

Analyses of the high frequency time series on example of Circadian rhythms in the Long-Tail Pocket mouse.

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Abstract

The present study aimed to examine the forecasting performance of various univariate approaches to forecast high-frequency time series which contains more than one seasonal pattern that must be taken into consideration.

To achieve the purpose of the study, we performed the following procedures with major findings. First, we investigated different frequencies and produced traditional forecasts for it using ts object. The limitation of ts objects for only low frequency data and also for using multiple seasonalities led us to make additional sets of forecasts by applying non-benchmark approaches with double seasonality using msts object with Fourier terms for each seasonal pattern. This indicates that forecast accuracy was able to be improved by incorporating two different types of seasonal patterns simultaneously.

Seasonality is the main component of time series, and the consideration of seasonality has become more important with the increasing frequency of time series produced in industry. Circadian rhythms in the Long-Tail Pocket mouse time series is a representative example of business time series data that has been collected every two minutes for a long period of time. The high frequency data, with time series containing closely spaced time intervals creates a necessity of creating accurate modeling approaches for modeling seasonal patterns of time series.

As each seasonal pattern has distinct periods and effects, it is not trivial to design a model that can capture multiple seasonal patterns at once. We found that fourier regression gives more accurate results and can detect multiple seasonalities better then etc and tbats and sarima.

Introduction

The investigating dataset is a temperature recording made at a 2 minutes interval for 83 days on a nocturnal mammal. The Long-tailed pocket mouse live in the South West of the United States and Northern Mexico. The data represents interest because it does contain periodicities in the behaviour of the animals in regards to circadian rhythm. The Circadian is the reaction to light and darkness of their environment. The experiments used separate equipment for each mammal to monitor the environment and so data for each animal could be varied. The data given here are the telemeter frequency temperature recordings for one Long-tailed pocket mouse from the first experiment, with higher counts indicating higher temperature. Since the interest was in periodicity, no effort was made to relate the telemeter frequency to actual temperature. The file that was collected had 59616 observations that correspond to the 83 days. The proportion of outliers in the original data has been 8% and outliers were removed.

Typically, multiple seasonality patterns are more likely to occur whenever the series has data with high frequency (for example, daily, hourly, half-hourly, and so on), as there are more options to aggregate the series to a lower frequency. This data is a typical example of multiple seasonality, which could have multiple seasonal patterns, as the observations taken every two minutes of the hours of the day, the day of the week. On the other hand, as the frequency of the series is lower (for example, monthly, quarterly, and so on), it is more likely to have only one dominant seasonal pattern as opposed to a high frequency series, as there are fewer aggregation options for another type of frequencies. Seven last days of data for investigation were selected.

Model Specification

7days of observations = 30(in one hour) * 24(hours) * 7(days) = 5040 observations

Having more observations per cycle unit, that is, high-frequency time series data, could potentially provide more insightful information about the series behavior as opposed

to a lower frequency time series data. However, this comes with the price of additional complexity, which therefore requires more effort in the analysis process.

Potentially, as mentioned previously, the series can have three different seasonal patterns.

1 week = 30*24 *7 =5040

Daily 30 * 24 = 720

Hourly = 30

2 hourly = 30*2=60

6 hourly = 30*6 = 180

10 hourly = 30 *10 = 330

Time series objects allow maximum frequency =350 and therefore we will use hourly seasonality with frequency = 30 and 10 hourly with frequency = 300. For multiple seasonalities we have to use ms object instead of ts object.

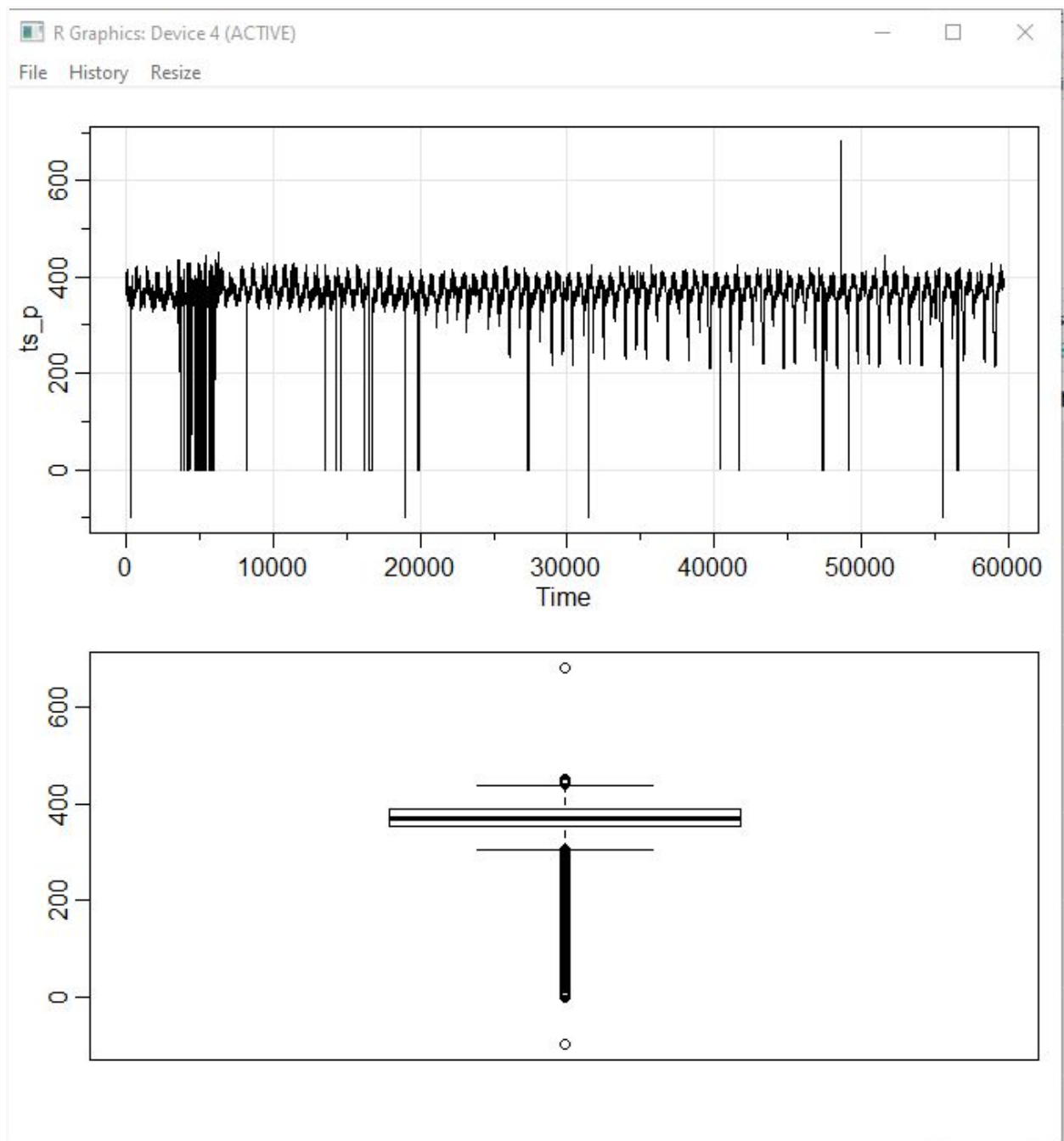
Original data

```
library(astsa)
library("forecast")
library("lubridate")
examine.mod <- function(mod.fit.obj, p, d, q, P=0, D=0, Q=0, S=-1, lag.max=24) {
  dev.new(width=6, height=6)
  par(mfrow=c(2,1))
  pacf(mod.fit.obj$fit$residuals, main="PACF of Residuals", lag.max)
  if ((P==0)&(D==0)&(Q==0)) {
    title(paste("Model: (", p, ",", d, ",", q, ")", sep=""), adj=0, cex.main=0.75)
  }
  else {
    title(paste("Model: (", p, ",", d, ",", q, ") (", P, ",", D, ",", Q, ") [", S, "]", sep=""), adj=0, cex.main=0.75)
  }

  std.resid <- mod.fit.obj$fit$residuals/sqrt(mod.fit.obj$fit$sigma2)
  hist(std.resid, main="Histogram of Standardized Residuals", xlab="Standardized Residuals",
  freq=FALSE)
  curve(expr=dnorm(x, mean=mean(std.resid), sd=sd(std.resid)), col="red", add=TRUE)
}
```

```
pformosu<-read.table(file =  
"C:/Users/inna/Desktop/DepaulClasses/ApplyMathClasses/Time_Series/Final Project/pformosu.txt")
```

```
row =3726  
col= 16  
pformosu_dat = numeric(col*row)  
count = 1  
for(i in 1:row)  
{  
  for(j in 1:col)  
  {  
    pformosu_dat[count]=pformosu[i,j]  
    count = count + 1  
  }  
}  
ts_p<-ts(pformosu_dat)  
ength(ts_p)  
dev.new()  
par(mfrow=c(2,1))  
tsplot(ts_p)  
boxplot(ts_p)$out
```



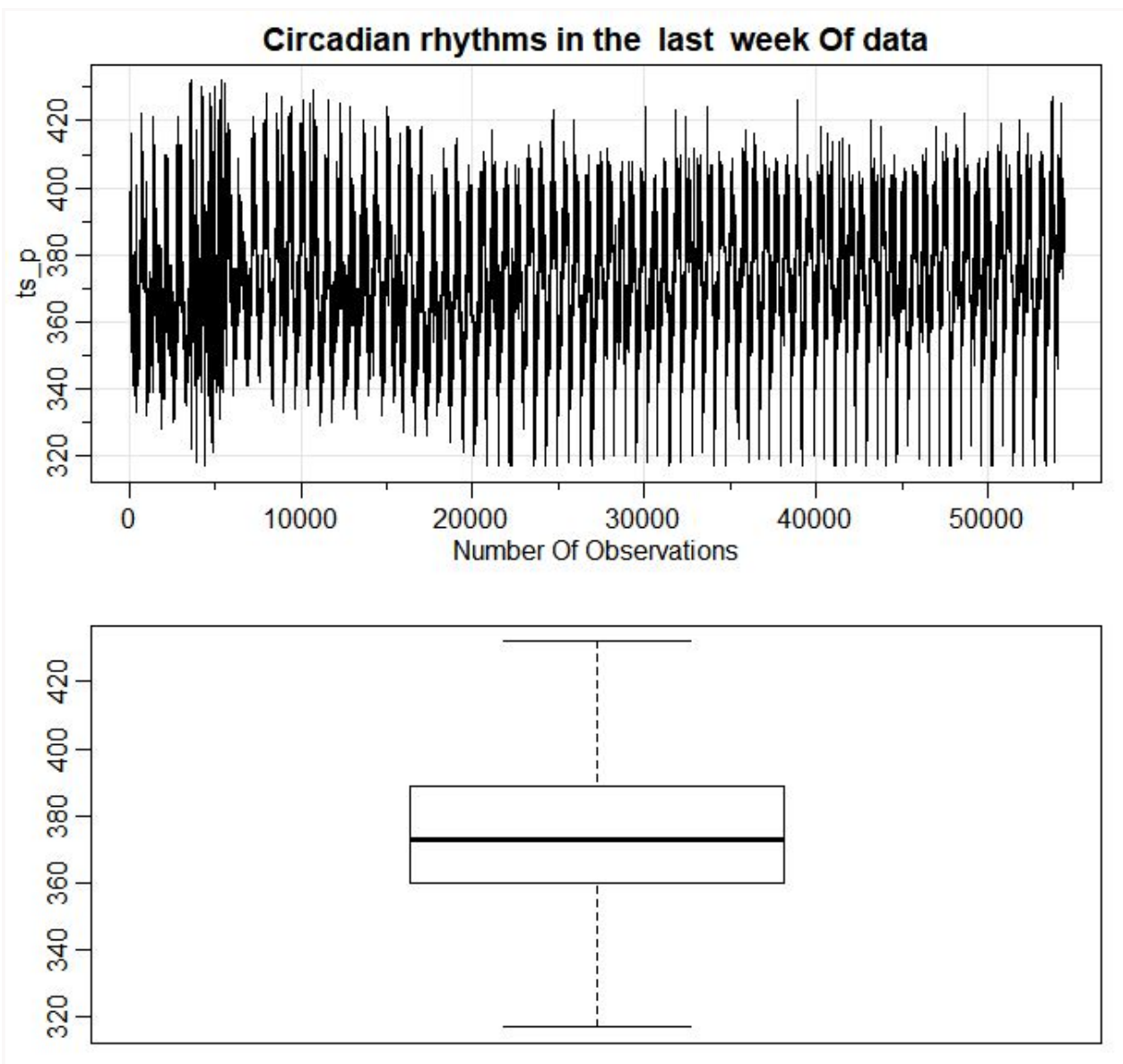
Removing outliers.

```
num_of_outliers = 0
repeat
{
  outliers <- boxplot(ts_p, plot=FALSE)$out
  v<-(which(ts_p %in% outliers))
  if(length(v)== 0)
    break
}
```

```

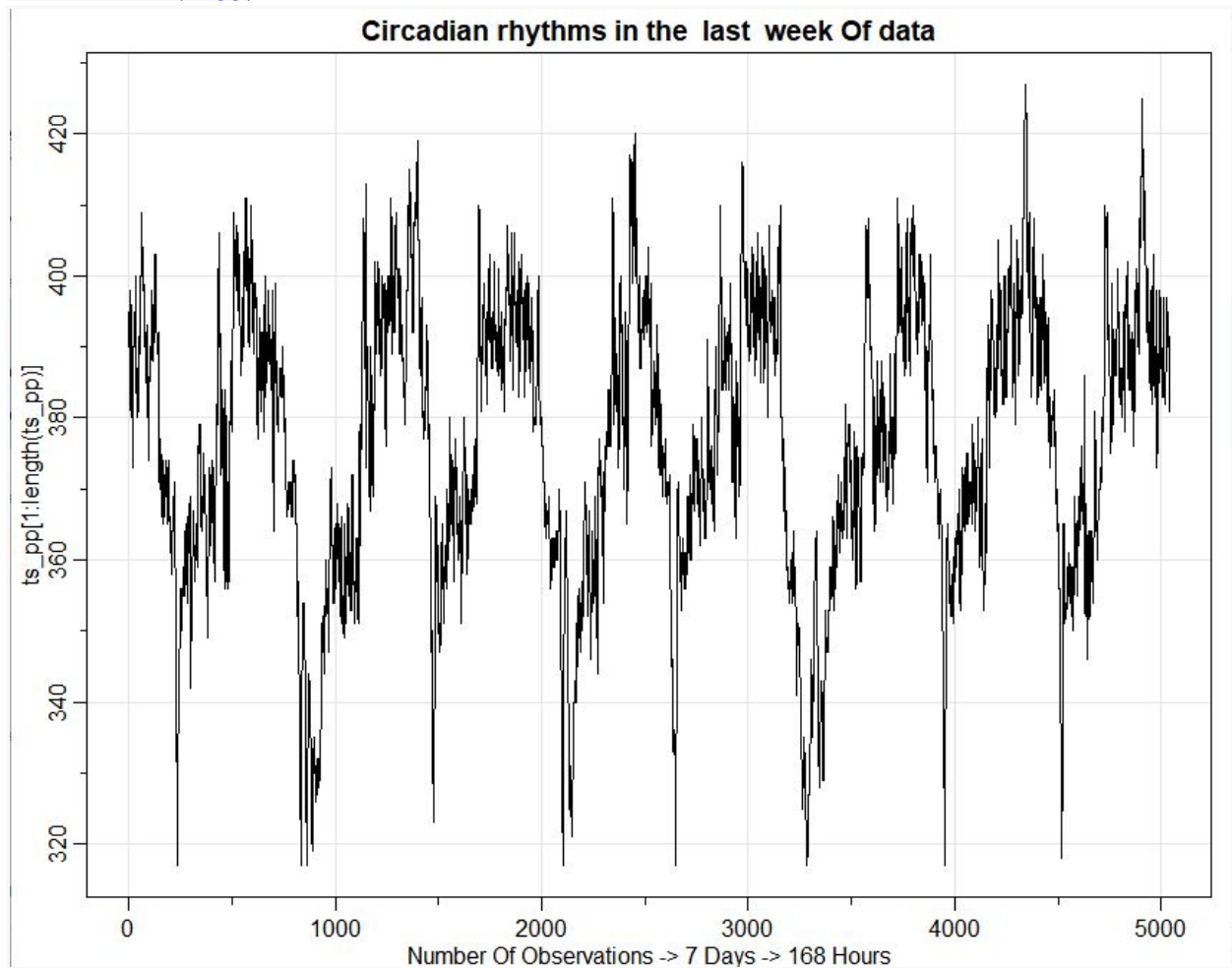
else
{
  ts_p<-ts_p[-v]
  num_of_outliers = num_of_outliers + length(v)
}
}
print(paste("Number Of Outliers removed = ", num_of_outliers))
print(paste("Length Of Time Series",length(ts_p)))
dev.new()
par(mfrow=c(2,1))
tsplot(ts_p,xlab="Number Of Observations",main="Circadian rhythms in the last week Of data")
boxplot(ts_p)$out

```

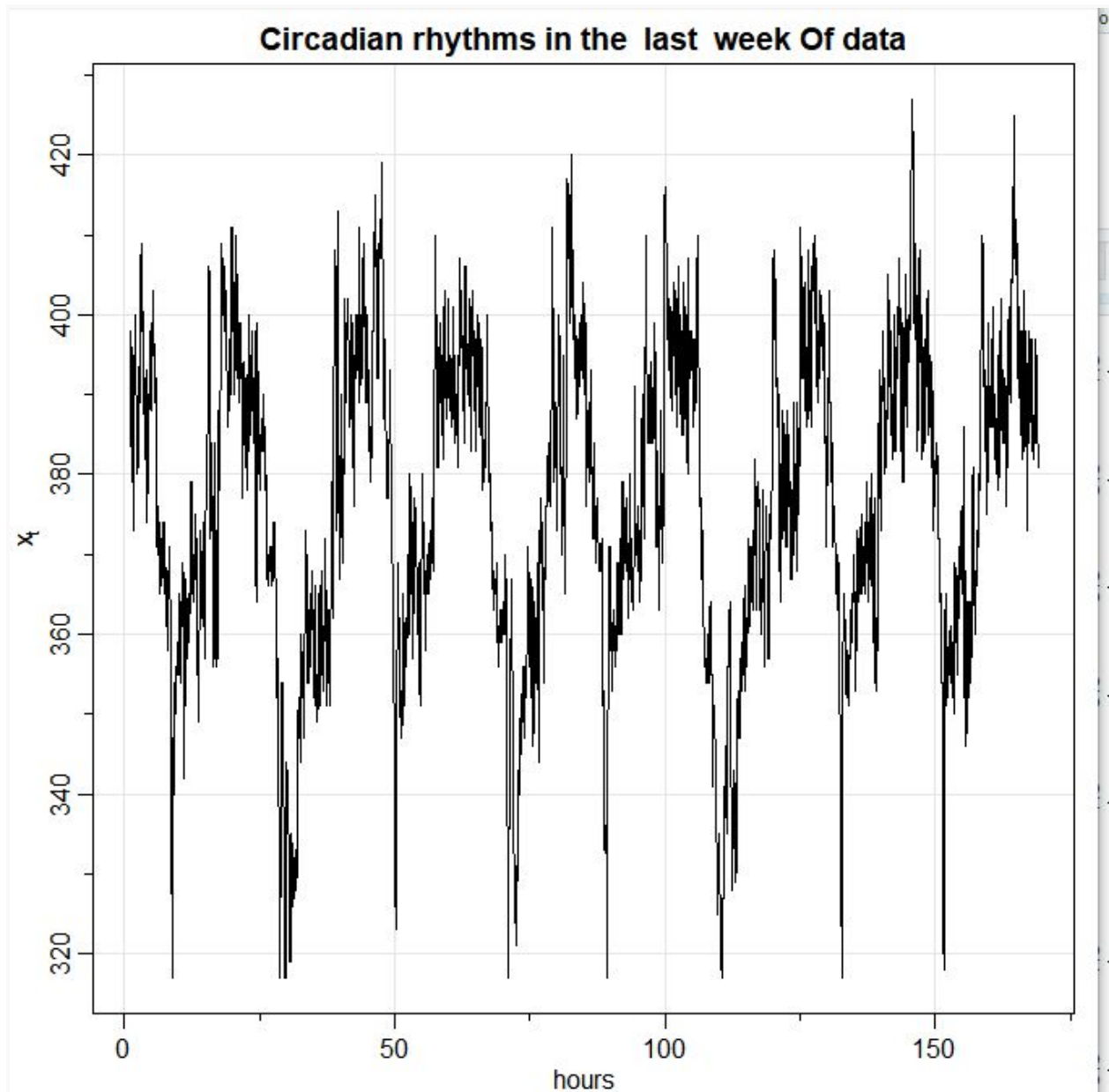


Getting one week of data and plotting it.

```
frequency <- 30  
one_day<- 30*24  
one_week<- 24*30*7  
ts_pp <- ts(ts_p[(length(ts_p)-(one_week) + 1):(length(ts_p))],frequency = frequency)  
dev.new()  
tsplot(ts_pp[1:length(ts_pp)],xlab="Number Of Observations -> 7 Days -> 168 Hours",  
      main="Circadian rhythms in the last week Of data")  
x<-as.numeric(ts_pp)
```

