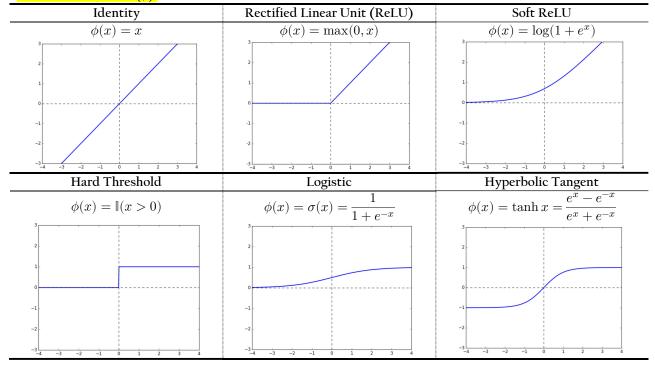
#### **Neural Networks**

Neuron/Unit: A model of a real neuron, which activates after a threshold of voltage is passed. Like logistic regression.

$$y(\mathbf{x}) = \phi(f(\mathbf{x})) = \phi(\mathbf{w} \cdot \mathbf{x} + b)$$

Activation Function  $(\phi)$ : The threshold that controls the activation of a neuron.



Black Box: In computing, a program where users don't care about inner workings, just using its inputs/outputs. Neural Network: A directed graph of neurons – has input layer of all  $\mathbf{x}_i$ , hidden layers, and output layer of all  $y_i(\mathbf{x})$ 

- Recurrent:
- A neural network that is a directed graph with cycles
- Feed-Forward:
- A neural network that is a directed acyclic graph
- Fully-Connected Layer: A neural network where all units in one layer connect to all units in the next layer
- Bottleneck Layer:
- A layer with much smaller dimensions than surrounding layers. "Loses" info
- Multilayer Perceptron: A multilayer neural network of fully-connected layers.

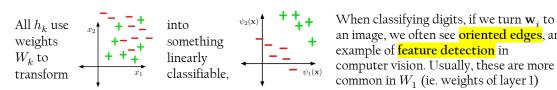
Assume for a layer, N input neurons  $(\mathbf{x} \in \mathbb{R}^N)$ , M output neurons  $(h(\mathbf{x}) \in \mathbb{R}^M)$ . Also,  $\phi$  is applied component-wise.

$$h(\mathbf{x}) = \phi(f(\mathbf{x})) = \phi(W\mathbf{x} + \mathbf{b}) = \phi\left(\begin{bmatrix} (w_1)_1 & \cdots & (w_1)_N \\ \vdots & \ddots & \vdots \\ (w_M)_1 & \cdots & (w_M)_N \end{bmatrix} \begin{bmatrix} x_1 \\ \vdots \\ x_N \end{bmatrix} + \begin{bmatrix} b \\ \vdots \\ b \end{bmatrix}\right)$$

A neural network is a composition of layer functions that can be implemented as a black box:

$$\begin{array}{ccc} h_1(\mathbf{x}) = \phi(W_1\mathbf{x} + \mathbf{b}_1) & \text{Regression:} \\ (h_2 \circ h_1)(\mathbf{x}) = \phi(W_2h_1(\mathbf{x}) + \mathbf{b}_2) & \text{Choose } \phi^*(x) = x \in \mathbb{R} \\ (h_3 \circ h_2 \circ h_1)(\mathbf{x}) = \phi(W_3(h_2 \circ h_1)(\mathbf{x}) + \mathbf{b}_3) & \text{Binary Classification:} \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ y(\mathbf{x}) = (h_K \circ \ldots \circ h_1)(\mathbf{x}) = \phi^*(W_K(h_{K-1} \circ \ldots \circ h_1)(\mathbf{x}) + \mathbf{b}_K) & \text{Choose } \phi^*(x) = \sigma(x) \in \mathbb{R} \\ \end{array}$$

Choose  $\phi^*(x) = \sigma(x) \in [0,1]$ For regression/binary classification,  $W_K = \mathbf{w}_k$  (one output)



an image, we often see oriented edges, an common in  $W_1$  (ie. weights of layer 1)



#### XOR, Binary Linear Classification, Feature Maps

$$\begin{split} f(\mathbf{x}) &= \mathbf{w} \cdot \psi(\mathbf{x}) \\ \psi(\mathbf{x}) &= (x_0, x_1, x_2, x_1 x_2) \end{split}$$

	$\mathbf{x}^{\mathrm{T}}$		$\psi(\mathbf{x})$				+
$x_0$	$x_1$	$x_2$	$\psi_0(\mathbf{x})$	$\psi_1(\mathbf{x})$	$\psi_2(\mathbf{x})$	$\psi_3(\mathbf{x})$	ι
1	0	0	1	0	0	0	0
1	0	1	1	0	1	0	1
1	1	0	1	1	0	0	1
1	1	1	1	1	1	1	0

$$\begin{split} f(\mathbf{x}_1) &= w_0(1) + w_1(0) + w_2(0) + w_1w_2(0) < 0 \\ f(\mathbf{x}_2) &= w_0(1) + w_1(0) + w_2(1) + w_1w_2(0) \ge 0 \\ f(\mathbf{x}_3) &= w_0(1) + w_1(1) + w_2(0) + w_1w_2(0) \ge 0 \\ f(\mathbf{x}_4) &= w_0(1) + w_1(1) + w_2(1) + w_1w_2(1) \ge 0 \end{split}$$

#### XOR, Neural Network

We set  $\phi(x) = \mathbb{I}(x > 0)$ .

Each edge is a  $w_i$  we must find.

Each I represents a  $x_0$  with weight  $w_0$  (AKA b)

$$\begin{array}{ll} h_1(\mathbf{x}) = \mathbb{I}(x_1(1) & +x_2(1) & +1(-0.5 \ ) > 0) \\ h_2(\mathbf{x}) = \mathbb{I}(x_1(1) & +x_2(1) & +1(-1.5 \ ) > 0) \\ y(\mathbf{x}) = \mathbb{I}(h_1(\vec{x})(1) + h_2(\vec{x})(-1) + 1(-0.5 \ ) > 0) \end{array}$$

Notice the connection to logic:

$$\begin{split} h_1(\mathbf{x}) &= \mathbb{I}(x_1 + x_2 > 0.5) \equiv x_1 \vee x_2 \\ h_2(\mathbf{x}) &= \mathbb{I}(x_1 + x_2 > 1.5) \equiv x_1 \wedge x_2 \\ y(\mathbf{x}) &= \mathbb{I}(h_1(\mathbf{x}) + \left(1 - h_2(\mathbf{x})\right) > 1.5) \\ &\equiv h_1(\mathbf{x}) \wedge \left(\neg h_2(\mathbf{x})\right) \equiv x_1 \text{ XOR } x_2 \end{split}$$

For 1-layer XOR on n inputs, use  $2^n$  neurons for all T/F combinations.

- Binary linear classification can't classify XOR without feature maps, which are hard to design, especially for very complex functions. Neural networks help find these feature maps for us.
- The hard threshold/0-1 loss  $\phi(x) = \mathbb{I}(x > 0)$  is not ideal for gradient descent as the derivative is almost all 0
- Use the smooth logistic function  $\phi(x) = \sigma(nx)$  (where n is large), which approaches the hard threshold.

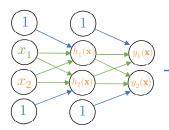
Depth: Number of layers in a neural network (shallow means little layers, deep means much layers, deep learning refers to training high-depth neural networks) Width: Number of units in a layer

Input Space: Set  $\mathcal{X}$  of all possible inputs

Target Space: Set  $\mathcal{T}$  of all possible targets

Hypothesis: A function  $f: \mathcal{X} \to \mathcal{T}$ 

**Hypothesis Space:** The set  $\mathcal{H}$  of all possible hypotheses f a given model can have



- Parametric Algorithms:  $\mathcal{H}$  is a set of all parameter combinations
- Non-Parametric Algorithms:  $\mathcal{H}$  is defined in terms of the training data
- Inductive Bias: Of an algorithm, its preference to choose some hypotheses  $f \in \mathcal{H}$  over others.
  - $\circ$  eg. For linear regression,  $\mathcal{H}$  is the set of linear functions
  - o eg. Ridge regression results in an inductive bias for smaller weights
- Expressivity: The variety of functions able to be represented by  $\mathcal{H}$ 
  - o  $\mathcal{H}_B \subseteq \mathcal{H}_A$  means model A is more expressive; it can represent any function  $f \in \mathcal{H}_B$
  - o eg. Neural networks where  $\phi(x) = x$  are <u>as expressive as linear regression</u>. Every layer  $h_k(x)$  is a linear function, which is equivalent to a single linear layer.
  - o eg. Multilayer feed-forward neural networks with nonlinear  $\phi$  are universal function approximators: they can approximate any function arbitrarily well.  $\forall f \colon \mathcal{X} \to \mathcal{T}, \exists f_i \in \mathcal{H}, f_i \to f$ 
    - This is not necessarily a good thing overfitting is a concern
    - Early Stopping: Stopping training just as generalization error begins to increase

No Free Lunch Theorem: If datasets were not naturally biased, no ML algorithm would be better than any other.

Let  $\star_n$  represent the n-th component of function  $\star$ , and  $\star^{(n)}$  represent a value in the n-th layer.

$$\begin{split} h_1(\mathbf{x}) &= \phi \left( W_1^{(1)} \cdot \mathbf{x} \right) = \phi \left[ (w_1)_0^{(1)} + (w_1)_1^{(1)} x_1 + (w_1)_2^{(1)} x_2 \right] \\ h_2(\mathbf{x}) &= \phi \left( W_2^{(1)} \cdot \mathbf{x} \right) = \phi \left[ (w_2)_0^{(1)} + (w_2)_1^{(1)} x_1 + (w_2)_2^{(1)} x_2 \right] \\ y_1(\mathbf{x}) &= \phi \left( W_1^{(2)} \cdot h(\mathbf{x}) \right) = \phi \left[ (w_1)_0^{(2)} + (w_1)_1^{(2)} h_1(\mathbf{x}) + (w_1)_2^{(2)} h_2(\mathbf{x}) \right] \\ y_2(\mathbf{x}) &= \phi \left( W_2^{(2)} \cdot h(\mathbf{x}) \right) = \phi \left[ (w_2)_0^{(2)} + (w_2)_1^{(2)} h_1(\mathbf{x}) + (w_2)_2^{(2)} h_2(\mathbf{x}) \right] \\ Note: \text{Use } \mathbf{x} &= (1, x_1, \dots, x_n) \text{ and } h(\mathbf{x}) = \left( 1, h_1(\mathbf{x}), \dots, h_n(\mathbf{x}) \right) \\ \text{since we squish } b \text{ into } W = (W_1, W_2) \end{split}$$

$$h(\mathbf{x}) = \phi \left( \begin{bmatrix} (w_1)_0 & (w_1)_1 & (w_2)_1 \\ (w_2)_0 & (w_2)_1 & (w_2)_2 \end{bmatrix}^{(1)} \begin{bmatrix} 1 \\ x_1 \\ x_2 \end{bmatrix} \right)$$

$$= \phi(W^{(1)}\mathbf{x})$$

$$y(\mathbf{x}) = \phi \left( \begin{bmatrix} (w_1)_0 & (w_1)_1 & (w_2)_1 \\ (w_2)_0 & (w_2)_1 & (w_2)_2 \end{bmatrix}^{(2)} \begin{bmatrix} 1 \\ h_1(\mathbf{x}) \\ h_2(\mathbf{x}) \end{bmatrix} \right)$$

$$= \phi \left( W^{(2)}h(\mathbf{x}) \right)$$

Calculating Univariate  $\frac{\partial \mathcal{L}}{\partial \mathbf{w}}$  the Normal Way Suppose we have the network

$$\begin{split} f &= f(\mathbf{x}) = \mathbf{w} \cdot \mathbf{x} + b \\ y &= y(\mathbf{x}) = \sigma \big( f(\mathbf{x}) \big) \\ \mathcal{L} &= \mathcal{L}(y(\mathbf{x}), t) = \frac{1}{2} (y(\mathbf{x}) - t)^2 \end{split}$$

The derivation for  $\sigma'(f(\mathbf{x}))$  is the exact same as for binary classification, logistic regression.

Calculating Univariate  $\frac{\partial \mathcal{L}}{\partial \mathbf{w}}$  the Efficient Way

$$\overline{y} = \frac{\partial \mathcal{L}}{\partial y} = \frac{\partial}{\partial y} \left[ \frac{1}{2} (y - t)^2 \right]$$

$$= y - t$$

$$\overline{f} = \frac{\partial \mathcal{L}}{\partial f} = \frac{\partial \mathcal{L}}{\partial y} \cdot \frac{\partial y}{\partial f}$$

$$= \overline{y} (y(1 - y))$$

$$\begin{split} \frac{\partial \mathcal{L}}{\partial \mathbf{w}} &= \frac{\partial \mathcal{L}}{\partial y} \cdot \frac{\partial y}{\partial f} \cdot \frac{\partial f}{\partial \mathbf{w}} \\ &= \frac{\partial}{\partial y} \left[ \frac{1}{2} (y - t)^2 \right] \frac{\partial}{\partial f} [\sigma(f)] \frac{\partial}{\partial \mathbf{w}} [(\mathbf{w} \cdot \mathbf{x} + b)] \\ &= \left[ (y - t) \frac{\partial}{\partial y} (y - t) \right] \sigma'(f) \mathbf{x} \\ &= (y - t) y (1 - y) \mathbf{x} \\ \frac{\partial \mathcal{L}}{\partial b} &= \frac{\partial \mathcal{L}}{\partial y} \cdot \frac{\partial y}{\partial f} \cdot \frac{\partial f}{\partial b} = (y - t) y (1 - y) \end{split}$$

$$\begin{split} \overline{\mathbf{w}} &= \frac{\partial \mathcal{L}}{\partial \mathbf{w}} = \frac{\partial \mathcal{L}}{\partial f} \cdot \frac{\partial f}{\partial \mathbf{w}} & \text{We found the loss function in the order} \\ &= \overline{f}\mathbf{x} & \text{We found derivatives in reverse order,} \\ &\bar{b} &= \frac{\partial \mathcal{L}}{\partial b} = \frac{\partial \mathcal{L}}{\partial f} \cdot \frac{\partial f}{\partial b} & \text{We found derivatives in reverse order,} \\ &= \overline{f} & \text{The notation } \overline{*} = \frac{\partial \mathcal{L}}{\partial *} \text{ is called error signal.} \end{split}$$

 $\otimes$  and  $\ominus$  are element-wise  $\times$  and -

Computational Graph: A directed graph where a function's most-direct dependencies point to it.

$$\begin{array}{c} \mathbf{x} & t \\ b & \mathbf{y}(\mathbf{x}) \longrightarrow \mathcal{L}(y(\mathbf{x}), t) \end{array}$$

**Topological Ordering:** Of a directed graph G, a linear ordering L where  $\forall u, v \in G$ , u points to v in  $G \Leftrightarrow u$  precedes v in L

Fan-out: The max number of children a node can have in a graph.

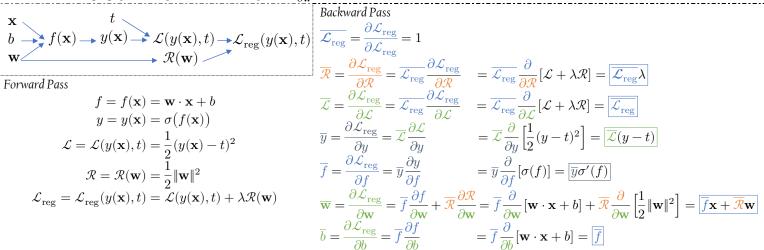
Backpropagation: A computationally efficient way of finding  $\frac{\partial \mathcal{L}}{\partial \mathbf{w}}$  to update the weights of a neural network. Algorithm:

Let  $v_1, \ldots, v_N$  be a topological ordering of the computation graph (ie. parents are before children).

