

# BIOS 7731 HW 7

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1. a) Chebyshev's inequality :  $P[|X| \geq a] \leq \frac{E[X^2]}{a^2}$

can be used to derive the Bernoulli Law:

$$P[|\bar{X} - p| \geq \varepsilon] \leq \frac{E[(\bar{X} - p)^2]}{\varepsilon^2} = \frac{\text{Var}(\sum_{i=1}^n X_i)}{n^2 \varepsilon^2} = \frac{p(1-p)}{n \varepsilon^2} \quad \checkmark$$

Taking this formula, we set  $\varepsilon = 0.1$ . Because we want the probability that  $\bar{X} < 0.4$  or  $\bar{X} > 0.6$  to be 10%, we set the right hand side of the Bernoulli Law equal to 0.1 and solve for  $n$ :

$$\frac{p(1-p)}{n \varepsilon^2} = 0.1 = \frac{(0.5)^2}{n(0.1)^2} \quad \checkmark$$

$$\text{So } \frac{1}{n} = \frac{(0.1)^3}{(0.5)^2} = \frac{0.001}{0.25}, \text{ so } n \text{ must be } \checkmark$$

$$\geq 250. \quad \checkmark$$

b) For the normal approximation, we start with the central limit theorem which tells us that

$$\bar{X} \longrightarrow N(p, \frac{p(1-p)}{n}) = N(0.5, \frac{0.25}{n}) \quad \checkmark$$

And

$$\frac{\sqrt{n}(\bar{X} - 0.5)}{\sqrt{0.25}} \sim N(0, 1) \quad \checkmark$$

Then we set:

$$1a \left. \begin{matrix} a \\ b \end{matrix} \right\} 10/10$$

$$2a \left. \begin{matrix} a \\ b \end{matrix} \right\} 10/10$$

$$3a \left. \begin{matrix} a \\ b \end{matrix} \right\} 4/10$$

$$4a \left. \begin{matrix} a \\ b \\ c \end{matrix} \right\} \frac{8.75}{10}$$

$$\frac{32.75}{40}$$

$$P\left(\frac{\sqrt{n}(\bar{x}-0.5)}{\sqrt{0.25}} \leq \frac{\sqrt{n}(0.4-0.5)}{\sqrt{0.25}}\right) = 0.9 \quad \checkmark$$

Then solve for  $n$ :

$$\frac{\sqrt{n}(0.1)}{\sqrt{0.25}} = 1.645 \quad \checkmark$$

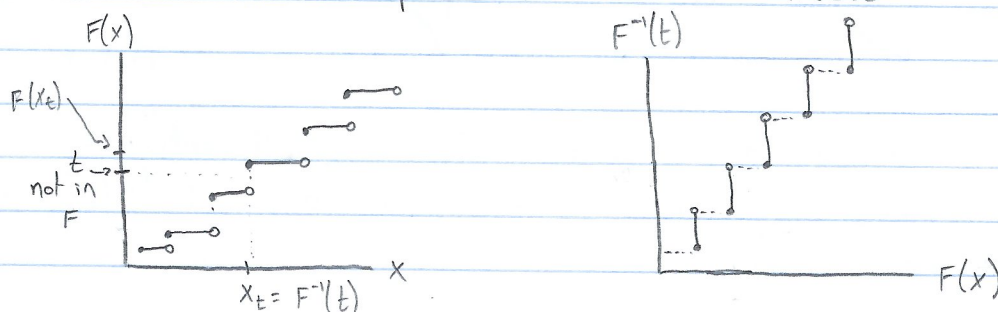
So with this approximation  $n \approx 68$ , which is much lower than the estimate based on Chebychev.

2. a) If we let  $Z(s) = s$ , then it's clear that  $Z_n$  converges in probability to  $Z$ . As  $n \rightarrow \infty$ , the interval of the indicator approaches 0, so the sequence converges to  $s = Z$ , because  $\lim_{n \rightarrow \infty} P(|Z_n - Z| \geq \epsilon) = 0$ .

b) Although the sequence converges in probability, it does not converge almost surely. There is no value of  $s$  such that  $Z_n(s) \rightarrow Z$ . For every possible value of  $s$ , the value of  $Z_n$  alternates between  $s$  and  $s+1$ . In other words,

$P(\lim_{n \rightarrow \infty} |Z_n - Z| < \epsilon) \neq 1$ . However, because the sequence converges in probability, this implies that a subsequence can be found that converges almost surely.

