Computational Statistics II

Unit C.1: Missing data problems

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Unit C.1

Main concepts

- Missing data problems;
- Data augmentation and Gibbs sampling;
- Connection with the EM algorithm.

Missing data problems

- In this unit we will take advantage of specific structures of the model to facilitate Bayesian computations.
- However, in most cases, this will involve the introduction of hidden features of the model, sometimes called latent variables, which might be interesting per se.
- Depending on the context, these latent quantities will have a precise meaning or they will be regarded as purely abstract objects.
- An obvious examples of latent components with a precise interpretation is the case of missing or censored observations.
- Key idea. If the complete data were available, computations would be easier. Besides, imputing the missing values could be interesting on its own.

Example: survival analysis with an exponential model

- Let $\mathbf{z} = (z_1, \dots, z_n)^\mathsf{T}$ be iid exponential random variables with rate parameter θ .
- If the prior $\theta \sim Ga(a, b)$, then thanks to conjugacy we get the following posterior

$$(\theta \mid \mathbf{z}) \sim \mathsf{Ga}\left(\mathbf{a} + \mathbf{n}, \mathbf{b} + \sum_{i=1}^{n} z_{i}\right).$$

- However, in many cases observations are censored, as in **Unit A.1**. In fact, we observe the values $t = (t_1, ..., t_n)^T$ which are either complete $(t_i = z_i)$ or censored $(t_i \le z_i)$.
- If the observations were all complete, then inference would be straightforward.
- Intuitively, what we are going to do is to sample the missing information from the corresponding conditional distribution in order to make inference about θ .

Data augmentation

- Let X be the observed data, following some distribution $\pi(X \mid \theta)$, i.e. the likelihood, with $\theta \in \Theta \subseteq \mathbb{R}^p$ being an unknown set of parameters.
- Let $\pi(\theta)$ be the prior distribution associated to θ and let $\pi(\theta \mid X)$ be the posterior.
- Let $z \in \mathcal{Z} \subseteq \mathbb{R}^q$ be a vector of latent variables.
- We assume that the likelihood function $\pi(X \mid \theta)$ can be written as the marginal distribution of a complete likelihood, namely

$$\pi(\mathbf{X} \mid \mathbf{\theta}) = \int_{\mathcal{Z}} \pi(\mathbf{X}, \mathbf{z} \mid \mathbf{\theta}) d\mathbf{z}.$$

■ Remark. We focus on continuous density w.r.t. the Lebesgue measure for the sake of notational simplicity, but the same idea applies more generally.

Data augmentation

- The quantity $\pi(X, z \mid \theta)$ is the complete or augmented likelihood.
- Within the Bayesian framework, we should treat the latent variables z as if they were an additional set of unknown parameters, leading to the augmented posterior

$$\pi(\theta, z \mid X) \propto \pi(X, z \mid \theta)\pi(\theta).$$

- In other words, we aim at sampling $(\theta^{(r)}, \mathbf{z}^{(r)})$ using MCMC from the joint posterior $\pi(\theta, \mathbf{z} \mid \mathbf{X})$, which can be performed using any of the strategies we have described.
- If one is interested only in the original parameters θ or in the latent dimensions z, then it suffices to ignore the other set of parameters.
- If the interest is in θ , the advantage of sampling from $\pi(\theta, \mathbf{z} \mid \mathbf{X})$ and then discarding \mathbf{z} rather than directly targeting $\pi(\theta \mid \mathbf{X})$ is that the augmented likelihood is typically more tractable than the original one.

