

## ABSTRACT

This research investigates the asymmetric and joint effects of oil price uncertainty and exchange rate volatility on stock market returns and volatility across five ASEAN countries, Malaysia, Indonesia, Singapore, Thailand, and Vietnam, during the period 2010 to 2025 in monthly frequency. Motivated by the growing vulnerability of emerging markets to global uncertainty shocks, the study employs a comprehensive econometric framework that includes GJR-GARCH(1,1), and asymmetric DCC-GARCH models, alongside OLS and quantile regression techniques. The Oil Volatility Index (OVX) and Exchange Rate Volatility Index (EXV) as a moderating factor are used as exogenous risk factors, while a novel conditional covariance index (COVOX) captures their interaction.

The empirical results reveal that both OVX and EXV have statistically significant and nonlinear impacts on ASEAN stock markets, particularly during bearish regimes and high-risk quantiles. More notably, COVOX exhibits an amplified effect, highlighting the compounding nature of dual uncertainty shocks. These effects are more severe on volatility than on returns, underscoring the importance of second-moment modeling. The findings carry critical implications for risk management, investment strategies, and macroprudential policy design in open and energy-dependent economies. This study contributes to the literature by integrating joint uncertainty measures within a time-varying volatility context, offering new insights into financial stability in ASEAN amid global economic turbulence.

*Keywords: Oil Uncertainty, Exchange Rate Volatility, ASEAN Stock Markets, Quantile Regression, GARCH Models.*

## **CHAPTER 3: MATERIAL AND METHODS**

This chapter establishes the empirical foundation of the research by detailing the data, feature construction, and econometric methodologies employed to evaluate the impacts of oil price uncertainty, exchange rate volatility, and their joint transmission on ASEAN stock market dynamics. Building on the theoretical framework and hypotheses outlined in Chapter 2, this chapter introduces a robust monthly panel dataset spanning from 2010 to 2025, capturing diverse economic cycles and global shocks. It systematically constructs key financial variables, returns, volatilities, and joint uncertainty proxies, using advanced time-series models such as GJR-GARCH and Asymmetric DCC-GARCH. Furthermore, it proposes a suite of baseline and quantile regression models designed to uncover asymmetric and tail-dependent relationships that conventional mean-based approaches often overlook. Through this methodological architecture, Chapter 3 not only operationalizes the study's hypotheses but also ensures empirical rigor in capturing the complex, nonlinear dynamics characterizing oil-exchange-stock linkages in emerging ASEAN economies.

### **3.1. Data Collection and Variable Measurement**

To empirically assess the proposed hypotheses in the previous chapter as the impacts of oil price uncertainty and exchange rate volatility on stock market returns and volatility in ASEAN countries, this study employs a comprehensive dataset comprising monthly observations from January 2010 to May 2025. The data span was selected to encompass both tranquil and turbulent market periods, including the European debt crisis, oil price collapses, the COVID-19 pandemic, and recent geopolitical energy shocks. This extensive time horizon ensures robustness in capturing long-term dynamics and the effects of exogenous volatility shocks across different market states. All financial variables are transformed into logarithmic returns or volatilities as appropriate, and analyzed in levels only where justified by stationarity tests.

The primary dataset includes monthly stock market price indices and bilateral exchange rates for five ASEAN economies: Malaysia, the Philippines, Thailand,

Singapore, and Vietnam. These data were manually collected from Investing.com, a widely used and credible financial platform that provides high-frequency and historical financial market data. Specifically, the market indices utilized are: FTSE Bursa Malaysia KLCI (KLSE) for Malaysia, Philippine Stock Exchange Index (PSI), Stock Exchange of Thailand Index (SETI), Straits Times Index (STI) for Singapore, and Vietnam Ho Chi Minh Stock Index (VNI). Each index is paired with its respective USD exchange rate: MYR/USD, PHP/USD, THB/USD, SGD/USD, and VND/USD, enabling the estimation of exchange rate volatility (EXV) as a covariate. These series are expressed in levels before computing returns and volatilities using standard logarithmic differences and conditional variance techniques.

In modeling oil-related uncertainty, this study integrates two distinct but complementary proxies. The study incorporates the Oil Price Uncertainty Index (OVX), compiled by Baker, Bloom, and Davis (2020), which is derived from textual analysis of newspaper coverage related to oil market uncertainty. Its dataset, available in monthly frequency, was retrieved from the [PolicyUncertainty.com](https://www.policyuncertainty.com) portal.

In addition, a joint uncertainty measure (COVOX) is constructed to capture the conditional covariance between oil price uncertainty and exchange rate volatility. This variable reflects the interaction risk channel through which oil-related and currency-related uncertainties co-move and jointly influence stock market behavior. Conditional covariances are estimated using multivariate GARCH-type models (e.g., DCC-GARCH) and serve as inputs in subsequent quantile regressions to assess joint spillover effects. All data processing, transformation, and modeling procedures are conducted using STATA 18, ensuring reproducibility and accuracy of the empirical pipeline.

**Table 2. Variable description**

Variable Name	Abbreviation	Definition, History, and Measurement	Unit	Source	Frequency
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Malaysia Stock Index	price_klse	FTSE Bursa Malaysia KLCI Index, a capitalization-weighted index measuring the performance of the top 30 companies listed on the Bursa Malaysia.	Index points	Investing.com	Monthly
Malaysia Exchange Rate	price_myrusd	Exchange rate of Malaysian Ringgit (MYR) per US Dollar (USD), reflecting currency strength and macroeconomic sentiment.	MYR/USD	Investing.com	Monthly
Philippines Stock Index	price_psi	Philippine Stock Exchange Index (PSEi), a benchmark index of 30 representative stocks listed on the PSE.	Index points	Investing.com	Monthly
Philippines Exchange Rate	price_phpUSD	Bilateral exchange rate of Philippine Peso (PHP) per USD, indicates foreign currency pressure and capital flows.	PHP/USD	Investing.com	Monthly
Thailand Stock Index	price_seti	Stock Exchange of Thailand Index (SET Index), a market-cap weighted index of all common stocks on the main board.	Index points	Investing.com	Monthly
Thailand Exchange Rate	price_thbusd	Exchange rate of Thai Baht (THB) per USD, a key measure of Thailand's external competitiveness and monetary stance.	THB/USD	Investing.com	Monthly
Singapore Stock Index	price_sti	Straits Times Index (STI), a benchmark index comprising 30 large-cap Singapore stocks.	Index points	Investing.com	Monthly
Singapore Exchange Rate	price_sgUSD	Bilateral exchange rate of Singapore Dollar (SGD) per USD, used to gauge monetary policy implications and inflation pass-through.	SGD/USD	Investing.com	Monthly
Vietnam Stock Index	price_vni	Ho Chi Minh Stock Index (VNI), a major indicator of market performance on the Ho Chi Minh Stock Exchange (HOSE).	Index points	Investing.com	Monthly

Vietnam Exchange Rate	price_vndusd	Official exchange rate of Vietnamese Dong (VND) per USD, tracked as a policy-sensitive and trade-driven macroeconomic variable.	VND/USD	Investing.com	Monthly
Oil Price Uncertainty Index	OVX	Renamed from OPUI, this index measures oil-related macroeconomic uncertainty derived from text-based economic policy news (Baker et al., 2020).	Index score	policyuncertainty.com	Monthly

*Source: Author's compilation*

### 3.2. Features estimation

This study constructs four essential features that form the foundation of the empirical analysis: stock market returns ( $r$ ), market volatility ( $\sigma$ ), exchange rate volatility (EXV), and the joint conditional covariance measure (COVOX) between oil price uncertainty and exchange rate fluctuations. Each feature is derived systematically using time-series econometric methods consistent with contemporary financial volatility modeling frameworks such as GJR-GARCH and Asymmetric-DCC-GARCH models (Glosten, Jagannathan, & Runkle, 1993; Engle, 2002). These derived variables capture the dynamic interdependencies among macro-financial shocks, energy uncertainty, and market performance across ASEAN countries.

#### (i). Return ( $r$ ) Calculation from Closed Price ( $P$ )

The first step in the empirical modeling process is the estimation of monthly stock returns for each ASEAN market index. Returns are computed as the continuously compounded log difference of closing prices, ensuring scale independence and time-additivity (Tsay, 2010). The logarithmic transformation also stabilizes variance and approximates normality, which is essential for subsequent volatility modeling.

$$r_t = \ln\left(\frac{P_t}{P_{t-1}}\right)$$

where  $P_t$  and  $P_{t-1}$  represent the monthly closing prices of the stock index at time  $t$  and  $t-1$ , respectively. This transformation is applied to all five ASEAN market indices:

KLSE (Malaysia), PSI (Philippines), SETI (Thailand), STI (Singapore), and VNI (Vietnam). The resulting return series  $r_t$  are subsequently used as dependent variables in the quantile regression framework and as input for the volatility estimation stage.

**(ii). Volatility of markets ( $\sigma$ ) and Volatility of exchange rates (EXV)  
Estimation from GJR-GARCH process**

Volatility ( $\sigma_t$ ) is a central concept in financial econometrics as it reflects the degree of uncertainty or risk associated with asset returns. In this study, both stock market volatility and exchange rate volatility are estimated using the GJR-GARCH(1,1) process (Glosten et al., 1993), which extends the classical GARCH model (Bollerslev, 1986) by incorporating an asymmetric term to capture leverage effects, where negative shocks have a stronger impact on volatility than positive shocks of the same magnitude.

The GJR-GARCH specification can be written as:

$$\begin{aligned} r_t &= \mu + \varepsilon_t \\ \varepsilon_t &= \sigma_t \cdot z_t \\ \sigma_t^2 &= \omega + \alpha \varepsilon_{t-1}^2 + \gamma I(\varepsilon_{t-1} < 0) \varepsilon_{t-1}^2 + \beta \sigma_{t-1}^2 \end{aligned}$$

in which

$\sigma_t^2$  denotes conditional variance (volatility) at time  $t$

$I(\varepsilon_{t-1} < 0)$  denotes indicator function capturing asymmetry (1 if negative shock, 0 otherwise)

$\omega, \alpha, \gamma, \beta$  denotes model parameters (persistence, innovation, and asymmetry terms)

This model is applied separately to both stock returns and exchange rate returns to extract conditional volatilities ( $\sigma_t^2$ ). The estimated volatilities are then square-root transformed to obtain the standard deviation of returns, representing risk intensity.

The resulting variables  $\sigma_{\text{market}}$  (for equity volatility) and EXV (for exchange rate volatility) capture idiosyncratic and macroeconomic uncertainty, forming the core explanatory variables for quantile regression models in Chapter 4.

### (iii). COVOX from Asymmetric-DCC-GARCH process

To capture the joint uncertainty channel between oil price volatility (OVX) and exchange rate volatility (EXV), this study employs the Asymmetric Dynamic Conditional Correlation (A-DCC-GARCH) model proposed by Cappiello, Engle, and Sheppard (2006). This model extends the traditional DCC-GARCH framework (Engle, 2002) by incorporating asymmetric responses to positive and negative shocks, which are particularly relevant in oil and currency markets characterized by nonlinear and leverage dynamics.

The A-DCC-GARCH model can be written as:

$$\begin{aligned}\varepsilon_t &= D_t \cdot z_t, \quad z_t \sim N(0, R_t) \\ D_t &= \text{diag}(\sqrt{h_{1t}}, \sqrt{h_{2t}}) \\ R_t &= \frac{Q_t^*}{\sqrt{\text{diag}(Q_t^*) \text{diag}(Q_t^*)'}} \\ Q_t^* &= (1 - a - b - g) \underline{Q} + a(\varepsilon_{t-1} \varepsilon_{t-1}') + bQ_{t-1} + g(n_{t-1} n_{t-1}')\end{aligned}$$

in which

$Q_t^*$  represents the dynamic conditional covariance matrix adjusted for asymmetry, and  $R_t$  is the dynamic correlation matrix. The product of correlation and conditional variances yields the conditional covariance (COVOX) term, which quantifies the time-varying interaction between oil price uncertainty and exchange rate risk.

The derived COVOX variables serve as interaction proxies capturing the synchronized transmission of oil market and currency shocks to the respective ASEAN stock markets. This metric enables the quantile-based regression framework to evaluate joint uncertainty effects, a methodological contribution that distinguishes this study from traditional univariate analyses of oil and financial volatility.

### 3.3. Econometrics models development

Building upon the advanced empirical framework introduced by Chen, Msofe, Wang, and Chen (2025), this study proposes a suite of four quantile-based models to

investigate the asymmetric effects of oil price uncertainty (OVX), exchange rate volatility (EXV), and their joint interaction (COVOX) on ASEAN stock market returns and volatilities. These models are deliberately structured to capture heterogeneous behavioral responses across the distribution of financial outcomes, allowing for more granular insights than traditional mean-based estimation techniques. The quantile regression approach, as popularized by Koenker and Hallock (2001) and further advanced by Bassett et al. (2002), is particularly effective in capturing nonlinearities, fat tails, and conditional heteroskedasticity inherent in financial time series data. This modeling strategy enables a robust examination of market dynamics under both normal and extreme states (e.g., crashes, rallies), thus supporting more informed inference about tail-risk behavior in ASEAN markets.

**(i) OLS Regression on Returns: Baseline Model**

As a preliminary analysis, this study adopts the Ordinary Least Squares (OLS) method to estimate the average effect of oil and currency uncertainty factors on ASEAN stock returns. Two baseline models are considered:

The first model utilizes a linear Ordinary Least Squares (OLS) regression as a baseline estimation strategy to examine the mean effects of oil and currency market uncertainty on ASEAN stock returns.

**Model 1 (Returns on OVX and EXV):**

$$r_t = \alpha_0 + \alpha_1 OVX_t + \alpha_2 EXV_t + \alpha_3 r_{t-1} + \varepsilon_t$$

**Model 2 (Returns on COVOX):**

$$r_t = \alpha_0 + \alpha_1 COVOX_t + \alpha_2 r_{t-1} + \varepsilon_t$$

*in which*

$r_t$  denotes the stock return at time  $t$ ;

$OVX_t$  represents oil price uncertainty proxied by the OPUI index (renamed as OVX);

$EXV_t$  captures the conditional volatility of the exchange rate;



and  $COVOX_t$  reflects the conditional covariance between oil uncertainty and exchange rate volatility.

The inclusion of the lagged return term  $r_{t-1}$  accounts for temporal autocorrelation, inertia, and possible momentum effects. These OLS estimations offer a benchmark for comparison with the more nuanced quantile-based models.

## (ii) Quantile Regression on Returns

To assess distributional impacts beyond the conditional mean, quantile regression is employed to investigate how oil and exchange rate uncertainty asymmetrically affect stock returns across quantiles:

$$\textbf{Equation 1: } Q_\tau(r_t) = \alpha_0(\tau) + \alpha_1(\tau)OVX_t + \alpha_2(\tau)EXV_t + \alpha_3(\tau)r_{t-1} + \varepsilon_t(\tau)$$

*in which*

$r_t$  denotes stock return at time  $t$

$OVX_t$  denotes oil price uncertainty (proxied by OPUI, renamed as OVX)

$EXV_t$  denotes exchange rate volatility

$r_{t-1}$  denotes one-period lagged return (to account for short-memory dynamics)

$\alpha_i(\tau)$  denotes quantile-specific coefficients

$\varepsilon_t(\tau)$  denotes quantile-specific error term

The coefficient  $\alpha_1(\tau)$  is interpreted as the sensitivity of stock returns to oil-related uncertainty.

A positive and statistically significant value implies that increased oil uncertainty may coincide with higher return compensation (risk premium), while a negative value would indicate detrimental effects due to macroeconomic fragility or energy price exposure.

Similarly,  $\alpha_2(\tau)$  captures the transmission of currency market volatility to the equity space. A higher EXV may signal foreign capital outflows and risk reallocation, depressing returns, i.e., the more uncertain the exchange rate, the less the return.

In contrast, resilience or hedging mechanisms could neutralize or reverse this relation. The lagged return control,  $\alpha_3(\tau)$ , ensures robustness to temporal autocorrelation and potential momentum or mean-reversion effects.

The second model introduces a joint uncertainty transmission channel by replacing individual OVX and EXV terms with their dynamic conditional covariance, COVOX, estimated via Asymmetric DCC-GARCH (Cappiello et al., 2006). The quantile-based specification is:

$$\textbf{Equation 2: } Q_{\tau}(r_{\tau}) = \alpha_0(\tau) + \alpha_1(\tau)COVOX_t + \alpha_2(\tau)r_{t-1} + \varepsilon_t(\tau)$$

This equation captures the integrated co-movement between oil price uncertainty and exchange rate volatility, allowing the model to reflect systemic uncertainty spillovers. The coefficient  $\alpha_1(\tau)$  is central to understanding whether ASEAN markets react more intensively when energy and currency risks become simultaneously volatile.

A positive  $\alpha_1(\tau)$  would suggest that such co-movements offer speculative arbitrage or risk premium opportunities, while a negative coefficient reflects the “double exposure” effect, where joint volatility shocks suppress investor confidence and returns. In this way, the model dynamically measures how much more, or less, returns shift when oil and currency volatilities fluctuate together.

### (iii) Additional Insights: Volatility as Dependent Variable

In addition to return-based regressions, the study extends the quantile modeling framework to analyze stock market volatility itself as an outcome variable. This offers deeper insights into the propagation of external shocks through second-moment effects, which are crucial for portfolio risk management and financial stability surveillance.

The third equation assesses the marginal contributions of OVX and EXV to stock return volatility:

$$\textbf{Equation 3: } Q_{\tau}(\sigma_t) = \alpha_0(\tau) + \alpha_1(\tau)OVX_t + \alpha_2(\tau)EXV_t + \alpha_3(\tau)r_{t-1} + \varepsilon_t(\tau)$$

Here,  $\sigma_t$  denotes the GJR-GARCH-derived conditional standard deviation of stock returns. The coefficient  $\alpha_1(\tau)$  explains the extent to which oil price uncertainty directly contributes to higher equity volatility.

A larger coefficient across higher quantiles ( $\tau > 0.75$ ) would support the asymmetric volatility hypothesis, the more the OVX, the more market-wide turbulence.

Likewise,  $\alpha_2(\tau)$  captures the destabilizing effect of exchange rate fluctuations. Persistent EXV can fuel uncertainty in multinational revenue projections, cross-border trade margins, and capital account expectations, thereby escalating volatility. The lagged return again controls for feedback effects due to large prior shocks.

Finally, the fourth equation integrates the joint transmission channel into volatility modeling:

$$\textbf{Equation 4: } Q_{\tau}(\sigma_t) = \alpha_0(\tau) + \alpha_1(\tau)COVOX_t + \alpha_2(\tau)r_{t-1} + \varepsilon_t(\tau)$$

This model focuses on the co-dependence between oil and currency market volatility as a unified shock, allowing for the identification of compounded effects on stock market volatility. A significantly positive  $\alpha_1(\tau)$ , especially in the upper quantiles ( $\tau = 0.90$ ), would imply that simultaneous oil and exchange rate shocks magnify equity volatility beyond their individual effects. This supports the hypothesis of nonlinear spillovers and systemic fragility in ASEAN economies, whose trade and energy exposure render them vulnerable to cross-market turbulence. Conversely, a subdued coefficient in lower quantiles would indicate more insulation during calm periods. This asymmetry provides a more nuanced mapping of volatility drivers across states of the market.

