. Probability of 95% confidence interval for β_E containing the true value for EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). All OMs had low observation error and contrast in fishing mortality. Vertical lines represent 95% confidence intervals.

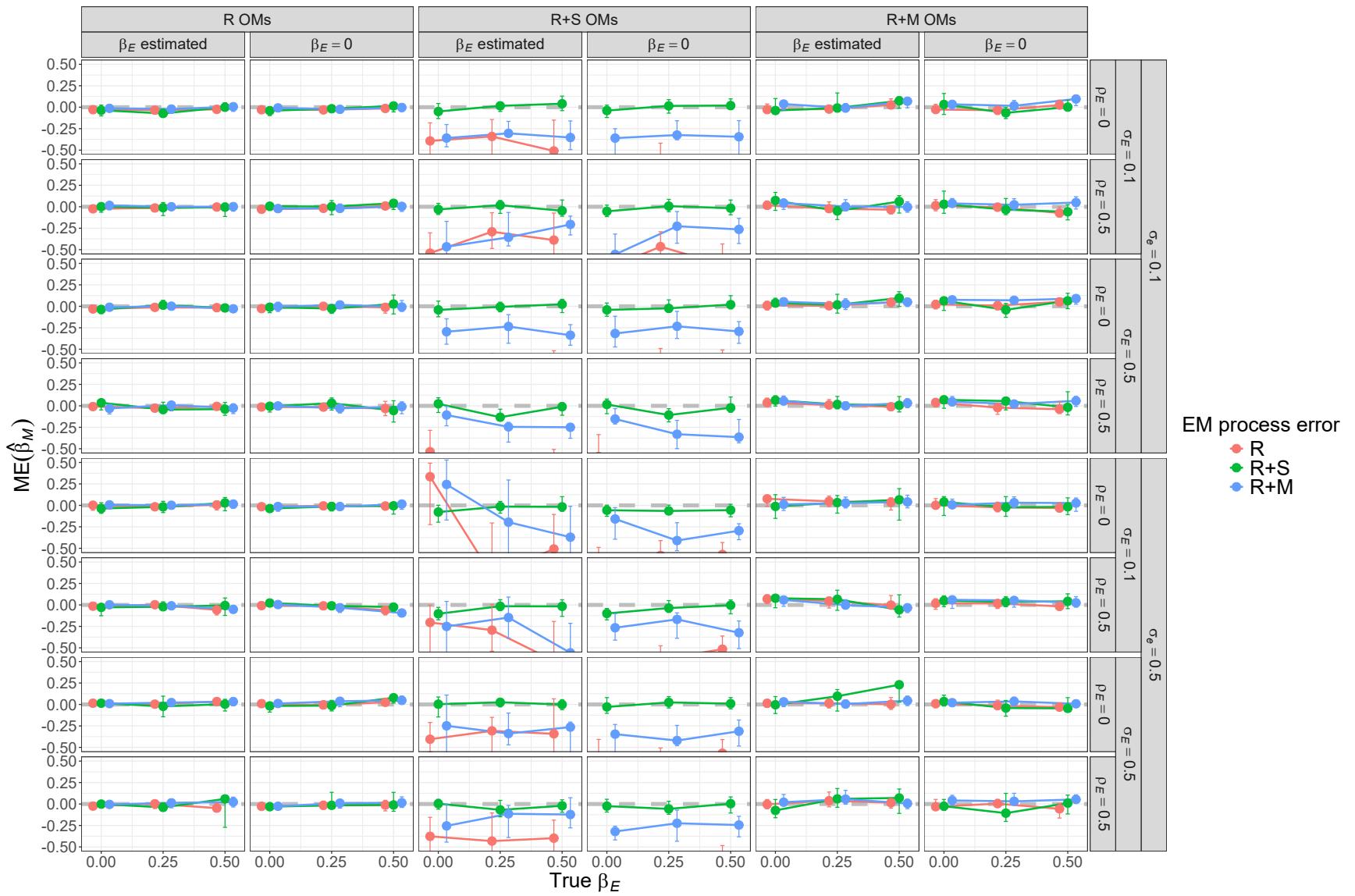


Fig. 5. Median error (ME) of estimates of β_M from fitting EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated). All OMs had low observation error and contrast in fishing mortality. Vertical lines represent 95% confidence intervals.

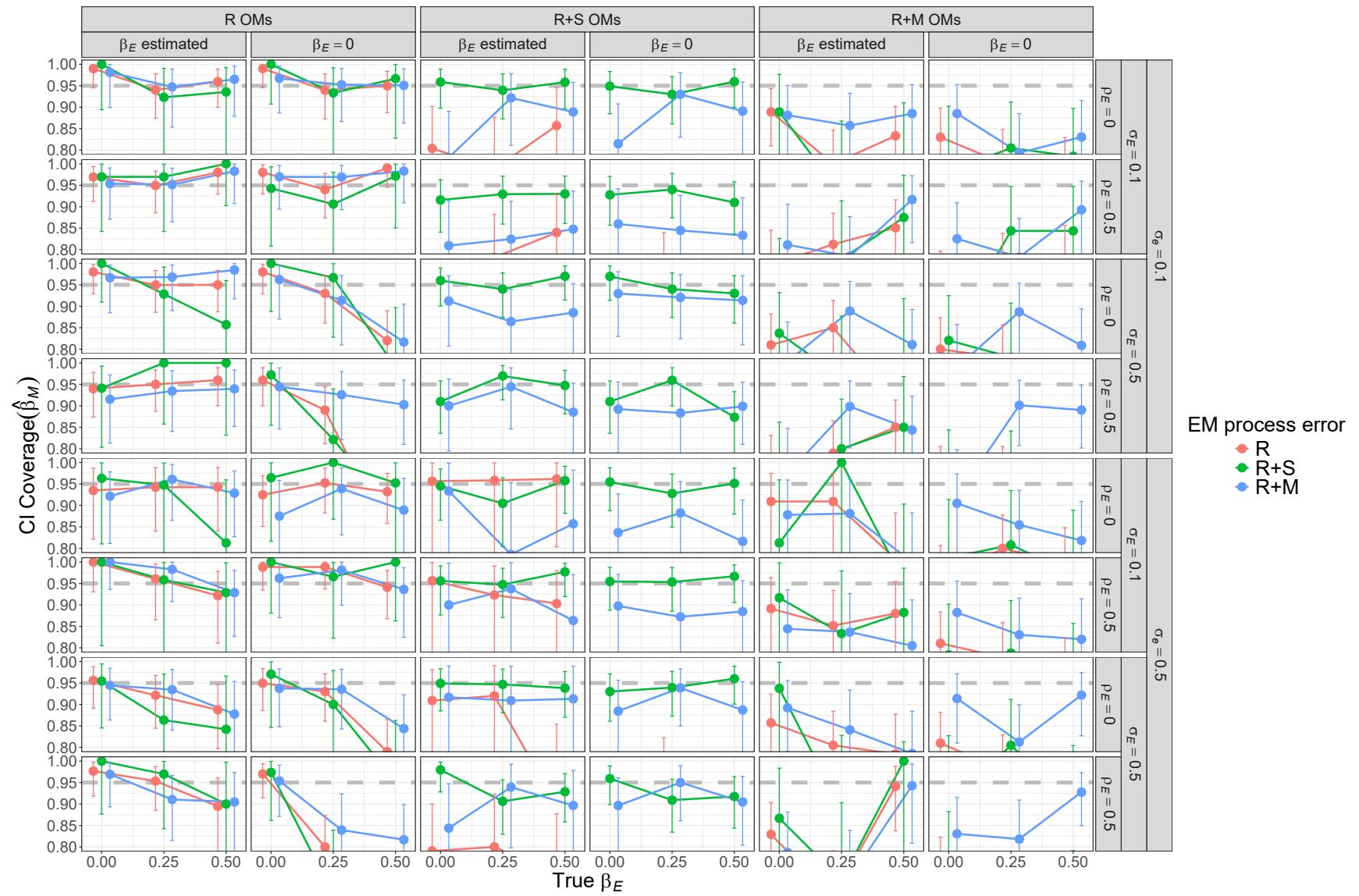


Fig. 6. Probability of 95% confidence interval for β_M containing the true value for EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated). All OMs had low observation error and contrast in fishing mortality. Vertical lines represent 95% confidence intervals.

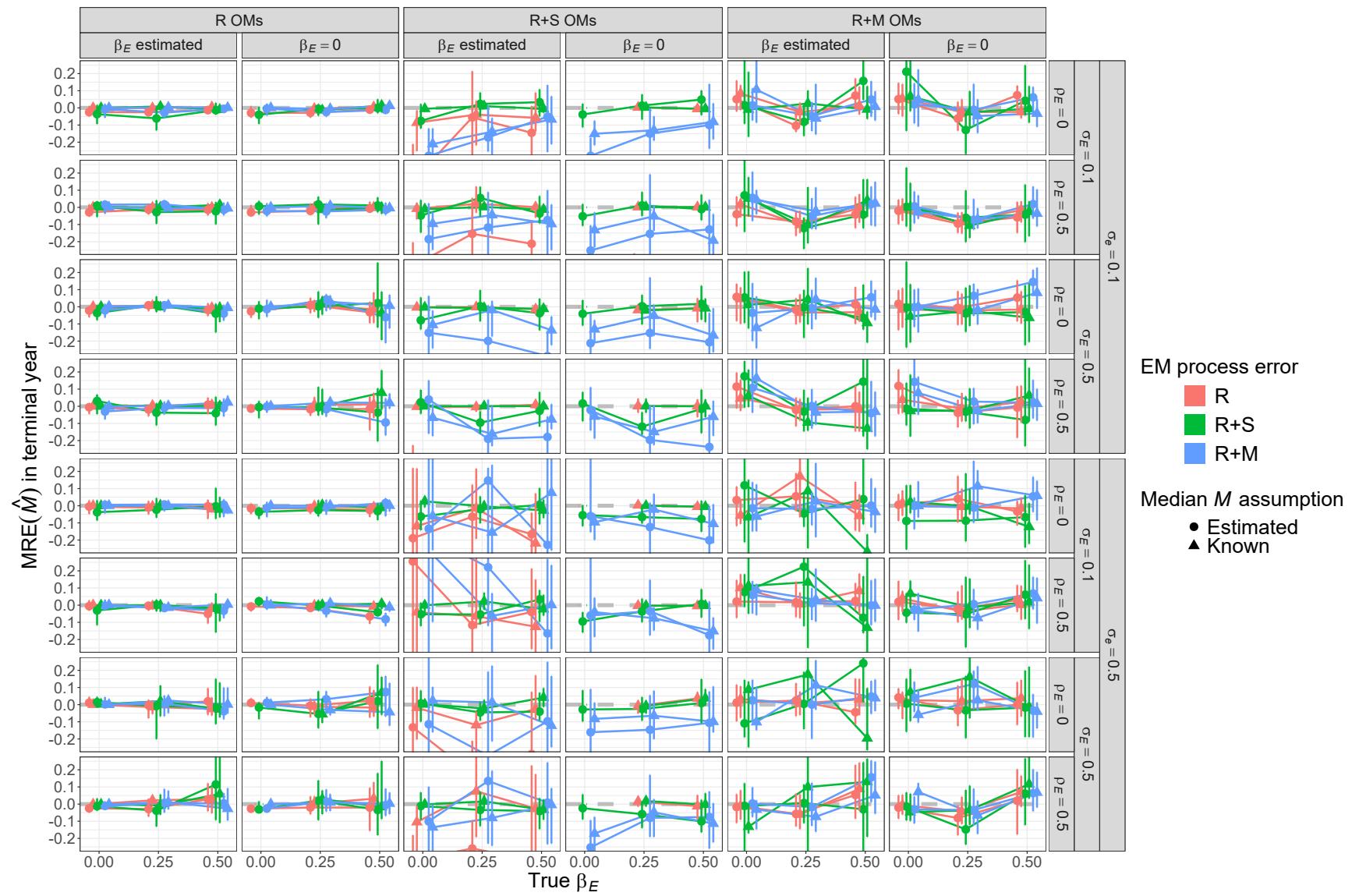


Fig. 7. Median relative error (MRE) of estimates of natural mortality rate in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known). All OMs had low observation error and contrast in fishing mortality.

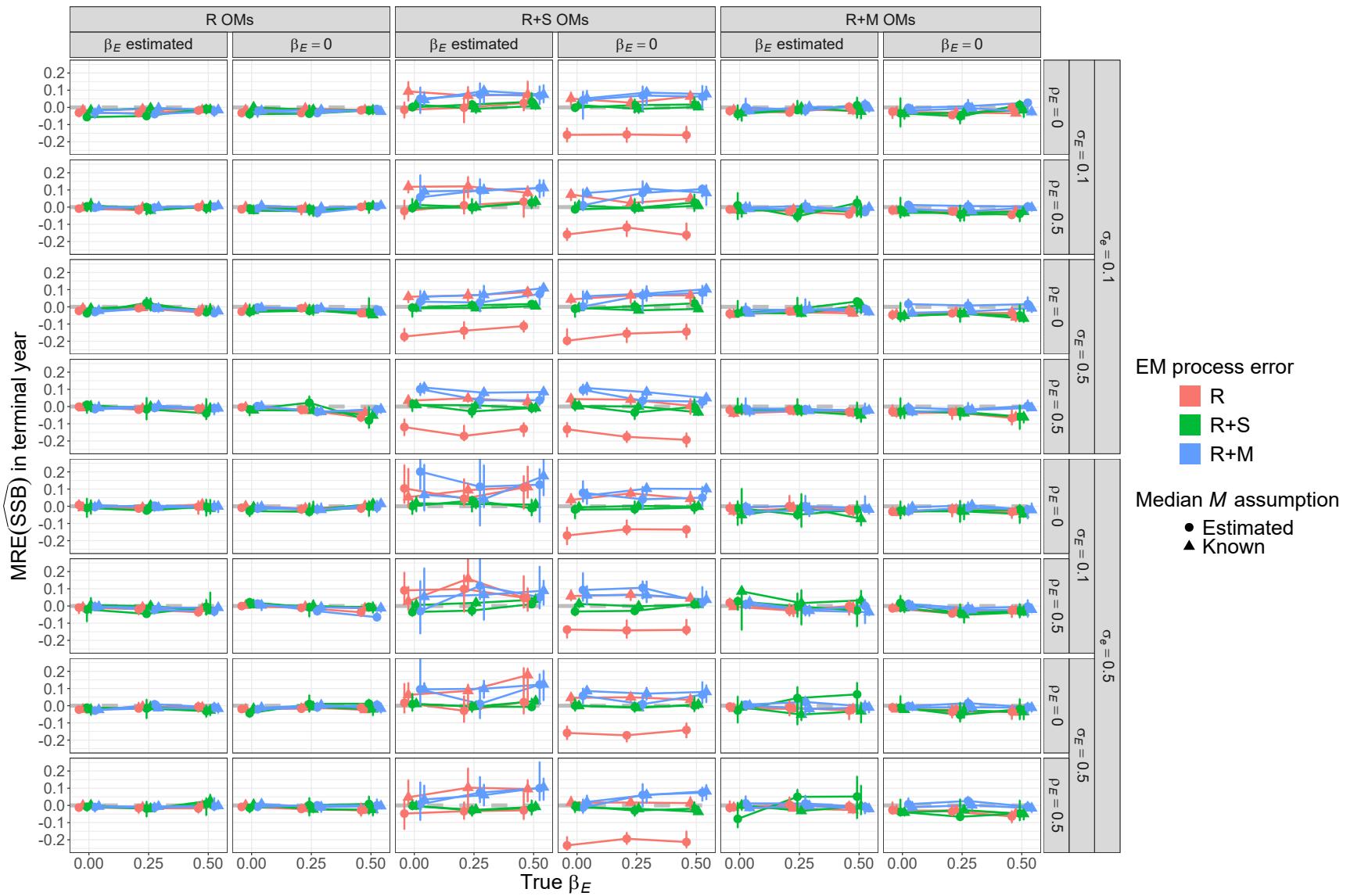


Fig. 8. Median relative error (MRE) of estimates of spawning stock biomass (SSB) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known). All OMs had low observation error and contrast in fishing mortality.

⁵⁶⁴ Supplemental Materials

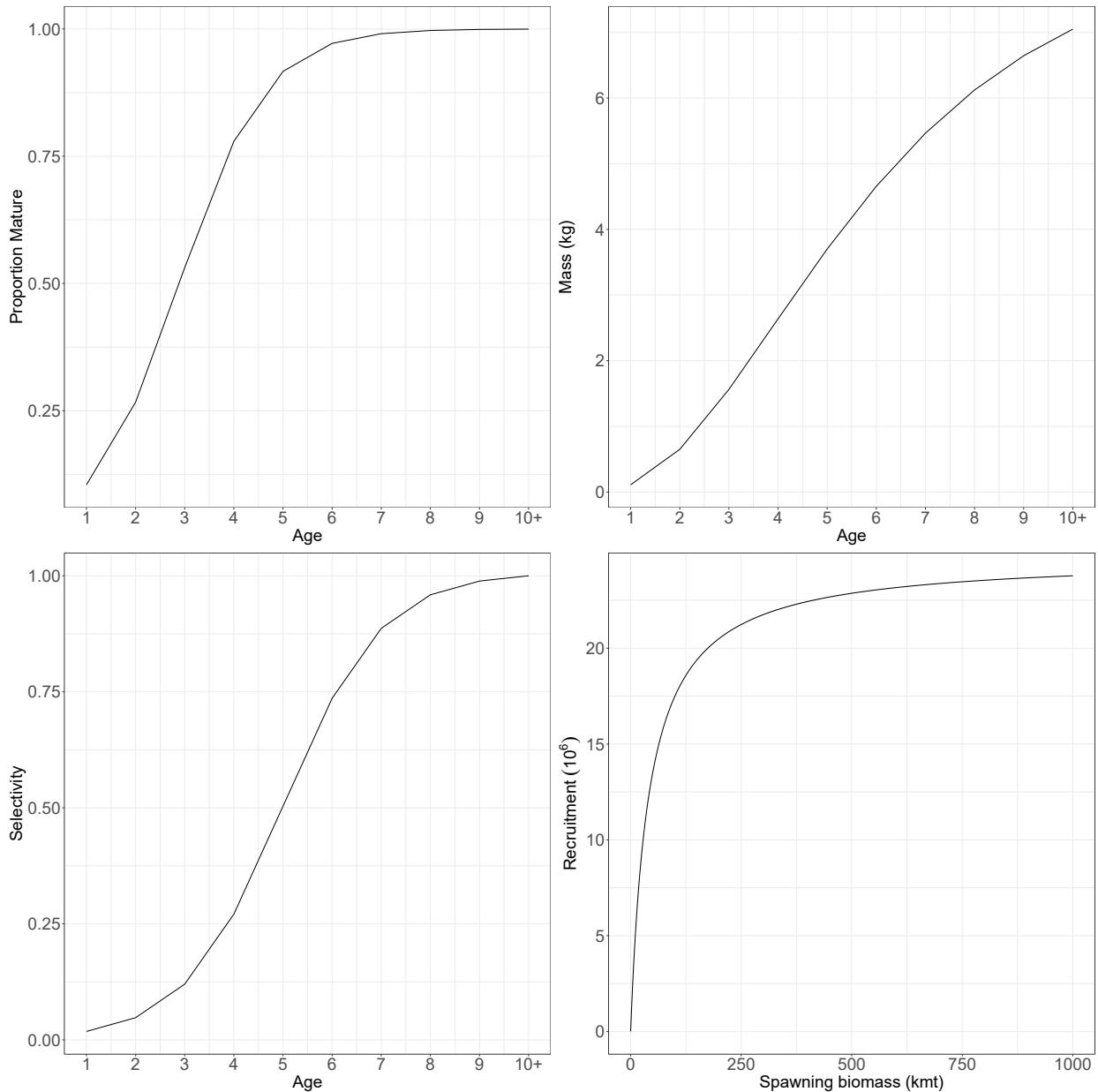


Fig. S1. The proportion mature at age, weight at age, fleet and index selectivity at age, and Beverton-Holt stock-recruit relationship assumed for the population in all operating models. For operating models with random effects on fleet selectivity, this represents the selectivity at the mean of the random effects.

Fig. S2. Example simulations of environmental covariate latent processes and observations with different levels of observation error, and different assumptions about variability of the latent process.

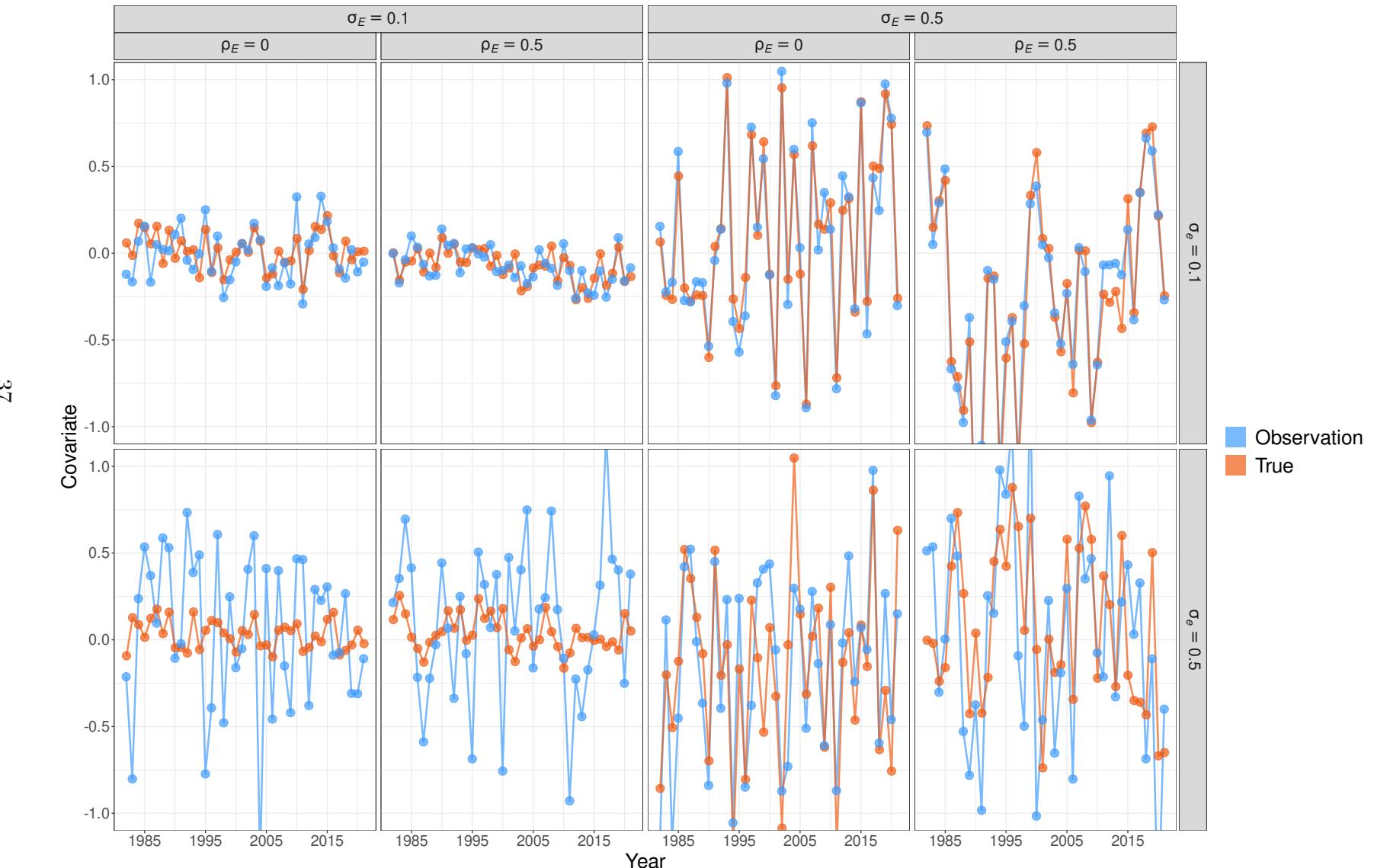
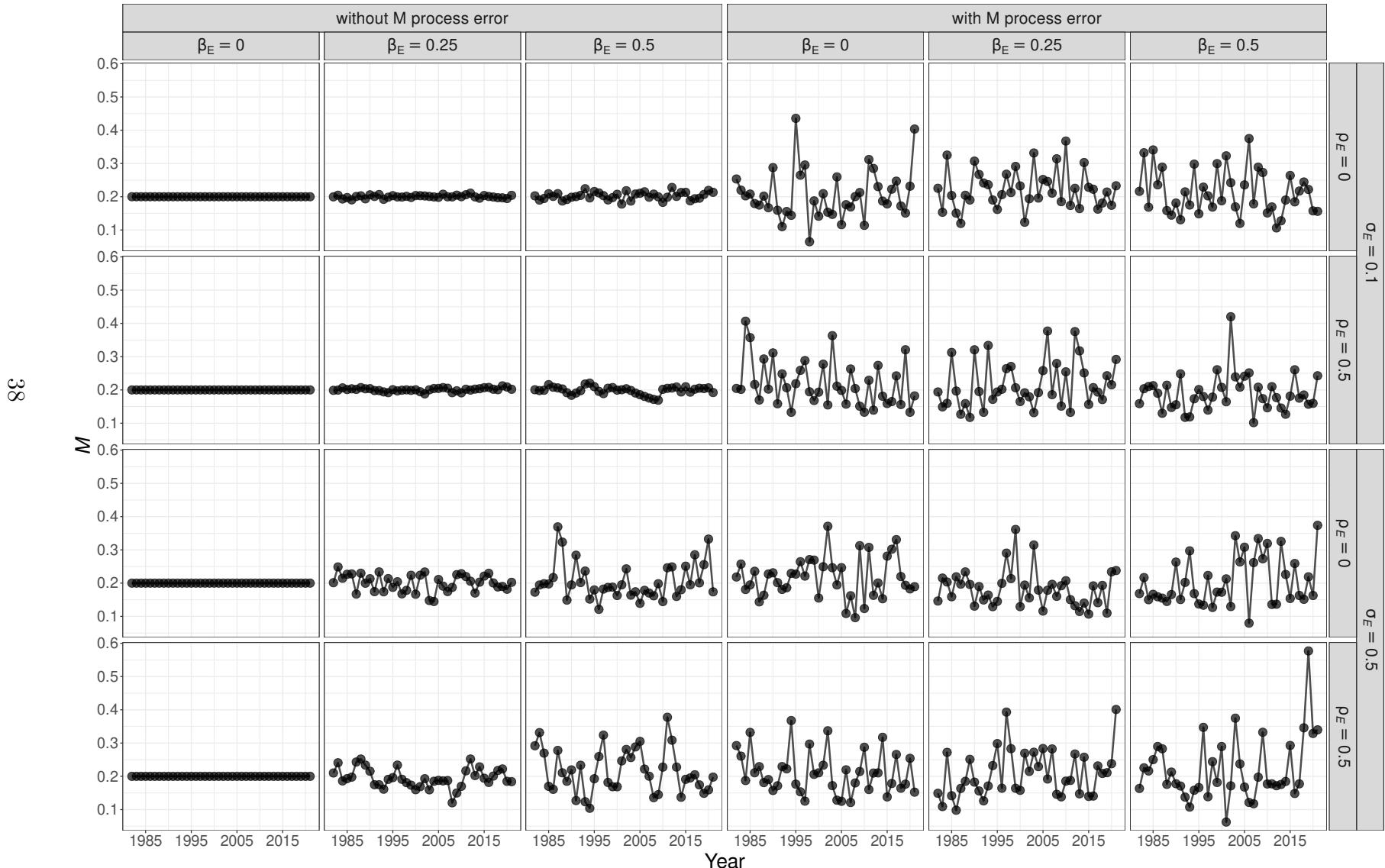


Fig. S3. Example simulations of annual natural mortality rates that may be a function of a temporally varying environmental covariate and autoregressive random effects.



565 **Convergence results**

566 For R OMs, R EMs converged frequently except when there was contrast in fishing and
567 covariate effects (β_E) were estimated and covariate observation uncertainty (σ_e) was high
568 (Figure S4). There was little difference in convergence for R EMs whether median M was
569 estimated or not. When R OMs had contrast in fishing, convergence of R+M EMs was
570 generally better when the EM estimated median M, whereas in OMs with constant fishing,
571 convergence was better when treating median M as known. Convergence of R+S EMs
572 was generally worse than other process error configurations for R OMs, and there was no
573 difference between EMs that estimated median M or not except when fishing pressure was
574 constant where better convergence occurred when M was treated as known. Effects of true
575 covariate effect size on convergence were apparent in some R OMs, but there did not appear
576 to be any consistent patterns.

577 For R+S OMs, R+M EMs generally showed the worse probability of convergence than EMs
578 with R and R+S configurations (Figure S5). Convergence probability was high for all R
579 and R+S EMs when covariate observation uncertainty was low. Although probability of
580 convergence was generally greater when covariate effects were not estimated, there were no
581 strong trends in convergence for any of the EMS with increased true effect size.

582 For R+M OMs, R+S EMs generally converged less frequently than those with R and R+M
583 configurations (Figure S6). Convergence probability for R OMs was generally the greatest
584 of the three process error configurations. As with R+S OMs, there were no strong trends in
585 convergence for any of the EMS with increased true covariate effect size. However, there was
586 also not strong effects of estimating the covariate effect on convergence except for some OMs
587 with higher covariate observation uncertainty. Like R OMs, when R+M OMs had contrast
588 in fishing, convergence of R+M EMs was generally better when the EM estimated median
589 M and the convergence was better when treating median M as known when the OM had
590 constant fishing.

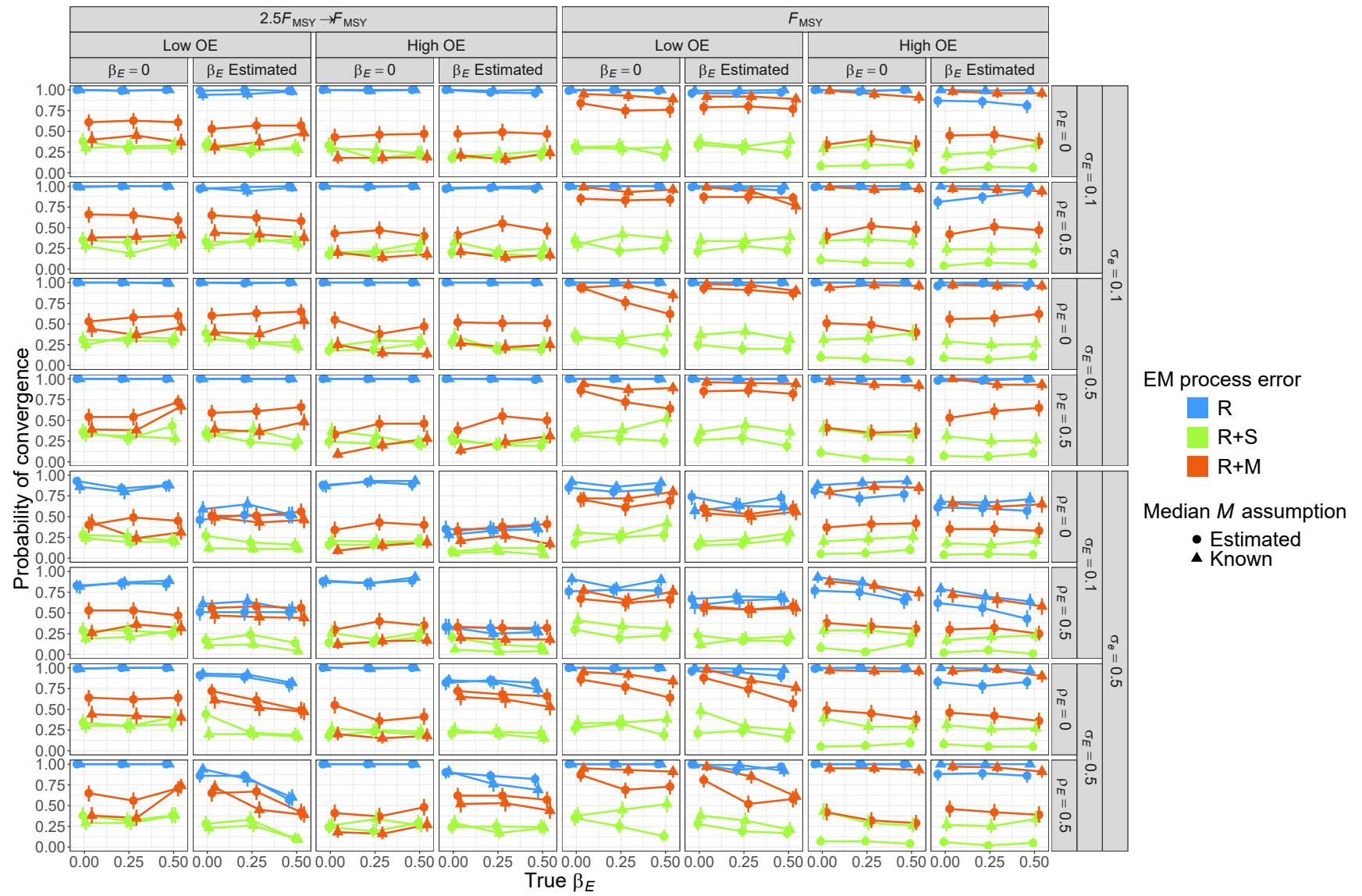


Fig. S4. Estimated probability of fits providing Hessian-based standard errors for EMs assuming alternative process error, that estimate or assume known median natural mortality, and that estimate or assume no covariate effect on median natural mortality when fitted to R OMs and three levels of true covariate effect on median natural mortality (x axis). Vertical lines represent 95% confidence intervals.