Data augmentation and Gibbs sampling

- Although in principle any MCMC strategy could be used to target $\pi(\theta, \mathbf{z} \mid \mathbf{X})$, the Gibbs sampling is often a natural choice.
- In fact, it is often the case that the following full conditional distributions are available in closed form. Moreover, they also have a nice interpretation.
- **Step 1**. Sample from the "posterior" of θ based on the complete likelihood, namely

$$\pi(\theta \mid X, z) \propto \pi(X, z \mid \theta)\pi(\theta).$$

■ Step 2. Impute the missing observations z by sampling from the full conditional

$$\pi(z \mid X, \theta) \propto \pi(X, z \mid \theta).$$

lacktriangledown Obviously, we are allowed to split $m{ heta}$ and $m{z}$ into blocks of parameters if this facilitate the Gibbs sampling.

Example: survival analysis with an exponential model

Recall the exponential model example with censored data t and censorship indicators $d = (d_1, \dots, d_n)^\mathsf{T}$. The original likelihood is therefore equal to

$$\pi(\boldsymbol{t}, \boldsymbol{d} \mid \theta) = \theta^{n_c} \exp \left\{ -\theta \sum_{i=1}^n t_i \right\}, \qquad n_c = \sum_{i=1}^n d_i.$$

- **Remark**. This is a toy example whose purpose is fixing ideas. Indeed, under a Gamma prior, the posterior distribution of θ using this likelihood is also available.
- In this setting, the latent variables z represent the complete survival times having exponential distribution, so that the complete likelihood is

$$\pi(\mathbf{z} \mid \theta) = \theta^n \exp \left\{ -\theta \sum_{i=1}^n z_i \right\}.$$

■ The Gibbs sampling therefore alternates the full conditional $\pi(\theta \mid z)$ and an imputation sampling step from $\pi(z \mid t, \theta)$. As for the latter, note that $(z_i - t_i \mid t_i, \theta) \stackrel{\text{ind}}{\sim} \mathsf{Exp}(\theta)$.

Connection with the EM algorithm

- A Gibbs sampling based on data augmentation strategies is strongly connected with the so-called expectation-maximization (EM) algorithm.
- The EM is a deterministic algorithm that aims at maximizing the likelihood (MLE) or the posterior distribution (MAP).
- \blacksquare Hence, as opposed to sampling strategies, the $\rm EM$ is widely used both within the frequentist and the Bayesian framework.
- Compared to other gradient-based maximizers, it leads to a monotonic sequence ⇒ the function always increases during the procedure.
- On the other hand, it requires a (tractable) augmented likelihood.

The EM algorithm (recap)

- The EM algorithm alternates between the following steps, which are reminiscent of those of the Gibbs sampling, as they involve similar quantities.
- Step 1 (Expectation). Let $\theta^{(r)}$ be the current value of the maximization procedure, then obtain the function

$$Q(\boldsymbol{\theta} \mid \boldsymbol{\theta}^{(r)}) = \log \pi(\boldsymbol{\theta}) + \mathbb{E}\{\log \pi(\boldsymbol{X}, \boldsymbol{z} \mid \boldsymbol{\theta}^{(r)})\},\$$

where the expectation is taken with respect to the conditional distribution $\pi(\mathbf{z} \mid \mathbf{X}, \boldsymbol{\theta})$.

Step 2 (Maximization). The new value of the procedure $\theta^{(r)}$ is obtained by maximizing the function

$$\arg\max_{oldsymbol{ heta}\in\Theta}\mathcal{Q}(oldsymbol{ heta}\midoldsymbol{ heta}^{(r)}).$$

■ In many practical cases, the E-step amounts at calculating $\mathbb{E}(z)$ and then plugging-in in the augmented log-likelihood. Indeed, $\log \pi(\mathbf{X}, z \mid \theta^{(r)})$ is often linear in z.

Data augmentation schemes

- Differently from other MCMC methods, there is no general recipe for finding useful data augmentation schemes.
- In principle, whenever the likelihood can be expressed in an integral form, this leads to a potential data augmentation mechanism.
- However, the resulting augmented likelihood has to be tractable (in some sense), otherwise the whole procedure is pointless.
- A crucial class of models that greatly benefit from data-augmentation schemes are mixture models, which are because they indeed deserve an entire course on their own.