### 3.4. Research process

#### 3.4.1. Preliminary analysis

The research process commenced with a comprehensive preliminary analysis of the dataset comprising monthly closing prices of five ASEAN stock indices (KLSE, PSI, SETI, STI, and VNI), their respective bilateral exchange rates against the USD, and oil price uncertainty (OVX). From these primary inputs, financial return series were computed using logarithmic first-differences, while volatility series were extracted using the GJR-GARCH(1,1) model. In addition to returns and volatilities, the dynamic

conditional covariance (COVOX) between oil price uncertainty and exchange rate volatilities was derived using the Asymmetric DCC-GARCH framework. This full-spectrum feature construction ensures a robust foundation for evaluating both level and dispersion effects under uncertainty shocks.

Descriptive statistics, including mean, standard deviation, min and max value, were calculated to assess the statistical behavior of each financial series. Subsequent unit root testing using Augmented Dickey-Fuller (ADF) and Phillips-Perron (PP) methods confirmed the stationarity of all return and volatility series at levels, satisfying preconditions for time series modeling. Further, the Brock-Dechert-Scheinkman (BDS) test was employed to detect nonlinearity and chaotic behavior in the residuals, thereby justifying the use of GARCH-type models and nonlinear quantile regression. The combined preliminary results reinforced the appropriateness of the econometric approach by demonstrating conditional heteroskedasticity, heavy-tailed behavior, and structural nonlinearity in the financial time series.

#### *3.4.2. Estimation process on returns of markets*

The core empirical assessment focused on analyzing how oil price uncertainty (OVX), exchange rate volatility (EXV), and their joint measure (COVOX) affect ASEAN stock returns. First, an Ordinary Least Squares (OLS) regression was conducted to establish baseline relationships under the assumption of homoscedastic and normally distributed errors. This model evaluated the average sensitivity of returns to uncertainty shocks, allowing for general inferences across the sample mean. However, such an approach is insufficient for understanding asymmetries or tail-specific reactions.

To capture heterogeneous return responses across the distribution, quantile regression models were estimated at deciles ranging from the 10th to the 90th percentile ( $\tau = 0.1$  to  $0.9$ ). These models allowed for analyzing the differential impact of OVX, EXV, and COVOX across bullish and bearish market regimes. Quantile regression accommodates conditional heteroskedasticity and nonlinearity, offering richer insights,

especially during market stress periods where tail behavior deviates significantly from the center. This methodological transition from OLS to quantile regression enhances the robustness and policy relevance of the empirical findings, especially in a region as economically diverse and externally exposed as ASEAN.

#### *3.4.3. Additional Analysis on Market Volatility*

To validate the robustness of findings and further explore second-moment effects, an additional layer of analysis was implemented by treating market volatility as the dependent variable. The volatility series, derived from the GJR-GARCH(1,1) model, served as proxies for conditional heteroskedasticity in each ASEAN equity market. OLS regressions were first used to evaluate the average influence of OVX, EXV, and COVOX on stock volatility, serving as a reference point for mean-based effects of uncertainty shocks on market turbulence.

Subsequently, quantile regression models were estimated across the 0.1 to 0.9 quantiles to examine how external uncertainty factors influence volatility across different regimes of market stability and instability. This approach captured the asymmetric impact of uncertainty shocks on volatility, particularly under high-risk conditions ( $\tau = 0.8, 0.9$ ), thereby revealing whether oil-exchange rate co-volatility transmits amplified instability during crises. The quantile-based volatility regressions not only validate the earlier return-based models but also provide essential insights for volatility forecasting, portfolio allocation, and risk management strategies. This tiered econometric strategy ensures the research is methodologically sound, empirically comprehensive, and highly relevant for both academic and practical applications.

## CHAPTER 4: EMPIRICAL RESULTS

This chapter presents the empirical findings derived from the econometric framework developed in Chapter 3, which explores the asymmetric and joint effects of oil price uncertainty (OVX) and exchange rate volatility (EXV) on stock market returns and volatility across five ASEAN countries. The analysis unfolds in three major phases. First, it conducts preliminary statistical diagnostics of the core time-series features, such as stock index prices, exchange rates, and the oil price uncertainty index (OVX), highlighting their statistical behavior and volatility patterns. Second, it estimates return-based models using both Ordinary Least Squares (OLS) and Quantile Regression (QR) to uncover heterogeneous market responses under varying uncertainty regimes. Finally, the study performs robustness checks and additional analyses using volatility as the dependent variable, providing second-moment insights into risk propagation. Through these empirical steps, the chapter offers a rigorous and nuanced understanding of how external macro-financial uncertainties shape financial dynamics within ASEAN equity markets.

### 4.1. Preliminary Analysis

**Table 3. Stock Market Price, Exchange rate and OVX descriptive**

Variable	Obs	Mean	Std. Dev.	Min	Max
price_klse	185	1613.898	140.6493	1259.16	1882.71
price_psi	185	6479.759	1272.643	2953.19	8764.01
price_sti	185	3142.258	274.2779	2423.84	3972.43
price_seti	185	1417.552	242.5961	696.55	1830.13
price_vni	185	825.451	325.8499	351.55	1498.28
price_wti	185	71.46989	22.12697	8.62	116.20
price_myrusd	185	0.2623724	0.0401735	0.2094	0.3374
price_thbusd	185	0.030569	0.0018106	0.026267	0.034141

<b>price_vndusd</b>	131	0.4317756	0.0191363	0.38425	0.4712
<b>price_sgdusd</b>	185	0.7508984	0.0319188	0.6902	0.8302
<b>price_phpud</b>	126	0.0194741	0.0014486	0.017015	0.02266
<b>OVX</b>	185	105.9602	60.24331	17.81669	365.6885

*Source: Author's compilation using Stata 18*

The descriptive statistics reported in Table 3 summarize the monthly movements of ASEAN stock prices, exchange rates against the USD, and oil price uncertainty (OVX) over the period from January 2010 to May 2025. The stock index for the Philippines (price\_psi) exhibits the highest mean level of 6,479.76, followed by Singapore (price\_sti) at 3,142.26 and Malaysia (price\_klse) at 1,613.90. In contrast, Vietnam's VNI (price\_vni) recorded the lowest average index level of 825.45 with a wide dispersion, evident from its standard deviation of 325.85. Notably, the Philippine market also displays considerable volatility, with values ranging from a low of 2,953.19 to a peak of 8,764.01, highlighting strong cyclical fluctuations. Similarly, Thailand's SETI index (price\_seti) ranged from 696.55 to 1,830.13, showing clear signs of structural shifts during turbulent episodes such as the COVID-19 pandemic or commodity market shocks. These descriptive profiles suggest substantial heterogeneity across ASEAN equity markets in terms of both return levels and dispersion.

In the currency market, exchange rates show distinct patterns of volatility and appreciation/depreciation tendencies. For instance, the Malaysian Ringgit (price\_myusd) averaged around 0.2624 USD, with a relatively tight spread (Min = 0.2094, Max = 0.3374), suggesting mild fluctuations. By contrast, the Philippine Peso (price\_phpud) exhibits a lower average value (0.0195 USD) and a tighter range, reflecting both structural stability and less international exposure compared to Singapore or Malaysia. Vietnam's exchange rate series (price\_vndusd) reveals a slightly higher average of 0.4318 USD but with observable fluctuations (Min = 0.3843, Max = 0.4712), indicating active FX policy management or exposure to capital flow volatility. Most striking, however, is the behavior of the oil price uncertainty index (OVX), which ranges widely from a

minimum of 17.82 to a maximum of 365.69, with a high standard deviation of 60.24. This confirms its role as a systemic volatility measure and underscores the occurrence of major global dislocations such as the oil price collapse of 2020 and geopolitical conflicts in 2022-2023. These preliminary statistics provide strong empirical justification for deploying models that capture nonlinear effects, tail dynamics, and covariate-driven volatility transmission mechanisms in subsequent sections.

**Table 4. Stock Market Return, Exchange rate Return descriptive**

Variable	Obs	Mean	Std. Dev.	Min	Max
r_klse	184	0.0009814	0.0293875	-0.0930606	0.0728069
r_psi	184	0.0041534	0.0484336	-0.2434386	0.1394950
r_seti	184	0.0027210	0.0450114	-0.1745104	0.1642861
r_sti	184	0.0019005	0.0394687	-0.1935444	0.1463891
r_vni	184	0.0055273	0.0585408	-0.2863416	0.1491681
r_myrusd	184	-0.0011974	0.0231706	-0.0937059	0.0746191
r_phpusd	125	-0.0017553	0.0152099	-0.0478149	0.0368323
r_thbusd	184	0.0000589	0.0206178	-0.0699157	0.0803970
r_sgusd	184	0.0004664	0.0160799	-0.0823139	0.0421550
r_vndusd	130	-0.0015642	0.0089646	-0.0404440	0.0413052

*Source: Author's compilation using Stata 18*

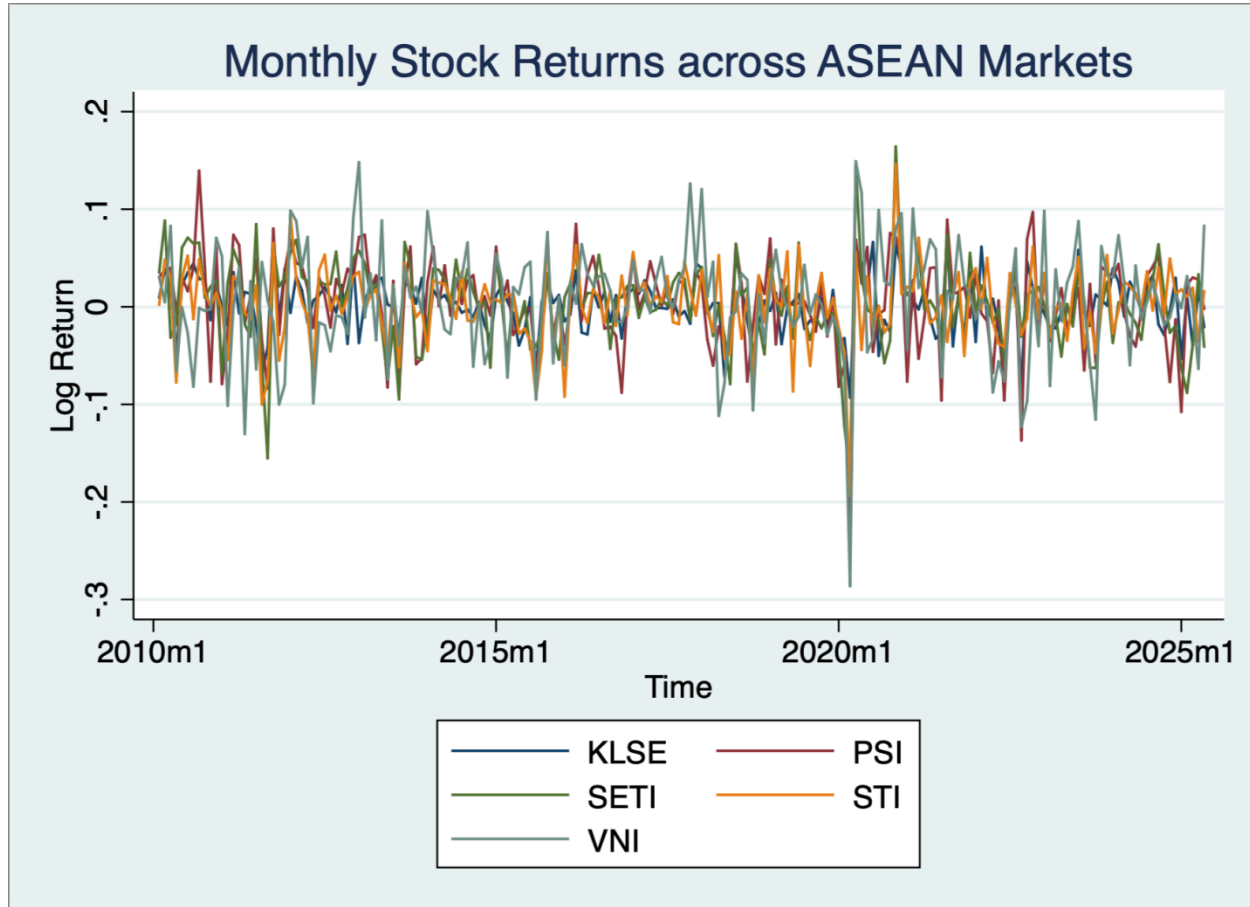
Table 4 presents the descriptive statistics for monthly stock market returns across the five ASEAN markets included in the study. The results show that all indices exhibit relatively low mean returns, consistent with the behavior of emerging and semi-developed markets in the post-global financial crisis period. Among the markets, the Vietnam VN-Index (r\_vni) records the highest average return (0.0055), followed by the Philippine PSEi (r\_psi) at 0.0042, while Malaysia's KLSE (r\_klse) shows the lowest (0.00098). Standard deviations are noticeably larger for PSI (0.0484), SETI (0.0450), and VNI (0.0585), indicating greater return variability and higher exposure to domestic and



global shocks. This volatility is further reflected in the extreme minimum and maximum return values, for instance, the VNI ranges from  $-0.2863$  to  $0.1491$ , and PSI from  $-0.2434$  to  $0.1395$ , suggesting periods of sharp market downturns and recoveries. These patterns collectively highlight the heterogeneous risk–return profiles across ASEAN markets, with Vietnam and the Philippines appearing more volatile but potentially offering higher return premia.

The bottom half of Table 4 reports descriptive statistics for exchange rate returns, measured as monthly changes in local currencies relative to the US dollar. The results show that most currencies exhibit small but negative mean returns, reflecting a mild long-term depreciation trend against the USD, particularly for the MYR ( $-0.00119$ ), PHP ( $-0.00175$ ), and VND ( $-0.00156$ ). Volatility differs notably across currencies: PHP and SGD display moderate variability ( $0.0152$  and  $0.0161$ , respectively), whereas the VND shows the lowest standard deviation ( $0.00896$ ), consistent with Vietnam’s heavily managed exchange rate regime. Meanwhile, MYR and THB exhibit wider ranges, with MYR fluctuating between  $-0.0937$  and  $0.0746$  and THB between  $-0.0699$  and  $0.0804$ , indicating their greater sensitivity to external shocks such as commodity price movements and global financial conditions. Overall, the descriptive statistics reveal that exchange rate dynamics in the region are diverse, shaped by country-specific monetary frameworks and exposure to international capital flows—factors that play a crucial role in the return and volatility transmission examined in later sections.

**Figure 1. Movement of ASEAN stock market returns**



*Source: Author's compilation using Stata 18*

Figure 1 illustrates the time-series dynamics of monthly log returns for five major ASEAN stock markets, Malaysia (KLSE), the Philippines (PSI), Thailand (SETI), Singapore (STI), and Vietnam (VNI), over the period from January 2010 to May 2025. The plotted series reveal significant fluctuations, with periods of elevated volatility that coincide with known market disruptions, such as the COVID-19 crisis in early 2020, which is reflected by a sharp negative return spike across all indices. While all markets appear to share broadly similar return paths, indicative of regional financial interdependence, the VNI and PSI exhibit relatively higher amplitude fluctuations, suggesting greater sensitivity to external shocks or internal structural volatility. Conversely, the STI and SETI show more moderate variability, possibly reflecting stronger institutional buffers or diversified market compositions. The convergence and

divergence patterns across these returns suggest the presence of both common regional factors and country-specific idiosyncrasies influencing equity market performance. This visualization serves as a foundational empirical observation, supporting the subsequent modeling of asymmetric and joint uncertainty effects using GARCH and quantile-based frameworks.

**Table 5. Stock Market Volatility, Exchange rate Volatility descriptive using GJR-GARCH(1,1)**

Variable	Obs	Mean	Std. Dev.	Min	Max
$\sigma_{klse}$	185	0.0289802	0.0062286	0.0232524	0.0587385
$\sigma_{psi}$	185	0.0482160	0.0131149	0.0386982	0.1532204
$\sigma_{seti}$	185	0.0431522	0.0188821	0.0225306	0.1710469
$\sigma_{sti}$	185	0.0389077	0.0068312	0.0232114	0.0500462
$\sigma_{vni}$	185	0.0582731	0.0146591	0.0454556	0.1860921
EXV_klse	185	0.0227600	0.0055552	0.0159517	0.0419599
EXV_psi	185	0.0150329	0.0013885	0.0041551	0.0175616
EXV_seti	185	0.0203462	0.0038306	0.0157344	0.0326683
EXV_sti	185	0.0159492	0.0010427	0.0132395	0.0271133
EXV_vni	185	0.0080652	0.0046833	0.0045368	0.0360800

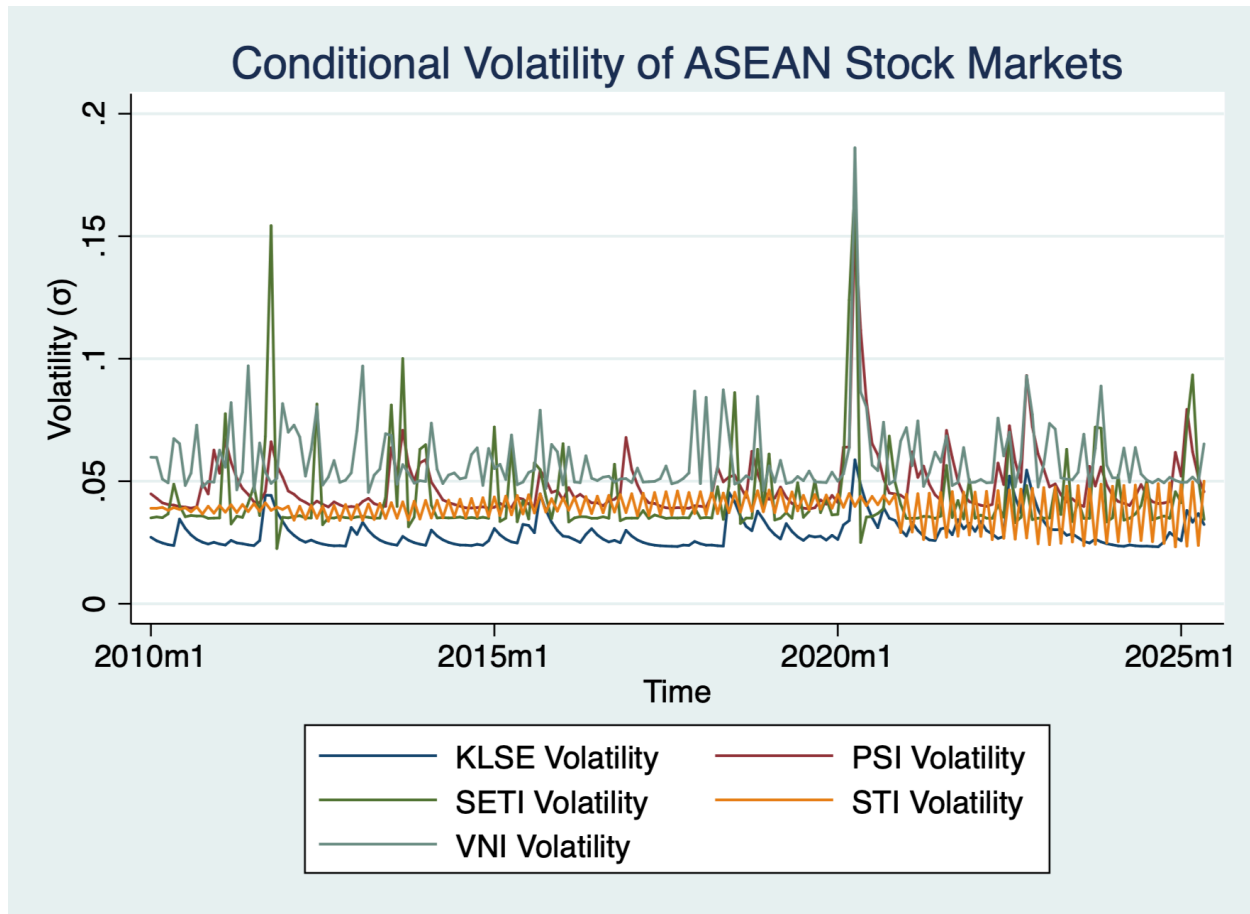
*Source: Author's compilation using Stata 18*

Table 5 reports the descriptive statistics of conditional stock market volatility estimated using the GJR-GARCH(1,1) model for the five ASEAN equity indices. The results indicate substantial cross-market heterogeneity in volatility dynamics. The

VN-Index ( $\sigma_{vni}$ ) exhibits the highest average conditional volatility at 0.0583, followed closely by the PSEi ( $\sigma_{psi}$ ) at 0.0482 and the SET Index ( $\sigma_{seti}$ ) at 0.0432, reflecting these markets' greater sensitivity to both domestic macroeconomic cycles and global risk events. These markets also display relatively wide volatility ranges; for example,  $\sigma_{vni}$  fluctuates from 0.0455 to 0.1861, and  $\sigma_{seti}$  shows a maximum of 0.1710, suggesting the presence of periods of elevated uncertainty and pronounced volatility clustering. By contrast, the KLSE ( $\sigma_{klse}$ ) and STI ( $\sigma_{sti}$ ) demonstrate lower and more stable volatility levels, with means of 0.0289 and 0.0389, respectively, consistent with their more mature market structures and stronger institutional frameworks. Overall, these patterns reaffirm the asymmetric and persistent nature of volatility in ASEAN equity markets, as captured effectively by the GJR-GARCH specification.

