

Midterm: 6.30-7.30 pm, Tuesday (tomorrow), IRC 2

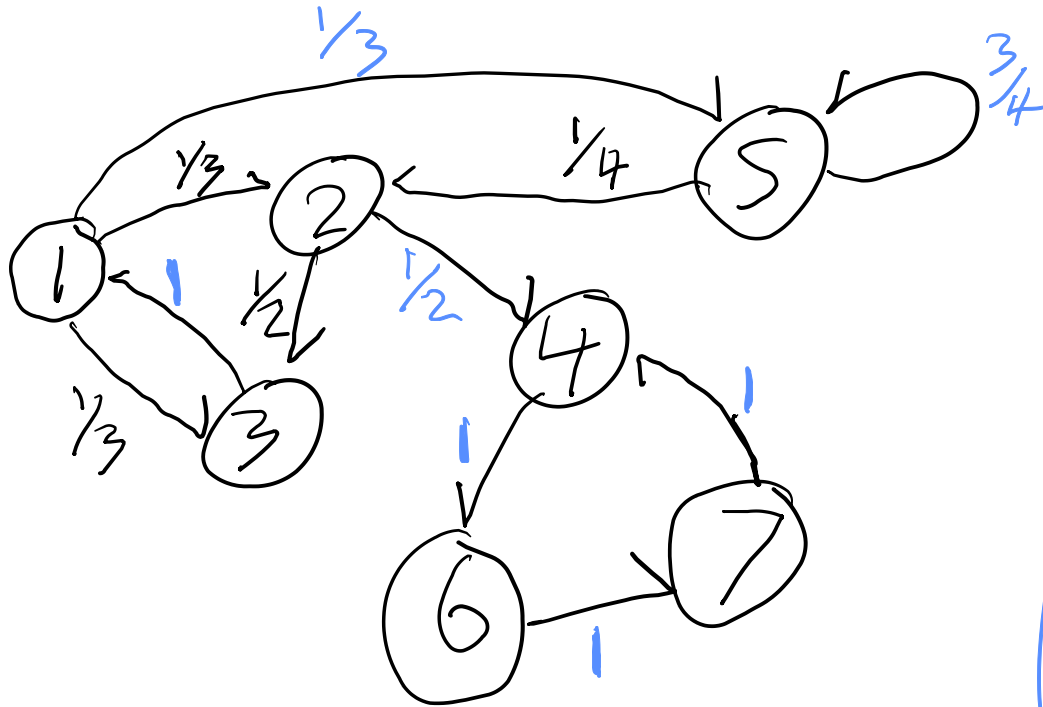
Wednesday: no class!

Lecture 19

Review

Problem 1

Fill in the missing probabilities on the following transition diagram, then determine the transition matrix.



	1	2	3	4	5	6	7
1		$\frac{1}{3}$	$\frac{1}{3}$		$\frac{1}{3}$		
			$\frac{1}{2}$	$\frac{1}{2}$			
		$\frac{1}{4}$			$\frac{3}{4}$		

Classify states:

Communicating classes: not closed

$\{1, 2, 3, 5\}$ transient, period = 1

$\{4, 6, 7\}$ recurrent, period = 3
closed

Starting in State 5, what is the probability that the MC visits State 6 within 6 steps?

$$A = \{ \text{visit 6 w/in 6 steps} \}$$
$$= \{ \text{visit 4 w/in 5 steps} \}$$

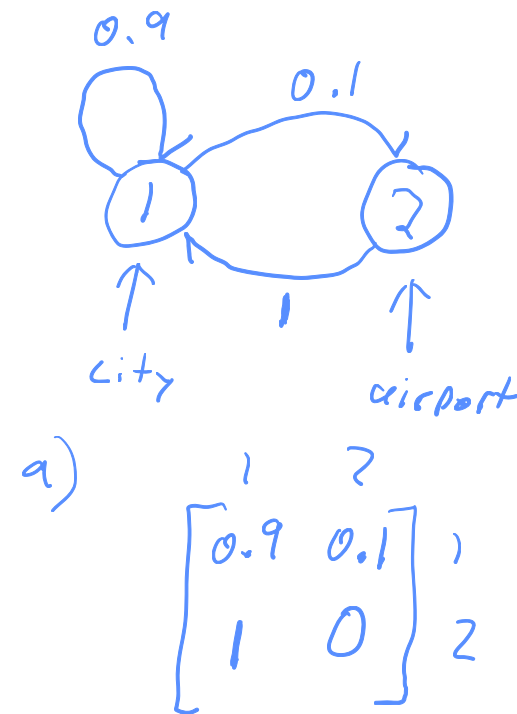
$$P(A) = \sum \text{prob. of path that get to 4 w/in 5 steps}$$

