Monday, Sep 23

Estimation of a Ratio and Ratio Estimators

Confusingly, there are two somewhat related problems that we will consider involving "ratios" in estimation.

- 1. The estimation of a ratio of totals of two variables (an "estimator of a ratio").
- 2. Using an estimated ratio to estimate a population mean or total (a "ratio estimator").

Today's lecture considers only the first problem — i.e., estimating a ratio of totals.

Estimation of a Ratio (of Totals)

Sometimes we are interested in estimating the ratio of the totals of two variables, defined as

$$R = \frac{\tau_y}{\tau_x} = \frac{\sum_{i=1}^{N} y_i}{\sum_{i=1}^{N} x_i}.$$

Note that because $\mu_y = \frac{1}{N} \sum_{i=1}^N y_i$ and $\mu_x = \frac{1}{N} \sum_{i=1}^N x_i$, then we can also write $R = \mu_y/\mu_x$.

Examples include the estimation of a proportions, rates, and densities.

Example: Consider the problem of estimating the *proportion* of trees that are of a particular species in a region of land divided into N plots.

| Quantity | Description |
|---|-------------|
| $\begin{matrix} y_i \\ x_i \\ R \end{matrix}$ | |

Example: Consider the problem of estimating the birth rate in a population of N villages.

| Quantity | Description |
|---|--|
| $\begin{matrix} y_i \\ x_i \\ R \end{matrix}$ | number of births in the past year in village i number of people in village i birth rate (number of births per person per year) |

Example: Consider the problem of estimating the *density* of errors in a stream of data divided into N chunks of varying size.

| Quantity | escription | |
|----------|--|--|
| · | number of errors in the chunk i size of chunk i error density (errors per unit size) | |

We will also see that sometimes estimators of means will take the form of a ratio. Examples include the estimation of the means of *domains*, and the estimation of means when elements are *clustered*.

Estimator for a Ratio

For a simple random sampling design, an estimator of R is

$$r = \frac{\hat{\tau}_y}{\hat{\tau}_x} = \frac{N\bar{y}}{N\bar{x}} = \frac{\bar{y}}{\bar{x}}.$$

Example: Consider a survey using a simple random sampling design with n = 10 villages to estimate birth rate, where y_i is the number of births in the *i*-th sampled village, and x_i is the number of people in the *i*-th sampled village.

| village | y_i | x_i |
|-----------------|-------|-------|
| a | 6 | 58 |
| b | 7 | 158 |
| $^{\mathrm{c}}$ | 5 | 82 |
| d | 9 | 177 |
| e | 6 | 188 |
| f | 1 | 9 |
| g | 4 | 106 |
| h | 4 | 178 |
| i | 4 | 110 |
| j | 8 | 91 |

We have that $\bar{y}=5.4$ and $\bar{x}=115.7$, so the estimated birth rate is then $r=5.4/115.7\approx 0.05$ births per person.

The estimated variance of the estimator r under simple random sampling is

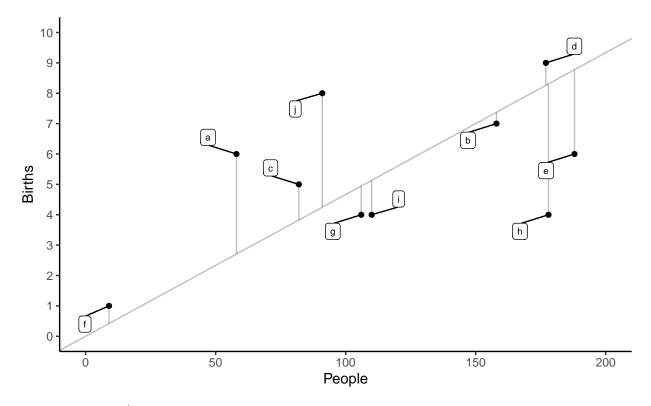
$$\hat{V}(r) = \left(\frac{1}{\mu_x^2}\right) \left(1 - \frac{n}{N}\right) \frac{s_r^2}{n} \quad \text{where} \quad s_r^2 = \frac{\sum_{i \in \mathcal{S}} (y_i - rx_i)^2}{n - 1}.$$

Note: If μ_x is unknown it can be replaced with \bar{x} .

Example: Returning to the birth rate problem, what is $\hat{V}(r)$ if N = 100? And what is the bound on the error of estimation?

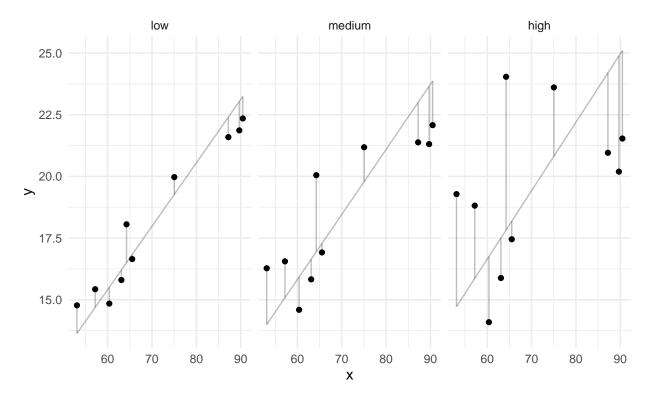
| village | y_i | x_i | $y_i - rx_i$ | $(y_i - rx_i)^2$ |
|--------------|-------|-------|--------------|------------------|
| a | 6 | 58 | 3.29 | 10.84 |
| b | 7 | 158 | -0.37 | 0.14 |
| \mathbf{c} | 5 | 82 | 1.17 | 1.38 |
| d | 9 | 177 | 0.74 | 0.55 |
| e | 6 | 188 | -2.77 | 7.7 |
| f | 1 | 9 | 0.58 | 0.34 |
| g | 4 | 106 | -0.95 | 0.9 |
| h | 4 | 178 | -4.31 | 18.56 |
| i | 4 | 110 | -1.13 | 1.29 |
| j | 8 | 91 | 3.75 | 14.08 |
| | | | | 55.76 |

Note that quantities shown in the table are rounded.



