

# MM 225 – AI and Data Science

## Day 19: Hypothesis Testing 1

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# Introduction

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Consider a population with distribution function  $F(\theta)$ , where  $\theta$  is unknown.

Let  $\theta_0$  be a specific known number.

Want to know whether  $\theta = \theta_0$

In statistical parlance: want to test the Null Hypothesis that  $\theta = \theta_0$  denoted by

$$H_0 : \theta = \theta_0$$

# The Critical Region

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Let  $X_1, X_2, \dots, X_n$  be a random sample of size  $n$  from  $F(\theta)$

Based on these  $n$  values we should decide if  $H_0$  is true or not

Want to find region  $C$  in  $n$ -dimensional space so that  $H_0$  is rejected if random sample  $(X_1, X_2, \dots, X_n)$  lies in region  $C$

Such a region  $C$  is called critical region.

Thus want to find a statistical test determined by the critical region  $C$  such that the test

Rejects  $H_0$  if  $(X_1, X_2, \dots, X_n) \in C$

And

Accepts  $H_0$  if  $(X_1, X_2, \dots, X_n) \notin C$

# Resulting Errors

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Want to test hypothesis  $H_0 : \theta = \theta_0$

Critical region has been found

Two types of error can result

Decision \ Reality	$H_0$ is true	$H_0$ is not true
Accept $H_0$	good	Error Type II
Reject $H_0$	Error Type I	good

Also known as

- Type I error = false negative
- Type II error = False positive

Decision \ Reality	$H_0$ is true	$H_0$ is not true
Accept $H_0$	good	Error Type II
Reject $H_0$	Error Type I $\leq \alpha$	good

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Would like to minimise both types of errors

Mathematically not possible

Classical approach is to

- fix  $\alpha$  at the minimum possible level and,
- Set up test so that probability of Type I error of the test is  $\leq \alpha$

# Significance Level and Power of Test

$$P[\text{type I error}] \leq \alpha \Rightarrow P[\text{Reject } H_0 | H_0 \text{ is true}] \leq \alpha$$

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$$\Rightarrow P[(X_1, X_2, \dots, X_n \in C) | H_0] \leq \alpha$$

$\alpha$  is called significance level of the test

**1-  $\alpha$  is called confidence level of the test**

$$P[\text{type II error}] = \beta \Rightarrow P[\text{Accept } H_0 | H_0 \text{ is not true}] = \beta$$

$$\Rightarrow P[\text{Reject } H_0 | H_0 \text{ is not true}] = 1 - \beta$$

$$\Rightarrow P[(X_1, X_2, \dots, X_n \in C) | H_0 \text{ is not true}] = 1 - \beta$$

**1-  $\beta$  is called Power of the test**

# Alternate Hypothesis

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For testing the null hypothesis  $H_0$  it is important to know what is meant by  $H_0$  is not true?

Example: Let  $H_0 : \theta = \theta_0$ . When  $H_0$  is not true, there are following three possibilities:

1.  $\theta \neq \theta_0$
2.  $\theta < \theta_0$
3.  $\theta > \theta_0$

Thus, it is important to explicitly state the alternative hypothesis, generally denoted by  $H_A$  or  $H_1$ , as definition of critical region C depends on the alternate hypothesis

# Classical approach

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1. Fix level of significance  $\alpha$
2. Clearly state the null hypothesis and alternate hypothesis in terms of population parameter  $\theta$
3. Choose an appropriate estimator for  $\theta$  using data  $(X_1, X_2, \dots, X_n)$  say  $d(X_1, X_2, \dots, X_n) = d(\mathbf{X})$
4. Define critical region  $C$ , where  $H_0$  is rejected
5. Calculate  $P[C | H_0] = \alpha$  to determine exact nature of  $C$
6. Give decision



# Case of $N(\mu, \sigma^2)$ , when $\sigma^2$ is known

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$X_1, X_2, \dots, X_n \sim iid N(\mu, \sigma^2)$ , and  $\sigma^2$  is known and  $\mu$  is unknown

