
Out-of-Distribution Generalization of In-Context Learning: A Low-Dimensional Subspace Perspective

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1 Introduction

The remarkable capability of ICL in Transformer-based large language models (LLMs) [1] has sparked both empirical [2–10] and theoretical research [7, 8, 11–15]. However, the generalization capabilities of ICL, particularly whether it can extend beyond its pre-training distribution, remain unclear. For example, Garg et al. [2] empirically showed that ICL is relatively robust to distribution shifts in several settings, as the performance of the Transformer closely matched that of the least-squares estimator on linear regression tasks. Zhang et al. [7] shared a similar conclusion, showing that while shifts in the features cannot be tolerated for a one-layer linear attention model, shifts in the regression weights can be handled well. However, Wang et al. [5] challenged these views, empirically demonstrating that ICL can only solve in-distribution tasks in general. These contrasting views, combined with the lack of a theoretical foundation, highlight the need for a rigorous characterization of the OOD generalization capabilities of ICL.

This work proposes a mathematical framework to demystify and quantify the OOD generalization capabilities of ICL. We theoretically study ICL on a single-layer linear attention model with linear regression, where the weight (or task) vectors are sampled from low-dimensional subspaces. This setup enables us to quantify the distribution shift in the task vectors via the principal angles between subspaces and to characterize the OOD test risk as a function of these angles. Then, we precisely identify the conditions on the pre-training task vectors under which the OOD test risk is either sensitive to or independent of these angles, thereby explaining both the limitations and capabilities of ICL. Specifically, we prove that when the training task vectors are drawn from a single r -dimensional subspace, ICL inevitably incurs test error as a function of the principal angle. On the other hand, when the training task vectors are drawn from a union of subspaces, we show that ICL incurs a test risk that is independent of the principal angles. Unlike the single-subspace setting, this result implies that ICL can generalize to any subspace within the span of the training subspaces, even regions with zero probability density under the training distribution. We hypothesize that this explains when ICL exhibits OOD generalization: the testing task vector lies within the span of the training task vectors.

2 Problem Setup and Theoretical Results

Problem Setup. We study a standard ICL task of predicting the next token. For training, we draw a feature and label pair (\mathbf{x}_i, y_i) as follows: let each feature vector be $\mathbf{x}_i \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$. For all $i \in [n + 1]$, we generate each label $y_i \in \mathbb{R}$ as such:

$$y_i = \mathbf{w}^\top \mathbf{x}_i + \eta_i \quad \text{where} \quad \mathbf{w} \sim \mathcal{N}(\mathbf{0}, \Sigma_s), \quad \eta_i \sim \mathcal{N}(0, \sigma^2), \quad (1)$$

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$\sigma \geq 0$ is the noise level, and $\Sigma_s \in \mathbb{R}^{d \times d}$ is the source task covariance matrix, i.e., the covariance of the training weight \mathbf{w} , which we often refer to as the “task vector”. Then, given $n + 1$ paired examples, we train the single-layer linear attention model in Equation (9) by solving Equation (10). We use g_{ATT}^* and \mathcal{W}^* to respectively denote the optimal model and weights according to this setup. Then, our main goal is to investigate how distribution shifts in the task vector affect the test risk of the optimal model g_{ATT}^* . To this end, at test time, we draw a feature and label pair $(\mathbf{x}_j, \tilde{y}_j)$ independent of the training data in a similar fashion: let each feature vector be $\mathbf{x}_j \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$. For all $j \in [m + 1]$, we generate the label $\tilde{y}_j \in \mathbb{R}$ according to each $\mathbf{x}_j \in \mathbb{R}^d$ as

$$\tilde{y}_j = \tilde{\mathbf{w}}^\top \mathbf{x}_j + \eta_j \quad \text{where } \tilde{\mathbf{w}} \sim \mathcal{N}(\mathbf{0}, \Sigma_t), \quad \eta_j \sim \mathcal{N}(0, \sigma^2), \quad (2)$$

and $\Sigma_t \in \mathbb{R}^{d \times d}$ is the target covariance matrix, i.e., the covariance for the task vector at test time. Next, we give forms to Σ_s and Σ_t to quantify the distribution shift from training to test time.

Suppose $d \gg r$, and let $\mathbf{U}_s \in \mathbb{R}^{d \times r}$ be an orthonormal basis for an r -dimensional subspace in \mathbb{R}^d . We parameterize Σ_s and Σ_t as follows:

$$\Sigma_s = \mathbf{U}_s \mathbf{U}_s^\top + \epsilon \cdot \mathbf{I}_d \quad \text{and} \quad \Sigma_t = \mathbf{U}_t \mathbf{U}_t^\top + \epsilon \cdot \mathbf{I}_d, \quad (3)$$

where $\epsilon > 0$ is a small constant to ensure invertibility, and \mathbf{U}_t is parameterized as [16, Section 3.8]:

$$\mathbf{U}_t = \mathbf{U}_s \cdot \cos(\Theta) + \mathbf{U}_{s,\perp} \cdot \sin(\Theta), \quad (4)$$

and $\mathbf{U}_{s,\perp} \in \mathbb{R}^{d \times r}$ is an r -dimensional orthonormal basis that is *orthogonal* to \mathbf{U}_s . For simplicity, we will assume all principal angles are equal, i.e., for all $i \in [r]$, $\theta_i = \theta$ for some $\theta \in [0, \frac{\pi}{2}]$ so that $\Theta = \theta \cdot \mathbf{I}_r$. Notice when $\theta = 0$, $\mathbf{U}_t = \mathbf{U}_s$, and when $\theta = \frac{\pi}{2}$, $\mathbf{U}_t = \mathbf{U}_{s,\perp}$. Hence, by parameterizing Σ_s and Σ_t using \mathbf{U}_s and \mathbf{U}_t , changing the value of θ allows us to control how aligned the testing covariance Σ_t is with the training covariance Σ_s . In other words, θ measures the distribution shift from training to testing. Our goal is to quantify the OOD test risk of g_{ATT}^* in terms of θ .

Main Results. In Proposition 1 (available in Appendix B), we prove that even with infinitely many samples, ICL with a single-layer linear attention model exhibits test risk with a non-negligible dependence on the shift between the covariance matrices Σ_t and Σ_s , as measured by θ . This also empirically holds for nonlinear models such as GPT-2 (see Figure 2), which demonstrates that ICL is not inherently robust to subspace shifts. However, consider the following covariance matrices:

$$\Sigma_s = \mathbf{U}_s \mathbf{U}_s^\top + \epsilon \cdot \mathbf{I}_d \quad \text{and} \quad \Sigma_{s,\perp} = \mathbf{U}_{s,\perp} \mathbf{U}_{s,\perp}^\top + \epsilon \cdot \mathbf{I}_d. \quad (5)$$

Then, instead of the training task vector in Equation (1), consider training g_{ATT} on prompts with labels $y_i = \mathbf{w}^\top \mathbf{x}_i + \eta_i$ whose task vector is drawn from a mixture of two Gaussians:

$$\mathbf{w} \sim \gamma \cdot \mathcal{N}(\mathbf{0}, \Sigma_s) + (1 - \gamma) \cdot \mathcal{N}(\mathbf{0}, \Sigma_{s,\perp}), \quad (6)$$

where γ is the mixture probability. The following result states that this training procedure mitigates dependence on θ , given that the prompt lengths are sufficiently large.

Theorem 1 (Test Risk under the Span of Covariance Matrices). *Let g_{ATT}^* denote the optimal linear attention model corresponding to the independent data setting in Equation (1), where the task vector now follows Equation (6) with $\gamma = 0.5$. For all $j \in [m + 1]$, suppose that the prompts at test time are constructed with features $\mathbf{x}_j \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$ and labels*

$$\tilde{y}_j = \tilde{\mathbf{w}}^\top \mathbf{x}_j + \eta_j, \quad \text{where } \tilde{\mathbf{w}} \sim \mathcal{N}(\mathbf{0}, \Sigma_t), \quad \eta_j \sim \mathcal{N}(0, \sigma^2),$$

and $\Sigma_t \in \mathbb{R}^{d \times d}$ is from Equation (3). For any $\theta \in [0, \frac{\pi}{2}]$ and $\delta \in (0, r)$, we have

$$\lim_{m,n \rightarrow \infty} \lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y}_{m+1} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] = \sigma^2. \quad (7)$$

This highlights an interesting property of Transformers: if the pre-training task vectors are drawn from a union of subspaces, then ICL can interpolate to the space between the subspaces. In other words, even if certain regions have zero probability density in the distribution over the training task vector, ICL can still generalize to those regions at test time, as long as they lie within the overall span of the training task vectors. We hypothesize this can explain why ICL can seemingly achieve OOD generalization: the test data actually lies within the span of the training data. Due to space limitations, we only present the main ideas and defer all other details to the Appendix.

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Appendix

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A Background: Single-Layer Linear Attention Model

The work by Ahn et al. [17] empirically showed many phenomena observed in vanilla Transformers can be replicated in Transformers with linear attention. These findings motivated other works [7, 14, 15] to use linear attention as a test-bed for studying ICL. Following these works, we consider a single-layer linear attention model for analysis. Let $\{\mathbf{x}, y\} \in \mathbb{R}^d \times \mathbb{R}$ denote a feature and label pair.

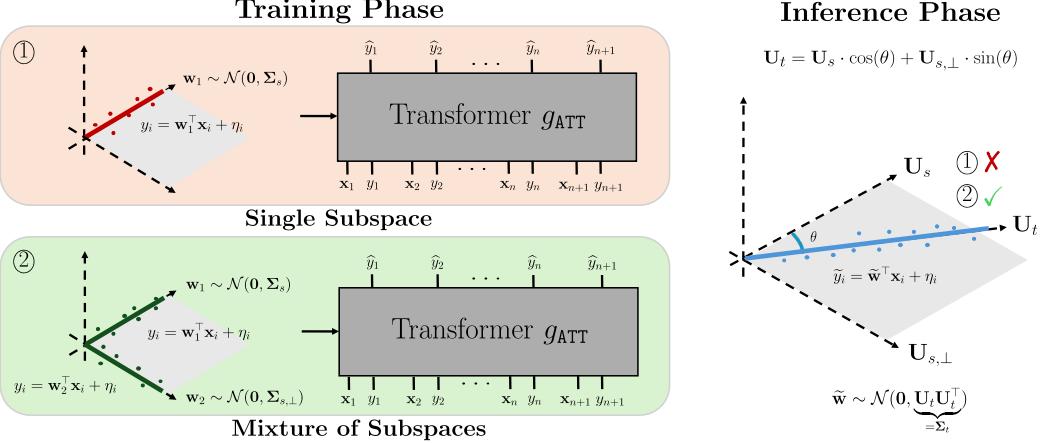


Figure 1: **Overview of this paper.** We consider two models: one trained with task vectors drawn from a single subspace, and one with task vectors drawn from a union of subspaces. At inference, we test both models using a task vector at an angle between two subspaces. The single subspace model fails to generalize under distribution shifts, while the latter generalizes across all angles.

Given $n + 1$ paired examples $\{\mathbf{x}_i, y_i\}_{i=1}^{n+1}$, we construct the training-time input prompt as such:

$$\mathbf{Z} = [\mathbf{z}_1 \quad \dots \quad \mathbf{z}_n \quad \mathbf{z}_{n+1}]^\top = \begin{bmatrix} \mathbf{x}_1 & \dots & \mathbf{x}_n & \mathbf{x}_{n+1} \\ y_1 & \dots & y_n & 0 \end{bmatrix}^\top \in \mathbb{R}^{(n+1) \times (d+1)},$$

Following Ahn et al. [14] and Mahankali et al. [18], we employ a causal mask to the prompt to ensure inputs cannot attend to their own labels:

$$\mathbf{Z}_{\mathcal{M}} = [\mathbf{z}_1 \quad \dots \quad \mathbf{z}_n \quad 0]^\top, \quad \text{where } \mathbf{z}_i = \begin{bmatrix} \mathbf{x}_i \\ y_i \end{bmatrix} \quad \text{and } \mathbf{z}_q = \begin{bmatrix} \mathbf{x}_{n+1} \\ 0 \end{bmatrix}. \quad (8)$$

The goal of ICL is to leverage the in-context examples $\{\mathbf{x}_i, y_i\}_{i=1}^n$ in the prompt $\mathbf{Z}_{\mathcal{M}}$ to predict the correct label y_{n+1} according to the query \mathbf{x}_{n+1} (equivalently \mathbf{z}_q). We input the prompt $\mathbf{Z}_{\mathcal{M}}$ and query \mathbf{z}_q into a (normalized) single head linear attention model to make the prediction \hat{y}_{n+1} :

$$\hat{y}_{n+1} = g_{\text{ATT}}(\mathbf{z}_q, \mathbf{Z}_{\mathcal{M}}) = \frac{1}{n} (\mathbf{z}_q^\top \mathbf{W}_Q \mathbf{W}_K^\top \mathbf{Z}_{\mathcal{M}}) \mathbf{Z}_{\mathcal{M}} \mathbf{W}_V \mathbf{p}, \quad (9)$$

where $\mathbf{W}_K, \mathbf{W}_Q, \mathbf{W}_V \in \mathbb{R}^{(d+1) \times (d+1)}$ are the key, query, and value weight matrices, respectively, and $\mathbf{p} \in \mathbb{R}^{d+1}$ is the linear prediction head. We denote $\mathcal{W} = \{\mathbf{W}_K, \mathbf{W}_Q, \mathbf{W}_V, \mathbf{p}\}$ as the collection of trainable weights corresponding to the linear attention model. We train the model g_{ATT} by minimizing the following expected squared loss with respect to the parameters \mathcal{W} :

$$\min_{\mathcal{W}} \mathcal{L}_{\text{ATT}}(\mathcal{W}), \quad \text{where } \mathcal{L}_{\text{ATT}}(\mathcal{W}) = \mathbb{E} \left[(y_{n+1} - g_{\text{ATT}}(\mathbf{z}_q, \mathbf{Z}_{\mathcal{M}}))^2 \right]. \quad (10)$$

