

Interaction/Moderation Effects Using Regression

Details behind interaction models. Use of centered, standardized and raw data in interaction models.

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February 1, 2025

Agenda

- Understand details behind interaction models.
- Understand details on use of centered, standardized and raw data in interaction models.
- Develop an interaction/moderation model and estimate it using STATA.

Do attitude towards immigrants differs from a low-unemployment year (2007) to a high-unemployment year (2008)?



Understanding Interaction/Moderation Effects

- **Linear Additive Models:** Assume that the effect of an independent variable on a dependent variable is **constant** across all values of other independent variables.
- **Non-Additive (Interaction) Models:** Allow the effect of one independent variable to vary depending on another variable, providing a more nuanced understanding.

Why Interaction Effects Matter

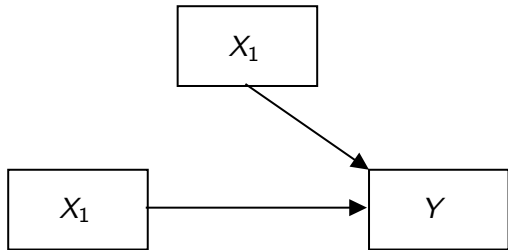
- Standard statistical models often assume **invariance** in relationships between variables.
- Interaction models help identify situations where this assumption **does not hold**, leading to more accurate conclusions.

Defining Interaction/Moderation Effects

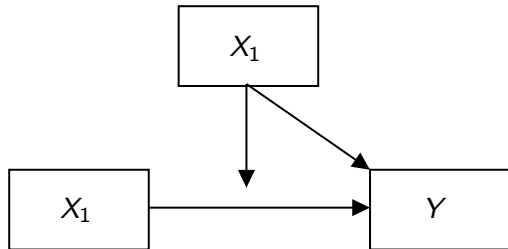
- Occurs when a **moderator variable** (X_1) influences the relationship between an **independent variable** (X_2) and a **dependent variable** (Y).
- This is demonstrated by a **significant change** in the effect size and/or direction of X_2 on Y at different values of X_1 .
- Interaction effects reveal **conditional relationships** that linear additive models may overlook.
- They are essential in empirical research to **capture complexity** in social, economic, and behavioral studies.

(See next figure for visual representation.)

Interaction/Moderation Effect Diagram



Main Effect



Interaction/Moderation Effect

Product-Term Approach

- The product-term approach involves creating a new variable (X_3) by multiplying two interacting variables ($X_1 \times X_2$) and including it in the regression model alongside X_1 and X_2 .
- This results in the following regression model:

$$E[Y_i] = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{1i} X_{2i} \quad (1)$$

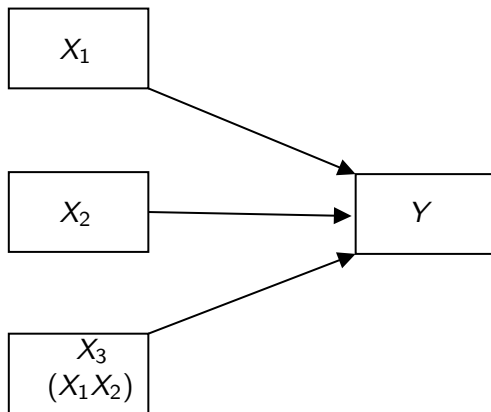
Interpreting the Product-Term Approach

- In an additive model, the equation is:

$$E[Y_i] = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} \quad (2)$$

- Here, β_1 and β_2 represent the main effects of X_1 and X_2 on Y , assuming their effects are constant.
- In the interaction model (Equation (1)), β_1 and β_2 now represent conditional effects, meaning their impact on Y varies depending on the value of the other variable.

