

5) a) $\min_{g(x)} E((Y - g(x))^2)$

$$\frac{\partial E}{\partial g(x)} = E(-2(Y - g(x))) = 0$$

$$\Rightarrow E(Y) = E(g(x))$$

$$\Rightarrow \boxed{E(Y|x) = g(x)}$$

c) $y = x^2 + z$ $x \sim N(0, 1)$
 $\hat{y} = x^2$ (wrt x) $z \sim N(0, 1)$

$$MSE = E((y - \hat{y})^2)$$

$$= \boxed{E(z^2) = 1}$$

$$MSE = 1$$

d) $\hat{y} = \alpha x + c$

$$E((y - \hat{y})^2) = E((x^2 + z - \alpha x - c)^2)$$

$$= E((x^2 + z - \alpha x - c)(x^2 + z - \alpha x - c)) = E \left[\begin{aligned} &x^4 + x^2 z - \alpha x^3 - c x^2 + z x^2 + z^2 - \alpha z x \\ &- c z - \alpha x^3 - \alpha z x + \alpha^2 x^2 + \alpha c x \\ &- c x^2 - c z + c \alpha x + c^2 \end{aligned} \right]$$

$$= E \left[x^4 + x^3(-2\alpha) + x^2[2z - 2c + \alpha^2] + x[-2\alpha z - 2\alpha c] + [z^2 - 2cz + c^2] \right]$$

$$\begin{aligned}
&= E(x^2) - 2\alpha E(xz) + E(x^2(2z - 2c + \alpha^2)) - 2\alpha E(x(c+z)) \\
&\quad + E(2z - 2cz + c^2) \\
&= 2E(x^2) - 2cE(x^2) + \alpha^2 E(x^2) - 2\alpha E(x) - 2\alpha E(xz) \\
&\quad + E(2z) - 2cE(z) + E(c^2) \\
&= -2c + \alpha^2 + 1 + c^2 = \boxed{1 + \alpha^2 + c^2 - 2c}
\end{aligned}$$

1) a) $y[k] = x^T[k] \theta + e[k]$

$$\hat{\sigma}_e^2 = \frac{\sum_{k=0}^{N-1} (e[k])^2}{N-p}$$

$$e = y - \hat{y} = y - P y = P^\perp y$$

$$\begin{aligned}
\hat{\sigma}_e^2 &= \frac{e e^T}{N-p} = \frac{(P^\perp y) (P^\perp y)^T}{N-p} = \frac{(P^\perp)^2 y^T y}{N-p} \quad P^\perp y = P^\perp e \\
&= \frac{(P^\perp)^2 e^2}{N-p} = \frac{\sum_{k=0}^{N-1} (P^\perp e[k])^2}{N-p} = \frac{\sum_{k=0}^{N-1} e^2[k]}{N-p}
\end{aligned}$$

$$\begin{aligned}
E(\hat{\sigma}_e^2) &= E\left(\frac{\sum_{k=0}^{N-1} e^2[k]}{N-p}\right) = \frac{\sum_{k=0}^{N-1} E(e^2[k])}{N-p} \\
&= \frac{N-p}{N-p} \sigma_e^2 = \boxed{\sigma_e^2}
\end{aligned}$$

$$a) \quad y[k] = -d_1 y[k-1] - \dots - d_p y[k-p] + e[k]$$

L-S estimator:-

$$\hat{\theta} = (\sum y y)^{-1} \sigma_e^2$$

Y-W estimator:-

$$\sigma_{yy}[l] = -d_1 \sigma_{yy}[l-1] - d_2 \sigma_{yy}[l-2] - \dots - d_p \sigma_{yy}[l-p] + \sigma_{ee}[l]$$

$$\Rightarrow \sigma_{yy}[l] + d_1 \sigma_{yy}[l-1] + \dots + d_p \sigma_{yy}[l-p] = \sigma_{ee}[l]$$

