Consider a permutation test (the ideas here apply broadly to all sampling-based tests of independence or conditional independence, including marginal permutation test, conditional permutation test, marginal randomization test, and conditional randomization test). Let B be the number of resamples. Let $T_1, \ldots, T_B$ be the ordered resampled test statistics. Let T^* be the test statistic computed on the raw data. Let I_i be the indicator that T^* is less than T_i, i.e.

$$I_i = \mathbb{I}\left(T^* \leq T_i\right).$$

We have the basic fact under the null hypothesis that

$$\mathbb{E}(I_i) = \mathbb{P}(T^* \leq T_i) = i/B.$$

We want to combine $I_1, \ldots, I_B$ in such a way so as to produce an e-value E (i.e., a non-negative variable E such that $\mathbb{E}(E) \leq 1$). To test the hypothesis, we check if $E > \beta$ for some threshold β.

We call $f : \mathbb{R}^B \to \mathbb{R}^{\geq 0}$ a "combining function" if f combines $I_1, \ldots, I_B$ in such a way so as to produce a valid E-value, i.e.,

$$f(I_1, \ldots, I_B) \geq 0; \quad \mathbb{E}\left[f(I_1, \ldots, I_B)\right] \leq 1.$$

Let $\mathcal{F}$ be the class of combining functions. An interesting subset of $\mathcal{F}$ is the set of linear combining functions. A function $g \in \mathcal{F}$ is a linear combining function if

$$g(I_1, \ldots, I_B) = c + \sum_{i=1}^{B} a_i I_i$$

for given scalars $a_1, \ldots, a_B$, and c. The linearity of expectation implies that

$$\mathbb{E}\left[g(I_1, \ldots, I_B)\right] = c + \sum_{i=1}^{n} a_i \mathbb{E}(I_i) = c + \frac{1}{B} \sum_{i=1}^{B} (a_i)(i).$$

Therefore, under the stricter requirement that $\mathbb{E}(E) = 1$, we have the constraints

i $\sum_{i=1}^{B} (a_i) i = B(1 - c)$

ii $c + \sum_{i=1}^{B} a_i I_i \geq 0$.

We give a couple examples of linear combining functions.

Example 1. Let $c = 0$ and $a_i = 1/i$ for all i. Then

$$\sum_{i=1}^{B} (1/i)i = B(1-0),$$

confirming constraint (i). Next,

$$0 + \sum_{i=1}^{B} \frac{I_i}{i} \geq 0,$$

confirming constraint (ii). Therefore,

$$E = \sum_{i=1}^{B} \frac{I_i}{i}$$

is a valid e-value. This is a left-tailed test: if T^* is small, then E is big, leading us to reject the null.

Example 2. The standard permutation test is a special case of the proposed framework. Let $c = 3/2 + 1/(2B)$ and $a_i = -1/B$ for all a_i. Verifying condition (i),

$$\begin{aligned}\sum_{i=1}^{B} (a_i)i = -\frac{1}{B}\sum_{i=1}^{B} i = -\frac{1}{B}\left(\frac{B^2+B}{2}\right) &= -(1/2)B - 1/2 \\ &= B(1-(3/2)-1/(2B)) = B(1-c).\end{aligned}$$

Next, verifying condition (ii),

$$(3/2) + 1/(2B) - (1/B)\sum_{i=1}^{B} I_i \geq (3/2) + 1/(2B) - 1 \geq 0.$$

Therefore,

$$E = 3/2 + 1/(2B) - \frac{1}{B}\sum_{i=1}^{B} I_i$$

is a valid e-value. Note that $p_B := \frac{1}{B}\sum_{i=1}^{B} I_i$ is simply the p-value corresponding to the right-sided permutation test. For given $\alpha \in (0,1)$, we reject the null hypothesis if and only if $p_B < \alpha$, which is equivalent to rejecting the null hypothesis if and only if

$$E > 3/2 + 1/(2B) - \alpha.$$

Therefore, the p-permutation test is a special case of the e-permutation test. We typically think of e-values as controlling FDR, but here e-values control type-I error.

