

# 1 Pseudo-code

## (1+1) EA

1. Choose  $s \in \{0, 1\}^n$  randomly.
2. Produce  $s'$  by flipping each bit of  $s$  with probability  $1/n$ .
3. Replace  $s$  by  $s'$  if  $f(s') \geq f(s)$ .
4. Repeat Steps 2 and 3 forever.

## RLS

1. Choose  $s \in \{0, 1\}^n$  randomly.
2. Produce  $s'$  from  $s$  by flipping **one** randomly chosen bit.
3. Replace  $s$  by  $s'$  if  $f(s') \geq f(s)$ .
4. Repeat Steps 2 and 3 forever.

# 2 Chernoff Bound Initialisation

## 2.1 Expected Ones Initialisation

Given that both algorithms initialise an  $n$ -set of binary string at random specifically  $X_i = \text{Unif}\{0, 1\}$  for  $i = 1, 2, \dots, n$ . Therefore the probability of getting 0 or 1 is the same as

$$P(X_i = 1) = P(X_i = 0) = \frac{1}{2}$$

The expected value of each  $X_i$  is

$$\mathbb{E}[X_i] = 0 \cdot \frac{1}{2} + 1 \cdot \frac{1}{2} = \frac{1}{2}$$

Suppose that  $X = X_1 + X_2 + \dots + X_n$ , then the expected value of  $X$  by using the *Linearity of Expectation* is

$$\mathbb{E}[X] = \sum_{i=1}^n \mathbb{E}[X_i] = \frac{n}{2}$$

If  $X$  is the summation of those values inside the binary string, then there will be expected to have  $\frac{n}{2}$  ones.

## 2.2 Chernoff

$$\forall \delta > 0 : P(X > (1 + \delta) \cdot \mathbb{E}[X]) < \left( \frac{e^\delta}{(1 + \delta)^{1+\delta}} \right)^{\mathbb{E}[X]} \quad (\star)$$

$$\forall 0 < \delta < 1 : P(X < (1 - \delta) \cdot \mathbb{E}[X]) < e^{-\mathbb{E}[X]\delta^2/2}$$

From the previous section, it's known that  $\mathbb{E}[X] = \frac{n}{2}$ , choose inequality  $(\star)$  and then substitute.

$$\begin{aligned} P\left(X > (1 + \delta) \cdot \frac{n}{2}\right) &< \left(\frac{e^\delta}{(1 + \delta)^{1+\delta}}\right)^{\frac{n}{2}} \\ \Rightarrow P\left(X > \left(\frac{1}{2} + \frac{\delta}{2}\right) \cdot n\right) &< \dots \end{aligned}$$

Let  $2\delta' = \delta$ , which is always inside the bound of  $\delta > 0$ . Now choose the domain of  $\delta$  as  $0 < \delta \leq 1$ . Therefore, the new domain will be  $0 < 2\delta' \leq 1 \Rightarrow 0 < \delta' \leq \frac{1}{2}$ , thus

$$\begin{aligned} P\left(X > \left(\frac{1}{2} + \delta'\right) \cdot n\right) &< \left(\frac{e^{2\delta'}}{(1+2\delta')^{1+2\delta'}}\right)^{\frac{n}{2}} \\ &\Rightarrow \dots < \rho^{\frac{n}{2}} \\ &\Rightarrow \dots < e^{\frac{n}{2} \ln \rho} \end{aligned}$$

Evaluate  $\rho = \frac{e^{2\delta'}}{(1+2\delta')^{1+2\delta'}}$

$$\ln \rho = 2\delta' - (1 + 2\delta') \ln(1 + 2\delta') < 0, \quad \text{Negative constant}$$

Therefore,

$$P\left(X > \left(\frac{1}{2} + \delta'\right) \cdot n\right) < e^{-\Omega(n)}$$

Then,

$$\begin{aligned} P\left(X > \frac{2n}{3}\right) &< e^{-\Omega(n)} \\ \Rightarrow 1 - P\left(X > \frac{2n}{3}\right) &\leq 1 - e^{-\Omega(n)} \\ \Rightarrow P\left(X \leq \frac{2n}{3}\right) &\leq 1 - e^{-\Omega(n)} \end{aligned}$$

This means that both algorithms with uniform random binary string initialisation will have a very high chance of having at most  $2n/3$  one-bits inside the initialisation. It's so high that with large  $n$  or large binary space it's almost guaranteed.

