

Model Selection Methods

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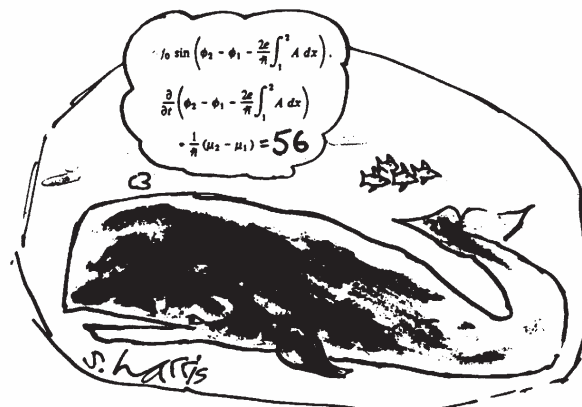
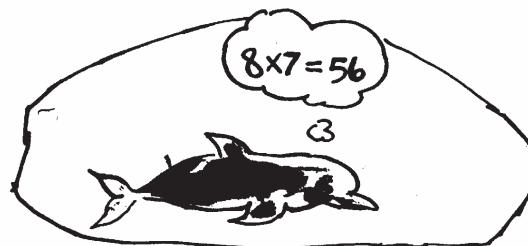
In collaboration with Mark Pitt

Amsterdam Workshop on Model Selection
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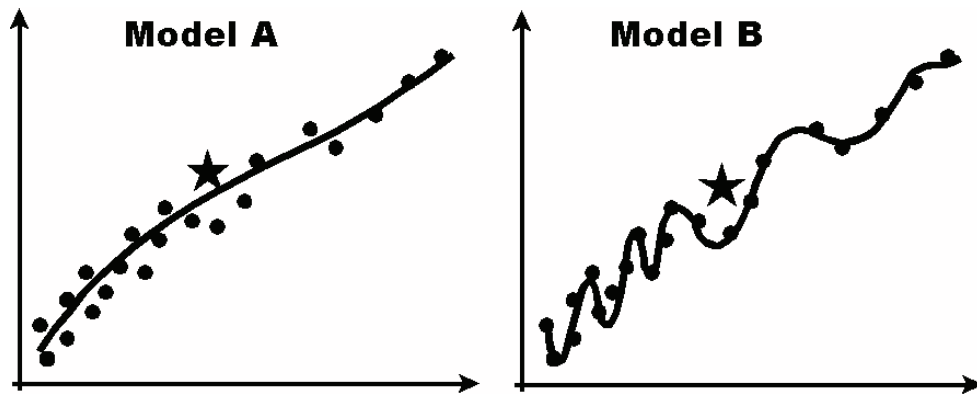
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Whale's Views of Model Selection



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Preview of Conclusion:

“Thou shall not select the **best-fitting** model but shall select the **best-predicting** model.”

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Overview

- **Part 1: Non-technical Introduction to Model Selection**
- **Part 2: “Technical” Tour of Model Selection Methods**
- **Part 3: Example Application**
- **Part 4: Conclusions**

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Part 1:

Non-technical Introduction to Model Selection

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Terminology

- **Model Selection**
- **Model Choice**
- **Model Comparison**



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What is a Quantitative Model?

- Mathematical instantiations of key assumptions and principles embodied in the theory from which it evolved.
- A formalization of a theory that enables the exploration of its operation.

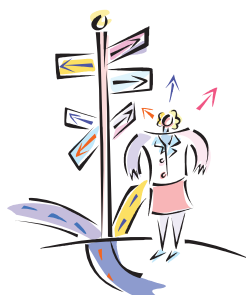
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Why Modeling?

- To infer the underlying structural properties of a mental process from behavioral data that were thought to have been generated by that process.

Often entertain **multiple models** as possible explanations of observed data



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Model Selection Problem

- **Q:** How should we choose between differing explanations (models) of data?
- **A:** Select the one, among candidate models, that “best” captures the underlying regularities.

How to identify such a model?

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Goodness of Fit (GOF) Measures as Methods of Model Selection

Examples of GOF measures:

- Percent Variance Accounted For (**PVAF**)

$$PVA = 100 * \left(\frac{\sum_{i=1}^m (obs_i - pred_i)^2}{\sum_{i=1}^m (obs_i - obs_{mean})^2} \right)$$

- Root Mean Square Deviation (**RMSD**)

$$RMSD = \sqrt{\frac{\sum_{i=1}^N (obs_i - pred_i)^2}{N}}$$

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Problems with GOF as a Model Selection Method

Data: **Noise** **&** **Regularity**
(sampling error) (underlying mental process)

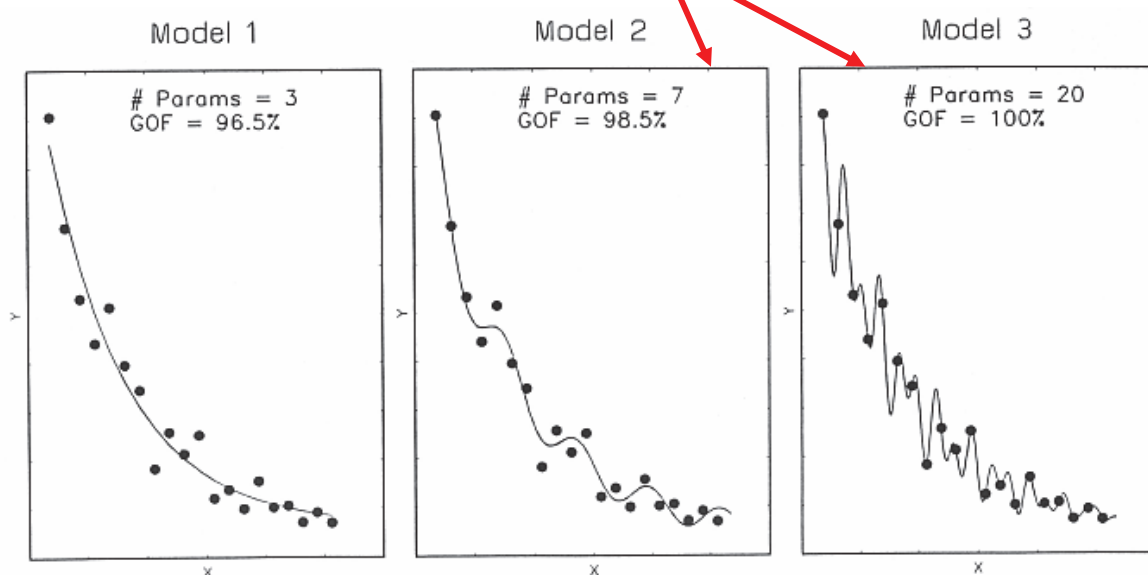
GOF = fit to noise + fit to regularity

Properties of the model that have nothing to do with its ability to capture the underlying regularities can improve fit.