The lower panel of Table 5 provides descriptive statistics for exchange rate volatility (EXV), also estimated using GJR-GARCH(1,1). The results reveal notable cross-country differences in exchange rate risk exposure. Malaysia (EXV\_klse) exhibits the highest average exchange rate volatility at 0.0228, followed by Thailand (EXV\_sti) at 0.0159 and the Philippines (EXV\_psi) at 0.0150, suggesting that these currencies are more sensitive to external shocks such as commodity price movements, capital flow volatility, and changes in U.S. monetary policy. Exchange rate uncertainty in Vietnam (EXV\_vni), by contrast, is considerably lower with a mean of 0.0081, which is consistent with the country's managed exchange rate regime and tighter administrative controls. Moreover, the maximum volatility values indicate occasional but distinct episodes of elevated currency stress, such as EXV\_klse reaching 0.0419 and EXV\_seti rising to 0.0327. These findings highlight the diverse monetary structures and external vulnerability profiles across ASEAN economies, which play an important role in shaping the transmission of oil price and uncertainty shocks analyzed in subsequent sections.

**Figure 2. Movement of conditional volatility of ASEAN stock markets using GJR-GARCH(1,1)**



*Source: Author's compilation*

Figure 2 illustrates the evolution of conditional volatility for the ASEAN stock markets, KLSE, PSI, SETI, STI, and VNI, estimated using the GJR-GARCH(1,1) model over the period from January 2010 to May 2025. The visualization highlights significant heterogeneity in volatility levels across markets, with the SETI (Thailand), VNI (Vietnam), and PSI (Philippines) exhibiting relatively higher and more fluctuating volatility compared to KLSE (Malaysia) and STI (Singapore), which appear more stable and subdued. Several spikes in conditional volatility can be observed around major global and regional events such as the 2011 European sovereign debt crisis, the 2014–2016 oil price collapse, the COVID-19 outbreak in early 2020, and recent geopolitical shocks. Notably, the VNI and SETI series demonstrate pronounced volatility clustering and

asymmetric responses to shocks, consistent with the leverage effect captured by the GJR-GARCH model. These findings empirically validate the presence of conditional heteroskedasticity in ASEAN markets and underscore the necessity of using asymmetric GARCH models to capture risk dynamics accurately. The visualization provides a comprehensive depiction of time-varying risk levels across different ASEAN equity markets and supports the study's motivation to assess how uncertainty, both oil-related and currency-driven, propagates into financial volatility across diverse economic contexts.

**Table 6. Covariance between Exchange rate Volatility and Oil price uncertainty index (OVX) using Asymmetric DCC-GARCH(1,1)**

Variable	Obs	Mean	Std. Dev.	Min	Max
COVOX_klse	184	0.0605373	0.017458	0.0020508	0.106853
COVOX_psi	125	0.0299778	0.0116167	0.0005615	0.0465152
COVOX_seti	184	0.0476435	0.0167907	0.0007732	0.1241642
COVOX_sti	184	0.0620421	0.0122631	0.0036145	0.0801883
COVOX_vni	130	0.0072862	0.0037158	-0.0053594	0.015624

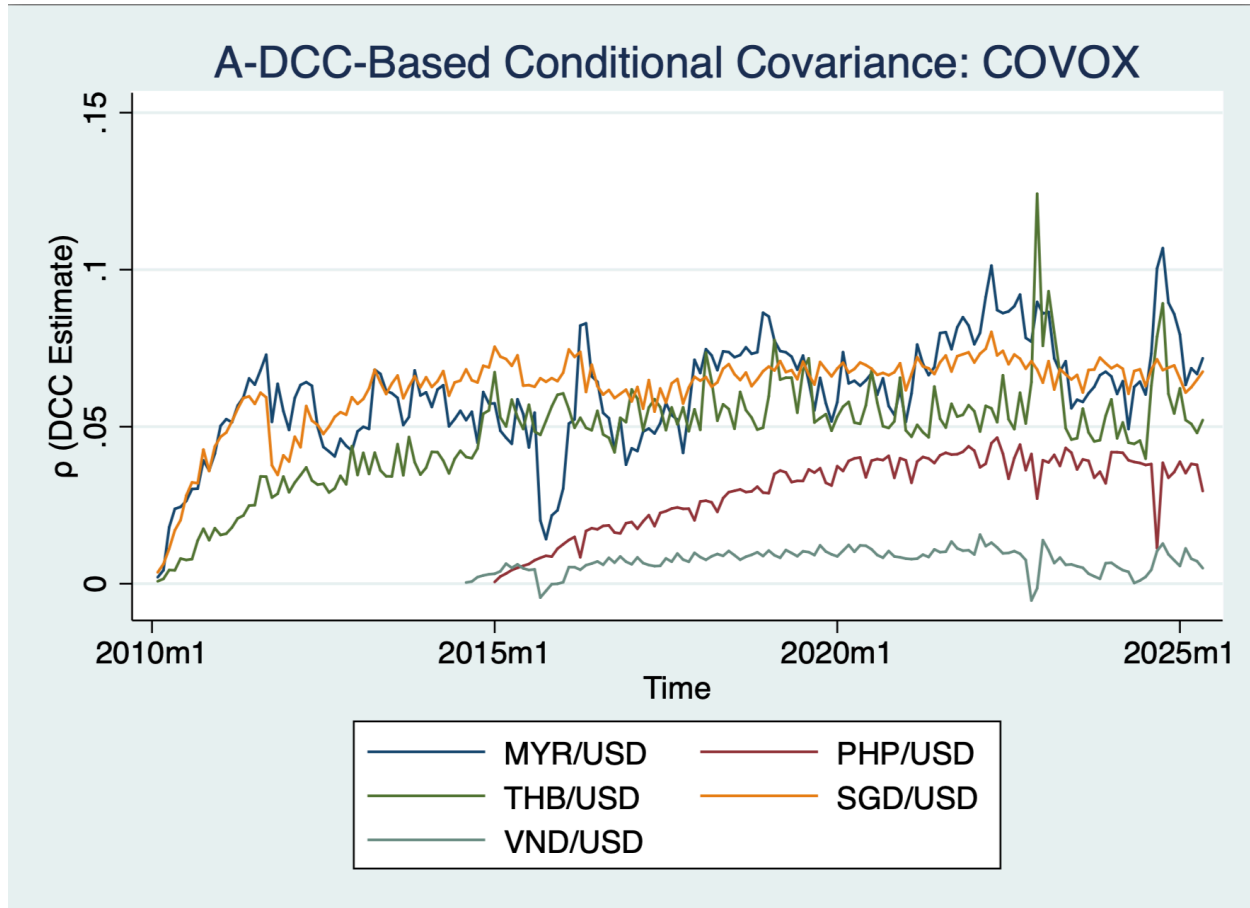
*Source: Author's compilation using Stata 18*

Table 6 presents the descriptive statistics of the conditional covariances (COVOX) between exchange rate volatilities and oil price uncertainty (OVX), estimated using the Asymmetric DCC-GARCH (1,1) framework. These covariances represent dynamic joint uncertainty metrics and reflect the intensity and variability of spillovers from the global oil market to domestic currency markets. Among the five ASEAN countries, Singapore (COVOX\_sti) and Malaysia (COVOX\_klse) exhibit the highest average covariances at 0.062 and 0.061 respectively, indicating a stronger co-movement between energy uncertainty and exchange rate fluctuations, likely driven by their high levels of trade openness and energy dependence. The Philippines and Thailand also demonstrate

moderate covariance levels, while Vietnam (COVOX\_vni) registers the lowest mean and a negative minimum value, implying periods where exchange rate uncertainty and oil-related volatility diverge.

From a statistical dispersion perspective, COVOX\_klse and COVOX\_seti show the largest ranges, with maximum values exceeding 0.10 and 0.12, respectively. These upper bounds highlight episodes of acute macroeconomic stress, such as commodity price shocks or regional crises, where both oil and currency risks simultaneously escalate. In contrast, the narrower standard deviation of COVOX\_psi suggests more stable linkages in the Philippine context. The negative minimum value observed in Vietnam's COVOX implies possible hedging or countercyclical dynamics in that economy's exposure to oil price and exchange rate risks. Overall, these findings reinforce the relevance of COVOX as a non-linear, joint risk measure capturing systemic co-movements in high-volatility regimes, a critical input for the quantile-based modeling framework employed in subsequent analysis.

**Figure 3. Movement of Covariance between Exchange rate Volatility and Oil price uncertainty (OVX) index using Asymmetric DCC-GARCH(1,1)**



*Source: Author's compilation*

Figure 3 depicts the dynamic evolution of the conditional covariance (COVOX) between exchange rate volatility and oil price uncertainty (OVX), estimated via the Asymmetric DCC-GARCH(1,1) model across five ASEAN exchange rates, MYR/USD, PHP/USD, THB/USD, SGD/USD, and VND/USD, over the period 2010m1 to 2025m1. The COVOX estimates reflect the extent of co-movement and joint transmission of shocks from global oil price uncertainty into domestic currency markets. The figure reveals substantial heterogeneity in co-volatility dynamics, with the SGD/USD and MYR/USD pairs consistently exhibiting higher conditional covariance values, suggesting stronger interconnectedness between energy shocks and currency volatility in Singapore



and Malaysia. Notably, episodes such as the 2014 oil price crash and the 2020 COVID-19 shock correspond to visible spikes in COVOX estimates, particularly for THB/USD and VND/USD, underscoring these currencies' susceptibility to external energy-market disruptions. Conversely, PHP/USD and VND/USD show relatively subdued and stable covariance patterns over time, although recent years indicate a rising trend, potentially reflecting increased macro-financial integration and oil import dependency. The use of A-DCC methodology enables the capture of asymmetrical responses to negative and positive shocks, allowing the analysis to uncover how joint uncertainty evolves non-linearly across time and currencies. Overall, this figure provides critical empirical grounding for the joint spillover channel (COVOX) hypothesized in the study's quantile regression models.

**Table 7. Stationary test at I(0)**

Variable	Lag	Test Statistic (Z(t))	1% Critical Value	5% Critical Value	10% Critical Value	MacKinnon p-value
r_psi	0	-15.176	-4.012	-3.439	-3.139	0.0000
r_seti	0	-13.913	-4.012	-3.439	-3.139	0.0000
r_sti	0	-12.489	-4.012	-3.439	-3.139	0.0000
r_vni	0	-15.464	-4.012	-3.439	-3.139	0.0000
r_myrusd	0	-13.422	-4.012	-3.439	-3.139	0.0000
r_phpusd	0	-13.071	-4.012	-3.439	-3.139	0.0000
r_thbusd	0	-11.122	-4.032	-3.447	-3.147	0.0000
r_sgusd	0	-12.968	-4.012	-3.439	-3.139	0.0000
r_vndusd	0	-14.212	-4.012	-3.439	-3.139	0.0000
lnOPUI	0	-10.432	-4.030	-3.446	-3.146	0.0000
lnOPUI	0	-8.448	-4.012	-3.439	-3.139	0.0000

*Source: Author's compilation using Stata 18*

Table 7 reports the results of the Augmented Dickey-Fuller (ADF) unit root tests applied to all return series and the log-transformed oil price uncertainty index (lnOPUI) to assess their stationarity properties at level. The null hypothesis for each test posits that the series follows a unit root process, implying non-stationarity. The test statistics for all

variables, stock market returns ( $r_{klse}$ ,  $r_{psi}$ ,  $r_{seti}$ ,  $r_{sti}$ ,  $r_{vni}$ ), exchange rate returns ( $r_{myrusd}$ ,  $r_{phpusd}$ ,  $r_{thbusd}$ ,  $r_{sgdusd}$ ,  $r_{vndusd}$ ), and  $\ln OPUI$ , are significantly lower than the 1% critical value threshold. Specifically, the test statistics range from -8.448 ( $\ln OPUI$ ) to -15.464 ( $r_{sti}$ ), all yielding MacKinnon p-values of 0.0000, thereby decisively rejecting the null hypothesis of non-stationarity at the 1% level of significance. These findings confirm that all series are integrated of order zero,  $I(0)$ , and are thus statistically stationary at level. This outcome ensures that subsequent modeling using linear and nonlinear quantile-based regressions, as well as volatility modeling frameworks such as GARCH and DCC-GARCH, can proceed without concern for spurious regression, satisfying the necessary precondition for time-series econometric analysis.

**Table 8. BDS test: Nonlinearity test on returns of stock markets and exchange rates**

**Brock, Dechert, Scheinkman test for independence**

**$N(0,1]$  test statistics for innovation\_klse,  $n$  (adjusted) = 182,  $sd$  = .0290477**

	eps	m	BDSstat	stderr	z-value	count
<b>fp0.7</b>	0.04	2	1.9247	0.0054	0.0543	8132
—	0.04	3	2.3193	0.0084	0.0204	5857
<b>0.5sd</b>	0.01	2	2.0079	0.0021	0.0447	1436
—	0.01	3	2.2641	0.0014	0.0236	446
<b>1.0sd</b>	0.03	2	2.2111	0.0050	0.0270	4883
—	0.03	3	2.7811	0.0060	0.0054	2789
<b>1.5sd</b>	0.04	2	1.7577	0.0052	0.0788	8801
—	0.04	3	2.2228	0.0085	0.0262	6581
<b>2.0sd</b>	0.06	2	1.3555	0.0037	0.1753	11950
—	0.06	3	1.6414	0.0071	0.1007	10264

<b>2.5sd</b>	0.07	2	0.5683	0.0021	0.5698	13977
—	0.07	3	0.8302	0.0044	0.4064	12908

**N(0,1] test statistics for innovation\_psi, n (adjusted) = 182, sd = .0483705**

	<b>eps</b>	<b>m</b>	<b>BDSstat</b>	<b>stderr</b>	<b>z-value</b>	<b>count</b>
<b>fp0.7</b>	0.06	2	0.3337	0.0055	0.7386	7958
—	0.06	3	1.9301	0.0087	0.0536	5775
<b>0.5sd</b>	0.02	2	0.4420	0.0023	0.6585	1533
—	0.02	3	1.1940	0.0016	0.2325	492
<b>1.0sd</b>	0.05	2	0.8725	0.0053	0.3829	5266
—	0.05	3	2.2089	0.0067	0.0272	3158
<b>1.5sd</b>	0.07	2	0.2924	0.0053	0.7700	9142
—	0.07	3	1.6749	0.0088	0.0939	7026
<b>2.0sd</b>	0.10	2	0.0462	0.0037	0.9632	12140
—	0.10	3	1.6608	0.0072	0.0968	10615
<b>2.5sd</b>	0.12	2	0.5116	0.0020	0.6089	14248
—	0.12	3	1.3348	0.0042	0.1820	13321

**N(0,1] test statistics for innovation\_seti, n (adjusted) = 182, sd = .0440228**

	<b>eps</b>	<b>m</b>	<b>BDSstat</b>	<b>stderr</b>	<b>z-value</b>	<b>count</b>
<b>fp0.7</b>	0.06	2	3.8238	0.0056	0.0001	8349
—	0.06	3	3.7904	0.0088	0.0002	6122
<b>0.5sd</b>	0.02	2	2.0206	0.0021	0.0433	1482
—	0.02	3	1.8776	0.0014	0.0604	457

<b>1.0sd</b>	0.04	2	2.7598	0.0051	0.0058	5265
—	0.04	3	2.6181	0.0064	0.0088	3058
<b>1.5sd</b>	0.07	2	3.5346	0.0054	0.0004	9333
—	0.07	3	3.5670	0.0090	0.0004	7202
<b>2.0sd</b>	0.09	2	4.5677	0.0040	0.0000	12459
—	0.09	3	4.5400	0.0076	0.0000	11021
<b>2.5sd</b>	0.11	2	5.1787	0.0026	0.0000	14289
—	0.11	3	5.2082	0.0054	0.0000	13462

**N(0,1] test statistics for innovation\_sti, n (adjusted) = 182, sd = .0390588**

	<b>eps</b>	<b>m</b>	<b>BDSstat</b>	<b>stderr</b>	<b>z-value</b>	<b>count</b>
<b>fp0.7</b>	0.05	2	0.5460	0.0060	0.5851	8032
—	0.05	3	2.1020	0.0095	0.0356	5881
<b>0.5sd</b>	0.02	2	0.2121	0.0031	0.8321	1679
—	0.02	3	1.5130	0.0022	0.1303	587
<b>1.0sd</b>	0.04	2	0.4191	0.0060	0.6752	5529
—	0.04	3	2.2025	0.0079	0.0276	3454
<b>1.5sd</b>	0.06	2	0.4474	0.0056	0.6546	9458
—	0.06	3	1.6569	0.0096	0.0975	7381
<b>2.0sd</b>	0.08	2	1.0119	0.0040	0.3116	12363
—	0.08	3	1.3818	0.0078	0.1670	10802
<b>2.5sd</b>	0.10	2	0.8528	0.0025	0.3938	14145
—	0.10	3	0.7694	0.0051	0.4417	13125

