

Lecture 3

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1 Singular Value Decomposition

Theorem 1.1: Singular Value DecompositionSuppose $\mathbf{A} \in \mathbb{R}^{m \times d}$ and

$$\mathbf{v}_i := \operatorname{argmax}_{\substack{\mathbf{v}: \mathbf{v} \perp \mathbf{v}_j, j \leq i \\ \|\mathbf{v}\|=1}} \|\mathbf{A}\mathbf{v}\|^2. \quad (1)$$

Then,

1. $V_k := \operatorname{span}\{\mathbf{v}_1, \mathbf{v}_2, \dots, \mathbf{v}_k\}$ is a subspace of dimension k that minimizes

$$\sum_{i=1}^m d(\mathbf{A}_i, \mathbf{v})^2. \quad (2)$$

That is, V_k is a least squares subspace.

2. $\mathbf{v}_1, \mathbf{v}_2, \dots, \mathbf{v}_d$ are right singular vectors of \mathbf{A} with corresponding left singular vectors $\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_d$ and singular values $\sigma_1, \sigma_2, \dots, \sigma_d$. Thus,

$$\mathbf{A} = \sum_{i=1}^d \sigma_i \mathbf{u}_i \mathbf{v}_i^\top. \quad (3)$$

Proof: The proof of 2 is an exercise (similar to \mathbf{v}_1 from last lecture). Here we prove 1 by induction on k . We showed $k = 1$ last time. Suppose V_{k-1} is a $(k-1)$ -dimensional least squares subspace. Let V'_k be a k -dimensional least squares subspace. Let $\mathbf{w}_1, \mathbf{w}_2, \dots, \mathbf{w}_k$ be an orthonormal basis of V'_k such that $\mathbf{w}_k \perp V_{k-1}$. Then, V'_k maximizes

$$\sum_{i=1}^k \|\mathbf{A}\mathbf{w}_i\|^2 \leq \sum_{i=1}^m d(\mathbf{A}_i, V_{k-1})^2 + \|\mathbf{A}\mathbf{w}_k\|^2. \quad (4)$$

Since \mathbf{w}_k was a candidate in the optimization of V_k ,

$$\sum_{i=1}^m d(\mathbf{A}_i, V_{k-1})^2 + \|\mathbf{A}\mathbf{w}_k\|^2 \leq \sum_{i=1}^m d(\mathbf{A}_i, V_{k-1})^2 + \|\mathbf{A}\mathbf{v}_k\|^2 \quad (5)$$

$$= \sum_{i=1}^m d(\mathbf{A}_i, V_k)^2. \quad (6)$$

So V_k is a k -dimensional least squares subspace. ■

2 SVD and Mixtures of Gaussians

Theorem 2.1: SVD and Mixtures of Gaussians

Suppose F is a mixture of Gaussians with weights w_i and $F_i = \mathcal{N}(\boldsymbol{\mu}_i, \sigma_i^2 \mathbf{I}_d)$. Consider the algorithm which projects the sample to its top- k SVD subspace V_k , then clusters in \mathbb{R}^k using distances (as in last lecture). Then,

1. V_k contains $\{\boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \dots, \boldsymbol{\mu}_k\}$.
2. The algorithm succeeds with probability $1 - \delta$ if

$$\|\boldsymbol{\mu}_i - \boldsymbol{\mu}_j\| \geq c \left(\log \frac{m}{\delta} \cdot k \right)^{1/4} \cdot \max(\sigma_i, \sigma_j). \quad (7)$$

Proof: 2 follows from 1 and last lecture. Here we prove 1. Suppose $k = 1$, then

$$\mathbf{v}_1 = \operatorname{argmax}_{\|\mathbf{v}\|=1} \mathbb{E}_F[(\mathbf{x}^\top \mathbf{v})^2] \quad (8)$$

$$= \mathbb{E} \left[((\mathbf{x} - \boldsymbol{\mu} + \boldsymbol{\mu}) \cdot \mathbf{v})^2 \right] \quad (9)$$

$$= \mathbb{E} \left[((\mathbf{x} - \boldsymbol{\mu}) \cdot \mathbf{v})^2 \right] + (\boldsymbol{\mu} \cdot \mathbf{v})^2 \quad (10)$$

$$= \sigma^2 + (\boldsymbol{\mu} \cdot \mathbf{v})^2 \quad (11)$$

$$\implies \mathbf{v} = \frac{\boldsymbol{\mu}}{\|\boldsymbol{\mu}\|}. \quad (12)$$

Because the Gaussians are spherical, the best k -dimensional subspace is any subspace containing $\boldsymbol{\mu}$. So for k Gaussians, we just contain all their means; this is optimal for each Gaussian individually, so it is optimal for their mixture. ■

Note that this strategy works for general Gaussians if the separation grows with the largest variance of the component Gaussians.

3 Linearly Independent Mixtures of Gaussians

What if we allow the Gaussians to overlap? Suppose we have a mixture F of k spherical Gaussians as above, and the only assumption we make is that $\boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \dots, \boldsymbol{\mu}_k$ are linearly independent. We can't use clustering here, but perhaps we can estimate the parameters of the model. Here, second moments will not suffice.

We will need the property of *isotropy* to proceed. F is isotropic if $\mathbb{E}_F[\mathbf{x}] = \mathbf{0}$ and $\mathbb{E}_F[\mathbf{x}\mathbf{x}^\top] = \mathbf{I}$.

Fact 3.1: Isotropy

Any distribution with bounded second moments can be made isotropic by an affine transformation.

Proof: Suppose $\mathbb{E}[\mathbf{x}] = \boldsymbol{\mu}$ and $\mathbb{E}[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^\top] = \mathbf{A} = \mathbf{B}\mathbf{B}^\top$. Let

$$\mathbf{y} = \mathbf{B}^{-1}(\mathbf{x} - \boldsymbol{\mu}) = \mathbf{A}^{-1/2}(\mathbf{x} - \boldsymbol{\mu}). \quad (13)$$

Then, $\mathbb{E}[\mathbf{y}] = \mathbf{0}$ and

$$\mathbb{E}[\mathbf{y}\mathbf{y}^\top] = \mathbf{B}^{-1}\mathbb{E}[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^\top](\mathbf{B}^{-1})^\top \quad (14)$$

$$= \mathbf{B}^{-1}\mathbf{B}\mathbf{B}^\top(\mathbf{B}^{-1})^\top \quad (15)$$

$$= \mathbf{I}. \quad (16)$$

■