

Statistical Thinking in Biology Research

Chapter 1

Terry Neeman and Timothée Bonnet

November 27, 2018

Acknowledgements and warning

Key ideas for today

- Statistics in biology is the study of biological variation

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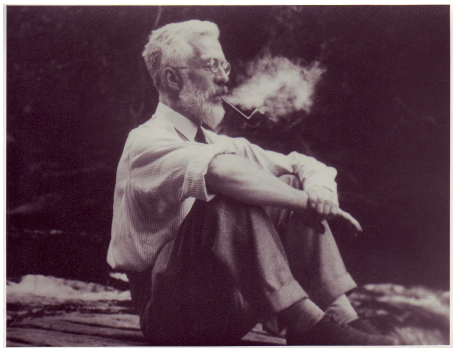
- Statistics in biology is the study of biological variation
- Statistical ideas about biological variation inform the design of experiments
- Statistical ideas about biological variation inform the analysis of experiments

Key ideas for today

- Statistics in biology is the study of biological variation
- Statistical ideas about biological variation inform the design of experiments
- Statistical ideas about biological variation inform the analysis of experiments
- Statistical thinking is an essential component of scientific thinking

A bit of history of statistical methods

R.A. Fisher: 1890-1962



Statistical Principles for Research Workers (1925)

A bit of history of statistical methods

R.A. Fisher: 1890-1962

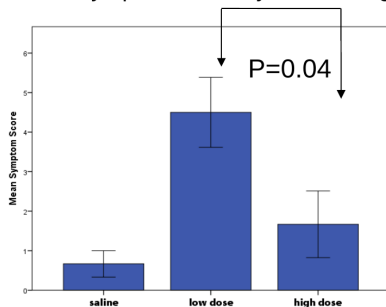


Statistical Principles for Research Workers (1925)

- 1 Cautionary tales from the front
- 2 Introduction to Statistical Modelling
- 3 Another look at essential steps

Message 1: A small p-value is not always evidence of a treatment effect

Mean symptom score by treatment group



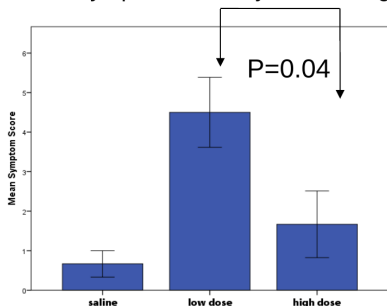
Vaccine challenge experiment:

- 6 mice/group (saline/low dose/high dose)
- All mice challenged with *Shigella*
- Followed for 14 days
- Outcome: Symptom score average Days 2 - 8

One-way ANOVA (post-hoc Bonferroni) $p=0.04$

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One-way ANOVA (post-hoc Bonferroni) $p=0.04$

Do you think the vaccine works? What is strange?

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Experimental design

The observed difference in outcome could be the result of:

- Cage effects
- Mouse strain effects

These effects are **CONFOUNDED** with treatment effect



Cage 1:
saline



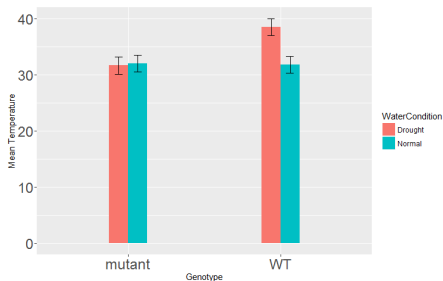
Cage 2:
Low Dose



Cage 3:
High Dose

Message 2: p-values from simple comparisons cannot tell us when differences are “different”

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Are temperature mechanisms modified in a genetically modified tomato plant?

- Genotypes: WT/mutant
- Water condition: Normal/Drought
- Leaf temperature measured

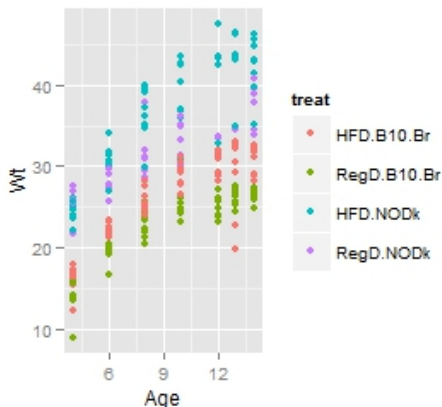
Comparisons made using t-tests

Evidence of difference + No evidence of difference \neq Evidence that differences are different.

Message 3: Interpreting experimental results needs more than t-tests

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Research question: Are mice susceptible to obesity when exposed to a high fat diet?



