

# Statistical Thinking in Biology Research

Terry Neeman and Timothee Bonnet

November 20, 2018

# Acknowledgements and warning

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- Statistics in biology is the study of biological variation

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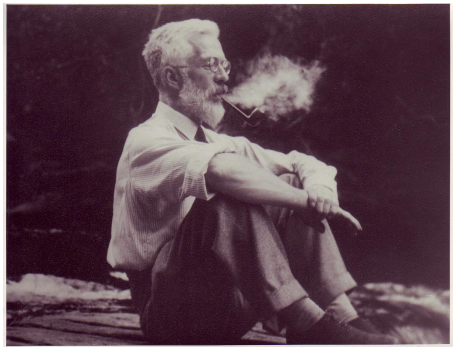
- Statistics in biology is the study of biological variation
- Statistical ideas about biological variation inform the design of experiments
- Statistical ideas about biological variation inform the analysis of experiments

# Key ideas for today

- Statistics in biology is the study of biological variation
- Statistical ideas about biological variation inform the design of experiments
- Statistical ideas about biological variation inform the analysis of experiments
- Statistical thinking is an essential component of scientific thinking

# A bit of history of statistical methods

R.A. Fisher: 1890-1962



Statistical Principles for Research Workers (1925)

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R.A. Fisher: 1890-1962



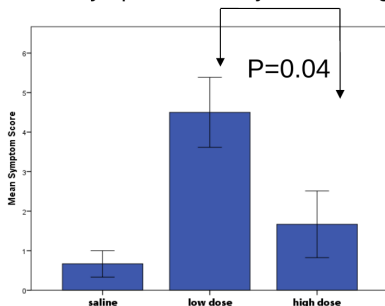
Statistical Principles for Research Workers (1925)



- 1 Cautionary tales from the front
- 2 Introduction to Statistical Modelling
- 3 Another look at essential steps

# Message 1: A small p-value is not always evidence of a treatment effect

Mean symptom score by treatment group



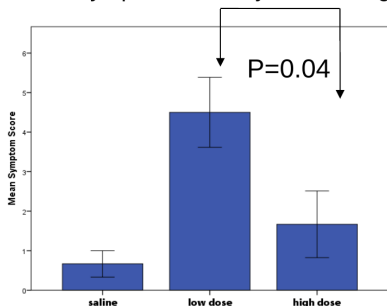
## Vaccine challenge experiment:

- 6 mice/group (saline/low dose/high dose)
- All mice challenged with *Shigella*
- Followed for 14 days
- Outcome: Symptom score average Days 2 - 8

One-way ANOVA (post-hoc Bonferroni)  $p=0.04$

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*Do you think the vaccine works? What is strange?*

# Message 1: A small p-value is not always evidence of a treatment effect

# Message 1: A small p-value is not always evidence of a treatment effect

## Experimental design

The observed difference in outcome could be the result of:

- Cage effects
- Mouse strain effects

These effects are **CONFOUNDED** with treatment effect



Cage 1:  
saline



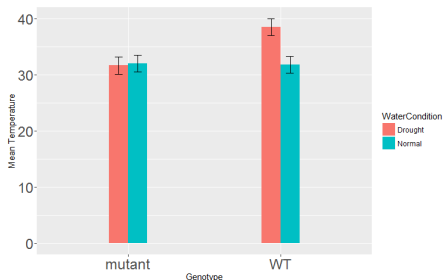
Cage 2:  
Low Dose



Cage 3:  
High Dose

Message 2: p-values from simple comparisons cannot tell us when differences are “different”

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Are temperature mechanisms modified in a genetically modified tomato plant?

- Genotypes: WT/mutant
- Water condition: Normal/Drought
- Leaf temperature measured

### Comparisons made using t-tests

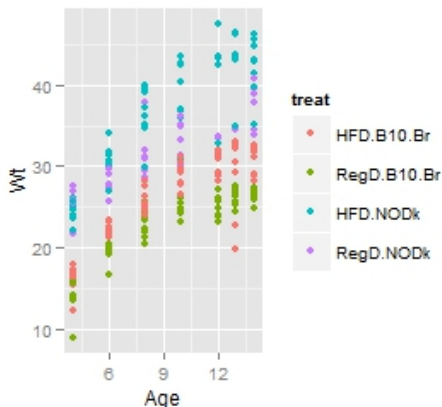
Evidence of difference + No evidence of difference  $\neq$  Evidence that differences are different.

# Message 3: Interpreting experimental results needs more than t-tests



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Research question: Are mice susceptible to obesity when exposed to a high fat diet?