We can extend this idea to control family-wise error rate.

Example 3. Suppose we are using CRT or CPT to test conditional independence. If we have misspecified the model for $X|Z$, then the resampled test statistics $T_1, \ldots, T_B$ will be off. In particular, qq-plots often reveal that in such settings the tail of the distribution is incorrect, leading to p-value inflation. We can use the proposed framework to downweight re-sampled test statistics in the tail of the distribution, possibly leading to better calibration. For example, consider again example 1, where previously we set $a_i = 1/i$ for all i. Suppose that we use (e.g.) a Gaussian kernel to assign the weights. Let

$$\hat{\mu} = \frac{1}{B}\sum_{i=1}^{B} T_i$$

be the average of the resampled test statistics. Next, let

$$K(t) = \frac{1}{\sqrt{2\pi}} e^{-(1/2)t^2}$$

be the Gaussian kernel. Define a_i as

$$a_i := (B/i)\frac{K(|T_i - \hat{\mu}|/h)}{\sum_{i=1}^{B} K(|T_i - \hat{\mu}|/h)}.$$

Again, set $c = 0$. We see that

$$\sum_{i=1}^{B}(a_i)i = \frac{\sum_{i=1}^{B} K(|T_i - \hat{\mu}|/h)}{\sum_{i=1}^{B} K(|T_i - \hat{\mu}|/h)} = B = B(1-c),$$

verifying the first condition. Moreover, $A_i \geq 0$ for all i, satisfying the second condition. It follows that

$$E := \sum_{i=1}^{B} a_i I_i$$

is a valid e-value. This e-value is less sensitive to the tails of the sampling distribution than the e-value from example 1.

Advantages to this approach. The proposed framework has several advantages (I think). First, because the outputs are e-values, they can be combined under arbitrary dependence to control FDR using the e-BH procedure. Second, the framework is flexible: we can choose among many different "combining functions" (including the easy-to-use linear variants) to achieve different objectives. Certain combining functions might have robustness properties, for example. Third, we easily can combine e-values across data splits, enabling us to run multiple train-test splits without worrying about combining the resulting p-values.

Open questions. What is the "best" combining function that we could use?

Additional properties of the I_is. Recall that $I_i = \mathbb{I}(T^* \leq T_i)$. We state several additional properties of the I_is. Let $k_1, k_d, \ldots, k_p \in \mathbb{N}$ be indexes such that

$$1 \leq k_1 \leq k_2 \leq \cdots \leq k_p \leq B.$$

Furthermore, let $r_1, r_2, \ldots, r_p \in \mathbb{N}$ be non-negative integers. We have that

$$I_{k_1}^{r_1} I_{k_2}^{r_2} \ldots I_{k_p}^{r_p} = I_{k_1}.$$

The reason is as follows. First, because I_{k_i} is Bernoulli,

$$I_{k_1}^{r_1} I_{k_2}^{r_2} \ldots I_{k_p}^{r_p} = I_{k_1} I_{k_2} \ldots, I_{k_p}.$$

Next, we take cases on T^*. First, suppose that $T^* \leq k_1$. Then $T^* \leq k_2, \ldots, k_p$. Therefore,

$$I_{k_1} I_{k_2} \ldots I_{k_p} = 1 \cdot 1 \ldots 1 = I_{k_1}.$$

Next, suppose that $T^* > k_1$. Then $I_{k_1} = 0$, and so

$$I_{k_1} I_{k_2} \ldots I_{k_p} = 0 \cdot I_{k_2} \ldots I_{k_p} = 0 = I_{k_1}.$$