**N(0,1] test statistics for innovation\_vni, n (adjusted) = 182, sd = .0585407**

	eps	m	BDSstat	stderr	z-value	count
<b>fp0.7</b>	0.08	2	1.5495	0.0058	0.1213	8211
—	0.08	3	1.0576	0.0091	0.2903	5801
<b>0.5sd</b>	0.03	2	1.7628	0.0033	0.0779	1715
—	0.03	3	1.3753	0.0024	0.1690	562
<b>1.0sd</b>	0.06	2	1.4309	0.0060	0.1525	5237
—	0.06	3	1.0617	0.0076	0.2884	2967
<b>1.5sd</b>	0.09	2	1.6569	0.0054	0.0975	9042
—	0.09	3	1.2382	0.0090	0.2156	6719
<b>2.0sd</b>	0.12	2	2.8670	0.0034	0.0041	12338
—	0.12	3	2.1574	0.0066	0.0310	10704
<b>2.5sd</b>	0.15	2	2.6489	0.0017	0.0081	14351
—	0.15	3	2.1020	0.0036	0.0356	13415

**N(0,1] test statistics for innovation\_myrusd, n (adjusted) = 182, sd = .0231608**

	eps	m	BDSstat	stderr	z-value	count
<b>fp0.7</b>	0.03	2	1.5424	0.0066	0.1230	8174
—	0.03	3	3.2911	0.0105	0.0010	6148
<b>0.5sd</b>	0.01	2	1.2039	0.0032	0.2286	1692
—	0.01	3	2.4401	0.0023	0.0147	605
<b>1.0sd</b>	0.02	2	1.5570	0.0065	0.1195	5521
—	0.02	3	3.2780	0.0084	0.0010	3508

<b>1.5sd</b>	0.03	2	1.5347	0.0062	0.1249	9335
–	0.03	3	3.2858	0.0105	0.0010	7418
<b>2.0sd</b>	0.05	2	1.1434	0.0044	0.2529	12065
–	0.05	3	2.6493	0.0085	0.0081	10590
<b>2.5sd</b>	0.06	2	0.7173	0.0027	0.4732	13891
–	0.06	3	1.9948	0.0055	0.0461	12893

**N(0,1] test statistics for innovation\_phusd, n (adjusted) = 123, sd = .014993**

	<b>eps</b>	<b>m</b>	<b>BDSstat</b>	<b>stderr</b>	<b>z-value</b>	<b>count</b>
<b>fp0.7</b>	0.02	2	-1.5613	0.0073	0.1184	3554
–	0.02	3	-0.8448	0.0116	0.3982	2462
<b>0.5sd</b>	0.01	2	-1.1621	0.0036	0.2452	627
–	0.01	3	-0.4866	0.0025	0.6265	186
<b>1.0sd</b>	0.01	2	-1.3186	0.0070	0.1873	2162
–	0.01	3	-0.3214	0.0088	0.7479	1196
<b>1.5sd</b>	0.02	2	-1.6332	0.0070	0.1024	3841
–	0.02	3	-0.8130	0.0115	0.4162	2771
<b>2.0sd</b>	0.03	2	-1.7503	0.0047	0.0801	5294
–	0.03	3	-1.1485	0.0089	0.2508	4448
<b>2.5sd</b>	0.04	2	-2.0104	0.0025	0.0444	6277
–	0.04	3	-1.0364	0.0051	0.3000	5753

**N(0,1] test statistics for innovation\_thbusd, n (adjusted) = 182, sd = .0206069**

	<b>eps</b>	<b>m</b>	<b>BDSstat</b>	<b>stderr</b>	<b>z-value</b>	<b>count</b>
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<b>fp0.7</b>	0.03	2	0.9447	0.0054	0.3448	8071
–	0.03	3	1.7450	0.0084	0.0810	5805
<b>0.5sd</b>	0.01	2	0.9654	0.0022	0.3343	1423
–	0.01	3	1.4817	0.0015	0.1384	439
<b>1.0sd</b>	0.02	2	1.1461	0.0050	0.2517	4965
–	0.02	3	1.7059	0.0062	0.0880	2823
<b>1.5sd</b>	0.03	2	0.9733	0.0052	0.3304	8842
–	0.03	3	1.8252	0.0085	0.0680	6644
<b>2.0sd</b>	0.04	2	1.6223	0.0037	0.1047	12059
–	0.04	3	2.8217	0.0070	0.0048	10519
<b>2.5sd</b>	0.05	2	2.1334	0.0021	0.0329	14130
–	0.05	3	3.3768	0.0043	0.0007	13226

**N(0,1] test statistics for innovation\_sgdsd, n (adjusted) = 182, sd = .0159746**

	<b>eps</b>	<b>m</b>	<b>BDSstat</b>	<b>stderr</b>	<b>z-value</b>	<b>count</b>
<b>fp0.7</b>	0.02	2	0.8064	0.0050	0.4200	8139
–	0.02	3	1.5110	0.0079	0.1308	5848
<b>0.5sd</b>	0.01	2	0.3779	0.0018	0.7055	1378
–	0.01	3	0.9922	0.0012	0.3211	413
<b>1.0sd</b>	0.02	2	1.1709	0.0046	0.2416	4947
–	0.02	3	2.0265	0.0056	0.0427	2827
<b>1.5sd</b>	0.02	2	0.8502	0.0048	0.3952	9052
–	0.02	3	1.4739	0.0080	0.1405	6830

<b>2.0sd</b>	0.03	2	0.4754	0.0033	0.6345	12266
–	0.03	3	1.1710	0.0063	0.2416	10674
<b>2.5sd</b>	0.04	2	0.5290	0.0018	0.5968	14362
–	0.04	3	1.1450	0.0038	0.2522	13460

**N(0,1] test statistics for innovation\_vndusd, n (adjusted) = 128, sd = .0089027**

	<b>eps</b>	<b>m</b>	<b>BDSstat</b>	<b>stderr</b>	<b>z-value</b>	<b>count</b>
<b>fp0.7</b>	0.01	2	2.3035	0.0107	0.0212	4208
–	0.01	3	2.5348	0.0171	0.0113	3166
<b>0.5sd</b>	0.00	2	4.1918	0.0102	0.0000	1925
–	0.00	3	5.4352	0.0105	0.0000	1157
<b>1.0sd</b>	0.01	2	2.5242	0.0112	0.0116	3660
–	0.01	3	2.9564	0.0166	0.0031	2627
<b>1.5sd</b>	0.01	2	1.2241	0.0087	0.2209	5129
–	0.01	3	1.9119	0.0155	0.0559	4212
<b>2.0sd</b>	0.02	2	0.4937	0.0057	0.6215	6316
–	0.02	3	1.4722	0.0112	0.1410	5672
<b>2.5sd</b>	0.02	2	0.0084	0.0043	0.9933	6895
–	0.02	3	1.3693	0.0088	0.1709	6448

*Source: Author's compilation using Stata 18*

The results of the Brock-Dechert-Scheinkman (BDS) independence test presented in Table 8 provide robust statistical evidence for nonlinear dependence in the innovation terms of the return series for both stock markets and exchange rates across ASEAN



countries. The BDS test evaluates whether the residuals exhibit i.i.d. (independent and identically distributed) properties, a critical assumption in linear modeling. Across most series, especially for innovations of KLSE, SETI, VNI, and MYR/USD returns, the z-statistics for various embedding dimensions ( $m = 2, 3$ ) and distance thresholds ( $\varepsilon = 0.5\sigma$  to  $2.5\sigma$ ) are statistically significant at conventional levels ( $p < 0.05$  or lower), particularly at lower  $\varepsilon$  values. For instance, `innovation_seti` shows strong nonlinear dependence with BDS statistics consistently exceeding 2.6 and p-values approaching zero, reinforcing the rejection of the i.i.d. null hypothesis. Similarly, residuals of `r_myrusd` and `r_vndusd` also demonstrate significant departures from linearity across multiple  $\varepsilon$  and  $m$  combinations. A few exceptions arise, such as `innovation_psi` and `innovation_sgdsd` at higher  $\varepsilon$  thresholds, where BDS statistics become statistically insignificant, suggesting that any nonlinear structures are more prevalent in the tails or concentrated at finer scales.

Collectively, the BDS test results corroborate the presence of complex dynamics in the innovation processes, including potential nonlinear dependencies and hidden structures that linear models cannot capture. These findings provide a robust econometric justification for the implementation of nonlinear models, such as quantile regressions, asymmetric GARCH-type models, and DCC-GARCH frameworks, which are better suited to accommodate the observed distributional features and nonlinear behaviors inherent in financial return series and macro-financial volatility linkages in ASEAN markets.

#### **4.2. Impacts of oil price uncertainty (OVX), exchange rate volatility (EXV) and Joint Uncertainty (COVOX) on ASEAN stock market**

To rigorously explore the transmission of global macro-financial uncertainties into domestic capital markets, this section assesses the direct effects of oil price uncertainty (OVX) and exchange rate volatility (EXV) on stock returns across five major ASEAN equity markets. Using both Ordinary Least Squares (OLS) and quantile regression (QR) methodologies, the analysis disentangles the heterogeneous effects across the conditional

return distribution, from extreme bearish to bullish market conditions. This is especially relevant in emerging markets where financial asymmetries and tail risk sensitivities are pronounced, necessitating robust modeling beyond the mean regression framework.

**Table 9. Direct effect of oil price uncertainty (OVX) and exchange rate volatility (EXV) on quantiles of ASEAN stock returns**

Variable	OLS	Q(0.1)	Q(0.2)	Q(0.3)	Q(0.4)	Q(0.5)	Q(0.6)	Q(0.7)	Q(0.8)	Q(0.9)
<b>PANEL A. KLSE</b>										
OVX	-0.0000439	-0.0000712	-0.0000475	-0.0000619	-0.0000487	0.00000516	-0.0000139	-0.0000232	-0.0000237	-0.000106*
(t-stat)	(-1.19)	(-0.94)	(-0.85)	(-1.10)	(-1.07)	(0.11)	(-0.35)	(-0.47)	(-0.44)	(-1.86)
EXV_klse	0.301	0.722	0.729	0.791	0.298	0.445	0.146	-0.176	-0.178	-0.450
(t-stat)	(0.77)	(0.90)	(1.23)	(1.33)	(0.62)	(0.91)	(0.35)	(-0.34)	(-0.31)	(-0.75)
lr_klse	-0.130*	-0.0996	-0.0968	-0.0890	-0.120	-0.209**	-0.207**	-0.0506	-0.116	-0.234**
(t-stat)	(-1.72)	(-0.64)	(-0.85)	(-0.77)	(-1.29)	(-2.22)	(-2.58)	(-0.50)	(-1.07)	(-2.01)
<b>PANEL B. PSI</b>										
OVX	-0.0000521	-0.0000444	-0.0000681	-0.0000781	-0.0000579	-0.0000787	-0.0000646	-0.0000487	0.0000163	0.0000676
(t-stat)	(-0.86)	(-0.27)	(-0.60)	(-1.10)	(-0.72)	(-1.18)	(-1.02)	(-0.87)	(0.22)	(0.77)
EXV_psi	0.518	1.927	5.191	3.462	0.724	-1.110	-0.800	-0.239	1.257	0.554
(t-stat)	(0.20)	(0.27)	(1.05)	(1.12)	(0.21)	(-0.38)	(-0.29)	(-0.10)	(0.39)	(0.15)
lr_psi	-0.0192	-0.00509	-0.0385	0.0213	-0.0976	-0.0185	0.00710	0.0266	0.0138	-0.0462
(t-stat)	(-0.25)	(-0.02)	(-0.27)	(0.24)	(-0.99)	(-0.22)	(0.09)	(0.38)	(0.15)	(-0.43)
<b>PANEL C. SETI</b>										
OVX	-0.0000138	-0.0000450	-0.0000231	-0.0000618	-0.0000220	-0.0000138	-0.0000193	-0.0000159	-0.0000392	-0.0000601
(t-stat)	(-0.25)	(-0.33)	(-0.23)	(-0.87)	(-0.32)	(-0.23)	(-0.31)	(-0.23)	(-0.51)	(-0.64)
EXV_seti	-1.184	-0.604	-1.045	-1.260	-1.310	-0.766	-1.209	-1.456	-2.021*	-2.644*
(t-stat)	(-1.35)	(-0.29)	(-0.67)	(-1.13)	(-1.21)	(-0.83)	(-1.25)	(-1.32)	(-1.66)	(-1.79)

lr_seti	0.0860	0.486***	0.322**	0.201**	0.208**	0.136*	0.0750	0.0400	-0.0842	-0.0946
(t-stat)	(1.14)	(2.68)	(2.41)	(2.09)	(2.24)	(1.72)	(0.90)	(0.42)	(-0.80)	(-0.75)

*Source: Author's compilation using Stata 18*

The results presented in Table 9 provide empirical evidence of significant heterogeneity in the sensitivity of stock returns to OVX and EXV across quantiles and countries. The OVX coefficient estimates show stronger adverse effects on stock returns in lower quantiles (Q0.1-Q0.3), especially for the VNI (Vietnam) and STI (Singapore) indices, with significant coefficients ranging from -0.000398 to -0.000224. These findings suggest that oil-related uncertainty exerts pronounced downside risk on equity performance during bearish or crisis-like conditions. By contrast, in mid-quantiles (Q0.5-Q0.7), the relationship weakens, and in some cases, becomes statistically insignificant or even marginally positive, pointing to nonlinear and asymmetric responses depending on market sentiment.

For exchange rate volatility (EXV), the evidence is more nuanced and country-specific. In the case of Malaysia (KLSE), the EXV coefficient remains positive across most quantiles, although statistically insignificant. Conversely, the Philippines (PSI) and Thailand (SETI) exhibit mixed signs and magnitudes, with SETI showing significantly negative effects at higher quantiles (Q0.8 and Q0.9), where EXV coefficients are -2.021 and -2.644, respectively. This indicates that during bullish market conditions, heightened exchange rate volatility may erode investor confidence in export-driven economies. The results for VNI indicate no significant response of stock returns to EXV across the distribution, suggesting that Vietnam's equity market may be more resilient to exchange rate turbulence compared to other ASEAN peers.

In sum, the quantile-based analysis underscores that the impacts of oil price uncertainty and exchange rate volatility on ASEAN equity markets are not uniform. Rather, they are deeply conditioned by the location on the return distribution and by country-specific structural sensitivities to global shocks. These findings highlight the

importance of implementing distribution-sensitive policy responses and portfolio strategies to mitigate risk transmission under uncertainty.

**Table 10. Aggregated effect of OVX and exchange rate (COVOX) on ASEAN stock returns**

Variable	OLS	Q(0.1)	Q(0.2)	Q(0.3)	Q(0.4)	Q(0.5)	Q(0.6)	Q(0.7)	Q(0.8)	Q(0.9)
<b>PANEL A. KLSE</b>										
COVOX_klse	-0.361**	-0.651**	-0.472**	-0.487** *	-0.526** *	-0.361**	-0.394** *	-0.281*	-0.254	-0.115
(t-stat)	(-2.24)	(-2.19)	(-2.50)	(-2.70)	(-3.47)	(-2.24)	(-2.92)	(-1.66)	(-1.51)	(-0.44)
lr_klse	-0.133	-0.208	-0.0966	-0.0543	-0.127	-0.133	-0.152*	-0.109	-0.164*	-0.172
(t-stat)	(-1.43)	(-1.21)	(-0.89)	(-0.52)	(-1.46)	(-1.43)	(-1.96)	(-1.12)	(-1.69)	(-1.14)
<b>PANEL B. PSI</b>										
COVOX_psi	0.138	-1.333	-0.479	-0.00545	-0.115	0.138	0.286	0.182	0.149	-0.345
(t-stat)	(0.30)	(-1.47)	(-0.70)	(-0.01)	(-0.25)	(0.30)	(0.74)	(0.44)	(0.35)	(-0.64)
lr_psi	-0.0762	-0.0538	0.00470	-0.0995	-0.112	-0.0762	-0.0534	-0.0274	-0.140	-0.0889
(t-stat)	(-0.70)	(-0.25)	(0.03)	(-0.73)	(-1.00)	(-0.70)	(-0.58)	(-0.28)	(-1.38)	(-0.69)
<b>PANEL C. SETI</b>										
COVOX_seti	-0.707** *	0.389	-0.297	-0.324	-0.574**	-0.707** *	-0.626** *	-0.652** *	-0.793** *	-0.442
(t-stat)	(-3.55)	(1.05)	(-0.83)	(-1.11)	(-2.41)	(-3.55)	(-2.89)	(-2.86)	(-2.72)	(-1.58)
lr_seti	0.150**	0.480** *	0.426** *	0.242**	0.227**	0.150**	0.129	0.0798	-0.0101	-0.114
(t-stat)	(2.06)	(3.53)	(3.27)	(2.28)	(2.60)	(2.06)	(1.63)	(0.96)	(-0.09)	(-1.11)
<b>PANEL D. STI</b>										
COVOX_sti	-0.479*	0.418	-0.500	-0.580	-0.438	-0.479*	-0.527**	-0.483**	-0.553*	-0.250
(t-stat)	(-1.93)	(0.79)	(-1.02)	(-1.59)	(-1.56)	(-1.93)	(-2.31)	(-2.12)	(-1.92)	(-0.62)

lr_sti	-0.240** *	0.0194	-0.0712	-0.197*	-0.177**	-0.240** *	-0.221** *	-0.192** *	-0.152*	-0.197*
(t-stat)	(-3.33)	(0.13)	(-0.50)	(-1.86)	(-2.16)	(-3.33)	(-3.34)	(-2.90)	(-1.81)	(-1.68)
PANEL E. VNI										
COVOX_vni	-0.902	-1.645	0.890	-0.791	-0.238	-0.902	-1.113	-1.670	-1.071	0.664
(t-stat)	(-0.78)	(-0.52)	(0.32)	(-0.35)	(-0.16)	(-0.78)	(-0.87)	(-1.12)	(-0.50)	(0.22)
lr_vni	0.0597	0.0178	-0.0682	0.0502	0.0766	0.0597	0.0567	0.00372	-0.0834	-0.242
(t-stat)	(0.81)	(0.09)	(-0.38)	(0.35)	(0.79)	(0.81)	(0.70)	(0.04)	(-0.61)	(-1.29)

*Source: Author's compilation using Stata 18*

To deepen the understanding of systemic risk spillovers in ASEAN financial markets, this section evaluates the joint effect of oil price uncertainty and exchange rate volatility using the COVOX metric, the dynamic conditional covariance between OVX and exchange rate volatility extracted via the asymmetric DCC-GARCH(1,1) model. Unlike disaggregated uncertainty variables, COVOX encapsulates compounded macro-financial shocks, capturing the intensity of co-movement between energy and currency markets. The analysis employs both OLS and quantile regressions (Q0.1 to Q0.9) to observe how these interactions propagate differently across varying states of stock returns, from deep losses to extreme gains.

