

108 pts
 Reading: Chapter 2 of *Elements of Information Theory*

Lecture 1B: Entropy

1. (10 pts) *Coin flips.*

- (a) (4 pts) The number X of tosses till the first head appears has the geometric distribution with parameter $p = 1/2$, where $P(X = n) = pq^{n-1}$, $n \in \{1, 2, \dots\}$. Hence the entropy of X is

$$\begin{aligned}
 H(X) &= - \sum_{n=1}^{\infty} pq^{n-1} \log(pq^{n-1}) \\
 &= - \left[\sum_{n=0}^{\infty} pq^n \log p + \sum_{n=0}^{\infty} npq^n \log q \right] \\
 &= \frac{-p \log p}{1-q} - \frac{pq \log q}{p^2} \\
 &= \frac{-p \log p - q \log q}{p} \\
 &= H(p)/p \text{ bits.}
 \end{aligned}$$

- (b) (2 pts) If $p = 1/2$, then $H(X) = 2$ bits.

- (c) (4 pts) Intuitively, it seems clear that the best questions are those that have equally likely chances of receiving a yes or a no answer. Consequently, one possible guess is that the most “efficient” series of questions is: Is $X = 1$? If not, is $X = 2$? If not, is $X = 3$? ... with a resulting expected number of questions equal to $\sum_{n=1}^{\infty} n(1/2^n) = 2$. This should reinforce the intuition that $H(X)$ is a measure of the uncertainty of X . Indeed in this case, the entropy is exactly the same as the average number of questions needed to define X , and in general $E(\# \text{ of questions}) \geq H(X)$. This problem has an interpretation as a source coding problem. Let 0=no, 1=yes, X =Source, and Y =Encoded Source. Then the set of questions in the above procedure can be written as a collection of (X, Y) pairs: $(1, 1)$, $(2, 01)$, $(3, 001)$, etc. . In fact, this intuitively derived code is the optimal (Huffman) code minimizing the expected number of questions.

2. (4 pts) We wish to find *all* probability vectors $\mathbf{p} = (p_1, p_2, \dots, p_n)$ which minimize

$$H(\mathbf{p}) = - \sum_i p_i \log p_i.$$

Now $-p_i \log p_i \geq 0$, with equality iff $p_i = 0$ or 1. Hence the only possible probability vectors which minimize $H(\mathbf{p})$ are those with $p_i = 1$ for some i and $p_j = 0, j \neq i$. There are n such vectors, i.e., $(1, 0, \dots, 0), (0, 1, 0, \dots, 0), \dots, (0, \dots, 0, 1)$, and the minimum value of $H(\mathbf{p})$ is 0. These points are the corners of the simplex.

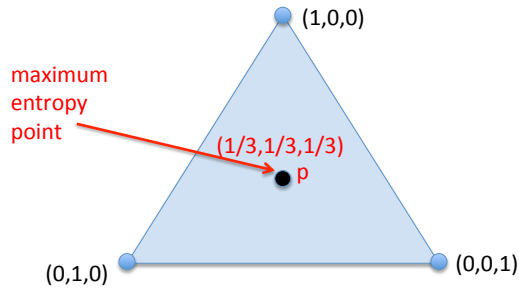


Figure 1: Illustration of minimum-entropy points (corners) for $n = 3$ simplex.

3. (12 pts) *Entropy of functions of a random variable.*

- (a) $H(X, g(X)) = H(X) + H(g(X)|X)$ by the chain rule for entropies.
- (b) $H(g(X)|X) = 0$ since for any particular value of X , $g(X)$ is fixed, and hence $H(g(X)|X) = \sum_x p(x) H(g(X)|X = x) = \sum_x 0 = 0$.
- (c) $H(X, g(X)) = H(g(X)) + H(X|g(X))$ again by the chain rule.
- (d) $H(X|g(X)) \geq 0$, with equality iff X is a function of $g(X)$, i.e., $g(\cdot)$ is one-to-one. Hence $H(X, g(X)) \geq H(g(X))$.

Combining parts (b) and (d), we obtain $H(X) \geq H(g(X))$.

