# A hierarchical Bayesian model of individual differences in the size-congruity effect

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MCLS Symposium: Current directions in symbolic number processing

# Size congruity effect

Typical laboratory task: choose the physically larger digit

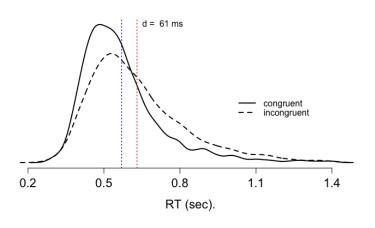
Congruent Incongruent

2 8

2 8

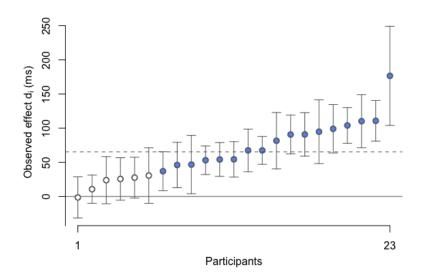
# Size congruity effect

Typical result – mean RT larger for incongruent trials



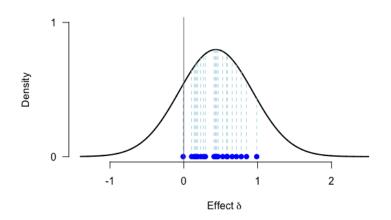
# Size congruity effect

What about individual effects?



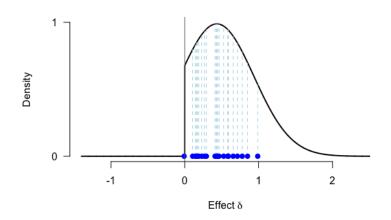
#### Individual differences

Suppose these observed effects  $d_i$  are drawn from population of *true* effects  $\delta$ . What is the structure of this population?



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## A new question

**Does everybody** exhibit the size congruity effect?

- if yes, then SCE is obligatory, resistant to strategic control, ...
- if *no*, then SCE is complex, malleable, ...

Importantly, both answers have downstream consequences for processing architecture of numerical cognition

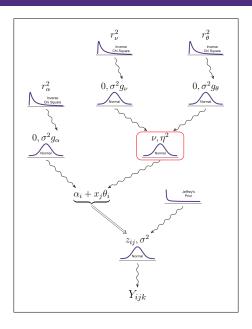
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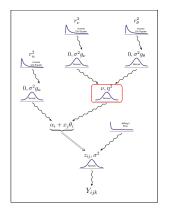
How to answer:

- build models of individual difference structures in SCE (e.g., Haaf & Rouder, 2017)
- adjudicate the models via Bayesian model comparison

## Hierarchical structure



#### Hierarchical structure



Basic idea (Haaf & Rouder, 2017):

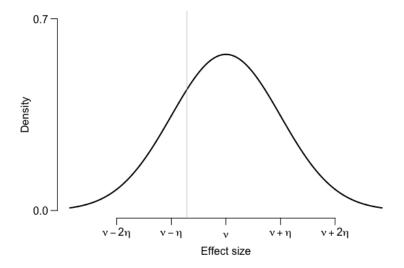
- model RTs as a random-effects linear model with effect parameter  $\theta_i$   $(i = 1 \dots, N)$
- assume (as baseline) that  $\theta_i$  is drawn from a normal distribution with mean  $\nu$  and variance  $\eta^2$
- use g-priors (Zellner & Siow, 1980), specifying a priori scale on variance of overall effect and individual variability around effect
- define competing models by constraining effect parameter θ<sub>i</sub>

# Four competing models

- 1. Unrestricted model,  $\mathcal{M}_u$
- 2. Positive-effects model,  $\mathcal{M}_+$
- 3. Common-effect model,  $\mathcal{M}_1$
- 4. Null-effect model,  $\mathcal{M}_0$