Experimental set-up:

- 37 mice: 16 NODk /21 WT
- Randomised to either regular or high fat diet
- Monitored for 14 weeks
- Outcome measure: Body weight (g)
- Experimental factors: Diet (2), Strain (2), Time (8)

Acknowledgements: Ainy Hussain, PhD student 2013

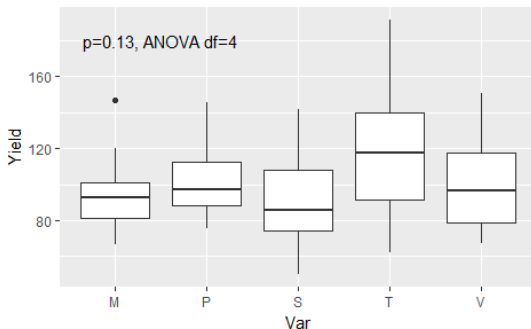
Message 4: Knowing how to combine information across subgroups can improve inference

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Comparing yield in five barley varieties (1930s)

Experimental factors: 5 varieties of barley, 6 locations, 2 time points.

Outcome measure: yield



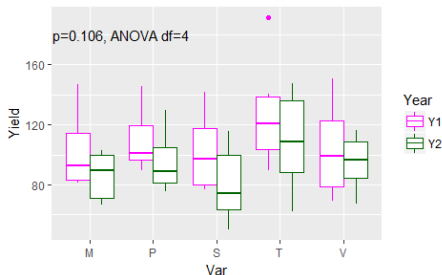
Acknowledgements: MASS R-package

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Controlling for other sources of variation:

- Controlling for year = comparing yield **WITHIN** years and combining these

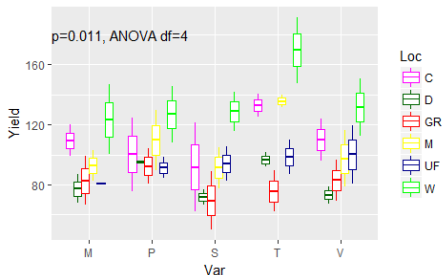
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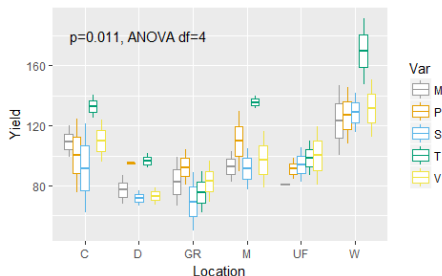


Controlling for other sources of variation:

- Control for year = compare yield WITHIN years and combine these
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Message 4: Knowing how to combine information across subgroups can improve inference



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Message 5: Knowing what factors contribute to the variation in outcome helps design experiments and analyses

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Research question: How does cold duration impact upon germination in alpine plant *A. glacialis*?



Experimental set-up:

- Seed collections from alpine region in Australia
- 3 Regions- low/high altitude
- 4 sets of Petri dishes
- 4 cabinet shelves
- Response - % germinated

What other factors are important to consider when comparing cold duration?

Summary

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- ④ Combining information across subgroups can improve inference. **A statistical model enables accumulation of evidence across experiments.**

Summary

- 1 A small p-value is not always evidence of a treatment effect. **Good experimental design matters.**
- 2 p-values from simple comparisons cannot tell us when differences are “different”. **For each question / comparison, a specific test**
- 3 Interpreting experimental results needs more than t-tests. **Need a statistical model of the experiment, matching scientific question.**
- 4 Combining information across subgroups can improve inference. **A statistical model enables accumulation of evidence across experiments.**
- 5 Knowing what factors contribute to the variation in outcome matters. **A statistical model allows one to incorporate effect of other factors in the analysis.**

- 1 Cautionary tales from the front
- 2 Introduction to Statistical Modelling
- 3 Another look at essential steps

Introduction to Statistical Modelling

- What is a statistical model?
- Modelling outcomes:
 - ▶ a summary of data
 - ▶ a prediction model
 - ▶ an explanatory model
- Model – may take many different functional forms
- Model – a conceptualization of the experiment

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ALWAYS BEGIN WITH A RESEARCH QUESTION

Key components of a statistical model of an experiment

- Outcome measure
 - ▶ Response variable
 - ▶ Measure of interest
- Experimental factors
 - ▶ Conditions that can be manipulated
 - ▶ Conditions of interest (e.g. genotype, gender)
 - ▶ Main questions: do the conditions impact upon the outcome measure?
- Blocking factors
 - ▶ Conditions (not of interest) that may impact upon the outcome measure
 - ▶ Sources of variation in the experiment that need to be controlled for
 - ▶ Clustering of experimental units

ALWAYS BEGIN WITH A RESEARCH QUESTION

Key Objectives of a statistical model of an experiment

- To compare the mean response of an organism/system to a set of different experimental conditions.
 - ▶ Obtain estimate of “Treatment effect”
 - ▶ Is this “effect” different in subgroups of interest?
- What are the most important factors influencing the mean response?
- Subsidiary question: how can we design our experiment in future to more efficiently test our hypotheses?