$$y[k] - \bar{y} = -d_1 (y[k-1] - \bar{y}) - \dots - d_p (y[k-p] - \bar{y}) + e[k]$$

$$\Rightarrow \sigma_{yy}[l] - \bar{y} \mu = -d_1 [\sigma_{yy}[l-1] - \bar{y} \mu] - \dots - d_p [\sigma_{yy}[l-p] - \bar{y} \mu] + \sigma_{ee}[l]$$

$$\Rightarrow (\sigma_{yy}[l] - \bar{y} \mu) + d_1 (\sigma_{yy}[l-1] - \bar{y} \mu) + \dots + d_p (\sigma_{yy}[l-p] - \bar{y} \mu) = \sigma_{ee}[l]$$

$$\Rightarrow \begin{bmatrix} \sigma_{yy}[0] - \bar{y} \mu & \sigma_{yy}[1] - \bar{y} \mu & \dots & \sigma_{yy}[p] - \bar{y} \mu \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{yy}[p] - \bar{y} \mu & \dots & \dots & \sigma_{yy}[0] - \bar{y} \mu \end{bmatrix} \begin{bmatrix} 1 \\ d_1 \\ \vdots \\ d_p \end{bmatrix} = \begin{bmatrix} \sigma_{ee}[0] \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

$$\Rightarrow \left(\sum y y - \bar{y} \bar{y} I_{\substack{p+1 \\ (p+1) \times (p+1)}} \right) \hat{\theta} = \sigma_e^2$$

$$\hat{\theta} = \left(\sum y y - \mu \bar{y} \frac{1}{(p+1) \times (p+1)} \right)^{-1} \sigma_e^2 \quad (\text{Y-W estimator})$$

$$\text{as } N \rightarrow \infty, \bar{y} = 0$$

$$\Rightarrow \hat{\theta} = \left(\sum y y \right)^{-1} \sigma_e^2 \quad (\text{Y-W estimator})$$

$$\text{as } N \rightarrow \infty \quad \boxed{\hat{\theta}_{Y-W} = \hat{\theta}_{L-S}}$$

$$y[k] = -d_1 y[k-1] + e[k]$$

$$E[y[k]] = E[-d_1 y[k-1]] + E[e[k]]$$

$$\Rightarrow (1+d_1)\mu = 0 \Rightarrow \boxed{\mu = 0}$$

$$y[k] y[k-1] = -d_1 y[k-1] y[k-1] + e[k] y[k-1]$$

$$\Rightarrow \cancel{\sigma_{yy}}[e] = -d_1 \cancel{\sigma_{yy}}[e-1] + \sigma_{ey}[e]$$

$$y[k] y[k] = -d_1 y[k-1] [-d_1 y[k-1] + e[k]] + e[k] y[k]$$

$$\Rightarrow \sigma_{yy}[0] = d_1^2 \sigma_{yy}[0] + \sigma_e^2$$

$$\Rightarrow \boxed{\sigma_y^2 = \frac{\sigma_e^2}{1-d_1^2}}$$

$$\hat{y}[k] = -d_1 \hat{y}[k-1]$$

$$e[k] = y[k] - \hat{y}[k] = y[k] + d_1 y[k-1]$$

$$p(y, \theta) = \frac{1}{\sqrt{2\pi\sigma_y^2}} \exp\left(\frac{-1}{2\sigma_y^2} y^2[0]\right) * \frac{1}{\sqrt{2\pi\sigma_x^2}} \exp\left(\frac{-1}{2\sigma_x^2} [y[1] + d_1 y[0]]^2\right)$$

$$= \frac{\sqrt{1-d_1^2}}{2\pi\sigma_x^2} \exp\left[\frac{-1}{2\sigma_x^2} \left[(1-d_1^2) y^2[0] + (y[1] + d_1 y[0])^2\right]\right]$$

