

ECN627 Assignment 1

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Problem 2.6

a)

$$\begin{aligned} E[Y] &= \sum_i y_i \Pr(Y = y_i) \\ &= 0 \times 0.068 + 1 \times 0.932 \\ &= 0.932 \end{aligned}$$

b)

$$\begin{aligned} E[\bar{Y}] &= \sum_i y_i (1 - \Pr(Y = y_i)) \\ &= 0 \times (1 - 0.068) + 1 \times (1 - 0.932) \\ &= 0.068 \\ &= 1 - E[Y] \end{aligned}$$

c)

$$\begin{aligned} E[Y|X=1] &= \sum_i y_i \Pr(Y = y_i|X=1) \\ &= 0 \times \Pr(Y=0|X=1) + 1 \times \Pr(Y=1|X=1) \\ &= 0 \times \frac{\Pr(Y=0 \cap X=1)}{\Pr(X=1)} + 1 \times \frac{\Pr(Y=1 \cap X=1)}{\Pr(X=1)} \\ &= 0 \times \frac{0.015}{0.361} + 1 \times \frac{0.346}{0.361} \\ &\approx 0.958 \end{aligned}$$

$$\begin{aligned} E[Y|X=0] &= \sum_i y_i \Pr(Y = y_i|X=0) \\ &= 0 \times \Pr(Y=0|X=0) + 1 \times \Pr(Y=1|X=0) \\ &= 0 \times \frac{\Pr(Y=0 \cap X=0)}{\Pr(X=0)} + 1 \times \frac{\Pr(Y=1 \cap X=0)}{\Pr(X=0)} \\ &= 0 \times \frac{0.053}{0.639} + 1 \times \frac{0.586}{0.639} \\ &\approx 0.917 \end{aligned}$$

d)

$$\Pr(Y=0 \cap X=1) = 0.015$$

$$\Pr(Y=0 \cap X=0) = 0.053$$

e)

$$\begin{aligned} \Pr(Y=0|X=1) &= \frac{\Pr(Y=0 \cap X=1)}{\Pr(X=1)} \\ &= \frac{0.015}{0.361} \\ &\approx 0.041 \end{aligned}$$

$$\begin{aligned} \Pr(Y=0|X=0) &= \frac{\Pr(Y=0 \cap X=0)}{\Pr(X=0)} \\ &= \frac{0.053}{0.639} \\ &\approx 0.083 \end{aligned}$$

f)

$$\begin{aligned}\Pr_Y(Y = 0) &= 0.068 & \Pr(Y = 0|X = 0) &\approx 0.083 & \Pr(Y = 0|X = 1) &\approx 0.042 \\ \Pr_Y(Y = 1) &= 0.932 & \Pr(Y = 1|X = 0) &\approx 0.917 & \Pr(Y = 1|X = 1) &\approx 0.984\end{aligned}$$

$$\begin{aligned}\Pr_X(X = 0) &= 0.639 & \Pr(X = 0|Y = 0) &\approx 0.779 & \Pr(X = 0|Y = 1) &\approx 0.629 \\ \Pr_X(X = 1) &= 0.361 & \Pr(X = 1|Y = 0) &\approx 0.221 & \Pr(X = 1|Y = 1) &\approx 0.371\end{aligned}$$

Since $\Pr_{X|Y}(X|Y) \neq \Pr_X(x), \Pr_X(x) > 0$ and $\Pr_{Y|X}(Y|X) \neq \Pr_Y(y), \Pr_Y(y) > 0 \forall x$ and y , employment status and educational achievement NOT independent.

Problem 2.8

$$\begin{aligned}\mu_Y &= 1 \\ \sigma_y^2 &= 4 \\ Z &= \frac{1}{2}(Y - 1) \\ &= \frac{1}{2}(Y) - \frac{1}{2}\end{aligned}$$

Hence,

$$\begin{aligned}E[Z] &= \frac{1}{2}\mu_Y - \frac{1}{2} \\ &= \frac{1}{2}(1) - \frac{1}{2} \\ &= 0\end{aligned}$$

and,

$$\begin{aligned}Var[Z] &= \left(\frac{1}{2}\right)^2 Var[Y] \\ &= \frac{1}{4}(4) \\ &= 1\end{aligned}$$