### 3 RLS Lower Bound

Given from the initialisation that there are expected to have  $2n/3$  one-bits, thus  $m = n/3$  zero-bits needed to be flipped. Each zero-bit has a chance of being flipped is  $1/n$ . We want to know what is the lower bound  $\Omega$  that will let flipping all distinct 0-bits at least once. This is a Collector's Coupon Theorem.

#### 3.1 Collector's Coupon Theorem

Suppose that we have  $n$  distinct cards and we draw each everytime with *replacement*. How many times do I expect to draw to collect all of those cards. This implies a similarity as our problem with RLS lower bound.

From [Wikipedia](#), let  $T$  be the total number of times to finish flipping all the 0-bits,  $t_i$  be the number of times when reaching to the  $i^{th}$  bit after  $i - 1$  bits, therefore the probability for reach. Then  $T = t_k + t_{k+1} + \dots + t_m$ . The probability of each 0-bit being picked is,

$$p_i = \frac{n - (i - 1)}{n}$$

Here comes the *significant* point, each  $t_i$  has a *geometric distribution*, recall the geometric distribution from [Wikipedia](#),

The probability distribution of the number  $X$  of Bernoulli trials needed to get one success.

So, each time  $t_i$  reflects a geometric distribution since we want to know how many times do we need to be able to pick that 0-bit, by the geometric distribution we can infer the expected value as  $\mathbb{E}[t_i] = 1/p_i$ , by the *linearity of the Expectation*,

$$\begin{aligned}
\mathbb{E}[T] &= \mathbb{E}[t_k] + \mathbb{E}[t_{k+1}] + \dots + \mathbb{E}[t_m] \\
&= \frac{1}{p_k} + \frac{1}{p_{k+1}} + \dots + \frac{1}{p_m} \\
&= \frac{n}{n - (k - 1)} + \frac{n}{n - k} + \dots + \frac{n}{n - (m - 1)} \\
&= n \left( \frac{1}{n - (k - 1)} + \frac{1}{n - k} + \dots + \frac{1}{n - (m - 1)} \right) \\
&= nH_{m-k} \\
&= n \ln(m - k) + C \\
&= n \ln(n/3 - k) + C \\
&= n \ln(n - 3k) + C \\
\mathbb{E}[T] &= \Omega(n \log n)
\end{aligned}$$

By using the Theorem, we can derive a lower bound for the RLS to be  $\Omega(n \log n)$ .

## 4 (1+1)EA Lower Bound

Given the initialisation, as for the RLS the lower bound is  $\Omega(n \log n)$ , so we are also trying to show that for the EA it's also  $\Omega(n \log n)$ . We may assume that everytime we flip the 0-bit to 1-bit it's will yield a improvement. We may choose  $t = (n - 1) \ln n$  iterations until optimum, the reason of choosing such  $t$  maybe from more complicated mathematical proof, since this will give a nice bound for when hitting an optimum.

### 4.1 Waterfall Probability

$$\begin{aligned}
1 - \frac{1}{n}, & \quad \text{the probability of a specific bit not being flipped.} \\
\left(1 - \frac{1}{n}\right)^t, & \quad \text{the probability of a specific bit not being flipped for all } t \text{ iterations} \\
1 - \left(1 - \frac{1}{n}\right)^t, & \quad \text{the probability of a specific bit being flipped for all } t \text{ iterations} \\
\left[1 - \left(1 - \frac{1}{n}\right)^t\right]^m, & \quad \text{the probability of all } m \text{ bits being flipped for all } t \text{ iterations} \\
1 - \left[1 - \left(1 - \frac{1}{n}\right)^t\right]^m, & \quad \text{the probability of all } m \text{ bits not being flipped for all } t \text{ iterations}
\end{aligned}$$

The last one is called the *failure probability*.