For inference, given $m + 1$ paired examples $\{\mathbf{x}_j, \tilde{y}_j\}_{j=1}^{m+1}$, we construct the input prompts as such:

$$\tilde{\mathbf{Z}}_{\mathcal{M}} = [\tilde{\mathbf{z}}_1 \quad \dots \quad \tilde{\mathbf{z}}_m \quad 0]^\top, \quad \text{where } \tilde{\mathbf{z}}_j = \begin{bmatrix} \mathbf{x}_j \\ \tilde{y}_j \end{bmatrix} \quad \text{and } \tilde{\mathbf{z}}_q = \begin{bmatrix} \mathbf{x}_{m+1} \\ 0 \end{bmatrix}.$$

Then, the inputs $\tilde{\mathbf{Z}}_{\mathcal{M}}$ and $\tilde{\mathbf{z}}_q$ are fed into the trained linear attention model to obtain a prediction for \hat{y}_{m+1} . Specifically, let $\mathcal{W}^* = \{\mathbf{W}_K^*, \mathbf{W}_Q^*, \mathbf{W}_V^*, \mathbf{p}^*\}$ denote the optimally trained linear attention model for minimizing the loss in Equation (10). We compute

$$\hat{y}_{m+1} = g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) = \frac{1}{m} (\tilde{\mathbf{z}}_q^\top \mathbf{W}_Q^* \mathbf{W}_K^{*\top} \tilde{\mathbf{Z}}_{\mathcal{M}}^\top) \tilde{\mathbf{Z}}_{\mathcal{M}} \mathbf{W}_V^* \mathbf{p}^*,$$

where we normalize by a factor of m instead of n . Doing so decouples the training and testing prompt lengths, which allows us to analyze the behavior of ICL under different conditions.

B Main Results

This section presents our main results in detail to support those discussed in the main text. We illustrate an overview of the setup in Figure 1.

B.1 Transformers Are Not Robust To Subspace Shifts

In this section, we consider the setup in Section 2, where we train a single-layer linear attention model according to Equation (1), and test the (optimal) model with Equation (2). We prove that even with infinitely many samples, ICL exhibits test risk with a non-negligible dependence on the shift between the covariance matrices Σ_t and Σ_s , as measured by θ . This result demonstrates that ICL is not inherently robust to subspace shifts.

Proposition 1 (Task Distribution Shift). *Let g_{ATT}^* denote the optimal linear attention model corresponding to the independent data setting in Equation (1). For all $j \in [m+1]$, suppose that the prompts at test time are constructed with features $\mathbf{x}_j \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$ and labels*

$$\tilde{y}_j = \tilde{\mathbf{w}}^\top \mathbf{x}_j + \eta_j, \quad \text{where } \tilde{\mathbf{w}} \sim \mathcal{N}(\mathbf{0}, \Sigma_t), \quad \eta_j \sim \mathcal{N}(0, \sigma^2),$$

and $\Sigma_t \in \mathbb{R}^{d \times d}$ is from Equation (3). Then, we have

$$\lim_{m \rightarrow \infty} \lim_{n \rightarrow \infty} \lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y}_{m+1} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] = r \sin^2(\theta) + \sigma^2, \quad (11)$$

where $\theta \in [0, \frac{\pi}{2}]$ are the r principal angles between $\mathbf{U}_s \in \mathbb{R}^{d \times r}$ and the test subspace $\mathbf{U}_t \in \mathbb{R}^{d \times r}$.

The proof is provided in Appendix G.1.2. We take $\epsilon \rightarrow 0$ for two reasons: (i) to eliminate any dependence on ϵ and isolate its effect on test risk as it is assumed to be a small constant, and (ii) to ensure that the covariance matrices are exactly low-rank. Then, in the asymptotic regime, our result reveals the following: when $\theta = 0$, the $\sin(\cdot)$ term vanishes, allowing perfect recovery up to the label noise variance. However, as θ increases from 0 to $\frac{\pi}{2}$, the test risk increases with respect to θ . At $\theta = \frac{\pi}{2}$, the test risk becomes exactly the rank of the covariance matrix. Notably, this represents the largest possible error in this setting, as a low-rank covariance matrix induces an error dependent on the rank rather than the ambient dimension, as observed in related work [2, 19].

The analysis involves deriving the test risk under an arbitrary distribution shift, assuming the linear attention model is parameterized by the optimal weights according to Equation (10). At the optimal weights, the model reduces to a single step of projected gradient descent (PGD) [13–15, 18]. Denoting $\mathbf{A} \in \mathbb{R}^{d \times d}$ as the PGD projection matrix that arises from the optimal weights, we sketch how the dependence on θ arises (assuming $\sigma = 0$ for simplicity):

$$\begin{aligned} \hat{y}_{m+1} &= g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) = \frac{1}{m} \mathbf{x}_{m+1}^\top \mathbf{A} \mathbf{X}^\top \mathbf{y} = \frac{1}{m} \mathbf{x}_{m+1}^\top \mathbf{A} \mathbf{X}^\top \mathbf{X} \tilde{\mathbf{w}} \quad (\text{Substitute } \mathbf{y} = \mathbf{X} \tilde{\mathbf{w}}) \\ &\rightarrow \mathbf{x}_{m+1}^\top \mathbf{U}_s \mathbf{U}_s^\top \tilde{\mathbf{w}} \quad (\text{Take } m, n \rightarrow \infty \text{ and } \epsilon \rightarrow 0) \\ &= \mathbf{x}_{m+1}^\top \mathbf{U}_s \mathbf{U}_s^\top \mathbf{U}_t \mathbf{g}, \quad (\tilde{\mathbf{w}} = \mathbf{U}_t \mathbf{g} \text{ for } \mathbf{g} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_r)) \end{aligned}$$

where $\mathbf{X} := [\mathbf{x}_1 \dots \mathbf{x}_m]^\top \in \mathbb{R}^{m \times d}$ and $\mathbf{y} := [\tilde{y}_1 \dots \tilde{y}_m]^\top \in \mathbb{R}^m$. By taking appropriate limits, it is easy to see that the dependence on θ arises from $\mathbf{U}_s^\top \mathbf{U}_t$, which reflects a rotation by an angle θ between the subspaces. Since $\mathbf{A} \rightarrow \mathbf{U}_s \mathbf{U}_s^\top$ in the asymptotic regime, PGD projects the data onto an “incorrect” subspace, thereby inducing an error proportional to θ in the test risk. Put differently, in cases in which $\Sigma_t \neq \Sigma_s$, ICL can generalize only if $\mathcal{R}(\Sigma_t) \subset \mathcal{R}(\Sigma_s)$.

In Figure 2, we present experiments corroborating Proposition 1 on both linear and nonlinear Transformers. Interestingly, our experiments show that both models incur the same test risk under the distribution shift when given enough in-context examples. This implies that the linear attention model can adequately capture the behavior in this setting, and that the observed error is not merely an artifact of using a linear model. Lastly, we assumed equal principal angles between the subspaces for simplicity, and defer the more general result to Proposition 2 in Appendix F.1.

B.2 Transformers Can Generalize to the Span When Trained on a Union of Subspaces

Previously, we observed that shifting the covariance matrix induces a dependence on θ in the test risk due to projection onto a misaligned subspace, implying that the training and testing data must

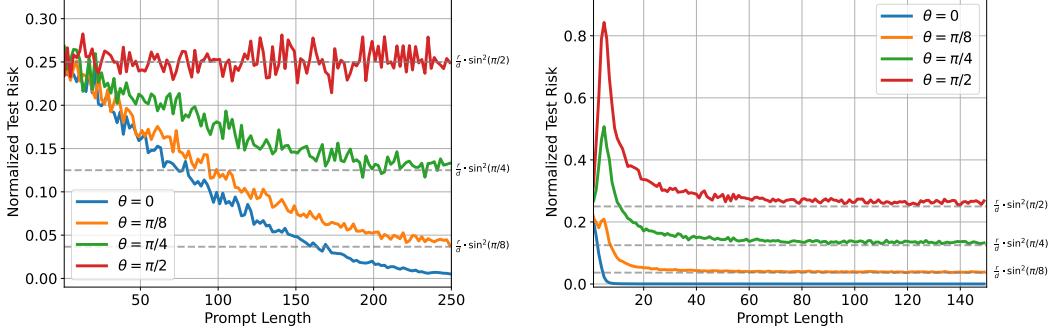


Figure 2: Plot of the normalized test risk for OOD linear regression as a function of the prompt length for a linear Transformer (left) and a nonlinear Transformer (right) under covariance shifts. As the covariance at test time shifts away from the covariance used at training time as a function of θ , the test risk exhibits a non-negligible dependence on θ for both the linear and nonlinear Transformer. Moreover, for both models, the test risk exactly matches the predicted risk from Proposition 1.

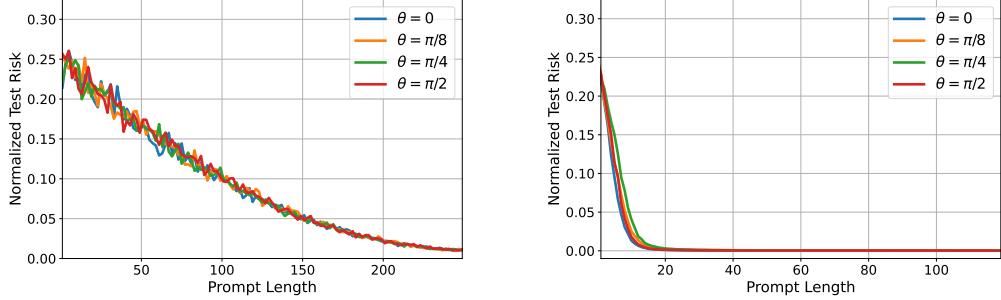


Figure 3: Plot of the test risk for OOD linear regression as a function of the prompt length for a linear Transformer (left) and a nonlinear Transformer (right). When the prompt length at test time is large enough, the test risk goes nearly to zero for all $\theta \in [0, \frac{\pi}{2}]$, corroborating Theorem 2 in that both linear and nonlinear Transformers can generalize to the span of the training task vectors at test-time.

span the same r -dimensional subspace. This raises the question: are there settings in ICL where the dependence on θ can be mitigated? In the main text, we showed that this dependence can be mitigated, roughly speaking, by introducing diversity into the training prompts. Specifically, we showed that by drawing task vectors from a union of subspaces, the projection matrix can better capture shifts in θ , allowing OOD generalization. In the following, we re-phrase Theorem 1 in the same format as Proposition 1.

Theorem 2 (Test Risk under the Span of Covariance Matrices). *Let g_{ATT}^* denote the optimal linear attention model corresponding to the independent data setting in Equation (1), where the task vector now follows Equation (6) with $\gamma = 0.5$. For all $j \in [m+1]$, suppose that the prompts at test time are constructed with features $\mathbf{x}_j \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$ and labels*

$$\tilde{y}_j = \tilde{\mathbf{w}}^\top \mathbf{x}_j + \eta_j, \quad \text{where } \tilde{\mathbf{w}} \sim \mathcal{N}(\mathbf{0}, \Sigma_t), \quad \eta_j \sim \mathcal{N}(0, \sigma^2),$$

and $\Sigma_t \in \mathbb{R}^{d \times d}$ is from Equation (3). For any $\theta \in [0, \frac{\pi}{2}]$ and $\delta \in (0, r)$, if

$$m \geq n > \frac{(2(r + \sigma^2) + 1)r}{\delta} - (2(r + \sigma^2) + 1), \quad (12)$$

then $\lim_{\epsilon \rightarrow 0} \mathbb{E} \left[(\tilde{y}_{m+1} - g_{ATT}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}))^2 \right] < \sigma^2 + \delta$.

