Desenvolvimento

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Modelo: $y_i = b_0 + b_1 x_i + \epsilon$, onde $\epsilon \sim N(0, \sigma^2)$ e $y_i \sim N(b_0 + b_1 x_i, \sigma^2)$.

Mostre que o estimador MSE = $\frac{SQE}{n-2}$ é não viciado para σ^2 .

$$E[MSE] = E\left[\frac{SQE}{n-2}\right] = \frac{1}{n-2}E[SQE] = \frac{1}{n-2}E\left[\sum_{i=1}^{n} (y_i - \hat{y}_i)^2\right]$$
$$= \frac{1}{n-2}E\left[\sum_{i=1}^{n} (y_i - \hat{b}_0 - \hat{b}_1 x_i)^2\right]$$

Como $\hat{b}_0 = \bar{y} - \hat{b}_1 \bar{x}$

$$E[MSE] = \frac{1}{n-2} E\left[\sum_{i=1}^{n} (y_i - \bar{y} + \hat{b}_1 \bar{x} - \hat{b}_1 x_i)^2\right] = \frac{1}{n-2} E\left\{\sum_{i=1}^{n} [(y_i - \bar{y}) - \hat{b}_1 (x_i - \bar{x})]^2\right\}$$

$$= \frac{1}{n-2} E\left\{\sum_{i=1}^{n} [(y_i - \bar{y})^2 - 2\hat{b}_1 (y_i - \bar{y})(x_i - \bar{x}) + \hat{b}_1^2 (x_i - \bar{x})^2]\right\}$$

$$= \frac{1}{n-2} E\left\{\sum_{i=1}^{n} (y_i - \bar{y})^2 - 2\hat{b}_1 \sum_{i=1}^{n} (y_i - \bar{y})(x_i - \bar{x}) + \hat{b}_1^2 \sum_{i=1}^{n} (x_i - \bar{x})^2\right\}$$

Como

$$\hat{b}_1 = \frac{\sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$
 [Mínimos quadrados]

então

$$\delta = \hat{b}_1 \sum_{i=1}^n (x_i - \bar{x})^2$$

Portanto

$$E[MSE] = \frac{1}{n-2} E\left[\sum_{i=1}^{n} (y_i - \bar{y})^2 - 2\hat{b}_1^2 \sum_{i=1}^{n} (x_i - \bar{x})^2 + \hat{b}_1^2 \sum_{i=1}^{n} (x_i - \bar{x})^2\right]$$
$$= \frac{1}{n-2} E\left[\sum_{i=1}^{n} (y_i - \bar{y})^2 - \hat{b}_1^2 \sum_{i=1}^{n} (x_i - \bar{x})^2\right]$$

Obs.:

$$\hat{b}_1 = \frac{S_{xy}}{S_{xx}}$$

Continuando

$$\begin{split} E[\text{MSE}] &= \frac{1}{n-2} \Biggl\{ \sum_{i=1}^n E[(y_i - \bar{y})^2] - S_{xx} \sum_{i=1}^n E[\hat{b}_1^2] \Biggr\} = \frac{1}{n-2} \Biggl\{ \sum_{i=1}^n E[(y_i - \bar{y})^2] - S_{xx} E[\hat{b}_1^2] \Biggr\} \\ &= \frac{1}{n-2} \Biggl\{ \sum_{i=1}^n E[(y_i - \bar{y})^2] - S_{xx} \bigl[Var(\hat{b}_1) + E^2(\hat{b}_1) \bigr] \Biggr\} \end{split}$$

Temos que

$$E^2(\hat{b}_1) = b_1^2, \quad Var(\hat{b}_1) = \frac{\sigma^2}{S_{rr}}$$

Logo

$$E[\text{MSE}] = \frac{1}{n-2} \left\{ \sum_{i=1}^{n} \underbrace{E[(y_i - \bar{y})^2]}_{\hat{\xi}} - S_{xx} \left[\frac{\sigma^2}{S_{xx}} + b_1^2 \right] \right\}$$

Em ξ , temos que

$$\bar{y} = \frac{1}{n} \sum_{i=1}^{n} (b_0 + b_1 x_i + \epsilon_i) = \frac{1}{n} \sum_{i=1}^{n} b_0 + \frac{1}{n} \sum_{i=1}^{n} b_1 x_i + \frac{1}{n} \sum_{i=1}^{n} \epsilon_i = b_0 + b_1 \bar{x} + \bar{\epsilon}$$

Portanto

$$\xi = E[(b_0 + b_1 x_i + \epsilon_i - b_0 - b_1 \bar{x} - \bar{\epsilon})^2] = E\{[b_1 (x_i - \bar{x}) + (\epsilon_i - \bar{\epsilon})]^2\}$$

$$= E[b_1^2 (x_i - \bar{x}) + 2b_1 (x_i - \bar{x})(\epsilon_i - \bar{\epsilon}) + (\epsilon_i - \bar{\epsilon})^2]$$

$$= E[b_1^2 (x_i - \bar{x})^2] + E[2b_1 (x_i - \bar{x})(\epsilon_i - \bar{\epsilon})] + E[(\epsilon_i - \bar{\epsilon})^2]$$

$$= b_1^2 (x_i - \bar{x})^2 + \underbrace{2b_1 (x_i - \bar{x})E[\epsilon_i - \bar{\epsilon}]}_{0} + E[(\epsilon_i - \bar{\epsilon})^2]$$