## Experimental set-up:

- 37 mice: 16 NODk /21 WT
- Randomised to either regular or high fat diet
- Monitored for 14 weeks
- Outcome measure: Body weight (g)
- Experimental factors: Diet (2), Strain (2), Time (8)

Acknowledgements: Ainy Hussain, PhD student 2013

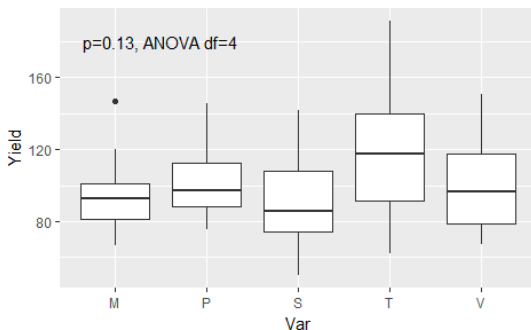
## Message 4: Knowing how to combine information across subgroups can improve inference

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Comparing yield in five barley varieties (1930s)

Experimental factors: 5 varieties of barley, 6 locations, 2 time points.

Outcome measure: yield



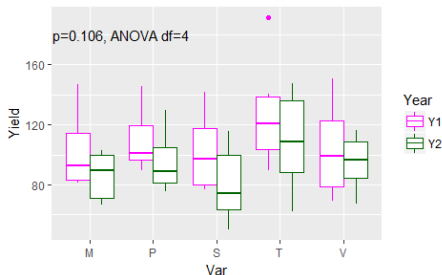
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## Controlling for other sources of variation:

- Controlling for year = comparing yield **WITHIN** years and combining these

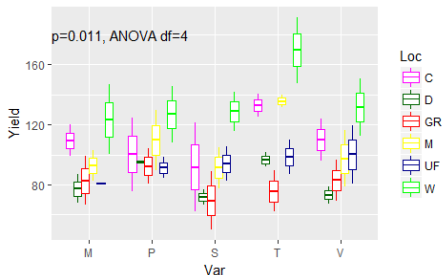
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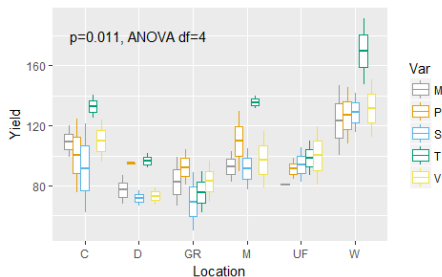


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# Message 5: Knowing what factors contribute to the variation in outcome helps design experiments and analyses

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Research question: How does cold duration impact upon germination in alpine plant *A. glacialis*?



## Experimental set-up:

- Seed collections from alpine region in Australia
- 3 Regions- low/high altitude
- 4 sets of Petri dishes
- 4 cabinet shelves
- Response - % germinated

**What other factors are important to consider when comparing cold duration?**



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- ④ Combining information across subgroups can improve inference. **A statistical model enables accumulation of evidence across experiments.**

# Summary

- 1 A small p-value is not always evidence of a treatment effect. **Good experimental design matters.**
- 2 p-values from simple comparisons cannot tell us when differences are “different”. **For each question / comparison, a specific test**
- 3 Interpreting experimental results needs more than t-tests. **Need a statistical model of the experiment, matching scientific question.**
- 4 Combining information across subgroups can improve inference. **A statistical model enables accumulation of evidence across experiments.**
- 5 Knowing what factors contribute to the variation in outcome matters. **A statistical model allows one to incorporate effect of other factors in the analysis.**

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# Introduction to Statistical Modelling

- What is a statistical model?
- Modelling outcomes:
  - ▶ a summary of data
  - ▶ a prediction model
  - ▶ an explanatory model
- Model – may take many different functional forms
- Model – a conceptualization of the experiment

# Introduction to Statistical Modelling

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ALWAYS BEGIN WITH A RESEARCH QUESTION



# Key components of a statistical model of an experiment

- Outcome measure
  - ▶ Response variable
  - ▶ Measure of interest
- Experimental factors
  - ▶ Conditions that can be manipulated
  - ▶ Conditions of interest (e.g. genotype, gender)
  - ▶ Main questions: do the conditions impact upon the outcome measure?
- Blocking factors
  - ▶ Conditions (not of interest) that may impact upon the outcome measure
  - ▶ Sources of variation in the experiment that need to be controlled for
  - ▶ Clustering of experimental units

ALWAYS BEGIN WITH A RESEARCH QUESTION

# Key Objectives of a statistical model of an experiment

- To compare the mean response of an organism/system to a set of different experimental conditions.
  - ▶ Obtain estimate of “Treatment effect”
  - ▶ Is this “effect” different in subgroups of interest?
- What are the most important factors influencing the mean response?
- Subsidiary question: how can we design our experiment in future to more efficiently test our hypotheses?