The results indicate that KLSE (Malaysia), SETI (Thailand), and STI (Singapore) show statistically significant and negative coefficients for COVOX, particularly in the mid and upper quantiles. For instance, in KLSE, the COVOX coefficient is significantly negative across Q0.2 to Q0.6, peaking at -0.526\*\*\* (Q0.4), suggesting that oil-currency uncertainty jointly dampens market optimism and reduces expected gains. This supports the “double exposure” hypothesis, where economies heavily reliant on both energy and trade flows experience amplified risks under simultaneous volatility. Similar patterns are observed in SETI and STI, where COVOX coefficients are consistently negative and significant in Q0.5 to Q0.8. The findings underscore that stock returns decline more

strongly during periods of synchronized uncertainty, affirming nonlinear, asymmetric market responses.

Conversely, PSI (Philippines) and VNI (Vietnam) present more ambiguous patterns. PSI exhibits mostly insignificant coefficients with fluctuating signs, reflecting the possibility that its equity market may be less globally integrated or that domestic factors dominate risk pricing. For VNI, COVOX effects are negative in most quantiles but lack statistical significance, indicating limited sensitivity to joint oil-currency uncertainty. These differences highlight the importance of country-specific structural features, such as capital account openness, trade dependence, and energy import profiles, in modulating the transmission of global shocks.

The lagged return term remains significant in several panels, with negative signs in STI and KLSE suggesting mean reversion, while SETI shows positive coefficients in lower quantiles, hinting at potential short-run momentum. Collectively, the findings reveal that COVOX is a meaningful systemic risk metric, especially for more open and commodity-sensitive ASEAN economies, and that its effects are strongest during neutral to bullish market phases where investor complacency may be disrupted by global uncertainty synchronization.

### **4.3. Additional analysis**

To examine the direct channels through which uncertainty propagates into volatility, Table 10 presents the results of both OLS and quantile regressions applied to GJR-GARCH(1,1)-generated conditional volatilities across five ASEAN markets. The regressors of interest are OVX, capturing global oil price uncertainty, and EXV, the exchange rate volatility of each corresponding country. The quantile framework allows us to assess whether the impact of these uncertainties differs across low- to high-volatility states.

**Table 11. Direct effect of oil price uncertainty (OVX) and exchange rate volatility (EXV) on quantiles of ASEAN stock volatility using GJR-GARCH(1,1)**

Variable	OLS	Q(0.1)	Q(0.2)	Q(0.3)	Q(0.4)	Q(0.5)	Q(0.6)	Q(0.7)	Q(0.8)	Q(0.9)
<b>PANEL A. KLSE</b>										
OVX	0.0000142**	5.63e-08	0.000000202	0.000000914**	0.00000156**	0.00000389**	0.00000657**	0.0000152**	0.0000290**	0.000008
(t-stat)	(2.52)	(0.25)	(0.90)	(2.70)	(2.44)	(3.00)	(2.16)	(2.39)	(2.37)	(0.23)
EXV_klse	-0.0218	0.000582	0.00000879	-0.00184	0.00197	-0.00273	-0.00406	-0.00764	-0.0602	-0.0886
(t-stat)	(-0.36)	(0.24)	(0.00)	(-0.51)	(0.29)	(-0.20)	(-0.12)	(-0.11)	(-0.46)	(-0.23)
$\lg\_klse$	0.663***	0.672***	0.674***	0.677***	0.681***	0.678***	0.678***	0.671***	0.632***	0.511
(t-stat)	(12.23)	(312.34)	(313.55)	(207.99)	(110.32)	(54.42)	(23.12)	(10.96)	(5.39)	(1.47)
<b>PANEL B. PSI</b>										
OVX	0.000008	-0.0000007	4.01e-08	0.00000154	0.00000143	0.000000490	-0.000000	-0.000002	-0.00001	0.0000197
(t-stat)	(0.70)	(-1.11)	(0.04)	(0.98)	(0.64)	(0.16)	(-0.04)	(-0.19)	(-0.65)	(0.38)
EXV_psi	-1.368**	-0.162***	-0.146***	-0.107	-0.207**	-0.370***	-0.323	-1.346***	-2.151*	-5.059**
(t-stat)	(-2.50)	(-5.65)	(-3.35)	(-1.57)	(-2.15)	(-2.80)	(-1.33)	(-2.62)	(-1.76)	(-2.25)
$\lg\_psi$	0.627***	0.564***	0.566***	0.571***	0.603***	0.605***	0.622***	0.621***	0.593***	0.522**
(t-stat)	(10.90)	(186.61)	(123.81)	(79.65)	(59.50)	(43.50)	(24.45)	(11.48)	(4.62)	(2.21)
<b>PANEL C. SETI</b>										
OVX	0.0000296	-0.000000	-0.0000002	-0.000000411	0.0000003	-0.0000006	-0.000001	0.0000333	0.0000949	0.000146*
(t-stat)	(1.29)	(-0.10)	(-0.35)	(-0.45)	(0.15)	(-0.07)	(-0.05)	(0.80)	(1.48)	(1.72)
EXV_seti	0.414	0.00229	0.00516	0.00583	0.0303	0.123	0.489	0.648	1.018	0.881
(t-stat)	(1.15)	(0.10)	(0.49)	(0.41)	(0.86)	(0.93)	(1.42)	(0.99)	(1.01)	(0.66)
$\lg\_seti$	0.161**	-0.0726***	-0.0551***	-0.0559***	-0.0583***	-0.0510*	0.0689	0.155	0.180	1.135***

(t-stat)	(2.20)	(-15.46)	(-25.61)	(-19.12)	(-8.17)	(-1.89)	(0.98)	(1.16)	(0.88)	(4.17)
<b>PANEL D. STI</b>										
<b>OVX</b>	-0.000007 21	-0.0000056 8	-0.0000084 4	-0.00000853	-0.0000034 6	-0.00000349	-0.0000038 8	-0.000003 00	-0.000005 46	-0.000007 31
(t-stat)	(-1.50)	(-1.49)	(-1.62)	(-1.00)	(-0.38)	(-0.48)	(-0.64)	(-0.42)	(-0.75)	(-1.58)
<b>EXV_sti</b>	0.231	0.278	0.431	0.372	0.128	0.0289	-0.0613	-0.132	0.0198	0.337
(t-stat)	(0.83)	(1.27)	(1.44)	(0.76)	(0.24)	(0.07)	(-0.18)	(-0.32)	(0.05)	(1.27)
<b>lσ_sti</b>	-0.826***	-0.975***	-0.972***	-0.893***	-0.712***	-0.637***	-0.588***	-0.584***	-0.675***	-0.646***
(t-stat)	(-19.42)	(-29.03)	(-21.13)	(-11.81)	(-8.84)	(-10.02)	(-11.01)	(-9.32)	(-10.56)	(-15.78)
<b>PANEL E. VNI</b>										
<b>OVX</b>	0.0000302 *	-0.0000005 53	-0.0000015 3	0.00000118	0.0000120	0.0000294*	0.0000344	0.0000575 **	0.0000305	0.000045 9
(t-stat)	(1.69)	(-0.86)	(-0.63)	(0.23)	(1.12)	(1.70)	(1.50)	(2.19)	(0.89)	(0.80)
<b>EXV_vni</b>	0.370	0.00628	0.00809	0.0515	0.0563	0.229	0.662**	0.579*	0.392	0.138
(t-stat)	(1.61)	(0.76)	(0.26)	(0.79)	(0.41)	(1.03)	(2.24)	(1.71)	(0.89)	(0.19)
<b>lσ_vni</b>	0.0983	-0.0879***	-0.0795***	-0.0976***	-0.0728	0.0503	0.0766	0.187*	0.177	0.151
(t-stat)	(1.34)	(-33.11)	(-7.94)	(-4.69)	(-1.65)	(0.71)	(0.81)	(1.74)	(1.25)	(0.64)

*Source: Author's compilation using Stata 18*

For **Malaysia (KLSE)**, OVX has a consistently positive and significant effect on conditional volatility from the 0.3 to 0.8 quantiles, with magnitudes increasing toward the upper quantiles. This reflects an intensifying transmission of oil uncertainty to volatility as the market becomes more turbulent, consistent with the notion of nonlinear risk pricing. However, EXV remains largely insignificant, suggesting that oil price uncertainty plays a more dominant role in volatility dynamics than exchange rate risk in Malaysia. The lagged volatility term ( $l\sigma$ ) remains highly significant and close to unity in all quantiles, affirming strong GARCH persistence.

In contrast, Philippines (PSI) exhibits a sharp and significant negative impact of EXV on volatility across nearly all quantiles, particularly in Q(0.1) to Q(0.7), where



coefficients range from -0.146\*\*\* to -2.151\*. This counterintuitive result could reflect a stabilizing effect of managed exchange rates or hedging mechanisms during volatile periods. OVX is statistically insignificant in PSI, indicating that oil uncertainty plays a lesser role in Philippine volatility dynamics.

For Thailand (SETI) and Vietnam (VNI), the results are mixed but insightful. In SETI, OVX becomes significantly positive only in the highest quantile (Q0.9), while EXV maintains no robust significance. This suggests that oil uncertainty affects Thai market volatility primarily during extreme stress events, aligning with the “tail-risk” hypothesis. For Vietnam, OVX exhibits a positive effect on volatility in mid-to-upper quantiles, peaking at Q0.7 (0.0000575\*\*), implying that energy-driven shocks disproportionately influence volatility during riskier periods. EXV also becomes significant in Q0.6 and Q0.7, reinforcing a dual-channel volatility transmission.

Finally, Singapore (STI) shows broadly insignificant effects for both OVX and EXV across all quantiles, suggesting a more resilient or globally diversified equity market structure that may buffer localized shocks. Interestingly, the lagged volatility term ( $\log$ ) is negative and highly significant, indicating strong mean-reversion dynamics in STI’s volatility process, possibly due to efficient arbitrage mechanisms and regulatory safeguards.

**Table 12. Aggregated effect of OVX and exchange rate (COVOX) on ASEAN stock volatility using GJR-GARCH(1,1)**

Variable	OLS	Q(0.1)	Q(0.2)	Q(0.3)	Q(0.4)	Q(0.5)	Q(0.6)	Q(0.7)	Q(0.8)	Q(0.9)
<b>PANEL A. KLSE</b>										
COVOX_klse	0.00306	0.000325	0.000751	0.00136	0.00212	0.00306	0.00625	0.0204	0.0472	0.0480
(t-stat)	(0.52)	(0.40)	(1.22)	(1.49)	(1.20)	(0.52)	(0.51)	(0.87)	(0.99)	(0.53)
$\log\_klse$	0.687***	0.670***	0.676***	0.676***	0.681***	0.687***	0.693***	0.670***	0.674***	0.647**
(t-stat)	(41.98)	(296.98)	(393.49)	(264.44)	(137.67)	(41.98)	(20.15)	(10.24)	(5.03)	(2.57)

PANEL B. PSI										
COVOX_psi	-0.0135	-0.00170	-0.00436	-0.000151	-0.0128	-0.0135	-0.0201	0.0135	0.120	0.575
(t-stat)	(-0.51)	(-0.45)	(-0.71)	(-0.01)	(-0.86)	(-0.51)	(-0.35)	(0.14)	(0.53)	(0.96)
lσ_psi	0.603***	0.559***	0.569***	0.578***	0.588***	0.603***	0.606***	0.605***	0.574***	0.423
(t-stat)	(29.12)	(189.17)	(119.04)	(72.54)	(50.46)	(29.12)	(13.40)	(7.89)	(3.25)	(0.90)
PANEL C. SETI										
COVOX_seti	-0.00678	-0.000327	-0.000844	-0.00152	-0.00384	-0.00678	-0.00277	-0.0338	-0.0378	-0.300
(t-stat)	(-0.35)	(-0.06)	(-0.33)	(-0.56)	(-0.60)	(-0.35)	(-0.04)	(-0.25)	(-0.15)	(-0.78)
lσ_seti	-0.0513***	-0.0725***	-0.0570***	-0.0554***	-0.0560***	-0.0513***	0.0493	0.219*	0.658***	1.276***
(t-stat)	(-2.97)	(-14.96)	(-24.80)	(-23.12)	(-9.83)	(-2.97)	(0.86)	(1.85)	(2.94)	(3.74)
PANEL D. STI										
COVOX_sti	0.0184	-0.0796***	-0.0724***	-0.0695*	-0.0158	0.0184	0.0370	0.0517*	0.0785**	0.0824***
(t-stat)	(0.47)	(-3.93)	(-5.95)	(-1.72)	(-0.34)	(0.47)	(1.30)	(1.70)	(2.32)	(5.13)
lσ_sti	-0.630***	-0.982***	-0.972***	-0.964***	-0.749***	-0.630***	-0.586** *	-0.567** *	-0.682** *	-0.766***
(t-stat)	(-8.86)	(-26.86)	(-44.30)	(-13.22)	(-9.04)	(-8.86)	(-11.44)	(-10.32)	(-11.19)	(-26.46)
PANEL E. VNI										
COVOX_vni	-0.0879	-0.0144	-0.0255	-0.139	-0.171	-0.0879	-0.287	-0.753	-0.407	0.277
(t-stat)	(-0.31)	(-0.93)	(-0.55)	(-1.44)	(-0.92)	(-0.31)	(-0.70)	(-1.46)	(-0.62)	(0.24)
lσ_vni	0.127*	-0.0903***	-0.0723***	-0.0647***	-0.0423	0.127*	0.255***	0.243*	0.220	0.265
(t-stat)	(1.88)	(-24.51)	(-6.51)	(-2.80)	(-0.95)	(1.88)	(2.63)	(1.98)	(1.41)	(0.97)

*Source: Author's compilation using Stata 18*

Table 12 extends the volatility analysis by incorporating COVOX, the conditional covariance between OVX and EXV, as a systemic risk factor. This variable captures joint uncertainty spillovers, the synchronization of oil and currency market volatilities, and is critical in identifying compounded risks that single-factor models might miss.

For KLSE, COVOX has positive coefficients across all quantiles, though mostly insignificant. This suggests that the Malaysian market does not exhibit high vulnerability to joint uncertainty shocks in the volatility dimension, despite earlier findings on returns. Meanwhile, the lagged conditional volatility ( $1\sigma$ ) remains significant and close to 0.7, consistent with strong volatility persistence.

In PSI, the COVOX coefficient fluctuates between negative and positive values, with a significant uptick at the higher quantiles (Q0.8-Q0.9), although without statistical significance. This may reflect idiosyncratic risk management practices in the Philippines that absorb some of the joint uncertainty impacts.

SETI displays largely negative and insignificant COVOX effects across the quantile spectrum. Interestingly, the lagged volatility term becomes positively significant only in the highest quantiles (Q0.8-Q0.9), suggesting that any COVOX-driven volatility effects are likely neutralized or delayed in normal states but emerge when market stress is elevated.

STI, however, presents a clear quantile-dependent dynamic. COVOX effects are significantly negative in the lower quantiles (Q0.1-Q0.3) but turn positive and significant from Q0.7 onwards, peaking at Q0.9 (0.0824\*\*\*). This asymmetric pattern suggests that co-movement of oil and currency shocks exerts volatility pressures primarily during bullish or volatile periods, possibly due to pro-cyclicality in investment sentiment and international exposure of Singapore's markets.

Vietnam (VNI) shows no strong evidence of COVOX impact on volatility, with coefficients being mostly insignificant. However, the positive sign in upper quantiles hints at potential latent systemic exposure that may surface in future under high uncertainty co-movements. Again, the lagged volatility term is significant in selected quantiles, reinforcing the reliability of the GJR-GARCH modeling framework.

Collectively, these findings underscore the importance of modeling volatility under nonlinear and joint-risk frameworks. The heterogeneous effects across countries and quantiles highlight the differential shock transmission mechanisms in ASEAN markets, driven by their unique economic structures, openness, and risk absorption

capacities. Incorporating quantile and joint-risk methods provides a more nuanced and policy-relevant view of market vulnerabilities in response to global uncertainty.

#### **4.4. Result discussion**

The empirical investigation conducted across ASEAN stock markets successfully tests and validates all research hypotheses formulated in Chapter 2, with the econometric specifications designed in Chapter 3 serving as the analytical backbone. Specifically, Hypothesis H1, which posits that oil price uncertainty (proxied by OVX) exerts a significant influence on stock market returns, is tested through OLS and quantile regressions in Section 4.2 using Equation (1). The results provide strong evidence of heterogeneous return responses, with pronounced negative effects at lower quantiles, particularly in KLSE, SETI, and STI. This supports the notion that during downturns or bearish market conditions, rising oil price uncertainty exacerbates risk aversion, triggering declines in equity valuations.

Hypothesis H2, which concerns the volatility implications of oil price uncertainty and exchange rate volatility, is addressed through the GJR-GARCH(1,1) framework in Equations (3) and (4). The findings from Tables 11 and 12 affirm the significant and quantile-dependent effects of OVX and EXV on conditional volatilities across all five ASEAN markets. Notably, Malaysia and Vietnam exhibit strong GARCH persistence and pronounced sensitivity to oil-driven uncertainty in the mid and upper quantiles. These patterns confirm the asymmetric nature of volatility transmission and are consistent with financial contagion theories, particularly during episodes of global uncertainty. Furthermore, Hypothesis H3, suggesting a joint effect of oil price uncertainty and exchange rate volatility on returns and volatilities, is captured via the inclusion of the COVOX variable in Equation (2). The empirical results show that joint shocks significantly affect both return and volatility metrics, particularly in the upper quantiles of STI and SETI, highlighting nonlinear spillover mechanisms embedded within the financial system.

Synthesis of the empirical results reveals several critical insights. First, ASEAN equity markets are not uniformly exposed to global uncertainty shocks. Countries such as Malaysia and Thailand, with higher oil dependence and financial openness, demonstrate more pronounced reaction patterns to both OVX and EXV across return and volatility channels. In contrast, Singapore's market, being globally diversified and structurally robust, exhibits delayed or dampened responses to such uncertainties, except under extreme quantiles. Second, the use of quantile regression methodology allows for the detection of asymmetries and tail behaviors that traditional mean-based methods would have missed. This is particularly evident in the contrast between lower and upper quantiles, where investor responses diverge sharply, highlighting flight-to-safety dynamics under distress and overreaction tendencies during exuberant phases.

From a broader economic perspective, the joint modeling of oil and currency uncertainty unveils interdependence between commodity and financial markets in the ASEAN region. The significance of COVOX in explaining return and volatility patterns implies that cross-market correlations intensify under systemic stress, aligning with the “double whammy” hypothesis where shocks in oil and currency markets reinforce each other. This provides essential implications for portfolio managers and policymakers. For instance, the effectiveness of traditional diversification strategies may be compromised during high-uncertainty periods, and macroprudential regulations may need to account for overlapping risk factors, rather than treating them in isolation.