- 1) Forward Pass: A loop  $1 \to N$  down the computational graph AKA forward in the ordering. For each  $v_{\text{child}}$ , express it in terms of ParentFunc $(v_{\text{child}})$
- Backward Pass: A loop N o 1 up the computational graph AKA backward in the ordering.

For each 
$$v_{\mathrm{parent}}$$
, calculate  $\overline{v_i} = \sum_{\mathrm{child} \in \mathrm{ChildrenFunc}(v_i)} \overline{v_{\mathrm{child}}} \frac{\partial v_{\mathrm{child}}}{\partial v_{\mathrm{parent}}}$ 

Backpropagation, Regularized Linear Regression  $\frac{\partial \mathcal{L}}{\partial \mathbf{w}}$ 



Backpropagation, Simple Neural Network  $\frac{\partial \mathcal{L}}{\partial \mathbf{w}}$ 

$$\begin{array}{c|c} W_1 & \mathbf{t} \\ \mathbf{x} & h(\mathbf{x}) & \mathcal{L}(y(\mathbf{x}), \mathbf{t}) \\ \mathbf{b}_1 & \mathbf{b}_2 & \overline{\mathcal{L}} = \frac{\partial \mathcal{L}}{\partial \mathcal{L}} = 1 \\ \hline \overline{\mathbf{y}} = \frac{\partial \mathcal{L}}{\partial \mathbf{y}} = \overline{\mathcal{L}} \end{array}$$

Forward Pass

$$\begin{split} \mathbf{f} &= f(\mathbf{x}) = W_1 \mathbf{x} + \mathbf{b}_1 \\ \mathbf{h} &= h(\mathbf{x}) = \sigma \big( f(\mathbf{x}) \big) \text{ (element-wise)} \\ \mathbf{y} &= y(\mathbf{x}) = W_2 h(\mathbf{x}) + \mathbf{b}_2 \\ \mathcal{L} &= \mathcal{L}(y(\mathbf{x}), \mathbf{t}) = \frac{1}{2} \|y(\mathbf{x}) - \mathbf{t}\|^2 \end{split}$$

For backward passes, keep track of the error signals' dimensions. Change the order of matrix multiplication if necessary.

Backward Pass 
$$\overline{\mathcal{L}} = \frac{\partial \mathcal{L}}{\partial \mathcal{L}} = 1$$

$$\overline{\mathbf{y}} = \frac{\partial \mathcal{L}}{\partial \mathbf{y}} = \overline{\mathcal{L}} \frac{\partial \mathcal{L}}{\partial \mathbf{y}} = \overline{\mathcal{L}} \frac{\partial}{\partial y} \left[ \frac{1}{2} \| \mathbf{y} - \mathbf{t} \|^2 \right] = \overline{\mathcal{L}} (\mathbf{y} - \mathbf{t})$$

$$\overline{W}_2 = \frac{\partial \mathcal{L}}{\partial W_2} = \overline{\mathbf{y}} \frac{\partial \mathcal{Y}}{\partial W_2} = \overline{\mathbf{y}} \frac{\partial}{\partial W_2} [W_2 \mathbf{h} + \mathbf{b}_2] = \overline{\mathbf{y}} \mathbf{h}^T$$

$$\overline{\mathbf{b}}_2 = \frac{\partial \mathcal{L}}{\partial \mathbf{b}_2} = \overline{\mathbf{y}} \frac{\partial \mathbf{y}}{\partial \mathbf{b}_2} = \overline{\mathbf{y}} \frac{\partial}{\partial \mathbf{b}_2} [W_2 \mathbf{h} + \mathbf{b}_2] = \overline{\mathbf{y}}$$

$$\overline{\mathbf{h}} = \frac{\partial \mathcal{L}}{\partial \mathbf{h}} = \overline{\mathbf{y}} \frac{\partial \mathbf{y}}{\partial \mathbf{h}} = \overline{\mathbf{y}} \frac{\partial}{\partial \mathbf{h}} [W_2 \mathbf{h} + \mathbf{b}_2] = \overline{\mathbf{W}}_2^T \overline{\mathbf{y}}$$

$$\overline{\mathbf{f}} = \frac{\partial \mathcal{L}}{\partial \mathbf{f}} = \overline{\mathbf{h}} \frac{\partial \mathbf{h}}{\partial \mathbf{f}} = \overline{\mathbf{h}} \frac{\partial}{\partial \mathbf{f}} [\sigma(\mathbf{f})] = \overline{\mathbf{h}} \otimes \sigma'(\mathbf{f}) \text{ (elementwise } \times \text{)}$$

$$\overline{W}_1 = \frac{\partial \mathcal{L}}{\partial \mathbf{W}_1} = \overline{\mathbf{f}} \frac{\partial \mathbf{f}}{\partial W_1} = \overline{\mathbf{f}} \frac{\partial}{\partial W_1} [W_1 \mathbf{x} + \mathbf{b}_1] = \overline{\mathbf{f}} \mathbf{x}^T$$

$$\overline{\mathbf{b}}_1 = \frac{\partial \mathcal{L}}{\partial \mathbf{b}_1} = \overline{\mathbf{f}} \frac{\partial \mathbf{f}}{\partial \mathbf{b}_1} = \overline{\mathbf{f}} \frac{\partial}{\partial \mathbf{b}_1} [W_1 \mathbf{x} + \mathbf{b}_1] = \overline{\mathbf{f}}$$

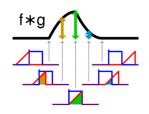
Convolution (\*): In functional analysis, a mathematical operator on functions (sometimes called signals for the application in signal processing). In the discrete case:

For $f, g: \mathbb{Z} \to \mathbb{R}$ ,	For $f, g: \mathbb{Z}^2 \to \mathbb{R}$ ,
$(f * g)(x) = \sum_{i=-\infty}^{\infty} f(i) g(x-i)$	$(f*g)(x,y) = \sum_{i=-\infty}^{\infty} \sum_{j=-\infty}^{\infty} f(i,j) g[x-i,y-j]$

Computationally, think of signals *f* and g like an <u>array</u>. We find the sum of all  $f[i] \times g. \, \text{reverse}()[i]$ 

Properties of Convolution (f \* g)(x) = (g \* f)(x) $(f * (\alpha q + \beta h))(x) = \alpha (f * g)(x) + \beta (f * h)(x)$ 

The continuous case replaces sum with an integral. Visually, think of it like reflecting g (left) and sliding it left to right across f. The area intersecting f and g in any given moment is f \* g.



g

See this video for better visual explanation. Visually, for the discrete case, consider the problem [2, -1, 1] \* [1, 1, 2]:

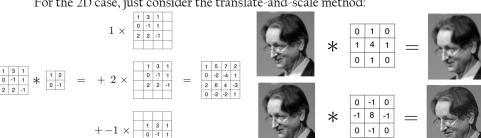
# Translate-And-Scale

Zero Padding: Treating finite signals as infinitely large by adding 0s where the domain's undefined

# Flip-And-Filter

(f\*q)(i) is equal to the sum of all possible  $f(i) \times g(i)$ obtained by sliding q(i)rightwards.

For the 2D case, just consider the translate-and-scale method:



For any cell in the kernel matrix, consider the image formed by translating every pixel in the original image by 1 (in the same direction as from the matrix's center cell to the given cell). Multiply the new image by the cell. Do this for every cell, and add the results.

Convolutional Network: A regularized type of multilayer perceptron specialized for image recognition

- Locally-Connected Layers: When neurons in one layer do not connect to every neuron in the next layer
- Convolution Layer: Layers where each set of neurons look at one image region, and weights are shared between all image locations. Can be stacked.

  - o Consists of filters, functions which, when convolved with the image, creates a feature map/channel
  - Shared weights allow the network to detect features in many locations of the image
- Note that W is linear; if the system is completely linear, no number of layers is more powerful than 1 layer.
  - Pooling Layer: Layers that deliberately lose some information to reduce computational load and add non-linearity and invariance
    - Max-Pooling: When neurons take the max of neurons from a pooling group in the previous layer.
  - $\circ$  Linear Rectification: A non-linear activation function-based transformation that adds a non-linearity, like ReLU,  $y_n(\mathbf{x}) = \max(0, h_n(\mathbf{x}))$
- Equivariance: A property of convolution layers translating inputs translates outputs by the same amount
- Invariance: A property of convolution layers translating inputs doesn't affect outputs

# Probabilistic Modelling

**Likelihood Function** (L): Given observed data  $\mathcal{D} = \{x_1, ..., x_N\}$ , a probability function L(...) (parameters vary) that retroactively describes the probability distribution of the "real" underlying data (where order of data matters).

- eg.  $L(\theta) = p_{X|\theta}(\mathcal{D}|\theta) = p_{X|\theta}(x_1,\dots,x_N|\theta)$ , where X is the "real" probability distribution of data,  $\theta$  is a parameter. This reads as "the probability that the data would be  $x_1,\dots,x_N$ , given we know  $\theta$ "
- Notation can be inconsistent.  $p(\theta) = p_{\theta}(\theta) = p_{\theta}(k)$ . I use k only if to differentiate between possible  $\theta$  values.

Log-Likelihood Function ( $\ell = \ln(L)$ ): Used in practice as it gives reasonably-sized numbers and simplifies mathege. Flip a coin N = 100 times. Let  $\mathcal{D} = \{x_1, \dots, x_N\}$ , where  $x_i \in \{0,1\}$  (1 = Heads, 0 = Tails)

Since coin flips are i.i.d., we model them as

 $X \sim \text{Bernoulli}(\theta)$  where  $\theta \in [0,1]$  is the chance of heads.

$$\begin{split} L(\theta) &= p_{X|\theta}(\mathcal{D}, \theta) = \prod_{n=1}^{N} \begin{cases} \theta & x_n = 1 \\ 1 - \theta & x_n = 0 \end{cases} \\ &= \prod_{n=1}^{N} \theta^{x_n} (1 - \theta)^{1 - x_n} \end{split}$$

To make the math easier, we use log-likelihood:

$$\begin{split} \ell(\theta) &= \ln \prod_{n=1}^N \theta^{x_n} (1-\theta)^{1-x_n} \\ &= \sum_{n=1}^N [x_n \ln \theta + (1-x_n) \ln (1-\theta)] \end{split}$$

Note the similarity to binary cross-entropy!  $\mathcal{L}(y(\vec{x}),t) = -t \ln y(\vec{x}) - (1-t) \ln (1-y(\vec{x}))$ 

**Maximum Likelihood Criterion:** Maximize L (set  $\nabla \ell = 0$ ) and plug in observed data  $\mathcal D$  to find  $\ell$ 's parameters.

• As  $\ln x$  is convex,  $\operatorname{argmax}(L) = \operatorname{argmax}(\ell)$ , and  $\ell$  is easier to work with. Solve directly or gradient descent!

eg. Flip a coin N=100 times. Suppose we observe  $N_{\rm H}=55, N_{\rm T}=45$ . What are the odds of heads/tails?

The likelihood function is the probability mass function of the number of heads/tails after N=100 coin flips.

$$\begin{split} \ell(\theta) &= \ln \prod_{n=1}^N \theta^{x_n} (1-\theta)^{1-x_n} \\ &= \sum_{n=1}^N [x_n \ln \theta + (1-x_n) \ln (1-\theta)] \end{split}$$

Maximum likelihood criterion says to obtain the most accurate  $\theta$ , maximize  $\ell(\theta)$ .

$$\begin{split} \bullet \quad & \frac{d\ell}{d\theta} = \frac{d}{d\theta} \left( \sum_{n=1}^{N} [x_n \ln \theta + (1 - x_n) \ln(1 - \theta)] \right) \\ & = \sum_{n=1}^{N} \left( \frac{x_n}{\theta} + \frac{1 - x_n}{1 - \theta} \right) \\ & = \sum_{n=1}^{N} \frac{x_n}{\theta} + \sum_{n=1}^{N} \frac{1 - x_n}{1 - \theta} \\ & = \frac{N_{\mathrm{H}}}{\theta} - \frac{N_{\mathrm{T}}}{1 - \theta} = 0 \\ & \qquad \qquad \therefore \theta^* = \frac{N_{\mathrm{H}}}{N_{\mathrm{H}} + N_{\mathrm{T}}} = 0.55 \end{split}$$

eg. Suppose we get a temperature distribution  $\mathcal{D}=\{-2.5,-9.9,-12.1,-8.9,-6.0,-4.8,2.4\}=\{x_1,\dots,x_N\}$  where N=7. Choose  $X\sim \operatorname{Normal}(\mu,\sigma^2)$  for simplicity. Need to find  $\mu$  and  $\sigma$ .

$$\begin{split} \ell(\mu,\sigma^2) &= \ln p_{X|\mu,\sigma^2}(\mathcal{D}|\mu,\sigma^2) \\ &= \ln \prod_{n=1}^N \left[ \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x_n-\mu)^2}{2\sigma^2}} \right] \\ &= \sum_{n=1}^N \ln \left[ \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x_n-\mu)^2}{2\sigma^2}} \right] \\ &= \sum_{n=1}^N \left[ \ln \left( \frac{1}{\sigma\sqrt{2\pi}} \right) + \ln \left( e^{-\frac{(x_n-\mu)^2}{2\sigma^2}} \right) \right] \\ &= -\sum_{n=1}^N \left[ \ln(\sigma) + \ln \sqrt{2\pi} + \frac{(x_n-\mu)^2}{2\sigma^2} \right] \end{split}$$

To find optimal  $\sigma$ ,  $\mu$  values, we differentiate with respect to both variables and set them to 0.

$$\frac{d\ell}{d\mu} = -\sum_{n=1}^{N} \left[ \frac{2(x_n - \mu)(-1)}{2\sigma^2} \right]$$

$$= \frac{1}{\sigma^2} \sum_{n=1}^{N} (x_n - \mu)$$

$$= \frac{1}{\sigma^2} \left[ \sum_{n=1}^{N} (x_n) - N\mu \right] = 0$$
Therefore, set  $\mu = \frac{1}{N} \sum_{n=1}^{N} x_n = -5.97$ .
$$d\ell = \sum_{n=1}^{N} \left[ 1 - (x_n - \mu)^2 \right]$$

$$\frac{d\ell}{d\sigma} = -\sum_{n=1}^{N} \left[ \frac{1}{\sigma} - \frac{(x_n - \mu)^2}{\sigma^3} \right]$$
$$= -\sum_{n=1}^{N} \left[ \frac{\sigma^2 - (x_n - \mu)^2}{\sigma^3} \right] = 0$$

Therefore, set  $N\sigma^2 = \sum_{n=1}^{N} (x_n - \mu)^2$ , so  $\sigma = 4.55$ .

Discriminative: A classification approach based on finding targets directly via training set data.

- Try to "learn"  $p(t|\mathbf{x})$  directly, finds a mapping  $y(\mathbf{x}) \approx t$ . Includes every model we've covered so far.
- $p(t|\mathbf{x})$  is informal, reads as "the probability that an input's target is t, given the input is  $\mathbf{x}$ "

Bayesian/Generative: A classification approach based on modelling the distribution of training set data in a target.