$$= \frac{\sqrt{1-d_1^2}}{2\pi\sigma_x^2} \exp\left[\frac{-1}{2\sigma_x^2} \left[y^2[0] + y^2[1] + 2y[1]y[0]d_1\right]\right]$$

$$= \frac{\sqrt{1-d_1^2}}{2\pi\sigma_x^2} \exp\left[\frac{-1}{2\sigma_x^2} \left[x^2[1] + x^2[2] + 2d_1 x[1]x[2]\right]\right]$$

$$L(y, \theta) = \frac{1}{2} \ln(1-d_1^2) - \ln(2\pi\sigma_x^2) - \frac{1}{2\sigma_x^2} \left[x^2[1] + x^2[2] + 2d_1 x[1]x[2]\right]$$

$$\frac{\partial L}{\partial \sigma_x^2} = \frac{-1}{\sigma_x^2} + \frac{d_1 x[1]x[2]}{(\sigma_x^2)^2} = 0$$

$$\Rightarrow \frac{1}{\sigma_x^2} = d_1 x[1]x[2]$$

$$\frac{\partial L}{\partial \sigma_x^2} = \frac{-1}{\sigma_x^2} + \frac{1}{2\sigma_x^4} \left[x^2[1] + x^2[2] + 2d_1 x[1]x[2]\right] = 0$$

$$\Rightarrow \sigma_x^4 = \frac{1}{2} \left[x^2[1] + x^2[2] + 2d_1 x[1]x[2]\right]$$

$$\frac{\partial L}{\partial d_1} = \frac{1(-2d_1)}{2(1-d_1^2)} - \frac{1}{2\sigma_e^2} [2x[1]x[2]] = 0$$

$$\Rightarrow \frac{d_1}{1-d_1^2} + \frac{x[1]x[2]}{\sigma_e^2} = 0$$

$$\Rightarrow (1-d_1^2)x[1]x[2] = -d_1\sigma_e^2$$

$$\Rightarrow x[1]x[2]d_1^2 - d_1\sigma_e^2 - x[1]x[2] = 0$$

$$\Rightarrow d_1^2 - \frac{d_1\sigma_e^2}{x[1]x[2]} - 1 = 0$$

$$d_1 = \frac{\sigma_e^2}{x[1]x[2]} \pm \sqrt{\left(\frac{\sigma_e^2}{x[1]x[2]}\right)^2 + 4}$$

$$= \frac{\sigma_e^2}{2x[1]x[2]} \pm \sqrt{\left(\frac{\sigma_e^2}{2x[1]x[2]}\right)^2 + 1}$$

$$u) \quad \hat{\theta} = (\phi^T \phi)^{-1} \phi^T y$$

$$= (\Sigma_{yy})^{-1} \sigma_e^2$$

$$\sigma_e^2 = \sigma_y^2 [1-d_1^2]$$

$$\Sigma_{yy} = \begin{bmatrix} x^2[0] & x[0]x[1] \\ x[0]x[1] & x^2[1] \end{bmatrix}$$

