

# TOPICS IN ECONOMETRICS AND STATISTICS

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# INTRODUCTION

INTRODUCTION

**MOTIVATION**

## Superconsistent estimation of points of impact in non-parametric regression with functional predictors

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**Summary.** Predicting scalar outcomes by using functional predictors is a classical problem in functional data analysis. In many applications, however, only specific locations or time points of the functional predictors have an influence on the outcome. Such 'points of impact' are typically unknown and must be estimated in addition to estimating the usual model components. We show that our points-of-impact estimator enjoys a superconsistent rate of convergence and does not require knowledge or pre-estimates of the unknown model components. This remarkable result facilitates the subsequent estimation of the remaining model components as shown in the theoretical part, where we consider the case of non-parametric models and the practically relevant case of generalized linear models. The finite sample properties of our estimators are assessed by means of a simulation study. Our methodology is motivated by data from a psychological experiment in which the participants were asked to rate their emotional state continuously while watching an affective video eliciting a varying intensity of emotional reactions.

**Keywords:** Emotional stimuli; Functional data analysis; Non-parametric regression; On-line video rating; Quasi-maximum-likelihood; Variable selection

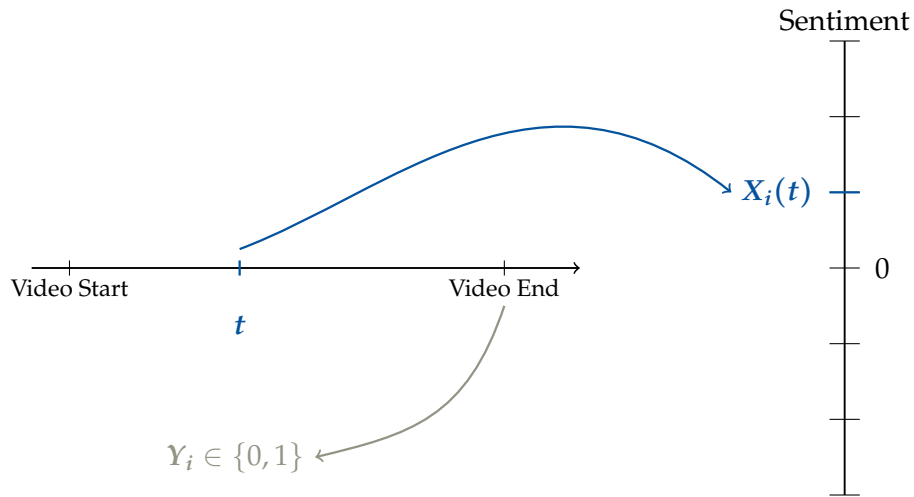
# THE SETTING

LIEBL ET AL. (2020)



