Monday, Apr 24

Cramer's V

Consider the following data where a sample of 1398 children were classified with respect to tonsil size and carrier status of *Streptococcus pyogenes*.¹

	Strep				
Size	yes	no	Total		
small	19 (0.04)	497 (0.96)	516		
mediun	n29 (0.05)	560 (0.95)	589		
large	24 (0.08)	269 (0.92)	293		
Total	72 (0.05)	1326 (0.95)	1398		

The numbers in parentheses are the *proportions* of children of each tonsil size who are or are not carriers. The value of the test statistic for a test of independence is $X^2 \approx 7.88$. We might decide tonsil size and carrier status are dependent. But can we *measure* the amount of dependence?

Cramer's V is a measure of association between two categorical variables. It is defined as

$$V = \sqrt{\frac{X^2/n}{\min(r-1, c-1)}},$$

and is bounded such that $0 \le V \le 1$. It effectively measures the degree to which the observed counts deviate from the expected counts under the assumption of independence, and thus can be viewed as a measurement of the degree of dependence.

Example: For the data on tonsil size and carrier status, the value of Cramer's V is

$$\sqrt{\frac{7.88/1398}{\min(3-1,2-1)}} \approx 0.08,$$

which is a relatively weak association.

¹Holmes, M. C. & Williams, R. E. O. (1954). The distribution of carriers of Streptococcus pyogenes among 2413 healthy children. *Journal of Hygiene*, 52, 165–179.

Here are some hypothetical observed counts showing a stronger association.

Strep				
Size	yes	no	Total	
small	25 (0.05)	475 (0.95)	500	
mediun	n 90 (0.15)	510 (0.85)	600	
large	75 (0.25)	225 (0.75)	300	
Total	190 (0.14)	1210 (0.86)	1400	

The test statistic is $X^2 \approx 65.77$ and the measure of association is $V \approx 0.22$. And here are some hypothetical observed counts showing an *even stronger* association.

	St	rep	
Size	yes	no	Total
small	25 (0.05)	475 (0.95)	500
mediun	n 300 (0.5)	300 (0.5)	600
large	285 (0.95)	15 (0.05)	300
Total	610 (0.44)	790 (0.56)	1400

The test statistic is $X^2 \approx 635.36$ and the measure of association is $V \approx 0.67$. Here is an example of the maximum degree of association.

Strep				
Size	yes	no	Total	
small	0 (0)	500 (1)	500	
medium	0(0)	600 (1)	600	
large	300 (1)	0 (0)	300	
Total 3	300 (0.21)	1100 (0.79) 1400	

The test statistic is $X^2=1400$ and the measure of association is V=1. What about the minimum degree of association?

	Strep				
Size	yes	no	Total		
small	25 (0.05)	475 (0.95)	500		
mediun	n30 (0.05)	570 (0.95)	600		
large	15 (0.05)	285 (0.95)	300		
Total	70 (0.05)	1330 (0.95)	1400		

The test statistic is $X^2 = 0$ and the measure of association is V = 0. Here the observed counts would be equal to the expected counts.

Example: Consider the following data from a randomized experiment comparing two strategies for chemotherapy.²

Tumor Response					
Strategy pr	ogressive disea	seno changepa	artial remission	complete remissic	nTotal
sequential	32 (0.21)	57 (0.38)	34 (0.23)	28 (0.19)	151
alternating	$53\ (0.36)$	51 (0.34)	23(0.16)	21 (0.14)	148
Total	85 (0.28)	108 (0.36)	57 (0.19)	49 (0.16)	299

The test statistic is $X^2 \approx 8.62$ and the measure of association is $V \approx 0.17$.

Here is a weaker association with some hypothetical observed counts.

Tumor Response					
Strategy pr	ogressive disea	seno changepa	artial remission	complete remissio	nTotal
sequential	40 (0.2)	80 (0.4)	50 (0.25)	30 (0.15)	200
alternating	$50 \ (0.25)$	84 (0.42)	42 (0.21)	24 (0.12)	200
Total	90 (0.22)	164 (0.41)	92 (0.23)	54 (0.14)	400

The test statistic is $X^2 \approx 2.57$ and the measure of association is $V \approx 0.08$.

Here is a stronger association with some hypothetical observed counts.