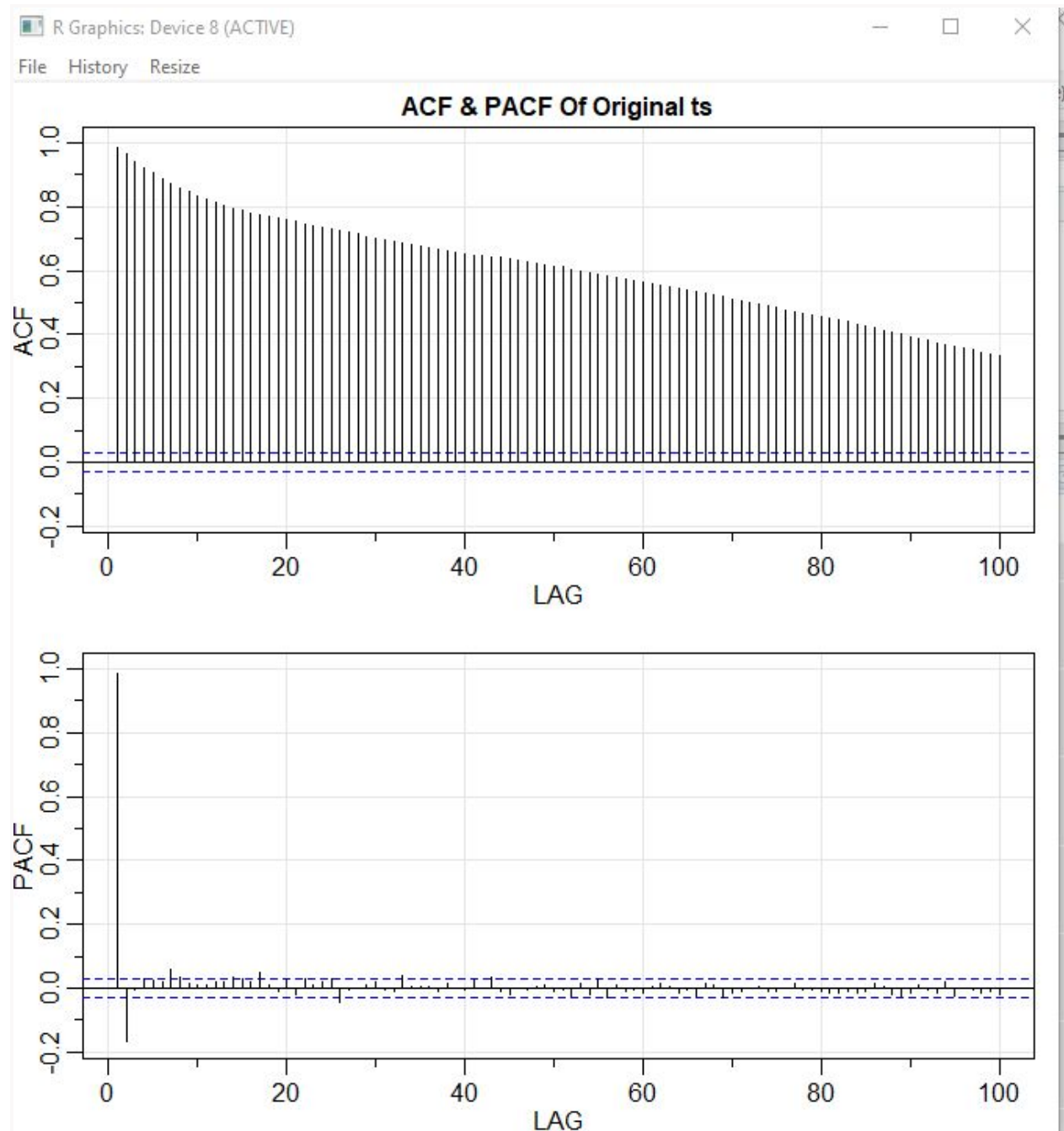


```
dev.new()  
tsplot(ts_pp, ylab=expression(x[t]),  
       xlab="hours", main="Circadian rhythms in the last week Of data")
```



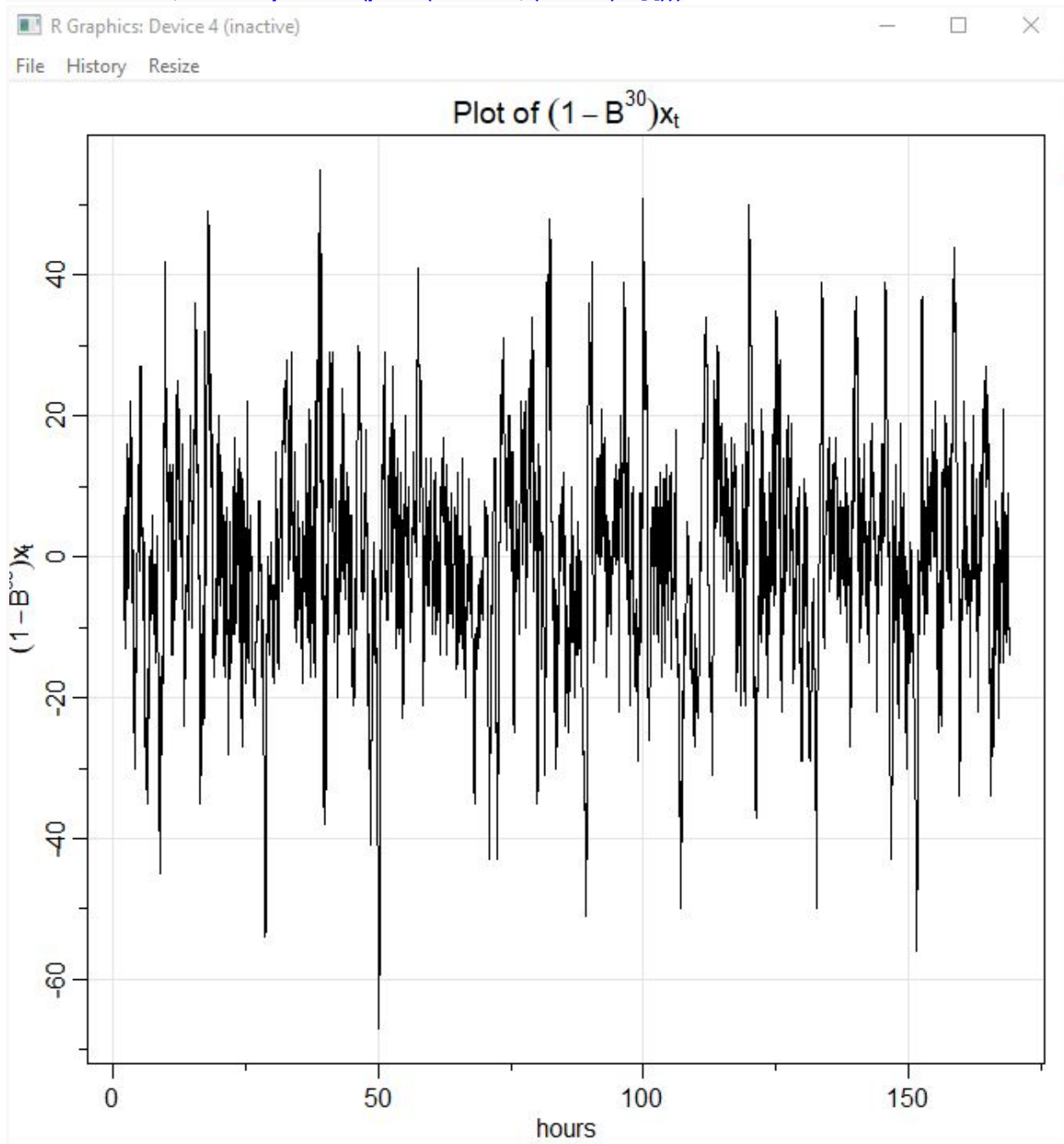

```
dev.new()
```

```
acf2(x, max.lag = length(ts_pp) - 1, main = "ACF & PACF Of Original ts")
```



We can see from the plot that ACF is decaying exponentially and PACF has 2 lags but we have to differentiate to get rid of seasonalities ($1 - B^{30}$)

```
# Plot of  $(1-B^{30})x_t$ 
dev.new()
tsplot(diff(ts_pp, lag=frequency, differences=1), ylab=expression((1-B^30)*x[t]),
       xlab="hours", main=expression(paste("Plot of ", (1-B^30)*x[t])))
```

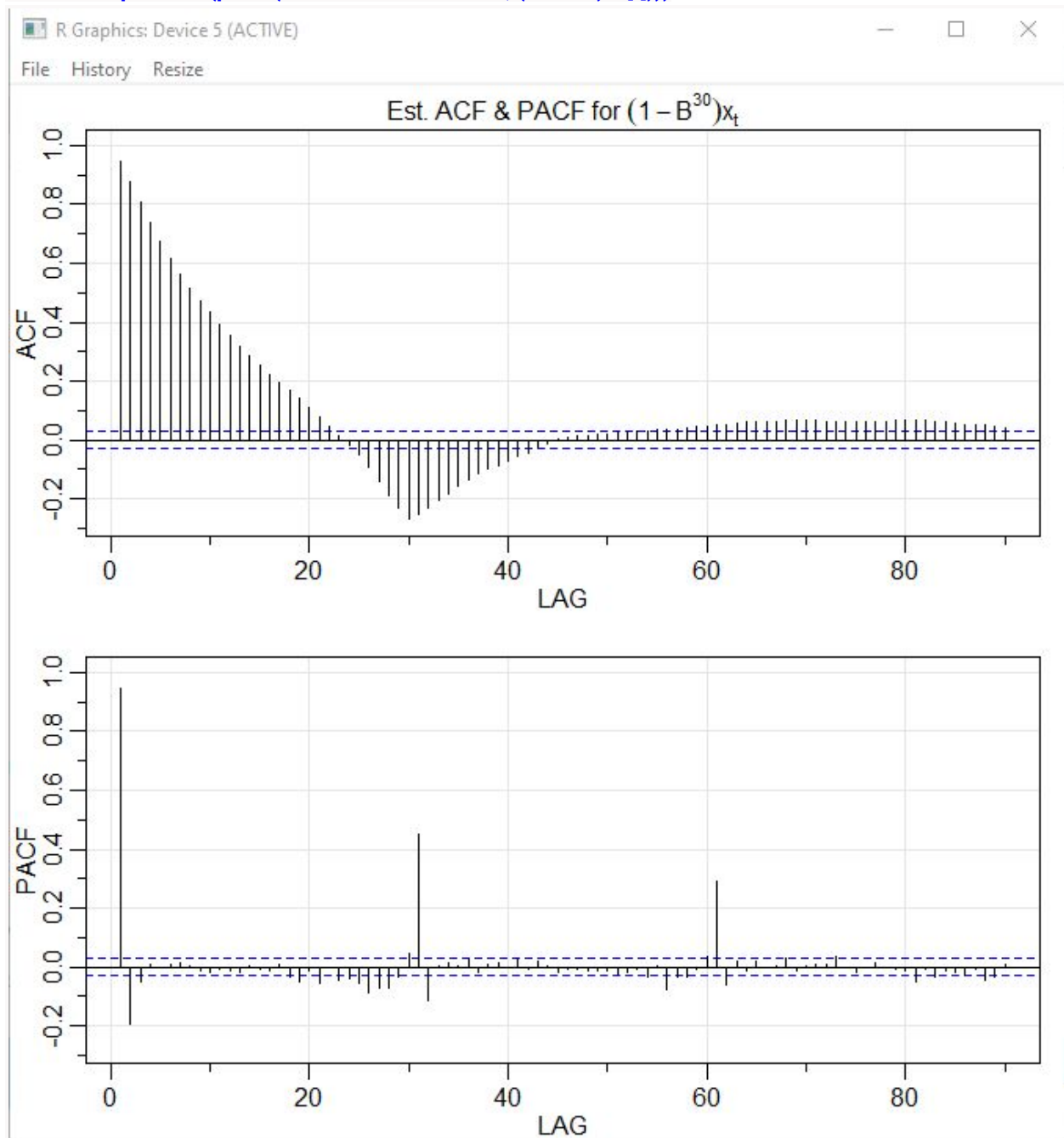


```
# ACF indicated cut of after 30 lag that suggests that ARIMA Q=1,
# PACF exponentially decay that suggests ARIMA P=0
# Diff =1 S=18  $(1-B^{30})x_t$ 
```

```
dev.new()
```

```
acf2(diff(x, lag=frequency, differences=1), max.lag=frequency*3,
```

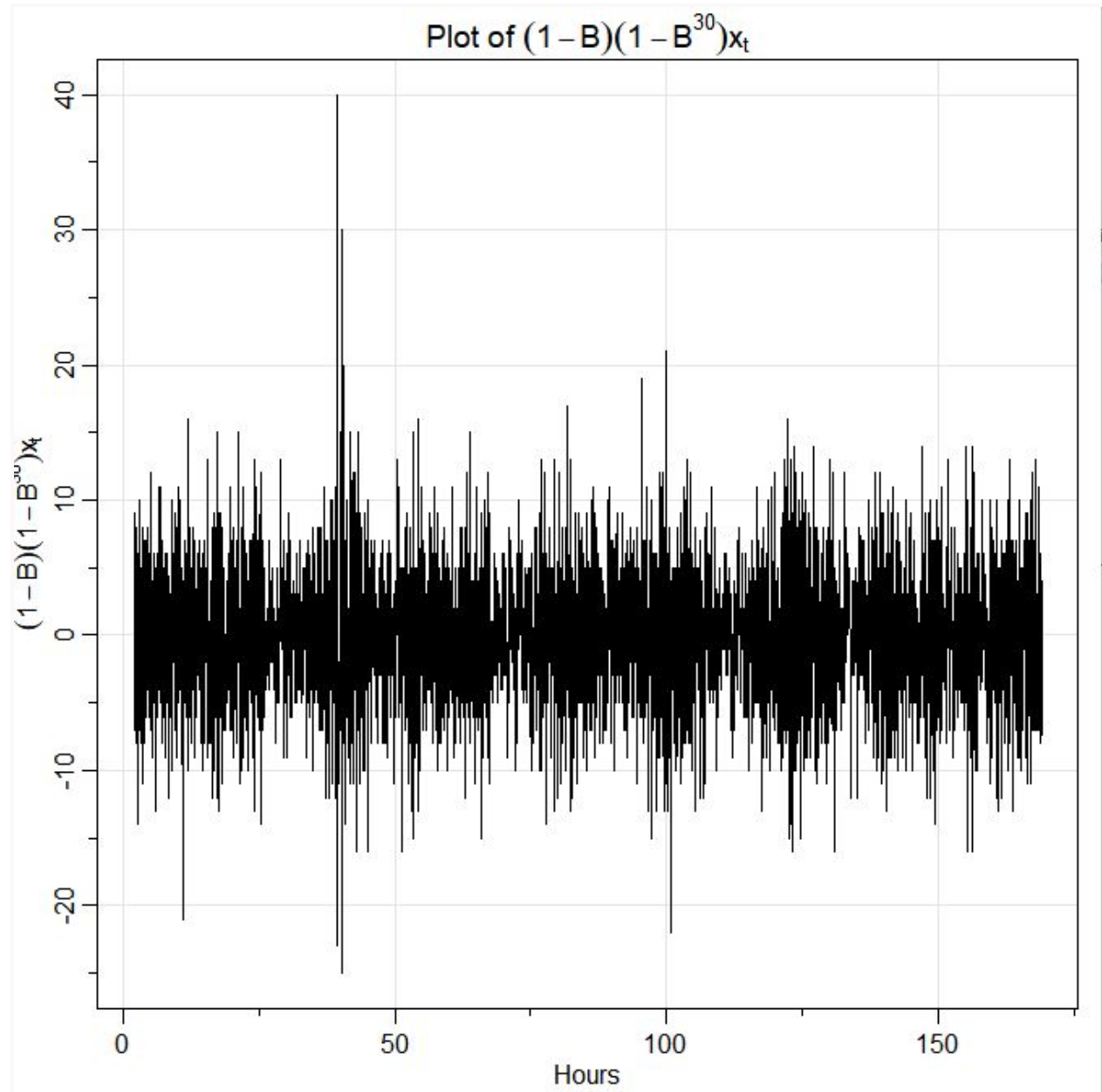
```
main=expression(paste("Est. ACF & PACF for ", (1-B^30)*x[t])))
```