Example 1: Does dark respiration differ between C3 and C4 plants?

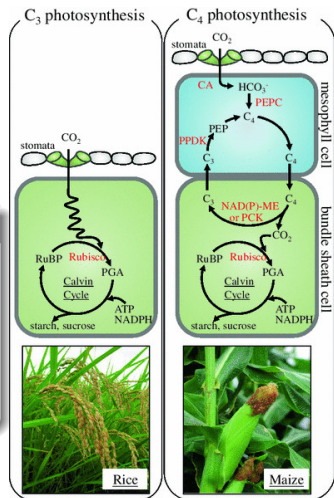
Outcome measure: dark respiration

Experimental factor: Plant type (C4/C3)

Data: 6 plants each of C4, C3

Can calculate

- Observed overall mean
- Observed mean C3 plants
- Observed mean C4 plants
- Variation around each mean



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Statistical model

Respiration = Mean for C3 + Difference C4-C3 * (is C4?) + Noise

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- Variation around each mean

Statistical model

Respiration = Mean for C3 + Difference C4-C3 * (is C4?) + Noise

response = *A* + *D* × *predictor* + ϵ

A and *D* are the model PARAMETERS.

We want to infer whether *D* is different from 0

Example 1: Does dark respiration differ between C3 and C4 plants?

$$\text{response} = A + D \times \text{predictor} + \epsilon$$

Can we separate the signal D from the noise ϵ ?

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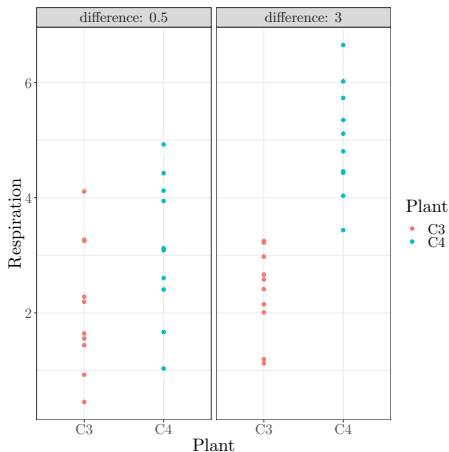
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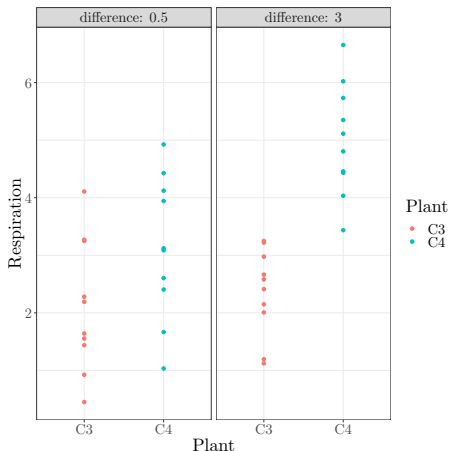
When can we know whether $D \neq 0$?



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Is it easier when the true difference is 0.5 or when it is 3 ?

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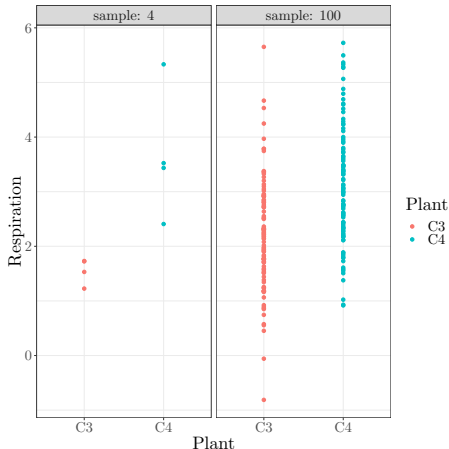


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1 Large true difference between the means

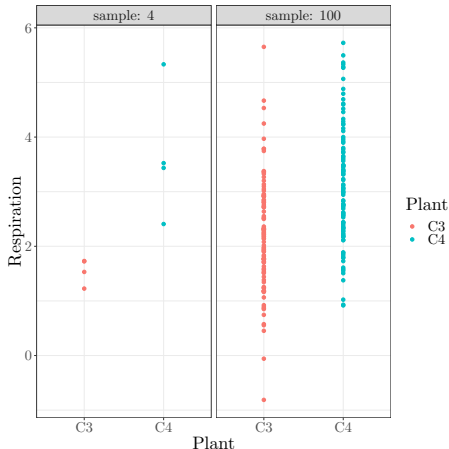
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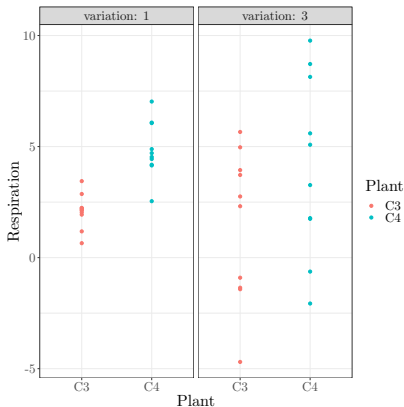