Lecture 1C: Relative Entropy

4. (4 pts) *Computing Relative Entropy for 2D p and q .*

Let $p(x, y)$ be given by

X/\mathcal{Y}	0	1
0	$\frac{1}{6}$	$\frac{7}{12}$
1	$\frac{1}{6}$	$\frac{1}{12}$

Let $q(x, y)$ be given by

X/\mathcal{Y}	0	1
0	$\frac{1}{4}$	$\frac{1}{2}$
1	$\frac{1}{12}$	$\frac{1}{6}$

Find $D(p||q)$.

Solution:

$$D(p||q) = \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} p(x, y) \log \left(\frac{p(x, y)}{q(x, y)} \right) \quad (1)$$

$$= \frac{1}{6} \log \frac{4}{6} + \frac{1}{6} \log \frac{12}{6} + \frac{7}{12} \log \frac{14}{12} + \frac{1}{12} \log \frac{6}{12} \quad (2)$$

$$= \frac{1}{6} \log \frac{2}{3} + \frac{1}{6} + \frac{7}{12} \log \frac{7}{6} - \frac{1}{12} \quad (3)$$

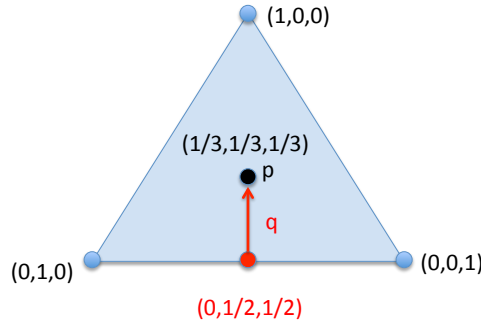
$$= 0.1156 \quad (4)$$

5. (16 pts) *Computing Relative Entropy for p and q on a line in the 3D Simplex.*

Let $p(x)$ and $q(x)$ be three-outcome PMFs with the possible outcomes $\mathcal{X} = \{a, b, c\}$ so that p and q lie on the 3D simplex which is a 2D triangle in 3D space. Furthermore, let the PMF for p be the point $(\frac{1}{3}, \frac{1}{3}, \frac{1}{3})$ and the PMF for q_λ be the point $(\frac{\lambda}{3}, \frac{3-\lambda}{6}, \frac{3-\lambda}{6})$ in the simplex.

- (a) (4 pts) As in lecture, draw a triangle representing the simplex, show the point p and the line segment that shows the trajectory of q as λ varies between 0 and 1.

Solution:



- (b) (4 pts) Find $D(p||q_\lambda)$ as a function of λ as λ varies between 0 and 1 and use MATLAB to make a nice plot of $D(p||q_\lambda)$ vs. λ . You may not be able to plot $D(p||q_\lambda)$ in MATLAB for values of λ near zero, but please evaluate (compute) what the value should be at $\lambda = 0$ (possibly infinity or a finite value).

Solution:

$$D(p||q_\lambda) = \sum_{x \in \mathcal{X}} p(x) \log \frac{p(x)}{q_\lambda(x)} \quad (5)$$

$$= \frac{1}{3} \log \frac{1}{\lambda} + \frac{2}{3} \log \frac{2}{3-\lambda} \quad (6)$$

See the curve for $D(p||q_\lambda) = \infty$ in Fig. ???. At $\lambda = 0$, $D(p||q_\lambda) = \infty$.

- (c) (4 pts) Find $D(q_\lambda||p)$ as a function of λ as λ varies between 0 and 1 and use MATLAB to make a nice plot of $D(q_\lambda||p)$ vs. λ . Include your plot from the previous part for comparison. You may not be able to plot $D(q_\lambda||p)$ for values of λ near zero, but please evaluate (compute) what the value should be at $\lambda = 0$ (possibly infinity or a finite value).

Solution:

$$D(q_\lambda||p) = \sum_{x \in \mathcal{X}} q_\lambda(x) \log \frac{q_\lambda(x)}{p(x)} \quad (7)$$

$$= \frac{\lambda}{3} \log \lambda + \frac{3-\lambda}{3} \log \frac{3-\lambda}{2} \quad (8)$$

See the curve for $D(q_\lambda||p)$ in Fig. ???. Note that $0 \log 0 = 0$ so at $\lambda = 0$, $D(p||q_\lambda) = \log \frac{3}{2} = 0.585$.