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Fit improves with more parameters
(i.e, **over-fitting**)



Model 1: $Y = ae^{-bX} + c$

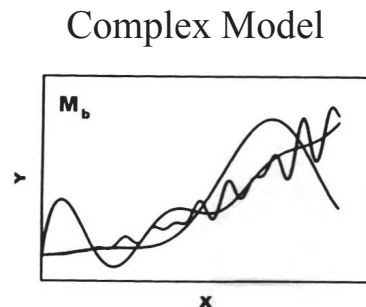
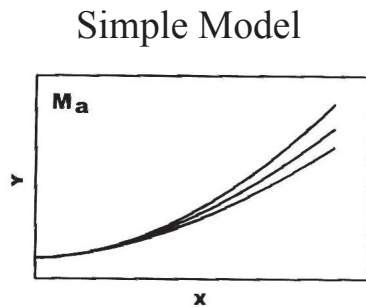
Model 2: $Y = ae^{-bX} + c + dX^{-e} \cdot \sin(f \cdot X + g)$

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Model Complexity

Complexity: A model's inherent **flexibility** that enables it to fit a wide range of data patterns



Complexity: # of parameters + functional form

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Complexity: More than number of parameters?

$$M1: y = ax + b$$

$$M2: y = ax^b$$

$$M3: y = \sin(\cos ax)^a \exp(-bx) / x^b$$

Are these all equally complex?

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Wanted:

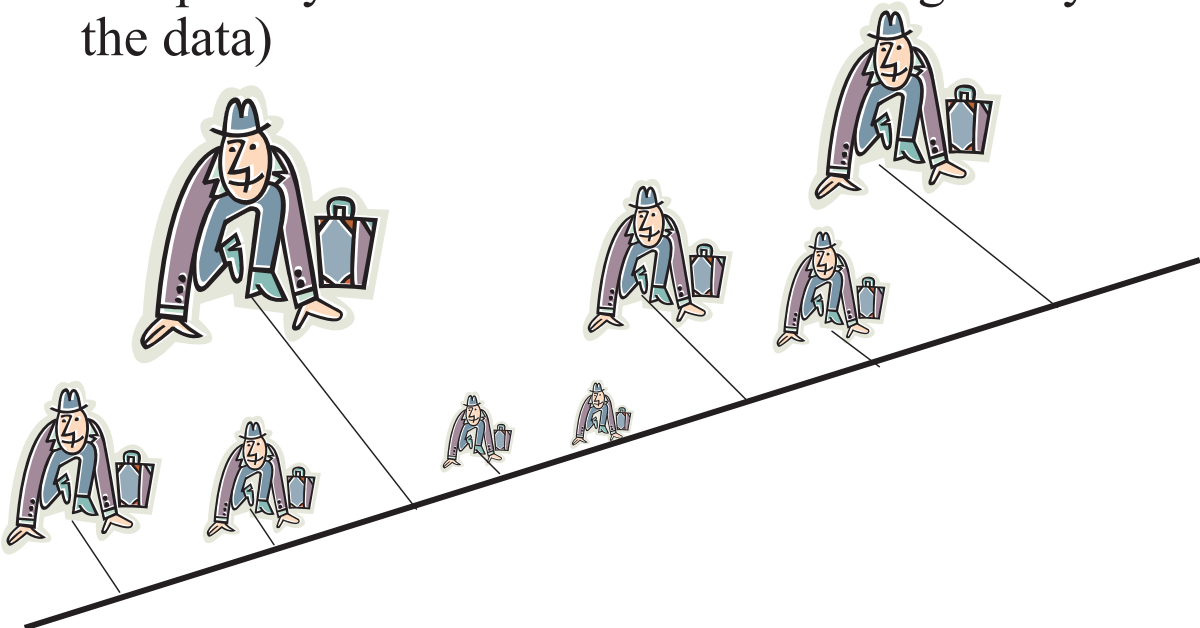
A Method of Model Selection that
Takes into Account Effects of
Complexity

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Placing Models on an Equal Footing

Penalize models for excess complexity (i.e., more complexity than is needed to fit the regularity in the data)



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Akaike Information Criterion (AIC) as a Method of Model Selection

Akaike (1973):

$$AIC = \underbrace{-2 \ln f(y | \hat{\theta})}_{\text{Goodness of fit (ML)}} + \underbrace{2k}_{\substack{\text{\# of parameters} \\ \downarrow \\ \text{Model Complexity}}}$$

The model that minimizes AIC should be preferred

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Bayesian Information Criterion (BIC)

Schwarz (1978):

$$BIC = \underbrace{-2 \ln f(y | \hat{\theta})}_{\text{Goodness of fit (ML)}} + \underbrace{k \ln n}_{\text{Model Complexity}}$$

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Selection Criteria as Formal Implementations of Principle of **Occam's Razor**



“Entities should not be multiplied beyond necessity”

(William of Occam, ca. 1290-1349)

“Select the **simplest** model that describes the data **sufficiently well.**”

$$AIC = -2 \ln f(y | \hat{\theta}) + 2k$$

$$BIC = -2 \ln f(y | \hat{\theta}) + k \ln n$$

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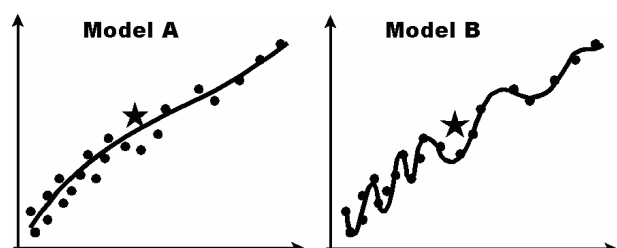
What Do AIC and BIC Measure?

They estimate a model's **generalizability** – the model's ability to fit all “future” data samples from the same underlying process, not just the current data sample.

Generalizability

= ‘**proximity**’ to underlying process

= **Predictive accuracy**



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“An important goal of scientific theorizing is to identify hypotheses that generate accurate predictions.”

“Overfitting is a sin precisely because it undermines the goal of predictive accuracy.”

(both from Hitchcock & Sober, 2004)

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Definition of generalizability

Formally, a model's generalizability may be defined as

$$E[D(M, T)] = \int D(f_M(y | \hat{\theta}), f_T(y)) f_T(y) dy$$

As mean discrepancy between the model of interest and the true model under some *discrepancy function* D satisfying

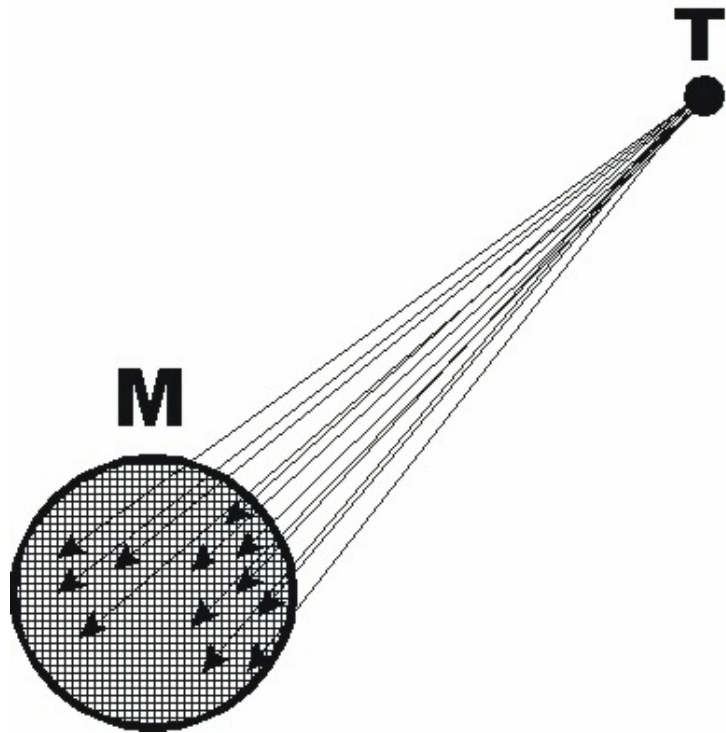
$$D(f, g) > D(f, f) = 0 \text{ for } f \neq g$$

