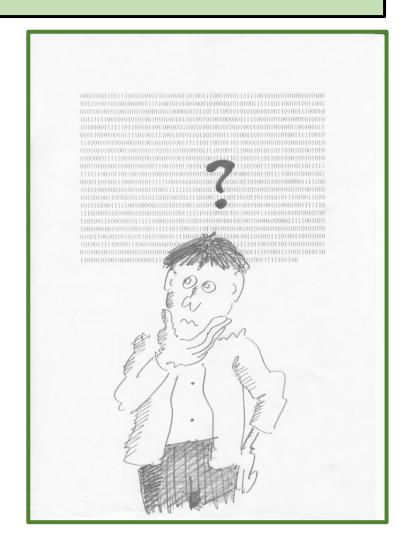
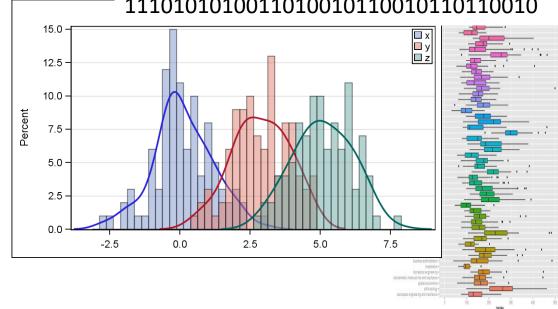
#### Statistics and data analysis

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# Independence and variations, convolution, computer age statistics



#### Negative Binomial Distribution

- In successive Bernoulli(p) instances, what is the distribution of the number of trials (in some versions failures) needed until the  $r^{\rm th}$  success. (the Geometric Distribution is equivalent to r=1)
- For this number to equal y we should have exactly r-1 successes in first y-1 trials, followed by a success

$$p(y) = {y-1 \choose r-1} p^r (1-p)^{y-r} \quad y = r, r+1, \dots$$

$$E(Y) = \frac{r}{p}$$

$$V(Y) = \frac{r(1-p)}{p^2}$$

#### Randomistan basketball, again

Players shoot synchronously.

#### Player 1:

Probability of scoring = p < 1/2Shoots until he has r successes.  $X_1$  is the attempt when that happened.

#### Player 2:

Probability of scoring = mp for some integer 1 < m so that mp < 1 Shoots until she has mr successes.  $X_2$  is the attempt when that happened.





- Which is higher  $E(X_1)$  or  $E(X_2)$ ?
- Which is higher  $V(X_1)$  or  $V(X_2)$ ?
- Placing a bet on  $X_1 > X_2$ ? (Player 2 is better)

$$E(X_1) = \frac{r}{p} = \frac{mr}{mp} = E(X_2)$$

$$V(X_1) = \frac{r(1-p)}{p^2} ? \frac{mr(1-mp)}{(mp)^2} = V(X_2)$$

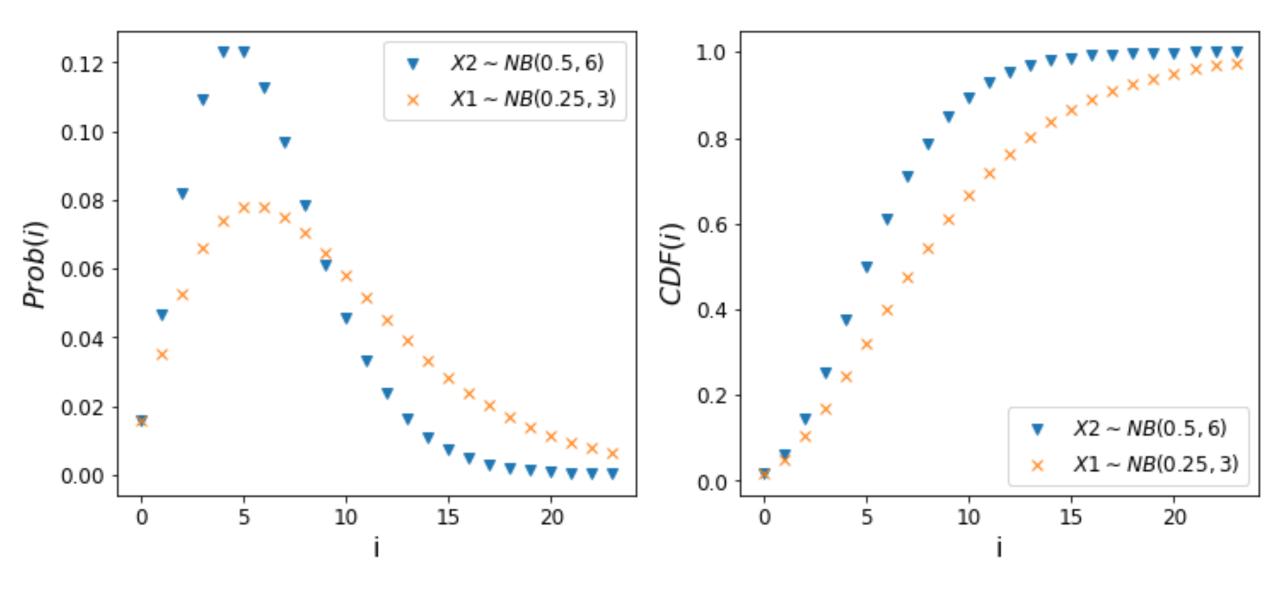
From: DeGroot, Probability and Statistics