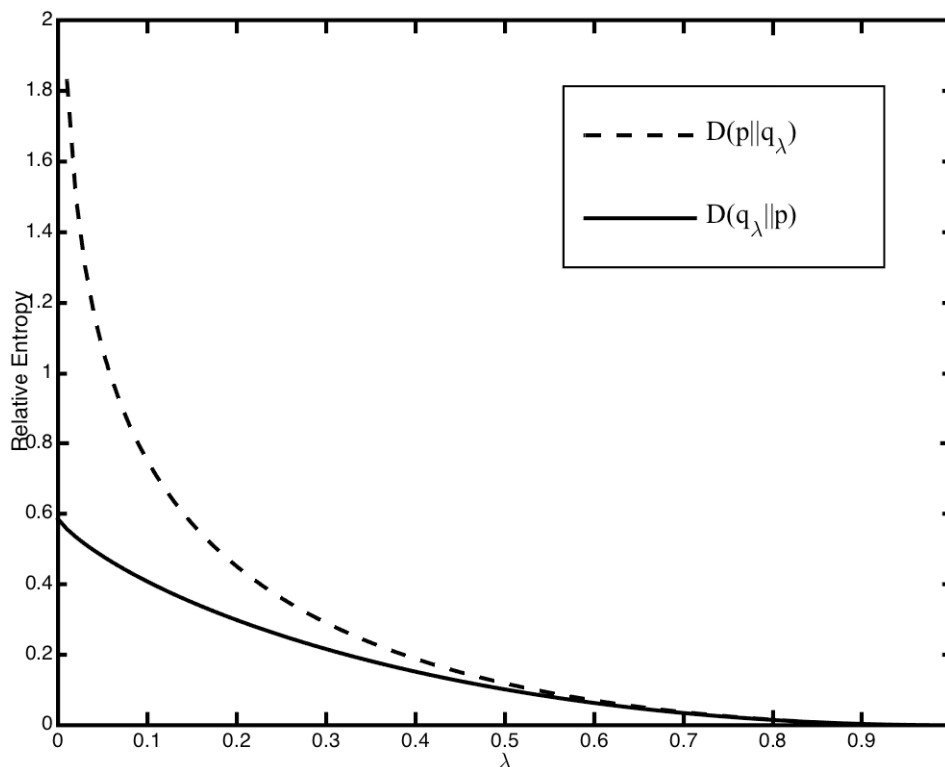


Figure 2: $D(p||q_\lambda)$ and $D(q_\lambda||p)$ vs. λ .

- (d) (4 pts) Discuss the differences between $D(p||q_\lambda)$ and $D(q_\lambda||p)$. We learned that one interpretation of $D(p||q)$ is that it is a penalty for using the wrong distribution for determining description length. How come this penalty is infinitely larger at $\lambda = 0$ in one case as compared to the other?

Solution: Looking at the curves in Fig. ?? we can see that certainly $D(p||q_\lambda) \neq D(q_\lambda||p)$, but the differences are negligible for λ near one and become infinite as λ approaches zero. Lets consider the $\lambda = 0$ points from the perspective of $D(p||q_\lambda)$ and $D(q_\lambda||p)$ being measures of the penalty of using the distribution on the right of the $||$ to determine description length when the distribution on the left of the

$||$ is the actual distribution producing the symbols to be compressed. In the q_λ distribution, the probability of $x = a$ goes to zero as $\lambda \rightarrow 0$. When q_λ is the true distribution this means that the penalty in compression only applies to $x = b$ and $x = c$ since $x = a$ happens with probability zero. The penalty is that we use the description length $-\log \frac{1}{3}$ which is longer than the description length $-\log \frac{1}{2}$ we should have used for $x = b$ and $x = c$. This is a finite penalty.

Now consider the case where p is the true distribution. The three outcomes $x = a$, $x = b$, and $x = c$ all happen with equal probability of $\frac{1}{3}$ but in the q_λ distribution, the probability of $x = a$ goes to zero as $\lambda \rightarrow 0$, which means that the appropriate description length for $x = a$ goes to infinity as $\lambda \rightarrow 0$. We end up using an infinitely long description length a third of the time which leads to an infinite description length and hence an infinite “penalty” $D(q_\lambda || p)$.

Lecture 1D: Mutual Information

6. (4 pts) *Mutual Information?*

Can the relative entropy computed in the previous problem be expressed as a mutual information? Explain fully.

