

# ECON 121 FA23 Problem Set 3

## Solution

### Question 2

16 percent of the sample reports being in fair or poor health, and 13 percent died before 2019. The sample has a median age of 49. (The mean age is less meaningful because age was top-coded at 85. You did not need to notice this.) 56 percent of the sample is female, perhaps surprisingly. This gender imbalance has two sources. First, men and women responded to the survey at different rates, so the gender imbalance would shrink when we use the sampling weights. Second, men die at higher rates than women, so the gender imbalance grows with age. For the outcome variables, 16 percent of the sample reports being in fair or poor health, while 13 percent dies by 2019.

```
# generate fair/poor health dummy
table(nhis2010$health)

##
## Excellent Very Good      Good      Fair      Poor
##      5953     7447     7012    2968     962

nhis2010$fpoor <- ifelse(nhis2010$health == "Fair" |
                           nhis2010$health == "Poor", 1, 0)

# summarize the dataset
summary(nhis2010)

##      sampweight          psu          hhnum          pernum
##  Min.   : 853   Min.   : 1.0   Min.   : 1   Min.   : 1.000
##  1st Qu.: 4339  1st Qu.:156.0  1st Qu.:10380  1st Qu.: 1.000
##  Median : 6879   Median :306.5   Median :21096   Median : 1.000
##  Mean   : 8214   Mean   :304.8   Mean   :21235   Mean   : 1.371
##  3rd Qu.:10712  3rd Qu.:460.0  3rd Qu.:31968  3rd Qu.: 2.000
##  Max.   :65899   Max.   :600.0   Max.   :43208   Max.   :12.000
##
##      age          male          marstat          white
##  Min.   :25.00  Min.   :0.0000  Married   :11724  Min.   :0.0000
##  1st Qu.:37.00  1st Qu.:0.0000  Widowed   : 2549  1st Qu.:0.0000
##  Median :49.00  Median :0.0000  Divorced   : 3988  Median :1.0000
##  Mean   :50.79  Mean   :0.4381  Separated  : 1003  Mean   :0.5764
##  3rd Qu.:63.00  3rd Qu.:1.0000  Never married: 5043  3rd Qu.:1.0000
##  Max.   :85.00  Max.   :1.0000  NA's       :  49  Max.   :1.0000
##
##      black          hisp          asian          other
##  Min.   :0.0000  Min.   :0.0000  Min.   :0.00000  Min.   :0.00000
##  1st Qu.:0.0000  1st Qu.:0.0000  1st Qu.:0.00000  1st Qu.:0.00000
##  Median :0.0000  Median :0.0000  Median :0.00000  Median :0.00000
##  Mean   :0.1611  Mean   :0.1824  Mean   :0.06249  Mean   :0.01757
##  3rd Qu.:0.0000  3rd Qu.:0.0000  3rd Qu.:0.00000  3rd Qu.:0.00000
##  Max.   :1.0000  Max.   :1.0000  Max.   :1.00000  Max.   :1.00000
```

```

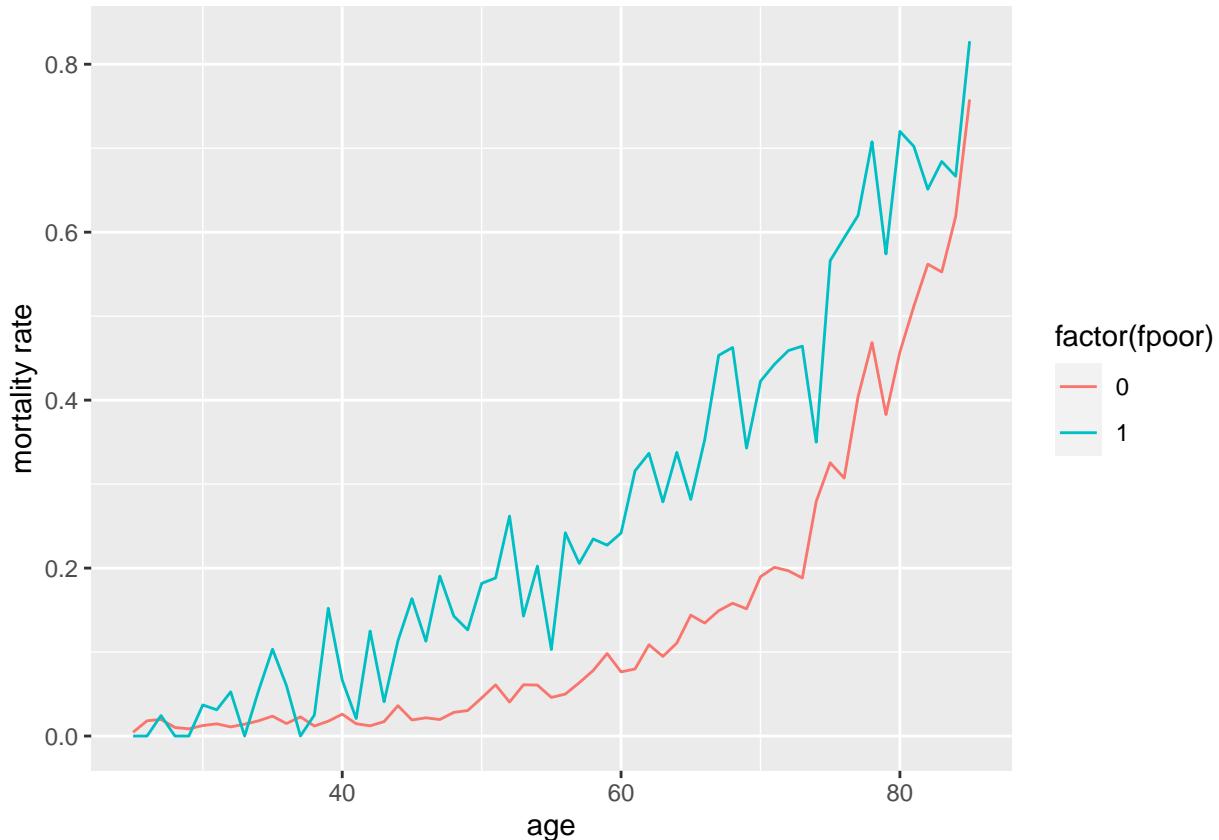
##          edyrs                  empstat
##  Min.   : 1.00   Working for pay at job/business   :13248
##  1st Qu.:13.00  Not in labor force                 : 8856
##  Median :14.00  Not employed                      :1451
##  Mean   :13.81  With job, but not at work           : 563
##  3rd Qu.:16.00  Working, w/out pay, at job/business: 224
##  Max.   :19.00  (Other)                           :    0
##  NA's    :119   NA's                            :   14
##          incfam                health        mort      bmi
##  $0 - $34,999   :9737  Excellent:5953  Min.   :0.0000  Min.   : 9.89
##  $35,000 - $49,999:3469  Very Good:7447  1st Qu.:0.0000  1st Qu.:23.72
##  $50,000 - $74,999:3849  Good   :7012   Median :0.0000  Median :26.69
##  $75,000 - $99,999:2334  Fair    :2968   Mean    :0.1289  Mean    :27.91
##  $100,000 and over:3634  Poor   : 962   3rd Qu.:0.0000  3rd Qu.:30.86
##  NA's       :1333  NA's    :    14  Max.   :1.0000  Max.   :87.84
##                               NA's    :362   NA's    :933
##          uninsured            cancerev      cheartdiev      heartattev
##  Min.   :0.0000  Min.   :0.000000  Min.   :0.000000  Min.   :0.00000
##  1st Qu.:0.0000  1st Qu.:0.000000  1st Qu.:0.000000  1st Qu.:0.00000
##  Median :0.0000  Median :0.000000  Median :0.000000  Median :0.00000
##  Mean   :0.1744  Mean   :0.09488   Mean   :0.05445   Mean   :0.03798
##  3rd Qu.:0.0000  3rd Qu.:0.000000  3rd Qu.:0.000000  3rd Qu.:0.00000
##  Max.   :1.0000  Max.   :1.000000  Max.   :1.000000  Max.   :1.00000
##  NA's    :62     NA's    :20      NA's    :58      NA's    :25
##          hypertenев      diabeticev      alc5upyr      smokev
##  Min.   :0.0000  Min.   :0.0000  Min.   : 0.00  Min.   :0.0000
##  1st Qu.:0.0000  1st Qu.:0.0000  1st Qu.: 0.00  1st Qu.:0.0000
##  Median :0.0000  Median :0.0000  Median : 0.00  Median :0.0000
##  Mean   :0.3571  Mean   :0.1272  Mean   :10.95  Mean   :0.4202
##  3rd Qu.:1.0000  3rd Qu.:0.0000  3rd Qu.: 2.00  3rd Qu.:1.0000
##  Max.   :1.0000  Max.   :1.0000  Max.   :365.00  Max.   :1.0000
##  NA's    :38     NA's    :16      NA's    :9745   NA's    :178
##          vig10fwk      hrsleep      asad
##  Min.   : 0.000  Min.   : 3.000  None of the time  :17381
##  1st Qu.: 0.000  1st Qu.: 6.000  A little of the time: 3428
##  Median : 0.000  Median : 7.000  Some of the time   : 2428
##  Mean   : 1.494  Mean   : 7.158  Most of the time   :  650
##  3rd Qu.: 2.000  3rd Qu.: 8.000  All of the time    :  302
##  Max.   :28.000  Max.   :22.000  NA's                   : 167
##  NA's    :309   NA's    :367
##          fpoor
##  Min.   :0.0000
##  1st Qu.:0.0000
##  Median :0.0000
##  Mean   :0.1614
##  3rd Qu.:0.0000
##  Max.   :1.0000
##  NA's   :14