Product-Term Interaction Diagram



```
*Load the dataset
use workout.dta, clear

*Generate the interaction term manually
gen healthage = health *age

*Regression with manually created interaction term
reg whours health age healthage

*Alternative: Using factor variable notation
reg whours c.health c.age c.health#c.age

*Simplified notation using ##
reg whours c.health##c.age
```

Interaction Between a Continuous Predictor and a Continuous Moderator

Population Regression Model:

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 (X_{1i} \times X_{2i}) + \varepsilon_i \quad (3)$$

Model Specification:

- X_1 : Predictor (Health Motivation)
- X_2 : Continuous moderator (Age)
- $E[Y_i]$: Expected number of hours spent working out in a month
- β_0 : Expected workout hours when both health motivation and age are zero
- β_1 : Effect of health motivation when age is zero
- β_2 : Effect of age when health motivation is zero
- β_3 : Change in the effect of health motivation on workout hours as age increases

Interpretation:

- The interaction term β_3 determines whether the relationship between health motivation and workout hours varies across different age levels.
- A positive β_3 suggests that the impact of health motivation on workout hours increases with age, while a negative β_3 suggests a decreasing effect.

Conditional Effects at Different Ages

$$E[Y_i] = (\beta_0 + a\beta_2) + (\beta_1 + a\beta_3)X_{1i} \quad (4)$$

- When $X_2 = 20$:

$$E[Y_i] = (\beta_0 + 20\beta_2) + (\beta_1 + 20\beta_3)X_{1i} \quad (5)$$

- When $X_2 = 30$:

$$E[Y_i] = (\beta_0 + 30\beta_2) + (\beta_1 + 30\beta_3)X_{1i} \quad (6)$$

- When $X_2 = 40$:

$$E[Y_i] = (\beta_0 + 40\beta_2) + (\beta_1 + 40\beta_3)X_{1i} \quad (7)$$

- When $X_2 = 50$:

$$E[Y_i] = (\beta_0 + 50\beta_2) + (\beta_1 + 50\beta_3)X_{1i} \quad (8)$$

- When $X_2 = 60$:

$$E[Y_i] = (\beta_0 + 60\beta_2) + (\beta_1 + 60\beta_3)X_{1i} \quad (9)$$

```
. *Regression with manually created interaction term
. reg whours health age healthage
```

Source		SS	df	MS	Number of obs	=	210
-----+-----					F(3, 206)	=	5.48
Model		808.220699	3	269.4069	Prob > F	=	0.0012
Residual		10119.7031	206	49.1247724	R-squared	=	0.0740
-----+-----					Adj R-squared	=	0.0605
Total		10927.9238	209	52.2867168	Root MSE	=	7.0089

whours		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----							
health		-1.796974	1.361392	-1.32	0.188	-4.481022	.8870748
age		-.4756232	.191777	-2.48	0.014	-.8537204	-.097526
healthage		.0639482	.0368331	1.74	0.084	-.00867	.1365665
_cons		27.6244	6.912041	4.00	0.000	13.99699	41.25181
-----+-----							

Conditional Effect of Health on Workout Hours at Different Ages

- We examine the conditional effect of health motivation on workout hours for individuals aged 16, 26, 36, 46, 56, 66, and 76.

```
margins , dydx(health) at(age=(16(10)76))
```

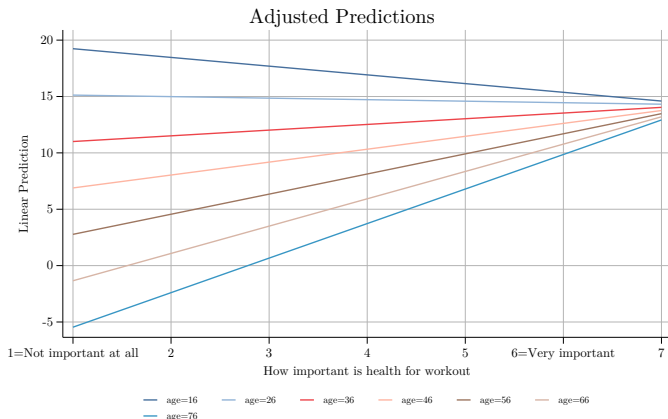
- The `dydx(health)` refers to the change-in-Y/change-in-X ratio.
- The `(10)` between 16 and 76 specifies the increment for age.