(e.g., Kullback-Liebler information distance)

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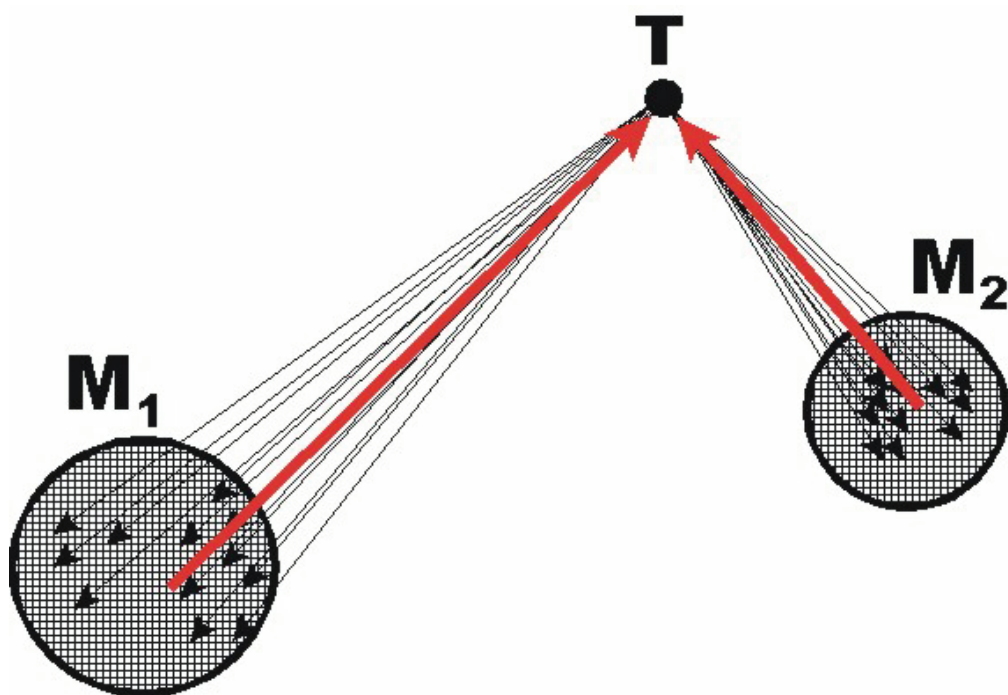
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“Geometric” Definition of Generalizabilty



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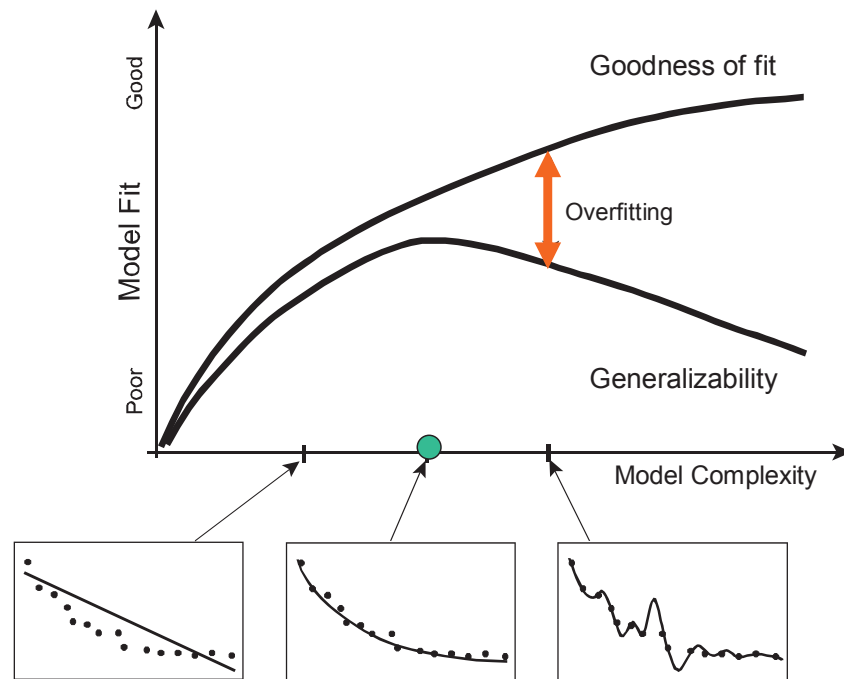
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Relationship between Goodness of Fit and Generalizability



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Part 2: “Technical” Tour of Model Selection Methods

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Selections Methods to be discussed

- AIC
- Cross-validation
- Bootstrap
- Bayesian Methods (Bayes Factor, BIC)
- **Minimum Description Length**

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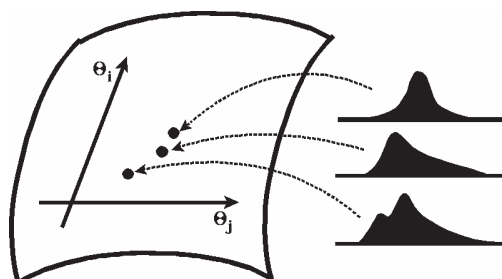
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Formal Definition of A Statistical Model

A model is defined as a parametric collection of probability distributions, indexed by model parameters:

$$M = \{f(y | \theta) | \theta \in \Omega\}$$

forming a *Riemannian manifold*, embedded in the space of probability distributions (Rao, 1945; Efron, 1975; Amari, 1980)



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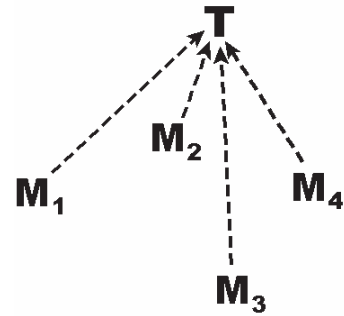
Akaike Information Criterion (AIC)

(Akaike, 1973)

AIC derived as asymptotic approximation of Kullback-Liebler information distance between the model of interest and the truth:

$$KL(M, T | x) = \int f_T(y) \ln \frac{f_T(y)}{f_M(y | \hat{\theta}(x))} dy$$

$$2 \cdot E[KL(M, T | x)] = 2 \cdot \int KL(M, T | x) f_T(x) dx$$
$$= \text{AIC} + (\text{higher order terms})$$



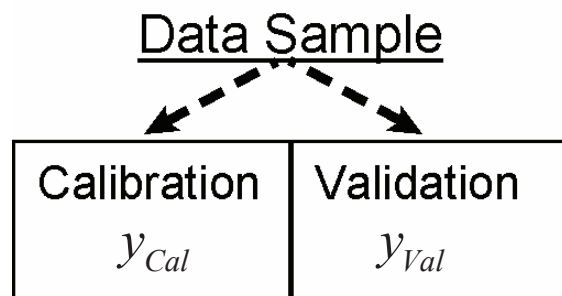
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Cross-validation (CV)

(Stone, 1974; Geisser, 1975)

