

Critique Papers on Causal Inferences

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Propensity Scores

Sullivan, A. L., & Field, S. (2013). Do preschool special education services make a difference in kindergarten reading and mathematics skills?: A propensity score weighting analysis. *Journal of School Psychology, 51*(2), 243–260.

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Summary

Sullivan and Field (2013) attempted to investigate the marginal benefit created by a special education program mandated by the US *Individuals with Disabilities Education Act* (IDEA). The 1986 amendments to the IDEA legislation imposed legal requirement on the states to create a sequence of intervention programs targeting three stages of development for children with needs: 1) from birth to 2-year-old, 2) 3-year to kindergarten entry, and 3) kindergarten to 21-year-and-11-month-old. This paper focused on the middle segment of the intervention sequence named preschool special education services and followed children longitudinally through the delivery period. The study began by measuring children's probability of being admitted into the special education program using a logit regression, and at the end of the intervention period, compared the average performance scores in maths and reading of the children who actually received the preschool special education treatment against the average scores of those who have not. Using the propensity score weighting technique, the authors interrogated the counterfactual question of “what if” the treatment were never applied and reached an conclusion that children in treatment group would have been better off academically had they not received any intervention at all. This disappointing result lent itself to an existing literature of similar negative findings evaluating special education effectiveness, cautioning the rosy objectives of preschool special education services originally marketed by policy makers.

Causal Question

Would the children who received special education services have been better off academically, on average, had they not received such services?

Validity

Internal and External Validity

Historically, internal validity referred to inferences about whether “the experimental treatments make a difference in this specific experimental instance” while external validity asked “to what populations, settings, treatment variables, and measurement variables can this effect be generalized” (Campbell & Stanley, 1963, p. 5). Cook and Campbell (1979) advanced the idea of internal validity to the question whether the covariation observed between the independent and dependent variables were resulted from a *causal* relationship, whereas external validity further asks whether such cause-effect relationship holds over certain variation in persons, settings, treatment variables, and measurement variables.

In order to support an inference that the observed covariation between A and B reflects a causal relationship, Shadish et al. (2002) prescribed a trifecta that 1) A preceded B in time, 2) A covaries with B , and 3) no other explanations for the relationship are plausible. It is too often the third strand that undermines the internal validity of inference making—the relationship between A and B is not causal because it could have occurred even in the absence of the treatment and that it could have led to the same outcomes that were observed for the treatment. Amongst the list of potential threats to internal validity (Shadish et al., 2002, pp. 54–61), maturation presents the strongest challenge to Sullivan and Field (2013). Children who have been identified as “in need” so early in life (“early onsetters”) can be reasonably believed to be in possession of different developmental profiles from children who showed needs later in life (“late onsetters”). As participants in the treatment group mature, gaps in academic performance may well emerge out of such delayed developmental trajectories with or without education services. It is therefore not preschool interventions that “caused” lower academic scores but the two-tier growth profiles that did. Failing to rule out such alternative, and rather plausible, explanation weakens the internal validity of inference made by the authors.

Weak external validity has also been acknowledged by the authors in Section 4.2 of the paper. Inferences can only been drawn, first of all, over children with mild to moderate impairments resultant from the sample deletion procedure; while it were the children with the most severe impairment that policy makers wished to monitor and retain (“interaction of the causal relationship with units” by Shadish et al. (2002)). Secondly, Sullivan and Field (2013)

evaluated only the academic performance of young children to the exclusion of other developmental markers such as motor-behavioural and social-affective skills—all vital policy objectives along with reading and maths scores, if not more important, for kindergarten-entry age children (“interaction of the causal relationship with outcomes”). Lastly, the averaging procedure in calculating ATT washed out important differences across race and socio-economic groups, important factors reported by prior literatures as non-ignorable (“context-dependent mediation”).

