

Machine Learning

Lecture 8 Regression

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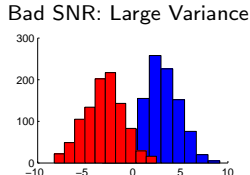
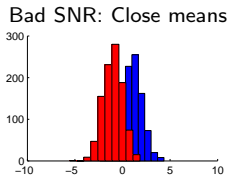
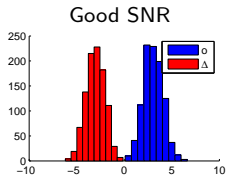
June 4, 2019



Signal-to-Noise Ratio in Classification Settings

A useful definition of Signal-to-Noise Ratio for classification is

$$\frac{\text{Between Class Variance}}{\text{Within Class Variance}} \quad (1)$$

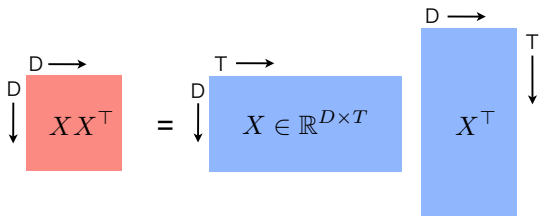


Covariance Matrices

Given T data points $\mathbf{x} \in \mathbb{R}^D$ in a data matrix $\mathbf{X} \in \mathbb{R}^{D \times T}$ the empirical estimate of the **covariance matrix** is defined as

$$\frac{1}{T} \mathbf{X} \mathbf{X}^\top \quad (2)$$

where we assume centered data, i.e. $\sum_{t=1}^T \mathbf{x}_t = 0$.

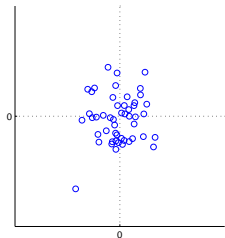


Correlated Data and Linear Mappings

Simulating correlated data can help understanding it

We generate uncorrelated data $\mathbf{x} \in \mathbb{R}^2$ drawn from a normal distribution $\mathbf{x} \sim \mathcal{N}(0, 1)$
We induce *correlations* by a diagonal scaling matrix D and a rotation matrix R

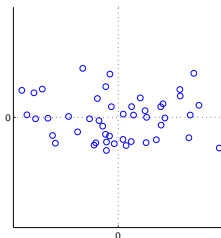
Uncorrelated



$$\mathbf{x} \sim \mathcal{N}(0, 1)$$

$$\mathbf{xx}^\top = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

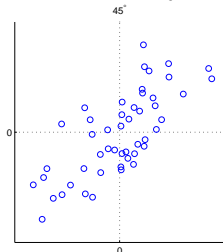
Uncorrelated, scaled



$$\begin{bmatrix} 3 & 0 \\ 0 & 1 \end{bmatrix} \mathbf{x}$$

$$\mathbf{xx}^\top = \begin{bmatrix} 9 & 0 \\ 0 & 1 \end{bmatrix}$$

Scaled, rotated by 45°



$$\begin{bmatrix} \cos(\phi) & -\sin(\phi) \\ \sin(\phi) & \cos(\phi) \end{bmatrix} \begin{bmatrix} 3 & 0 \\ 0 & 1 \end{bmatrix} \mathbf{x}$$

$$\mathbf{xx}^\top = \begin{bmatrix} 5 & 4 \\ 4 & 5 \end{bmatrix}$$



Fisher's Linear Discriminant Analysis

$$\operatorname{argmax}_w \frac{\mathbf{w}^\top \mathbf{S}_B \mathbf{w}}{\mathbf{w}^\top \mathbf{S}_W \mathbf{w}} \quad (3)$$

$$\mathbf{S}_B = \underbrace{(\boldsymbol{\mu}_+ - \boldsymbol{\mu}_-)}_{\text{Distance Class Means}} (\boldsymbol{\mu}_+ - \boldsymbol{\mu}_-)^{\top} \quad (4)$$

$$\begin{aligned} \mathbf{S}_W = & \sum_{i \in \mathcal{Y}_+1} (\mathbf{x}_i - \boldsymbol{\mu}_+)(\mathbf{x}_i - \boldsymbol{\mu}_+)^{\top} \\ & + \sum_{j \in \mathcal{Y}_{-1}} (\mathbf{x}_j - \boldsymbol{\mu}_-) \underbrace{(\mathbf{x}_j - \boldsymbol{\mu}_-)^{\top}}_{\text{Distance from Class Mean}} \end{aligned} \quad (5)$$