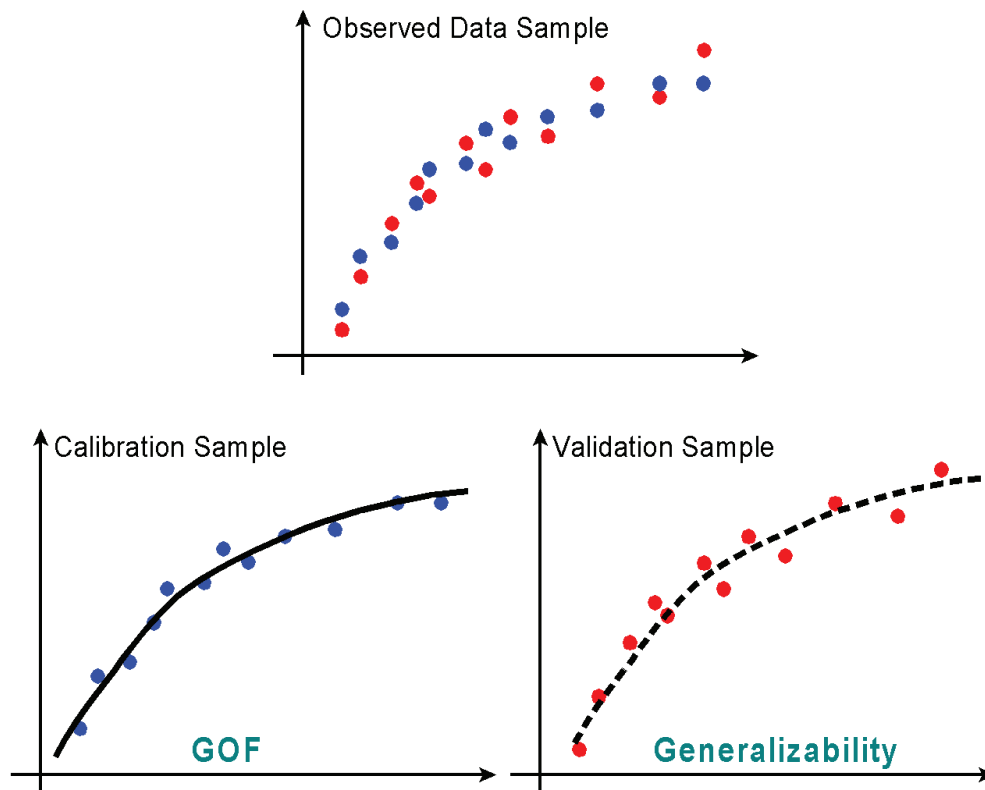
- Sampling-based method of estimating generalizability
- No explicit measure of model complexity, unlike AIC



$$CV = -\ln f(y_{Val} | \hat{\theta}(y_{Cal}))$$

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Features of CV

- Pros
 - Easy to use
 - Sensitive to functional form as well as number of parameters
 - Asymptotically equivalent to AIC
- Cons
 - Sensitive to the partitioning used
 - Averaging over multiple partitions
 - *Leave-one-out CV*, instead of *split-half CV*
 - Instability of the estimate due to “loss” of data

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Bootstrap Model Selection (BMS)

(Efron, 1983; Shao, 1996)

Similar to CV,

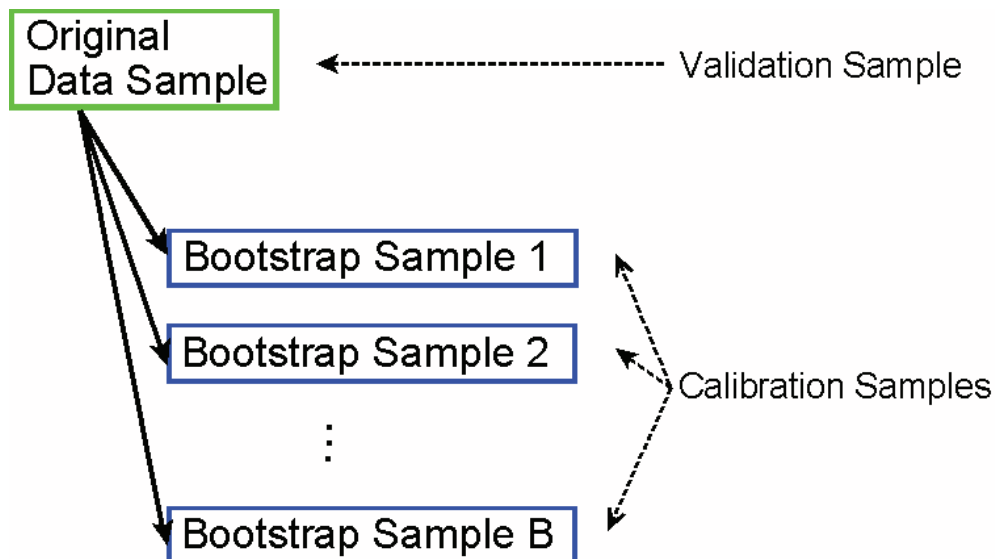
- **Res**ampling-based method of estimating generalizability
- No explicit measure of model complexity

Unlike CV,

- Full use of data sample in estimating generalizability

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$$BMS = -\frac{1}{B} \sum_{i=1}^N \ln f(y_{Original} | \hat{\theta}(y_{Boot_i}))$$

CV with n , not $n/2$

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Bayesian Methods

(Kass & Raftery, 1995)



- In Bayesian model selection, each model is evaluated based on its *marginal likelihood* defined as

$$P(y_{obs} | M_j) = \int f(y_{obs} | \theta, M_j) \pi(\theta | M_j) d\theta, \quad j = 1, 2, \dots, J$$

- Model selection is then based on the ratio of two marginal likelihoods or *Bayes factor (BF)*

$$BF_{ij} \equiv \frac{P(y_{obs} | M_i)}{P(y_{obs} | M_j)}$$

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- Under the assumption of equal model priors, BF is reduced to the *posterior odds*:

$$\begin{aligned} BF_{ij} &= \frac{P(y_{obs} | M_i)}{P(y_{obs} | M_j)} \\ &= \frac{P(M_i | y_{obs})}{P(M_j | y_{obs})} \frac{P(M_i)}{P(M_j)} \quad (\text{from Bayes rule}) \\ &= \frac{P(M_i | y_{obs})}{\underline{P(M_j | y_{obs})}} \end{aligned}$$

- Therefore, the model that maximizes marginal likelihood is the one with highest probability of being “true” given observed data

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Features of Bayes Factor

- Pros
 - No optimization (i.e., no maximum likelihood)
 - No explicit measure of model complexity
 - No overfitting, by averaging likelihood function across parameters
- Cons
 - Issue of parameter prior (virtue or vice?)
 - Non-trivial computations requiring numerical integration

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BIC as an approximation of BF

- A large sample approximation of the marginal likelihood yields the easily-computable *Bayesian information criterion (BIC)*:

$$\begin{aligned} -2\ln P(y_{obs} | M_j) &= -2\ln \int f(y_{obs} | \theta, M_j) \pi(\theta | M_j) d\theta \\ &= \underbrace{-2\ln f(y_{obs} | \hat{\theta}, M_j) + k \ln n + (higher\ order\ terms)}_{BIC} \end{aligned}$$

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Selections Methods to be discussed

- AIC
- Cross-validation
- Bootstrap
- Bayesian Methods (Bayes Factor, BIC)
- **Minimum Description Length**

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Minimum Description Length (MDL)

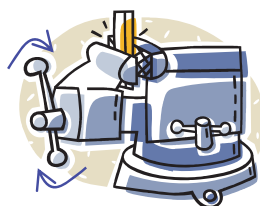
(Rissanen, 1978, 1989, 1996, 2001; Hansen & Yu, 2001)



- Algorithmic coding theory
- Models and data as compressible codes
- Regularities (redundancy) can be used to compressed the data

The MDL Principle:

“The best model is the one that provides the shortest description length of the data in bits by “compressing” the data as tightly as possible.”