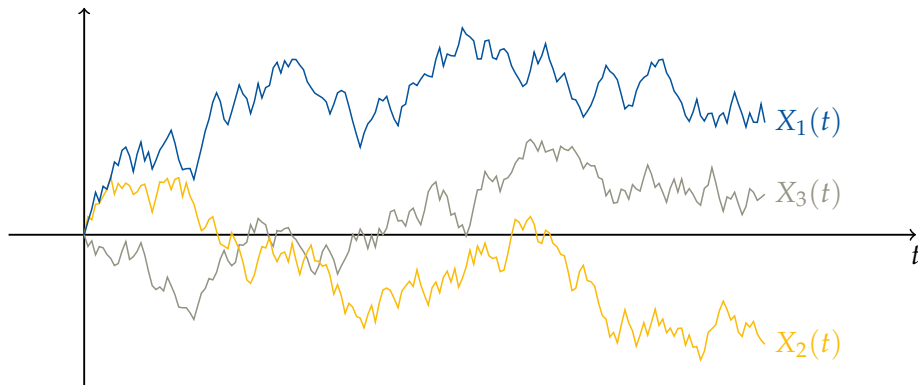
Figure 1: Snapshot of YouTube video

# THE SETTING



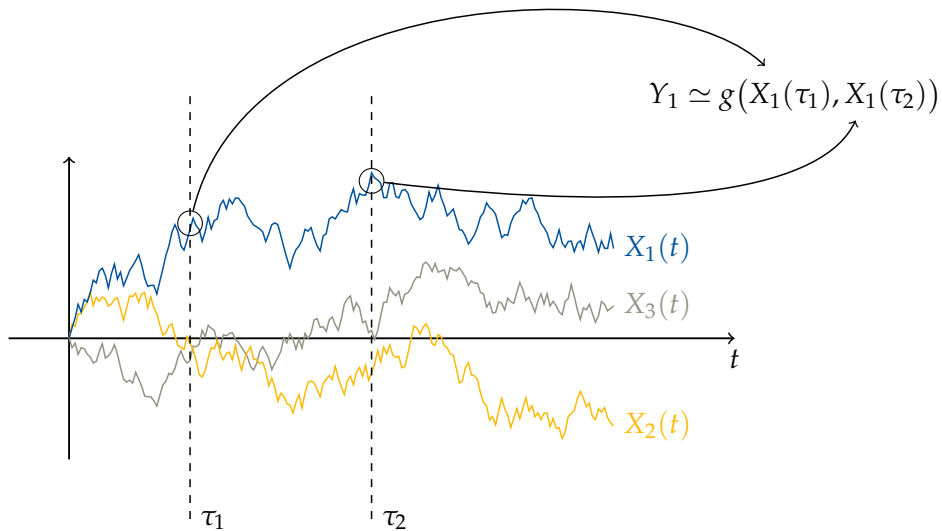
# THE DATA

## REGRESSORS





# THE PROBLEM



INTRODUCTION

**THE MODEL**

## THE MODEL

$$Y_i = g(X_i(\tau_1), \dots, X_i(\tau_S)) + \epsilon_i$$

- $X_i = \{X_i(t) : t \in [0, 1]\}$  square integrable process
- $\mathbb{E}[\epsilon_i \mid X_i(t)] = 0$  exogeneity
- $g : \mathbb{R}^S \rightarrow \mathbb{R}$  *unknown* link function
- $\tau_1, \dots, \tau_S \in [0, 1]$  *unknown* points of impact

# APPROACH

1. stage   Estimate  $S$  and  $\tau_1, \dots, \tau_S$
2. stage   Estimate link function  $g$

*Focus here:* 1. stage

# HEURISTICS

# PLAN

- Simple linear model  $Y_i = \beta_0 + \sum_{s=1}^S \beta_s X_i(\tau_s) + \epsilon_i$
- Gaussian Process Regressors  $X_i \sim \mathcal{GP}(0, \sigma)$

**Task:** Think of how to identify/estimate  $\tau_s$  for different  $\sigma$

**Kernels:** White, Brownian Motion, Matern, RBF

**First insight:** Investigate object  $\mathbb{E} [Y_i X_i(t)]$

# BUILDING INTUITION

What is the connection between the kernel and a sample trajectory?

$\implies$  Visualization

## GENERAL REPRESENTATION

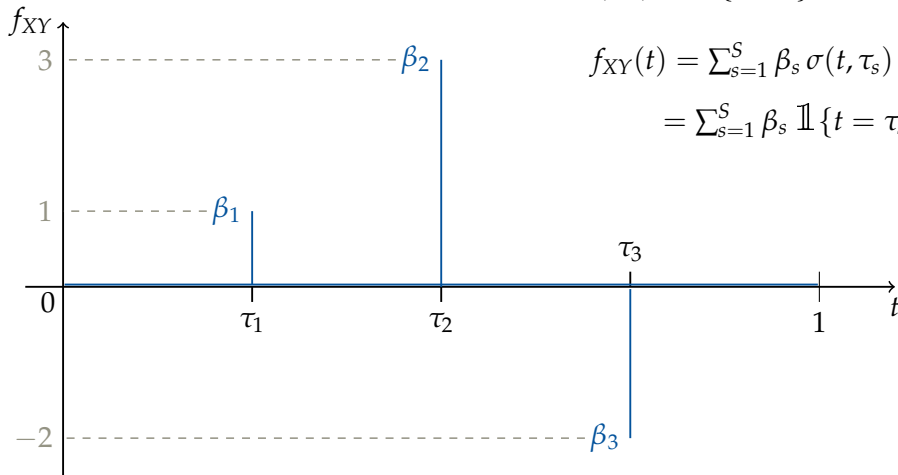
$$\begin{aligned} f_{XY}(t) &\stackrel{def}{=} \mathbb{E} [Y_i X_i(t)] \\ &= \mathbb{E} \left[ \left( \beta_0 + \sum_{s=1}^S \beta_s X_i(\tau_s) + \epsilon_i \right) X_i(t) \right] \\ &= \sum_{s=1}^S \beta_s \mathbb{E} [X_i(\tau_s) X_i(t)] \\ &= \sum_{s=1}^S \beta_s \sigma(t, \tau_s) \end{aligned}$$