# Example 1: Does dark respiration differ between C3 and C4 plants?

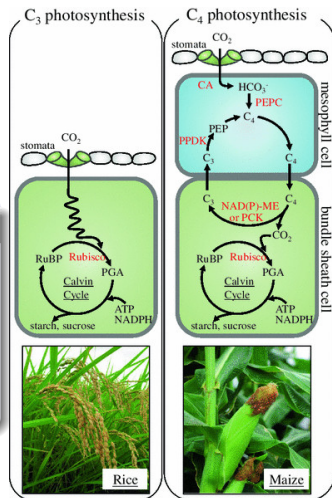
Outcome measure: dark respiration

Experimental factor: Plant type (C4/C3)

Data: 6 plants each of C4, C3

## Can calculate

- Observed overall mean
- Observed mean C3 plants
- Observed mean C4 plants
- Variation around each mean



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## Statistical model

Respiration = Mean for C3 + Difference C4-C3 \* (is C4?) + Noise

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## Can calculate

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## Statistical model

Respiration = Mean for C3 + Difference C4-C3 \* (is C4?) + Noise

*response* = *A* + *D* × *predictor* +  $\epsilon$

*A* and *D* are the model PARAMETERS.

We want to infer whether *D* is different from 0

# Example 1: Does dark respiration differ between C3 and C4 plants?

$$\text{response} = A + D \times \text{predictor} + \epsilon$$

Can we separate the signal  $D$  from the noise  $\epsilon$ ?

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## T-test

- Outcome is a continuous variable
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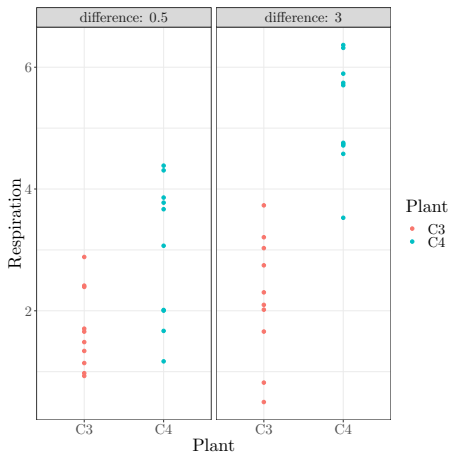
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$$t = \frac{D}{\text{Variation of } \epsilon} \times \frac{\text{Sample Size}}{\sqrt{2}}$$



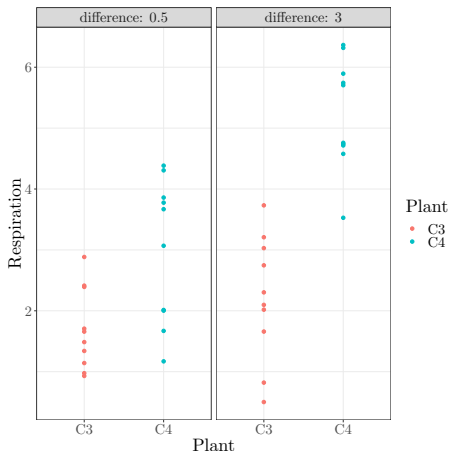
# When can we know whether $D \neq 0$ ?



$$t = \frac{D}{\text{Variation of } \epsilon} \times \frac{\text{Sample Size}}{\sqrt{2}}$$

Is it easier when the true difference is 0.5 or when it is 3 ?

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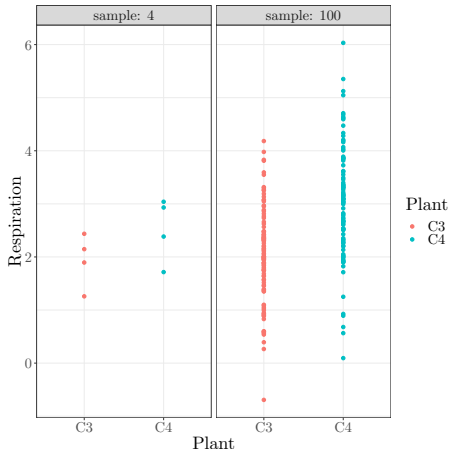


