### **Discrete Distributions**

### Bernoulli (p)

Models the probability of an event succeeding or failing, with x=1 as a success and x=0 as fail. There is only one trial.

Given 
$$x = 0, 1$$
;  $0 \le p \le 1$ 

- p is the probability of getting selected trait
- Has p probability of being 1 and 1-p probability of being 0

$$P(X=x) = p^x (1-p)^{1-x} \ ext{CDF} = egin{cases} 0 & x < 0 \ 1-p & 0 \leq x < 1 \ 1 & x \geq 1 \end{cases} \ \mu = p \ \sigma^2 = p(1-p) \ M(t) = (1-p) + pe^t \end{cases}$$

## Binomial (n, p)

Models the probability of getting x successes out of n trials with p probability of each trial succeeding.

Given 
$$x = 0, 1, 2, ..., n; 0 \le p \le 1$$

- $oldsymbol{\cdot}$  p is probability of selecting a particular trait
- n is number of samples in a round of sampling
- · Predicts probability of getting certain number of chosen trait in sample set

$$egin{aligned} P(X=x) &= inom{n}{x} p^x (1-p)^{n-x} \ \mu &= np \ \sigma^2 &= np (1-p) \ M(t) &= (pe^t + (1-p))^n \end{aligned}$$

# **Discrete Uniform** (N)

Models the chance of x as a result when all numbers from 1 to N are equally likely.

Given 
$$x = 1, 2, ..., N; N = 1, 2, ...$$

- N is the largest possible sample
- All numbers from 1 to N are equally likely

$$egin{aligned} P(X=x) &= 1/N \ ext{CDF} &= rac{\lfloor x 
floor + 1}{N} \ \mu &= rac{N+1}{2} \ \sigma^2 &= rac{(N+1)(N-1)}{12} \ M(t) &= rac{1}{N} \sum_{i=1}^N e^{it} \end{aligned}$$

# Geometric (p)

Models chance of taking  $\boldsymbol{x}$  trials to get a success, with each success having  $\boldsymbol{p}$  chance of succeeding.

Given 
$$x = 1, 2, ...; 0 \le p \le 1$$

- $oldsymbol{\cdot}$  p is probability of getting certain trait
- Predicts number of samples needed to get a sample of particular trait

$$egin{aligned} P(X=x) &= p(1-p)^{x-1} \ ext{CDF} &= egin{cases} 1 - (1-p)^{\lfloor x 
floor} & x \geq 0 \ 0 & x < 0 \end{cases} \ \mu &= 1/p \ \sigma^2 &= rac{1-p}{p^2} \ M(t) &= rac{pe^t}{1-(1-p)e^t} \end{aligned}$$

# Hypergeometric (N, K, M)

Predicts likelihood of selecting x successes after K trials of selecting from a population of N with M samples within that are successes without replacement.

Given 
$$x = 0, 1, 2, \dots, K; \quad M - (N - K) \le x \le M; \quad N, M, K = 0, 1, 2, \dots$$

ullet N is the population size

- M is the number of samples in the population with a certain trait
- K number of samples taken in a round of sampling
- Predicts the likelihood of selecting X samples of type M after selecting K samples from population N

$$egin{aligned} P(X=x) &= rac{inom{M}{x}inom{N-M}{K-x}}{inom{N}{K}} \ \mu &= KM/N \ \sigma^2 &= rac{KM(N-M)(N-K)}{N^2(N-1)} \end{aligned}$$

# **Negative Binomial** (r, p)

Predicts likelihood of sampling r successes of p probability after x+r samples

Given 
$$x = 0, 1, 2, ...; 0 \le p \le 1$$

- p is the probability of getting a particular trait in one sample
- r is the desired number of samples with a particular trait
- Predicts number of likelihood of getting r samples of trait after X+r samples

$$egin{align} P(X=x) &= inom{r+x-1}{x} p^r (1-p)^x \ \mu &= rac{r(1-p)}{p} \ \sigma^2 &= rac{r(1-p)}{p^2} \ M(t) &= ig(rac{p}{1-(1-p)e^t}ig)^r \ \end{pmatrix}$$

## **Poisson Distribution** $(\lambda)$

Assumes chances of an event happening in a short time is proportional to a large time. Predicts likelihood of x events happening in the next unit if  $\lambda$  events happen on average per unit where the unit does not need to be time.

