

# On the Asymptotic Validity of a Fully Sequential Elimination Procedure for Indifference-Zone Ranking and Selection with Tight Bounds on Probability of Correct Selection

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We prove the validity of the sequential elimination IZ procedure proposed by Frazier [1] when  $\delta$  goes to 0.

## Asymptotic Validity when the Variances are Known

Without loss of generality, suppose that the true means of the systems are indexed so that  $\mu_k \geq \mu_{k-1} \geq \dots \geq \mu_1$ . We suppose that samples from system  $x \in \{1, \dots, k\}$  are normally distributed and independent, over time and across alternatives. We also suppose that the algorithm ends in at most  $R(\delta) \in \mathbb{N}$  iterations, and  $R(\delta) \rightarrow \infty$  as  $\delta \rightarrow 0$  with probability 1. Furthermore,  $R^{1/2}\delta \rightarrow \Delta$  with probability 1 where  $\infty > \Delta > 0$  with probability 1. We also define  $\lambda_z^2 := \max_{i \in \{1, \dots, k\}} \{\lambda_i^2\}$ .

**Lemma 1.** If  $x \in \{1, \dots, k\}$  and  $t \in [0, 1]$ , then

$$C_x(\delta, t) := \frac{Y_{\text{ceil}\left(\frac{\lambda_x^2}{\lambda_z^2}(n_0+tR)\right),x} - \frac{\lambda_x^2}{\lambda_z^2}(n_0+tR)\mu_x}{\frac{\lambda_x^2}{\lambda_z^2}\sqrt{R}} \Rightarrow W_x(t)$$

where  $Y_{n,x} = \sum_{i=1}^n X_{xi}$  and  $W_x$  is a standard Brownian motion.

**Proof.** By the Lindeberg-Lévy central limit theorem and the fact that  $\frac{ceil(nt)}{n} \rightarrow t$ ,

$$\frac{Y_{ceil(nt),x} - ceil(nt)(\mu_x)}{\lambda_x \sqrt{n}} \Rightarrow W_x(t).$$

Consequently,

$$\frac{Y_{ceil(nt),x} - nt(\mu_x)}{\lambda_x \sqrt{n}} \Rightarrow W_x(t)$$

because  $\frac{nt - ceil(nt)}{\lambda_x \sqrt{n}} \rightarrow 0$ . Thus

$$\frac{Y_{ceil\left(\frac{\lambda_x^2}{\lambda_z^2}tR\right),x} - \frac{\lambda_x^2}{\lambda_z^2}tR(\mu_x)}{\frac{\lambda_x^2}{\lambda_z^2}\sqrt{R}} \Rightarrow W_x(t)$$

as  $\delta \rightarrow 0$ , and then

$$\frac{Y_{ceil\left(\frac{\lambda_x^2}{\lambda_z^2}(n_0+tR)\right),x} - \frac{\lambda_x^2}{\lambda_z^2}(n_0+tR)\mu_x}{\frac{\lambda_x^2}{\lambda_z^2}\sqrt{R}} \Rightarrow W_x(t).$$

■

**Proposition 1.** We have that for  $t$  sufficiently large

$$\begin{aligned} \frac{W_{tx}}{n_{tx}} &= \frac{\lambda_x^2}{\lambda_z} \sqrt{R} \frac{\left(C_x\left(\delta, \frac{t}{R}\right) + \frac{\lambda_x^2}{\lambda_z^2}(n_0+tR)\mu_x\right)}{\frac{\lambda_x^2}{\lambda_z^2}(n_0+t)} \\ &= \lambda_z \sqrt{R} \frac{\left(C_x\left(\delta, \frac{t}{R}\right) + \frac{\lambda_x^2}{\lambda_z^2}(n_0+t)\mu_x\right)}{(n_0+t)}. \end{aligned}$$

Now, we consider the set  $D[0,1]^k$  which is the set of functions from  $[0,1]$  to  $\mathbb{R}^k$  that are right-continuous and have left-hand limits. We'll use the Skorokhod metric  $d$  on  $D[0,1]^k$ :

$$d(X, Y) = \inf_{\lambda \in \Lambda} \{\|\lambda - I\| \vee \|X - Y \circ \lambda\|\}$$

where  $\Lambda$  is the set of strictly increasing, continuous mappings of  $[0,1]$  onto itself, and  $\|\cdot\|$  is the uniform norm.

