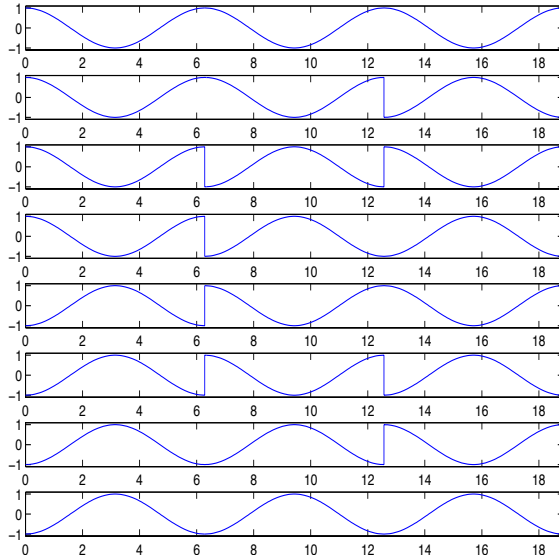


Week 3

Random processes

10.2.3 Each of the sample functions should encode one of the sequences $\{(0, 0, 0), (1, 0, 0), \dots, (1, 1, 1)\}$.



(For these pictures, I have chosen: $T = 2\pi$, and $f_0 = \frac{1}{T}$. But feel free to choose other values for T and f_0 if you like.)

10.10.3 To find the autocorrelation function of the random process $W(t) = X \cos(2\pi f_0 t) + Y \sin(2\pi f_0 t)$ (with uncorrelated RV's X and Y) we have to compute:

$$\begin{aligned}
 R_W(t, \tau) &= E[W(t)W(t + \tau)] & (3.1) \\
 &= E[(X \cos(2\pi f_0 t) + Y \sin(2\pi f_0 t))(X \cos(2\pi f_0(t + \tau)) + Y \sin(2\pi f_0(t + \tau)))] \\
 &= E[X \cos(2\pi f_0 t)X \cos(2\pi f_0(t + \tau)) + X \cos(2\pi f_0 t)Y \sin(2\pi f_0(t + \tau)) + \\
 &\quad Y \sin(2\pi f_0 t)X \cos(2\pi f_0(t + \tau)) + Y \sin(2\pi f_0 t)Y \sin(2\pi f_0(t + \tau))] \\
 &= E[X^2] \cos(2\pi f_0 t) \cos(2\pi f_0(t + \tau)) + E[XY] \cos(2\pi f_0 t) \sin(2\pi f_0(t + \tau)) + \\
 &\quad E[XY] \sin(2\pi f_0 t) \cos(2\pi f_0(t + \tau)) + E[Y^2] \sin(2\pi f_0 t) \sin(2\pi f_0(t + \tau))] & (3.2)
 \end{aligned}$$

Because X and Y are uncorrelated $Cov[X, Y] = E[XY] - E[X]E[Y] = 0$, and because the expected value $E[X] = E[Y] = 0$, we know $E[XY] = 0$. Furthermore is given that

$\text{Var}[X] = E[X^2] - E[X]^2 = \sigma^2$, and therefore $E[X^2] = E[Y^2] = \sigma^2$. Combining gives:

$$R_W(t, \tau) = \sigma^2 (\cos(2\pi f_0 t) \cos(2\pi f_0(t + \tau)) + \sin(2\pi f_0 t) \sin(2\pi f_0(t + \tau)))$$

From Math Fact B.2 we know:

$$\begin{aligned} \cos(A) \cos(B) &= \frac{1}{2} [\cos(A - B) + \cos(A + B)] \\ \sin(A) \sin(B) &= \frac{1}{2} [\cos(A - B) - \cos(A + B)] \end{aligned}$$

Substituting $A = 2\pi f_0 t$ and $B = 2\pi f_0(t + \tau)$, rewriting, and noticing that the terms $\cos(A + B)$ cancel, we obtain:

$$R_W(t, \tau) = \sigma^2 \cos(-2\pi f_0 \tau) = \sigma^2 \cos(2\pi f_0 \tau) \quad (3.3)$$

and we see that the autocovariance function is *independent* on t .