Solution: Yes. In fact the relative entropy computed above is exactly $I(X; Y)$ since it turns out that $q(x, y) = p(x)p(y)$.

7. (12 pts) *Example of joint entropy*

- (a) $H(X) = \frac{2}{3} \log \frac{3}{2} + \frac{1}{3} \log 3 = 0.918 \text{ bits} = H(Y)$.
- (b) $H(X|Y) = \frac{1}{3} H(X|Y=0) + \frac{2}{3} H(X|Y=1) = 0.667 \text{ bits} = H(Y|X)$.
- (c) $H(X, Y) = 3 \times \frac{1}{3} \log 3 = 1.585 \text{ bits}$.
- (d) $H(Y) - H(Y|X) = 0.251 \text{ bits}$.
- (e) $I(X; Y) = H(Y) - H(Y|X) = 0.251 \text{ bits}$.
- (f) See Figure 2.2 in *Elements of Information Theory*.

8. (8 pts) *Mutual Information and the Weather*

(a)

$$I(P_s; W) = H(P_s) - H(P_s|W) = 0 - 0 = 0 \quad (9)$$

(b)

$$I(P_w; W) = H(W) - H(W|P_w) = H(.9) - .75 \times H(0) - .25 \times H(.4) = 0.2262 \quad (10)$$

- (c) Wendy, provides the most information. In fact, Stormy provides no information at all.
- (d) Plant your tulip bulbs when Wendy forecasts rain.

Lecture 2A: Convexity

9. (6 pts) *Concavity of entropy*

(a) Show that $\log x$ is concave in x for positive x . **Solution:**

$$\frac{d^2}{dx^2} \log_2 x = \frac{d^2}{dx^2} (\log_2 e) \ln x \quad (11)$$

$$= (\log_2 e) \frac{d}{dx} x^{-1} \quad (12)$$

$$= -(\log_2 e) x^{-2}, \quad (13)$$

which is negative for positive x so $\log x$ is concave in x for positive x .

(b) Show that $x \log x$ is convex in x for positive x . **Solution:**

$$\frac{d^2}{dx^2} x \log_2 x = \frac{d^2}{dx^2} (\log_2 e) x \ln x \quad (14)$$

$$= (\log_2 e) \frac{d}{dx} (\ln x + 1) \quad (15)$$

$$= (\log_2 e) x^{-1}, \quad (16)$$

which is positive for positive x so $x \log x$ is convex in x for positive x .

(c) Use the second derivative to show that $H(p) = -p \log p - (1-p) \log(1-p)$ is concave in p for $0 \leq p \leq 1$.

Solution: While one can essentially refer to slide 44 of lecture 2 as follows: $H(p) = \log |\mathcal{X}| - D(p||u)$ so the convexity of relative entropy implies the concavity of entropy, this exercise required that you show the concavity by differentiation as in the other two parts.

$$\frac{d^2}{dp^2} -p \log p - (1-p) \log(1-p) = -(\log_2 e) (p^{-1} + (1-p)^{-1}), \quad (17)$$

which is negative for $0 \leq p \leq 1$ so $H(p)$ is concave in p for $0 \leq p \leq 1$.

Lecture 2B: Jensen's Inequality and its Applications

10. (4 pts) *Maximum entropy.* What is the maximum value of $H(p_1, \dots, p_n) = H(\mathbf{p})$ as \mathbf{p} ranges over the set of n -dimensional probability vectors? Find all \mathbf{p} 's which achieve this maximum.

Solution: Slides 16-17 of Lecture 2 showed that entropy is upper bounded by $H(X) \leq \log |\mathcal{X}|$ because $H(p) = \log |\mathcal{X}| - D(p||u)$. We can achieve $H(X) = \log |\mathcal{X}|$ with a uniform distribution (i.e. all probabilities equal to $|\mathcal{X}|^{-1}$). This is the only distribution that achieves the maximum value since any other distribution will have a nonzero $D(p||u)$.

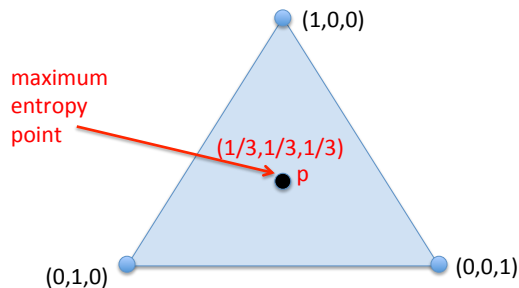


Figure 3: Illustration of maximum entropy point for $n = 3$ simplex.