This property is useful because it allows us to more easily evaluate expressions of the form $\left(\sum_{i=1}^{n} I_i\right)^r$ for $r \in \mathbb{N}$. As a warmup, we consider $r = 1$:

$$\sum_{i=1}^{B} I_i = a_i^{(1)} I_i,$$

where $a_1^{(1)} = 1$. Next, consider $r = 2$:

$$\left(\sum_{i=1}^{B} I_i\right)^2 = \sum_{i=1}^{B}\sum_{j=1}^{B} I_i I_j = \sum_{i=1}^{B} I_i^2 + 2\sum_{i=1}^{B-1}\sum_{j=i+1}^{B} I_i I_j = \sum_{i=1}^{B} I_i + 2\sum_{i=1}^{B-1}\sum_{j=i+1}^{B} I_i$$
$$= \sum_{i=1}^{B} I_i + \sum_{i=1}^{B-1} 2(B-i) I_i = \sum_{i=1}^{B} I_i + \sum_{i=1}^{B} 2(B-i) I_i = \sum_{i=1}^{B} (2B-2i+1) I_i = \sum_{i=1}^{B} a_i^{(2)} I_i,$$

where $a_i^{(2)} = 2B - 2i + 1$.

Consider the multinomial theorem:

$$\left(\sum_{i=1}^{B} I_i\right)^r = \sum_{k_1+k_2+\cdots+k_B=r} \binom{r}{k_1, k_2, \ldots, k_B} \prod_{i=1}^{B} I_i^{k_i},$$

where

$$\binom{r}{k_1, k_2, \ldots, k_B} = \frac{r!}{k_1! k_2! \ldots k_B!}.$$

Let

$$C(r, B) := \left\{ (k_1, k_2, \ldots, k_B) \in \{1, \ldots, r\}^B : \sum_{i=1}^{B} k_i = r \right\},$$

i.e., the set of tuples of length B of integers from 1 to r such that the elements of the tuple sum to r. Let $\tau : C(r, B) \to B$ be defined by

$$\tau(k_1, \ldots, k_B) = \min\{i : \{1, \ldots, B\} : k_i \geq 1\},$$

i.e., τ is the minimal nonzero index of a given configuration. The inverse τ^{-1} of τ is the set of configurations with a given minimal nonzero index. That is, for $s \in \{1, \ldots, B\}$,

$$\tau^{-1}(s) = \left\{ (0, \ldots, 0, k_s, k_{s+1}, \ldots, k_B) : \sum_{i=s}^{B} k_i = r \right\}.$$

Therefore, the multinomial theorem reduces to

$$\left(\sum_{i=1}^{B} I_i\right)^r = \sum_{(k_1,\ldots,k_B):k_1+\cdots+k_B=r} \binom{r}{k_1,k_2,\ldots,k_B} I_{\tau(k_1,\ldots,k_B)}$$
$$= \sum_{s=1}^{B} \sum_{(0,\ldots,0,k_s,\ldots,k_B):k_s\geq 1,k_s+\cdots+k_B=r} \binom{r}{k_1,k_2,\ldots,k_B} I_s.$$

We consider the above sum:

$$\sum_{(0,\ldots,0,k_s,\ldots,k_B):k_s\geq 1,k_s+\cdots+k_B=r} \binom{r}{k_1,k_2,\ldots,k_B}$$
$$= \sum_{(k_1,\ldots,k_{B-s+1}):k_1\geq 1,k_1+\cdots+k_{B-s+1}=r} \binom{r}{k_1,k_2,\ldots,k_B}$$
$$= \sum_{j=1}^{r} \sum_{l_1+\cdots+l_{B-s}=r-j} \binom{r}{j,l_1,\ldots,l_{B-s}}$$
$$= \sum_{j=1}^{r} \sum_{l_1+\cdots+l_{B-s}=r-j} \frac{r!}{j!l_1!\ldots l_{B-s}!}$$
$$= \sum_{j=1}^{r} \sum_{l_1+\cdots+l_{B-s}=r-j} \frac{r(r-1)\ldots(r-j+1)(r-j)!}{j!l_1!\ldots l_{B-s}!}$$
$$= \sum_{j=1}^{r} \sum_{l_1+\cdots+l_{B-s}=r-j} \frac{r!}{(r-j)!j!}\binom{r-j}{l_1,\ldots,l_{B-s}}$$
$$= \sum_{j=1}^{r} \frac{r!}{(r-j)!j!} \sum_{l_1,\ldots,l_{B-s}} \binom{r-j}{l_1,\ldots,l_{B-s}}$$
$$= \sum_{j=1}^{r} \binom{r}{j}(B-s)^{r-j} = \sum_{j=1}^{r} \binom{r}{j}(B-s)^{r-j}1^j$$
$$= \sum_{j=0}^{r} \binom{r}{j}(B-s)^{r-j}1^j - 1(B-s)^r = (B-s+1)^r - (B-s)^r := M(B,r,s).$$