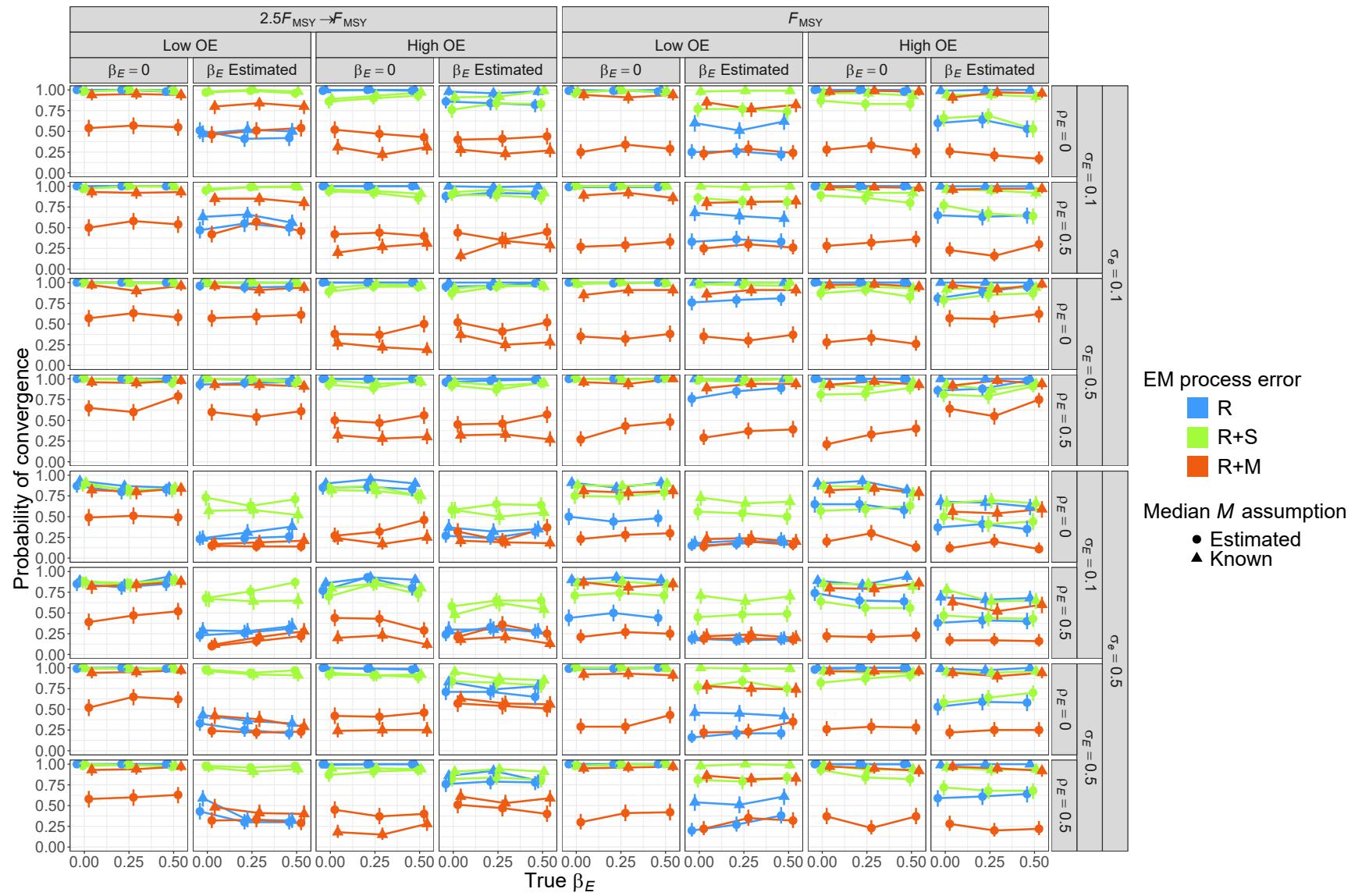


Fig. S5. Estimated probability of fits providing Hessian-based standard errors for EMs assuming alternative process error, that estimate or assume known median natural mortality, and that estimate or assume no covariate effect on median natural mortality when fitted to R+S OMs and three levels of true covariate effect on median natural mortality (x axis). Vertical lines represent 95% confidence intervals.

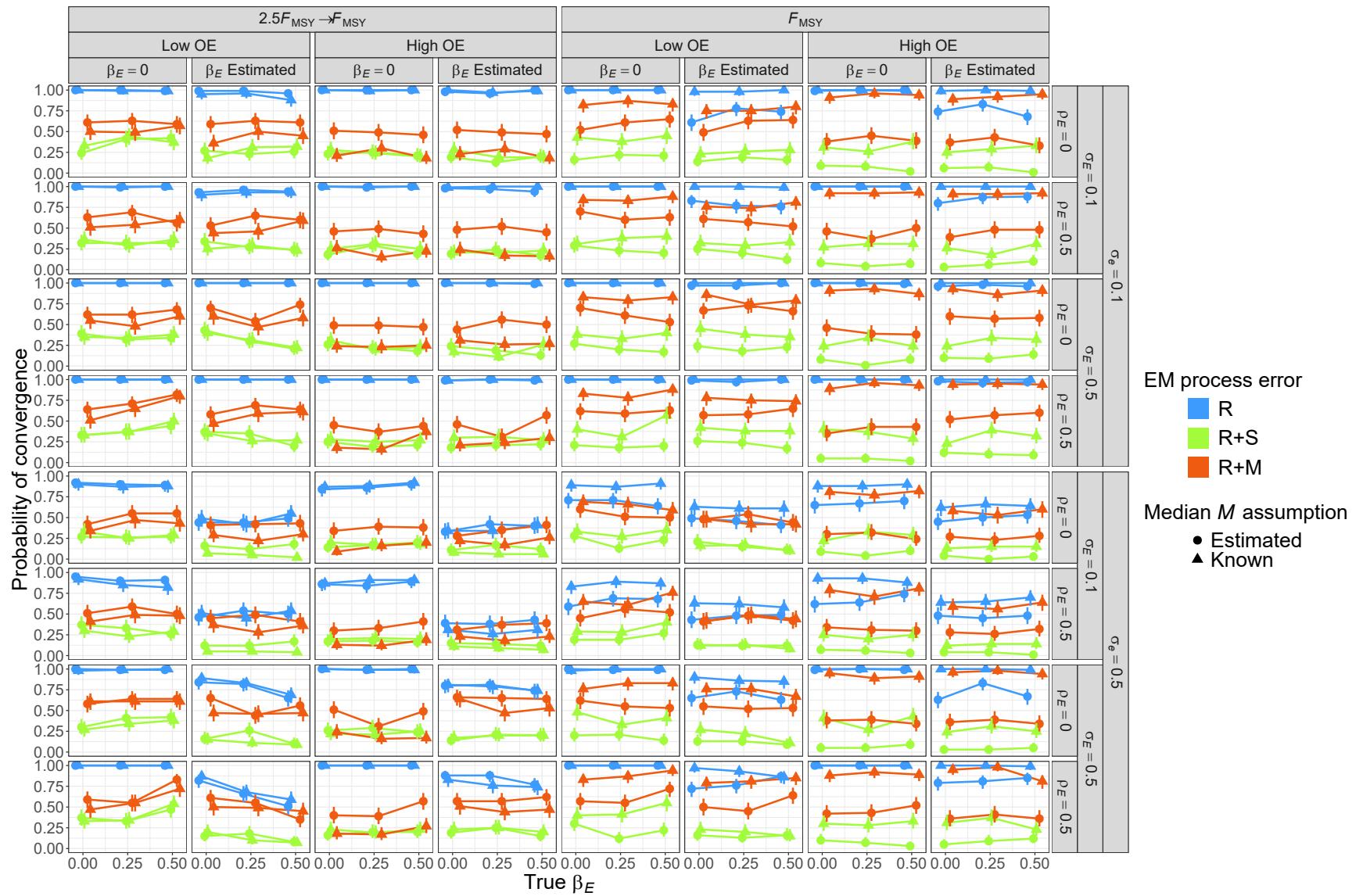


Fig. S6. Estimated probability of fits providing Hessian-based standard errors for EMs assuming alternative process error, that estimate or assume known median natural mortality, and that estimate or assume no covariate effect on median natural mortality when fitted to R+M OMs and three levels of true covariate effect on median natural mortality (x axis). Vertical lines represent 95% confidence intervals.

591 **AIC results**

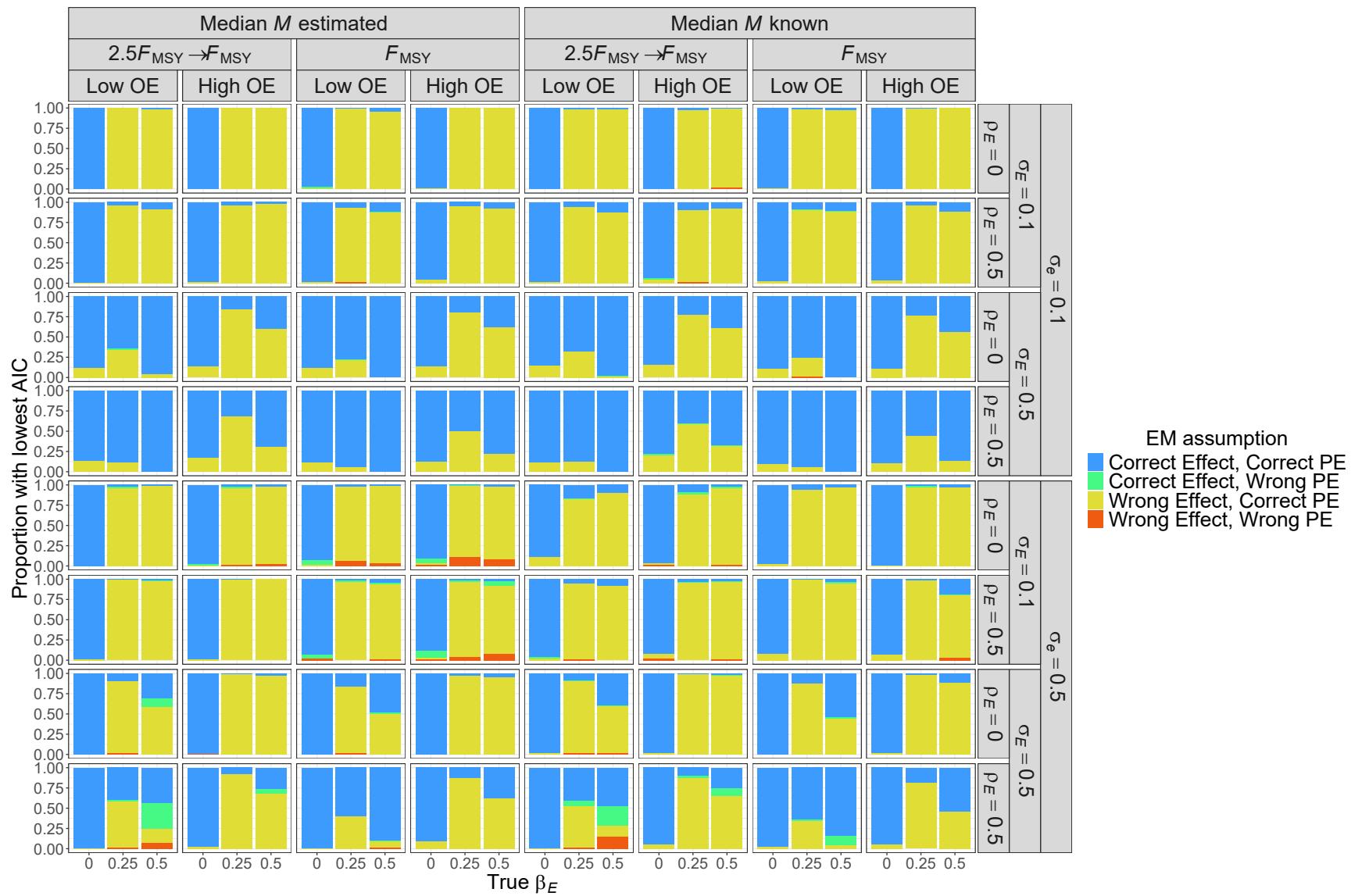


Fig. S7. Proportion of simulated data sets for R OMs where the EM type (treatment of environmental covariate and assumed process error type) had the lowest AIC.

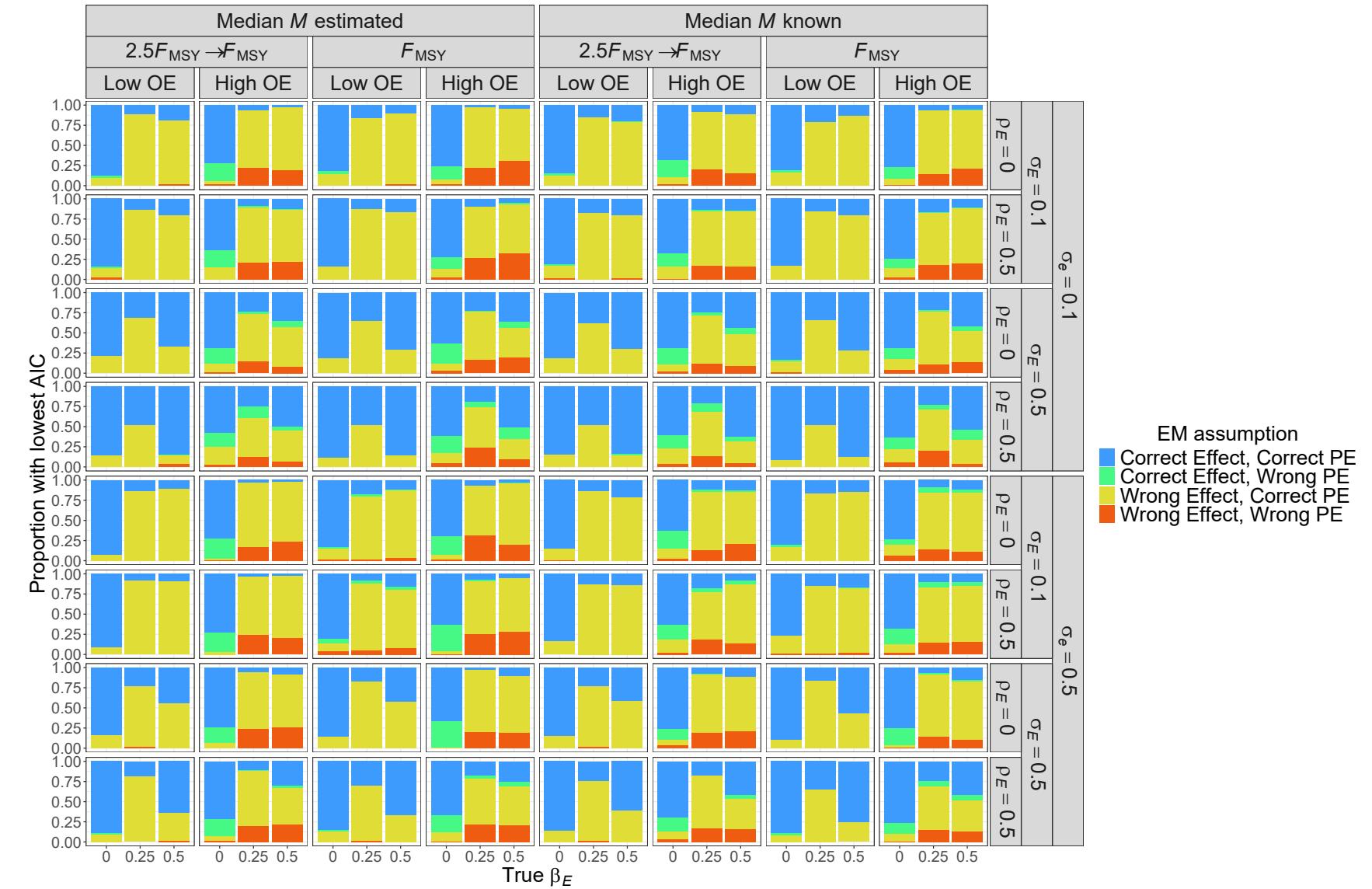


Fig. S8. Proportion of simulated data sets for R+S OMs where the EM type (treatment of environmental covariate and assumed process error type) had the lowest AIC.

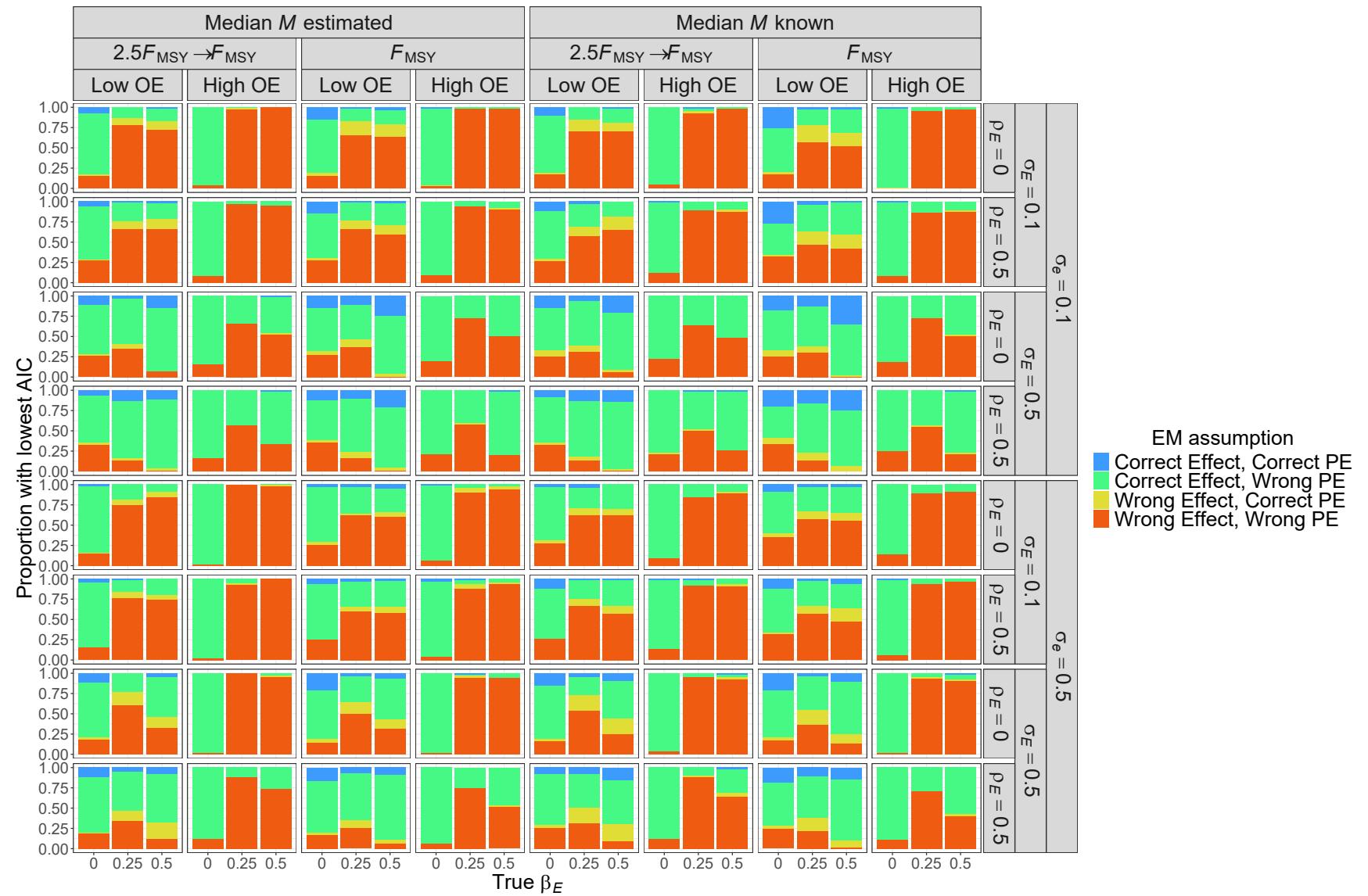


Fig. S9. Proportion of simulated data sets for R+M OMs where the EM type (treatment of environmental covariate and assumed process error type) had the lowest AIC.

592 Covariate effect bias

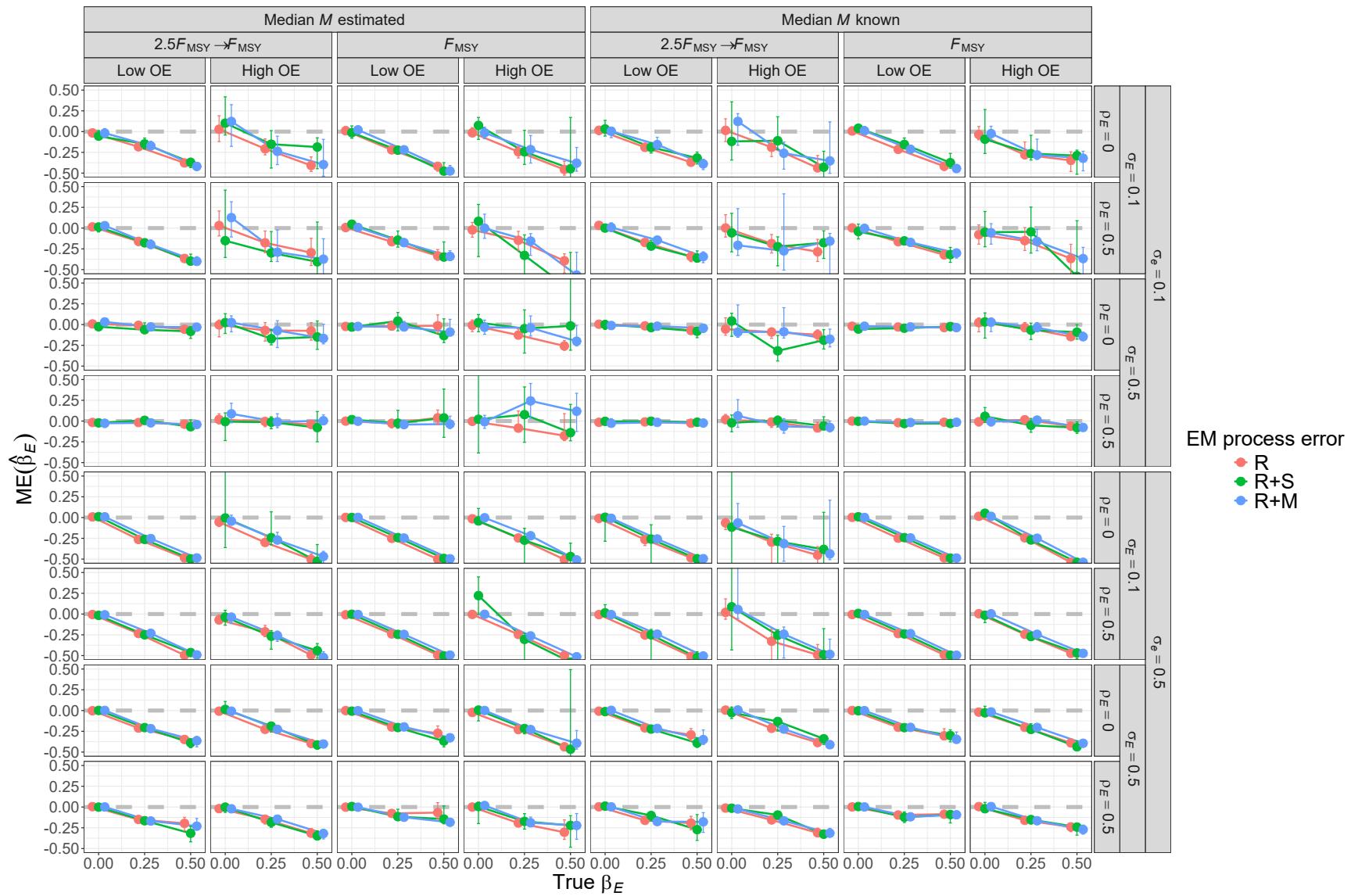


Fig. S10. For R OMs, median error (ME) of estimates of environmental effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). Vertical lines represent 95% confidence intervals.

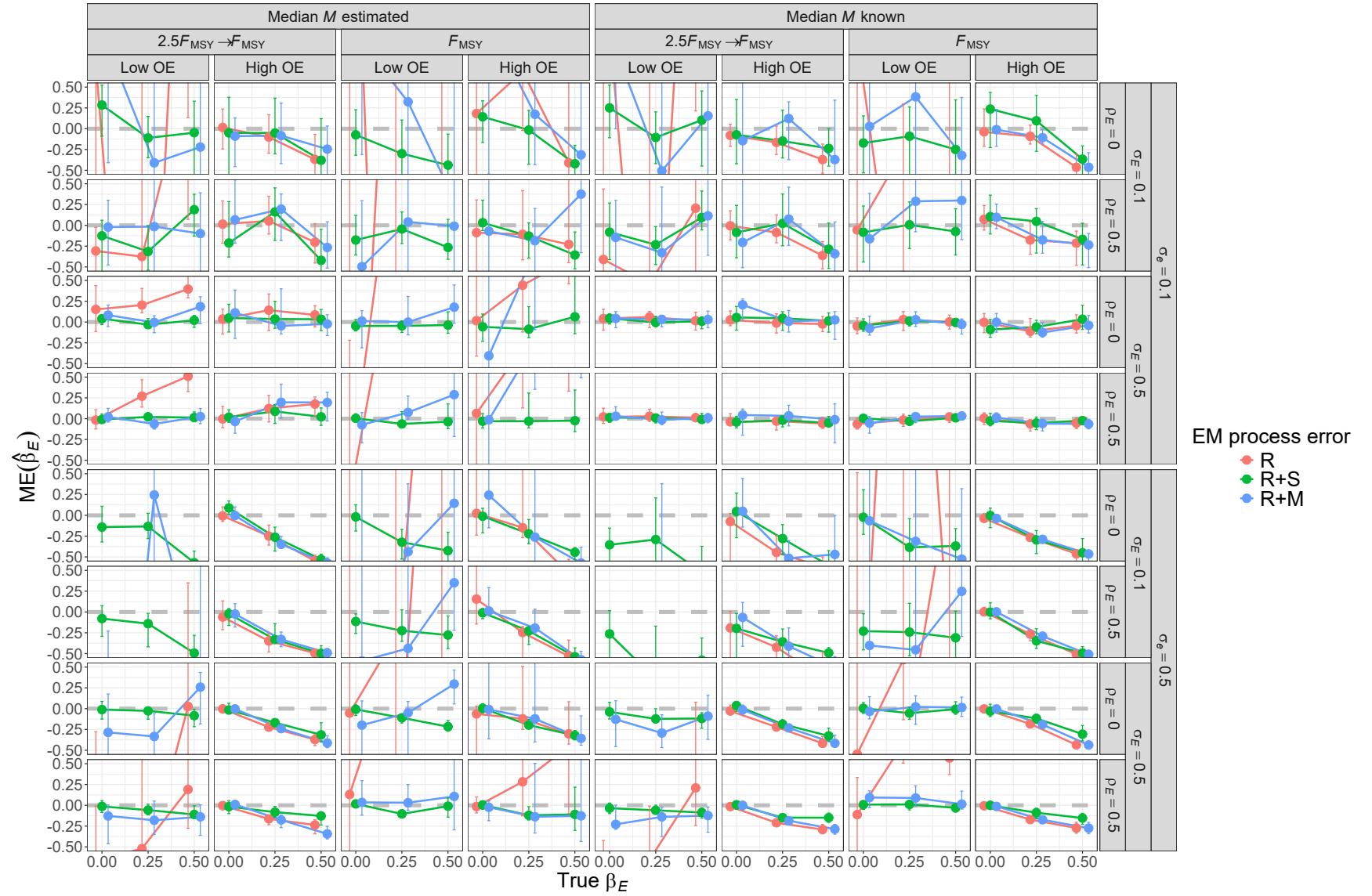


Fig. S11. For R+S OMs, median error (ME) of estimates of environmental effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). Vertical lines represent 95% confidence intervals.

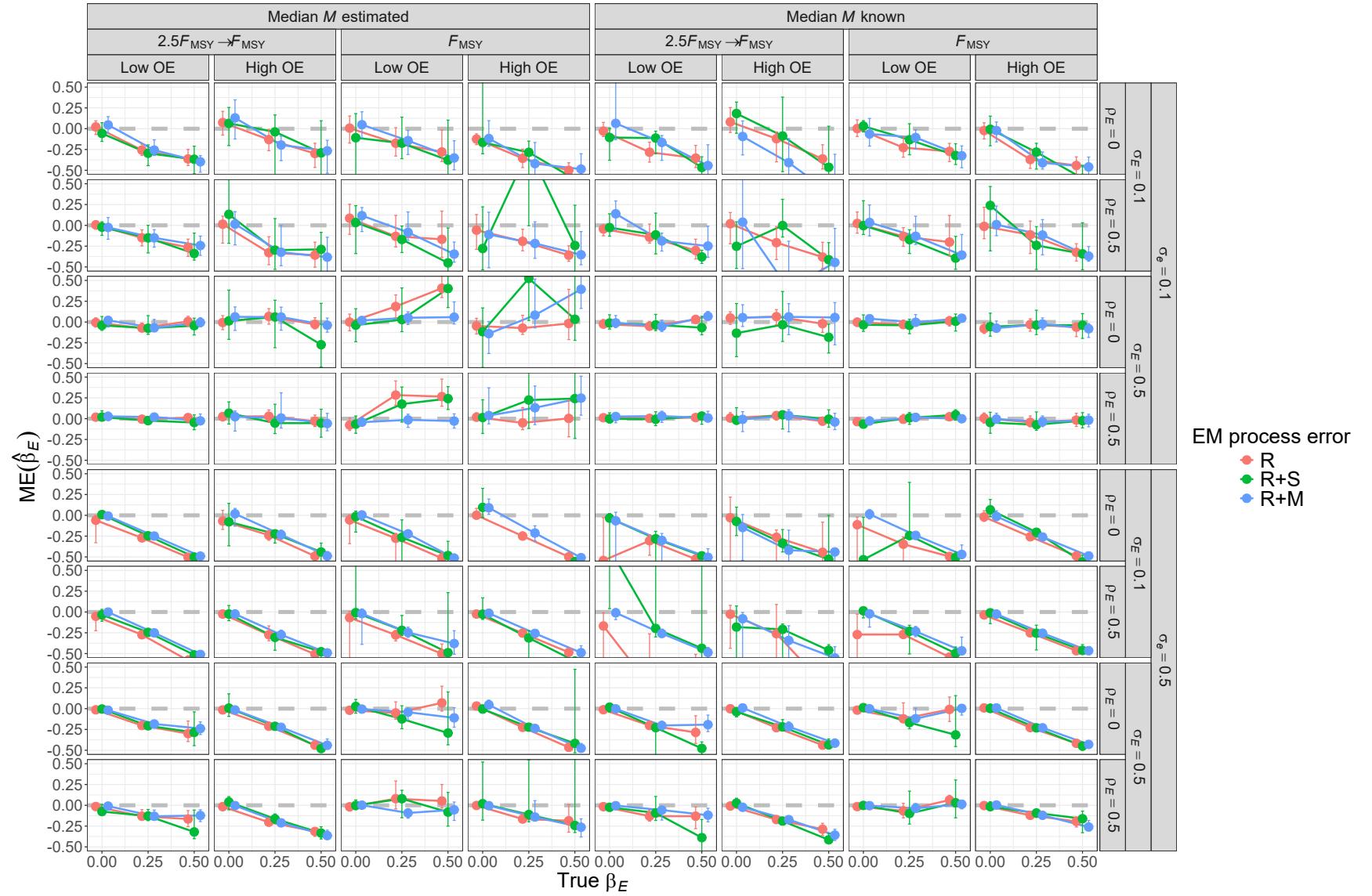


Fig. S12. For R+M OMs, median error (ME) of estimates of environmental effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). Vertical lines represent 95% confidence intervals.

