# COMS W4701: Artificial Intelligence

Lecture 10: Bayesian Networks

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### Review: Inference in HMMs

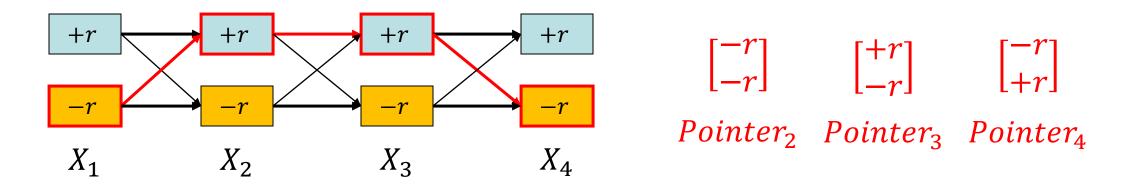
- Forward algorithm for state estimation: Find  $f_t = P(X_t|e_{1:t})$

- Viterbi algorithm to find most likely sequence of states:  $argmax_{x_{1},t} P(x_{1:t}|e_{1:t})$
- Compute the joint distribution  $m_t = \max_t P(x_{1:t-1}, X_t, e_{1:t})$
- Start from  $m_0 = P(X_0)$  and compute:

$$m'_{t+1}(x_{t+1}) = \max_{x_t} P(x_{t+1}|x_t) m_t(x_t)$$
 $m_{t+1} = O_{t+1} m'_{t+1}$ 

Start from 
$$m_0 = P(X_0)$$
 and compute: 
$$m'_{t+1}(x_{t+1}) = \max_{x_t} \frac{P(x_{t+1}|x_t)m_t(x_t)}{x_t} = \begin{pmatrix} \max(T_{11}m_t(1), T_{12}m_t(2), ..., T_{1n}m_t(n)) \\ \max(T_{21}m_t(1), T_{22}m_t(2), ..., T_{2n}m_t(n)) \\ \vdots \\ \max(T_{n1}m_t(1), T_{n2}m_t(2), ..., T_{nn}m_t(n)) \end{pmatrix}$$

### Review: Viterbi Algorithm



- Also need to keep track of most likely prior state when computing each  $m{m}_{t+1}$ 

$$Pointer_{t+1}(x_{t+1}) = \underset{x_t}{\operatorname{argmax}} P(x_{t+1}|x_t) m_t(x_t)$$

- $\underset{x_T}{\operatorname{argmax}} m_T$  is the last state in most likely sequence
- Backward pass: Follow pointers backward to iteratively obtain each prior state

## Today

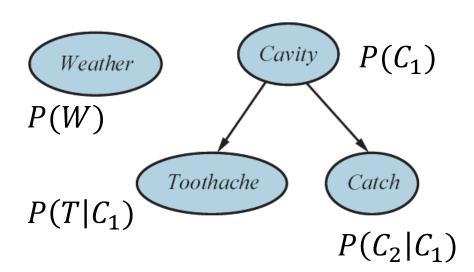
 Last time: Hidden Markov models help us reason about belief states that evolve over space or time while incorporating observations

 This time: Bayesian networks are general graph structures that can represent arbitrary relationships and independences

Naïve Bayes models and maximum likelihood learning

### **Bayesian Networks**

- Recall: A set of RVs is described by their joint distribution
- Bayesian network: A directed acyclic graph (DAG) representation of a distribution
- Each node corresponds to a random variable
- Each edge indicates influence or correlation
- May also be causation, but not always



- Parameters of the Bayes net: A conditional probability table for each node
- The table for a node  $X_i$  contains the values  $P(X_i|parents(X_i))$

### Joint Distribution

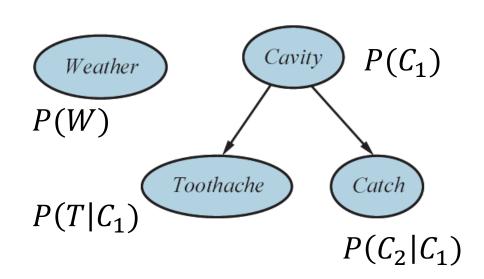
Recall that the chain rule gives us a general way to compute joint distributions:

$$P(x_1, \dots, x_n) = P(x_1)P(x_2|x_1)\cdots P(x_n|x_1, \dots, x_{n-1}) = \prod_{i=1}^{n} P(x_i|x_1, \dots, x_{i-1})$$