11. (8 pts) *Drawing with and without replacement.* Intuitively, it is clear that if the balls are drawn with replacement, the number of possible choices for the i -th ball is larger, and therefore the conditional entropy is larger. But computing the conditional distributions is slightly involved. It is easier to compute the unconditional entropy.
- With replacement. In this case the conditional distribution of each draw is the same for every draw. Thus

$$X_i = \begin{cases} \text{red} & \text{with prob. } \frac{r}{r+w+b} \\ \text{white} & \text{with prob. } \frac{w}{r+w+b} \\ \text{black} & \text{with prob. } \frac{b}{r+w+b} \end{cases} \quad (18)$$

and therefore

$$H(X_i | X_{i-1}, \dots, X_1) = H(X_i) \quad (19)$$

$$= \log(r+w+b) - \frac{r}{r+w+b} \log r - \frac{w}{r+w+b} \log w - \frac{b}{r+w+b} \log b \quad (20)$$

- Without replacement. The unconditional probability of the i -th ball being red is still $r/(r+w+b)$, etc. Thus the unconditional entropy $H(X_i)$ is still the same as with replacement. The conditional entropy $H(X_i | X_{i-1}, \dots, X_1)$ is less than the unconditional entropy (We showed that as an application of Jensen's inequality.), and therefore the entropy of drawing without replacement is lower.

Lecture 2C: Markov Chains and the Data Processing Inequality

12. (10 pts) *Conditional Mutual Information.*

- (a) (5 pts) Show that if $X \rightarrow Y \rightarrow Z$ forms a Markov chain, $I(X; Y | Z) \leq I(X; Y)$.

$$I(X; Y, Z) = I(X; Z) + I(X; Y | Z) \quad (21)$$

$$= I(X; Y) + I(X; Z | Y) \quad (22)$$

For a Markov chain $X \rightarrow Y \rightarrow Z$, $I(X; Z|Y) = 0$. Thus

$$I(X; Y|Z) = I(X; Y) - I(X; Z) \quad (23)$$

$$\leq I(X; Y). \quad (24)$$

- (b) (5 pts) Is it always true that $I(X; Y|Z) \leq I(X; Y)$ (i.e even for every case where $X \rightarrow Y \rightarrow Z$ does not form a Markov chain? No. Consider this example, which is also given in the text. Let X, Y be independent fair binary random variables and let $Z = X + Y$. In this case we have that,

$$I(X; Y) = 0$$

and,

$$I(X; Y | Z) = H(X | Z) = 1/2.$$

So $I(X; Y) < I(X; Y | Z)$. Note that in this case X, Y, Z are not Markov.

13. (10 pts) *Find the gap.*

You know that for $X \rightarrow Y \rightarrow Z$, $I(X; Z) \leq I(Y; Z)$. Find the exact value of the gap between these mutual informations. i.e. Find $I(Y; Z) - I(X; Z)$ for the Markov chain $X \rightarrow Y \rightarrow Z$.

For full credit your answer must be a single information theoretic expression such as an entropy, a mutual information, or a conditional mutual information.

The answer is $I(Y; Z) - I(X; Z) = I(Y; Z|X)$ for a Markov chain.

Following the proof of Theorem 2.8.1 on pages 32-33, one approach is to write a mutual information two ways via the chain rule for mutual information:

$$I(X, Y; Z) = I(X; Z) + I(Y; Z|X) \quad (25)$$

$$= I(Y; Z) + I(X; Z|Y). \quad (26)$$

Realizing that $I(X; Z|Y) = 0$ for Markov chains completes the proof.

Another technique is the following:

$$I(X; Z) = H(Z) - H(Z|X) \quad (27)$$

$$= H(Z) - H(Z|X) + H(Z|Y) - H(Z|Y) \quad (28)$$

$$= I(Y; Z) - \left(H(Z|X) - H(Z|Y) \right) \quad (29)$$

$$= I(Y; Z) - \left(H(Z|X) - H(Z|Y, X) \right) \quad \text{Since } X \rightarrow Y \rightarrow Z. \quad (30)$$

$$= I(Y; Z) - I(Y; Z|X) \quad (31)$$