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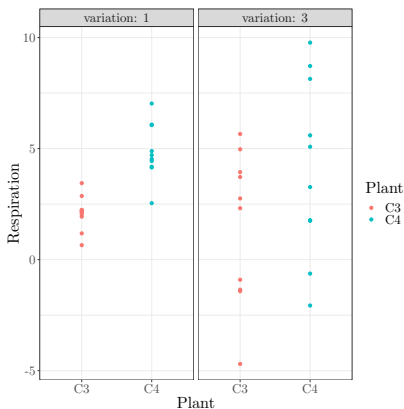
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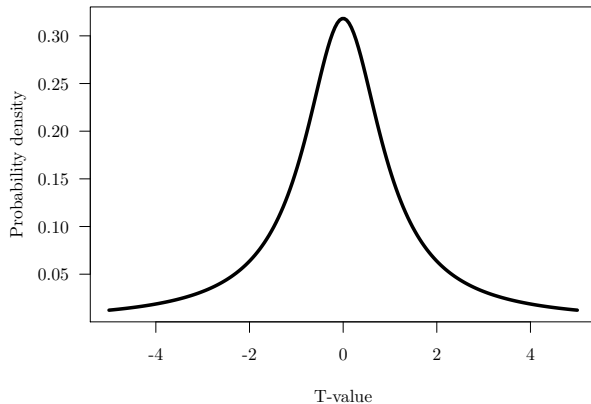
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What makes t large:

- 1 Large true difference between the means
- 2 Large sample size
- 3 Small unexplained variation

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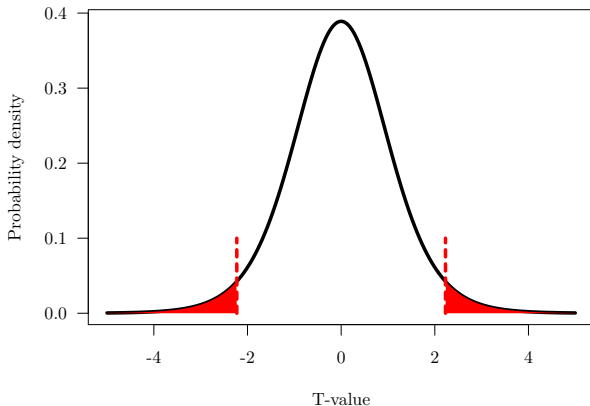
Expected t-values when $D = 0$



When can we know whether $D \neq 0$?

p-value: probability (area under curve) of getting a value as extreme as what you observed, when $D=0$

Expected t-values when $D = 0$



But really, what is a p-value?

Candy practical

- You got a pack of 20 candies with a mix of Halloween and Fruit candies

How to?

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- And so on, until 5 candies. You wonder if you have been cheated.
- Are there more Halloween than Fruit candies in that pack?
- You decide to use statistics to find out

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- Draw 5 candies out of the pack

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- Write down how many Halloween candies

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How to?

- Draw 5 candies out of the pack
- Write down how many Halloween candies
- How often is it 5?

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- How often is it 5?
- Estimate the p-value for the test “candies have same frequency”
- Redo the experiment in R, using random sampling (`rbinom`)
- What is the correct null-distribution?

Back to C3/C4 plants. Analyse real data in R

1. Set working directory (`setwd(' / ')`) or create a R-project
2. Load and check data

```
resp <- read.csv("d_respiration.csv")  
str(resp)  
View(resp)
```

3. Visualize data

```
library(ggplot2)  
ggplot(resp, aes(Plant_type, rrarea, colour=Plant_type))+  
  geom_point()+facet_wrap(~Variation)
```

Fit a t-test in R: `t.test()`

Subset data by Variation (High and Low)

```
resp_H <- subset(resp, Variation == "High")  
resp_L <- subset(resp, Variation == "Low")
```

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```

Compare C3 and C4 plants in “High Variation” subset

```
t.test(rrarea~Plant_type, data=resp_H, var.equal=TRUE)
```


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```

```
Two Sample t-test  
data: rrarea by Plant_type  
t = -0.93776, df = 10, p-value = 0.3705  
alternative hypothesis: true difference in means is not equal to 0  
95 percent confidence interval:  
 -1.7619349  0.7181446  
sample estimates:  
mean in group C3 mean in group C4  
      2.720021      3.241916
```