By definition, a joint distribution in a Bayes net can be written as

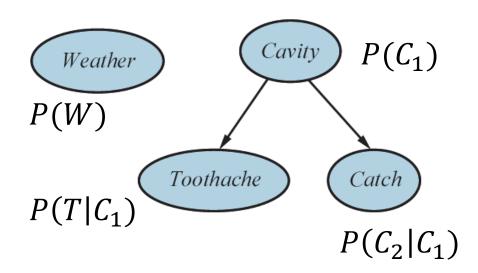
$$P(x_1, ..., x_n) = \prod_{i=1}^n P(x_i | parents(X_i))$$

- This can simplify a lot of computation!
- Example:  $P(w, c_1, t, c_2)$ =  $P(w)P(c_1)P(t|c_1)P(c_2|c_1)$



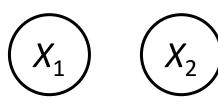
## **Topological Ordering**

- We asserted that  $P(x_1,...,x_n) = \prod_{i=1}^n P(x_i|parents(X_i))$
- This implies that a node is conditionally independent of predecessors given its parents
- Numbering must correspond to topological ordering: parents before children
- One possible ordering: W,  $C_1$ , T,  $C_2$
- Another one:  $C_1$ ,  $C_2$ , T, W
- Another one:  $C_1$ ,  $C_2$ , W, T
- Many more...



### **Constructing Bayes Nets**

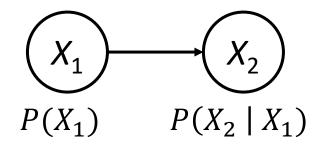
Bayes net representations are not unique!



 $P(X_2)$ 

 $P(X_1)$ 

- Example: 2 (independent) coin flips
- Can use either  $P(X_2)$  or  $P(X_2|X_1)$
- Independence is deduced from probabilities
- No path guarantees independence between two RVs
- Presence of a path may indicate correlation between two RVs
- May also be independent, depending on probabilities
- Ex:  $P(X_2|X_1)$  is storing redundant information



$X_1$	$X_2$	$P(X_2 X_1)$
+X	+x	0.5
+X	-X	0.5
-X	+x	0.5
-X	-X	0.5

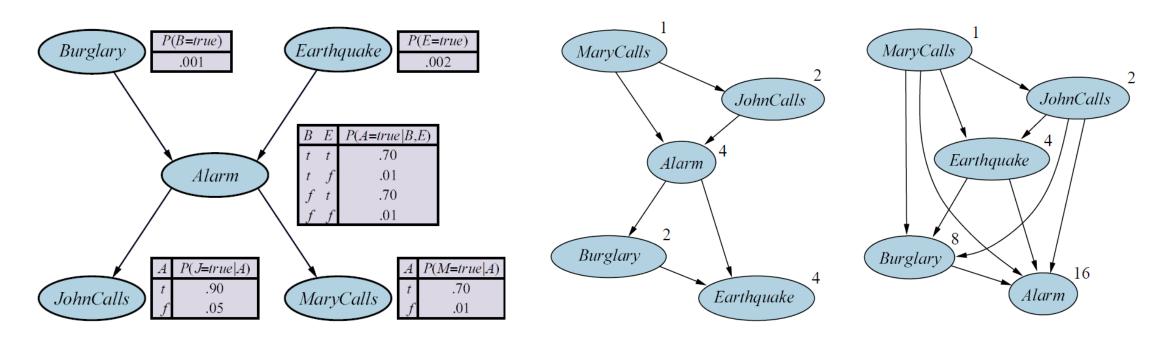
### Compactness

- When constructing a Bayes net, compact structures are more efficient
- The more independences that we can assert, the fewer the connections and parameters that we need to store

- Consider a distribution of n RVs, each with domain size d
- Full joint distribution is specified by a table of  $d^n$  numbers
- If any node has at most k parents, then its distribution is  $O(d^{k+1})$
- We only need  $O(n \times d^{k+1})$  parameters to store all information