```

### Question 3

5-year mortality is higher for people with fair/poor health than for people with good/very good/excellent health. Thus, self-reported health status is predictive of mortality. In both groups, 5-year mortality rises non-linearly with age.

```
tbl <-  
  nhis2010 |>  
  drop_na(age, fpoor, mort) |>  
  group_by(age, fpoor) |>  
  summarise(mort = mean(mort))  
  
## `summarise()` has grouped output by 'age'. You can override using the `.groups`  
## argument.  
ggplot(tbl, aes(x = age, y = mort, color = factor(fpoor))) +  
  geom_line() +  
  labs(x="age", y="mortality rate")
```

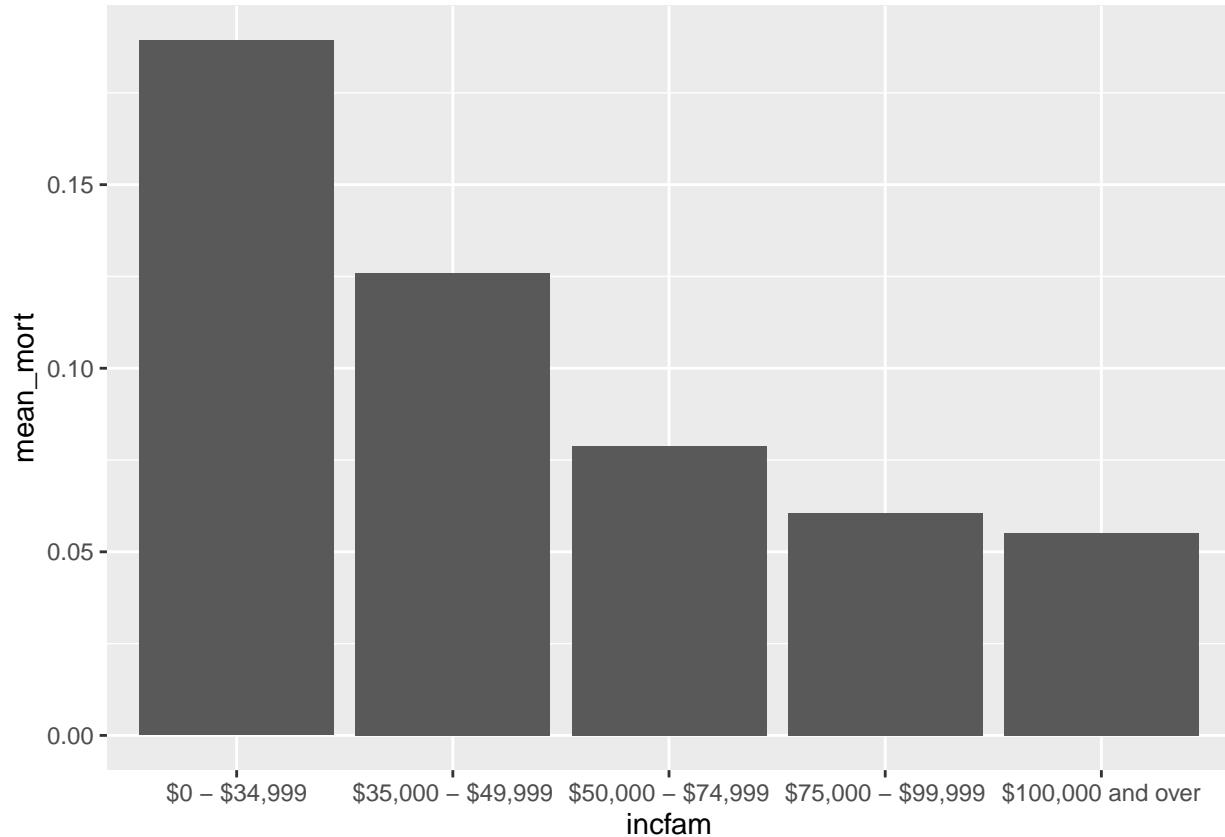


## Question 4

Rates of mortality and fair/poor health decline with family income. The same general pattern holds for education as well, although individuals with post- graduate education do not appear to be in worse health than college graduates.

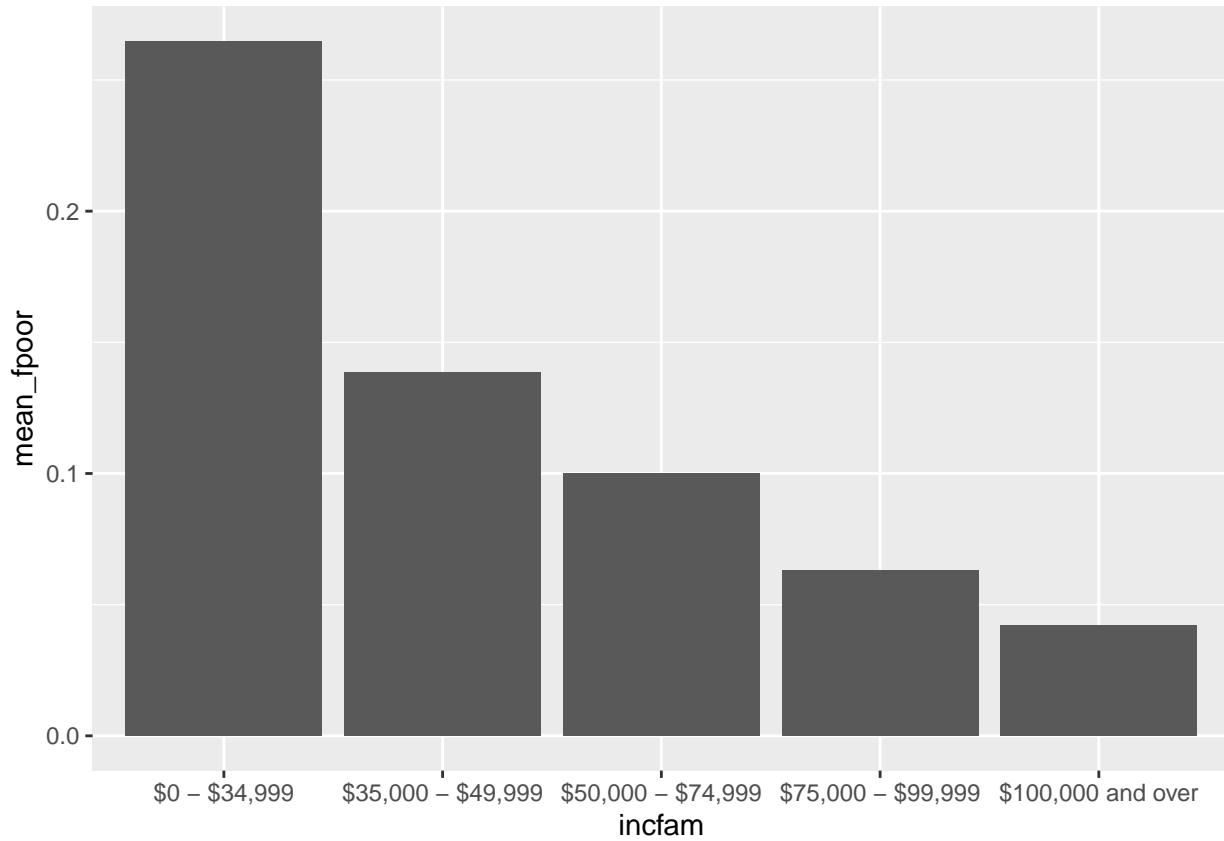
```
# Mortality by income
tbl_mort_income <-
  nhis2010 |>
  drop_na(incfam, mort) |>
  group_by(incfam) |>
  summarize(mean_mort = mean(mort))

ggplot(tbl_mort_income, aes(x = incfam, y = mean_mort)) +
  geom_col()
```