- Try to "learn"  $p(\mathbf{x}|t)$ , calculate  $p(t|\mathbf{x})$  via Bayes Rule
- $p(\mathbf{x}|t)$  is informal, reads as "the probability that the input looks like  $\mathbf{x}$ , given its target is t"

Naïve Bayes: Generative classifier that assumes all  $x_i$  in x are conditionally independent given t (ie. covariance is 0) eg. An email spam classifier, with N emails in training set, where  $t \in \{0,1\}$ ,  $\mathbf{x} \in \{0,1\}^D$ , and  $x_d \in \{0,1\}$  represents whether the d-th word (of D total words from all emails) appears inside an email

$$\begin{split} & \text{Finding log-likelihood,} \\ & L(\theta) = p_{X|\theta}(\mathcal{D}|\theta) \\ & = \prod_{n=1}^N p_{X|\theta}((x_n)_1, \dots (x_n)_D, t_n|\theta) \\ & \ell(\theta) = \sum_{n=1}^N \ln p(\mathbf{x}_n, t_n) \end{split}$$

For now, I omitted  $\theta$  for visual simplicity. How do we expand  $p(\mathbf{x}_n, t_n)$ ?

$$\begin{split} \overrightarrow{p}(\mathbf{x}_n, t_n) &= p(t_n) \times p((x_n)_1 | t_n) \\ &\times p((x_n)_2 | (x_n)_1, t_n) \\ &\times p((x_n)_3 | (x_n)_1, (x_n)_2, t_n) \\ &\times \dots \\ &\times p((x_n)_D | (x_n)_1, \dots, (x_n)_{D-1}, t_n) \end{split}$$

This is computationally infeasible, so in Naïve Bayes, we assume

$$p(\mathbf{x}_n,t_n)=p(t_n)\times\prod_{d=1}^Dp((x_n)_d|t_n)$$
 That is,  $p(\text{word }i\text{ appears}|t_n)$  doesn't affect  $p(\text{word }j\text{ appears}|t_n).$ 

Although the Naïve Bayes assumption is terrible, it simplifies our math.

$$\begin{split} \ell(\theta) &= \sum_{n=1}^{N} \ln p(\mathbf{x}_n, t_n) \\ &= \sum_{n=1}^{N} \ln [p(t_n) \cdot p(\mathbf{x}_n | t_n)] \\ &= \sum_{n=1}^{N} \ln [p(t_n) \cdot p(\mathbf{x}_n | t_n)] \\ &= \sum_{n=1}^{N} \ln [p(t_n) \cdot p((x_n)_1 | t_n) \cdot \dots \cdot p((x_n)_D | t_n)] \end{split} \\ \Rightarrow \sum_{n=1}^{N} \ln [p(t_n) \cdot p((x_n)_1 | t_n) \cdot \dots \cdot p((x_n)_D | t_n)] \end{split}$$

Optimizing Left Term

Let's define  $\phi = p(t_n = 1)$ .

$$\begin{aligned} \operatorname{Recall that} t_n &\in \{0,1\}, \\ p(t_n) &= \begin{cases} \phi & t_n = 1 \\ 1 - \phi & t_n = 0 \\ = \phi^{t_n} (1 - \phi)^{1 - t_n} \end{aligned} \end{aligned} \\ &= \sum_{n=1}^N \ln[\phi^{t_n} (1 - \phi)^{1 - t_n}] \\ &= \sum_{n=1}^N \left[ t_n \ln \phi + (1 - t_n) \ln(1 - \phi) \right] \end{aligned} \\ &= \sum_{n=1}^N \left[ t_n \ln \phi + (1 - t_n) \ln(1 - \phi) \right] \end{aligned} \\ \overset{d\ell_1}{=} \sum_{n=1}^N \frac{t_n}{\phi} + \sum_{n=1}^N \frac{1 - t_n}{1 - \phi} \\ &= \frac{N_{\operatorname{spam}}}{\phi} + \frac{N_{\operatorname{not spam}}}{1 - \phi} = 0 \end{aligned}$$
 
$$\vdots \phi^* = \frac{N_{\operatorname{spam}}}{N_{\operatorname{spam}}} + N_{\operatorname{not spam}} \end{aligned}$$
 That is the percentage of spam emains

Optimizing Right Term

Let's define 
$$\phi_k = p((x_n)_d = 1 | t_n = k)$$
 where  $k \in \{0,1\}$ . 
$$p((x_n)_d | t_n) = \begin{cases} \phi_k & (x_n)_d = 1 \\ 1 - \phi_k & (x_n)_d = 0 \end{cases}$$
 
$$= \phi_k^{(x_n)_d} (1 - \phi_k)^{1 - (x_n)_d}$$
 
$$\therefore \ell_2(\phi_k) = \ln \prod_{n=1}^N p((x_n)_d | t_n)$$
 
$$= \sum_{n=1}^N [(x_n)_d \ln \phi_k + (1 - (x_n)_d) \ln (1 - \phi_k)]$$
 
$$= \sum_{n=1}^N (1 - t_n) [(x_n)_d \ln \phi_0 + (1 - (x_n)_d) \ln (1 - \phi_0)] + \sum_{n=1}^N t_n [(x_n)_d \ln \phi_1 + (1 - (x_n)_d) \ln (1 - \phi_1)]$$

Differentiating with respect to  $\phi_0$  and  $\phi_1$ ,

$$\begin{split} \frac{d\ell_2}{d\phi_0} &= \sum_{n=1}^N (1-t_n) \left[ \frac{(x_n)_d}{\phi_0} - \frac{1-(x_n)_d}{1-\phi_0} \right] = 0 \\ \sum_{n=1}^N (1-t_n) [(x_n)_d (1-\phi_0) - (1-(x_n)_d)\phi_0] &= 0 \\ \sum_{n=1}^N (1-t_n) [(x_n)_d - \phi_0] &= 0 \\ \sum_{n=1}^N (1-t_n) (x_n)_d &= \phi_0 \sum_{n=1}^N (1-t_n) \\ \phi_0^* &= \frac{\sum_{n=1}^N (1-t_n) (x_n)_d}{\sum_{n=1}^N (1-t_n)} \\ &= \frac{N_{\text{dth word,not spam}}}{N_{\text{not spam}}} \end{split}$$

$$\begin{split} \frac{d\ell_2}{d\phi_1} &= \sum_{n=1}^N t_n \left[ \frac{(x_n)_d}{\phi_1} - \frac{1 - (x_n)_d}{1 - \phi_1} \right] = 0 \\ \sum_{n=1}^N t_n [(x_n)_d (1 - \phi_1) - (1 - (x_n)_d) \phi_1] &= 0 \\ \sum_{n=1}^N t_n [(x_n)_d - \phi_1] &= 0 \\ \sum_{n=1}^N t_n (x_n)_d &= \phi_1 \sum_{n=1}^N t_n \\ \phi_1^* &= \frac{\sum_{n=1}^N t_n (x_n)_d}{\sum_{n=1}^N t_n} \\ &= \frac{N_{\text{dth word,spam}}}{N_{\text{spam}}} \end{split}$$

"The percentage of spams containing the *d*-th word"

Now that we know optimal  $p(t_n)$  and  $p((x_n)_d|t_n)$  values, at test time, we apply Bayes' Rule to find  $p(t|\mathbf{x})$ .

$$p(t=k|\mathbf{x}) = \frac{p(t=k)p(\mathbf{x}|t=k)}{p(\mathbf{x})} = \frac{p(t=k)p(\mathbf{x}|t=k)}{\sum_{k'=0}^{1}p(t=k')p(\mathbf{x}|t=k')} = \frac{p(t=k)\prod_{d=1}^{D}p(x_d|t=k)}{\sum_{k'=0}^{1}p(t=k')\prod_{d=1}^{D}p(x_d|t=k')}$$
 Recall that the ideal  $p(t=k)$  is  $\frac{N_{\text{spam}}}{N_{\text{spam}}+N_{\text{not spam}}}$ , and the ideal  $p(x_d=1|t=k)$  is  $\frac{N_{\text{dth word, spam/no spam}}}{N_{\text{spam/no spam}}}$ . Predict the  $t \in \{0,1\}$  that results in the highest  $p(t|\mathbf{x})$ ; the denominator is constant and can be omitted.

For computational speed, use the  $x=e^{\ln x}$  rule to turn the product into a summation:  $p(t=k)e^{\sum_{d=1}^D \ln p(x_d|t=k)}$ 

- Works with any probability distribution
- Efficient 1 pass of data at train-time,  $\mathcal{O}(D)$  at test-time but inaccurate, due to the independence assumption

Data Sparsity: Lack of data. Causes log to fail, overfitting in maximum likelihood. Solve with prior distributions.

- eg. If  $p(x_d|t=k)=0$  in our dataset, then  $\ln p(x_d|t=k)=-\infty$ .
- eg. Flip a coin twice, get H twice, then maximum likelihood says  $\theta = \frac{2}{2+0} = 1$

Prior Distribution $oldsymbol(p(oldsymbol{ heta})oldsymbol)$	Likelihood $(L(\theta) = p(\mathcal{D} \theta))$	Posterior Distribution $(p(\theta \mathcal{D}))$
Our beliefs about parameters before	The likelihood of observing a	Our updated beliefs about parameters
observing data.	given configuration of data,	after observing a configuration of data.
<ul> <li>Uninformative Prior: A prior distribution that assumes the least possible.</li> <li>Usually, it's something computationally convenient (uniform distribution, Bernoulli)</li> </ul>	given parameters.	$p(\theta \mathcal{D}) = \frac{p(\theta)p(\mathcal{D} \theta)}{\int_{-\infty}^{\infty} p(\theta')p(\mathcal{D} \theta')  d\theta'}$ $\propto p(\theta)p(\mathcal{D} \theta)$ We rarely compute the denominator.

Maximum A-Posteriori (MAP) Estimation: Maximizing the posterior (rather than the likelihood in maximum likelihood). Usually apply log first: it simplifies computations and doesn't affect the argmax (as  $\ln x$  is monotonic).

$$\operatorname*{argmax}_{\theta \in [0,1]} p(\theta | \mathcal{D}) = \operatorname*{argmax}_{\theta \in [0,1]} \ln p(\theta | \mathcal{D})$$

**Normalization:** Multiplying a f(x) by a constant to make it a probability density function (ie.  $\int_{-\infty}^{\infty} f(x) dx = 1$ ). Beta Distribution: Continuous distribution  $X \sim \text{Beta}(a, b)$  common in machine learning. See this for visualizations.

$$f_X(x) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} x^{a-1} (1-x)^{b-1} \quad \text{for } x \in [0,1]$$

$$f_X(x)$$
 is only defined on  $x \in [0,1]$ , 0 elsewhere  $f_X(x) = \frac{a}{x}$ 

 $\triangleright \mathbb{E}[X] = \frac{a}{a+b}$ 

 $\triangleright a = b = 1$  is the uniform distribution

 $\Gamma(x)=\int_{0}^{\infty}t^{x-1}e^{-t}dt=(x-1)!$  is the normalizer (ignore it)

 $\triangleright$  High values of a, b mean larger, narrower peaks.

eg. Flip a coin. We don't know how biased the coin is, so heta can be anything from 0 to 1.

Let the prior  $p(\theta)$  follow the beta distribution,

$$\begin{array}{l} \theta {\sim} \mathrm{Beta}(a,b) \\ p(\theta) \propto \theta^{a-1} (1-\theta)^{b-1} \end{array}$$

From before, we know the likelihood  $L(\theta)$  is

$$L(\theta) = p(\mathcal{D}|\theta) = \prod_{n=1}^N \theta^{x_n} (1-\theta)^{1-x_n} = \theta^{N_{\mathrm{H}}} (1-\theta)^{N_{\mathrm{T}}}$$

Therefore, calculating the posterior,

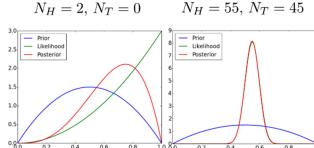
$$\begin{split} & \therefore p(\theta|\mathcal{D}) \propto p(\theta) f(\mathcal{D}|\theta) \\ & \propto [\theta^{a-1} (1-\theta)^{b-1}] [\theta^{N_{\mathrm{H}}} (1-\theta)^{N_{\mathrm{T}}}] \\ & = \theta^{N_{\mathrm{H}} + a - 1} (1-\theta)^{N_{\mathrm{T}} + b - 1} \end{split}$$

If you normalize this, this is equivalent to

$$\begin{split} \theta | \mathcal{D} \sim & \text{Beta}(N_{\text{H}} + a, N_{\text{T}} + b). \\ & \therefore \mathbb{E}[\theta | \mathcal{D}] = \frac{N_{\text{H}} + a}{N_{\text{H}} + N_{\text{T}} + a + b} \end{split}$$

Pseudo-Count: Numbers added to observations as to change a generated probability. In this case, it's a, bConjugates: Expressions represented by the same functional form. In this case, it's  $p(\theta)$  and  $p(\theta|\mathcal{D})$ .

Large data setting 
$$N_H = 55, N_T = 45$$



Finding the MAP estimation of  $\theta$ ,

$$\begin{split} \frac{d}{d\theta} \ln p(\theta|\mathcal{D}) &= \frac{d}{d\theta} \ln[\theta^{N_{\mathrm{H}}+a-1} (1-\theta)^{N_{\mathrm{T}}+b-1}] \\ &= \frac{d}{d\theta} [(N_{\mathrm{H}}+a-1) \ln \theta \\ &\quad + (N_{\mathrm{T}}+b-1) \ln (1-\theta)] \\ &= \frac{N_{\mathrm{H}}+a-1}{\theta} - \frac{N_{\mathrm{T}}+b-1}{1-\theta} = 0 \\ &\therefore \theta = \frac{N_{\mathrm{H}}+a-1}{N_{\mathrm{H}}+N_{\mathrm{T}}+a+b-2} \end{split}$$

Multivariate Random Variables	Linear Algebra Formulas
$\begin{split} \mathbb{E}[X] &= \begin{bmatrix} E[X_1] \\ \vdots \\ E[X_d] \end{bmatrix} \\ \operatorname{Var}[X] &= \mathbb{E}[(X - \mathbb{E}[X])(X - \mathbb{E}[X])^{\mathrm{T}}] \\ &= \mathbb{E}[XX^{\mathrm{T}}] - \mathbb{E}[X]\mathbb{E}[X]^{\mathrm{T}} \\ \operatorname{Cov}[X, Y] &= \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])^{\mathrm{T}}] \\ &= \mathbb{E}[XY^{\mathrm{T}}] - \mathbb{E}[X]\mathbb{E}[Y]^{\mathrm{T}} \end{split}$	$\frac{\partial}{\partial \vec{x}} [\mathbf{x}^{\mathrm{T}} A^{-1} \mathbf{x}] = 2A^{-1} \mathbf{x}$ $\frac{\partial}{\partial A} [\vec{x}^{\mathrm{T}} A^{-1} \vec{x}] = -A^{-1} \mathbf{x} \mathbf{x}^{\mathrm{T}} A^{-1}$ $\frac{\partial \det(A)}{\partial A} = \det(A) (A^{-1})^{\mathrm{T}}$

Normal/Gaussian Distribution: A continuous distribution. Popular in machine learning for simple calculations

Central Limit Theorem: The sum of many independent random variables is roughly Gaussian