Fisher's Linear Discriminant Analysis

Setting the first derivative of eq. 3 to zero, we obtain the generalized eigenvalue equation

$$\mathbf{S}_B \mathbf{w} = \mathbf{S}_W \mathbf{w} \lambda \quad (6)$$

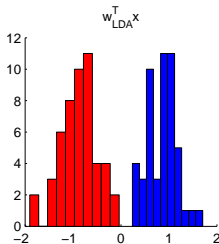
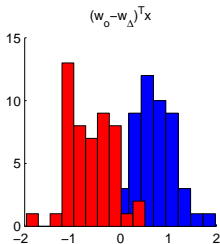
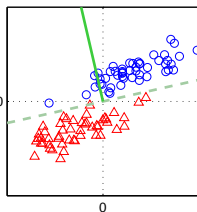
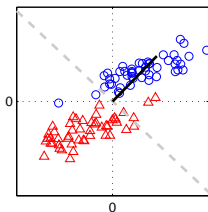
Left multiplying with \mathbf{S}_W^{-1} yields

$$\begin{aligned} \mathbf{S}_W^{-1} \mathbf{S}_B \mathbf{w} &= \mathbf{S}_W^{-1} \mathbf{S}_W \mathbf{w} \lambda \\ \mathbf{S}_W^{-1} (\mu_+ - \mu_-) \underbrace{(\mu_+ - \mu_-)^\top \mathbf{w}}_{\beta} &= \mathbf{w} \quad (7) \\ \mathbf{w} &\propto \mathbf{S}_W^{-1} (\mu_+ - \mu_-) \end{aligned}$$

→ Fisher's LDA first *decorrelates* the data followed by nearest centroid classification



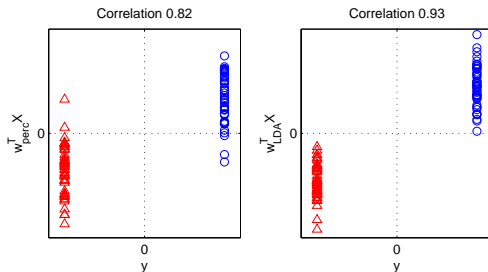
Fisher Linear Discriminant Analysis



Fisher Linear Discriminant Analysis

Another view on LDA:

Maximizing the correlation between class labels y and $w^T X$:



What if our labels are not $\in \{-1, +1\}$ but $\in \mathbb{R}$?



From Classification to Regression

$y \in \{-1, +1\}$	$y \in \mathbb{R}$
Classification	Regression

The most popular and best understood type of regression is **linear regression** using a *least-squares cost function*.



Linear Regression

Target variable $y \in \mathbb{R}$ is modeled as a **linear combination**
 $w \in \mathbb{R}^N$ of N regressors $\phi(\mathbf{x}) \in \mathbb{R}^N$

$$y = \mathbf{w}^\top \phi(\mathbf{x}) \quad (8)$$

where $\phi(\cdot)$ is one or more (potentially non-linear) function on \mathbf{x} .

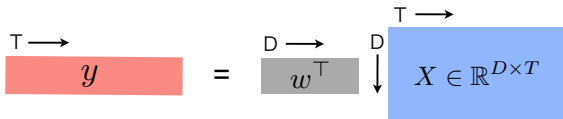
For the sake of simplicity we assume $\phi(\mathbf{x}) = \mathbf{x}$.



Linear Regression

Let T be the number of samples, so $\mathbf{y} \in \mathbb{R}^{1 \times T}$ and $\mathbf{X} \in \mathbb{R}^{N \times T}$.
The Linear Regression model in matrix notation then becomes

$$\mathbf{y} = \mathbf{w}^\top \mathbf{X}. \quad (9)$$



Linear Regression

The most popular loss function to optimize w is the **least-square error** [Gauß, 1809; Legendre, 1805]

$$\mathcal{E}_{lsq}(w) = \sum_{t=1}^T (y_t - \mathbf{w}^\top \mathbf{X}_t)^2 \quad (10)$$



C.F. Gauß (1777-1855)



A.M. Legendre (1752-1833)