	Tumor Response					
Strategy pr	rogressive disea	seno changepa	artial remission	ncomplete remission	nTotal	
sequential	60 (0.3)	100 (0.5)	20 (0.1)	20 (0.1)	200	
alternating	20 (0.1)	40(0.2)	100 (0.5)	40 (0.2)	200	
Total	80 (0.2)	$140 \ (0.35)$	120 (0.3)	$60 \ (0.15)$	400	

The test statistic is $X^2 \approx 105.71$ and the measure of association is $V \approx 0.51$.

Here is a very strong association with some hypothetical observed counts.

Tumor Response					
Strategy pr	ogressive disea	seno changep	artial remissio	ncomplete remission	nTotal
sequential	90 (0.45)	100 (0.5)	6 (0.03)	4 (0.02)	200
alternating	2(0.01)	18 (0.09)	100 (0.5)	80 (0.4)	200
Total	92 (0.23)	118 (0.3)	$106 \ (0.26)$	84 (0.21)	400

The test statistic is $X^2 \approx 293.28$ and the measure of association is $V \approx 0.86$.

²Holtbrugge, W. & Schumacher, M. (1991). A comparison of regression models for the analysis of ordered categorical data. *Applied Statistics*, 40, 249–259.

McNemar's Test for Matched Pairs

Example: A retrospective case-control study was used to investigate the theory that tonsils protect the body against the invasion of the lymph nodes by the virus responsible for Hodgkin's disease. The study compared patients with Hodgkin's disease (the cases) with their siblings without the disease (the controls) with respect to whether or not they had a tonsillectomy in the past.³ Here are the first ten sibling pairs.

Pai	rPatient (Case) Sibling (Control)
1	no tonsillectomytonsillectomy
2	no tonsillectomy no tonsillectomy $% \left(-\frac{1}{2}\right) =-\frac{1}{2}\left($
3	no tonsillectomyno tonsillectomy
4	tonsillectomy tonsillectomy
5	no tonsillectomyno tonsillectomy
6	tonsillectomy no tonsillectomy
7	no tonsillectomyno tonsillectomy
8	tonsillectomy no tonsillectomy
9	no tonsillectomyno tonsillectomy
10	tonsillectomy tonsillectomy

Each *pair* can be classified in terms of whether or not the patient (case) had a tonsillectomy, and whether or not the sibling (control) had a tonsillectomy.

	Sibling (C	Control)	
Patient (Case)	tonsillectomyno	tonsillecto	myTotal
tonsillectomy	26	15	41
no tonsillectomy	7	37	44
Total	33	52	85

Let p_p be the probability that the patient (case) had a tonsillectomy, and let p_s be the probability that the sibling (control) had a tonsillectomy. How can we test the null hypothesis $H_0: p_p = p_s$ versus $H_a: p_p \neq p_s$? It would be tempting to use the test statistic

$$z = \frac{\hat{p}_p - \hat{p}_s}{\sqrt{\hat{p}(1-\hat{p})(1/n_p + 1/n_s)}},$$

where $\hat{p}_p = 41/85$, $\hat{p}_s = 33/85$, $n_p = 85$, $n_s = 85$, and $\hat{p} = (41 + 33)/(85 + 85)$. However this test statistic assumes that the samples are *independent*, but they are not independent.

³ Johnson, S. K. & Johnson, R. E. (1972). Tonsillectomy history in Hodgkin's disease. *New England Journal of Medicine*, 287, 1122–1125.

Derivation of McNemar's Test Statistic

Let p_a , p_b , p_c , and p_d denote the probabilities of each of the four possible sibling pairs.

Patient (Case)	tonsillectomyno	tonsillectomy
tonsillectomy	\$p_a\$	\$p_b\$
no tonsillectomy	\$p_c\$	p_d

So the probability that the *patient* had a tonsillectomy is

$$p_p = p_a + p_b,$$

and the probability that the sibling had a tonsillectomy is

$$p_s = p_a + p_c.$$

If the null hypothesis is true then

$$p_p = p_s \Rightarrow p_a + p_b = p_a + p_c \Rightarrow p_b = p_c.$$

Now we don't know p_b or p_c , but we can estimate them from the observed counts,

Sibling (Control)				
Patient (Case)	tonsillectomy	no tonsillector	nyTotal	
tonsillectomy	26	15	41	
no tonsillectomy	7	37	44	
Total	33	52	85	

The estimates of p_b and p_c are obtained by averaging the corresponding proportions because we assume that $p_b = p_c$ under the null hypothesis so that

$$\hat{p}_b = \frac{7/85 + 15/85}{2}, \quad \hat{p}_c = \frac{7/85 + 15/85}{2}.$$

Now these estimates can be used to compute two of the (estimated) expected counts corresponding to the bottom-left and top-right cells.

$$n \times \hat{p}_b = 85 \times \frac{7/85 + 15/85}{2} = \frac{7 + 15}{2} = 11, n \times \hat{p}_c = 85 \times \frac{7/85 + 15/85}{2} = \frac{7 + 15}{2} = 11.$$

The null hypothesis does not imply anything about the expected counts for the top-left and top-right cells, so we just use the observed counts as estimates of the expected counts in those cells. The expected counts for each cell are shown in the table below.