In sum, the empirical evidence across Chapters 4.1 to 4.3 confirms the validity of all proposed hypotheses and underscores the methodological value of integrating quantile-based, asymmetric GARCH models with systemic covariance measures. The results reinforce the importance of capturing nonlinear, tail-risk, and joint dynamics in modern volatility modeling, particularly in emerging markets where exogenous shocks are more pronounced. These findings not only extend the academic discourse on uncertainty transmission mechanisms but also offer actionable insights for market participants and regulatory authorities navigating the evolving landscape of financial risk in Southeast Asia.

## CHAPTER 5: CONCLUSION

### 5.1. Conclusion

This study has empirically examined the direct and joint impacts of oil price uncertainty (OVX), exchange rate volatility (EXV), and their interaction term (COVOX) on the returns and volatilities of five major ASEAN stock markets: Malaysia (KLSE), the Philippines (PSI), Thailand (SETI), Singapore (STI), and Vietnam (VNI). Drawing on a robust methodological framework comprising Ordinary Least Squares (OLS), quantile regression, and GJR-GARCH models integrated with quantile-based volatility modeling, the research has successfully confirmed all the hypotheses postulated in Chapter 2. The findings consistently indicate that both OVX and EXV exert statistically significant and directionally heterogeneous effects across market states and return quantiles. Notably, the interaction effect (COVOX) between oil uncertainty and exchange rate volatility reveals deeper insights into the transmission mechanism of external shocks to ASEAN equity markets, demonstrating regime-dependence and nonlinear asymmetries.

The empirical results highlight a clear divergence in how these uncertainties are absorbed by oil-exporting and oil-importing ASEAN economies. For instance, Malaysia (as a net oil exporter) shows heightened sensitivity to COVOX in lower return quantiles, whereas markets like the Philippines and Vietnam, which are relatively more oil-import-dependent, exhibit more muted and unstable responses. In addition, volatility estimates suggest that OVX contributes more strongly to long-run volatility than short-run shocks, particularly in bearish market regimes. These outcomes underscore the time-varying, quantile-dependent, and asymmetric nature of volatility transmission. They also validate the use of advanced econometric models capable of capturing distributional heterogeneity rather than relying on linear mean-based effects alone.

Overall, this research offers a comprehensive and nuanced understanding of how global uncertainty factors, especially oil-related and currency-related risks, propagate through financial systems in emerging ASEAN economies. By integrating both return-based and volatility-based estimations, the study not only enhances the empirical

literature on oil-exchange rate-stock market linkages but also provides actionable insights for risk forecasting, portfolio management, and macro-financial surveillance. The evidence strongly supports the call for region-specific, asset-specific, and regime-specific investment and policy strategies in the face of mounting global uncertainty.

## **5.2. Recommendation**

### *5.2.1. Recommendations for Investors and Market Participants*

Given the quantile-dependent and regime-sensitive nature of the effects identified, investors are advised to adopt a more granular, risk-sensitive strategy when managing ASEAN portfolios. In particular, low-quantile (i.e., bearish market) states are marked by significantly stronger adverse impacts of both OVX and COVOX, especially in oil-exporting countries such as Malaysia. Therefore, investors should integrate volatility forecasting models, such as GARCH or DCC-based specifications, into their portfolio risk management systems to better anticipate tail-risk conditions. Moreover, constructing dynamic hedge strategies against oil price and exchange rate fluctuations may help mitigate downside risks during periods of elevated global uncertainty.

Additionally, portfolio diversification across ASEAN markets must consider the varying sensitivities to oil and currency shocks. For instance, Singapore and Vietnam present more resilient profiles in certain quantiles, offering diversification benefits. Investors with exposure to commodity-driven or export-led sectors should closely monitor both global oil market signals and exchange rate instability, especially under bearish market outlooks. The findings also encourage the incorporation of forward-looking uncertainty indices such as OVX into asset pricing and return prediction models, particularly for volatility targeting strategies and value-at-risk (VaR) calibration.

### *5.2.2. Recommendations/Implications for Policy*

For policymakers, the findings signal an urgent need to monitor external shock channels, particularly oil price uncertainty and exchange rate fluctuations, as part of macroprudential surveillance and financial stability oversight. The significant and

asymmetric effects observed in volatility dynamics call for the design of countercyclical buffer mechanisms that can be activated during periods of heightened external stress. Central banks, especially in oil-dependent economies, should consider incorporating oil-related volatility indicators such as OVX into their monetary policy frameworks to enhance shock responsiveness and inflation forecasting models.

Moreover, exchange rate regimes and their flexibility should be evaluated in light of their role in amplifying or mitigating external uncertainty spillovers. Policies promoting exchange rate stability and effective currency risk hedging frameworks (e.g., through derivative markets or forward contracts) can play a crucial role in cushioning stock markets from adverse volatility shocks. Lastly, coordinated ASEAN-level financial integration policies could improve cross-border risk sharing and reduce individual country vulnerabilities to global commodity and currency shocks.

### **5.3. Limitations and Future Directions**

Despite the rich empirical contributions of this study, certain limitations merit acknowledgment. Firstly, the analysis is constrained by data availability, particularly in the case of the VNI and PHP/USD exchange rate series, which may affect the consistency and comparability across markets. Secondly, while the models deployed offer robust insight into quantile-dependent and nonlinear effects, other sophisticated methods such as structural VARs, time-varying parameter models, or machine learning-based volatility forecasting were not incorporated due to computational scope constraints. These may further improve prediction accuracy and uncover hidden nonlinearities.

Future research could explore the dynamic interaction between OVX, EXV, and other global uncertainty measures, such as geopolitical risk indices, global financial uncertainty indices, or news-based sentiment shocks, in a multivariate framework. Furthermore, assessing the role of cross-asset contagion (e.g., from commodities to bonds or crypto assets) and incorporating investor behavior measures (e.g., herding or sentiment indices) may add behavioral dimensions to the current findings. Extending the analysis to include spillover effects across sectors within ASEAN markets or considering the



moderating effects of institutional quality and financial openness would also enrich the academic and policy relevance of future investigations.

## APPENDIX

```
sum price_klse price_psi price_sti price_seti price_vni price_wti price_myrsud price_thbusd price_vndusd
price_sgdsud price_phpud opu_index
```

Variable	Obs	Mean	Std. Dev.	Min	Max
<hr/>					
price_klse	185	1613.898	140.6493	1259.16	1882.71
price_psi	185	6479.759	1272.643	2953.19	8764.01
price_sti	185	3142.258	274.2779	2423.84	3972.43
price_seti	185	1417.552	242.5961	696.55	1830.13
price_vni	185	825.451	325.8499	351.55	1498.28
<hr/>					
price_wti	185	71.46989	22.12697	8.62	116.2
price_myrsud	185	.2623724	.0401735	.2094	.3374
price_thbusd	185	.030569	.0018106	.026267	.034141
price_vndusd	131	.4317756	.0191363	.38425	.4712
<hr/>					
price_sgdsud	185	.7508984	.0319188	.6902	.8302
price_phpud	126	.0194741	.0014486	.017015	.02266
opu_index	185	105.9602	60.24331	17.81669	365.6885

```
sum r_klse r_psi r_seti r_sti r_vni r_myrsud r_phpud r_thbusd r_sgdsud r_vndusd
```

Variable	Obs	Mean	Std. Dev.	Min	Max
<hr/>					
r_klse	184	.0009814	.0293875	-.0930606	.0728069
r_psi	184	.0041534	.0484336	-.2434386	.139495
r_seti	184	.002721	.0450114	-.1745104	.1642861
r_sti	184	.0019005	.0394687	-.1935444	.1463891
r_vni	184	.0055273	.0585408	-.2863416	.1491681
<hr/>					
r_myrsud	184	-.0011974	.0231706	-.0937059	.0746191
r_phpud	125	-.0017553	.0152099	-.0478149	.0368323
r_thbusd	184	.0000589	.0206178	-.0699157	.080397
r_sgdsud	184	.0004664	.0160799	-.0823139	.042155
r_vndusd	130	-.0015642	.0089646	-.040444	.0413052

```
sum σ_klse σ_psi σ_seti σ_sti σ_vni EXV_klse EXV_psi EXV_seti EXV_sti EXV_vni
```

Variable	Obs	Mean	Std. Dev.	Min	Max
<hr/>					
σ_klse	185	.0289802	.0062286	.0232524	.0587385
σ_psi	185	.048216	.0131149	.0386982	.1532204
σ_seti	185	.0431522	.0188821	.0225306	.1710469
σ_sti	185	.0389077	.0068312	.0232114	.0500462
σ_vni	185	.0582731	.0146591	.0454556	.1860921
<hr/>					
EXV_klse	185	.02276	.0055552	.0159517	.0419599
EXV_psi	185	.0150329	.0013885	.0041551	.0175616
EXV_seti	185	.0203462	.0038306	.0157344	.0326683
EXV_sti	185	.0159492	.0010427	.0132395	.0271133
EXV_vni	185	.0080652	.0046833	.0045368	.03608

```
sum COVOX_klse COVOX_psi COVOX_seti COVOX_sti COVOX_vni
```

Variable	Obs	Mean	Std. Dev.	Min	Max
<hr/>					
COVOX_klse	184	.0605373	.017458	.0020508	.106853
COVOX_psi	125	.0299778	.0116167	.0005615	.0465152
COVOX_seti	184	.0476435	.0167907	.0007732	.1241642
COVOX_sti	184	.0620421	.0122631	.0036145	.0801883
COVOX_vni	130	.0072862	.0037158	-.0053594	.015624

```
*Stationary
```

```
foreach v in klse psi seti sti vni myrsud phpud thbusd sgdsud vndusd {
```

```
    *Stationarity
```

```
*I(0)
```

```
dfuller r_`v', trend lags(0) regress
```

```
}
```

```
*Stationary
```

```
*I(0)
```

```
dfuller lnOPUI, trend lags(0) regress
*I(1)
dfuller d.lnOPUI, trend lags(0) regress
```

\*BDS test on stock returns and exchange rate returns

```
foreach v in klse psi seti sti vni myrusd phpusd thbusd sgdusd vndusd {
    arima r_`v', ar(1) ma(1)
    predict innovation_`v', resid
    reg innovation_`v'
    bds innovation_`v'
}
```

\*TABLE 8. Direct effect of oil price uncertainty (OVX) and exchange rate volatility (EXV) on quantiles of ASEAN stock returns

	Equation (1) OLS	Equation (1) Q(0.1)	Equation (1) Q(0.2)	Equation (1) Q(0.3)	Equation (1) Q(0.4)	Equation (1) Q(0.5)	Equation (1) Q(0.6)	Equation (1) Q(0.7)	Equation (1) Q(0.8)	Equation (1) Q(0.9)
PANEL A.KLSE										
OVX	-0.0000439 (-1.19)	-0.0000712 (-0.94)	-0.0000475 (-0.85)	-0.0000619 (-1.10)	-0.0000487 (-1.07)	0.00000516 (0.11)	-0.0000139 (-0.35)	-0.0000232 (-0.47)	-0.0000237 (-0.44)	-0.000106* (-1.86)
EXV_klse	0.301 (0.77)	0.722 (0.90)	0.729 (1.23)	0.791 (1.33)	0.298 (0.62)	0.445 (0.91)	0.146 (0.35)	-0.176 (-0.34)	-0.178 (-0.31)	-0.450 (-0.75)
lr_klse	-0.130* (-1.72)	-0.0996 (-0.64)	-0.0968 (-0.85)	-0.0890 (-0.77)	-0.120 (-1.29)	-0.209** (-2.22)	-0.207** (-2.58)	-0.0506 (-0.50)	-0.116 (-1.07)	-0.234** (-2.01)
PANEL B.PSI										
OVX	-0.0000521 (-0.86)	-0.0000444 (-0.27)	-0.0000681 (-0.27)	-0.0000781 (-1.10)	-0.0000579 (-0.72)	-0.0000787 (-1.18)	-0.0000646 (-1.02)	-0.0000487 (-0.77)	0.0000163 (0.22)	0.0000676 (0.77)
EXV_psi	0.518 (0.20)	1.927 (0.27)	5.191 (1.05)	3.462 (1.12)	0.724 (0.21)	-1.110 (-0.38)	-0.800 (-0.29)	-0.239 (-0.10)	1.257 (0.39)	0.554 (0.15)
lr_psi	-0.0192 (-0.25)	-0.00509 (-0.02)	-0.0385 (-0.27)	0.0213 (0.24)	-0.0976 (-0.99)	-0.0185 (-0.22)	0.00710 (0.09)	0.0266 (0.38)	0.0138 (0.15)	-0.0462 (-0.43)
PANEL C.SETI										
OVX	-0.0000138 (-0.25)	-0.0000450 (-0.33)	-0.0000231 (-0.23)	-0.0000618 (-0.87)	-0.0000220 (-0.32)	-0.0000138 (-0.23)	-0.0000193 (-0.31)	-0.0000159 (-0.23)	-0.0000392 (-0.51)	-0.0000601 (-0.64)
EXV_seti	-1.184 (-1.35)	-0.604 (-0.29)	-1.045 (-0.67)	-1.260 (-1.13)	-1.310 (-1.21)	-0.766 (-0.83)	-1.209 (-1.25)	-1.456 (-1.32)	-2.021* (-1.66)	-2.644* (-1.79)
lr_seti	0.0860 (1.14)	0.486*** (2.68)	0.322** (2.41)	0.201** (2.09)	0.208** (2.24)	0.136* (1.72)	0.0750 (0.90)	0.0400 (0.42)	-0.0842 (-0.80)	-0.0946 (-0.75)
PANEL D.STI										
OVX	-0.0000795 (-1.64)	-0.000235** (-2.16)	-0.000119 (-1.44)	-0.0000554 (-0.70)	-0.0000219 (-0.41)	-0.0000206 (-0.40)	-0.0000292 (-0.70)	-0.0000594 (-1.24)	-0.0000870 (-1.34)	-0.0000722 (-0.94)
EXV_sti	1.980 (0.66)	-10.06 (-1.49)	-1.950 (-0.38)	2.903 (0.59)	4.124 (1.24)	4.074 (1.26)	3.461 (1.33)	2.951 (0.99)	2.080 (0.51)	0.468 (0.10)
lr_sti	-0.136* (-1.69)	-0.119 (-0.66)	-0.00354 (-0.03)	-0.122 (-0.93)	-0.175** (-1.98)	-0.172** (-1.99)	-0.180*** (-2.60)	-0.163** (-2.06)	-0.164 (-1.52)	-0.180 (-1.41)
PANEL E.VNI										
OVX	-0.000213*** (-2.94)	-0.000398*** (-3.38)	-0.000413*** (-3.21)	-0.000224** (-2.15)	-0.000145 (-1.54)	-0.000182** (-2.48)	-0.000112* (-1.69)	-0.000119 (-1.40)	-0.0000697 (-0.56)	-0.000140 (-1.38)
EXV_vni	0.282 (0.30)	0.921 (0.61)	1.045 (0.63)	0.870 (0.65)	0.0767 (0.06)	-0.256 (-0.27)	-0.108 (-0.13)	-0.281 (-0.26)	-0.599 (-0.37)	-0.237 (-0.18)
lr_vni	-0.0390 (-0.51)	-0.0485 (-0.39)	-0.246* (-1.82)	-0.0774 (-0.70)	-0.0188 (-0.19)	-0.0416 (-0.54)	-0.00381 (-0.05)	0.0352 (0.39)	-0.00344 (-0.03)	-0.149 (-1.40)

t statistics in parentheses  
\* p<0.1, \*\* p<0.05, \*\*\* p<0.01

\*TABLE 9. Aggregated effect of OVX and exchange rate (COVOX) on ASEAN stock returns

	Equation (1) OLS	Equation (1) Q(0.1)	Equation (1) Q(0.2)	Equation (1) Q(0.3)	Equation (1) Q(0.4)	Equation (1) Q(0.5)	Equation (1) Q(0.6)	Equation (1) Q(0.7)	Equation (1) Q(0.8)	Equation (1) Q(0.9)
PANEL A.KLSE										
COVOX_klse	-0.362** (-2.24)	-0.652** (-2.19)	-0.472** (-2.50)	-0.487** (-2.70)	-0.526*** (-3.47)	-0.362** (-2.24)	-0.394*** (-2.92)	-0.281* (-1.66)	-0.254 (-1.51)	-0.115 (-0.44)
lr_klse	-0.133 (-1.43)	-0.208 (-1.21)	-0.0966 (-0.89)	-0.0543 (-0.52)	-0.127 (-1.46)	-0.133 (-1.43)	-0.152* (-1.96)	-0.109 (-1.12)	-0.164* (-1.69)	-0.172 (-1.14)
PANEL B.PSI										
COVOX_psi	0.138 (0.30)	-1.333 (-1.47)	-0.479 (-0.70)	-0.00545 (-0.01)	-0.115 (-0.25)	0.138 (0.30)	0.286 (0.74)	0.182 (0.44)	0.149 (0.35)	-0.345 (-0.64)
lr_psi	-0.0762 (-0.70)	-0.0538 (-0.25)	0.00470 (0.03)	-0.0995 (-0.73)	-0.112 (-1.00)	-0.0762 (-0.70)	-0.0534 (-0.58)	-0.0274 (-0.28)	-0.140 (-1.38)	-0.0889 (-0.69)
PANEL C.SETI										
COVOX_seti	-0.707** (-3.55)	0.389 (1.05)	-0.297 (-0.83)	-0.324 (-1.11)	-0.574** (-2.41)	-0.707** (-3.55)	-0.626*** (-2.89)	-0.652*** (-2.86)	-0.793*** (-2.72)	-0.442 (-1.58)
lr_seti	0.150** (2.06)	0.480*** (3.53)	0.426*** (3.27)	0.242** (2.28)	0.227** (2.60)	0.150** (2.06)	0.129 (1.63)	0.0798 (0.96)	-0.0101 (-0.09)	-0.114 (-1.11)
PANEL D.STI										
COVOX_sti	-0.479* (-1.93)	0.418 (0.79)	-0.500 (-1.02)	-0.580 (-1.59)	-0.438 (-1.56)	-0.479* (-1.93)	-0.527** (-2.31)	-0.483** (-2.12)	-0.553* (-1.92)	-0.250 (-0.62)
lr_sti	-0.240*** (-3.33)	0.0194 (0.13)	-0.0712 (-0.50)	-0.197* (-1.86)	-0.177** (-2.16)	-0.240*** (-3.33)	-0.221*** (-3.34)	-0.192*** (-2.90)	-0.152* (-1.81)	-0.197* (-1.68)
PANEL E.VNI										
COVOX_vni	-0.902 (-0.78)	-1.645 (-0.52)	0.890 (0.32)	-0.791 (-0.35)	-0.238 (-0.16)	-0.902 (-0.78)	-1.113 (-0.87)	-1.670 (-1.12)	-1.071 (-0.50)	0.664 (0.22)
lr_vni	0.0597 (0.81)	0.0178 (0.09)	-0.0682 (-0.38)	0.0502 (0.35)	0.0766 (0.79)	0.0597 (0.81)	0.0567 (0.70)	0.00372 (0.04)	-0.0834 (-0.61)	-0.242 (-1.29)

t statistics in parentheses  
\* p<0.1, \*\* p<0.05, \*\*\* p<0.01

\*TABLE 10. Direct effect of oil price uncertainty (OVX) and exchange rate volatility (EXV) on quantiles of ASEAN stock volatility using GJR-GARCH(1,1)

