

For each of the following problems, provide your answer and show the steps taken to solve the problem.

Problem 1. Maximum Likelihood Estimation (50 points) Given a dataset  $\{x_1, x_2, \dots, x_N\}$  of size  $N$ , derive the maximum likelihood estimate (as a function of  $x_1, \dots, x_N$ ) for: (a) The lower and upper limits,  $a$  and  $b$ , of a uniform distribution,

$$f(x; a, b) = \{$$

$$\begin{aligned} & \frac{1}{b-a}, & \text{if } a \leq x \leq b \\ & 0, & \text{otherwise} \end{aligned}$$

1.

(assuming each  $x_i \in \mathbb{R}$ ). Show all of your work. (25 points)

to find the upper limit of a uniform distribution we have:

$$\begin{aligned} b_{MLE} &= \operatorname{argmax}_b P_b(x_1, x_2, \dots, x_N) \\ &= \operatorname{argmax}_b \prod_{i=1}^N \frac{1}{b-a} \\ &= \operatorname{argmax}_b \left( \frac{1}{b-a} \right)^N \end{aligned}$$

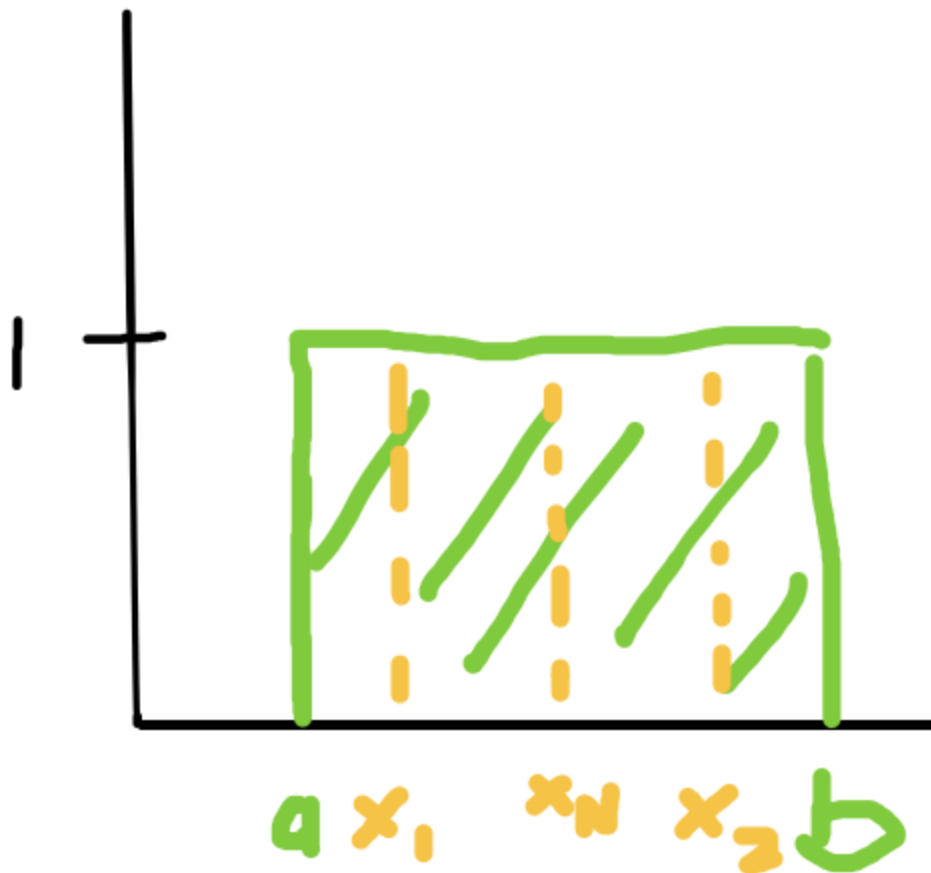
From this expression we can see that the larger  $b$ , the larger the maximum likelihood estimation. Therefore, we can conclude that  $b = \max(x_{\{i\}})$

Similarly for the lower limit,

$$\begin{aligned} a_{MLE} &= \operatorname{argmax}_a P_a(x_1, x_2, \dots, x_N) \\ &= \operatorname{argmax}_a \prod_{i=1}^N \frac{1}{b-a} \\ &= \operatorname{argmax}_a \left( \frac{1}{b-a} \right)^N \end{aligned}$$

The smaller  $a$  the larger the MLE, Therefore, we can conclude that  $a = \min(x_i)$

To test this, we can draw the uniform distribution and some  $x$  values:



uniform distribution

If some value of  $x$  was greater than  $b$ , the probability would be  $0$  and if some value of  $x$  was less than  $a$ , the probability would also be  $0$ . Therefore to maximize the likelihood, we want all values of  $x$  to fall within the range  $a$  and  $b$ .

b. The  $\lambda$  parameter of a Poisson distribution,

$$f(x; \lambda) = \begin{cases} \frac{e^{-\lambda} \lambda^x}{x!}, & \text{if } x \geq 0 \\ 0, & \text{if } x < 0 \end{cases}$$