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resp_H <- subset(resp, Variation == "High")  
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Two Sample t-test

data: rrarea by Plant_type

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sample estimates:

mean in group C3	mean in group C4
2.720021	3.241916

Fit a t-test in R: `t.test()`

Compare C3 and C4 plants in “Low Variation” subset

```
t.test(rrarea~Plant_type, data=resp_L, var.equal=TRUE)
```

Fit an anova in R: aov()

```
aov1 <- aov(rrarea~Plant_type, data=resp_H)  
summary(aov1)
```

Fit an anova in R: aov()

```
aov1 <- aov(rrarea~Plant_type, data=resp_H)
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```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Plant_type	1	0.817	0.8171	0.879	0.37
Residuals	10	9.292	0.9292		

Fit an anova in R: aov()

```
aov1 <- aov(rrarea~Plant_type, data=resp_H)
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Fit a linear model in R: `lm()`

```
lm1<-lm(rrarea ~ Plant_type, data = resp_L)  
summary(lm1)
```

Fit a linear model in R: `lm()`

```
lm1<-lm(rrarea ~ Plant_type, data = resp_L)
summary(lm1)
```

```
lm(formula = rrarea ~ Plant_type, data = resp_H)
```

Residuals:

Min	1Q	Median	3Q	Max
-1.7380	-0.4201	-0.1437	0.6706	1.6754

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	2.7200	0.3935	6.912	4.13e-05 ***
Plant_typeC4	0.5219	0.5565	0.938	0.37

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.9639 on 10 degrees of freedom

Multiple R-squared: 0.08083, Adjusted R-squared: -0.01109

F-statistic: 0.8794 on 1 and 10 DF, p-value: 0.3705

Fit a linear model in R: `lm()`

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lm1<-lm(rrarea ~ Plant_type, data = resp_L)
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Residuals:

	Min	1Q	Median	3Q	Max
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Fit a linear model in R: `lm()`

```
library(emmeans)
emmeans(lm1, ~Plant_type)
```

Plant_type	emmean	SE	df	lower.CL	upper.CL
C3	2.720021	0.3935305	10	1.843180	3.596861
C4	3.241916	0.3935305	10	2.365076	4.118757

Confidence level used: 0.95

Fit a linear model in R: `lm()`

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Plant_type	emmean	SE	df	lower.CL	upper.CL
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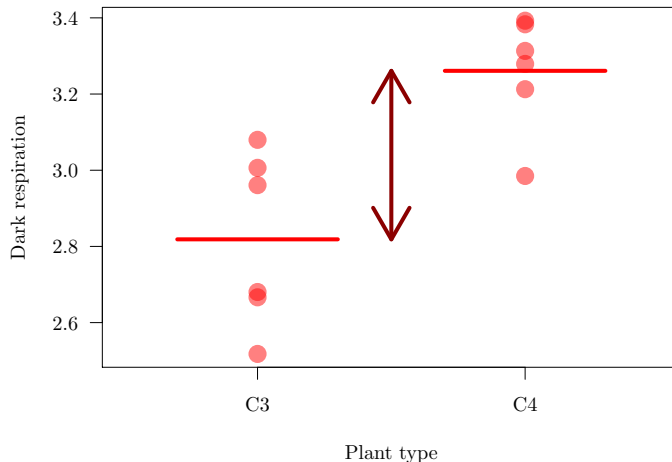
Confidence level used: 0.95