		Delta-method					
		dy/dx	Std. Err.	t	P> t	[95% Conf. Interval]	
health							
	_at						
	1	-.7738018	.8284528	-0.93	0.351	-2.407135	.8595316
	2	-.1343195	.5564723	-0.24	0.810	-1.231431	.9627917
	3	.5051629	.452023	1.12	0.265	-.3860216	1.396347
	4	1.144645	.6085426	1.88	0.061	-.0551248	2.344415
	5	1.784128	.8986986	1.99	0.048	.0123013	3.555954
	6	2.42361	1.231394	1.97	0.050	-.0041401	4.85136
	7	3.063092	1.57998	1.94	0.054	-.0519124	6.178097

Results:

- The effect of health on workout hours is not statistically significant for ages 16, 26, and 36.
- The effect is statistically significant ($p < 0.1$) for ages 46, 56, 66, and 76.
- The magnitude of this effect increases with age.
- At age 76, each unit increase in health motivation leads to an average increase of 3 workout hours per month.

```
*Calculate the mean-Y at all six values of health for each age 16, 26,  
36, 46, 56, 66 and 76  
margins, at(health=(1(1)6) age=(16(10)76))  
marginsplot, noci x(health) recast(line) xtitle("How important is health  
for workout")
```



Interaction Between a Continuous Predictor and a Dummy Moderator

Population Regression Model:

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 (X_{1i} \times X_{2i}) + \varepsilon_i \quad (10)$$

Model Specification:

- X_1 : Predictor (Age, measured in years)
- X_2 : Dummy-variable moderator (Gender, 0 = Women, 1 = Men)
- $E[Y_i]$: Expected number of hours spent working out in a month
- β_1 : Coefficient on age for women
- β_2 : Coefficient on gender when age is zero
- β_3 : Difference in slope of age between men and women

Interpretation:

- β_3 determines whether the effect of age on workout hours is stronger or weaker for men compared to women.
- Moving from $X_2 = 0$ (women) to $X_2 = 1$ (men), the slope coefficient on age changes by β_3 .

Conditional Effects:

$$E[Y_i] = \beta_0 + \beta_1 X_{1i} + \beta_2(0) + \beta_3(X_{1i} \times 0) \quad (11)$$

$$E[Y_i] = \beta_0 + \beta_1 X_{1i} \quad (12)$$

When $X_2 = 0$ (for women), the effect of age on workout hours is given by β_1 .

```
. *Estimate whether the effect of age on whours is different for women and men
. reg whours c.age i.gender c.age#i.gender
```

Source		SS	df	MS	Number of obs	=	210
-----+-----					F(3, 206)	=	8.49
Model		1202.50438	3	400.834794	Prob > F	=	0.0000
Residual		9725.41943	206	47.2107739	R-squared	=	0.1100
-----+-----					Adj R-squared	=	0.0971
Total		10927.9238	209	52.2867168	Root MSE	=	6.871

-----+-----						
whours		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
age		-.0317262	.0546461	-0.58	0.562	-.1394635 .0760111
gender						
men		10.61049	3.25435	3.26	0.001	4.19439 17.02659
gender#c.age						
men		-.2135282	.0789673	-2.70	0.007	-.3692158 -.0578406
_cons		13.09677	2.290885	5.72	0.000	8.580188 17.61336
-----+-----						

```
. *Estimate the margins
. margins, dydx(age) at(gender=(0 1))
```

```
Average marginal effects      Number of obs      =          210
Model VCE      : OLS
```