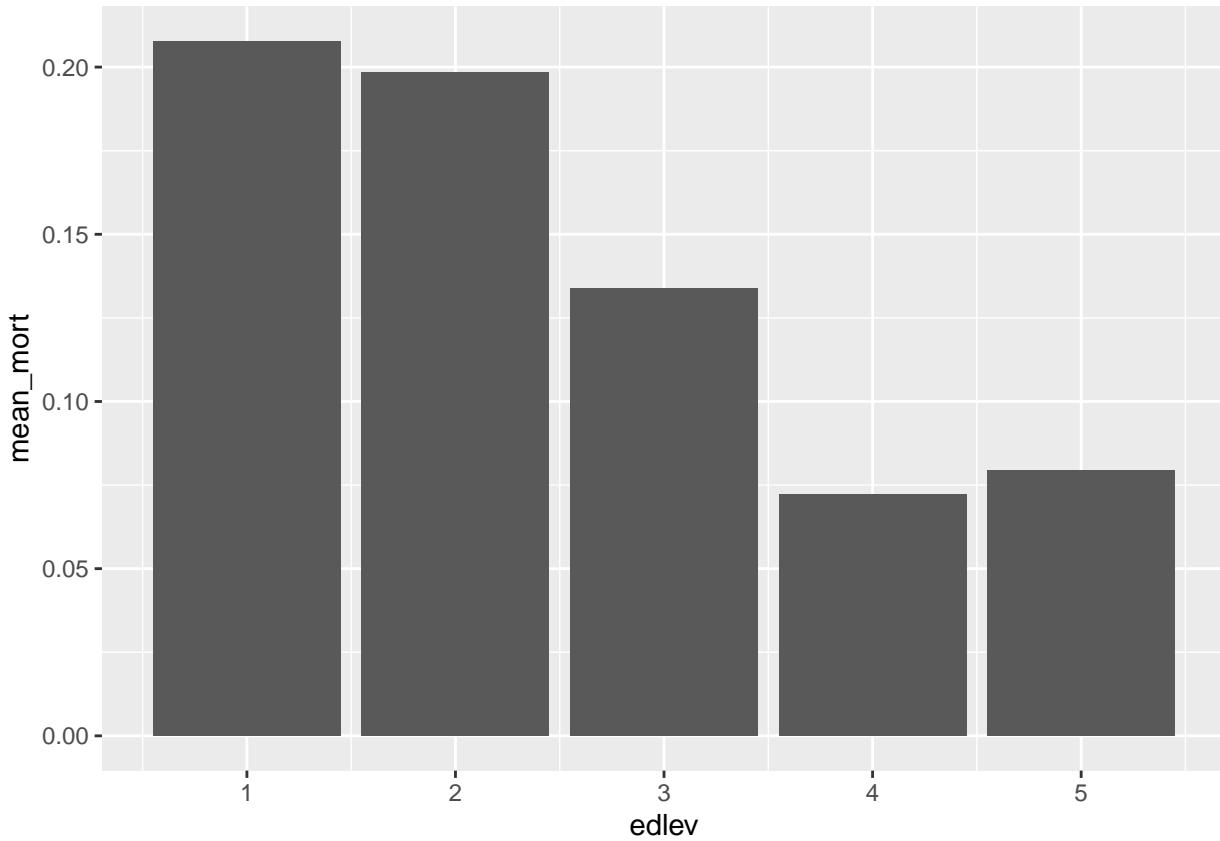
```
# Health by income
tbl_health_income <-
  nhis2010 |>
  drop_na(incfam, fpoor) |>
  group_by(incfam) |>
  summarize(mean_fpoor = mean(fpoor))

ggplot(tbl_health_income, aes(x = incfam, y = mean_fpoor)) +
  geom_col()
```



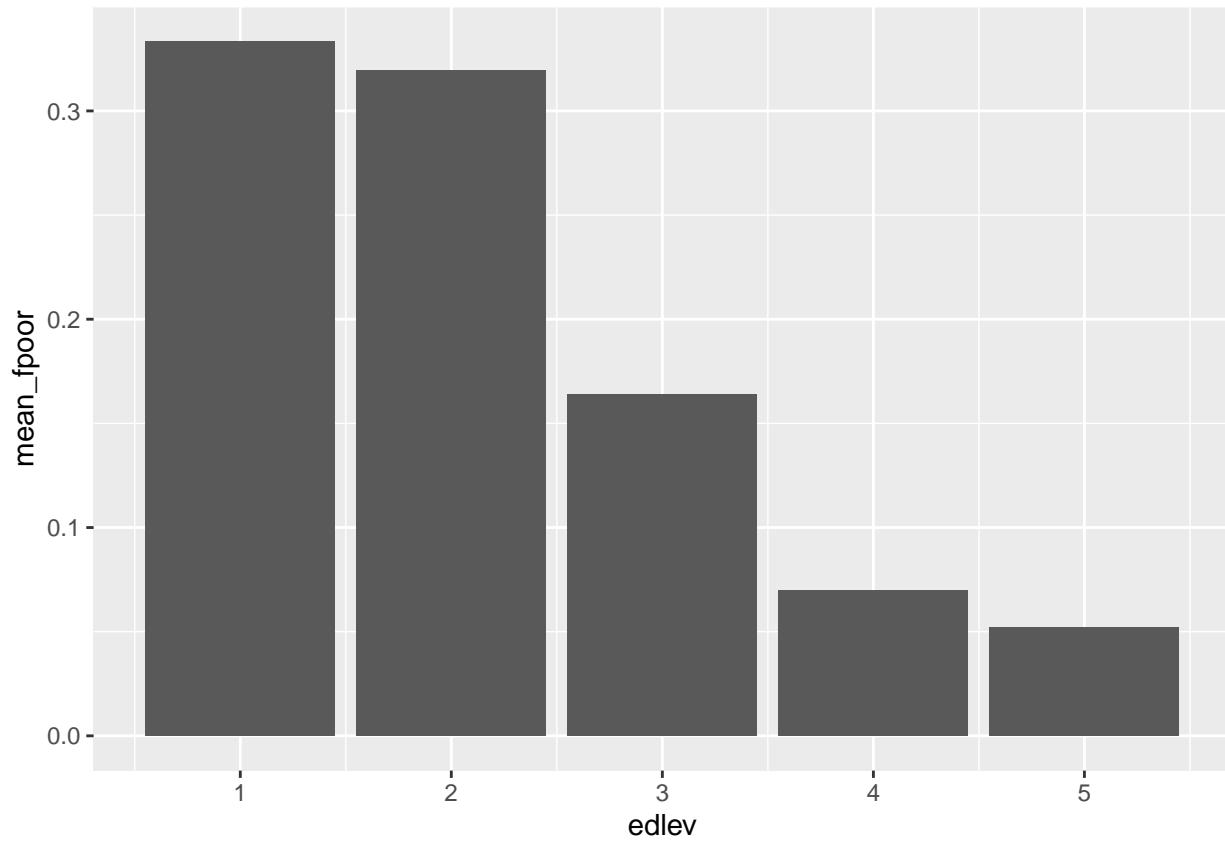
```
# Mortality by education
tbl_mort_education <-
  nhis2010 |>
  drop_na(edyrs, mort) |>
  mutate(edlev = case_when((edyrs<12) ~ 1, # code edlev as
                           (edyrs==12) ~ 2, # numeric so the
                           (edyrs>=13 & edyrs<15) ~ 3, # graph is ordered
                           (edyrs==16) ~ 4, # correctly
                           (edyrs>=17) ~ 5)) |>
  group_by(edlev) |>
  summarize(mean_mort = mean(mort))

ggplot(tbl_mort_education, aes(x = edlev, y = mean_mort)) +
  geom_col()
```



```
# Health by education
tbl_health_education <-
  nhis2010 |>
  drop_na(edyrs, fpoor) |>
  mutate(edlev = case_when((edyrs<12) ~ 1, # code edlev as
                           (edyrs==12) ~ 2, # numeric so the
                           (edyrs>=13 & edyrs<15) ~ 3, # graph is ordered
                           (edyrs==16) ~ 4, # correctly
                           (edyrs>=17) ~ 5)) |>
  group_by(edlev) |>
  summarize(mean_fpoor = mean(fpoor))

ggplot(tbl_health_education, aes(x = edlev, y = mean_fpoor)) +
  geom_col()
```