Problem 2.10

(a)

$$Y \sim \mathcal{N}(1, 4)$$

$$\Pr(Y \leq y) = \Pr\left(Z \leq \frac{y - \mu_Y}{\sigma_Y}\right)$$

$$\Pr(Y \leq 3) = \Pr\left(Z \leq \frac{3 - 1}{2}\right)$$

$$= \Pr(Z \leq 1)$$

$$= \Phi(1)$$

$$= 0.84134$$

(b)

$$Y \sim \mathcal{N}(3, 9)$$

$$\Pr(Y > y) = 1 - \Pr\left(Z \leq \frac{y - \mu_Y}{\sigma_Y}\right)$$

$$\Pr(Y > 0) = 1 - \Pr\left(Z \leq \frac{0 - 3}{3}\right)$$

$$= 1 - \Pr(Z \leq -1)$$

$$= 1 - \Pr(Z > 1)$$

$$= 1 - (1 - \Pr(Z \leq 1))$$

$$= \Pr(Z \leq 1)$$

$$= \Phi(1)$$

$$= 0.84134$$

(c)

$$\begin{aligned} Y &\sim \mathcal{N}(50, 25) \\ \Pr(y_1 \leq Y \leq y_2) &= \Pr\left(\frac{y_1 - \mu_Y}{\sigma_Y} \leq Z \leq \frac{y_2 - \mu_Y}{\sigma_Y}\right) \\ \Pr(40 \leq Y \leq 52) &= \Pr\left(\frac{40 - 50}{5} \leq Z \leq \frac{52 - 50}{5}\right) \\ &= \Pr(-2 \leq Z \leq 0.4) \\ &= \Pr(Z \leq 0.4) - \Pr(Z \leq -2) \\ &= \Pr(Z \leq 0.4) - \Pr(Z \geq 2) \\ &= \Pr(Z \leq 0.4) - (1 - \Pr(Z \leq 2)) \\ &= \Pr(Z \leq 0.4) + \Pr(Z \leq 2) - 1 \\ &= \Phi(0.4) + \Phi(2) - 1 \\ &= 0.65542 + 0.97725 - 1 \\ &= 0.63267 \end{aligned}$$

(d)

$$\begin{aligned} Y &\sim \mathcal{N}(5, 2) \\ \Pr(6 \leq Y \leq 8) &= \Pr\left(\frac{6 - 5}{\sqrt{2}} \leq Z \leq \frac{8 - 5}{\sqrt{2}}\right) \\ &= \Pr\left(\frac{1}{\sqrt{2}} \leq Z \leq \frac{3}{\sqrt{2}}\right) \\ &= \Pr\left(Z \leq \frac{3}{\sqrt{2}}\right) - \Pr\left(Z \leq \frac{1}{\sqrt{2}}\right) \\ &= \Phi\left(\frac{3}{\sqrt{2}}\right) - \Phi\left(\frac{1}{\sqrt{2}}\right) \\ &\approx 0.98300 - 0.76115 \\ &\approx 0.22185 \end{aligned}$$

Problem 2.11

(a)

$$\begin{aligned} Y &\sim \chi_4^2 \\ \Pr(Y \leq 7.78) &= \Pr(Z_1^2 + Z_2^2 + Z_3^2 + Z_4^2 \leq 7.78) \\ &= 0.90 \end{aligned}$$

(b)

$$\begin{aligned} Y &\sim \chi_4^{10} \\ \Pr(Y > 18.31) &= 1 - \Pr(Y \leq 18.31) \\ &= 1 - \Pr\left(\sum_{i=1}^{10} Z_i^2 \leq 18.31\right) \\ &= 1 - 0.95 \\ &= 0.05 \end{aligned}$$

(c)

$$\begin{aligned} Y &\sim F_{10,\infty} \\ \Pr(Y > 1.83) &= 1 - \Pr(Y \leq 1.83) \\ &= 1 - 0.95 \\ &= 0.05 \end{aligned}$$

d)

The $F_{m,\infty}$ is the distribution of the χ_m^2 random variable divided by m .

