

STAT 431 — Applied Bayesian Analysis — Course Notes

Improper Priors

Fall 2022

Improper Densities

An **improper density** $f(\mathbf{x})$ satisfies

- ▶ $f(\mathbf{x}) \geq 0$ for all \mathbf{x}
- ▶ $\sum_{\mathbf{x}} f(\mathbf{x}) = \infty$ (discrete) or
 $\int f(\mathbf{x}) d\mathbf{x} = \infty$ (continuous)
(i.e. it cannot be *normalized* to a **proper** density)

When an improper density is used as a (“noninformative”) prior density, it is called an **improper prior**.

Eg: Population proportion θ (binomial model)

$$\pi(\theta) = \frac{1}{\theta(1-\theta)} \quad 0 < \theta < 1$$

(like a beta density with “ $\alpha = 0$ ” and “ $\beta = 0$ ”)

Improper because

$$\int_0^1 \pi(\theta) d\theta = \infty$$

[Draw density ...]

Warning: Improper priors may lead to improper posteriors!

Eg: (continued)

If $Y \sim \text{Binomial}(n, \theta)$ has prior

$$\pi(\theta) = \frac{1}{\theta(1-\theta)} \quad 0 < \theta < 1$$

and either $y = 0$ or $y = n$, then can show that $p(\theta | y)$ is improper!

Alternative: Use “vague” priors — proper, but close to improper.

Eg: Use $\text{Beta}(\alpha, \beta)$ with α and β “near” zero.

Example: Mean-Only Normal Sample

$$Y_1, \dots, Y_n \mid \mu \sim iid \text{ Normal}(\mu, \sigma^2) \quad \sigma^2 \text{ known}$$

Recall likelihood

$$\propto e^{-\frac{n}{2\sigma^2} (\mu - \bar{y})^2}$$

Try improper “flat” prior

$$\pi(\mu) \propto 1 \quad -\infty < \mu < \infty$$

We must check that the posterior is proper ...

$$p(\mu \mid \mathbf{y}) \propto f(\mathbf{y} \mid \mu) \underbrace{\pi(\mu)}_{\propto 1} \propto e^{-\frac{n}{2\sigma^2} (\mu - \bar{y})^2}$$

Recognize as the kernel of $\text{Normal}(\bar{y}, \sigma^2/n)$ (why?), so the posterior is indeed proper:

$$\mu \mid \mathbf{y} \sim \text{Normal}(\bar{y}, \sigma^2/n)$$

Note: The posterior mean is \bar{y} and the posterior standard deviation is the (frequentist) standard error of \bar{y} .

Can show that, under this improper prior,

- ▶ credible intervals for μ are the same as confidence intervals
- ▶ the posterior probability of a *one-sided* H_0 is the same as a p -value
(Not true for the two-sided case.)
- ▶ this posterior is the limit as $\tau_0^2 \rightarrow 0$ (equiv. $m \rightarrow 0$) of the posterior when the $\text{Normal}(\mu_0, 1/\tau_0^2)$ prior is used

Notation: For a flat prior, write, e.g.

$$\mu \sim 1 d\mu$$

Eg: Jevons's Coin Data

Posterior under flat prior:

$$\mu \mid \mathbf{y} \sim \text{Normal}(\mu_1 = \bar{y} = 7.8730, \sigma_1^2 = \sigma^2/n \approx 0.0001194)$$

(so $\sigma_1 \approx 0.01093$)

[Draw density w/ probability limits ...]

So an approx. 95% credible interval is

$$\mu_1 \pm 1.96 \sigma_1 \approx (7.8516, 7.8944)$$

It excludes all values meeting the min. legal weight (7.9379).

Indeed, for $H_0 : \mu \geq 7.9379$ we find

$$\text{Prob}(H_0 \mid \mathbf{y}) = 1 - \Phi\left(\frac{7.9379 - \mu_1}{\sigma_1}\right) \approx 10^{-9}$$

(same as a p -value, in this case)

Now consider randomly selecting another coin of the same kind (minted before 1830). Its (random) weight will be Y^* .

Using the flat prior and previous formulas for the posterior predictive distribution,

$$Y^* \mid \mathbf{y} \sim \text{Normal}(7.8730, 0.0001194 + (0.05353)^2)$$

The posterior predictive standard deviation works out to be about 0.05463.

The posterior predictive prob. that *this coin* is of legal weight:

$$\begin{aligned} \text{Prob}(Y^* \geq 7.9379 \mid \mathbf{y}) &\approx 1 - \Phi\left(\frac{7.9379 - 7.8730}{0.05463}\right) \\ &\approx 0.1174 \end{aligned}$$

Remarks

- ▶ An improper prior may be hard to interpret subjectively, since it can't produce probabilities for θ .

Perhaps interpret as equivalent to a series of proper priors whose posteriors converge to the posterior produced by the improper prior.

(e.g. as $\sigma_0^2 \rightarrow \infty$, $\text{Normal}(0, \sigma_0^2)$ priors produce posteriors converging to the flat prior's posterior)

- ▶ Representing the condition of “no prior information” often requires considering improper priors (“noninformative” priors — next).