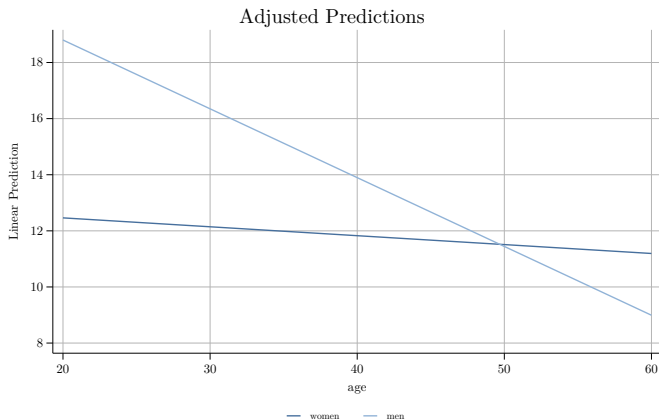
```
Expression      : Linear prediction, predict()
dy/dx w.r.t.    : age
```

```
1._at          : gender          =          0
```

```
2._at          : gender          =          1
```

		Delta-method				
		dy/dx	Std. Err.	t	P> t	[95% Conf. Interval]
age						
	_at					
	1	-.0317262	.0546461	-0.58	0.562	-.1394635 .0760111
	2	-.2452544	.0570056	-4.30	0.000	-.3576435 -.1328652

```
*Plot the difference between these two coefficients \beta_3  
margins, at(age=(20(10)60) gender=(0 1))  
marginsplot, noci x(age) recast(line)
```



Interaction Between a Dummy Predictor and a Dummy Moderator

Population Regression Model:

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 (X_{1i} \times X_{2i}) + \varepsilon_i \quad (13)$$

Model Specification:

- X_1 : Dummy predictor (Gender, 0 = Women, 1 = Men)
- X_2 : Dummy moderator (Marital Status, 0 = Married, 1 = Single)
- $E[Y_i]$: Expected number of hours spent working out in a month
- β_0 : Expected workout hours for married women
- β_1 : Effect of gender on workout hours for married individuals
- β_2 : Effect of marital status on workout hours for women
- β_3 : Difference in the effect of gender on workout hours between married and single individuals

Conditional Effects:

- When $X_2 = 0$ (Married):

$$E[Y_i] = \beta_0 + \beta_1 X_{1i} \quad (14)$$

- When $X_2 = 1$ (Single):

$$E[Y_i] = (\beta_0 + \beta_2) + (\beta_1 + \beta_3) X_{1i} \quad (15)$$

Interpretation:

- The interaction term β_3 determines whether the effect of gender on workout hours differs by marital status.
- A positive β_3 suggests that the effect of gender on workout hours is stronger for single individuals, while a negative β_3 suggests a weaker effect.

. *Estimate the interaction between a dummy predictor (gender) and a dummy moderator (marital status)

. reg whours i.gender i.marital i.gender#i.marital

Source		SS		df	MS	Number of obs	=
		210					
-----+-----						F(3, 206)	=
		4.55					
Model		678.978763		3	226.326254	Prob > F	=
		0.0041					
Residual		10248.945		206	49.7521604	R-squared	=
		0.0621					
-----+-----						Adj R-squared	=
		0.0485					
Total		10927.9238		209	52.2867168	Root MSE	=
		7.0535					

whours		Coef.	Std. Err.	t	P> t	[95% Conf.
						Interval]

```
. *Estimate the margins
. margins, dydx(gender) at(marital=(0 1))
```

```
Conditional marginal effects      Number of obs      =
      210
```

```
Model VCE      : OLS
```

```
Expression      : Linear prediction, predict()
```

```
dy/dx w.r.t.    : 1.gender
```

```
1._at           : marital      =          0
```

```
2._at           : marital      =          1
```