## Question 5

Because age and education have non-linear relationships with health, I include a series of dummy variables for categories. I use the education categories from above, and 10-year age intervals.

For both outcomes and for all three models, the results show that mortality and fair/poor health decline with income, decline with education, and rise with age. One surprising result is that conditional on the socioeconomic variables, racial gaps in mortality are small and insignificant. There are larger racial gaps in fair/poor health. Another surprising result is that Hispanics have low mortality risk (conditional on the other covariates).

The linear probability results are similar to the probit and logit average marginal effects, although the similarity is much stronger for fair/poor health than for mortality. You did not need to comment on the reason in your response, but the larger difference in the case of mortality is probably due to the fact that mortality risk is exceptionally low across much of the age distribution, so that the marginal effect is calculated in the flatter part of the CDF.

```
# Generate age and education categories
nhis2010 <-
  nhis2010 %>%
  mutate(agecat = floor(age/10)*10,
        edlev = case_when((edyrs<12) ~ 1,
                           (edyrs==12) ~ 2,
                           (edyrs>=13 & edyrs<15) ~ 3,
                           (edyrs==16) ~ 4,
                           (edyrs>=17) ~ 5))

# Mortality analyses
ols_mort <-
  feols(mort ~ incfam + factor(edlev) + factor(agecat) +
         black + hisp + asian + other,
        data = nhis2010, vcov = 'hetero')

## NOTE: 1,701 observations removed because of NA values (LHS: 362, RHS: 1,419).
summary(ols_mort)

## OLS estimation, Dep. Var.: mort
## Observations: 22,655
## Standard-errors: Heteroskedasticity-robust
##                               Estimate Std. Error     t value   Pr(>|t|)
## (Intercept)            0.073432  0.008105  9.059683 < 2.2e-16 ***
## incfam$35,000 - $49,999 -0.032004  0.005956 -5.372999 7.8202e-08 ***
## incfam$50,000 - $74,999 -0.054827  0.005510 -9.949666 < 2.2e-16 ***
## incfam$75,000 - $99,999 -0.057139  0.006096 -9.373067 < 2.2e-16 ***
## incfam$100,000 and over -0.057600  0.005674 -10.152461 < 2.2e-16 ***
## factor(edlev)2       0.010445  0.014579  0.716441 4.7373e-01
## factor(edlev)3      -0.018857  0.007588 -2.485204 1.2955e-02 *
## factor(edlev)4      -0.034375  0.008383 -4.100574 4.1356e-05 ***
## factor(edlev)5      -0.032430  0.009195 -3.527038 4.2107e-04 ***
## factor(agecat)30    0.013449  0.003125  4.304022 1.6843e-05 ***
## factor(agecat)40    0.028204  0.003677  7.670231 1.7860e-14 ***
## factor(agecat)50    0.077196  0.005002 15.433618 < 2.2e-16 ***
## factor(agecat)60    0.151830  0.006785 22.376300 < 2.2e-16 ***
## factor(agecat)70    0.325477  0.010798 30.143234 < 2.2e-16 ***
## factor(agecat)80    0.639263  0.013159 48.578255 < 2.2e-16 ***
## black                 -0.002238  0.005944 -0.376543 7.0652e-01
```

```

## hisp          -0.046483  0.004988 -9.319681 < 2.2e-16 ***
## asian         -0.037678  0.006817 -5.527279 3.2882e-08 ***
## other         -0.000834  0.014384 -0.057952 9.5379e-01
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## RMSE: 0.284184 Adj. R2: 0.269659

probit_mort <-
  feglm(mort ~ incfam + factor(edlev) + factor(agecat) +
    black + hisp + asian + other,
    data = nhis2010, vcov = 'hetero', family = 'probit')

## NOTE: 1,701 observations removed because of NA values (LHS: 362, RHS: 1,419).
summary(probit_mort)

## GLM estimation, family = binomial(link = "probit"), Dep. Var.: mort
## Observations: 22,655
## Standard-errors: Heteroskedasticity-robust
##                               Estimate Std. Error     z value Pr(>|z|)
## (Intercept)           -1.932348  0.085988 -22.472354 < 2.2e-16 ***
## incfam$35,000 - $49,999 -0.190400  0.036212 -5.257877 1.4573e-07 ***
## incfam$50,000 - $74,999 -0.357315  0.040155 -8.898419 < 2.2e-16 ***
## incfam$75,000 - $99,999 -0.391555  0.051932 -7.539749 4.7088e-14 ***
## incfam$100,000 and over -0.421553  0.046507 -9.064244 < 2.2e-16 ***
## factor(edlev)2        0.049972  0.073592  0.679050 4.9711e-01
## factor(edlev)3        -0.092318  0.039459 -2.339613 1.9304e-02 *
## factor(edlev)4        -0.233724  0.052412 -4.459321 8.2220e-06 ***
## factor(edlev)5        -0.198558  0.059864 -3.316843 9.1041e-04 ***
## factor(agecat)30       0.249896  0.086916  2.875136 4.0385e-03 **
## factor(agecat)40       0.495115  0.083473  5.931454 3.0026e-09 ***
## factor(agecat)50       0.941066  0.080090 11.750129 < 2.2e-16 ***
## factor(agecat)60       1.304036  0.079502 16.402509 < 2.2e-16 ***
## factor(agecat)70       1.821506  0.080705 22.570043 < 2.2e-16 ***
## factor(agecat)80       2.617116  0.084105 31.117342 < 2.2e-16 ***
## black                 -0.000115  0.035422 -0.003237 9.9742e-01
## hisp                  -0.334561  0.043756 -7.646122 2.0713e-14 ***
## asian                 -0.299337  0.068069 -4.397580 1.0946e-05 ***
## other                  0.018750  0.096302  0.194701 8.4563e-01
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Log-Likelihood: -6,114.4 Adj. Pseudo R2: 0.287882
## BIC: 12,419.4 Squared Cor.: 0.275609

avg_slopes(probit_mort) # probit marginal effects

## Warning: The `agecat` variable is treated as a categorical (factor) variable, but
##   the original data is of class numeric. It is safer and faster to convert
##   such variables to factor before fitting the model and calling a
##   `marginaleffects` function.
##
##   This warning appears once per session.

##                               Contrast Estimate Std. Error     z Pr(>|z|)
## agecat 30 - 20            1.01e-02  0.00323  3.12624  0.00177
## agecat 40 - 20            2.61e-02  0.00374  6.98816 < 0.001