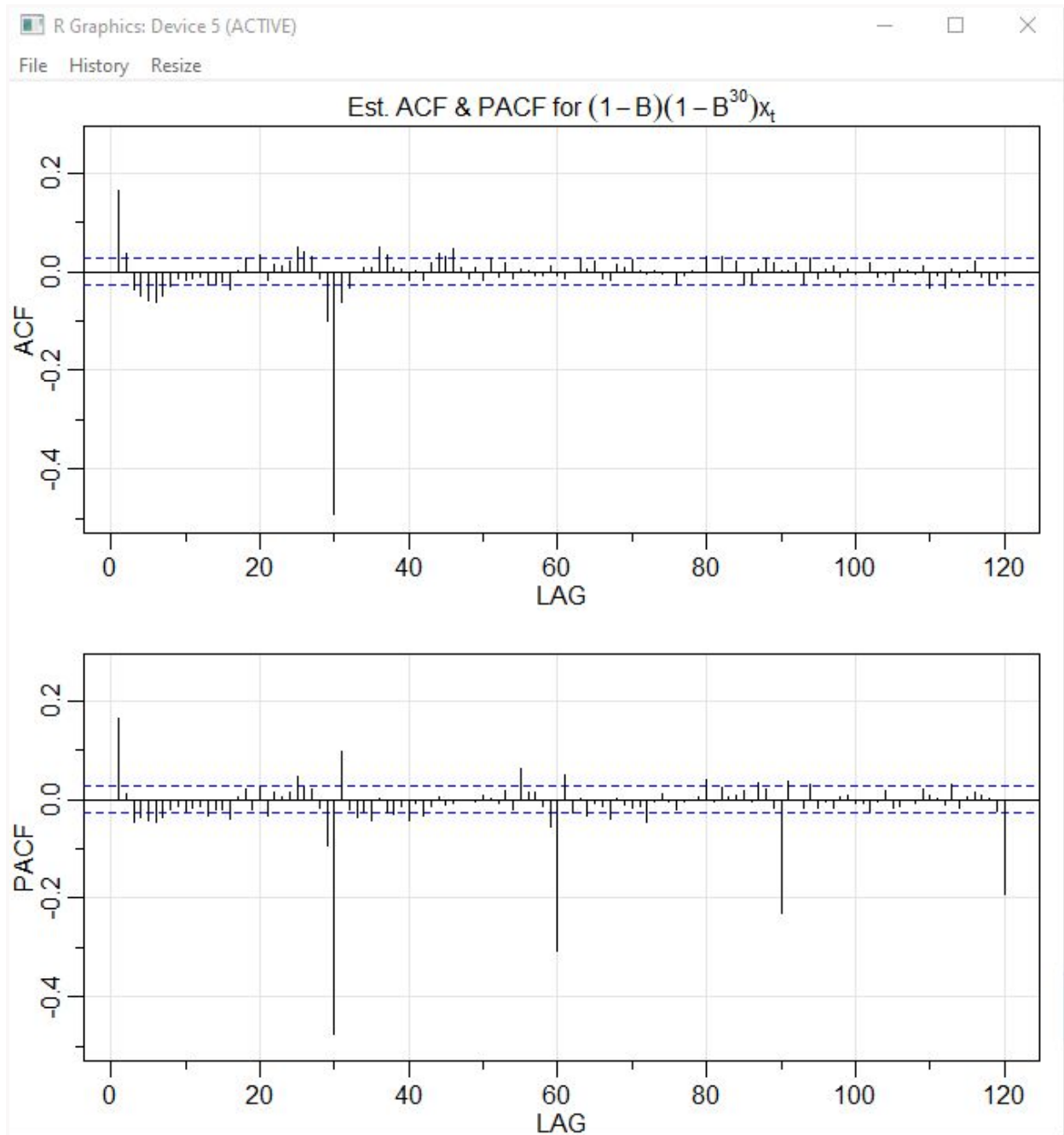


PACF exponentially decaying at lags 30,60 indicating that $P = 0$

ACF seasonal pick at Lag = 30 and no other seasonal picks at lags 60 , 90
Indication that Q = 1, giving us the initial estimation for ARIMA(0,1,1)30.
But graph has some little trend so we have to remove the trend using non seasonal difference (1-B)

```
ev.new()  
tsplot(diff(diff(ts_pp, lag=frequency, differences=1)),  
       ylab=expression((1-B)(1-B^30)*x[t]),  
       xlab="Hours", main=expression(paste("Plot of ",  
       (1-B)(1-B^30)*x[t])))
```



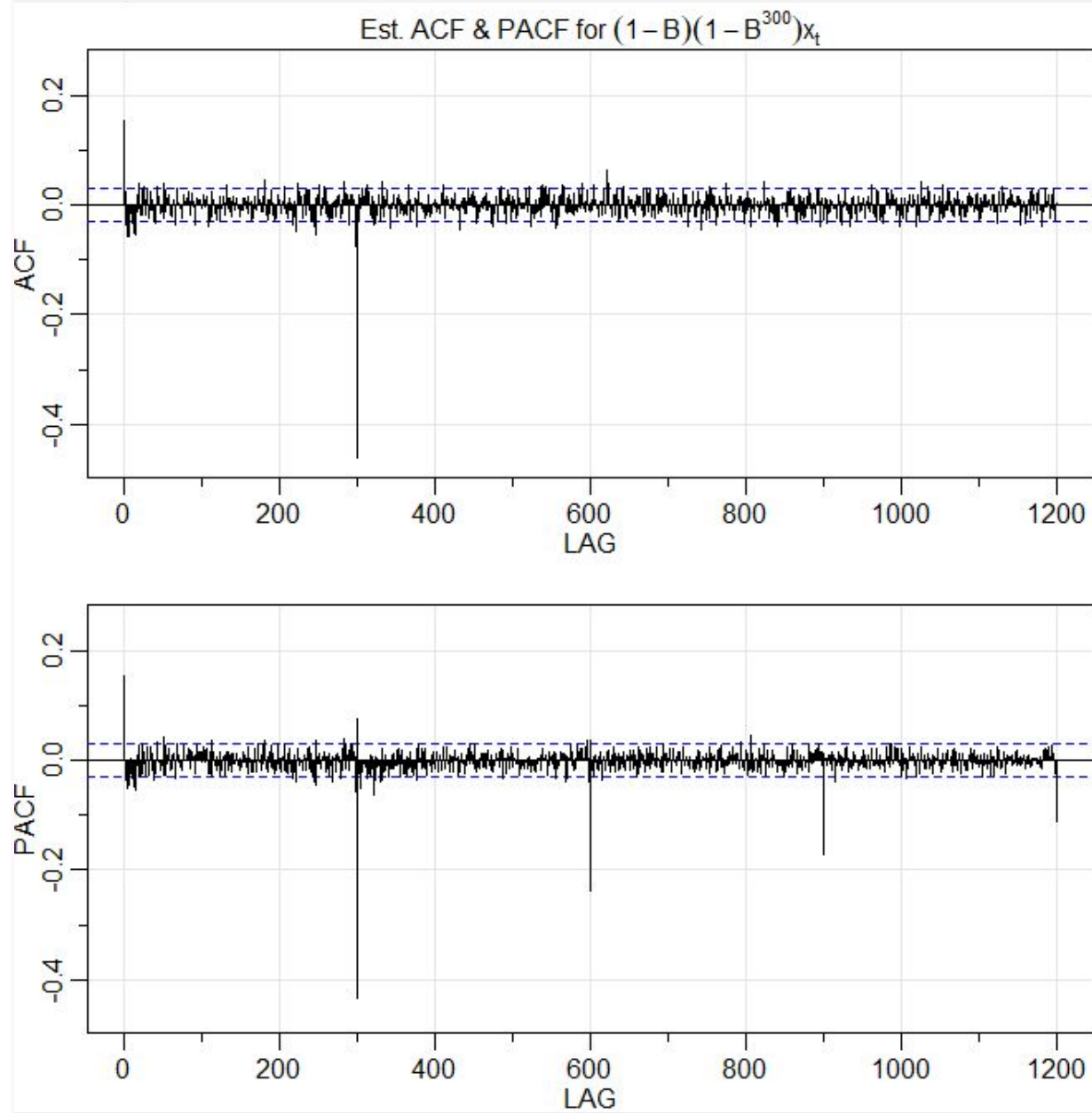


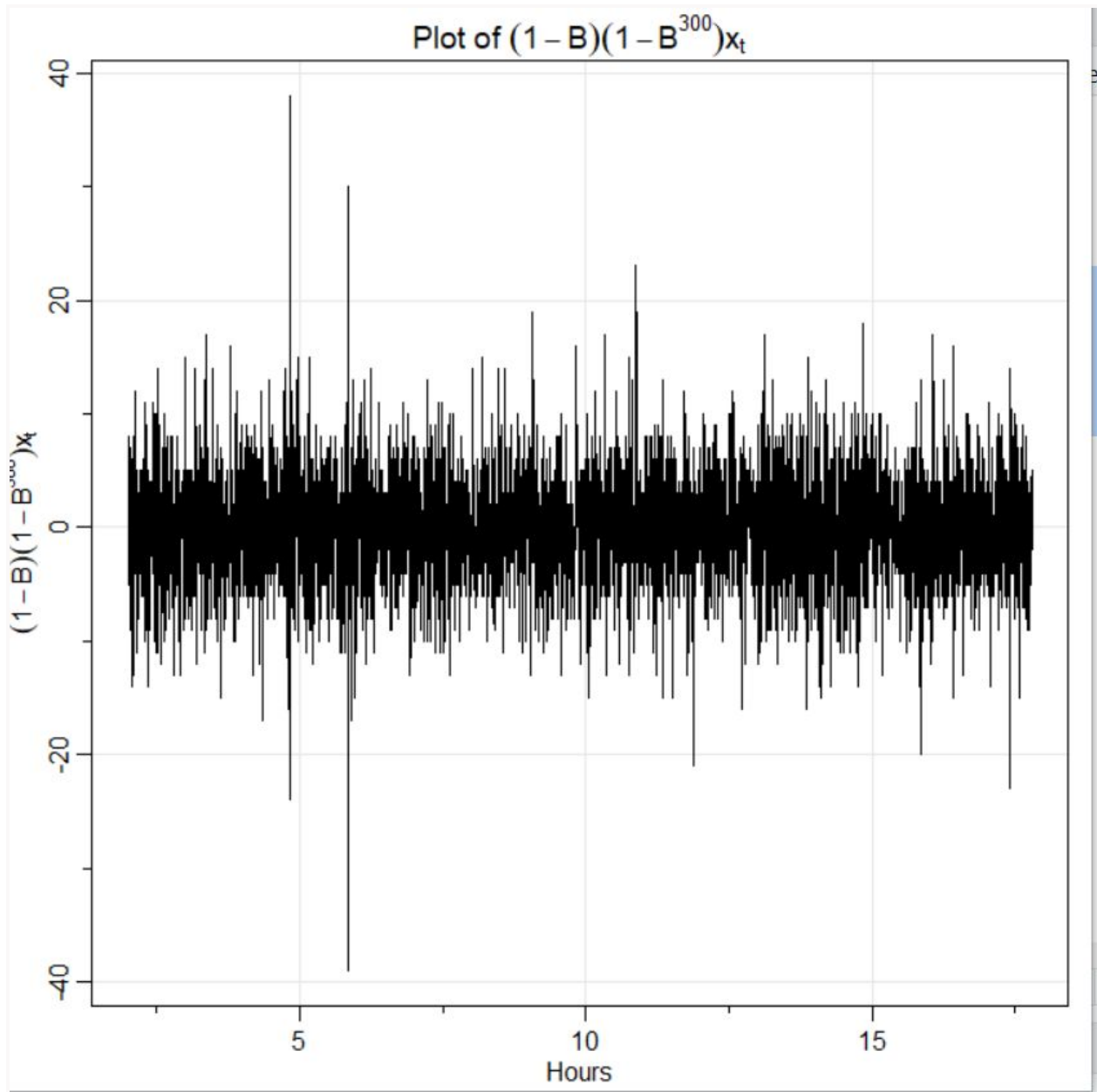
ACF seasonal $Q = 1$, non-seasonal $q = 1-5$

PACF decaying at lags 30,60,90 , and clusters around each lag indication that seasonal $P = 0$,
nonseasonal $p = 1-5$

ARIMA(p=(0-5),d=1,q=(0-5)),(P=0,D=1,Q=1)[30]

Trying ACF frequency = 300





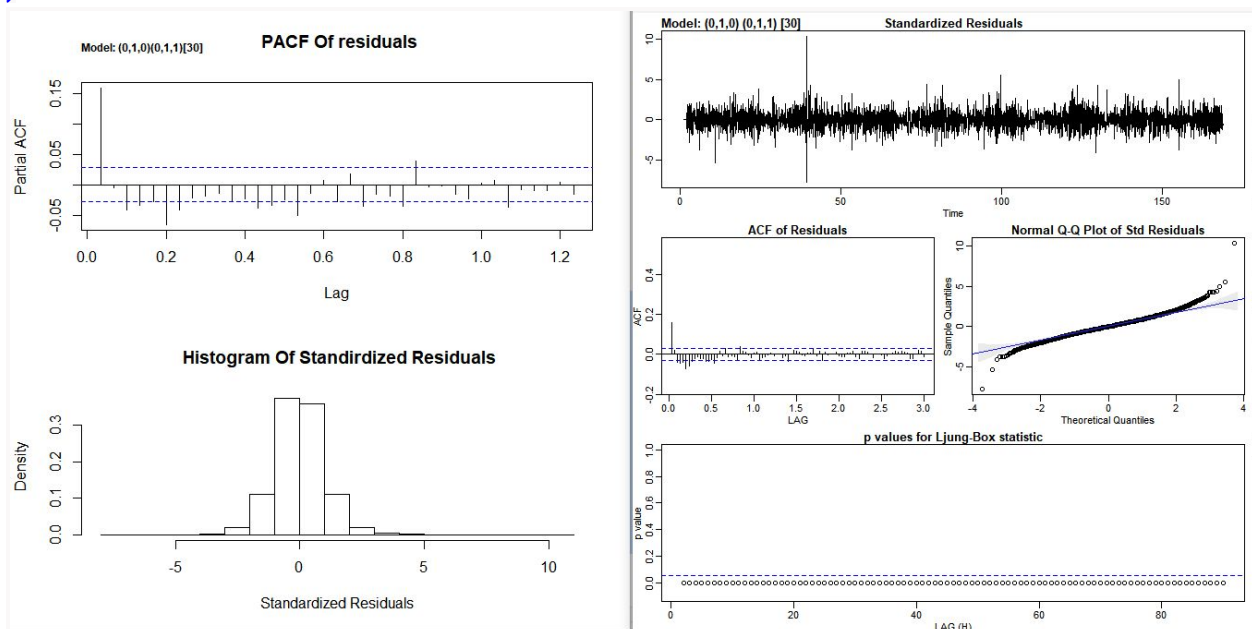
ACF shows seasonal spike at lag 300 and non-seasonal at some cluster around lag 1
 PACF shows seasonal exponential decay at 300,600 .. and also cluster around lag one
 Also indicating **ARIMA(p=(0-5),d=1,q=(0-5)),(P=0,D=1,Q=1)[300]**
ARIMA(p=(0-5),d=1,q=(0-5),P=0,D=1,Q=1,S=30)

Fitting and Diagnostics

Trying ARIMA($p=(0)$, $d=1$, $q=(0)$, $P=0$, $D=1$, $Q=1$, $S=30$)

```
dev.new()
mod.fit1<- sarima(ts_pp,p=0,d=1,q=0,P=0,D=1,Q=1,S=30)
mod.fit1

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$residuals,main = " PACF Of residuals")
title(paste("Model: (",
  p = 0, ",",
  d = 1, ",",
  q = 0, ")(",
  P = 0, ",",
  D = 1, ",",
  Q = 1, ")[30]",
  sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$residuals / sqrt(mod.fit1$sigma2)
hist(std.resid1,main = " Histogram Of Standardized Residuals",
  xlab='Standardized Residuals',
  freq = FALSE
)
```




```

$`fit`

call:
stats::arima(x = xdata, order = c(p, d, q), seasonal = list(order = c(P, D,
  Q), period = S), include.mean = !no.constant, transform.pars = trans, fixed = fixed,
  optim.control = list(trace = trc, REPORT = 1, reltol = tol))

Coefficients:
          sma1
        -1.0000
s.e.      0.0152

sigma^2 estimated as 12.15:  log likelihood = -13438.07,  aic = 26880.13

$degrees_of_freedom
[1] 5008

$table
      Estimate      SE  t.value p.value
sma1        -1 0.0152 -65.6435      0

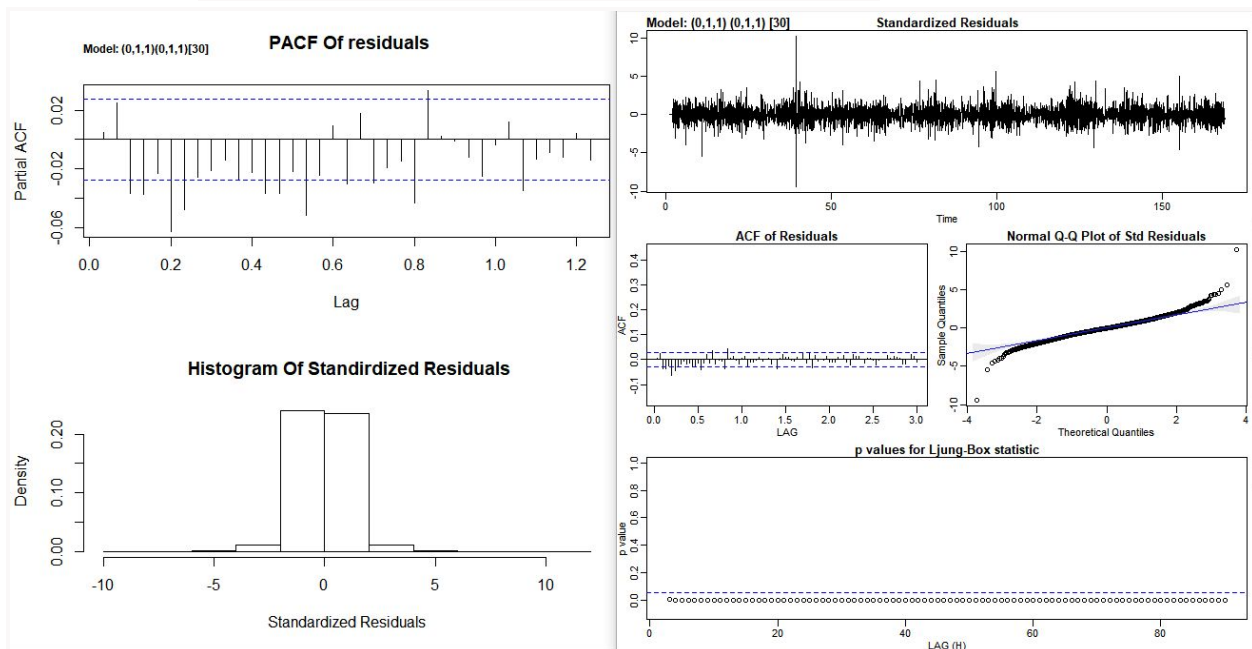
$AIC
[1] 5.335477

$AICC
[1] 5.335477

$BIC
[1] 5.338065