The proof technique is similar to that of Proposition 1 and is available in Appendix G.1.3. Moreover, we can generalize this result to a mixture of $K > 2$ subspaces; see Appendix F.2. For Theorem 2,

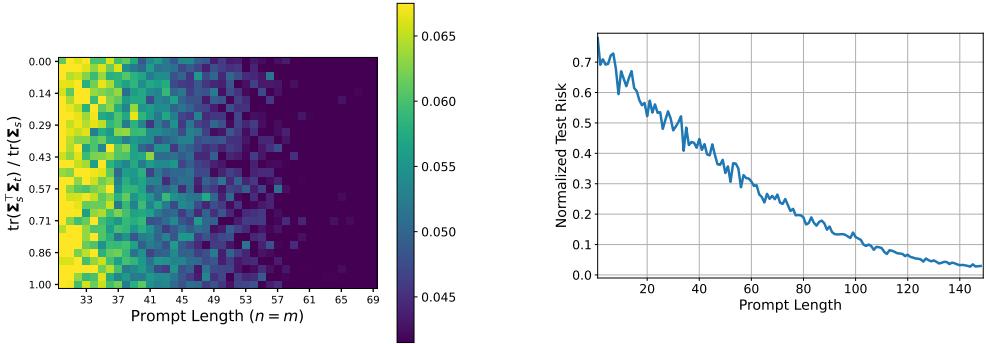


Figure 4: Left: Phase plot of the test risk as we vary the angle between Σ_s and Σ_t and the prompt length with $m = n$ for a linear attention model trained with a mixture of Gaussians. The test risk is low across all angle shifts, and decreases further as the prompt length increases. Right: Plot of the test risk as a function of the prompt length for a case in which $\Sigma_s \neq \Sigma_t$ but with $\theta = 0$, following the OOD example in Gatmiry et al [20]. This serves to explain why ICL can seemingly do OOD generalization as observed in the literature.

we similarly sketch how θ becomes mitigated in the test risk. Consider the case where $\delta \rightarrow 0$, i.e., $m, n \rightarrow \infty$. Then, we can simplify the linear attention model as such (again assuming $\sigma = 0$):

$$\hat{y}_{m+1} = g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \rightarrow \mathbf{x}_{m+1}^\top \mathbf{U}_{2r} \mathbf{U}_{2r}^\top \tilde{\mathbf{w}} = \mathbf{x}_{m+1}^\top \mathbf{U}_{2r} \mathbf{U}_{2r}^\top \mathbf{U}_t \mathbf{g},$$

where again $\mathbf{g} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_r)$ and $\mathbf{U}_{2r} = [\mathbf{U}_s \quad \mathbf{U}_{s,\perp}]$. Since $\mathcal{R}(\mathbf{U}_t) \subset \mathcal{R}(\mathbf{U}_{2r})$ for all $\theta \in [0, \frac{\pi}{2}]$, the trained model perfectly recovers \hat{y}_{m+1} . In Figure 3, we present results on linear Transformers and GPT-2 that corroborate our theory. In both models, the test risk approaches zero for all $\theta \in [0, \frac{\pi}{2}]$, meaning there is no dependence on θ . The only noticeable difference is the linear attention model requires a longer prompt length to reach near-zero risk, which is also highlighted by our theory.

Overall, this highlights an interesting property of Transformers: if the training task vectors are drawn from a union of subspaces, then ICL can interpolate to the space between the subspaces. In other words, even if certain regions have zero probability density in the distribution over the training task vector, ICL can still generalize to those regions at test time, as long as they lie within the overall span of the training task vectors. We hypothesize this can explain why ICL can seemingly achieve OOD generalization: the test data actually remains within the span of the training distribution.

C Experimental Results

Experimental Setup. Unless otherwise stated, the experimental setup is as follows: for both the linear and nonlinear Transformer, we consider noiseless linear regression, and set $d = 20$, $r = 5$, and $\epsilon = 10^{-6}$. To construct the train and test subspaces, we sample an orthogonal matrix $\mathbf{U} \in \mathbb{R}^{d \times d}$ uniformly at random, set \mathbf{U}_s to be the first r columns of \mathbf{U} , and set $\mathbf{U}_{s,\perp}$ to be the second r columns. Given this setup, we typically consider a mixture of $K = 2$ subspaces for the experiments.

For the experiments with the linear Transformer, we plug in the optimal weights according to their respective settings (e.g., optimal weights using a single subspace or a mixture of subspaces) and set $m = n = 250$. For the nonlinear Transformer, following Garg et al. [2], we use a small GPT-2 model with 6 layers, 4 heads, and a 128-dimensional embedding space. We append a learnable linear transformation to map the vector predicted by the model to a scalar. We use a learning rate of $\eta = 10^{-4}$, batch size 64, prompt lengths $m = n = 120$, and train for 100K iterations.

C.1 More Results on Linear Function Classes

Linear Regression. Previously, we presented results on the test risk as a function of the prompt length. In Figure 4 (left), we present a phase plot of the test risk as a function of both $\text{Tr}(\Sigma_s^\top \Sigma_t) / \text{Tr}(\Sigma_s)$ (which measures the angle between two covariance matrices) and the prompt length on linear attention with task vectors drawn from a mixture of two Gaussians. Similar to Figure 3, the test risk is low for all values of $m = n$, and it decreases further as the prompt length

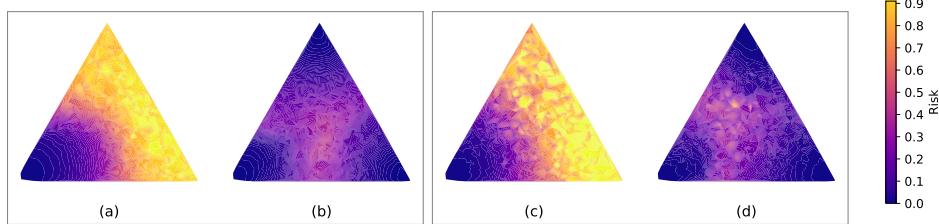


Figure 5: Visualization of the generalization behavior of Transformers for learning nonlinear function classes in-context. Each corner of a triangle represents a one-dimensional subspace spanned by ψ_1 (bottom left), ψ_2 (bottom right), or ψ_3 (top), with all possible convex combinations given by the interior. In all cases, we show the risk when evaluated at different points in $\text{span}(\{\psi_1, \psi_2, \psi_3\})$ for the appropriate function space. (a) Train on prompts drawn from $\text{span}(\{\psi_1^C\})$. (b) Train on prompts drawn from $\text{span}(\{\psi_1^C\}) \cup \text{span}(\{\psi_2^C\}) \cup \text{span}(\{\psi_3^C\})$. (c) Train on prompts drawn from $\text{span}(\{\psi_1^H\})$. (d) Train on prompts drawn from $\text{span}(\{\psi_1^H\}) \cup \text{span}(\{\psi_2^H\}) \cup \text{span}(\{\psi_3^H\})$.

increases. Note that the largest possible normalized test risk in this setting is $r/d = 0.25$, so the test risk is still considered low even when the prompt length is small.

In Section B.2, we discussed how apparent abilities of ICL to perform OOD generalization arises when the test task lies within the span of the training task vectors. Here, we present an extra experiment to support this claim, using the example from Gatmiry et al. [20], with $d = 5$, $\Sigma_s = \mathbf{I}_5$ and $\Sigma_t = \mathbf{V}\Lambda_t\mathbf{V}^\top$, where $\mathbf{V} \in \mathbb{R}^{5 \times 5}$ is a random orthogonal matrix and $\Lambda_t = \text{Diag}(1, 1, 1/2, 1/4, 1)$. In Figure 4 (right), we observe that the test risk approaches zero given enough samples. This implies that our result may help explain many observations of OOD generalization in ICL and offers a unifying perspective on findings reported in the literature.

C.2 Beyond Linear Function Classes

Finally, we demonstrate that the theoretical findings in Appendix B extend to *nonlinear* function classes. Specifically, we look at two function spaces, namely $L^2([0, 1])$ and $L^2(\mathbb{R}, e^{-x^2/2}/\sqrt{2\pi} dx)$, i.e., square-integrable functions under the uniform and Gaussian measures respectively, which model rich sets of signals observed in real-world data. For the former, we construct an orthonormal basis via cosines, i.e., $\psi_n^C(x) = (1/\sqrt{2}) \cos(n\pi x)$ for $n \in \mathbb{N}$. For the latter, we construct an orthonormal basis via Hermite polynomials:

$$\psi_n^H(x) = \frac{(-1)^n}{\sqrt{n!}} e^{x^2/2} \frac{d^n(e^{-x^2/2})}{dx^n} \quad \text{for } n \in \mathbb{N}.$$

As described in previous sections, we consider two settings: observing instances of a single (one-dimensional) subspace, as well as for a union of three (one-dimensional) subspaces. As before, we draw the function coefficients from standard multivariate Gaussian. We draw the inputs from the distribution appropriate to the function space measure, i.e., $x \sim \mathcal{U}([0, 1])$ for $L^2([0, 1])$ and $x \sim \mathcal{N}(0, 1)$ for $L^2(\mathbb{R}, e^{-x^2/2}/\sqrt{2\pi} dx)$. All other details are identical to previous (nonlinear) Transformer experiments. The results are shown in Figure 5. As shown in panels (a) and (c) of Figure 5, we see that Transformers are not robust to subspace shifts for either function class, with increasing test risk with respect to the subspace angle from the train subspace, in accordance with Proposition 1. On the other hand, as shown in panels (b) and (d) of Figure 5, we have the generalization behavior described by Theorem 3, where training on the mixture of subspaces results in low risk in the space spanned by the basis vectors.

D Discussion

In this work, we analyzed the OOD generalization capabilities of ICL by studying a single-layer linear attention model with linear regression, where the task vector was parameterized by low-dimensional subspaces. We uncovered two key properties of ICL: (i) it is not inherently robust to subspace shifts, and (ii) it can generalize to the span of covariance matrices if trained on a union

of subspaces. We also provided insights into how LoRA can be used to model distribution shifts, and showed how our findings extend to nonlinear function classes. One limitation of this work is that the analysis focuses on single-layer linear attention, as in prior studies; a promising direction for future research is to extend the analysis to multi-layer nonlinear Transformers.

E Related Work

ICL on Transformers with Linear Attention. There is abundant research on ICL that analyzes single-layer linear attention models. Below, we survey several works most relevant to our work; like ours, many of them focus on linear regression settings, where for all $i \in [n + 1]$:

$$y_i = f(\mathbf{x}_i) = \mathbf{w}^\top \mathbf{x}_i + \eta, \quad \text{where } \mathbf{w} \sim \mathcal{N}(\mathbf{0}_d, \Sigma_{\mathbf{w}}), \quad \mathbf{x}_i \sim \mathcal{N}(\mathbf{0}_d, \Sigma_{\mathbf{x}}),$$

and η is additive Gaussian noise. As previously mentioned, Zhang et al. [7] studied the training dynamics of a single-layer linear attention model on the population loss for a linear regression ICL task. Specifically, assuming $\Sigma_{\mathbf{w}} = \mathbf{I}_d$ and an arbitrary $\Sigma_{\mathbf{x}}$, they showed the model weights converge to a globally optimal solution under gradient flow, despite the non-convex objective. They also provide closed-form expressions for the model weights at the global minima. A follow-up work [21] considered a linear regression task with $\mathbf{w} \sim \mathcal{N}(\mu_{\mathbf{w}}, \Sigma_{\mathbf{w}})$ and a linear Transformer model (a linear attention layer followed by a two-layer linear network). They showed a single linear attention layer incurs a sub-optimal risk that depends on $\mu_{\mathbf{w}}$, but adding a linear network allows the model to achieve the Bayes optimal risk.

Other works [13–15, 21–23] study the underlying learning algorithms that linear attention models implement when learning linear functions in-context. Specifically, for a single linear attention layer, Von Oswald et al. [13] demonstrated the existence of model weights that implement a single step of GD on a mean-squared error loss. They further showed empirically that the weights of a trained linear attention layer closely align with those that implement a GD step. Follow-up works [14, 15, 22] rigorously proved the equivalence between a single step of preconditioned gradient descent (PGD) with zero initialization and the weights of a single-layer linear attention model under the population loss. Specifically, Ahn et al. [14] theoretically showed when $\Sigma_{\mathbf{w}} = \mathbf{I}_d$ and $\Sigma_{\mathbf{x}}$ is arbitrary, the single-layer linear attention model learns a preconditioning matrix that is dependent on $\Sigma_{\mathbf{x}}$. Li et al. [15] generalized this result by considering an arbitrary $\Sigma_{\mathbf{w}}$ in addition to $\Sigma_{\mathbf{x}}$ — they showed the learned preconditioning matrix depends on both $\Sigma_{\mathbf{x}}$ and $\Sigma_{\mathbf{w}}$. Finally, [21] showed when $\mathbf{w} \sim \mathcal{N}(\mu_{\mathbf{w}}, \Sigma_{\mathbf{w}})$, a linear attention layer followed by a linear network implements a PGD step while *learning* the initialization. While our work builds on the fact that a single-layer linear attention model implements PGD, our goal differs from these prior works: we study how ICL under this model can generalize out-of-distribution.

Empirical Observations on ICL’s OOD Generalization. As part of their study, Garg et al. [2] empirically observed Transformer-based ICL is robust to a number of distribution shifts, such as between the train and test distributions of the features \mathbf{x}_i , as well as between the features \mathbf{x}_i and query \mathbf{x}_q . These observations inspired an extensive line of empirical work studying ICL’s ability to generalize to OOD tasks [3–6, 9, 10, 24]. To our knowledge, [4, 5] are the most closely related with our setting. Specifically, these works consider sampling tasks from a mixture of *function class* distributions, e.g., f is sampled from the class of dense linear functions with probability $\gamma \in (0, 1)$, or from the class of sparse linear functions with probability $1 - \gamma$. Yadlowskey et al. [4] showed when Transformers are trained for ICL on a mixture of function classes, ICL cannot generalize well to function classes not present in the training mixture. Wang et al. [5] argue if the test task is not in the training mixture, Transformers select a task from the training mixture that minimizes the test error. In contrast, our work assumes that the target function is sampled from a mixture of low-dimensional subspaces in a fixed function space. In other words, the mixture distribution from which we sample is always within a *single* function class. We emphasize this is different from sampling from a mixture of *multiple* function class distributions.