```

```

##  agecat 50 - 20          7.83e-02  0.00494 15.84384 < 0.001
##  agecat 60 - 20          1.50e-01  0.00657 22.83655 < 0.001
##  agecat 70 - 20          3.02e-01  0.01050 28.81040 < 0.001
##  agecat 80 - 20          6.02e-01  0.01441 41.73642 < 0.001
##  asian   1 - 0           -3.92e-02  0.00782 -5.01889 < 0.001
##  black   1 - 0           -1.69e-05  0.00522 -0.00324 0.99742
##  edlev   2 - 1           8.28e-03  0.01234 0.67111 0.50215
##  edlev   3 - 1           -1.44e-02  0.00633 -2.27731 0.02277
##  edlev   4 - 1           -3.43e-02  0.00778 -4.41678 < 0.001
##  edlev   5 - 1           -2.96e-02  0.00882 -3.35708 < 0.001
##  hisp    1 - 0           -4.49e-02  0.00527 -8.51433 < 0.001
## incfam $100,000 and over - $0 - $34,999 -6.17e-02  0.00624 -9.89274 < 0.001
## incfam $35,000 - $49,999 - $0 - $34,999 -3.08e-02  0.00568 -5.43109 < 0.001
## incfam $50,000 - $74,999 - $0 - $34,999 -5.38e-02  0.00566 -9.51247 < 0.001
## incfam $75,000 - $99,999 - $0 - $34,999 -5.81e-02  0.00690 -8.42709 < 0.001
## other   1 - 0           2.79e-03  0.01442 0.19321 0.84680

##      S    2.5 %  97.5 %
##  9.1  0.00376  0.01642
## 38.4  0.01879  0.03344
## 185.4 0.06863  0.08801
## 381.0 0.13720  0.16297
## 603.9 0.28184  0.32298
##     Inf  0.57328  0.62977
## 20.9 -0.05454 -0.02391
## 0.0 -0.01025  0.01022
## 1.0 -0.01590  0.03246
## 5.5 -0.02682 -0.00201
## 16.6 -0.04959 -0.01910
## 10.3 -0.04692 -0.01233
## 55.7 -0.05524 -0.03456
## 74.2 -0.07398 -0.04951
## 24.1 -0.04196 -0.01971
## 68.9 -0.06493 -0.04274
## 54.6 -0.07163 -0.04460
## 0.2 -0.02547  0.03105
##
## Columns: term, contrast, estimate, std.error, statistic, p.value, s.value, conf.low, conf.high
## Type: response

logit_mort <-
  feglm(mort ~ incfam + factor(edlev) + factor(agecat) +
    black + hisp + asian + other,
    data = nhis2010, vcov = 'hetero', family = 'logit')

## NOTE: 1,701 observations removed because of NA values (LHS: 362, RHS: 1,419).

summary(logit_mort) # logit results

## GLM estimation, family = binomial(link = "logit"), Dep. Var.: mort
## Observations: 22,655
## Standard-errors: Heteroskedasticity-robust
##                               Estimate Std. Error   z value Pr(>|z|)
## (Intercept)            -3.805985  0.209324 -18.182252 < 2.2e-16 ***
## incfam$35,000 - $49,999 -0.329330  0.066824 -4.928348 8.2928e-07 ***
## incfam$50,000 - $74,999 -0.666063  0.076015 -8.762250 < 2.2e-16 ***

```

```

## incfam$75,000 - $99,999 -0.752075  0.100595 -7.476249 7.6474e-14 ***
## incfam$100,000 and over -0.787600  0.090405 -8.711926 < 2.2e-16 ***
## factor(edlev)2          0.092624  0.135017  0.686018 4.9270e-01
## factor(edlev)3          -0.167995  0.071146 -2.361266 1.8213e-02 *
## factor(edlev)4          -0.434090  0.097586 -4.448303 8.6551e-06 ***
## factor(edlev)5          -0.359353  0.111136 -3.233459 1.2230e-03 **
## factor(agecat)30         0.605196  0.221428  2.733156 6.2731e-03 **
## factor(agecat)40         1.197880  0.210748  5.683950 1.3162e-08 ***
## factor(agecat)50         2.141488  0.201712 10.616549 < 2.2e-16 ***
## factor(agecat)60         2.821297  0.199807 14.120125 < 2.2e-16 ***
## factor(agecat)70         3.686436  0.200528 18.383665 < 2.2e-16 ***
## factor(agecat)80         4.979329  0.204565 24.341044 < 2.2e-16 ***
## black                    -0.014310  0.066288 -0.215869 8.2909e-01
## hisp                     -0.647340  0.083904 -7.715234 1.2076e-14 ***
## asian                    -0.623562  0.128269 -4.861373 1.1657e-06 ***
## other                    0.041697  0.184577  0.225908 8.2127e-01
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Log-Likelihood: -6,116.8  Adj. Pseudo R2: 0.287604
##                 BIC: 12,424.2  Squared Cor.: 0.275127

```

`avg_slopes(logit_mort) # logit marginal effects`

##	##	Term	Contrast	Estimate	Std. Error	z	Pr(> z )
##	##	agecat 30 - 20		0.00955	0.00316	3.019	0.00253
##	##	agecat 40 - 20		0.02602	0.00372	6.996	< 0.001
##	##	agecat 50 - 20		0.07899	0.00494	15.984	< 0.001
##	##	agecat 60 - 20		0.15079	0.00655	23.020	< 0.001
##	##	agecat 70 - 20		0.29723	0.01036	28.681	< 0.001
##	##	agecat 80 - 20		0.59276	0.01481	40.013	< 0.001
##	##	asian 1 - 0		-0.04361	0.00773	-5.643	< 0.001
##	##	black 1 - 0		-0.00114	0.00527	-0.216	0.82872
##	##	edlev 2 - 1		0.00829	0.01222	0.678	0.49789
##	##	edlev 3 - 1		-0.01417	0.00616	-2.300	0.02145
##	##	edlev 4 - 1		-0.03445	0.00777	-4.432	< 0.001
##	##	edlev 5 - 1		-0.02901	0.00884	-3.283	0.00103
##	##	hisp 1 - 0		-0.04672	0.00540	-8.660	< 0.001
##	##	incfam \$100,000 and over - \$0 - \$34,999	-0.06216	0.00648	-9.593	< 0.001	
##	##	incfam \$35,000 - \$49,999 - \$0 - \$34,999	-0.02897	0.00570	-5.081	< 0.001	
##	##	incfam \$50,000 - \$74,999 - \$0 - \$34,999	-0.05412	0.00575	-9.412	< 0.001	
##	##	incfam \$75,000 - \$99,999 - \$0 - \$34,999	-0.05987	0.00709	-8.438	< 0.001	
##	##	other 1 - 0	0.00336	0.01501	0.224	0.82287	
##	##	S 2.5 % 97.5 %					
##	##	8.6 0.00335 0.01574					
##	##	38.5 0.01873 0.03331					
##	##	188.6 0.06931 0.08868					
##	##	387.1 0.13795 0.16363					
##	##	598.5 0.27692 0.31754					
##	##	Inf 0.56372 0.62179					
##	##	25.8 -0.05875 -0.02846					
##	##	0.3 -0.01147 0.00919					
##	##	1.0 -0.01567 0.03224					
##	##	5.5 -0.02625 -0.00210					
##	##	16.7 -0.04968 -0.01921					

```

##      9.9 -0.04633 -0.01169
##      57.6 -0.05730 -0.03615
##     70.0 -0.07486 -0.04946
##    21.3 -0.04015 -0.01780
##    67.5 -0.06539 -0.04285
##    54.8 -0.07377 -0.04596
##     0.3 -0.02605  0.03277
##
## Columns: term, contrast, estimate, std.error, statistic, p.value, s.value, conf.low, conf.high
## Type: response

oddsratio <- exp(coef(logit_mort)) # logit odds ratios
ci <- exp(confint(logit_mort))
cbind(oddsratio, ci)

##                                     oddsratio      2.5 %      97.5 %
## (Intercept)                 0.02223729  0.01475383  0.03351652
## incfam$35,000 - $49,999    0.71940576  0.63109333  0.82007624
## incfam$50,000 - $74,999    0.51372696  0.44261710  0.59626118
## incfam$75,000 - $99,999    0.47138737  0.38703579  0.57412276
## incfam$100,000 and over   0.45493515  0.38106297  0.54312807
## factor(edlev)2            1.09704919  0.84197568  1.42939632
## factor(edlev)3            0.84535799  0.73532767  0.97185263
## factor(edlev)4            0.64785375  0.53507169  0.78440794
## factor(edlev)5            0.69812808  0.56148249  0.86802851
## factor(agecat)30          1.83161153  1.18673525  2.82691594
## factor(agecat)40          3.31308713  2.19201709  5.00750946
## factor(agecat)50          8.51209139  5.73242693  12.63962030
## factor(agecat)60          16.79862430 11.35528025 24.85132661
## factor(agecat)70          39.90237508 26.93451630 59.11372304
## factor(agecat)80          145.37674878 97.35736292 217.08064447
## black                     0.98579235  0.86568717  1.12256089
## hisp                      0.52343649  0.44406313  0.61699731
## asian                     0.53603155  0.41687685  0.68924389
## other                     1.04257899  0.72609993  1.49699910

# Fair/poor health analyses
ols_fpoor <-
  feols(fpoor ~ incfam + factor(edlev) + factor(agecat) +
         black + hisp + asian + other,
        data = nhis2010, vcov = 'hetero')

## NOTE: 1,426 observations removed because of NA values (LHS: 14, RHS: 1,419).

summary(ols_fpoor)

## OLS estimation, Dep. Var.: fpoor
## Observations: 22,930
## Standard-errors: Heteroskedasticity-robust
##                               Estimate Std. Error   t value Pr(>|t|) 
## (Intercept)                 0.241144  0.011218  21.496888 < 2.2e-16 ***
## incfam$35,000 - $49,999   -0.093818  0.007343 -12.776999 < 2.2e-16 ***
## incfam$50,000 - $74,999   -0.115876  0.006791 -17.062306 < 2.2e-16 ***
## incfam$75,000 - $99,999   -0.142442  0.007160 -19.894249 < 2.2e-16 ***
## incfam$100,000 and over   -0.154658  0.006617 -23.372571 < 2.2e-16 ***
## factor(edlev)2            -0.014593  0.019831  -0.735868 4.6182e-01

