

作业9. 习题1. 解(1). 若 $x_1 \geq c$, 则 $\forall k \in \{1, 2, \dots, n\} f(x_k; c, \theta) = \frac{1}{\theta^n} \cdot e^{\frac{nc}{\theta} - (\frac{1}{\theta} + \frac{1}{\theta^2} - 1)x_1}$.

此时 $f(x_1; c, \theta) \cdot f(x_2; c, \theta) \cdots f(x_n; c, \theta) = (\frac{1}{\theta})^n \cdot e^{\frac{nc}{\theta}}$.

若 $x_1 < c$, $\exists k \in \{1, 2, \dots, n\} f(x_k; c, \theta) = 0$

此时 $f(x_1; c, \theta) \cdots f(x_n; c, \theta) = 0$

$$\text{故 } L(c, \theta) = \begin{cases} \frac{1}{\theta^n} \cdot \exp\left\{\frac{nc}{\theta} - \frac{1}{\theta} \cdot \sum_{i=1}^n x_i\right\}, & c \leq x_1 \\ 0, & c > x_1 \end{cases}$$

利用极大似然法, 可以看到 $L(c, \theta)$ 关于 c 单调递增 而 $c \leq x_1$

故 $\hat{c} = x_1$ 由 $\nabla_{\theta} L(c, \theta) = 0$ 可得

$$-\frac{n}{\theta^n \cdot \theta} \cdot \exp\left\{\frac{nc}{\theta} - \frac{1}{\theta} \sum_{i=1}^n x_i\right\} + \frac{1}{\theta^n} \cdot \exp\left\{\frac{nc}{\theta} - \frac{1}{\theta} \sum_{i=1}^n x_i\right\} \cdot \left(-\frac{nc}{\theta^2} + \frac{1}{\theta^2} \sum_{i=1}^n x_i\right) = 0$$

$$\text{可得 } \hat{\theta} = \frac{\sum_{i=1}^n x_i}{n} - \hat{c} = \bar{x} - x_1$$

$$12). E_{c, \theta}(x) = \int_c^{+\infty} \frac{t}{\theta} e^{-\frac{t-c}{\theta}} dt = \theta e^{\frac{c}{\theta}} \cdot \int_{\frac{c}{\theta}}^{+\infty} k e^{-k} dk \quad \text{取 } k = \frac{t}{\theta}$$

$$= \theta e^{\frac{c}{\theta}} \left[(0 - (-\frac{c}{\theta} e^{-\frac{c}{\theta}})) + (0 - (-e^{-\frac{c}{\theta}})) \right] = c + \theta$$

$$E_{c, \theta}(x^2) = \int_c^{+\infty} \frac{t^2}{\theta} e^{-\frac{t-c}{\theta}} dt = \theta^2 e^{\frac{c}{\theta}} \cdot \int_{\frac{c}{\theta}}^{+\infty} k^2 e^{-k} dk \quad \text{取 } k = \frac{t}{\theta}$$

$$= \theta^2 e^{\frac{c}{\theta}} \cdot \left(-k^2 e^{-k} \Big|_{\frac{c}{\theta}}^{+\infty} + 2 \int_{\frac{c}{\theta}}^{+\infty} k e^{-k} dk \right)$$

$$= \theta^2 e^{\frac{c}{\theta}} \left(\frac{c^2}{\theta^2} e^{-\frac{c}{\theta}} + \frac{2(c+\theta)}{\theta} e^{-\frac{c}{\theta}} \right) = c^2 + 2c\theta + 2\theta^2$$

$$E_{c, \theta}(x) = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x} \quad E_{c, \theta}(x^2) = \frac{1}{n} \sum_{i=1}^n x_i^2$$

$$\theta^2 = E_{c, \theta}(x^2) - [E_{c, \theta}(x)]^2 = \frac{1}{n} \sum_{i=1}^n x_i^2 - \left(\frac{1}{n} \sum_{i=1}^n x_i \right)^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$$

$$\hat{\theta} = \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2}$$

这个利用公式
 $D(x) = E(x^2) - [E(x)]^2$

$$\hat{c} = \bar{x} - \hat{\theta} = \bar{x} - \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2}$$