### Example: Alarm Network

- Boolean random variables: Each table has  $2^k$  parameters
- All three structures can correctly represent same joint distribution
- Optimal structure has 10 parameters in total; others have 13 and 31



### Conditional Independences

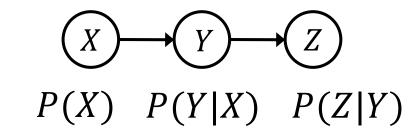
- Given a set of evidence variables, what conditional independences arise?
- We know that a node is independent of its "ancestors" given all its parents

 More generally, two nodes are conditionally independent if they are dseparated with colliders in all paths between them

- Here, "paths" do not depend on edge directionality!
- We identify three scenarios, which can generalize to any situation

### Causal Chain

 Nodes in a causal chain (edges pointing in same direction) generally not independent



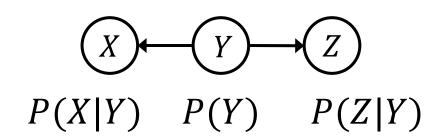
- But an observed node acts a collider
- Nodes on one side become independent of nodes on the other

$$X \longrightarrow Y \longrightarrow Z$$

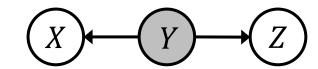
Proof: 
$$P(x,z|y) = \frac{P(x)P(y|x)P(z|y)}{P(y)}$$
$$= \frac{P(x)P(z|y)}{P(y)} \frac{P(x|y)P(y)}{P(x)} = P(z|y)P(x|y)$$

### Common Cause

 Nodes connected by a common cause (arc heads pointing away from each other) are generally not independent



- But an observed node cause acts a collider
- Nodes on one side become independent of nodes on the other



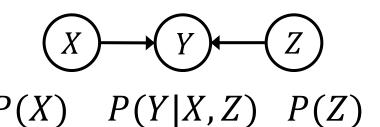
• Proof: 
$$P(x,z|y) = \frac{P(x,y,z)}{P(y)} = \frac{P(y)P(x|y)(z|y)}{P(y)} = P(x|y)P(z|y)$$

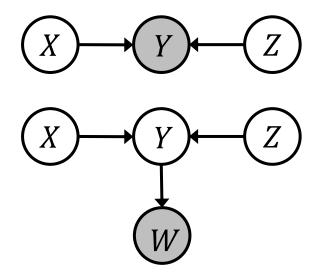
### Common Effect

- A common effect (arc heads pointing toward each other) between two nodes acts a collider
- Nodes on either side are independent

But observing the effect or any of its descendants removes the collider!

 Nodes are all connected; independence no longer guaranteed

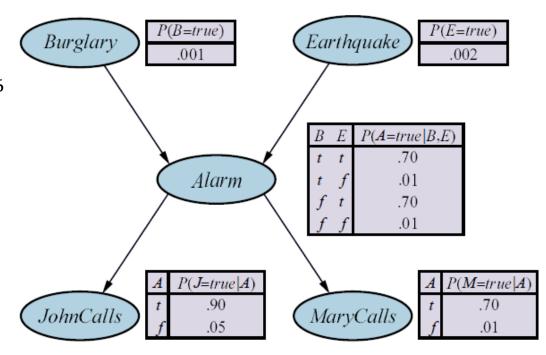




### Example: Common Effect

$$P(+b,+e) = P(+b)P(+e) = 0.001 \times 0.002 = 2 \times 10^{-6}$$

$$P(+b,+e|+a) = \frac{P(+b,+e,+a)}{P(+a)}$$
$$= \frac{P(+b)P(+e)P(+a|+b,+e)}{\sum_{b,e} P(b)P(e)P(+a|b,e)}$$

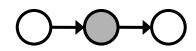


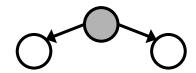
$$= \frac{0.001 \times 0.002 \times 0.7}{.001(.002)(.7) + .001(.998)(.01) + .999(.002)(.7) + .999(.998)(.01)}$$
$$= 1.23 \times 10^{-4}$$

### **D-Separation**

- Question: Are two nodes  $X_i$  and  $X_j$  independent conditioned on a set of nodes E?
- First identify "conditioned" nodes (shade them in)
- If at least one path between  $X_i$  and  $X_j$  is unblocked (no colliders): not conditionally independent
- If every path between  $X_i$  and  $X_j$  contains a collider: conditionally independent!

