

# Problem Set #4: Unsupervised Learning and Reinforcement Learning

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## 1. EM for Supervised Learning

(a) We have the log-likelihood of the parameters as following:

$$\begin{aligned}
 \ell(\phi, \theta_0, \theta_1) &= \log \prod_{i=1}^m p(y^{(i)} | x^{(i)}; \phi, \theta_0, \theta_1) \\
 &= \log \prod_{i=1}^m p(y^{(i)} | x^{(i)}, z^{(i)}; \theta_0, \theta_1) p(z^{(i)} | x^{(i)}; \phi) \\
 &= \sum_{i=1}^m \left( \log p(y^{(i)} | x^{(i)}, z^{(i)} = 0; \theta_0) + \log p(z^{(i)} = 0 | x^{(i)}; \phi) \right) + \\
 &\quad \sum_{i=1}^m \left( \log p(y^{(i)} | x^{(i)}, z^{(i)} = 1; \theta_1) + \log p(z^{(i)} = 1 | x^{(i)}; \phi) \right)
 \end{aligned}$$

Taking the gradient of the log-likelihood w.r.t  $\theta_0$  and setting it equal to zero, we have:

$$\begin{aligned}
 0 &= \nabla_{\theta_0} \ell(\phi, \theta_0, \theta_1) \\
 &= \nabla_{\theta_0} \sum_{i=1}^m \log p(y^{(i)} | x^{(i)}, z^{(i)} = 0; \theta_0) \\
 &= \nabla_{\theta_0} \sum_{i=1}^m \log \frac{1}{\sqrt{2\pi}\sigma} \exp \left( \frac{-(y^{(i)} - \theta_0^T x^{(i)})^2}{2\sigma^2} \right) \\
 &= \nabla_{\theta_0} \sum_{i=1, z^{(i)}=0}^m (y^{(i)} - \theta_0^T x^{(i)})^2
 \end{aligned}$$

We can easily see that this is just the least squares problem on a subset of the data. Then, we can obtain the maximum likelihood estimate of  $\theta_0$ :

$$\theta_0 = (X_{z=0}^T X_{z=0})^{-1} X_{z=0}^T \vec{y}_{z=0}$$

where  $X_{z=0}$  and  $\vec{y}_{z=0}$  are design matrix and vector created by examples where  $z^{(i)} = 0$ . Similarly, the maximum likelihood estimate of  $\theta_1$  is:

$$\theta_1 = (X_{z=1}^T X_{z=1})^{-1} X_{z=1}^T \vec{y}_{z=1}$$

And for the case of  $\phi$ , we have that:

$$\begin{aligned}
 \nabla_{\phi} \ell(\phi, \theta_0, \theta_1) &= \nabla_{\phi} \sum_{i=1}^m \log p(z^{(i)} = 0 | x^{(i)}; \phi) + \sum_{i=1}^m \log p(z^{(i)} = 1 | x^{(i)}; \phi) \\
 &= \nabla_{\phi} \sum_{i=1}^m (1 - z^{(i)}) (1 - g(\phi^T x^{(i)})) + \sum_{i=1}^m z^{(i)} g(\phi^T x^{(i)})
 \end{aligned}$$

which is the standard form of logistic regression problem. Thus, we can obtain the gradient and Hessian:

$$\begin{aligned}
 \nabla_{\phi} \ell(\phi, \theta_0, \theta_1) &= X^T (\vec{z} - \vec{h}), \text{ where } \vec{h}_i = h(x^{(i)}) = g(\phi^T x^{(i)}) \\
 H &= \sum_{i=1}^m -h(x^{(i)}) (1 - h(x^{(i)})) x^{(i)} x^{(i)T}
 \end{aligned}$$

(b) When  $z$  is a latent variable, the log-likelihood is:

$$\begin{aligned}
\ell(\phi, \theta_0, \theta_1) &= \log \prod_{i=1}^m p(y^{(i)} | x^{(i)}; \phi, \theta_0, \theta_1) \\
&= \log \prod_{i=1}^m \sum_{z^{(i)}} p(y^{(i)} | x^{(i)}, z^{(i)}; \theta_0, \theta_1) p(z^{(i)} | x^{(i)}; \phi) \\
&= \sum_{i=1}^m \log \sum_{z^{(i)}} p(y^{(i)} | x^{(i)}, z^{(i)}; \theta_0, \theta_1) p(z^{(i)} | x^{(i)}; \phi) \\
&= \sum_{i=1}^m \log \left( (1 - g(\phi^T x^{(i)})) \frac{1}{\sqrt{2\pi}\sigma} \exp\left(\frac{-(y^{(i)} - \theta_0^T x^{(i)})^2}{2\sigma^2}\right) + \right. \\
&\quad \left. g(\phi^T x^{(i)}) \frac{1}{\sqrt{2\pi}\sigma} \exp\left(\frac{-(y^{(i)} - \theta_1^T x^{(i)})^2}{2\sigma^2}\right) \right)
\end{aligned}$$