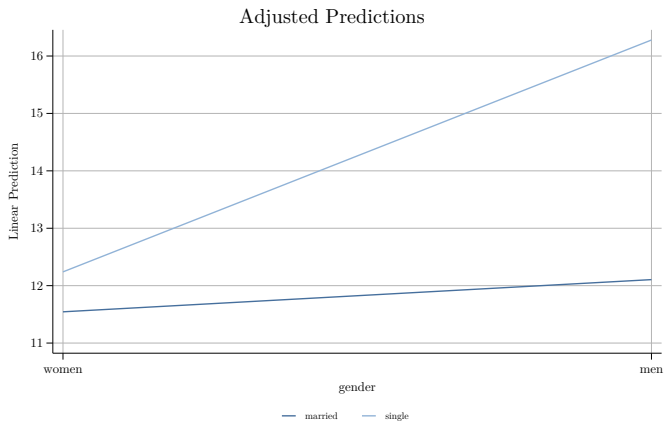
```
-----
```

			Delta-method				
		dy/dx	Std. Err.	t	P> t	[95% Conf.	
		Interval]					

```
-----+-----
```

```
0.gender        | (base outcome)
```

```
*Plot the difference between these two coefficients \beta_3  
margins, at(gender=(0 1) marital=(0 1))  
marginsplot, noci x(gender) recast(line)
```



Interaction Between a Continuous Predictor and a Polytomous Moderator

Population Regression Model:

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{3i} + \beta_4 (X_{1i} \times X_{2i}) + \beta_5 (X_{1i} \times X_{3i}) + \varepsilon_i \quad (16)$$

Model Specification:

- X_1 : Continuous predictor (Age)
- X_2 : First dummy-variable moderator (University education, 0 = No, 1 = Yes)
- X_3 : Second dummy-variable moderator (More than university education, 0 = No, 1 = Yes)
- $E[Y_i]$: Expected number of hours spent working out in a month
- β_0 : Expected workout hours for individuals with secondary/high school education when age is zero
- β_1 : Effect of age on workout hours for individuals with secondary/high school education
- β_2 : Difference in workout hours between university-educated individuals and those with secondary/high school education when age is zero
- β_3 : Difference in workout hours between individuals with more than university education and those with secondary/high school education when age is zero
- β_4 : Difference in the effect of age on workout hours between university-educated individuals and those with secondary/high school education
- β_5 : Difference in the effect of age on workout hours between individuals with more than university education and those with secondary/high school education

Conditional Effects:

- When $X_2 = 0$ and $X_3 = 0$ (Secondary/High School):

$$E[Y_i] = \beta_0 + \beta_1 X_{1i} \quad (17)$$

- When $X_2 = 1$ and $X_3 = 0$ (University):

$$E[Y_i] = (\beta_0 + \beta_2) + (\beta_1 + \beta_4) X_{1i} \quad (18)$$

- When $X_2 = 0$ and $X_3 = 1$ (More than University):

$$E[Y_i] = (\beta_0 + \beta_3) + (\beta_1 + \beta_5) X_{1i} \quad (19)$$

Interpretation:

- The interaction terms β_4 and β_5 determine whether the effect of age on workout hours differs by education level.
- A positive β_4 suggests that age has a stronger effect on workout hours for university-educated individuals compared to those with secondary/high school education.
- A positive β_5 suggests that age has a stronger effect on workout hours for individuals with more than university education compared to those with secondary/high school education.

```
. *Estimate whether the effect of age on whours is dependent on educational level
. reg whours c.age i.educ c.age#i.educ
```

Source	SS	df	MS	Number of obs	=	210
				F(5, 204)	=	6.27
Model	1455.73963	5	291.147926	Prob > F	=	0.0000
Residual	9472.18418	204	46.4322754	R-squared	=	0.1332
				Adj R-squared	=	0.1120
Total	10927.9238	209	52.2867168	Root MSE	=	6.8141

	whours	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
	age	-.2190014	.0579104	-3.78	0.000	-.3331811	-.1048217
	educ						
	university	-10.46198	4.363082	-2.40	0.017	-19.06449	-1.859459
more than	university	-13.36174	4.018536	-3.33	0.001	-21.28493	-5.438548
	educ#c.age						
	university	.2014475	.102947	1.96	0.052	-.001529	.404424
more than	university	.2564962	.0985739	2.60	0.010	.062142	.4508504
	_cons	23.00372	2.148345	10.71	0.000	18.76791	27.23953

```
. *Estimate the margins
. margins, dydx(age) at(educ=(1 2 3))
```

```
Average marginal effects      Number of obs      =      210
Model VCE      : OLS
```

```
Expression      : Linear prediction, predict()
dy/dx w.r.t.    : age
```