(assuming each  $x_i \geq 0$ ). Show all of your work. (25 points) Hints: (i) plotting some sample data may be helpful and calculus should not be required (a); (ii) maximizing the log likelihood provides the same parameter values and often provides a simpler path to a solution (b); (iii)  $\log(ab) = \log a + \log b$ ; (iv)  $\log e^a = a$ .

$$l(\lambda|x) = p(x|\lambda) = \prod_{i=1}^N p(x_i|\lambda) = \prod_{i=1}^N e^{-\lambda} \frac{\lambda^{x_i}}{x_i!}$$

Then, take the log likelihood

$$L(\lambda|x) = \log l(\lambda|x) = \sum_{i=1}^N \log p(x_i|\lambda)$$

$$= \sum_{i=1}^N \log e^{-\lambda} \frac{\lambda^{x_i}}{x_i!}$$

$$= \sum_{i=1}^N \log e^{-\lambda} + \log \lambda^{x_i} - \log x_i!$$

$$= -\lambda \sum_{i=1}^N 1 + x \log \lambda - \log x_i!$$

$$MLE = \lambda^* = \operatorname{argmax}_{\lambda} L(\lambda|x)$$

$$= -\lambda \sum_{i=1}^N 1 + x \log \lambda - \log x_i!$$

Problem 2. Bayesian Parameter Estimation (50 points) The density function of an exponential distribution is given by  $f_{\lambda}(x) = \lambda e^{-\lambda x}$ . The MLE for the parameter  $\lambda$  can be calculated as  $\lambda = \frac{n}{\sum_i x_i}$ .

We will now consider Bayesian parameter estimation for this distribution.

- a. Using a prior distribution from the Gamma distribution,  $f_{\alpha,\beta}(\lambda) = \frac{\beta^{\alpha} \lambda^{\alpha-1} e^{-\lambda\beta}}{\Gamma(\alpha)}$ , with parameters  $\alpha$  and  $\beta$ , show that the posterior distribution for  $\lambda$ , after updating using three datapoints  $x_1, x_2, x_3$ , is also a Gamma distribution and show its new parameter values,  $\alpha_0$  and  $\beta_0$ , in terms of  $\alpha, \beta, x_1, x_2$ , and  $x_3$ . (25 points)

$$p(x_1, x_2, x_3 | \lambda) = \lambda^3 e^{-\lambda(x_1+x_2+x_3)}$$

$$p(\lambda | x_1, x_2, x_3) \propto p(x_1, x_2, x_3 | \lambda) \cdot p(\lambda)$$

substituting the prior and likelihood functions we get:

$$p(\lambda | x_1, x_2, x_3) \propto \frac{\beta^{\alpha}}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\lambda\beta} \cdot \lambda^3 e^{-\lambda(x_1+x_2+x_3)}$$

which simplifies to :

$$p(\lambda | x_1, x_2, x_3) \propto \frac{\beta^{\alpha}}{\Gamma(\alpha)} \lambda^{\alpha+2-1} e^{-\lambda(\beta+x_1+x_2+x_3)}$$

and we get the parameter values as:

$$\alpha_0 = \alpha + 2$$

$$\beta_0 = \beta + x_1 + x_2 + x_3$$

b. If our prior parameters are  $\alpha = 2$  and  $\beta = 1$ , and our data sample consists of  $x_1 = 3.7$ ,  $x_2 = 4.5$ ,  $x_3 = 4.8$ : Compute the posterior probability of a new datapoint  $x_4 = 3.8$  under the fully Bayesian estimation of  $\lambda$ . You can either leave your answer in terms of the Gamma function, or provide the exact answer. (13 points) Hints: (i) You shouldn't have to solve the complicated integral; (ii) Since the Gamma distribution normalizes to 1,  $\int_0^\infty \lambda^{\alpha-1} e^{-\lambda\beta} d\lambda = \frac{\Gamma(\alpha)}{\beta^\alpha}$ ; (iii) The Gamma function is related to the factorial function as  $\Gamma(x) = (x-1)!$  for positive integers  $x$ .

$$\alpha_0 = 2 + 2 = 4$$

$$\beta_0 = 1 + 3.7 + 4.5 + 4.8 = 14$$

To compute  $x_4$ , plug  $\alpha_0$ ,  $\beta_0$  and  $x_4$  into the density function

$$f(\lambda \mid \alpha_0, \beta_0) = \frac{\Gamma(\alpha_0)}{\beta_0^{\alpha_0}} \lambda^{\alpha_0-1} e^{-\beta_0 \lambda}$$

so we get:

$$f(\lambda \mid 4, 14) = \frac{\Gamma(4)}{14^4} \cdot (3.8)^{4-1} e^{-14 \cdot 3.8}$$

- c. If we have the same prior and datapoints as in (b), what is the probability of a new datapoint  $x_4 = 3.8$  using maximum a posteriori estimation of  $\lambda$ ? (12 points) Hint: (i) The mode of the Gamma distribution (i.e., the  $\lambda$  that attains its maximal probability) is  $\alpha - 1$