For each E-step, set:

$$\begin{aligned}
Q_i(z^{(i)}) &= p(z^{(i)} | x^{(i)}, y^{(i)}; \phi, \theta_0, \theta_1) \\
&= \frac{p(y^{(i)} | x^{(i)}, z^{(i)}; \theta_0, \theta_1) p(z^{(i)} | x^{(i)}; \phi)}{\sum_z p(y^{(i)} | x^{(i)}, z; \theta_0, \theta_1) p(z | x^{(i)}; \phi)}
\end{aligned}$$

In the M-step, we define:

$$w_j^{(i)} = p(z^{(i)} = j | x^{(i)}, y^{(i)}; \phi, \theta_j), \quad j = 0, 1$$

Taking the gradient of the lower bound  $\sum_i \sum_{z^{(i)}} Q_i(z^{(i)}) \log \frac{p(y^{(i)} | x^{(i)}, z^{(i)}; \phi, \theta_0, \theta_1) p(z^{(i)} | x^{(i)}; \phi)}{Q_i(z^{(i)})}$  of the log-likelihood w.r.t  $\theta_0$ , we have:

$$\begin{aligned}
&\nabla_{\theta_0} \sum_i \sum_{j=0}^1 w_j^{(i)} \log \frac{p(y^{(i)} | x^{(i)}, z^{(i)} = j; \phi, \theta_0, \theta_1) p(z^{(i)} = j | x^{(i)}; \phi)}{w_j^{(i)}} \\
&= \nabla_{\theta_0} \sum_i \left( w_0^{(i)} \log \frac{1}{w_0^{(i)}} (1 - g(\phi^T x^{(i)})) \frac{1}{\sqrt{2\pi}\sigma} \exp\left(\frac{-(y^{(i)} - \theta_0^T x^{(i)})^2}{2\sigma^2}\right) \right) \\
&= \frac{1}{2\sigma^2} \nabla_{\theta_0} \sum_i -w_0^{(i)} (y^{(i)} - \theta_0^T x^{(i)})^2
\end{aligned}$$

Setting this gradient equal to zero, we simply have:

$$\nabla_{\theta_0} \sum_i -w_0^{(i)} (y^{(i)} - \theta_0^T x^{(i)})^2 = 0$$

which is the weighted linear regression problem. And thus, it has the solution:

$$\theta_0 = (X_{z=0}^T W X_{z=0})^{-1} X_{z=0}^T W \vec{y}_{z=0},$$

where  $W_{ii} = w_0^{(i)}$ . Likewise, we have that:

$$\theta_1 = (X_{z=1}^T W X_{z=1})^{-1} X_{z=1}^T W \vec{y}_{z=1}$$

For the case of  $\phi$ , we have:

$$\begin{aligned}
&\nabla_{\phi} \sum_i \sum_{j=0}^1 w_j^{(i)} \log \frac{p(y^{(i)} | x^{(i)}, z^{(i)} = j; \phi, \theta_0, \theta_1) p(z^{(i)} = j | x^{(i)}; \phi)}{w_j^{(i)}} \\
&= \nabla_{\phi} \sum_i \left( w_0^{(i)} \log(1 - g(\phi^T x^{(i)})) + w_1^{(i)} \log g(\phi^T x^{(i)}) \right)
\end{aligned}$$

which is in the form of a logistic regression problem (since  $w_0^{(i)} = 1 - w_1^{(i)}$ ). Therefore, the gradient and Hessian are obtained by:

$$\nabla_{\phi} \sum_i \sum_{j=0}^1 w_j^{(i)} \log \frac{p(y^{(i)} | x^{(i)}, z^{(i)} = j; \phi, \theta_0, \theta_1) p(z^{(i)} = j | x^{(i)}; \phi)}{w_j^{(i)}} = X^T (\vec{w} - \vec{h})$$