Linear Regression

To minimize the least-squares loss function in eq. 10

$$\mathcal{E}_{lsq}(\mathbf{w}) = (\mathbf{y} - \mathbf{w}^\top \mathbf{X})^2 = \mathbf{y}\mathbf{y}^\top - 2\mathbf{w}^\top \mathbf{X}\mathbf{y}^\top + \mathbf{w}^\top \mathbf{X}\mathbf{X}^\top \mathbf{w}$$

we compute derivative w.r.t. \mathbf{w}



Linear Regression

To minimize the least-squares loss function in eq. 10

$$\mathcal{E}_{lsq}(\mathbf{w}) = (\mathbf{y} - \mathbf{w}^\top \mathbf{X})^2 = \mathbf{y}\mathbf{y}^\top - 2\mathbf{w}^\top \mathbf{X}\mathbf{y}^\top + \mathbf{w}^\top \mathbf{X}\mathbf{X}^\top \mathbf{w}$$

we compute derivative w.r.t. \mathbf{w}

$$\frac{\partial \mathcal{E}_{lsq}(\mathbf{w})}{\partial \mathbf{w}} = -2\mathbf{X}\mathbf{y}^\top + 2\mathbf{X}\mathbf{X}^\top \mathbf{w} \quad (11)$$



Linear Regression

To minimize the least-squares loss function in eq. 10

$$\mathcal{E}_{lsq}(\mathbf{w}) = (\mathbf{y} - \mathbf{w}^\top \mathbf{X})^2 = \mathbf{y}\mathbf{y}^\top - 2\mathbf{w}^\top \mathbf{X}\mathbf{y}^\top + \mathbf{w}^\top \mathbf{X}\mathbf{X}^\top \mathbf{w}$$

we compute derivative w.r.t. \mathbf{w}

$$\frac{\partial \mathcal{E}_{lsq}(\mathbf{w})}{\partial \mathbf{w}} = -2\mathbf{X}\mathbf{y}^\top + 2\mathbf{X}\mathbf{X}^\top \mathbf{w} \quad (11)$$

set it to zero and solve for \mathbf{w}

$$-2\mathbf{X}\mathbf{y}^\top + 2\mathbf{X}\mathbf{X}^\top \mathbf{w} = 0$$

$$\mathbf{X}\mathbf{X}^\top \mathbf{w} = \mathbf{X}\mathbf{y}^\top$$

$$\mathbf{w} = (\underbrace{\mathbf{X}\mathbf{X}^\top}_{\text{Cov. Mat.}})^{-1} \mathbf{X}\mathbf{y}^\top \quad (12)$$



Linear Regression for Vector Labels

Prediction of vector-valued labels $\mathbf{y} \in \mathbb{R}^M$
is called **Multiple Linear Regression**:

For a measurement $\mathbf{X} \in \mathbb{R}^{N \times T}$, $\mathbf{Y} \in \mathbb{R}^{M \times T}$ the MLR model is

$$\mathbf{Y} = \mathbf{W}^T \mathbf{X} \quad (13)$$

where $\mathbf{W}^T \in \mathbb{R}^{M \times N}$ is a **linear mapping** from data to labels.



Linear Regression for Vector Labels

Given Data $\mathbf{X} \in \mathbb{R}^{N \times T}$ and labels $\mathbf{Y} \in \mathbb{R}^{M \times T}$
the error function for multiple linear regression is

$$\mathcal{E}_{MLR}(\mathbf{W}) = \sum_{m=1}^M (\mathbf{Y}_m - \mathbf{W}_m^T \mathbf{X})^2 \quad (14)$$

where \mathbf{Y}_m denotes the m -th output dimension
and \mathbf{W}_m the corresponding weight vector



Linear Regression for Vector Labels

Given Data $\mathbf{X} \in \mathbb{R}^{N \times T}$ and labels $\mathbf{Y} \in \mathbb{R}^{M \times T}$
the error function for multiple linear regression is

$$\mathcal{E}_{MLR}(\mathbf{W}) = \sum_{m=1}^M (\mathbf{Y}_m - \mathbf{W}_m^\top \mathbf{X})^2 \quad (14)$$

where \mathbf{Y}_m denotes the m -th output dimension
and \mathbf{W}_m the corresponding weight vector

Eq. 14 is minimized by

$$\mathbf{W} = (\mathbf{X}\mathbf{X}^\top)^{-1}\mathbf{X}\mathbf{Y}^\top \quad (15)$$



Linear Discriminant Analysis and Linear Regression

Remember the solution to linear discriminant analysis

$$\mathbf{w}_{LDA} \propto \mathbf{S}^{-1}(\boldsymbol{\mu}_+ - \boldsymbol{\mu}_-)$$