(e)

$$\begin{aligned} Y &\sim \chi_1^2 \\ \Pr(Y \leq 1.0) &= \Pr(Z_1^2 \leq 1.0) \\ \Pr(Y \leq 1.0) &= \Pr\left(-1.0 \leq \sqrt{Z_1^2} \leq 1.0\right) \\ &= \Pr(Z \leq 1.0) - \Pr(Z \leq -1.0) \\ &= \Pr(Z \leq 1.0) - \Pr(Z > 1.0) \\ &= \Pr(Z \leq 1.0) - (1 - \Pr(Z \leq 1.0)) \\ &= 2(0.84134) - 1 \\ &= 0.68268 \end{aligned}$$

Problem 2.14

$$\begin{aligned} \bar{Y} &\sim \mathcal{N}\left(\mu_Y, \frac{\sigma_Y^2}{n}\right) \\ \bar{Y} &\sim \mathcal{N}\left(100, \frac{43}{n}\right) \end{aligned}$$

(a)

$$\begin{aligned} \bar{Y} &\sim \mathcal{N}\left(100, \frac{43}{100}\right), n \geq 30 \\ \Pr(\bar{Y} \leq 101) &= \Pr\left(Z \leq \frac{101 - 100}{\sqrt{\frac{43}{100}}}\right) \\ &= \Pr(Z \leq 1.52) \\ &= 0.93574 \end{aligned}$$

(b)

$$\begin{aligned}\bar{Y} &\sim \mathcal{N}\left(100, \frac{43}{165}\right), n \geq 30 \\ \Pr(\bar{Y} > 98) &= 1 - \Pr\left(Z \leq \frac{98 - 100}{\sqrt{\frac{43}{165}}}\right) \\ &= 1 - \Pr\left(Z \leq \frac{98 - 100}{\sqrt{\frac{43}{165}}}\right) \\ &= 1 - \Pr(Z \leq -3.92) \\ &= 1 - \Pr(Z > 3.92) \\ &= 1 - (1 - \Pr(Z \leq 3.92)) \\ &= \Pr(Z \leq 3.92) \\ &= 0.99996\end{aligned}$$

(c)

$$\begin{aligned}\bar{Y} &\sim \mathcal{N}\left(100, \frac{43}{64}\right), n \geq 30 \\ \Pr(101 \leq \bar{Y} \leq 103) &= \Pr\left(Z \leq \frac{103 - 100}{\sqrt{\frac{43}{64}}}\right) - \Pr\left(Z \leq \frac{101 - 100}{\sqrt{\frac{43}{64}}}\right) \\ &= \Pr(Z \leq 3.66) - \Pr(Z \leq 1.22) \\ &= 0.99987 - 0.88877 \\ &= 0.1111\end{aligned}$$

Problem 2.23

$$E[Y|X](x) = E[Y|X = x]$$

$$\text{Corr}(X, Y) = 0$$

$$X \sim \mathcal{N}(1, 0)$$

$$Z \sim \mathcal{N}(1, 0)$$

$$Y = X^2 + Z$$

(a)

$$\begin{aligned}E[Y] &= E[X^2 + Z] \\E[E[Y|X]] &= E[X^2] + E[Z] \\E[E[Y|X]] &= E[X^2] + 1\end{aligned}$$

Hence,

$$E[Y|X] = X^2$$

(b)

$$\begin{aligned}\mu_Y &= E[Y] \\&= E[E[Y|X]] \\&= E[X^2]\end{aligned}$$

recall,

$$\begin{aligned}\text{Var}(X) &= E[X^2] - (E[X])^2 \\(\sigma_x)^2 &= E[X^2] - (1)^2 \\0 &= E[X^2] - (1)^2 \\E[X^2] &= 1\end{aligned}$$

Hence,

$$\mu_Y = 1$$

(c)

$$\begin{aligned}Cov(X, Y) &= E[(X - E[X])(Y - E[Y])] \\&= E[(XY - XE[Y] - YE[X] + E[X]E[Y])] \\&= E[XY - XE[Y] - Y + E[Y]] \\&= E[XY - X - Y + 1] \\&= E[XY] - E[X] - E[Y] + 1 \\&= E[XY] - 1 - 1 + 1 \\0 &= E[XY] - 1 \\1 &= E[XY]\end{aligned}$$

(d)

$$\begin{aligned}\rho_{XY} = Corr(X, Y) &= \frac{Cov(X, Y)}{\sigma_x \sigma_Y} \\0 &= \frac{Cov(X, Y)}{\sigma_x \sigma_Y} \\0 &= Cov(X, Y)\end{aligned}$$