$$H = \sum_{i=1}^m -h(x^{(i)}) \left(1 - h(x^{(i)})\right) x^{(i)} x^{(i)T},$$

where  $\vec{h}_i = h(x^{(i)}) = g(\phi^T x^{(i)})$

## 2. Factor Analysis and PCA

(a) We have that:

$$\begin{aligned} z &\sim \mathcal{N}(0, I) \\ x|z &\sim \mathcal{N}(Uz, \sigma^2 I) \end{aligned}$$

is equivalent to:

$$\begin{aligned} z &\sim \mathcal{N}(0, I) \\ \epsilon &\sim \mathcal{N}(0, \sigma^2 I) \\ x &= Uz + \epsilon \end{aligned}$$

where  $z, \epsilon$  are independent. And thus, we can easily obtain:

$$\mathbb{E}x = \mathbb{E}(Uz + \epsilon) = U\mathbb{E}z + \mathbb{E}\epsilon = 0$$

The joint distribution over  $(x, z)$  can be defined as:

$$\begin{aligned} \begin{bmatrix} z \\ x \end{bmatrix} &\sim \mathcal{N}(\mu, \Sigma), \text{ where:} \\ \mu &= \begin{bmatrix} \mu_z \\ \mu_x \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \\ \Sigma &= \begin{bmatrix} \Sigma_{zz} & \Sigma_{zx} \\ \Sigma_{xz} & \Sigma_{xx} \end{bmatrix} \end{aligned}$$

Using the fact about the independence between  $z, \epsilon$ , we have that:

$$\begin{aligned} \Sigma_{zz} &= \mathbb{E}(z - \mu_z)(z - \mu_z)^T \\ &= \mathbb{E}zz^T = \text{Cov}(z) = I \\ \Sigma_{zx} &= \mathbb{E}(z - \mu_z)(x - \mu_x)^T \\ &= \mathbb{E}(z(Uz + \epsilon)^T) \\ &= \mathbb{E}(zz^T U^T + z\epsilon^T) \\ &= \mathbb{E}(zz^T U^T) + \mathbb{E}z\mathbb{E}\epsilon^T \\ &= U^T \\ \Sigma_{xz} &= \Sigma_{zx}^T = U \\ \Sigma_{xx} &= \mathbb{E}(x - \mu_x)(x - \mu_x)^T \\ &= \mathbb{E}(Uz + \epsilon)(Uz + \epsilon)^T \\ &= \mathbb{E}(Uzz^T U^T + Uz\epsilon^T + \epsilon z^T U^T + \epsilon\epsilon^T) \\ &= U\mathbb{E}(zz^T)U^T + U\mathbb{E}z\mathbb{E}\epsilon^T + \mathbb{E}\epsilon\mathbb{E}z^T U^T + \mathbb{E}(\epsilon\epsilon^T) \\ &= UU^T + \sigma^2 I \end{aligned}$$

Putting everything together, we therefore have that:

$$\begin{bmatrix} z \\ x \end{bmatrix} \sim \mathcal{N}\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} I & U^T \\ U & UU^T + \sigma^2 I \end{bmatrix}\right)$$

And also,  $z|x \sim \mathcal{N}(\mu_{z|x}, \Sigma_{z|x})$ , where:

$$\begin{aligned}\mu_{z|x} &= \mu_z + \Sigma_{zx} \Sigma_{xx}^{-1} (x - \mu_x) \\ &= U^T (UU^T + \sigma^2 I)^{-1} x \\ &= \frac{U^T x}{U^T U + \sigma^2} \\ \Sigma_{z|x} &= \Sigma_{zz} - \Sigma_{zx} \Sigma_{xx}^{-1} \Sigma_{xz} \\ &= I - U^T (UU^T + \sigma^2 I)^{-1} U \\ &= 1 - \frac{U^T U}{U^T U + \sigma^2}\end{aligned}$$