Linear Discriminant Analysis and Linear Regression

Remember the solution to linear discriminant analysis

$$\begin{aligned}\mathbf{w}_{LDA} &\propto \mathbf{S}^{-1}(\boldsymbol{\mu}_+ - \boldsymbol{\mu}_-) \\ &\propto \mathbf{S}^{-1}(\underbrace{(+1)}_y \underbrace{(1/N_{+1})}_{\gamma_1} \sum_{i \in \mathcal{Y}_{+1}} \mathbf{x}_i + \underbrace{(-1)}_y \underbrace{(1/N_{-1})}_{\gamma_2} \sum_{j \in \mathcal{Y}_{-1}} \mathbf{x}_j)\end{aligned}$$



Linear Discriminant Analysis and Linear Regression

Remember the solution to linear discriminant analysis

$$\begin{aligned}\mathbf{w}_{LDA} &\propto \mathbf{S}^{-1}(\boldsymbol{\mu}_+ - \boldsymbol{\mu}_-) \\ &\propto \mathbf{S}^{-1}(\underbrace{(+1)}_y \underbrace{(1/N_{+1})}_{\gamma_1} \sum_{i \in \mathcal{Y}_{+1}} \mathbf{x}_i + \underbrace{(-1)}_y \underbrace{(1/N_{-1})}_{\gamma_2} \sum_{j \in \mathcal{Y}_{-1}} \mathbf{x}_j) \\ &\propto \mathbf{S}^{-1} \mathbf{X} \mathbf{y}^\top \text{ assuming } N_{+1} = N_{-1}\end{aligned}$$



Linear Discriminant Analysis and Linear Regression

Remember the solution to linear discriminant analysis

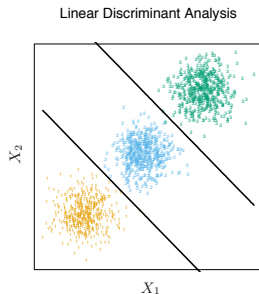
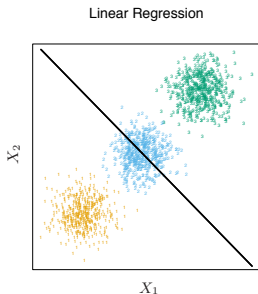
$$\begin{aligned}
 \mathbf{w}_{LDA} &\propto \mathbf{S}^{-1}(\boldsymbol{\mu}_+ - \boldsymbol{\mu}_-) \\
 &\propto \mathbf{S}^{-1} \left(\underbrace{(+1)}_y \underbrace{(1/N_{+1})}_{\gamma_1} \sum_{i \in \mathcal{Y}_{+1}} \mathbf{x}_i + \underbrace{(-1)}_y \underbrace{(1/N_{-1})}_{\gamma_2} \sum_{j \in \mathcal{Y}_{-1}} \mathbf{x}_j \right) \\
 &\propto \mathbf{S}^{-1} \mathbf{X} \mathbf{y}^\top \text{ assuming } N_{+1} = N_{-1}
 \end{aligned}$$

LDA	Linear Regression
$\mathbf{w} \propto \mathbf{S}^{-1} \mathbf{X} \mathbf{y}^\top$	$\mathbf{w} = (\mathbf{X} \mathbf{X}^\top)^{-1} \mathbf{X} \mathbf{y}^\top$



Classification by Linear Regression?

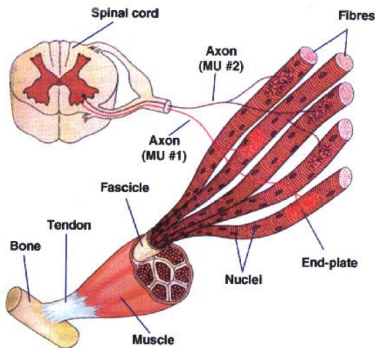
So when coding the class as a boolean multivariate label vector, can we do classification with linear regression?