```

Trying ARIMA(p=0,d=1,q=1,P=0,D=1,Q=1,S=30)



```
> mod.fit1
$fit

Call:
stats::arima(x = xdata, order = c(p, d, q), seasonal = list(order = c(P, D,
  Q), period = S), include.mean = !no.constant, transform.pars = trans, fixed = fixed,
  optim.control = list(trace = trc, REPORT = 1, reltol = tol))

Coefficients:
      ma1      sma1
    0.1538   -1.0000
s.e.  0.0135    0.0126

sigma^2 estimated as 11.85:  log likelihood = -13376.12,  aic = 26758.24

$degrees_of_freedom
[1] 5007

$tttable
      Estimate      SE  t.value p.value
ma1    0.1538  0.0135  11.3692      0
sma1   -1.0000  0.0126 -79.6657      0

$AIC
[1] 5.311282

$AICC
[1] 5.311282

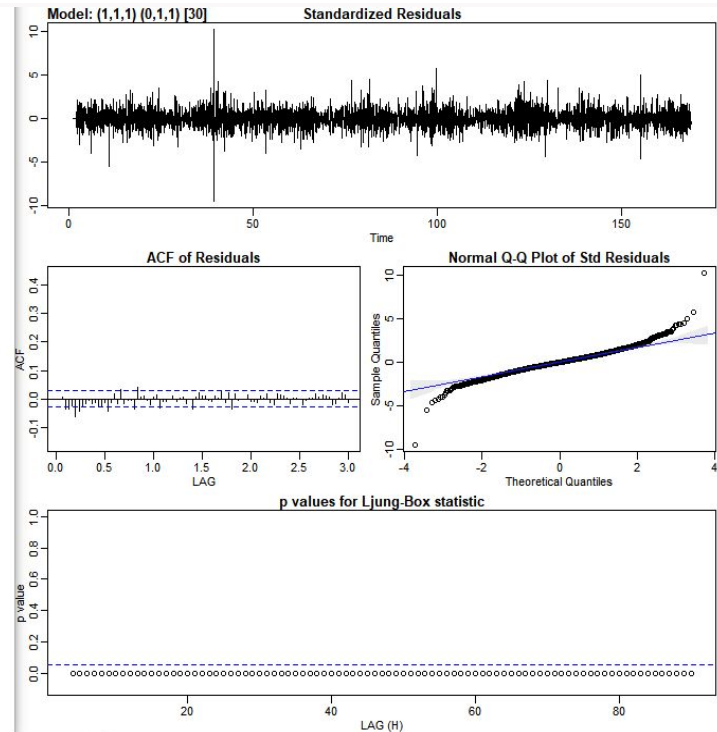
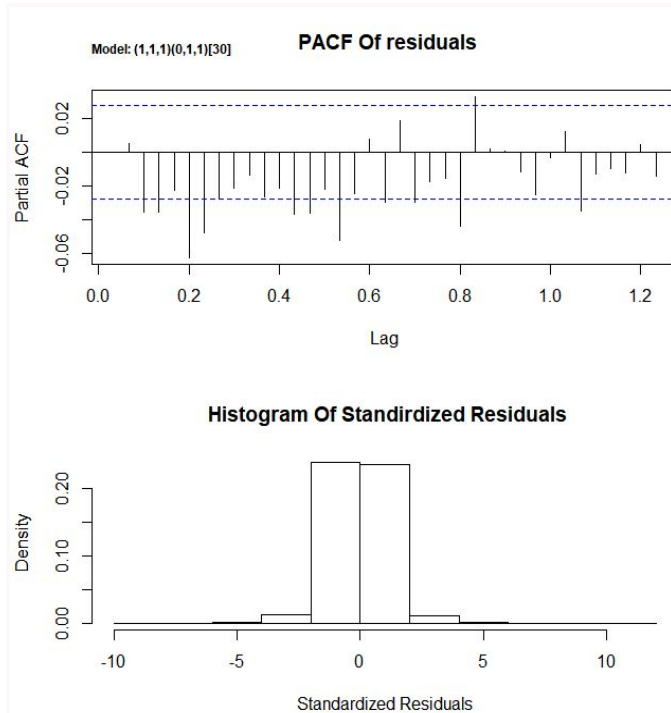
$BIC
[1] 5.315164
```

ARIMA(p=0,d=1,q=1,P=0,D=1,Q=1,S=30)

model seems like better than the previous because **BIC = 5.315164** compare to

ARIMA(p=0,d=1,q=0,P=0,D=1,Q=1,S=30) **BIC = 5.338065**

Trying ARIMA(p=1,d=1,q=1,P=0,D=1,Q=1,S=30)



```
> mod.fit1
$`fit`

Call:
stats::arima(x = xdata, order = c(p, d, q), seasonal = list(order = c(P, D,
  Q), period = S), include.mean = !no.constant, transform.pars = trans, fixed = fixed,
  optim.control = list(trace = trc, REPORT = 1, reltol = tol))

Coefficients:
      ar1      ma1      sma1
    0.1370  0.0219 -1.0000
s.e.  0.0727  0.0730  0.0129

sigma^2 estimated as 11.84:  log likelihood = -13374.49,  aic = 26756.98

$degrees_of_freedom
[1] 5006

$table
      Estimate      SE  t.value p.value
ar1    0.1370  0.0727   1.8833  0.0597
ma1    0.0219  0.0730   0.2995  0.7646
sma1   -1.0000  0.0129  -77.7387  0.0000

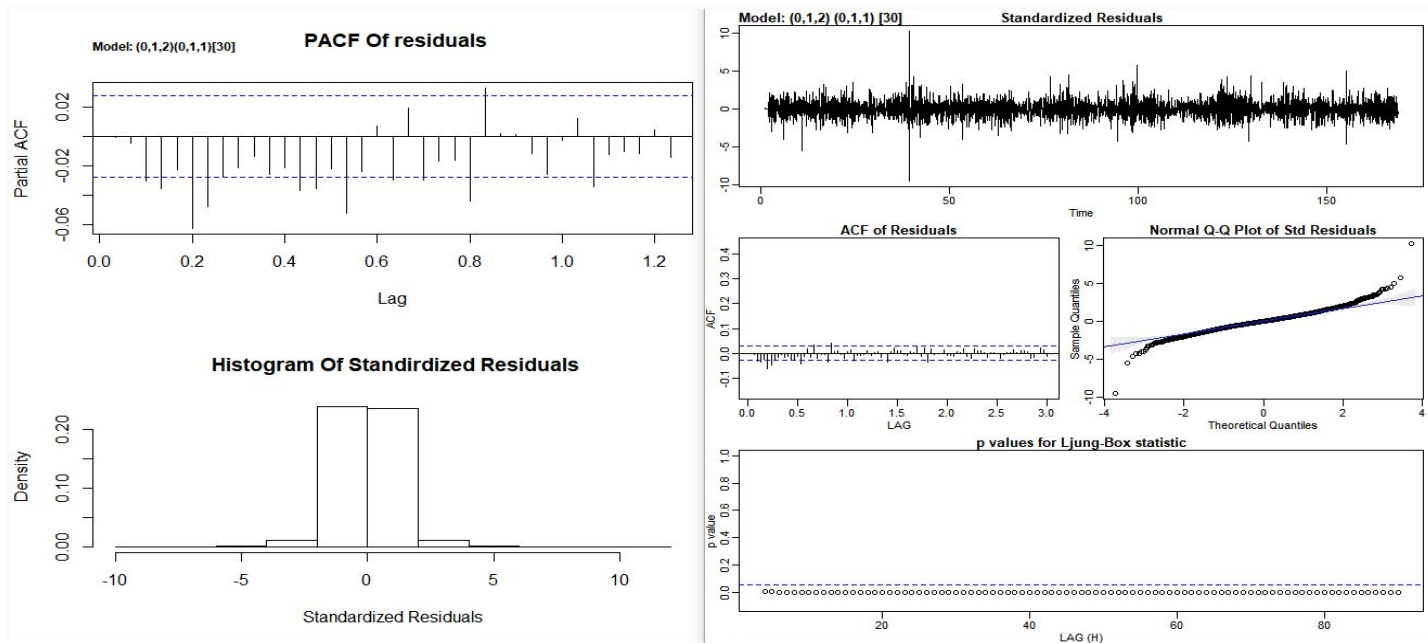
$AIC
[1] 5.311032

$AICC
[1] 5.311033

$BIC
[1] 5.316208
```

The ARIMA(p=1,d=1,q=1,P=0,D=1,Q=1,S=30) model has p-value for ma1 > 0.05 and this makes **ARIMA(p=0,d=1,q=1,P=0,D=1,Q=1,S=30)** is better out of all previous models

Trying ARIMA(p=0,d=1,q=2,P=0,D=1,Q=1,S=30)



Coefficients:

	ma1	ma2	sma1
	0.1599	0.0320	-1.000
s.e.	0.0142	0.0147	0.013

sigma^2 estimated as 11.84: log likelihood = -13373.75, aic = 26755.5

\$degrees_of_freedom

[1] 5006

\$ttable

	Estimate	SE	t.value	p.value
ma1	0.1599	0.0142	11.2785	0.0000
ma2	0.0320	0.0147	2.1813	0.0292
sma1	-1.0000	0.0130	-76.7035	0.0000

\$AIC

[1] 5.310738

\$AICc

[1] 5.310739

\$BIC

[1] 5.315914

The ARIMA($p=0, d=1, q=2, P=0, D=1, Q=1, S=30$) model has aal p-value for ma1 < 0.05 and this makes
And BIC=5.315914

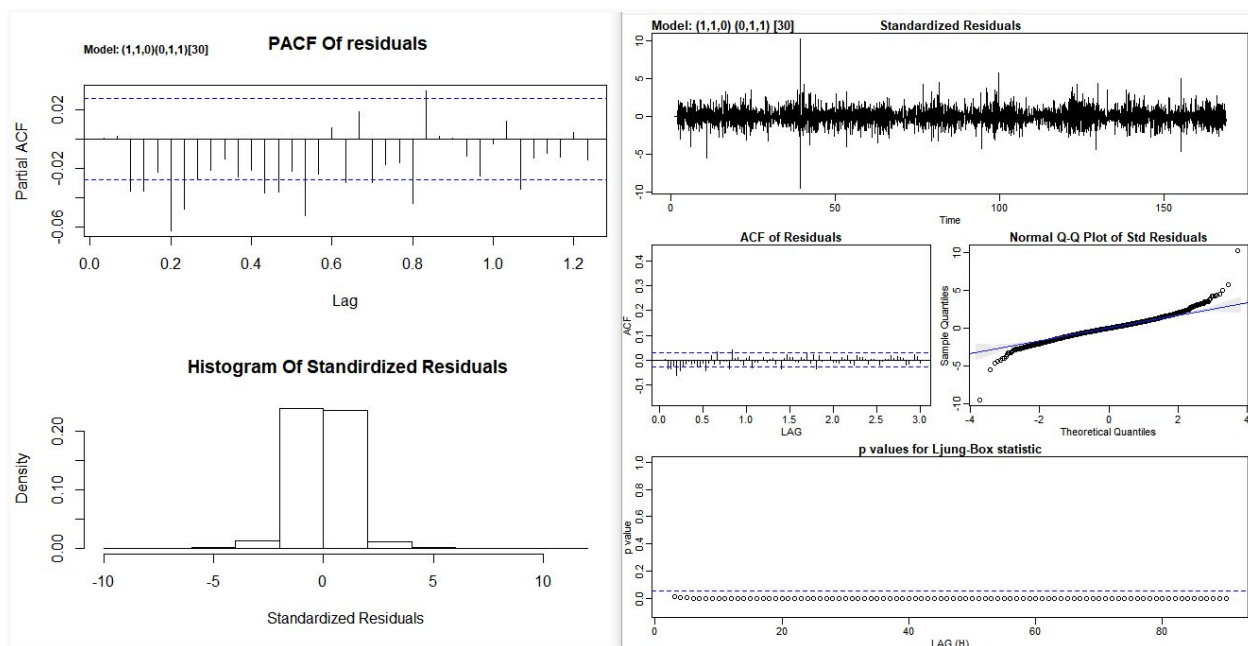
ARIMA($p=0, d=1, q=1, P=0, D=1, Q=1, S=30$) BIC = 5.315164 < 5.315914 = BIC

ARIMA($p=0, d=1, q=2, P=0, D=1, Q=1, S=30$) which makes the best out of all previous models

ARIMA($p=0, d=1, q=1, P=0, D=1, Q=1, S=30$)

When Trying ARIMA($p=0, d=1, q=(3 \text{ or } 4), P=0, D=1, Q=1, S=30$) values of BIC is increasing and therefore
the best model now is still **ARIMA($p=0, d=1, q=1, P=0, D=1, Q=1, S=30$)**

Trying ARIMA(p=1,d=1,q=0,P=0,D=1,Q=1,S=30)



Coefficients:

	ar1	sma1
	0.1583	-1.0000
s.e.	0.0139	0.0129

sigma^2 estimated as 11.84: log likelihood = -13374.53, aic = 26755.07

\$degrees_of_freedom
[1] 5007

	Estimate	SE	t.value	p.value
ar1	0.1583	0.0139	11.3460	0
sma1	-1.0000	0.0129	-77.3007	0

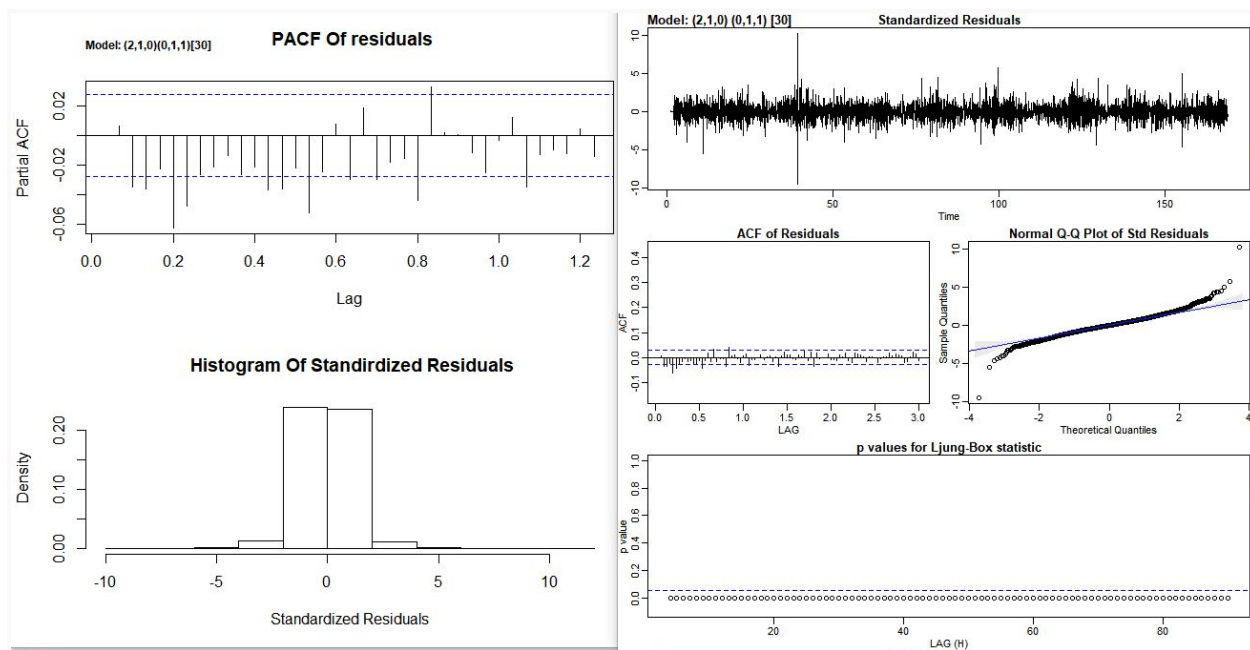
\$AIC
[1] 5.310653

\$AICC
[1] 5.310653

\$BIC
[1] 5.314535

The ARIMA(p=1,d=1,q=0,P=0,D=1,Q=1,S=30) model has aal p-value for ma1 < 0.05 and
And BIC=5.314535 < BIC = 5.315164 of ARIMA(p=0,d=1,q=1,P=0,D=1,Q=1,S=30)
This makes **ARIMA(p=1,d=1,q=0,P=0,D=1,Q=1,S=30)** better model

Trying ARIMA(p=2,d=1,q=0,P=0,D=1,Q=1,S=30)



```
$ttable
      Estimate      SE  t.value p.value
ar1    0.1591 0.0141  11.2615  0.0000
ar2   -0.0052 0.0141   -0.3683  0.7127
sma1   -1.0000 0.0128 -77.9575  0.0000

$AIC
[1] 5.311023

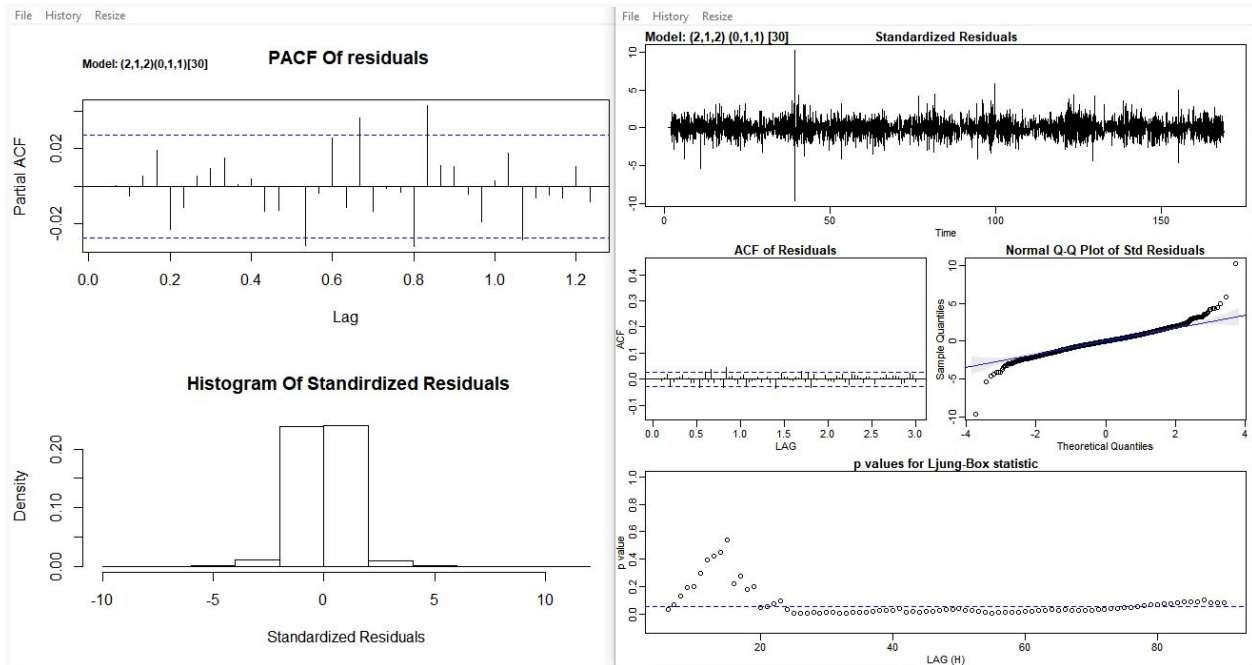
$AICC
[1] 5.311024

$BIC
[1] 5.316199

>
> dev.new()
```

Ar2 pvalue >0.05 which makes the **ARIMA(p=1,d=1,q=0,P=0,D=1,Q=1,S=30)** better model

Trying ARIMA(p=2,d=1,q=2,P=0,D=1,Q=1,S=30)



```

Coefficients:
      ar1      ar2      ma1      ma2      sma1
s.e.  0.0682  0.0654  0.0722  0.0707  0.0102

sigma^2 estimated as 11.64:  log likelihood = -13332.36,  aic = 26676.72

$degrees_of_freedom
[1] 5004

$table
      Estimate      SE  t.value p.value
ar1    1.3022  0.0682  19.0957  0.0000
ar2   -0.3833  0.0654  -5.8623  0.0000
ma1   -1.1594  0.0722 -16.0541  0.0000
ma2    0.2056  0.0707   2.9097  0.0036
sma1   -1.0000  0.0102 -97.6950  0.0000

$AIC
[1] 5.295102

$AICC
[1] 5.295104

$BIC
[1] 5.302866
  
```

p values <0.05 which makes the ARIMA(p=2,d=1,q=2,P=0,D=1,Q=1,S=30) BIC=5.302866 < BIC=5.314535 ARIMA(p=1,d=1,q=0,P=0,D=1,Q=1,S=30) which makes **ARIMA(p=2,d=1,q=2,P=0,D=1,Q=1,S=30) better than all above models**

Trying ARIMA(p=3,d=1,q=3,P=0,D=1,Q=1,S=30)