⁵⁹³ Covariate effect standard error estimation bias

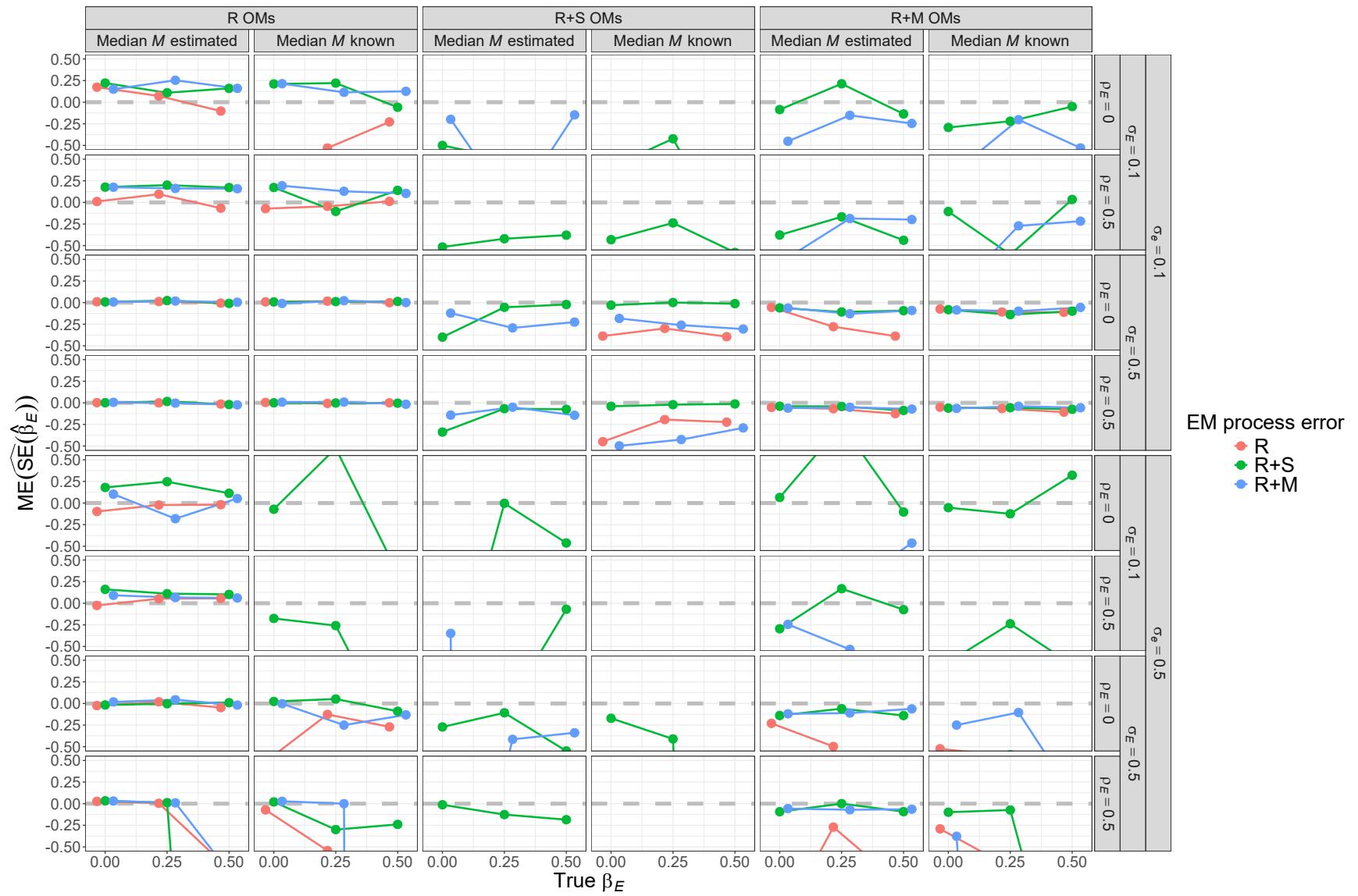


Fig. S13. Median error (ME) of Hessian-based estimates of standard error for covariate effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). All OMs had low observation error and contrast in fishing mortality. True standard error is defined as the mean of the standard error estimates across converged fits to simulated data sets for a given OM scenario.

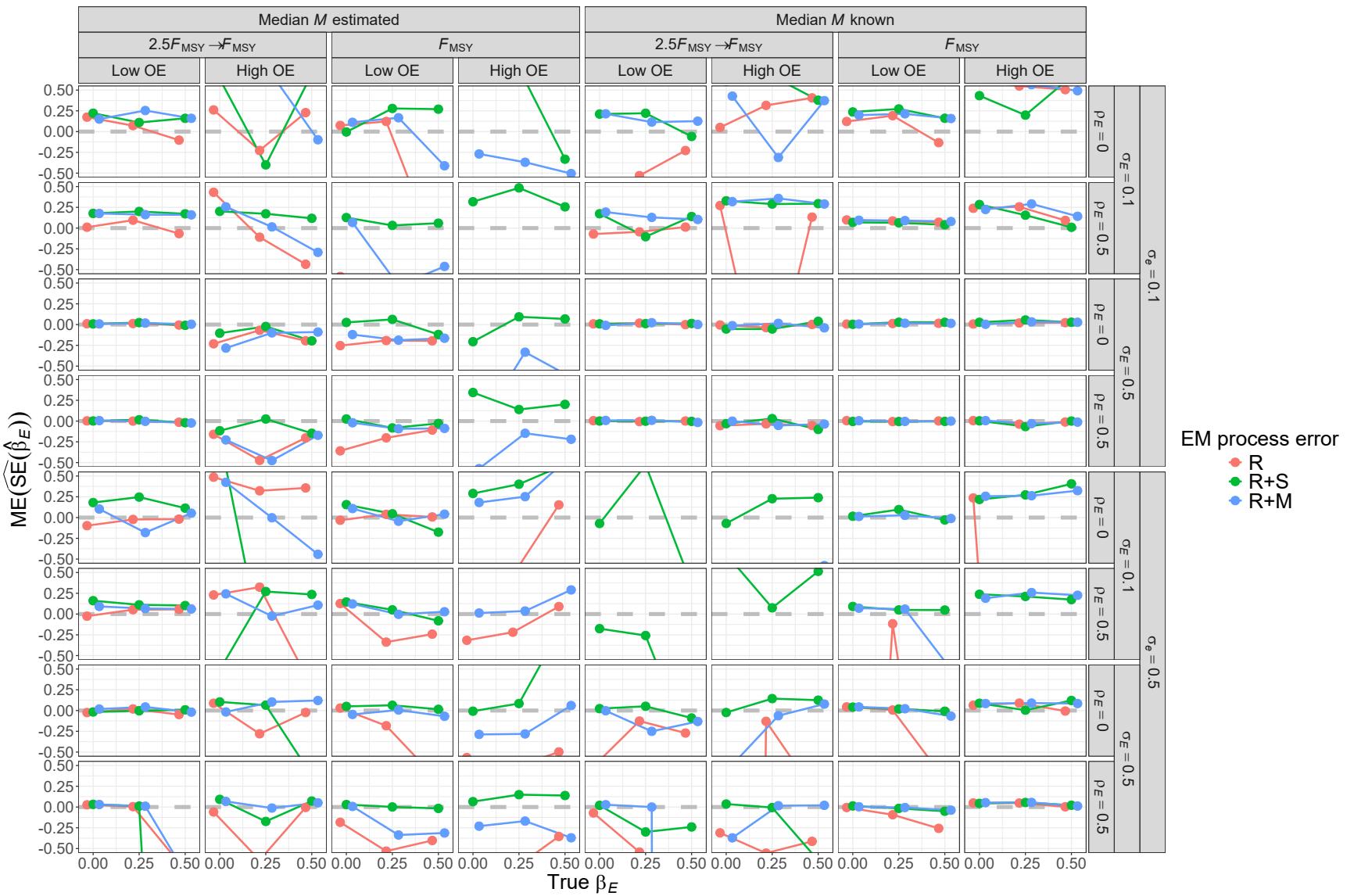


Fig. S14. For R OMs, median error (ME) of Hessian-based estimates of standard error for covariate effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). True standard error is defined as the mean of the standard error estimates across converged fits to simulated data sets for a given OM scenario.

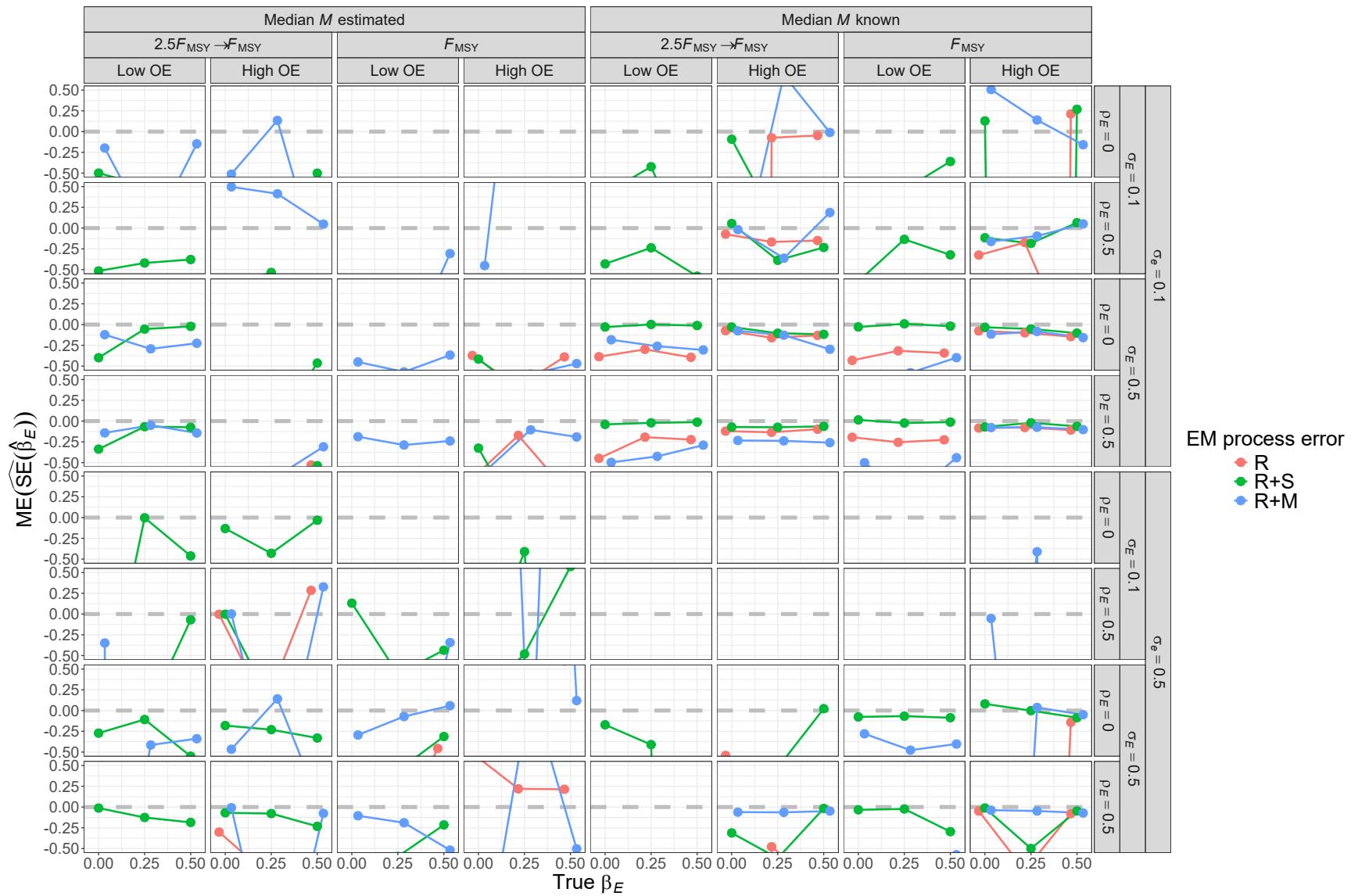


Fig. S15. For R+S OMs, median error (ME) of Hessian-based estimates of standard error for covariate effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). True standard error is defined as the mean of the standard error estimates across converged fits to simulated data sets for a given OM scenario.

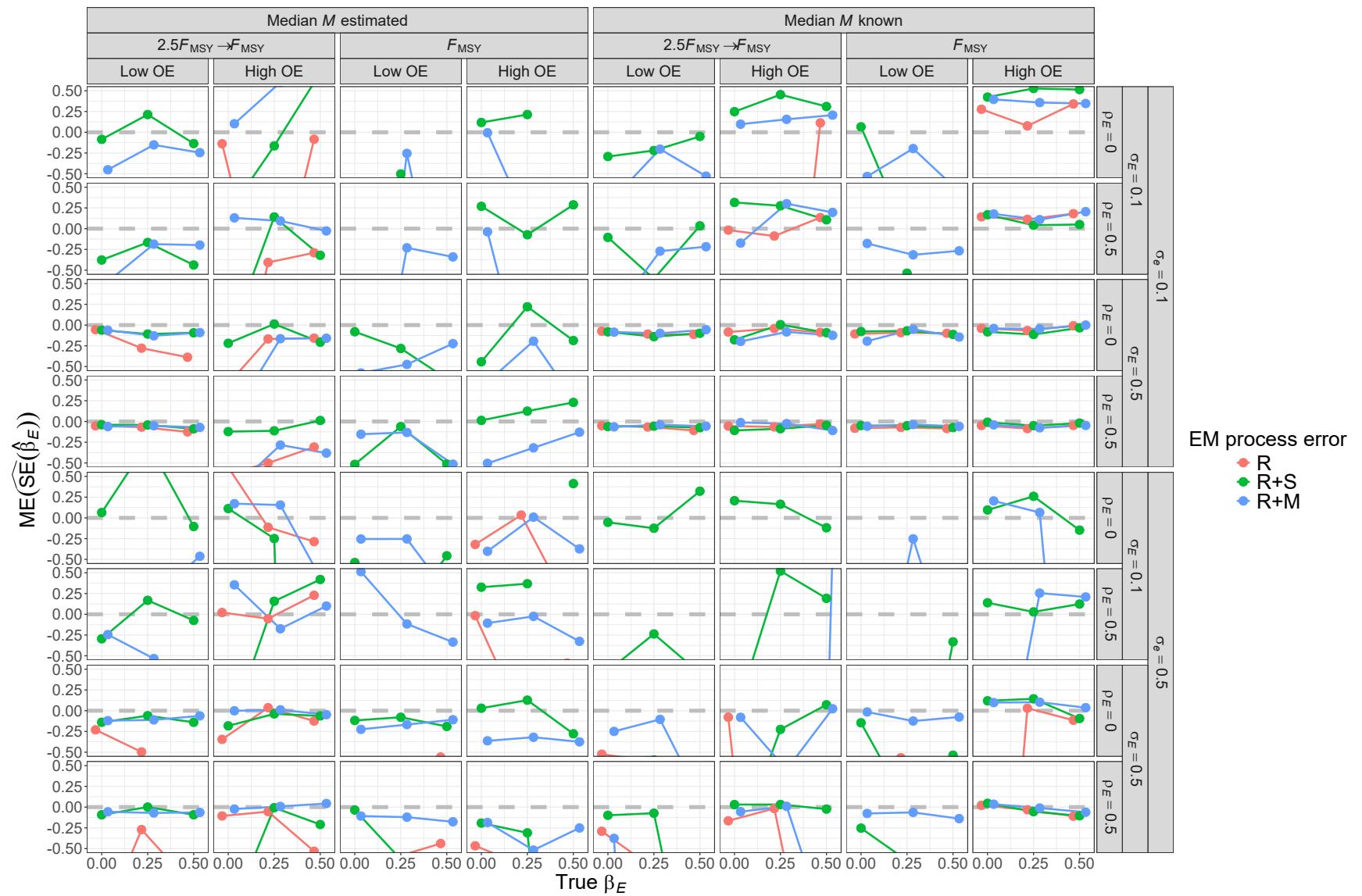


Fig. S16. For R+M OMs, median error (ME) of Hessian-based estimates of standard error for covariate effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). True standard error is defined as the mean of the standard error estimates across converged fits to simulated data sets for a given OM scenario.

⁵⁹⁴ Covariate effect confidence interval coverage

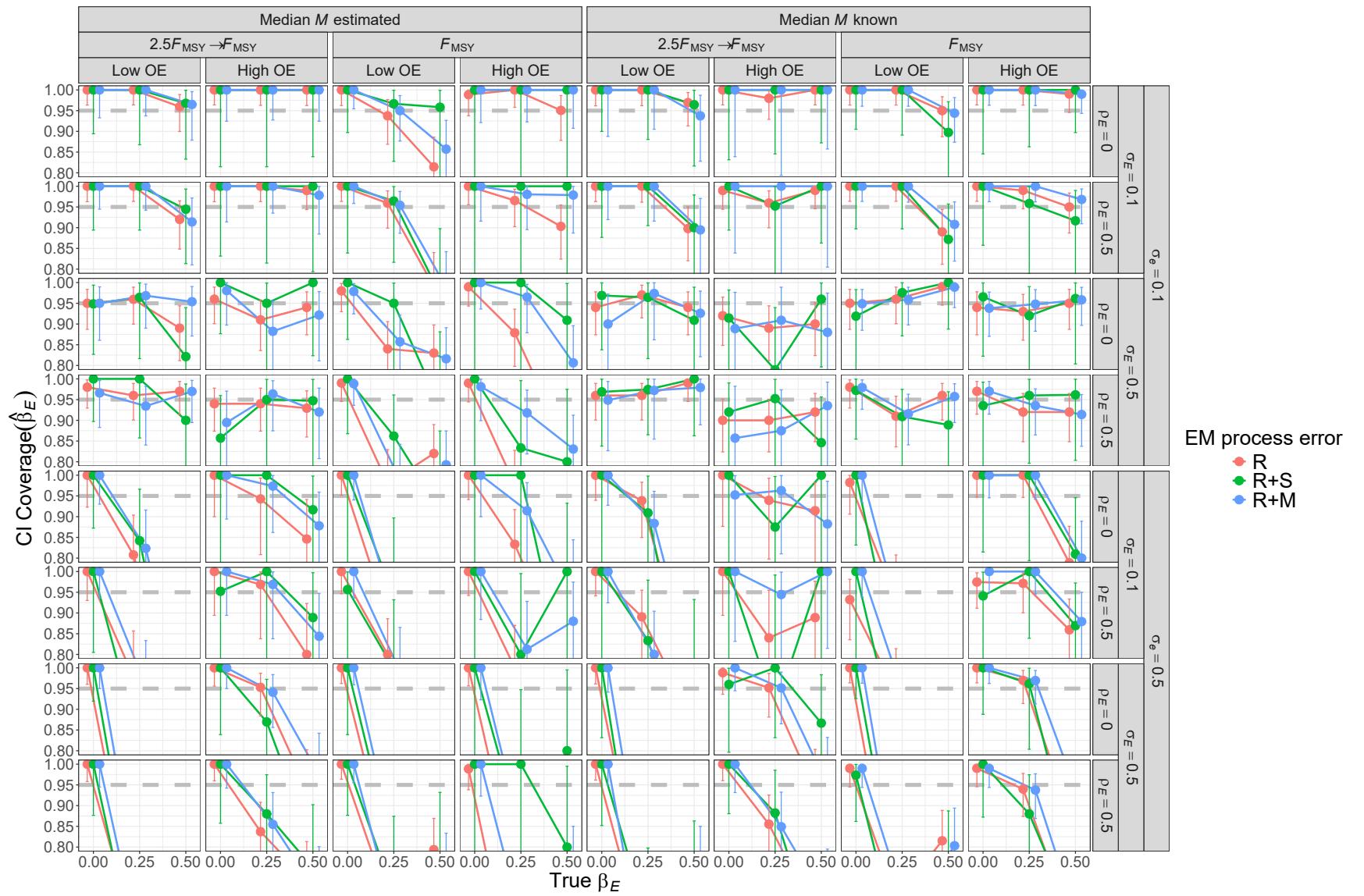


Fig. S17. For R OMs, probability of 95% confidence interval for β_E containing the true value for EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). Vertical lines represent 95% confidence intervals.

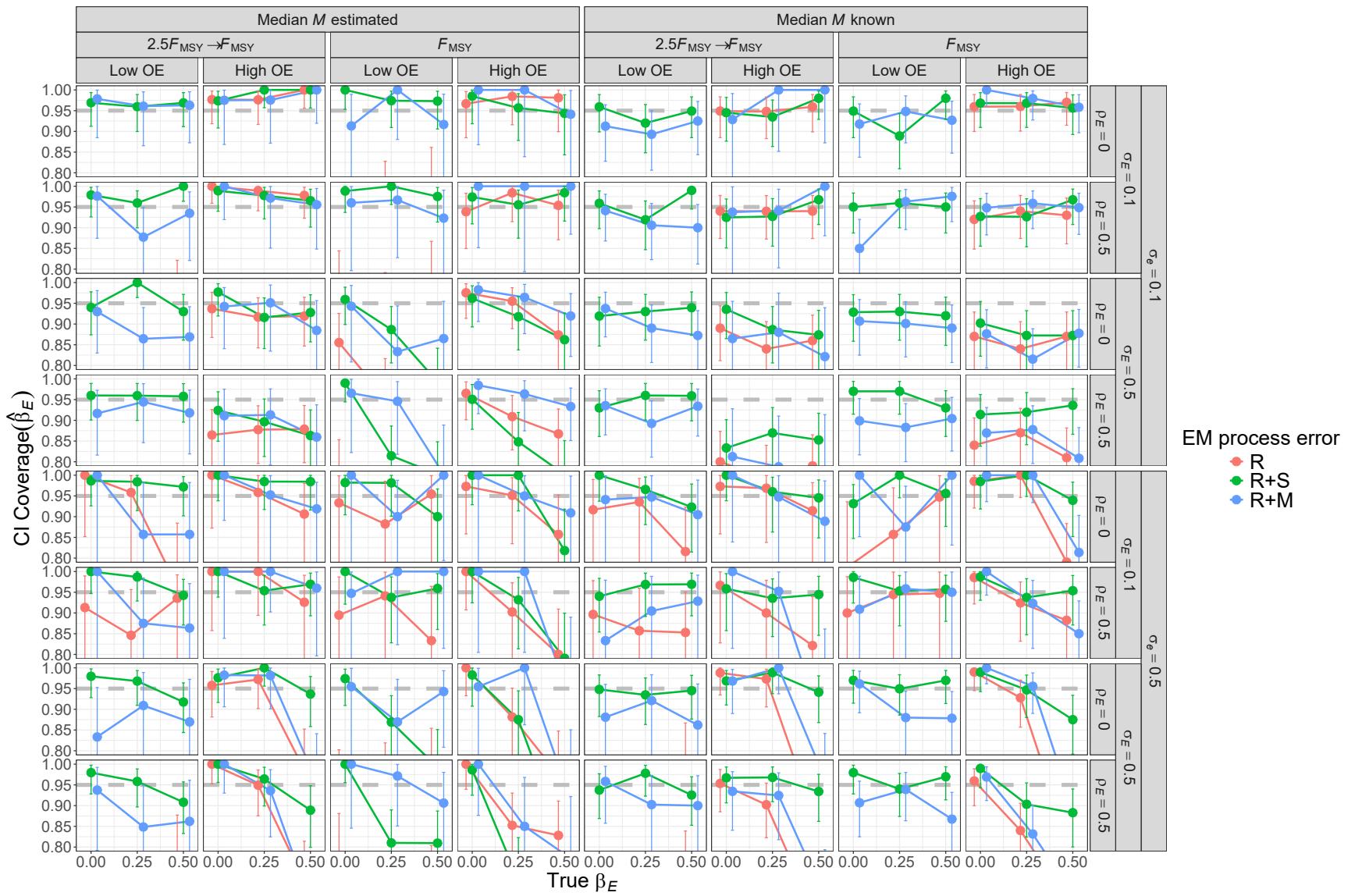


Fig. S18. For R+S OMs, probability of 95% confidence interval for β_E containing the true value for EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). Vertical lines represent 95% confidence intervals.

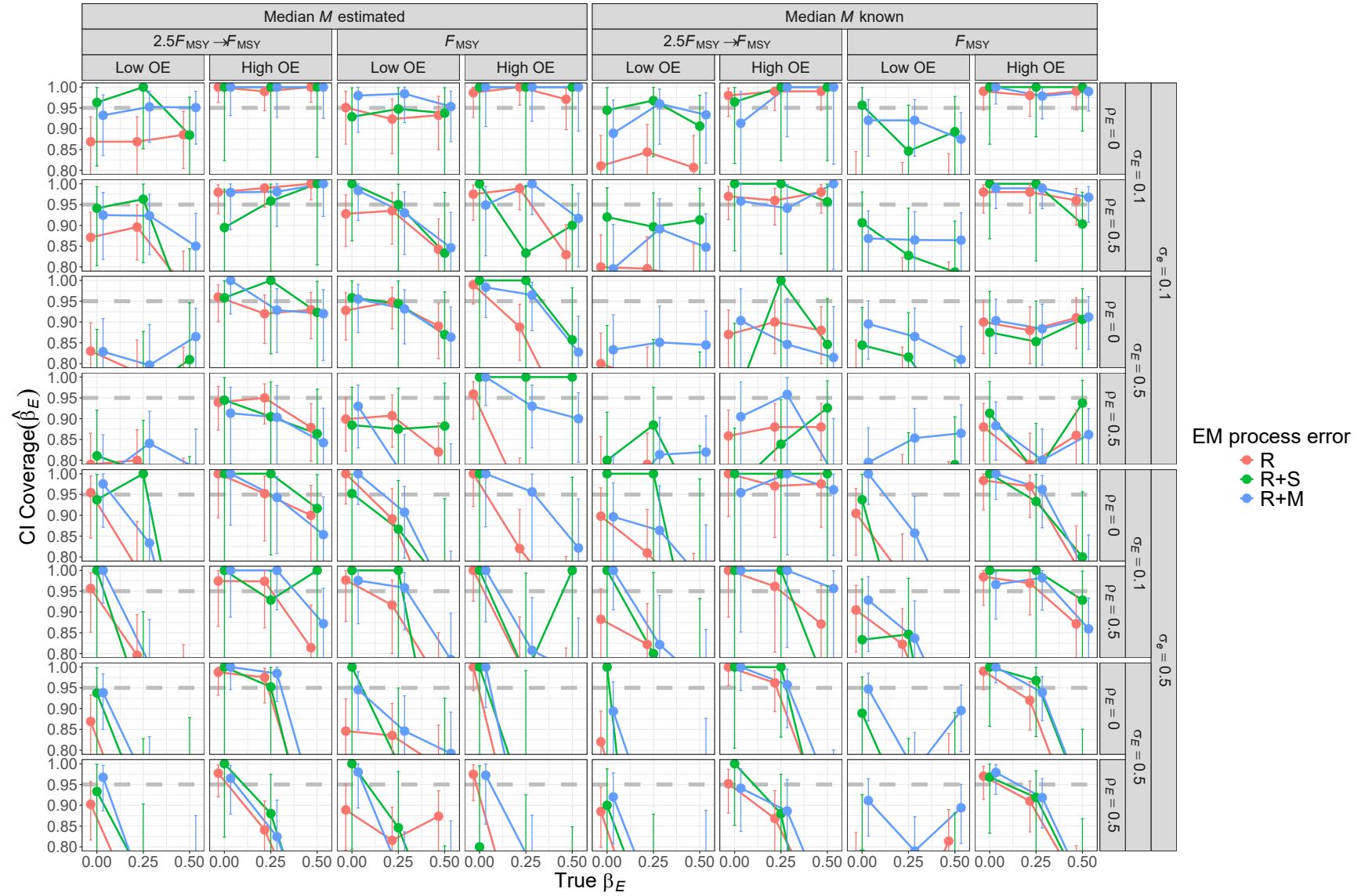


Fig. S19. For R+M OMs, probability of 95% confidence interval for β_E containing the true value for EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). Vertical lines represent 95% confidence intervals.

595 Covariate effect RMSE

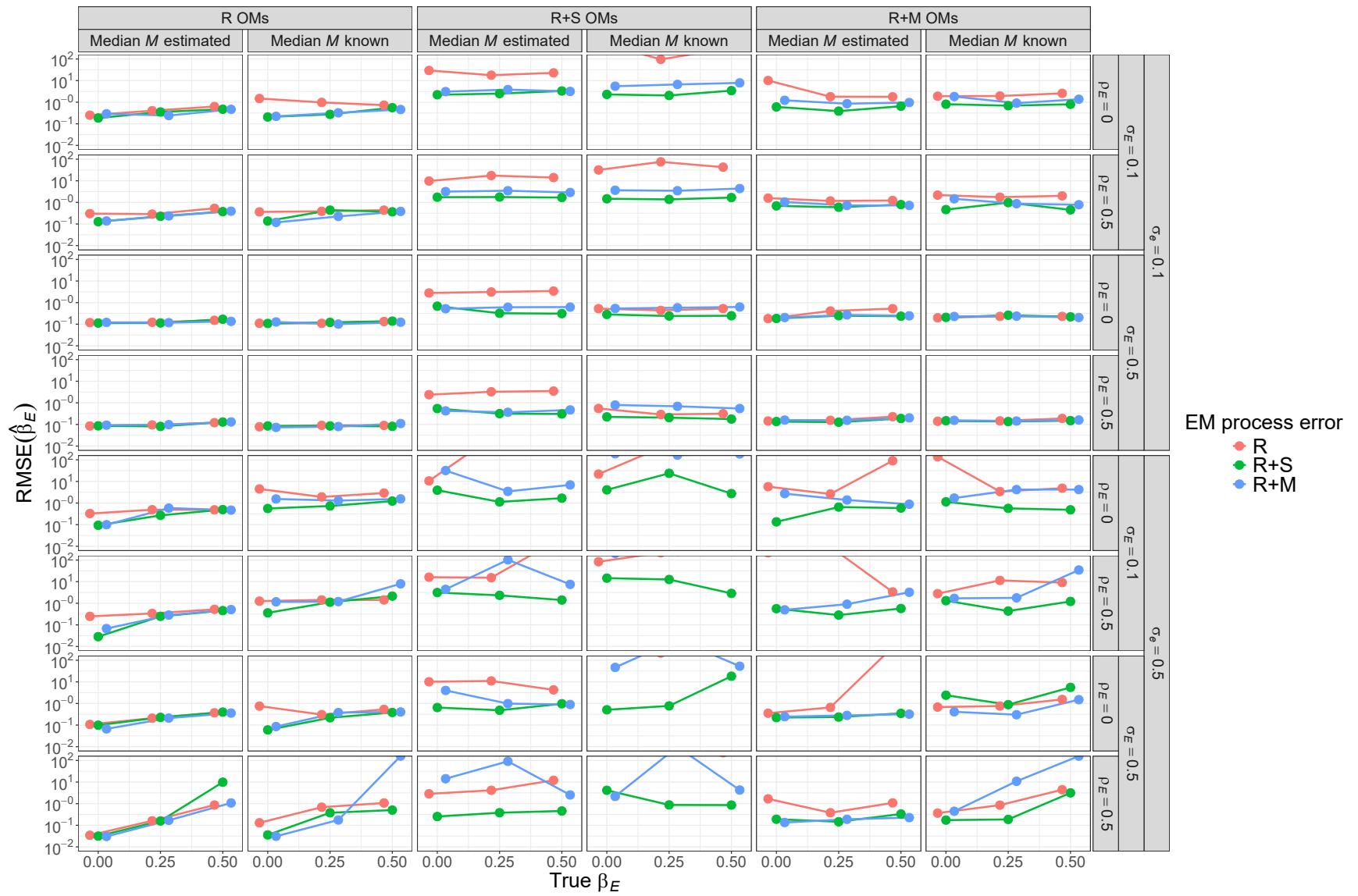


Fig. S20. Root mean square error (RMSE) of estimates of covariate effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated). All OM had low observation error and contrast in fishing mortality.