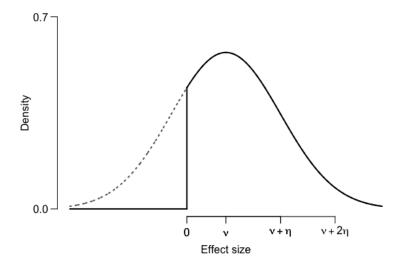
## **Unrestricted model**

 $\mathcal{M}_u: \theta_i \sim \mathsf{Normal}(\nu, \eta^2)$ 



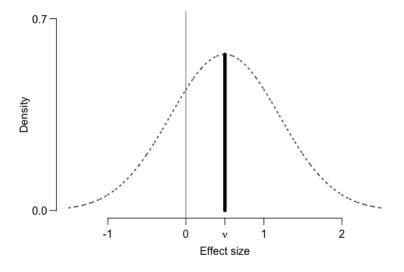
## Positive-effects model

•  $\mathcal{M}_+$ :  $\theta_i \sim \text{Normal}_+(\nu, \eta^2)$ 



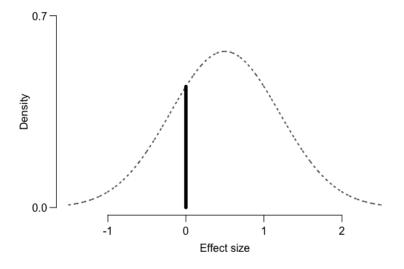
## Common-effect model

•  $\mathcal{M}_1: \theta_i = \nu$ 



## Null-effect model

•  $\mathcal{M}_0$  :  $\theta_i = 0$ 



## Model comparison

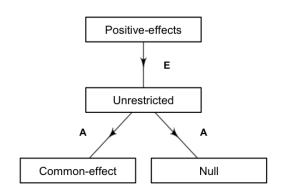
We use the **Bayes factor**, which indexes how well the observed data **Y** are predicted under one model relative to another:

$$B_{ab} = \frac{p(\mathbf{Y} \mid \mathcal{M}_a)}{p(\mathbf{Y} \mid \mathcal{M}_b)}$$

Ex: suppose  $B_{ab} = 10$ . This means:

- ullet the observed data are 10 times more likely under  $\mathcal{M}_a$  compared to  $\mathcal{M}_b$
- "10-to-1 evidence for  $\mathcal{M}_a$  over  $\mathcal{M}_b$ "

So 
$$B_{ab} = \frac{p(\mathbf{Y} \mid \mathcal{M}_a)}{p(\mathbf{Y} \mid \mathcal{M}_b)}$$
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$$p(\mathbf{Y} \mid \mathcal{M}) = \int_{\boldsymbol{\xi} \in \Xi} p(\mathbf{Y} \mid \boldsymbol{\xi}, \mathcal{M}) p(\boldsymbol{\xi} \mid \mathcal{M}) d\boldsymbol{\xi}$$

So 
$$B_{ab} = \frac{p(\mathbf{Y} \mid \mathcal{M}_a)}{p(\mathbf{Y} \mid \mathcal{M}_b)}$$
. How do we compute this?

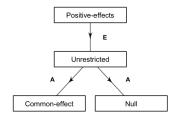
$$p(\mathbf{Y} \mid \mathcal{M}) = \int_{\boldsymbol{\xi} \in \Xi} p(\mathbf{Y} \mid \boldsymbol{\xi}, \mathcal{M}) p(\boldsymbol{\xi} \mid \mathcal{M}) d\boldsymbol{\xi}$$

Problem: for our models  $\mathcal{M}$ , the parameter vectors  $\boldsymbol{\xi}$  look like

$$\boldsymbol{\xi} = (\mu, \sigma^2, \nu, \alpha_1, \dots, \alpha_N, \theta_1, \dots, \theta_N, g_\alpha, g_\nu, g_\theta)$$

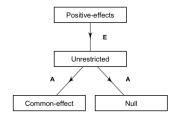
so the integral is carried out in  $\mathbb{R}^{2N+6}$ .

For N = 35, this would be a 76-dimensional integral!