3. a) Let  $x_t = F^{-1}(t)$ . By the right continuity of  $F$ ,  $F(x_t) \geq t$  (proved in class). As a visual demonstration plot  $F(x)$  and its inverse:

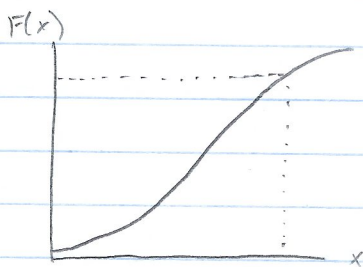


I realize the plot of  $F^{-1}(t)$  is bizarre notation, but in that plot solid lines indicate areas where the function is undefined. I did this to make the connection between the plots more obvious. If we pick a value of  $t$  such that  $F$  is discontinuous at that point, then mark the corresponding point  $x_t = F^{-1}(t)$  on the  $x$ -axis, it's clear that  $F(x_t) > t$ . If  $t$  is in the range of  $F$ , we simply have  $F(F^{-1}(t)) = F(x) = t$  by definition of the inverse CDF. But you still need to show equality iff  $t$  is in range of  $F$ . - see solution

it is defined at these points - its equal to the limit from the left because the CDF is nondecreasing and right continuous.

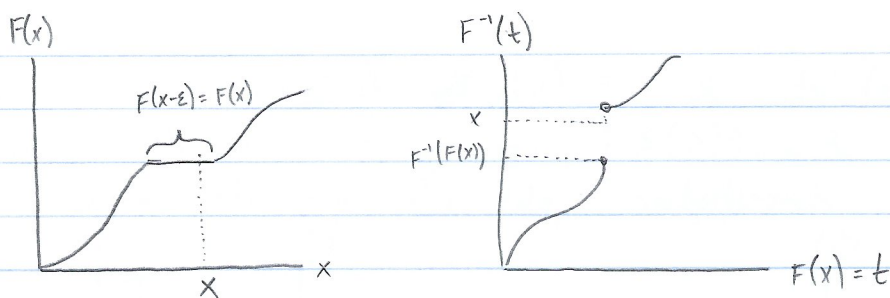
(-2)

b) This is similar to above, but the idea is that the inequality only applies when there are discontinuities. ✓  
If  $F(x)$  is defined for all  $x$  in  $(-\infty, \infty)$ , and strictly increasing, it naturally follows that  $F^{-1}(F(x)) = x$  for all  $x$  by the definition of the inverse CDF:



However, if there are any discontinuities or flat sections in the CDF, then the inverse CDF will be undefined there.  
→ equal to the limit from the left.





Then, because the inverse CDF is left-continuous, the inequality "kicks in" and we have  $F^{-1}(F(x)) \leq x$  (see plots above). This is the general idea, but you need to show it. -4  
 (see plots above). need to show it. see solution

BD 5.3.13. a) From the Delta method we know that if  $\sqrt{n}(g(Y_n) - g(\theta)) \rightarrow N(0, \sigma^2 g'(\theta)^2)$  if  $\leftarrow$  need to remark that  $g$  has to be continuous (which it is) -1/4  
 $\sqrt{n}(Y_n - \theta)$  converges to  $N(0, \sigma^2)$  in distribution.  
 So, we can rewrite  $\sqrt{s_n} - \sqrt{n}$  as:

$$\sqrt{s_n} - \sqrt{n} = \sqrt{n} \left( \frac{\sqrt{s_n}}{\sqrt{n}} - 1 \right) \quad \checkmark$$

In this case  $g(x) = \sqrt{x}$ , so by the Delta method:

$$\sqrt{n} \left( \frac{\sqrt{s_n}}{\sqrt{n}} - 1 \right) \rightarrow N\left(0, \overset{2}{\sigma^2} \left( \frac{1}{2\sqrt{\theta}} \right)^2\right) = N\left(0, \frac{1}{2}\right) \quad \text{(-1/2)}$$

b) From a) we know that  $\sqrt{n}(\frac{s_n}{n} - 1) \xrightarrow{P} N(0, \sigma^2)$ .  
 So we can set  $g(x) = \sqrt{2x}$ . So,

