

## Set Functions

(events)

Collection  $\mathcal{B}$  of subsets also called

$\sigma$ -Field

## Remark

$A \subset \Omega$ , we repeat experiments  $N$  times

→ rel. frequency  $f_A = \frac{|\{A\}|}{N}$  ← # of times  $A$  occurred

→  $f_A \geq 0$  AND  $f_{\Omega} = 1$

Let  $A_1, A_2$  with  $A_1 \cap A_2 = \emptyset$ , then

$$f_{A_1 \cup A_2} = f_{A_1} + f_{A_2}$$

## Definition

$\Omega$ : Sample Space

$\mathcal{B}$ : set of events

$P$ : real-valued functions on  $\mathcal{B}$

$P$  must satisfy:

①  $P(A) \geq 0, \forall A \in \mathcal{B}$

②  $P(\Omega) = 1$

③ If  $\{A_n\}$  sequence of events in  $\mathcal{B}$  with

$A_m \cap A_n = \emptyset \quad \forall m \neq n$ , then :

$$P\left(\bigcup_{n=1}^{\infty} A_n\right) = \sum_{n=1}^{\infty} P(A_n)$$

A collection  $\mathcal{B} = \{A_n\}, n \in \mathbb{N}$  is **exhaustive** if

$$\bigcup_{i=1}^{\infty} A_i = \Omega, \text{ mutually exclusive AND exh.}$$

collections of events form a **partition** on  $\Omega$ .

### Theorem 1.3.1

$$A \in \mathcal{B}, P(A) = 1 - P(A^c)$$

Proof:

$$\Omega = A \cup A^c, A \cap A^c = \emptyset$$

$$\Rightarrow P(\Omega) = P(A) + P(A^c) = 1$$

$$\Leftrightarrow P(A) = 1 - P(A^c) \quad \square //$$

### Theorem 1.3.2

$$P(\emptyset) = 0$$

Proof:

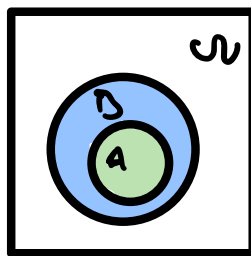
$$\text{Let } A = \emptyset, \text{ thus } A^c = \Omega$$

From  $P(\Omega) = P(A) + P(A^c) = 1$  follows, that

$$P(\Omega) = P(\emptyset) + P(\Omega)$$

$$1 = 1 + P(\emptyset)$$

$$\Leftrightarrow 1 - 1 = P(\emptyset) = 0 \quad \square //$$



### Theorem 1.3.3

$A, B$  events s.t.  $A \subset B$ , then  $P(A) \leq P(B)$

Proof

$$B = A \cup (A^c \cap B)$$

$$A = A \cap (A^c \cap B) = \emptyset \quad \text{---} \quad = \delta \geq 0$$

$$\Rightarrow P(B) = P(A) + P(A^c \cap B), \text{ thus}$$

$$P(B) = P(A) + \delta$$

$$\Rightarrow P(B) \geq P(A) \quad \square //$$

### Theorem 1.3.4

For each  $A \in \mathcal{B}$ ,  $0 \leq P(A) \leq 1$

Proof  $\emptyset \subset A \subset \Omega$  is trivial

$$\Rightarrow \emptyset \subset A \subset \Omega \Leftrightarrow P(\emptyset) \leq P(A) \leq P(\Omega)$$

$$\Leftrightarrow 0 \leq P(A) \leq 1 \quad \square //$$

### Theorem 1.3.5

Let  $A, B \in \mathcal{B}$ , then

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

Proof

$$A \cup B = A \cup (A^c \cap B)$$

$$B = (A \cap B) \cup (A^c \cap B)$$

$$\Rightarrow P(A \cup B) = P(A) + \underline{P(A^c \cap B)}, \quad (1)$$

$$\downarrow P(B) = P(A \cap B) + P(A^c \cap B) \quad (2)$$

$$(2) \quad P(A^c \cap B) = P(B) - P(A \cap B)$$

In (1)