```

```

## factor(edlev)3      -0.122761  0.009691 -12.668072 < 2.2e-16 ***
## factor(edlev)4      -0.160108  0.010542 -15.187872 < 2.2e-16 ***
## factor(edlev)5      -0.175008  0.010850 -16.129124 < 2.2e-16 ***
## factor(agecat)30     0.029379  0.006457  4.550053 5.3908e-06 ***
## factor(agecat)40     0.075986  0.007024  10.817946 < 2.2e-16 ***
## factor(agecat)50     0.158984  0.007890  20.150903 < 2.2e-16 ***
## factor(agecat)60     0.161410  0.008581  18.809394 < 2.2e-16 ***
## factor(agecat)70     0.165609  0.010816  15.311735 < 2.2e-16 ***
## factor(agecat)80     0.167053  0.013467  12.404813 < 2.2e-16 ***
## black                 0.059795  0.007383  8.099259 5.8030e-16 ***
## hisp                  -0.011195  0.006998 -1.599758 1.0967e-01
## asian                 0.010366  0.008622  1.202340 2.2924e-01
## other                 0.059887  0.018916  3.165876 1.5482e-03 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## RMSE: 0.346702  Adj. R2: 0.114901

probit_fpoor <-
  feglm(fpoor ~ incfam + factor(edlev) + factor(agecat) +
         black + hisp + asian + other,
        data = nhis2010, vcov = 'hetero', family = 'probit')

## NOTE: 1,426 observations removed because of NA values (LHS: 14, RHS: 1,419).
summary(probit_fpoor)

## GLM estimation, family = binomial(link = "probit"), Dep. Var.: fpoor
## Observations: 22,930
## Standard-errors: Heteroskedasticity-robust
##                               Estimate Std. Error   z value Pr(>|z|)
## (Intercept)             -0.989583  0.054493 -18.159933 < 2.2e-16 ***
## incfam$35,000 - $49,999 -0.354259  0.031778 -11.147949 < 2.2e-16 ***
## incfam$50,000 - $74,999 -0.487437  0.033262 -14.654253 < 2.2e-16 ***
## incfam$75,000 - $99,999 -0.694225  0.046671 -14.874991 < 2.2e-16 ***
## incfam$100,000 and over -0.837998  0.045435 -18.444001 < 2.2e-16 ***
## factor(edlev)2          -0.022330  0.060059 -0.371802 0.71004013
## factor(edlev)3          -0.371939  0.031113 -11.954437 < 2.2e-16 ***
## factor(edlev)4          -0.627346  0.044523 -14.090392 < 2.2e-16 ***
## factor(edlev)5          -0.745788  0.055902 -13.340930 < 2.2e-16 ***
## factor(agecat)30         0.173580  0.051364  3.379405 0.00072643 ***
## factor(agecat)40         0.456868  0.049945  9.147343 < 2.2e-16 ***
## factor(agecat)50         0.825459  0.048763  16.928093 < 2.2e-16 ***
## factor(agecat)60         0.831443  0.050001  16.628616 < 2.2e-16 ***
## factor(agecat)70         0.815085  0.053535  15.225199 < 2.2e-16 ***
## factor(agecat)80         0.826804  0.059335  13.934393 < 2.2e-16 ***
## black                   0.236192  0.028609  8.255853 < 2.2e-16 ***
## hisp                   -0.002173  0.032986 -0.065863 0.94748712
## asian                  0.043994  0.050757  0.866760 0.38607379
## other                  0.271903  0.073617  3.693507 0.00022118 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Log-Likelihood: -8,764.8  Adj. Pseudo R2: 0.136072
## BIC: 17,720.3  Squared Cor.: 0.124998

avg_slopes(probit_fpoor) # probit marginal effects

```

```

##          Term      Contrast Estimate Std. Error      z Pr(>|z|)
## agecat 30 - 20           0.022485  0.00642  3.5034 <0.001
## agecat 40 - 20           0.070748  0.00705 10.0379 <0.001
## agecat 50 - 20           0.156646  0.00779 20.1151 <0.001
## agecat 60 - 20           0.158254  0.00840 18.8396 <0.001
## agecat 70 - 20           0.153874  0.00984 15.6360 <0.001
## agecat 80 - 20           0.157007  0.01207 13.0041 <0.001
## asian   1 - 0            0.009484  0.01112  0.8532  0.394
## black   1 - 0            0.053479  0.00687  7.7874 <0.001
## edlev   2 - 1           -0.006358  0.01703 -0.3734  0.709
## edlev   3 - 1           -0.093892  0.00857 -10.9580 <0.001
## edlev   4 - 1           -0.142910  0.01014 -14.0939 <0.001
## edlev   5 - 1           -0.161457  0.01093 -14.7755 <0.001
## hisp    1 - 0            -0.000461  0.00699 -0.0659  0.947
## incfam $100,000 and over - $0 - $34,999 -0.163387  0.00686 -23.8344 <0.001
## incfam $35,000 - $49,999 - $0 - $34,999 -0.086220  0.00721 -11.9625 <0.001
## incfam $50,000 - $74,999 - $0 - $34,999 -0.111974  0.00693 -16.1649 <0.001
## incfam $75,000 - $99,999 - $0 - $34,999 -0.145021  0.00767 -18.8980 <0.001
## other   1 - 0            0.063888  0.01892  3.3765 <0.001
##          S      2.5 %  97.5 %
## 11.1  0.00991  0.0351
## 76.3  0.05693  0.0846
## 296.5 0.14138  0.1719
## 260.6 0.14179  0.1747
## 180.7 0.13459  0.1732
## 126.0 0.13334  0.1807
## 1.3   -0.01230  0.0313
## 47.1   0.04002  0.0669
## 0.5   -0.03973  0.0270
## 90.4   -0.11069 -0.0771
## 147.4  -0.16278 -0.1230
## 161.7  -0.18287 -0.1400
## 0.1   -0.01416  0.0132
## 414.7  -0.17682 -0.1500
## 107.1  -0.10035 -0.0721
## 192.8  -0.12555 -0.0984
## 262.2  -0.16006 -0.1300
## 10.4   0.02680  0.1010
##
## Columns: term, contrast, estimate, std.error, statistic, p.value, s.value, conf.low, conf.high
## Type: response
logit_fpoor <-
  feglm(fpoor ~ incfam + factor(edlev) + factor(agecat) +
         black + hisp + asian + other,
        data = nhis2010, vcov = 'hetero', family = 'logit')

## NOTE: 1,426 observations removed because of NA values (LHS: 14, RHS: 1,419).
summary(logit_fpoor)

## GLM estimation, family = binomial(link = "logit"), Dep. Var.: fpoor
## Observations: 22,930
## Standard-errors: Heteroskedasticity-robust