Want to test  $H_0 : \mu = \mu_0$  vs.  $H_A : \mu \neq \mu_0$  (*Two sided Alternative*)

1. Let  $\alpha$  be fixed
2.  $H_0 : \mu = \mu_0$  vs.  $H_A : \mu \neq \mu_0$
3. It is shown that  $E(\bar{X}) = \mu$ ,  $\bar{X}$  is estimator
4.  $H_0$  can be rejected if  $\bar{X}$  is not in the close vicinity of  $\mu_0$ , hence
$$C = \{X_1, X_2, \dots, X_n \mid |\bar{X} - \mu_0| > c\}$$

Note that  $Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \sim N(0,1)$ , when  $H_0$  is true

# Case of $N(\mu, \sigma^2)$ , when $\sigma^2$ is known

4.  $H_0$  can be rejected if  $\bar{X}$  is not in the close vicinity of  $\mu_0$ , hence

$$C = \{X_1, X_2, \dots, X_n \mid |\bar{X} - \mu_0| > c\}$$

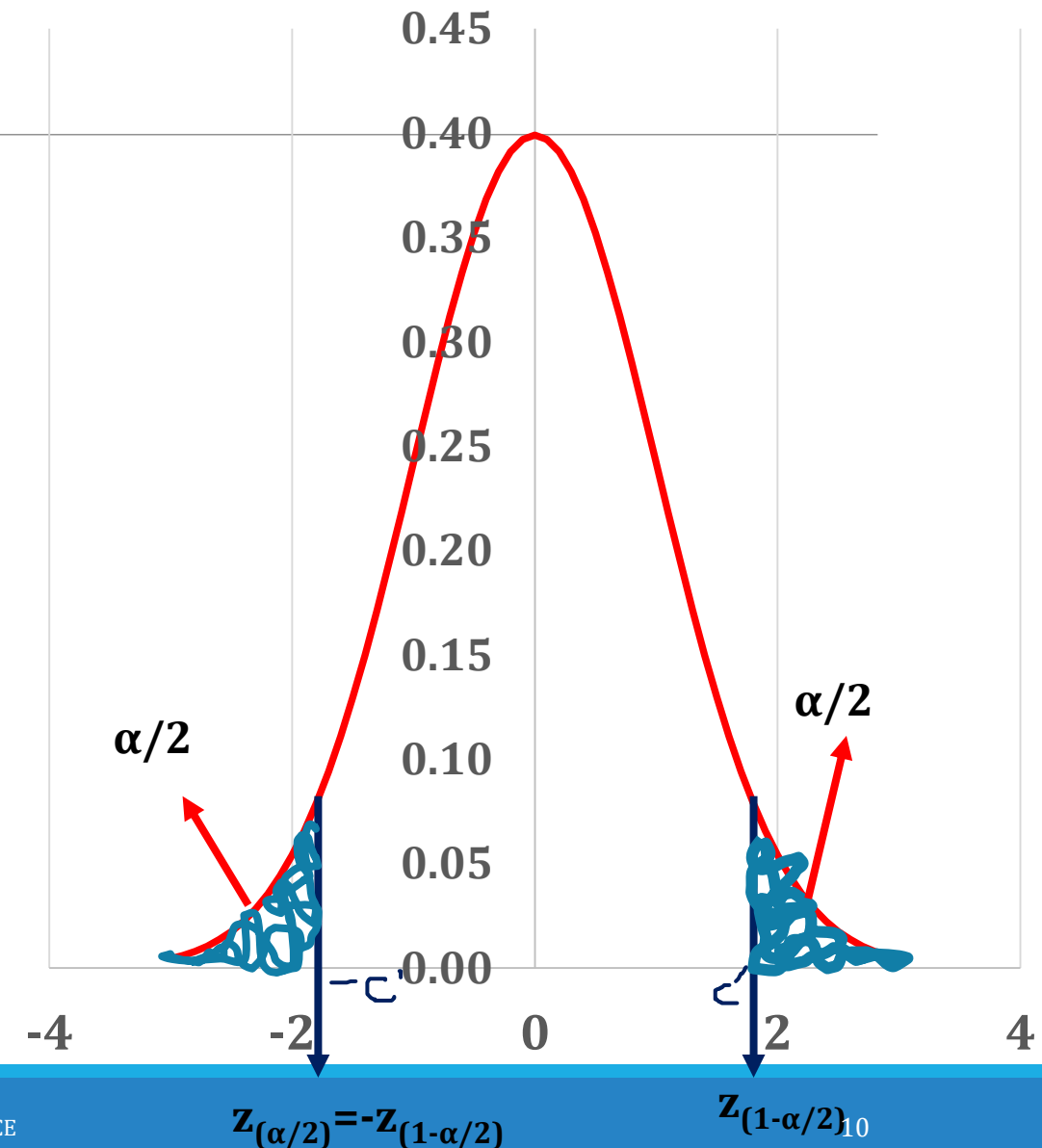
Note that  $Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \sim N(0,1)$ , when  $H_0$  is true