- $N_0 = 0$
- For s>0,  $N_t$  and  $N_{t+s}-N_t$  are independent random variables
- ullet  $N_s$  and  $N_{t+s}-N_t$  are the same distribution
- $\lim_{t\downarrow 0}rac{P(N_t=1)}{t}=\lambda$  (Average rate of events will be  $\lambda$  everywhere)

•  $\lim_{t\downarrow 0}rac{P(N_t>1)}{t}=0$  (Can't have two events at the same time)

Given 
$$x = 0, 1, 2, \ldots$$
;  $0 \le \lambda$ 

- $\lambda$  is the number of times on average an event will happen within an interval
- Predicts number of times an event will happen within an interval
- Approximates the Binomial Distribution

$$egin{aligned} P(X=x) &= rac{e^{-\lambda}\lambda^x}{x!} \ \mu &= \lambda \ \sigma^2 &= \lambda \ M(t) &= e^{\lambda(e^t-1)} \end{aligned}$$

### **Continuous Distributions**

# Beta $(\alpha, \beta)$

Often used to model proportions that lie on (1,0)

Given 
$$0 \le x \le 1$$
;  $\alpha > 0$ ;  $\beta > 0$ 

$$egin{align} f(x) &= rac{x^{lpha-1}(1-x)^{eta-1}}{B(lpha,eta)} \ \mu &= rac{lpha}{lpha+eta} \ \sigma^2 &= rac{lphaeta}{(lpha+eta)^2(lpha+eta+1)} \ M(t) &= 1 + \sum\limits_{k=1}^{\infty}rac{t^k}{k!}\prod\limits_{r=0}^{k-1}rac{a+r}{lpha+eta+r} \ \end{array}$$

# Cauchy $(\theta, \sigma)$

Symmetric, bell shaped, and has no mean, often used as an extreme edge case for statistical theory, but also represents the ratio between two normal distributions among other unsuspecting things.

Given 
$$-\infty < x < \infty; \quad -\infty < \theta < \infty; \quad \sigma > 0$$
  $f(x) = rac{1}{\pi\sigma(1+(rac{x- heta}{\sigma})^2)}$ 

# Chi squared (p)

Given  $0 \le x < \infty$ ; p = 1, 2, 3, ...

$$egin{aligned} f(x) &= rac{x^{p/2-1}e^{-x/2}}{\Gamma(p/2)2^{p/2}} \ \mu &= p \ \sigma^2 &= 2p \ M(t) &= (rac{1}{1-2t})^{p/2} \end{aligned}$$

# **Double Exponential** $(\mu, \sigma)$

Given by reflecting the exponential around its mean. Does not have a bell shape.

Given 
$$-\infty < x < \infty; \quad -\infty < \mu < \infty; \quad \sigma > 0$$
  $f(x) = rac{e^{-|x-\mu|/\sigma}}{2\sigma}$ 

$$\mu = \mu$$

$$\sigma^2=2\sigma^2$$

$$M(t)=rac{e^{\mu t}}{1-(\sigma t)^2}$$

# **Exponential** $\beta$

Given  $0 \le x < \infty$ ;  $\beta > 0$ 

$$egin{aligned} f(x) &= rac{e^{-x/eta}}{eta} \ \mu &= eta \ \sigma^2 &= eta^2 \ M(t) &= rac{1}{1-eta t} \end{aligned}$$

# $\mathsf{F}\left(v_{1},v_{2} ight)$

Given  $0 \leq \infty$ ;  $v_1, v_2 = 1, 2, 3, \dots$ 

$$egin{aligned} f(x) &= rac{\Gamma(rac{v_1+v_2}{2})}{\Gamma(rac{v_1}{2})\Gamma(rac{v_2}{2})} ig(rac{v_1}{v_2}ig)^{v_1/2} rac{x^(v_1-2)/2}{(1+rac{v_1x}{v_2})^{(v_1+v_2)/2}} \ \mu &= rac{v_2}{v_2-2} \ \sigma^2 &= 2ig(rac{v_2}{v_2-2}ig)^2 rac{v_1+v_2-2}{v_1(v_2-4)} \ EX^n &= rac{\Gamma(rac{v_1+2n}{2})\Gamma(rac{v_2-2n}{2})}{\Gamma(v_1/2)\Gamma(v_2/2)} ig(rac{v_2}{v_1}ig)^n \quad ; n < rac{v_2}{2} \end{aligned}$$

# Gamma $(\alpha, \beta)$

 $\alpha$  being the shape parameter and  $\beta$  being the scale parameter, it is also a generalization of the exponential, erlang, weibull, and chi-squared distributions.

Given 
$$0 \leq x < \infty; \quad \alpha, \beta > 0$$
  $f(x) = rac{x^{lpha - 1}e^{-x/eta}}{\Gamma(lpha)eta^lpha}$   $\mu = lphaeta$   $\sigma^2 = lphaeta^2$   $M(t) = (rac{1}{1-eta t})^lpha$ 

# Logistic $(\mu, \beta)$

Given 
$$-\infty < x < \infty; \quad -\infty < \mu < \infty; \quad \beta > 0$$
 
$$f(x) = \frac{e^{-(x-\mu)/\beta}}{\beta(1+e^{-(x-\mu)/\beta)})^2}$$
 
$$\mu = \mu$$
 
$$\sigma^2 = \frac{\pi^2\beta^2}{3}$$
 
$$M(t) = e^{\mu t}\Gamma(1+\beta t)$$

# **Lognormal** $(\mu, \sigma^2)$

Distribution where the log of a variable is normally distributed. Similar in appearance to the gamma distribution.