题2. 解: (1)

$$L(\theta) = f(x_1; \theta) \cdots f(x_n; \theta)$$

$$H(\theta) = \ln L(\theta) = \sum_{k=1}^n \ln f(x_k; \theta) = -n \ln \theta + \frac{1-\theta}{\theta} \sum_{i=1}^n \ln x_i$$

$$\nabla H(\theta) = 0 \text{ 可得 } -\frac{n}{\theta} + (-\frac{1}{\theta^2}) (\sum_{i=1}^n \ln x_i) = 0 \text{ 得 } \hat{\theta} = -\frac{1}{n} \sum_{i=1}^n \ln x_i$$

故 θ 的极大似然估计量为 $\hat{\theta} = -\frac{1}{n} \sum_{i=1}^n \ln x_i$

12) 证明: $E(\cdot)$ 对 $V: i \in \{1, \dots, n\}$

$$E(\ln x_i) = \int_0^1 \ln x_i \cdot \frac{1}{\theta} x_i^{\frac{1-\theta}{\theta}} dx_i$$

$$= \frac{1}{\theta} \cdot \int_0^1 \ln x \cdot x^{\frac{1-\theta}{\theta}} dx = \frac{1}{\theta} \cdot \int_0^1 \ln x d(x^\theta)$$

$$= \ln x \cdot x^{\frac{\theta}{\theta}} \Big|_0^1 - \int_0^1 x^{\frac{\theta}{\theta}-1} dx = -\theta$$

$$E(\hat{\theta}) = E\left(-\frac{1}{n} \sum_{i=1}^n \ln x_i\right) = -\frac{1}{n} \cdot \sum_{i=1}^n E(\ln x_i) = -\frac{1}{n} \cdot (-n\theta) = \theta$$

故 $\hat{\theta}$ 为 θ 的无偏估计

习题3. 解: 样本的条件概率函数

$$q(x|M) = \pi(x_1|M) \cdots \pi(x_n|M)$$

$$= \frac{1}{5^n (\sqrt{2\pi})^n} \exp\left\{-\frac{1}{2\delta^2} \sum_{i=1}^n (x_i - \mu)^2\right\}$$

$$\text{联合密度函数 } h(M|x) = \frac{q(x|M) \cdot \pi(M)}{\int_{-\infty}^{+\infty} q(x|M) \cdot \pi(M) dM}$$

$$\text{可得 } h(M|x) \propto \exp\left[-\frac{1}{2\delta^2} \sum_{i=1}^n (x_i - M)^2\right] \exp\left[-\frac{1}{2\delta_M^2} (M - M_0)^2\right]$$

$$\text{我们要使 } h(M|x) \propto \exp\left[\frac{-(a-b)^2}{2c^2}\right] \quad a \text{ 可以被 } M \text{ 替换}$$

找到 b, c , 这样就可以找到 $M|x \sim N(b, c^2)$

(结构类似)

$$\exp\left[-\frac{1}{2\delta^2} \sum_{i=1}^n (x_i - M)^2 - \frac{1}{2\delta_M^2} (M - M_0)^2\right] = \quad (\text{双顶})$$

作为一个之前根本没有过概率论的小白, 结合自己的理解, 根据看到的方法, 尝试一下.