	7 1
Univariate	Multivariate
$X \sim \text{Normal}(\mu, \sigma^2)$	$X{\sim}\mathrm{Normal}(oldsymbol{\mu},\Sigma)$
$\mu = \mathbb{E}[x]$ $\sigma^2 = \text{Var}[x]$	$ \begin{aligned} \boldsymbol{\mu} &= \mathbb{E}[\mathbf{x}] & \boldsymbol{\Sigma} &= \operatorname{Cov}(\mathbf{x}, \mathbf{x}) \\ &= \begin{bmatrix} \boldsymbol{\mu}_1 \\ \vdots \\ \boldsymbol{\mu}_d \end{bmatrix} & = \mathbb{E}[(\vec{x} - \vec{\mu})(\vec{x} - \vec{\mu})^{\mathrm{T}}] \\ &= \begin{bmatrix} \sigma_{1,1}^2 & \sigma_{1,2} & \cdots & \sigma_{1,d} \\ \sigma_{2,1} & \sigma_{2,2}^2 & \cdots & \sigma_{2,d} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{d,1} & \sigma_{d,2} & \cdots & \sigma_{d,d}^2 \end{bmatrix} \end{aligned} $
$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{\frac{-(x-\mu)^2}{2\sigma^2}}$	$f_X(\mathbf{x}) = \frac{1}{\det(\Sigma)^{\frac{1}{2}} (2\pi)^{\frac{d}{2}}} e^{-\frac{1}{2} (\mathbf{x} - \mathbf{\mu})^\mathrm{T} \Sigma^{-1} (\mathbf{x} - \mathbf{\mu})}$
Translate $N(0,1)$ by $\mu$ units right, stretch by factor of $\sigma$ to get $N(\mu, \sigma^2)$	Translate $N(0,I)$ in direction $\mu$ , scale by $\Sigma^{\frac{1}{2}}$ (ie. $\Sigma^{\frac{1}{2}}\Sigma^{\frac{1}{2}}=\Sigma$ ) to get $N(\mu,\Sigma)$

- $\Sigma$  is a symmetric and positive semi-definite matrix (see next section)
- $\mathbf{x}_n$  is uncorrelated with  $\mathbf{x}_m$  if covariance matrix  $\Sigma$  is diagonal (ie. all non-diagonal values are 0)
  - o Gaussian Naïve Bayes assumes that  $\Sigma$  is diagonal

eg. Let  $\mathcal{D} = \{(-2.5, -7.5), (-9.9, -14.9), (-12.1, -17.5), (-8.9, -13.9), (-6.0, -11.1)\} = \{x_1, \dots, x_N\}$  where N=5, a temperature distribution of highs/lows. Choose  $X\sim \mathrm{Normal}(\mu,\Sigma)$  for simplicity. Find  $\mu,\Sigma$ 

$$\begin{split} \ell(\mathbf{\mu}, \Sigma) &= \ln \prod_{n=1}^{N} p_{X|\mathbf{\mu}, \Sigma}(\mathcal{D}|\mathbf{\mu}, \Sigma) \\ &= \ln \prod_{n=1}^{N} \left[ \frac{1}{\det(\Sigma)^{\frac{1}{2}} (2\pi)^{\frac{d}{2}}} e^{-\frac{1}{2} (\mathbf{x}_n - \mathbf{\mu})^{\mathrm{T}} \Sigma^{-1} (\mathbf{x}_n - \mathbf{\mu})} \right] \\ &= \sum_{n=1}^{N} \ln \left[ \frac{1}{\det(\Sigma)^{\frac{1}{2}} (2\pi)^{\frac{d}{2}}} e^{-\frac{1}{2} (\mathbf{x}_n - \mathbf{\mu})^{\mathrm{T}} \Sigma^{-1} (\mathbf{x}_n - \mathbf{\mu})} \right] \\ &= \sum_{n=1}^{N} \left[ -\ln \left( \det(\Sigma)^{\frac{1}{2}} (2\pi)^{\frac{d}{2}} \right) + \ln e^{-\frac{1}{2} (\mathbf{x}_n - \mathbf{\mu})^{\mathrm{T}} \Sigma^{-1} (\mathbf{x}_n - \mathbf{\mu})} \right] \\ &= \sum_{n=1}^{N} \left[ -\ln \det(\Sigma)^{\frac{1}{2}} (2\pi)^{\frac{d}{2}} \right) + \ln e^{-\frac{1}{2} (\mathbf{x}_n - \mathbf{\mu})^{\mathrm{T}} \Sigma^{-1} (\mathbf{x}_n - \mathbf{\mu})} \right] \\ &= \sum_{n=1}^{N} \left[ -\ln \det(\Sigma)^{\frac{1}{2}} -\ln (2\pi)^{\frac{d}{2}} - \frac{1}{2} (\mathbf{x}_n - \mathbf{\mu})^{\mathrm{T}} \Sigma^{-1} (\mathbf{x}_n - \mathbf{\mu}) \right] \\ &= -\frac{1}{2} \sum_{n=1}^{N} [\ln \det(\Sigma) + d \ln 2\pi + (\mathbf{x}_n - \mathbf{\mu})^{\mathrm{T}} \Sigma^{-1} (\mathbf{x}_n - \mathbf{\mu})] \end{split}$$
 This formula is called empirical mean, where mean is based off past results rather than a theoretical distribution. The formula below is empirical covariance.

Taking the derivative with respect to  $\mu$ ,

$$\begin{split} \frac{\partial \ell}{\partial \vec{\mu}} &= -\frac{1}{2} \sum_{n=1}^{N} \frac{\partial}{\partial \vec{\mu}} [(\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu})] \\ &= -\sum_{n=1}^{N} \boldsymbol{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu}) = 0 \\ & \\ & \\ & \\ & \\ & \approx (-1.576, -2.596) \end{split}$$

mean is based off past results rather than a theoretical distribution. The formula below is empirical covariance.

Taking the derivative with respect to  $\Sigma$ ,

$$\begin{split} \frac{\partial \ell}{\partial \Sigma} &= -\frac{1}{2} \sum_{n=1}^{N} \frac{\partial}{\partial \Sigma} \left[ \ln \det(\Sigma) + (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \Sigma^{-1} (\mathbf{x}_n - \boldsymbol{\mu}) \right] \\ &= -\frac{1}{2} \sum_{n=1}^{N} \left[ \frac{\partial \ln \det(\Sigma)}{\partial \det(\Sigma)} \cdot \frac{\partial \det(\Sigma)}{\partial \Sigma} + \frac{\partial}{\partial \Sigma} (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \Sigma^{-1} (\mathbf{x}_n - \boldsymbol{\mu}) \right] \\ &= -\frac{1}{2} \sum_{n=1}^{N} \left[ \frac{1}{\det(\Sigma)} \cdot \det(\Sigma) \left( \Sigma^{-1} \right)^{\mathrm{T}} - \Sigma^{-1} (\mathbf{x}_n - \boldsymbol{\mu}) (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \Sigma^{-1} \right] \\ &= -\frac{1}{2} \sum_{n=1}^{N} \left[ \Sigma^{-1} - \Sigma^{-1} (\mathbf{x}_n - \boldsymbol{\mu}) (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \Sigma^{-1} \right] \\ &= -\frac{1}{2} \sum_{n=1}^{N} \Sigma^{-1} \left[ 1 - (\mathbf{x}_n - \boldsymbol{\mu}) (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \Sigma^{-1} \right] = 0 \\ &\sum_{n=1}^{N} \left[ 1 - (\mathbf{x}_n - \boldsymbol{\mu}) (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \Sigma^{-1} \right] = 0 \\ &\sum_{n=1}^{N} \left[ 1 - (\mathbf{x}_n - \boldsymbol{\mu}) (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \Sigma^{-1} \right] \Sigma = 0 \\ &\sum_{n=1}^{N} \left[ \Sigma - (\mathbf{x}_n - \boldsymbol{\mu}) (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \right] = 0 \\ &\sum_{n=1}^{N} \left[ \Sigma - (\mathbf{x}_n - \boldsymbol{\mu}) (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \right] = 0 \\ &\sum_{n=1}^{N} \left[ \Sigma - (\mathbf{x}_n - \boldsymbol{\mu}) (\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}} \right] = 0 \end{split}$$

Linear regression is just  $t | \mathbf{x} \sim \text{Normal}(\mathbf{w} \cdot \mathbf{x}, 1)$ with extra constants (which don't affect argmax)

 $L_2$  regularization is just  $\mathbf{w} \sim \text{Normal}(0, \eta I)$  with extra constants (where  $\eta \in \mathbb{R}$  is like  $\frac{1}{\lambda}$  from ridge regression)

$$\begin{split} \ln p(t|\mathbf{x}) &= \ln \prod_{n=1}^{N} \left( \frac{1}{\sigma \sqrt{2\pi}} e^{\frac{-(t_n - \mathbf{w} \cdot \mathbf{x})^2}{2(1)^2}} \right) & \ln p(\mathbf{w}) = \ln \left( \frac{1}{\det(\eta I)^{\frac{1}{2}} (2\pi)^{\frac{d}{2}}} e^{-\frac{1}{2} (\mathbf{w} - 0)^{\mathrm{T}} (\eta I)^{-1} (\mathbf{w} - 0)} \right) \\ &= -\sum_{n=1}^{N} \left( \ln \sigma + \frac{1}{2} \ln 2\pi + \frac{(\mathbf{w} \cdot \mathbf{x} - t_n)^2}{2} \right) & = -\frac{1}{2} \ln (\det \eta I) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \mathbf{w}^{\mathrm{T}} (\eta I)^{-1} \mathbf{w} \\ &= \mathrm{Constant} - \frac{1}{2} \sum_{n=1}^{N} (\mathbf{w} \cdot \mathbf{x} - t_n)^2 & = \mathrm{Constant} - \frac{1}{2\eta} \|\mathbf{w}\|^2 \end{split}$$

Finding the argmax of -x (here) is equivalent to finding the argmin of x (linear regression)

#### Gaussian Discriminant Analysis: Classification algorithm that finds the best Gaussian fit to data $p(\mathbf{x}|t)$

$$p(\mathbf{x}|t) = f_X(\mathbf{x})$$

Each target has a separate  $\Sigma_k$  and  $\mu_k$ , which are the parameters to find using maximum likelihood.

eg. Multivariate Gaussian discriminant analysis for binary classification.  $t \in \{0,1\}$ 

Let  $\phi_k$  be percentage of inputs with target  $k \in \{0,1\}$ . At test-time, for any new query point,

$$\phi_k = \frac{1}{N} \sum_{n=1}^N \mathbb{I}(t_n = k)$$

Define  $\overrightarrow{\mu_k}$ ,  $\Sigma_k$  as empirical mean/covariance of all inputs with target k. Obtainable via max-likelihood.

$$\begin{split} \mathbf{\mu}_k &= \frac{\sum_{n=1}^N \mathbb{I}(t_n = k)\mathbf{x}_n}{\sum_{n=1}^N \mathbb{I}(t_n = k)} \\ \Sigma_k &= \frac{\sum_{n=1}^N \mathbb{I}(t_n = k)(\mathbf{x}_n - \mathbf{\mu}_k)(\mathbf{x}_n - \mathbf{\mu}_k)^\mathrm{T}}{\sum_{n=1}^N \mathbb{I}(t_n = k)} \end{split}$$

Using Bayes' rule (and logarithms on both sides).

$$\ln p(t = k|\mathbf{x}) = \ln \frac{p(t = k)p(\mathbf{x}|t = k)}{p(\mathbf{x})}$$

$$= \ln p(t=k) + \left[ -\frac{1}{2}\ln \det \Sigma_k - \frac{d}{2}\ln 2\pi - \frac{1}{2}(\mathbf{x} - \mathbf{\mu}_k)^{\mathrm{T}} \Sigma_k^{-1}(\mathbf{x} - \mathbf{\mu}_k) + k \ln \phi_k + (1-k) \ln (1-\phi_k) \right] - \ln p(\mathbf{x})$$

We also know that

 $p(t = k) = \begin{cases} \phi_1 & k = 1\\ 1 - \phi_1 & k = 0\\ 1 - \phi_0 & k = 0\\ 1 - \phi_0 & k = 1\\ 0 = \phi_k^k (1 - \phi_k)^{1-k} \end{cases}$ 

 $p(\mathbf{x}|t) = f_X(x) = \frac{1}{\det(\Sigma)^{\frac{1}{2}}(2\pi)^{\frac{d}{2}}} e^{-\frac{1}{2}(\mathbf{x}_n - \mathbf{\mu}_k)^{\mathrm{T}} \Sigma^{-1}(\mathbf{x}_n - \mathbf{\mu}_k)}$ 

In probabilistic modelling, the decision boundary is the solution to  $p(t = 0|\mathbf{x}) = 0.5 = p(t = 1|\mathbf{x})$ .

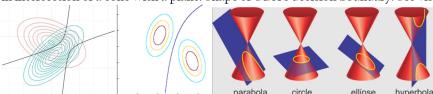
$$\ln p(t=0|\mathbf{x}) = \ln p(t=1|\mathbf{x})$$

$$-\frac{1}{2}\ln\det\Sigma_{0} - \frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_{0})^{\mathrm{T}}\Sigma_{0}^{-1}(\mathbf{x} - \boldsymbol{\mu}_{0}) + \ln(1 - \phi_{0}) = -\frac{1}{2}\ln\det\Sigma_{1} - \frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_{1})^{\mathrm{T}}\Sigma_{1}^{-1}(\mathbf{x} - \boldsymbol{\mu}_{1}) + \ln\phi_{1}$$

The terms without  $\vec{x}$  will be constants, so we simplify the equation to

$$(\mathbf{x} - \boldsymbol{\mu}_0)^T \boldsymbol{\Sigma}_0^{-1} (\mathbf{x} - \boldsymbol{\mu}_0) = (\mathbf{x} - \boldsymbol{\mu}_1)^T \boldsymbol{\Sigma}_1^{-1} (\mathbf{x} - \boldsymbol{\mu}_1) + \text{Constant}$$
  
$$\mathbf{x}^T \boldsymbol{\Sigma}_0^{-1} \mathbf{x} - 2\boldsymbol{\mu}_0^T \boldsymbol{\Sigma}_0^{-1} \mathbf{x} = \mathbf{x}^T \boldsymbol{\Sigma}_1^{-1} \mathbf{x} - 2\boldsymbol{\mu}_1^T \boldsymbol{\Sigma}_1^{-1} \mathbf{x} + \text{Constant}$$

Conic Section: An intersection of a cone with a plane. Shape of GDA's decision boundary. See visualizations here.



	parabola circle ellipse h	yperbola
Assumption	Result	Decision Boundary
Shared Covariances	$\mathbf{x}^{\mathrm{T}} \Sigma^{-1} \mathbf{x} - 2\mathbf{\mu}_{0}^{\mathrm{T}} \Sigma^{-1} \mathbf{x} = \mathbf{x}^{\mathrm{T}} \Sigma^{-1} \mathbf{x} - 2\mathbf{\mu}_{1}^{\mathrm{T}} \Sigma^{-1} \mathbf{x} + \text{Constant}$	
(ie. $\Sigma_0 = \Sigma_1 = \Sigma$ )	$\mathbf{\mu}_0^{\mathrm{T}} \Sigma^{-1} \mathbf{x} = \mathbf{\mu}_1^{\mathrm{T}} \Sigma^{-1} \mathbf{x} + \mathrm{Constant}$	
! ! !	$(\boldsymbol{\mu}_0 - \boldsymbol{\mu}_1)^{\mathrm{T}} \Sigma^{-1} \mathbf{x} = \mathrm{Constant}$	
i ! !	This is a linear decision boundary! Logistic regression!	
Naïve Bayes	Contours only vary via scaling, cannot be rotated. Still produces	
(ie. Diagonal $\Sigma$ )	conic section decision boundaries.	
  - 	For high-dimensional $\vec{x} \in \mathbb{R}^D$ , reduces size of $\Sigma_k$ from $D^2K$	
  - 	parameters to $D$ parameters.	
Spherical/Isotropic	Contours will be spheres, and decision boundaries are spherical. If	
Covariances	spheres are the same size (ie. $\sigma_0=\sigma_1$ ), then decision boundary is	
(ie. $\Sigma = \sigma^2 I$ )	a line bisecting $\overrightarrow{\mu_0}$ and $\overrightarrow{\mu_1}$ .	
 	$C = (\mathbf{x} - \mathbf{\mu}_0)^{\mathrm{T}} \Sigma^{-1} (\mathbf{x} - \mathbf{\mu}_0) - (\mathbf{x} - \mathbf{\mu}_1)^{\mathrm{T}} \Sigma^{-1} (\mathbf{x} - \mathbf{\mu}_1)$	
  -  -	$= (\mathbf{x} - \mathbf{\mu}_0)^{\mathrm{T}} (\sigma^2 I)^{-1} (\mathbf{x} - \mathbf{\mu}_0) - (\mathbf{x} - \mathbf{\mu}_1)^{\mathrm{T}} (\sigma^2 I)^{-1} (\mathbf{x} - \mathbf{\mu}_1)$	X 9
	$= \frac{1}{\sigma^2} [\ \mathbf{x} - \mathbf{\mu}_0\ ^2 - \ \mathbf{x} - \mathbf{\mu}_1\ ^2]$	

Gaussian Discriminant Analysis, Shared Covariances	Logistic Regression
<ul> <li>Stronger modelling assumption that class-conditional data is multivariate Gaussian</li> <li>If true, GDA is the most efficient model</li> <li>If not, then poorer prediction accuracy</li> </ul>	<ul> <li>Usually better than GDA for non-Gaussian distributions (ie. vast majority of distributions</li> </ul>
➤ Handles missing features better  Full Covariances (acc 0.805)	Shared Covariance (acc 0.717)
Naive Bayes (acc 0.780)	Logistic regression (acc 0.722)

# Principal Component Analysis

Field: A set F where the addition, subtraction, multiplication, and division of any elements of F are in F (eg. ℝ, ℕ, ℂ) Vector Space: Over a field  $\mathbb{F}$ , the set V defined by two operations,  $\mathbf{u}, \mathbf{v} \in V, \alpha \in \mathbb{F} \Rightarrow \mathbf{u} + \mathbf{v} \in V, \alpha \mathbf{v} \in V$ , where:

$$(\alpha + \beta)\mathbf{v} = \alpha\mathbf{v} + \beta\mathbf{v}$$
  
 $(\alpha\beta)\mathbf{v} = \alpha(\beta\mathbf{v})$ 

$$(\mathbf{u} + \mathbf{v}) + \mathbf{w} = \mathbf{u} + (\mathbf{v} + \mathbf{w})$$

$$\triangleright$$
  $\mathbf{0} \in V$   
 $\triangleright$   $-\mathbf{v} \in V$ 

Affine Space: Informally, a vector space without any concept of origin.