Veja que

$$\begin{split} E[(\epsilon_i - \bar{\epsilon})^2] &= E[\epsilon_i^2 - 2\epsilon_1 \bar{\epsilon} + \bar{\epsilon}^2] = E[\epsilon_i^2] - 2E[\epsilon_i \bar{\epsilon}] + E[\bar{\epsilon}^2] \\ &= \sigma^2 - 2E\left[\epsilon_i \frac{1}{n} \sum_i \epsilon_i\right] + E[\bar{\epsilon}^2] = \sigma^2 - 2E\left[\frac{1}{n}\epsilon_i^2\right] + E[\bar{\epsilon}^2] \\ &= \sigma^2 - \frac{2}{n}E[\epsilon_i^2] + E[\bar{\epsilon}^2] = \sigma^2 - \frac{2}{n}\sigma^2 + E[\bar{\epsilon}^2] \end{split}$$

Só nos resta lidar com

$$E[\bar{\epsilon}^2] = E\left[\left(\frac{1}{n}\sum_{i}\epsilon_i\right)^2\right] = E\left[\frac{1}{n^2}\sum_{i}\sum_{j}\epsilon_i\epsilon_j\right]$$

Como os erros são independentes, temos que

$$E[\epsilon_i \epsilon_j] = E[\epsilon_i] E[\epsilon_j] = 0, \quad i \neq j$$

ao passo que

$$E[\epsilon_i^2] = \sigma^2$$

Logo

$$E[\bar{\epsilon}^2] = \frac{1}{n^2} \left[\sum_i E[\epsilon_i^2] + \sum_{i \neq j} E[\epsilon_i \epsilon_j] \right] = \frac{1}{n^2} n\sigma^2 = \frac{1}{n} \sigma^2$$

Em ξ , teremos

$$\xi = b_1^2 (x_i - \bar{x})^2 + E[(\epsilon_i - \bar{\epsilon})^2] = b_1^2 (x_i - \bar{x})^2 + \sigma^2 - \frac{1}{n} \sigma^2$$

que, no início

$$E[MSE] = \frac{1}{n-2} \left\{ \sum_{i=1}^{n} \left[b_1^2 (x_i - \bar{x})^2 + \sigma^2 - \frac{1}{n} \sigma^2 \right] - S_{xx} \left[\frac{\sigma^2}{S_{xx}} + b_1^2 \right] \right\}$$

$$= \frac{1}{n-2} \left\{ b_1^2 S_{xx} + n\sigma^2 - \sigma^2 - \sigma^2 - b_1^2 S_{xx} \right\} = \frac{1}{n-2} (n\sigma^2 - 2\sigma^2)$$

$$= \frac{1}{n-2} (n-2)\sigma^2 = \sigma^2$$

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Mostre que

$$t(\hat{\beta}) = \frac{\hat{\beta} - \beta}{S_{e}} \sqrt{\sum (x_{i} - \bar{x})^{2}} \sim t_{n-2}$$

Desenvolvimento:

Temos que, se $Z \sim N(0,1)$ e $Y \sim \mathcal{X}_v^2$, então

$$\frac{Z}{\sqrt{Y/v}} \sim t_v$$

Sabemos que

$$\hat{\beta} \sim N\left(\beta, \frac{\sigma_e^2}{\sum (x_i - \bar{x})^2}\right)$$

Portanto

$$t(\hat{\beta}) = \left(\frac{\hat{\beta} - \beta}{\sigma_e} \sqrt{\sum (x_i - \bar{x})^2}\right) \frac{\sigma_e}{S_e} = \frac{Z}{\frac{S_e}{\sigma_e}}$$

Veja que

$$\left(\frac{S_e}{\sigma_e}\right)^2 = \frac{S_e^2}{\sigma_e^2} = \frac{\sum (y_i - \hat{y}_i)^2}{\sigma_e^2} \frac{1}{n-2}$$

Entretanto

$$\sum (y - \hat{y})^2 = \sum \hat{\varepsilon}_i^2, \qquad \hat{\varepsilon}_i \sim N(0, \sigma_e^2) e^{\frac{\hat{\varepsilon}_i}{\sigma_e}} \sim N(0, 1)$$

Logo

$$\frac{\sum \varepsilon_i^2}{\sigma_e^2} = \sum \frac{\varepsilon_i^2}{\sigma_e^2} = \sum \left(\frac{\varepsilon_i}{\sigma_e}\right)^2 = \sum Z^2 \sim \mathcal{X}_{n-2}^2$$

Finalmente

$$\frac{S_e^2}{\sigma_e^2} = \frac{Y}{n-2} \Rightarrow \frac{S_e}{\sigma_e} = \sqrt{\frac{Y}{n-2}}$$

tal que

$$t(\hat{\beta}) = \frac{Z}{\sqrt{\frac{Y}{n-2}}} \sim t_{n-2}$$