由上页, $h(M|x) \propto \exp\left[-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - M)^2\right] \cdot \exp\left[-\frac{1}{2\sigma_M^2} (M - M_0)^2\right]$

我们可以更进一步地去相似, 只要满足 消掉的式子与M无关

对 $-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - M)^2 - \frac{1}{2\sigma_M^2} (M - M_0)^2$ 而言. 去掉任何与M相关的式子,

可得 $-\frac{1}{2\sigma^2} \sum_{i=1}^n (M^2 - 2Mx_i) - \frac{1}{2\sigma_M^2} (M^2 - 2MM_0)$

即 $M^2 \cdot \left(-\frac{n}{2\sigma^2} - \frac{1}{2\sigma_M^2}\right) + M \left(\frac{\sum_{i=1}^n x_i}{\sigma^2} + \frac{M_0}{\sigma_M^2}\right)$

$= -\frac{1}{2} \left[M^2 \left(\frac{n}{\sigma^2} + \frac{1}{\sigma_M^2}\right) - 2M \left(\frac{\sum_{i=1}^n x_i}{\sigma^2} + \frac{M_0}{\sigma_M^2}\right) \right]$

取 $A^2 = \frac{n}{\sigma^2} + \frac{1}{\sigma_M^2}$ $B = \frac{\sum_{i=1}^n x_i}{\sigma^2} + \frac{M_0}{\sigma_M^2}$
 $\hookrightarrow = -\frac{1}{2} [A^2 M^2 - 2BM] = -\frac{1}{2} \cdot A^2 \left(M - \frac{B}{A^2}\right)^2 + \frac{B^2}{2A^2}$

由于A, B皆与M无关, 故 $\frac{B^2}{2A^2}$ 与M无关
 故该式 $\propto -\frac{1}{2} A^2 \left(M - \frac{B}{A^2}\right)^2 = -\frac{1}{2} \cdot \frac{(M - \frac{B}{A^2})^2}{\frac{1}{A^2}}$

即 $h(M|x) \propto \exp\left[-\frac{1}{2} \frac{(M - \frac{B}{A^2})^2}{\frac{1}{A^2}}\right]$

即 $M|x \sim N\left(\frac{B}{A^2}, \frac{1}{A^2}\right) = N\left(\frac{\frac{\sum_{i=1}^n x_i}{\sigma^2} + \frac{M_0}{\sigma_M^2}}{\frac{n}{\sigma^2} + \frac{1}{\sigma_M^2}}, \frac{1}{\frac{n}{\sigma^2} + \frac{1}{\sigma_M^2}}\right)$

这题做得真过瘾~

习题4. 解: $F(\alpha, \beta) \equiv X \sim \Gamma(\alpha, \beta)$ 即为 $f(x) = \frac{1}{\beta^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-\frac{x}{\beta}}$, $x > 0$

(这个能否在题目中告诉我公式呢??)

$\pi(\lambda) \propto \lambda^{\alpha-1} e^{-\beta\lambda}$

$g(x|\lambda) = P(x_1|x) \cdots P(x_n|x)$
 $\quad \quad \quad x_1 + \cdots + x_n$

$= \frac{\lambda^n}{x_1! \cdots x_n!} e^{-n\lambda}$

$\propto \lambda^{x_1 + \cdots + x_n} \cdot e^{-n\lambda}$

这个在lec 20 PPT
 还写错了:!

泊松分布
 $f(x) = e^{-\lambda} \frac{\lambda^x}{x!}$
 $x \geq 0$

$h(\lambda|x) = \frac{g(x|\lambda) \cdot \pi(\lambda)}{\int_{-\infty}^{+\infty} g(x|\lambda) \cdot \pi(\lambda)}$
 $\propto g(x|\lambda) \cdot \pi(\lambda)$

$$h(\lambda|x) \propto g(x|\lambda) \cdot \pi(\lambda) = \lambda^{x_1 + \dots + x_n} \cdot e^{-n\lambda} \cdot \lambda^{\alpha-1} \cdot e^{-\beta\lambda}$$

$$= \lambda^{x_1 + \dots + x_n + \alpha - 1} e^{-n\lambda - \beta\lambda}$$

$$\lambda|x \sim \Gamma(x_1 + \dots + x_n + \alpha, n + \beta)$$

λ 的后验期望估计 $\hat{\lambda} = \frac{x_1 + \dots + x_n + \alpha}{n + \beta}$