We can show that $\hat{V}(r) \approx 0.00004165567$, which gives a bound on the error of estimation of $B \approx 0.013$.

How does the relationship between y_i and x_i affect the variance of r?



Note that the lines in the figures below have intercepts of zero (i.e., they pass through the origin) and slopes of r.

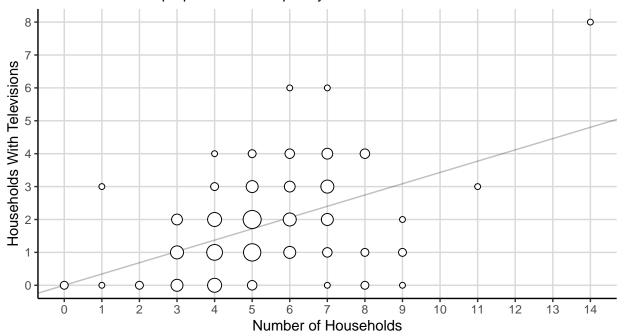
Example: Consider a survey to estimate the proportion of households with televisions in Des Moines, Iowa, in 1951. The sampling units are *blocks* of households.

| block | tv | households | |
|-------|---------------------|------------|--|
| 1 | 1 | 3 | |
| 2 | 0 | 8 | |
| 3 | 1 | 8 | |
| 4 | 0 | 7 | |
| 5 | 4 | 7 | |
| : | : | : | |
| 132 | 2 | 4 | |

We can compute $\bar{y} \approx 1.8$ and $\bar{x} \approx 5.23$. The estimate of the proportion of households with TVs is $r \approx 0.34$.

Households with Televisions by Number of Households

Note: Point area is proportional to frequency.



We have $\sum_{i\in\mathcal{S}}(y_i-rx_i)^2\approx 210.63$, and we know that N=9460 and that $\tau_x=56296$ households, so $\mu_x=56296/9460\approx 5.95$. We can show that $\hat{V}(r)\approx 0.0003$ and that the bound on the error of estimation is $B\approx 0.04$.

Estimation of a μ With Clusters of Elements

Example: Consider a sampling design where the elements are plot, y_i is tree biomass in the *i*-th plot, and x_i is the number of trees in the *i*-th plot. Here are the data for a simple random sampling design that selected n = 10 plots.

| Plot | Biomass | Trees |
|------|---------|-------|
| 1 | 125 | 4 |
| 2 | 185 | 7 |
| 3 | 82 | 4 |
| 4 | 152 | 8 |
| 5 | 263 | 9 |
| 6 | 52 | 2 |
| 7 | 131 | 5 |
| 8 | 201 | 8 |
| 9 | 114 | 5 |
| 10 | 148 | 5 |

Note that mean volume per tree (μ) is

$$\mu = \frac{\sum_{i=1}^{N} y_i}{\sum_{i=1}^{N} x_i} = \frac{\text{total biomass for all trees}}{\text{total number of trees}},$$

which can be estimated by the ratio $r = \bar{y}/\bar{x}$. Here we have that $\hat{\mu} = r = \bar{y}/\bar{x} \approx 25.5$ units of biomass per tree.

Note: We will see this estimator again when we discuss one-stage cluster sampling.

Estimation of Domain Means

Under simple random sampling, the estimator of μ_d can be written as

$$\bar{y}_d = \frac{\sum_{i \in \mathcal{S}} y_i'}{\sum_{i \in \mathcal{S}} x_i},$$

where

 $y_i' = \begin{cases} y_i, & \text{if the } i\text{-th element is in the domain,} \\ 0, & \text{if the } i\text{-th element is not in the domain,} \end{cases}$

and

$$x_i = \begin{cases} 1, & \text{if the i-th element is in the domain,} \\ 0, & \text{if the i-th element is not in the domain.} \end{cases}$$

Note that $\sum_{i \in S} x_i = n_d$ (i.e., the number of elements in the sample that are in the domain).

Example: Here is how y'_i and x_i would be defined for a sample of n = 10.

The variance of \bar{y}_d under simple random sampling tends to be a bit larger than if we had used the domains in a stratified random sampling design because n_d is random rather than fixed by design.

| Domain | y_i | y_i' | x_i |
|--------|-------|--------|-------|
| yes | 9 | 9 | 1 |
| yes | 5 | 5 | 1 |
| yes | 6 | 6 | 1 |
| no | 5 | 0 | 0 |
| yes | 2 | 2 | 1 |
| no | 8 | 0 | 0 |
| no | 3 | 0 | 0 |
| no | 2 | 0 | 0 |
| yes | 4 | 4 | 1 |
| yes | 9 | 9 | 1 |