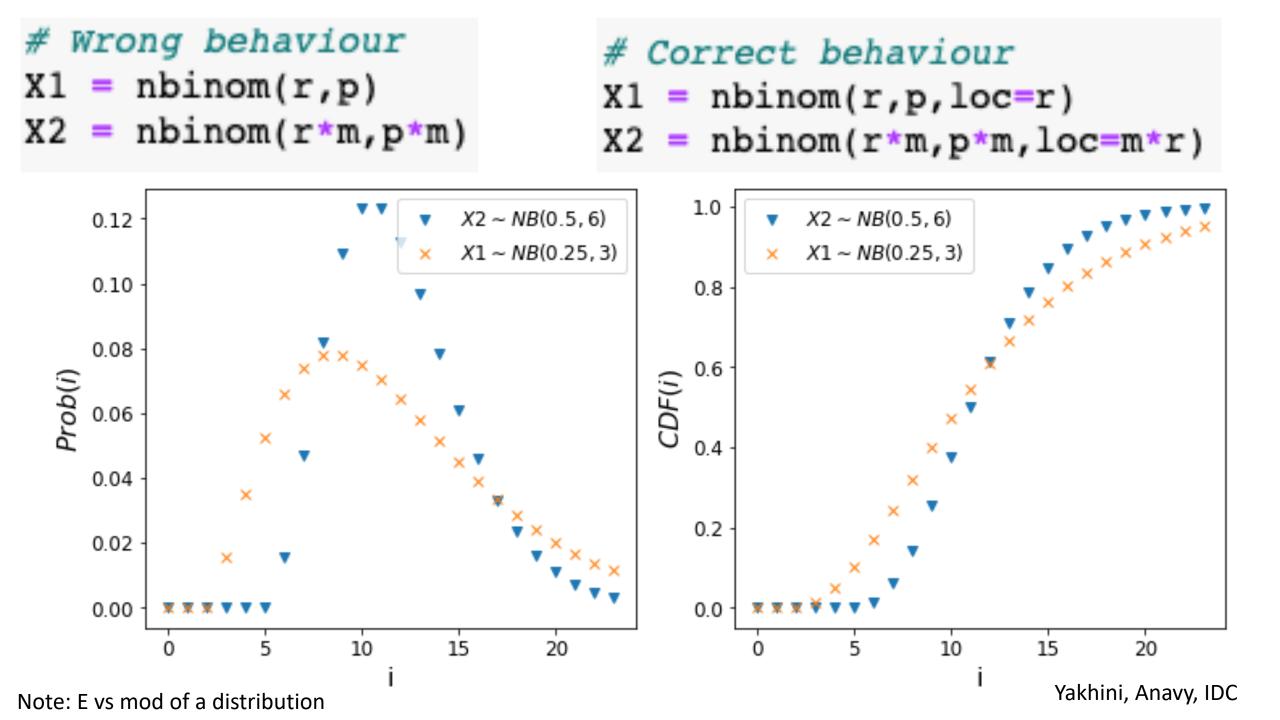
#### scipy.stats.nbinom

```
from scipy.stats import nbinom import numpy as np from matplotlib import pyplot as plt
```

```
r = 3
p = 0.25
m = 2
```

```
X1 = nbinom(r,p)
X2 = nbinom(r*m,p*m)
i = range(0, int(np.round(2*r/p, 0)))
p X1 i = X1.pmf([xx for xx in i])
p X2 i = X2.pmf([xx for xx in i])
plt.figure(figsize=(12,5))
plt.subplot(1,2,1)
plt.plot(i,p_X2_i,'v',label="$X2\sim NB({{{0}}}},{{{1}}})$".format(p*m,r*m))
plt.plot(i, p X1 i, 'x', label="$X1\sim NB({{{0}}}}, {{{1}}})$".format(p,r))
plt.xlabel("i",fontsize=16)
plt.ylabel('$Prob(i)$',fontsize=16)
plt.legend()
```