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Information Theoretic Motivation

MDL can be motivated from a communication game:

Task: A sender tries to transmit data to a receiver



How many bits are needed to allow the receiver to fully reconstruct the data?

Goal: What is the most efficient (shortest) coding strategy?

MDL idea: "Find a code (i.e., model) that takes advantage of the structure in the data, thereby requiring fewer bits to describe the data."

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The Basic Idea

Seq 1: 000100100010001...000100010001

Seq 2: 011101001101001...100110100101

(Coded Seq 1): *"for i= 1 to 100; print '0001'; next; halt"*

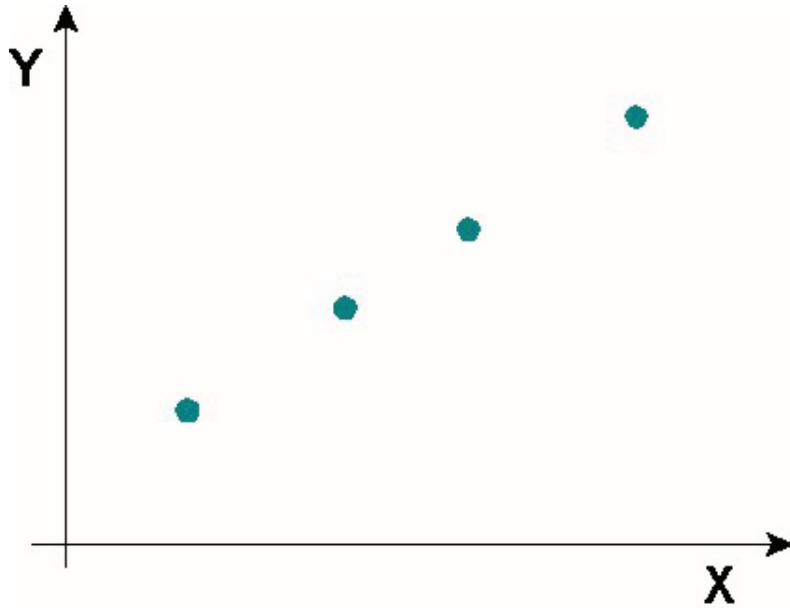
(Coded Seq 2): (not compressible! -- coin tossing outcomes)

- More regularity or redundancy in Seq 1 than Seq 2
- Shorter description for Seq 1 than Seq 2

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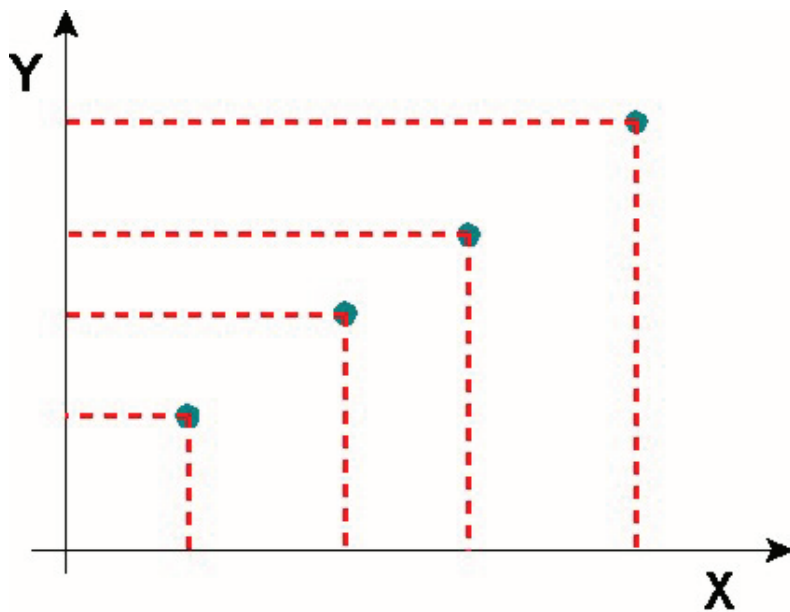
How to *describe* data?



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Raw method (no compression)



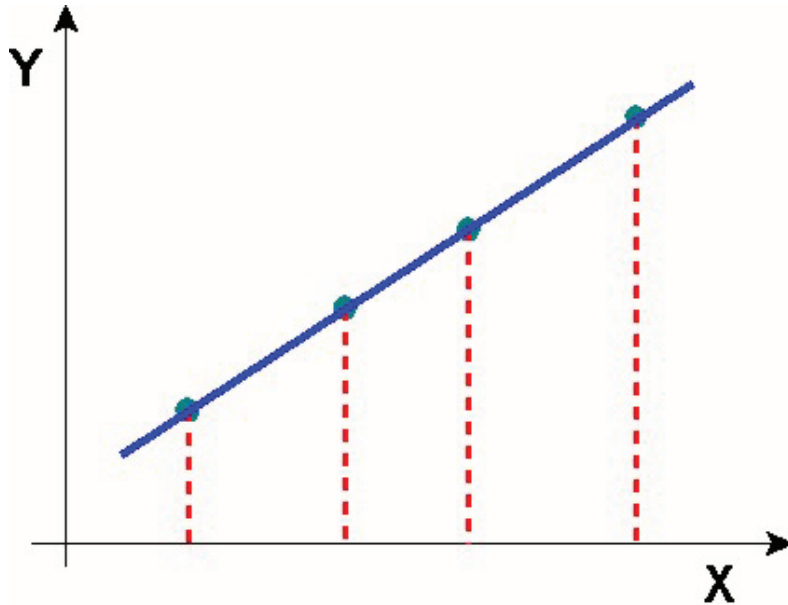
Overall description length (ODL1):

$$\text{ODL1} = \text{DL}(X_1, Y_1) + \text{DL}(X_2, Y_2) + \dots + \text{DL}(X_n, Y_n)$$

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Regularity-based Method (compressed)



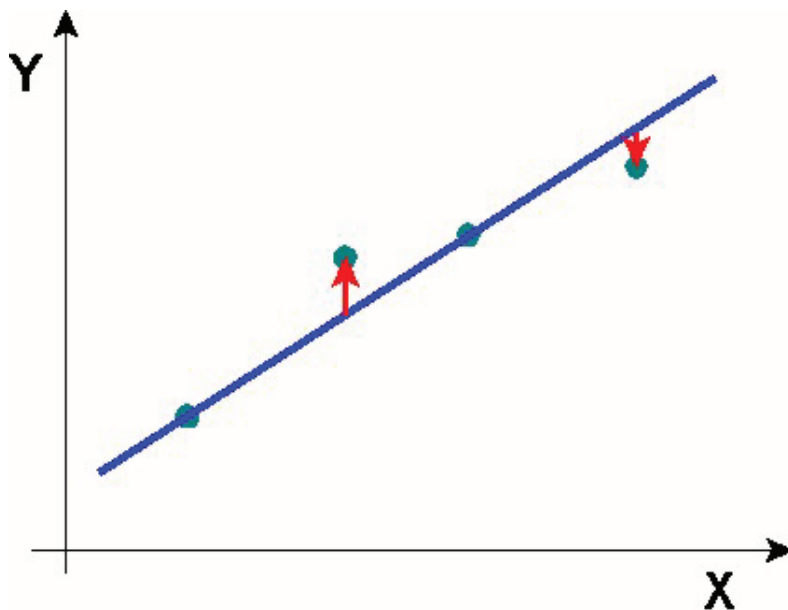
Overall description length (ODL2):

$$OLD2 = DL(X_1) + DL(X_2) + \dots + DL(X_n) + DL(Y_i = aX_i + b, i=1, \dots, n)$$

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How about ***noisy*** data?