Trying ARIMA(p=3,d=1,q=2,P=0,D=1,Q=1,S=30)

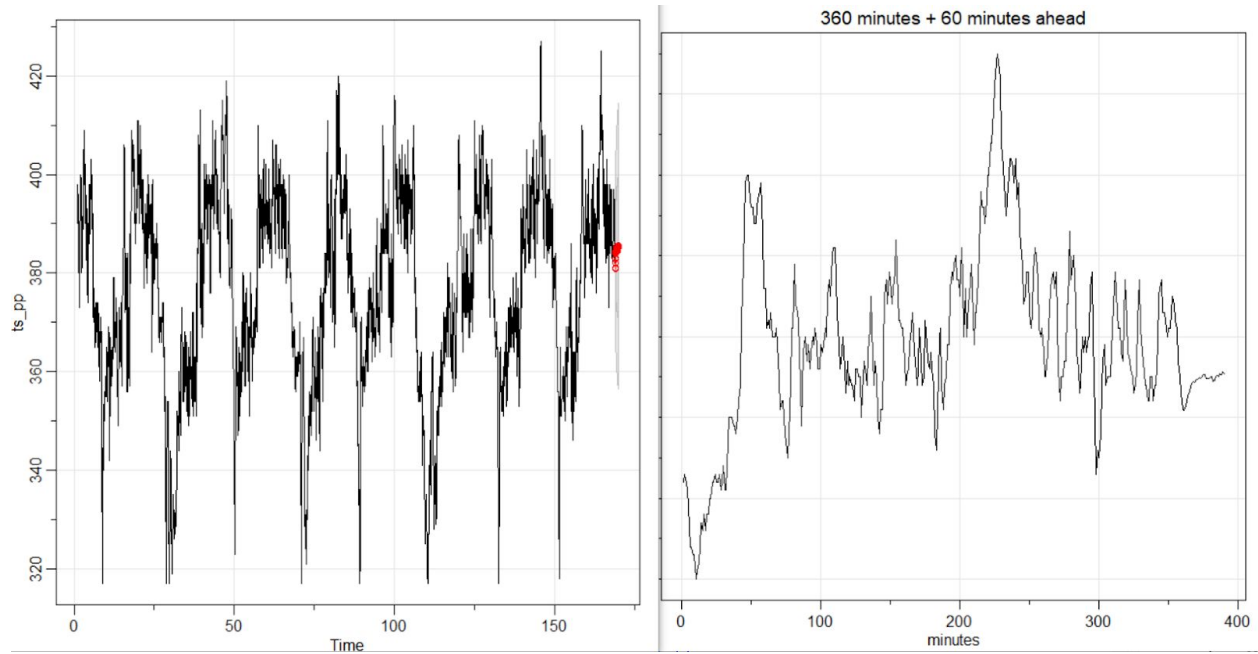
ARIMA(p=2,d=1,q=3,P=0,D=1,Q=1,S=30)

P-values >0.05 and therefore the best selected model is

ARIMA(p=2,d=1,q=2,P=0,D=1,Q=1,S=30)

Forecasting

```
ahead_1_hours <-30  
dev.new()  
fore.mod <- sarima.for(ts_pp, n.ahead=ahead_1_hours,  
  p=2, d=1, q=2, P=0, D=1, Q=1, S=30, plot.all=TRUE)
```

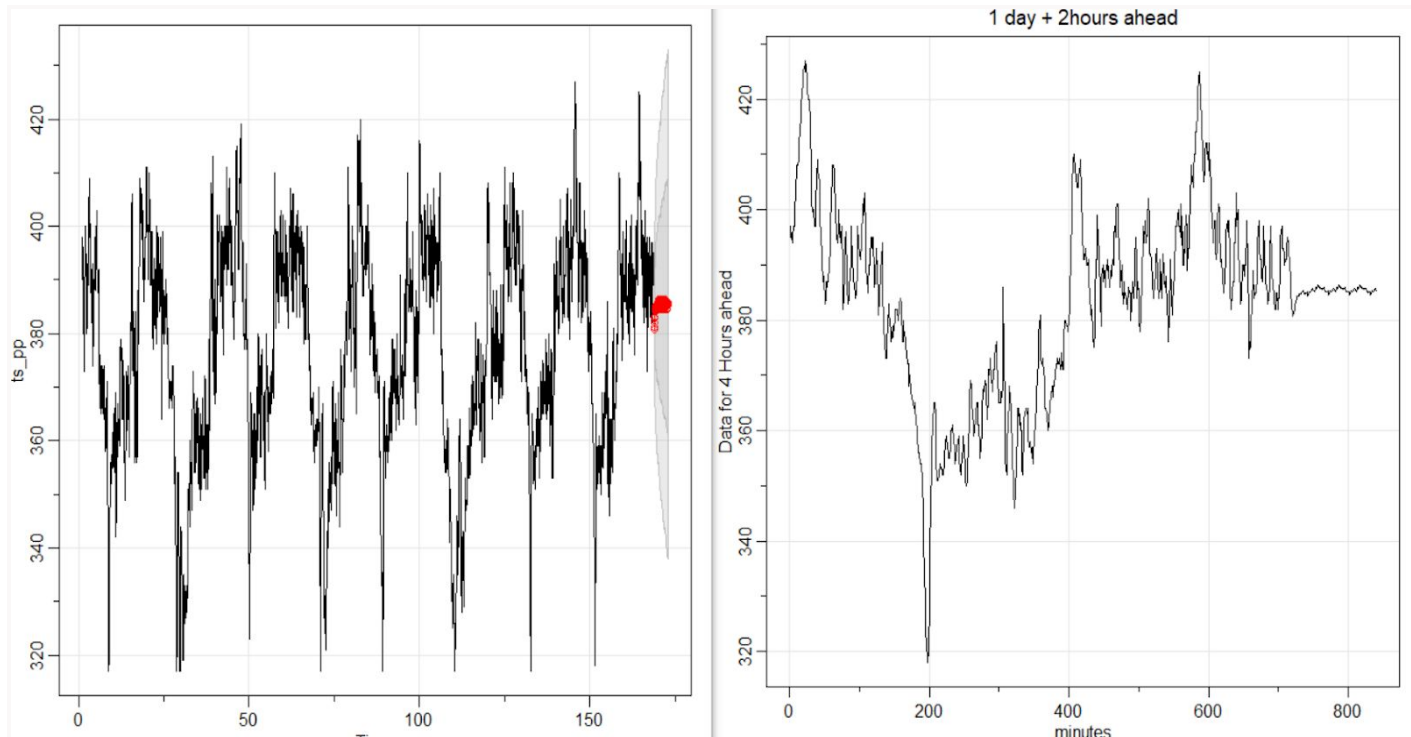


```

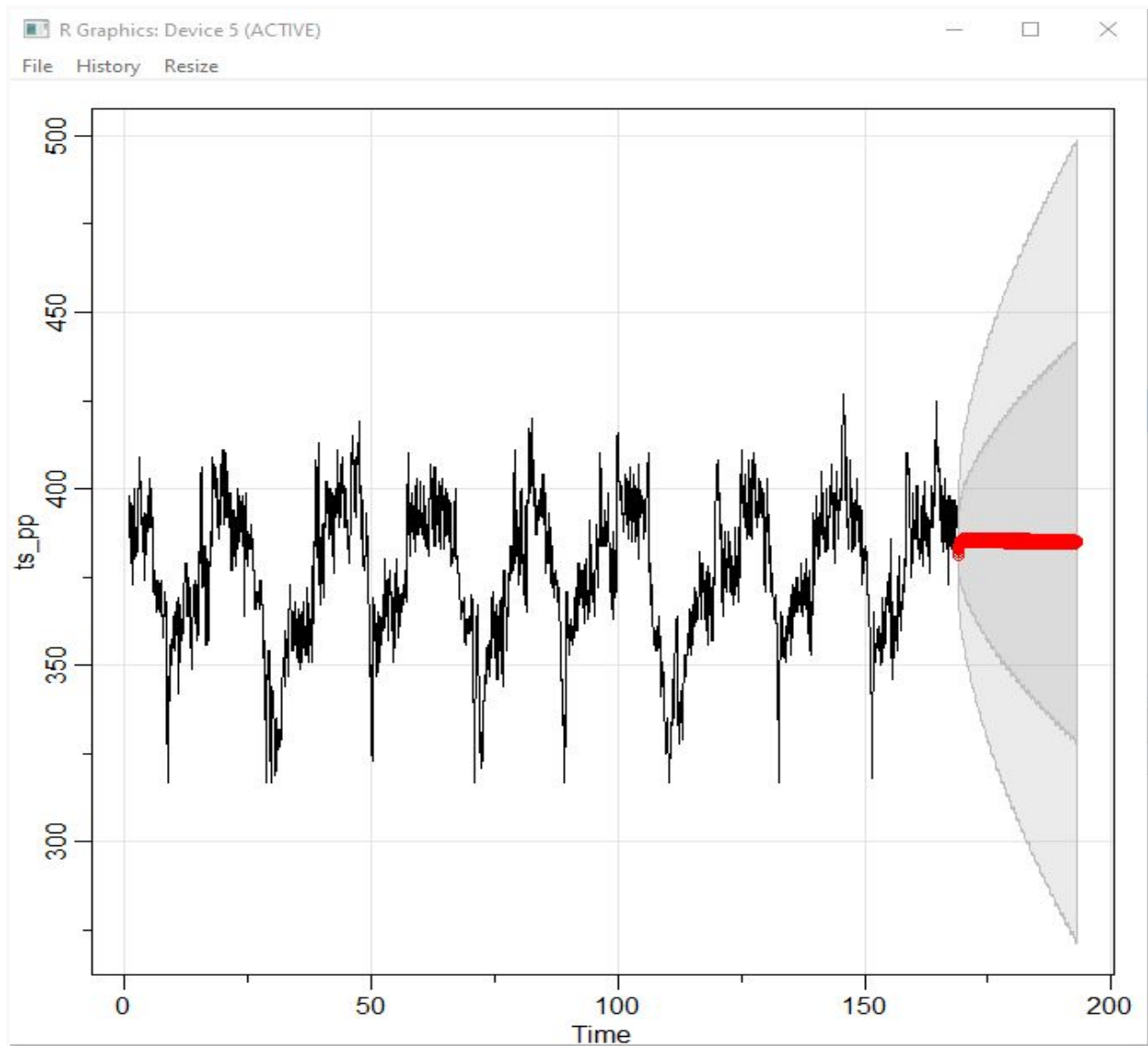
ahead_4_hours <- 4*30
dev.new()
fore.mod <- sarima.for(ts_pp, n.ahead=ahead_4_hours,
                      p=2, d=1, q=2, P=0, D=1, Q=1, S=30, plot.all=TRUE)

dev.new()
ts_pp_4_hour <- c(ts_pp[(length(ts_pp)-6*ahead_4_hours + 1):(length(ts_pp))],
                  fore.mod$pred[1:length(fore.mod$pred)])
tsplot(ts_pp_4_hour,
       ylab="Data for 4 Hours ahead",
       xlab="minutes", main=expression("1 day + 2hours ahead"))

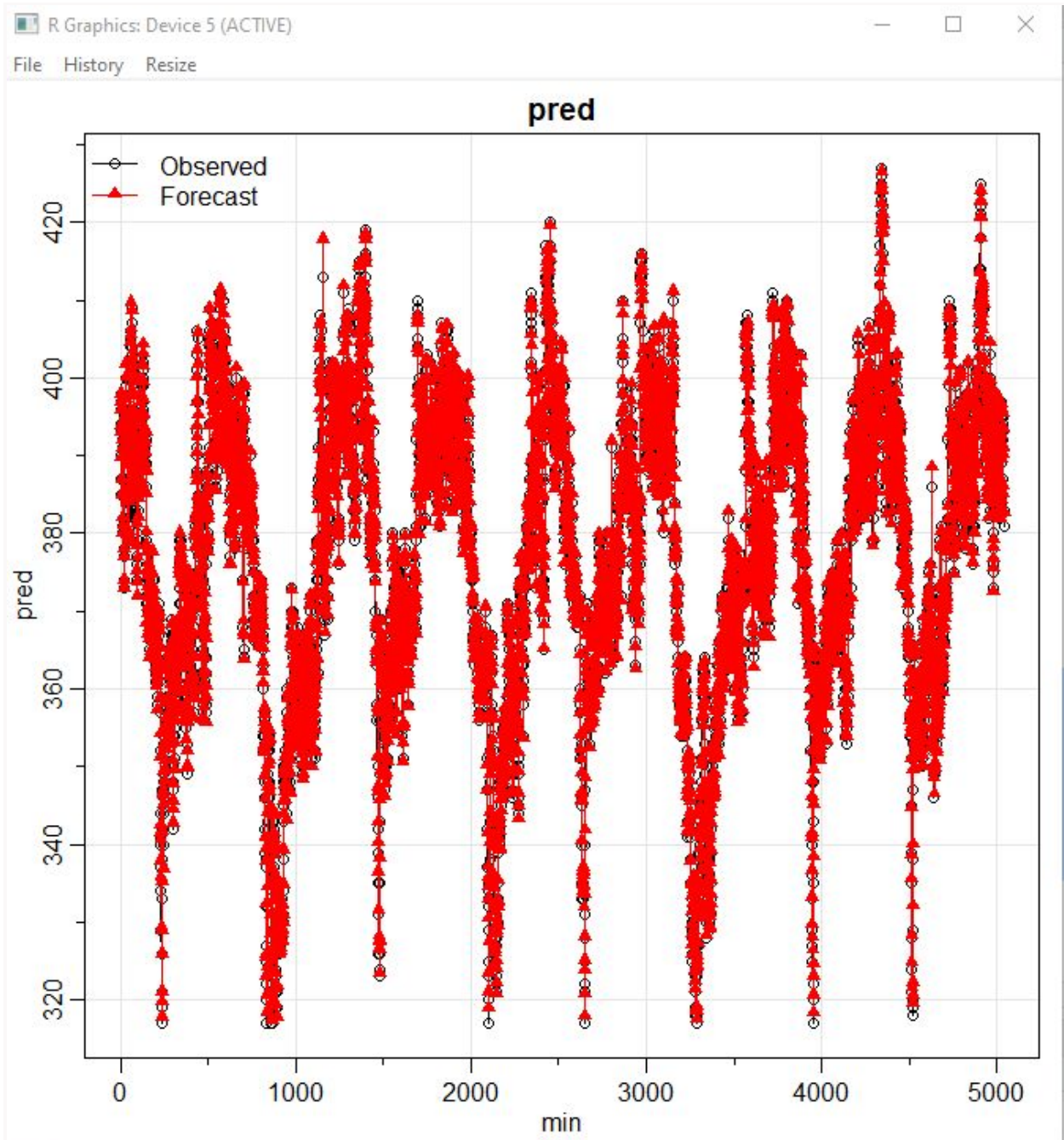
```



```
ahead_24_hours <- 24*30  
dev.new()  
fore.mod <- sarima.for(ts_pp, n.ahead=ahead_4_hours,  
                        p=2, d=1, q=2, P=0, D=1, Q=1, S=30, plot.all=TRUE)
```

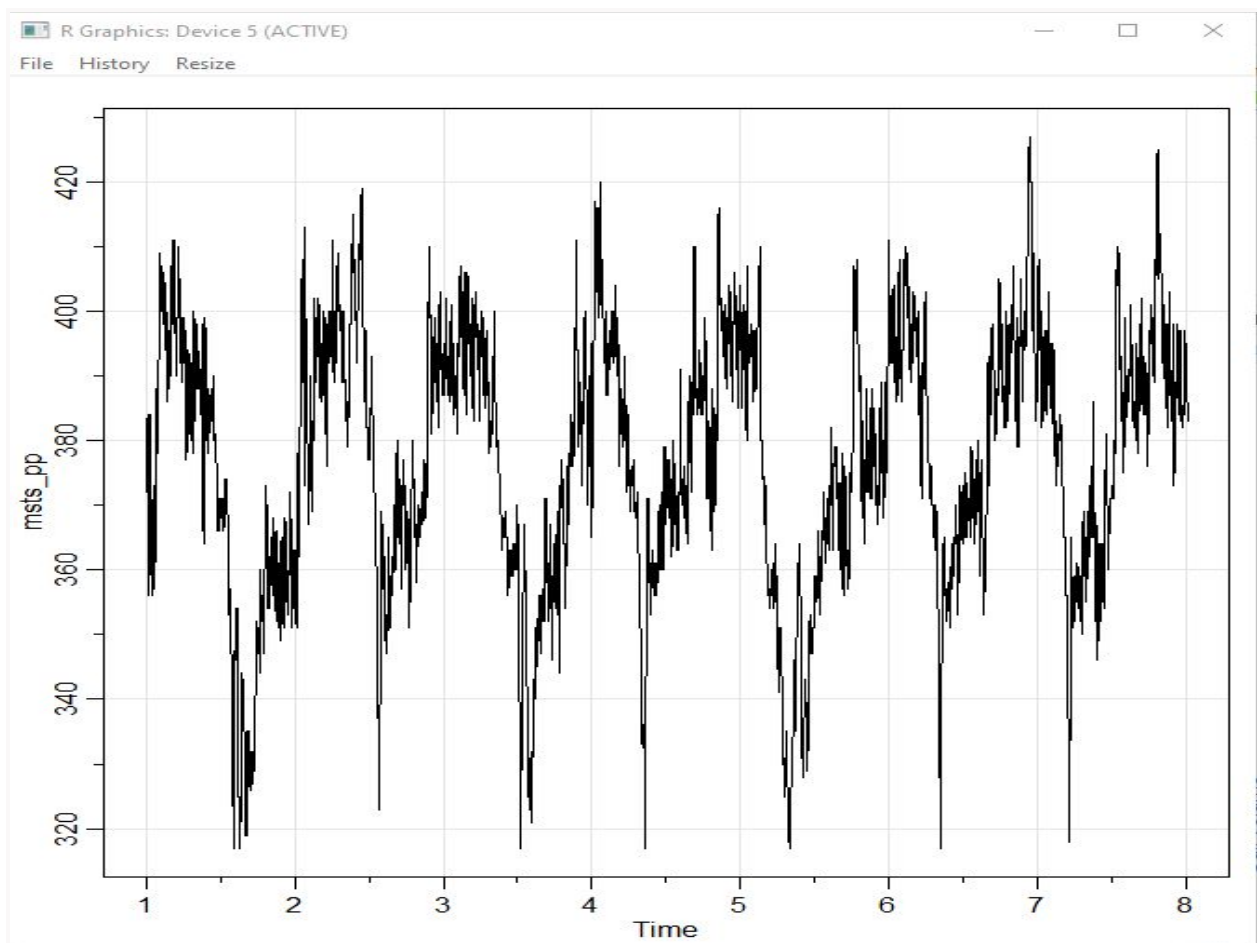


```
pred.mod <- ts(x - mod.fit.110.011$fit$residuals)
dev.new()
tsplot(ts_pp[1:length(ts_pp)], ylab="pred", xlab="min", type="o", main="pred")
lines(pred.mod, col="red", type="o", pch=17)
legend("topleft", legend=c("Observed", "Forecast"), lty=c("solid", "solid"), col=c("black",
"red"), pch=c(1, 17), bty="n")
```



