Discrete Mixture Models

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1 Introduction

This project deals with estimating a mixture model for two multimodal distributions for M different categories:

$$p(k_1, \cdots, k_M | p_1, \cdots, p_M) = \frac{n!}{\prod_l k_l!} \prod_{l=1}^M p_l^{k_l}, \qquad \sum_l^M p_l = 1 \text{ and } \sum_l^M k_l = n$$

$$p(k_1, \cdots, k_M | q_1, \cdots, q_M) = \frac{n!}{\prod_l k_l!} \prod_{l=1}^M q_l^{k_l}, \qquad \sum_l^M q_l = 1 \text{ and } \sum_l^M k_l = n$$

Consider a mixing process of iid realizations from each process $\{k_i(1)^{(p)}\}, \cdots, \{k_i(N)^{(p)}\}$ from p and $\{k_i(1)^{(q)}\}, \cdots, \{k_i(N)^{(q)}\}$ from q. At each time instance $t=1, \cdots, N$, we observe $\{k_i(t)^{(p)}\}$ with probability α and $\{k_i(t)^{(q)}\}$ with probability $(1-\alpha)$. In this project, we focus on M=2, the binomial case.

Because M=2, let $k_2=n-k_1$, $p_1=p$, $p_2=1-p$, $q_1=q$, $q_2=1-q$. Then, the probability mass function for the first bonomial distribution is given by:

$$\phi_1(k_1(i)) = p(k_1(i)|p) = \binom{n}{k_1(i)} p^{k_1(i)} (1-p)^{n-k_1(i)}$$

The probability mass function for the second binomial distribution is given by:

$$\phi_0(k_1(i)) = p(k_1(i)|q) = \binom{n}{k_1(i)} q^{k_1(i)} (1-q)^{n-k_1(i)}$$

Assuming a fixed observation length n at every time instance, the probability mass function for the mixture model is given by:

$$p(k_1(1), \dots, k_1(N) | \alpha, p, q) = \prod_{i=1}^{N} \left[\alpha \binom{n}{k_1(i)} p^{k_1(i)} (1-p)^{n-k_1(i)} + (1-\alpha) \binom{n}{k_1(i)} q^{k_1(i)} (1-q)^{n-k_1(i)} \right]$$
(1)

The parameter vector for this problem is given by:

$$\theta = \begin{bmatrix} \alpha \\ p \\ q \end{bmatrix}$$

2 FIM and CRLB

3 Maximum Likelihood and Expectation-Maximization

3.1 Maximum Likelihood Equations

Given Equation 1, let the likelihood be

$$L(\theta) = p(k_1(1), \cdots, k_1(N)|\theta)$$

Then, the log-likelihood is

$$\log L(\theta) = \sum_{i=1}^{N} \log \left[\alpha \binom{n}{k_1(i)} p^{k_1(i)} (1-p)^{n-k_1(i)} + (1-\alpha) \binom{n}{k_1(i)} q^{k_1(i)} (1-q)^{n-k_1(i)} \right]$$

The maximum log-likelihood is given by

$$\hat{\theta} = \operatorname*{argmax}_{\theta} \log L(\theta)$$

Therefore, the maximum likelihood equations are given by:

$$\frac{\delta \log p}{\delta \theta} = \begin{bmatrix} \sum_{i=1}^{N} \frac{\phi_1(k_1(i)) - \phi_0(k_1(i))}{\alpha \phi_1(k_1(i)) + (1 - \alpha) \phi_0(k_1(i))} \\ \sum_{i=1}^{N} \frac{\alpha \binom{n}{k_1(i)} \binom{k_1(i)p^{k_1(i) - 1}(1 - p)^{n - k_1(i)} - p^{k_1(i)}(n - k_1(i))(1 - p)^{n - k_1(i) - 1}}{\alpha \phi_1(k_1(i)) + (1 - \alpha) \phi_0(k_1(i))} \\ \sum_{i=1}^{N} \frac{(1 - \alpha) \binom{n}{k_1(i)} \binom{k_1(i)p^{k_1(i) - 1}(1 - q)^{n - k_1(i)} - q^{k_1(i)}(n - k_1(i))(1 - q)^{n - k_1(i) - 1}}{\alpha \phi_1(k_1(i)) + (1 - \alpha) \phi_0(k_1(i))} \end{bmatrix} = \mathbf{0}$$

