Section I: Maximum Likelihood Estimation

450C

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Overview

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- 2. Likelihood: An Intuition
- 3. Likelihood: A Recipe
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- 5. MLE and Uncertainty

Likelihood: An Intuition

• In previous classes, we asked:

"What is the effect of X on Y?"

"Given the (known) DGP, what is the probability of observing our sample?"

- In a clean causal inference setting, we know what the data-generating process is because we are in charge of assigning treatment (in an experiment) or assume that it is as-if random.
- But not all questions lend themselves to clean causal inference.
- What if we don't know the DGP?

"How can we best describe the process that generated this data?"

"Given our observed sample, what is the DGP?"

• When we don't know how X and Y are related and want to describe that relationship.

Likelihood: An Intuition (cont'd)

- Previously: given our **model**, how likely are the **results** that we observe?
- Now: given our **results**, how likely is the **model** that we assume?
- This approach yields powerful solutions to some questions
- Some methodological debates yielded dead ends (OLS v logit/probit)

Likelihood: A Recipe

1. Set up distribution that we assume has generated the data in question

$$Pr(y_i) = f(y_i)$$

2. Write down likelihood function -- the joint probability of observing all events under the assumed distribution

$$L(heta|y_i) = f(y_i| heta)$$

$$L(heta|\mathbf{y}) = \prod f(heta|y_i)$$

Likelihood: A Recipe (cont'd)

- 3. Refactor so that we can take the logs more easily;
- 4. Take the logs so we have the log-likelihood function:

$$\ell(\theta|\mathbf{y}) = \log(L(\theta|\mathbf{y}))$$

5. Find parameters that maximise log-likelihood:

$$rac{\partial \ell(heta|\mathbf{y})}{\partial heta_1} = 0
ightarrow heta_1^*$$

$$rac{\partial \ell(heta|\mathbf{y})}{\partial heta_2} = 0
ightarrow heta_2^*$$

6. Derive Fisher information to calculate variance of MLE estimate:

$$I_n(\theta) = -\mathbf{H}(L(\theta^*))$$

$$I_n(heta_1) = -rac{\partial^2 L(heta)}{\partial heta_1^2}(heta^*)$$

Likelihood: An Example

- Motivation: Suppose that we have N elections with two parties (A,B).
- A's vote share in the last five elections (y_samp) was:

```
## [1] 51.88486 51.50774 44.50988 44.34797 36.01733
```

- Can we describe the underlying data-generating process?
- What's our best guess?

Breakout Activity I: Guessing

How might we best model A's vote share?

What is your best guess for the vote share in the next election?

Normal MLE Estimation

- Let's assume: A's vote share is drawn i.i.d. from a normal distribution with (unknown) mean μ and (unknown) variance σ^2 .
- This gives us enough structure to proceed with MLE.
- If we **knew** mean and variance, we could calculate the probability of observing any value:

$$f(x) = \mathrm{pdf}(\mathcal{N}(\mu, \sigma^2))$$

• But we don't. So we have to make our best guess.

• (Step 2). We ask: what is the likelihood of observing any set combination of parameters (μ , σ^2), given the data that we observe?

$$egin{aligned} L(\mu,\sigma^2|\mathbf{y}) &= \prod f(y_i|\mu,\sigma^2) \ &= \prod rac{\exp(-rac{(y_i-\mu)^2}{2\sigma^2})}{\sqrt{2\pi\sigma^2}} \end{aligned}$$

• (Step 3). Refactoring.

$$L(\mu,\sigma^2|\mathbf{y}) = rac{\exp(-\sumrac{(y_i-\mu)^2}{2\sigma^2})}{(2\pi)^{n/2}\sigma^{2n/2}}$$

• (Step 4). Taking the logs.

$$\ell(\mu,\sigma^2|\mathbf{y}) = -\sum rac{(y_i-\mu)^2}{2\sigma^2} - rac{n}{2}\mathrm{log}(\sigma^2) + C$$

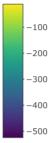
- Now we can plug in any combination of candidate values for μ and σ^2 into this function and we get a score.
- We have a nicely defined function \rightarrow time for some coding!

