Given $0 \le x < \infty$: $\alpha, \beta > 0$

 $f(x)=rac{x^{lpha-1}e^{-x/eta}}{\Gamma(lpha)eta^lpha}$ $\mu = \alpha\beta$

 $\sigma^2 = \alpha \beta^2$ $M(t) = (\frac{1}{1-\beta t})^{\alpha}$

Logistic (μ, β)

 $\mbox{Given} \ -\infty < x < \infty; \quad -\infty < \mu < \infty; \quad \beta > 0 \label{eq:continuous}$

 $f(x)=rac{e^{-(x-\mu)/eta}}{eta(1+e^{-(x-\mu)/eta)})^2}$

 $\mu = \mu$ $\sigma^2 = \frac{\pi^2 \beta^2}{3}$

 $M(t) = e^{\mu t} \Gamma(1+\beta t)$

Lognormal (μ, σ^2)

Distribution where the log of a variable is normally distributed. Similar in appearance to the gamma

 $\text{Given } 0 \leq x < \infty; \quad -\infty < \mu < \infty; \quad \sigma > 0$

$$f(x) = rac{1}{\sqrt{2\pi}\sigma} rac{e^{-[\log x - \mu]^2/(2\sigma^2)}}{x}$$
 $\mu = e^{\mu + (\sigma^2/2)}$
 $\sigma^2 = e^{2(\mu + \sigma^2)} - e^{2\mu + \sigma^2}$

 $EX^n = e^{n\mu + n^2\sigma^2/2}$

Normal (μ, σ^2)

Can be used to approximate many different kinds of distributions as their population sizes increase.

Given $-\infty < x < \infty$; $-\infty < \mu < \infty$; $\sigma > 0$

 $f(x)=rac{e^{-(x-\mu)^2/(2\sigma^2)}}{\sqrt{2\pi}\sigma}$

 $\mu = \mu$ $\sigma^2 = \sigma^2$

 $M(t)=e^{\mu t+\sigma^2t^2/2}$

Paretto (α, β)

Given $a < x < \infty; \quad \alpha, \beta > 0$

$$f(x) = \frac{\beta \alpha^{\beta}}{x^{\beta+1}}$$
 $\mu = \frac{\beta \alpha}{x^{\beta}} : \beta > 1$

$$\begin{split} \mu &= \frac{\beta\alpha}{\beta-1} \quad ; \beta > 1 \\ \sigma^2 &= \frac{\beta\alpha^2}{(\beta-1)^2(\beta-2)} \quad ; \beta > 2 \end{split}$$

T(v)

Given $-\infty < x < \infty; \quad v = 1, 2, 3, \dots$

 $f(x) = \frac{\Gamma(\frac{v+1}{2})}{\Gamma(\frac{v}{2})} \frac{1}{\sqrt{v\pi}} \frac{1}{(1+(\frac{x^2}{2r}))^{(v+1)/2}}$

 $\mu=0$; v>1

$$\begin{split} \sigma^2 &= \frac{v}{v-2} &\quad ; v > 2 \\ MX^n &= \begin{cases} \frac{\Gamma(\frac{v+1}{2})\Gamma(\frac{v-n}{2})}{\sqrt{\pi}\Gamma(v/2)} v^{n/2} & n < v; n \text{ is even} \\ 0 & n < v; n \text{ is odd} \end{cases} \end{split}$$

 $\mathrm{T}(v) = rac{N(0,1)}{\sqrt{rac{\chi_v^2}{n}}}$

Uniform (a, b)

All values between a and b are evenly distributed and x has equal chance of landing anywhere on that range.

Given $a \leq x \leq b$

- a is the lower bound of the distribution
- All values between a and b are equally distributed

 $f(x) = \frac{1}{b-a}$ $\mu = \frac{b+a}{2}$ $\sigma^2 = \frac{\frac{2}{(b-a)^2}}{12}$

 $M(t)=rac{e^{bt}-e^{at}}{t(b-a)}$

Weibull (γ, β)

 $\text{ Given } 0 \leq x < \infty; \quad \gamma, \beta > 0$

 $f(x) = rac{\gamma}{eta} x^{\gamma-1} e^{-x^{\gamma}/eta}$

 $\mu = eta^{1/\gamma} \Gamma(1 + \frac{1}{\gamma})$

 $\sigma^2 = eta^{2/\gamma} (\Gamma(1+rac{2}{\gamma}) - \Gamma^2(1+rac{1}{\gamma}))$