Overall description length (ODL3):

$$OLD3 = OLD2 + DL(\text{deviations})$$

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Formally, the description length of data consists of two components:

- **DL(M)**: Description length of the model itself
- **DL(D|M)**: Description length of the data when encoded with the help of the model

Overall description length (OVD):

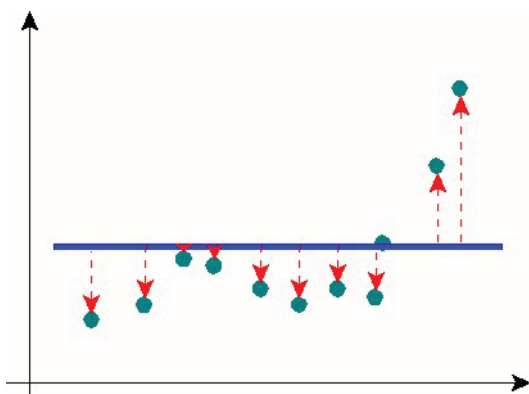
$$\text{OVD} = \text{DL(M)} + \text{DL(D|M)}$$

(expected patterns) (deviations)
 (model complexity) (fit)

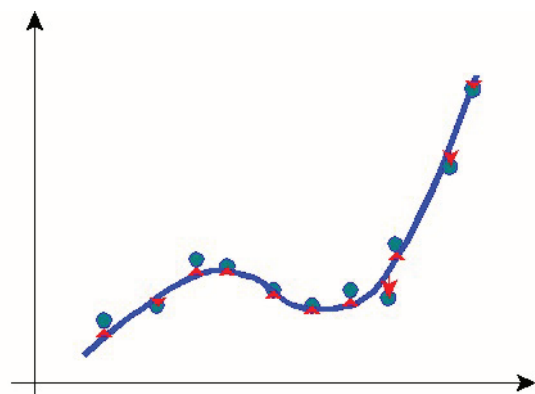
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$$\text{M1: } y = \theta_0 + e$$



$$\text{M2: } y = \sum_{i=0}^k \theta_i x^i + e$$



	DL(M)	DL(D M)	OVD
M1	2.0	14.5	16.5 bits
M2	7.8	3.1	10.9 bits

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Two Implementations of MDL for Parametric Models

- FIA (Fisher Information Approximation; Rissanen, 1996)
- NML (Normalized Maximum Likelihood; Rissanen, 2001)

$$FIA = -\ln f(y | \hat{\theta}) + \frac{k}{2} \ln \frac{n}{2\pi} + \ln \int \sqrt{|I(\theta)|} d\theta$$

$$NML = -\ln \frac{f(y | \hat{\theta})}{\int f(z | \hat{\theta}(z)) dz}$$

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Fisher Information Approximation (FIA)

Rissanen (1996):

$$FIA = \underbrace{-\ln f(y | \hat{\theta})}_{\text{Goodness of fit (ML)}} + \underbrace{\frac{k}{2} \ln \frac{n}{2\pi} + \ln \int \sqrt{|I(\theta)|} d\theta}_{\text{Model Complexity}}$$

The model that minimizes MDL should be preferred

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$$FIA = \underbrace{-\ln f(y | \hat{\theta})}_{\text{- Goodness of fit}} + \underbrace{\frac{k}{2} \ln \frac{n}{2\pi} + \ln \int \sqrt{|I(\theta)|} d\theta}_{\text{Model Complexity}}$$

{ - Goodness of fit}

{ Model Complexity}

$$\frac{k}{2} \ln \frac{n}{2\pi} + \ln \int \sqrt{\det I(\theta)} d\theta$$

Complexity due to number
of parameters (k)
(e.g., AIC, BIC)

Complexity due to *functional
form* of the model equation

Complexity: More than the number of parameters?

$$M1: y = ax + b$$

$$M2: y = ax^b$$

$$M3: y = \sin(\cos ax)^a \exp(-bx) / x^b$$

Are these all equally complex?



The geometry of the [space of probability distributions](#) provides a well-justified and intuitive framework of model complexity, the central concept in model selection.

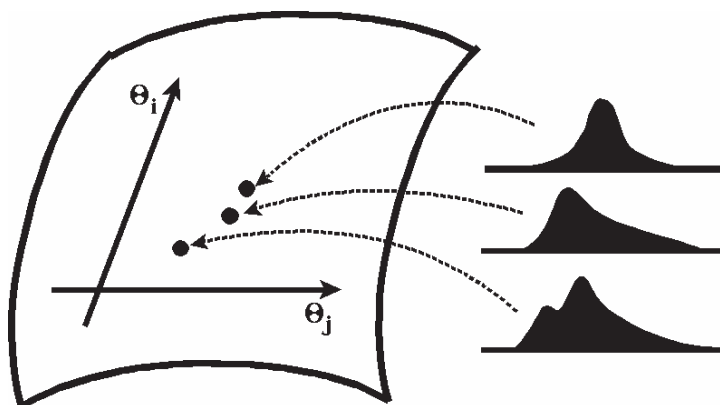
In this approach, we construct “geometric” complexity of a model by [counting](#) the number of different distributions it indexes.

(e.g.) Data space = {a, b, c, d, e, f, g, h}
Model A = {a, c, d} vs Model B = {b, d, e, g, h}

Space of Probability Distributions



The family of probability distributions forms a Riemannian manifold in which “similar” distributions are mapped to “nearby” points ([information geometry](#); Rao, 1945; Efron, 1975; Amari, 1980).





A **distance metric** that measures 'dissimilarity' between two neighboring distributions is defined as

$$ds^2 = \sum_{i,j} g_{ij}(\theta) d\theta_i d\theta_j$$

where g_{ij} is the *Riemannian metric tensor* of the form:

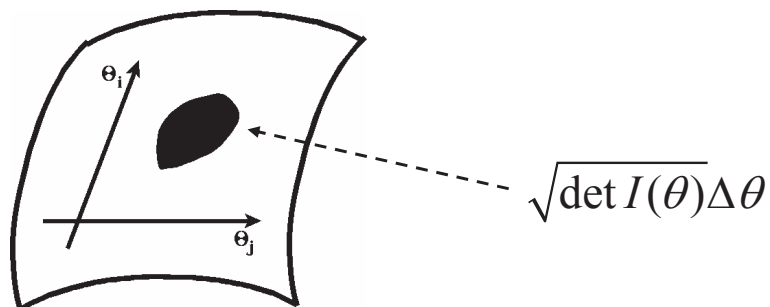
$$g_{ij}(\theta) = -E \left[\frac{\partial^2 \ln f(y|\theta)}{\partial \theta_i \partial \theta_j} \right]$$

which is the *Fisher information matrix*, $I(\theta)$.

Complexity and *Riemannian volume*



In a geometric context, model complexity should be related to the volume the associated manifold occupies in the space of distributions:

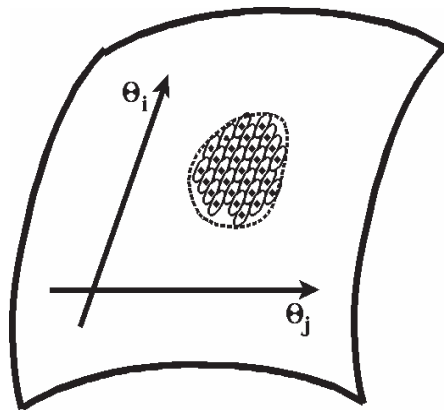


which is known as the **Riemannian volume** in differential geometry.