No, for more than 2 classes, this can lead to poor classification



Application Example: Myoelectric Control of Prostheses



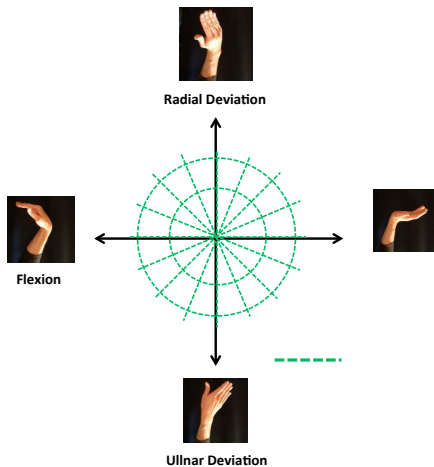
Neurons activate muscles via electric discharges
Electric activity can be measured non-invasively



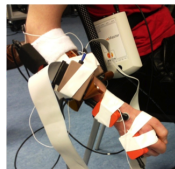
State-of-the-art hand prosthesis
Only 2 degrees of freedom are controlled
(open/close, rotate)
Controlled by muscle activity



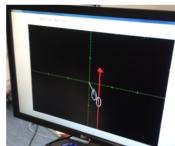
Acquisition of Training Data



Experimental Paradigm

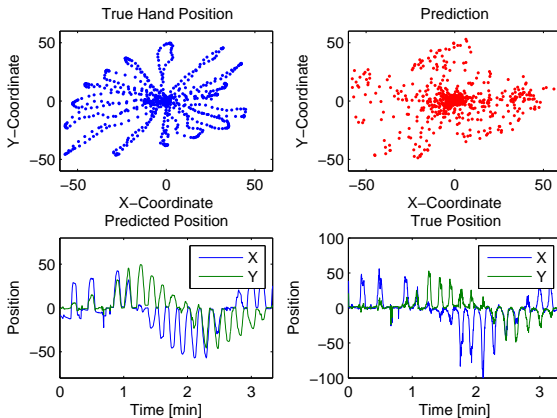


Motion Capture System

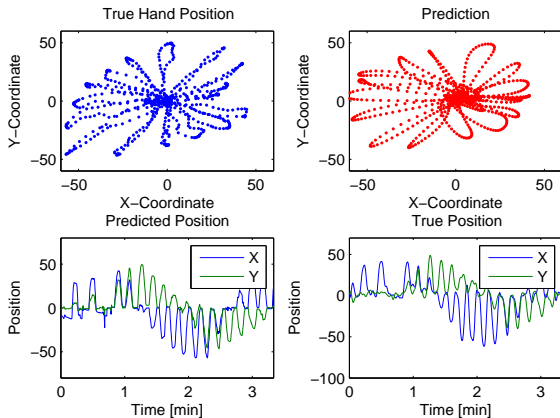


Visual Feedback

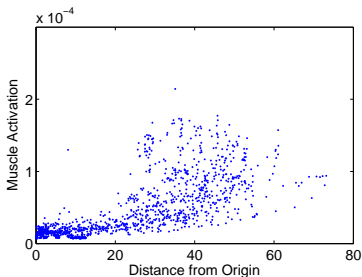
Results Linear Regression



Results Linear Regression - Smoothed



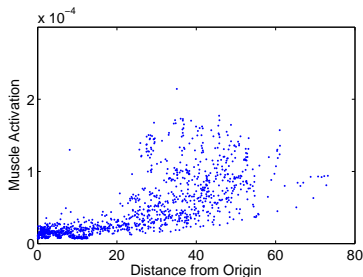
Linear Regression



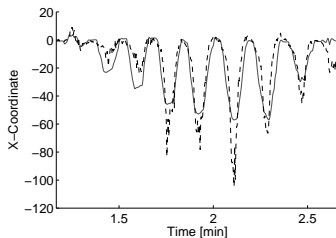
Hand position is a *non-linear* function of muscle activation



Linear Regression



Hand position is a *non-linear* function of muscle activation

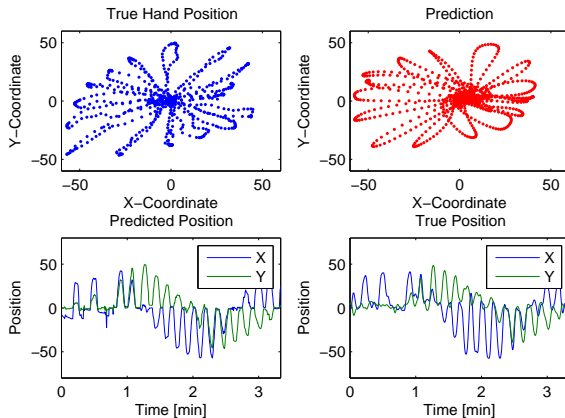


Weak muscle activation
→ True Handposition (dashed)
overestimated (grey)

