

Goal

Does local ancestry affect the gene network?

Two nodes and their correlations

Consider two genes k_1, k_2 and the corresponding gene expression vector of y_{k_1} and y_{k_2} , each of length n where n is the number of subjects. We also have two local ancestry vectors x_{k_1} and x_{k_2} . The temporary question of interest is whether the correlation between y_{k_1} and y_{k_2} is driven by either x_{k_1} or x_{k_2} .

Without loss of generality, let's see how much the correlation between the two expression levels changes with respect to the values of x_{k_1} . Each element of the local ancestry vector x takes one of three values : 0, 1, or 2. When we divide the n samples into three groups based on local ancestry each with size n_0, n_1, n_2 , we can observe three correlation coefficients: ρ_0, ρ_1 , and ρ_2 . Using Fisher's transformation, we consider three normally distributed z values for each group : z_0, z_1 , and z_2 .

$$z_i = \frac{1}{2} \ln \left(\frac{1 + \rho_i}{1 - \rho_i} \right) \sim N \left(\frac{1}{2} \ln \left(\frac{1 + \tilde{\rho}_i}{1 - \tilde{\rho}_i} \right), \frac{1}{\sqrt{n_i - 3}} \right)$$

where $\tilde{\rho}_i$ is the underlying true correlation coefficient for the group with local ancestry i .

We assume that z_0, z_1 , and z_2 have linear relationship. We formulate the model like below.

$$\begin{bmatrix} z_0 \\ z_1 \\ z_2 \end{bmatrix} \sim MVN \left(\begin{bmatrix} \mu \\ \mu + \beta \\ \mu + 2\beta \end{bmatrix}, \begin{bmatrix} \frac{1}{n_0-3} & 0 & 0 \\ 0 & \frac{1}{n_1-3} & 0 \\ 0 & 0 & \frac{1}{n_2-3} \end{bmatrix} \right)$$

with hypotheses

$$H_0 : \beta = 0$$

$$H_A : \beta \neq 0$$

For notational convenience, we define the following matrix and vectors.

$$\Sigma = \begin{bmatrix} \frac{1}{n_0-3} & 0 & 0 \\ 0 & \frac{1}{n_1-3} & 0 \\ 0 & 0 & \frac{1}{n_2-3} \end{bmatrix} = \begin{bmatrix} \sigma_0^2 & 0 & 0 \\ 0 & \sigma_1^2 & 0 \\ 0 & 0 & \sigma_2^2 \end{bmatrix}$$

$$\boldsymbol{\mu} = \begin{bmatrix} \mu \\ \mu \\ \mu \end{bmatrix}, \quad \boldsymbol{\mu}_\beta = \begin{bmatrix} \mu \\ \mu + \beta \\ \mu + 2\beta \end{bmatrix}, \quad \mathbf{z} = \begin{bmatrix} z_0 \\ z_1 \\ z_2 \end{bmatrix}$$

Also for notational convenience, we introduce Φ and Λ .

$$\Phi = \sigma_0^2 \sigma_1^2 + \sigma_1^2 \sigma_2^2 + \sigma_2^2 \sigma_0^2$$

$$\Lambda = \sigma_1^2 \sigma_2^2 z_0 + \sigma_0^2 \sigma_2^2 z_1 + \sigma_0^2 \sigma_1^2 z_2$$

Now, we set up the log likelihood of the two cases: when β is 0 and when β is the MLE estimator.

$$\ell_{\beta=0} = -\frac{1}{2} \log |2\pi\Sigma| - \frac{1}{2} ((\mathbf{z} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{z} - \boldsymbol{\mu})) \quad (1)$$

$$\frac{\partial \ell_{\beta=0}}{\partial \mu} = \frac{z_0 - \mu}{\sigma_0^2} + \frac{z_1 - \mu}{\sigma_1^2} + \frac{z_2 - \mu}{\sigma_2^2} = 0 \quad (2)$$

$$\Rightarrow \hat{\mu}_{0,MLE} = \frac{\sigma_1^2 \sigma_2^2 z_0 + \sigma_2^2 \sigma_0^2 z_1 + \sigma_0^2 \sigma_1^2 z_2}{\sigma_0^2 \sigma_1^2 + \sigma_1^2 \sigma_2^2 + \sigma_2^2 \sigma_0^2} = \frac{\Lambda}{\Phi} \quad (3)$$

is the MLE estimator for the μ under the null hypothesis.

For the alternative hypothesis,

$$\begin{aligned}\ell_{\beta \neq 0} &= -\frac{1}{2} \log |2\pi \Sigma| - \frac{1}{2} ((\mathbf{z} - \boldsymbol{\mu}_\beta)^T \Sigma^{-1} (\mathbf{z} - \boldsymbol{\mu}_\beta)) \\ \frac{\partial \ell_{\beta \neq 0}}{\partial \mu} &= \frac{z_0 - \mu}{\sigma_0^2} + \frac{z_1 - \mu - \beta}{\sigma_1^2} + \frac{z_2 - \mu - 2\beta}{\sigma_2^2} \\ \Rightarrow \hat{\mu}_{A,MLE} &= \frac{\sigma_1^2 \sigma_2^2 z_0 + \sigma_2^2 \sigma_0^2 (z_1 - \beta) + \sigma_0^2 \sigma_1^2 (z_2 - 2\beta)}{\sigma_0^2 \sigma_1^2 + \sigma_1^2 \sigma_2^2 + \sigma_2^2 \sigma_0^2} = \frac{\Lambda - \sigma_2^2 \sigma_0^2 \beta - 2\sigma_0^2 \sigma_1^2 \beta}{\Phi}\end{aligned}$$

is the MLE estimator for μ under the alternative hypothesis. Now we plug in the above to get the MLE estimator for β .

$$\begin{aligned}\frac{\partial \ell_{\beta \neq 0}}{\partial \beta} &= \frac{z_1 - \mu - \beta}{\sigma_1^2} + \frac{2(z_2 - \mu - 2\beta)}{\sigma_2^2} \\ \Rightarrow \sigma_2^2(z_1 - \hat{\mu}_{A,MLE} - \hat{\beta}_{A,MLE}) + 2\sigma_1^2(z_2 - \hat{\mu}_{A,MLE} - 2\hat{\beta}_{A,MLE}) &= 0 \\ \Rightarrow \frac{\Phi z_1 - \Lambda + (\sigma_0^2 \sigma_2^2 + 2\sigma_0^2 \sigma_1^2 - \Phi) \hat{\beta}_{A,MLE}}{\Phi \sigma_1^2} + \frac{2(\Phi z_2 - \Lambda + (\sigma_0^2 \sigma_2^2 + 2\sigma_0^2 \sigma_1^2 - 2\Phi) \hat{\beta}_{A,MLE})}{\Phi \sigma_2^2} &= 0 \\ \Rightarrow \hat{\beta}_{A,MLE} &= \frac{(\sigma_2^2 + 2\sigma_1^2) \Lambda - (\sigma_2^2 z_1 + 2\sigma_1^2 z_2) \Phi}{(\sigma_2^2 + 2\sigma_1^2)(\sigma_0^2 \sigma_2^2 + 2\sigma_0^2 \sigma_1^2 - 2\Phi)}\end{aligned}$$

Now we can perform the likelihood ratio test and see if β is significantly different from 0. Under the null,

$$-2 \log \left(\frac{N_3(\mathbf{z}; \hat{\boldsymbol{\mu}}, \Sigma)}{N_3(\mathbf{z}; \hat{\boldsymbol{\mu}}_\beta, \Sigma)} \right) \sim \chi_1^2$$

and we can compute the corresponding p -value.