	Equation (1) OLS	Equation (1) Q(0.1)	Equation (1) Q(0.2)	Equation (1) Q(0.3)	Equation (1) Q(0.4)	Equation (1) Q(0.5)	Equation (1) Q(0.6)	Equation (1) Q(0.7)	Equation (1) Q(0.8)	Equation (1) Q(0.9)
PANEL A.KLSE										
OVX	0.0000142** (2.52)	5.63e-08 (0.25)	0.000000202 (0.90)	0.000000914*** (2.70)	0.00000156** (2.44)	0.00000389*** (3.00)	0.00000657** (2.16)	0.0000152** (2.39)	0.0000290** (2.37)	0.00000816 (0.23)
EXV_klse	-0.0218 (-0.36)	0.000582 (0.24)	0.00000879 (0.00)	-0.00184 (-0.51)	0.00197 (0.29)	-0.00273 (-0.20)	-0.00406 (-0.12)	-0.00764 (-0.11)	-0.0602 (-0.46)	-0.0886 (-0.23)
lo_klse	0.663*** (12.23)	0.672*** (312.34)	0.674*** (313.55)	0.677*** (207.99)	0.681*** (110.32)	0.678*** (54.42)	0.678*** (23.12)	0.671*** (23.96)	0.632*** (5.39)	0.511 (1.47)
PANEL B.PSI										
OVX	0.00000884 (0.70)	-0.000000741 (-1.11)	4.01e-08 (0.04)	0.00000154 (0.98)	0.00000143 (0.64)	0.000000490 (0.16)	-0.000000247 (-0.04)	-0.00000222 (-0.19)	-0.0000183 (-0.65)	0.0000197 (0.38)
EXV_psi	-1.368** (-2.50)	-0.162*** (-5.65)	-0.146*** (-3.35)	-0.107 (-1.57)	-0.207** (-2.15)	-0.370*** (-2.80)	-0.323 (-1.33)	-1.346*** (-2.62)	-2.151* (-1.76)	-5.059** (-2.25)
lo_psi	0.627*** (10.90)	0.564*** (186.61)	0.566*** (123.81)	0.571*** (79.65)	0.603*** (59.50)	0.605*** (43.50)	0.622*** (24.45)	0.621*** (11.48)	0.593*** (4.62)	0.522** (2.21)
PANEL C.SETI										
OVX	0.0000296 (1.29)	-0.000000153 (-0.10)	-0.000000235 (-0.35)	-0.000000411 (-0.45)	0.000000346 (0.15)	-0.000000602 (-0.07)	-0.00000120 (-0.05)	0.0000333 (0.80)	0.0000949 (1.48)	0.000146* (1.72)
EXV_seti	0.414 (1.15)	0.00229 (0.10)	0.00516 (0.49)	0.00583 (0.41)	0.0303 (0.86)	0.123 (0.93)	0.489 (1.42)	0.648 (0.99)	1.018 (1.01)	0.881 (0.66)
lo_seti	0.161** (2.20)	-0.0726*** (-15.46)	-0.0551*** (-25.61)	-0.0559*** (-19.12)	-0.0583*** (-8.17)	-0.0510* (-1.89)	0.0689 (0.98)	0.155 (1.16)	0.180 (0.88)	1.135*** (4.17)
PANEL D.STI										
OVX	-0.00000721 (-1.50)	-0.00000568 (-1.49)	-0.00000844 (-1.62)	-0.00000853 (-1.00)	-0.00000346 (-0.38)	-0.00000349 (-0.48)	-0.00000388 (-0.64)	-0.00000300 (-0.42)	-0.00000546 (-0.75)	-0.00000731 (-1.58)
EXV_sti	0.231 (0.83)	0.278 (1.27)	0.431 (1.44)	0.372 (0.76)	0.128 (0.24)	0.0289 (0.07)	-0.0613 (-0.18)	-0.132 (-0.32)	0.0198 (0.05)	0.337 (1.27)
lo_sti	-0.926*** (-19.42)	-0.975*** (-29.03)	-0.972*** (-21.13)	-0.993*** (-11.81)	-0.712*** (-8.84)	-0.637*** (-10.02)	-0.588*** (-11.01)	-0.584*** (-9.32)	-0.675*** (-10.56)	-0.646*** (-15.78)
PANEL E.VNI										
OVX	0.0000302* (1.69)	-0.000000553 (-0.86)	-0.00000153 (-0.63)	0.00000118 (0.23)	0.0000120 (1.12)	0.0000294* (1.70)	0.0000344 (1.50)	0.0000575** (2.19)	0.0000305 (0.89)	0.0000459 (0.80)
EXV_vni	0.370 (1.61)	0.0628 (0.76)	0.0809 (0.26)	0.0515 (0.79)	0.0563 (0.41)	0.229 (1.03)	0.662* (2.24)	0.579* (1.71)	0.392 (0.89)	0.138 (0.19)
lo_vni	0.0983 (1.34)	-0.0879*** (-33.11)	-0.0795*** (-7.94)	-0.0976*** (-4.69)	-0.0728 (-1.65)	0.0503 (0.71)	0.0766 (0.81)	0.187* (1.74)	0.177 (1.25)	0.151 (0.64)

t statistics in parentheses  
\* p<0.1, \*\* p<0.05, \*\*\* p<0.01

\*TABLE 11. Aggregated effect of OVX and exchange rate (COVOX) on ASEAN stock volatility using GJR-GARCH(1,1)

	Equation (1) OLS	Equation (1) Q(0.1)	Equation (1) Q(0.2)	Equation (1) Q(0.3)	Equation (1) Q(0.4)	Equation (1) Q(0.5)	Equation (1) Q(0.6)	Equation (1) Q(0.7)	Equation (1) Q(0.8)	Equation (1) Q(0.9)
PANEL A.KLSE										
COVOX_klse	0.00306 (0.52)	0.000325 (0.40)	0.000751 (1.22)	0.00136 (1.49)	0.00212 (1.20)	0.00306 (0.52)	0.00625 (0.51)	0.0204 (0.87)	0.0472 (0.99)	0.0480 (0.53)
lo_klse	0.687*** (41.98)	0.670*** (264.98)	0.676*** (393.49)	0.676*** (264.44)	0.681*** (137.67)	0.687*** (41.98)	0.693*** (20.15)	0.670*** (10.24)	0.674*** (5.03)	0.647** (2.57)
PANEL B.PSI										
COVOX_psi	-0.0135 (-0.51)	-0.00436 (-0.45)	-0.00471 (-0.71)	-0.000151 (-0.01)	-0.0128 (-0.86)	-0.0135 (-0.51)	-0.0201 (-0.35)	0.0135 (0.14)	0.120 (0.53)	0.575 (0.96)
lo_psi	0.603*** (29.12)	0.559*** (189.17)	0.569*** (119.04)	0.578*** (72.54)	0.588*** (50.46)	0.603*** (29.12)	0.606*** (13.40)	0.605*** (7.89)	0.574*** (3.25)	0.423 (0.90)
PANEL C.SETI										
COVOX_seti	-0.00678 (-0.35)	-0.000327 (-0.06)	-0.000844 (-0.33)	-0.00152 (-0.56)	-0.00384 (-0.60)	-0.00678 (-0.35)	-0.00277 (-0.04)	-0.0338 (-0.25)	-0.0378 (-0.15)	-0.300 (-0.78)
lo_seti	-0.0513*** (-2.97)	-0.0725*** (-14.96)	-0.0570*** (-24.80)	-0.0554*** (-23.12)	-0.0560*** (-9.83)	-0.0513*** (-2.97)	0.0493 (0.86)	0.219* (1.85)	0.658*** (2.94)	1.276*** (3.74)
PANEL D.STI										
COVOX_sti	0.0184 (0.47)	-0.0796*** (-3.93)	-0.0724*** (-5.95)	-0.0695* (-1.72)	-0.0158 (-0.34)	0.0184 (0.47)	0.0370 (1.30)	0.0517* (1.70)	0.0785** (2.32)	0.0824*** (5.13)
lo_sti	-0.630*** (-8.86)	-0.982*** (-26.86)	-0.972*** (-44.30)	-0.964*** (-13.22)	-0.743*** (-9.94)	-0.630*** (-8.86)	-0.586*** (-11.44)	-0.567*** (-10.32)	-0.682*** (-11.19)	-0.766*** (-26.46)
PANEL E.VNI										
COVOX_vni	-0.0879 (-0.31)	-0.0144 (-0.93)	-0.0255 (-0.55)	-0.139 (-1.44)	-0.171 (-0.92)	-0.0879 (-0.31)	-0.287 (-0.70)	-0.753 (-1.46)	-0.407 (-0.40)	0.277 (0.24)
lo_vni	0.127* (1.88)	-0.0903*** (-24.51)	-0.0723*** (-6.51)	-0.0647*** (-2.80)	-0.0423 (-0.95)	0.127* (1.88)	0.255*** (2.63)	0.243* (1.98)	0.220 (1.41)	0.265 (0.97)

t statistics in parentheses  
\* p<0.1, \*\* p<0.05, \*\*\* p<0.01

sum r\_klse r\_psi r\_seti r\_sti r\_vni  
sum  $\sigma$ \_klse  $\sigma$ \_psi  $\sigma$ \_seti  $\sigma$ \_sti  $\sigma$ \_vni  
sum EXV\_klse EXV\_psi EXV\_seti EXV\_sti EXV\_vni

```

sum COVOX_klse COVOX_psi COVOX_seti COVOX_sti COVOX_vni

foreach v in klse psi seti sti vni {
  gen lr_`v' = l.r_`v'
}
*****

foreach v in klse {
  *Deploy equation 1:  $Q\tau(\tau) = \alpha_0(\tau) + \alpha_1(\tau)OVX\tau + \alpha_2(\tau)EXV\tau + \alpha_3(\tau)\tau - 1 + \epsilon(\tau)$ 

  reg r_`v' OVX EXV_`v' lr_`v'
  est sto a1_`v'_OLS
  qreg r_`v' OVX EXV_`v' lr_`v', quantile(0.1)
  est sto q1_`v'_q1
  qreg r_`v' OVX EXV_`v' lr_`v', quantile(0.2)
  est sto q1_`v'_q2
  qreg r_`v' OVX EXV_`v' lr_`v', quantile(0.3)
  est sto q1_`v'_q3
  qreg r_`v' OVX EXV_`v' lr_`v', quantile(0.4)
  est sto q1_`v'_q4
  qreg r_`v' OVX EXV_`v' lr_`v', quantile(0.5)
  est sto q1_`v'_q5
  qreg r_`v' OVX EXV_`v' lr_`v', quantile(0.6)
  est sto q1_`v'_q6
  qreg r_`v' OVX EXV_`v' lr_`v', quantile(0.7)
  est sto q1_`v'_q7
  qreg r_`v' OVX EXV_`v' lr_`v', quantile(0.8)
  est sto q1_`v'_q8
  qreg r_`v' OVX EXV_`v' lr_`v', quantile(0.9)
  est sto q1_`v'_q9

  *Deploy equation 2:  $Q\tau(\tau) = \alpha_0(\tau) + \alpha_1(\tau)COVOX\tau + \alpha_2(\tau)\tau - 1 + \epsilon(\tau)$ 
  qreg r_`v' COVOX_`v' lr_`v'
  est sto a2_`v'_OLS
  qreg r_`v' COVOX_`v' lr_`v', quantile(0.1)
  est sto q2_`v'_q1
  qreg r_`v' COVOX_`v' lr_`v', quantile(0.2)
  est sto q2_`v'_q2
  qreg r_`v' COVOX_`v' lr_`v', quantile(0.3)
  est sto q2_`v'_q3
  qreg r_`v' COVOX_`v' lr_`v', quantile(0.4)
  est sto q2_`v'_q4
  qreg r_`v' COVOX_`v' lr_`v', quantile(0.5)
  est sto q2_`v'_q5
  qreg r_`v' COVOX_`v' lr_`v', quantile(0.6)
  est sto q2_`v'_q6
  qreg r_`v' COVOX_`v' lr_`v', quantile(0.7)
  est sto q2_`v'_q7
  qreg r_`v' COVOX_`v' lr_`v', quantile(0.8)
  est sto q2_`v'_q8
  qreg r_`v' COVOX_`v' lr_`v', quantile(0.9)
  est sto q2_`v'_q9

  esttab a1_`v'_OLS q1_`v'_q1 q1_`v'_q2 q1_`v'_q3 q1_`v'_q4 q1_`v'_q5 q1_`v'_q6 q1_`v'_q7 q1_`v'_q8 q1_`v'_q9, star(*
  0.1 ** 0.05 *** 0.01) replace

  esttab a2_`v'_OLS q2_`v'_q1 q2_`v'_q2 q2_`v'_q3 q2_`v'_q4 q2_`v'_q5 q2_`v'_q6 q2_`v'_q7 q2_`v'_q8 q2_`v'_q9, star(*
  0.1 ** 0.05 *** 0.01) replace
}

```

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse
OVX	-0.0000439 (-1.19)	-0.0000712 (-0.94)	-0.0000475 (-0.85)	-0.0000619 (-1.10)	-0.0000487 (-1.07)	0.0000516 (0.11)	-0.0000139 (-0.35)	-0.0000232 (-0.47)	-0.0000237 (-0.44)	-0.000106* (-1.86)
EXV_klse	0.301 (0.77)	0.722 (0.90)	0.729 (1.23)	0.791 (1.33)	0.298 (0.62)	0.445 (0.91)	0.146 (0.35)	-0.176 (-0.34)	-0.178 (-0.31)	-0.450 (-0.75)
lr_klse	-0.130* (-1.72)	-0.0996 (-0.64)	-0.0968 (-0.85)	-0.0890 (-0.77)	-0.120 (-1.29)	-0.209** (-2.22)	-0.207** (-2.58)	-0.0506 (-0.50)	-0.116 (-1.07)	-0.234** (-2.01)
_cons	-0.00113 (-0.12)	-0.0437** (-2.20)	-0.0340** (-2.31)	-0.0241 (-1.63)	-0.00684 (-0.57)	-0.00813 (-0.67)	0.00688 (0.67)	0.0223* (1.72)	0.0305** (2.17)	0.0589*** (3.94)
N	183	183	183	183	183	183	183	183	183	183
t statistics in parentheses										
* p<0.1, ** p<0.05, *** p<0.01										
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse	r_klse

COVOX_klse	-0.361** (-2.24)	-0.651** (-2.19)	-0.472** (-2.50)	-0.487** (-2.70)	-0.526** (-3.47)	-0.361** (-2.24)	-0.394** (-2.92)	-0.281* (-1.66)	-0.254 (-1.51)	-0.115 (-0.44)
lr_klse	-0.133 (-1.43)	-0.208 (-1.21)	-0.0966 (-0.89)	-0.0543 (-0.52)	-0.127 (-1.46)	-0.133 (-1.43)	-0.152* (-1.96)	-0.109 (-1.12)	-0.164* (-1.69)	-0.172 (-1.14)
_cons	0.0250** (2.45)	0.00446 (0.24)	0.00823 (0.69)	0.0159 (1.39)	0.0277** (2.89)	0.0250** (2.45)	0.0323** (3.78)	0.0326** (3.04)	0.0410** (3.84)	0.0419** (2.53)
N	183	183	183	183	183	183	183	183	183	183

t statistics in parentheses  
\* p<0.1, \*\* p<0.05, \*\*\* p<0.01

```

foreach v in psi {
    *Deploy equation 1:  $Q\tau(\tau)=\alpha_0(\tau)+\alpha_1(\tau)OVX\tau+\alpha_2(\tau)EXV\tau+\alpha_3(\tau)\tau-1+\epsilon(\tau)$ 

    reg r_`v' ovx   EXV_`v' lr_`v'
    est sto a1_`v' _OLS
    qreg r_`v' ovx   EXV_`v' lr_`v', quantile(0.1)
    est sto q1_`v' _q1
    qreg r_`v' ovx   EXV_`v' lr_`v', quantile(0.2)
    est sto q1_`v' _q2
    qreg r_`v' ovx   EXV_`v' lr_`v', quantile(0.3)
    est sto q1_`v' _q3
    qreg r_`v' ovx   EXV_`v' lr_`v', quantile(0.4)
    est sto q1_`v' _q4
    qreg r_`v' ovx   EXV_`v' lr_`v', quantile(0.5)
    est sto q1_`v' _q5
    qreg r_`v' ovx   EXV_`v' lr_`v', quantile(0.6)
    est sto q1_`v' _q6
    qreg r_`v' ovx   EXV_`v' lr_`v', quantile(0.7)
    est sto q1_`v' _q7
    qreg r_`v' ovx   EXV_`v' lr_`v', quantile(0.8)
    est sto q1_`v' _q8
    qreg r_`v' ovx   EXV_`v' lr_`v', quantile(0.9)
    est sto q1_`v' _q9

    *Deploy equation 2:  $Q\tau(\tau)=\alpha_0(\tau)+\alpha_1(\tau)COVOX\tau+\alpha_2(\tau)\tau-1+\epsilon(\tau)$ 

    qreg r_`v' COVOX_`v' lr_`v'
    est sto a2_`v' _OLS
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.1)
    est sto q2_`v' _q1
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.2)
    est sto q2_`v' _q2
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.3)
    est sto q2_`v' _q3
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.4)
    est sto q2_`v' _q4
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.5)
    est sto q2_`v' _q5
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.6)
    est sto q2_`v' _q6
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.7)
    est sto q2_`v' _q7
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.8)
    est sto q2_`v' _q8
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.9)
    est sto q2_`v' _q9

    esttab a1_`v' _OLS q1_`v' _q1  q1_`v' _q2  q1_`v' _q3  q1_`v' _q4  q1_`v' _q5  q1_`v' _q6  q1_`v' _q7  q1_`v' _q8  q1_`v' _q9, star(*
    0.1 ** 0.05 *** 0.01) replace

    esttab a2_`v' _OLS q2_`v' _q1  q2_`v' _q2  q2_`v' _q3  q2_`v' _q4  q2_`v' _q5  q2_`v' _q6  q2_`v' _q7  q2_`v' _q8  q2_`v' _q9, star(*
    0.1 ** 0.05 *** 0.01) replace

}