**Projection:** Of  $\mathbf{x} \in \mathbb{R}^D$  on subspace  $\mathcal{S}$  (K-dimensional, in  $\mathbb{R}^D$ , where K < D), the nearest point of  $\mathcal{S}$  to  $\mathbf{x}$ .

$$\triangleright$$
 For  $S = \operatorname{span}(\mathbf{u})$ ,

$$\operatorname{proj}_{\mathcal{S}}(\mathbf{x}) = \operatorname{vcomp}_{\overrightarrow{u}}(\mathbf{x}) = (\mathbf{x} \cdot \mathbf{u})\mathbf{u}$$

$$ightharpoonup$$
 For  $S = \operatorname{span}(\mathbf{u}_1, \dots, \mathbf{u}_K)$ ,

$$\operatorname{proj}_{\mathcal{S}}(\mathbf{x}) = \sum_{k=1}^{K} \operatorname{vcomp}_{\mathbf{u}_{k}}(\mathbf{x}) = \sum_{k=1}^{K} (\mathbf{x} \cdot \mathbf{u}_{k}) \mathbf{u}_{k}$$

o In vector form, we have 
$$\operatorname{proj}_{\mathcal{S}}(\mathbf{x}) = UU^{\mathrm{T}}\mathbf{x}$$

where 
$$U = [\mathbf{u}_1 \quad \cdots \quad \mathbf{u}_K]$$
 is  $D \times K$ 

For 
$$S =$$
 an affine space,

$$\mathrm{proj}_{\mathcal{S}}(\mathbf{x}) = \sum_{k=1}^K ig( (\mathbf{x} - \mathbf{\mu}) \cdot \mathbf{u}_k ig) \mathbf{u}_k + \mathbf{\mu}$$
 where  $\vec{\mu} \in \mathcal{S}$  is an arbitrary origin

where 
$$\vec{\mu} \in \mathcal{S}$$
 is an arbitrary origin

In vector form, we have 
$$\operatorname{proj}_{\mathcal{S}}(\mathbf{x}) = UU^{\mathrm{T}}(\mathbf{x} - \mathbf{\mu}) + \mathbf{\mu}$$

where 
$$U = \begin{bmatrix} \mathbf{u}_1 & \cdots & \mathbf{u}_K \end{bmatrix}$$
 is  $D \times K$ 

**Reconstruction:** The term  $\mathbf{x} = \text{proj}_{\mathcal{S}}(\mathbf{x}) \in \mathcal{S}$ 

• Note that **x** and **x** have the same means

Representation/Code: A  $K \times 1$  vector **z** representing a mapping of data to a space that's easier to manipulate/visualize

- In the above case,  $\mathbf{z} = U^{\mathrm{T}}\mathbf{x}$  for vector spaces,  $\mathbf{z} = U^{\mathrm{T}}(\mathbf{x} \mathbf{\mu})$  for affine spaces
- If  $K \ll D$ , it's computationally cheaper to work with the representation than **x**

Representation Learning: When a learning algorithm tries to learn a representation

Dimensionality Reduction: Mapping data to a lower-dimensional space

Saves computation/memory, reduces overfitting, and can allow visualization (if reduced to 3D)

Orthonormal: Two vectors  $\mathbf{u}, \mathbf{v}$  if both are unit vectors and orthogonal (ie.  $\mathbf{u} \cdot \mathbf{v} = 0$ )

Orthogonal Matrix: Matrix Q if 
$$Q^{T}Q = QQ^{T} = I$$
, AKA  $Q^{T} = Q^{-1}$ 

• Q is orthogonal  $\Leftrightarrow Q'$ s columns are orthogonal

Eigenvector: Of square matrix A, vector  $\vec{x}$  such that  $A\mathbf{x} = \lambda \mathbf{x}$  for some  $\lambda$ , the corresponding eigenvalue.

- Matrices of size  $D \times D$  have at most D unique eigenvalues
- Symmetric matrices have D linearly independent eigenvectors  $\mathbf{x}_d$  (where  $\lambda_d \in \mathbb{R}$ ) which form a basis for  $\mathbb{R}^D$  (i.e. their span is  $\mathbb{R}^D$ ) and can be chosen to be orthonormal (by changing the magnitude of the  $\mathbf{x}_d$ ).

Spectral Decomposition: AKA eigendecomposition, a way to factor a symmetric  $D \times D$  matrix A

$$A = Q \Lambda Q^{\mathrm{T}} = Q \Lambda Q^{-1}$$
 
$$\blacktriangleright \ \ Q = \begin{bmatrix} \mathbf{q}_1 & \cdots & \mathbf{q}_D \end{bmatrix} \text{ is an orthogonal matrix } (\mathbf{q}_d \text{ are } A \text{'s eigenvectors})$$
 
$$\blacktriangleright \ \ \Lambda = \begin{bmatrix} \lambda_1 & 0 & 0 \\ 0 & \ddots & 0 \\ 0 & 0 & \lambda_D \end{bmatrix} \text{ is a diagonal matrix } (\lambda_d \text{ are } A \text{'s eigenvalues})$$

Let 
$$q_d$$
 be  $A$ 's eigenvectors, 
$$A\mathbf{q}_d = \lambda_d\mathbf{q}_d$$
 Therefore

$$AQ = Q\Lambda$$
$$A = Q\Lambda Q^{-1}$$

We can convert between  ${\mathcal E}$  and eigenvector coordinates Q:

$$\begin{split} [\mathbf{x}]_{\mathcal{E}} &= Q^{\mathrm{T}}[\mathbf{x}]_{\mathcal{Q}} \\ [\mathbf{x}]_{\mathcal{Q}} &= Q[\mathbf{x}]_{\mathcal{E}} \end{split}$$

The transformation A rescales individual dimensions by a factor  $\lambda_d$ :

$$\begin{split} [\mathbf{x}]_{\mathcal{E}} &= [x_1]_{\mathcal{Q}} \mathbf{q}_1 + \dots + [x_D]_{\mathcal{Q}} \mathbf{q}_D \\ A[\mathbf{x}]_{\mathcal{E}} &= [x_1]_{\mathcal{Q}} A \mathbf{q}_1 + \dots + [x_D]_{\mathcal{Q}} A \mathbf{q}_D \\ &= [x_1]_{\mathcal{Q}} \lambda_1 \mathbf{q}_1 + \dots + [x_D]_{\mathcal{Q}} \lambda_D \mathbf{q}_D \end{split}$$

All symmetric matrices are diagonal under the right coordinate system.

Quadratic Form: A function  $f(\mathbf{x}) = \mathbf{x}^T A \mathbf{x}$  where A is symmetric

- Positive Definite:
- $A \succ 0$ , if  $\mathbf{x}^T A \mathbf{x} > 0$  for all  $\mathbf{x} \neq \mathbf{0}$
- Positive Semi-Definite: A > 0, if  $\mathbf{x}^T A \mathbf{x} \ge 0$  for all  $\mathbf{x} \ne \mathbf{0}$
- Negative Semi-Definite:  $A \leq 0$ , if  $\mathbf{x}^T A \mathbf{x} < 0$  for all  $\mathbf{x} \neq \mathbf{0}$
- Negative Definite: Indefinite:
- $A \prec 0$ , if  $\mathbf{x}^T A \mathbf{x} < 0$  for all  $\mathbf{x} \neq \mathbf{0}$ If  $\mathbf{x}^T A \mathbf{x}$  can be positive or negative









Singular/Degenerate: A non-invertible matrix.

Matrix Square Root: Of a positive semi-definite (PSD) matrix A, the matrix  $A^{\frac{1}{2}}$  where  $A^{\frac{1}{2}}A^{\frac{1}{2}}=A$ 

$$\triangleright AA^{\mathrm{T}}$$
 is PSD

 $ightharpoonup A ext{ is PSD} \Rightarrow \mathbb{E}[A] ext{ is PSD}$ 

$$\blacktriangleright \ \ A \text{ is PD} \Leftrightarrow \lambda_d > 0$$

$$ightharpoonup A ext{ is PSD} \Leftrightarrow \lambda_d \geq 0$$

For symmetric matrices,

For symmetric matrices, 
$$A^k = Q \Lambda^k Q^{\mathrm{T}}$$
 
$$A^{-1} = Q \Lambda^{-1} Q^{\mathrm{T}}$$
 
$$A^{\frac{1}{2}} = Q \Lambda^{\frac{1}{2}} Q^{\mathrm{T}} \text{ (if } A \text{ is PSD)}$$
 
$$\det A = \prod_{d=1}^D \lambda_d$$

- ➤ Non-negative linear combinations of PSD matrices are PSD
- $\triangleright$  A is PD  $\Leftrightarrow$  contour lines of  $\mathbf{x}^{\mathrm{T}}A\mathbf{x}$  is elliptical (principal axes of ellipses aligned with eigenvectors)
- $\triangleright$  A is PD and diagonal  $\Leftrightarrow$  contours lines of  $\vec{x}^T A \vec{x}$  is elliptical (principle axes of ellipses aligned with axes)

#### Determinant Properties

Frobenius Norm: Norm generalized to a matrix, 
$$\|A\|_F = \sqrt{\sum_{m=1}^M \sum_{n=1}^N A_{n,m}^2}$$

Principal Component Analysis (PCA): Unsupervised learning algorithm that performs linear dimensionality reduction to find the most "important" dimensions (ie. principle components) in the data.

- Set  $\mu \in \mathbb{R}^D$  to the empirical mean of the  $\mathbf{x}_n$  to center our data (it <u>simplifies</u> computations).
- Choose the most important matrix U (size  $D \times K$ ) with orthonormal columns with two <u>equivalent criteria</u>:

$$\begin{aligned} & \operatorname{argmin}_{U} \frac{1}{N} \textstyle \sum_{n=1}^{N} \lVert \mathbf{x}_{n} - \underline{\mathbf{x}}_{n} \rVert^{2} \\ & \operatorname{argmax}_{U} \frac{1}{N} \textstyle \sum_{n=1}^{N} \lVert \underline{\mathbf{x}}_{n} - \underline{\boldsymbol{\mu}} \rVert^{2} \end{aligned}$$

Maximizing reconstruction variance

$$\operatorname{argmax}_{U} \frac{1}{N} \sum_{n=1}^{N} \|\mathbf{x}_{n} - \mathbf{\mu}\|^{2}$$

The goal is finding an accurate low-dimensional ( $\mathbb{R}^K$ ) representation of data ( $\mathbb{R}^D$ )

	` '	1 , ,
Let $X$ be a $N \times D$ matrix representing	$X = (\mathbf{x}_1^{\mathrm{T}}, \dots, \mathbf{x}_N^{\mathrm{T}})$	The reconstruction error can be vectorized into
all data as vectors $\mathbf{x}_n \in \mathbb{R}^D$	, 1	$\frac{1}{N} \sum_{n=1}^{N} \ \mathbf{x}_{n} - \underline{\mathbf{x}}_{n}\ ^{2} = \frac{1}{N} \sum_{n=1}^{N} \ \mathbf{x}_{n} - U\mathbf{z}_{n}\ ^{2}$
Let $Z$ be a $N \times K$ matrix representing	$Z = (\mathbf{z}_1^{\mathrm{T}}, \dots, \mathbf{z}_N^{\mathrm{T}})$	$\frac{1}{N} \sum_{n=1}^{N} \ \mathbf{x}_n - \underline{\mathbf{x}}_n\  - \frac{1}{N} \sum_{n=1}^{N} \ \mathbf{x}_n - U\mathbf{z}_n\ $
simplified data as vectors $\mathbf{z}_n \in \mathbb{R}^K$		$= \frac{1}{N} \ X - ZU^{\mathrm{T}}\ _{F}^{2}$
Let $U$ be a $D \times K$ matrix	$U = [\mathbf{u}_1  \cdots  \mathbf{u}_K]$	
	•	So $X \approx ZU^{\mathrm{T}} \otimes Z$ is a rank- $K$ approximation of $X$ .

- K is a hyperparameter. For K = 1, where  $U = \mathbf{u}$ ,
  - $\circ \quad \text{We can show } \tfrac{1}{N} \sum_{n=1}^N \lVert \mathbf{x}_n \boldsymbol{\mu} \rVert^2 = \mathbf{a}^{\mathrm{T}} \Lambda \mathbf{a} \text{ where } \mathbf{a} = Q^{\mathrm{T}} \mathbf{u}$
  - We can show the optimal choice is  ${f a}={f e}_1,$  where  $\{{f e}_1,\ldots,{f e}_D\}$  is standard basis, AKA  ${f u}={f q}_1$ 
    - We can show the i-th principle component is the eigenvector associated with the i-th largest  $\lambda_d$  (ie. The *i*-th principal component is the *i*-th eigenvalue of  $\Sigma$ )
  - Components of **z** are uncorrelated with each other (ie. the covariance matrix is diagonal)

$$\begin{split} \operatorname{Cov}(\mathbf{z}) &= \operatorname{Cov} \left( U^{\operatorname{T}}(\mathbf{x} - \boldsymbol{\mu}) \right) \\ &= U^{\operatorname{T}} \operatorname{Cov}(\mathbf{x}) U \text{ (as } \operatorname{Cov}(A\mathbf{x}) = A \operatorname{Cov}(\mathbf{x}) A^{\operatorname{T}} \text{ and } \boldsymbol{\mu} \text{ is constant)} \\ &= U^{\operatorname{T}} \Sigma U \\ &= U^{\operatorname{T}} Q \Lambda Q^{\operatorname{T}} U \text{ (by spectral decomposition)} \\ &= [I \quad 0] \Lambda [I \quad 0]^{\operatorname{T}} \text{ (since columns of } U, Q \text{ orthonormal and can be rearranged)} \\ &= \operatorname{top left} \ K \times K \text{ block of } \Lambda \end{split}$$

Singular-Value Decomposition (SVD): Generalization of spectral decomposition to any real  $N \times D$  matrix X.