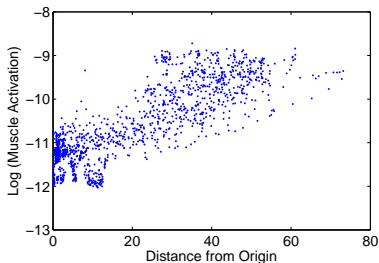
Strong muscle activation
→ True Handposition **underestimated**



Results Linear Regression - Smoothed and Log Features



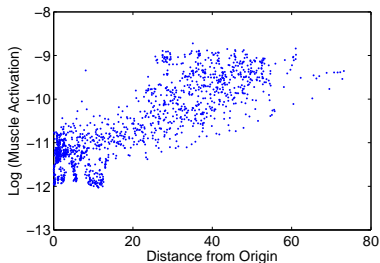
Linear(ized) Regression



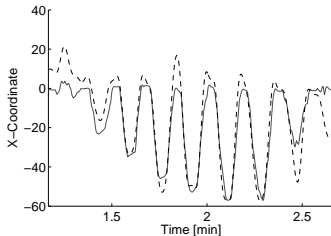
Hand position is *almost linearly* related to log of muscle activation



Linear(ized) Regression



Hand position is *almost linearly* related to log of muscle activation



Weak muscle activation
→ Handposition *less underestimated*

Strong muscle activation
→ Handposition *less overestimated*



Regularization

Often it is important to **control the complexity** the solution w .

This is done by constraining the \mathcal{L}_p -norm of w .

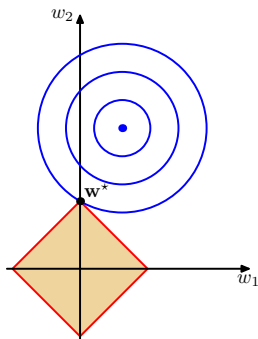
So we add to the least-squares error minimization the constraint

$$||\mathbf{w}||_p = \lambda \quad (16)$$

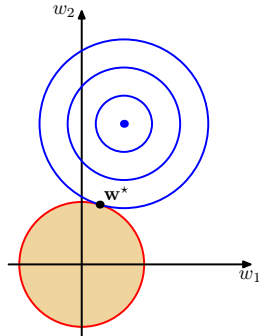


Regularization

What does this mean geometrically?



$$\|\mathbf{w}\|_1 = \lambda$$



$$\|\mathbf{w}\|_2 = \lambda$$

The least squares error $\mathcal{E}(\mathbf{w})$ is the same on the blue circles

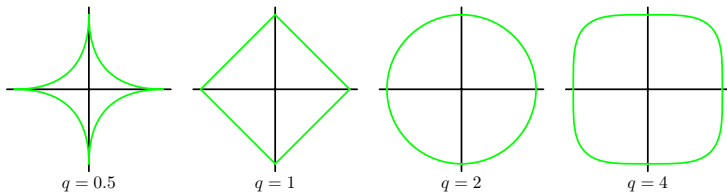
The \mathcal{L}_p -norm of \mathbf{w} is indicated by the red line

The optimal \mathbf{w} is on the intersection of the constraint and the error



Other Regularizers

Other norms than \mathcal{L}_1 -norm and \mathcal{L}_2 -norm are



Why \mathcal{L}_1 -norm and \mathcal{L}_2 -norm?

- The \mathcal{L}_2 -norm is analytically tractable:

$$\frac{\partial \|\mathbf{w}\|_2^2}{\partial \mathbf{w}} = \frac{\partial \sqrt{\mathbf{w}^\top \mathbf{w}}^2}{\partial \mathbf{w}} = 2\mathbf{w}$$

→ Very popular and known as Tikhonov Regularization, Weight Decay, Shrinkage, Ridge Regression, ...

- The \mathcal{L}_1 -norm is not differentiable (at 0)
 - But it has the nice property of leading to **sparse solutions**
- Known as: Lasso, Sparse [insert any method here], ...



