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MLSP - Assignment 1

# Solution 1 - Induction in PCA

→ We know that, In order to maximize the variance of 1-dim Projection  $y = w^T x$  of D-dimensional data  $x$ , we need to find  $w = v_1$ , such that  $v_1$  is eigen vector corresponding to highest eigen value of the Covariance matrix

$$S = \frac{1}{N} \sum_{n=1}^N (x_n - \bar{x})(x_n - \bar{x})^T$$

→ Let's assume that the variance of M-dimensional projection  $y_M = w_M^T x$  is maximized by  $W = [v_1, v_2, \dots, v_M]$  where  $v_1, v_2, \dots, v_M$  are orthogonal eigen vectors of  $S$  corresponding to the M - largest eigen values.

→ To maximize the variance of M+1 dimension projection  $y_{M+1} = w_{M+1}^T x$  we need to prove that  $w_{M+1} = [w_M, v_{M+1}]$  where  $v_{M+1}$  is eigen vector <sup>of S</sup> corresponding to ~~largest~~  $(M+1)^{th}$  largest eigen value.

The variance of Projected data is

$$S_y = \frac{1}{N} \sum_{n=1}^N (y_n - \bar{y})^2$$

$$\frac{1}{N} \sum_{n=1}^N (w_{M+1}^T x_n - w_{M+1}^T \bar{x})^2$$

$$\left\{ y_{M+1} = w_{M+1}^T x_n \right\}$$

$$\frac{1}{N} \sum_{n=1}^N w_{M+1}^T (x_n - \bar{x})(x_n - \bar{x})^T w_{M+1}$$

$$\frac{1}{N} w_{M+1}^T \sum_{n=1}^N (x_n - \bar{x})(x_n - \bar{x})^T w_{M+1}$$

$$w_{M+1}^T \underbrace{\frac{1}{N} \sum_{n=1}^N (x_n - \bar{x})(x_n - \bar{x})^T}_S w_{M+1}$$

$$w_{M+1}^T \frac{1}{N} \sum_{n=1}^N (x_n - \bar{x})(x_n - \bar{x})^T w_{M+1}$$

we can write it as.

$$\sum_{i=1}^{M+1} U_i^T S U_i$$

$$\sum_{i=1}^{M+1} U_i^T S U_i$$

$U_i \rightarrow$  eigen vector.  
Corresponding to eigenvalue  $\lambda_i$

We have to maximize the

$$\max_{U_i} \sum_{i=1}^{M+1} U_i^T S U_i$$

$$\sum_{i=1}^M U_i^T S U_i + U_{M+1}^T S U_{M+1} \quad \text{--- (1)}$$

We already know that this term is maximized by ~~for~~  $M$ -largest eigen vectors corresponding to  $M$ -largest eigen values, therefore we need to maximize the ~~the~~ second term.  $U_{M+1}^T S U_{M+1}$  such that

$$U_{M+1}^T U_{M+1} = 1 \text{ and } U_{M+1}^T U_i = 0 \text{ for } 1 \leq i \leq M$$

$\rightarrow$  Solving this constrained optimization problem.

Writing the Lagrangian of the problem.

$$\mathcal{L}(\lambda_1, \lambda_2, \dots, \lambda_{M+1}) = U_{M+1}^T S U_{M+1} - \lambda_{M+1} (U_{M+1}^T U_{M+1} - 1)$$

$$\nabla \mathcal{L}(\lambda_1, \lambda_2, \dots, \lambda_{M+1}) = 2 S U_{M+1} - \lambda_{M+1} U_{M+1} = 0$$

$$= S U_{M+1} = \lambda_{M+1} U_{M+1}$$

from this we can say that  $(M+1)^{\text{th}}$  term of  $W_{M+1}$  is the eigen vector corresponding to  $(M+1)^{\text{th}}$  largest eigen value of  $S$

and it is orthogonal to all the eigenvectors  $U_i$   $\forall 1 \leq i \leq M$ .

and projector vector  $W_{M+1} [U_1, U_2, \dots, U_{M+1}]$  is

first ~~first~~  $(M+1)$  largest eigen vector corresponding to first  $(M+1)$  largest eigen values.