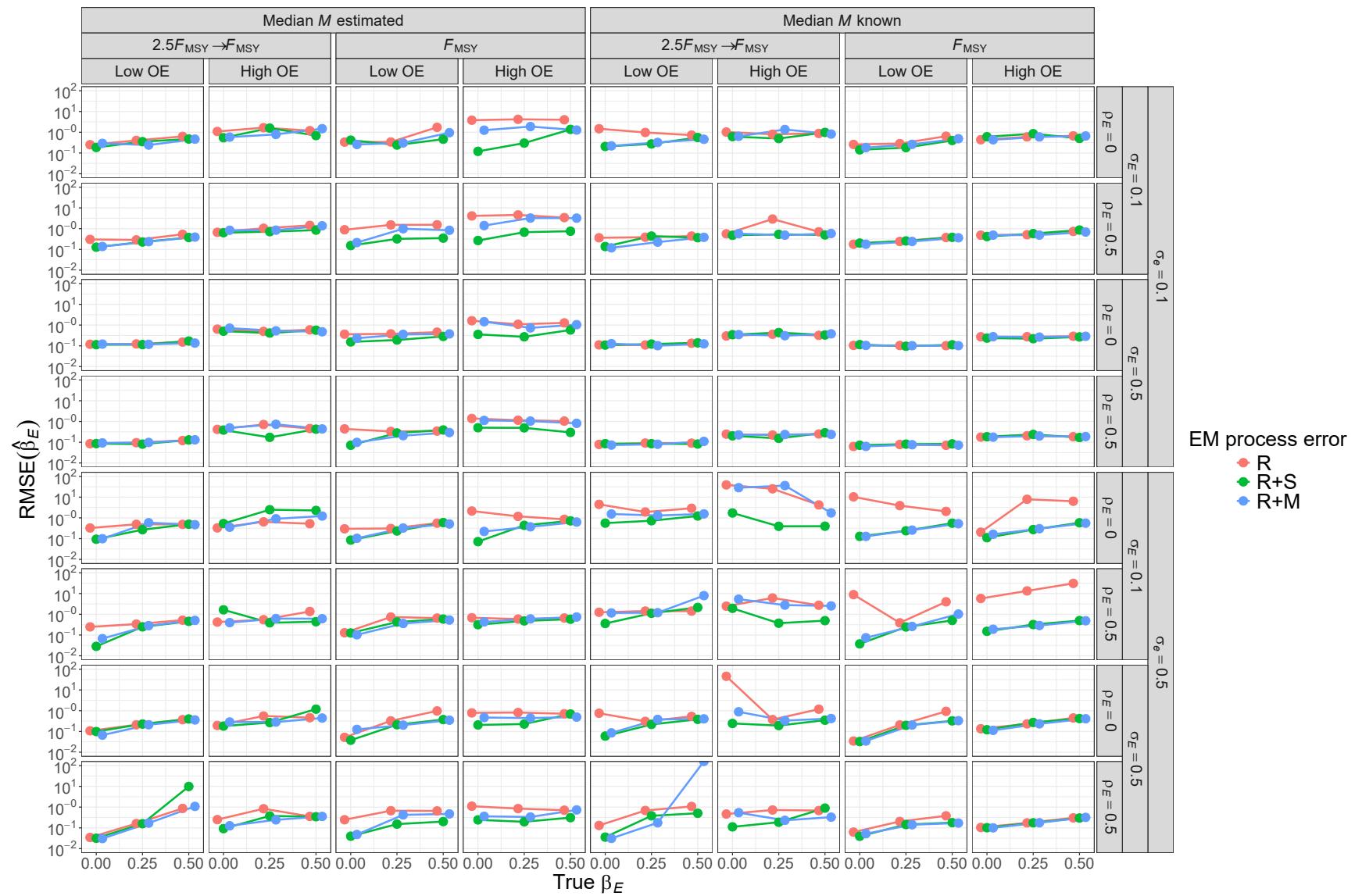


Fig. S21. For R OMs, root mean square error (RMSE) of estimates of covariate effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated).

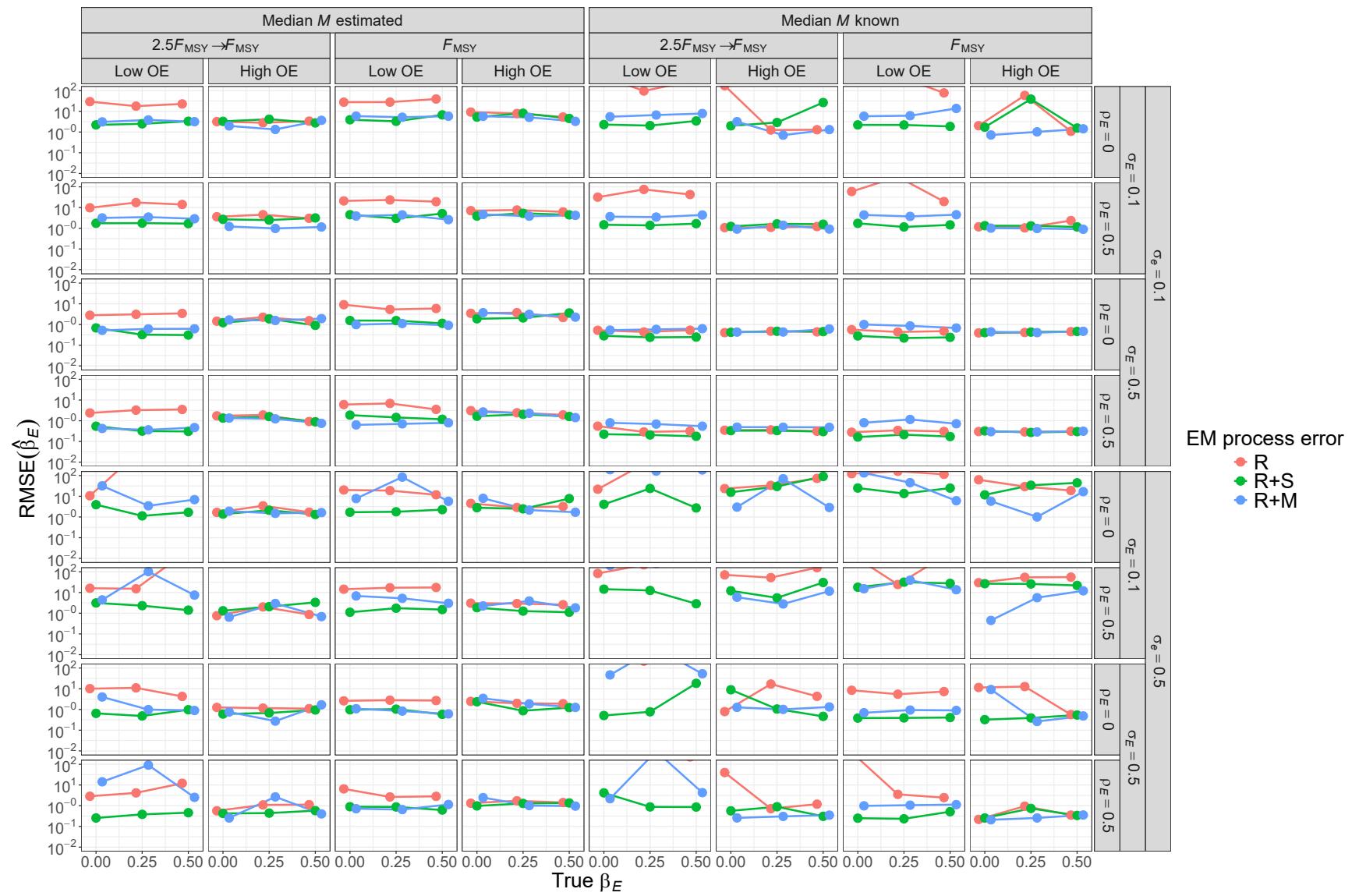


Fig. S22. For R+S OMs, root mean square error (RMSE) of estimates of covariate effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated).

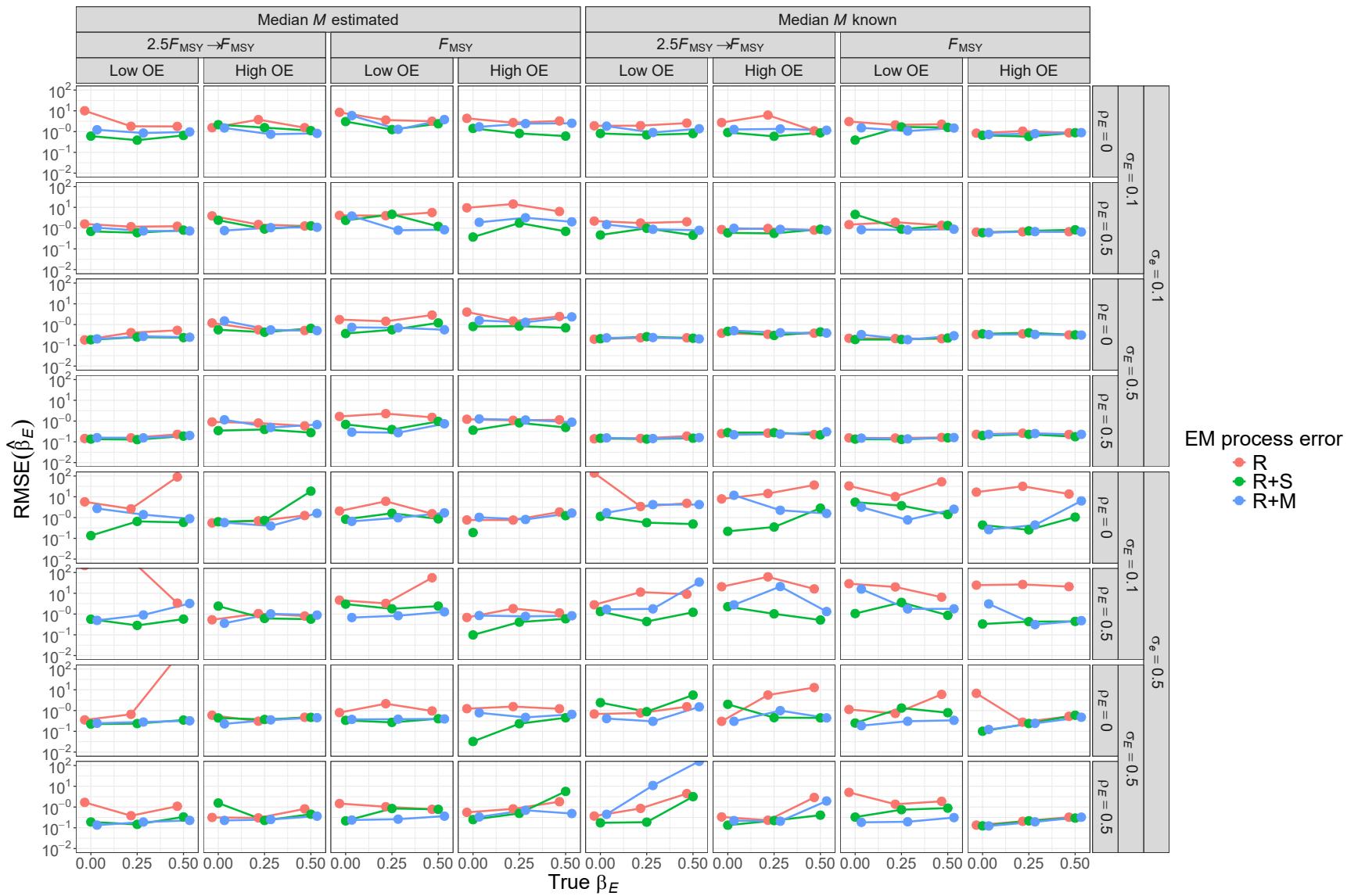


Fig. S23. For R+M OMs, root mean square error (RMSE) of estimates of covariate effect on natural mortality β_E from fitting EMs with alternative process error assumptions and treatment of median natural mortality (e_M^β known or estimated).

⁵⁹⁶ Covariate effect estimate and standard error example

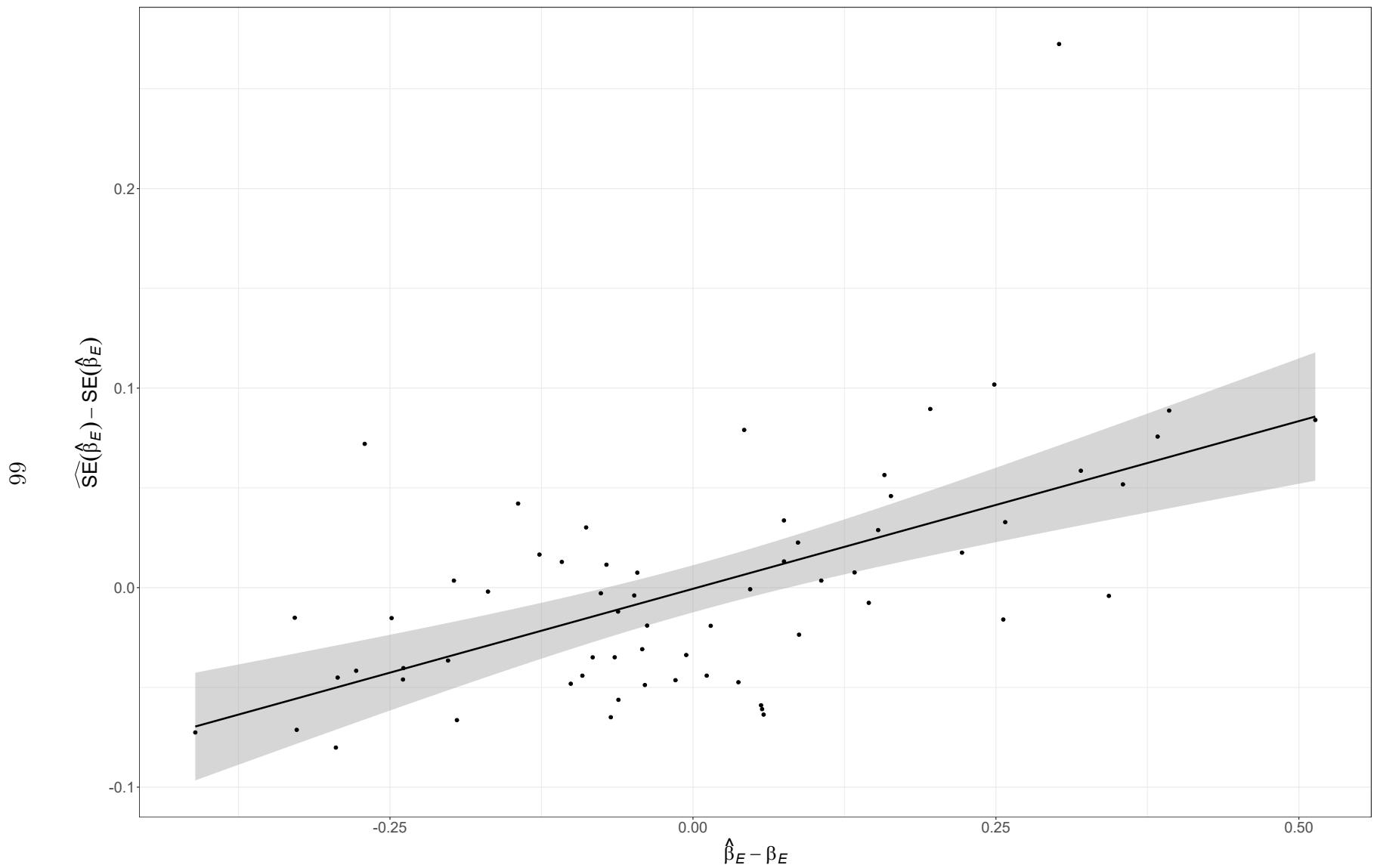


Fig. S24. Positive correlation of covariate effect estimates and Hessian-based standard error estimates for EM that also estimates the median natural mortality parameter and has correct R+M process error assumption fitted to simulated data from the OM with R+M process errors, temporal contrast in fishing pressure, low observation uncertainty for both population (*LowOE*) and covariate observations ($\sigma_e = 0.1$), high and uncorrelated temporal variability in the true covariate ($\sigma_E = 0.5$ and $\rho_E = 0$), and the strongest covariate effect on natural mortality ($\beta_E = 0.5$).

₅₉₇ Median Natural mortality parameter bias

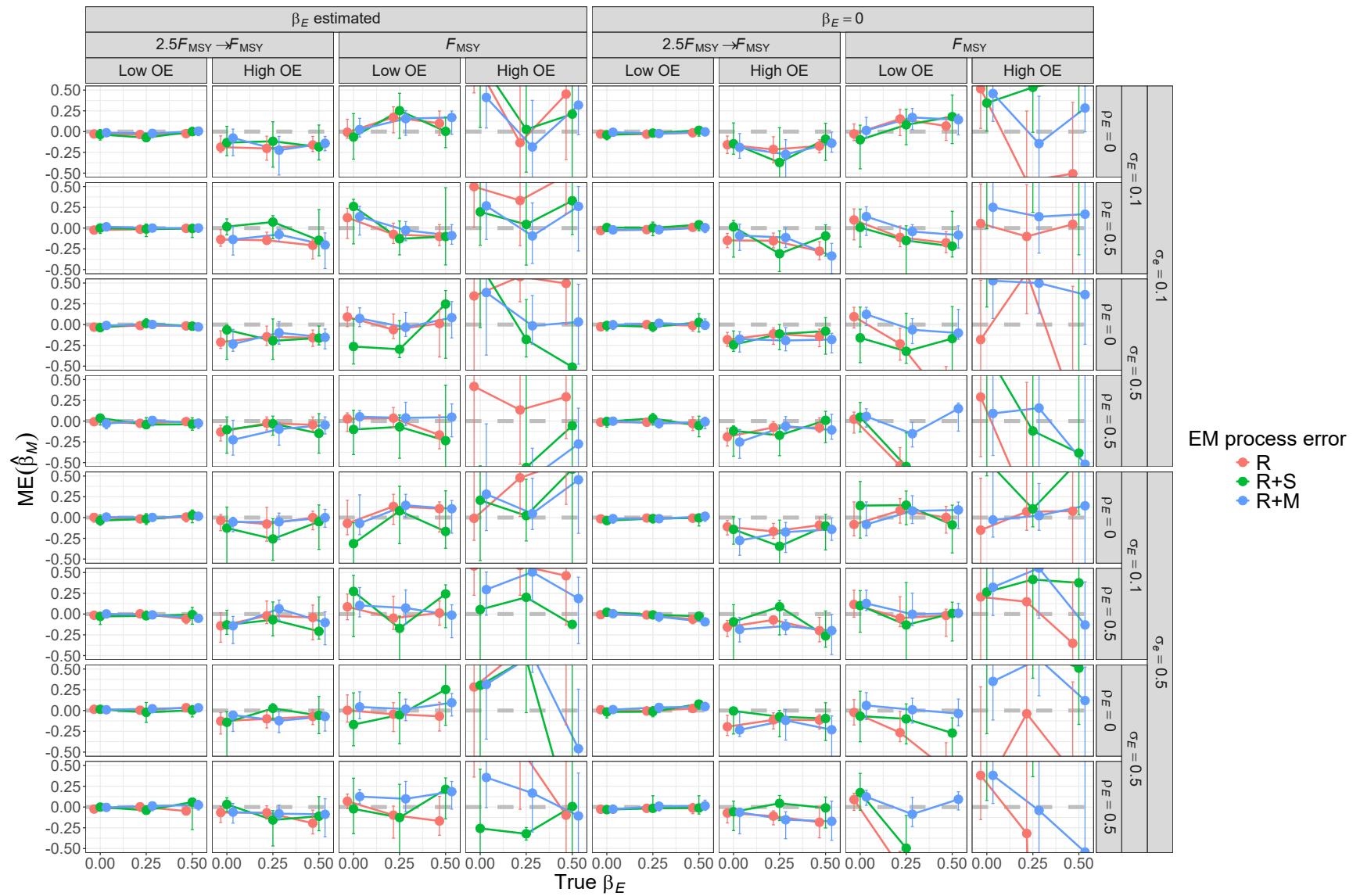


Fig. S25. For R OMs, median error (ME) of estimates of β_M from fitting EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated). Vertical lines represent 95% confidence intervals.

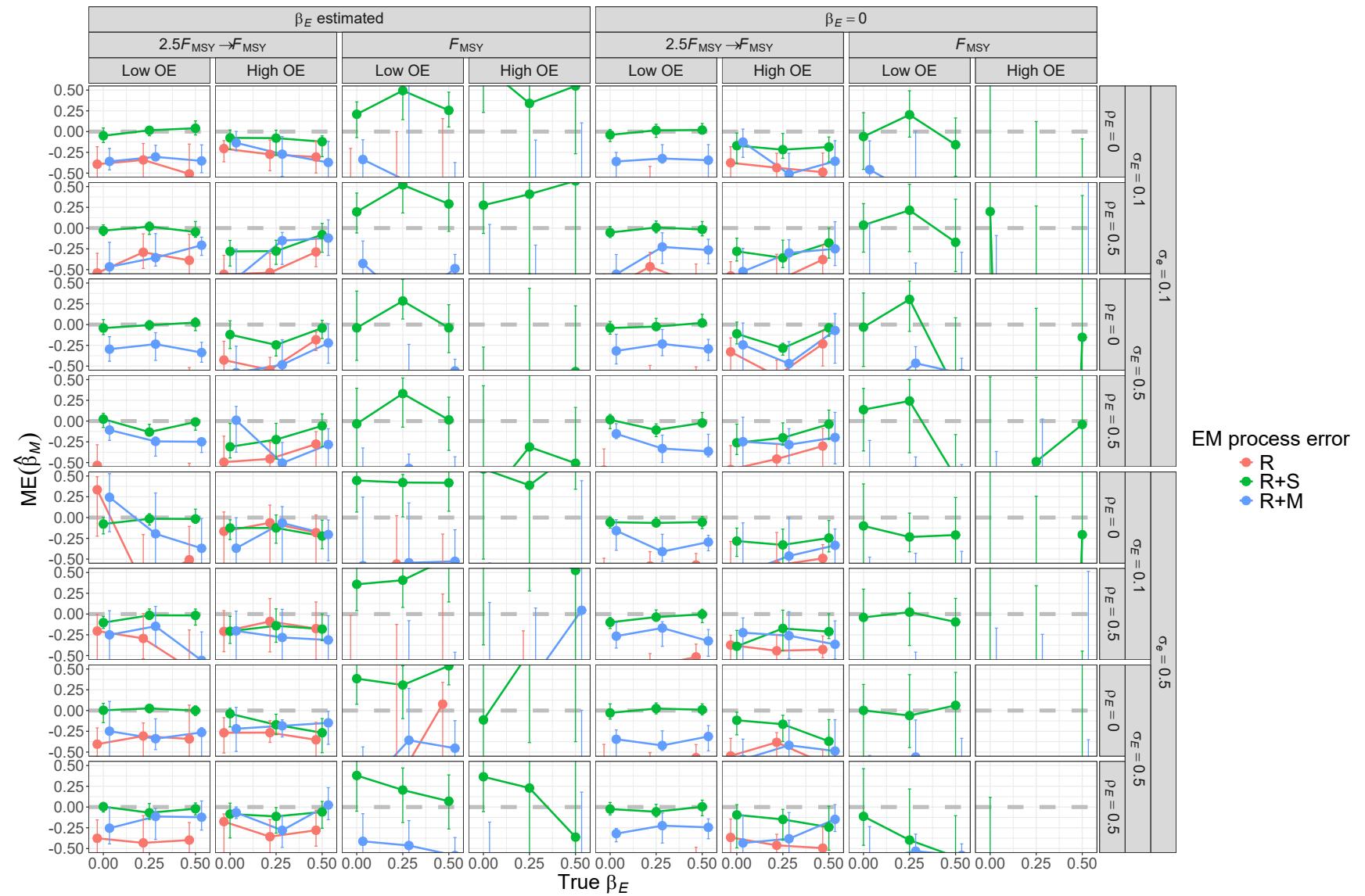


Fig. S26. For R+S OMs, median error (ME) of estimates of β_M from fitting EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated). Vertical lines represent 95% confidence intervals.

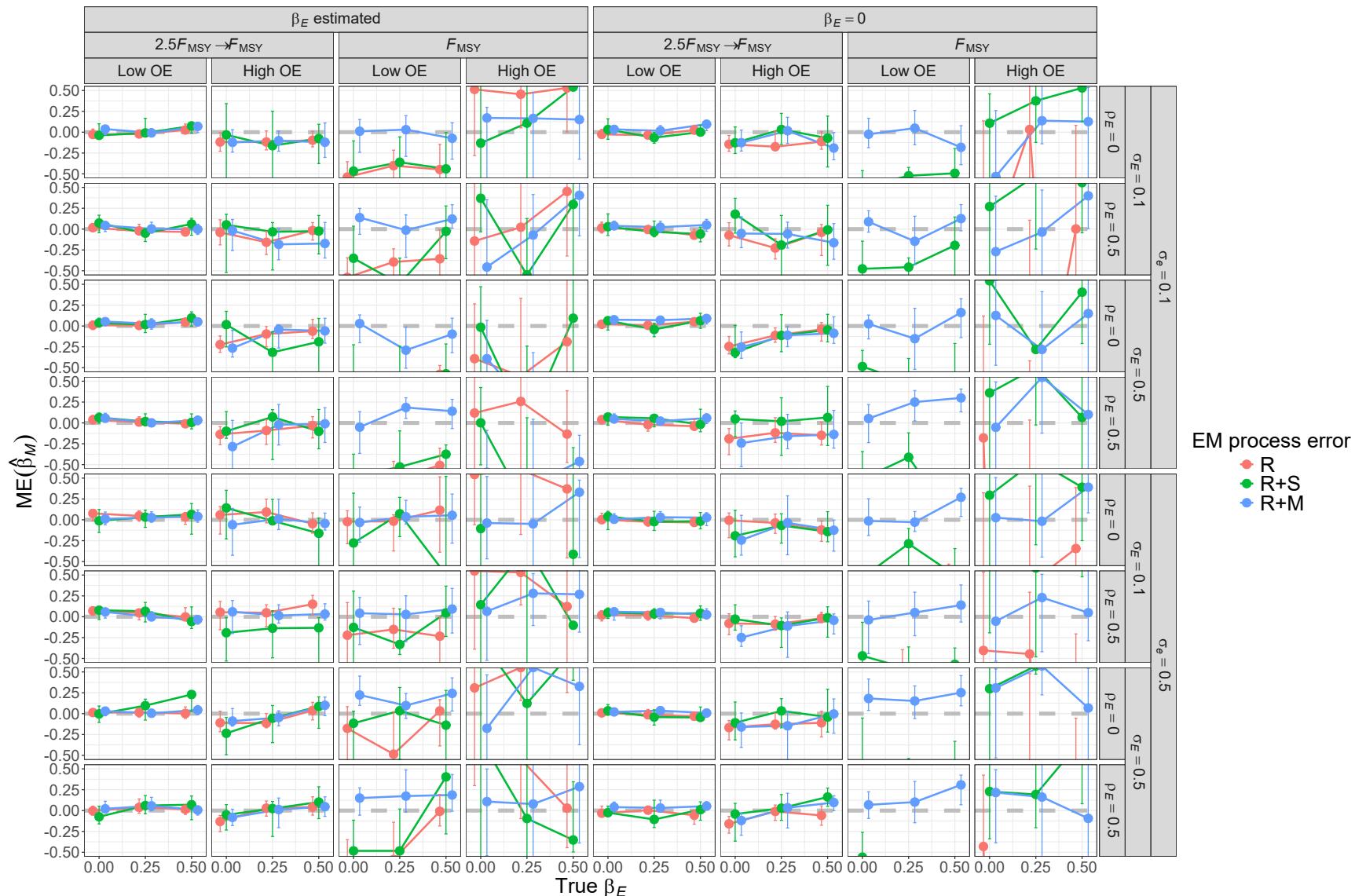


Fig. S27. For R+M OMs, median error (ME) of estimates of β_M from fitting EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated). Vertical lines represent 95% confidence intervals.

⁵⁹⁸ Median natural mortality parameter standard error estimation bias

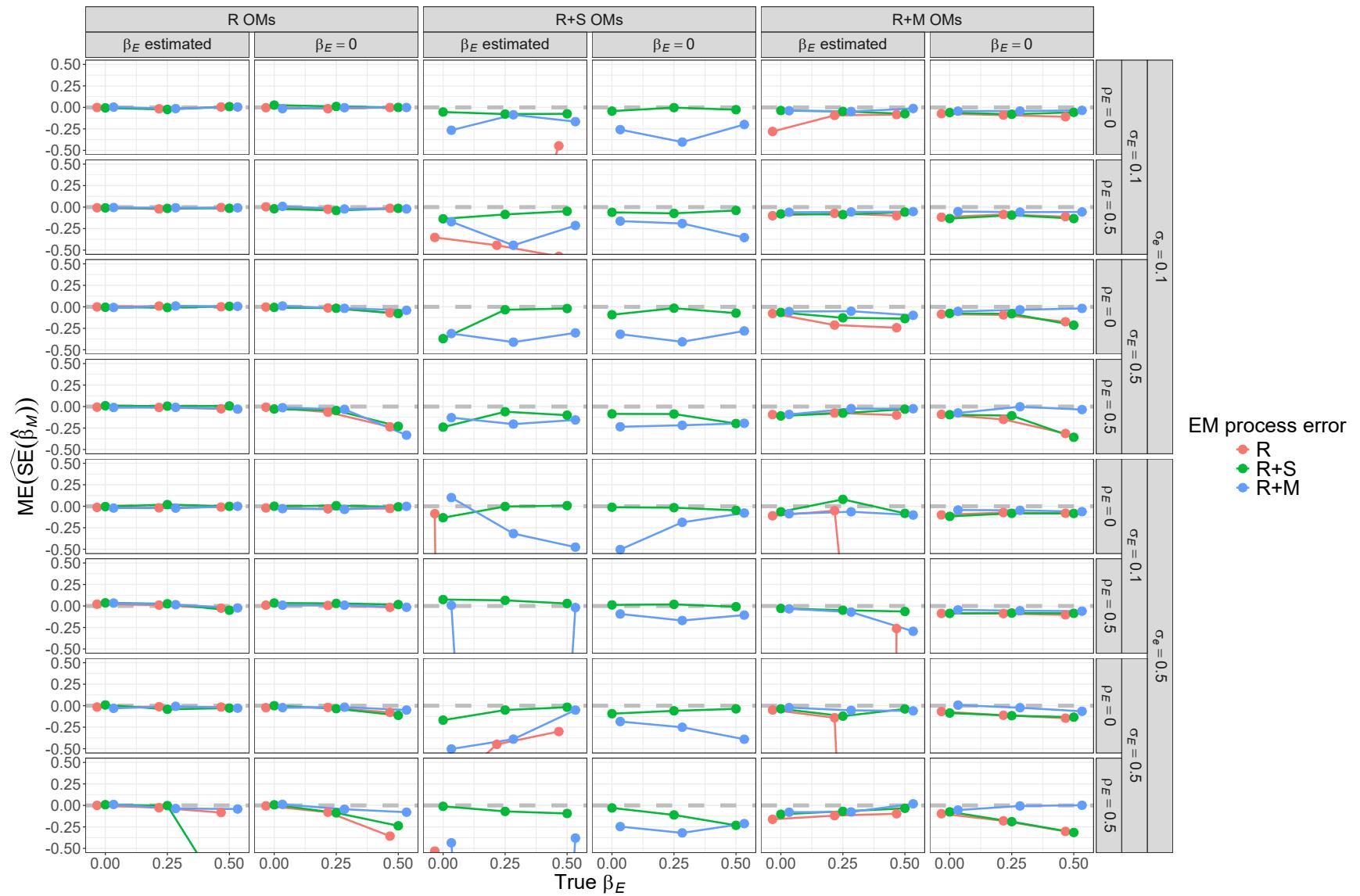


Fig. S28. Median error (ME) of Hessian-based estimates of standard error for median natural mortality parameter β_M from fitting EMs with alternative process error assumptions and treatment of the covariate effect ($\beta_E = 0$ or estimated). All OMs had low observation error and contrast in fishing mortality. True standard error is defined as the mean of the standard error estimates across converged fits to simulated data sets for a given OM scenario.