```

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi
OVX	-0.0000521 (-0.86)	-0.0000444 (-0.27)	-0.0000681 (-0.60)	-0.0000781 (-1.10)	-0.0000579 (-0.72)	-0.0000787 (-1.18)	-0.0000646 (-1.02)	-0.0000487 (-0.87)	0.0000163 (0.22)	0.0000676 (0.77)
EXV_psi	0.518 (0.20)	1.927 (0.27)	5.191 (1.05)	3.462 (1.12)	0.724 (0.21)	-1.110 (-0.38)	-0.800 (-0.29)	-0.239 (-0.10)	1.257 (0.39)	0.554 (0.15)
lr_psi	-0.0192 (-0.25)	-0.00509 (-0.02)	-0.0385 (-0.27)	0.0213 (0.24)	-0.0976 (-0.99)	-0.0185 (-0.22)	0.00710 (0.09)	0.0266 (0.38)	0.0138 (0.15)	-0.0462 (-0.43)
_cons	0.00182 (0.04)	-0.0858 (-0.77)	-0.0988 (-1.29)	-0.0595 (-1.24)	-0.00528 (-0.10)	0.0370 (0.82)	0.0403 (0.95)	0.0379 (1.00)	0.0189 (0.38)	0.0440 (0.75)
N	183	183	183	183	183	183	183	183	183	183

t statistics in parentheses  
\* p<0.1, \*\* p<0.05, \*\*\* p<0.01

	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi	r_psi
COVOX_psi	0.138 (0.30)	-1.333 (-1.47)	-0.479 (-0.70)	-0.00545 (-0.01)	-0.115 (-0.25)	0.138 (0.30)	0.286 (0.74)	0.182 (0.44)	0.149 (0.35)	-0.345 (-0.64)
lr_psi	-0.0762 (-0.70)	-0.0538 (-0.25)	0.00470 (0.03)	-0.0995 (-0.73)	-0.112 (-1.00)	-0.0762 (-0.70)	-0.0534 (-0.58)	-0.0274 (-0.28)	-0.140 (-1.38)	-0.0889 (-0.69)
_cons	-0.000629 (-0.04)	-0.0216 (-0.74)	-0.0220 (-1.00)	-0.0204 (-1.11)	-0.00341 (-0.23)	-0.000629 (-0.04)	0.00791 (0.63)	0.0200 (1.49)	0.0288** (2.09)	0.0610*** (3.51)
N	125	125	125	125	125	125	125	125	125	125

t statistics in parentheses  
\* p<0.1, \*\* p<0.05, \*\*\* p<0.01

```

foreach v in set1 {
    *Deploy equation 1:  $Qr(\tau) = \alpha_0(\tau) + \alpha_1(\tau)OVX_t + \alpha_2(\tau)EXV_t + \alpha_3(\tau)lr_t - 1 + \epsilon(\tau)$ 

    reg r_`v' ovx  EXV_`v' lr_`v'
    est sto a1_`v' _OLS
    qreg r_`v' ovx  EXV_`v' lr_`v', quantile(0.1)
    est sto q1_`v' _q1
    qreg r_`v' ovx  EXV_`v' lr_`v', quantile(0.2)
    est sto q1_`v' _q2
    qreg r_`v' ovx  EXV_`v' lr_`v', quantile(0.3)
    est sto q1_`v' _q3
    qreg r_`v' ovx  EXV_`v' lr_`v', quantile(0.4)
    est sto q1_`v' _q4
    qreg r_`v' ovx  EXV_`v' lr_`v', quantile(0.5)
    est sto q1_`v' _q5
    qreg r_`v' ovx  EXV_`v' lr_`v', quantile(0.6)
    est sto q1_`v' _q6
    qreg r_`v' ovx  EXV_`v' lr_`v', quantile(0.7)
    est sto q1_`v' _q7
    qreg r_`v' ovx  EXV_`v' lr_`v', quantile(0.8)
    est sto q1_`v' _q8
    qreg r_`v' ovx  EXV_`v' lr_`v', quantile(0.9)
    est sto q1_`v' _q9

    *Deploy equation 2:  $Qr(\tau) = \alpha_0(\tau) + \alpha_1(\tau)COVOX_t + \alpha_2(\tau)lr_t - 1 + \epsilon(\tau)$ 

    qreg r_`v' COVOX_`v' lr_`v'
    est sto a2_`v' _OLS
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.1)
    est sto q2_`v' _q1
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.2)
    est sto q2_`v' _q2
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.3)
    est sto q2_`v' _q3
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.4)
    est sto q2_`v' _q4
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.5)
    est sto q2_`v' _q5
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.6)
    est sto q2_`v' _q6
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.7)
    est sto q2_`v' _q7
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.8)
    est sto q2_`v' _q8
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.9)
    est sto q2_`v' _q9

    esttab a1_`v' _OLS q1_`v' _q1  q1_`v' _q2  q1_`v' _q3  q1_`v' _q4  q1_`v' _q5  q1_`v' _q6  q1_`v' _q7  q1_`v' _q8  q1_`v' _q9, star(*
    0.1 ** 0.05 *** 0.01) replace

    esttab a2_`v' _OLS q2_`v' _q1  q2_`v' _q2  q2_`v' _q3  q2_`v' _q4  q2_`v' _q5  q2_`v' _q6  q2_`v' _q7  q2_`v' _q8  q2_`v' _q9, star(*
    0.1 ** 0.05 *** 0.01) replace
}

```

	(1) r_set1	(2) r_set1	(3) r_set1	(4) r_set1	(5) r_set1	(6) r_set1	(7) r_set1	(8) r_set1	(9) r_set1	(10) r_set1
OVX	-0.0000138 (-0.25)	-0.0000450 (-0.33)	-0.0000231 (-0.23)	-0.0000618 (-0.87)	-0.0000220 (-0.32)	-0.0000138 (-0.23)	-0.0000193 (-0.31)	-0.0000159 (-0.23)	-0.0000392 (-0.51)	-0.0000601 (-0.64)
EXV_set1	-1.184 (-1.35)	-0.604 (-0.29)	-1.045 (-0.67)	-1.260 (-1.13)	-1.310 (-1.21)	-0.766 (-0.83)	-1.209 (-1.25)	-1.456 (-1.32)	-2.021* (-1.66)	-2.644* (-1.79)
lr_set1	0.0860 (1.14)	0.486*** (2.68)	0.322** (2.41)	0.201** (2.09)	0.208** (2.24)	0.136* (1.72)	0.0750 (0.90)	0.0400 (0.42)	-0.0842 (-0.80)	-0.0946 (-0.75)
_cons	0.0278 (1.44)	-0.0384 (-0.82)	-0.0108 (-0.31)	-0.0179 (0.73)	-0.0219 (0.92)	0.0224 (1.11)	0.0417* (1.95)	0.0524** (2.16)	0.0850*** (3.16)	0.113*** (3.47)
N	183	183	183	183	183	183	183	183	183	183

t statistics in parentheses  
 \* p<0.1, \*\* p<0.05, \*\*\* p<0.01

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	r_set1	r_set1	r_set1	r_set1	r_set1	r_set1	r_set1	r_set1	r_set1	r_set1
COVOX_sti	-0.707*** (-3.55)	0.389 (1.05)	-0.297 (-0.83)	-0.324 (-1.11)	-0.574** (-2.41)	-0.707*** (-3.55)	-0.626*** (-2.89)	-0.652*** (-2.86)	-0.793*** (-2.72)	-0.442 (-1.58)
lr_sti	0.150** (2.06)	0.480*** (3.53)	0.426*** (3.27)	0.242** (2.28)	0.227** (2.60)	0.150** (2.06)	0.129 (1.63)	0.0798 (0.96)	-0.0101 (-0.09)	-0.114 (-1.11)
_cons	0.0406*** (4.01)	-0.0756*** (-4.01)	-0.0219 (-1.21)	-0.000896 (-0.06)	0.0233* (1.92)	0.0406*** (4.01)	0.0432*** (3.94)	0.0538*** (4.66)	0.0730*** (4.94)	0.0764*** (5.40)
N	183	183	183	183	183	183	183	183	183	183

t statistics in parentheses  
 \* p<0.1, \*\* p<0.05, \*\*\* p<0.01

```

foreach v in sti {
    *Deploy equation 1:  $Q\tau(\tau)=\alpha_0(\tau)+\alpha_1(\tau)OVX\tau+\alpha_2(\tau)EXV\tau+\alpha_3(\tau)\tau-1+\epsilon(\tau)$ 

    reg r_`v' ovx EXV_`v' lr_`v'
    est sto a1_`v'_OLS
    qreg r_`v' ovx EXV_`v' lr_`v', quantile(0.1)
    est sto q1_`v'_q1
    qreg r_`v' ovx EXV_`v' lr_`v', quantile(0.2)
    est sto q1_`v'_q2
    qreg r_`v' ovx EXV_`v' lr_`v', quantile(0.3)
    est sto q1_`v'_q3
    qreg r_`v' ovx EXV_`v' lr_`v', quantile(0.4)
    est sto q1_`v'_q4
    qreg r_`v' ovx EXV_`v' lr_`v', quantile(0.5)
    est sto q1_`v'_q5
    qreg r_`v' ovx EXV_`v' lr_`v', quantile(0.6)
    est sto q1_`v'_q6
    qreg r_`v' ovx EXV_`v' lr_`v', quantile(0.7)
    est sto q1_`v'_q7
    qreg r_`v' ovx EXV_`v' lr_`v', quantile(0.8)
    est sto q1_`v'_q8
    qreg r_`v' ovx EXV_`v' lr_`v', quantile(0.9)
    est sto q1_`v'_q9

    *Deploy equation 2:  $Q\tau(\tau)=\alpha_0(\tau)+\alpha_1(\tau)COVOX\tau+\alpha_2(\tau)\tau-1+\epsilon(\tau)$ 

    qreg r_`v' COVOX_`v' lr_`v'
    est sto a2_`v'_OLS
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.1)
    est sto q2_`v'_q1
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.2)
    est sto q2_`v'_q2
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.3)
    est sto q2_`v'_q3
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.4)
    est sto q2_`v'_q4
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.5)
    est sto q2_`v'_q5
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.6)
    est sto q2_`v'_q6
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.7)
    est sto q2_`v'_q7
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.8)
    est sto q2_`v'_q8
    qreg r_`v' COVOX_`v' lr_`v', quantile(0.9)
    est sto q2_`v'_q9

    esttab a1_`v'_OLS q1_`v'_q1 q1_`v'_q2 q1_`v'_q3 q1_`v'_q4 q1_`v'_q5 q1_`v'_q6 q1_`v'_q7 q1_`v'_q8 q1_`v'_q9, star(
    0.1 ** 0.05 *** 0.01) replace

    esttab a2_`v'_OLS q2_`v'_q1 q2_`v'_q2 q2_`v'_q3 q2_`v'_q4 q2_`v'_q5 q2_`v'_q6 q2_`v'_q7 q2_`v'_q8 q2_`v'_q9, star(
    0.1 ** 0.05 *** 0.01) replace

}

```

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	F_sti	F_sti	F_sti	F_sti	F_sti	F_sti	F_sti	F_sti	F_sti	F_sti
OVX	-0.0000795 (-1.64)	-0.000235** (-2.16)	-0.000119 (-1.44)	-0.0000554 (-0.70)	-0.0000219 (-0.41)	-0.0000206 (-0.40)	-0.0000292 (-0.70)	-0.0000594 (-1.24)	-0.0000870 (-1.34)	-0.0000722 (-0.94)
EXV_sti	1.980 (0.66)	-10.06 (-1.49)	-1.950 (-0.38)	2.903 (0.59)	4.124 (1.24)	4.074 (1.26)	3.461 (1.33)	2.951 (0.99)	2.080 (0.51)	0.468 (0.10)
lr_sti	-0.136*	-0.119	-0.00354	-0.122	-0.175**	-0.172**	-0.180***	-0.163**	-0.164	-0.180



	(-1.69)	(-0.66)	(-0.03)	(-0.93)	(-1.98)	(-1.99)	(-2.60)	(-2.06)	(-1.52)	(-1.41)
_cons	-0.0210 (-0.43)	0.140 (1.29)	0.0153 (0.19)	-0.0521 (-0.66)	-0.0631 (-1.18)	-0.0558 (-1.07)	-0.0389 (-0.93)	-0.0204 (-0.43)	0.00606 (0.09)	0.0470 (0.61)
N	183	183	183	183	183	183	183	183	183	183
t statistics in parentheses										
* p<0.1, ** p<0.05, *** p<0.01										
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	r_sti	r_sti	r_sti	r_sti	r_sti	r_sti	r_sti	r_sti	r_sti	r_sti
COVOX_sti	-0.479* (-1.93)	0.418 (0.79)	-0.500 (-1.02)	-0.580 (-1.59)	-0.438 (-1.56)	-0.479* (-1.93)	-0.527** (-2.31)	-0.483** (-2.12)	-0.553* (-1.92)	-0.250 (-0.62)
lr_sti	-0.240*** (-3.33)	0.0194 (0.13)	-0.0712 (-0.50)	-0.197* (-1.86)	-0.177** (-2.16)	-0.240*** (-3.33)	-0.221*** (-3.34)	-0.192*** (-2.90)	-0.152* (-1.81)	-0.197* (-1.68)
_cons	0.0361** (2.29)	-0.0748** (-2.22)	0.00725 (0.23)	0.0261 (1.13)	0.0248 (1.39)	0.0361** (2.29)	0.0463*** (3.20)	0.0519*** (3.59)	0.0645*** (3.53)	0.0598** (2.34)
N	183	183	183	183	183	183	183	183	183	183
t statistics in parentheses										
* p<0.1, ** p<0.05, *** p<0.01										

```
foreach v in vni {
    *Deploy equation 1:  $Qr(\tau) = \alpha_0(\tau) + \alpha_1(\tau)OVX_t + \alpha_2(\tau)EXV_t + \alpha_3(\tau)r_t - 1 + \epsilon(\tau)$ 
```

```
reg r_`v' ovx exv_`v' lr_`v'
est sto a1_`v'_OLS
qreg r_`v' ovx exv_`v' lr_`v', quantile(0.1)
est sto q1_`v'_q1
qreg r_`v' ovx exv_`v' lr_`v', quantile(0.2)
est sto q1_`v'_q2
qreg r_`v' ovx exv_`v' lr_`v', quantile(0.3)
est sto q1_`v'_q3
qreg r_`v' ovx exv_`v' lr_`v', quantile(0.4)
est sto q1_`v'_q4
qreg r_`v' ovx exv_`v' lr_`v', quantile(0.5)
est sto q1_`v'_q5
qreg r_`v' ovx exv_`v' lr_`v', quantile(0.6)
est sto q1_`v'_q6
qreg r_`v' ovx exv_`v' lr_`v', quantile(0.7)
est sto q1_`v'_q7
qreg r_`v' ovx exv_`v' lr_`v', quantile(0.8)
est sto q1_`v'_q8
qreg r_`v' ovx exv_`v' lr_`v', quantile(0.9)
est sto q1_`v'_q9
```

```
    *Deploy equation 2:  $Qr(\tau) = \alpha_0(\tau) + \alpha_1(\tau)COVOX_t + \alpha_2(\tau)r_t - 1 + \epsilon(\tau)$ 
```

```
qreg r_`v' COVOX_`v' lr_`v'
est sto a2_`v'_OLS
qreg r_`v' COVOX_`v' lr_`v', quantile(0.1)
est sto q2_`v'_q1
qreg r_`v' COVOX_`v' lr_`v', quantile(0.2)
est sto q2_`v'_q2
qreg r_`v' COVOX_`v' lr_`v', quantile(0.3)
est sto q2_`v'_q3
qreg r_`v' COVOX_`v' lr_`v', quantile(0.4)
est sto q2_`v'_q4
qreg r_`v' COVOX_`v' lr_`v', quantile(0.5)
est sto q2_`v'_q5
qreg r_`v' COVOX_`v' lr_`v', quantile(0.6)
est sto q2_`v'_q6
qreg r_`v' COVOX_`v' lr_`v', quantile(0.7)
est sto q2_`v'_q7
qreg r_`v' COVOX_`v' lr_`v', quantile(0.8)
est sto q2_`v'_q8
qreg r_`v' COVOX_`v' lr_`v', quantile(0.9)
est sto q2_`v'_q9
```

```
esttab a1_`v'_OLS q1_`v'_q1 q1_`v'_q2 q1_`v'_q3 q1_`v'_q4 q1_`v'_q5 q1_`v'_q6 q1_`v'_q7 q1_`v'_q8 q1_`v'_q9, star(*
0.1 ** 0.05 *** 0.01) replace
```

```
esttab a2_`v'_OLS q2_`v'_q1 q2_`v'_q2 q2_`v'_q3 q2_`v'_q4 q2_`v'_q5 q2_`v'_q6 q2_`v'_q7 q2_`v'_q8 q2_`v'_q9, star(*
0.1 ** 0.05 *** 0.01) replace
```

```
}
```

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
	r_vni	r_vni	r_vni	r_vni	r_vni	r_vni	r_vni	r_vni	r_vni	r_vni
OVX	-0.000213*** (-2.94)	-0.000398** (-3.38)	-0.000413*** (-3.21)	-0.000224** (-2.15)	-0.000145 (-1.54)	-0.000182** (-2.48)	-0.000112* (-1.69)	-0.000119 (-1.40)	-0.0000697 (-0.56)	-0.000140 (-1.38)
EXV_vni	0.282 (0.30)	0.921 (0.61)	1.045 (0.63)	0.870 (0.65)	0.0767 (0.06)	-0.256 (-0.27)	-0.108 (-0.13)	-0.281 (-0.26)	-0.599 (-0.37)	-0.237 (-0.18)

lr_vni	-0.0390 (-0.51)	-0.0485 (-0.39)	-0.246* (-1.82)	-0.0774 (-0.70)	-0.0188 (-0.19)	-0.0416 (-0.54)	-0.00381 (-0.05)	0.0352 (0.39)	-0.00344 (-0.03)	-0.149 (-1.40)
_cons	0.0258** (2.24)	-0.0326* (-1.74)	-0.00625 (-0.31)	0.000136 (0.01)	0.0137 (0.92)	0.0306*** (2.62)	0.0329*** (3.11)	0.0435*** (3.21)	0.0590*** (3.00)	0.0960*** (5.97)
N	183	183	183	183	183	183	183	183	183	183
t statistics in parentheses * p<0.1, ** p<0.05, *** p<0.01										
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
COVX_vni	r_vni -0.902 (-0.78)	r_vni -1.645 (-0.52)	r_vni 0.890 (0.32)	r_vni -0.791 (-0.35)	r_vni -0.238 (-0.16)	r_vni -0.902 (-0.78)	r_vni -1.113 (-0.87)	r_vni -1.670 (-1.12)	r_vni -1.071 (-0.50)	r_vni 0.664 (0.22)
lr_vni	0.0597 (0.81)	0.0178 (0.09)	-0.0682 (-0.38)	0.0502 (0.35)	0.0766 (0.79)	0.0597 (0.81)	0.0567 (0.70)	0.00372 (0.04)	-0.0834 (-0.61)	-0.242 (-1.29)
_cons	0.0207** (2.20)	-0.0595** (-2.30)	-0.0443* (-1.93)	-0.00246 (-0.14)	0.00259 (0.21)	0.0207** (2.20)	0.0300*** (2.88)	0.0428*** (3.53)	0.0542*** (3.08)	0.0730*** (3.03)
N	130	130	130	130	130	130	130	130	130	130
t statistics in parentheses * p<0.1, ** p<0.05, *** p<0.01										