$\Downarrow$

$$\Rightarrow P(A \cup B) = P(A) + P(B) - P(A \cap B) \quad \square //$$

Examples 1.3.1

$\Omega$ , two dies are thrown  $P((x,y)) = \frac{1}{36}$

$$C_1 = \{(1,1), (2,1), (3,1), (4,1), (5,1)\}$$

$$C_2 = \{(1,2), (2,2), (3,2)\}$$

$$P(C_1) = \frac{5}{36}, \quad P(C_2) = \frac{3}{36}, \quad P(C_1 \cup C_2) = \frac{8}{36}$$

$$\text{AND } P(C_1 \cap C_2) = P(\emptyset) = 0 //$$

1.3.2

Two coins tossed

$$\Omega = \{(H,H), (H,T), (T,H), (T,T)\}$$

$P$  assigns  $\frac{1}{4}$  to each outcome.

$$C_1 = \{(H, H), (H, T)\}$$

$$C_2 = \{(H, H), (T, H)\}$$

$$P(C_1) = P(C_2) = \frac{1}{2}$$

$$P(C_1 \cap C_2) = \frac{1}{4}$$

$$P(C_1 \cup C_2) = \frac{1}{2} + \frac{1}{2} - \frac{1}{4} = \frac{3}{4} //$$

Finite sample space

$$\Omega = \{x_1, \dots, x_m\}$$

$$0 \leq p_i \leq 1, i \in \{1, 2, \dots, m\} \text{ AND } \sum_{i=1}^m p_i = 1$$

$$P(A) = \sum_{x_i \in A} p_i$$

Definition 1.3.2

Equilikely case,  $\Omega = \{x_1, \dots, x_m\}$

$$P(A) = \sum_{x_i \in A} \frac{1}{m} = \frac{|A|}{m}$$

(for "fair" experiments)

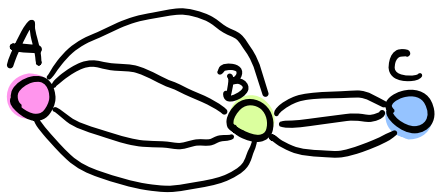
# Counting Rules 1.3.1

## mn-rule

$$A = \{x_1, \dots, x_m\}$$

$$B = \{y_1, \dots, y_n\}$$

There exist  $m \cdot n$  ordered pairs



4 ways  $A \rightarrow B$

3 ways  $B \rightarrow C$

By mn-rule:  $4 \cdot 3 = 12$  ways  $A \rightarrow C$

## Example

Drivers licence = 3 capital letters & 3 Numbers

gives  $26^3 \cdot 10^3$  possibilities.

$k$ -tuples of  $A$ :

$$n \cdot n \cdot \dots \cdot n = n^k$$

$k$ -tuples (no repetition)

$$\overset{1st}{n} \cdot \overset{2nd}{(n-1)} \cdot \dots \cdot \overset{kth}{(n-(k-1))}$$

these tuples: PERMUTATIONS

$\rightarrow P_k^n$  (# of  $k$ -perm's From an  $n$  Element set)

$$P_k^n = n \cdot (n-1) \cdot \dots \cdot (n-(k-1))$$

$$= \frac{n!}{(n-k)!}$$

Example (Birthday Problem)

$n$  people,  $n < 365$

$A =$  <sup>at least</sup> two people share birthday

$365^n$  possible  $n$ -tuples with

a probability of  $365^{-n} = \frac{1}{365^n}$

$A^c =$  all Birthdays are distinct

+ number of  $n$ -tuples in  $A^c$ , which is

$P_{n, 365}^{365}$

Thus,  $P(A) = 1 - \frac{P_n^{365}}{365^n}$

If  $n=2$ ,  $P(A) = 1 - \frac{365 \cdot 364}{365^2} = 0,0027$   
 $= 0,27\%$  //

Now suppose order is irrelevant

→ DONT COUNT PERMUTATIONS

→ COUNT NUMBER OF SUBSETS  
(of  $k$  elements)

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$

also:  $\binom{n}{k} = C_k^n$

$$(a+b)^n = \sum_{k=0}^n \binom{n}{k} a^k b^{n-k}$$

Example (Poker)