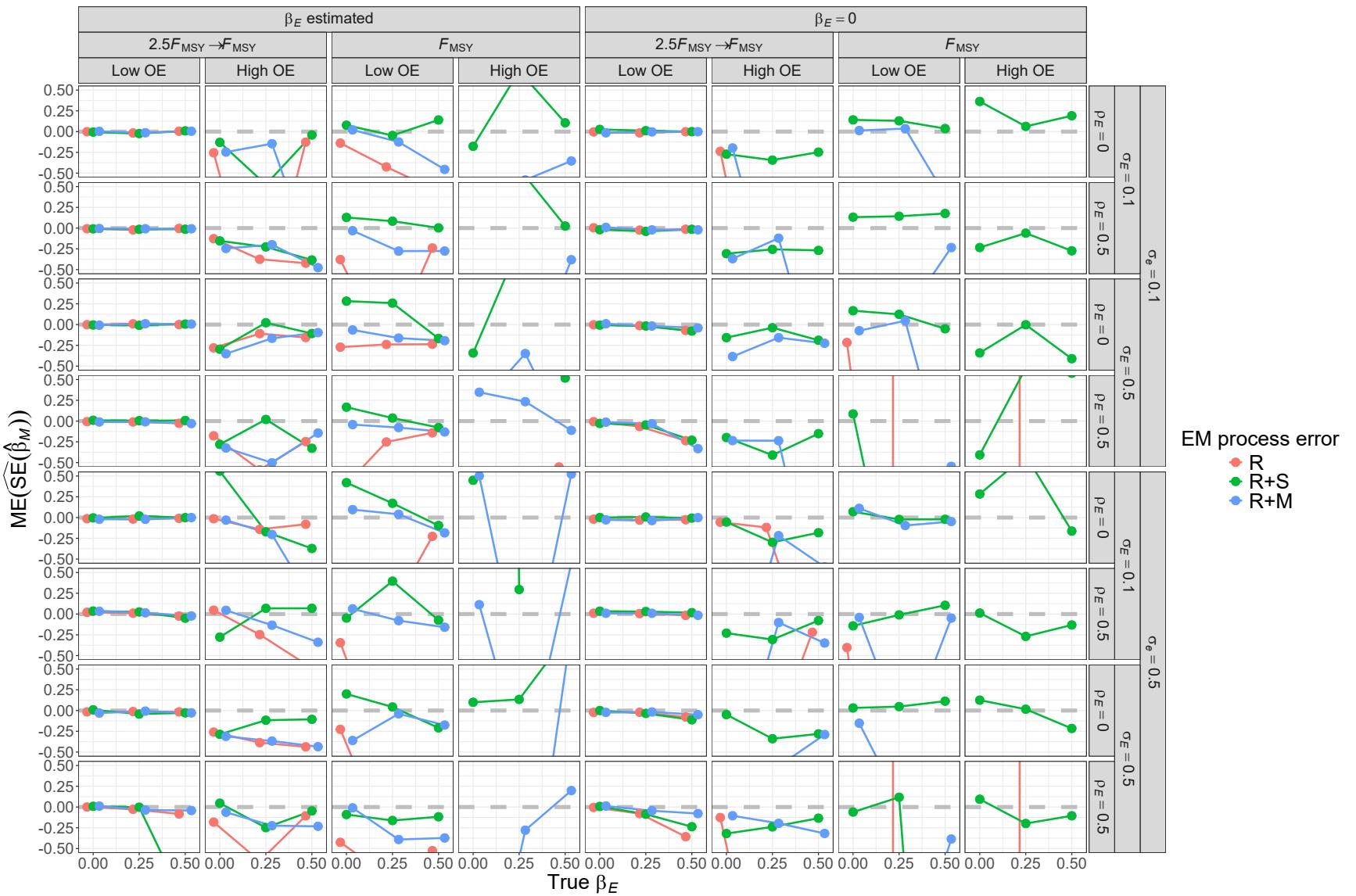


Fig. S29. For R OMs, median error (ME) of Hessian-based estimates of standard error for median natural mortality parameter β_M from fitting EMs with alternative process error assumptions and treatment of the covariate effect ($\beta_E = 0$ or estimated). True standard error is defined as the mean of the standard error estimates across converged fits to simulated data sets for a given OM scenario.

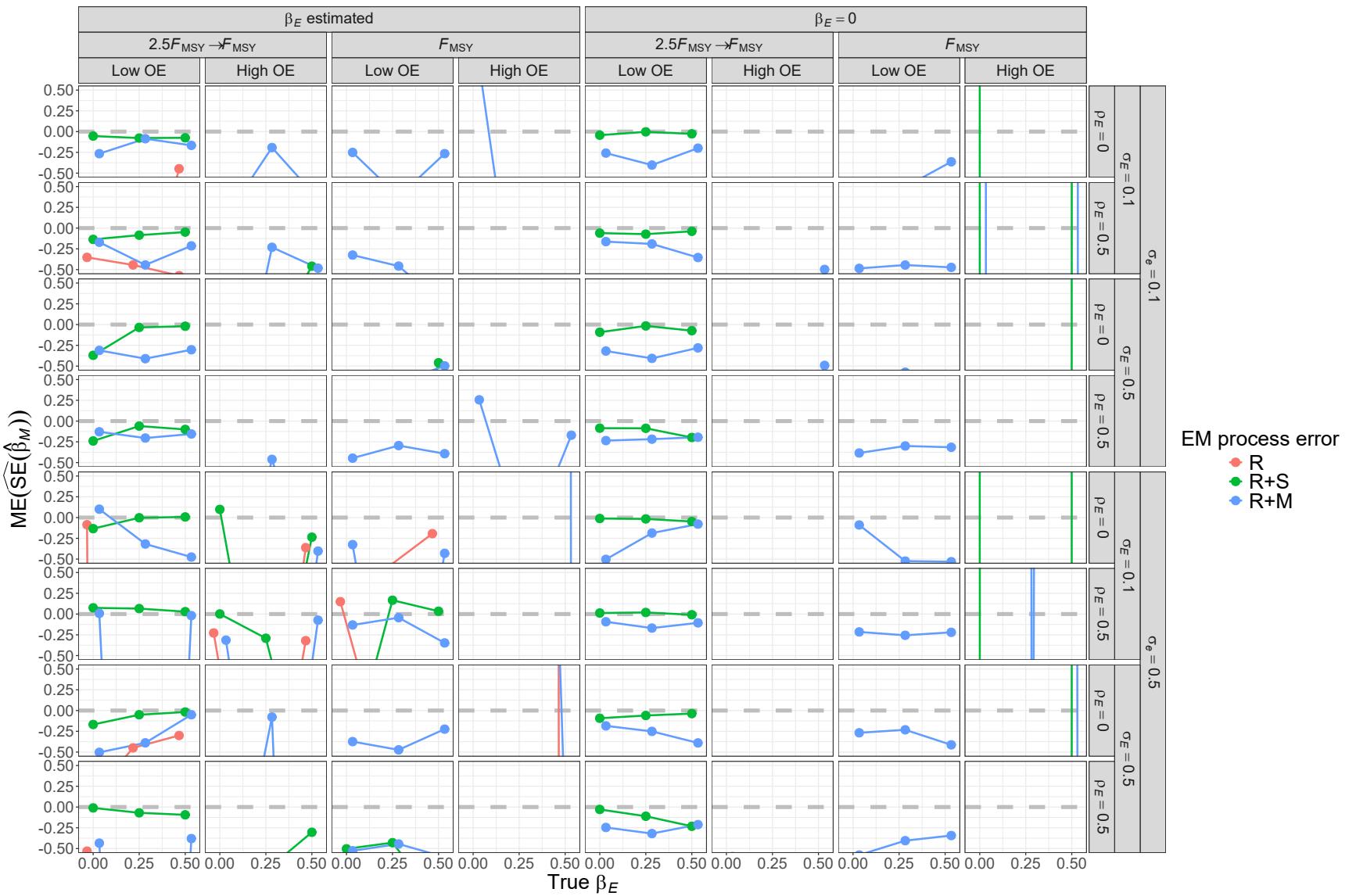


Fig. S30. For R+S OMs, median error (ME) of Hessian-based estimates of standard error for median natural mortality parameter β_M from fitting EMs with alternative process error assumptions and treatment of the covariate effect ($\beta_E = 0$ or estimated). True standard error is defined as the mean of the standard error estimates across converged fits to simulated data sets for a given OM scenario.

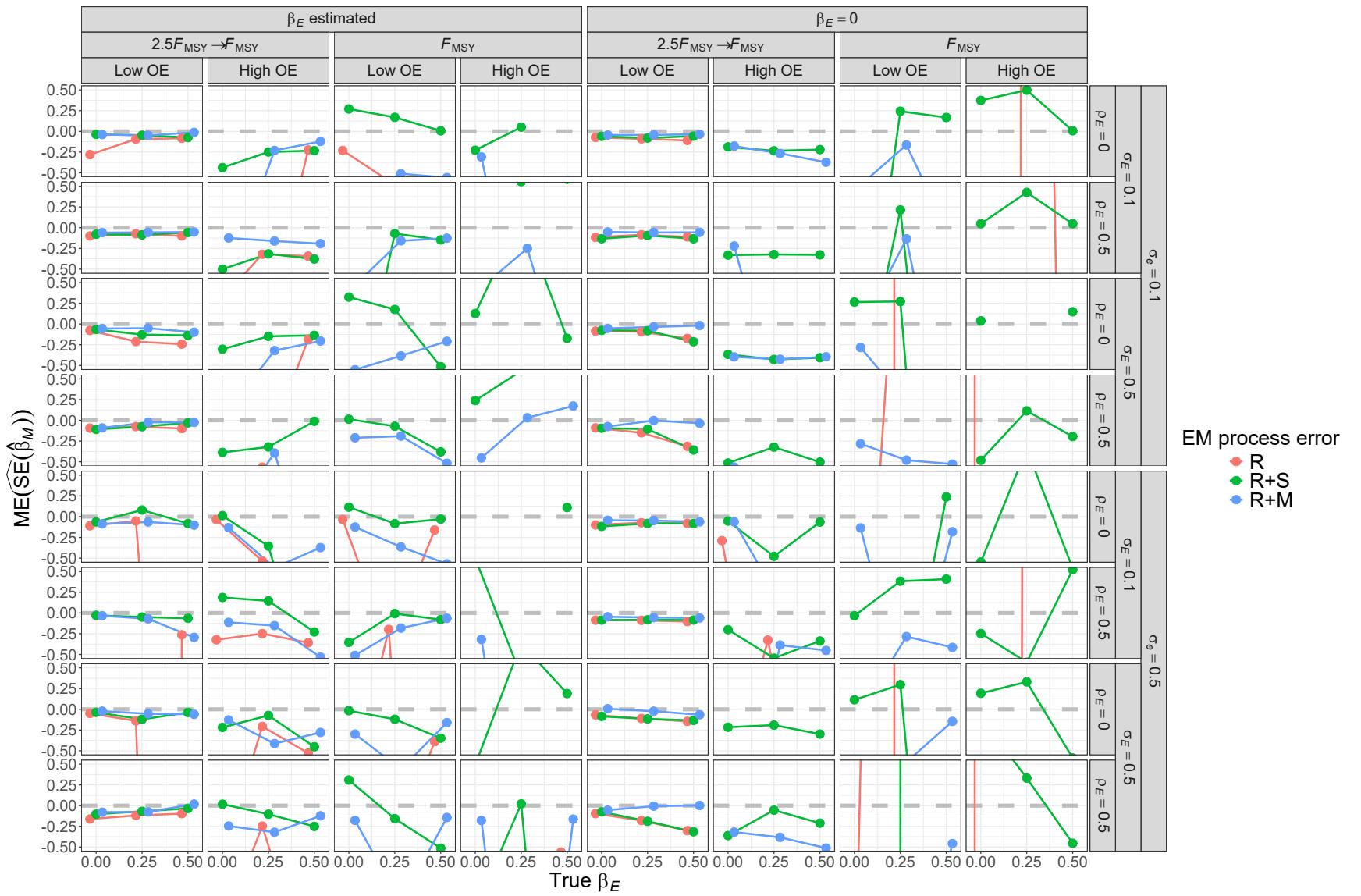


Fig. S31. For R+M OMs, median error (ME) of Hessian-based estimates of standard error for median natural mortality parameter β_M from fitting EMs with alternative process error assumptions and treatment of the covariate effect ($\beta_E = 0$ or estimated). True standard error is defined as the mean of the standard error estimates across converged fits to simulated data sets for a given OM scenario.

⁵⁹⁹ Median Natural mortality parameter confidence interval coverage

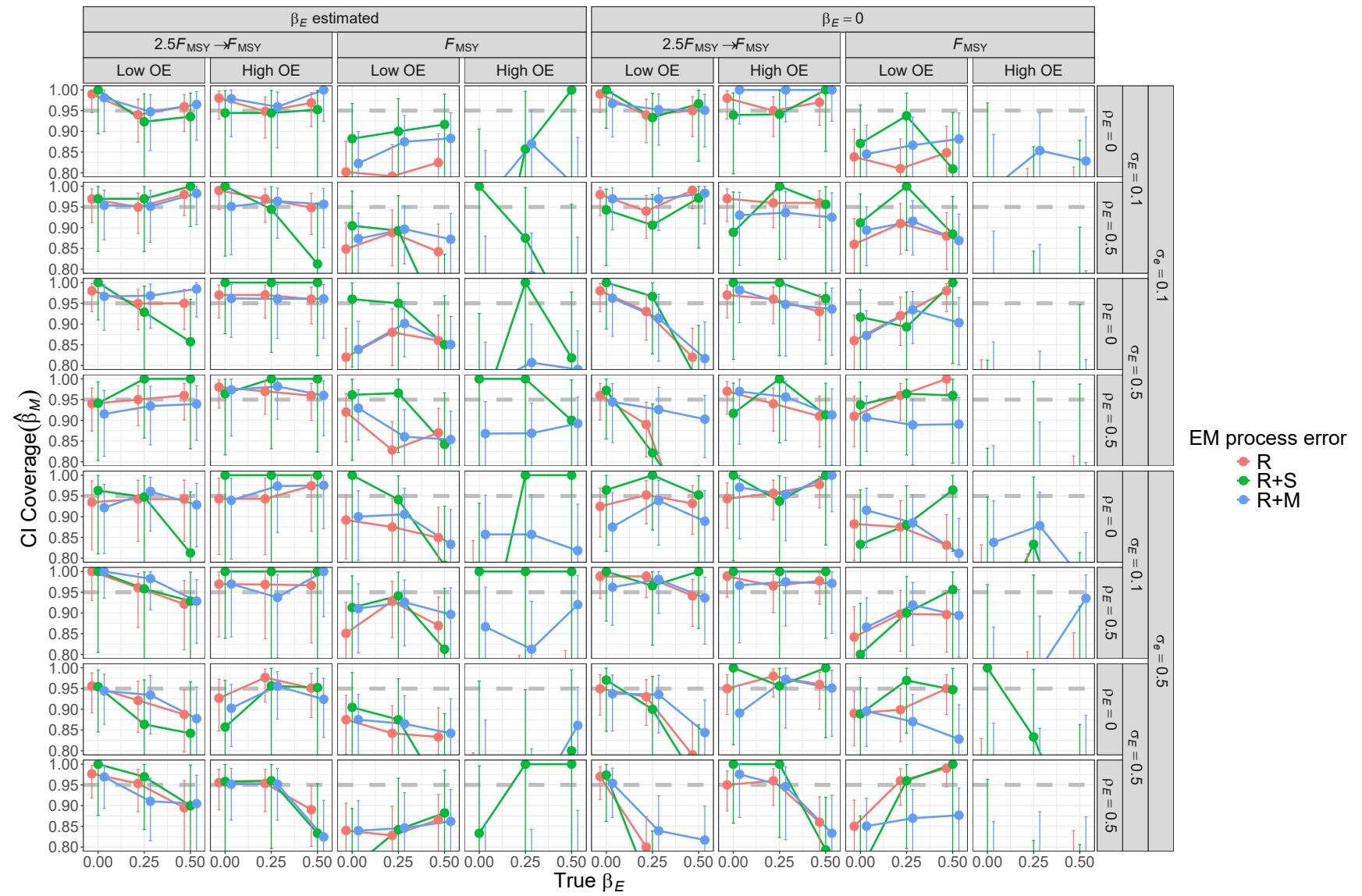


Fig. S32. For R OMs, probability of 95% confidence interval for β_M containing the true value for EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated). Vertical lines represent 95% confidence intervals.

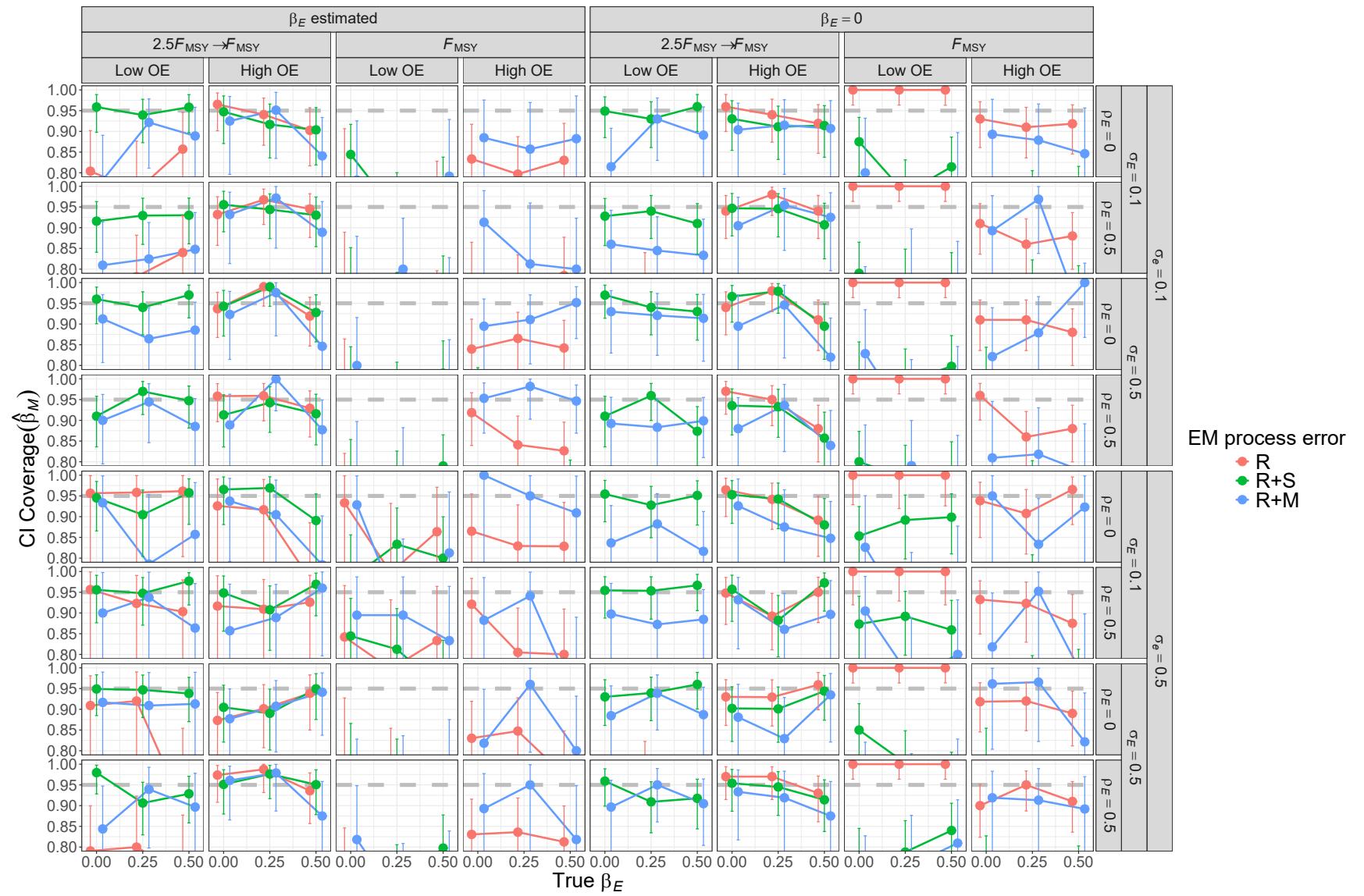


Fig. S33. For R+S OMs, probability of 95% confidence interval for β_M containing the true value for EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated). Vertical lines represent 95% confidence intervals.

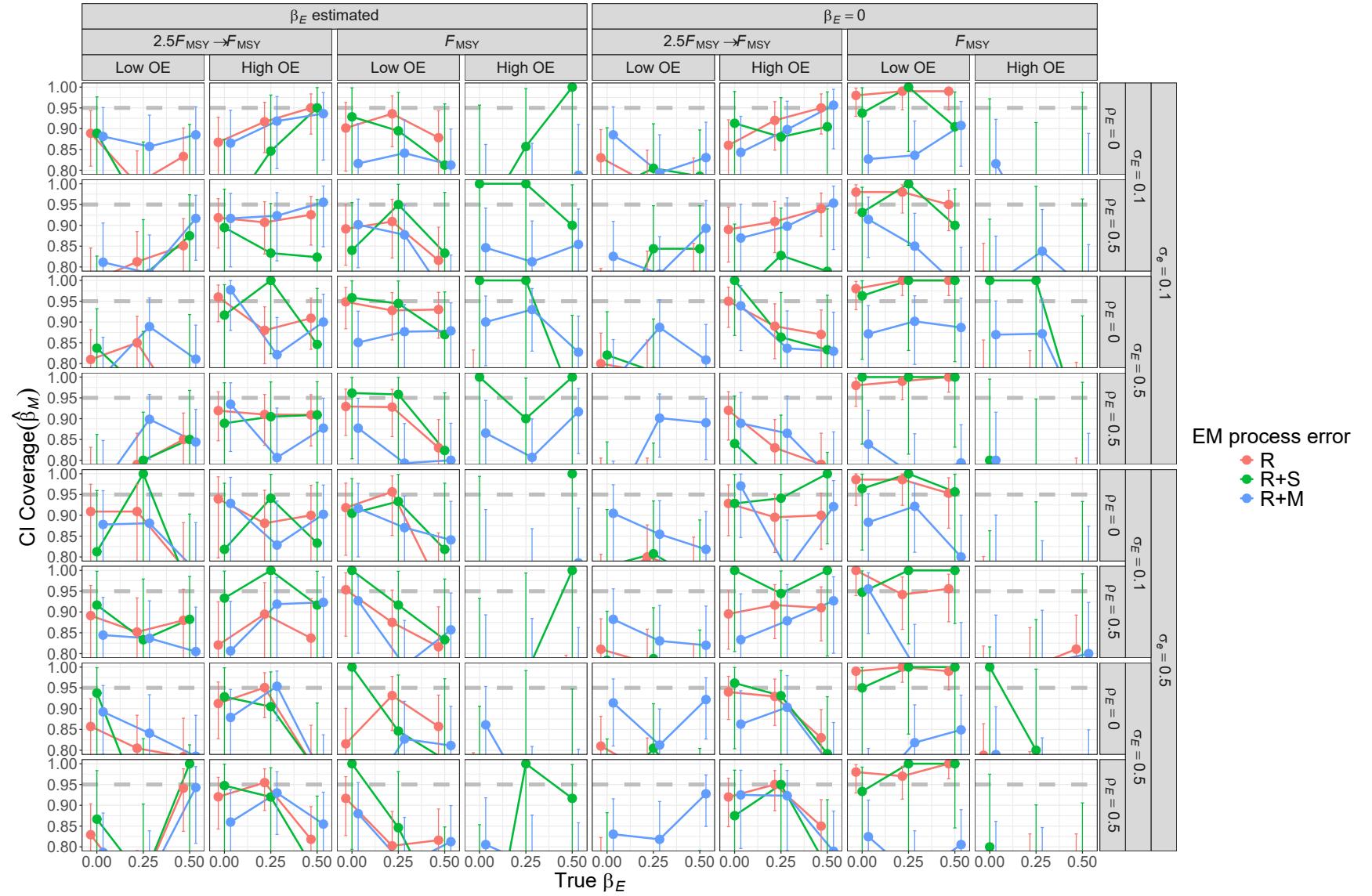


Fig. S34. For R+M OMs, probability of 95% confidence interval for β_M containing the true value for EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated). Vertical lines represent 95% confidence intervals.

600 Median natural mortality parameter estimate and standard error

601 example

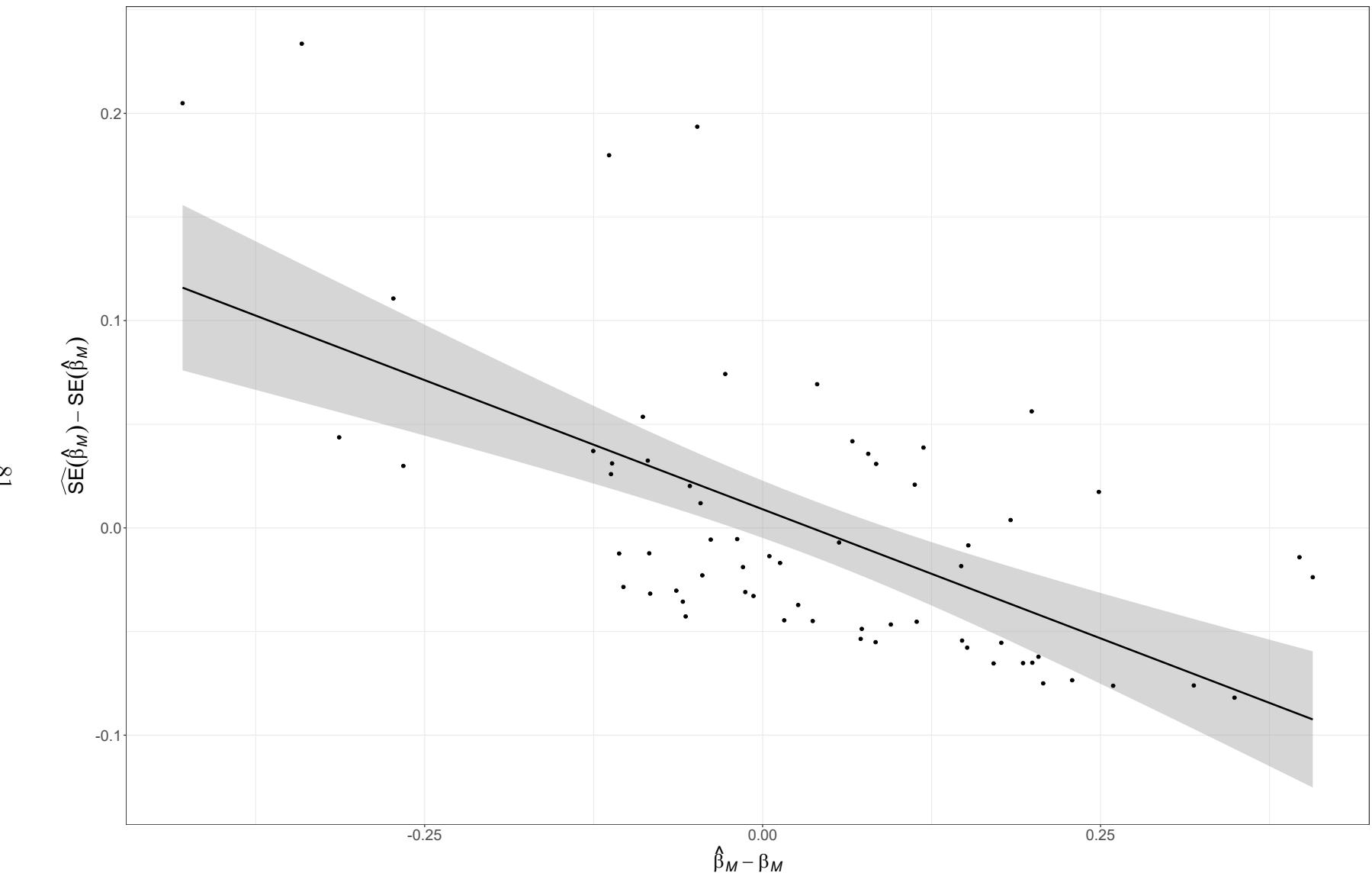


Fig. S35. Negative correlation of β_M estimates and Hessian-based standard error estimates for EM that also estimates the covariate effect and has correct R+M process error assumption fitted to simulated data from the OM with R+M process errors, temporal contrast in fishing pressure, low observation uncertainty for both population (*LowOE*) and covariate observations ($\sigma_e = 0.1$), high and uncorrelated temporal variability in the true covariate ($\sigma_E = 0.5$ and $\rho_E = 0$), and the strongest covariate effect on natural mortality ($\beta_E = 0.5$).

₆₀₂ Median Natural mortality parameter RMSE

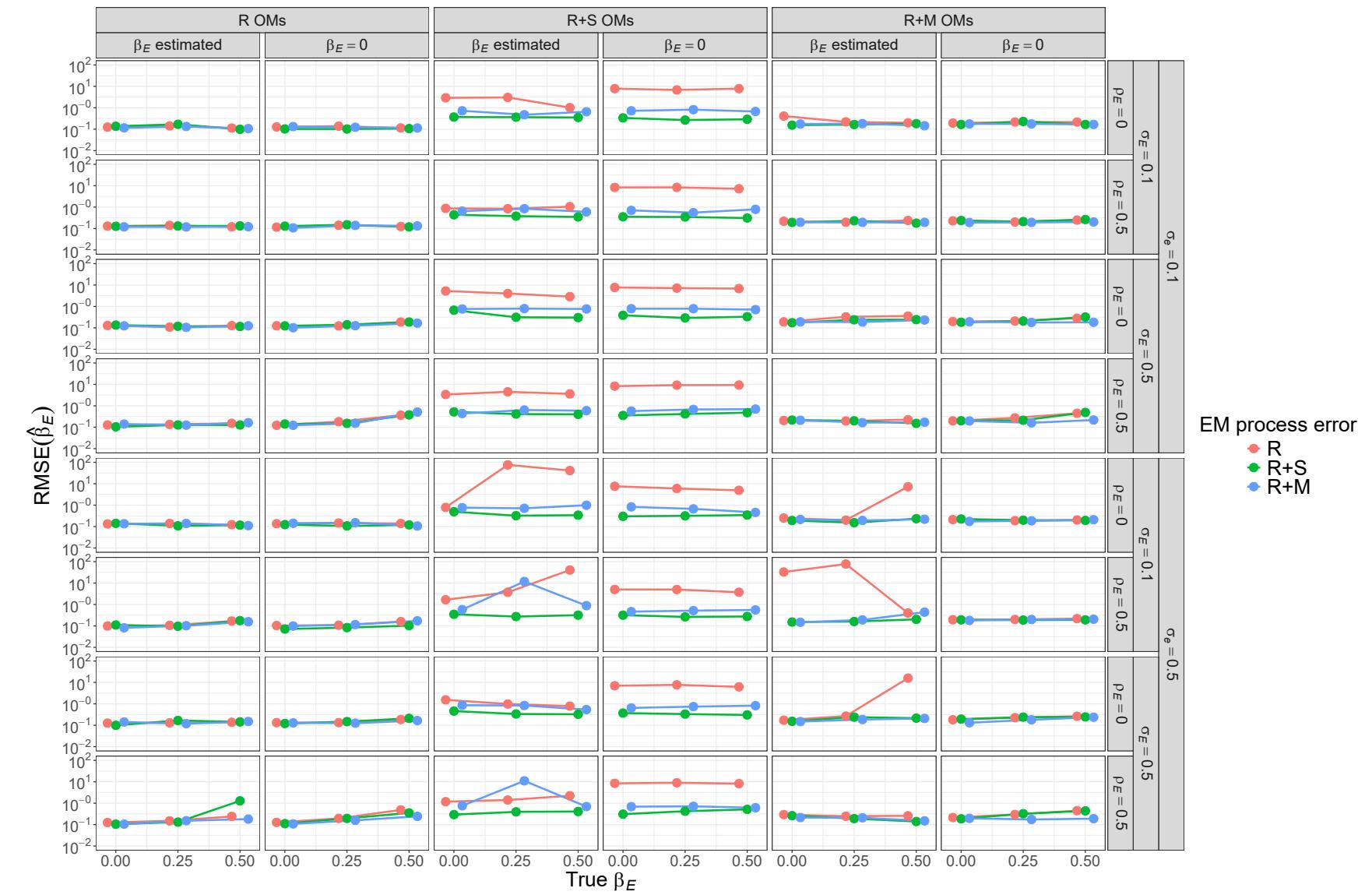


Fig. S36. Root mean square error (RMSE) of estimates of β_M from fitting EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated). All OMs had low observation error for population observations and contrast in fishing mortality.