Therefore,

$$\left(\sum_{i=1}^{B} I_i\right)^r = \sum_{s=1}^{B} M(B,r,s) I_s.$$

The number $M(B,r,s)$ can be quite large. However, normalizing the sum by B neutralizes this problem:

$$\begin{aligned}\left(\frac{1}{B}\sum_{i=1}^{B} I_i\right)^r &= (1/B)^r \sum_{i=1}^{B} \left[(B-i+1)^r - (B-i)^r\right] I_i \\ &= \sum_{i=1}^{B} \left[(1-i/B+1/B)^r - (1-i/B)^r\right] I_i.\end{aligned}$$

All terms are easy to compute.

1 Checking correctness

Let us check this formula for specific values of r. First, setting $r=1$, we have $M(B,1,s) = (B-s+1)-(B-s) = 1$. Therefore,

$$\left(\sum_{i=1}^{B} I_s\right)^1 = \sum_{i=1}^{B} I_s.$$

Next, setting $r=2$, we have

$$M(B,1,s) = (B+s-1)^2 - (B-s)^2 = 2B-2i+1.$$

Therefore,

$$\left(\sum_{i=1}^{B} I_i\right)^2 = \sum_{i=1}^{B} [2B-2i+1] I_i,$$

which matches what we have above.

2 Polynomials

We easily can compute a polynomial of $\sum_{i=1}^{B} I_i$. We have that

$$\begin{aligned}
\left(\sum_{i=1}^{B} I_i - x_0\right)^l &= \sum_{k=0}^{l} \binom{l}{k} \left(\sum_{i=1}^{B} I_i\right)^k x_0^{l-k} \\
&= x_0^l + \sum_{k=1}^{l} \binom{l}{k} \left(\sum_{i=1}^{B} I_i\right)^k x_0^{l-k} \\
&= x_0^l + \sum_{k=1}^{l} \binom{l}{k} \left[\sum_{i=1}^{B} M(B,k,i) I_i\right] x_0^{l-k} \\
&= x_0^l + \sum_{k=1}^{l} \sum_{i=1}^{B} x_0^{l-k} \binom{l}{k} M(B,k,i) I_i \\
&= x_0^l + \sum_{i=1}^{B} \left[\sum_{k=1}^{l} x_0^{l-k} \binom{l}{k} M(B,k,i)\right] I_i
\end{aligned}$$

Finally, for an integer r and polynomial coefficients $a_0, a_1, \ldots, a_r \in \mathbb{R}$, we have that

$$\begin{aligned}
\sum_{l=0}^{r} a_l \left(S - x_0\right)^l &= a_0 + \sum_{l=1}^{r} a_l (S - x_0)^l \\
&= a_0 + \sum_{l=1}^{r} a_l \left[x_0^l + \sum_{i=1}^{B} \left[\sum_{k=1}^{l} x_0^{l-k} \binom{l}{k} M(B,k,i)\right] I_i\right] \\
&= \left(a_0 + \sum_{l=1}^{r} a_l x_0^l\right) + \sum_{i=1}^{B} \left[\sum_{l=1}^{r} \sum_{k=1}^{l} a_l x_0^{l-k} \binom{l}{k} M(B,k,i)\right] I_i.
\end{aligned}$$