$|D| = 52$  we draw spades card

$$E_1 = \spadesuit$$

$$P(E_1) = \frac{13}{52} = \frac{1}{4},$$

$$E_2 = \text{KING}$$

$$P(E_2) = \frac{4}{52} = \frac{1}{13},$$

A "hand" consist of 5 cards, thus

$\binom{52}{5}$  "hands" are playable.

$$\text{Prob. of a hand: } P(H;) = \frac{1}{\binom{52}{5}}$$

$$E_1 = \text{FLUSH}$$

$$\binom{4}{1} = 4 \text{ Suits AND}$$

$$\binom{13}{5} = 1287 \text{ hands for each suit}$$



Thus, we obtain:

$$P(E_1) = \frac{\binom{4}{1} \binom{13}{5}}{\binom{52}{5}} = \frac{4 \cdot 1287}{2598960} = 0,00198$$

(straight flush prob. is included here of course)

$E_2$  = THREE OF A KIND (only 3 of a kind)

• The "kind" can have  $\binom{13}{1} = 13$  kinds

• The "three" has  $\binom{4}{3}$  ways.

(From ♠3 ♣3 ♦3 ♥3) ← we take 3

• The other "two cards"  $\binom{12}{2}$  and one from

each of the last two kinds:  $\binom{4}{1} \binom{4}{1}$  ways

$$4 \{ \spadesuit \clubsuit \heartsuit \} \times \{ \spadesuit \clubsuit \heartsuit \}$$

Thus  $P(E_2)$  : (exact 3 of a kind)

$$P(E_2) = \frac{\binom{13}{1} \binom{4}{3} \binom{12}{2} \binom{4}{1}^2}{\binom{52}{5}} = 0,0211$$

$\approx 2,11\%$

$E_3$  = exactly 3 KINGS and 2 QUEENS

$$P(E_3) = \frac{\binom{4}{3} \binom{4}{2}}{\binom{52}{5}} = 0,0000093$$

# Additional Properties of Probability

## Theorem 1.3.6

$\{C_n\}$  nondecreasing event sequence :

$$\lim_{n \rightarrow \infty} P(C_n) = P(\lim_{n \rightarrow \infty} C_n) = P\left(\bigcup_{n=1}^{\infty} C_n\right)$$

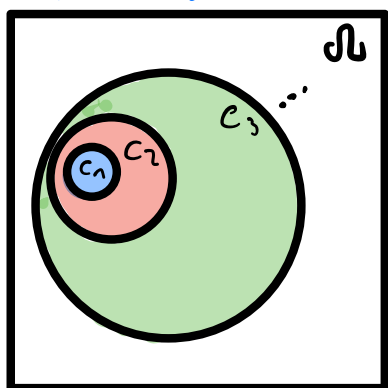
$\{C_n\}$  increasing :

$$\lim_{n \rightarrow \infty} P(C_n) = P(\lim_{n \rightarrow \infty} C_n) = P\left(\bigcap_{n=1}^{\infty} C_n\right)$$

## Proof

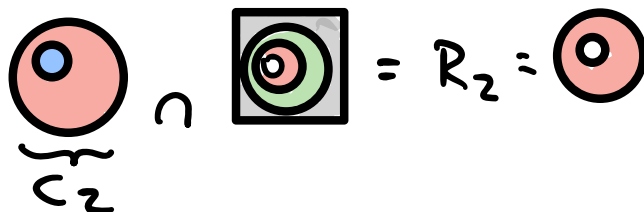
Sets are called rings. Define :

$$R_1 = C_1, n > 1 \quad R_n = C_n \cap C_{n-1}^c$$

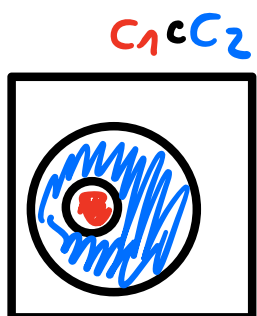


$$C_1 = R_1$$

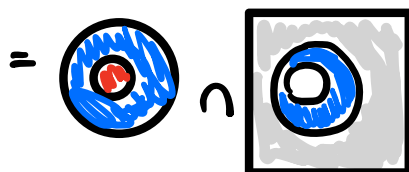
$$R_2 = C_2 \cap C_1^c$$



$$\Rightarrow \bigcup_{n=1}^{\infty} C_n = \bigcup_{n=1}^{\infty} R_n \quad \wedge \quad R_m \cap R_n = \emptyset, m \neq n$$