[1] -52.77709

Let's set up some more code to plot the likelihood for every combination of μ and σ^2 .

```
mu rg \leftarrow seq(30, 70, by = 0.5)
sigma2_rg \leftarrow seq(3, 60, by = 0.5)
viz_df ← expand.grid(mu = mu_rg, sigma2 = sigma2_rg) %>%
   as.data.frame %>%
   rowwise() %>%
  mutate(likelihood = likelihood normal(v samp, mu, sigma2))
max_row ← which.max(viz_df$likelihood)
viz_df[max_row, ]
## Source: local data frame [1 x 3]
## Groups: <by row>
##
## # A tibble: 1 x 3
        mu sigma2 likelihood
     <dbl> <dbl>
                       <dbl>
## 1 45.5 34
                       -11.3
```

```
viz_mat 		 viz_df %>%
  pivot_wider(names_from = mu, values_from = likelihood) %>%
  dplyr::select(-sigma2) %>% as.matrix
plot_ly(x = mu_rg, y = sigma2_rg, z = viz_mat,
  type="surface")
```



- We can also find the parameter combination that optimises the likelihood algebraically.
- Recall from Wednesday's lecture:

$$egin{aligned} \mu^* &= rac{\sigma y_i}{n} = ar{y} \ \sigma^{2*} &= rac{1}{n} \sum (y_i - ar{y})^2 \end{aligned}$$

• Again, this is something that we can implement in code.

Breakout activity II

Implement the two functions for the MLE of mean and variance in R and compute the estimated mean and variance of the MLE normal for y_samp.

• What if I told you that the data were generated with:

$$\mu=50, \sigma^2=25$$

- Our MLE estimate only takes the "sample" from the DGP.
- We can't make any assumptions about the parameters in the DGP: that's the thing we're trying to estimate using MLE!
- But because of the convergence in distribution, we can still infer how likely the observed MLE estimate is if we assume a true parameter θ_0 .

- ullet Let's create M samples with size n from our true DGP.
- For each of these samples, we calculate the MLE estimate and its variance.

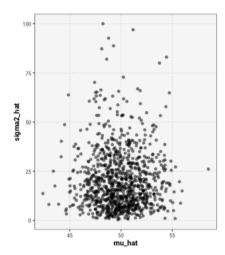
```
get_mean_variance 
  function(n){
   y_samp 
   rnorm(n, 50, 5)
   y_mean 
  mean(y_samp)
   y_sigma 
   1/n * sum((y_samp - y_mean)^2)
   return(c(y_mean, y_sigma))
}

n 
  5
m 
  1000

rep_vec 
  1:m
names(rep_vec) 
  1:m

samp_df 
  map_dfr(rep_vec, 
  get_mean_variance(n)) %>%
  t %>%
  as.data.frame
```

```
ggplot(samp_df, aes(V1, V2)) +
  geom_point(alpha = .5) +
  theme_tn() +
  labs(x = "mu_hat", y = "sigma2_hat")
```



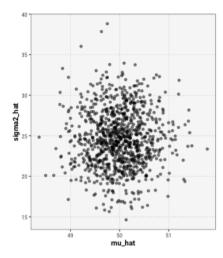
This is a two-dimensional distribution. We can characterise its (empirical) mean and variance.

What do these quantities correspond to?

What happens if we increase the sample size?

```
n ← 100
samp_df ← map_dfr(rep_vec, ~ get_mean_variance(n)) %>%
    t %>%
    as.data.frame
```

```
ggplot(samp_df, aes(V1, V2)) +
  geom_point(alpha = .5) +
  theme_tn() +
  labs(x = "mu_hat", y = "sigma2_hat")
```



12.4 mle_var

What do these quantities correspond to?

24.6

2

Recall the asymptotic property of MLE estimators as $n \to \infty$:

$$p(\hat{\mu},\hat{\sigma}^2) \stackrel{d}{ o} ext{MVN}\Big((ar{y},rac{1}{n}\sum(y_i-ar{y})^2),egin{bmatrix} rac{\sigma^2}{n} & 0 \ 0 & rac{2(\sigma^2)^2}{n} \end{bmatrix}\Big)$$

Since we know the true parameters...

- ullet We have problems when n=5
- Mean ($\mu=50$) and Variance ($\sigma^2=25$) parameters are correctly estimated with n=100
- "Sampling" uncertainty of these parameters falls with sample size
- Variance of sampling distribution converges to 25/100=0.25 and $(2*25^2)/100=12.5$, respectively

Summary

- Generic recipe for a how to think about likelihood.
 - Decide on model
 - \circ Write down likelihood f'n: how likely is θ given the observed data?
 - Refactor and take the logs
 - \circ Maximise w.r.t. θ (take first derivative)
 - Derive second derivative / Hessian for variance
- Applied to normal distribution (both with algebra and with code)
- Thinking about uncertainty and inference in the context of MLE