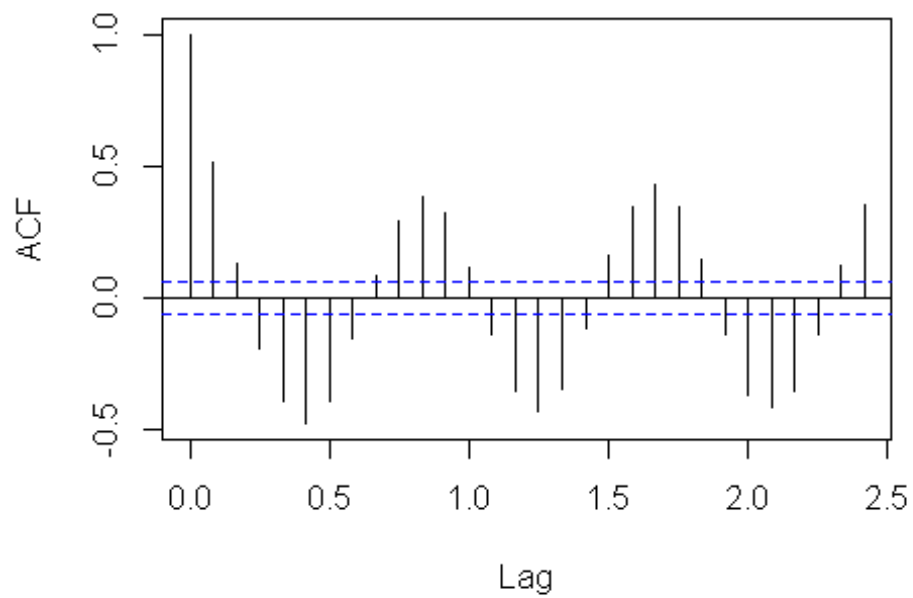
4b.R

Vishal Subbiah

Sun Nov 09 15:07:30 2014

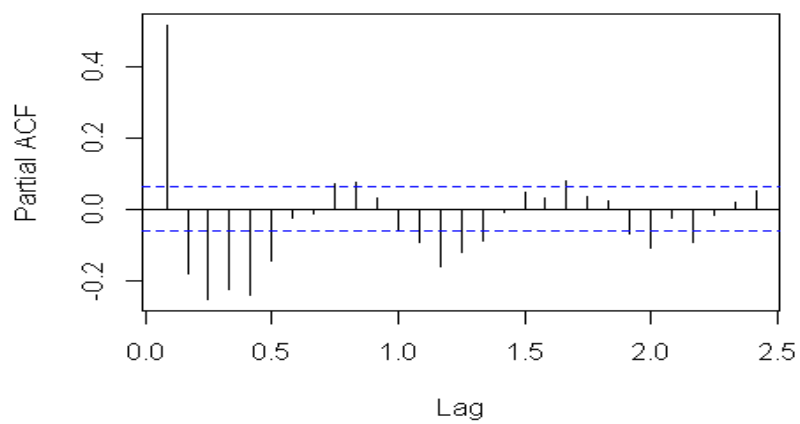
```
load('sarima_data.Rdata')  
#plot(yk)  
#acf(yk,xLab="Lag",yLab='ACF')  
#pacf(yk)  
vk=diff(yk)  
acf(vk)
```