Typically, we will have $a_l = f^{(l)}(x_0)/l!$, where f is a p-to-e calibrator. This follows from the following rth order Taylor expansion:

$$f\left(\sum_{i=1}^{B} I_i\right) \approx \left(a_0 + \sum_{l=1}^{r} a_l x_0^l\right) + \sum_{i=1}^{B} \left[\sum_{l=1}^{r} \sum_{k=1}^{l} \frac{f^{(l)}(x_0) x_0^{l-k} \binom{l}{k} M(B,k,i)}{l!}\right] I_i.$$

Of course, the above is a linear combination of the I_is. We define

$$W(f, x_0, r, B, i) = \sum_{l=1}^{r} \sum_{k=1}^{l} \frac{f^{(l)}(x_0) x_0^{l-k} \binom{l}{k} M(B,k,i)}{l!}.$$

Therefore, we can write

$$f\left(\sum_{i=1}^{B} I_i\right) \approx \left(a_0 + \sum_{l=1}^{r} a_l x_0^l\right) + \sum_{i=1}^{B} W(f, x_0, r, B, i) T_i.$$

In practice, we could compute the W terms ahead of running the algorithm. We even could save these terms in within the package for common choices of f, x_0, r, and B (e.g., $f(p) = 2p^{-1}$, $x_0 = 1/2$, $r = 20$, $B = 1000$).

3 p-to-e calibrators

A decreasing function $f : [0, 1] \to [0, \infty]$ is a p-to-e calibrator if and only if $\int_0^1 f \leq 1$. It is admissible if and only if f is upper semicontinuous, $f(0) = \infty$, and $\int_0^1 f = 1$. An example of a p-to-e calibrator is

$$f_\kappa(p) = \kappa p^{\kappa-1}.$$

We derive the Taylor series of f. First, we compute the derivatives of f. The first derivative is

$$f_\kappa^{(1)}(p) = \kappa(\kappa - 1)p^{\kappa-2}.$$

Meanwhile, the second derivative is

$$f_\kappa^{(2)}(p) = \kappa(\kappa - 1)(\kappa - 2)p^{\kappa-3}.$$

In general, we have that

$$f_\kappa^{(m)}(p) = \prod_{j=0}^{m} (\kappa - j)p^{\kappa-m-1}.$$

Therefore, the Taylor series of f at a is

$$f_\kappa(p) = \sum_{m=0}^{\infty} \frac{f^{(m)}(a)}{m!}(p - a)^m.$$

4 Order stat probability

Let $Y, T_1, \ldots, T_B \sim F$. Suppose that F is supported on the real line (the bounded case is similar). Let $X = T_{(B)}$ be the maximum over the T_is.

The cumulative density function F_X of X is $F_X(x) = [F(x)]^n$. Further, $F_Y(y) = F(y)$. We have that

$$\mathbb{P}(Y \leq X) = \int_{-\infty}^{\infty} \int_{y}^{\infty} f(y)f(x)dxdy = \int_{-\infty}^{\infty} f(y) \left[\int_{y}^{\infty} f(x)dx \right] dy$$

$$= \int_{-\infty}^{\infty} f(y) \left[F_X(\infty) - F_X(y) \right] dy = \int_{-\infty}^{\infty} f(y)[1 - (F(y))^n]dy$$

$$= \int_{-\infty}^{\infty} f(y) - f(y)F(y)^n dy = F(y) - \frac{1}{n+1}[F(y)]^{n+1}\Big|_{-\infty}^{\infty} = 1 - \frac{1}{n+1} = \frac{n}{n+1}.$$