$$\begin{aligned} 5. \quad P \left[ X_1, X_2, \dots, X_n \mid \left| \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \right| > c' \right] &= \alpha \\ &= P[Z > c'] = \frac{\alpha}{2} \\ \therefore c' &= z_{1-\alpha/2} \end{aligned}$$

6. If  $\left| \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \right| > z_{1-\alpha/2}$  then reject  $H_0$

$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$  is called “TEST STATISTIC” and  $z_{1-\alpha/2}$  is called critical value of the test

Shaded region is the critical region C



# Example

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An aerospace industry is interested in buying certain super alloy rods from a foundry. The industry has been told that the super alloy would have yield strength of 1110 MPa with standard deviation 110 MPa. The industry takes a random sample of size 100 from the supplied lot and finds that the average yield strength to be 1129 MPa. Should the company accept the supply?

$\mu_0 = 1110$  MPa,  $\sigma = 110$  MPa,  $n = 100$  and  $\bar{x} = 1129$  MPa

Let us follow the steps of classical Hypothesis testing process

1. Let us fix  $\alpha = 0.05$
2.  $H_0 : \mu = 1110 \text{ MPa}$  vs.  $H_A : \mu \neq 1110 \text{ MPa}$
3. Statistic of interest is  $\bar{X}$  and  $\bar{x} = 1129$
4.  $C = \{X_1, X_2, \dots, X_n \mid \left| \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \right| > z\}$  when  $H_0$  is true