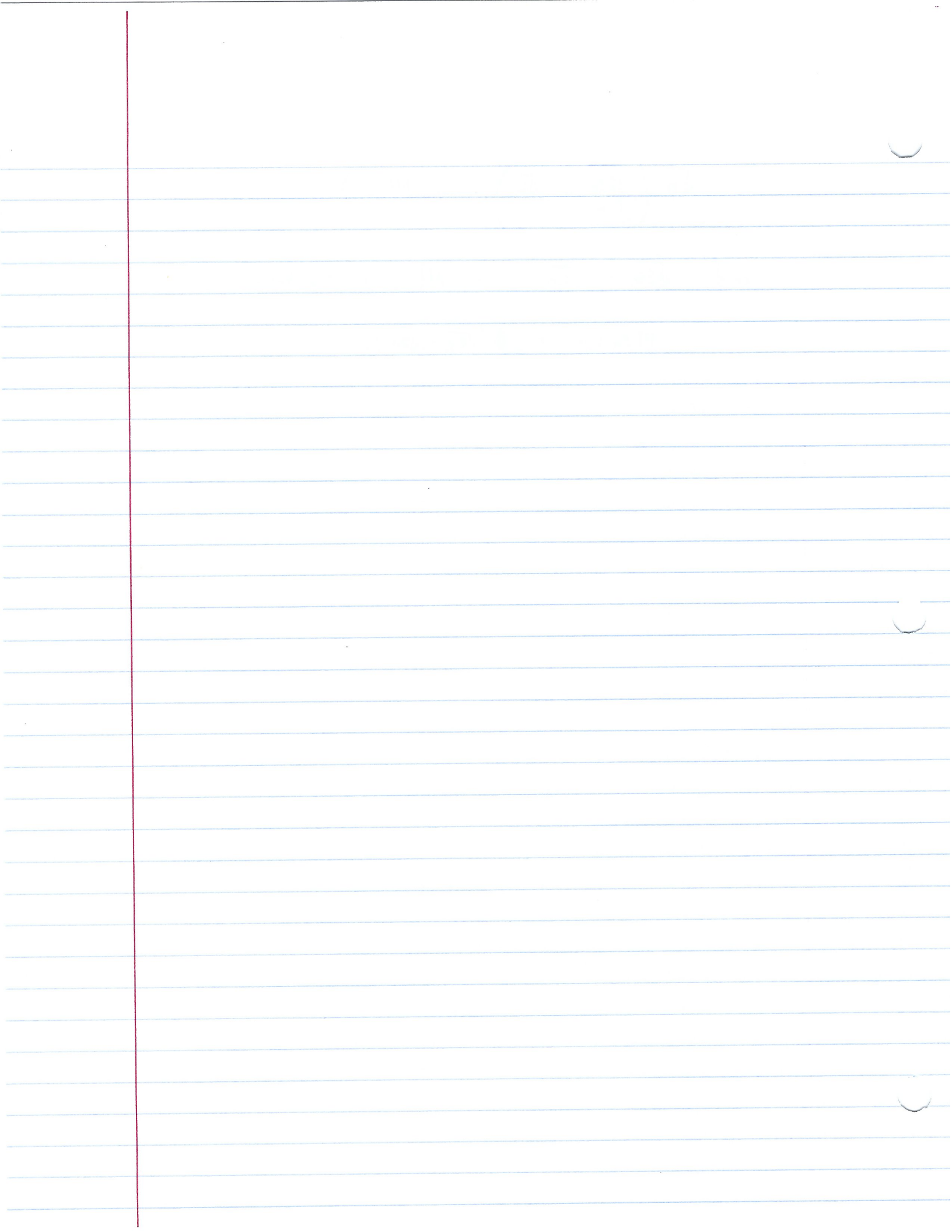
$$\begin{aligned} \sqrt{n} \left( \sqrt{\frac{2s_n}{n}} - \sqrt{2\theta} \right) &\rightarrow N\left(0, \sigma^2 g'(\theta)^2\right) \\ &\text{From (a) } \sqrt{s_n} - \sqrt{n} \rightarrow N\left(0, \frac{1}{2}\right) \quad \text{(-1/2)} \\ &\rightarrow N\left(0, \frac{\sigma^2}{\sqrt{2\theta}}\right) = N\left(0, \frac{1}{\sqrt{2}}\right) \\ &\Rightarrow \text{So } \sqrt{2}(\sqrt{s_n} - \sqrt{n}) \rightarrow N\left(0, \frac{1}{2}\right) \\ &= N(0, 1) \end{aligned}$$

Thus, we know that

$$\sqrt{n} \left( \sqrt{\frac{2s_n}{n}} - \sqrt{2} \right) \rightarrow N(0, 1)$$

and  $\sqrt{2s_n} - \sqrt{2n} \rightarrow N(0, 1)$ . Therefore,

$$P(S_n \leq x) \approx \Phi(\sqrt{2x} - \sqrt{2n}).$$



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## BD 5.3.13

c)

For each n (degrees of freedom), find the critical value. Plug the critical value and n into the two approximations from above and compare.

```
# Vector of n
n = c(5,10,25)
# Get critical values
q9 = qchisq(0.9,df=n)
q99 = qchisq(0.99,df=n)
# CLT approximations
clt9 = pnorm((q9-n)/sqrt(2*n))
clt99 = pnorm((q99-n)/sqrt(2*n))
# Part b
b9 = pnorm(sqrt(2*q9)-sqrt(2*n))
b99 = pnorm(sqrt(2*q99)-sqrt(2*n))
# Make tables
t9 = cbind(n,b9,clt9)
kable(t9,caption = "90th %ile",col.names = c("n","Part b)","CLT"),
      digits = 3)
```

Table 1: 90th %ile

n	Part b)	CLT
5	0.872	0.910
10	0.881	0.910
25	0.889	0.908

```
t99 = cbind(n,b99,clt99)
kable(t99,caption = "99th %ile",col.names = c("n","Part b)","CLT"),
      digits = 3)
```

Table 2: 99th %ile

n	Part b)	CLT
5	0.99	0.999
10	0.99	0.998
25	0.99	0.997

For the  $x_{0.90}$  case, the approximation from part b) slightly underestimates the probability while the CLT approach slightly overestimates. Both seem to perform well, however, and get very close to the correct value as  $n$  increases. In the  $x_{0.99}$  case, the approximation from part b) is correct for every value of  $n$ , while the CLT approximation is too large. The CLT approximation again improves as  $n$  increases, but I think the approximation from part b) is better overall. ✓