IV example

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0. College in County

Card (1995)¹ was interested in the causal effect of schooling years, educ, on log wages, lwage. However, we can imagine that several variables will affect both the independent (wages) and dependent (schooling) variables and, thus, that OLS will yield biased estimators. Instead, Card (1995) decided to instrumentalize schooling with a dummy variable that takes value 1 when an individual lived in a county in which there was a 4-year college, nearc4, while controlling for several available common covariates: years of experience, exper, and whether or not a person is black, black, lives in the southern US, south, is married, married, and lives in an urban area (Standard Metropolitan Statistical Area), smsa.

1. Loading packages and data

```
# Packages
#install.packages("AER")
                             # function iv req
#install.packages("haven")
library(AER)
library(haven)
# Data
Card1995 <- read dta("https://raw.github.com/scunning1975/mixtape/master/card.dta")
# Subset variables of interest
Card1995 <-
  Card1995[c("lwage",
             "educ",
                                                              #treatment(endogenous)
             "exper", "black", "south", "married", "smsa",
                                                              #covariates
             "nearc4")]
                                                              #instrument
```

2. Regressions

2.1. Bivariate regression

A first bivariate log-lin regression suggests that increases by 1 year of education predict increments of 5,2% in wages.

^{*}Cunningham, Scott (2021). Causal Inference: The Mixtape

¹Card, David (1995). "Aspects of Labour Economics: Essays in Honour of John Vanderkamp." In. University of Toronto Press.

```
ols1 <- lm(lwage ~ educ, data = Card1995)
summary(ols1)
##
## Call:
## lm(formula = lwage ~ educ, data = Card1995)
##
## Residuals:
##
       Min
                 1Q
                      Median
                                   3Q
                                           Max
## -1.73799 -0.27764 0.02373 0.28839
                                       1.46080
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 5.57088
                          0.03883 143.47
                                            <2e-16 ***
## educ
                                            <2e-16 ***
               0.05209
                          0.00287
                                    18.15
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.4214 on 3008 degrees of freedom
## Multiple R-squared: 0.09874,
                                   Adjusted R-squared: 0.09844
## F-statistic: 329.5 on 1 and 3008 DF, p-value: < 2.2e-16
```

2.2. OLS with controls

As you know, we can add available control variables to account for differences between groups that might confound the effect of schooling. This step corrects our prior prediction by 2,1 percent points.

```
##
## Call:
## lm(formula = lwage ~ educ + exper + black + south + married +
##
       smsa, data = Card1995)
##
## Residuals:
       Min
                  10
                     Median
                                    30
                                            Max
## -1.59924 -0.23035 0.01812 0.23046 1.36797
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
                          0.063740 79.437
## (Intercept) 5.063317
                                              <2e-16 ***
                          0.003482 20.438
## educ
               0.071173
                                              <2e-16 ***
## exper
               0.034152
                           0.002214 15.422
                                              <2e-16 ***
                                    -9.426
## black
              -0.166027
                           0.017614
                                              <2e-16 ***
               -0.131552
                          0.014969 -8.788
                                              <2e-16 ***
## south
              -0.035871
                           0.003401 -10.547
                                              <2e-16 ***
## married
               0.175787
                          0.015458 11.372
## smsa
                                              <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.3702 on 2996 degrees of freedom
     (7 observations deleted due to missingness)
```

```
## Multiple R-squared: 0.305, Adjusted R-squared: 0.3036
## F-statistic: 219.2 on 6 and 2996 DF, p-value: < 2.2e-16</pre>
```