#### Randomistan basketball story

- Which is higher  $E(X_1)$  or  $E(X_2)$ ?
- Which is higher  $V(X_1)$  or  $V(X_2)$ ?
- Placing a bet on  $X_1 > X_2$ ? (Player 2 is better)

```
r = 3
p = 0.25
m = 2
mean_X1, var_X1 = nbinom.stats(r,p,loc=r)
mean_X2, var_X2 = nbinom.stats(r*m,p*m,loc=m*r)
print(f'E(X1_1) = {mean_X1}, Var(X1_1) = {var_X1}')
print(f'E(X1_2) = {mean_X2}, Var(X1_2) = {var_X2}')

E(X1_1) = 12.0, Var(X1_1) = 36.0
E(X1_2) = 12.0, Var(X1_2) = 12.0
```

#### How to assess betting on the players?

• Placing a bet on  $X_1 > X_2$ ? (Player 2 is better)

We can choose a player and bet on whether they succeed before 8, before 20. Which player should we prefer?

Calculate 
$$P(X_1 \le 8)$$
 and  $P(X_2 \le 8)$ 

Calculate  $P(X_1 \le 20)$  and  $P(X_2 \le 20)$ 

```
v1 = 8

v2 = 20

f_X1_v1 = X1.cdf(v1)

f_X2_v1 = X2.cdf(v1)

f_X1_v2 = X1.cdf(v2)

f_X2_v2 = X2.cdf(v2)
```

```
P(X1 <= 8) = 0.32

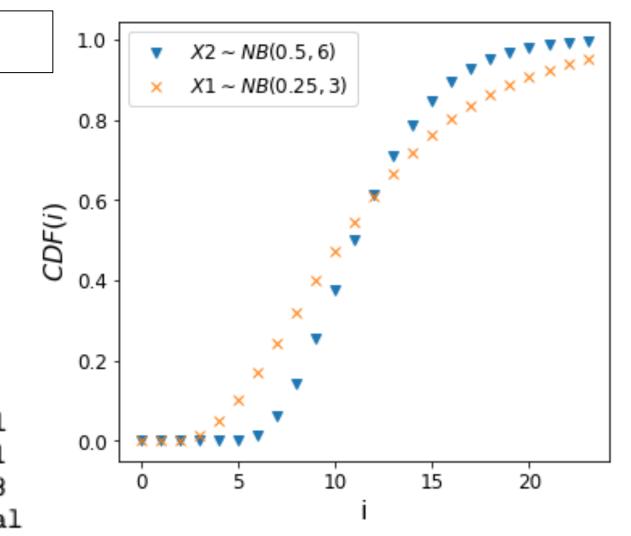
P(X2 <= 8) = 0.14

X1 wins on 8 trial

P(X1 <= 20) = 0.91

P(X2 <= 20) = 0.98

X2 wins on 20 trial
```



# Who completes the task earlier? Computer age statistics

Calculate  $P(X_1 > X_2)$ 

Lower Bound:

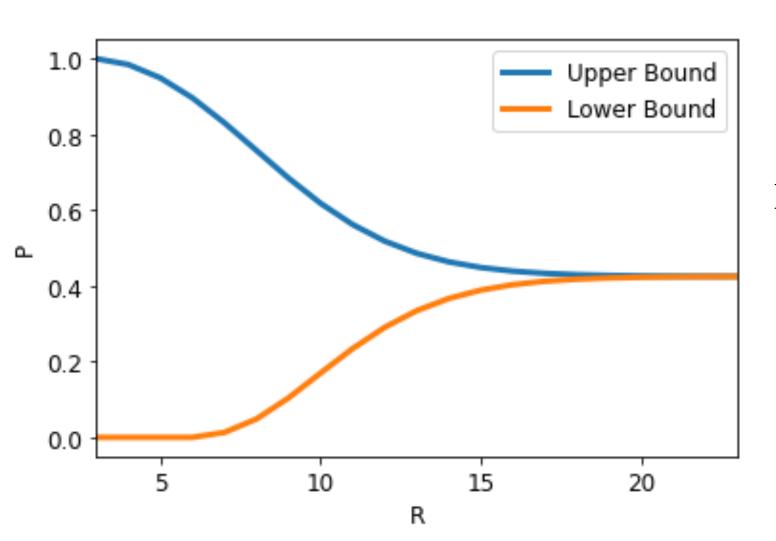
$$\begin{split} P(X_1 > X_2) &= \sum_{y=m \cdot k}^{\inf} P(X_2 = y) P(X_1 > y) \geq \\ \sum_{y=m \cdot k}^{R} P(X_2 = y) P(X_1 > y) &= \sum_{y=m \cdot k}^{R} P(X_2 = y) (1 - CDF_{X_1}(y)) \end{split}$$

Upper Bound:

$$P(X_1 > X_2) = 1 - P(X_2 \ge X_1) = 1 - \sum_{x=k}^{\inf} P(X_1 = x) P(X_2 \ge x) \le 1 - \sum_{x=k}^{R} P(X_1 = x) P(X_2 \ge x) = 1 - \sum_{x=k$$

### $P(X_1 > X_2)$

Calculate P(X > Y)



$$P(X_1 > X_2) \in [0.4246, 0.4251]$$

The RV T counts the number of observations required to see at least m=1 users from each country of the n=100. How many visits will it take if every visit comes from each of the countries with equal probabilities and independent of all previous visits?

$$T = X_1 + X_2 + X_3 + ... + X_i + ... + X_{99} + X_{100}$$

Where the random variable  $X_i$  counts the number of visits, after the first i-1 countries are in, until the i-th country is also in.

We saw

$$E(T) = E(X_1 + X_2 + X_3 + \dots + X_i + \dots + X_{99} + X_{100}) = \sum_{i=1}^{100} E(X_i)$$

Note that  $X_i \sim Geom\left(p_i = \frac{100-i+1}{100}\right)$  and we therefore have  $E(X_i) = \frac{1}{p_i} = \frac{100}{100-i+1}$  So:

$$E(T) = 100\left(1 + \frac{1}{2} + \frac{1}{3} + \frac{1}{4} + \dots + \frac{1}{100}\right)$$

$$\underline{E(T)} = nH(n) \sim n \ln n.$$

How can we use Chebyshev's inequality to get a bound on:

$$P(T > nH(n) + cn)$$
?

$$P(|X - \mu| \ge \lambda) \le \frac{V(X)}{\lambda^2}$$

$$\left| P(|X - \mu| \ge \lambda) \le \frac{V(X)}{\lambda^2} \right|$$

$$P(T \ge E(T) + \lambda) \le P(|T - E(T)| \ge \lambda)$$

$$P(T \ge nH(n) + cn) \le \frac{V(T)}{c^2 n^2}$$

What about the Variance of T?

$$Var(T) = Var(X_1 + X_2 + X_3 + \dots + X_i + \dots + X_{99} + X_{100}) = \sum_{i=1}^{n} Var(X_i)$$

 $X_i \sim Geom\left(p_i = \frac{100-i+1}{100}\right)$  and we therefore have  $Var(X_i) = \frac{1-p_i}{p_i^2}$  So:

$$Var(T) = \sum_{i=1}^{100} \frac{1 - p_i}{p_i^2} = \sum_{i=1}^{100} \frac{1}{p_i^2} - \sum_{i=1}^{100} \frac{1}{p_i} = 100^2 \sum_{i=1}^{100} \frac{1}{i^2} - 100 \sum_{i=1}^{100} \frac{1}{i}$$