(b) For the E-step, we have that:

$$\begin{aligned}Q_i(z^{(i)}) &= p(z^{(i)}|x^{(i)}; U) \\ &= \frac{1}{(2\pi)^{k/2} |\Sigma_{z^{(i)}|x^{(i)}}|^{1/2}} \exp \left( -\frac{1}{2} (z^{(i)} - \mu_{z^{(i)}|x^{(i)}})^T \Sigma_{z^{(i)}|x^{(i)}}^{-1} (z^{(i)} - \mu_{z^{(i)}|x^{(i)}}) \right)\end{aligned}$$

where  $\mu_{z^{(i)}|x^{(i)}}$  and  $\Sigma_{z^{(i)}|x^{(i)}}$  can be computed as above. For the M-step, we need to maximize:

$$\begin{aligned}& \sum_{i=1}^m \int_{z^{(i)}} Q_i(z^{(i)}) \log \frac{p(x^{(i)}, z^{(i)}; U)}{Q_i(z^{(i)})} dz^{(i)} \\ &= \sum_{i=1}^m \mathbb{E}_{z^{(i)} \sim Q_i} \left[ \log p(x^{(i)}|z^{(i)}; U) + \log p(z^{(i)}) - \log Q_i(z^{(i)}) \right]\end{aligned}$$

Dropping terms that do not depend on the parameter  $U$  and taking the gradient w.r.t  $U$ , we have that:

$$\begin{aligned}& \nabla_U \sum_{i=1}^m \mathbb{E}_{z^{(i)} \sim Q_i} \left[ \log p(x^{(i)}|z^{(i)}; U) \right] \\ &= \nabla_U \sum_{i=1}^m \mathbb{E} \left[ \log \frac{1}{(2\pi)^{n/2} \sigma} - \frac{1}{2\sigma^2} (x^{(i)} - Uz^{(i)})^T (x^{(i)} - Uz^{(i)}) \right] \\ &= \frac{1}{2\sigma^2} \nabla_U \sum_{i=1}^m \mathbb{E} \left( x^{(i)T} Uz^{(i)} + z^{(i)T} U^T x^{(i)} - z^{(i)T} U^T U z^{(i)} \right) \\ &= \frac{1}{\sigma^2} \sum_{i=1}^m \mathbb{E} \left( x^{(i)} z^{(i)T} - Uz^{(i)} z^{(i)T} \right) \\ &= \frac{1}{\sigma^2} \left[ \sum_{i=1}^m x^{(i)} \mathbb{E} [z^{(i)T}] - U \sum_{i=1}^m \mathbb{E} [z^{(i)} z^{(i)T}] \right]\end{aligned}$$

Setting this equal to zero and solving for  $U$ , we have that:

$$U = \left( \sum_{i=1}^m x^{(i)} \mathbb{E}_{z^{(i)} \sim Q_i} [z^{(i)T}] \right) \left( \sum_{i=1}^m \mathbb{E}_{z^{(i)} \sim Q_i} [z^{(i)} z^{(i)T}] \right)^{-1} \quad (1)$$

From our definition of  $Q_i$  being Gaussian with mean  $\mu_{z^{(i)}|x^{(i)}}$  and covariance  $\Sigma_{z^{(i)}|x^{(i)}}$ , we easily find:

$$\begin{aligned}\mathbb{E}_{z^{(i)} \sim Q_i} [z^{(i)T}] &= \mu_{z^{(i)}|x^{(i)}} \\ \mathbb{E}_{z^{(i)} \sim Q_i} [z^{(i)} z^{(i)T}] &= \mu_{z^{(i)}|x^{(i)}}^T \mu_{z^{(i)}|x^{(i)}} + \Sigma_{z^{(i)}|x^{(i)}}\end{aligned}$$

Substituting this back into equation (1), we get the M-step update for  $U$ :

$$U = \left( \sum_{i=1}^m x^{(i)} \mu_{z^{(i)}|x^{(i)}} \right) \left( \sum_{i=1}^m \mu_{z^{(i)}|x^{(i)}}^T \mu_{z^{(i)}|x^{(i)}} + \Sigma_{z^{(i)}|x^{(i)}} \right)^{-1}$$

(c) When  $\sigma^2 \rightarrow 0$ , we have that:

$$\begin{aligned}\mu_{z|x} &\rightarrow \frac{U^T x}{U^T U} \\ \Sigma_{z|x} &\rightarrow 0\end{aligned}$$

From the definition of  $w$  that  $w_i = \mu_{z^{(i)}|x^{(i)}}$  and results obtained from part (a), we have:

$$w = \frac{XU}{U^T U}$$

From what we have got from part (b), we have that:

$$\begin{aligned}U &= \left( \sum_{i=1}^m x^{(i)} \mu_{z^{(i)}|x^{(i)}} \right) \left( \sum_{i=1}^m \mu_{z^{(i)}|x^{(i)}} \mu_{z^{(i)}|x^{(i)}}^T + \Sigma_{z^{(i)}|x^{(i)}} \right)^{-1} \\ &= \left( \sum_{i=1}^m x^{(i)} w_i \right) \left( \sum_{i=1}^m w_i w_i^T \right)^{-1} \\ &= \frac{X^T w}{w^T w}\end{aligned}$$