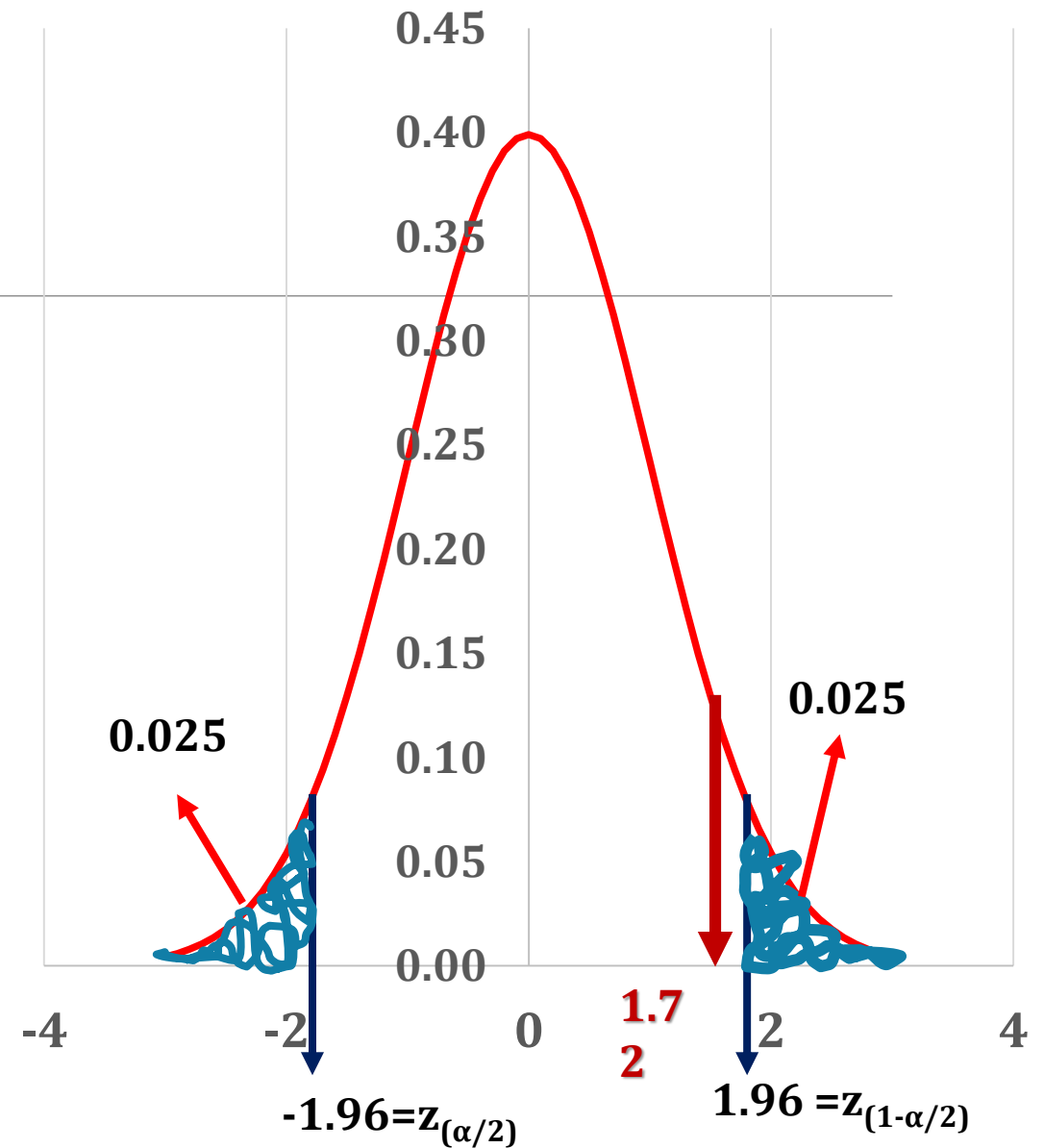
$$5. \quad P\left(\left|\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}\right| > z\right) = \alpha = 0.05, \text{ as } Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \sim N(0, 1)$$

$$P\left(\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} > z_{1-\alpha/2}\right) = \frac{\alpha}{2} = 0.025$$

From Standard Normal table  $z_{1-\alpha/2} = z_{0.975} = 1.96$  is the critical value

$$\frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} = \frac{1129 - 1110}{110} \sqrt{100} = 1.72 \text{ is not } > 1.96$$

6. Decision :  $H_0$  cannot be rejected. So the industry should accept the supplied lot



# Case of $N(\mu, \sigma^2)$ , when $\sigma^2$ is unknown

## $\sigma^2$ is known

$X_1, X_2, \dots, X_n \sim iid N(\mu, \sigma^2)$ , and  $\sigma^2$  is known and  $\mu$  is unknown

Want to test  $H_0 : \mu = \mu_0$  vs.  $H_A : \mu \neq \mu_0$

1. Let  $\alpha$  be fixed
2.  $H_0 : \mu = \mu_0$  vs.  $H_A : \mu \neq \mu_0$
3. It is shown that  $E(\bar{X}) = \mu$ ,  $\bar{X}$  is estimator
4.  $H_0$  can be rejected if  $\bar{X}$  is not in the close vicinity of  $\mu_0$ , hence

$$C = \{X_1, X_2, \dots, X_n \mid |\bar{X} - \mu_0| > c\}$$

Note that  $Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \sim N(0,1)$ , when  $H_0$  is true