$$\text{response} = A + D \times \text{predictor} + \epsilon$$

Compare the output from t.test, aov and lm

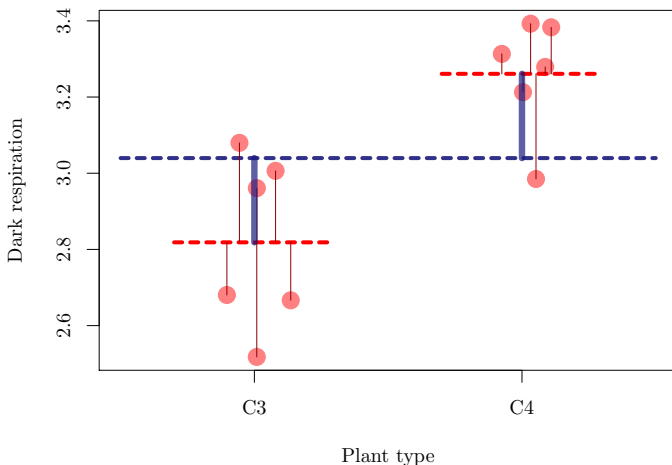
Three equivalent ways to look at data

T-test, focus on difference between two means



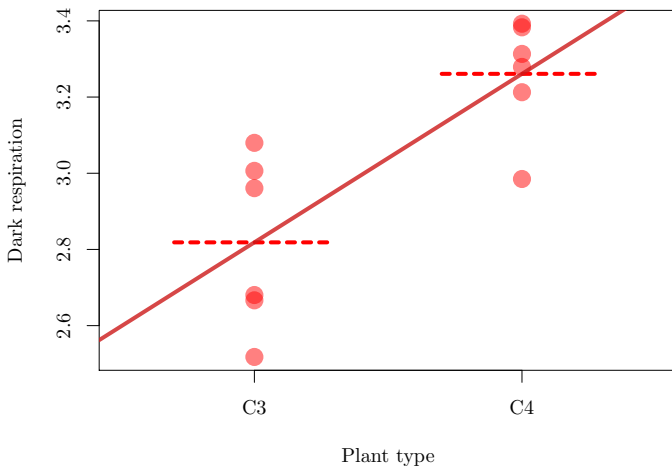
Three equivalent ways to look at data

ANOVA, focus on variation within VS. between



Three equivalent ways to look at data

Linear regression, focus on rate of change



All is one. . .

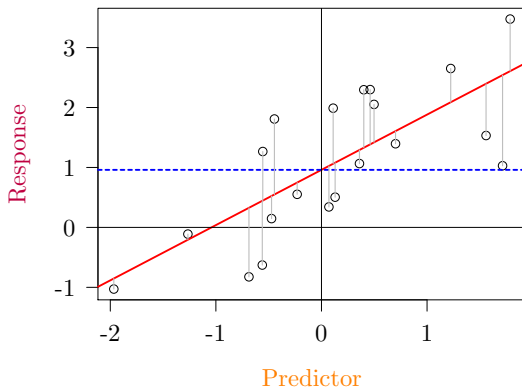
All is one. . .

...but `lm()` rules (IMH)

- t-test, ANOVA, regression and others can be mathematically equivalent
- In R, `lm()` and related functions can do them all. . .
- ...and much more!

Focus on linear models

$$\text{Response} = \text{Intercept} + \text{Slope} \times \text{Predictor} + \text{Error}$$



A simple linear model

$$\text{Response} = \text{Intercept} + \text{Slope} \times \text{Predictor} + \text{Error}$$

```
lm(response ~ 1 + predictor1 + predictor2, data=data)
```

equivalent to

```
lm(response ~ predictor1 + predictor2, data=data)
```

equivalent to

```
lm(response ~ predictor2 + predictor1, data=data)
```

- Intercept can be explicit or implicit
- Can remove intercept with `... ~ 0 + ...`
- Error is implicit
- Feed the option `data=` to keep code short, reliable and flexible
- Order of predictors do not matter

- 1 Cautionary tales from the front
- 2 Introduction to Statistical Modelling
- 3 Another look at essential steps