Solution 2 b To prove  $\frac{\delta}{\delta A} \text{Tr}(AB) = 2B - \text{diag}(B)$ .

$$\frac{\delta \text{Tr}(AB)}{\delta A} = \begin{bmatrix} \frac{\delta \text{Tr}(AB)}{\delta A_{11}} & \frac{\delta \text{Tr}(AB)}{\delta A_{12}} & \dots & \frac{\delta \text{Tr}(AB)}{\delta A_{1n}} \\ \vdots & & & \\ \frac{\delta \text{Tr}(AB)}{\delta A_{n1}} & \dots & \dots & \frac{\delta \text{Tr}(AB)}{\delta A_{nn}} \end{bmatrix} \quad \text{--- (1)}$$

$\text{Tr}(AB)$  can be written as.

$$\sum_{i=1}^n \sum_{j \neq i}^n a_{ij} b_{ji} + \sum_{i=1}^n a_{ii} b_{ii}$$

We can break the first term into two parts.

~~$\sum_{i=1}^n \sum_{j \neq i}^n a_{ij} b_{ji}$~~   $a_{ij} b_{ji} + a_{ji} b_{ij}$  for a specific  $(i, j)$  and

they will be same due to symmetry. therefore we can say that ~~there is~~

$$\frac{\delta \text{Tr}(AB)}{\delta A_{ij}} = 2B_{ji} \quad \text{for non diagonal}$$

$$\frac{\delta \text{Tr}(AB)}{\delta A_{ii}} = B_{ii} \quad \text{for diagonal term}$$

from (1)

$$\frac{\delta \text{Tr}(AB)}{\delta A} = \begin{bmatrix} B_{11} & 2B_{12} & 2B_{13} & \dots & 2B_{1n} \\ \vdots & & & & \\ 2B_{n1} & \dots & \dots & \dots & B_{nn} \end{bmatrix} \quad \left\{ \begin{array}{l} \text{Comment:} \\ [B_{ij} = B_{ji}] \end{array} \right\}$$

We can write

$$\begin{bmatrix} 2B_{11} - B_{11} & 2B_{12} & 2B_{13} & \dots & 2B_{1n} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 2B_{n1} & \dots & \dots & \dots & 2B_{nn} - B_{nn} \end{bmatrix}$$

$$= \begin{bmatrix} 2B_{11} & 2B_{12} & \dots & \dots & 2B_{1n} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 2B_{n1} & \dots & \dots & \dots & 2B_{nn} \end{bmatrix} - \begin{bmatrix} B_{11} & 0 & 0 & \dots & 0 \\ 0 & B_{22} & & & \\ \vdots & & \ddots & & \\ 0 & & & & \\ 0 & & & & \end{bmatrix}$$

$$\Rightarrow 2B - \text{diag}(B)$$

Solution a. To prove  $\frac{\delta \log(|A|)}{\delta A} = 2A^{-1} - \text{diag}(A^{-1})$

$$\rightarrow \frac{\delta \log(|A|)}{\delta A} = \frac{1}{|A|} \frac{\delta A}{\delta \det(A)}$$

$$\frac{\delta \det(A)}{\delta A} = \begin{bmatrix} \frac{\delta \det(A)}{\delta A_{11}} & \frac{\delta \det(A)}{\delta A_{12}} & \dots & \dots & \frac{\delta \det(A)}{\delta A_{1n}} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \frac{\delta \det(A)}{\delta A_{in}} & \dots & \dots & \dots & \frac{\delta \det(A)}{\delta A_{nn}} \end{bmatrix}$$



Solution 3. Covariance  $\mathbf{Q}_y$  of the whitened input output.

$$\mathbf{S}_T^y = \frac{1}{N} \sum_{n=1}^N \mathbf{y}_n \mathbf{y}_n^T \quad \text{--- (1)}$$

To show  $\mathbf{S}_T^y = \mathbf{I}$  where  $\mathbf{I}$  is identity matrix

We know that  $\mathbf{y}_n = \mathbf{\Lambda}^{-1/2} \mathbf{W}^T (\mathbf{x} - \boldsymbol{\mu})$  where  $\mathbf{W}$  is eigen vector &  $\mathbf{\Lambda}$  is eigen value diagonal matrix.