 $EX^n = \beta^{n/\gamma}\Gamma(1 + \frac{n}{\gamma})$

Multivariable Distributions

Covariance and Correlation

$$Cov(X, Y) = E((X - \mu_X)(Y - \mu_Y))$$

 $\rho_{XY} = \frac{Cov(X, Y)}{\sigma_X \sigma_Y}$

Multinomial Distribution

Given $\vec{\mu} \in \mathfrak{R}^k; \quad \Sigma \in \mathfrak{R}^{k^2}; \quad k \in \mathbb{N}$

 $f(\vec{x}) = (2\pi)^{-k/2} \det(\mathbf{\Sigma})^{-1/2} \exp\left(-\frac{1}{2}(\vec{x} - \vec{\mu})^T \mathbf{\Sigma}^{-1}(\vec{x} - \vec{\mu})\right)$

 $\sigma^2 = \Sigma$

 $M(\vec{t}) = \exp\left(\vec{\mu}^T \vec{t} + \frac{1}{2} \vec{t}^T \mathbf{\Sigma} \vec{t}\right)$

Bivariate case $(\mu_X, \mu_Y, \sigma_X^2, \sigma_Y^2, \rho)$

 $f(x,y) = \frac{1}{2\pi\sigma_{Y}\sigma_{Y}\sqrt{1-\sigma^{2}}} \exp\left(-\frac{1}{2(1-\rho^{2})}\left(\left(\frac{\vec{x}-\vec{\mu}_{X}}{\sigma_{X}}\right)^{2} - 2\rho\left(\frac{\vec{x}-\vec{\mu}_{x}}{\sigma_{X}}\right)\left(\frac{\vec{y}-\vec{\mu}_{Y}}{\sigma_{Y}}\right) + \left(\frac{\vec{y}-\vec{\mu}_{Y}}{\sigma_{Y}}\right)^{2}\right)\right)$

 $\vec{\mu} = \langle \mu_X, \mu_Y \rangle$ $\Sigma = \begin{bmatrix} \sigma_X^2 & \rho \sigma_X \sigma_Y \\ \rho \sigma_X \sigma_Y & \sigma_Y^2 \end{bmatrix}$

Families of Distributions

Exponential Family

Exponential Family

Any statistical distribution or family of distributions that can fit into the form:

 $f(x|\theta) = h(x)c(\theta) \exp \left(\sum_{i=1}^{k} w_i(\theta)t_i(x)\right)$

Binomial

$$\binom{n}{x}(1-p)^n \exp\left(\log(\frac{p}{1-p})x\right)$$

$$\frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{\mu^2}{2\sigma^2}\right) \exp\left(-\frac{x^2}{2\sigma^2} + \frac{\mu x}{\sigma^2}\right)$$

Location Scale Family

Location Scale Family

Any statistical distribution or family of distributions that can fit into the form: $g(x|\mu, \sigma) = \frac{1}{\sigma} f(\frac{x-\mu}{\sigma})$

Probability Inequalities

Chebychev's Inequality

Let X be a random variable and let g(x) be a non-negative function. Then, for any r>0

$$P(g(X) \ge r) \le \frac{Eg(X)}{r}$$

Normal Probability Inequality

With Z as a normal distribution

Very similar to a binomial, except there is more than one possible outcome per trial — as compared to success or failure in the binomial.

· Taking the marginal of any of the possible outcomes results in a regular binomial

Multivariable Normal $(\vec{\mu}, \Sigma)$

$$P(|Z| \geq t) \leq \sqrt{\frac{e}{\pi}} \frac{e^{-t^2/2}}{t}$$
 for all $t > 0$

Random Samples

Properties of the sample

Mean: $ar{X} = rac{X_1 + \ldots + X_n}{n} = rac{1}{n} \sum_{i=1}^n X_i$

Variance: $S^2 = \frac{1}{n-1} \sum_{i=1}^{n} (X_i - \bar{X})^2$

 $E(\sum_{i=1}^{n}g(X_{i}))=nEg(X_{1})$

 $\operatorname{Var}(\sum_{i=1}^{n}g(X_{i}))=n\operatorname{Var}g(X_{1})$

Properties of properties of the sample of random variables

 $E(\bar{X}) = \mu$ $\mathrm{Var}(ar{X}) = rac{\sigma^2}{n}$

Sample distributions of common distributions

Normal (μ, σ^2)

 $\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$

 S^2 is independent from $ar{X}$

Gamma (α, β)

 $\bar{X} \sim \operatorname{Gamma}\left(n\alpha, \frac{\beta}{n}\right)$

Cauchy

 $\operatorname{Cauchy}(0, \sigma_1) + \ldots + \operatorname{Cauchy}(0, \sigma_n) = \operatorname{Cauchy}(0, \sum_{i=1}^{n} \sigma_i)$

Chi-squared

$$\operatorname{Chi}(a_1) + \ldots + \operatorname{Chi}(a_n) = \operatorname{Chi}(\sum_{i=1}^n a_i)$$