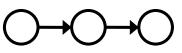
Collider

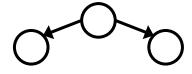


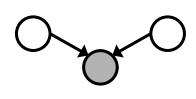


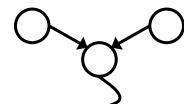


No collider







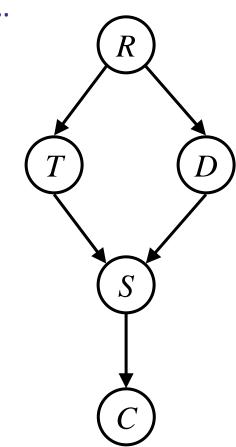


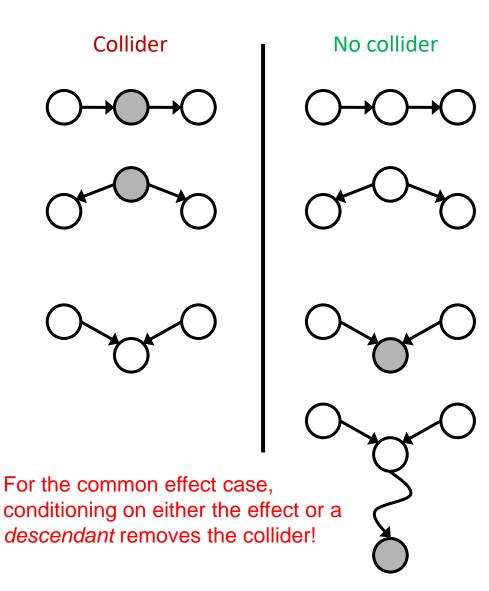
For the common effect case, conditioning on either the effect or a descendant removes the collider!

### **Example: D-Separation**

#### Which nodes are independent...

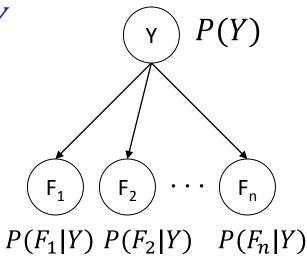
- Given *S*?
- Given R?
- Given T or D?
- Given T and D?
- Given R and S?
- Given R and C?





### Naïve Bayes Models

- In a naïve Bayes model, there is an underlying "cause" variable Y
- Y then influences a number of different effects  $F_i$
- Given Y, all the  $F_i$  are conditionally independent!
- Joint distribution:  $P(y, f_1, ..., f_n) = P(y) \prod_{i=1}^n P(f_i|y)$



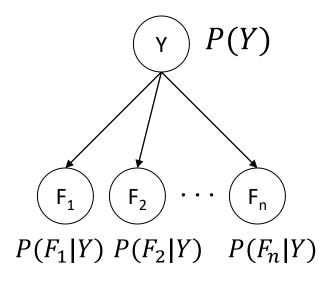
How to infer the cause given observed effects?

$$P(Y|f_1,...,f_n) = \frac{P(Y,f_1,...,f_n)}{P(f_1,...,f_n)} \propto_Y P(Y) \prod_{i=1}^n P(f_i|Y)$$

### Classification

- Naïve Bayes can be used for classification tasks
- Given a data input, predict an output label or class
- Inputs are features  $f_i$  and output is most likely class y
- Predicted class:

$$y = \underset{y}{\operatorname{argmax}} P(y|f_1, \dots, f_n) = \underset{y}{\operatorname{argmax}} P(y) \prod_{i=1}^n P(f_i|y)$$



- We don't even need to normalize—exact probabilities are not required
- Problem: If we have too many features, product will underflow

### Log Probabilities

- As with Viterbi, NB classification only looks at relative ordering of likelihoods, not the actual probabilities
- Logarithm function preserves ordering and can address underflow issues

$$\underset{y}{\operatorname{argmax}} P(y) \prod_{i=1}^{n} P(f_i|y) = \underset{y}{\operatorname{argmax}} \log \left( P(y) \prod_{i=1}^{n} P(f_i|y) \right)$$

Log of a product = sum of logs

$$\log\left(P(y)\prod_{i=1}^{n}P(f_{i}|y)\right) = \log(P(y)) + \sum_{i=1}^{n}\log(P(f_{i}|y))$$

• Since probabilities  $\leq 1$ , log probabilities are always  $\leq 0$ 

### Classification Examples

0

??