$$Var(T) = n^2 \sum_{i=1}^{n} \frac{1}{i^2} - nH(n)$$

$$Var(T) = n^{2} \sum_{i=1}^{n} \frac{1}{i^{2}} - nH(n) < n^{2} \sum_{i=1}^{n} \frac{1}{i^{2}} < n^{2} \frac{\pi^{2}}{6}$$

$$P(T \ge nH(n) + cn) \le \frac{V(T)}{c^{2}n^{2}} < \frac{\pi^{2}}{6c^{2}}$$

Taking n = 100:

$$c = 2$$
:  $P(T \ge 518 + 200) < \frac{\pi^2}{24} = 0.4112$   
 $c = 3$ :  $P(T \ge 518 + 300) < \frac{\pi^2}{54} = 0.1828$   
 $c = 4$ :  $P(T \ge 518 + 400) < \frac{\pi^2}{96} = 0.1028$ 

#### Computer age statistics

We can calculate the true value of the variance and use this for the Chebyshev bound:

$$Var(T) = n^2 \sum_{i=1}^{n} \frac{1}{i^2} - nH(n)$$
$$P(T \ge nH(n) + cn) \le \frac{V(T)}{c^2 n^2}$$

$$P(T \ge nH(n) + cn) \le \frac{V(T)}{c^2 n^2}$$

Taking n = 100: Var(T) = 16449

$$c = 2$$
:  $P(T \ge 518 + 200) \le 0.3958$ 

$$c = 3$$
:  $P(T \ge 518 + 300) \le 0.1759$ 

$$c = 4$$
:  $P(T \ge 518 + 400) \le 0.0989$ 

```
v = single coupon variance(100)
print(f'Using exact Variance')
print(f'P(T_100>718) \le \{v/4/100**2 : .4f\}')
print(f'P(T 100>818) \le \{v/9/100**2 :.4f\}')
print(f'P(T 100>918) \le \{v/16/100**2 :.4f\}')
print(f'Using upper bound on the variance')
print(f'P(T 100>718) <= {math.pi ** 2 / 6 / 4 :.4f}')</pre>
print(f'P(T 100>818) \le \{math.pi ** 2 / 6 / 9 :.4f\}')
print(f'P(T 100>918) <= {math.pi ** 2 / 6 / 16 :.4f}')
```

```
Using exact Variance
P(T 100>718) \le 0.3958
P(T 100>818) \le 0.1759
P(T 100>918) \le 0.0989
Using upper bound on the variance
P(T 100>718) \le 0.4112
P(T 100>818) \le 0.1828
P(T 100>918) \le 0.1028
```

How can we calculate the probability directly: P(T > nH(n) + cn)?

$$T = X_1 + X_2 + X_3 + ... + X_i + ... + X_{99} + X_{100}$$

### Sums of independent random variables

Let X and Y be two independent random variables. Let Z = X + Y . Then

$$P(Z=z) = \sum_{i=-\infty}^{\infty} P(X=i)P(Y=z-i)$$

For continuous random variables, the density function of Z is:

$$h(z) = \int_{-\infty}^{\infty} f(t)g(z-t)dt$$

#### Computer age statistics

We can use convolutions to compute the actual FULL (or rather – the interesting part) distribution of  $T_N$ :

$$P(T_N = k) = \sum_{\substack{i = -\infty \\ k-1}}^{\infty} P(G_N = i) P(T_{N-1} = k - i)$$
$$= \sum_{i=1}^{\infty} P(G_N = i) P(T_{N-1} = k - i)$$

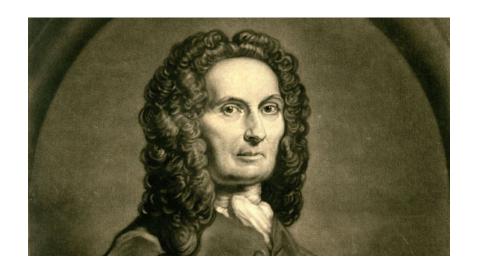
where 
$$G_S \sim Geo(p = \frac{N-s+1}{N})$$
.

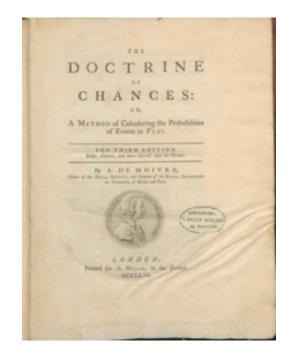
We need to initialize this with  $P(T_1 = 1) = 1$  and 0 for all other values.