$$= P(5 \rightarrow 2 \rightarrow 4) + P(5 \rightarrow 5 \rightarrow 2 \rightarrow 4) + P(5 \rightarrow 5 \rightarrow 5 \rightarrow 2 \rightarrow 4)$$
$$+ P(5 \rightarrow 2 \rightarrow 3 \rightarrow 1 \rightarrow 2 \rightarrow 4) + P(5 \rightarrow 5 \rightarrow 5 \rightarrow 2 \rightarrow 4)$$

$$= \left[\frac{1}{4} \cdot \frac{1}{2} + \frac{3}{4} \cdot \frac{1}{4} \cdot \frac{1}{2} + \left(\frac{3}{4}\right)^2 \cdot \frac{1}{4} \cdot \frac{1}{2} + \left(\frac{3}{4}\right)^3 \cdot \frac{1}{4} \cdot \frac{1}{2} + \frac{1}{4} \cdot \frac{1}{2} \cdot \frac{1}{3} \cdot \frac{1}{2} \right]$$

Smith drives a taxi that serves the city and the airport. A trip that originates in the city has a destination in the city with probability 0.9, and has the airport as destination with probability 0.1. A trip that originates in the airport always goes to the city.

- (2 marks) Let X_n denote Smith's location (city or airport) after his n^{th} trip. This defines a Markov chain. What is its transition matrix?
- (2 marks) Determine the stationary distribution of the Markov chain.
- (1 mark) What fraction of trips originate in the city, in the long run?
- (2 marks) What fraction of all trips are trips from the city to the airport?
- (3 marks) Smith makes an average profit of \$8 for trips that remain in the city, and an average profit of \$12 for trips that involve the airport. What is his overall average profit per trip?



b) Reversible, Balance eqns:

$$\pi_1 \cdot 0.1 = \pi_2 \cdot 1, \quad 1 = \pi_1 + \pi_2 = \pi_1 + 0.1\pi_1 = 1.1\pi_1,$$

$$\Rightarrow \boxed{\pi_1 = \frac{1}{1.1} = \frac{10}{11}, \quad \pi_2 = \frac{1}{11}}$$

c) Note MC is irreducible & finite state \Rightarrow recurrent
By "big thm", frac of trips originating in (1) = $\pi_1 = \frac{10}{11}$

d) (proportion of trips from city) $\cdot P(\text{city to airport})$

$$= \pi_1 \cdot p_{12} = \frac{10}{11} \cdot 0.1 = \boxed{\frac{1}{11}}$$

e) $8 \cdot P(\text{city to city}) + 12 \cdot P(\text{city} \rightarrow \text{airport or airport} \rightarrow \text{city})$

$$= 8 \cdot \pi_1 \cdot p_{11} + 12 \cdot (\pi_1 \cdot p_{12} + \pi_2 \cdot p_{2,1})$$

$$= \boxed{8 \cdot \frac{10}{11} \cdot \frac{9}{10} + 12 \cdot \left(\frac{10}{11} \cdot \frac{1}{10} + \frac{1}{11} \cdot 1 \right)}$$

One hundred balls, some of them black and some of them white, are in an urn. At each time step, a ball is chosen from the urn uniformly at random, and is replaced by a white ball with probability p and by a black ball with probability $1 - p$ (independently for each time). Here $0 < p < 1$. Let X_n denote the number of white balls after the n^{th} replacement.

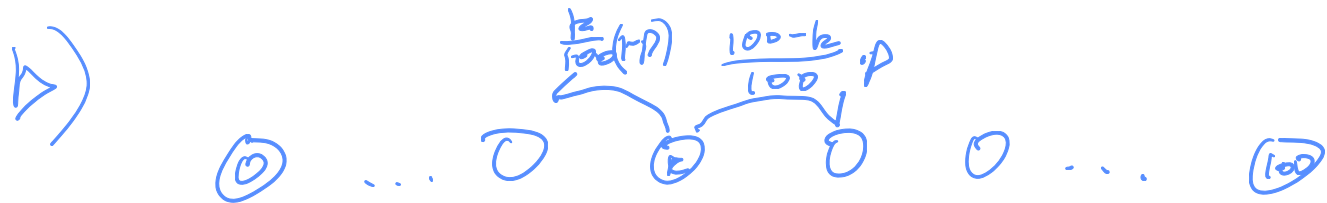
- (a) (3 marks) Find the transition probabilities for this Markov chain.
- (b) (5 marks) Using any method, determine the stationary distribution of the Markov chain.
- (c) (2 marks) Suppose that there are initially only black balls in the urn. How long will it take, on average, until there are again only black balls in the urn?