General approach

1. Scientific question

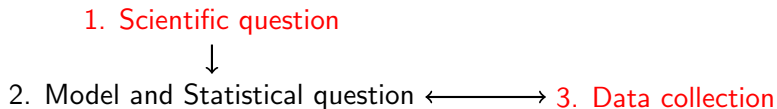
General approach

1. Scientific question

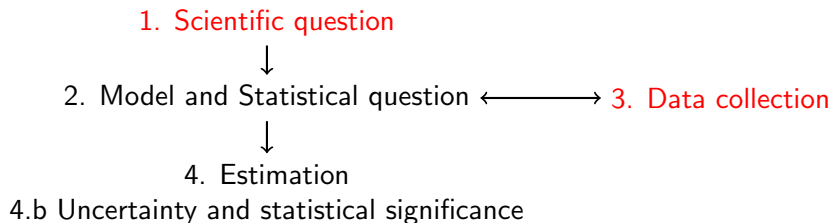


2. Model and Statistical question

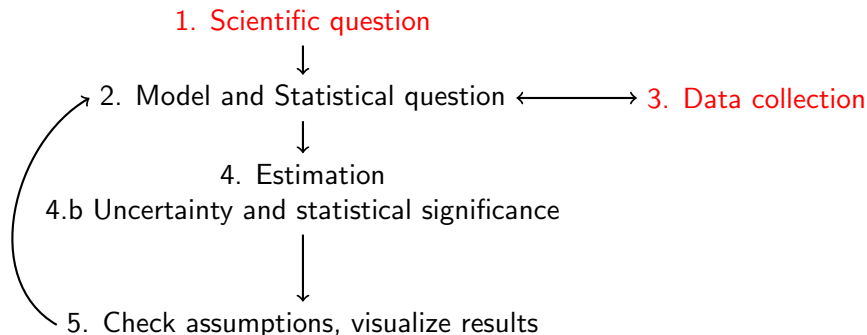
General approach



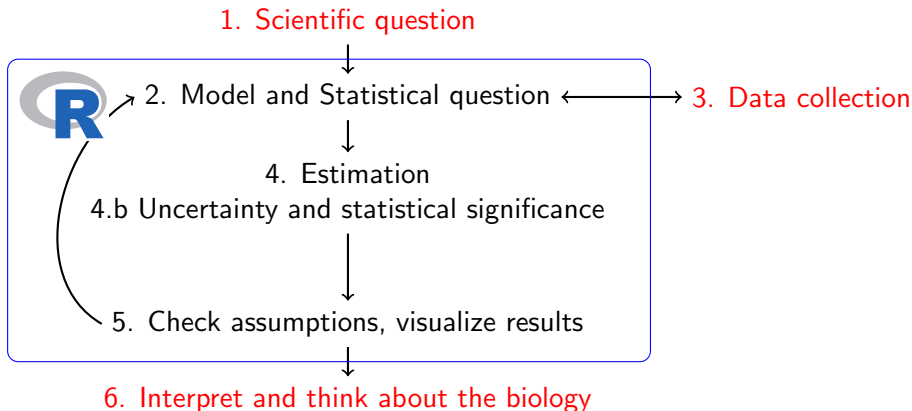
General approach



General approach



General approach



Back to C3/C4

```
lmL<-lm(rrarea ~ Plant_type, data = resp_L)  
summary(lmL)
```

Back to C3/C4

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lmL<-lm(rrarea ~ Plant_type, data = resp_L)
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```
...
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  2.81857    0.07856   35.878 6.72e-12 ***
Plant_typeC4  0.44235    0.11110    3.982  0.00259 **  ---
...
```

Estimation:

$$\text{response} = A + D \times \text{predictor} + \epsilon$$

$$A = ?, D = ?$$

Back to C3/C4

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$$A = 2.81857, D = 0.44235$$

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Estimation:

response = $A + D \times \text{predictor} + \epsilon$

$A = 2.81857$, $D = 0.44235$

Uncertainty:

For D SE= 0.11110 ; p-value=0.00259

Back to C3/C4

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What do we do next?

Check assumptions, visualize results

Check assumptions, visualize results

Linear model basic assumptions

- Predictor not perfectly correlated

Risk: Model won't run, unstable convergence, or huge SE

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Risk: bias estimates (underestimate with Gaussian error)
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Risk: Poor predictions
- Homoscedasticity (constant error variance)
Risk: Over-optimistic uncertainty, unreliable predictions

Check assumptions, visualize results

Linear model basic assumptions

- Predictor not perfectly correlated
Risk: Model won't run, unstable convergence, or huge SE
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Risk: bias estimates (underestimate with Gaussian error)
- Gaussian error distribution
Risk: Poor predictions
- Homoscedasticity (constant error variance)
Risk: Over-optimistic uncertainty, unreliable predictions
- Independence of error
Risk: Bias and over-optimistic uncertainty

Check assumptions, visualize results

Assessing model assumptions in R:

```
lmL<-lm(rrarea ~ Plant_type, data = resp_L)
plot(lmL)
summary(lmL)
```

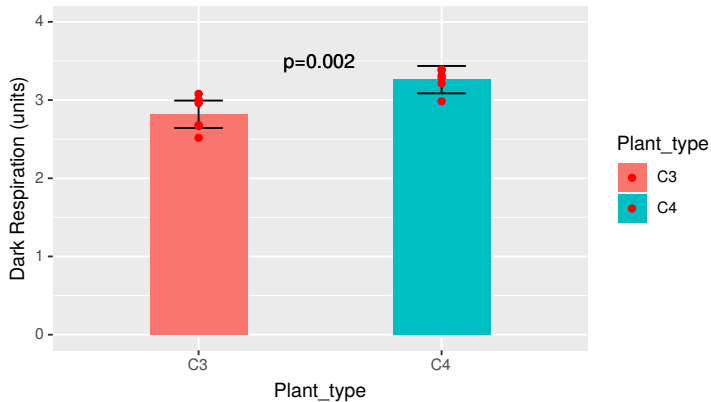
Check assumptions, visualize results

Visualize and report results