If  $U$  remains unchanged after the update, it must satisfy:

$$\begin{aligned}U &= \frac{X^T \frac{XU}{U^T U}}{\frac{U^T X^T XU}{U^T U}} \\ &= X^T XU \frac{U^T U}{U^T X^T XU} \\ &= \Sigma U \frac{m^2 U^T U}{U^T X^T XU} \\ &= \Sigma U \frac{1}{\lambda}\end{aligned}$$

### 3. PCA and ICA for Natural Images

#### 4. Convergence of Policy Iteration

$$\begin{aligned}V'(s) &= B(V)(s) \\ &= R(s) + \gamma \sum_{s' \in \mathcal{S}} P_{s\pi(s)}(s') V(s')\end{aligned}$$

(a) Since  $V_1(s) < V_2(s), \forall s \in \mathcal{S}$ , we have:

$$\begin{aligned}B(V_1)(s) &= R(s) + \gamma \sum_{s' \in \mathcal{S}} P_{s\pi(s)}(s') V_1(s') \\ &\leq R(s) + \gamma \sum_{s' \in \mathcal{S}} P_{s\pi(s)}(s') V_2(s') \\ &= B(V_2)(s)\end{aligned}$$

(b) For any  $V$ , we have:

$$\begin{aligned}\|B^\pi(V) - V^\pi\|_\infty &= \max_{s \in \mathcal{S}} \left| R(s) + \gamma \sum_{s' \in \mathcal{S}} P_{s\pi(s)}(s') V(s') - R(s) - \gamma \sum_{s' \in \mathcal{S}} P_{s\pi(s)}(s') V^\pi(s') \right| \\ &= \max_{s' \in \mathcal{S}} \gamma \left| \sum_{s' \in \mathcal{S}} P_{s\pi(s)}(s') (V(s') - V^\pi(s')) \right| \\ &\leq \max_{s' \in \mathcal{S}} \gamma |V(s') - V^\pi(s')| \\ &= \|V - V^\pi\|_\infty\end{aligned}$$

In the third step, we use the fact that  $\sum_{s' \in \mathcal{S}} P_{s\pi(s)} = 1$  and  $\sum_i \alpha_i x_i \leq \max_i x_i$  if  $\sum_i \alpha_i = 1, \alpha_i \geq 0$

(c) Since  $\pi'(s) = \arg \max_{a \in \mathcal{A}} \sum_{s' \in \mathcal{S}} P_{sa}(s') V^\pi(s')$ , we have:

$$\begin{aligned} V^\pi(s) &= R(s) + \gamma \sum_{s' \in \mathcal{S}} P_{s\pi(s)}(s') V^\pi(s') \\ &\leq R(s) + \gamma \sum_{s' \in \mathcal{S}} P_{s\pi'(s)}(s') V^\pi(s') \\ &= B^{\pi'}(V^\pi)(s) \end{aligned}$$

By part (a), applying  $B^{\pi'}$  on both sides of the inequality, we have:

$$V^\pi(s) \leq B^{\pi'}(V^\pi)(s) \Rightarrow B^{\pi'}(V^\pi)(s) \leq B^{\pi'}(B^{\pi'}(V^\pi))(s)$$

Using the result obtained from part (b), we have:

$$\begin{aligned} V^\pi(s) &\leq B^{\pi'}(V^\pi)(s) \\ &\leq B^{\pi'}(B^{\pi'}(V^\pi))(s) \\ &\leq B^{\pi'}(B^{\pi'}(\dots B^{\pi'}(V^\pi) \dots))(s) \\ &= V^{\pi'}(s) \end{aligned}$$

(d) As result obtained from part (c), the policies are monotonically improving. On the other hand, there are a finite number of possible policies. Hence, at some point, the policy's gonna stop updating, or in other words, we will produce a  $\pi' = \pi$ . If  $\pi' = \pi$ , we have:

$$V^\pi = V^{\pi'} = R(s) + \gamma \max_{a \in \mathcal{A}} \sum_{s' \in \mathcal{S}} P_{sa}(s') V^\pi(s')$$

or  $V^\pi = V^*$ , and thus,  $\pi = \pi^*$

## 5. Reinforcement Learning: The Mountain Car