## 4.2 Some Calculus

For whatever algorithms time complexity analysis, we always let  $n$  explodes or in another word let  $n \rightarrow \infty$ . Before that, let,

$$\frac{1}{h(n)} = \left(1 - \frac{1}{n}\right)^t$$

Then,

$$\begin{aligned} \lim_{n \rightarrow \infty} 1 - \left[1 - \frac{1}{h(n)}\right]^{n/3} \\ 1 - \left[\lim_{n \rightarrow \infty} \left(1 - \frac{1}{h(n)}\right)^n\right]^{1/3} \quad (\circ) \end{aligned}$$

Now, we need to analyse does  $h(n)$  grows therefore  $1/h(n)$  will diminish and also will the power of  $n$  catch up with the diminishing. Let's analyse the  $h(n)$  first,

$$\begin{aligned} \lim_{n \rightarrow \infty} h(n) &= \lim_{n \rightarrow \infty} \frac{1}{\left(1 - \frac{1}{n}\right)^t} \\ &= \frac{1}{\lim_{n \rightarrow \infty} \left(1 - \frac{1}{n}\right)^{(n-1) \ln n}}, \quad \text{Using the Quotient Limit Law} \\ &= \frac{1}{\lim_{n \rightarrow \infty} \left[\left(1 - \frac{1}{n}\right)^{(n-1)}\right]^{\ln n}} \\ &= \frac{1}{\lim_{n \rightarrow \infty} [e^{-1}]^{\ln n}}, \quad \text{Classic Limits Problem check on [Math Exchange](#)} \\ &= \frac{1}{n^{-1}} \\ &= n \end{aligned}$$

After analysis of  $h(n)$  it yields  $n$ , therefore with  $(\circ)$  as  $n \rightarrow \infty$  it seems to yields back the Original Limits problem. So,

$$\text{From } (\circ) \Rightarrow 1 - e^{-1/3}$$

We have found that the failure probability is constant which is  $\Omega(1)$

### Observation

The limit only approach the value from the side, in this case we only consider approach  $\infty$  from the left to right, therefore as  $n$  is slowly increasing but never reach the value thus an observation arises

$$\lim_{n \rightarrow \infty} \left(1 - \frac{1}{h(n)}\right)^n < e^{-1}$$

From  $(\circ) \Rightarrow \geq 1 - e^{-1/3}$ , Minus sign therefore change the equality sign

Recall that time complexity for the EA is the number of iterations until reaching the optimum, same notation as  $T$  is the total number of iterations until optimum, and we let  $t = (n - 1) \ln n$  iterations until optimum. Therefore if,

- $\mathbf{T} \leq \mathbf{t}$  meaning sucessfully reaching optimum.
- $\mathbf{T} > \mathbf{t}$  meaning failed to reaching optimum.

Then the probability of failing to reach optimum is

$$P(T > t) \geq 1 - e^{-1/3} = \Omega(1)$$

## Explanation

The reason of choosing the failure probability instead of successful probability for the lower bound is because we give a candidate of  $t$  iterations, the number of iterations we want until it's the iteration of optimum. But then after some maths, it is shown that it's not sufficient enough for  $t$  iteration with a probability therefore we need some more iterations, which is the lower bound. It's the buffer of how much more is needed.

## 4.3 Tying Everything Together

After going through the hazard we are blessed with the result of a *survival function* as  $1 - e^{-1/3}$ , one of the properties we can harness is the monotonic decreasing essence. From [Wikipedia](#)

*Every survival function is monotonically decreasing.*

### Using the Tail sum formula

$$\begin{aligned}
\mathbb{E}[T] &= \sum_{k=1}^{n=\infty} P(T > k) \\
&= \sum_{k=1}^t P(T > k) + \sum_{k=t}^{\infty} P(T > k) \\
&\geq \sum_{k=1}^t P(T > k) \\
&\geq \sum_{k=1}^t P(T > t), \quad \text{Survival function property} \\
&\geq \sum_{k=1}^t \Omega(1) \\
&\geq t\Omega(1) \\
&\geq (n-1)\ln(n)\Omega(1) \\
\mathbb{E}[T] &\geq \Omega(n \log n)
\end{aligned}$$

Therefore the lower bound for  $(1+1)$ EA is  $\Omega(n \log n)$