		$A = A \triangle D$	
A	$N \times N$	Orthonormal columns of left singular vectors	Left Singular Vectors: Eigenvectors of $XX^{\mathrm{T}}$
$\sum$	$N \times D$	Diagonal matrix of singular values	Singular Values: Square roots of eigenvalues
		conventionally sorted in descending order	of $XX^{\mathrm{T}}$ and $X^{\mathrm{T}}X$
B	$D \times D$	Orthonormal columns of right singular vectors	Right Singular Vectors: Eigenvectors of $X^{T}X$

- Since  $\Sigma$  is not guaranteed to be square, set values in the "extra" rows/columns to 0s.
- It is possible to show the first n singular vectors are the first n principal components (ie.  $Z = A\Sigma$ )

Derivation – Equivalence of minimizing reconstruction error, maximizing reconstructing variance

Recall that

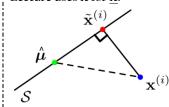
$$\begin{split} \mathbf{z} &= U^{\mathrm{T}}(\mathbf{x} - \mathbf{\mu}) \\ \underline{\mathbf{x}} &= U\mathbf{z} + \mathbf{\mu} \\ &= UU^{\mathrm{T}}(\mathbf{x} - \mathbf{\mu}) + \mathbf{\mu} \end{split}$$

Remember that U has orthonormal columns, so

$$U^{\mathrm{T}}U = I$$

 $\mathbf{x}$  is similar to  $\mathbf{x}$ , but because  $UU^{\mathrm{T}} = I_K, \mathbf{\underline{x}}$  has 0 in some components.

Lecture uses  $\tilde{\mathbf{x}}$  for  $\mathbf{x}$ .



We will show  $\mathbf{x} - \boldsymbol{\mu}$  and  $\mathbf{x} - \mathbf{x}$  are orthogonal.

$$\begin{aligned} (\underline{\mathbf{x}} - \boldsymbol{\mu})^{\mathrm{T}} (\underline{\mathbf{x}} - \mathbf{x}) &= (U\mathbf{z})^{\mathrm{T}} (U\mathbf{z} + \boldsymbol{\mu} - \mathbf{x}) \\ &= \mathbf{z}^{\mathrm{T}} U^{\mathrm{T}} (U\mathbf{z} - (\mathbf{x} - \boldsymbol{\mu})) \\ &= \mathbf{z}^{\mathrm{T}} U^{\mathrm{T}} U\mathbf{z} - \mathbf{z}^{\mathrm{T}} U^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}) \\ &= \mathbf{z}^{\mathrm{T}} \mathbf{z} - (U\mathbf{z})^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}) \\ &= (U^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}))^{\mathrm{T}} U^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}) - (UU^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}))^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}) \\ &= (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} U U^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}) - (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}) \\ &= (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}) - (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} (\mathbf{x} - \boldsymbol{\mu}) \\ &= 0 \end{aligned}$$

Thus there's a right triangle between  $\mathbf{x}$ ,  $\mathbf{x}$ , and  $\mathbf{\mu}$ , so by Pythagorean theorem,

$$\begin{split} \|\underline{\mathbf{x}} - \mathbf{\mu}\|^2 + \|\mathbf{x} - \underline{\mathbf{x}}\|^2 &= \|\mathbf{x} - \mathbf{\mu}\|^2 \\ \frac{1}{N} \sum_{n=1}^N &\|\underline{\mathbf{x}}_n - \mathbf{\mu}\|^2 + \frac{1}{N} \sum_{n=1}^N &\|\mathbf{x}_n - \underline{\mathbf{x}}_n\|^2 = \frac{1}{N} \sum_{n=1}^N &\|\mathbf{x}_n - \mathbf{\mu}\|^2 \\ \text{LS can be expressed in terms of } U, \text{ while RS can't, so RS is a constant. Thus} \end{split}$$

$$\frac{1}{N} \sum_{n=1}^{N} \lVert \mathbf{x}_n - \mathbf{\mu} \rVert^2 = \text{Constant} - \frac{1}{N} \sum_{n=1}^{N} \lVert \mathbf{x}_n - \underline{\mathbf{x}}_n \rVert^2$$

Reconstruction Variance = Constant - Reconstruction Error

Thus, maximizing LS is the same as minimizing RS.

Derivation – Optimization for principal component analysis, K=1

Remember that

$$\begin{aligned} & \underline{\mathbf{x}}, \mathbf{\mu} \in \mathcal{S} \\ & \mathbf{x} \in \mathbb{R}^D \\ & \mathbf{z} = U^{\mathrm{T}} \mathbf{x} \in \mathbb{R}^K \\ & U \text{ is } D \times K \end{aligned}$$

First, we note that

$$\begin{aligned} \|\underline{\mathbf{x}} - \mathbf{\mu}\|^2 &= \|U\mathbf{z}\|^2 \\ &= (U\mathbf{z})^{\mathrm{T}} U\mathbf{z} \\ &= \mathbf{z}^{\mathrm{T}} U^{\mathrm{T}} U\mathbf{z} \\ &= \mathbf{z}^{\mathrm{T}} \mathbf{z} \\ &= \|\mathbf{z}\|^2 \end{aligned}$$

In the case K = 1, the  $\operatorname{matrix} U \text{ of orthogonal} \\$ columns is a unit vector **u** and z is a scalar.

$$\begin{split} z &= (\mathbf{x} - \mathbf{\mu}) \cdot \mathbf{u} \\ z^2 &= \|\underline{\mathbf{x}} - \mathbf{\mu}\|^2 \end{split}$$

Let  $\{\mathbf{e}_D, \dots, \mathbf{e}_D\}$  be the standard basis.

We will then simplify reconstruction variance and maximize it,

$$\begin{split} \frac{1}{N} \sum_{n=1}^{N} & \|\mathbf{x}_{n} - \mathbf{\mu}\|^{2} = \frac{1}{N} \sum_{n=1}^{N} z_{n}^{2} \\ &= \frac{1}{N} \sum_{n=1}^{N} ((\mathbf{x}_{n} - \mathbf{\mu}) \cdot \mathbf{u})^{2} \\ &= \frac{1}{N} \sum_{n=1}^{N} \mathbf{u}^{\mathrm{T}} (\mathbf{x}_{n} - \mathbf{\mu}) (\mathbf{x}_{n} - \mathbf{\mu})^{\mathrm{T}} \mathbf{u} \quad (\text{as } (\mathbf{u} \cdot \mathbf{v})^{2} = \mathbf{u}^{\mathrm{T}} \mathbf{v} \mathbf{v}^{\mathrm{T}} \mathbf{u}) \\ &= \mathbf{u}^{\mathrm{T}} \left[ \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_{n} - \mathbf{\mu}) (\mathbf{x}_{n} - \mathbf{\mu})^{\mathrm{T}} \right] \mathbf{u} \\ &= \mathbf{u}^{\mathrm{T}} \Sigma \mathbf{u} \quad (\text{where } \Sigma \text{ is } D \times D) \\ &= \mathbf{u}^{\mathrm{T}} Q \Lambda Q^{\mathrm{T}} \mathbf{u} \quad (\text{by spectral decomposition, } \Sigma \text{ is symmetric}) \\ &= \mathbf{a}^{\mathrm{T}} \Lambda \mathbf{a} \quad (\text{where } \mathbf{a} = Q^{\mathrm{T}} \mathbf{u} \in \mathbb{R}^{D}) \\ &= [a_{1} \quad \cdots \quad a_{D}] \begin{bmatrix} \lambda_{1} a_{1} \\ \vdots \\ \lambda_{D} a_{D} \end{bmatrix} \\ &= \sum_{d=1}^{D} \lambda_{d} a_{d}^{2} \end{split}$$

When factoring  $\Sigma$  into  $Q\Lambda Q^{\mathrm{T}}$ , we arrange  $\Lambda$  such that the  $\lambda_d$  are ordered:

$$\begin{split} & \varLambda_{1,1} \geq \cdots \geq \varLambda_{D,D}^{u} \\ & \lambda_{1} \geq \cdots \geq \lambda_{D} \end{split}$$

Note that 
$$\mathbf{a}$$
 is a unit vector:  $\|\mathbf{a}\|^2 = \mathbf{a}^{\mathrm{T}}\mathbf{a}$ 

$$|\mathbf{r}| = \mathbf{a}^{\mathrm{T}} \mathbf{a}$$

$$= \mathbf{u}^{\mathrm{T}} Q Q^{\mathrm{T}} \mathbf{u}$$

$$= \mathbf{u}^{\mathrm{T}} \mathbf{u}$$

$$= \|\mathbf{u}\|^2 = 1$$

To maximize variance, set  $\mathbf{a} = \mathbf{e}_1$ , and  $\mathbf{a} = Q^{\mathrm{T}}\mathbf{u} \Rightarrow \mathbf{u} = Q\mathbf{a} = \mathbf{q}_1$ .

The *i*th principal component is the eigenvector with the *i*th largest eigenvalue of  $\Sigma$ .

**Sparse Matrix:** A matrix in which the majority of values are 0.

Partially Observed Matrix: A matrix in which not all values are known (eg. matrix of users to movie ratings).

Matrix Completion: The task of filling in missing entries of a partially observed matrix

Observed: A variable that is known (eg.  $\mathbf{x}$ ). It is partially observed if only some components are known.

Latent Factor/Hidden Variable: An unobserved variable  $\mathbf{z} \in \mathbb{R}^K$  where  $\mathbf{z} = f(\mathbf{x})$  for some  $f: \mathbb{R}^D \to \mathbb{R}^K$  with K < D**Latent Factor Models:** Models that tries to find K latent factors of all  $\mathbf{x} \in \mathbb{R}^D$  and predict outputs as a function of those K latent factors (K is a hyperparameter).

Formally, we find representation  $\mathbf{z} \in \mathbb{R}^K$  of  $\mathbf{x} \in \mathbb{R}^D$ 

Block Coordinate Descent/Alternating Minimization: An iterative algorithm that optimizes two variables x, y where on alternating steps, one of x, y is set constant while the other is optimized.

Alternating Least Squares: Alternating minimization using least squares distance

eg. Matrix R of N users to D movie ratings, partially observed as most users don't watch the majority of available movies. We want to predict missing entries; predicted ratings of never seen-before movies for a user.

Let  $\mathbf{u}_n \in \mathbb{R}^K$  be the latent factors of the n-th user Let  $\mathbf{z}_d \in \mathbb{R}^K$  be the latent factors of the d-th movie Let  $\mathbf{z}_d \in \mathbb{R}^K$  be the latent factors of the d-th movie The n-th user's rating for the d-th movie is  $R_{n,d} \approx \mathbf{u}_n \cdot \mathbf{z}_d$   $\frac{1}{N} \|R - ZU^{\mathrm{T}}\|_F^2 = \frac{1}{N} \sum_{n=1}^N \sum_{d=1}^D (R_{n,d} - \mathbf{u}_n \cdot \mathbf{z}_d)^2$ In matrix form,  $R \approx UZ^{\mathrm{T}}$  where

$$U = \begin{bmatrix} \mathbf{u}_1^{\mathrm{T}} \\ \vdots \\ \mathbf{u}_N^{\mathrm{T}} \end{bmatrix} \text{ is } N \times K$$
$$Z = \begin{bmatrix} \mathbf{z}_1^{\mathrm{T}} \\ \vdots \\ \mathbf{z}_D^{\mathrm{T}} \end{bmatrix} \text{ is } D \times K$$

We can write reconstruction error as

$$\frac{1}{N}\|R - ZU^{\mathrm{T}}\|_F^2 = \frac{1}{N} \sum_{n=1}^N \sum_{d=1}^D (R_{n,d} - \mathbf{u}_n \cdot \mathbf{z}_d)^2$$

But most values of  $R_{n,d}$  are blank! Thus let

$$R' = \{(n, d): R_{n, d} \text{ is observed}\}$$

Then we minimize error on R' with

$$\operatorname{argmin}_{U,Z} \left( \frac{1}{2} \sum_{(n,d) \in R'} (R'_{n,d} - \mathbf{u}_n \cdot \mathbf{z}_d)^2 \right)$$

And use the best U, Z to extrapolate to  $R \setminus R'$ .

- Previously,  $\mathbf{z} = U^{\mathrm{T}}(\mathbf{x} \mathbf{\mu})$ , so  $\mathbf{z}$  depended on U and  $\mathrm{argmin}_{U,Z} = \mathrm{argmin}_U$ . Now, we find  $\mathrm{argmin}_{U,Z}$
- Unfortunately,  $\frac{1}{2}\sum_{(n,d)\in R'}(R'_{n,d}-\mathbf{u}_n\cdot\mathbf{z}_d)^2$  is non-convex

We can re-write the objective function as

$$\frac{1}{2}\sum_{(n,d)\in R'}(R'_{n,d}-\mathbf{u}_n\cdot\mathbf{z}_d)^2=\frac{1}{2}\sum_{n=1}^N\sum_{d\in\{d:(n,d)\in R'\}}(R'_{n,d}-\mathbf{u}_n\cdot\mathbf{z}_d)^2=\frac{1}{2}\sum_{d=1}^D\sum_{n\in\{n:(n,d)\in R'\}}(R'_{n,d}-\mathbf{u}_n\cdot\mathbf{z}_d)^2$$
 For simplicity, define  $D_n=\{d:(n,d)\in R'\}$ 

$$A = \sum_{d \in D} (R'_{n,d} - \mathbf{u}_n^{\mathrm{T}} \mathbf{z}_d)^2 \qquad \qquad \text{For simplicity, define } N_d = \{n: (n,d) \in R'\} \\ B = \sum_{n \in \{n: (n,d) \in R'\}} (R'_{n,d} - \mathbf{z}_d^{\mathrm{T}} \mathbf{u}_n)^2 \\ \text{ bends only on } \mathbf{u}_n. \text{ Taking the derivative,} \qquad \text{This term depends only on } \mathbf{z}_d. \text{ Taking the derivative,}$$

This term depends only on  $\mathbf{u}_n$ . Taking the derivative,

$$\begin{split} \frac{\partial}{\partial \mathbf{u}_n} A &= -\sum_{d \in D_n} (R'_{n,d} - \mathbf{u}_n^\mathrm{T} \mathbf{z}_d) (\mathbf{z}_d) \\ &= 0 \\ \sum_{d \in D_n} R'_{n,d} \mathbf{z}_d &= \sum_{d \in D_n} (\mathbf{u}_n^\mathrm{T} \mathbf{z}_d) \mathbf{z}_d^\mathrm{T} \\ &= \mathbf{u}_n^\mathrm{T} \sum_{d \in D_n} \mathbf{z}_d \mathbf{z}_d^\mathrm{T} \\ \mathbf{u}_n^\mathrm{T} &= \sum_{d \in D_n} R'_{n,d} \mathbf{z}_d \left(\sum_{d \in D_n} \mathbf{z}_d \mathbf{z}_d^\mathrm{T}\right)^{-1} \\ \mathbf{u}_n &= \left(\sum_{d \in D_n} \mathbf{z}_d \mathbf{z}_d^\mathrm{T}\right)^{-1} \sum_{d \in D_n} R'_{n,d} \mathbf{z}_d \end{split}$$

$$\begin{split} \frac{\partial}{\partial \mathbf{u}_n} B &= -\sum_{n \in N_d} (R'_{n,d} - \mathbf{z}_d^{\mathrm{T}} \mathbf{u}_n) (\mathbf{u}_n) \\ &= 0 \\ \sum_{n \in N_d} R'_{n,d} \mathbf{u}_n &= \sum_{n \in N_d} (\mathbf{z}_d^{\mathrm{T}} \mathbf{u}_n) \mathbf{u}_n^{\mathrm{T}} \\ &= \mathbf{z}_d^{\mathrm{T}} \sum_{n \in N_d} \mathbf{u}_n \mathbf{u}_n^{\mathrm{T}} \\ \mathbf{z}_d^{\mathrm{T}} &= \sum_{n \in N_d} R'_{n,d} \mathbf{u}_n \left( \sum_{n \in N_d} \mathbf{u}_n \mathbf{u}_n^{\mathrm{T}} \right)^{-1} \\ \mathbf{z}_d &= \left( \sum_{n \in N_d} \mathbf{u}_n \mathbf{u}_n^{\mathrm{T}} \right)^{-1} \sum_{n \in N_d} R'_{n,d} \mathbf{u}_n \end{split}$$