\mathcal{L}_2 -norm Regularization: Ridge Regression

For the euklidian norm $p = 2$ the error function is then

$$\mathcal{E}_{RR}(\mathbf{w}) = (\mathbf{y} - \mathbf{w}^\top \mathbf{X})^2 + \lambda \|\mathbf{w}\|_2^2 \quad (17)$$



Ridge Regression

Computing the derivative w.r.t. \mathbf{w} yields

$$\frac{\partial \mathcal{E}_{RR}(\mathbf{w})}{\partial \mathbf{w}} = -2\mathbf{xy}^\top + 2\mathbf{XX}^\top \mathbf{w} + \lambda 2\mathbf{w}. \quad (18)$$

Setting eq. 18 to zero and rearranging terms the optimal \mathbf{w} is

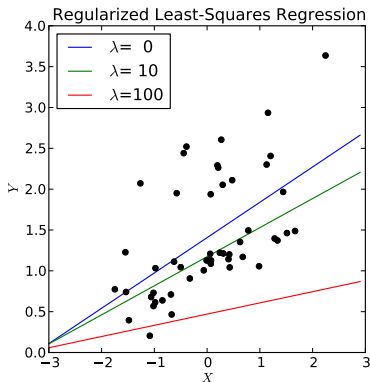
$$\begin{aligned} 2\mathbf{XX}^\top \mathbf{w} + \lambda 2\mathbf{w} &= 2\mathbf{Xy}^\top \\ (\mathbf{XX}^\top + \lambda \mathbf{I})\mathbf{w} &= \mathbf{Xy}^\top \\ \mathbf{w} &= (\mathbf{XX}^\top + \lambda \mathbf{I})^{-1} \mathbf{Xy}^\top \end{aligned} \quad (19)$$

[Hoerl and Kennar, 1970; Tychonoff, 1943]



Effect of λ on \mathbf{w}

Increasing λ will *shrink* coefficients of \mathbf{w} to zero



(Multi-)Linear (Ridge) Regression Algorithm

Algorithm 1 (Multi-)Linear (Ridge) Regression

Require: $\mathbf{x}_i \in \mathbb{R}^U$, $\mathbf{y}_i \in \mathbb{R}^V$, ridge λ

Ensure: Weight matrix \mathbf{W} for linear mapping of $\mathbb{R}^U \rightarrow \mathbb{R}^V$

1: Include offset parameters (row vector of N ones)

2: $\mathbf{X} = \begin{bmatrix} \mathbf{1} \\ \mathbf{X} \end{bmatrix}$

3: $\mathbf{W} = (\mathbf{X}\mathbf{X}^\top + \lambda\mathbf{I})^{-1}\mathbf{X}\mathbf{Y}^\top$



\mathcal{L}_1 -norm Regularization: The Lasso

Choosing the norm $p = 1$ for regression
is called the **Lasso** [Tibshirani, 1996]

The error function is

$$\mathcal{E}_{RR}(\mathbf{w}) = (\mathbf{y} - \mathbf{w}^\top \mathbf{X})^2 + \lambda \|\mathbf{w}\|_1 \quad (20)$$

There is no closed form solution for this.



Lasso with First Order Stochastic Gradient Descent

Algorithm 2 Lasso

Require: $\mathbf{x}_1, \dots, \mathbf{x}_N \in \mathbb{R}^D$, $\mathbf{y}_1, \dots, \mathbf{y}_N \in \mathbb{R}^1$, maximum \mathcal{L}_1 norm λ , step size η

Ensure: Weight vector \mathbf{w} for linear mapping of $\mathbb{R}^D \rightarrow \mathbb{R}^1$

```
1: Include offset parameters (row vector of  $N$  ones)
2:  $\mathbf{X} = \begin{bmatrix} \mathbf{1} \\ \mathbf{X} \end{bmatrix}$ 
3: # Initialize  $\mathbf{w} = \mathbf{0}/D$ 
4: for  $i = 1$  to  $N_{it}$  do
5:   # Draw random data point  $\mathbf{x}_i$  and label  $y_i$ 
6:    $\mathbf{v} = \mathbf{w}_-$ 
7:    $\mathbf{u} = \mathbf{w}_+$ 
8:    $v_d \leftarrow \max(v_d - \eta/i(\lambda + (y_i - \mathbf{w}^\top \mathbf{x}_i)\mathbf{x}_{i,d}), 0)$ 
9:    $u_d \leftarrow \max(u_d - \eta/i(\lambda - (y_i - \mathbf{w}^\top \mathbf{x}_i)\mathbf{x}_{i,d}), 0)$ 
10:   $\mathbf{w} = \mathbf{u} - \mathbf{v}$ 
11: end for
```

taken from [Bottou, 2010]