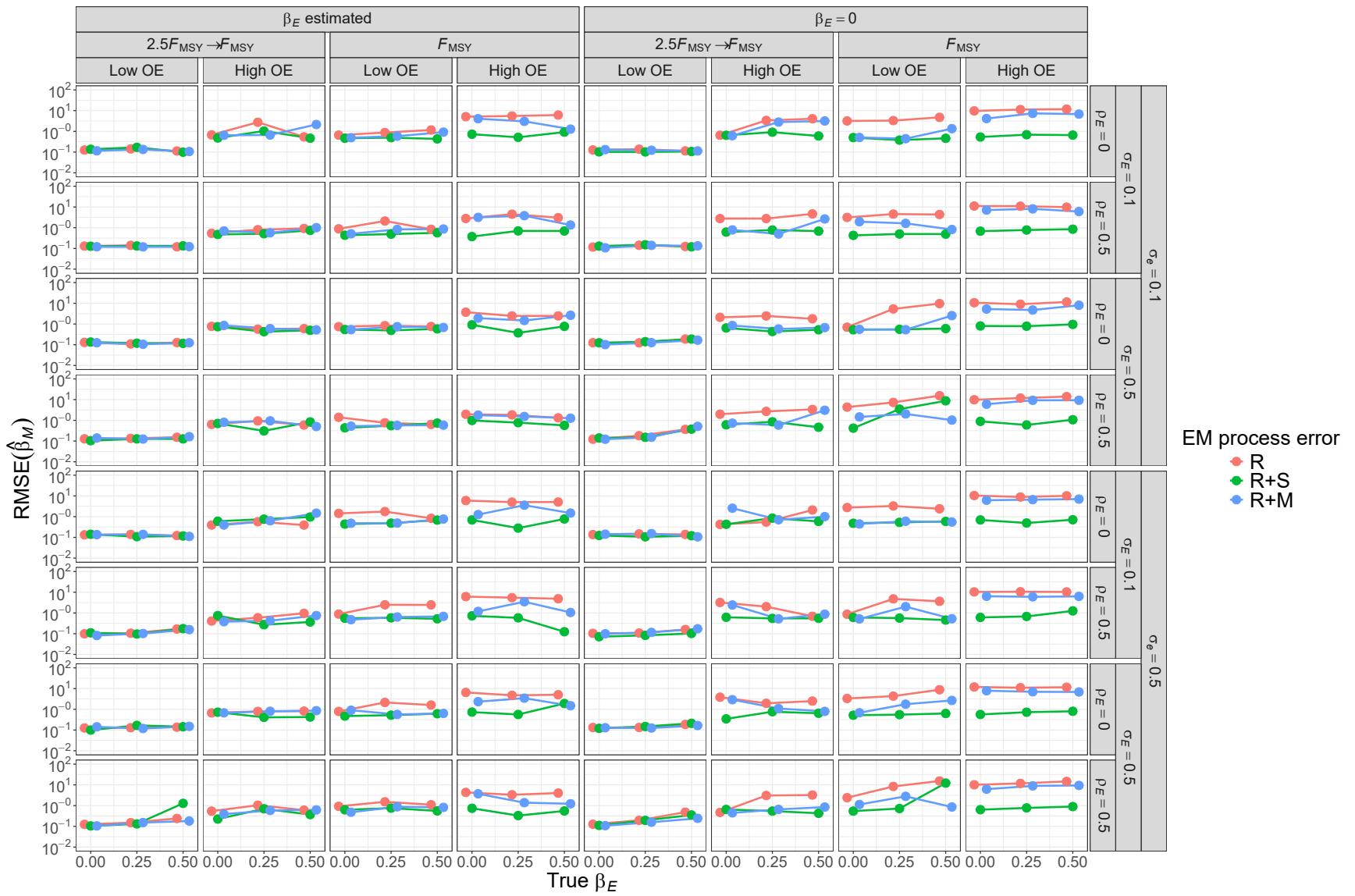


Fig. S37. For R OMs, root mean square error (RMSE) of estimates of β_M from fitting EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated).

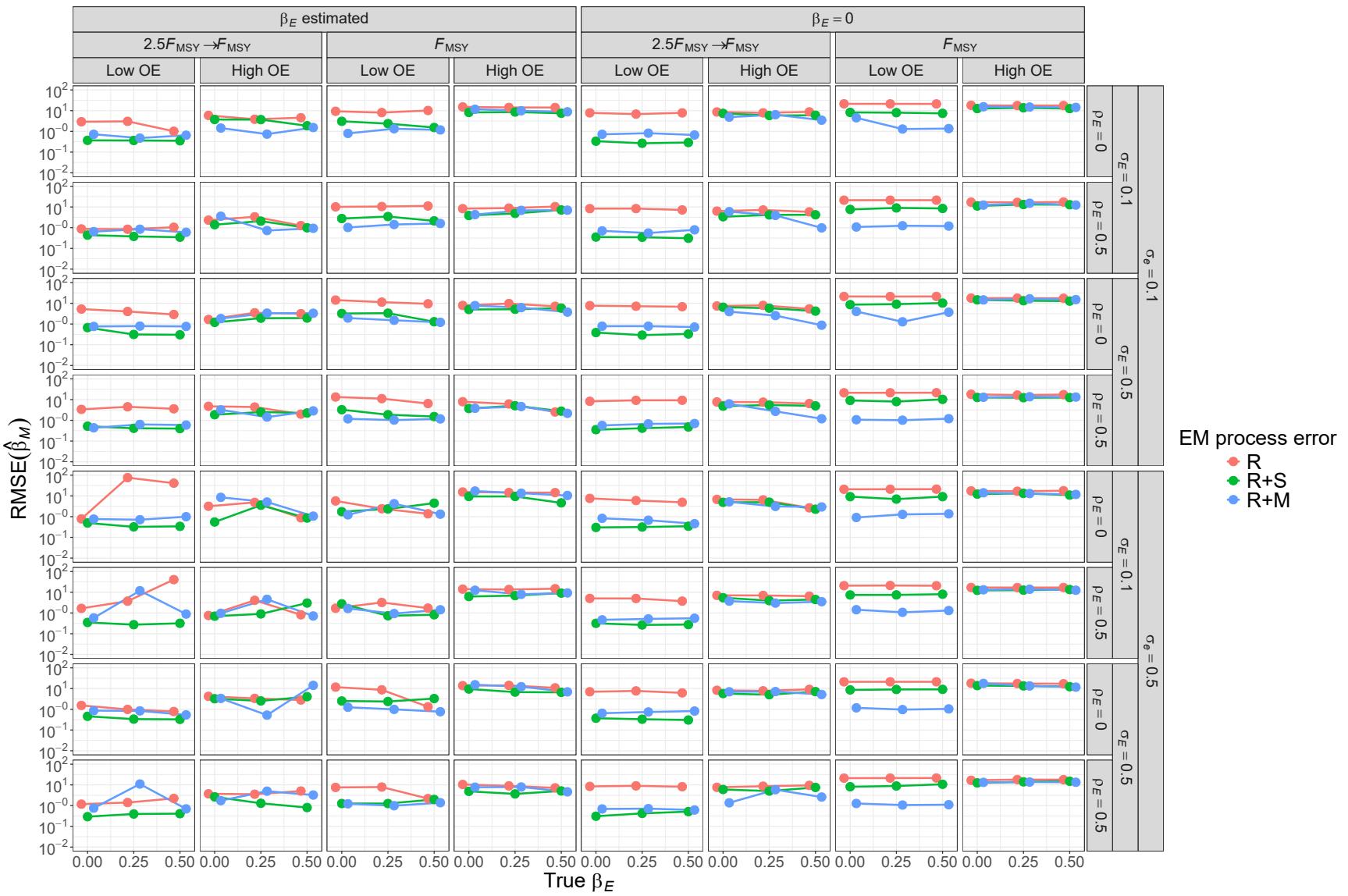


Fig. S38. For R+S OMs, root mean square error (RMSE) of estimates of β_M from fitting EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated).

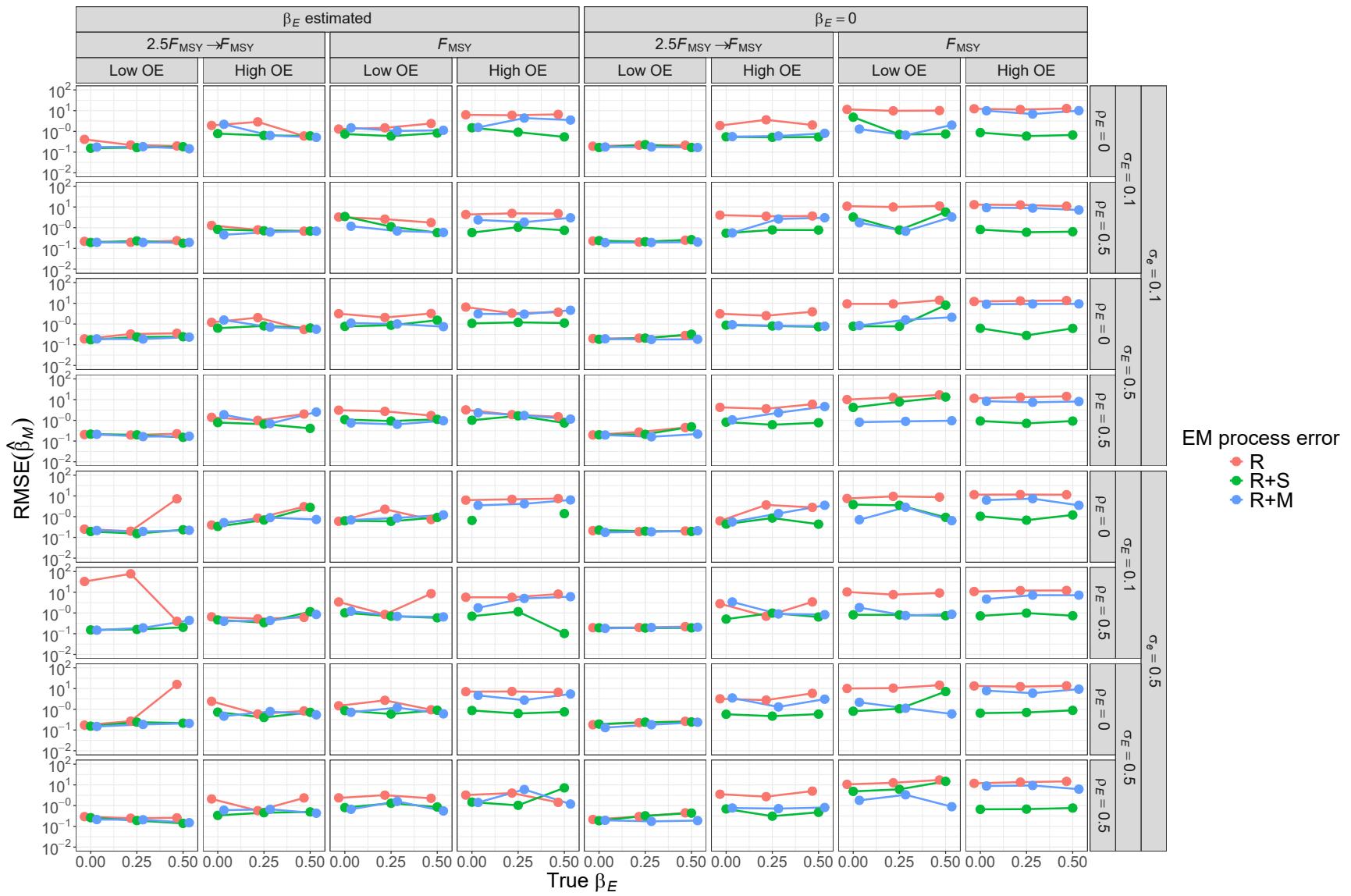


Fig. S39. For R+M OMs, root mean square error (RMSE) of estimates of β_M from fitting EMs with alternative process error assumptions and treatment of covariate effect ($\beta_E = 0$ or estimated).

⁶⁰³ Terminal year natural mortality bias

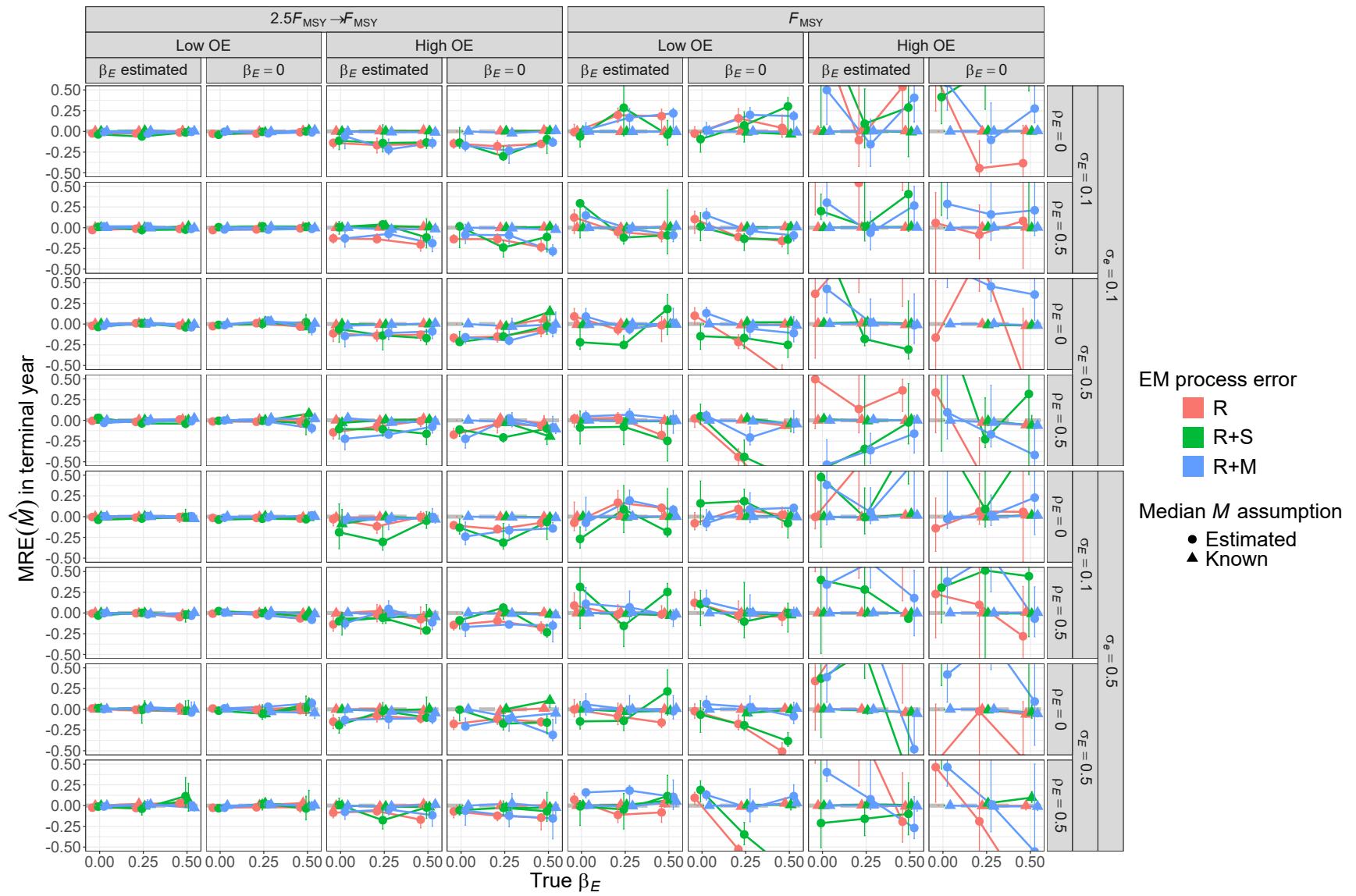


Fig. S40. For R OMs, median relative error (MRE) of estimates of natural mortality rate in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

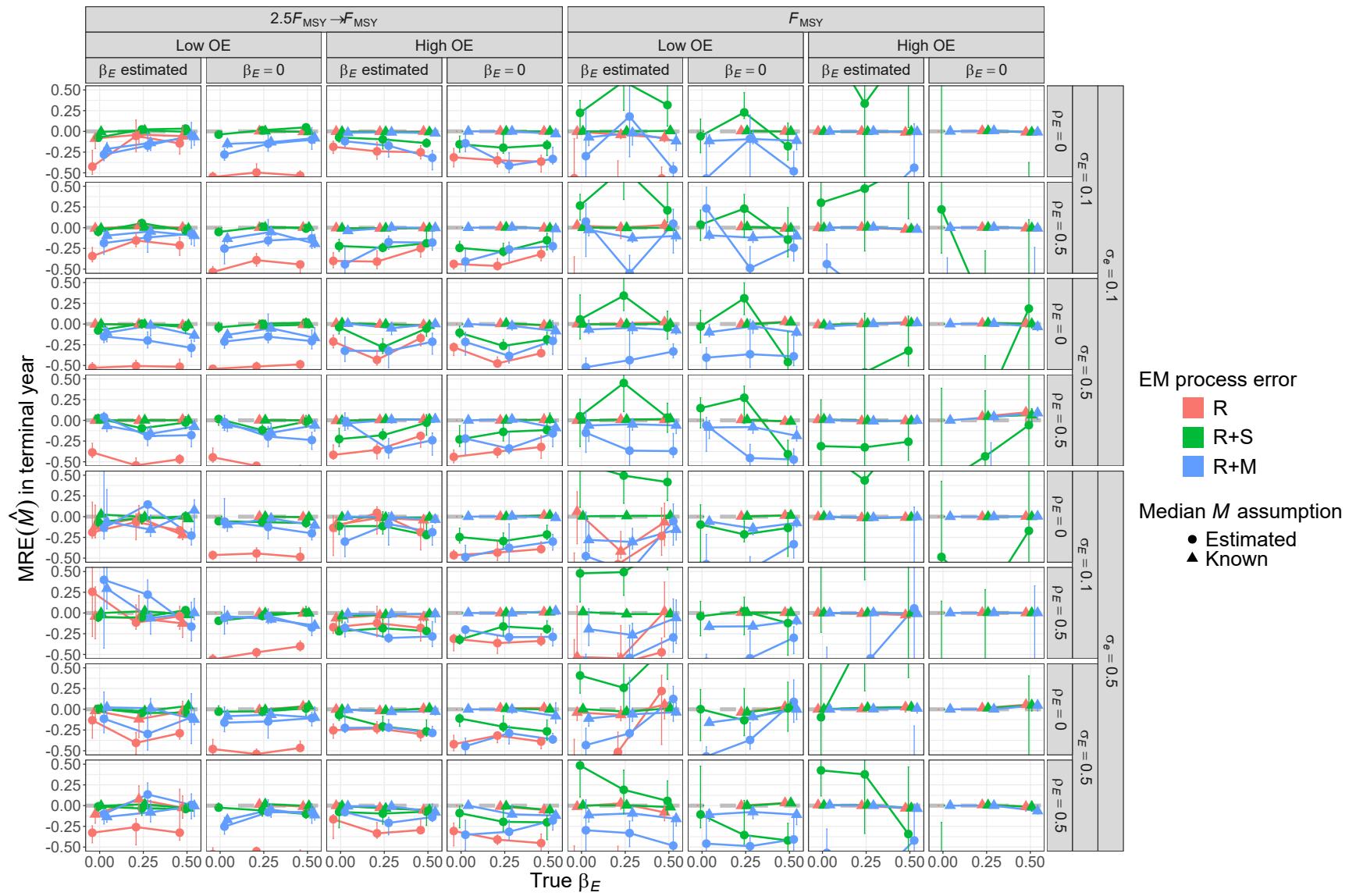


Fig. S41. For R+S OMs, median relative error (MRE) of estimates of natural mortality rate in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

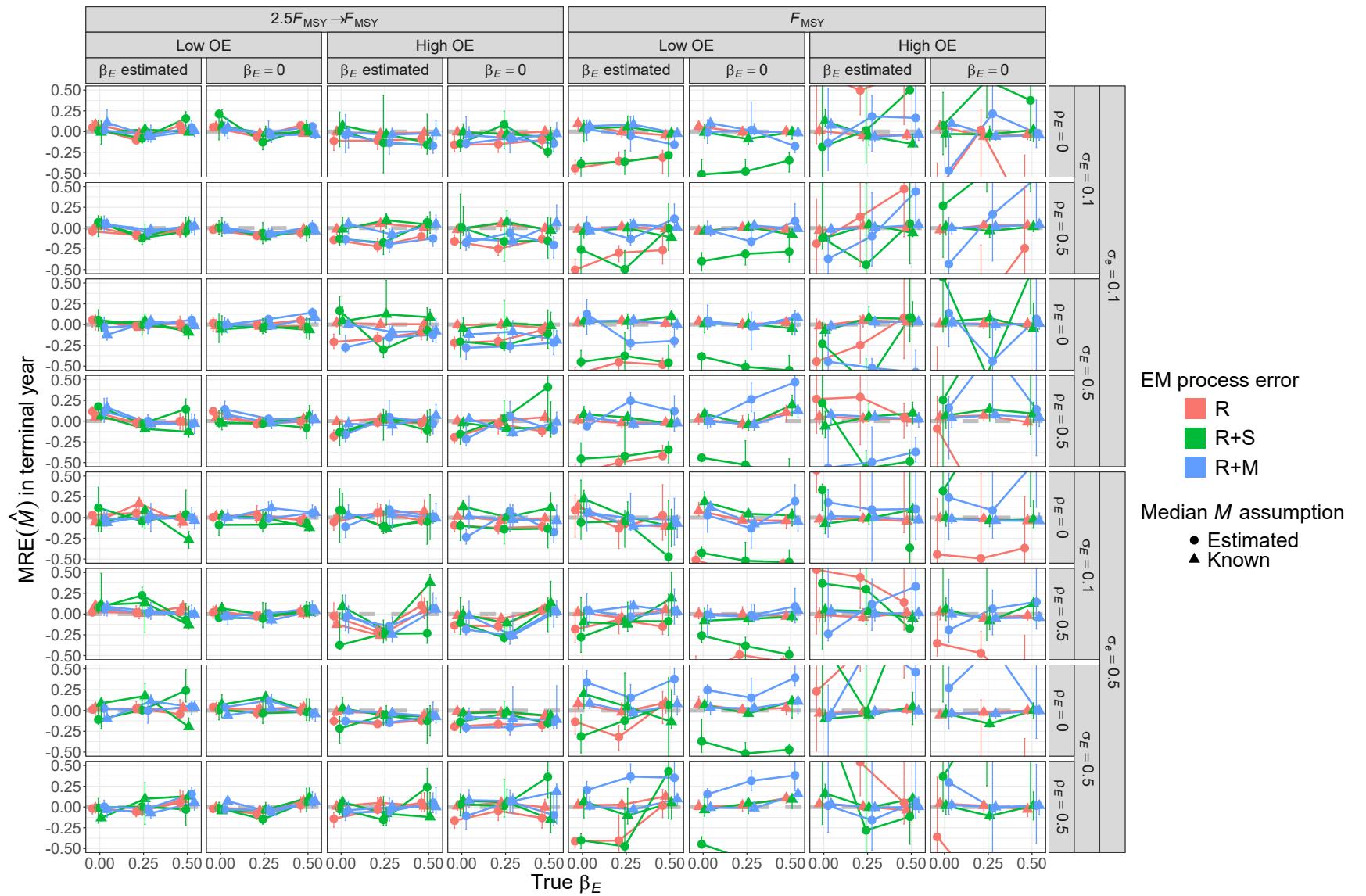


Fig. S42. For R+M OMs, median relative error (MRE) of estimates of natural mortality rate in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

⁶⁰⁴ Terminal year natural mortality RMSE

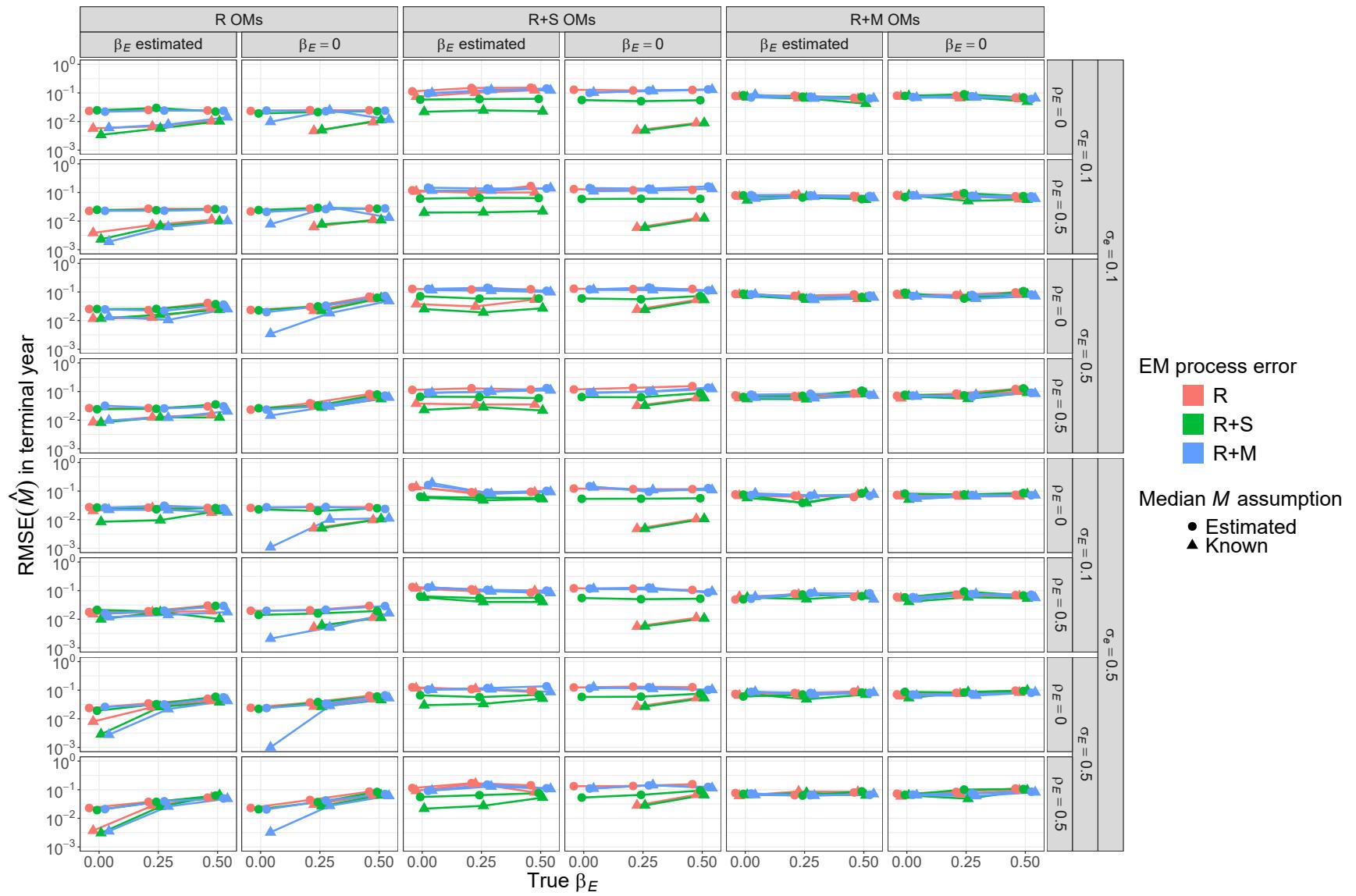


Fig. S43. Root mean square error (RMSE) of estimates of natural mortality rate in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known). All OMs had low population observation error and contrast in fishing mortality.

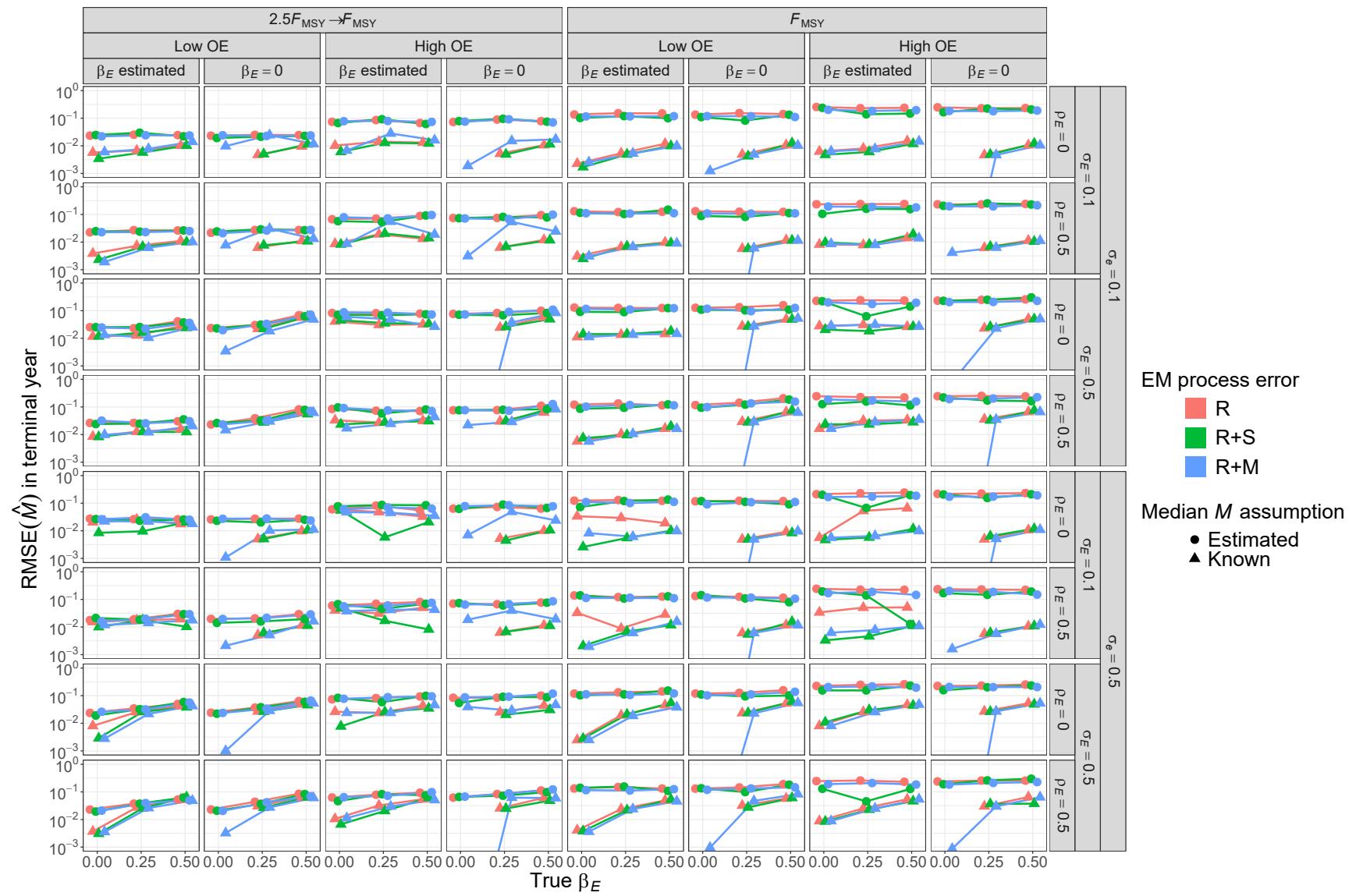


Fig. S44. For R OMs, root mean square error (RMSE) of estimates of natural mortality rate in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