Theoretical Studies on ICL’s OOD Generalization. The above empirical observations motivated theoretical studies on ICL’s OOD generalization ability. Under their setting, Zhang et al. [7] studied how a trained single linear attention layer handles various distribution shifts. Assuming the model weights were at the global minima of Equation (10), they derived a closed-form expression for the prediction \hat{y}_q for a given query \mathbf{x}_q and in-context examples $(\mathbf{x}_1, \mathbf{w}^\top \mathbf{x}_1, \dots, \mathbf{x}_m, \mathbf{w}^\top \mathbf{x}_m)$. Using

this expression for \hat{y}_q , they concluded a trained linear attention model is robust to task and query shifts, but cannot tolerate feature shifts well.

Other works have studied *nonlinear* models and function classes. For instance, [8] considered a binary classification ICL task. They showed a sufficiently trained single-layer, single-head Transformer model (one softmax attention layer followed by a two-layer perceptron) can achieve arbitrarily small generalization error when the inference-time features are *linear combinations* of the training features. Another work [25] assumed the function to learn in-context was $f(\mathbf{x}) = \mathbf{w}^\top g(\mathbf{x}) + \eta$, where $g(\mathbf{x}) = (g_1(\mathbf{x}), \dots, g_\ell(\mathbf{x}))$ is an arbitrary feature mapping. They showed if \mathbf{w} has iid, zero mean, unit variance entries at train time, and $\|\mathbf{w}\|_2$ is bounded at inference time, a trained single-layer, multi-head softmax attention model generalizes well under *any* shift in \mathbf{w} . Again, our paper differs from these works by studying when ICL can and cannot perform OOD generalization, particularly by using low-dimensional subspaces to parameterize the covariance matrices.

Learning Functions with Low-Dimensional Structure In-Context. To the best of our knowledge, the work by [19] is the only most related work that also considers learning functions with low-dimensional structures. In their setting, the function to learn in-context is a single-index model $f(\mathbf{x}) = \sigma(\mathbf{w}^\top \mathbf{x}) + \eta$, where $\sigma(\cdot)$ is a nonlinear link function, \mathbf{w} is drawn from a low-dimensional subspace, and η is additive noise. We only consider linear functions $f(\mathbf{x}) = \mathbf{w}^\top \mathbf{x} + \eta$ in our analysis, but also assume \mathbf{w} is sampled from a low-dimensional distribution. In our experiments, we sample nonlinear functions from subspaces of the *function space*, which differs from sampling the function *parameters* from a subspace of Euclidean space. Furthermore, our goal is to use such a parameterization to study OOD generalization, whereas the main focus of [19] is to examine whether ICL can solve such functions at all.

F Additional Results

In this section, we present additional results to supplement those presented in the main text. All experiments were run using either a Macbook Pro with an Apple M2 Pro Chip or a NVIDIA A100 GPU.

Appendix F.1 presents an additional theoretical result for Proposition 1 for when the principal angles are different. Additionally, in Appendix F.2, we present another result where we generalize the mixture of two Gaussians from Theorem 2 to a mixture of $K \geq 2$ Gaussians.

F.1 Result with Different Principal Angles

In Proposition 1, we assumed that all of the r principal angles between the subspaces $\mathbf{U}_s \in \mathbb{R}^{d \times r}$ and $\mathbf{U}_t \in \mathbb{R}^{d \times r}$ were all the same, i.e., $\theta_i = \theta \in [0, \frac{\pi}{2}]$, for simplicity. In Proposition 2, we relax this requirement and present a result where the angles are not necessarily the same.

Proposition 2 (Task Distribution Shift with Different Angles). *Let g_{ATT}^* denote the optimal linear attention model corresponding to the independent data setting in Equation (1). For all $j \in [m+1]$, suppose that the prompts at test time are constructed with features $\mathbf{x}_j \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$ and labels*

$$\tilde{y}_j = \tilde{\mathbf{w}}^\top \mathbf{x}_j + \eta_j, \quad \text{where } \tilde{\mathbf{w}} \sim \mathcal{N}(\mathbf{0}, \Sigma_t) \quad \text{and} \quad \eta_j \sim \mathcal{N}(0, \sigma^2),$$

with covariance matrix $\Sigma_t \in \mathbb{R}^{d \times d}$ from Equation (3). Then, we have

$$\lim_{m \rightarrow \infty} \lim_{n \rightarrow \infty} \lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y}_{m+1} - g_{ATT}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] = \sum_{i=1}^r \sin^2(\theta_i) + \sigma^2, \quad (13)$$

where $\theta_i \in [0, \frac{\pi}{2}]$ is the i -th principal angle between the train subspace $\mathbf{U}_s \in \mathbb{R}^{d \times r}$ and the test subspace $\mathbf{U}_t \in \mathbb{R}^{d \times r}$.

Recall that the test risk presented in Proposition 1 was $r \sin^2(\theta) + \sigma^2$. It is easy to see that if we set $\theta_i = \theta$, then the test risk in Proposition 2 recovers the risk in Proposition 1, i.e., $\sum_{i=1}^r \sin^2(\theta_i) = r \sin^2(\theta)$.

F.2 Generalization Beyond a Mixture of Two Gaussians

We now discuss how ICL can achieve OOD generalization when $\mathbf{w} \in \mathbb{R}^d$ is sampled from a mixture of K low-rank Gaussians for any $K \geq 2$. Let $\mathbf{U} = [\mathbf{u}_1 \dots \mathbf{u}_d] \in \mathbb{R}^{d \times d}$ be an orthonormal basis for \mathbb{R}^d . Assuming $d > Kr$, let $\mathbf{U}_{s,k} = [\mathbf{u}_{(k-1) \cdot r+1} \dots \mathbf{u}_{kr}] \in \mathbb{R}^{d \times r}$ for all $k \in [K]$. Note $\mathbf{U}_{s,k}^\top \mathbf{U}_{s,l} = \mathbf{0}_{r \times r}$ for all $k \neq l$.

We assume the training task $\mathbf{w} \in \mathbb{R}^d$ is sampled as such:

$$\mathbf{w} \sim \sum_{k=1}^K \gamma_k \cdot \mathcal{N}(\mathbf{0}, \Sigma_{s,k}), \text{ where } \Sigma_{s,k} = \mathbf{U}_{s,k} \mathbf{U}_{s,k}^\top + \epsilon \cdot \mathbf{I}_d \text{ and } \sum_{k=1}^K \gamma_k = 1. \quad (14)$$

Then, let $\bar{\mathbf{U}}_t$ be an arbitrary orthonormal basis for an r -dimensional subspace that lies in the span of $[\mathbf{U}_{s,1} \dots \mathbf{U}_{s,K}]$, i.e.,

$$\bar{\mathbf{U}}_t = \sum_{k=1}^K \alpha_k \mathbf{U}_{s,k} \text{ for some } \{\alpha_k\}_{k=1}^K \text{ s.t. } \sum_{k=1}^K \alpha_k^2 = 1, \quad (15)$$

where the last constraint on $\{\alpha_k\}_{k=1}^K$ ensures $\bar{\mathbf{U}}_t$ is an orthonormal basis. Similar to the $K = 2$ case, we show when tested on $\tilde{\mathbf{w}} \sim \mathcal{N}(\mathbf{0}, \bar{\Sigma}_t)$ with $\bar{\Sigma}_t = \bar{\mathbf{U}}_t \bar{\mathbf{U}}_t^\top + \epsilon \cdot \mathbf{I}_d$, the trained model can generalize to this previously unseen subspace $\bar{\mathbf{U}}_t$.

Theorem 3. *Let g_{ATT}^* denote the optimal linear attention model corresponding to the independent data setting in Equation (1), where the task vector is drawn from Equation (14) with $\gamma_k = \frac{1}{K}$ for all $k \in [K]$. For all $j \in [m+1]$, suppose that the prompts at test time are constructed with features $\mathbf{x}_j \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$ and labels*

$$\tilde{y}_j = \tilde{\mathbf{w}}^\top \mathbf{x}_j + \eta_j, \text{ where } \tilde{\mathbf{w}} \sim \mathcal{N}(\mathbf{0}, \bar{\Sigma}_t) \text{ and } \eta_j \sim \mathcal{N}(0, \sigma^2),$$

where $\bar{\Sigma}_t = \bar{\mathbf{U}}_t \bar{\mathbf{U}}_t^\top + \epsilon \cdot \mathbf{I}_d$ and $\bar{\mathbf{U}}_t$ is defined in Equation (15). For any $\{\alpha_k\}_{k=1}^K$ s.t. $\sum_{k=1}^K \alpha_k^2 = 1$ and $\delta \in (0, r)$, if

$$m \geq n > \frac{(K(r + \sigma^2) + 1)r}{\delta} - (K(r + \sigma^2) + 1), \quad (16)$$

$$\text{then } \lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{ATT}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] < \sigma^2 + \delta.$$

The proof is deferred to Appendix G.1.4. Similar to Theorem 2, if the linear attention model is trained on task vectors that lie in a union of K subspaces, it can generalize well to *any* region within the span of the K subspaces, even if those regions have zero probability density during training. We note setting $K = 2$, $\alpha_1 = \cos(\theta)$, and $\alpha_2 = \sin(\theta)$ perfectly recovers Theorem 2.

G Deferred Proofs

This section presents all deferred proofs and is organized as follows: Section G.1 contains all proofs related to shifts in the task vector $\mathbf{w} \in \mathbb{R}^d$, and Appendix G.2 provides auxiliary results used to support both the task and feature shift proofs.

G.1 Proofs for Task Shifts

G.1.1 Supporting Results

We first derive an expression for the test risk under a general distribution shift for the task vector.

Lemma 1 (Test Risk under General Task Distribution Shift). *Let g_{ATT}^* denote the optimal linear attention model corresponding to the independent data setting in Equation (1). For all $j \in [m+1]$, suppose that the prompts at test time are constructed with features $\mathbf{x}_i \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$ and labels*

$$\tilde{y}_j = \tilde{\mathbf{w}}^\top \mathbf{x}_j + \eta_j, \text{ where } \tilde{\mathbf{w}} \sim \mathcal{N}(\mathbf{0}, \Sigma_t) \text{ and } \eta_j \sim \mathcal{N}(0, \sigma^2).$$

Then,

$$\mathbb{E} \left[\left(\tilde{y}_{m+1} - g_{ATT}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] = M_t - \text{Tr}(\Sigma_t \mathbf{A}) + \frac{M_t}{m} \text{Tr}(\mathbf{A}^\top \mathbf{A}) - \text{Tr}(\Sigma_t \mathbf{A}) + \frac{m+1}{m} \text{Tr}(\mathbf{A} \Sigma_t \mathbf{A}),$$

where $M_t = \text{Tr}(\Sigma_t) + \sigma^2$.

Proof. Recall at inference time,

$$\tilde{\mathbf{Z}}_{\mathcal{M}} = [\tilde{\mathbf{z}}_1 \ \dots \ \tilde{\mathbf{z}}_m \ \mathbf{0}]^\top = \begin{bmatrix} \mathbf{x}_1 & \dots & \mathbf{x}_m & \mathbf{0} \\ \tilde{y}_1 & \dots & \tilde{y}_m & 0 \end{bmatrix}^\top \quad \text{and} \quad \tilde{\mathbf{z}}_q = \begin{bmatrix} \mathbf{x}_{m+1} \\ 0 \end{bmatrix} := \begin{bmatrix} \mathbf{x}_q \\ 0 \end{bmatrix}. \quad (17)$$

Then, let us define

$$\mathbf{X}_{te} := [\mathbf{x}_1 \ \mathbf{x}_2 \ \dots \ \mathbf{x}_m]^\top, \quad \mathbf{y}_{te} := [\tilde{y}_1 \ \tilde{y}_2 \ \dots \ \tilde{y}_m]^\top, \quad \boldsymbol{\eta}_{te} := [\eta_1 \ \eta_2 \ \dots \ \eta_m]^\top,$$

and $\eta_q := \eta_{m+1}$. Note $\mathbf{y}_{te} = \mathbf{X}_{te} \tilde{\mathbf{w}} + \boldsymbol{\eta}_{te}$. By Lemma 2, we have

$$g_{ATT}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) = \frac{1}{m} \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{y}_{te} = \mathbf{x}_q^\top \underbrace{\left(\frac{1}{m} \mathbf{A} \mathbf{X}_{te}^\top \mathbf{y}_{te} \right)}_{:= \tilde{\mathbf{w}}},$$

where $\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \Sigma_s \right)^{-1}$. By plugging this into the risk and linearity of expectation,

$$\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q + \eta_q - \tilde{\mathbf{w}}^\top \mathbf{x}_q)^2 \right] = \underbrace{\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q + \eta_q)^2 \right]}_{(a)} - 2 \underbrace{\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q + \eta_q) (\mathbf{x}_q^\top \tilde{\mathbf{w}}) \right]}_{(b)} + \underbrace{\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q)^2 \right]}_{(c)}. \quad (18)$$

It suffices to analyze each individual term.