2.3. 2SLS

educ

However, the inclusion of controls does not solve the problem of endogeinity in educ. What if there are other unobservable variables biasing our estimate? If nearc4 satisfies relevance (i.e., our first stage coefficient is strong), independence (i.e., that individuals' reasons for being settled in a county are related to whether or not nearc4 is 1 or 0 and, thus, it is as good as random), and exclusion (i.e., that nearc4 is associated with lwage, only through educ—or at least is uncorrelated with any unobservable variable biasing estimates), we can use it as imperfect instrument of educ.

```
iv_reg = ivreg(
  #Second stage
  lwage ~ educ +
    exper + black + south + married + smsa |
    #First stage (treatment omitted)
   nearc4 +
    exper + black + south + married + smsa,
  #Data
  data = Card1995)
summary(iv_reg)
##
## Call:
## ivreg(formula = lwage ~ educ + exper + black + south + married +
##
       smsa | nearc4 + exper + black + south + married + smsa, data = Card1995)
##
## Residuals:
##
        Min
                       Median
                  1Q
                                    30
                                            Max
## -1.81301 -0.23805 0.01766 0.24727
                                        1.32278
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept)
               4.162476
                           0.849590
                                     4.899 1.01e-06 ***
## educ
                0.124164
                           0.049956
                                      2.485 0.01299 *
## exper
                0.055588
                           0.020286
                                     2.740 0.00618 **
                           0.050741 -2.280 0.02268 *
## black
               -0.115686
                           0.023244 -4.869 1.18e-06 ***
## south
               -0.113165
## married
               -0.031975
                           0.005087 -6.286 3.73e-10 ***
## smsa
                0.147707
                           0.030895
                                     4.781 1.83e-06 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.3843 on 2996 degrees of freedom
## Multiple R-Squared: 0.2513, Adjusted R-squared: 0.2498
## Wald test: 139.8 on 6 and 2996 DF, p-value: < 2.2e-16
# Compare coefficients and SE from OLS and IV
out_ols2 <- summary(ols2)</pre>
out_iv_reg <- summary(iv_reg)</pre>
out_ols2$coefficients[,1:2]
                  Estimate Std. Error
## (Intercept) 5.06331654 0.063740191
```

0.07117285 0.003482405

```
## exper
                0.03415182 0.002214445
               -0.16602745 0.017613671
## black
## south
               -0.13155177 0.014969061
## married
               -0.03587071 0.003401161
## smsa
                0.17578712 0.015457781
out_iv_reg$coefficients[,1:2]
##
                  Estimate Std. Error
                4.16247588 0.849590453
## (Intercept)
## educ
                0.12416424 0.049955802
## exper
                0.05558822 0.020286089
## black
               -0.11568555 0.050741489
               -0.11316470 0.023243878
## south
## married
               -0.03197537 0.005086886
## smsa
                0.14770651 0.030895151
```

LATE is 12,4% (that is 5.3 percent points or 75% larger than our OLS estimate).

The external validity of our LATE (i.e., whether or not LATE is (approximately) equal to TOT) depends on our assumptions. We can interpret this result the following way: among compliers in our sample (i.e., those who extend their education because of living in a county with a college, for example, because it reduces the cost of schooling), an extra year of schooling increases wages 12,4% on average. If we can argue that this effect is indeed independent from being a complier, we might suggest that this is an unbiased estimator of years of education on wages for the whole population (as opposed to the OLS).

3. Standard Errors

Manual estimation of IV only takes into account the regression error from the second stage, i.e., it ignores the regression error from the first stage and, thus, **provides incorrect standard errors**. Modern software (e.g., R and Stata) provides robust standard errors automatically, so... even if the coefficients are the same, let sofware do the hard calculations for you!!

See the example below:

```
# Manual First stage
fs <- lm(educ ~
           nearc4 +
           exper + black + south + married + smsa,
         data = Card1995)
summary(fs)
##
## Call:
  lm(formula = educ ~ nearc4 + exper + black + south + married +
##
       smsa, data = Card1995)
##
## Residuals:
##
       Min
                1Q Median
                                 3Q
                                        Max
  -7.6308 -1.4454 -0.0526
                            1.2986
                                     6.3449
##
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
                            0.13075 128.727 < 2e-16 ***
## (Intercept) 16.83070
## nearc4
                0.32728
                            0.08242
                                      3.971 7.33e-05 ***
## exper
               -0.40443
                            0.00894 -45.238 < 2e-16 ***
```

```
0.09053 -10.467 < 2e-16 ***
## black
              -0.94753
## south
               -0.29735
                          0.07906 -3.761 0.000173 ***
## married
               -0.07269
                          0.01775 -4.096 4.31e-05 ***
               0.42090
                          0.08487
                                    4.959 7.47e-07 ***
## smsa
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.937 on 2996 degrees of freedom
     (7 observations deleted due to missingness)
## Multiple R-squared: 0.4774, Adjusted R-squared: 0.4764
## F-statistic: 456.1 on 6 and 2996 DF, p-value: < 2.2e-16
# Add a column with predicted treatment from first-stage
Card1995$educ_pred <- with(Card1995, 16.83069650 + 0.32728259*nearc4 -
                             0.40443401*exper - 0.94752809*black -
                             0.29735279*south - 0.07269361*married +
                             0.42089454*smsa)
# Plotting First-Stage
with(Card1995, plot(educ,educ_pred,xlab="educ",ylab="educ_pred",pch=3))
```