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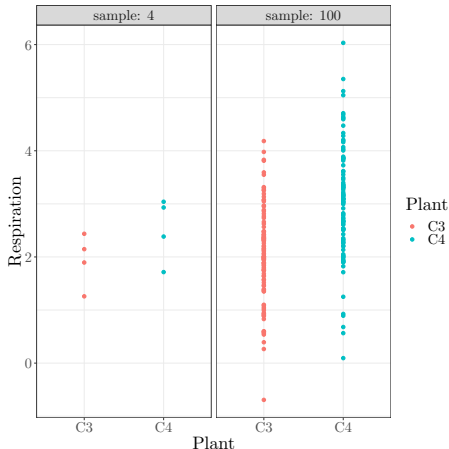
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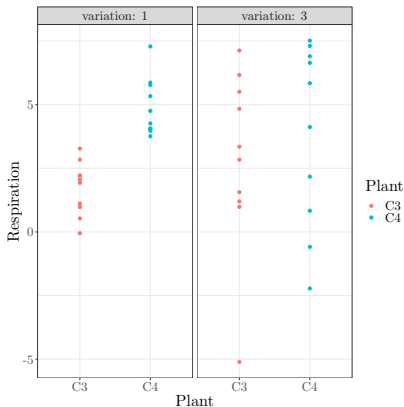


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- 1 Large true difference between the means
- 2 Large sample size

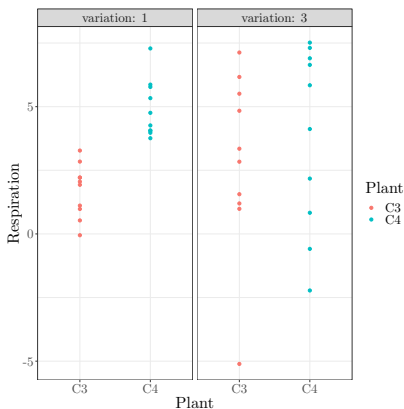
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Is it easier when unexplained variation is 1 or when it is 3?

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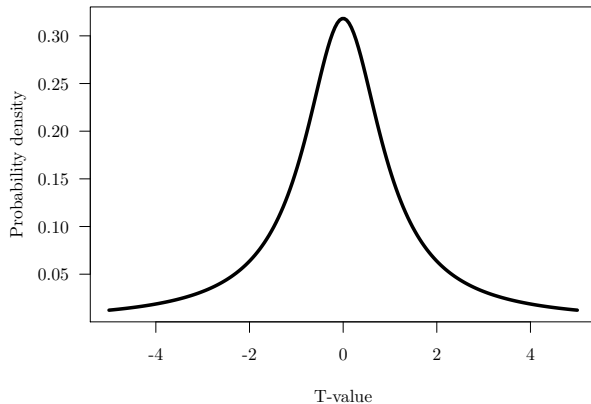
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## What makes $t$ large:

- 1 Large true difference between the means
- 2 Large sample size
- 3 Small unexplained variation

# When can we know whether $D \neq 0$ ?

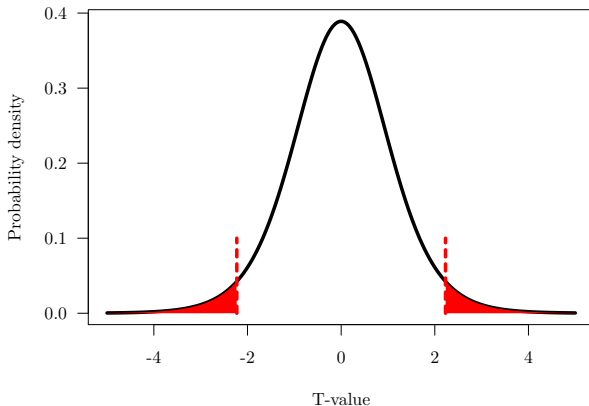
Expected t-values when  $D = 0$



# When can we know whether $D \neq 0$ ?

**p-value:** probability (area under curve) of getting a value as extreme as what you observed, when  $D=0$

Expected t-values when  $D = 0$





# But really, what is a p-value?

## Candy practical

You got 5 Halloween candies out of the bag. Does the bag contain more Halloween than normal candies?