**Lemma 2.** Suppose  $\Delta$  is given. Let  $Y \in D[0, 1]^k$  and  $t \in [0, 1]$

$$q_{tx}^{Y,\delta}(A) : = \exp\left(\delta\beta_{tR}\lambda_z\sqrt{R}\frac{Y_x(t)}{(n_0+tR)}\right) / \sum_{x' \in A} \exp\left(\delta\beta_{tR}\lambda_z\sqrt{R}\frac{Y_{x'}(t)}{(n_0+tR)} + \beta_{tR}\delta^2 I_{x',x}\right)$$

where  $\beta_{tR} = \frac{(n_0+tR)}{\lambda_z^2}$  and

$$I_{x',x} = \begin{cases} 0 & \text{if } x' = x \\ 1 & \text{if } x' > x \\ -1 & \text{if } x' < x \end{cases}$$

We then define

$$\begin{aligned} T_{Y,\delta}^0 &= 0 \\ A_0^{Y,\delta} &= \{1, \dots, k\} \\ P_0^{Y,\delta} &= P^* \\ T_{Y,\delta}^{n+1} &= \inf\left\{t \in [T_{Y,\delta}^n, 1] : \min_{x \in A_n^{Y,\delta}} q_{tx}^{Y,\delta}(A_n^{Y,\delta}) \leq c \text{ or } \max_{x \in A_n^{Y,\delta}} q_{tx}^{Y,\delta}(A_n^{Y,\delta}) \geq P_n^{Y,\delta}\right\} \\ Z_{n+1}^{Y,\delta} &\in \arg \min_{x \in A_n^{Y,\delta}} q_{T_{Y,\delta}^{n+1},x}^{Y,\delta}(A_n^{Y,\delta}) \\ A_{n+1}^{Y,\delta} &= A_n - \{Z_{n+1}^{Y,\delta}\} \\ P_{n+1}^{Y,\delta} &= P_n^{Y,\delta} / \left(1 - \min_{x \in A_n^{Y,\delta}} q_{T_{Y,\delta}^{n+1},x}^{Y,\delta}(A_n^{Y,\delta})\right). \end{aligned}$$

Now, let

$$M_{Y,\delta} = \inf\left\{n = 1, \dots, k-1 : \max_{x \in A_{n-1}^{Y,\delta}} q_{T_{Y,\delta}^n,x}^{Y,\delta}(A_n^{Y,\delta}) \geq P_{n-1}^{Y,\delta}\right\}$$

and

$$f(Y, \delta) = \begin{cases} 1 & \text{if } k \in \arg \max_{x \in A_{M-1}^{Y,\delta}} \frac{\frac{\lambda_k^2}{\lambda_z} \left(Y_k(T_{Y,\delta}^{M_{Y,\delta}})\right)}{\lambda_z^2 \sqrt{R}} + \delta \frac{\lambda_k^2}{\lambda_z^2 \sqrt{R}} \left(n_0 + T_{Y,\delta}^{M_{Y,\delta}} R\right) \geq \frac{\frac{\lambda_x^2}{\lambda_z} \left(Y_x(T_{Y,\delta}^{M_{Y,\delta}})\right)}{\lambda_z^2 \sqrt{R}} \\ 0 & \text{otherwise} \end{cases}.$$

Now, we also define

$$q_{tx}^Y(A) := \exp\left(\Delta \frac{Y_x(t)}{\lambda_z}\right) / \sum_{x' \in A} \exp\left(\Delta \frac{Y_{x'}(t)}{\lambda_z} + t \frac{\Delta^2}{\lambda_z^2} I_{x',x}\right)$$

$$\begin{aligned} T_Y^0 &= 0 \\ A_0^Y &= \{1, \dots, k\} \\ P_0^Y &= P^* \\ T_Y^{n+1} &= \inf \{t \in [T_Y^n, 1] : \min_{x \in A_n^Y} q_{tx}^Y(A_n^Y) \leq c \text{ or } \max_{x \in A_n^Y} q_{tx}^Y(A_n^Y) \geq P_n^Y\} \\ Z_{n+1}^Y &\in \arg \min_{x \in A_n^Y} q_{T_Y^{n+1},x}^Y(A_n^Y) \\ A_{n+1}^Y &= A_n - \{Z_{n+1}^Y\} \\ P_{n+1}^Y &= P_n^{Y,\delta} / \left(1 - \min_{x \in A_n^Y} q_{T_Y^{n+1},x}^Y(A_n^Y)\right). \end{aligned}$$