Substituting  $\mathbf{y}_n$  value in (1).

$$\begin{aligned} \mathbf{S}_T^y &= \frac{1}{N} \sum_{n=1}^N \mathbf{\Lambda}^{-1/2} \mathbf{W}^T (\mathbf{x} - \boldsymbol{\mu}) \cdot \left( \mathbf{\Lambda}^{-1/2} \mathbf{W}^T (\mathbf{x} - \boldsymbol{\mu}) \right)^T \\ &= \frac{1}{N} \sum_{n=1}^N \mathbf{\Lambda}^{-1/2} \mathbf{W}^T (\mathbf{x} - \boldsymbol{\mu}) (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{W} (\mathbf{\Lambda}^{-1/2})^T \\ &= \mathbf{\Lambda}^{-1/2} \mathbf{W}^T \underbrace{\frac{1}{N} \sum_{n=1}^N (\mathbf{x} - \boldsymbol{\mu}) (\mathbf{x} - \boldsymbol{\mu})^T}_{\mathbf{S}} \mathbf{W} (\mathbf{\Lambda}^{-1/2})^T \quad \left\{ \mathbf{\Lambda}^{-1/2} \text{ is diagonal} \right\} \quad \text{--- (2)} \end{aligned}$$

Since  $\mathbf{S}$  is the covariance matrix & it is symmetric we can write it in the form of eigen decomposition

$$\mathbf{S} = \mathbf{W} \mathbf{\Lambda} \mathbf{W}^T \quad \& \{ \mathbf{W}^T \mathbf{W} = \mathbf{I} \}$$

therefore we can write (2) as, by orthogonality of eigen vectors

$$\mathbf{\Lambda}^{-1/2} \underbrace{\mathbf{W}^T \mathbf{W}}_{\mathbf{I}} \mathbf{\Lambda} \underbrace{\mathbf{W}^T \mathbf{W}}_{\mathbf{I}} \mathbf{\Lambda}^{-1/2}$$

$$\mathbf{\Lambda}^{-1/2} \mathbf{\Lambda} \mathbf{\Lambda}^{-1/2} = \mathbf{I}$$

hence proved.

Solution to. We first define  $S_T = S_B + S_W$ .

where  $S_T$  is Total Variance. which is  $I$  in LDA

$S_B$  is B/w Class Variance

$S_W$  is within Class Variance

The first LDA Projection vector  $w$  is eigen vector of  $S_W^{-1} S_B$  corresponding to largest eigen value and  $S_W^{-1} S_B$  is positive definite matrix.

Assume  $S_W$  is invertible matrix

$$S_W^{-1} S_B U_i = \lambda_i U_i \quad \text{--- (1)} \quad U_i \rightarrow \text{eigen vector } (w = U_i)$$

from (1) we can write  $S_B = S_T - S_W$ .

$$S_B = I - S_W$$

Substituting it in (1)

$$S_W^{-1} (I - S_W) U_i = \lambda_i U_i$$

$$(S_W^{-1} - I) U_i = \lambda_i U_i$$

$$S_W^{-1} U_i = \lambda_i U_i + U_i$$

$$S_W^{-1} U_i = (\lambda_i + 1) U_i$$

$$S_W U_i = \left( \frac{1}{\lambda_i + 1} \right) U_i$$

where  $\left( \frac{1}{\lambda_i + 1} \right)$  is eigen ~~value~~ value of  $S_W$ .

from this we have shown that the eigen vector  $U_i$  is corresponding to minimum <sup>of  $S_W$</sup>  eigen value. Since  $\lambda_i$  was magnitude of

maximum for  $S_W^{-1} S_B$ .

hence proved.