```

```

##                               Estimate Std. Error   z value Pr(>|z|)
## (Intercept)                 -1.751317  0.103826 -16.867805 < 2.2e-16 ***
## incfam$35,000 - $49,999    -0.629661  0.057292 -10.990381 < 2.2e-16 ***
## incfam$50,000 - $74,999    -0.881977  0.062021 -14.220517 < 2.2e-16 ***
## incfam$75,000 - $99,999    -1.310141  0.092879 -14.105905 < 2.2e-16 ***
## incfam$100,000 and over   -1.609333  0.093835 -17.150666 < 2.2e-16 ***
## factor(edlev)2            -0.056218  0.101020 -0.556497 5.7787e-01
## factor(edlev)3            -0.640555  0.053029 -12.079315 < 2.2e-16 ***
## factor(edlev)4            -1.118474  0.082504 -13.556632 < 2.2e-16 ***
## factor(edlev)5            -1.360534  0.109570 -12.417022 < 2.2e-16 ***
## factor(agecat)30           0.343186  0.102040  3.363262 7.7027e-04 ***
## factor(agecat)40           0.884451  0.098060  9.019491 < 2.2e-16 ***
## factor(agecat)50           1.557965  0.095008 16.398234 < 2.2e-16 ***
## factor(agecat)60           1.548108  0.096927 15.971821 < 2.2e-16 ***
## factor(agecat)70           1.502036  0.102043 14.719707 < 2.2e-16 ***
## factor(agecat)80           1.496835  0.110544 13.540656 < 2.2e-16 ***
## black                      0.412738  0.050307  8.204395 2.3176e-16 ***
## hisp                       -0.013578  0.058954 -0.230319 8.1784e-01
## asian                      0.068276  0.093771  0.728118 4.6654e-01
## other                      0.495755  0.130652  3.794460 1.4797e-04 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Log-Likelihood: -8,763.6  Adj. Pseudo R2: 0.136188
##                  BIC: 17,718.0  Squared Cor.: 0.125558
avg_slopes(logit_fpoor) # logit marginal effects

```

```

##
##      Term                           Contrast Estimate Std. Error       z Pr(>|z|)
## agecat 30 - 20                     0.02259   0.00643   3.515 <0.001
## agecat 40 - 20                     0.07194   0.00710  10.137 <0.001
## agecat 50 - 20                     0.16085   0.00785  20.487 <0.001
## agecat 60 - 20                     0.15932   0.00839  18.980 <0.001
## agecat 70 - 20                     0.15227   0.00963  15.814 <0.001
## agecat 80 - 20                     0.15148   0.01156  13.099 <0.001
## asian  1 - 0                      0.00824   0.01149   0.717  0.473
## black   1 - 0                      0.05257   0.00682   7.712 <0.001
## edlev  2 - 1                      -0.00913   0.01629  -0.560  0.575
## edlev  3 - 1                      -0.09130   0.00829 -11.009 <0.001
## edlev  4 - 1                      -0.14111   0.01004 -14.051 <0.001
## edlev  5 - 1                      -0.16085   0.01093 -14.711 <0.001
## hisp   1 - 0                      -0.00161   0.00697  -0.231  0.817
## incfam $100,000 and over - $0 - $34,999 -0.16624   0.00682 -24.393 <0.001
## incfam $35,000 - $49,999 - $0 - $34,999 -0.08621   0.00719 -11.993 <0.001
## incfam $50,000 - $74,999 - $0 - $34,999 -0.11255   0.00696 -16.173 <0.001
## incfam $75,000 - $99,999 - $0 - $34,999 -0.14773   0.00767 -19.266 <0.001
## other  1 - 0                      0.06613   0.01928   3.430 <0.001
##      S  2.5 % 97.5 %
##      11.2 0.00999 0.0352
##      77.8 0.05803 0.0858
##      307.4 0.14546 0.1762
##      264.4 0.14287 0.1758
##      184.7 0.13339 0.1711
##      127.8 0.12881 0.1741
##      1.1 -0.01429 0.0308

```

```

##   46.2  0.03921  0.0659
##    0.8 -0.04106  0.0228
##   91.2 -0.10755 -0.0750
##  146.6 -0.16079 -0.1214
##  160.3 -0.18228 -0.1394
##    0.3 -0.01528  0.0121
## 434.2 -0.17960 -0.1529
## 107.7 -0.10030 -0.0721
## 193.0 -0.12619 -0.0989
## 272.3 -0.16276 -0.1327
##   10.7  0.02835  0.1039
##
## Columns: term, contrast, estimate, std.error, statistic, p.value, s.value, conf.low, conf.high
## Type: response
oddsratio <- exp(coef(logit_fpoor)) # logit odds ratios
ci <- exp(confint(logit_fpoor))
cbind(oddsratio, ci)

##                               oddsratio      2.5 %     97.5 %
## (Intercept)           0.1735452 0.1415910 0.2127108
## incfam$35,000 - $49,999 0.5327724 0.4761838 0.5960857
## incfam$50,000 - $74,999 0.4139636 0.3665805 0.4674713
## incfam$75,000 - $99,999 0.2697820 0.2248818 0.3236470
## incfam$100,000 and over 0.2000210 0.1664191 0.2404076
## factor(edlev)2        0.9453334 0.7755258 1.1523218
## factor(edlev)3        0.5269996 0.4749762 0.5847211
## factor(edlev)4        0.3267780 0.2779876 0.3841318
## factor(edlev)5        0.2565237 0.2069480 0.3179755
## factor(agecat)30       1.4094312 1.1539515 1.7214729
## factor(agecat)40       2.4216534 1.9982197 2.9348150
## factor(agecat)50       4.7491453 3.9422525 5.7211914
## factor(agecat)60       4.7025646 3.8889286 5.6864283
## factor(agecat)70       4.4908240 3.6767768 5.4851032
## factor(agecat)80       4.4675262 3.5972622 5.5483278
## black                  1.5109489 1.3690788 1.6675202
## hisp                   0.9865136 0.8788639 1.1073491
## asian                  1.0706612 0.8909104 1.2866786
## other                  1.6417367 1.2708436 2.1208741