Count only “distinguishable” distributions



The Riemannian volume measure is related to the *local density* of ‘*distinguishable*’ probability distributions indexed by the model.

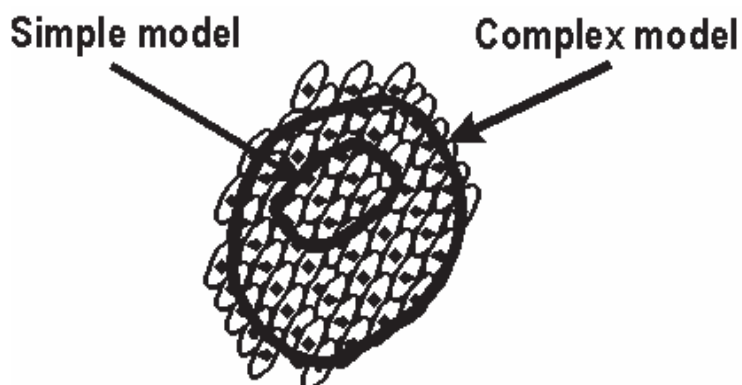


Overall volume: $V(f) = \int d\theta \sqrt{\det I(\theta)}$

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Simple vs complex models: An information geometric view



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Distributions close to the truth



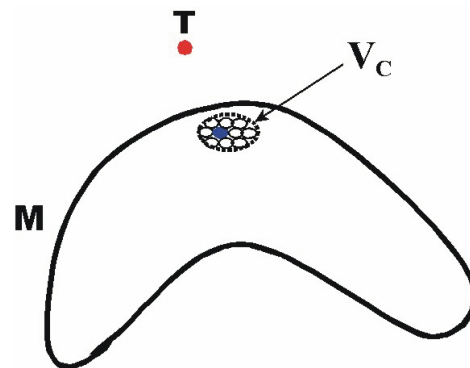
A good model should contain many distinguishable distributions that come **close to the truth**, in the sense.

C : a subset of distributions s.t.

$$f(y|\theta) \approx f(y|\hat{\theta})$$

The Riemannian volume of such region is obtained as:

$$V_C = \int dV_C \approx \left(\frac{2\pi}{n} \right)^{k/2}$$



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Model complexity as volume ratio



The **log volume ratio**, $V(f)/V_C$, gives

$$\ln \left(\frac{V(f)}{V_C} \right) = \frac{k}{2} \ln \frac{n}{2\pi} + \ln \int \sqrt{\det I(\theta)} d\theta$$

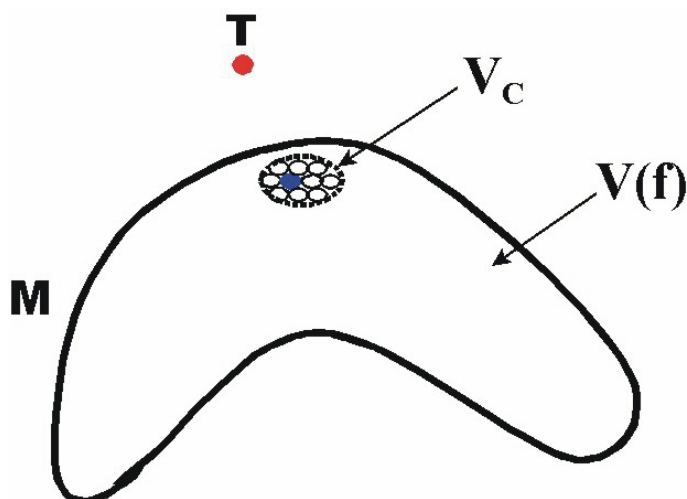
“geometric complexity”

→ **Geometric complexity** turns out to be equal to the complexity term of **Fisher Information Approximation** (FIA: Rissanen, 1996):

$$FIA = -\ln f(y|\hat{\theta}) + \frac{k}{2} \ln \frac{n}{2\pi} + \ln \int \sqrt{|I(\theta)|} d\theta$$

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$$\text{ModelComplexity}_{\text{FIA}} = \ln \left(\frac{V(f)}{V_c} \right)$$

From this view, a complex model is one containing many different distributions **overall** ($V(f)$) but relatively few ones **close to the truth** (V_c)

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Normalized Maximum Likelihood (NML)

$$NML = -\ln \frac{f(y | \hat{\theta})}{\int f(z | \hat{\theta}(z)) dz}$$

$$= -\ln \frac{\text{ML value of current data}}{\text{Sum of all ML values of all possible data}}$$

From the NML viewpoint, a good model is the one that gives relatively high ML only for current observations but low ML values for other data patterns.

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NML as Minimax Coding Strategy

$$NML = -\ln \frac{f(y | \hat{\theta})}{\int f(z | \hat{\theta}(z)) dz}$$

- NML derived as minus logarithm of a probability distribution that minimizes the maximum distance between the desired distribution and the best-fit member of the model family.

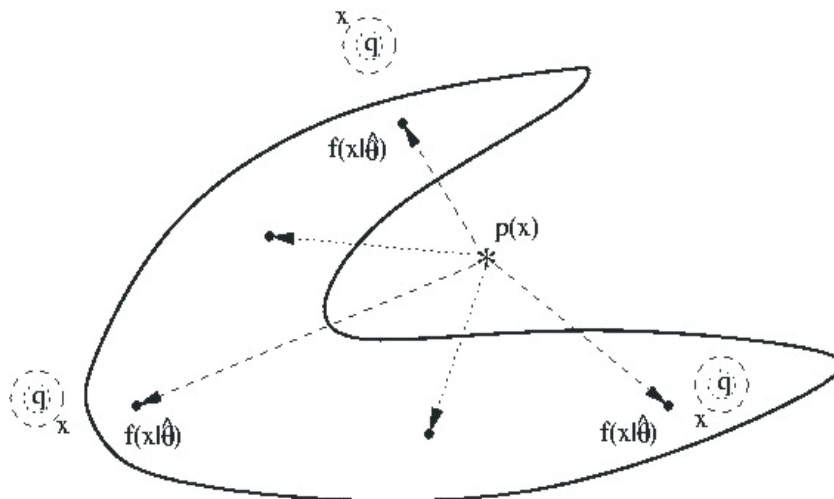
$$NML = -\ln p^*(y)$$

$$\text{where } p^*(y) \triangleq \arg \inf_p \sup_q E^q \left[\ln \frac{f(x | \hat{\theta})}{p(x)} \right]$$