3.2 EM Algorithm

Let $\mathbf{k_1} = k_1(1), \cdots, k_1(N)$ be the observed data. We introduce membership variables $\mathbf{y} = y(i), \cdots, y(N)$ (hidden data) such that $p(y(i) = c) = \alpha_c$. Because we only have two classes, then

$$p(y(i) = 1) = \alpha$$

 $p(y(i) = 0) = 1 - p(y(i) = 1) = 1 - \alpha$

The joint probability mass function is given by:

$$p(\mathbf{k_1}, \mathbf{y}|\theta) = \prod_{i=1}^{N} p(k_1(i), y(i)|\theta)$$
$$= \prod_{i=1}^{N} p(k_1(i)|y(i))p(y(i)|\theta)$$
$$= \prod_{i=1}^{N} \prod_{l=1}^{M} (\phi_l(k_1(i)\alpha_l))^{I(y(i)=l)}$$

Let Q be an auxiliary function such that

$$Q(\theta, \theta') = \operatorname{E} \left[\log p(\mathbf{k_1}, \mathbf{y} | \theta) | \mathbf{k_1}, \theta' \right]$$

$$= \operatorname{E} \left[\sum_{i=1}^{N} \sum_{l=1}^{M} I(y(i) = l) (\log \phi_l(k_1(i)) + \log \alpha_l) | \mathbf{k_1}, \theta' \right]$$

$$= \sum_{i=1}^{N} \sum_{l=1}^{M} \operatorname{E} \left[I(y(i) = l) | \mathbf{k_1}, \theta' \right] (\log \phi_l(k_1(i)) + \log \alpha_l)$$

$$= \sum_{i=1}^{N} \sum_{l=1}^{M} p(y(i) = c | k_1(i), \theta') (\log \phi_l(k_1(i)) + \log \alpha_l)$$

In the E-step of the EM algorithm, we compute $p(y(i) = c | k_1(i), \theta')$. In our case of M = 2 the E-step is given by:

$$p(y(i) = 1 | k_1(i), \theta') = \frac{\alpha \binom{n}{k_1(i)} p^{k_1(i)} (1 - p)^{n - k_1(i)}}{\alpha \binom{n}{k_1(i)} p^{k_1(i)} (1 - p)^{n - k_1(i)} + (1 - \alpha) \binom{n}{k_1(i)} q^{k_1(i)} (1 - q)^{n - k_1(i)}}$$

$$p(y(i) = 0 | k_1(i), \theta') = \frac{(1 - \alpha) \binom{n}{k_1(i)} q^{k_1(i)} (1 - q)^{n - k_1(i)}}{\alpha \binom{n}{k_1(i)} p^{k_1(i)} (1 - p)^{n - k_1(i)} + (1 - \alpha) \binom{n}{k_1(i)} q^{k_1(i)} (1 - q)^{n - k_1(i)}}$$

In the M-step, we compute

$$\theta^{k+1} = \underset{\theta}{\operatorname{argmax}} Q(\theta, \theta^k)$$

$$= \underset{\theta}{\operatorname{argmax}} \sum_{i=1}^{N} \sum_{l=1}^{M} p(y(i) = c | k_1(i), \theta^k) (\log \phi_l(k_1(i)) + \log \alpha_l)$$

Then, we get:

$$\begin{split} &\alpha_1^{k+1} = \alpha^{k+1} = \frac{1}{N} \sum_{i=1}^N p(y(i) = 1 | k_1(i), \theta^k) \\ &\alpha_0^{k+1} = 1 - \alpha^{k+1} \\ &p^{k+1} = \frac{\sum_{i=1}^N p(y(i) = 1 | k_1(i), \theta^k) k_1(i)}{n \sum_{i=1}^N p(y(i) = 1 | k_1(i), \theta^k)} \\ &q^{k+1} = \frac{\sum_{i=1}^N p(y(i) = 0 | k_1(i), \theta^k) k_1(i)}{n \sum_{i=1}^N p(y(i) = 0 | k_1(i), \theta^k)} \end{split}$$

4 Method of Moments