## $\sigma^2$ is unknown

$X_1, X_2, \dots, X_n \sim iid N(\mu, \sigma^2)$ , and  $\sigma^2$  is unknown and  $\mu$  is unknown

Want to test  $H_0 : \mu = \mu_0$  vs.  $H_A : \mu \neq \mu_0$

1. Let  $\alpha$  be fixed
2.  $H_0 : \mu = \mu_0$  vs.  $H_A : \mu \neq \mu_0$
3. It is shown that  $E(\bar{X}) = \mu$ ,  $\bar{X}$  is estimator
4.  $H_0$  can be rejected if  $\bar{X}$  is not in the close vicinity of  $\mu_0$ , hence

$$C = \{X_1, X_2, \dots, X_n \mid |\bar{X} - \mu_0| > c\}$$

Note that  $T = \frac{\bar{X} - \mu_0}{s/\sqrt{n}} \sim t(n-1)$ , when  $H_0$  is true

# Case of $N(\mu, \sigma^2)$ , when $\sigma^2$ is unknown

$\sigma^2$  is known

$$5. \quad P \left[ X_1, X_2, \dots, X_n \mid \left| \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \right| > c' \right] = \alpha$$

$$= P[Z > c'] = \frac{\alpha}{2}$$

$$\therefore c' = z_{1-\alpha/2}$$

$$6. \quad \text{If } \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} > z_{1-\alpha/2} \text{ then reject } H_0$$

$\sigma^2$  is unknown

$$5. \quad P \left[ X_1, X_2, \dots, X_n \mid \left| \frac{\bar{X} - \mu_0}{s/\sqrt{n}} \right| > c' \right] = \alpha$$

$$= P[t > c'] = \frac{\alpha}{2}$$

$$\therefore c' = t_{(n-1), \alpha/2}$$

$$6. \quad \text{If } \frac{\bar{X} - \mu_0}{s/\sqrt{n}} > t_{(n-1), \alpha/2} \text{ then reject } H_0$$

# Example: $\sigma^2$ is unknown

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An aerospace industry is interested in buying certain super alloy rods from a Maxfoundry. The industry has been told that the super alloy would have yield strength of 1110 MPa. The industry takes a random sample of size 100 from the supplied lot and finds that the average yield strength to be 1129 MPa with standard deviation of 112 MPa. Should the company accept the supply?

$$\mu_0 = 1110 \text{ MPa}, n = 100 \text{ and } \bar{x} = 1129 \text{ MPa}, s = 112 \text{ MPa}$$

Let us once again follow the steps of classical Hypothesis testing process

1. Let us fix  $\alpha = 0.05$
2.  $H_0 : \mu = 1110 \text{ MPa}$  vs.  $H_A : \mu \neq 1110 \text{ MPa}$
3. Statistic of interest is  $\bar{X} = 1129$
4. As  $\sigma^2$  is unknown  $C = \{X_1, X_2, \dots, X_n \mid \left| \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \right| > c\}$  when  $H_0$  is true

$$5. \quad P\left(\left| \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \right| > c\right) = \alpha = 0.05, \text{ as } t = \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \sim t_{(n-1)}$$

$$P\left(\frac{\bar{X} - \mu_0}{S/\sqrt{n}} > t_{1-\alpha/2, (n-1)}\right) = \frac{\alpha}{2} = 0.025$$

From t probability table  $t_{(n-1), 1-\alpha/2} = t_{99, 0.975} \approx 1.98$  is the critical value

$$\frac{\bar{X} - \mu_0}{S/\sqrt{n}} = \frac{1129 - 1110}{112} \sqrt{100} = 1.69 \text{ is not } > 1.98$$

6. Decision :  $H_0$  cannot be rejected. So the industry should accept the supplied lot



Thank you...