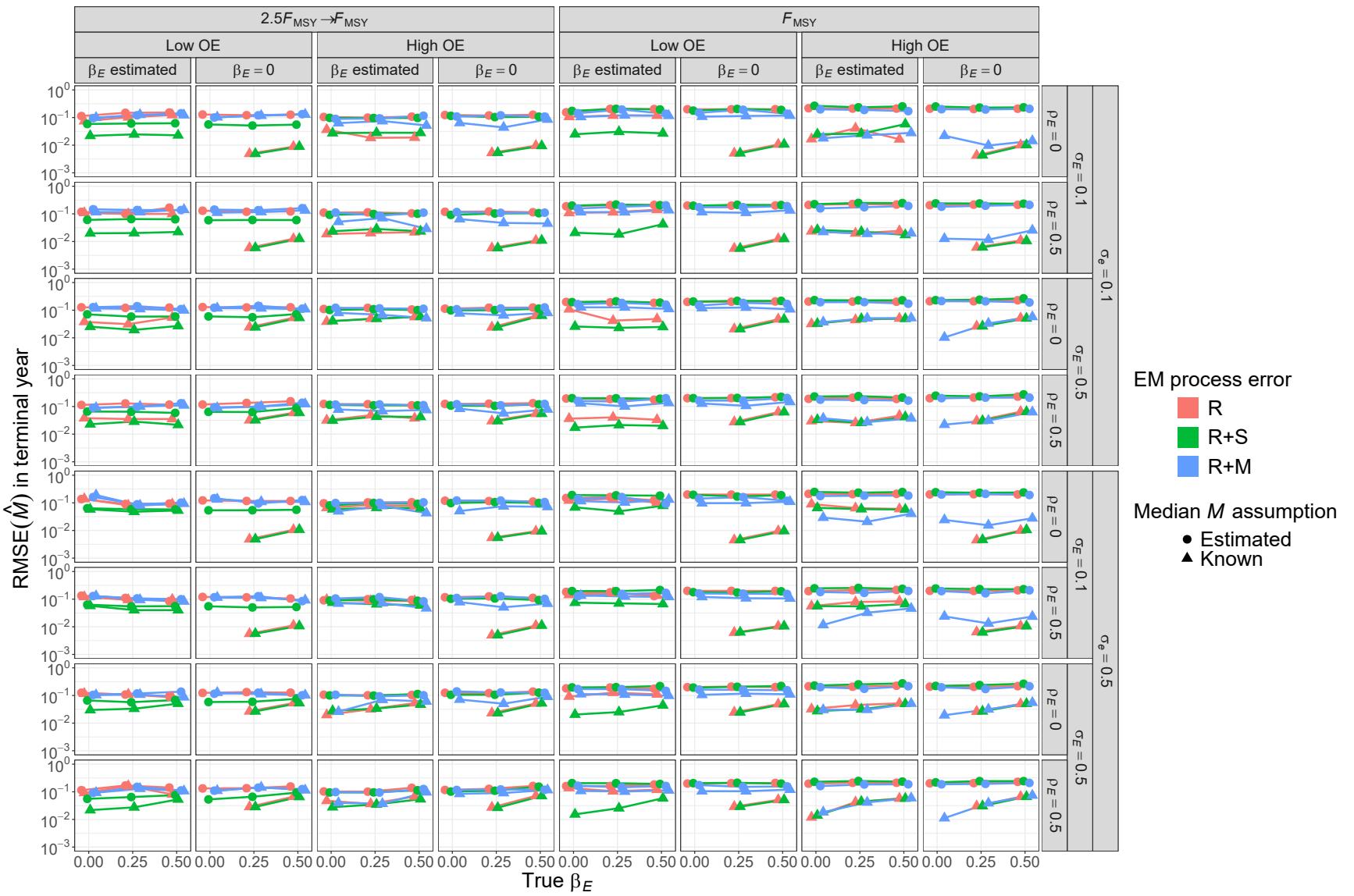


Fig. S45. For R+S OMs, root mean square error (RMSE) of estimates of natural mortality rate in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

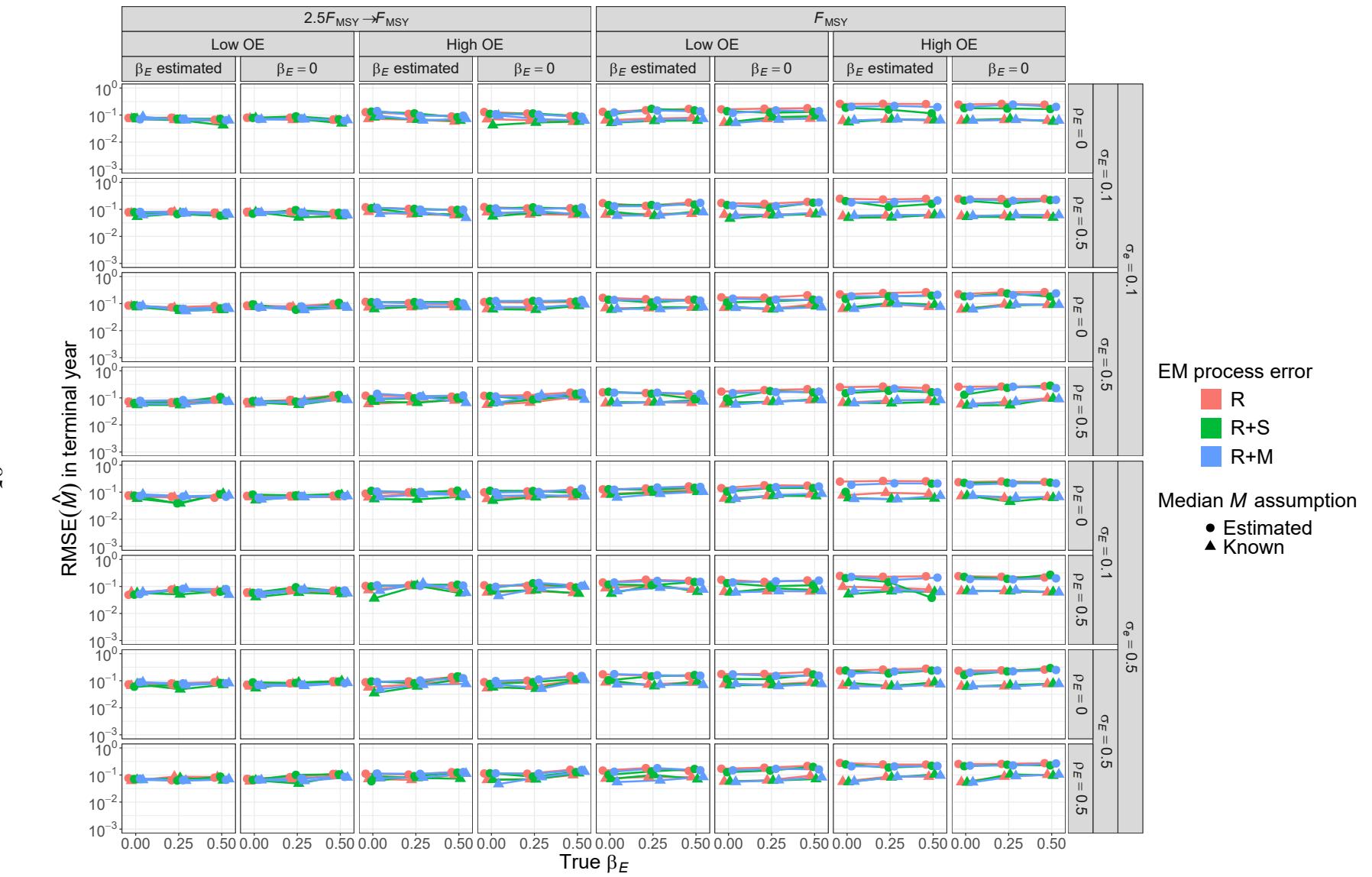


Fig. S46. For R+M OMs, root mean square error (RMSE) of estimates of natural mortality rate in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

605 Terminal year spawning stock biomass bias

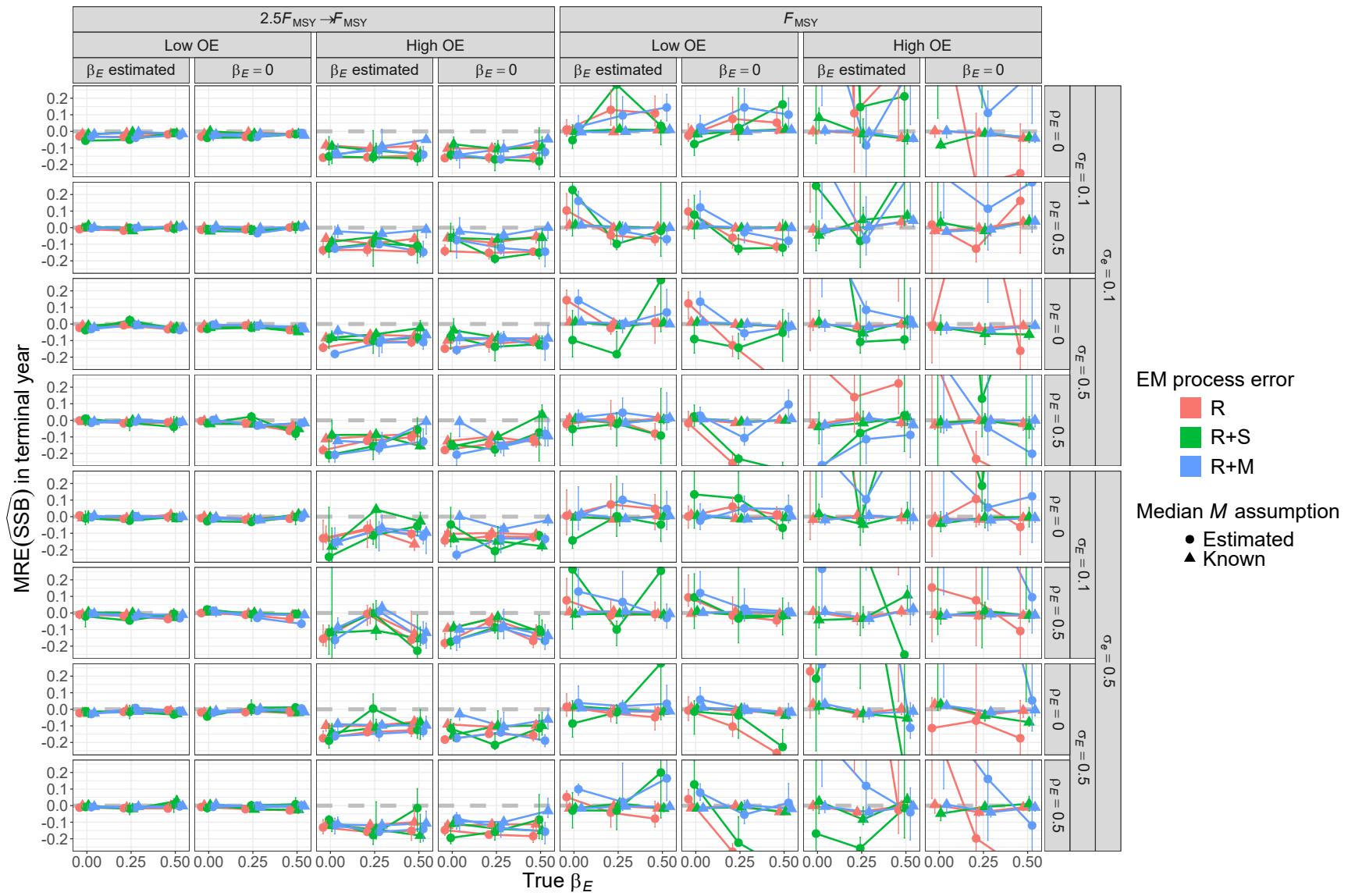


Fig. S47. For R OMs, median relative error (MRE) of estimates of spawning stock biomass (SSB) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

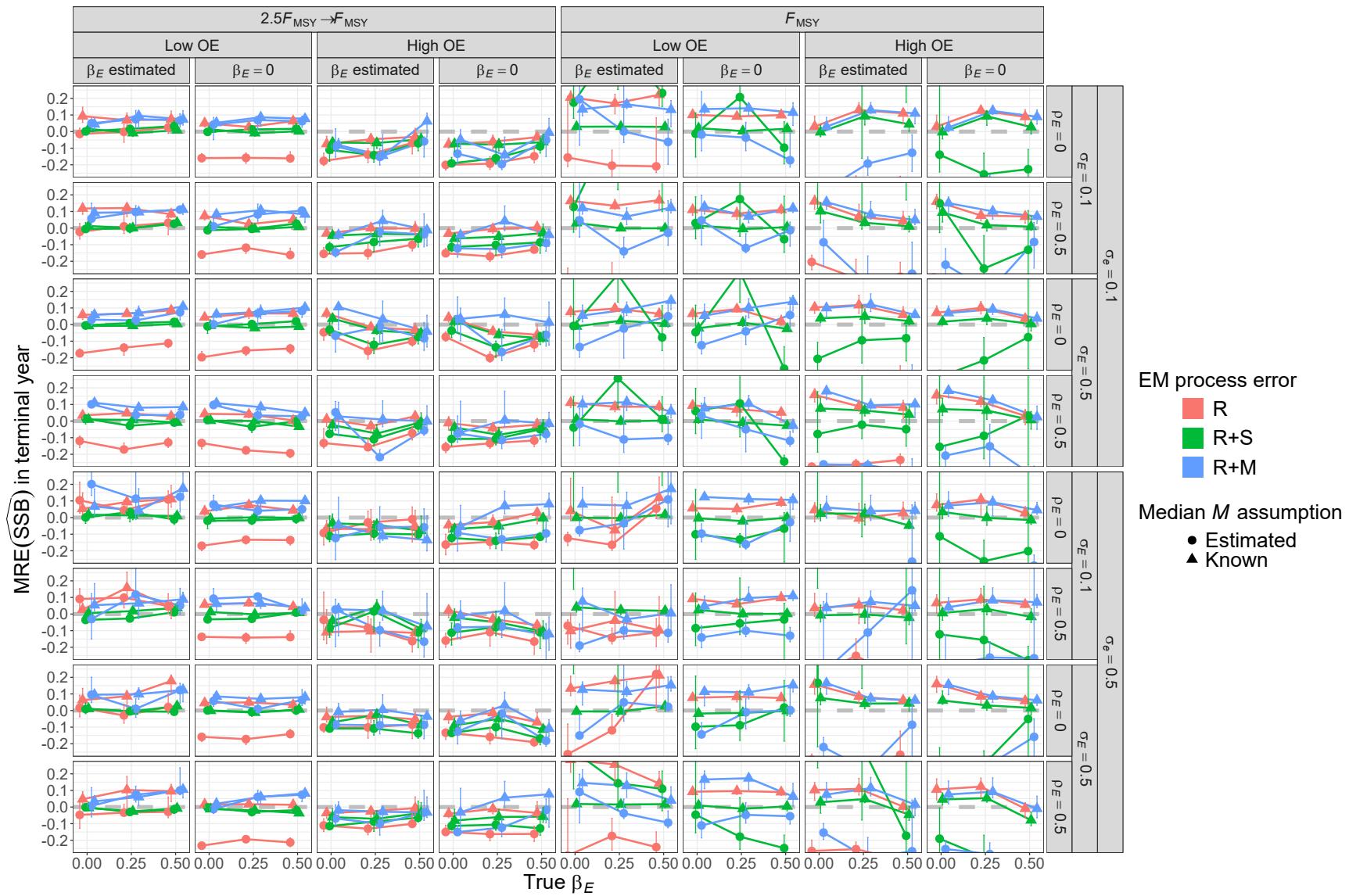


Fig. S48. For R+S OMs, median relative error (MRE) of estimates of spawning stock biomass (SSB) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

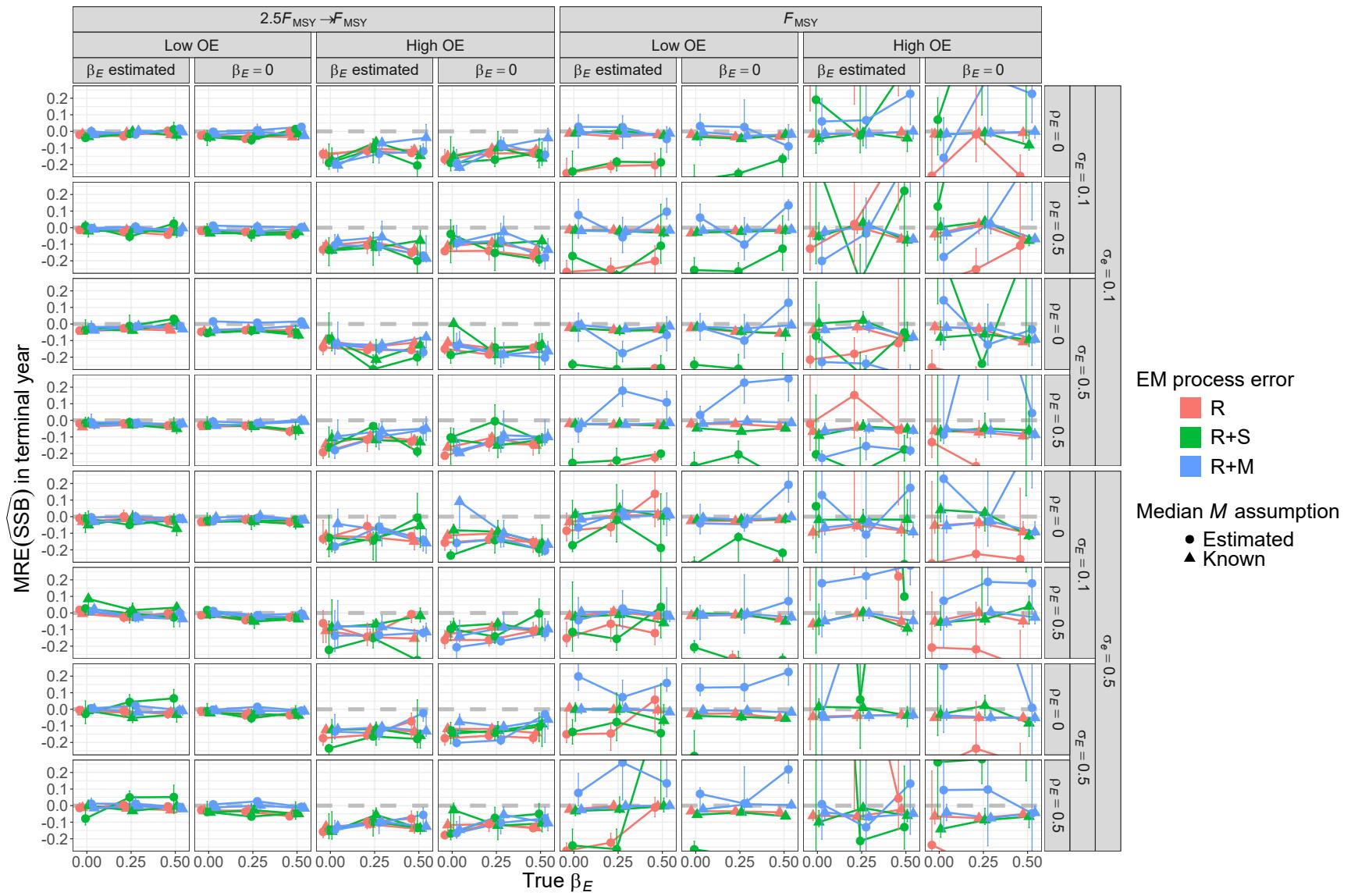


Fig. S49. For R+M OMs, median relative error (MRE) of estimates of spawning stock biomass (SSB) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

606 Terminal year spawning stock biomass RMSE

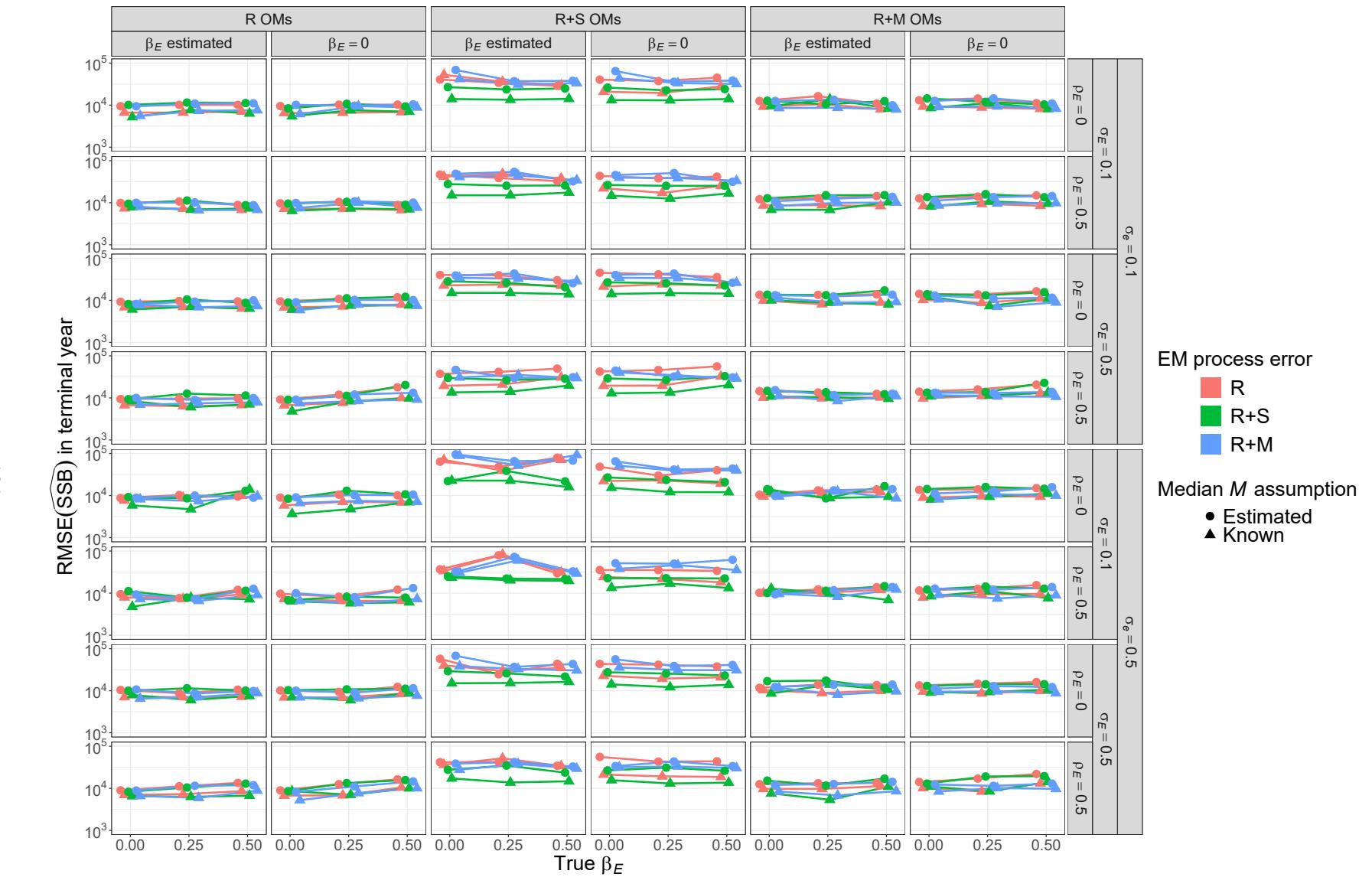


Fig. S50. Root mean square error (RMSE) of estimates of spawning stock biomass in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known). All OMs had low population observation error and contrast in fishing mortality.

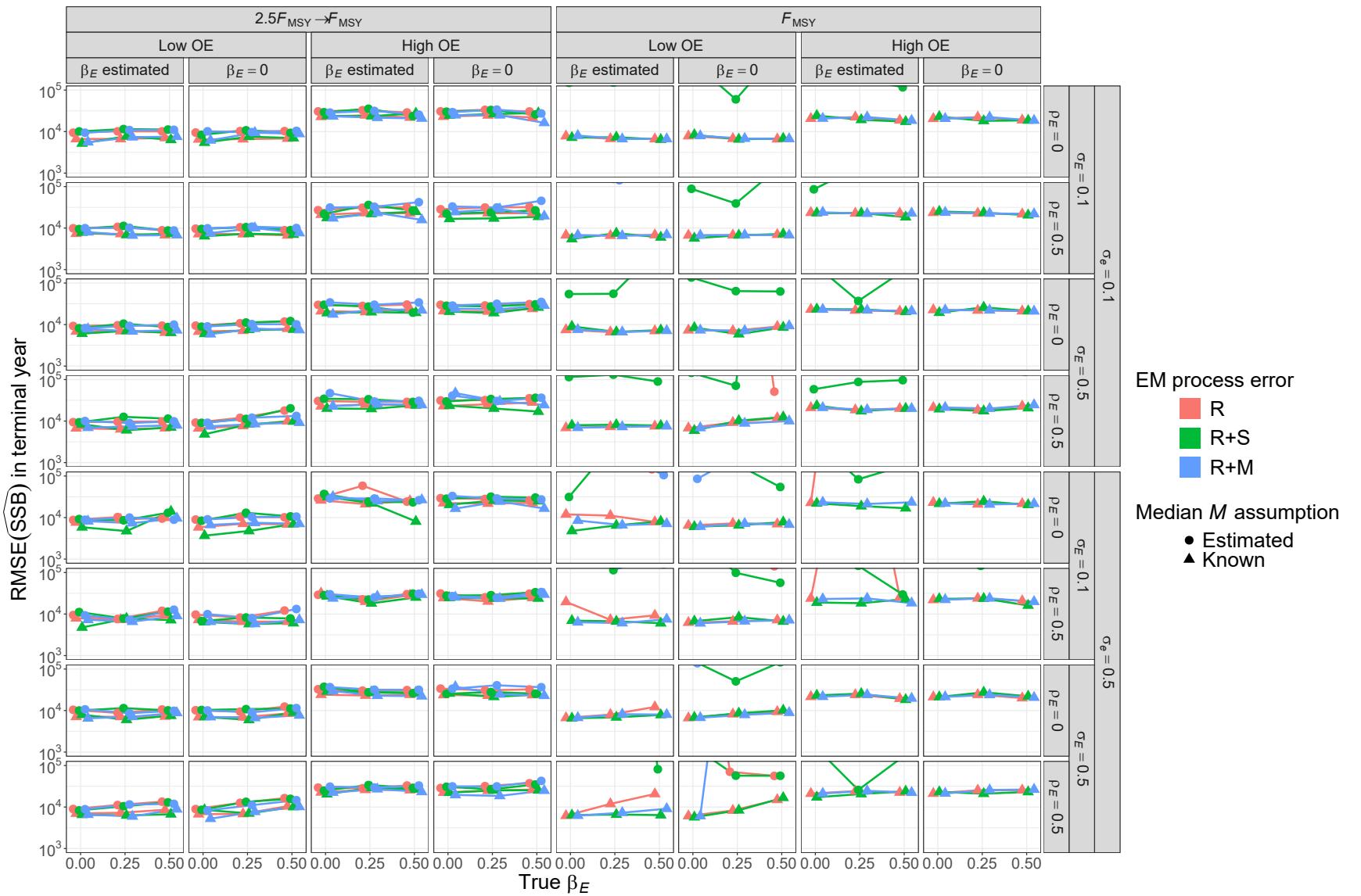


Fig. S51. For R OMs, root mean square error (RMSE) of estimates of spawning stock biomass in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

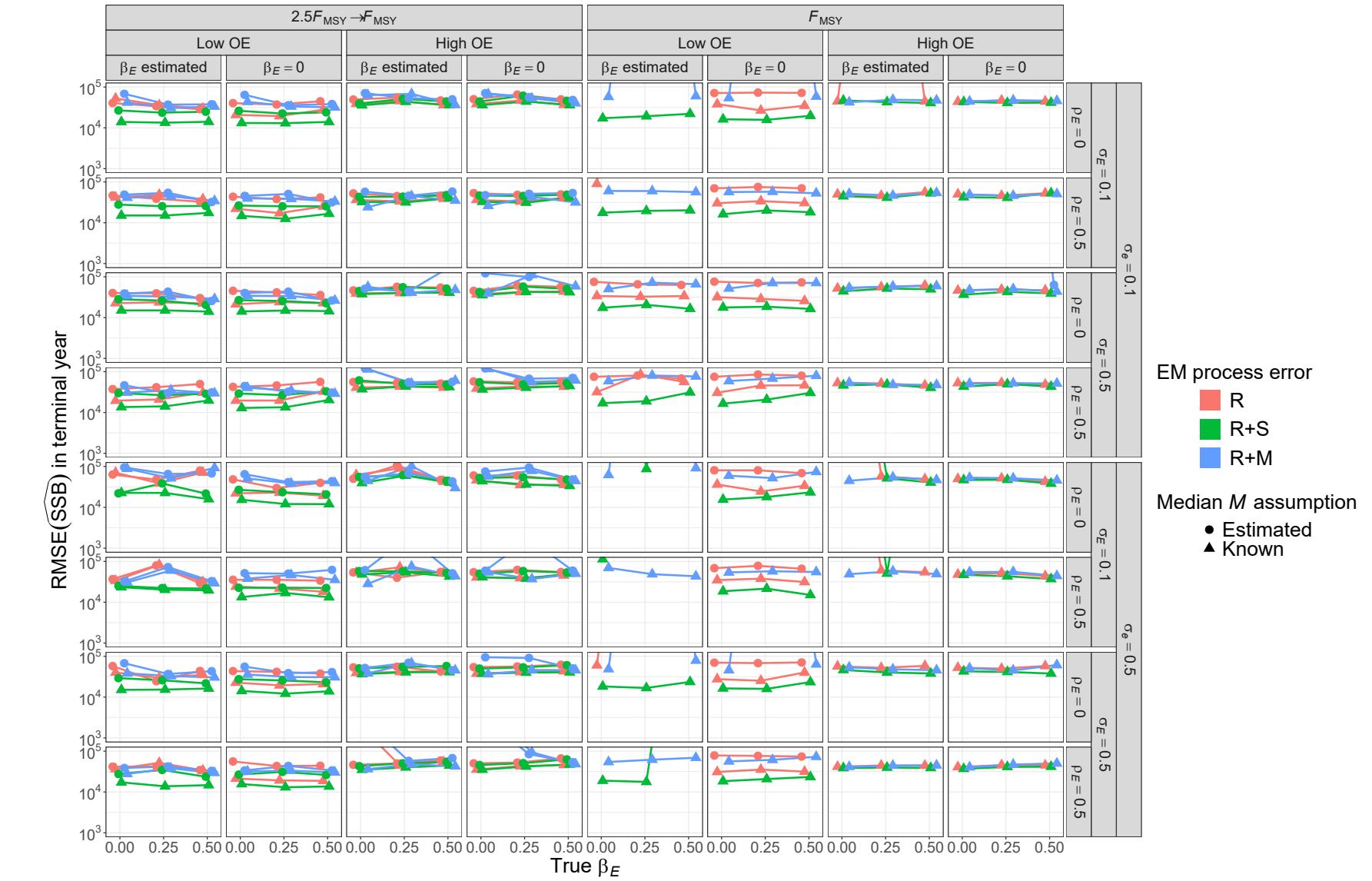


Fig. S52. For R+S OMs, root mean square error (RMSE) of estimates of spawning stock biomass in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

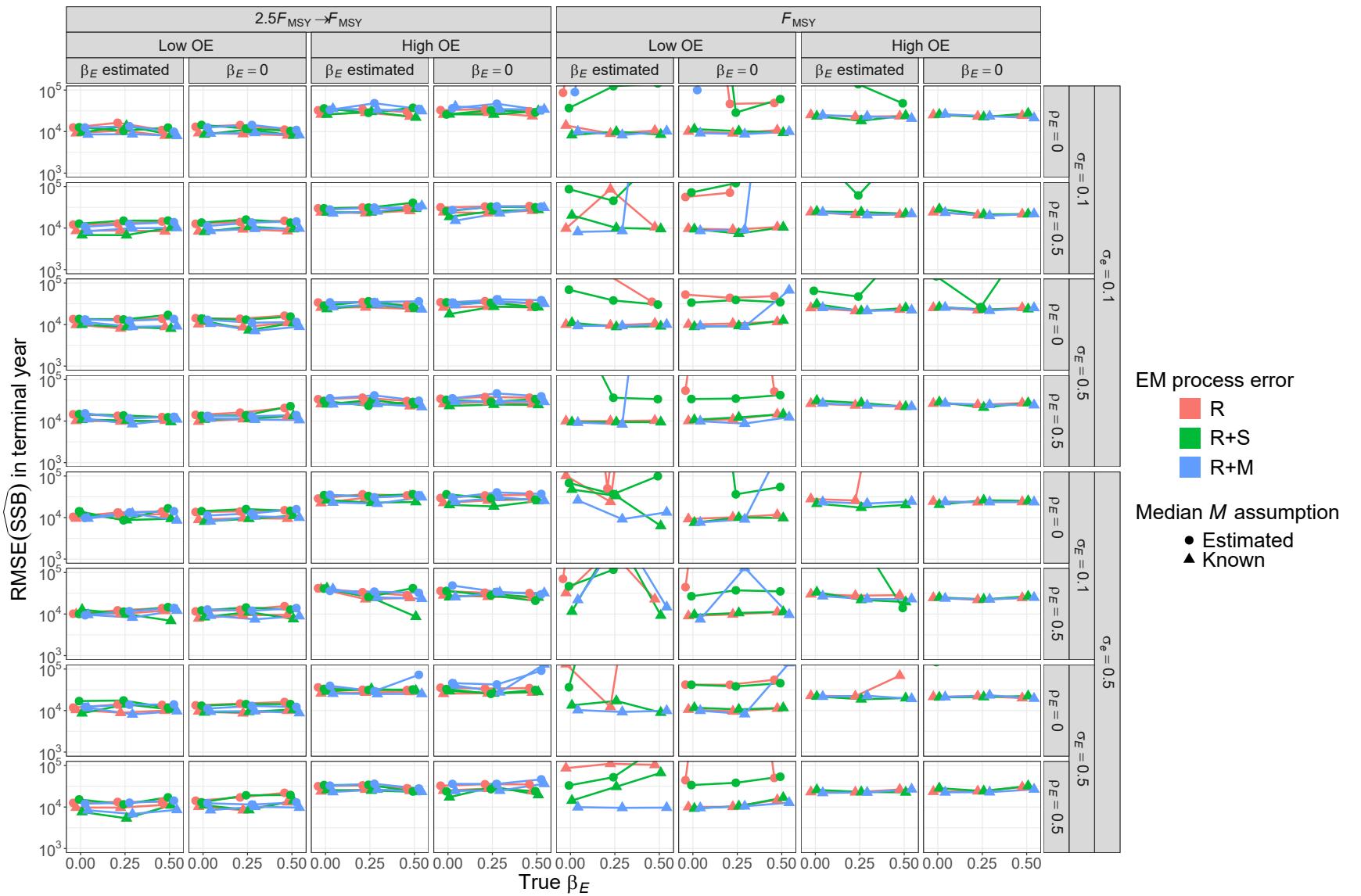


Fig. S53. For R+M OMs, root mean square error (RMSE) of estimates of spawning stock biomass in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

⁶⁰⁷ Terminal year fishing mortality bias

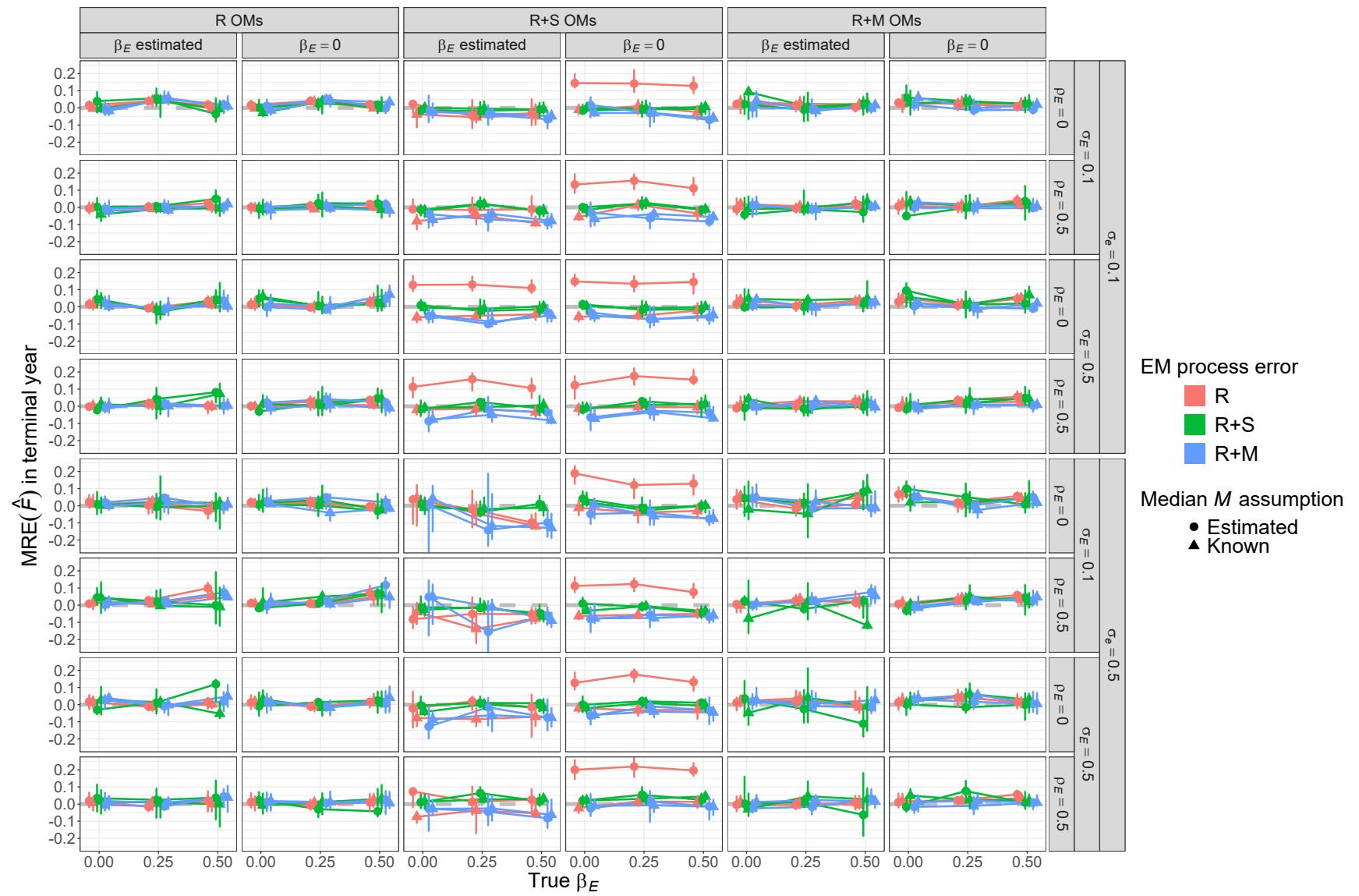


Fig. S54. Median relative error (MRE) of estimates of fully-selected fishing mortality (F) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known). All OMs had low population observation error and contrast in fishing mortality.