Construct Validity

Construct validity concerns itself with the degree of agreement between the concept the researchers intended to understand (e.g., academic performance) and the procedure as well as instrument they employed to capture and measure such concept (e.g., sum scores in maths and reading tests). Amongst the various threats proposed by Shadish et al. (2002, pp. 72–81), Sullivan and Field (2013) were particularly susceptible to “inadequate explication of constructs” and “construct confounding”. The concept of academic performance can be thought as the end result of a sequence of social activities: academic input (I)–academic processing(P)–academic output(O). When academic scores were low, one is unable to ascertain whether it was the result of inferior teaching (xPO), lack of learning skills (IxO), or inability to demonstrate or document learning outcome to observers (IPx). Although both xPO and IPx may show up as low academic scores, the “causal pathways” cannot be more different—a situation not assisted by the ECLS-B early reading and math batteries used in Sullivan and Field (2013) since none was designed to locate the source(s) of academic deficiency.

Construct confounding occurs when the concept under investigation has not been careful separated from other related concepts. Sullivan and Field (2013) clearly wished to study “academic performance” of young children but such construct covary particularly strongly for this age group with attention span and sociability. It is not unreasonable to conjecture that recipients of the special education program may not develop the above-mentioned skills at the same pace as their counterparts. Effectively, the ECLS-B batteries employed by Sullivan and Field (2013) were capturing young children’s short attention spans and under-developed social skills and presenting them as inferior academic performance. Both inadequate explication of constructs and construct confounding have,

therefore, weakened this study’s construct validity, undermining its inference of “special education causing lower academic performance”.

Statistical Conclusion Validity

By Cook and Campbell (1979), statistical conclusion validity refers to the appropriateness of statistical techniques employed by the researcher for the purposes of inferring whether the presumed independent and dependent variables indeed covary. The propensity score weighting technique employed by Sullivan and Field (2013) successfully circumvented many pitfalls summarised by Shadish et al. (2002, pp. 45–52) except for the “restriction of range” threat to statistical conclusion validity. Due to the necessity of constructing a region of common support, children are purposefully excluded from analyses if their probabilities of being accepted into the special education program fall outside of the 1% to 82% range. This practice is especially concerning for the above-82% group since these are the young children with demonstrated need for urgent education assistance. Under the law of diminishing marginal returns, it is more than probable that it is this most-in-need group that would have responded best and most rapidly to special education interventions. The wholesale omission of this positive outcome pool may have well contributed to the underestimation of the project effectiveness.

Appropriateness of Methods

Sullivan and Field (2013) largely followed the propensity score analysis procedure prescribed by Imbens and Rubin (2015) and Imbens (2015) in assessing causal effects. At the first stage DESIGN, the authors established sufficient overlapping by discarding some units from the original sample in order to establish the region of common support; the second stage SUPPLEMENTARY ANALYSIS, however, appeared to be lacking in Sullivan and Field (2013) where the plausibility of unconfoundedness shall be further addressed through pseudo-average treatment effect on the pseudo-outcome for trimmed sample (see Imbens, 2015, pp. 383–384); such absence would cast doubt on any result in the third stage ANALYSIS over the source of average treatment effect.

One highlight on methodology is the Bayesian approach to AAT sampling weights. Sullivan and Field (2013) correctly pointed out the “curse of dimensionality” when computing w_i for $D_i = 0$ cases and provided sufficient derivation through Bayes formula in reaching the

form

$$w_i = \frac{\mathbb{P}(D = 1|z_i)}{1 - \mathbb{P}(D = 1|z_i)} \cdot \frac{\mathbb{P}(D = 0)}{\mathbb{P}(D = 1)}, \text{ for } D_i = 0.$$

The authors, however, stayed short of advocating for a wider adoption of this approach to resampling weights but gave in to conventional literature in order to maximise comparability. This weighting formula overcome the peculiarity of the conventional scheme (only the first term in the formula above) summing to twice the size of the treated subsample and provided a more intuitive formation of summing to the sample size. Stronger advocacy can be expected from continuing research in popularising Sullivan and Field’s (2013) weighting formulation.