An algorithm for this would look something like: while not converged:

$$\begin{aligned} &\text{for } n \text{ in } \text{range}(1,N+1) \text{:} & \quad \mathbf{u}_n \leftarrow \left(\sum_{d \in D_n} \mathbf{z}_d \mathbf{z}_d^{\mathrm{T}}\right)^{-1} \sum_{d \in D_n} R'_{n,d} \mathbf{z}_d \\ &\text{for } d \text{ in } \text{range}(1,D+1) \text{:} & \quad \mathbf{z}_d \leftarrow \left(\sum_{n \in N_d} \mathbf{u}_n \mathbf{u}_n^{\mathrm{T}}\right)^{-1} \sum_{n \in N_d} R'_{n,d} \mathbf{u}_n \end{aligned}$$

Autoencoder: A feed-forward neural network that outputs a  $\mathbb{R}^K$  approximation of an input  $\mathbf{x} \in \mathbb{R}^D$ 

- ullet We want a lower-dimensional representation, so K < D
- Uses a bottleneck layer of *K* neurons (hyperparameter)
- Mapping to 2D allows visualization
- Helps learn abstract features in unlabelled data that can be applied to a supervised task

Linear Autoencoder: Autoencoder made up of linear functions

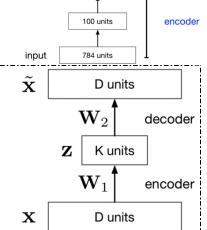
eg. Consider one hidden layer,  $\phi(x) = x$ , and  $\mathcal{L}(\mathbf{x}, \underline{\mathbf{x}}) = \|\mathbf{x} - \underline{\mathbf{x}}\|^2$ 

The reconstruction is a linear function,  $\underline{\mathbf{x}} = W_2 W_1 \mathbf{x}$ .

We know  $\underline{\mathbf{x}} \in \operatorname{span}(W_2 \text{ columns})$  as  $\underline{\mathbf{x}}$  is the result of a product with  $W_2$ 

The optimal K-dimensional linear subspace, reconstruction error-wise, is the subspace from PCA whose transformation matrix is U.

Optimal weights are just principal components! Set  $W_1=U^{\rm T}, W_2=U$ 



784 units

100 units

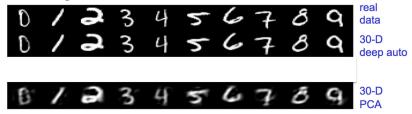
20 units

decoder

reconstruction

code vector

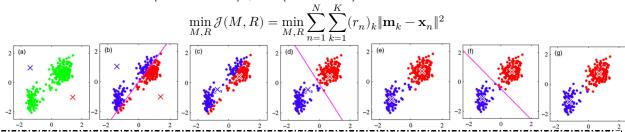
Nonlinear Autoencoder: Autoencoders that project data onto a nonlinear manifold (image of decoder). Can learn much more powerful codes compared to linear autoencoders



# Clustering

Multimodal: A distribution with multiple modes (local regions of high probability mass/density) **k-Means Clustering:** Unsupervised learning technique to fit N unlabelled data points into spherical clusters

- There are *K* clusters (hyperparameter), each with a center (the mean of all cluster points)
- Goal is to assign data points  $(\mathbf{x}_n \in \mathbb{R}^D)$  to clusters that minimize distances to cluster centers  $(\mathbf{m}_k \in \mathbb{R}^D)$ 
  - o Let  $\mathbf{r}_n \in \mathbb{R}^K$  be a one-hot vector where  $(r_n)_k = \mathbb{I}(\mathbf{x}_n \text{ assigned to cluster } k)$ . o Let  $M = [\mathbf{m}_1 \ \cdots \ \mathbf{m}_K]^\mathrm{T}, R = [\mathbf{r}_1 \ \cdots \ \mathbf{r}_N]^\mathrm{T}$



### Algorithm

- Randomly initialize cluster centers  $\mathbf{m}_k$
- Alternate between two steps until convergence (ie. data point assignments don't change):
  - **Assignment:** Assign data points to their closest clusters minimize with respect to R

$$\mathrm{argmin}_{\mathbf{r}_n} \sum_{k=1}^K (r_n)_k \|\mathbf{m}_k - \mathbf{x}_n\|^2 = \begin{bmatrix} \mathbb{I} \left( \mathrm{argmin}_{k \in \{1, \dots, K\}} \|\mathbf{m}_k - \mathbf{x}_n\|^2 = 1 \right) \\ \vdots \\ \mathbb{I} \left( \mathrm{argmin}_{k \in \{1, \dots, K\}} \|\mathbf{m}_k - \mathbf{x}_n\|^2 = K \right) \end{bmatrix} \text{ (closest cluster center to } \mathbf{x}_n)$$

b. Refitting: Move cluster centers to mean of its members – minimize with respect to M

$$\mathrm{argmin}_{\mathbf{m}_{k^*}} \sum_{n=1}^{N} \sum_{k=1}^{K} (r_n)_k \|\mathbf{m}_k - \mathbf{x}_n\|^2 = \frac{\sum_{n=1}^{N} (r_n)_{k^*} \mathbf{x}_n}{\sum_{n=1}^{N} (r_n)_{k^*}} \text{ (mean of all } \mathbf{x}_n \text{ in class } k^*)$$

- K-means will converge as cost is guaranteed to reduce each iteration, and number of cluster assignments is finite
- Unstable,  $\mathcal{J}$  is non-convex, no guarantee of global minimum.
- Used in vector quantization, technique for compressing data
- Used in image segmentation, technique for finding superpixels ("perceptual pixels")

















Optimization – Derivation of Iterative Solution, K-Means Clustering

$$\begin{split} \frac{\partial}{\partial \mathbf{m}_{k^*}} \sum_{n=1}^N \sum_{k=1}^K (r_n)_k \| \mathbf{m}_k - \mathbf{x}_n \|^2 &= \frac{\partial}{\partial \mathbf{m}_{k^*}} \sum_{n=1}^N (r_n)_{k^*} \| \mathbf{m}_{k^*} - \mathbf{x}_n \|^2 \text{ (since } \mathbf{r}_n \text{ is one-hot)} \\ &= 2 \sum_{n=1}^N (r_n)_{k^*} (\mathbf{m}_{k^*} - \mathbf{x}_n) = 0 \\ & \mathbf{m}_{k^*} \sum_{n=1}^N (r_n)_{k^*} = \sum_{n=1}^N (r_n)_{k^*} \mathbf{x}_n \\ & \therefore \mathbf{m}_{k^*} = \frac{\sum_{n=1}^N (r_n)_{k^*} \mathbf{x}_n}{\sum_{n=1}^N (r_n)_{k^*}} \end{split}$$

Soft k-Means Clustering: A variant of k-means that makes soft assignments (between 0 to 1) of data to clusters

Introduce hyperparameter  $\beta$  where as  $\beta \to \infty$ , soft k-means becomes k-means

#### Algorithm

- 1) Randomly initialize cluster centers  $\mathbf{m}_k$
- 2) Alternate between two steps until convergence (ie. data point assignments don't change):
  - Assignment: Assign data points to their closest clusters optimize for R

$$\mathbf{r}_n = \operatorname{softmax} \left( -\beta \begin{bmatrix} \|\mathbf{m}_1 - \mathbf{x}_n\|^2 \\ \vdots \\ \|\mathbf{m}_K - \mathbf{x}_n\|^2 \end{bmatrix} \right)$$

**Refitting:** Move cluster centers to mean of its members – optimize for M

$$\mathbf{m}_{k^*} = \frac{\sum_{n=1}^{N} (r_n)_{k^*} \mathbf{x}_n}{\sum_{n=1}^{N} (r_n)_{k^*}}$$

Identifiable: A probability model that given infinitely much data, can theoretically get true model parameters. **Latent Variable Model:** A probabilistic interpretation of latent factors Z,

$$p(\mathbf{x}) = \sum_{z \in Z} p_{\mathcal{D},Z}(\mathbf{x},z) = \sum_{z \in Z} p_{\mathcal{D}|Z}(\mathbf{x}|z) p_Z(z)$$

 $p(\mathbf{x}) = \sum_{z \in Z} p_{\mathcal{D},Z}(\mathbf{x},z) = \sum_{z \in Z} p_{\mathcal{D}|Z}(\mathbf{x}|z) p_Z(z)$  **Mixture Model:** A latent variable model where  $p_Z(z)$  is a caategorical distribution (ie.  $Z = \{z_1, \dots, z_N\}$  is finite).

- Assumes data was generated by
  - Choosing a mixture component randomly (ie.  $p_Z(z)$ )
  - Sampling from the distribution associated with that component (ie.  $p_{D|Z}(\mathbf{x}|z)$ )
- - **Mixing Proportions/Coefficients:** Vector  $\eta$  with  $\eta_1 + \cdots + \eta_K = 1$  and  $\eta_k \geq 0$ .
  - Parameters associated with component distributions

Gaussian Mixture Model (GMM): The mixture model

$$\begin{aligned} z_n &\sim \text{Categorical}(\mathbf{\eta}) \text{ (for some categorical distribution)} \\ \mathbf{x}_n | z_n &\sim \text{Normal}(\mathbf{\mu}_{z_n}, \boldsymbol{\Sigma}_{z_n}) \end{aligned}$$

$$p_{X|\theta}(\mathbf{x}|\theta) = \sum_{k=1}^{K} \frac{\mathbf{\eta}_k}{\det(\Sigma_k)^{\frac{1}{2}} (2\pi)^{\frac{D}{2}}} e^{-\frac{1}{2}(\mathbf{x} - \mathbf{\mu}_k)^{\mathrm{T}} \Sigma_k^{-1} (\mathbf{x} - \mathbf{\mu}_k)} \quad \left( \text{where } \theta = (\mathbf{\mu}, \mathbf{\eta}, \Sigma), \ \mathbf{\eta} = \{\mathbf{\eta}_1, \dots, \mathbf{\eta}_K\} \\ \Sigma = \{\Sigma_1, \dots, \Sigma_K\} \right)$$

- GMMs, and even diagonal GMMs (where the  $\Sigma_k$  are diagonal) are universal density approximators
- GMMs are not identifiable, infinitely many ways to fit data
  - eg. You can fit noise as just noise or as a rapidly varying function with no noise

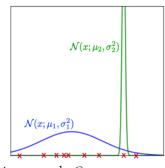
$$\circ \quad \text{ eg. } \theta_1=(\pmb{\mu}_1,\pmb{\eta}_1,\Sigma), \theta_2=(\pmb{\mu}_2,\pmb{\eta}_2,\Sigma) \text{ is the same as } \theta_1=(\pmb{\mu}_2,\pmb{\eta}_2,\Sigma), \theta_2=(\pmb{\mu}_1,\pmb{\eta}_1,\Sigma)$$

We can attempt to find parameters with maximum likelihood:

o Let 
$$X = [\mathbf{x}_1 \quad \cdots \quad \mathbf{x}_N]^{\mathrm{T}}, \mathbf{z} = [z_1 \quad \cdots \quad z_N]^{\mathrm{T}}$$

$$\begin{split} \ln p(X|\theta) &= \ln \prod_{n=1}^{N} p_{X_n|\theta}(\mathbf{x}_n|\theta) \\ &= \sum_{n=1}^{N} \ln \left( \sum_{k=1}^{K} \frac{\mathbf{\eta}_k}{\det(\boldsymbol{\Sigma}_k)^{\frac{1}{2}} (2\pi)^{\frac{D}{2}}} e^{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^{\mathrm{T}} \boldsymbol{\Sigma}_k^{-1} (\mathbf{x} - \boldsymbol{\mu}_k)} \right) \end{split}$$

- This isn't simplifiable, so the derivative is very complicated. No closed-form solution.
- Difficult to optimize with gradient descent due to non-convexity, enforcing non-negativity constraint on  $\pi_k$ , PSD constraint on  $\Sigma_k$ , and  $\frac{\partial}{\partial \Sigma_{-}}$  is computationally slow
- Singularity: A tiny-variance Gaussian component fitted to one training example. Can create arbitrarily high training likelihoods (think of it like K-means, where you give one point its own class).



Assuming we know  $\theta$ , we can find  $p(z_k=1|\mathbf{x}_n)$ using Bayes' Rule (similar to posterior):

$$p(z_k=1|\mathbf{x}_n) = \frac{p(z_k=1)p(\mathbf{x}|z_k=1)}{\sum_{i=0}^1 p(z_k=i)p(\mathbf{x}|z_k=i)}$$

Assuming we know  $\mathbf{z}$ , we can find the optimal  $\theta$  by maximizing joint log-likelihood,

sing Bayes' Rule (similar to posterior): 
$$p(z_k = 1 | \mathbf{x}_n) = \frac{p(z_k = 1)p(\mathbf{x}|z_k = 1)}{\sum_{i=0}^1 p(z_k = i)p(\mathbf{x}|z_k = i)} \quad \text{maximizing joint log-likelihood,} \\ \ln p(X, \mathbf{z}|\theta) = \sum_{n=1}^N \ln p_{X_n, \mathbf{z}|\theta}(\mathbf{x}_n, z_n|\theta) \\ = \sum_{n=1}^N \left[ \ln p_{\mathbf{z}|\theta}(z_n|\theta) + \ln p_{X_n|\theta}(\mathbf{x}_n|\theta) \right] \quad \text{The first term is found computationally.}$$

The first term is found computationally. Substitute the formula for GMM into the second term.

# Expectation-Maximization (E-M): Alternating minimization algorithm for ${f z}$ and ${f heta}$

#### Algorithm

- 1) Randomly initialize  $\mathbf{z}, \theta$
- 2) Alternate between two steps until convergence
  - **Expectation Step:** Assign each  $\mathbf{x}_n$  a soft probability of coming from each distribution find  $\mathbf{z}$  $(r_n)_k = p(z_k = 1 | \mathbf{x}_n, \theta)$
  - **Maximization Step:** Find the best Gaussian fits to their members optimize for  $\theta$

$$\eta_k = \frac{1}{N} \sum_{n=1}^{N} (r_n)_k \,, \qquad \mu_k = \frac{\sum_{n=1}^{N} (r_n)_k \, \mathbf{x}_n}{\sum_{n=1}^{N} (r_n)_k} \,, \qquad \Sigma_k = \frac{\sum_{n=1}^{N} (r_n)_k \, (\mathbf{x}_n - \boldsymbol{\mu}_k) (\mathbf{x}_n - \boldsymbol{\mu}_k)^{\mathrm{T}}}{\sum_{n=1}^{N} (r_n)_k}$$

(c) 2

If mixture coefficients are  $\eta_k = \frac{1}{K}$ , covariances are diagonal,  $\Sigma_k = \sigma I$  for any  $\sigma > 0$ , E-M is soft k-means.