# Back to C3/C4 plants. Analyse real data in R

1. Set working directory (`setwd(' / ')`) or create a R-project
2. Load and check data

```
resp <- read.csv("d_respiration.csv")  
str(resp)  
View(resp)
```

3. Visualize data

```
library(ggplot2)  
ggplot(resp, aes(Plant_type, rrarea, colour=Plant_type))+  
  geom_point()+facet_wrap(~Variation)
```

# Fit a t-test in R: `t.test()`

## Subset data by Variation (High and Low)

```
resp_H <- subset(resp, Variation == "High")  
resp_L <- subset(resp, Variation == "Low")
```

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## Compare C3 and C4 plants in “High Variation” subset

```
t.test(rrarea~Plant_type, data=resp_H, var.equal=TRUE)
```

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```
Two Sample t-test  
data: rrarea by Plant_type  
t = -0.93776, df = 10, p-value = 0.3705  
alternative hypothesis: true difference in means is not equal to 0  
95 percent confidence interval:  
 -1.7619349  0.7181446  
sample estimates:  
mean in group C3 mean in group C4  
      2.720021      3.241916
```

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```

# Fit a t-test in R: `t.test()`

**Compare C3 and C4 plants in “Low Variation” subset**

```
t.test(rrarea~Plant_type, data=resp_L, var.equal=TRUE)
```

# Fit an anova in R: aov()

```
aov1 <- aov(rrarea~Plant_type, data=resp_H)  
summary(aov1)
```



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aov1 <- aov(rrarea~Plant_type, data=resp_H)
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```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Plant_type	1	0.817	0.8171	0.879	0.37
Residuals	10	9.292	0.9292		

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aov1 <- aov(rrarea~Plant_type, data=resp_H)
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$$\text{response} = A + D \times \text{predictor} + \epsilon$$

# Fit a linear model in R: `lm()`

```
lm1<-lm(rrarea ~ Plant_type, data = resp_L)  
summary(lm1)
```

# Fit a linear model in R: `lm()`

```
lm1<-lm(rrarea ~ Plant_type, data = resp_L)
summary(lm1)
```

```
lm(formula = rrarea ~ Plant_type, data = resp_H)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-1.7380	-0.4201	-0.1437	0.6706	1.6754

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	2.7200	0.3935	6.912	4.13e-05 ***
Plant_typeC4	0.5219	0.5565	0.938	0.37

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.9639 on 10 degrees of freedom

Multiple R-squared: 0.08083, Adjusted R-squared: -0.01109

F-statistic: 0.8794 on 1 and 10 DF, p-value: 0.3705

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lm1<-lm(rrarea ~ Plant_type, data = resp_L)
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# Fit a linear model in R: `lm()`

```
library(emmeans)
emmeans(lm1, ~Plant_type)
```

Plant_type	emmean	SE	df	lower.CL	upper.CL
C3	2.720021	0.3935305	10	1.843180	3.596861
C4	3.241916	0.3935305	10	2.365076	4.118757

Confidence level used: 0.95

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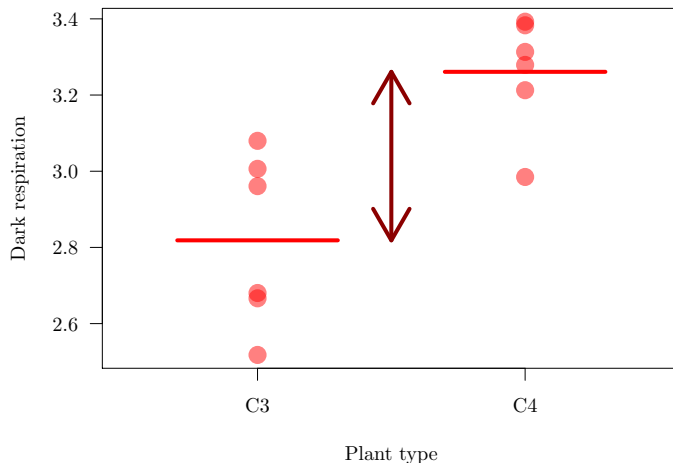
$$\text{response} = A + D \times \text{predictor} + \epsilon$$