## WHITE KERNEL

$$\sigma(s, t) = \mathbb{1}\{s = t\}$$

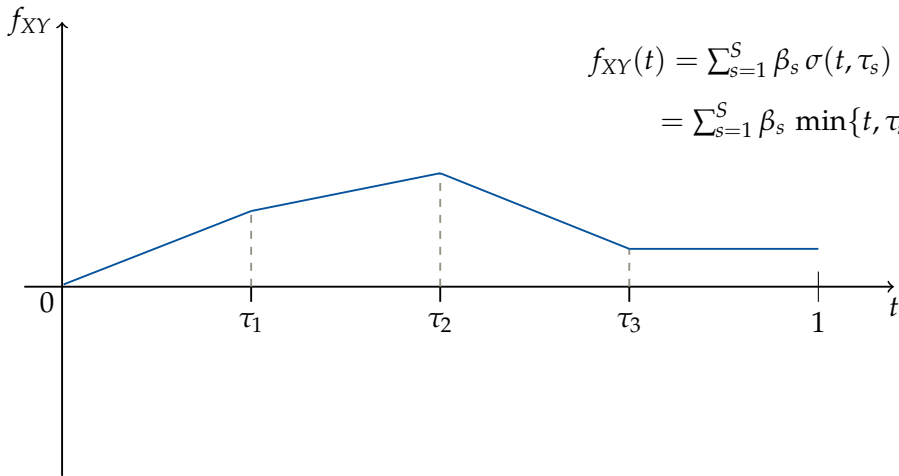
$$\begin{aligned} f_{XY}(t) &= \sum_{s=1}^S \beta_s \sigma(t, \tau_s) \\ &= \sum_{s=1}^S \beta_s \mathbb{1}\{t = \tau_s\} \end{aligned}$$