Conclusion

Sullivan and Field (2013) made a good attempt to apply the propensity score technique for causal assessment of preschool special education data. Despite some omissions in statistical procedure, what limited this paper’s impact on policy was *not* the econometric methodology it employed but the weak inferential validity. Since “[v]alidity is a property of inferences [...] *not* a property of designs or methods,” (Shadish et al., 2002, p. 34, emphasis in original text) no amount of technical sophistication is capable of compensating for validity, or the lack thereof. Sullivan and Field’s (2013) result shall be interpreted narrowly based on *this* particular round of study, using *this* sub-sample to quantify *this* particular sub-set of outcome measures, subject to *these* particular restrictions, omissions and commissions, and based on *these* many statistical assumptions which may or may not have been met. A naïve interpretation of “special education *causes* even worse academic outcomes” shall be rejected out right. After all, an absence of evidence is *not* the evidence of the absence, and the total social return generated by early-life education intervention programs shall be contextualised in the general equilibrium analysis framework (e.g., through estimating the multiplier effect) rather than a partial one. Nevertheless, Sullivan and Field (2013) had made contribution to both the propensity score methodology and to the substantive debate over the direction and magnitude of the effectiveness of one social project.

Instrumental Variables

Hanandita, W., & Tampubolon, G. (2014). Does poverty reduce mental health? An instrumental variable analysis. *Social Science & Medicine*, 113, 59–67.
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Summary

Hanandita and Tampubolon (2014) investigated the causal relationship between poverty and mental health decline using an instrumental variable (IV) approach in order to overcome the endogeneity problem. Using a sample size of 577,548 across 440 districts in Indonesia and precipitation anomaly as the IV, the authors were able to quantify the expenditure/income elasticity of mental disorders as -0.62 —a result five times stronger than that of the non-IV approach and robust to various stress tests. Moreover, income inequality also appeared to carry explanatory power to mental health concerns in addition to that of poverty, suggesting both the position (quantity of income) and the shape (distribution of income) of the income curve as policy variables worth pursuing for the betterment of population mental welfare.

Causal Question

Does poverty cause poor mental health?

Validity

Since the previous critique paper has documented a typology of validity, subsequent analyses would not repeat such content but focus on the application on the paper Hanandita and Tampubolon (2014).

Statistical Conclusion and Internal Validity

Statistical conclusion validity concerns itself with whether the presumed cause and effect covary and how strongly, whereas internal validity asks whether the observed covariation is causal in nature. Hanandita and Tampubolon (2014) carefully avoided all threats to statistical conclusion validity put forward by Shadish et al. (2002, Table 2.2, p. 45). The endogeneity problem, for example, would have violated the Gauss-Markov assumption of $\mathbb{E}(\varepsilon_i | \mathbf{x}_i) = 0$; Hanandita and Tampubolon (2014) not only restored

independent error condition through the introduction of an IV (see, for example, Greene (2018) Chapter 8 for a technical discussion of IV), but also reported the magnitude of underestimation due to such assumption violation. By ensuring the IV to be uncorrelated with mental health but highly correlated with income, this study introduced an appropriate circuit breaker to the infinite feedback loop between poverty and mental health conditions, clearly suggesting the covariation between the two had indeed been causal and the arrow of causation points from income to mental health, not the other way around.

Construct and External Validity

Both construct and external validity deals with generalisation. In addressing the construct validity, Section 4.3 of Hanandita and Tampubolon (2014) has been careful in distinguishing expenditure from income, and reported the observed deterioration in mental health condition as a response to reduction in consumption expenditure, in order to separate permanent income changes from intermittent income shock. Since this study used large dataset collected at national level, interactions of the causal relationship with both settings and outcomes can be minimum (see Table 3.2, Shadish et al., 2002, p. 87), therefore delivering strong external validity.

Appropriateness of Methods

The instrumental variable approach adopted by Hanandita and Tampubolon (2014) served their research purpose (to overcome endogeneity problem) and claim (poverty causes mental health decline) well. The last paragraph in Section 1 of the paper paid particular justification to the key assumptions behind the IV method, namely relevance condition, validity condition and exclusion restriction and revisited the suitability of these assumptions in the third paragraph of Section 6, admitting that “[t]he quality of an instrumental variabel estimation is only as good as its story” (p. 65). Although untestable, the proposal put forward by the authors that precipitation anomaly was a random assignment procedure perfectly uncorrelated with the outcome variable (mental health condition) but covaried strongly with input variables (income) due to large proportion of the Indonesian labour force being employed in a rain-dependent agricultural sector, was a convincing one.