- Spam filter: Inputs are emails, class/label is spam/ham
- Emails contain features that will help classify
  - Words: "wire money", "bank account", "supplements"
  - Text patterns: all caps, unusual font, dollar signs
  - Other flags: Sender not in contacts, geographic origin
- Digit recognition: Inputs are pixels, class/label is a digit
- Pixel grids contain features that will help classify
  - Whether specific pixels are ON or OFF
  - Patterns: Size, number of loops, number of lines

### Example: Spam Filtering

- Labels  $Y = \{\text{spam, ham}\}$
- Features  $W_i$  for position i in the email over common dictionary words
- "Bag-of-words" model: Each  $W_i$  is independent, identically distributed
  - No dependence on ordering or other nearby words!

Word	P(Y)	Hey	would	you	like	to	lose	weight
P(w spam)	0.33333	0.00002	0.00069	0.00881	0.00086	0.01517	0.00008	0.00016
P(w ham)	0.66666	0.00021	0.00084	0.00304	0.00083	0.01339	0.00002	0.00002

- $\log P(spam) + \sum \log P(w|spam) = -53.3$

## Supervised Learning

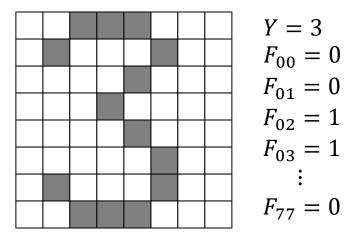
- Where do the probabilities (parameters) of our models come from?
- **Supervised learning**: Given *training data* with input-output pairs  $(x_1, y_1), ..., (x_N, y_N)$  generated by an unknown function f, find a *hypothesis* to best approximate f
- Expressiveness-complexity tradeoff: The more expressive the hypothesis space, the higher the computational complexity of searching for a good hypothesis
- Bias-variance tradeoff: More complex hypotheses that fit training data well, or simpler hypotheses that generalize to new data better?

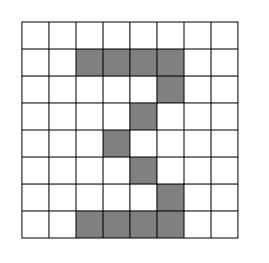
## Learning Naïve Bayes Models

- If our data are identical and independently distributed (i.i.d.), the likelihood of all of them occurring given a hypothesis h is  $P(data|h) = \prod_j P(data_j|h)$
- Maximum likelihood learning: Find the hypothesis that maximizes P(data|h)
- More specifically, find the hypothesis parameters  $\theta$  that maximize  $L(\theta) = P(data|h_{\theta})$
- For Naïve Bayes, the parameters are the individual model probabilities P(y),  $P(f_i|y)$
- $L(\theta)$  is maximized when the probabilities are chosen by simply counting the occurrences of each event in the data

## Example: Digit Recognition

- Labels  $Y = \{0,1,2,3,4,5,6,7,8,9\}$
- Binary features  $F_{ij}$  for each pixel (i,j) in the input image
- Suppose we have two training examples of 3s