Now, let

$$M_Y = \inf \left\{ n = 1, \dots, k-1 : \max_{x \in A_{n-1}^Y} q_{T_Y^n, x}^Y(A_n^Y) \geq P_{n-1}^Y \right\}$$

and

$$g(Y) = \begin{cases} 1 & \text{if } k \in \arg \max_{x \in A_{M-1}^Y} \frac{Y_k(T_Y^{M_Y})}{\lambda_z^2} + \Delta \frac{1}{\lambda_z} T_Y^{M_Y} \geq \frac{\frac{\lambda_x^2}{\lambda_z^2} (Y_x(T_Y^{M_Y}))}{\frac{\lambda_k^2}{\lambda_z^2}} \\ 0 & \text{otherwise} \end{cases}.$$

**Proof.** We have that if  $\hat{\tau}_M$  is the continuous version of  $\frac{\tau_M}{R}$ , then

$$f(C(\delta, t), \delta) = \begin{cases} 1 & \text{if } k \in \arg \max_{x \in A_{M-1}^{Y,\delta}} \frac{\lambda_x^2}{\lambda_z^2} \sqrt{R} \left( C_x(\delta, \hat{\tau}_M) + \frac{\lambda_x^2}{\lambda_z^2} (n_0 + \hat{\tau}_M R) \mu_x \right) \\ 0 & \text{otherwise} \end{cases}.$$

By lemma 1,

$$C(\delta, t) \Rightarrow W(t).$$

**Theorem.** If samples from system  $x \in \{1 \dots, k\}$  are normally distributed and independent, over time and across alternatives, then  $\lim_{\delta \rightarrow 0} Pr \{ \text{BIZ selects } k \} \geq P^*$  provided  $\mu_k \geq \mu_{k-1} + \delta$ . We also suppose that the algorithm ends in at most  $R(\delta) \in \mathbb{N}$  iterations, and  $R(\delta) \rightarrow \infty$  as  $\delta \rightarrow 0$  with probability 1. Furthermore,  $\sqrt{R}\delta \rightarrow \Delta$  with probability 1 where  $\infty > \Delta > 0$  with probability 1. We also suppose  $B_1 = \dots = B_k = 1$ .

**Proof.** Let

$$\begin{aligned} T_k(\delta) &= \min \left\{ t \in \left\{ \frac{\tau_{k-1}}{R}, \dots, 1 \right\} : \max_x q_{tR,x}(A) \geq P \right\}. \\ &= \min \left\{ t \in \left\{ \frac{\tau_{k-1}}{R}, \dots, 1 \right\} : \max_x \exp \left( \delta \beta_{tR} \frac{W_{tR,x}}{n_{tR,x}} \right) \geq P \sum_{x' \in A} \exp \left( \delta \beta_{tR} \frac{W_{tR,x'}}{n_{tR,x'}} \right) \right\}. \end{aligned}$$

Let

$$f(Y, \delta) = \begin{cases} 1 & \text{if } \frac{\lambda_k^2}{\lambda_z n_{tR,k}} \sqrt{R} Y_k(T_{Y,\delta}) + \delta \geq \frac{\lambda_x^2}{\lambda_z n_{tR,x}} \sqrt{R} Y_x(T_{Y,\delta}) \quad \forall x \in A \\ 0 & \text{otherwise} \end{cases}$$

where

$$\begin{aligned} T_{Y,\delta} &= \inf \{ t \in [0, 1] : \exists x \text{ s.t. } \exp \left( \frac{\lambda_x^2}{\lambda_z n_{tR,x}} \sqrt{R} \delta \beta_{tR} Y_x(t) \right) \left( \frac{1}{P} - 1 \right) \geq \\ &\quad \sum_{x' < x} \exp \left( \left( \frac{\lambda_{x'}^2}{\lambda_z n_{tR,x'}} \sqrt{R} Y_{x'}(t) \delta \beta_{tR} - \delta^2 \beta_{tR} \right) \right) + \\ &\quad \sum_{x > x'} \exp \left( \left( \frac{\lambda_{x'}^2}{\lambda_z n_{tR,x'}} \sqrt{R} Y_{x'}(t) \delta \beta_{tR} + \delta^2 \beta_{tR} \right) \right) \} \end{aligned}$$