Series vk



```
pacf(vk)
```

Series vk

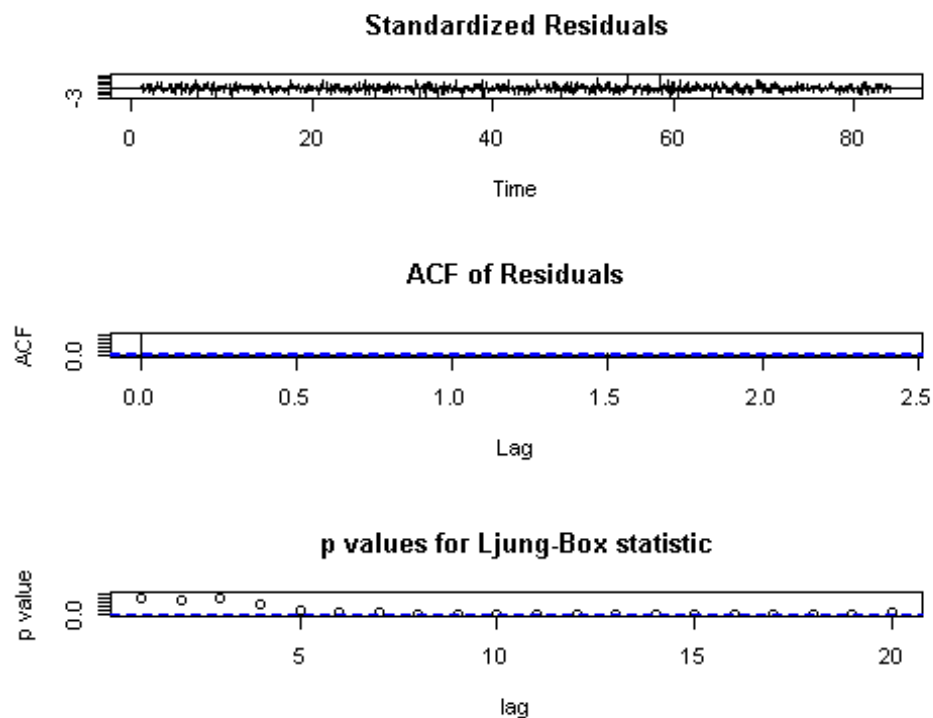


```

mod_sarima =
arima(vk,order=c(3,0,5),seasonal=list(order=c(0,0,1),period=12),include.me
an=F)
print(mod_sarima)

##
## Call:
## arima(x = vk, order = c(3, 0, 5), seasonal = list(order = c(0, 0, 1),
period = 12),
##      include.mean = F)
##
## Coefficients:
##
## Warning in sqrt(diag(x$var.coef)): NaNs produced
##
##      ar1      ar2      ar3      ma1      ma2      ma3      ma4      ma5
##      1.7802 -1.2625  0.1622 -1.4609  0.693  0.1477  0.0978 -0.0958
## s.e.  0.0265  0.0429  0.0265      NaN      NaN  0.0096  0.0059  0.0012
##      sma1
##      -0.0432
## s.e.  0.0341
##
## sigma^2 estimated as 1.073:  log likelihood = -1456.68,  aic = 2933.37
tsdiag(mod_sarima,gof.lag=20)

```



1c.R

```
vk1=arima.sim(n=1000,model=list(ar=c(1.1,-0.28)))
ols_1=ar.ols(vk1)
yw_1=ar.yw(vk1)
vk2=arima.sim(n=10000,model=list(ar=c(1.1,-0.28)))
ols_2=ar.ols(vk2)
yw_2=ar.yw(vk2)
```

3.R

```
data(cmort)
vko=cmort[3:508]
vk1=cmort[2:507]
vk2=cmort[1:506]
mod_ar2_lr=lm(vko ~ I(vk1)+ I(vk2))
vk_yw=ar.yw(cmort)
vk_ols=ar.ols(cmort)
```

4a.R

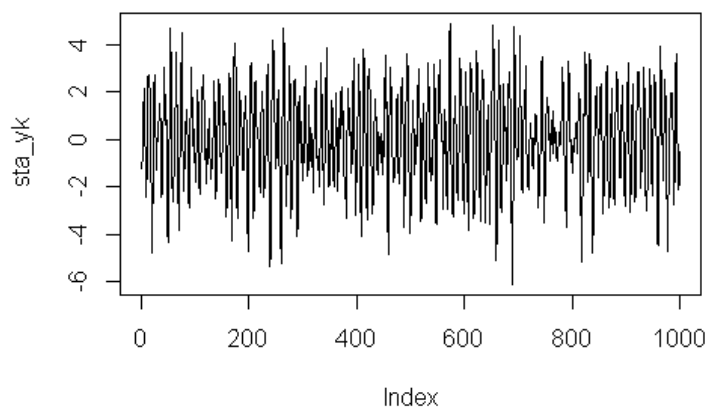
```
load('sarima_data.Rdata')
yk_1=stl(yk,s.window="periodic")
#plot(yk,type='l')
sea_yk=yk_1$time.series[1:1000,1]# seasonal
```

```
tre_yk=yk_1$time.series[1:1000,2]# trend  
sta_yk=yk_1$time.series[1:1000,3]# stationary
```