Analyzing (a). We first evaluate $\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q + \eta_q)^2 \right]$. First, we note

$$\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q + \eta_q)^2 \right] = \mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q)^2 \right] + 2\mathbb{E} [\eta_q \tilde{\mathbf{w}}^\top \mathbf{x}_q] + \mathbb{E} [\eta_q^2] = \mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q)^2 \right] + \sigma^2,$$

so it suffices to analyze $\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q)^2 \right]$. By law of total expectation and the fact that $\tilde{\mathbf{w}}, \mathbf{x}_q$ are independent,

$$\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q)^2 \right] = \mathbb{E}_{\tilde{\mathbf{w}}} \left[\mathbb{E}_{\mathbf{x}_q} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q)^2 \mid \tilde{\mathbf{w}} \right] \right].$$

Conditioned on $\tilde{\mathbf{w}}$, $\tilde{\mathbf{w}}^\top \mathbf{x}_q \sim \mathcal{N}(0, \|\tilde{\mathbf{w}}\|^2)$, so $\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q)^2 \mid \tilde{\mathbf{w}} \right] = \text{Var}(\tilde{\mathbf{w}}^\top \mathbf{x}_q \mid \tilde{\mathbf{w}}) = \|\tilde{\mathbf{w}}\|^2$. Therefore,

$$\mathbb{E}_{\tilde{\mathbf{w}}} \left[\mathbb{E}_{\mathbf{x}} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q)^2 \mid \tilde{\mathbf{w}} \right] \right] = \mathbb{E} [\|\tilde{\mathbf{w}}\|^2] = \text{Tr}(\mathbb{E} [\tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top]) = \text{Tr}(\Sigma_t).$$

Therefore,

$$\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q + \eta_q)^2 \right] = \text{Tr}(\Sigma_t) + \sigma^2.$$

Analyzing (b). Next, we analyze $\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q + \eta_q) (\mathbf{x}_q^\top \tilde{\mathbf{w}}) \right]$. We first note

$$\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q + \eta_q) (\mathbf{x}_q^\top \tilde{\mathbf{w}}) \right] = \mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q) (\mathbf{x}_q^\top \tilde{\mathbf{w}}) \right] + \underbrace{\mathbb{E} [\eta_q \mathbf{x}_q^\top \tilde{\mathbf{w}}]}_{=0} = \mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q) (\mathbf{x}_q^\top \tilde{\mathbf{w}}) \right],$$

so it suffices to analyze $\mathbb{E} [(\tilde{\mathbf{w}}^\top \mathbf{x}_q) (\mathbf{x}_q^\top \tilde{\mathbf{w}})]$. Substituting $\tilde{\mathbf{w}} := \frac{1}{m} \mathbf{A} \mathbf{X}_{te}^\top \mathbf{y}_{te} = \frac{1}{m} \mathbf{A} \mathbf{X}_{te}^\top (\mathbf{X}_{te} \tilde{\mathbf{w}} + \boldsymbol{\eta}_{te})$ yields

$$\begin{aligned} \mathbb{E} [(\tilde{\mathbf{w}}^\top \mathbf{x}_q) (\mathbf{x}_q^\top \tilde{\mathbf{w}})] &= \frac{1}{m} \mathbb{E} [\tilde{\mathbf{w}}^\top \mathbf{x}_q \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top (\mathbf{X}_{te} \tilde{\mathbf{w}} + \boldsymbol{\eta}_{te})] \\ &= \frac{1}{m} \left(\mathbb{E} [\tilde{\mathbf{w}}^\top \mathbf{x}_q \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}}] + \mathbb{E} [\tilde{\mathbf{w}}^\top \mathbf{x}_q \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \boldsymbol{\eta}_{te}] \right) \\ &= \frac{1}{m} \left(\mathbb{E} [\tilde{\mathbf{w}}^\top \mathbf{x}_q \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}}] + \underbrace{\mathbb{E} [\tilde{\mathbf{w}}^\top \mathbf{x}_q \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top] \mathbb{E} [\boldsymbol{\eta}_{te}]}_{=0} \right) \\ &= \frac{1}{m} \mathbb{E} [\tilde{\mathbf{w}}^\top \mathbf{x}_q \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}}] = \frac{1}{m} \mathbb{E} [\text{Tr} (\tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top \mathbf{x}_q \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te})] \\ &= \frac{1}{m} \text{Tr} (\mathbb{E} [\tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top \mathbf{x}_q \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te}]) \\ &= \frac{1}{m} \text{Tr} \left(\underbrace{\mathbb{E} [\tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top]}_{\Sigma_t} \underbrace{\mathbb{E} [\mathbf{x}_q \mathbf{x}_q^\top]}_{\mathbf{I}_d} \mathbf{A} \underbrace{\mathbb{E} [\mathbf{X}_{te}^\top \mathbf{X}_{te}]}_{m \cdot \mathbf{I}_d} \right) = \text{Tr} (\Sigma_t \mathbf{A}), \end{aligned}$$

where again $\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \Sigma_s^{-1} \right)^{-1}$.

Analyzing (c). Finally, we analyze $\mathbb{E} [(\mathbf{x}_q^\top \tilde{\mathbf{w}})^2]$:

$$\begin{aligned} \mathbb{E} [(\mathbf{x}_q^\top \tilde{\mathbf{w}})^2] &= \frac{1}{m^2} \mathbb{E} [(\mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top (\mathbf{X}_{te} \tilde{\mathbf{w}} + \boldsymbol{\eta}_{te}))^2] = \frac{1}{m^2} \mathbb{E} [(\mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}} + \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \boldsymbol{\eta}_{te})^2] \\ &= \frac{1}{m^2} \left(\mathbb{E} [(\mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}})^2] + 2 \underbrace{\mathbb{E} [(\mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}})(\mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \boldsymbol{\eta}_{te})]}_{=0} + \mathbb{E} [(\mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \boldsymbol{\eta}_{te})^2] \right) \\ &= \frac{1}{m^2} \left(\underbrace{\mathbb{E} [\mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top \mathbf{X}_{te}^\top \mathbf{X}_{te} \mathbf{A}^\top \mathbf{x}_q]}_{(d)} + \underbrace{\mathbb{E} [\mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \boldsymbol{\eta}_{te} \boldsymbol{\eta}_{te}^\top \mathbf{X}_{te} \mathbf{A}^\top \mathbf{x}_q]}_{(e)} \right). \end{aligned}$$

We first focus on (d):

$$\begin{aligned} \mathbb{E} [\mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top \mathbf{X}_{te}^\top \mathbf{X}_{te} \mathbf{A}^\top \mathbf{x}_q] &= \mathbb{E} [\text{Tr} (\mathbf{x}_q \mathbf{x}_q^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top \mathbf{X}_{te}^\top \mathbf{X}_{te} \mathbf{A}^\top)] \\ &= \text{Tr} \left(\underbrace{\mathbb{E} [\mathbf{x}_q \mathbf{x}_q^\top]}_{\mathbf{I}_d} \mathbf{A} \mathbb{E} [\mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top \mathbf{X}_{te}^\top \mathbf{X}_{te}] \mathbf{A}^\top \right) \\ &= \mathbb{E} [\text{Tr} (\mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top \mathbf{X}_{te}^\top \mathbf{X}_{te} \mathbf{A}^\top)] = \mathbb{E} [\text{Tr} (\tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top \mathbf{X}_{te}^\top \mathbf{X}_{te} \mathbf{A}^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te})] \\ &= \text{Tr} \left(\underbrace{\mathbb{E} [\tilde{\mathbf{w}} \tilde{\mathbf{w}}^\top]}_{\Sigma_t} \mathbb{E} [\mathbf{X}_{te}^\top \mathbf{X}_{te} \mathbf{A}^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te}] \right) = \mathbb{E} [\text{Tr} (\Sigma_t \mathbf{X}_{te}^\top \mathbf{X}_{te} \mathbf{A}^\top \mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te})] \\ &= \mathbb{E} [\text{Tr} (\mathbf{A} \mathbf{X}_{te}^\top \mathbf{X}_{te} \Sigma_t^{1/2} \Sigma_t^{1/2} \mathbf{X}_{te}^\top \mathbf{X}_{te} \mathbf{A}^\top)] := \mathbb{E} [\text{Tr} (\tilde{\mathbf{X}}_{te}^\top \bar{\mathbf{X}}_{te} \bar{\mathbf{X}}_{te}^\top \tilde{\mathbf{X}}_{te})], \end{aligned}$$

where $\tilde{\mathbf{X}}_{te}^\top := \mathbf{A} \mathbf{X}_{te}^\top$ and $\bar{\mathbf{X}}_{te}^\top := \Sigma_t^{1/2} \mathbf{X}_{te}^\top$. Note $\tilde{\mathbf{X}}_{te}^\top = [\mathbf{A} \mathbf{x}_1 \dots \mathbf{A} \mathbf{x}_m] := [\tilde{\mathbf{x}}_1 \dots \tilde{\mathbf{x}}_m]$ where $\tilde{\mathbf{x}}_i := \mathbf{A} \mathbf{x}_i \stackrel{iid}{\sim} \mathcal{N}(\mathbf{0}_d, \mathbf{A} \mathbf{A}^\top)$, and $\bar{\mathbf{X}}_{te}^\top = [\Sigma_t^{1/2} \mathbf{x}_1 \dots \Sigma_t^{1/2} \mathbf{x}_m] := [\bar{\mathbf{x}}_1 \dots \bar{\mathbf{x}}_m]$ where $\bar{\mathbf{x}}_i := \Sigma_t^{1/2} \mathbf{x}_i \stackrel{iid}{\sim} \mathcal{N}(\mathbf{0}_d, \Sigma_t)$. We can express $\tilde{\mathbf{X}}_{te}^\top \bar{\mathbf{X}}_{te}$ and $\bar{\mathbf{X}}_{te}^\top \tilde{\mathbf{X}}_{te}$ as such:

$$\tilde{\mathbf{X}}_{te}^\top \bar{\mathbf{X}}_{te} = \sum_{i=1}^m \tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^\top \quad \text{and} \quad \bar{\mathbf{X}}_{te}^\top \tilde{\mathbf{X}}_{te} = \sum_{j=1}^m \bar{\mathbf{x}}_j \tilde{\mathbf{x}}_j^\top.$$

Therefore,

$$\begin{aligned} \mathbb{E}\left[\operatorname{Tr}\left(\tilde{\mathbf{X}}_{te}^{\top} \bar{\mathbf{X}}_{te} \bar{\mathbf{X}}_{te}^{\top} \tilde{\mathbf{X}}_{te}\right)\right] &= \operatorname{Tr}\left(\mathbb{E}\left[\tilde{\mathbf{X}}_{te}^{\top} \bar{\mathbf{X}}_{te} \bar{\mathbf{X}}_{te}^{\top} \tilde{\mathbf{X}}_{te}\right]\right) \\ &= \operatorname{Tr}\left(\sum_{i=1}^m \sum_{j=1}^m \mathbb{E}\left[\tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^{\top} \bar{\mathbf{x}}_j \tilde{\mathbf{x}}_j^{\top}\right]\right) \\ &= \operatorname{Tr}\left(\sum_{i=1}^m \sum_{j \neq i} \mathbb{E}\left[\tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^{\top} \bar{\mathbf{x}}_j \tilde{\mathbf{x}}_j^{\top}\right]\right)+\operatorname{Tr}\left(\sum_{i=1}^m \mathbb{E}\left[\tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^{\top} \bar{\mathbf{x}}_i \tilde{\mathbf{x}}_i^{\top}\right]\right) \end{aligned}$$

We first consider the case when $i \neq j$. In this setting, \mathbf{x}_i and \mathbf{x}_j are independent, so

$$\mathbb{E}\left[\tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^{\top} \bar{\mathbf{x}}_j \tilde{\mathbf{x}}_j^{\top}\right]=\mathbb{E}\left[\tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^{\top}\right] \mathbb{E}\left[\bar{\mathbf{x}}_j \tilde{\mathbf{x}}_j^{\top}\right]=\mathbf{A} \underbrace{\mathbb{E}\left[\mathbf{x}_i \mathbf{x}_i^{\top}\right]}_{\mathbf{I}_d} \Sigma_t \underbrace{\mathbb{E}\left[\mathbf{x}_j \mathbf{x}_j^{\top}\right]}_{\mathbf{I}_d} \mathbf{A}^{\top}=\mathbf{A} \Sigma_t \mathbf{A}^{\top}.$$

Therefore,

$$\operatorname{Tr}\left(\sum_{i=1}^m \sum_{j \neq i} \mathbb{E}\left[\tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^{\top} \bar{\mathbf{x}}_j \tilde{\mathbf{x}}_j^{\top}\right]\right)=m \cdot(m-1) \cdot \operatorname{Tr}\left(\mathbf{A} \Sigma_t \mathbf{A}^{\top}\right).$$

We now consider the case where $i=j$:

$$\begin{aligned} \operatorname{Tr}\left(\sum_{i=1}^m \mathbb{E}\left[\tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^{\top} \bar{\mathbf{x}}_i \tilde{\mathbf{x}}_i^{\top}\right]\right) &= \sum_{i=1}^m \mathbb{E}\left[\operatorname{Tr}\left(\tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^{\top} \bar{\mathbf{x}}_i \tilde{\mathbf{x}}_i^{\top}\right)\right] \\ &= \sum_{i=1}^m \mathbb{E}\left[\tilde{\mathbf{x}}_i^{\top} \tilde{\mathbf{x}}_i \bar{\mathbf{x}}_i^{\top} \bar{\mathbf{x}}_i\right]=\sum_{i=1}^m \mathbb{E}\left[(\mathbf{x}_i^{\top} \mathbf{A}^{\top} \mathbf{A} \mathbf{x}_i)(\mathbf{x}_i^{\top} \Sigma_t \mathbf{x}_i)\right] \\ &\stackrel{(i)}{=} m \cdot\left(2 \operatorname{Tr}\left(\mathbf{A} \Sigma_t \mathbf{A}^{\top}\right)+\operatorname{Tr}\left(\mathbf{A}^{\top} \mathbf{A}\right) \operatorname{Tr}\left(\Sigma_t\right)\right), \end{aligned}$$

where (i) is because for $\mathbf{a} \sim \mathcal{N}(\mathbf{0}_d, \mathbf{I}_d)$ and fixed $\mathbf{Q}, \mathbf{R} \in \mathbb{R}^{d \times d}, \mathbb{E}\left[(\mathbf{a}^{\top} \mathbf{Q} \mathbf{a})(\mathbf{a}^{\top} \mathbf{R} \mathbf{a})\right]=\operatorname{Tr}\left(\mathbf{Q}(\mathbf{R}+\mathbf{R}^{\top})\right)+\operatorname{Tr}(\mathbf{Q}) \operatorname{Tr}(\mathbf{R})$ (see Section 8.2.4 in [26]).