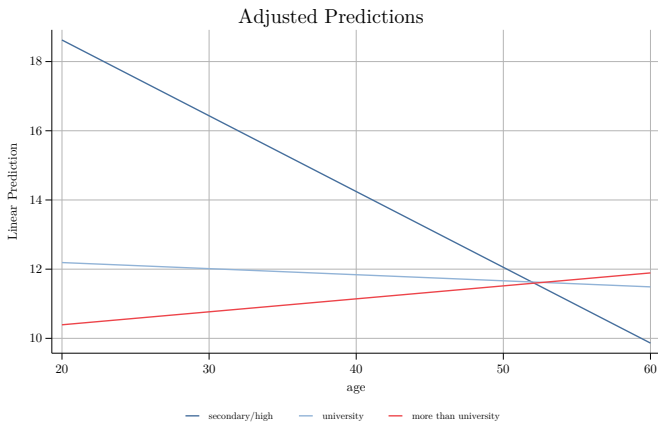
```
1._at      : educ      =      1
2._at      : educ      =      2
3._at      : educ      =      3
```

		Delta-method				
		dy/dx	Std. Err.	t	P> t	[95% Conf. Interval]
age						
	_at					
	1	-.2190014	.0579104	-3.78	0.000	-.3331811 -.1048217
	2	-.0175539	.0851144	-0.21	0.837	-.1853707 .1502629
	3	.0374948	.0797696	0.47	0.639	-.1197838 .1947734

`*Plot the difference between these three coefficients`

`margins, at(age=(20(10)60) educ=(1 2 3))`

`marginsplot, noci x(age) recast(line)`



Significant versus Non-Significant Interaction

Consideration of Interaction Terms:

- Up to now, we have assumed that interaction terms are statistically significant.
- However, what should be done if an interaction term is statistically non-significant?

Best Practices in Model Specification:

- If an interaction was hypothesized a priori (before data collection), it should remain in the model even if non-significant.
- In social sciences, particularly in non-experimental research, interactions are often examined a posteriori (after data collection).
- In such cases, exclusion of non-significant interactions is recommended to ensure a parsimonious and less complex model.

Focus on Interpretation:

- Interaction models emphasize the interpretation of interaction terms and simple (conditional) effects.
- Main effects of variables involved in interaction terms should not be the primary focus.

Why Consider Transformations?

- So far, we have worked with raw (untransformed) data to understand interactions.
- In an interaction model, coefficients reflect slopes when the moderator is zero.
- However, if the moderator does not have zero in its scale, interpretation becomes difficult.

How Centring Works:

- To improve interpretability, the moderator/predictor variable can be centred.
- The most common approach is centring at the mean: subtracting the mean value from each observation.
- With centred data, coefficients reflect the slope at the mean value of the moderator.
- The coefficient on the interaction term remains unchanged by centring.

How Standardization Works:

- Another transformation is the z-score standardization.
- This transformation subtracts the mean and divides by the standard deviation.
- Similar to centring, standardized coefficients reflect slopes at the mean.
- Interpretation shifts from raw units to standard deviations.

Choosing Between Raw, Centred, and Standardized Data

General Recommendation:

- Unless there is a specific reason, we recommend working with raw data.
- This approach maintains interpretability in the original metric.
- Raw data also provides flexibility in using prediction equations.
- When using categorical variables, centring or standardization is usually unnecessary.

- Laffers, L. (2021). *Draft poznámok k predmetu Moderná Aplikovaná regresia 1*. UMB Banská Bystrica.
- Mehmetoglu, M. and Jakobsen, T. G. (2022). *Applied Statistics using Stata: a Guide for the Social Sciences*. Sage.