10.4.1 For Y_k to be iid, it should have identical distribution for different k , and it should be independent. Each Y_k is the sum of two identical independent Gaussian random variables, so each Y_k has the same pdf. Next, Y_k is independent of Y_l when $l \neq k$ because they do not share any samples of X_k .

10.4.2 Again, each W_k is the sum of two identical independent Gaussian random variables, so they have the same pdf. But variables W_k and W_{k-1} share the sample X_{k-1} , so W_k and W_{k-1} are *not* independent.

4.11.1 Looking at the joint pdf:

$$F_{X,Y}(x, y) = ce^{-x^2/8 - y^2/18} \quad (3.4)$$

it looks suspiciously like a Gaussian distribution. Looking at Definition 4.17, pg 191 of the book, we see that this definition contains μ_X , μ_Y , σ_1 , σ_2 and ρ .

When we can identify σ_1 , σ_2 and ρ , the constant c can be computed as:

$$c = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \quad (3.5)$$

To find σ_1 and $E[X] = \mu_1$, we have to solve:

$$\frac{1}{2} \left(\frac{x - E[X]}{\sigma_1} \right)^2 = \frac{x^2}{8} \rightarrow E[X] = 0, \quad \text{and} \quad \sigma_1 = \sqrt{4} = 2 \quad (3.6)$$

$$\frac{1}{2} \left(\frac{y - E[Y]}{\sigma_2} \right)^2 = \frac{y^2}{18} \rightarrow E[Y] = 0, \quad \text{and} \quad \sigma_2 = \sqrt{9} = 3 \quad (3.7)$$

Because there is no cross term with $x \cdot y$, we have to conclude that $\rho = 0$. Solving c gives:

$$c = \frac{1}{2\pi\sqrt{8}\sqrt{18}\sqrt{1-0}} = \frac{1}{12\pi} \quad (3.8)$$

And because $\rho = 0$ the variables X and Y are uncorrelated, which means for a Gaussian that they are also independent.

4.11.4 (a) When the two random variables X and Y are iid continuous uniform between -50 and 50, it can actually happen that the archer misses the circular target completely! (For instance, when $x = 49$ and $y = 49$.) Because X and Y are independent, we can easily give the joint pdf:

$$f_{X,Y}(x, y) = f_X(x)f_Y(y) = \begin{cases} 10^{-4} & -50 \leq x \leq 50, -50 \leq y \leq 50 \\ 0 & \text{otherwise} \end{cases} \quad (3.9)$$

(both X and Y are uniformly distributed between -50 and 50). Therefore the probability of bullseye is:

$$P[A] = P[X^2 + Y^2 \leq 2^2] = 10^{-4} \cdot \pi 2^2 = 0.0013 \quad (3.10)$$

(b) When $f_{X,Y}(x, y)$ is uniform over the circular area (and X and Y are not independent anymore!), the density becomes the inverse of the area:

$$f_{X,Y}(x, y) = \begin{cases} \frac{1}{\pi 50^2} & x^2 + y^2 \leq 50^2 \\ 0 & \text{otherwise} \end{cases} \quad (3.11)$$

Then the probability of bullseye becomes:

$$P[A] = P[X^2 + Y^2 \leq 2^2] = \frac{\pi 2^2}{\pi 50^2} = 0.0016 \quad (3.12)$$

(c) When X and Y are independent Gaussian distributions with mean 0 and variance σ^2 , we have

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-x^2/2\sigma^2}, \quad f_Y(y) = \frac{1}{\sqrt{2\pi}\sigma} e^{-y^2/2\sigma^2}, \quad (3.13)$$

and the joint probability density becomes:

$$f_{X,Y}(x, y) = \frac{1}{2\pi\sigma^2} e^{-(x^2+y^2)/2\sigma^2}, \quad (3.14)$$

For the probability of bullseye we have to compute an interesting integral:

$$P[A] = P[X^2 + Y^2 \leq 2^2] \quad (3.15)$$

$$= \int_{x^2+y^2 \leq 2^2} f_{X,Y}(x, y) dx dy \quad (3.16)$$

$$= \frac{1}{2\pi\sigma^2} \int_{x^2+y^2 \leq 2^2} e^{-(x^2+y^2)/2\sigma^2} dx dy \quad (3.17)$$