Given 
$$0 \leq x < \infty; \quad -\infty < \mu < \infty; \quad \sigma > 0$$
 
$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \frac{e^{-(\log x - \mu)^2/(2\sigma^2)}}{x}$$
 
$$\mu = e^{\mu + (\sigma^2/2)}$$
 
$$\sigma^2 = e^{2(\mu + \sigma^2)} - e^{2\mu + \sigma^2}$$
 
$$EX^n = e^{n\mu + n^2\sigma^2/2}$$

# Normal $(\mu, \sigma^2)$

Can be used to approximate many different kinds of distributions as their population sizes increase.

Given 
$$-\infty < x < \infty; \quad -\infty < \mu < \infty; \quad \sigma > 0$$

$$egin{aligned} f(x) &= rac{e^{-(x-\mu)^2/(2\sigma^2)}}{\sqrt{2\pi}\sigma} \ \mu &= \mu \ \sigma^2 &= \sigma^2 \end{aligned}$$

### Paretto $(\alpha, \beta)$

Given  $a < x < \infty$ ;  $\alpha, \beta > 0$ 

$$egin{aligned} f(x) &= rac{eta lpha^eta}{x^{eta+1}} \ \mu &= rac{eta lpha}{eta-1} \quad ; eta > 1 \ \sigma^2 &= rac{eta lpha^2}{(eta-1)^2(eta-2)} \quad ; eta > 2 \end{aligned}$$

# t(v)

Given  $-\infty < x < \infty$ ;  $v = 1, 2, 3, \dots$ 

$$egin{align} f(x) &= rac{\Gamma(rac{v+1}{2})}{\Gamma(rac{v}{2})} rac{1}{\sqrt{v\pi}} rac{1}{(1+(rac{x^2}{v}))^{(v+1)/2}} \ \mu &= 0 \quad ; v > 1 \ \sigma^2 &= rac{v}{v-2} \quad ; v > 2 \ MX^n &= egin{cases} rac{\Gamma(rac{n+1}{2})\Gamma(rac{v-n}{2})}{\sqrt{\pi}\Gamma(v/2)} v^{n/2} & n < v; n ext{ is even} \ 0 & n < v; n ext{ is odd} \end{cases}$$

## Uniform (a, b)

All values between a and b are evenly distributed and x has equal chance of landing anywhere on that range.

Given  $a \le x \le b$ 

- a is the lower bound of the distribution
- b is the upper bound
- All values between a and b are equally distributed

$$f(x) = rac{1}{b-a} \ \mu = rac{b+a}{2} \ \sigma^2 = rac{(b-a)^2}{12} \ M(t) = rac{e^{bt}-e^{at}}{t(b-a)}$$

# Weibull $(\gamma, \beta)$

Given  $0 \le x < \infty$ ;  $\gamma, \beta > 0$ 

$$egin{aligned} f(x) &= rac{\gamma}{eta} x^{\gamma-1} e^{-x^{\gamma}/eta} \ \mu &= eta^{1/\gamma} \Gamma(1+rac{1}{\gamma}) \ \sigma^2 &= eta^{2/\gamma} (\Gamma(1+rac{2}{\gamma}) - \Gamma^2(1+rac{1}{\gamma})) \ EX^n &= eta^{n/\gamma} \Gamma(1+rac{n}{\gamma}) \end{aligned}$$

### **Families of Distributions**

# **Exponential Family**



#### **Exponential Family**

Any statistical distribution or family of distributions that can fit into the form:

$$f(x| heta) = h(x)c( heta) \exp\left(\sum\limits_{i=1}^k w_i( heta)t_i(x)
ight)$$

#### **Binomial**

$$\binom{n}{x}(1-p)^n \exp\left(\log(\frac{p}{1-p})x\right)$$

#### **Normal**

$$rac{1}{\sqrt{2\pi}\sigma} \mathrm{exp}\left(-rac{\mu^2}{2\sigma^2}
ight) \mathrm{exp}\left(-rac{x^2}{2\sigma^2}+rac{\mu x}{\sigma^2}
ight)$$

## **Location Scale Family**

#### **■** Def

#### **Location Scale Family**

Any statistical distribution or family of distributions that can fit into the

form:

$$g(x|\mu,\sigma)=rac{1}{\sigma}f(rac{x-\mu}{\sigma})$$