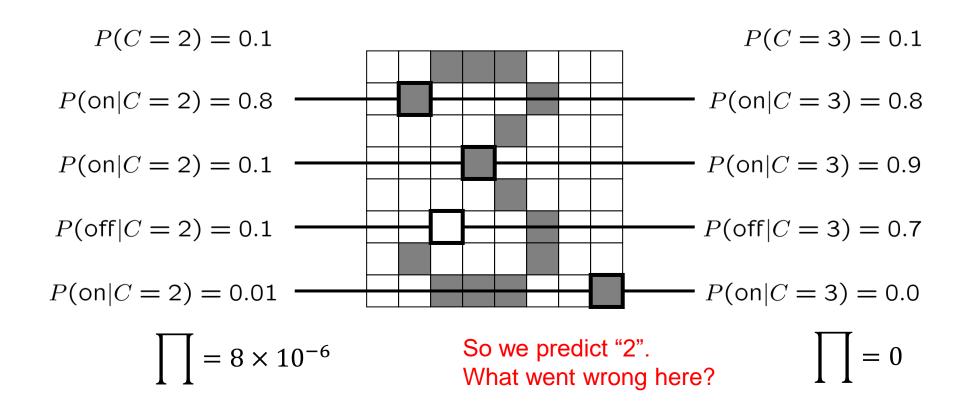


$$Y = 3$$
  $P(F_{00} = 1 | Y = 3) = 0$   
 $F_{00} = 0$   $P(F_{01} = 1 | Y = 3) = 0$   
 $P(F_{02} = 1 | Y = 3) = 0.5$   
 $F_{12} = 1$   $P(F_{13} = 1 | Y = 3) = 1$   
 $P(F_{13} = 1 | Y = 3) = 1$   
 $P(F_{14} = 1 | Y = 3) = 1$   
 $P(F_{15} = 1 | Y = 3) = 0$ 

In practice, we will be using much more than just 2 training samples

## Overfitting: Digit Recognition

Suppose we learned a model to classify 2 or 3, but our training data has no
 3s with the bottom-right pixel on



### Overfitting: Spam Detection

 We will not see every dictionary word in training data—0s all over the place in the model!

 If any word appears "ham" data but never in "spam", then any email with that word will be classified as "ham" (and vice-versa)

Dear Sir/Madam, I would like to seriously offer you an exclusive opportunity to win \$1M by simply sending us your bank info" -> HAM P(W|spam)

south-west: 0
nation: 0
morally: 0
nicely: 0
extent: 0
seriously: 0

P(W|ham)

screens : 0
minute : 0
guaranteed : 0
\$205.00 : 0
delivery : 0
signature : 0

### **Additive Smoothing**

- Problem: If a specific class/feature does not appear in training data, probability assumed to be 0!
  - This will automatically zero out  $P(Y, F_1, ..., F_n) = P(Y) \prod_i P(F_i | Y)$
  - To avoid overfitting, we need to **generalize** our models
- Idea: Smooth our estimates by adding a "pseudocount" k

$$\widehat{\theta} = \frac{c(x) + k}{N + k|X|}$$

$$\widehat{\theta} = \frac{c(x)}{N}$$
Original maximum likelihood estimator (average)
$$\widehat{\theta} = \frac{1}{|X|}$$
Uniform prior over all possibilities of  $X$  (ignore all samples)

## **Example: Additive Smoothing**

Estimate probabilities of red, blue, and green marbles







Color	$\hat{P}_{MLE}$ (color) = $\hat{P}_{k0}$ (color)
Red	0.667
Blue	0.333
Green	0

Color	$\widehat{P}_{k1}(\operatorname{color})$
Red	(2+1)/(3+3)=0.5
Blue	(1+1)/(3+3)=0.333
Green	(0+1)/(3+3)=0.167

Color	$\widehat{P}_{k100}$ (color)
Red	(2+100)/(3+300)=0.337
Blue	(1+100)/(3+300)=0.333
Green	(0+100)/(3+300)=0.330

$$\widehat{\theta} = \frac{c(x) + k}{N + k|X|}$$

$$k \to 0$$

$$\widehat{\theta} = \frac{c(x)}{N}$$

Original maximum likelihood estimator (average)

$$\widehat{\theta} = \frac{1}{|\lambda|}$$

Uniform prior over all possibilities of *X* (ignore all samples)

### Summary

- Bayesian networks graphically encode independence assumptions about joint distributions in a compact way
- D-separation rules can help infer additional independences given evidence
- Naïve Bayes models are simple ways of thinking about cause-effect or class-label relationships
- Supervised learning of model parameters can be done using maximum likelihood learning (for Naïve Bayes, simple counting)