The model building process was also appropriate. The authors ran their IV models against their baseline counterpart (i.e., models without IV); this comparison revealed a

five-fold increase in the estimated effect of poverty on mental health due to the introduction of IV, incidentally revealing the magnitude of underestimation of the naïve regression approach.

Other methodological considerations also enhanced Hanandita and Tampubolon (2014)'s credibility. The authors explored both linear (linear and LPM) and non-linear (Poisson and Probit) configurations of their models to show the reported results were unlikely to be an mere artifact of the chosen functional forms. Correctional procedures such as the incorporation of sampling weights and clustering also safeguarded variance estimates. Centring of continuous variables such as log per capita household expenditure and the Gini coefficient also enhanced interpretability of the numeric results.

Conclusion

Hanandita and Tampubolon (2014) had delivered a carefully designed study to the social science community. They elevated their research enquiry from a correlational endeavour to a causal one not only to satisfy one's methodological curiosity but to provide a conclusive response to the policy choice that "if causal links between wealth and health were confirmed, society would likely benefit from more universal access to health care and redistributive economic policy. Yet, if such causal links were rebutted, resources would be better spent on influencing health knowledge, preferences, and ultimately the behavior of individuals." (Stowasser et al. (2011) as cited in Hanandita and Tampubolon (2014)). The causal evidence presented by this paper would facilitate policy actions by updating scientific believes towards the former option and contribute to the betterment of mental health project in Indonesia and developing countries at large.

Fixed-effect Regression

White, M. P., Alcock, I., Wheeler, B. W., & Depledge, M. H. (2013). Would you be happier living in a greener urban area? A fixed-effects analysis of panel data. *Psychological Science*, 24(6), 920–928. <https://doi.org/10.1177/0956797612464659>

Summary

White et al. (2013) attempted to examine the causal relationship between green space and individuals' happiness using STATA's panel data analysis tool. By regressing self-reported mental health levels and life satisfaction scores on green space along with other regional- and individual-level control variables, the authors reported a small but statistically significant effect greenery plays in improving residents' happiness and used these statistics as evidence to further their advocacy for including more green space in urban design.

Causal Question

“[W]hether the same people would be happier (i.e., show higher well-being and lower mental distress) when living in areas with more green space than in areas with less green space.” (p. 921)

Validity

Statistical Conclusion Validity

The statistical conclusion validity asks whether the presumed cause and effect covary and how strongly they do so. Table 2 of White et al. (2013, p. 925) documented the regression coefficients (both unstandardised b and standardised β) of their fixed-effect analyses. It is first of all worth noticing that the magnitude of β for **Green space** is small albeit statistically significant—a *prima facie* evidence of a weak marginal effect attributable purely to green space. Such weak relationship between greenery and individual happiness has been further shadowed by individual character variables such as marital status (over 3 times stronger for both General Health Questionnaire (GHQ) and Life Satisfaction Survey (LSS) measures) and unemployment (5 and 2.5 times stronger in the opposite direction for GHQ and LSS respectively). Most other variables failed to reach statistical significance for both GHQ and LSS outcomes. This inconvenient fact exposed the weak statistical conclusion

validity of the current study: the key variable of advocacy **Green space** barely made into the winning list and remain as the least competitive explanatory variable.

This low statistical power validity is further threatened by violations of statistical assumptions. Input variables chosen by the authors, such as income, employment, education, and crime at LSOA-level and education, household income, employment status at individual-level, formed strong multicollinearity with each other since they all reflect the socioeconomic status of the research subjects; when one variable moves, other variables tend to follow closely, defeating the *ceteris paribus* interpretation. A second major violation is on the random sampling assumption: this study could essentially be framed as a multilevel model where instances of samplings were nested in individuals and individuals were nested in LSOA-levels. Even by discarding the repeated measure layer (by degrading a longitudinal study into a cross-sectional design), the authors paid no consideration to the clustering nature of individuals nested in LSOA areas, inflating any subsequent analyses. Given the already small β reported for **Green space**, one must wonder whether such effect would survive a clustering correction.

Internal Validity

The internal validity asks whether the observed covariation between, green space (A) and happiness (B) in this study, is indeed a causal one by testing whether A preceded B in time, A covaries with B (only barely, if at all, as reasoned in the previous paragraphs) and no other explanations for the relationship could be plausible. Ambiguous temporal precedence came to play particularly strongly to this paper. A declaration that “sad people moving to green space, then *because of and only because of* this green, they transformed themselves into happier ones” attracts onerous burden of proof, where the authors made minimum effort to fulfil. The authors further failed to rule out “happy people self-select into moving to greener spaces” as a plausible possibility; or thirdly, another variable (e.g., SES) contributed to both individuals’ happiness and their propensity to reside near green areas. In summary, White et al. (2013) made least effort in addressing their internal validity at all.

Construct Validity

The construct validity concerns the agreement between the study operations and the constructs used to describe those operations. White et al. (2013) followed up people’s decision

on, or the lack thereof, moving into greener living spaces without realising that this variable consists of two sub-constructs: both the *willingness* and *ability* to relocate next to greenery. Assuming the general population to be either indifferent or positive towards green space (a reasonable assumption in industrialised economies such as England used in this study but probably less so in impoverished jungles where locals are more eager to trade trees for economic prosperity), what White et al. (2013) picked up was mainly the second sub-construct, the ability to relocate to greener space. In a world where housing prices are dictated by the golden rule of real estate “Location, location, location!”, the author’s operation “proximity to green” had metamorphosed to a measure of wealth, which then quickly formed multicollinearity with other SES variables described in the “statistical conclusion validity” section. Ultimately, the author mixed the construct of “a green space” with “(being able to) live in a green space” by stealth and produced a bunch of statistics that measured neither one nor the other.

External Validity

The external validity examines inferences about the extent to which a causal relationship withstands tests across various settings. As green space is scarce in industrialised England but relatively abundant in neighbouring Scotland—such change in scarcity could swiftly eliminate any reported causal effect between green living and happiness. Finally, since the marginal effects discovered by White et al. (2013) were so small, the authors switched to the “one standard deviation above/below” interpretation without realising that individuals sitting 1 *SD* apart in a standardised model share very few similarities any longer. The β coefficients only serve the *marginal* effect interpretation at and around the tangent point where linear approximation would be appropriate. Extrapolating these neighbourhood estimations to 1 *SD* away can only be applied at the authors’ own risks.

Appropriateness of Methods

It remained opaque why the authors started with a longitudinal design and at some unspecified point regressed to a cross-sectional study. The promised advantages of panel data, such as controlling for autocorrelation and heteroscedasticity, were as a result unrealised in full. As previously discussed in the “statistical conclusion validity” section, a multilevel modelling approach would be better suited to address the nesting nature, i.e., non-random

sampling, of the dataset. White et al. (2013) attempted to drive the heavy vehicle of panel data analysis and abandoned the apparatus in the middle of the field while a more appropriate methodological instrument was never considered.

Conclusion

White et al. (2013) crumbles in front of validity examinations. In fact, the author's hesitation over the validity of their study was marked by their repeated use of "may be", "could be" or "perhaps" over 25 times over a six-page paper excluding statistical tables, with a final confession "Causality cannot thus be assumed." (p. 927). The study design was ambitious to start with but quickly went off the boil, delivering a set of weak and non-results that inspired neither methodology nor policy formation. It mixed two constructs, green spaces and the ability to live next to green spaces, by stealth and failed to acknowledge the key role SES has been playing throughout their study as confound. They violated a myriad of statistical assumptions in their analyses and made large distance extrapolations using software outputs that were only supposed to be interpreted locally. No policy decision, as a result, should ever be carried out based on this paper, not even in England where the data were sourced, let alone any other jurisdictions.

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