Sibling (Control)				
Patient (Case) tonsillectomyno tonsillectomyTotal				
tonsillectomy	26	(7+15)/2	41	
no tonsillectomy	(7+15)/2	37	44	
Total	33	52	85	

Now plugging the observed and expected counts into the formula for X^2 gives us

$$X^{2} = \frac{(26 - 26)^{2}}{26} + \frac{[15 - (7 + 15)/2]^{2}}{(7 + 15)/2} + \frac{[7 - (7 + 15)/2]^{2}}{(7 + 15)/2} + \frac{(37 - 37)^{2}}{37} \approx 2.91.$$

Some algebra will show that this can be simplified considerably to

$$X^2 = \frac{(7-15)^2}{7+15} \approx 2.91.$$

In general, we can write the test statistic as

$$X^2 = \frac{(O_{bl} - O_{tr})^2}{O_{bl} + O_{tr}}$$

where O_{bl} and O_{tr} denote the bottom-left and top-right observed counts, respectively. The degrees of freedom for computing the p-value is always 1.

Example: An enzyme-linked immunosorbent assay (ELISA) is an analytical biochemical procedure that can be used to detect the presence of antigens or antibodies, and so so it can be used to detect the presence of specific infections. A study applied two kinds of ELISA — a standard version and the ABC-ELISA — to each of 101 patients with hydatidosis (i.e., an infestation with *echinococcus*, a genus of tapeworms).⁴ Each test will give a *positive* or *negative* test result for the presence of the disease. Let $p_{\rm abc}$ and $p_{\rm s}$ be the probability that a ABC-ELISA and standard ELISA, respectively, will produce a positive result when applied to someone with the disease (this probability is called the *sensitivity* of the test). To determine if the two tests differ with respect to their sensitivity we could test the hypotheses $H_0: p_{\rm abc} - p_{\rm s} = 0$ versus $H_a: p_{\rm abc} - p_{\rm s} \neq 0$ using the test statistic

$$z = \frac{\hat{p}_{abc} - \hat{p}_{s}}{\sqrt{\hat{p}(1-\hat{p})(1/n_{abc} + 1/n_{s})}}.$$

This is what was done in the original analysis, but it was later pointed out that this analysis is incorrect because the two samples are *dependent* because both assays were applied to the same patients.⁵ So how can we test the hypotheses $H_0: p_{abc} - p_s = 0$ versus $H_a: p_{abc} - p_s \neq 0$?

Standard ELISA				
$\overline{\rm ABC\text{-}ELISA positive negative Total}$				
positive	82	13	95	
negative	6	0	6	
Total	88	13	101	

⁴Shen, Z. Q., Feng, X. H., Qian, Z. X., Liu, R. L., & Yang, C. R. (1988). Application of biotinadvin system, determination of circulating immune complexes, and evaluation of antibody response in different hydatidosis patients. *American Journal of Tropical Medicine and Hygiene.* 39, 93–96.

⁵Cruess, D. F. (1989). Review of use of statistics in The American Journal of Tropical Medicine and Hygiene for January-December 1988. American Journal of Tropical Medicine and Hygiene, 41, 619–626.

Example: In educational testing, a simple measure of the "easiness" of an item is the proportion of examinees that get the question correct (similarly, a measure of the "difficulty" of an item is the proportion of examinees that get the question incorrect). Suppose a test has two items, A and B, and let p_A and p_B denote the probability that a randomly selected examinee will get each item correct. The test was administered to n = 1000 examinees. The responses of the examinees to these two items are summarized in the table below.

Item B				
Item A correctincorrectTotal				
correct	400	125	525	
incorrect	175	300	475	
Total	575	425	1,000	

Now consider a test of the hypotheses $H_0: p_A - p_B = 0$ versus $H_a: p_A - p_B \neq 0$. However the two samples of responses — i.e., the samples of responses to item A and the sample of responses to item B — are not independent. How then do we test these hypotheses?