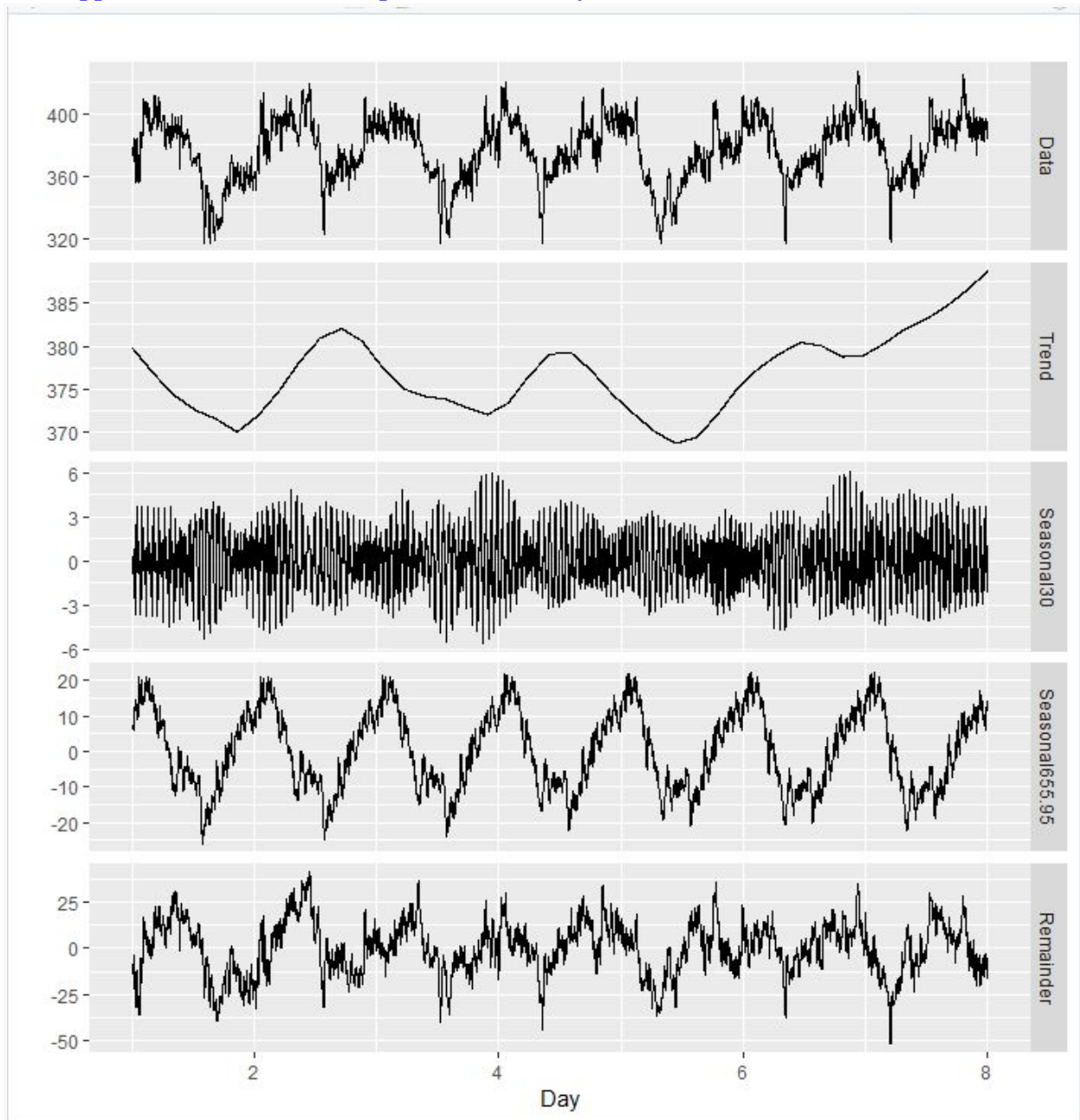
Discussion

```
On_hour <- 30  
one_day <- 30*24  
one_week <- 30*24*7  
frequency <- c(30,one_day)  
msts_pp <- msts(ts_p[(length(ts_p)-(one_week) + 1):(length(ts_p))],seasonal.periods =  
frequency)  
dev.new()tsplot(msts_pp)
```



The `mstl()` function is a variation on `stl()` designed to deal with multiple seasonality. It will return multiple seasonal components, as well as a trend and remainder component.

```
msts_pp %>% mstl() %>% autoplot() + xlab("Day")
```



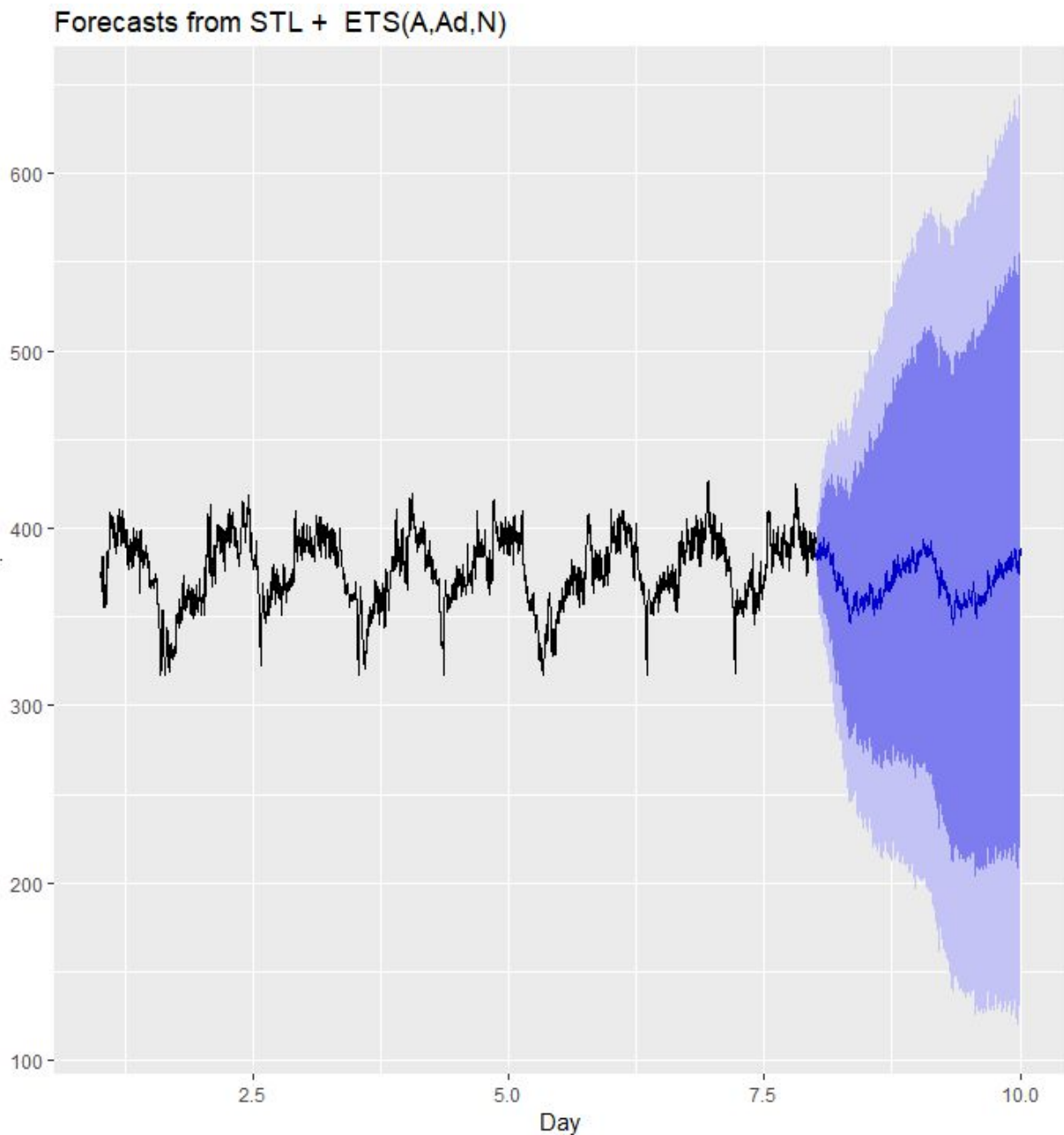
There are two seasonal patterns shown, one for the time of minutes in hour (the third panel), and one for the time of day (the fourth panel). To properly interpret this graph, it is important to notice the vertical scales. In this case,

the trend and the daily seasonality have relatively narrow ranges compared to the other components, because there is little trend seen in the data, and the daily seasonality is weak.

The decomposition can also be used in forecasting, with each of the seasonal components forecast using a seasonal naïve method, and the seasonally adjusted data forecasting using ETS (or some other user-specified method).

The `stlf()` function will do this automatically.

```
msts_pp %>% stlf() %>%  
autoplot() + xlab("Day")
```



Dynamic harmonic regression with multiple seasonal periods

Because there are multiple seasonalities, we need to add Fourier terms for each seasonal period. In this case, the seasonal periods are 30 and 656, so the Fourier terms are of the form

$$\sin(2\pi kt/30), \cos(2\pi kt/30), \sin(2\pi kt/656), \cos(2\pi kt/656)$$

For $k = 1, 2, \dots$

The `fourier()` function can generate these for you.

We will fit a dynamic harmonic regression model with an ARMA error structure. The total number of Fourier terms for each seasonal period have been chosen to minimise the AICc. We will use a log transformation (`lambda=0`) to ensure the forecasts and prediction intervals remain positive.

```
fit <- auto.arima(msts_pp, seasonal=FALSE, lambda=0, xreg=fourier(msts_pp, K=c(14,65)))
fit
```

```
> fit <- auto.arima(msts_pp, seasonal=FALSE, lambda=0,
+                   xreg=fourier(msts_pp, K=c(14,65)))
> fit
Series: msts_pp
Regression with ARIMA(3,1,1) errors
Box Cox transformation: lambda= 0
```

Coefficients:		ar1	ar2	ar3	ma1	S1-30	C1-30	S2-30	C2-30	S3-30	C3-30	S4-30	C4-30	S5-30	C5-30	S6-30	C6-30	S7-30	C7-30
s.e.		1.0682	-0.1358	-0.0397	-0.9440	9e-04	-1e-04	-1e-04	2e-04	-3e-04	6e-04	2e-04	2e-04	1e-04	0e+00	2e-04	-3e-04	3e-04	1e-04
		0.0187	0.0216	0.0154	0.0117	1e-03	1e-03	6e-04	6e-04	4e-04	4e-04	3e-04	3e-04	2e-04	2e-04	2e-04	2e-04	1e-04	1e-04
		0.0053	0.0037	0.0037	0.0029	0.0029	0.0025	0.0025	0.0022	0.0022	0.0020	0.0020	0.0019	0.0019	0.0019	0.0019	0.0018	0.0018	0.0017
		0e+00	-1e-04	-1e-04	0e+00	-2e-04	0e+00	0e+00	-1e-04	1e-04	0e+00	0e+00	-1e-04	-1e-04	-2e-04	0.0069	0.0343	0.0034	
s.e.		1e-04	1e-04	1e-04	1e-04	1e-04	1e-04	1e-04	1e-04	1e-04	1e-04	1e-04	1e-04	1e-04	1e-04	0.0102	0.0102	0.0053	
		C2-656	S3-656	C3-656	S4-656	C4-656	S5-656	C5-656	S6-656	C6-656	S7-656	C7-656	S8-656	C8-656	S9-656	C9-656	S10-656		
		0.0039	0.0059	-0.0076	-0.0012	0.0018	0.0009	-0.0032	0.0000	-0.0012	-0.0033	-0.0013	0.0038	0.0017	0.0013	-0.0007	-0.0005		
s.e.		0.0033	0.0037	0.0037	0.0029	0.0029	0.0025	0.0025	0.0022	0.0022	0.0020	0.0020	0.0019	0.0019	0.0019	0.0018	0.0018	0.0017	
		C10-656	S11-656	C11-656	S12-656	C12-656	S13-656	C13-656	S14-656	C14-656	S15-656	C15-656	S16-656	C16-656	S17-656	C17-656			
		-0.0010	0.0000	-0.0006	0.0015	-0.0022	-0.0009	0.0035	0.0024	-0.0023	0.0003	0.0003	-0.0025	0.0017	0.0005	0.0006			
s.e.		0.0017	0.0016	0.0016	0.0015	0.0014	0.0014	0.0014	0.0014	0.0013	0.0013	0.0013	0.0013	0.0013	0.0012	0.0012	0.0012		
		S18-656	C18-656	S19-656	C19-656	S20-656	C20-656	S21-656	C21-656	S22-656	C22-656	S23-656	C23-656	S24-656	C24-656	S25-656			
		-0.0002	0.0008	0.0024	0.0005	-0.0010	0.0024	-0.0009	-0.0016	-6e-04	-1e-04	5e-04	-8e-04	6e-04	-0.0011	6e-04			
s.e.		0.0012	0.0012	0.0011	0.0011	0.0011	0.0011	0.0011	0.0011	1e-03	1e-03	1e-03	1e-03	1e-03	0.0010	9e-04			
		C25-656	S26-656	C26-656	S27-656	C27-656	S28-656	C28-656	S29-656	C29-656	S30-656	C30-656	S31-656	C31-656	S32-656	C32-656			
		-0.0011	-9e-04	8e-04	-4e-04	-6e-04	0.0019	-3e-04	-8e-04	0e+00	-7e-04	-0.0013	1e-04	9e-04	-3e-04	8e-04			
s.e.		0.0009	9e-04	9e-04	9e-04	9e-04	0.0008	8e-04	8e-04	8e-04	8e-04	0.0008	8e-04	8e-04	8e-04	8e-04			
		S33-656	C33-656	S34-656	C34-656	S35-656	C35-656	S36-656	C36-656	S37-656	C37-656	S38-656	C38-656	S39-656	C39-656	S40-656			
		4e-04	-2e-04	-7e-04	-8e-04	-1e-03	-4e-04	-6e-04	-2e-04	0e+00	5e-04	1e-04	1e-03	-0.0013	-0.0015	0e+00			

```

s.e.      7e-04      7e-04      7e-04      7e-04      7e-04      7e-04      7e-04      7e-04      7e-04      7e-04      6e-04      6e-04      0.0006      0.0006      6e-04
C40-656 S41-656 C41-656 S42-656 C42-656 S43-656 C43-656 S44-656 C44-656 S45-656 C45-656 S46-656 C46-656 S47-656 C47-656
1e-04      0e+00      -2e-04      -7e-04      6e-04      -4e-04      -1e-04      3e-04      -8e-04      -5e-04      -1e-04      -2e-04      4e-04      2e-04      5e-04
s.e.      6e-04      6e-04      6e-04      6e-04      6e-04      6e-04      6e-04      6e-04      6e-04      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04
S48-656 C48-656 S49-656 C49-656 S50-656 C50-656 S51-656 C51-656 S52-656 C52-656 S53-656 C53-656 S54-656 C54-656 S55-656
0e+00      -4e-04      2e-04      -5e-04      0e+00      2e-04      -4e-04      -6e-04      -1e-04      -4e-04      1e-04      -8e-04      -2e-04      -3e-04      -5e-04
s.e.      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04      5e-04      4e-04
C55-656 S56-656 C56-656 S57-656 C57-656 S58-656 C58-656 S59-656 C59-656 S60-656 C60-656 S61-656 C61-656 S62-656 C62-656
-1e-04      -2e-04      -1e-04      1e-04      3e-04      -2e-04      3e-04      5e-04      0e+00      2e-04      -1e-04      -1e-04      1e-04      -5e-04      -1e-04
s.e.      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04
S63-656 C63-656 S64-656 C64-656 S65-656 C65-656
6e-04      1e-04      4e-04      6e-04      -1e-04      4e-04
s.e.      4e-04      4e-04      4e-04      4e-04      4e-04      4e-04

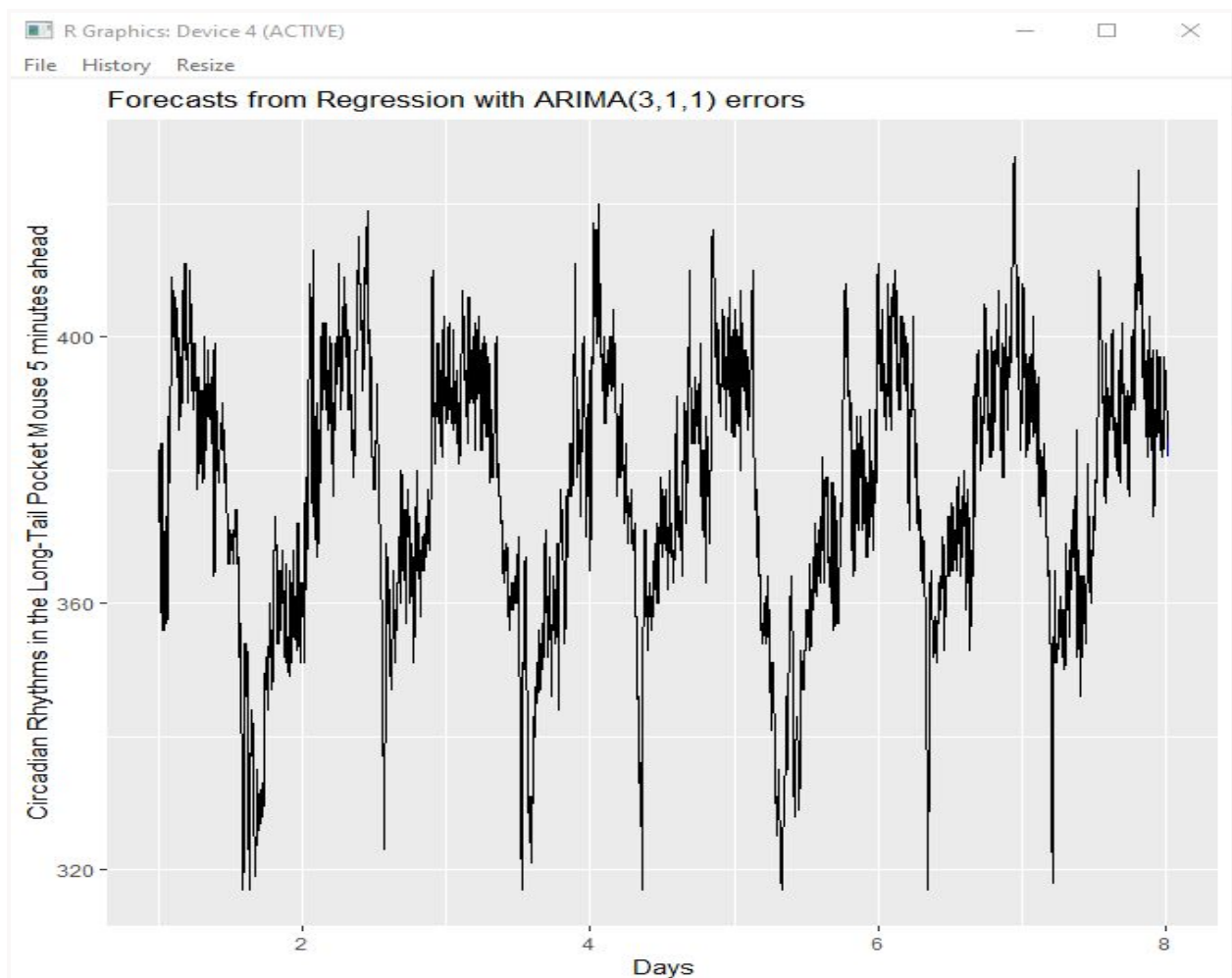
sigma^2 estimated as 8.103e-05: log likelihood=15190.22
AIC=-30054.44 AICC=-30042.36 BIC=-29006.08