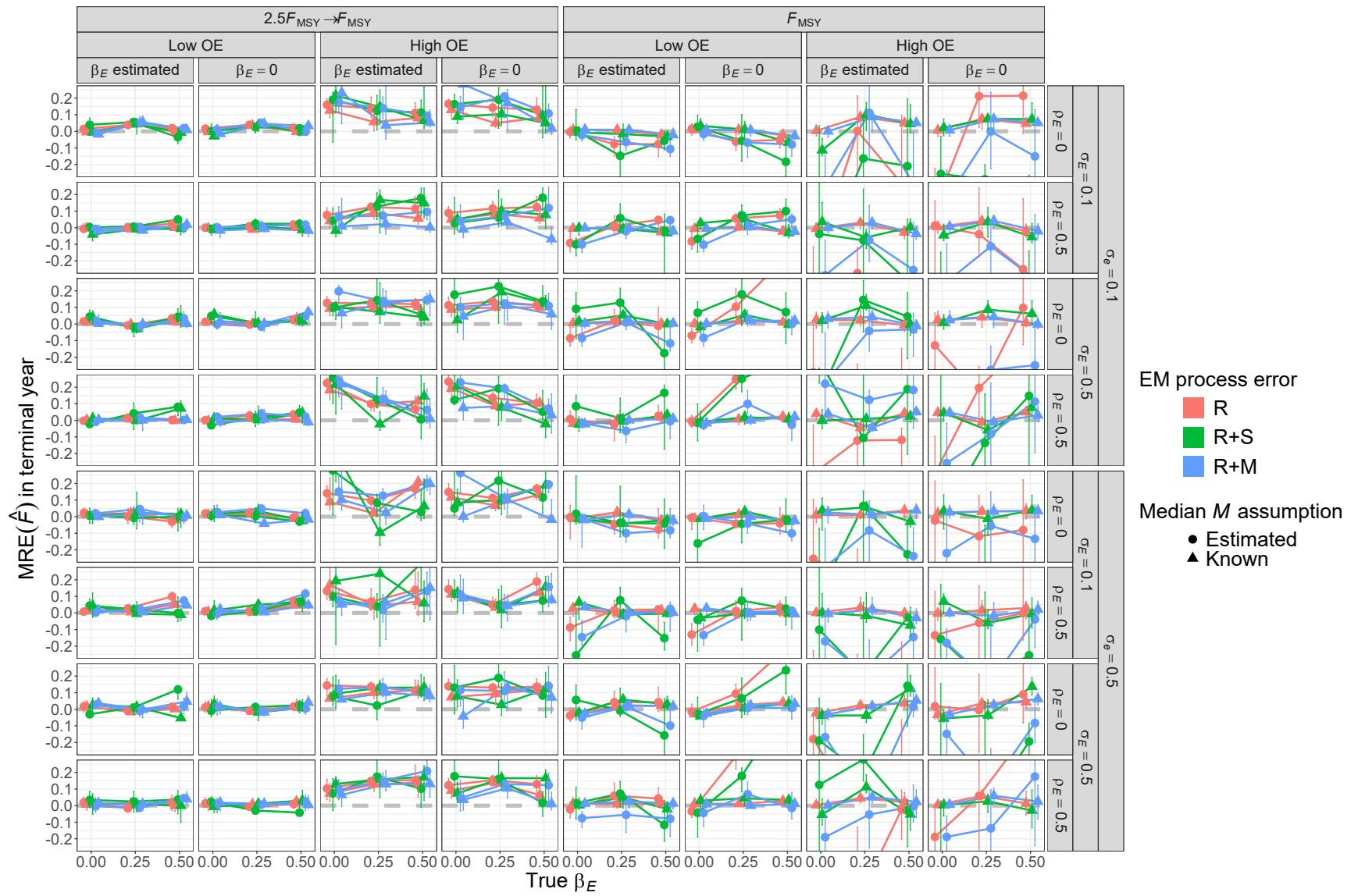


Fig. S55. For R OMs, median relative error (MRE) of estimates of fully-selected fishing mortality (F) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known). All OMs had low observation error and contrast in fishing mortality.

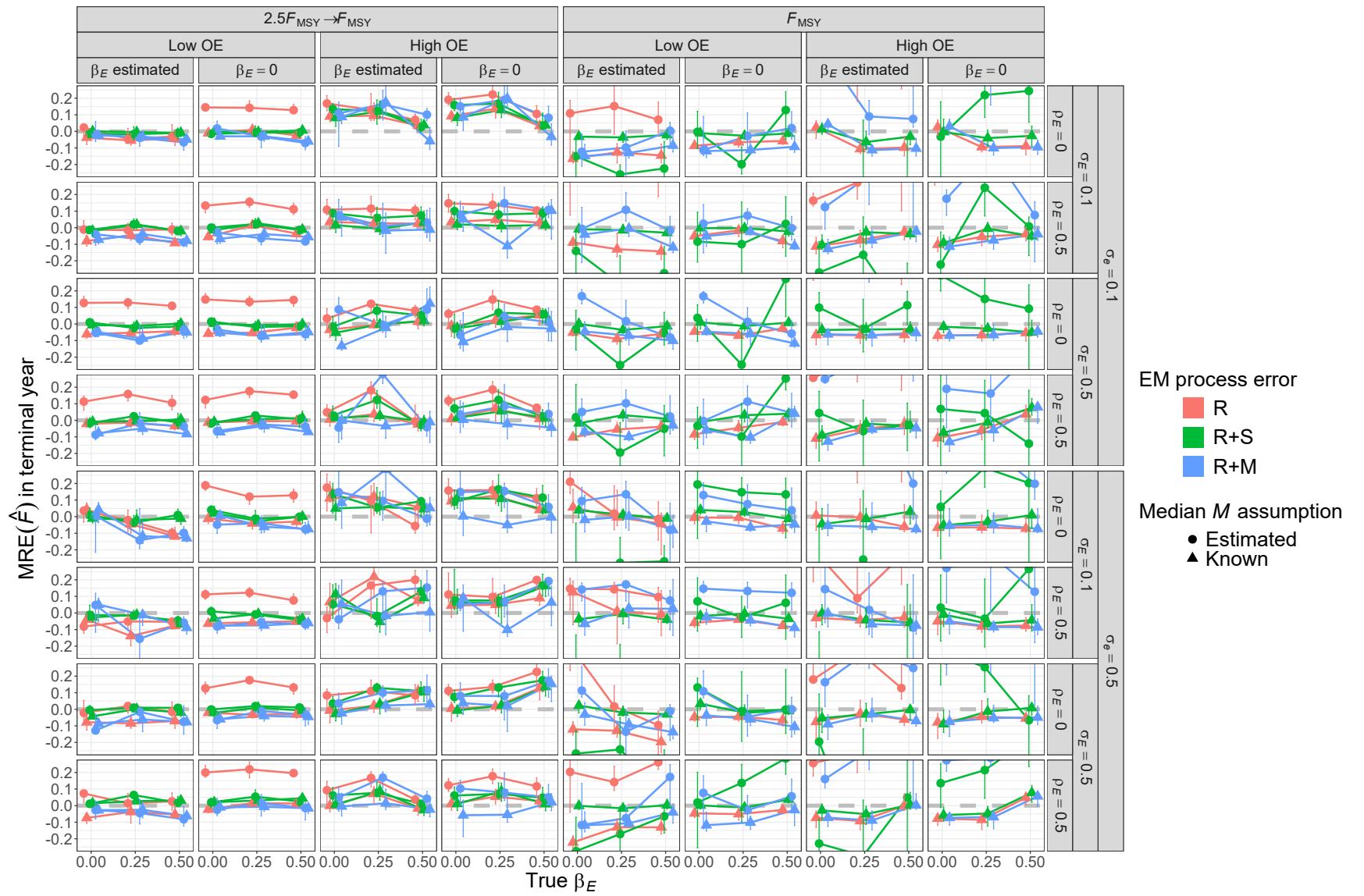


Fig. S56. For R+S OMs, median relative error (MRE) of estimates of fully-selected fishing mortality (F) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known). All OMs had low observation error and contrast in fishing mortality.

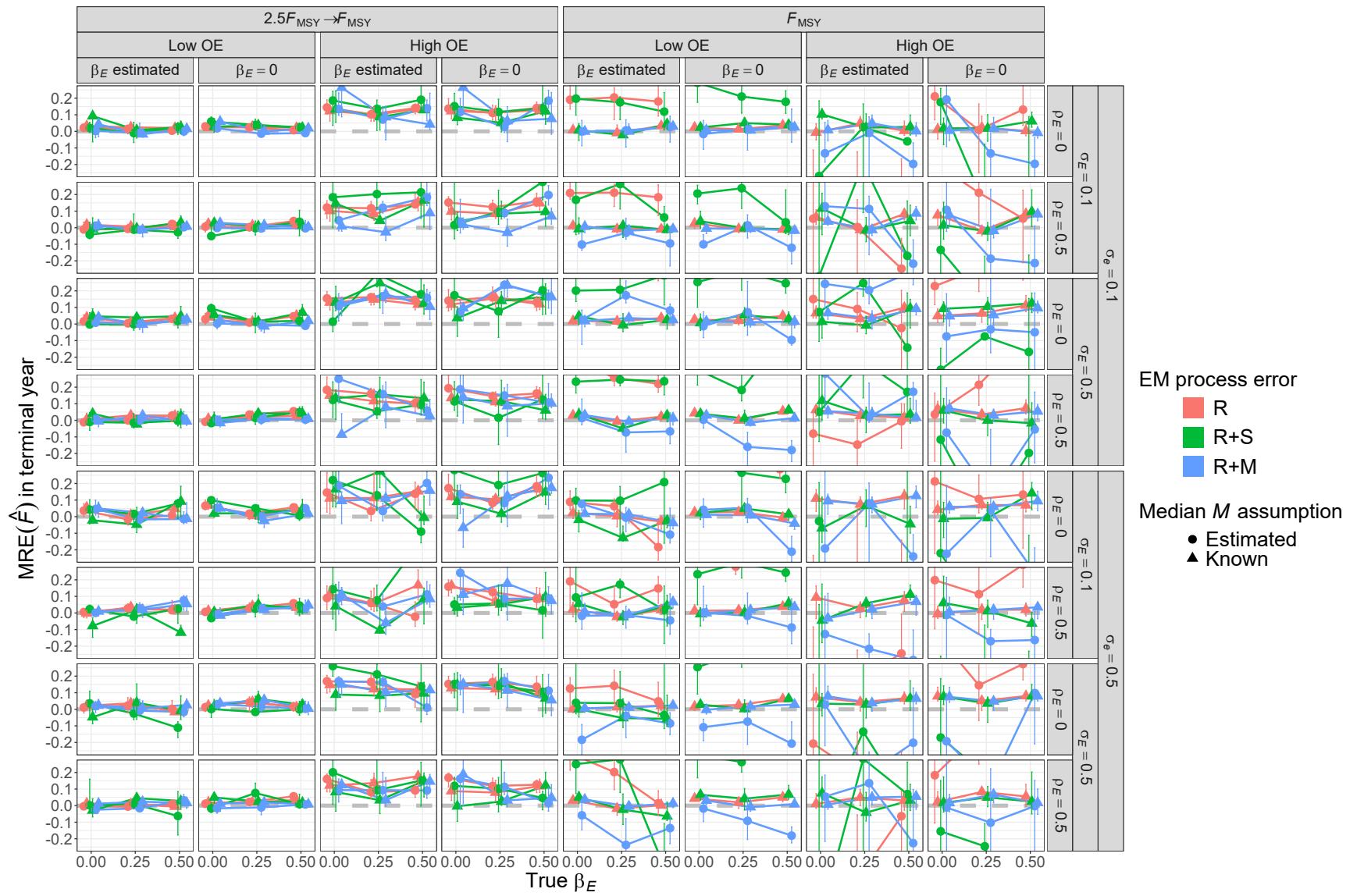


Fig. S57. For R+M OMs, median relative error (MRE) of estimates of fully-selected fishing mortality (F) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known). All OMs had low observation error and contrast in fishing mortality.

⁶⁰⁸ Terminal year fishing mortality RMSE

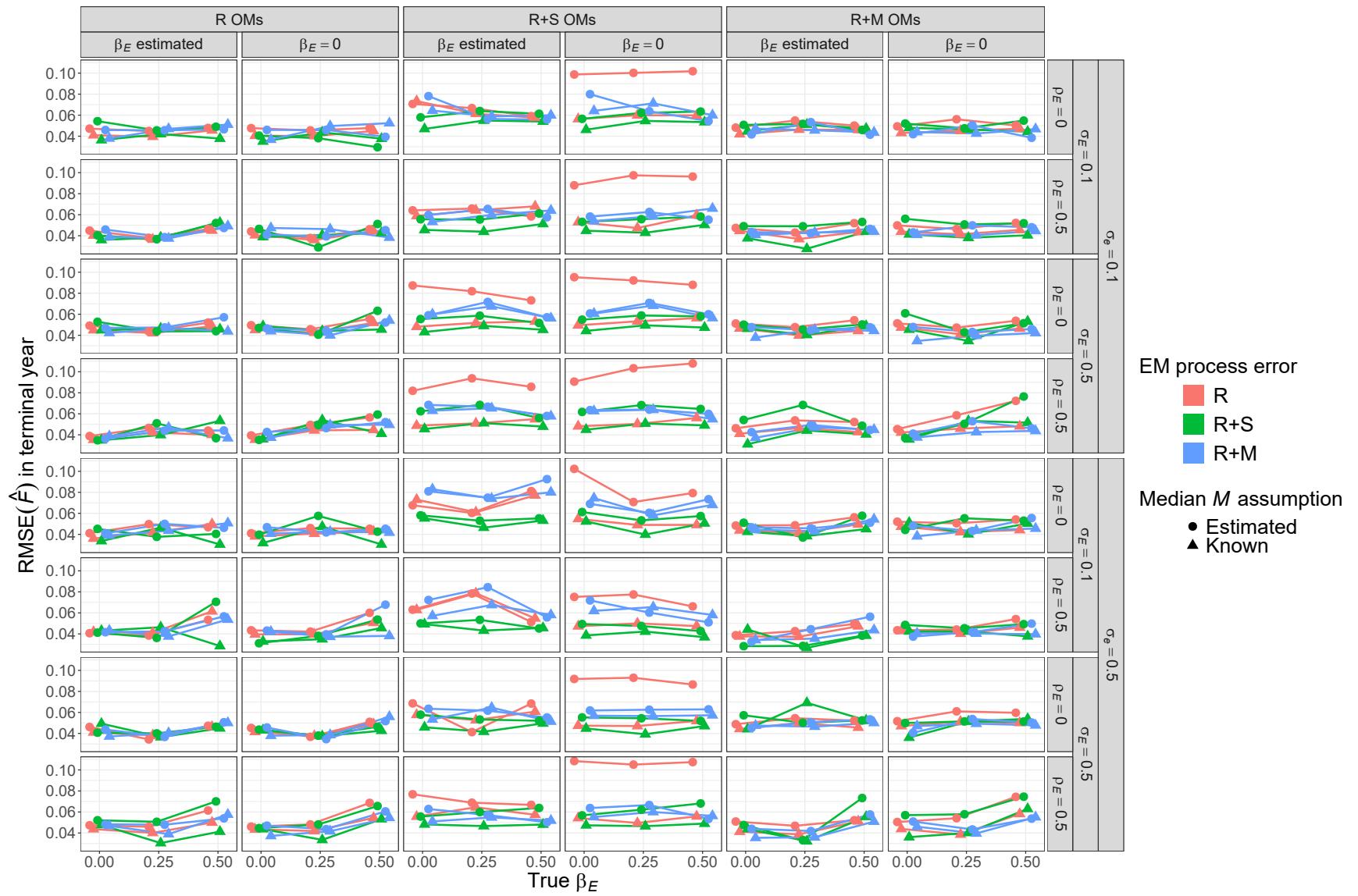


Fig. S58. Root mean square error (RMSE) of estimates of fully-selected fishing mortality (F) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known). All OMs had low population observation error and contrast in fishing mortality.

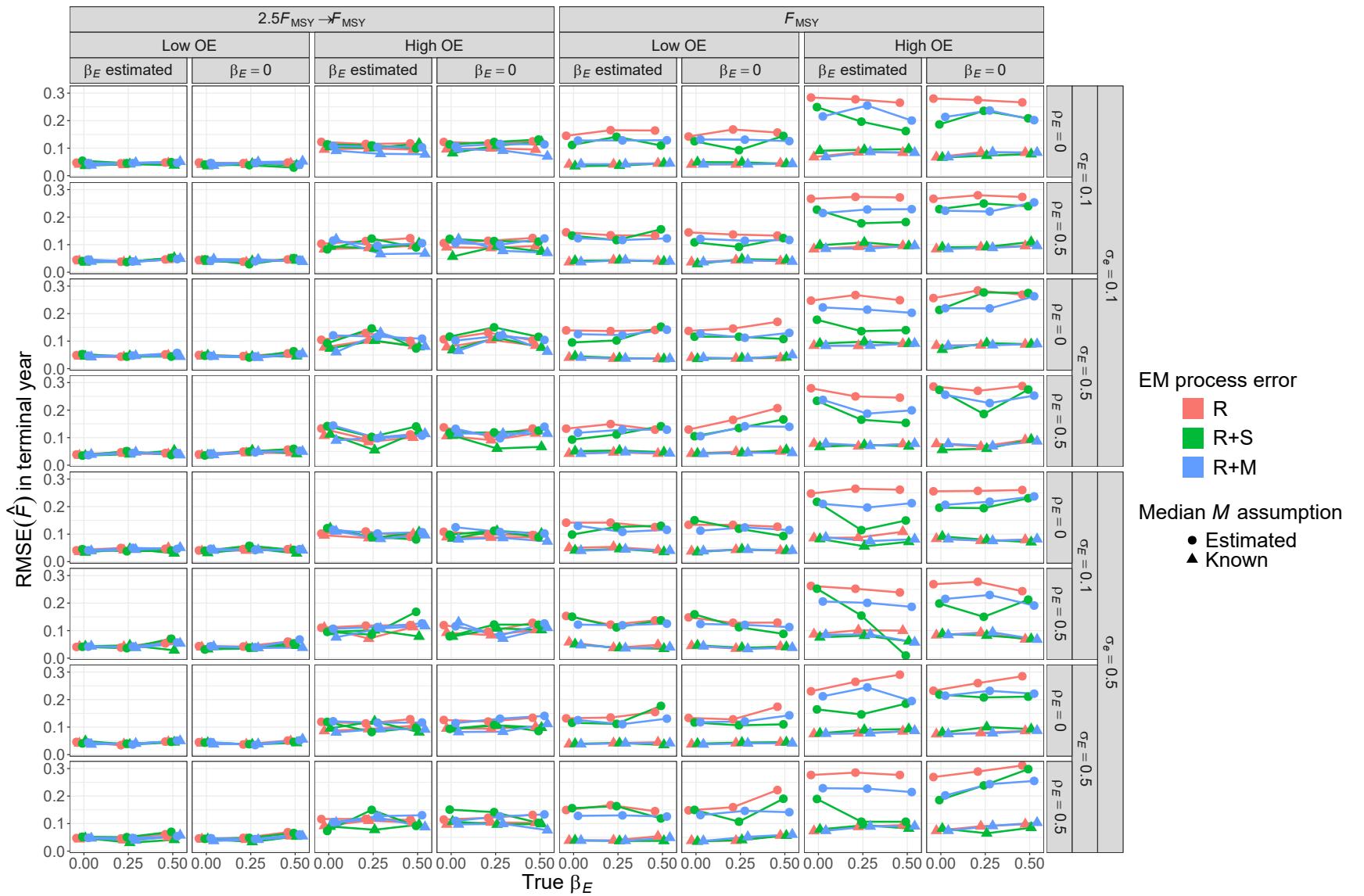


Fig. S59. For R OMs, root mean square error (RMSE) of estimates of fully-selected fishing mortality (F) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

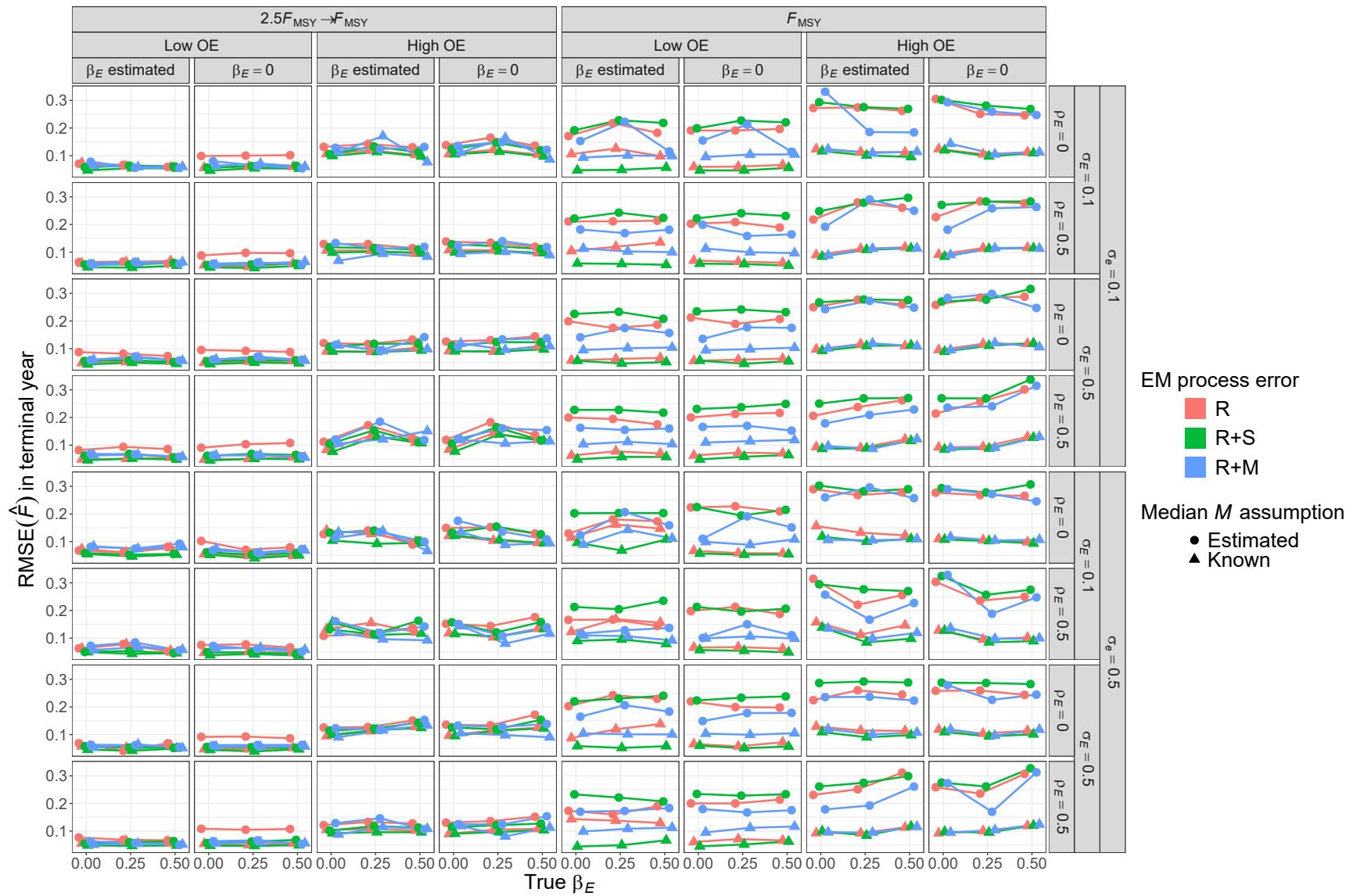


Fig. S60. For R+S OMs, root mean square error (RMSE) of estimates of fully-selected fishing mortality (F) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).

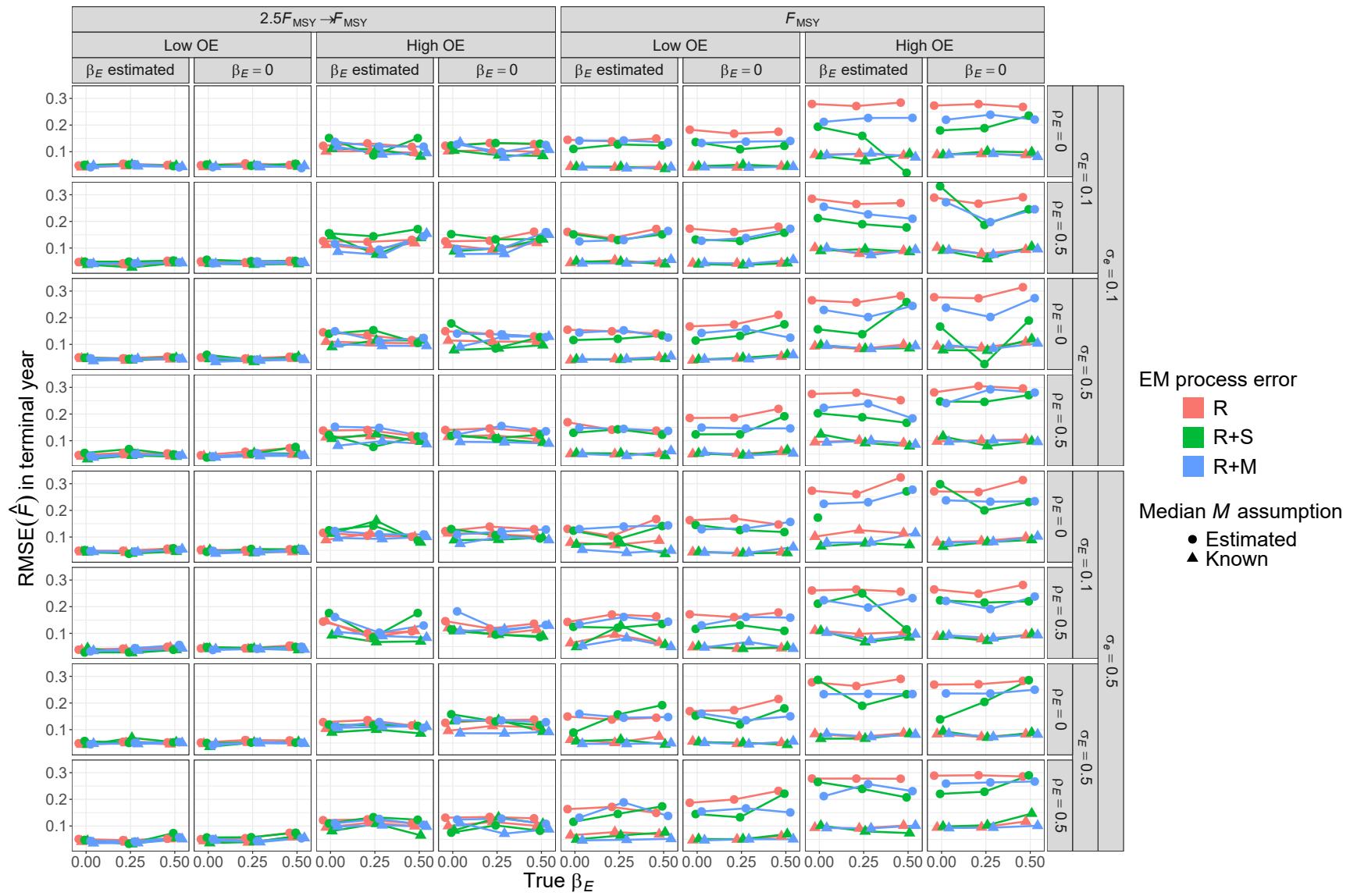


Fig. S61. For R+M OMs, root mean square error (RMSE) of estimates of fully-selected fishing mortality (\hat{F}) in the terminal year for EMs with alternative process error assumptions, treatment of covariate effect ($\beta_E = 0$ or estimated), and treatment of median natural mortality parameter (β_M estimated or known).