## Exact coupon collector waiting time

```
exact P(T_100>718) = 0.12
exact P(T_100>818) = 0.05
exact P(T_100>918) = 0.02
```

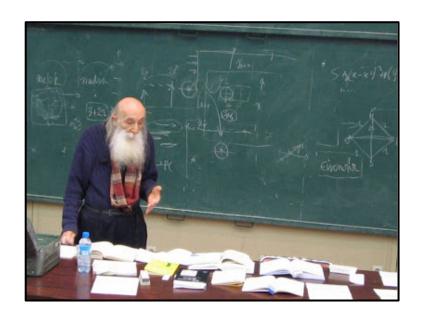
# Sample/coupon collector and computers: Historical notes





- The  $T_N$  discussion dates back to Abraham de Moivre 1667 (France) 1754 (England)
- Rigorously treated by William Feller in the 1940s
- Variants are still being studied as an active field of research

Jean Paul Benzecri, French statistician (1932-2019): "It is unthinkable to use methods conceived before the invention of the computer. Statistics will have to be completely rewritten!" Stated in 1965.



#### Pairwise independence

A set of random variables  $(X_1, X_2, ..., X_n)$  is said to be pairwise independent if any two random variables  $X_i$  and  $X_i$  are independent.

Recall – a set of random variables as above is called (collectively or mutually) independent if

$$\forall (x_1, x_2, ..., x_n)$$

$$P(X_1 = x_1, X_2 = x_2, ..., X_n = x_n) = \prod_{i=1}^n P(X_i = x_i)$$

#### Equivalence?

- Does collective independence imply pairwise independence?
- Does pairwise independence imply collective independence?

#### Var of a sum?

Pairwise independence is sufficient for the linearity of variances.

Let X and Y two independent Bernoulli w  $p=\frac{1}{2}$ . Let Z=XOR(X,Y).

We work in  $\Omega = \{0,1\}^3$ .

We have the following joint probability mass function:

X	Υ	Z	P
0	0	0	0.25
0	1	1	0.25
1	0	1	0.25
1	1	0	0.25

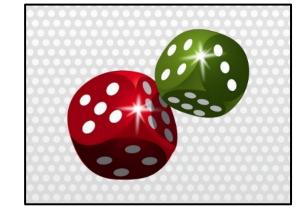
#### X + Y + Z vs Binom(0.5,3)

$$E(X + Y + Z) = ?$$
  
  $V(X + Y + Z) = ?$ 

X	Υ	Z	P
0	0	0	0.25
0	1	1	0.25
1	0	1	0.25
1	1	0	0.25

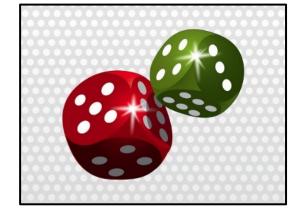
#### Covariance and independence

$$P(X = x) = 1/3 \text{ for } x = -1, 0, 1$$
,  $Y = X^2$ 



Roll a die n times, Y counts the number of 5's

- What is the distribution of *Y*?
- Can you treat this as a coin? What is p?
- What is E(Y)?
- What is V(Y)?