```

```

dev.new()
minutes_10<- 5
fit %>%
  forecast(xreg=fourier(msts_pp, K=c(14,65), minutes_10)) %>%
  autoplot(include=one_week) + ylab("Circadian Rhythms in the Long-Tail Pocket Mouse 10
minutes ahead") + xlab("Days")

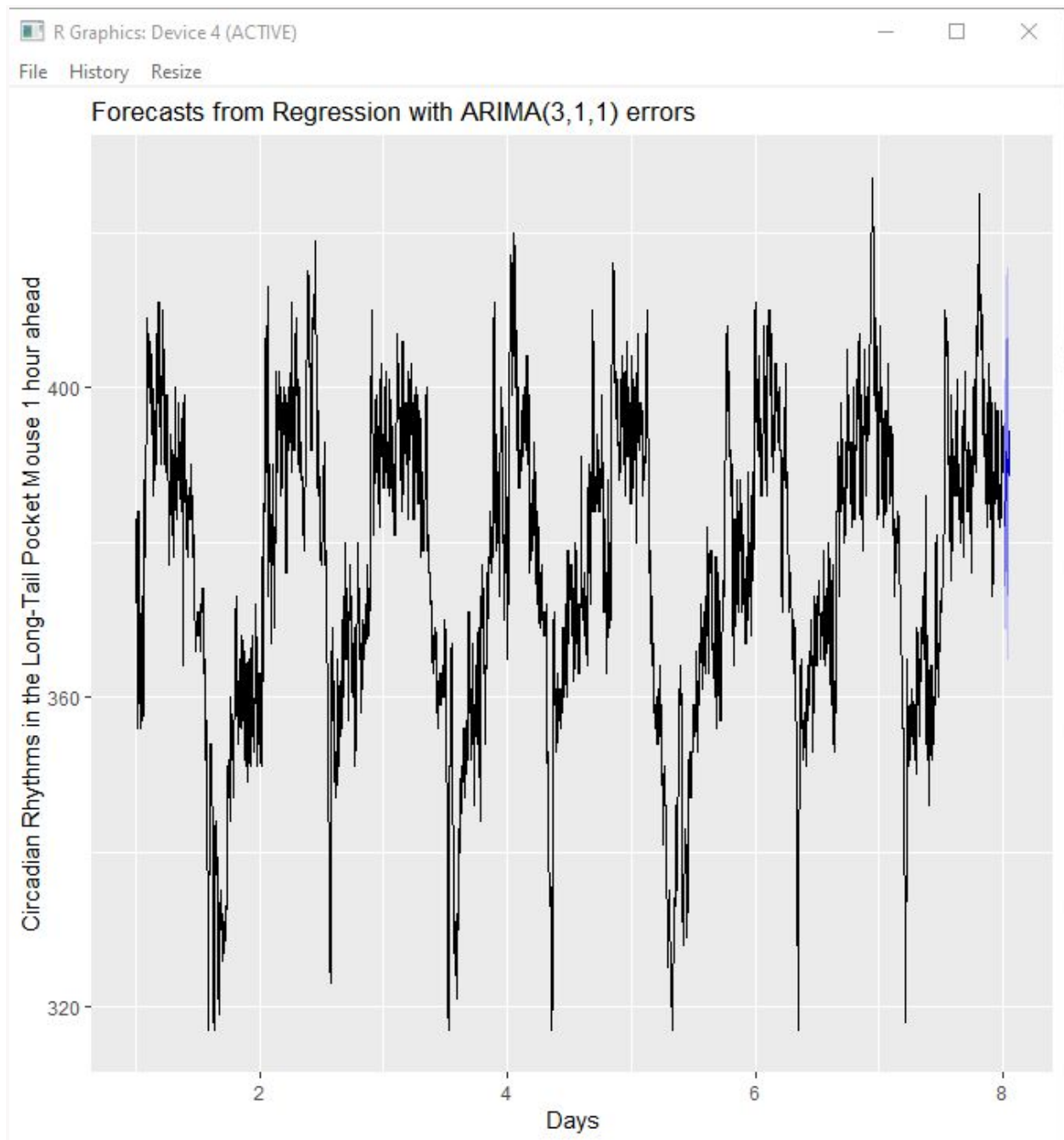
```



```

dev.new()
hour_1<- 30
fit %>%
  forecast(xreg=fourier(msts_pp, K=c(14,65), h=hour_1)) %>%
  autoplot(include=one_week) + ylab("Circadian Rhythms in the Long-Tail Pocket Mouse 1
hour ahead") + xlab("Days")

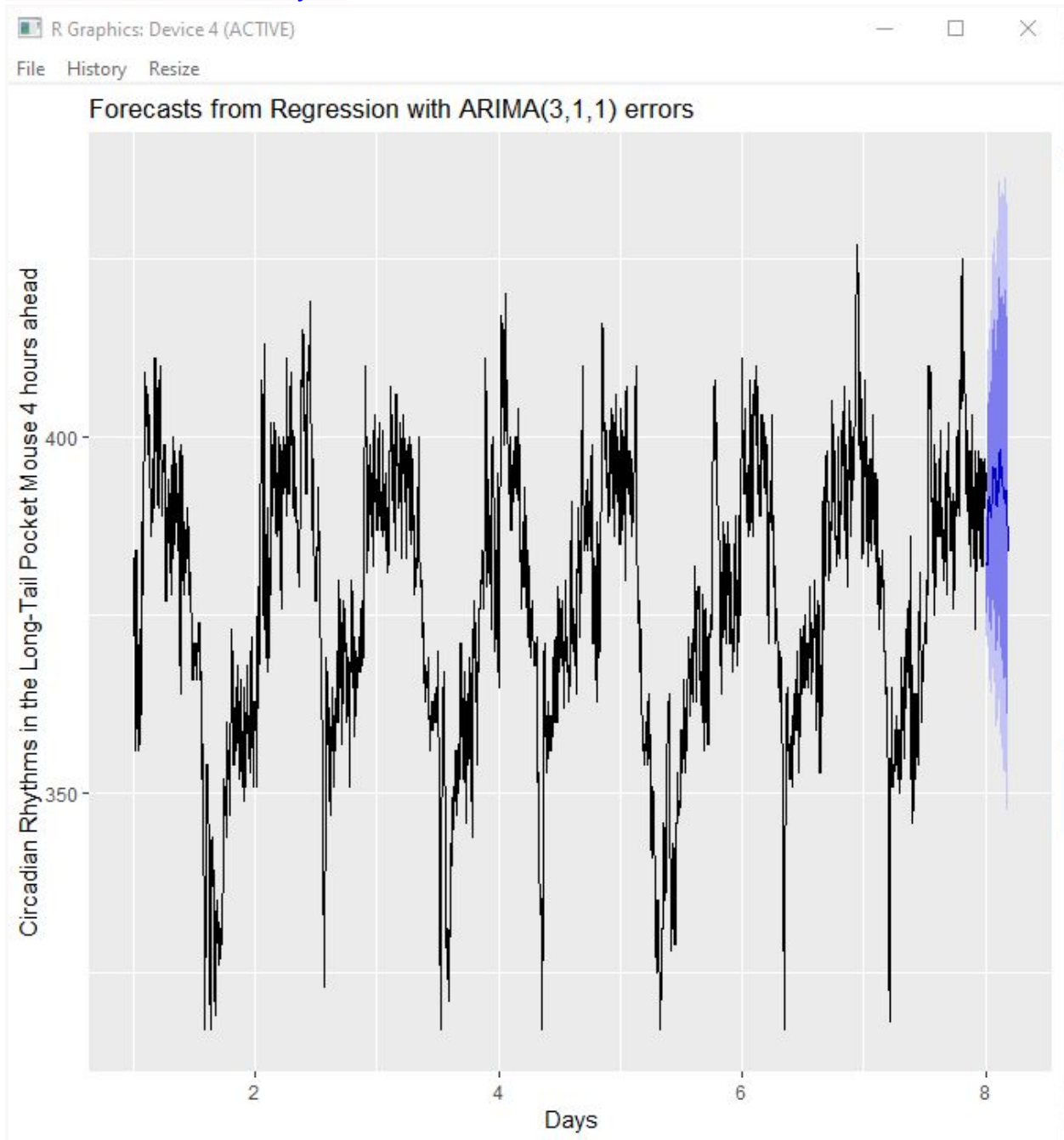
```



```

dev.new()
hours_4<- 30 *4
fit %>%
  forecast(xreg=fourier(msts_pp, K=c(14,65), h=hours_4)) %>%
  autoplot(include=one_week) + ylab("Circadian Rhythms in the Long-Tail Pocket Mouse 4
hours ahead") + xlab("Days")

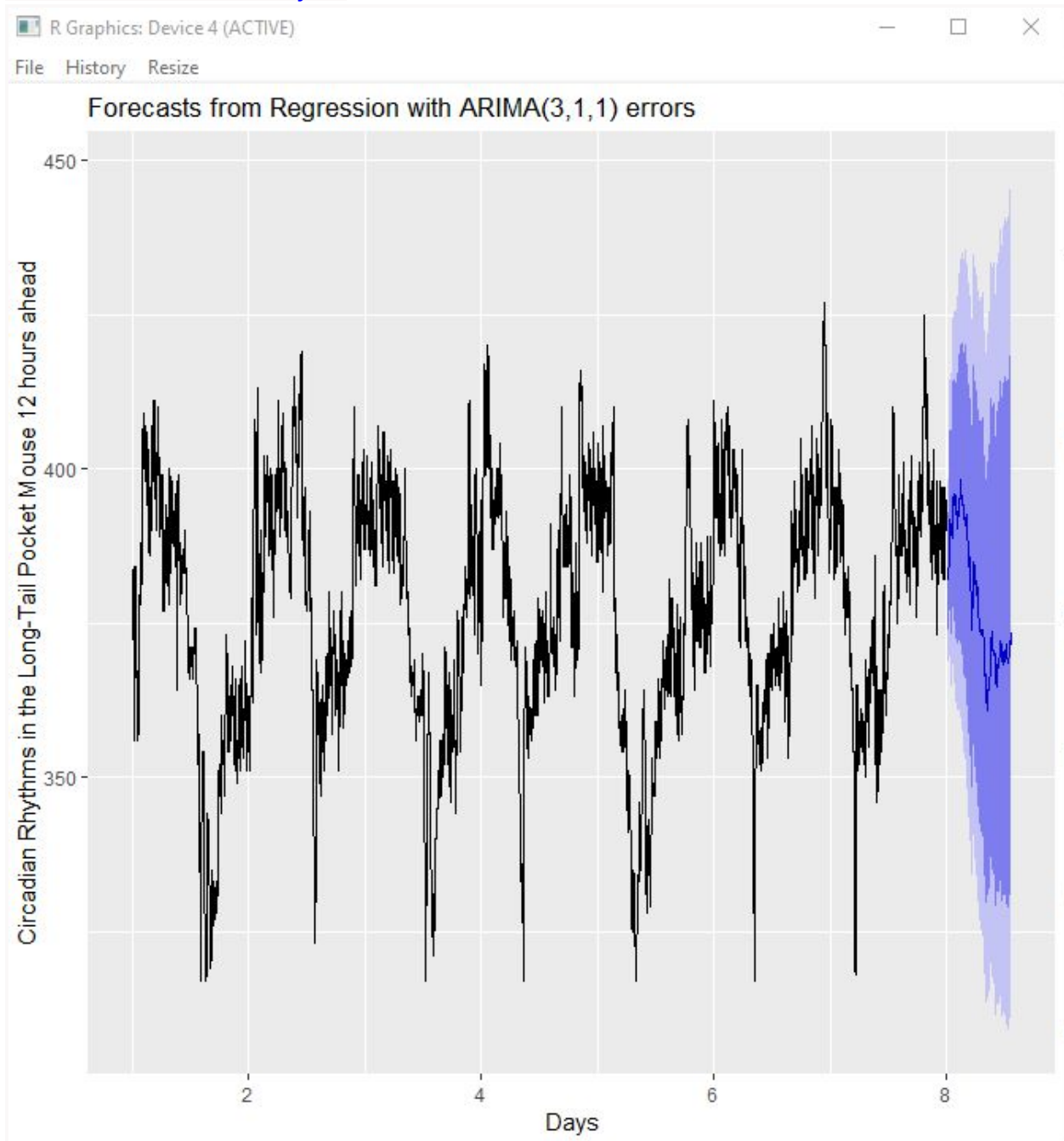
```



```

dev.new()
hours_12<- 30 * 12
fit %>%
  forecast(xreg=fourier(msts_pp, K=c(14,65), h=hours_12)) %>%
  autoplot(include=one_week) + ylab("Circadian Rhythms in the Long-Tail Pocket Mouse 12
hours ahead") + xlab("Days")

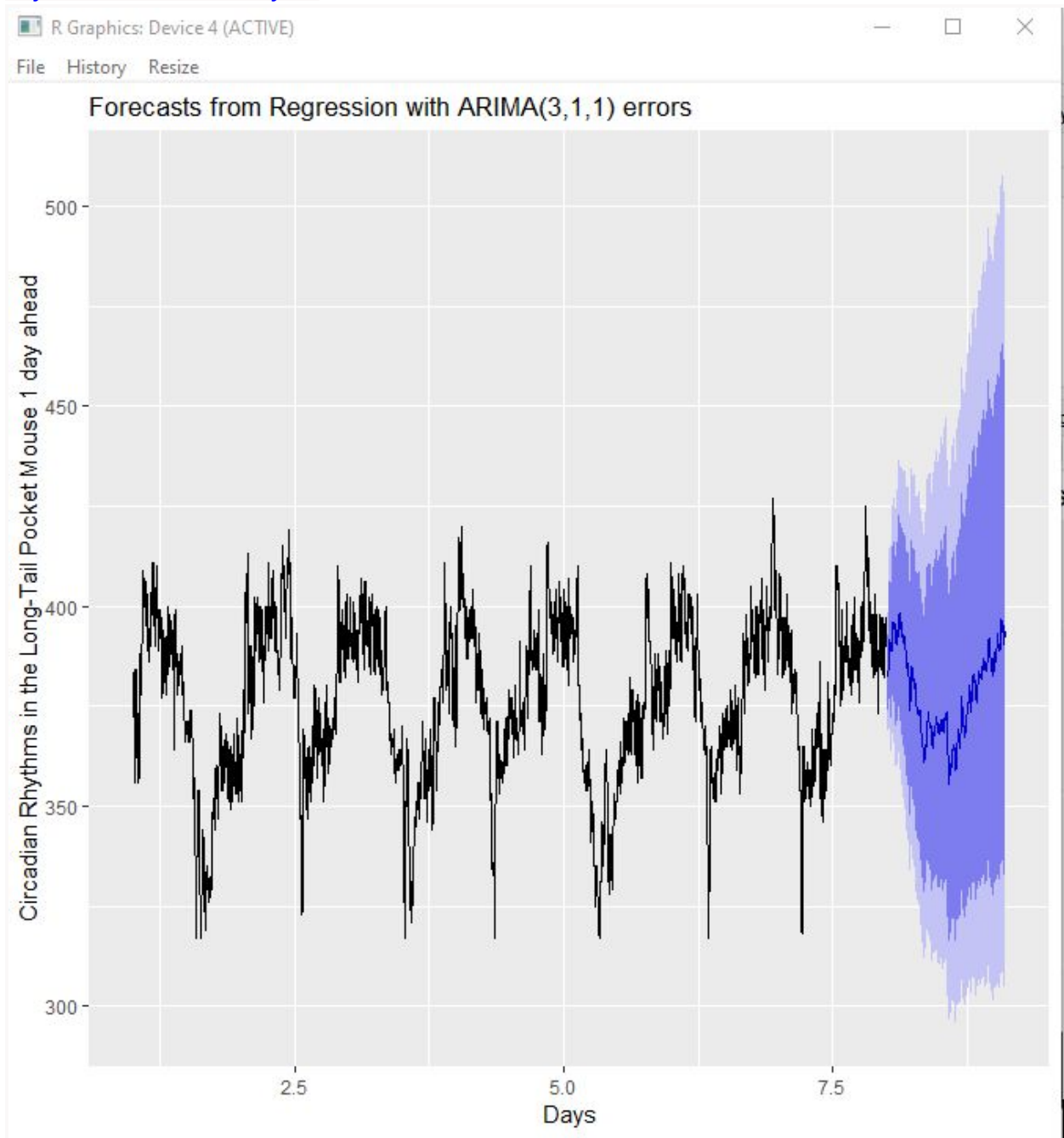
```



```

dev.new()
day_1 <- 30 * 24
fit %>%
  forecast(xreg=fourier(msts_pp, K=c(14,65), h=day_1)) %>%
  autoplot(include=one_week) + ylab("Circadian Rhythms in the Long-Tail Pocket Mouse 1
day ahead") + xlab("Days")

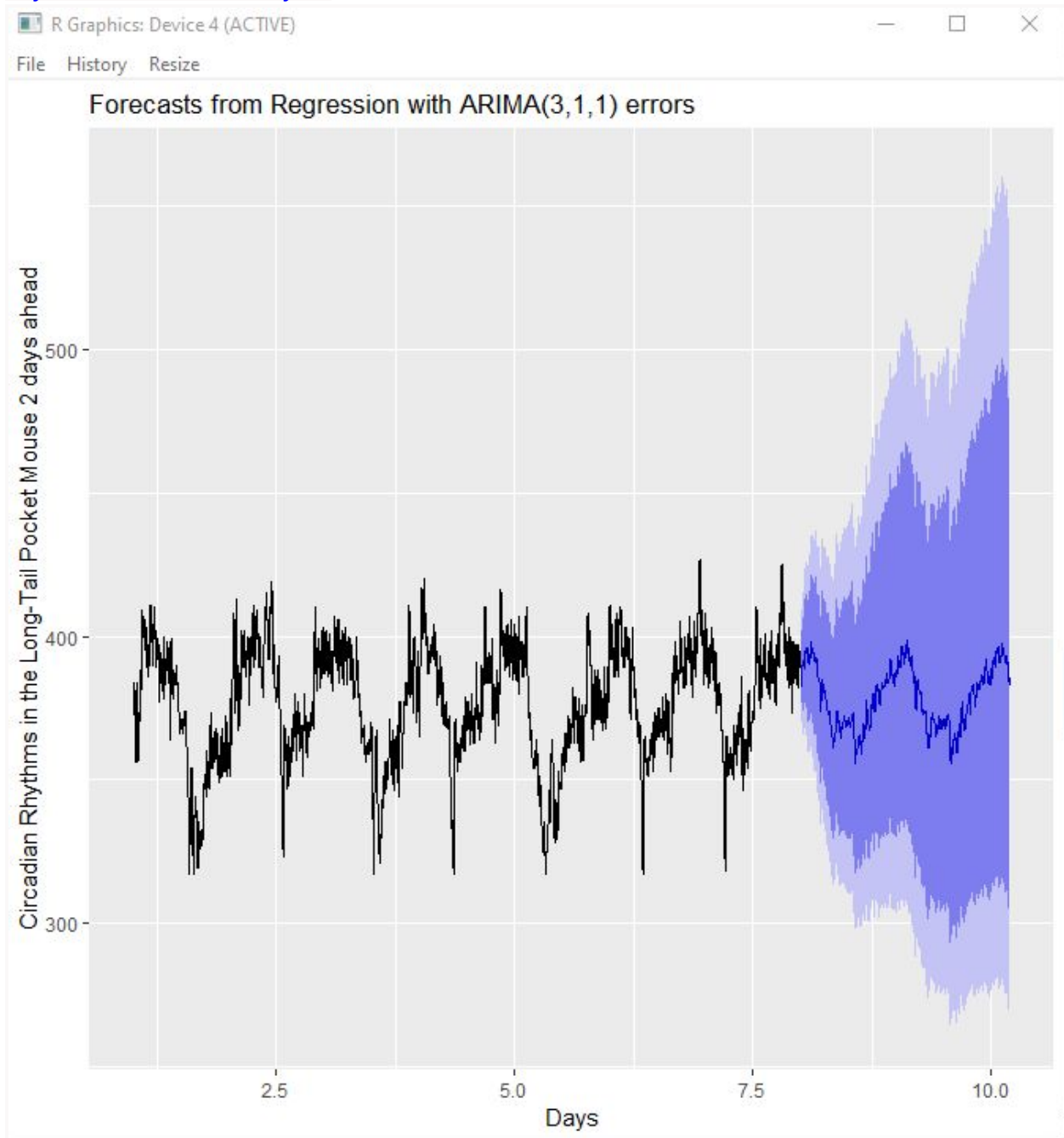
```