$$\begin{aligned}
 P_{k,k+1} &= P(X_1 = k+1 \mid X_0 = k) = P(\text{\# of white balls goes up by 1} \mid X_0 = k) \\
 &= P(\text{choosing a black ball, replacing w/ a white ball} \mid X_0 = k) \\
 &= \frac{100-k}{100} \cdot p
 \end{aligned}$$

\uparrow choose black \uparrow replace w/ white

$$P_{k,k-1} = \frac{k}{100} \cdot (1-p)$$

$$P_{k,k} = 1 - P_{k,k+1} - P_{k,k-1} = 1 - \frac{100-k}{100} \cdot p - \frac{k}{100} \cdot (1-p)$$



Reversible. Detailed balance:

$$\pi_k \cdot P_{k,k+1} = \pi_{k+1} P_{k+1,k}$$

$$\sum \pi_k = 1$$

$$\pi_k \cdot \frac{100-k}{100} \cdot p = \pi_{k+1} \cdot \frac{k+1}{100} \cdot (1-p)$$

$$\Rightarrow \pi_{k+1} = \frac{100-k}{k+1} \cdot \frac{p}{1-p} \cdot \pi_k$$

$$\pi_1 = \frac{100-0}{1} \cdot \frac{p}{1-p} \cdot \pi_0 = 100 \cdot \frac{p}{1-p} \cdot \pi_0$$

$$\pi_2 = \frac{100-1}{2} \cdot \frac{p}{1-p} \cdot \pi_1 = \frac{100 \cdot 99}{2} \cdot \left(\frac{p}{1-p}\right)^2 \cdot \pi_0$$

⋮

$$\pi_k = \frac{100!}{(100-k)! k!} \cdot \left(\frac{p}{1-p}\right)^k \cdot \pi_0$$

$$\text{Now, } 1 = \sum \pi_k = \sum \binom{100}{k} \cdot p^k \cdot \frac{1}{(1-p)^k} \cdot \pi_0$$

$$= \sum \binom{100}{k} \cdot p^k \cdot \underbrace{(1-p)^{100-k}}_{\frac{(1-p)^{100}}{(1-p)^k}} \cdot \frac{\pi_0}{(1-p)^{100}}$$

so
series
sums
to 1

$$= \sum P(\text{Bin}(100, p) = k) = 1$$

$$\Rightarrow \boxed{\pi_k = P(\text{Bin}(100, p) = k) = \binom{100}{k} \cdot p^k (1-p)^{100-k}}$$

c) Chain is irreducible and recurrent,
thus, by "big thm" $E \left[\begin{matrix} \text{return time to} \\ 0 \end{matrix} \right]$

Note: Recurrent \Rightarrow pos
recurrent since MC
is finite state

$$\begin{aligned} &= \frac{1}{\pi_0} = \frac{1}{\binom{100}{0} \cdot p^0 (1-p)^{100}} \\ &= \boxed{\frac{1}{(1-p)^{100}}} \end{aligned}$$

Sheila goes to an alien casino with n coins. In the first round of play, she

1. Flips each of the n coins three times, and gives all of the coins that weren't HHH to the casino.
2. For each coin that was HHH, she rolls a q -sided die (with sides labeled 1, 2, ..., q). For each roll, the casino gives her that many coins. *(she also keeps HHH coin)*

After the first round, she keeps playing again and again, always with all of her coins. At most, how many sides does the die have?

Suppose $n=1$. Then this is a branching process. Recall: $P(\text{extinction})=1 \iff E[Z] \leq 1$

$$E[Z] = E\left[\frac{Z}{2} \mid HHH\right] \cdot P(HHH) + E\left[\frac{Z}{2} \mid HHH^c\right] \cdot P(HHH^c)$$

\uparrow
 descendant law

$$= \left(1 + E[\text{die}]\right) \frac{1}{8} = \left(1 + \frac{q}{2}\right) \frac{1}{8} \leq 1$$

$$1 + \frac{q+1}{2} \leq 8 \Rightarrow \frac{q+1}{2} \leq 7 \Rightarrow q+1 \leq 14 \Rightarrow \boxed{q \leq 13}$$

\nwarrow
casino

For $n > 1$ each coin she starts w/ becomes its own independent branching process.

$$P(\text{all go extinct}) = P(1 \text{ goes extinct})^n$$

$$\Rightarrow \text{As before } P(\text{all go extinct}) = 1 \iff \mathbb{E} Z \leq 1$$

$$\text{Again; } \boxed{q \leq 13}$$