Random Kitchen Sinks

- What if the dependencies between features $\phi(\mathbf{x})$ and labels y are non-linear and we don't know the non-linearity?
- We can use algorithms that learn a non-linear function
 - Multilayer Perceptrons
 - Kernel Methods (next lectures)
- Use Linear Regression and **random basis functions** [Rahimi and Recht, 2008]
- This trick is called *Random Kitchen Sinks*

<http://www.keysduplicated.com/~ali/random-features/>



Random Kitchen Sinks

Algorithm 3 Random Kitchen Sinks

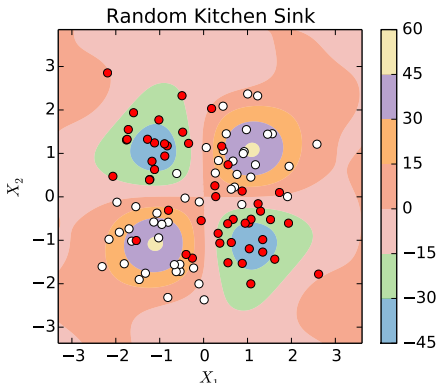
Require: $\mathbf{x}_i, \dots, \mathbf{x}_N \in \mathbb{R}^D$, $\mathbf{y}_i, \dots, \mathbf{y}_N \in \mathbb{R}^1$, regularizer λ , number of features F

Ensure: Weight vector \mathbf{a} , random basis \mathbf{W}

- 1: # Training
 - 2: $\mathbf{W} \in \mathbb{R}^{D \times F} \sim \mathcal{N}(0, 1)$
 - 3: $\mathbf{Z} = e^{i\mathbf{W}^\top \mathbf{x}}$
 - 4: $\mathbf{a} = (\mathbf{I}\lambda + \mathbf{Z}\mathbf{Z}^\top)^{-1} \mathbf{Z}\mathbf{y}^\top$
 - 5: # Testing
 - 6: $\hat{\mathbf{y}} = \mathbf{a}^\top e^{i\mathbf{W}^\top \mathbf{x}}$
-



Random Projections for Solving XOR



Summary

Linear Regression

- is a generic framework for prediction

- straightforwardly extends to vector labels

- can be made more robust by constraining $\|\mathbf{w}\|_p$

 - $p = 1$: Sparse solutions, no closed form solution

 - $p = 2$: Analytically tractable

- Non-linear dependencies:

 - Explicitly model non-linearity (if possible)

 - Random Projections



References

- C. M. Bishop. *Pattern Recognition and Machine Learning (Information Science and Statistics)*. Springer US, 2007.
- L. Bottou. Large-scale machine learning with stochastic gradient descent. In Y. Lechevallier and G. Saporta, editors, *Proceedings of the 19th International Conference on Computational Statistics (COMPSTAT'2010)*, pages 177–187, Paris, France, 2010. Springer.
- C. F. Gauß. *Theoria motus corporum coelestium in sectionibus conicis solem ambientium*. Göttingen, 1809.
- T. Hastie, R. Tibshirani, and J. H. Friedman. *The Elements of Statistical Learning*. 2003.
- A. E. Hoerl and R. W. Kennar. Ridge regression: Applications to nonorthogonal problems. *Technometrics*, 12(1):69–82, 1970.
- A.-M. Legendre. *Nouvelles méthodes pour la détermination des orbites des comètes*, chapter Sur la methode des moindres quarrés. Firmin Didot, <http://imgbase-scd-ulp.u-strasbg.fr/displayimage.php?pos=-141297>, 1805.
- K. P. Murphy. *Machine Learning: A Probabilistic Perspective*. Adaptive Computation and Machine Learning. The MIT Press, 1 edition, 2012. ISBN 0262018020, 9780262018029.
- A. Rahimi and B. Recht. Weighted Sums of Random Kitchen Sinks: Replacing minimization with randomization in learning. In D. Koller, D. Schuurmans, Y. Bengio, L. Bottou, D. Koller, D. Schuurmans, Y. Bengio, and L. Bottou, editors, *NIPS*, pages 1313–1320. MIT Press, 2008.
- R. Tibshirani. Regression shrinkage and selection via the lasso. *Journal of the Royal Statistical Society, Series B*, 58:267–288, 1996.
- A. N. Tychonoff. On the stability of inverse problems. *Doklady Akademii Nauk SSSR*, 39(5):195–198, 1943.