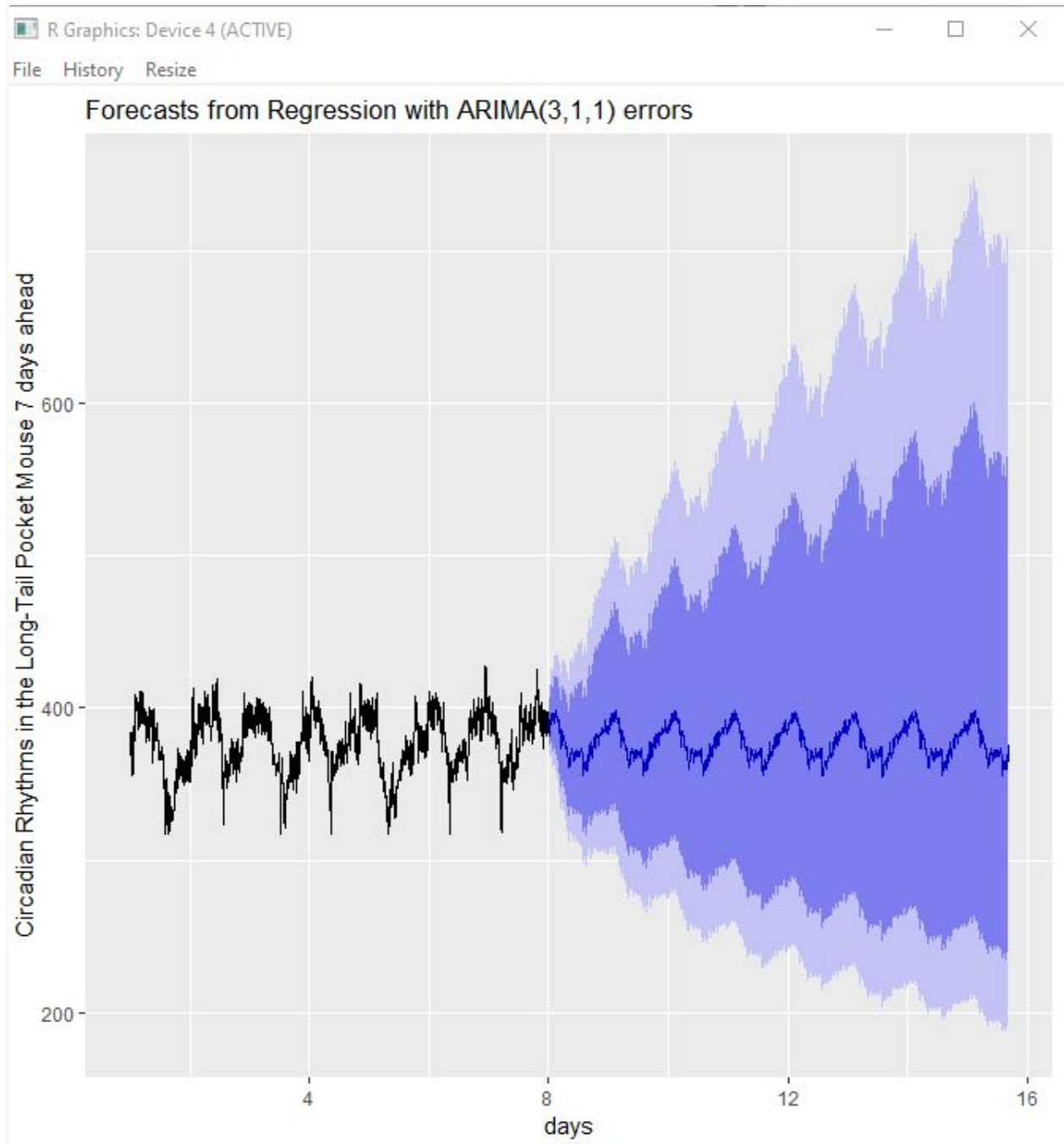

```

dev.new()
day_2<- 30 * 24 * 2
fit %>%
  forecast(xreg=fourier(msts_pp, K=c(14,65), h=day_2)) %>%
  autoplot(include=one_week) + ylab("Circadian Rhythms in the Long-Tail Pocket Mouse 2
days ahead") + xlab("Days")

```



```
dev.new()
week_1<- 30 * 24 * 7
fit %>%
  forecast(xreg=fourier(msts_pp, K=c(14,65), h=week_1)) %>%
  autoplot(include=one_week) + ylab("Circadian Rhythms in the Long-Tail Pocket Mouse 7 days ahead") + xlab("days")
```



Bibliography

David Benson, Forecasting Daily Data with Multiple Seasonality in R
<http://www.dbenson.co.uk/Rparts/subpages/forecastR/>

E. E. Holmes, M.D. Scheuerell, and E. J. Ward 2020-02-03,
Analysis for Fisheries and Environmental Sciences

Rob J Hyndman, George Athanasopoulos,
Forecasting Principles And Practice
DataSet <http://www.statsci.org/data/general/pformos5.html>

Appendix

```
library(astsa)
library("forecast")
library("lubridate")
examine.mod <- function(mod.fit.obj, p, d, q, P=0, D=0, Q=0, S=-1, lag.max=24) {
  dev.new(width=6, height=6)
  par(mfrow=c(2,1))
  pacf(mod.fit.obj$fit$residuals, main="PACF of Residuals", lag.max)
  if ((P==0)&(D==0)&(Q==0)) {
    title(paste("Model: (", p, ", ", d, ", ", q, ")", sep=""), adj=0, cex.main=0.75)
  }
  else {
    title(paste("Model: (", p, ", ", d, ", ", q, ") (", P, ", ", D, ", ", Q, ") [", S, "]", sep=""), adj=0,
    cex.main=0.75)
  }

  std.resid <- mod.fit.obj$fit$residuals/sqrt(mod.fit.obj$fit$sigma2)
  hist(std.resid, main="Histogram of Standardized Residuals", xlab="Standardized
Residuals", freq=FALSE)
  curve(expr=dnorm(x, mean=mean(std.resid), sd=sd(std.resid)), col="red", add=TRUE)
}

pformosu<-read.table(file =
"C:/Users/inna/Desktop/DepaulClasses/ApplyMathClasses/Time_Series/Final
Project/pformosu.txt")

row =3726
col= 16
pformosu__dat = numeric(col*row)
count = 1
for(i in 1:row)
{
  for(j in 1:col)
  {
    pformosu__dat[count]=pformosu[i,j]
    count = count + 1
  }
}
```

```

ts_p<-ts(pformosu_dat)
dev.new()
par(mfrow=c(2,1))
tsplot(ts_p)
boxplot(ts_p)$out
#####
num_of_outliers = 0
repeat
{
  outliers <- boxplot(ts_p, plot=FALSE)$out
  v<-(which(ts_p %in% outliers))
  if(length(v)== 0)
    break
  else
  {
    ts_p<-ts_p[-v]
    num_of_outliers = num_of_outliers + length(v)
  }
}
print(paste("Number Of Outliers removed = ", num_of_outliers))
print(paste("Length Of Time Series",length(ts_p)))
dev.new()
par(mfrow=c(2,1))
tsplot(ts_p,xlab="Number Of Observations",main="Circadian rhythms in the last week Of
data")
boxplot(ts_p)$out

#####
frequency <- 30
one_day<- 30*24
one_week<- 24*30*7
ts_pp <- ts(ts_p[(length(ts_p)-(one_week) + 1):(length(ts_p))],frequency = frequency)
dev.new()
tsplot(ts_pp,xlab="Number Of Observations -> 7 Days -> 168 Hours",
      main="Circadian rhythms in the last week Of data")
x<-as.numeric(ts_pp)

#####
dev.new()
tsplot(ts_pp, ylab=expression(x[t]),
      xlab="hours", main="Circadian rhythms in the last week Of data")

dev.new()
acf2(x, max.lag = length(ts_pp) - 1,main = "ACF & PACF Of Original ts")
dev.new()
acf2(x, max.lag = 100,main = "ACF & PACF Of Original ts")

```

```
#####
# Plot of  $(1-B^{30})x_t$ 
dev.new()
tsplot(diff(ts_pp, lag=frequency, differences=1), ylab=expression((1-B^30)*x[t]),
       xlab="hours", main=expression(paste("Plot of ", (1-B^30)*x[t])))

# ACF indicated cut of after 330 lag that suggests that ARIMA Q=1,
# PACF exponentially decay that suggests ARIMA P=0
# Diff =1 S=18  $(1-B^{30})x_t$ 
dev.new()
acf2(diff(x, lag=frequency, differences=1), max.lag=frequency*5,
     main=expression(paste("Est. ACF & PACF for ", (1-B^30)*x[t])))

#####
# Plot of  $(1-B)(1-B^{30})x_t$ 
dev.new()
tsplot(diff(diff(ts_pp, lag=frequency, differences=1)),
       ylab=expression((1-B)(1-B^30)*x[t]),
       xlab="Hours", main=expression(paste("Plot of ",
                                           (1-B)(1-B^30)*x[t])))

# ACF and PACF of  $(1-B)(1-B^{30})x_t$ 
dev.new()
acf2(diff(diff(x[1:length(x)], lag=frequency, differences=1)),
     max.lag=frequency,
     main=expression(paste("Est. ACF & PACF for ", (1-B)(1-B^30)*x[t])))
#####
#####
# First Estimate ARIMA(p=0,d=1,q=0,P=0,D=1,Q=1,S=30) #####

dev.new()
mod.fit1<- sarima(ts_pp,p=0,d=1,q=0,P=0,D=1,Q=1,S=30)
mod.fit1

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$fit$residuals,main = " PACF Of residuals")
title(paste("Model: (",
           p = 0, ", ",
           d = 1, ", ",
           q = 0, ")(",
           P = 0, ", ",
           D = 1, ", ",
           Q = 1, ")[30]",
           sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$fit$residuals / sqrt(mod.fit1$fit$sigma2)
hist(std.resid1,main = " Histogram Of Standirdized Residuals",
```

```

    xlab='Standardized Residuals',
    freq = FALSE
)
### # First Estimate ARIMA(p=0,d=1,q=1,P=0,D=1,Q=1,S=30) #####

dev.new()
mod.fit1<- sarima(ts_pp,p=0,d=1,q=1,P=0,D=1,Q=1,S=30)
mod.fit1

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$fit$residuals,main = " PACF Of residuals")
title(paste("Model: (",
  p = 0, ", ",
  d = 1, ", ",
  q = 1, ")(",
  P = 0, ", ",
  D = 1, ", ",
  Q = 1, ")[30]",
  sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$fit$residuals / sqrt(mod.fit1$fit$sigma2)
hist(std.resid1,main = " Histogram Of Standardized Residuals",
  xlab='Standardized Residuals',
  freq = FALSE
)
### ARIMA(p=1,d=1,q=1,P=0,D=1,Q=1,S=30)
dev.new()
mod.fit1<- sarima(ts_pp,p=1,d=1,q=1,P=0,D=1,Q=1,S=30)
mod.fit1

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$fit$residuals,main = " PACF Of residuals")
title(paste("Model: (",
  p = 1, ", ",
  d = 1, ", ",
  q = 1, ")(",
  P = 0, ", ",
  D = 1, ", ",
  Q = 1, ")[30]",
  sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$fit$residuals / sqrt(mod.fit1$fit$sigma2)
hist(std.resid1,main = " Histogram Of Standardized Residuals",
  xlab='Standardized Residuals',
  freq = FALSE
)
#### ARIMA(p=0,d=1,q=2,P=0,D=1,Q=1,S=30)
dev.new()

```

```

mod.fit1<- sarima(ts_pp,p=0,d=1,q=2,P=0,D=1,Q=1,S=30)
mod.fit1

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$fit$residuals,main = " PACF Of residuals")
title(paste("Model: (",
  p = 0, ", ",
  d = 1, ", ",
  q = 2, ")",
  P = 0, ", ",
  D = 1, ", ",
  Q = 1, ")[30]",
  sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$fit$residuals / sqrt(mod.fit1$fit$sigma2)
hist(std.resid1,main = " Histogram Of Standardized Residuals",
  xlab='Standardized Residuals',
  freq = FALSE
)
#### ARIMA(p=0,d=1,q=3,P=0,D=1,Q=1,S=30)
dev.new()
mod.fit1<- sarima(ts_pp,p=0,d=1,q=3,P=0,D=1,Q=1,S=30)
mod.fit1

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$fit$residuals,main = " PACF Of residuals")
title(paste("Model: (",
  p = 0, ", ",
  d = 1, ", ",
  q = 3, ")",
  P = 0, ", ",
  D = 1, ", ",
  Q = 1, ")[30]",
  sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$fit$residuals / sqrt(mod.fit1$fit$sigma2)
hist(std.resid1,main = " Histogram Of Standirdized Residuals",
  xlab='Standardized Residuals',
  freq = FALSE
)
#### ARIMA(p=1,d=1,q=0,P=0,D=1,Q=1,S=30)
dev.new()
mod.fit1<- sarima(ts_pp,p=1,d=1,q=0,P=0,D=1,Q=1,S=30)
mod.fit1

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$fit$residuals,main = " PACF Of residuals")

```

```

title(paste("Model: (",
  p = 1, ", ",
  d = 1, ", ",
  q = 0, ")(",
  P = 0, ", ",
  D = 1, ", ",
  Q = 1, ")[30]",
  sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$fit$residuals / sqrt(mod.fit1$fit$sigma2)
hist(std.resid1,main = " Histogram Of Standirdized Residuals",
  xlab='Standardized Residuals',
  freq = FALSE
)
#### ARIMA(p=2,d=1,q=0,P=0,D=1,Q=1,S=30)
dev.new()
mod.fit1<- sarima(ts_pp,p=2,d=1,q=0,P=0,D=1,Q=1,S=30)
mod.fit1

```

```

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$fit$residuals,main = " PACF Of residuals")
title(paste("Model: (",
  p = 2, ", ",
  d = 1, ", ",
  q = 0, ")(",
  P = 0, ", ",
  D = 1, ", ",
  Q = 1, ")[30]",
  sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$fit$residuals / sqrt(mod.fit1$fit$sigma2)
hist(std.resid1,main = " Histogram Of Standardized Residuals",
  xlab='Standardized Residuals',
  freq = FALSE
)

```

```

#### ARIMA(p=2,d=1,q=2,P=0,D=1,Q=1,S=30)
dev.new()
mod.fit1<- sarima(ts_pp,p=2,d=1,q=2,P=0,D=1,Q=1,S=30)
mod.fit1

```

```

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$fit$residuals,main = " PACF Of residuals")
title(paste("Model: (",
  p = 2, ", ",
  d = 1, ", ",
  q = 2, ")(",

```

```

P = 0, "",
D = 1, "",
Q = 1, "")[30]",
sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$fit$residuals / sqrt(mod.fit1$fit$sigma2)
hist(std.resid1,main = " Histogram Of Standardized Residuals",
     xlab='Standardized Residuals',
     freq = FALSE
)

#### ARIMA(p=2,d=1,q=3,P=0,D=1,Q=1,S=30)
dev.new()
mod.fit1<- sarima(ts_pp,p=2,d=1,q=3,P=0,D=1,Q=1,S=30)
mod.fit1

dev.new()
par(mfrow=c(2,1))
pacf(mod.fit1$fit$residuals,main = " PACF Of residuals")
title(paste("Model: (",
            p = 3, "",
            d = 1, "",
            q = 2, ")(",
            P = 0, "",
            D = 1, "",
            Q = 1, "")[30]",
            sep=""),adj=0,cex.main=0.75)
std.resid1 <- mod.fit1$fit$residuals / sqrt(mod.fit1$fit$sigma2)
hist(std.resid1,main = " Histogram Of Standardized Residuals",
     xlab='Standardized Residuals',
     freq = FALSE
)

##### Forecasting #####

dev.new()
mod.fit.110.011 <- sarima(x, p=2,d=1,q=2,P=0,D=1,Q=1,S=30)
mod.fit.110.011
examine.mod(mod.fit.110.011, 2,1,2, 0,1,1, 30)

ahead_1_hours <-30
dev.new()
fore.mod <- sarima.for(ts_pp, n.ahead=ahead_1_hours,
                      p=2, d=1, q=2, P=0, D=1, Q=1, S=30, plot.all=TRUE)

dev.new()
ts_pp_1_hour <- c(ts_pp[(length(ts_pp)-12*ahead_1_hours + 1):(length(ts_pp))],
                 fore.mod$pred[1:length(fore.mod$pred)])
tsplot(ts_pp_1_hour,
       ylab="Data for 1 Hour ahead",

```



```

      xlab="minutes", main=expression("360 minutes + 60 minutes ahead"))
#####
ahead_4_hours <- 4*30
dev.new()
fore.mod <- sarima.for(ts_pp, n.ahead=ahead_4_hours,
  p=2, d=1, q=2, P=0, D=1, Q=1, S=30, plot.all=TRUE)

dev.new()
ts_pp_4_hour <- c(ts_pp[(length(ts_pp)-6*ahead_4_hours + 1):(length(ts_pp))],
  fore.mod$pred[1:length(fore.mod$pred)])
tsplot(ts_pp_4_hour,
  ylab="Data for 4 Hours ahead",
  xlab="minutes", main=expression("1 day + 2hours ahead"))
#####
ahead_24_hours <- 24*30
dev.new()
fore.mod <- sarima.for(ts_pp, n.ahead=ahead_24_hours,
  p=2, d=1, q=2, P=0, D=1, Q=1, S=30, plot.all=TRUE)

dev.new()
ts_pp_24_hour <- c(ts_pp[(length(ts_pp)-2*ahead_24_hours + 1):(length(ts_pp))],
  fore.mod$pred[1:length(fore.mod$pred)])
tsplot(ts_pp_24_hour,
  ylab="Data for 24 Hours ahead",
  xlab="minutes", main=expression("4 day + 24 hours ahead"))
#####
ahead_24_hours <- 24*30
dev.new()
fore.mod <- sarima.for(ts_pp, n.ahead=ahead_24_hours, p=2, d=1, q=2, P=0, D=1, Q=1,
S=30, plot.all=TRUE)

pred.mod <- ts(x - mod.fit.110.011$fit$residuals)

dev.new()
tsplot(ts_pp, ylab="pred", xlab="min", type="o", main="Predicted Model fitting with
Data")
lines(pred.mod, col="red", type="o", pch=17)
legend("topleft", legend=c("Observed", "Forecast"), lty=c("solid", "solid"), col=c("black",
"red"),
  pch=c(1, 17), bty="n")

```