#### A = analytic approach

- Zellner & Siow (1980); Rouder et al. (2012)
- place g-priors on individual intercepts and effect parameters
- everything except the g-parameters integrates symbolically
- g-parameters can be well approximated with MCMC sampling
- techniques coded into BayesFactor package in R



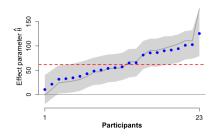
#### E = encompassing approach

- Klugkist et al. (2005)
- generalization of Savage-Dickey density ratio

• 
$$B_{+u} = \frac{P(\theta > 0 \mid \mathbf{Y}, \mathcal{M}_u)}{P(\theta > 0 \mid \mathcal{M}_u)}$$

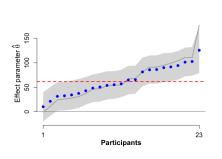
 probabilities computed as fraction of MCMC samples from unrestricted model that are **positive** for all individuals (both in the prior and posteriori)

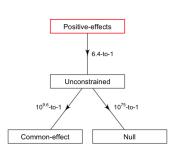
# Results - Exp 1



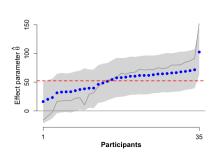
- Red line = estimated effect  $\theta$  from  $\mathcal{M}_1$
- Blue dots = individual effect estimates  $\theta_i$
- Gray line = estimates from mean differences d<sub>i</sub>
- Gray area = 95% credible intervals

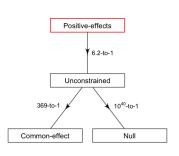
# Results - Exp 1





# Results - Exp 2





# Sensitivity to prior specifications

#### Experiment 1:

$r_{ u}$	$r_{\theta}$	$\mathcal{M}_0$	$\mathcal{M}_1$	$\mathcal{M}_+$	$\mathcal{M}_{\it u}$
$\frac{1}{6}$ (50 ms)	$\frac{1}{10}$ (30 ms)	5.6e-77	4.2e-11	*	0.16
$\frac{1}{12}$ (25 ms)	$\frac{1}{20}$ (15 ms)	1.8e-76	7.7e-11	*	0.16
$\frac{1}{12}$ (25 ms)	$\frac{1}{5}$ (60 ms)	1.9e-77	8.1e-12	*	0.05
$\frac{1}{3}$ (100 ms)	$\frac{1}{20}$ (15 ms)	1.9e-76	1.9e-10	*	0.39
$\frac{1}{3}$ (100 ms)	$\frac{1}{5}$ (60 ms)	3.0e-77	3.1e-11	*	0.16

Note: Bayes factors computed against the "winning" model, denoted by  ${}^{*}$ 

# Sensitivity to prior specifications

#### Experiment 2:

$r_{ u}$	$r_{\theta}$	$\mathcal{M}_0$	$\mathcal{M}_1$	$\mathcal{M}_+$	$\mathcal{M}_{\it u}$
$\frac{1}{6}$ (50 ms)	$\frac{1}{10}$ (30 ms)	4.3e-41	0.0004	*	0.17
$\frac{1}{12}$ (25 ms)	$\frac{1}{20}$ (15 ms)	1.2e-40	0.0007	*	0.16
$\frac{1}{12}$ (25 ms)	$\frac{1}{5}$ (60 ms)	2.6e-41	0.0002	*	0.06
$\frac{1}{3}$ (100 ms)	$\frac{1}{20}$ (15 ms)	1.4e-40	0.0017	*	0.42
$\frac{1}{3}$ (100 ms)	$\frac{1}{5}$ (60 ms)	4.1e-41	0.0005	*	0.19

Note: Bayes factors computed against the "winning" model, denoted by \*

## **Summary points**

**Does everybody** exhibit the size congruity effect?

- if yes, then SCE is obligatory, resistant to strategic control, ...
- if *no*, then SCE is complex, malleable, ...

Importantly, both answers have downstream consequences for processing architecture of numerical cognition

 what does this say about early vs. late interaction debate (e.g., Faulkenberry et al., 2016; Sobel et al., 2016; 2017; Faulkenberry, Vick, & Bowman, in press)

# **Summary points**

#### Some other benefits:

- Bayes factors easy to interpret
- hierarchical structure removes trial noise from individual estimates
- common effect (CE) model provides important self-check:
  - if CE model is best, is our design adequate to capture individual differences
- Might be good approach to disentangle competing theories of mental arithmetic
  - Does everyone exhibit size-by-format interaction?
  - Does everyone reflect fast counting in small addition problems?

# Thank you!

- Thanks to Tarleton Office of Research and Innovation for funding!
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