(d)

# Reinforcement Learning

Agent: The learning model that is controlling an entity in some program.

Sequential Decision Making: When agents choose from many actions which affect future possibilities.

Action  $(A_t)$ : Set of everything the agent can do at time t.

State  $(S_t)$ : Set of all possible descriptions of the program at time t. (assume this is fully observable, AKA knowable)

Action Space (A): Set of all possible action. Can be finite or infinite (assume it's finite for now)

**State Space** ( $\mathcal{S}$ ): Set of all possible states. Can be discrete or continuous.

**Transition Probability** ( $\mathcal{P}$ ): The probability  $\mathcal{P}: \mathcal{S}|\mathcal{S}, \mathcal{A} \to [0,1]$  of transforming to a new state given an action. Describes the **dynamics** of the environment. If known, we can **plan**; otherwise, we do **trial-and-error learning Policy** ( $\pi$ ): A function  $\pi$  describing how agents choose actions, given the state.

- **Deterministic Policy:** The function  $\pi: \mathcal{S} \to \mathcal{A}$  "Given a state, output an action"
- Stochastic Policy: The function  $\pi: \mathcal{S} \times \mathcal{A} \to [0,1]$  "Given a state & action, output the prob. of choosing the action"
  - O Can be interpreted as  $\pi: \mathcal{S} \to [0,1]^{|\mathcal{A}|}$  "Given a state, output a vector of probabilities for each action"
  - Can be interpreted as a probability function  $\pi: \mathcal{A}|\mathcal{S} \to [0,1]$
  - $\circ$  The probability of taking actions  $a_1, \ldots, a_T$ , given corresponding states  $s_1, \ldots, s_t$ , is

$$p(s_1, a_1, \dots, s_T, a_T) = p(s_1)[\pi(a_1|s_1)\mathcal{P}(s_2|s_1, a_1)] \cdots [\pi(a_{T-1}|s_{T-1})\mathcal{P}(s_T|s_{T-1}, a_{T-1})]\pi(a_T|s_T)$$

**Reward Signal (** $R_t$ **):** A function outputting a value that determines how "good" a decision is at time t.

- Stochastic rewards are  $R_t \sim \mathcal{R}(\mathcal{S}, \mathcal{A}|S_t, A_t)$  for some  $\mathcal{R}: \mathcal{S} \times \mathcal{A} \to \mathbb{R}$
- Deterministic rewards  $R_t = \mathcal{R}(S_t, A_t)$  for some  $\mathcal{R}: \mathcal{S} \times \mathcal{A} \to \mathbb{R}$  (assume this for now)

**Return**  $(G_t)$ : Reward signal, but accounting for long-term reward. We should maximize  $\mathbb{E}[G_t]$ .

- Undiscounted: Return that places no weight on future reward,  $G_t = R_t + R_{t+1} + R_{t+2} + \cdots$
- Discounted: Return that places differing weights on future reward,  $G_t = R_t + \gamma R_{t+1} + \gamma^2 R_{t+2} + \cdots$ 
  - o Discount Factor: Hyperparameter  $0 \le \gamma < 1$  affecting weights of rewards now vs. later

Markov Assumption: The assumption that  $S_{t+1}$  depends on  $S_t$  but nothing before it.

Markov Decision Process (MDP): Mathematical framework for reinforcement learning, tuple  $(S, A, P, R, \gamma)$ 

 $\triangleright$   $\mathcal{S}$ : State space

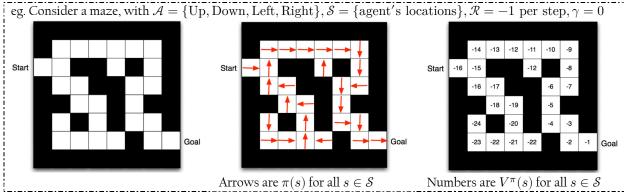
- $\triangleright$   $\mathcal{R}$ : Immediate reward functions
- $\triangleright$  A: Action space
- $\triangleright$   $\gamma$ : Discount factor
- $\triangleright$   $\mathcal{P}$ : Transition probabilities

Value Function  $(V^{\pi})$ : For policy  $\pi$  at state  $s^* \in \mathcal{S}$ , function  $V^{\pi} : \mathcal{S} \to \mathbb{R}$  measuring the desirability of being in state  $s^*$ . Equal to expected gain based on state,

$$\begin{split} V^{\pi}(s^*) &= \mathbb{E}_{\pi}[G_t|S_t = s^*] \\ &= \mathbb{E}_{\pi}\left[\sum_{k=0}^{\infty} \gamma^k R_{t+k} \left| S_t = s^* \right| \right] \\ &= \sum_{a \in A_t} \pi(a|s^*) \left[\mathcal{R}(s^*, a) + \gamma \sum_{s \in S_{t+1}} \mathcal{P}(s|s^*, a) \mathbb{E}_{\pi}[G_{t+1}|\mathcal{S}_{t+1} = s] \right] \\ &= \sum_{a \in A_t} \underbrace{\pi(a|s^*)}_{\substack{\text{prob of action } a}} \left[\underbrace{\mathcal{R}(s^*, a)}_{\substack{\text{initial gain}}} + \gamma \sum_{s \in S_{t+1}} \mathcal{P}(s|s^*, a) V^{\pi}(s) \right] \quad (\pi(a|s^*) \text{ is one-hot if deterministic}) \end{split}$$

State-Action Value Function/Q-Function ( $Q^{\pi}$ ): For policy  $\pi$  at state  $s^* \in \mathcal{S}$ , given action  $a^* \in \mathcal{A}$ , function  $Q^{\pi} : \mathcal{S} \times \mathcal{A} \to \mathbb{R}$  measuring the desirability of an action, given a state.

$$\begin{split} Q^{\pi}(s^*, a^*) &= \mathbb{E}_{\pi} \left[ \sum_{k=0}^{\infty} \gamma^k R_{t+k} \left| S_t = s^*, A_t = a^* \right| \right. \\ &= \underbrace{\mathcal{R}(s^*, a^*)}_{\text{initial gain}} + \underbrace{\gamma \sum_{s \in S_{t+1}} \mathcal{P}(s | a^*, s^*) \sum_{a \in A_{t+1}} \pi(a | s) Q^{\pi}(s, a)}_{\text{future gain}} \right. \\ &\left. (\pi(a | s) \text{ is one-hot if deterministic}) \right. \end{split}$$



Bellman Equation: A type of equation that returns the "value" of action a at time t as the sum of the action's immediate reward and the "value" of future actions as a result of a.

- Includes  $V^{\pi}$  and  $Q^{\pi}$

$$V^{\pi} \text{ and } Q^{\pi} \text{ can be expressed in terms of each other. It's easier to find } V^{\pi} \text{ from } Q^{\pi} \text{ (as you iterate through } a, \text{not } s)$$
 
$$V^{\pi}(s^{*}) = \sum_{a \in A_{t}} \pi(a|s^{*})Q^{\pi}(s^{*}, a) \qquad \qquad Q^{\pi}(s^{*}, a^{*}) = \mathcal{R}(s^{*}a^{*}) + \gamma \sum_{s \in S_{t+1}} \mathcal{P}(s|s^{*}, a^{*})V^{\pi}(s)$$

Bellman Backup/Update Operator: On vectors of all states/actions  $s^* \in \mathcal{S}, a^* \in \mathcal{A}$ , the operator  $T^{\pi}$  where

$$\begin{split} (T^\pi V)(s^*) &= V^\pi \big(V(s^*)\big) \\ (T^\pi Q)(s^*,a^*) &= Q^\pi \big(Q(s^*,a^*)\big) \end{split}$$

Where V,Q are analogous to  $V^{\pi},Q^{\pi}$ , and all functions are applied element-wise.

Fixed Point: A point that does not change under a transformation; that is, a point x such that f(x) = xContraction Map: Function  $f: \mathbb{R}^n \to \mathbb{R}^n$  if  $||f(\mathbf{x}_1) - f(\mathbf{x}_2)|| \le \alpha ||\mathbf{x}_1 - \mathbf{x}_2||$  for some  $0 \le \alpha < 1$ 

- For some point  $\mathbf{x}_1$  contraction map  $f, (f \circ ... \circ f)(\mathbf{x}_1) \to \mathbf{x}_2$  for some unique fixed point  $\mathbf{x}_2$  o This is because  $\|(\underbrace{f \circ ... \circ f}_{k \text{ times}})(\mathbf{x}_1) \mathbf{x}_2\| \le \alpha^k \|\mathbf{x}_1 \mathbf{x}_2\| \to 0$  as  $k \to \infty$

Dynamic Programming: An optimization and programming method based on breaking problems into nested subproblems and recursively solving them. Requires its problems to fit Bellman equations.

Assuming a deterministic policy, we choose the action that maximizes the desirability of the resulting state (ie. maximize the Q-function).

$$\begin{split} \pi^*(s) &= \operatorname{argmax}_{a \in A_t} Q^{\pi}(s, a) \\ &\therefore Q^{\pi^*}(s^*, a^*) = \mathcal{R}(s, a) + \gamma \sum_{s \in S_{t+1}} \mathcal{P}(s|a^*, s^*) Q^{\pi^*}\big(s, \pi^*(s)\big) \\ &= \mathcal{R}(s, a) + \gamma \sum_{s \in S_{t+1}} \mathcal{P}(s|a^*, s^*) \max_{a \in A_{t+1}} Q^{\pi^*}(s, a) \\ &\therefore (T^{\pi^*}Q)(s^*, a^*) = Q^{\pi^*}\big(Q(s^*, a^*)\big) \end{split}$$

 $T^{\pi^*}Q$  happens to be a **contraction map** (see proof below).

Value Iteration: The idea of applying  $T^{\pi^*}$ iteratively until the result converges to an ideal fixed point:

$$Q^{\pi^*}\big(Q(s^*,a^*)\big) = Q(s^*,a^*)$$

Use the pseudocode below:

initialize Q randomly while not converged:

$$Q \leftarrow T^{\pi^*}Q$$

**Planning:** Creating a MDP where the dynamics (ie.  $\mathcal{P}$ ) are known

- Value iteration requires knowing  $\mathcal{P}$  and explicitly representing Q as a vector
  - $|\mathcal{S}|$  can be extremely large or infinite
  - $|\mathcal{A}|$  can also be continuous and/or infinite

**Learning:** Creating a MDP where the dynamics (ie.  $\mathcal{P}$ ) are unknown

- Learn a model of the environment, and do planning in this model (ie. model-based reinforcement learning)
- Learn a value function
- Learn a policy directly
- Deal with large |S| using function approximations!

Optimization – Proof that  $T^{\pi^*}$  is a Contraction Map

$$\begin{split} |(T^{\pi^*}Q_1)(s^*,a^*) - (T^{\pi^*}Q_2)(s^*,a^*)| &= \\ \left| \left[ \mathcal{R}(s^*,a^*) + \gamma \sum_{s \in S_{t+1}} \mathcal{P}(s|s^*,a^*) \max_{a \in A_{t+1}} Q_1(s^*,a^*) \right] - \left[ \mathcal{R}(s^*,a^*) + \gamma \sum_{s \in S_{t+1}} \mathcal{P}(s|s^*,a^*) \max_{a \in A_{t+1}} Q_2(s^*,a^*) \right] \right| \\ &= \gamma \left| \sum_{s \in S_{t+1}} \mathcal{P}(s|s^*,a^*) \left[ \max_{a \in A_{t+1}} Q_1(s^*,a^*) - \max_{a \in A_{t+1}} Q_2(s^*,a^*) \right] \right| \\ &\leq \gamma \sum_{s \in S_{t+1}} \mathcal{P}(s|s^*,a^*) \left[ \max_{a \in A_{t+1}} |Q_1(s^*,a^*) - Q_2(s^*,a^*)| \right] \text{ (by triangle inequality?)} \\ &= \gamma \max_{a \in A_{t+1}} |Q_1(s^*,a^*) - Q_2(s^*,a^*)| \sum_{s \in S_{t+1}} \mathcal{P}(s|s^*,a^*) \\ &= \gamma \max_{a \in A_{t+1}} |Q_1(s^*,a^*) - Q_2(s^*,a^*)| \text{ (transition probabilities add to 1)} \\ &= \gamma \|Q_1(s^*,a^*) - Q_2(s^*,a^*)\|_{\infty} \text{ (the $\infty$-norm is defined as } \max\{x_1,\dots,x_n\}) \end{split}$$

Since  $s^*, a^*$  was arbitrary, this holds for any  $s^*, a^*$ , so  $\max_{s \in S_t, a \in A_t} |T^{\pi^*}Q_1 - T^{\pi^*}Q_2| \leq \gamma \max_{a \in A_t+1} |Q_1 - Q_2|$  $\|T^{\pi^*}Q_1 - T^{\pi^*}Q_2\|_{\infty} \le \gamma \|Q_1 - Q_2\|_{\infty}$ 

Monte Carlo: An estimating technique based on taking random samples

To estimate  $\mu = \mathbb{E}[X]$ , one can repeatedly sample X and update  $\mu \leftarrow \mu + \alpha(X - \mu)$  with learning rate  $\alpha$ 

Exploration-Exploitation Tradeoff: The tradeoff between choosing what seems to be the best path versus trying out new paths, which occurs when we don't know  $\mathcal P$  and must find the best choices ourselves.

$$\bullet \quad \text{ $\epsilon$-Greedy Policy:} \text{ Stochastic policy } \pi(a,s) = \begin{cases} \operatorname{argmax}_{a \in A_t} Q^\pi(s,a) & \operatorname{probability } 1 - \epsilon \\ \operatorname{random action} & \operatorname{probability } \epsilon \end{cases} \text{ (usually $\epsilon = \frac{1}{20}$)}$$

**Q-Learning:** The process of finding the best Q without knowledge of  $\mathcal{P}$ 

- Initialize Q(s, a) for all  $(s, a) \in \mathcal{S} \times \mathcal{A}$
- Agent starts at state  $S_0$
- For each time step  $t \in \mathbb{N}$ ,
  - o Choose  $A_t$  according to  $\epsilon$ -greedy policy
  - Observe  $S_{t+1}$  and  $R_t$ , update Q based on results as if it were a Monte Carlo sample:

$$\begin{split} Q(S_{t+1},A_{t+1}) \leftarrow Q(S_t,A_t) + \alpha \left[ R_t + \gamma \max_{a \in \mathcal{A}} Q(S_{t+1},a) - Q(S_t,A_t) \right] \\ & \quad \text{Note that the policy $\pi$ isn't in this formula! $\pi$ only determines what states we visit.} \end{split}$$

- Temporal Difference Learning: Updating prior predictions to match later predictions

Off-Policy: An algorithm that doesn't rely on  $\pi$  to converge to the optimal Q-function – we can pick any  $\pi$  we want On-Policy: An algorithm that relies on  $\pi$  to converge to the optimal Q-function (eg. policy gradient, not in CSC311)

Tabular Representation: Representation of data as a table/matrix.

- We've stored Q as a  $|\mathcal{S}| \times |\mathcal{A}|$  matrix impractical, doesn't share structure between related states Parameterized Function: A lower-dimensional approximation function  $\mathbb{R}^n \to \mathbb{R}^m$  with n < m
  - Can be done with a neural network; called deep Q-learning

#### EXTRA NOTES AFTER STUDYING FOR EXAMS

- Dropout, sparsity, capacity, momentum, cosine distance, saturation concepts taught in previous offerings of CSC311
- Though validation set is used for hyperparameters, it's okay to train optimization hyperparameters (learning rate, batch size) on the training set – they determine convergence rather than overfitting.
- Testing set is used for generalization, so that's why we can't use it for anything else