Now we have to do a trick, a coordinate transform in polar coordinates: $r^2 = x^2 + y^2$ and $dx dy = r dr d\theta$, and we integrate:

$$P[A] = \frac{1}{2\pi\sigma} \int_0^2 \int_0^{2\pi} e^{-r^2/2\sigma^2} r dr d\theta \quad (3.18)$$

$$= \frac{1}{\sigma^2} \int_0^2 r e^{-r^2/2\sigma^2} dr \quad (3.19)$$

$$= \left[-e^{-r^2/2\sigma^2} \right]_0^2 = 1 - e^{-4/200} = 0.020 \quad (3.20)$$

10.5.1 The arrivals of new telephone calls can be modelled by a Poisson process. The rate $\lambda = 4$ is given, and therefore our PRM is defined:

$$P_{N(T)}(n) = \begin{cases} (4T)^n e^{-4T} / n! & n = 0, 1, 2, 3, \dots \\ 0 & \text{otherwise} \end{cases} \quad (3.21)$$

(a) Now we only have to fill in:

$$P_{N(1)}(0) = (4)^0 e^{-4} / 0! = e^{-4} \quad (3.22)$$

(b) and (c) Similarly:

$$P_{N(1)}(4) = (4)^4 e^{-4} / 4! = 10.67 e^{-4} \quad (3.23)$$

$$P_{N(2)}(2) = (8)^2 e^{-8} / 2! = 32 e^{-8} \quad (3.24)$$

10.5.6 It is given that the response time T is an exponential random variable with mean 8. That means that the pdf is:

$$f_T(t) = \begin{cases} \frac{1}{8}e^{-t/8} & t \geq 0 \\ 0 & \text{otherwise} \end{cases} \quad (3.25)$$

(a) The probability of a response time larger than 4:

$$P[T \geq 4] = 1 - P[T < 4] = 1 - \int_{-\infty}^4 f_T(t) dt \quad (3.26)$$

$$= 1 - \int_0^4 \frac{1}{8}e^{-t/8} dt = 1 + \left[e^{-t/8} \right]_0^4 \quad (3.27)$$

$$= 1 + e^{-4/8} - 1 = e^{-1/2} \quad (3.28)$$

(b) The conditional probability is (because when $T \geq 13$ then T is also always larger than 5):

$$P[T \geq 13 | T \geq 5] = \frac{P[T \geq 13, T \geq 5]}{P[T \geq 5]} = \frac{P[T \geq 13]}{P[T \geq 5]} \quad (3.29)$$

Now we can compute:

$$P[T \geq 13] = 1 - \int_0^{13} \frac{1}{8}e^{-t/8} dt = e^{-13/8} \quad (3.30)$$

$$P[T \geq 5] = 1 - \int_0^5 \frac{1}{8}e^{-t/8} dt = e^{-5/8} \quad (3.31)$$

$$(3.32)$$

so therefore:

$$P[T \geq 13 | T \geq 5] = \frac{e^{-13/8}}{e^{-5/8}} = e^{-1} \quad (3.33)$$

(c) This seems simple: we have a sequence of arrivals, and their interarrival time is exponential. So the $N(t)$ should be a Poisson process:

$$P[N(t) = n] = \begin{cases} \left(\frac{t}{8}\right)^n e^{-t/8} / n! & n = 0, 1, 2, \dots \\ 0 & \text{otherwise} \end{cases} \quad (3.34)$$

But there is one tricky detail here: it is given that the first query is made at time zero. So the Poisson process does not start with $N(t) = 0$ counts, but with $N(t) = 1$ counts. We have to shift the Poisson distribution by one count:

$$P[N(t) = n] = \begin{cases} \left(\frac{t}{8}\right)^{n-1} e^{-t/8} / (n-1)! & n = 1, 2, \dots \\ 0 & \text{otherwise} \end{cases} \quad (3.35)$$

Note that for $t > 0$ we will always have that $N(t) \geq 1$.