# Compare the output from t.test, aov and lm



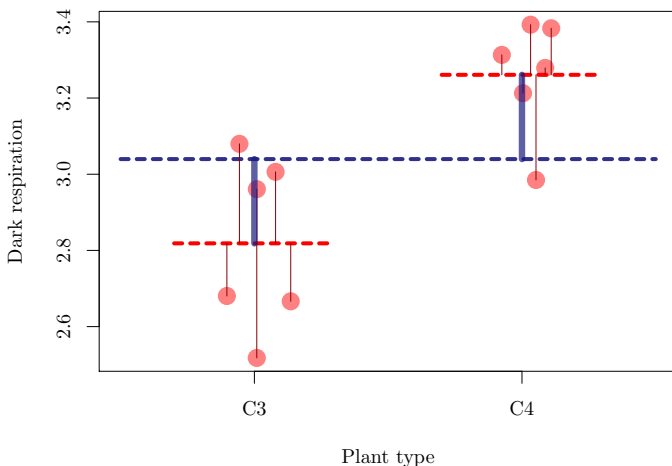
# Three equivalent ways to look at data

T-test, focus on difference between two means



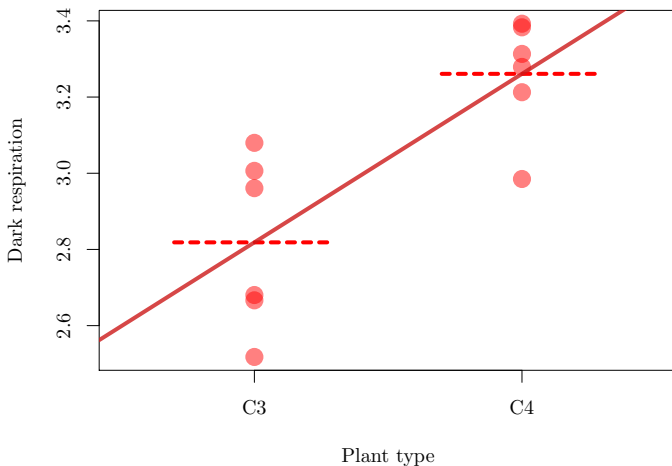
# Three equivalent ways to look at data

ANOVA, focus on variation within VS. between



# Three equivalent ways to look at data

Linear regression, focus on rate of change



# All is one. . .

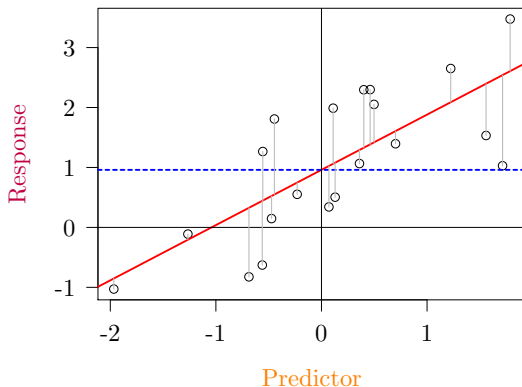
# All is one. . .

## ...but `lm()` rules (IMH)

- t-test, ANOVA, regression and others can be mathematically equivalent
- In R, `lm()` and related functions can do them all. . .
- ...and much more!

# Focus on linear models

$$\text{Response} = \text{Intercept} + \text{Slope} \times \text{Predictor} + \text{Error}$$



# A simple linear model

$$\text{Response} = \text{Intercept} + \text{Slope} \times \text{Predictor} + \text{Error}$$

```
lm(response ~ 1 + predictor1 + predictor2, data=data)
```

equivalent to

```
lm(response ~ predictor1 + predictor2, data=data)
```

equivalent to

```
lm(response ~ predictor2 + predictor1, data=data)
```

- Intercept can be explicit or implicit
- Can remove intercept with `... ~ 0 + ...`
- Error is implicit
- Feed the option `data=` to keep code short, reliable and flexible
- Order of predictors do not matter

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# General approach

## 1. Scientific question

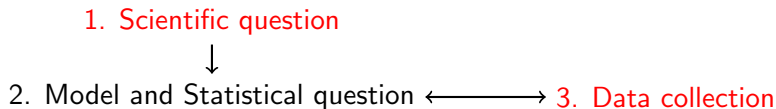
# General approach

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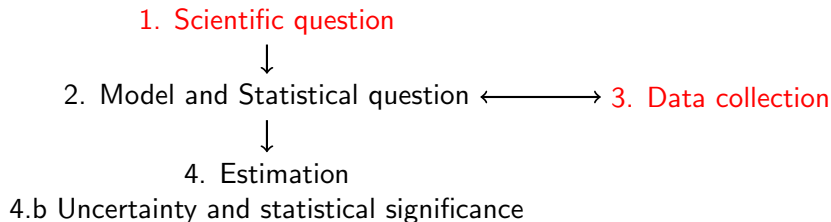


