# Sta 531 HW5 (The one due right after spring break)

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#### (1) Some weird Bernoulli stuff

Let  $Z_i$  be iid  $N(\alpha, \sigma^2)$  and  $X_1, \ldots, X_n$  be independent Bernoulli random variables given by:

$$X_i = \begin{cases} 0 & Z_i \le u \\ 1 & Z_i > u \end{cases}$$

#### (a) Likelihood

$$L(p) = \prod_{i=1}^{n} p^{x_i} (1-p)^{1-x_i}$$
$$= p^{\sum x_i} (1-p)^{\sum (1-x_i)}$$
$$= p^{\sum x_i} (1-p)^{n-\sum x_i}$$

And, as for p,

$$p = Pr(X_i = 1) = Pr(Z_i > u) = \int_u^\infty N(x; \alpha, \sigma^2) = \Phi\left(\frac{\alpha - u}{\sigma}\right),$$

where  $\Phi$  denotes the cdf of the normal distribution.

#### (b) Complete-Data Log Likelihood

The log-likelihood of the  $z_i$ 's is the standard normal log-likelihood.

$$l^{C}(\alpha, \sigma^{2}) = -\frac{n}{2}\log(2\pi\sigma^{2}) - \frac{1}{2\sigma^{2}}\sum(Z_{i} - \alpha)^{2}$$

We then take the expectation conditional on the  $x_i$ 's and use the book's notation to denote this as Q.

$$\begin{split} Q(\alpha, \sigma^2) &= E\left[ -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum (Z_i - \alpha)^2 \mid x \right] \\ &= -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum E\left[ (Z_i - \alpha)^2 \mid x_i \right] \\ &= -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum \left[ E(Z_i^2 \mid x_i) - 2\alpha E(Z_i \mid x_i) + \alpha^2 \right] \end{split}$$

#### (c) The M-step

We want

$$\arg \max_{\alpha} Q = \arg \max_{\alpha} \sum_{\alpha} \left[ 2\alpha E(Z_i \mid x_i, \alpha_{(j)}^2, \sigma_{(j)}^2) - E(Z_i^2 \mid x_i, \alpha_{(j)}^2, \sigma_{(j)}^2) - \alpha^2 \right]$$

$$= \arg \max_{\alpha} \sum_{\alpha} \left[ 2\alpha E(Z_i \mid x_i, \alpha_{(j)}^2, \sigma_{(j)}^2) \right] - n\alpha^2$$

$$= \frac{1}{n} \sum_{\alpha} E(Z_i \mid x_i, \alpha_{(j)}^2, \sigma_{(j)}^2)$$

$$= \hat{\alpha}_{(j+1)}$$

where we obtain the last equality by setting the derivative with respect to  $\alpha$  equal to 0. We'll do the same for  $\sigma^2$ .

$$0 = -\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum (v_i - 2\alpha t_i + \alpha^2)$$
$$\sigma^2 = \frac{1}{n} \left[ \sum v_i - 2\alpha \sum t_i + n\alpha^2 \right]$$
$$= \frac{1}{n} \left[ \sum v_i - (\frac{1}{n} \sum t_i)^2 \right] \quad \text{(from above)}$$

- (2) Why we gotta classify the animals? It's 2018, let them be who they want.
- (a) Likelihood

$$L(Y_1 \mid \theta) = \frac{197!}{125! \ 18! \ 20! \ 34!} \left(\frac{2+\theta}{4}\right)^{125} \left(\frac{1-\theta}{4}\right)^{18} \left(\frac{1-\theta}{4}\right)^{20} \left(\frac{\theta}{4}\right)^{34}$$
$$L(Y_2 \mid \theta) = \frac{20!}{14! \ 0! \ 1! \ 5!} \left(\frac{2+\theta}{4}\right)^{14} \left(\frac{1-\theta}{4}\right) \left(\frac{\theta}{4}\right)^{5}$$

#### (b) Newton-Raphson

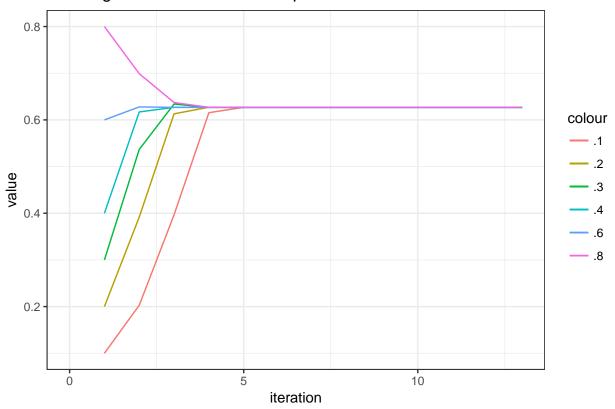
Note that when we take the log-likelihood, we will ignore the constant at the front and instead maximize:

$$l(Y_1 \mid \theta) = 125 \log \left(\frac{2+\theta}{4}\right) + 18 \log \left(\frac{1-\theta}{4}\right) + 20 \log \left(\frac{1-\theta}{4}\right) + 34 \log \left(\frac{\theta}{4}\right)$$

```
likeB = function(theta){
    125*log((2+theta)/4) + 18*log((1-theta)/4) + 20*log((1-theta)/4) + 34*log(theta/4)
}
dlikeB = function(theta){
    125/4 * 4/(2+theta) - 18/4 * 4/(1-theta) - 20/4 * 4/(1-theta) + 34/4 * 4/theta
}
d2likeB = function(theta){
    -125/(2+theta)^2 - 18/(1-theta)^2 - 20/(1-theta)^2 - 34/(theta)^2
}
resB = NULL
for (theta_0 in c(.1,.2,.3,.4,.6,.8)){
    x = c(theta_0,rep(NA, 100))
    for (j in 1:100){
        x[j+1] = x[j] - dlikeB(x[j])/d2likeB(x[j])
    }
    resB = cbind(resB, x)
}
```

```
library(dplyr)
resB = data.frame(iteration = 1:101, resB)
colnames(resB) = c("iteration", "x.1", "x.2", "x.3", "x.4", "x.6", "x.8")
library(ggplot2)
ggplot(data = resB, aes(x = iteration)) +
  geom_line(aes(y = x.1, color = ".1")) +
  geom_line(aes(y = x.2, color = ".2")) +
  geom_line(aes(y = x.3, color = ".3")) +
  geom_line(aes(y = x.4, color = ".4")) +
  geom_line(aes(y = x.6, color = ".6")) +
  geom_line(aes(y = x.8, color = ".8")) +
  theme_bw() +
  xlim(0,13) +
  ggtitle("Convergence of the Newton-Rapshon") +
  ylab("value")
```

## Convergence of the Newton-Rapshon

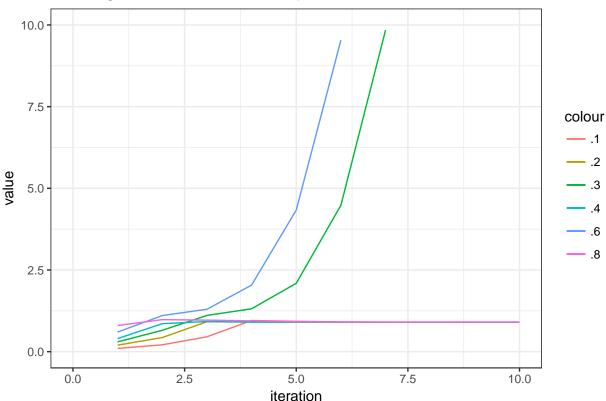


#### (c) Newton-Raphson again

```
likeC = function(theta){
   14*log((2+theta)/4) + 1*log((1-theta)/4) + 5*log(theta/4)
}
dlikeC = function(theta){
   14/4 * 4/(2+theta) - 1/4 * 4/(1-theta) + 5/4 * 4/theta
}
d2likeC = function(theta){
   -14/(2+theta)^2 - 1/(1-theta)^2 - 5/(theta)^2
```

```
}
resC = NULL
for (theta_0 in c(.1,.2,.3,.4,.6,.8)){
  x = c(theta_0, rep(NA, 100))
  for (j in 1:100){
    x[j+1] = x[j] - dlikeC(x[j])/d2likeC(x[j])
  resC = cbind(resC, x)
}
library(dplyr)
resC = data.frame(iteration = 1:101, resC)
colnames(resC) = c("iteration", "x.1", "x.2", "x.3", "x.4", "x.6", "x.8")
library(ggplot2)
ggplot(data = resC, aes(x = iteration)) +
  geom_line(aes(y = x.1, color = ".1")) +
  geom_line(aes(y = x.2, color = ".2")) +
  geom\_line(aes(y = x.3, color = ".3")) +
  geom\_line(aes(y = x.4, color = ".4")) +
  geom_line(aes(y = x.6, color = ".6")) +
  geom_line(aes(y = x.8, color = ".8")) +
  theme_bw() +
  xlim(0,10) +
  ylim(0,10) +
  ggtitle("Convergence of the Newton-Rapshon") +
  ylab("value")
```

## Convergence of the Newton-Rapshon

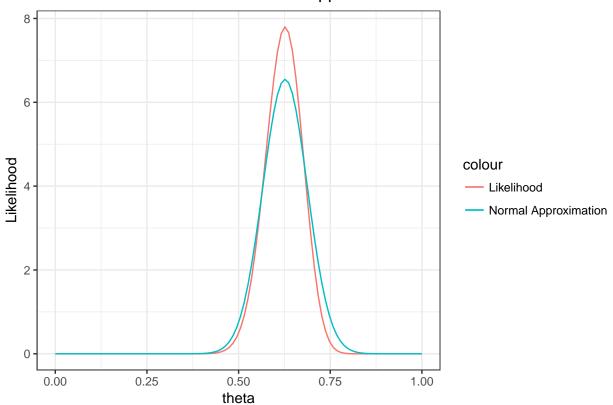


We see that 0.6 and 0.3 starting points do not lead to convergence of the Newton-Raphson algorithm in this

case.

### (d) Assymptotic Normality

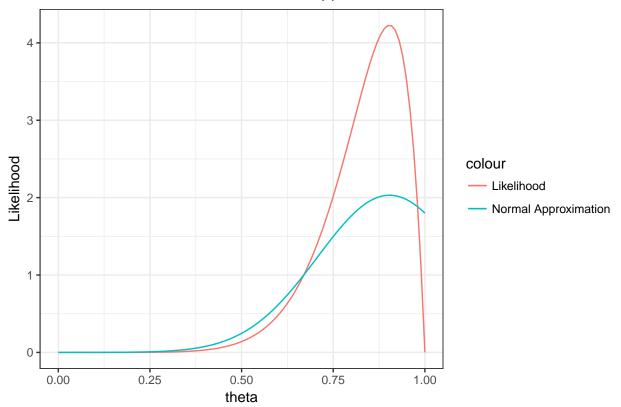
## Normalized Likelihood and Normal Approximation



This normal approximation looks pretty good.

#### (d) Assymptotic Normality round 2

## Normalized Likelihood and Normal Approximation



The normal approximation does not look so good here.