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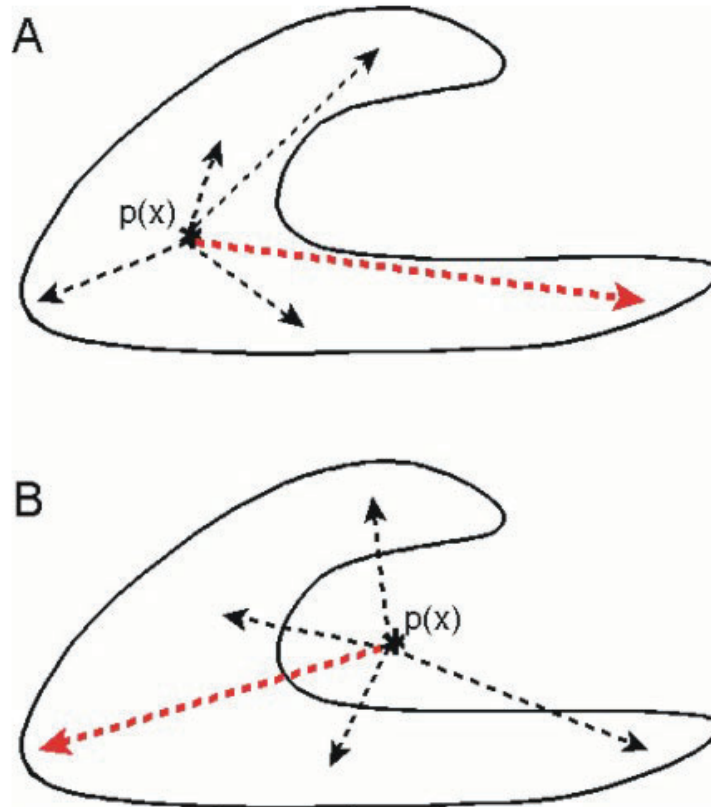
Minimax Problem in a Model Manifold



$p^*(x)$: “Universal” distribution in the sense that it can mimic the behavior of the entire model class of probability distributions.

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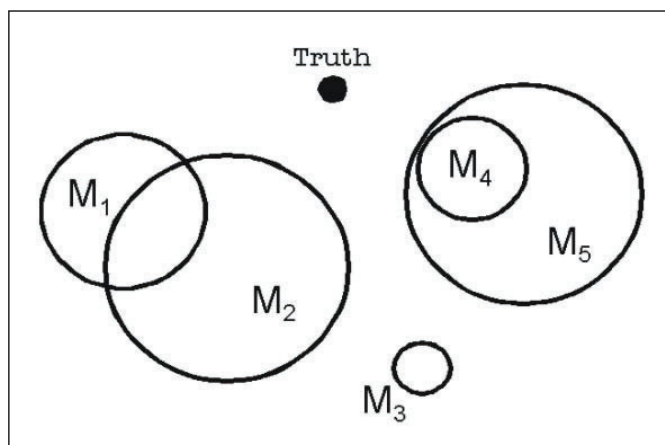


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Derivation NML as a solution to the minimax strategy does not require that:

- Models be nested within one another;
- None of the models be “true”;
- NML solution be a member of the model family



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Model complexity: A NML view

$$NML = -\ln f(y | \hat{\theta}) + \ln \int f(z | \hat{\theta}(z)) dz$$

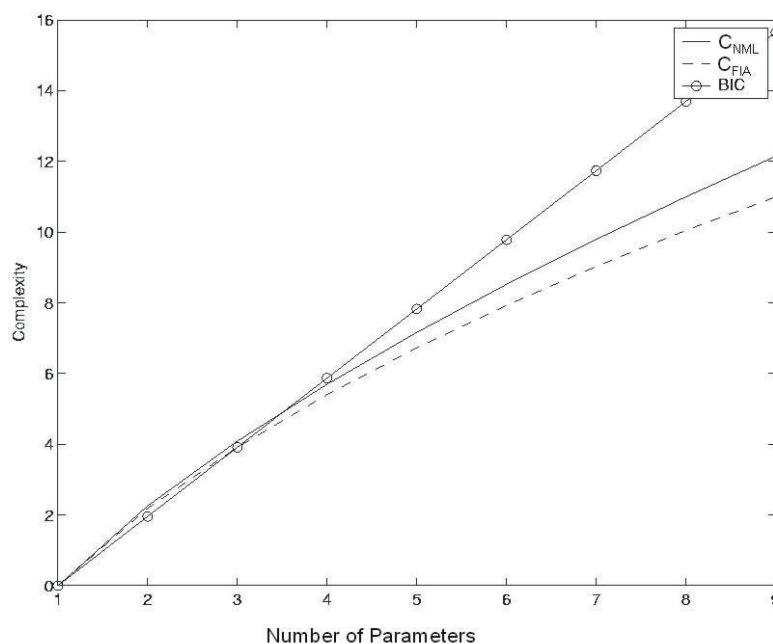
C_{NML}:

- Normalizing constant of NML distribution
- Minimax distance achieved
- Sum of all “best” (ML) fits
- Sensitive to number of parameters, sample size, functional form, experimental design, etc.

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Complexity Comparison



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Other model selection methods

- ICOMP (Bozdogan, 1990)
- RMSEA (Steiger, 1990)
- AGFI (Jöreskog & Sörbom, 1986)
- NIC (Murata et al, 1994)
- DIC (Spiegelhalter et al, 2002)
- FIC (Claeskens & Hjort, 2003)
-
-
-

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Overview

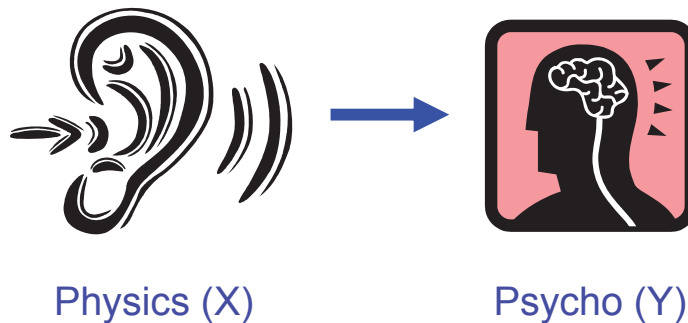
- **Part 1: Non-technical Introduction to Model Selection**
- **Part 2: “Technical” Tour of Model Selection Methods**
- **Part 3: Example Application**
- **Part 4: Conclusions**

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Example Application in Psychophysics

Models of psychophysics describe the relationship between physical dimensions (e.g., tone intensity) and their psychological counterparts (e.g., loudness).



$$Y = f(X, \theta)$$

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Psychophysics models

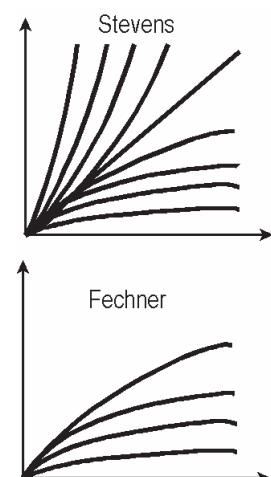
Stevens law: $y = ax^b$

Fechner's law: $y = a \log(x + b)$

Complexity difference:

$$C_{\text{MDL,Stevens}} - C_{\text{MDL,Fechner}} = 3.8$$

The difference in complexity must be due to the effects of functional form



$$FIA = -\ln f(y | \hat{\theta}) + \frac{k}{2} \ln \frac{n}{2\pi} + \ln \int \sqrt{|I(\theta)|} d\theta$$

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Model Recovery Simulation (% recovery)

Selection Method	Data From:	Stevens	Fechner
	Model Fitted:		
AIC (BIC)	Stevens	100	63
	Fechner	0	37
FIA	Stevens	99	2
	Fechner	1	98

Model Recovery Simulation (% recovery)

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Conclusions

- Models should be evaluated based on **generalizability**, not on **goodness of fit**



“Thou shall not select the **best-fitting** model but shall select the **best-predicting** model.”

- Other non-statistical but *very important* selection criteria:
 - Plausibility
 - Interpretability
 - Explanatory adequacy
 - Falsifiability

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“All models are wrong, but some are useful.”

(G. P. E. Box, 1978)

“Model selection methods can help identify *useful* models, in the sense of *predictive accuracy* or *generalizability*.”

(J.I.M.)



Thank You for Your Attention!

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