We now focus on (e):

$$\begin{aligned} \mathbb{E}\left[\mathbf{x}_q^{\top} \mathbf{A} \mathbf{X}_{te}^{\top} \boldsymbol{\eta}_{te} \boldsymbol{\eta}_{te}^{\top} \mathbf{X}_{te} \mathbf{A}^{\top} \mathbf{x}_q\right] &= \mathbb{E}\left[\operatorname{Tr}\left(\mathbf{x}_q \mathbf{x}_q^{\top} \mathbf{A} \mathbf{X}_{te}^{\top} \boldsymbol{\eta}_{te} \boldsymbol{\eta}_{te}^{\top} \mathbf{X}_{te} \mathbf{A}^{\top}\right)\right] \\ &= \operatorname{Tr}\left(\underbrace{\mathbb{E}\left[\mathbf{x}_q \mathbf{x}_q^{\top}\right]}_{\mathbf{I}_d} \mathbf{A} \mathbb{E}\left[\mathbf{X}_{te}^{\top} \boldsymbol{\eta}_{te} \boldsymbol{\eta}_{te}^{\top} \mathbf{X}_{te}\right] \mathbf{A}^{\top}\right)=\operatorname{Tr}\left(\mathbb{E}\left[\mathbf{A} \mathbf{X}_{te}^{\top} \boldsymbol{\eta}_{te} \boldsymbol{\eta}_{te}^{\top} \mathbf{X}_{te} \mathbf{A}^{\top}\right]\right) \\ &:=\operatorname{Tr}\left(\mathbb{E}\left[\tilde{\boldsymbol{\eta}}_{te} \tilde{\boldsymbol{\eta}}_{te}^{\top}\right]\right), \end{aligned}$$

where $\tilde{\boldsymbol{\eta}}_{te}:=\mathbf{A} \mathbf{X}_{te}^{\top} \boldsymbol{\eta}_{te}=\tilde{\mathbf{X}}_{te}^{\top} \boldsymbol{\eta}_{te}$. Note the columns of $\tilde{\mathbf{X}}_{te}^{\top}$ are iid Gaussian with covariance $\mathbf{A} \mathbf{A}^{\top}$. By Corollary 6 in [27], $\tilde{\boldsymbol{\eta}}_{te} \sim \text { GAL }_d\left(2 \sigma^2 \mathbf{A} \mathbf{A}^{\top}, \mathbf{0}_d, m / 2\right)$, where $\text { GAL }_p(\boldsymbol{\Sigma}, \boldsymbol{\mu}, s)$ denotes a p -dimensional *multivariate generalized asymmetric Laplace distribution* with mean $s \boldsymbol{\mu}$ and covariance $s(\boldsymbol{\Sigma}+\boldsymbol{\mu} \boldsymbol{\mu}^{\top})$ (Definition 1 and Proposition 2 in [27]). Therefore,

$$\operatorname{Tr}\left(\mathbb{E}\left[\tilde{\boldsymbol{\eta}}_{te} \tilde{\boldsymbol{\eta}}_{te}^{\top}\right]\right)=\operatorname{Tr}\left(\operatorname{Cov}\left(\tilde{\boldsymbol{\eta}}_{te}\right)\right)=m \sigma^2 \operatorname{Tr}\left(\mathbf{A} \mathbf{A}^{\top}\right).$$

Adding (a), (b), and (c). Adding the expressions for (a), (b), and (c), where (c) = (d) + (e), yields and combining like terms yields the following expression:

$$\begin{aligned} \mathbb{E}\left[\left(\tilde{\mathbf{w}}^{\top} \mathbf{x}_q+\eta_q-\hat{\mathbf{w}}^{\top} \mathbf{x}_q\right)^2\right] &= \underbrace{\operatorname{Tr}(\boldsymbol{\Sigma}_t)+\sigma^2}_{=(a)}-\underbrace{2 \operatorname{Tr}(\boldsymbol{\Sigma}_t \mathbf{A})}_{=(b)} \\ &+ \underbrace{\frac{1}{m^2}\left(\underbrace{m(m-1) \operatorname{Tr}\left(\mathbf{A} \boldsymbol{\Sigma}_t \mathbf{A}^{\top}\right)+2 m \operatorname{Tr}\left(\mathbf{A} \boldsymbol{\Sigma}_t \mathbf{A}^{\top}\right)+m \operatorname{Tr}\left(\boldsymbol{\Sigma}_t\right) \operatorname{Tr}\left(\mathbf{A}^{\top} \mathbf{A}\right)}_{=(d)}+\underbrace{m \sigma^2 \operatorname{Tr}\left(\mathbf{A}^{\top} \mathbf{A}\right)}_{=(e)}\right)}_{=(c)} . \end{aligned}$$

Combining like terms yields

$$\begin{aligned}\mathbb{E} \left[(\tilde{\mathbf{w}}^\top \mathbf{x}_q + \eta_q - \hat{\mathbf{w}}^\top \mathbf{x}_q)^2 \right] &= \left(\frac{1}{m} \text{Tr}(\mathbf{A}^\top \mathbf{A}) + 1 \right) \left(\text{Tr}(\Sigma_t) + \sigma^2 \right) - 2 \text{Tr}(\Sigma_t \mathbf{A}) + \frac{m+1}{m} \text{Tr}(\mathbf{A} \Sigma_t \mathbf{A}^\top) \\ &= M_t - \text{Tr}(\Sigma_t \mathbf{A}) + \frac{M_t}{m} \text{Tr}(\mathbf{A}^\top \mathbf{A}) - \text{Tr}(\Sigma_t \mathbf{A}) + \frac{m+1}{m} \text{Tr}(\mathbf{A} \Sigma_t \mathbf{A}),\end{aligned}$$

which is exactly Equation (18). This completes the proof. \square

G.1.2 Proof of Proposition 1

Proof. For simplicity, we denote $\tilde{y} := \tilde{y}_{m+1}$. Recall $\mathbf{U} := [\mathbf{U}_s \quad \mathbf{U}_{s,\perp} \quad \mathbf{U}_{2r,\perp}] \in \mathbb{R}^{d \times d}$, where $\mathbf{U}_s, \mathbf{U}_{s,\perp} \in \mathbb{R}^{d \times r}$ and $\mathbf{U}_{2r,\perp} \in \mathbb{R}^{d \times (d-2r)}$ all have orthonormal columns, while $\mathbf{U}_s^\top \mathbf{U}_{\perp,s} = \mathbf{0}_{r \times r}$ and $\mathbf{U}_s^\top \mathbf{U}_\perp = \mathbf{U}_{s,\perp}^\top \mathbf{U}_{2r,\perp} = \mathbf{0}_{r \times (d-2r)}$. We re-write Σ_s as such:

$$\Sigma_s = \mathbf{U}_s \mathbf{U}_s^\top + \epsilon \mathbf{I}_d = \mathbf{U} \begin{bmatrix} \mathbf{I}_r & \\ & \mathbf{0}_{(d-r) \times (d-r)} \end{bmatrix} \mathbf{U}^\top + \epsilon \mathbf{I} = \mathbf{U} \begin{bmatrix} (1+\epsilon) \mathbf{I}_r & \\ & \epsilon \mathbf{I}_{d-r} \end{bmatrix} \mathbf{U}^\top.$$

Note this is a valid eigendecomposition of Σ_s . Thus, by Lemma 5, we have

$$\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \Sigma_s^{-1} \right)^{-1} = \mathbf{U} \Lambda \mathbf{U}^\top, \quad (19)$$

where

$$\Lambda = \begin{bmatrix} \frac{n(1+\epsilon)}{(n+1)\epsilon+M_s} \cdot \mathbf{I}_r & \\ & \frac{n\epsilon}{(n+1)\epsilon+M_s} \cdot \mathbf{I}_{d-r} \end{bmatrix} := \begin{bmatrix} \nu_1 \mathbf{I}_r & \\ & \nu_2 \mathbf{I}_{d-r} \end{bmatrix}.$$

and $M_s = \text{Tr}(\Sigma_s) + \sigma^2$.

By Lemma 1 (and omitting the subscripts in the expectation),

$$\mathbb{E} \left[(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}))^2 \right] = \left(\frac{1}{m} \text{Tr}(\mathbf{A}^\top \mathbf{A}) + 1 \right) \left(\text{Tr}(\Sigma_t) + \sigma^2 \right) - 2 \text{Tr}(\Sigma_t \mathbf{A}) + \frac{m+1}{m} \text{Tr}(\mathbf{A} \Sigma_t \mathbf{A}^\top). \quad (20)$$

We simplify the remaining $\text{Tr}(\cdot)$ terms using Equation (19).

Simplifying $\text{Tr}(\mathbf{A})$ and $\text{Tr}(\mathbf{A}^\top \mathbf{A})$. Directly from Equation (19):

$$\text{Tr}(\mathbf{A}) = r \cdot \nu_1 + (d-r) \cdot \nu_2 \quad \text{and} \quad \text{Tr}(\mathbf{A}^\top \mathbf{A}) = \text{Tr}(\mathbf{A}^2) = r \cdot \nu_1^2 + (d-r) \cdot \nu_2^2,$$

where $\mathbf{A}^2 = \mathbf{U} \Lambda^2 \mathbf{U}^\top$.

Simplifying $\text{Tr}(\Sigma_t \mathbf{A})$ and $\text{Tr}(\mathbf{A} \Sigma_t \mathbf{A}^\top)$. First note $\text{Tr}(\mathbf{A} \Sigma_t \mathbf{A}^\top) = \text{Tr}(\Sigma_t \mathbf{A}^2)$. We first focus on $\text{Tr}(\Sigma_t \mathbf{A})$:

$$\begin{aligned}\Sigma_t \mathbf{A} &= (\mathbf{U}_t \mathbf{U}_t^\top + \epsilon \mathbf{I}_d) \mathbf{U} \Lambda \mathbf{U}^\top = \mathbf{U}_t \mathbf{U}_t^\top \mathbf{U} \Lambda \mathbf{U}^\top + \epsilon \mathbf{U} \Lambda \mathbf{U}^\top \\ &\implies \text{Tr}(\Sigma_t \mathbf{A}) = \text{Tr}(\mathbf{U}_t^\top \mathbf{U} \Lambda \mathbf{U}^\top \mathbf{U}_t) + \epsilon \text{Tr}(\mathbf{A}).\end{aligned}$$

Recall we defined \mathbf{U}_t in Equation (4) as follows:

$$\mathbf{U}_t = \mathbf{U}_s \cos(\Theta) + \mathbf{U}_{s,\perp} \sin(\Theta).$$

Therefore:

$$\mathbf{U}_t^\top \mathbf{U} = (\mathbf{U}_s \cos(\Theta) + \mathbf{U}_{s,\perp} \sin(\Theta))^\top [\mathbf{U}_s \quad \mathbf{U}_{s,\perp} \quad \mathbf{U}_\perp] = [\cos(\Theta) \quad \sin(\Theta) \quad \mathbf{0}_{d \times d-2r}],$$

and thus,

$$\begin{aligned}\text{Tr}(\mathbf{U}_t^\top \mathbf{U} \Lambda \mathbf{U}^\top \mathbf{U}_t) &= \text{Tr} \left([\cos(\Theta) \quad \sin(\Theta) \quad \mathbf{0}_{d \times (d-2r)}] \begin{bmatrix} \nu_1 \mathbf{I}_r & & \\ & \nu_2 \mathbf{I}_r & \\ & & \nu_2 \mathbf{I}_{d-2r} \end{bmatrix} \begin{bmatrix} \cos(\Theta) \\ \sin(\Theta) \\ \mathbf{0}_{(d-2r) \times d} \end{bmatrix} \right) \\ &= \text{Tr} \left(\begin{bmatrix} \nu_1 \cos^2(\Theta) & & \\ & \nu_2 \sin^2(\Theta) & \\ & & \mathbf{0}_{(d-2r) \times (d-2r)} \end{bmatrix} \right) = r \cdot \nu_1 \cdot \cos^2(\theta) + r \cdot \nu_2 \cdot \sin^2(\theta),\end{aligned}$$

where we used the fact that the principal angles are all equal to θ . Using a similar argument,

$$\text{Tr}(\Sigma_t^\top \mathbf{A}^2) = r \cdot \nu_1^2 \cdot \cos^2(\theta) + r \cdot \nu_2^2 \cdot \sin^2(\theta) + \epsilon \text{Tr}(\mathbf{A}^2)$$

Simplifying the Test Risk. Substituting the expressions for the $\text{Tr}(\cdot)$ terms into Equation (20) yields

$$\begin{aligned}\mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] &= \left(\frac{1}{m} (r\nu_1^2 + (d-r)\nu_2^2) + 1 \right) (r + \epsilon d + \sigma^2) \\ &\quad - 2(r\nu_1 \cos^2(\theta) + r\nu_2 \sin^2(\theta) + (r\nu_1 + (d-r)\nu_2)\epsilon) \\ &\quad + \frac{m+1}{m} (r\nu_1^2 \cos^2(\theta) + r\nu_2^2 \sin^2(\theta) + (r\nu_1^2 + (d-r)\nu_2^2)\epsilon)\end{aligned}$$

