

Exercise E

Setup: $E[y_i] = \mu_i = g^{-1}(\beta x_i)$, with $\rho = 1$.

Case $y_i \sim \text{ED}(\mu_i, \phi)$. This is a Gaussian distribution, with $y_i \sim N(\beta x_i, \phi)$.
known. *Common link is the identity*

$$e(\beta) = -\frac{1}{2\phi} \sum_{i=1}^m (y_i - \beta x_i)^2$$

$$e^*(\beta) = \frac{\partial}{\partial \beta} e(\beta) = \frac{1}{\phi} \sum_{i=1}^m x_i (y_i - \beta x_i) \Rightarrow \hat{\beta} = \frac{\sum_{i=1}^m x_i y_i}{\sum_{i=1}^m x_i^2}$$

$$I(\beta) = \text{var}(e^*(\beta)) = \frac{1}{\phi^2} \sum_{i=1}^m x_i^2 \text{var}(y_i - \beta x_i) = \frac{\cancel{\phi}}{\phi^2} \sum_{i=1}^m x_i^2 = \frac{m}{\phi} \bar{x}_2.$$

They coincide $= -e^{**}(\beta)$ *"* ϕ

Because of the properties of the Gaussian distribution,

$\hat{\beta} = \frac{1}{m \bar{x}_2} \cdot \sum_{i=1}^m x_i y_i$ is a Gaussian (linear combination of Gaussians), with

mean and variance

$$E[\hat{\beta}] = \frac{1}{m \bar{x}_2} \sum_{i=1}^m x_i E[y_i] = \beta \frac{m \bar{x}_2}{m \bar{x}_2} = \beta$$

" βx_i

$$\text{var}(\hat{\beta}) = \frac{1}{m^2 \bar{x}_2^2} \sum_{i=1}^m x_i^2 \text{var}(y_i) = \phi \frac{\cancel{m \bar{x}_2}}{m \bar{x}_2^2} = \frac{\phi}{m \bar{x}_2} = I^{-1}$$

$$\Rightarrow \hat{\beta} \sim N\left(\beta, \frac{\phi}{m \bar{x}_2}\right), \text{ and } \hat{\mu}_i = \hat{\beta} x_i \sim N\left(\beta x_i, \frac{\phi x_i^2}{m \bar{x}_2}\right)$$

C.I. for μ_i is $\hat{\mu}_i \pm z_{1-\alpha/2} \sqrt{\frac{\phi x_i^2}{m \bar{x}_2}}$
this is known, so the interval is exact.

Case $y_i \sim \text{ED}(\mu_i, \mu_i)$. This is a Poisson, so that $y_i \sim \text{Poisson}(\mu_i)$, $\mu_i = e^{\beta x_i} = \alpha^{x_i}$.

$$e(\beta) = \sum_{i=1}^m y_i \beta x_i - e^{\beta x_i}$$

$$e^*(\beta) = \sum_{i=1}^m x_i y_i - x_i e^{\beta x_i} \Rightarrow \hat{\beta} \text{ solves } \sum_{i=1}^m x_i y_i = \sum_{i=1}^m x_i e^{\hat{\beta} x_i}$$

$$I(\beta) = \text{var}(e^*(\beta)) = \sum_{i=1}^m x_i^2 \text{var}(y_i) = \sum_{i=1}^m x_i^2 e^{\beta x_i} = x^T W x.$$

$$\hat{\beta} \underset{\substack{\uparrow \\ \text{asymptotically}}}{\sim} N(\beta, (\sum_{i=1}^m x_i^2 e^{\beta x_i})^{-1})$$

$$\Rightarrow \hat{\eta}_i = \hat{\beta} x_i \sim N\left(\beta x_i, \frac{x_i^2}{\sum_{i=1}^m x_i^2 e^{\beta x_i}}\right) \text{ therefore the confidence interval is}$$

$$CI = \hat{\eta}_i \pm z_{1-\alpha/2} \sqrt{\frac{x_i^2}{\sum_{i=1}^m x_i^2 e^{\beta x_i}}} \quad (\text{Approximate CI})$$

The CI for μ_i can be simply obtained by exponentiating the above terms, obtaining

$$\exp\left(\hat{\eta}_i \pm z_{1-\alpha/2} \sqrt{\frac{x_i^2}{\sum_{i=1}^m x_i^2 e^{\beta x_i}}}\right) \quad (\text{Approximate CI}).$$

This is not Wald.

Case $y_i \sim \text{ED}(\mu_i, \mu_i(1-\mu_i))$. This is a binary variable $y_i \sim \text{Ber}(\mu_i)$, with

$$\text{logit}(\mu_i) = \log\left(\frac{\mu_i}{1-\mu_i}\right) = \beta x_i.$$

$$e(\beta) = \sum_{i=1}^m y_i \beta x_i - \log(1 + e^{\beta x_i})$$

$$e^*(\beta) = \sum_{i=1}^m x_i y_i - x_i \frac{e^{\beta x_i}}{1 + e^{\beta x_i}} \Rightarrow \hat{\beta} \text{ solves } \sum_{i=1}^m x_i y_i = \sum_{i=1}^m x_i \frac{e^{\hat{\beta} x_i}}{1 + e^{\hat{\beta} x_i}}.$$

$$I(\beta) = \sum_{i=1}^m x_i^2 \frac{e^{\beta x_i}}{(1 + e^{\beta x_i})^2} = x^T W x$$

$$\hat{\beta} \underset{\text{asymptotically}}{\sim} N\left(\beta, \left(\sum_{i=1}^n x_i^2 \frac{e^{\beta x_i}}{(1+e^{\beta x_i})^2}\right)^{-1}\right) \Rightarrow \hat{\eta}_i \sim N(x_i \beta, x_i^2 \text{var}(\beta))$$

$= \text{var}(\beta)$

$$CI = \hat{\eta}_i \pm z_{1-\alpha/2} \sqrt{x_i^2 \text{var}(\hat{\beta})} \quad (\text{Approximate CI})$$

The CI for μ_i can be simply obtained by transforming the above terms, obtaining

$$\frac{e^{\hat{\eta}_i \pm z_{1-\alpha/2} \sqrt{x_i^2 \text{var}(\hat{\beta})}}}{1 + e^{\hat{\eta}_i \pm z_{1-\alpha/2} \sqrt{x_i^2 \text{var}(\hat{\beta})}}} \quad (\text{Approximate CI}).$$

This is not Wald.