

1 Metropolis: The Bounded Normal Mean

We choose uniform distribution on $[-2, 2]$ as the proposal distribution. The algorithm for generating samples from posterior is shown as below.

- **Step 1.** Start with arbitrary θ_0
- **Step 2.** At stage n , generate proposal θ' from $Unif(-2, 2)$.
- **Step 3.** Determine θ_{n+1} as
 - $\theta_{n+1} = \theta'$ with probability $\rho(\theta_n, \theta')$;
 - $\theta_{n+1} = \theta_n$ with probability $1 - \rho(\theta_n, \theta')$.
- **Step 4.** Set $n = n + 1$ and go to **Step 2**.

We derive the acceptance ratio as follows.

As we know that the observations are taken from normal distribution with precision $\tau = 1/4$, that is $f(y|\theta) \propto \frac{1}{2} \exp\{-\frac{1}{8}(y - \theta)^2\}$, then the target density function $\pi(\theta|X)$ is proportional to

$$\begin{aligned} \pi(\theta|\mathbf{y}) &\propto f(\mathbf{y}|\theta)\pi(\theta) \\ &\propto \prod_{i=1}^n \exp\left\{-\frac{1}{8}(y_i - \theta)^2\right\} \frac{1}{2} \cos^2\left(\frac{\pi\theta}{4}\right) \\ &= \exp\left\{-\frac{1}{8}\sum_{i=1}^n (y_i - \theta)^2\right\} \frac{1}{2} \cos^2\left(\frac{\pi\theta}{4}\right) \end{aligned}$$

As we choose independent proposal distribution, then the proposed θ' and θ_n are independent. In this case, we have

$$q(\theta_n|\theta') = \frac{1}{4}, \quad q(\theta'|\theta_n) = \frac{1}{4}.$$

Thus we compute the acceptance ratio as

$$\rho(\theta_n, \theta') = \min\left\{1, \frac{\pi(\theta') q(\theta_n|\theta')}{\pi(\theta_n) q(\theta'|\theta_n)}\right\} = \min\left\{1, \frac{\exp\left\{-\frac{1}{8}\sum_{i=1}^n (y_i - \theta')^2\right\} \frac{1}{2} \cos^2\left(\frac{\pi\theta'}{4}\right)}{\exp\left\{-\frac{1}{8}\sum_{i=1}^n (y_i - \theta_n)^2\right\} \frac{1}{2} \cos^2\left(\frac{\pi\theta_n}{4}\right)}\right\}.$$

- (a) We choose the proposal as the uniform distribution over $[-2, 2]$, which means we have a Metropolis random walk. By simulation, we obtain Figure 1.
- (b) Bayesian estimator is -0.397 and the 95% credible set is $[-1.442, 0.750]$.

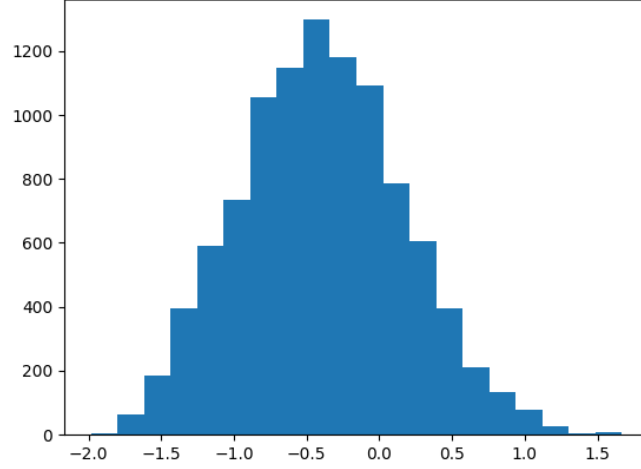


Figure 1: One simulation for the posterior.

2 Gibbs Sampler and High/Low Protein Diet in Rats

First, we denote that $\mathbf{y}_1 = (y_{1i})_{i=1}^{12\top}$ and $\mathbf{y}_2 = (y_{2i})_{i=1}^{7\top}$ as vectors representing the 12 observations under high protein and 7 observations under lower protein, respectively.

(a) For θ_1 , we have that

$$\begin{aligned}\pi(\theta_1, \tau_1, \mathbf{y}_1) &= f(\mathbf{y}_1 | \theta_1, \tau_1) \pi(\theta_1) \pi(\tau_1) \\ &\propto \exp \left(-\frac{\tau_1}{2} \sum_{i=1}^{12} (y_{1i} - \theta_1)^2 - \frac{1}{200} (\theta_1 - \theta_{10})^2 - 4\tau_1 \right) \tau_1^{5.01}.\end{aligned}$$

Therefore, for θ_1 , we have

$$\pi(\theta_1 | \tau_1, \mathbf{y}_1) \propto \exp \left(-\frac{12\tau_1 + 1/100}{2} \left(\theta_1 - \frac{1440\tau_1 + 110/100}{12\tau_1 + 1/100} \right)^2 \right),$$

which means $[\theta_1 | \tau_1, \mathbf{y}_1] \sim \mathcal{N}(\frac{1440\tau_1 + 110/100}{12\tau_1 + 1/100}, \frac{1}{12\tau_1 + 1/100})$. Similarly, we obtain $[\tau_1 | \theta_1, \mathbf{y}_1] \sim \mathcal{Ga}(6.01, 4 + \frac{\sum_{i=1}^{12} (y_{1i} - \theta_1)^2}{2})$.

Furthermore, we obtain the sampler for the $[\theta_2 | \tau_2, \mathbf{y}_2] \sim \mathcal{N}(\frac{707\tau_2 + 110/100}{7\tau_2 + 1/100}, \frac{1}{7\tau_2 + 1/100})$ and $[\tau_2 | \theta_2, \mathbf{y}_2] \sim \mathcal{Ga}(6.01, 4 + \frac{\sum_{i=1}^7 (y_{2i} - \theta_2)^2}{2})$.

(b) By simulation, we obtain that $\theta_1 - \theta_2 = 19$ and the proportion of positive differences equals to 1.

(c) The credible set is $[18.989, 19.010]$, which does not contain 0.

A Code for Problem 1

```
import numpy as np
import matplotlib.pyplot as plt
import matplotlib
import numpy.random as rand

matplotlib.use('TkAgg')
rand.seed(100)

class Metrolis_hasting(object):
    def post(self, a, x):
        #if not isinstance(x, float):
        #    raise Exception("Input Value Error!")
        f = np.exp(-1/8*np.sum((a-x)**2)) * 1/2 * np.cos(np.pi*x/4)**2
        return f

    def proposal(self, x, y):
        return 1

    def Update(self, a, x):
        y = rand.uniform(-2,2,1)
        ratio = min(1, self.post(a, y)*self.proposal(x, y)
                    / (self.post(a, x)*self.proposal(y, x)))
        accept_ratio = rand.uniform(0,1,1)
        if accept_ratio <= ratio:
            return y
        return x

a = np.array([-3, -3, 4, -7, 0, 4])
x=0
Sampler = Metrolis_hasting()

## warm_up
for _ in range(500):
    x = Sampler.Update(a, x)

## statistics
result = np.array([])
for _ in range(10000):
    x = Sampler.Update(a, x)
```

```
result = np.append(result , np.array(x) , 0)

estimate = np.mean(result)
result_sort = np.sort(result)
print(estimate)
print(result_sort[249] , result_sort[9749])
plt.hist(result , bins=20)
plt.show()
```

B Code for Problem 2

```
import numpy as np
import matplotlib.pyplot as plt
import matplotlib
import numpy.random as rand
matplotlib.use('TkAgg')

rand.seed(10)

class Gibbs(object):
    def normal(self, a, tau):
        #if not isinstance(x, float):
        #    raise Exception("Input Value Error!")
        p = len(a)
        f = (np.sum(a)*tau+110/100)/(p*tau+1/100) + np.sqrt(1/(p*tau+1/100))*
        return f

    def gamma(self, a, x):
        par = np.sum((a-x)**2)/2
        f = rand.gamma(6.01, 4+par)
        return f

a_1 = np.array([134,146,104,119,124,161,107,83,113,129,97,123])
a_2 = np.array([70,118,101,85,107,132,94])
theta_1=0
theta_2=0
tau_1 = 10
tau_2 = 10
Sampler = Gibbs()
### warm_up

for _ in range(500):
    theta_1 = Sampler.normal(a_1, tau_1)
    theta_2 = Sampler.normal(a_2, tau_2)
    tau_1 = Sampler.gamma(a_1, theta_1)
    tau_2 = Sampler.gamma(a_2, theta_2)

result_1 = np.array([])
result_2 = np.array([])
result_diff = np.array([])
```

```
count=0
### statistics
for _ in range(10000):
    theta_1 = Sampler.normal(a_1, tau_1)
    theta_2 = Sampler.normal(a_2, tau_2)
    tau_1 = Sampler.gamma(a_1, theta_1)
    tau_2 = Sampler.gamma(a_2, theta_2)
    if theta_1>theta_2:
        count += 1
    result_1 = np.append(result_1, np.array(theta_1), 0)
    result_2 = np.append(result_2, np.array(theta_2), 0)
    result_diff = np.append(result_diff, np.array(theta_1-theta_2), 0)

print(count/10000)

estimate_1 = np.mean(result_1)
estimate_2 = np.mean(result_2)
result_sort = np.sort(result_diff)

print(estimate_1)
print(estimate_2)
print(result_sort[249], result_sort[9749])
```