```

## Question 6

It is possible to use the `hypotheses()` function with coefficients on the categories of factor variables, but it's easier to work with dummy variables. For pedagogical purposes, I will generate income and education category dummies, and then I will re-run the model.

```

nthis2010 <-
  nthis2010 %>%
  mutate(inc_35_50 = ifelse(incfam=="$35,000 - $49,999",1,0),
        inc_50_75 = ifelse(incfam=="$50,000 - $74,999",1,0),
        inc_75_100 = ifelse(incfam=="$75,000 - $99,999",1,0),
        inc_gt_100 = ifelse(incfam=="$100,000 and over",1,0),
        ed_12 = ifelse(edlev==2,1,0),
        ed_13_15 = ifelse(edyrs>12 & edyrs<16,1,0),
        ed_16 = ifelse(edyrs==16,1,0),
        ed_gt16 = ifelse(edyrs>16,1,0))

logit_mort2 <-
  feglm(mort ~ inc_35_50 + inc_50_75 + inc_75_100 + inc_gt_100 +
         ed_12 + ed_13_15 + ed_16 + ed_gt16 + factor(agecat) +
         black + hisp + asian + other,
        data = nthis2010, vcov = 'hetero', family = 'logit')

## NOTE: 1,701 observations removed because of NA values (LHS: 362, RHS: 1,419).
summary(logit_mort2)

## GLM estimation, family = binomial(link = "logit"), Dep. Var.: mort
## Observations: 22,655
## Standard-errors: Heteroskedasticity-robust
##           Estimate Std. Error   z value Pr(>|z|)
## (Intercept) -3.805985  0.209324 -18.182252 < 2.2e-16 ***
## inc_35_50    -0.329330  0.066824 -4.928348 8.2928e-07 ***
## inc_50_75    -0.666063  0.076015 -8.762250 < 2.2e-16 ***
## inc_75_100   -0.752075  0.100595 -7.476249 7.6474e-14 ***
## inc_gt_100   -0.787600  0.090405 -8.711926 < 2.2e-16 ***
## ed_12        0.092624  0.135017  0.686018 4.9270e-01
## ed_13_15     -0.167995  0.071146 -2.361266 1.8213e-02 *
## ed_16        -0.434090  0.097586 -4.448303 8.6551e-06 ***
## ed_gt16      -0.359353  0.111136 -3.233459 1.2230e-03 **
## factor(agecat)30 0.605196  0.221428  2.733156 6.2731e-03 **
## factor(agecat)40 1.197880  0.210748  5.683950 1.3162e-08 ***
## factor(agecat)50 2.141488  0.201712 10.616549 < 2.2e-16 ***
## factor(agecat)60 2.821297  0.199807 14.120125 < 2.2e-16 ***
## factor(agecat)70 3.686436  0.200528 18.383665 < 2.2e-16 ***
## factor(agecat)80 4.979329  0.204565 24.341044 < 2.2e-16 ***
## black       -0.014310  0.066288 -0.215869 8.2909e-01
## hisp        -0.647340  0.083904 -7.715234 1.2076e-14 ***
## asian       -0.623562  0.128269 -4.861373 1.1657e-06 ***
## other        0.041697  0.184577  0.225908 8.2127e-01
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Log-Likelihood: -6,116.8  Adj. Pseudo R2: 0.287604
##                 BIC: 12,424.2  Squared Cor.: 0.275127

```

This model is the same as the one above. We just coded the categorical variables as dummies. The difference

in log odds between Groups A and B is given by:

```
hypotheses(logit_mort2, "asian - black - ed_16 - inc_gt_100 = 0")  
  
##                                     Term Estimate Std. Error      z Pr(>|z|)    S  
##  asian - black - ed_16 - inc_gt_100 = 0     0.612      0.187 3.27  0.00108 9.9  
##  2.5 % 97.5 %  
##  0.245   0.98  
##  
## Columns: term, estimate, std.error, statistic, p.value, s.value, conf.low, conf.high
```

Since this is positive, we conclude that the poorer, less-educated Asian group have higher mortality risk than richer, more-educated Black group. The difference is statistically significant at the 5% level. If we exponentiate this difference, we get:

```
exp(.612)
```

```
## [1] 1.844116
```

which implies that the odds of dying are 84 percent higher for the poorer, less-educated, Asian group. You did not need to state this quantity in your answer.

It's likely that this model is not the best for testing differences between these groups. It would be better to include interactions of race and income.

### **Question 7**

We probably should not interpret these results as causal. One problem is that there are many confounding variables that we do not observe but may jointly determine health and income, for instance place of birth. Another problem is that there may be reverse causality, i.e. health may affect income.

## Question 8

I use the logit model again, and I exponentiated the coefficients for interpretability. I control for insurance status, smoking status, exercise, bacon consumption, and obesity. To keep the samples the same in the regressions with and without the additional control variables, I run the long regression and the short regression on the same sample, which required subsetting the data first. This was not required.

Smoking, exercise, and obesity predicted mortality: ever smoking doubled the odds of death, ever exercising reduced the odds by 38%, ever binge drinking raised the odds by 20%, and obesity raised the odds by 17%. In contrast, uninsurance did not significantly associate with mortality. The patterns explain part of the socioeconomic gradient in health. After controlling for these variables, the odds ratio on the highest income category rose from 0.44 to 0.51 and that on the >16 years of education dummy rose from 0.65 to 0.87. Because the odds ratios are

moving closer to 1, mortality gaps are smaller after controlling for health behaviors. Health behavior explains a larger share of the education-mortality relationship than the income-mortality relationship.

```
# Recode behavior variables as 0/1 dummies
nhis2010 <-
  nhis2010 %>%
    mutate(exev = ifelse(vig10fwk > 0, 1, 0), # ever exercise
          bingev = ifelse(alc5upyr > 0, 1, 0), # ever binge drink
          obese = ifelse(bmi >= 30, 1, 0) ) # obese if bmi>=30

# Subset data to non-missing obs
nhis_subset <-
  nhis2010 %>%
    drop_na(mort, incfam, edlev, agecat, white, black, hisp,
            uninsured, smokev, exev, bingev, obese)

# Run regression without health behaviors, report odds ratios and CIs
logit_model1 <-
  feglm(mort ~ incfam + factor(edlev) + factor(agecat) +
         black + hisp + asian + other,
        data = nhis_subset, vcov = 'hetero', family = 'logit')
cbind(exp(coef(logit_model1)), exp(confint(logit_model1)))

##                                exp(coef(logit_model1))      2.5 %      97.5 %
## (Intercept)                  0.02748355  0.01661693  0.04545637
## incfam$35,000 - $49,999     0.66322885  0.54218269  0.81129942
## incfam$50,000 - $74,999     0.46099871  0.37185553  0.57151177
## incfam$75,000 - $99,999     0.48557368  0.37244428  0.63306598
## incfam$100,000 and over    0.44455744  0.35282746  0.56013589
## factor(edlev)2              0.96438136  0.60932015  1.52634275
## factor(edlev)3              0.77210731  0.59801353  0.99688331
## factor(edlev)4              0.59568643  0.43914635  0.80802747
## factor(edlev)5              0.65164663  0.46880033  0.90580851
## factor(agecat)30             1.57269409  0.95956029  2.57760429
## factor(agecat)40             2.63158819  1.63964737  4.22362549
## factor(agecat)50             6.78294289  4.33401583  10.61563132
## factor(agecat)60             12.36096673  7.92124302  19.28908102
## factor(agecat)70             27.68028609  17.61740927  43.49097111
## factor(agecat)80             168.38644050 104.27751900 271.90897534
## black                        1.26354391  1.03731708  1.53910818
## hisp                         0.52159771  0.39581901  0.68734489
## asian                        0.57683989  0.37377527  0.89022544
## other                        0.89862068  0.51457521  1.56929271
```

```
# Run regression with health behaviors
logit_model2 <-
  feglm(mort ~ incfam + factor(edlev) + factor(agecat) +
    black + hisp + asian + other +
    uninsured + smokev + exev + bingev + obese,
    data = nhis_subset, vcov = 'hetero', family = 'logit')
cbind(exp(coef(logit_model2)), exp(confint(logit_model2)))
```

	exp(coef(logit_model2))	2.5 %	97.5 %
## (Intercept)	0.01833452	0.01075589	0.03125309
## incfam\$35,000 - \$49,999	0.69135905	0.56413112	0.84728058
## incfam\$50,000 - \$74,999	0.48912590	0.39364126	0.60777201
## incfam\$75,000 - \$99,999	0.52185988	0.39624955	0.68728844
## incfam\$100,000 and over	0.50827402	0.40075975	0.64463179
## factor(edlev)2	1.00298351	0.64435663	1.56120985
## factor(edlev)3	0.84234194	0.65295651	1.08665728
## factor(edlev)4	0.75235950	0.55353964	1.02259131
## factor(edlev)5	0.87234154	0.62405405	1.21941323
## factor(agecat)30	1.53659614	0.93473555	2.52598470
## factor(agecat)40	2.45293857	1.52163988	3.95422577
## factor(agecat)50	5.80128711	3.68565060	9.13134093
## factor(agecat)60	10.47090258	6.64464479	16.50047582
## factor(agecat)70	23.79172112	14.92407194	37.92838820
## factor(agecat)80	163.19634302	99.23440763	268.38520035
## black	1.26439402	1.03558439	1.54375854
## hisp	0.56489548	0.42772439	0.74605731
## asian	0.62696144	0.40244087	0.97674136
## other	0.80265160	0.45323793	1.42143793
## uninsured	0.90131106	0.71650301	1.13378675
## smokev	2.00578259	1.72220459	2.33605451
## exev	0.61552368	0.52115416	0.72698143
## bingev	1.19625658	1.00757484	1.42027147
## obese	1.17465753	1.00388539	1.37447994