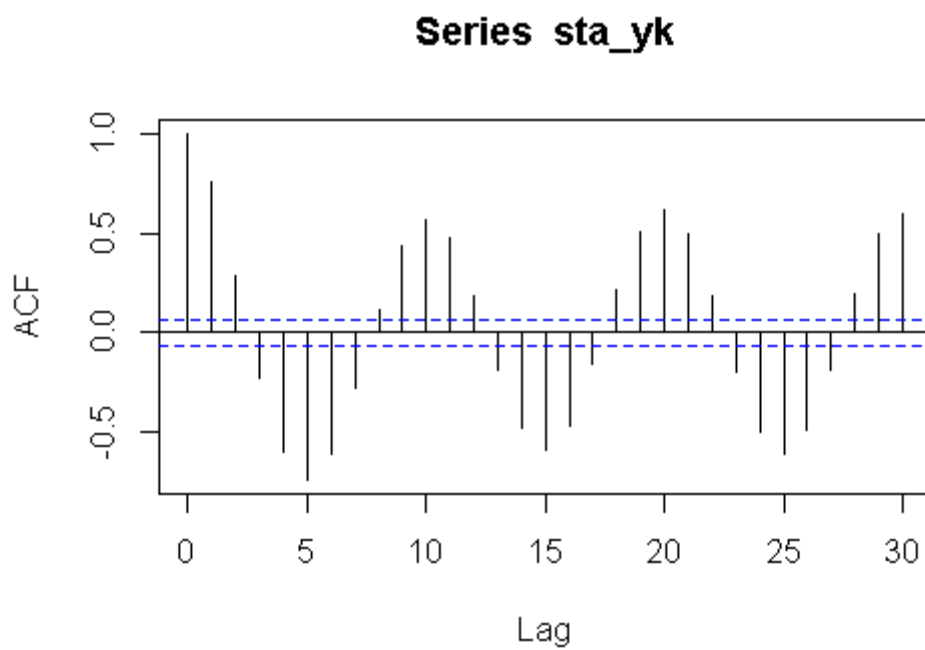
```
#plot(sea_yk,type='l')
```

```
#plot(tre_yk,type='l')
```

```
plot(sta_yk,type='l')
```

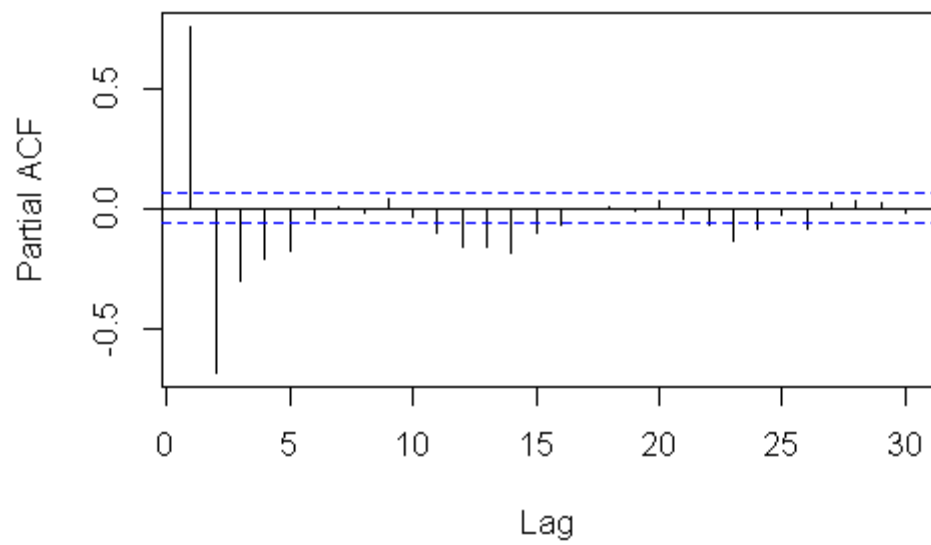


```
acf(sta_yk)
```



```
pacf(sta_yk)
```

Series sta_yk



```
vk=arma(sta_yk,order=c(2,5))
```

```
acf(vk$residuals[10:1000], type="covariance") # residuals is white
```

Series vk\$residuals[10:1000]