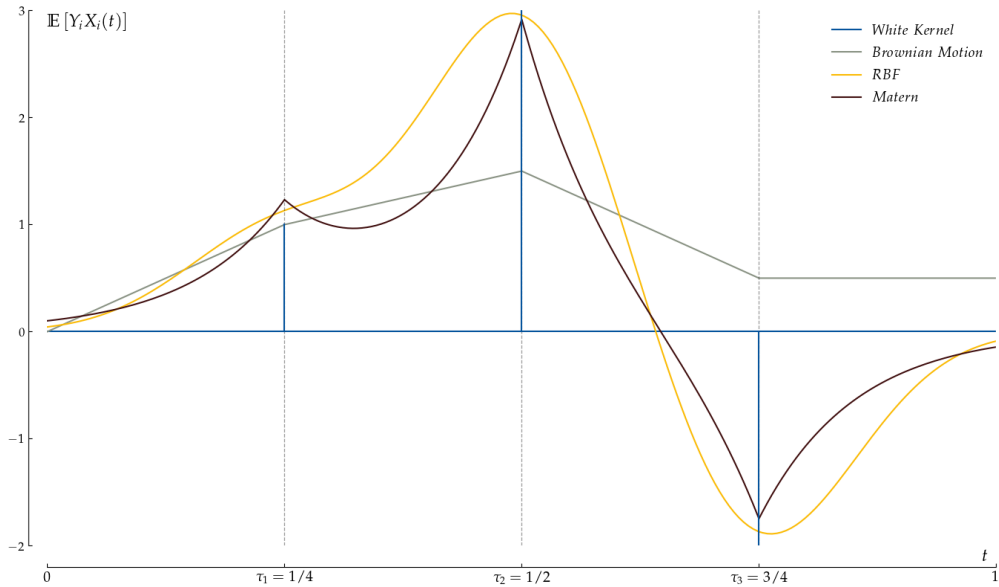


# BROWNIAN MOTION

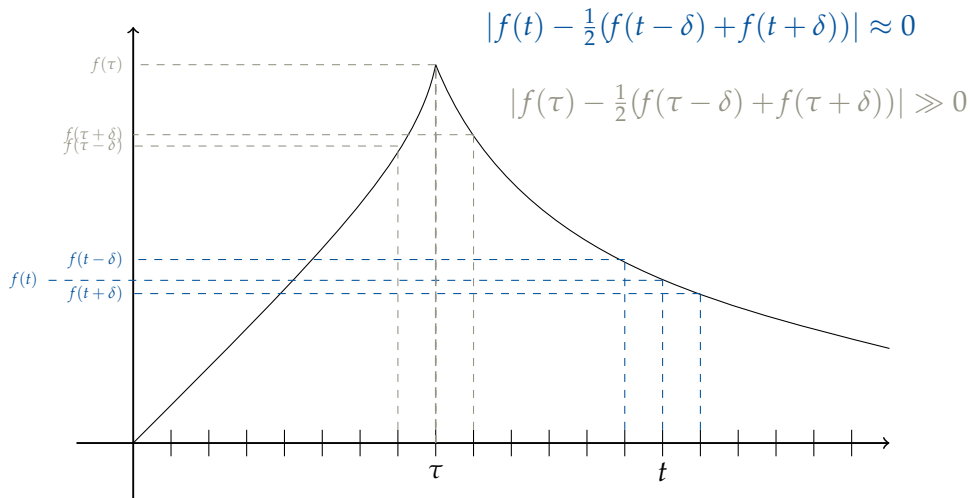
$$\sigma(s, t) = \min\{s, t\}$$

$$\begin{aligned} f_{XY}(t) &= \sum_{s=1}^S \beta_s \sigma(t, \tau_s) \\ &= \sum_{s=1}^S \beta_s \min\{t, \tau_s\} \end{aligned}$$



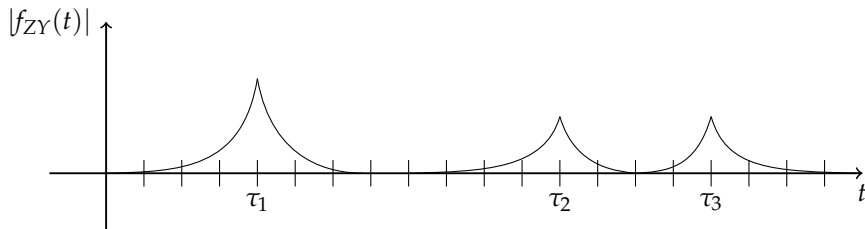


## HOW DO WE FIND $\tau$ ?



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$$f_{ZY}(t) \stackrel{\text{def}}{=} f_{XY}(t) - \frac{1}{2}(f_{XY}(t - \delta) + f_{XY}(t + \delta))$$



## ESTIMATION?

- $Z_{\delta,i}(t) := X_i(t) - (X_i(t + \delta) + X_i(t - \delta)) / 2$
- $\implies \mathbb{E} [Z_{\delta,i}(t) Y_i] = f_{XY}(t) - \frac{1}{2} (f_{XY}(t + \delta) + f_{XY}(t - \delta)) = f_{ZY}(t)$
- Estimate by:  $\hat{f}_{ZY}(t) = n^{-1} \sum_{i=1}^n Z_{\delta,i}(t) Y_i$
- But what about  $\delta$ ?  $\implies$  Visualization

# IDENTIFICATION AND ESTIMATION

IDENTIFICATION AND ESTIMATION

**IDENTIFICATION**



## THEOREM 1

Let  $X_i$  be a Gaussian process and  $g : \mathbb{R}^S \rightarrow \mathbb{R}$  be an arbitrary function with continuous partial derivatives almost everywhere such that

$$0 < \left| \mathbb{E} \left[ \frac{\partial}{\partial x_s} g(X_i(\tau_1), \dots, X_i(\tau_S)) \right] \right| < \infty.$$

Define

$$\vartheta_s = \mathbb{E} \left[ \frac{\partial}{\partial x_s} g(X_i(\tau_1), \dots, X_i(\tau_S)) \right]$$

then,

$$f_{XY}(t) = \mathbb{E} [X_i(t)Y_i] = \sum_{s=1}^S \vartheta_s \sigma(t, \tau_s)$$

## ASSUMPTION 1

Given the kernel  $\sigma$ , there exists open  $\Omega \supset [0, 1]^3$  and twice continuously differentiable function  $\omega : \Omega \rightarrow \mathbb{R}$  as well as some  $\kappa \in (0, 2)$  such that  $\forall s, t \in [0, 1]$

$$\sigma(s, t) = \omega(s, t, |s - t|^\kappa).$$

Moreover,  $0 < \inf \{c(t) : t \in [0, 1]\}$ , where

$$c(t) = -\frac{\partial}{\partial z} \omega(t, t, z)|_{z=0}.$$

Can show that, under Assumption 1 and Theorem 1, as  $\delta \rightarrow 0$

$$f_{ZY}(t) = \mathbb{E} [Z_{\delta,i}(t)Y_i] = \begin{cases} \mathcal{O}(\delta^\kappa) & , \text{if } t = \tau_s \text{ for some } \tau_s \in \{\tau_1, \dots, \tau_S\} \\ \mathcal{O}(\delta^2) & , \text{else} \end{cases}$$

$$\frac{1}{n} \sum_{i=1}^n Z_{\delta,i}(t)Y_i - \mathbb{E} [Z_{\delta,i}(t)Y_i] = \mathcal{O}_{\mathbb{P}} \left( \sqrt{\delta^\kappa/n} \right)$$

As seen in visualization, need sensible choice of  $\delta$

- $\delta$  too small (e.g.  $\delta^\kappa \sim n^{-1}$ )  $\implies$  estimation noise dominates
- $\delta$  too big  $\implies$  cannot distinguish between neighboring points

IDENTIFICATION AND ESTIMATION

**ESTIMATION**

## ESTIMATION

- $X_i$  observed over  $p$  equidistant points  $t_j$  in  $[0, 1]$
- Possibly  $p \gg n$
- Estimate  $\{\tau_s\}$  using local maxima of  $|\hat{f}_{ZY}(t)| = |n^{-1} \sum_{i=1}^n Z_{\delta,i}(t) Y_i|$
- Algorithm 1: Given  $\delta > 0$  determine  $\hat{\tau}_1, \dots, \hat{\tau}_{M_\delta}$  with  $M_\delta \in \mathbb{N}$

Finally,

$$\hat{S} = \min \left\{ \ell \in \mathbb{N}_0 : \left| \frac{\sum_{i=1}^n Z_{\delta,i}(\hat{\tau}_{\ell+1}) Y_i}{\{\sum_{i=1}^n Z_{\delta,i}(\hat{\tau}_{\ell+1})^2\}^{1/2}} \right| < \lambda \right\},$$

and we select only the first  $\hat{\tau}_s$ , where  $s = 1, \dots, \hat{S}$ .

## ALGORITHM 1

- 1: compute  $\hat{f}_{XY}(t_j) = \sum_i X_i(t_j)Y_i/n$ , for all  $j = 1, \dots, p$
- 2: choose  $\delta > 0$  s.t.  $\exists k_\delta \in \mathbb{N}$  with  $1 \leq k_\delta < (p-1)/2$  and  $\delta = k_\delta/(p-1)$
- 3: define  $\mathcal{J}_\delta = \{k_\delta + 1, \dots, p - k_\delta\}$  and set  $\ell = 1$
- 4: compute  $\hat{f}_{ZY}(t_j) = \hat{f}_{XY}(t_j) - (\hat{f}_{XY}(t_j + \delta) + \hat{f}_{XY}(t_j - \delta))/2$ , for all  $j \in \mathcal{J}_\delta$
- 5: **while**  $|\mathcal{J}_\delta| \neq \emptyset$  **do**
- 6:     estimate  $\hat{\tau}_\ell = \operatorname{argmax} \left\{ |\hat{f}_{ZY}(t_j)| : \text{for } t_j \text{ with } j \in \mathcal{J}_\delta \right\}$
- 7:     update  $\mathcal{J}_\delta \leftarrow \mathcal{J}_\delta \setminus [\hat{\tau}_\ell - \sqrt{\delta}, \hat{\tau}_\ell + \sqrt{\delta}]$
- 8:     update  $\ell \leftarrow \ell + 1$
- 9: **return**  $\{\hat{\tau}_\ell\}$

# ASYMPTOTICS

Can we say something about the convergence rate of  $\hat{\tau}_s$ ?

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Assumption 1 + 2 and Theorem 1 + 2  $\implies$  superconsistent rates



## ASSUMPTION 2

(a)  $X_1, \dots, X_n \stackrel{iid}{\sim} X$ , where  $X$  is a Gaussian process

(b)  $\exists 0 < \sigma_{|y|} < \infty$  s.t.  $\forall m \geq 1 : \mathbb{E} [|Y_i|^{2m}] \leq 2^{m-1} m! (\sigma_{|y|})^{2m}$

Condition (b) is fulfilled, for example, if  $Y_i$  is bounded, as in the logistic regression case, or if (a) holds and the errors  $\epsilon_i$  are sub-Gaussian and  $g$  has bounded partial derivatives.

## THEOREM 2

(UNDER ASSUMPTION 1, 2 AND THEOREM 1)

If  $\delta = \delta_n \rightarrow 0$  as  $n \rightarrow \infty$  with  $n\delta^\kappa / |\log \delta| \rightarrow \infty$  and  $\delta^\kappa / n^{-\kappa+1} \rightarrow 0$ , then

(i)  $\max_{\ell=1,\dots,\hat{S}} \min_{s=1,\dots,S} |\hat{\tau}_\ell - \tau_s| = \mathcal{O}_{\mathbb{P}}(n^{-1/\kappa})$

(ii)  $\exists D \in (0, \infty)$  such that when algorithm 1 is applied with threshold

$$\lambda = \lambda_n = A \sqrt{\sigma_{|y|}^2 / n \log(1/\delta)}, \text{ with } A > D, \text{ and } \delta^2 = \mathcal{O}(n^{-1}), \text{ then,}$$

$$\mathbb{P} [\hat{S} = S] \rightarrow 1 \text{ as } n \rightarrow \infty.$$

## GENERALIZATIONS AND EXTENSIONS

# GENERALIZATIONS

- Non-Gaussian processes
  - Build on framework of elliptical processes (heavy-tailed, logistic, t)

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- Non-Gaussian processes
  - Build on framework of elliptical processes (heavy-tailed, logistic, t)
- Generalizing covariance assumption
  - Do not need rough process with non-diffb. kernel at diagonal, rather need the kernel to be fewer times differentiable at the diagonal than off the diagonal

# EXTENSIONS

- Non-IID data (time-series)

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- Non-IID data (time-series)
- Higher-dim. index set, e.g.  $[0, 1]^2$  instead of  $[0, 1]$
- Sparsely sampled data



# REFERENCES

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# CONTACT

*Slides and codes are hosted on GitHub.  
For further questions contact me via email.*



`github.com/timmens/topics-metrics-2021`



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