Substituting the expressions for ν_1 and ν_2 and taking $\epsilon \rightarrow 0$ results in the following:

$$\begin{aligned}\lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] &= \left(\frac{rn^2}{m(n+1+r+\sigma^2)^2} + 1 \right) (r + \sigma^2) \\ &\quad - \frac{2rn \cos^2(\theta)}{n+1+r+\sigma^2} + \frac{(m+1)rn^2 \cos^2(\theta)}{m(n+1+r+\sigma^2)^2}\end{aligned}$$

Subsequently taking $m, n \rightarrow \infty$ yields

$$\lim_{m \rightarrow \infty} \lim_{n \rightarrow \infty} \lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] = r + \sigma^2 - r \cos^2(\theta) = r \sin^2(\theta) + \sigma^2,$$

which completes the proof. \square

G.1.3 Proof of Theorem 2

Proof. For simplicity, we denote $\tilde{y} := \tilde{y}_{m+1}$. Recall that by Lemma 1 and Lemma 3, we have

$$\begin{aligned}\mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] &= \left(\frac{1}{m} \text{Tr}(\mathbf{A}^\top \mathbf{A}) + 1 \right) (\text{Tr}(\boldsymbol{\Sigma}_t) + \sigma^2) \\ &\quad - 2 \text{Tr}(\boldsymbol{\Sigma}_t \mathbf{A}) + \frac{m+1}{m} \text{Tr}(\mathbf{A} \boldsymbol{\Sigma}_t \mathbf{A}^\top),\end{aligned}\tag{21}$$

where $\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \boldsymbol{\Sigma}^{-1} \right)^{-1}$, $M_s = \text{Tr}(\boldsymbol{\Sigma}) + \sigma^2$ with $\boldsymbol{\Sigma} = \gamma \boldsymbol{\Sigma}_s + (1-\gamma) \boldsymbol{\Sigma}_{s,\perp}$. First, we simplify $\boldsymbol{\Sigma}$ as such:

$$\begin{aligned}\boldsymbol{\Sigma} &= \gamma \boldsymbol{\Sigma}_s + (1-\gamma) \boldsymbol{\Sigma}_{s,\perp} \\ &= \gamma (\mathbf{U}_s \mathbf{U}_s^\top + \epsilon \cdot \mathbf{I}_d) + (1-\gamma) (\mathbf{U}_{s,\perp} \mathbf{U}_{s,\perp}^\top + \epsilon \cdot \mathbf{I}_d) \\ &= \mathbf{U} \begin{bmatrix} \gamma(1+\epsilon) \cdot \mathbf{I}_r & \gamma \epsilon \cdot \mathbf{I}_{d-r} \end{bmatrix} \mathbf{U}^\top + \mathbf{U} \begin{bmatrix} (1-\gamma)\epsilon \cdot \mathbf{I}_r & (1-\gamma)(1+\epsilon) \cdot \mathbf{I}_r & (1-\gamma)\epsilon \cdot \mathbf{I}_{d-2r} \end{bmatrix} \mathbf{U}^\top \\ &= \mathbf{U} \begin{bmatrix} (\gamma+\epsilon) \cdot \mathbf{I}_r & (\epsilon-\gamma+1) \cdot \mathbf{I}_r & \epsilon \cdot \mathbf{I}_{d-2r} \end{bmatrix} \mathbf{U}^\top,\end{aligned}$$

and so we have

$$M_s = \text{Tr}(\boldsymbol{\Sigma}) + \sigma^2 = r + \epsilon d + \sigma^2 \quad \text{and} \quad \boldsymbol{\Sigma}^{-1} = \mathbf{U} \begin{bmatrix} \frac{1}{\gamma+\epsilon} \cdot \mathbf{I}_r & & \\ & \frac{1}{\epsilon-\gamma+1} \cdot \mathbf{I}_r & \\ & & \frac{1}{\epsilon} \cdot \mathbf{I}_{d-2r} \end{bmatrix} \mathbf{U}^\top.$$

Then, by Lemma 5, we have

$$\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \boldsymbol{\Sigma}^{-1} \right)^{-1} = \mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^\top,\tag{22}$$

where

$$\boldsymbol{\Lambda} = \begin{bmatrix} \frac{n(\gamma+\epsilon)}{(n+1)(\gamma+\epsilon)+M_s} \cdot \mathbf{I}_r & & \\ & \frac{n(\epsilon-\gamma+1)}{(n+1)(\epsilon-\gamma+1)+M_s} \cdot \mathbf{I}_r & \\ & & \frac{n\epsilon}{(n+1)\epsilon+M_s} \cdot \mathbf{I}_{d-2r} \end{bmatrix} := \begin{bmatrix} \nu_1 \cdot \mathbf{I}_r & & \\ & \nu_2 \cdot \mathbf{I}_r & \\ & & \nu_3 \cdot \mathbf{I}_{d-2r} \end{bmatrix}.$$

We simplify the $\text{Tr}(\cdot)$ terms using Equation (22).

Simplifying $\text{Tr}(\mathbf{A})$ and $\text{Tr}(\mathbf{A}^\top \mathbf{A})$. Directly from Equation (22):

$$\text{Tr}(\mathbf{A}) = r\nu_1 + r\nu_2 + (d - 2r)\nu_3, \quad \text{and} \quad \text{Tr}(\mathbf{A}^\top \mathbf{A}) = \text{Tr}(\mathbf{A}^2) = r\nu_1^2 + r\nu_2^2 + (d - 2r)\nu_3^2.$$

Simplifying $\text{Tr}(\Sigma_t \mathbf{A})$ and $\text{Tr}(\mathbf{A} \Sigma_t \mathbf{A}^\top)$. First note $\text{Tr}(\mathbf{A} \Sigma_t \mathbf{A}^\top) = \text{Tr}(\Sigma_t \mathbf{A}^2)$. We first focus on $\text{Tr}(\Sigma_t \mathbf{A})$:

$$\begin{aligned} \Sigma_t \mathbf{A} &= (\mathbf{U}_t \mathbf{U}_t^\top + \epsilon \mathbf{I}_d) \mathbf{U} \Lambda \mathbf{U}^\top = \mathbf{U}_t \mathbf{U}_t^\top \mathbf{U} \Lambda \mathbf{U}^\top + \epsilon \mathbf{U} \Lambda \mathbf{U}^\top \\ &\implies \text{Tr}(\Sigma_t \mathbf{A}) = \text{Tr}(\mathbf{U}_t^\top \mathbf{U} \Lambda \mathbf{U}^\top \mathbf{U}_t) + \epsilon \text{Tr}(\mathbf{A}). \end{aligned}$$

Recall $\mathbf{U}_t = \mathbf{U}_s \cos(\Theta) + \mathbf{U}_{s,\perp} \sin(\Theta)$, and so we have

$$\mathbf{U}_t^\top \mathbf{U} = (\mathbf{U}_s \cos(\Theta) + \mathbf{U}_{s,\perp} \sin(\Theta))^\top [\mathbf{U}_s \quad \mathbf{U}_{s,\perp} \quad \mathbf{U}_\perp] = [\cos(\Theta) \quad \sin(\Theta) \quad \mathbf{0}_{d \times d-2r}],$$

$$\begin{aligned} \text{Tr}(\mathbf{U}_t^\top \mathbf{U} \Lambda \mathbf{U}^\top \mathbf{U}_t) &= \text{Tr}\left([\cos(\Theta) \quad \sin(\Theta) \quad \mathbf{0}_{d \times (d-2r)}] \begin{bmatrix} \nu_1 \mathbf{I}_r & & \\ & \nu_2 \mathbf{I}_r & \\ & & \nu_3 \mathbf{I}_{d-2r} \end{bmatrix} \begin{bmatrix} \cos(\Theta) \\ \sin(\Theta) \\ \mathbf{0}_{(d-2r) \times d} \end{bmatrix}\right) \\ &= \text{Tr}\left(\begin{bmatrix} \nu_1 \cos^2(\Theta) & & \\ & \nu_2 \sin^2(\Theta) & \\ & & \mathbf{0}_{(d-2r) \times (d-2r)} \end{bmatrix}\right) = r \cdot \nu_1 \cdot \cos^2(\theta) + r \cdot \nu_2 \cdot \sin^2(\theta), \end{aligned}$$

where we used the fact that the principal angles are all equal to θ . Using a similar argument,

$$\text{Tr}(\Sigma_t^\top \mathbf{A}^2) = r \cdot \nu_1^2 \cdot \cos^2(\theta) + r \cdot \nu_2^2 \cdot \sin^2(\theta) + \epsilon \text{Tr}(\mathbf{A}^2)$$

Simplifying the Test Risk. Substituting the expressions for the $\text{Tr}(\cdot)$ terms into Equation (21) yields

$$\begin{aligned} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] &= \left(\frac{1}{m} (r\nu_1^2 + r\nu_2^2 + (d - 2r)\nu_3^2) + 1 \right) (r + \epsilon d + \sigma^2) \\ &\quad - 2(r\nu_1 \cos^2(\theta) + r\nu_2 \sin^2(\theta) + (r\nu_1 + r\nu_2 + (d - 2r)\nu_3) \epsilon) \\ &\quad + \frac{m+1}{m} (r\nu_1^2 \cos^2(\theta) + r\nu_2^2 \sin^2(\theta) + (r\nu_1^2 + r\nu_2^2 + (d - 2r)\nu_3^2) \epsilon). \end{aligned}$$

Then, taking $\epsilon \rightarrow 0$:

$$\begin{aligned} \lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] &= \\ &\left(\frac{1}{m} \left(r \left(\frac{\gamma n}{\gamma(n+1) + r + \sigma^2} \right)^2 + r \left(\frac{(1-\gamma)n}{(1-\gamma)(n+1) + r + \sigma^2} \right)^2 + 1 \right) \right) (r + \sigma^2) \\ &\quad - 2 \left(\frac{r\gamma n \cos^2(\theta)}{\gamma(n+1) + r + \sigma^2} + \frac{r(1-\gamma)n \sin^2(\theta)}{(1-\gamma)(n+1) + r + \sigma^2} \right) \\ &\quad + \frac{m+1}{m} \left(r \cos^2(\theta) \left(\frac{\gamma n}{\gamma(n+1) + r + \sigma^2} \right)^2 + r \sin^2(\theta) \left(\frac{(1-\gamma)n}{(1-\gamma)(n+1) + r + \sigma^2} \right)^2 \right). \end{aligned}$$

Substituting $\gamma = 0.5$ and combining like terms yields

$$\lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] = r + \sigma^2 + \frac{m+1+2(r+\sigma^2)}{m} \cdot \frac{rn^2}{(n+1+2(r+\sigma^2))^2} - \frac{2rn}{n+1+2(r+\sigma^2)}.$$

Now suppose $n \leq m$. Then, we have

$$\lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] \leq r + \sigma^2 - \frac{rn}{n+1+2(r+\sigma^2)}$$

Upper bounding this by $\sigma^2 + \delta$ for some $\delta \in (0, r)$, then solving for n , yields the following result. For any $\delta \in (0, r)$, if

$$m \geq n > \frac{(2(r+\sigma^2)+1)r}{\delta} - (2(r+\sigma^2)+1),$$

then $\lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] < \sigma^2 + \delta$, which completes the proof. \square

G.1.4 Proof of Theorem 3

Proof. The proof is similar to that of Theorem 2. Again let $\tilde{y} := \tilde{y}_{m+1}$. By Lemma 1 and Lemma 4, we have

$$\mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] = \left(\frac{1}{m} \text{Tr}(\mathbf{A}^\top \mathbf{A}) + 1 \right) \left(\text{Tr}(\bar{\Sigma}_t) + \sigma^2 \right) - 2 \text{Tr}(\bar{\Sigma}_t \mathbf{A}) + \frac{m+1}{m} \text{Tr}(\mathbf{A} \bar{\Sigma}_t \mathbf{A}^\top), \quad (23)$$

where $\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \boldsymbol{\Sigma}^{-1} \right)^{-1}$, $M_s = \text{Tr}(\boldsymbol{\Sigma}) + \sigma^2$, and $\boldsymbol{\Sigma} = \sum_{k=1}^K \gamma_k \boldsymbol{\Sigma}_{s,k}$.

Let $\mathbf{U} := [\mathbf{U}_{s,1} \ \mathbf{U}_{s,2} \ \dots \ \mathbf{U}_{s,K} \ \mathbf{U}_\perp]$, where $\mathbf{U}_\perp \in \mathbb{R}^{d \times (d-Kr)}$ completes the orthonormal basis for \mathbb{R}^d . By Lemma 5,

$$\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \boldsymbol{\Sigma}^{-1} \right)^{-1} = \mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^\top,$$

where

$$\boldsymbol{\Lambda} = \begin{bmatrix} \nu_1 \mathbf{I}_r & & & \\ & \ddots & & \\ & & \nu_K \mathbf{I}_r & \\ & & & \nu_{K+1} \mathbf{I}_{d-Kr} \end{bmatrix}$$

with $\nu_k = \frac{n(\gamma_k + \epsilon)}{(n+1)(\gamma_k + \epsilon) + M_s}$ for all $k \in [K]$, and $\nu_{K+1} = \frac{n\epsilon}{(n+1)\epsilon + r + \epsilon d + \sigma^2}$.