$$R_2 = C_2 \cap C_1^c$$






$$= \text{Ring } R_2$$


$$\Rightarrow R_1 \not\subset R_2$$

Thus  $R_m \cap R_n = \emptyset$ , also

$$P(R_n) = P(C_n) - P(C_{n-1}) \quad (*)$$

$$\downarrow$$

$$P\left(\overset{C_n}{\text{Ring } R_n} \setminus \overset{C_{n-1}}{\text{Ring } R_{n-1}}\right) = P(\text{Ring } R_n) - P(\text{Ring } R_{n-1})$$




This yields:

$$\begin{aligned} P(\lim_{n \rightarrow \infty} C_n) &= P\left(\bigcup_{n=1}^{\infty} C_n\right) \\ &= P\left(\bigcup_{n=1}^{\infty} R_n\right) \\ &= \sum_{n=1}^{\infty} P(R_n) \\ &= \lim_{n \rightarrow \infty} \sum_{j=1}^n P(R_j) \\ &\quad \downarrow (*) \end{aligned}$$

$$= \lim_{n \rightarrow \infty} \left( P(C_1) + \sum_{j=2}^n P(C_j) - P(C_{j-1}) \right)$$

$$= \lim_{n \rightarrow \infty} P(C_n) \quad \square //$$

### Theorem 1.3.7 (Boole's Inequality)

$\{C_n\}$  sequence of events :

$$P\left(\bigcup_{n=1}^{\infty} C_n\right) \leq \sum_{n=1}^{\infty} P(C_n)$$

Proof

$$\text{Let } D_n = \bigcup_{i=1}^n C_i \quad (*)$$

$\Rightarrow \{D_n\}$  is increasing sequence

upto  $\bigcup_{n=1}^{\infty} C_n$ .

Also:

$$\forall j: D_j = D_{j-1} \cup C_j \quad (*)$$

Hence:

$$P(D_j) \leq P(D_{j-1}) + P(C_j)$$

$$\Rightarrow P(D_j) - P(D_{j-1}) \leq P(C_j)$$

Using  $P(C_1) = P(D_1)$  gives

$$P\left(\bigcup_{n=1}^{\infty} C_n\right) = P\left(\bigcup_{n=1}^{\infty} D_n\right)$$

$$= \lim_{n \rightarrow \infty} \left( P(D_1) + \sum_{j=2}^n P(D_j) - P(D_{j-1}) \right)$$

$$\leq \lim_{n \rightarrow \infty} \sum_{j=1}^n P(C_j) = \sum_{n=1}^{\infty} P(C_n) \quad \square //$$

### Remark 1.3.2 (Inclusion - Exclusion)

$$P(C_1 \cup C_2 \cup C_3) = P_1 - P_2 + P_3$$

$$P_1 = P(C_1) + P(C_2) + P(C_3)$$

$$P_2 = P(C_1 \cap C_2) + P(C_1 \cap C_3) + P(C_2 \cap C_3)$$

$$P_3 = P(C_1 \cap C_2 \cap C_3)$$

Generalized:

$$P(C_1 \cup \dots \cup C_k) = P_1 - P_2 + \dots + (-1)^{k+1} P_k$$

with  $P_i$  being the sum of the  $P$ 's of all possible intersections involving  $i$  sets.

It holds  $P_1 \geq \dots \geq P_k$

$$P_1 = P(C_1) + \dots + P(C_k) \geq P(C_1 \cup \dots \cup C_k)$$

1.3.7

For  $k=2$

$$1 \geq P(C_1 \cup C_2) = P(C_1) + P(C_2) - P(C_1 \cap C_2)$$

Donferroni's inequality  $\swarrow$  this gives

$$P(C_1 \cap C_2) \geq P(C_1) + P(C_2) - 1$$

Other useful inequalities come from this:

$$P_1 \geq P(C_1 \cup \dots \cup C_k) \geq P_1 - P_2$$

AND

$$P_1 - P_2 + P_3 \geq P(C_1 \cup \dots \cup C_k) \geq P_1 - P_2 + P_3 - P_4$$