```
lm1.results<-summary(emmeans(lm1,~Plant_type))

ggplot(lm1.results,aes(Plant_type,emmean, fill=Plant_type))+
  geom_bar(stat="identity", width=.4)+
  geom_errorbar(aes(ymin =lm1.results$lower.CL,
ymax = lm1.results$upper.CL), width=.2)+
  ylim(0,4)+
  geom_point(data=resp_L, aes(x=Plant_type, y=rrarea), color="red")+
  labs(y = "Dark Respiration (units)")+
  geom_text(aes(x=1.5, y=3.5, label="p=0.002"))
```

Check assumptions, visualize results



Another example

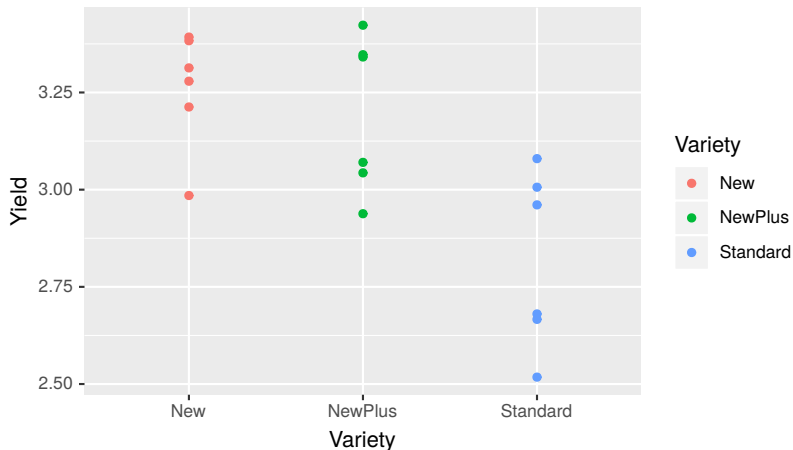
Compare wheat yields between 3 varieties

- Outcome measure: Tonnes/hectare
- Experimental factor: Variety (new/newPLUS/standard)
- Data: 6 plots/ variety



How many parameters in this model?

Results from Wheat Yield Experiment with 3 Varieties



Analyse these data in R

```
wheat2<-read.csv("wheat yield PLUS.csv")  
str(wheat2) #check data types for each variable  
View(wheat2) #View data  
ggplot(wheat2, aes(Variety, Yield, colour=Variety)) +  
  geom_point()
```

Sample analysis in R: 1-way ANOVA

```
aov1=aov(Yield~Variety, data = wheat2)
summary(aov1)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Variety	2	0.6820	0.3410	8.951	0.00276 **
Residuals	15	0.5714	0.0381		

ANOVA

- Compares means between TWO or MORE GROUPS
- Relies on F-statistic = $\frac{\text{Between-groups variance}}{\text{Within groups variance}} = \frac{\text{Explained}}{\text{Unexplained}}$
- One test for significance of all groups

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- One test for significance of all groups

But which groups are different???

Using emmeans to extract group estimates

```
emmeans(aov1, ~Variety)
```

Variety	emmean	SE	df	lower.CL	upper.CL
Standard	2.81	0.079	15	2.64	2.98
New	3.26	0.079	15	3.09	3.43
NewPlus	3.19	0.079	15	3.02	3.36

Confidence level used: 0.95

Using emmeans to compare groups

```
emmeans(aov1, pairwise~Variety)
```

```
$contrasts
```

contrast	estimate		SE	df	t.ratio	p.value
Standard - New	-0.44	0.112	15	-3.925	0.0036	
Standard { NewPlus	-0.37	0.112	15	-3.330	0.0120	
New - NewPlus	0.06	0.112	15	0.595	0.8248	

P value adjustment: tukey method for comparing a family of 3 estimates

Equivalent with lm()

```
lm2<-lm(Yield ~ Variety, data = wheat2)
anova(lm2)
summary(lm2)
```

Analysis of Variance Table

Response: Yield

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Variety	2	0.68203	0.34101	8.9513	0.002764 **
Residuals	15	0.57145	0.03810		

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	2.81857	0.07968	35.372	7.26e-16 ***
VarietyNew	0.44235	0.11269	3.925	0.00135 **
VarietyNewPlus	0.37529	0.11269	3.330	0.00457 **

Equivalent with `lm()`

```
emmeans(lm2, pairwise~Variety)
```

\$emmeans

Variety	emmean	SE	df	lower.CL	upper.CL
Standard	2.81	0.0796	15	2.64	2.98
New	3.26	0.0796	15	3.09	3.43
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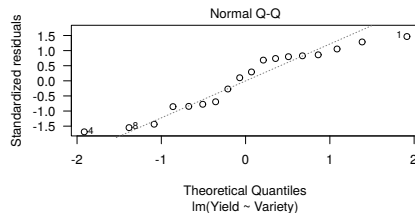
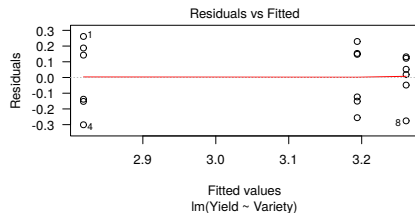
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New - NewPlus	0.067	0.112	15	0.595	0.8248

P value adjustment: tukey method for comparing a family of 3 estimates

Assessing model assumptions for `lm()`

```
plot(lm2)
```



Summary of results with post hoc comparisons:

