# **Introduction to Statistical Method**

# **Comparing Two Means and Two Variances**

# **Comparing Two Means - A Point Estimator**

We have two populations with different means  $\mu_1$  and  $\mu_2$ , the goal is to estimate the difference  $\mu_1 - \mu_2$  by taking a sample from each population in independent way.

Natural point estimator:  $\mu_1 - \mu_2 := \hat{\mu_1} - \hat{\mu_2} = \overline{X_1} - \overline{X_2}$ 

To determine confidence intervals and to test hypothesis we need to know the distribution  $\overline{X}_1 - \overline{X}_2$ 

#### **Theorem**

The  $\overline{X}_1$  and  $\overline{X}_2$  be the sample means based on independent random samples of size  $n_1$  and  $n_2$  drawn from normal distributions with mean  $\mu_1$  and  $\mu_2$  and variance  $\sigma_1^2$  and  $\sigma_2^2$ .

The  $\overline{X}_1-\overline{X}_2$  is normal with mean  $\mu_1-\mu_2$  and variance  $\sigma_1^2/n_1+\sigma_2^2/n_2$ 

$$rac{\overline{X}_1-\overline{X}_2-(\mu_1-\mu_2)}{\sqrt{\sigma_1^2/n_1+\sigma_2^2/n_2}}$$
 is standard normal random variable.

(Central Limit Theorem allows us to apply this result even to non-normal populations if we have really large sample sizes)

### **OC Curve Application**

$$d=rac{|\mu_1-\mu_2|}{\sqrt{\sigma_1^2+\sigma_2^2}}$$
 , if  $n=n_1=n_2$  , unchanged, else  $n=rac{\sigma_1^2+\sigma_2^2}{\sigma_1^2/n_1+\sigma_2^2/n_2}$ 

If the variances are unknown, we need some more sophisticated methods.

The unknown variances are equal, situation is much easier.

# **Comparing Two Variances**

Consider test types of this:

•  $H_0: \sigma_1^2 = \sigma_2^2$ ,  $H_1: \sigma_1^2 > \sigma_2^2$  (right-tailed test) •  $H_0: \sigma_1^2 = \sigma_2^2$ ,  $H_1: \sigma_1^2 \neq \sigma_2^2$  (two-tailed test)

we move the  $\sigma$  to one side to consider only about the quotient.

 $(n-1)S^2/\sigma^2$  follows a chi-squared distribution with n-1 degree of freedom.

If the variance is put into quotient, it is easier to handle

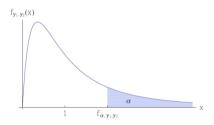
#### F-Distribution

 $X_{\gamma_1}^2$  and  $X_{\gamma_2}^2$  are independent chi-squared random variables with  $\gamma_1$  and  $\gamma_2$  degrees of freedom.

Random variable  $F_{\gamma_1,\gamma_2}=\frac{X_{\gamma_1}^2/\gamma_1}{X_{\gamma_2}^2/\gamma_2}$  is said to follow a F-distribution with  $\gamma_1$  and  $\gamma_2$  degree of freedom.

$$P[F_{\gamma_1,\gamma_2} < x] = P[1/F_{\gamma_1,\gamma_2} > 1/x] = 1 - P[F_{\gamma_2,\gamma_1} < 1/x]$$

Then 
$$f_{\gamma_1,\gamma_2}(x)=\gamma_1^{\gamma_1/2}\gamma_2^{\gamma_2/2}rac{\Gammaig(rac{\gamma_1+\gamma_2}{2}ig)}{\Gammaig(rac{\gamma_1}{2}ig)\Gammaig(rac{\gamma_2}{2}ig)}rac{x^{\gamma_1/2-1}}{(\gamma_1x+\gamma_2)^{(\gamma_1+\gamma_2)/2}}$$
 for  $x\geq 0$ 



Define 
$$f_{lpha,\gamma_1,\gamma_2}$$
 by  $P[F_{\gamma_1,\gamma_2}>f_{lpha,\gamma_1,\gamma_2}]=lpha$ 

Then 
$$\begin{aligned} 1-\alpha &= P[F_{\gamma_1,\gamma_2} \geq f_{1-\alpha,\gamma_1,\gamma_2}] \\ &= 1-P[F_{\gamma_1,\gamma_2} < f_{1-\alpha,\gamma_1,\gamma_2}] \\ &= P[F_{\gamma_2,\gamma_1} < 1/f_{1-\alpha,\gamma_1,\gamma_2}] \\ &= 1-P[F_{\gamma_2,\gamma_1} \geq 1/f_{1-\alpha,\gamma_1,\gamma_2}] \end{aligned} \text{ also we can see } \alpha = P[F_{\gamma_2,\gamma_1} \geq f_{\alpha,\gamma_2,\gamma_1}].$$

So 
$$f_{1-lpha,\gamma_1,\gamma_2}\cdot f_{lpha,\gamma_2,\gamma_1}=1$$
 .

#### Remark

Let  $S_1^2$  and  $S_2^2$  be sample variance based on independent random samples of size  $n_1$  and  $n_2$  from normal populations with means  $\mu_1$  and  $\mu_2$  and variance  $\sigma_1^2$  and  $\sigma_2^2$ 

If  $\sigma_1^2=\sigma_2^2$  then the statistic  $S_1^2/S_2^2$  follows F-distribution with  $n_1-1$  and  $n_2-1$  distribution.

Since 
$$F_{n_1-1,n_2-1}=rac{[(n_1-1)S_1^2/\sigma_1^2]/(n_1-1)}{[(n_2-1)S_2^2/\sigma_2^2]/(n_2-1)}=rac{\sigma_2^2S_1^2}{\sigma_1^2S_2^2}$$
 , so it is trivial to require  $\sigma_1^2=\sigma_2^2$ 

#### F-Test

We can derive F-Test from F-distribution that:

$$H_0:\sigma_1=\sigma_2$$
 based on  $F_{n_1-1,n_2-1}=rac{S_1^2}{S_2^2}$  is a F-Test

We reject  $H_0$  at significance level lpha

- ullet in favor of  $H_1:\sigma_1>\sigma_2$  if  $rac{S_1^2}{S_2^2}>f_{lpha,n_1-1,n_2-1}$
- ullet in favor of  $H_1:\sigma_1<\sigma_2$  if  $rac{S_2^{\overline{2}}}{S_1^2}>f_{lpha,n_2-1,n_1-1}$
- ullet in favor of  $H_1:\sigma_1
  eq\sigma_2$  if  $rac{\hat{S_1^2}}{\hat{S_2^2}}>f_{lpha/2,n_1-1,n_2-1}$  or  $rac{S_2^2}{S_1^2}>f_{lpha/2,n_2-1,n_1-1}$

When testing to see whether two population variances are equal for the purpose of comparing their means, one hopes to not reject  $H_0$ .

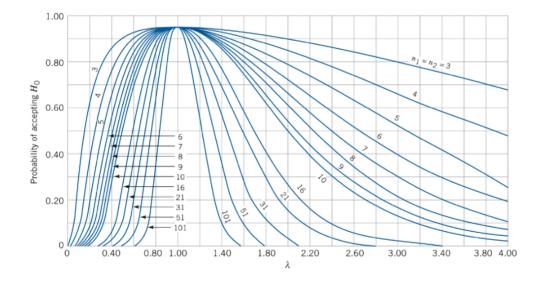
If  $H_0$  is not rejected, one can assume that the variances are in fact equal and continue with the test for equality for means.

In this case, a small Type II error  $\beta$  is more important than  $\alpha$  small.

#### **OC Curves for F-Test**

For case  $n=n_1=n_2$ , the OC curves plotting eta against the parameter  $\lambda=rac{\sigma_1}{\sigma_2}.$ 

The curves are for both one- two- sided alternatives.



### **Comparing Two Means - Equal Variances**

$$rac{\overline{X}_1-\overline{X}_2-(\mu_1-\mu_2)}{\sqrt{\sigma_1^2/n_1+\sigma_2^2/n_2}}$$
 follows standard normal distribution.

We now want to estimate  $\sigma^2$ .

The pooled estimator is  $S_p^2=rac{(n_1-1)S_1^2+(n_2-1)S_2^2}{n_1+n_2-2}$ 

$$X_{n_1+n_2-2}^2 = rac{(n+1+n_2-2)S_p^2}{\sigma^2} = rac{(n_1-1)S_1^2}{\sigma^2} + rac{(n_2-1)S_2^2}{\sigma^2}$$

Furthermore, 
$$T_{n_1+n_2-2}=rac{Z}{\sqrt{X_{n_1+n_2-2}^2/(n_1+n_2-2)}}=rac{(\overline{X}_1-\overline{X}_2)-(\mu_1-\mu_2)}{\sqrt{S_p^2(1/n_1+1/n_2)}}$$
 follows T-distribution

with  $n_1 + n_2 - 2$  degree of freedom.

So the 
$$100(1-lpha)\%$$
 confidence interval for  $\mu_1-\mu_2$  is  $(\overline{X}_1-\overline{X}_2)\pm t_{lpha/2,n_1+n_2-2}\sqrt{S_p^2(1/n_1+1/n_2)}$ 

# **Pooled T-Test - Variance Equal**

Let  $X_1^{(i)} ... X_{n_i}^{(i)}$ , i = 1, 2 be random samples of size  $n_i$  from two normal distributions with means  $\mu_i$  and identical  $\sigma^2$ .

 $S_p^2$  be the pooled sample variance and  $(\mu_1 - \mu_2)_0$  a null value for difference of means.

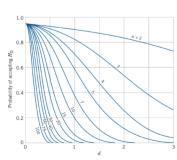
Then Test 
$$H_0: \mu_1-\mu_2=(\mu_1-\mu_2)_0$$
 based on  $T_{n_1+n_2-2}=\dfrac{(\overline{X}_1-\overline{X}_2)-(\mu_1-\mu_2)}{\sqrt{S_p^2(1/n_1+1/n_2)}}$  is a pooled test for equality of means.

We reject  $H_0$  at significance level  $\alpha$ 

- in favor of  $H_1: \mu_1 \mu_2 \neq (\mu_1 \mu_2)_0$  if  $|T_{n_1+n_2-2}| > t_{\alpha/2,n_1+n_2-2}$
- ullet in favor of  $H_1: \mu_1-\mu_2>(\mu_1-\mu_2)_0$  if  $T_{n_1+n_2-2}>t_{lpha,n_1+n_2-2}$
- in favor of  $H_1: \mu_1 \mu_2 < (\mu_1 \mu_2)_0$  if  $T_{n_1 + n_2 2} < -t_{\alpha, n_1 + n_2 2}$

# **OC Curves T Test - Variance Equal**

Equal variance  $\sigma^2$  and equal sample size  $n_1=n_2=n$ ,  $d=\frac{|\mu_1-\mu_2|}{2\sigma}$ , we must use the modified sample size  $n^*=2n-1$ . The  $\sigma$  can be substitute with an estimated one or express the deviation in terms of  $\sigma$ .



### **Unequal Variances**

$$\frac{\overline{X}_1 - \overline{X}_2 - (\mu_1 - \mu_2)}{\sqrt{\sigma_1^2/n_1 + \sigma_2^2/n_2}}$$
 for the unequal variance  $\sigma_1$  and  $\sigma_2$ , we can estimate the variance to get the statistic  $(\overline{X}_1 - \overline{X}_2) = (\mu_1 - \mu_2)$ 

$$T_{\gamma} = \frac{(\overline{X}_1 - \overline{X}_2) - (\mu_1 - \mu_2)_0}{\sqrt{S_1^2/n_1 + S_2^2/n_2}} \text{ where the } \gamma \text{ for the degree of freedom is } \gamma = \frac{(S_1^2/n_1 + S_2^2/n_2)^2}{\frac{(S_1^2/n_1)^2}{n_1 - 1} + \frac{(S_2^2/n_2)^2}{n_2 - 1}}$$

# **Pooled T-Test - Variances Unequal**

We reject  $H_0$  at significance level lpha

- ullet in favor of  $H_1: \mu_1-\mu_2
  eq (\mu_1-\mu_2)_0$  if  $|T_\gamma|>t_{lpha/2,\gamma}$
- ullet in favor of  $H_1: \mu_1-\mu_2>(\mu_1-\mu_2)_0$  if  $T_\gamma>t_{lpha,\gamma}$
- ullet in favor of  $H_1: \mu_1-\mu_2<(\mu_1-\mu_2)_0$  if  $T_\gamma<-t_{lpha,\gamma}$

### **Paired T-Test**