Simplifying $\text{Tr}(\bar{\Sigma}_t)$. We can write $\text{Tr}(\bar{\Sigma}_t)$ as such:

$$\text{Tr}(\bar{\Sigma}_t) = \text{Tr}(\bar{\mathbf{U}}_t \bar{\mathbf{U}}_t^\top) + \epsilon \text{Tr}(\mathbf{I}_d) = r + \epsilon d.$$

Simplifying $\text{Tr}(\mathbf{A})$ and $\text{Tr}(\mathbf{A}^\top \mathbf{A})$. We can write $\text{Tr}(\mathbf{A})$ and $\text{Tr}(\mathbf{A}^\top \mathbf{A})$ as such:

$$\text{Tr}(\mathbf{A}) = r \sum_{k=1}^K \nu_k + (d - Kr) \nu_{K+1} \quad \text{and} \quad \text{Tr}(\mathbf{A}^\top \mathbf{A}) = \text{Tr}(\mathbf{A}^2) = r \sum_{k=1}^K \nu_k^2 + (d - Kr) \nu_{K+1}^2.$$

Simplifying $\text{Tr}(\bar{\Sigma}_t \mathbf{A})$ and $\text{Tr}(\mathbf{A} \bar{\Sigma}_t \mathbf{A}^\top)$. Note $\text{Tr}(\mathbf{A} \bar{\Sigma}_t \mathbf{A}^\top) = \text{Tr}(\bar{\Sigma}_t \mathbf{A}^2)$. We first focus on $\text{Tr}(\bar{\Sigma}_t \mathbf{A})$:

$$\begin{aligned} \bar{\Sigma}_t \mathbf{A} &= (\bar{\mathbf{U}}_t \bar{\mathbf{U}}_t^\top + \epsilon \mathbf{I}_d) \mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^\top = \bar{\mathbf{U}}_t \bar{\mathbf{U}}_t^\top \mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^\top + \epsilon \mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^\top \\ \implies \text{Tr}(\bar{\Sigma}_t \mathbf{A}) &= \text{Tr}(\bar{\mathbf{U}}_t^\top \mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^\top \bar{\mathbf{U}}_t) + \epsilon \text{Tr}(\mathbf{A}). \end{aligned}$$

Recall $\bar{\mathbf{U}}_t = \sum_{k=1}^K \alpha_k \mathbf{U}_{s,k}$ where $\sum_{k=1}^K \alpha_k^2 = 1$, and so we have

$$\bar{\mathbf{U}}_t^\top \mathbf{U} = \left(\sum_{k=1}^K \alpha_k \mathbf{U}_k \right)^\top [\mathbf{U}_{s,1} \ \dots \ \mathbf{U}_{s,K} \ \mathbf{U}_\perp] = \begin{bmatrix} \alpha_1 \mathbf{I}_r & & & \\ & \ddots & & \\ & & \alpha_K \mathbf{I}_r & \\ & & & \mathbf{0}_{(d-Kr) \times (d-Kr)} \end{bmatrix}$$

Thus,

$$\text{Tr}(\bar{\mathbf{U}}_t^\top \mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^\top \bar{\mathbf{U}}_t) = \text{Tr} \left(\begin{bmatrix} \alpha_1^2 \nu_1 \mathbf{I}_r & & & \\ & \ddots & & \\ & & \alpha_K^2 \nu_K \mathbf{I}_r & \\ & & & \mathbf{0}_{(d-Kr) \times (d-Kr)} \end{bmatrix} \right) = r \sum_{k=1}^K \alpha_k^2 \nu_k$$

Using a similar argument,

$$\text{Tr}(\bar{\Sigma}_t^\top \mathbf{A}^2) = r \sum_{k=1}^K \alpha_k^2 \nu_k^2 + \epsilon \text{Tr}(\mathbf{A}^2).$$

Simplifying the test risk. Substituting the expressions for the $\text{Tr}(\cdot)$ terms into Equation (23) yields

$$\begin{aligned}\mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] &= \left(\frac{1}{m} \left(r \sum_{k=1}^K \nu_k^2 + (d - Kr) \nu_{K+1}^2 \right) + 1 \right) (r + \epsilon d + \sigma^2) \\ &\quad - 2 \left(r \sum_{k=1}^K \alpha_k^2 \nu_k + \left(r \sum_{k=1}^K \nu_k + (d - Kr) \nu_{K+1} \right) \epsilon \right) \\ &\quad + \frac{m+1}{m} \left(r \sum_{k=1}^K \alpha_k^2 \nu_k^2 + \left(r \sum_{k=1}^K \nu_k^2 + (d - Kr) \nu_{K+1}^2 \right) \epsilon \right).\end{aligned}$$

Taking $\epsilon \rightarrow 0$ results in the following expression for the test risk:

$$\begin{aligned}\lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] &= r + \sigma^2 + \frac{(r + \sigma^2)r}{m} \sum_{k=1}^K \left(\frac{\gamma_k n}{\gamma_k(n+1) + r + \sigma^2} \right)^2 \\ &\quad - 2r \sum_{k=1}^K \frac{\alpha_k^2 \gamma_k n}{\gamma_k(n+1) + r + \sigma^2} + \frac{(m+1)r}{m} \sum_{k=1}^K \left(\frac{\alpha_k \gamma_k n}{\gamma_k(n+1) + r + \sigma^2} \right)^2\end{aligned}$$

Substituting $\gamma_k = \frac{1}{K}$ for all $k \in [K]$ and combining like terms yields

$$\lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] = r + \sigma^2 + \frac{m+1+K(r+\sigma^2)}{m} \cdot \frac{rn^2}{(n+1+K(r+\sigma^2))^2} - \frac{2rn}{n+1+K(r+\sigma^2)}.$$

Now suppose $n \leq m$. Then, we have

$$\lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] \leq r + \sigma^2 - \frac{rn^2}{n+1+K(r+\sigma^2)}.$$

Upper bounding this by $\sigma^2 + \delta$ for some $\delta \in (0, r)$, then solving for n , yields the following result. For any $\delta \in (0, r)$, if

$$m \geq n > \frac{(K(r+\sigma^2)+1)r}{\delta} - (K(r+\sigma^2)+1),$$

then $\lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\left(\tilde{y} - g_{\text{ATT}}^*(\tilde{\mathbf{z}}_q, \tilde{\mathbf{Z}}_{\mathcal{M}}) \right)^2 \right] < \sigma^2 + \delta$, which completes the proof. \square

G.2 Auxiliary Results

G.2.1 Optimal Linear Attention Weights

We first provide results on the form of the weights matrices after training a single-layer linear attention model on the loss in Equation (10). The following results are largely inspired by Theorem 1 in [15], but are slightly different since we consider a normalization factor of $1/n$ in our linear attention model.

Lemma 2 (Optimal Attention Weights [15]). *Consider the independent data model in Equation (1) where the task vector is drawn from $\mathbf{w} \sim \mathcal{N}(\mathbf{0}, \Sigma_s)$, and let $n \in \mathbb{N}$ denote the in-context prompt length used at training. Then, the optimal linear attention weights obtained by minimizing the loss in Equation (10) are given by*

$$\mathbf{W}_K^* = \mathbf{W}_V^* = \mathbf{I}_{d+1}, \quad \mathbf{W}_Q^* = \begin{bmatrix} \mathbf{A} & \mathbf{0}_d \\ \mathbf{0}_d^\top & 0 \end{bmatrix}, \quad \text{and} \quad \mathbf{v}^* = \begin{bmatrix} \mathbf{0}_d \\ 1 \end{bmatrix}, \quad (24)$$

where $\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \Sigma_s^{-1} \right)^{-1}$ and $M_s = \text{Tr}(\Sigma_s) + \sigma^2$, with empirical risk $\mathcal{L}_s^* = M_s - \text{Tr}(\Sigma_s \mathbf{A})$.

Proof. The proof is the same as that of Theorem 1 in [15] by absorbing the $1/n$ factor into \mathbf{W}_Q . \square

Lemma 3 (Optimal Attention Weights for Mixture of 2 Gaussians). *Consider the independent data model in Equation (1) where the task vector is drawn from $\mathbf{w} \sim \gamma \cdot \mathcal{N}(\mathbf{0}, \Sigma_s) + (1-\gamma) \cdot \mathcal{N}(\mathbf{0}, \Sigma_{s,\perp})$ for some $\gamma \in (0, 1)$. Let $n \in \mathbb{N}$ denote the in-context prompt length used at training. Define $\Sigma = \gamma \cdot \Sigma_s + (1-\gamma) \cdot \Sigma_{s,\perp}$. Then, the optimal linear attention weights obtained by minimizing the loss in Equation (10) are given by*

$$\mathbf{W}_K^* = \mathbf{W}_V^* = \mathbf{I}_{d+1}, \quad \mathbf{W}_Q^* = \begin{bmatrix} \mathbf{A} & \mathbf{0}_d \\ \mathbf{0}_d^\top & 0 \end{bmatrix}, \quad \text{and} \quad \mathbf{v}^* = \begin{bmatrix} \mathbf{0}_d \\ 1 \end{bmatrix}, \quad (25)$$

where $\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \Sigma^{-1} \right)^{-1}$ and $M_s = \text{Tr}(\Sigma) + \sigma^2$, with empirical risk $\mathcal{L}_s^* = M_s - \text{Tr}(\Sigma \mathbf{A})$.

Proof. It is straightforward to see that if $\mathbf{w} \sim \gamma \cdot \mathcal{N}(\mathbf{0}, \Sigma_s) + (1-\gamma) \cdot \mathcal{N}(\mathbf{0}, \Sigma_{s,\perp})$, then

$$\Sigma := \text{Cov}(\mathbf{w}) = \gamma \cdot \Sigma_s + (1-\gamma) \cdot \Sigma_{s,\perp}.$$

Then, the proof from Lemma 2 follows verbatim by using Σ instead of Σ_s . \square

Lemma 4 (Optimal Attention Weights for Mixture of K Gaussians). *Consider the independent data model in Equation (1) where the task vector is drawn from $\mathbf{w} \sim \sum_{k=1}^K \gamma_k \cdot \mathcal{N}(\mathbf{0}, \Sigma_{s,k})$ with $\gamma_k \in (0, 1)$*

for all $k \in [K]$ and $\sum_{k=1}^K \gamma_k = 1$. Let $n \in \mathbb{N}$ denote the in-context prompt length used at training.

Define $\Sigma = \sum_{k=1}^K \gamma_k \cdot \Sigma_{s,k}$. Then, the optimal linear attention weights obtained by minimizing the loss in Equation (10) are given by

$$\mathbf{W}_K^* = \mathbf{W}_V^* = \mathbf{I}_{d+1}, \quad \mathbf{W}_Q^* = \begin{bmatrix} \mathbf{A} & \mathbf{0}_d \\ \mathbf{0}_d^\top & 0 \end{bmatrix}, \quad \text{and} \quad \mathbf{v}^* = \begin{bmatrix} \mathbf{0}_d \\ 1 \end{bmatrix}, \quad (26)$$

where $\mathbf{A} = \left(\frac{n+1}{n} \mathbf{I}_d + \frac{M_s}{n} \Sigma^{-1} \right)^{-1}$ and $M_s = \text{Tr}(\Sigma) + \sigma^2$, with empirical risk $\mathcal{L}_s^* = M_s - \text{Tr}(\Sigma \mathbf{A})$.

Proof. The proof is equivalent to that of Lemma 3 by letting $\Sigma = \sum_{k=1}^K \gamma_k \Sigma_{s,k}$ instead. \square

G.2.2 Miscellaneous Results

Lemma 5. *Let $0 \prec \Sigma \in \mathbb{R}^{d \times d}$ and $c, k > 0$ be constants. Then,*

$$(c \cdot \mathbf{I}_d + k \cdot \Sigma^{-1})^{-1} = \mathbf{V} \begin{bmatrix} \frac{\lambda_1}{c \cdot \lambda_1 + k} & 0 & \dots & 0 \\ 0 & \frac{\lambda_2}{c \cdot \lambda_2 + k} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \frac{\lambda_d}{c \cdot \lambda_d + k} \end{bmatrix} \mathbf{V}^\top, \quad (27)$$

where $\mathbf{V} \in \mathbb{R}^{d \times d}$ is an orthonormal matrix whose columns are eigenvectors of Σ , and λ_i is the i^{th} largest eigenvalue of Σ .

Proof. Since $\Sigma \succ 0$, there exists an eigendecomposition $\Sigma = \mathbf{V} \Lambda \mathbf{V}^\top$ such that \mathbf{V} is an orthonormal matrix and Λ is a diagonal matrix consisting of the real, positive eigenvalues of Σ , denoted as $\lambda_1, \lambda_2, \dots, \lambda_d$. Thus,

$$\begin{aligned} \Sigma^{-1} = \mathbf{V} \Lambda^{-1} \mathbf{V}^\top &\implies c \cdot \mathbf{I}_d + k \cdot \Sigma^{-1} = \mathbf{V} \underbrace{\begin{bmatrix} c + \frac{k}{\lambda_1} & 0 & \dots & 0 \\ 0 & c + \frac{k}{\lambda_2} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & c + \frac{k}{\lambda_d} \end{bmatrix}}_{\tilde{\Lambda}} \mathbf{V}^\top \\ &\implies (c \cdot \mathbf{I}_d + k \cdot \Sigma^{-1})^{-1} = \mathbf{V} \tilde{\Lambda}^{-1} \mathbf{V}^\top, \end{aligned}$$

which completes the proof. \square

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