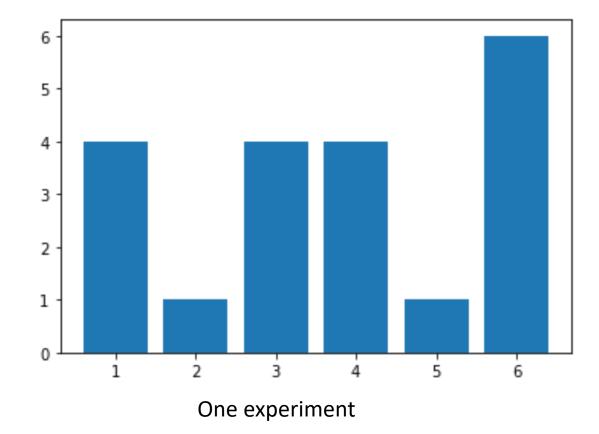


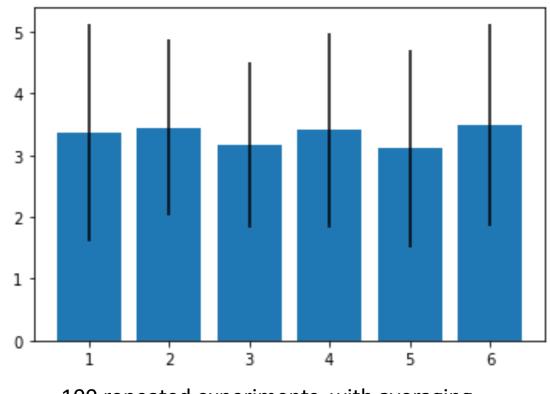
Now roll a die n times and define  $X = (X_1, X_2, ..., X_6)$  where  $X_i$  counts the number of i's

$$X = (X_1, X_2, ..., X_d) \sim \text{Multinomial}(n, p)$$

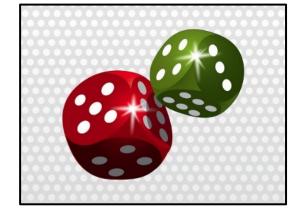
- What is *d*?
- What is *p*?
- What is  $E(X_i)$ ?
- What is  $V(X_i)$ ?

$$X = (X_1, X_2, \dots, X_6) \sim \text{Multinomial}\left(20, \left(\frac{1}{6}, \dots, \frac{1}{6}\right)\right)$$





100 repeated experiments, with averaging



Now roll a die n times and define  $X=(X_1,X_2,\dots,X_6)$  where  $X_i$  counts the number of i's

$$X = (X_1, X_2, ..., X_d) \sim \text{Multinomial}(n, p)$$

- Are these random variables collectively independent?
- Pairwise independent?

#### Multinomial Distribution - covariances

Let  $X \sim \text{MNom}(N, P)$ ,  $X = (X_1, X_2, ..., X_d)$ . What is  $Cov(X_i, X_j)$ ?

$$Var(X_i + X_j) = V(X_i) + 2Cov(X_i, X_j) + V(X_j)$$

Now observe that  $X_i + X_j \sim \text{Binom}(N, p_i + p_j)$  and therefore we can write:

$$2Cov(X_i, X_j) = Var(X_i + X_j) - V(X_i) - V(X_j) =$$

$$= N[(p_i + p_j)(1 - p_i - p_j) - p_i(1 - p_i) - p_j(1 - p_j)]$$

$$= -2Np_i p_j$$

#### Multinomials, example

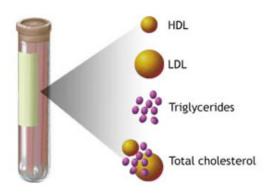
Let 
$$X \sim MNom(N, P)$$
,  $X = (X_1, X_2, ..., X_{10})$ 

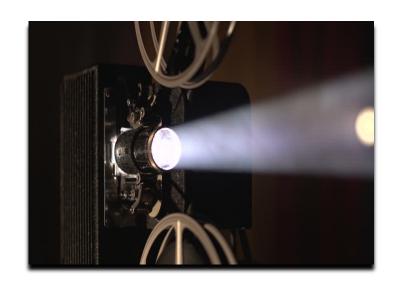
Also assume that P is uniform  $\frac{1}{10}$ 

What is 
$$Cov(X_1 + X_2, X_7 + X_8)$$
?

#### Conditional independence

- Is the blood cholesterol level of a person independent of the number of movies watched by that person so far?
- No they are both related to the age of the person.
- But they are conditionally independent given the age.
- Presumably ..., socioeconomic and behavioral factors ignored ...
- Notation:  $X \perp Y \mid C$





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# Conditional Independence - Definition

Two random variables X and Y are conditionally independent given a third rv C if for all pairs (x, y) AND for all possible values c of C, we have:

$$P((X = x \land Y = y) | C = c) = P((X = x) | C = c) \cdot P((Y = y) | C = c)$$

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Conditionally independent but not independent?

Independent but not conditionally independent?

#### Summary

- Computer age statistics:
  - + Comparing negative binomials
  - Coupon collector exact calculations
  - + Independence and convolutions
- Mutual indpce vs lower order indpce
- Multinomials
- Conditional independence