Now, in the case that $r = n$, we have that $F_X(x) = [F(x)]^n$, and the problem is easy (see above). We now consider the more general case. We have the basic identity

$$\int_{-\infty}^{\infty} f(y)F^m(y)dy = \frac{1}{m+1}.$$

We want to show that

$$\mathbb{P}(Y \leq X_r) = 1 - \frac{n+1-r}{n+1} = \frac{r}{n+1}.$$

For a given function $F_{X_r}(y)$, let $L(F_{X_r}(y))$ denote the linear operation

$$L(F_{X_r}(y)) = \int_{-\infty}^{\infty} f(y)F_{X_r}(y)dy.$$

We equivalently want to show that

$$L(F_{X_r}(y)) = \frac{n+1-r}{n+1},$$

where

$$F_{X_r}(y) = \sum_{j=r}^{n} \binom{n}{j} [F(y)]^j [1 - F(y)]^{n-j}$$

For example, when $r = n$, $F_{X_n}(y) = [F(y)]^n$, so $L(F_{X_n}(y)) = \frac{1}{n+1}$. Next, when $r = n - 1$,

$$F_{X_{n-1}}(y) = \binom{n}{n-1} [F(y)]^{n-1} [1 - F(y)] + F_{X_n}(y).$$

Considering the first piece, we have that

$$
\begin{aligned}
L(n[F(y)]^{n-1}[1-F(y)]) = nL(F^{n-1}(y) - F^{n}(y)) &= n\left(\frac{1}{n} - \frac{1}{n+1}\right) \\
&= 1 - \frac{n}{n+1} = \frac{1}{n+1}.
\end{aligned}
$$

Combining both pieces,

$$L(F_{X_{n-1}}(y)) = \frac{1}{n+1} + \frac{1}{n+1} = \frac{2}{n+1}.$$

Next, we examine

$$F_{X_{n-2}}(y) = \binom{n}{n-2}[F(y)]^{n-2}[1-F(y)]^{2} + F_{X-1}(y).$$

Focusing on the first piece, we have that

$$
\begin{aligned}
F^{n-2}(y)[1-F(y)]^{2} = F^{n-2}(y)(1 - 2F(y) + F^{2}(y)) \\
= F^{n-2}(y) - 2F^{n-1}(y) + F^{n}(y).
\end{aligned}
$$

Applying the linear operator and multiplying by the binomial coefficient,

$$
\begin{aligned}
\binom{n}{n-1}\left[L(F^{n-2}(y) - 2F^{n-1}(y) + F^{n}(y))\right] = \binom{n}{n-1}\left[\frac{1}{n-1} - \frac{2}{n} + \frac{1}{n+1}\right] \\
= \frac{1}{n+1}.
\end{aligned}
$$

Therefore, $L(F_{X-2}(y)) = 3/(n+1)$, as desired. Next, let $r \in \{1, \ldots, n\}$ be arbitrary. We want to show that

$$\binom{n}{r} L\left(F^{r}(y)[1-F(y)]^{n-r}\right) = \frac{1}{n+1}.$$

The binomial theorem implies

$$[1-F(y)]^{n-r} = \sum_{k=0}^{n-r} \binom{n-r}{k} F^{k}(y)(-1)^{k},$$

and so

$$F^r(y)[1-F(y)]^{n-r} = \sum_{k=0}^{n-r} \binom{n-r}{k} F^{k+r}(y)(-1)^k,$$

implying

$$\binom{n}{r} L(F^r(y)[1-F(y)]^{n-r}) = \binom{n}{r} \sum_{k=0}^{n-r} \binom{n-r}{k} \frac{(-1)^k}{k+r+1} := Q(n,r).$$

Surry 2004 Corollary 2.2 implies

$$\sum_{k=1}^{n-r} \binom{n-r}{k} \frac{(-1)^k}{k+r+1} = \frac{1}{(n+1)\binom{n}{r}}.$$

Therefore,

$$\binom{n}{r} L(F^r(y)[1-F(y)]^{n-r}) = \frac{1}{n+1}.$$

We conclude that

$$L(F_{X_r}(y)) = \frac{n+1-r}{n+1},$$

implying

$$\mathbb{P}(Y \leq X_r) = \frac{r}{n+1},$$

as desired.