2. Model and Statistical question

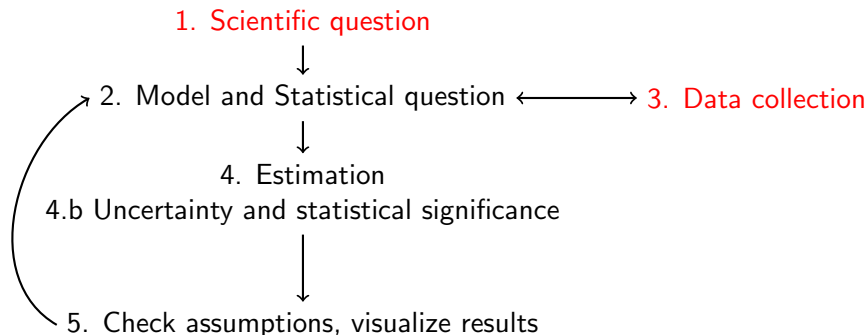
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# General approach

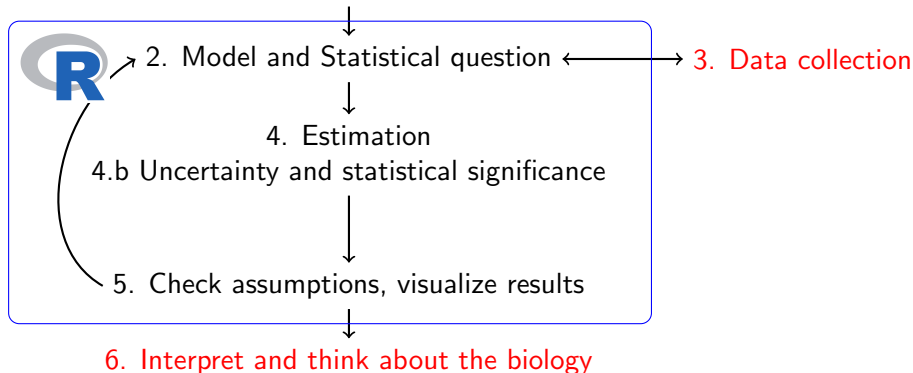


# General approach



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1. Scientific question



## Back to C3/C4

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lmL<-lm(rrarea ~ Plant_type, data = resp_L)  
summary(lmL)
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              Estimate Std. Error t value Pr(>|t|)
(Intercept)  2.81857    0.07856  35.878 6.72e-12 ***
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```

## Estimation:

$$\text{response} = A + D \times \text{predictor} + \epsilon$$

$$A = ?, D = ?$$



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## Uncertainty:

For  $D$  SE= 0.11110 ; p-value=0.00259

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## What do we do next?

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- Gaussian error distribution  
*Risk: Poor predictions*
- Homoscedasticity (constant error variance)  
*Risk: Over-optimistic uncertainty, unreliable predictions*
- Independence of error  
*Risk: Bias and over-optimistic uncertainty*

# Check assumptions, visualize results

Assessing model assumptions in R:

```
lmL<-lm(rrarea ~ Plant_type, data = resp_L)
plot(lmL)
summary(lmL)
```

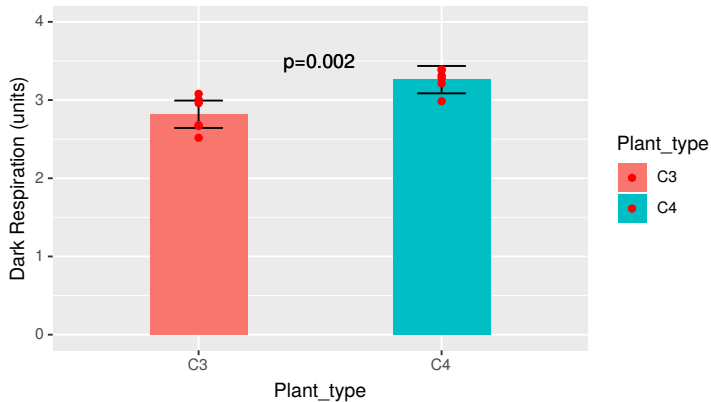
# Check assumptions, visualize results

## Visualize and report results

```
lm1.results<-summary(emmeans(lm1,~Plant_type))

ggplot(lm1.results,aes(Plant_type,emmean, fill=Plant_type))+
  geom_bar(stat="identity", width=.4)+
  geom_errorbar(aes(ymin =lm1.results$lower.CL,
ymax = lm1.results$upper.CL), width=.2)+
  ylim(0,4)+
  geom_point(data=resp_L, aes(x=Plant_type, y=rrarea), color="red")+
  labs(y = "Dark Respiration (units)")+
  geom_text(aes(x=1.5, y=3.5, label="p=0.002"))
```