Suppose without loss of generality that F is supported on the real line (the proof is similar in the case of bounded F). For notational simplicity, let $Y = T^*$. Denote the density of F by f (i.e., $f = F'$). Denote the density of Y, T_i, and (Y, T_i) by $f_Y = f$, f_{T_i}, and f_{Y,T_i} respectively. Standard distributional theory on order statistics indicates that the CDF F_{T_i} of T_i is given by

$$F_{T_i}(t) = \sum_{j=i}^{B} \binom{B}{j} [F(t)]^j [1-F(t)]^{B-j}.$$

Meanwhile, the CDF F_Y of Y is given by $F_Y = F$. We seek to show that $\mathbb{P}(Y \leq T_i) = i/(B+1)$. We have that

$$\mathbb{P}(Y \leq T_i) = \int_{-\infty}^{\infty} \int_{y}^{\infty} f_{Y,T_i}(y,t)dtdy = \int_{-\infty}^{\infty} \int_{y}^{\infty} f_Y(y) f_{T_i}(t)dtdy$$

$$= \int_{-\infty}^{\infty} f(y) \left[\int_{y}^{\infty} f_{T_i}(t)dt \right] dy = \int_{-\infty}^{\infty} f(y) \left[F_{T_i}(\infty) - F_{T_i}(y) \right] dy = 1 - \int_{-\infty}^{\infty} f(y) F_{T_i}(y) dy. \tag{1}$$

The problem reduces to evaluating the integral in the final equality of (1). We have that

$$\begin{aligned}
\int_{-\infty}^{\infty} f(y) T_{T_i}(y) dy &= \int_{-\infty}^{\infty} f(y) \sum_{j=i}^{B} \binom{B}{j} F^j(y)[1-F(y)]^{B-j} dt \\
&= \int_{-\infty}^{\infty} f(y) \sum_{j=i}^{B} \binom{B}{j} F^j(y) \sum_{k=0}^{B-j} \binom{B-j}{k} (-1)^k F^k(y) dt \\
&= \sum_{j=i}^{B} \binom{B}{j} \sum_{k=0}^{B-j} \binom{B-j}{k} (-1)^k \int_{-\infty}^{\infty} f(y) F^{j+k}(y) dy \\
&= \sum_{j=i}^{B} \binom{B}{j} \sum_{k=0}^{B-j} \binom{B-j}{k} (-1)^k \left[\frac{1}{1+j+k} F^{j+k+1}(y)|_{-\infty}^{\infty} \right] \\
&= \sum_{j=i}^{B} \binom{B}{j} \sum_{k=0}^{B-j} \binom{B-j}{k} (-1)^k \left[\frac{1}{1+j+k} \right]. \quad (2)
\end{aligned}$$

[?] (Corollary 2.2) showed that

$$\sum_{k=0}^{B-j} (-1)^k \binom{B-j}{k} \frac{1}{1+j+k} = \frac{1}{(B+1)\binom{B}{j}}.$$

Substituting the above into (2), we obtain.

$$\sum_{j=i}^{B} \binom{B}{j} \sum_{k=0}^{B-j} \binom{B-j}{k} (-1)^k \left[\frac{1}{1+j+k} \right] =$$

$$\sum_{j=i}^{B} \binom{B}{j} \frac{1}{(B+1)\binom{B}{j}} = \sum_{j=i}^{B} \frac{1}{B+1} = \frac{B-i+1}{B+1}. \quad (3)$$

Finally, combining (1), (2), and (3), we conclude that

$$\mathbb{P}(Y \leq T_i) = \mathbb{P}(T^* \leq T_i) = \frac{B+1}{B+1} - \frac{B-i+1}{B+1} = \frac{i}{B+1}.$$