# Check assumptions, visualize results



# Another example

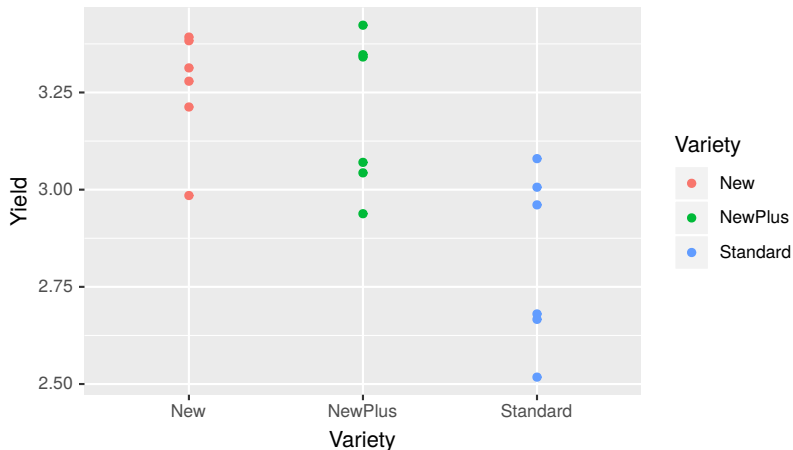
## Compare wheat yields between 3 varieties

- Outcome measure: Tonnes/hectare
- Experimental factor: Variety (new/newPLUS/standard)
- Data: 6 plots/ variety



How many parameters in this model?

# Results from Wheat Yield Experiment with 3 Varieties



# Analyse these data in R

```
wheat2<-read.csv("wheat yield PLUS.csv")  
str(wheat2) #check data types for each variable  
View(wheat2) #View data  
ggplot(wheat2, aes(Variety, Yield, colour=Variety)) +  
  geom_point()
```

# Sample analysis in R: 1-way ANOVA

```
aov1=aov(Yield~Variety, data = wheat2)
summary(aov1)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Variety	2	0.6820	0.3410	8.951	0.00276 **
Residuals	15	0.5714	0.0381		

## ANOVA

- Compares means between TWO or MORE GROUPS
- Relies on F-statistic =  $\frac{\text{Between-groups variance}}{\text{Within groups variance}} = \frac{\text{Explained}}{\text{Unexplained}}$
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*But which groups are different???*

# Using emmeans to extract group estimates

```
emmeans(aov1, ~Variety)
```

Variety	emmean	SE	df	lower.CL	upper.CL
Standard	2.81	0.079	15	2.64	2.98
New	3.26	0.079	15	3.09	3.43
NewPlus	3.19	0.079	15	3.02	3.36

Confidence level used: 0.95

# Using emmeans to compare groups

```
emmeans(aov1, pairwise~Variety)
```

```
$contrasts
```

contrast	estimate	SE	df	t.ratio	p.value
Standard - New	-0.44	0.112	15	-3.925	0.0036
Standard { NewPlus	-0.37	0.112	15	-3.330	0.0120
New - NewPlus	0.06	0.112	15	0.595	0.8248

P value adjustment: tukey method for comparing a family of 3 estimates

# Equivalent with lm()

```
lm2<-lm(Yield ~ Variety, data = wheat2)
anova(lm2)
summary(lm2)
```

## Analysis of Variance Table

Response: Yield

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Variety	2	0.68203	0.34101	8.9513	0.002764 **
Residuals	15	0.57145	0.03810		

## Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	2.81857	0.07968	35.372	7.26e-16 ***
VarietyNew	0.44235	0.11269	3.925	0.00135 **
VarietyNewPlus	0.37529	0.11269	3.330	0.00457 **

# Equivalent with `lm()`

```
emmeans(lm2, pairwise~Variety)
```

\$emmeans

Variety	emmean	SE	df	lower.CL	upper.CL
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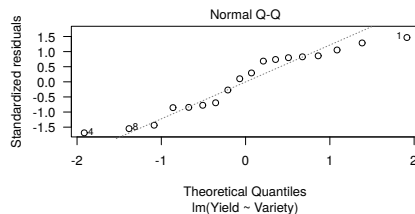
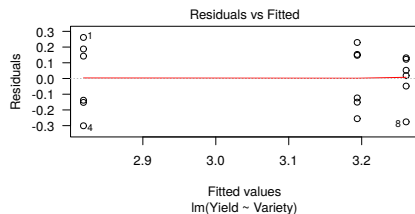
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# Assessing model assumptions for `lm()`

```
plot(lm2)
```



# Summary of results with post hoc comparisons:

