

Regularization methods in multiple regression

Malgorzata Bogdan

University of Wroclaw

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$X_{n \times p}$ - matrix of regressors

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$$-X'Y + (X'X + \gamma I)b = 0 \Leftrightarrow b = (X'X + \gamma I)^{-1}X'Y$$

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$$Tr[M] = \sum_{i=1}^p \lambda_i(M), \text{ where } \lambda_1(M), \dots, \lambda_n(M) \text{ are eigenvalues of } M$$

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Eigenvalues of M

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$$\hat{P}E = RSS + 2\sigma^2 \sum_{i=1}^p \frac{\lambda_i(X'X)}{\lambda_i(X'X) + \gamma}$$

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$$\begin{aligned} E(\hat{\beta}_i - \beta_i)^2 &= E\left(\frac{1}{1+\gamma}\beta_i - \beta_i + \frac{1}{1+\gamma}Z_i\right)^2 \\ &= \frac{\gamma^2}{(1+\gamma)^2}\beta_i^2 + \frac{\sigma^2}{(1+\gamma)^2} \end{aligned}$$

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$$\gamma < \frac{2p\sigma^2}{\|\beta\|^2 - p\sigma^2}$$

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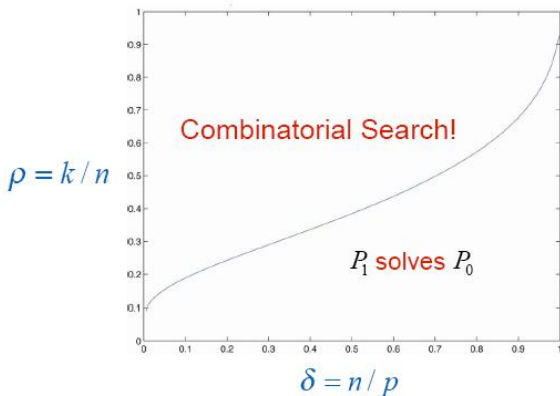
Basis Pursuit can recover β if k is small enough.

Transition curve (Donoho and Tanner, 2005)

Let's assume that $p \rightarrow \infty$, $n/p \rightarrow \delta$ and $k/n \rightarrow \epsilon$.

If X_{ij} are iid $N(0, \tau^2)$ then the probability that BP recovers β converges to 1 if $\epsilon < \rho(\delta)$ and to 0 if $\epsilon > \rho(\delta)$, where $\rho(\delta)$ is the *transition curve*.

Phase Transition: (l_1, l_0) equivalence



Victoria Stodden

Department of Statistics, Stanford University

Noisy case - multiple regression

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The convex function $f(x)$ attains a minimum at x_0 if and only if $0 \in \partial_f(x_0)$.

LASSO for the orthogonal design $X'X = I$

$$\beta^{LS} = Y'X, \quad \|Y - Xb\|^2 + \lambda \|b\|_1 = Y'Y + \sum_{i=1}^p f_i(b_i)$$

$$f_i(x) = x^2 - 2\beta_i^{LS}x + \lambda|x|$$

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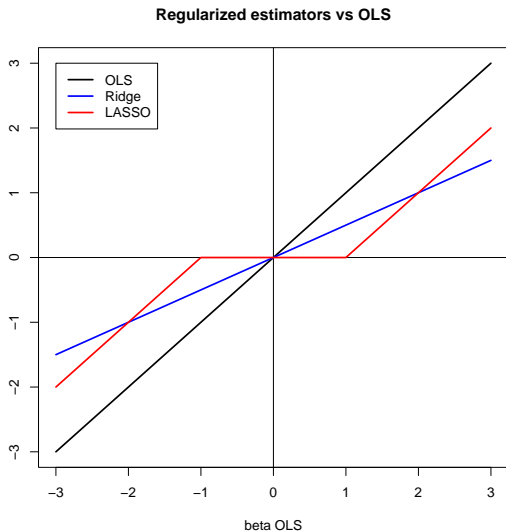
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$$\lambda\partial_{|x|}(x_0) = \begin{cases} \lambda & \text{for } x_0 > 0 \\ -\lambda & \text{for } x_0 < 0 \\ < -\lambda, \lambda > & \text{for } x_0 = 0 \end{cases}$$

LASSO for the orthogonal design

$$\hat{\beta}_i^L = \begin{cases} \beta_i^{LS} - \lambda/2 & \text{when } \beta_i^{LS} > \lambda/2 \\ -\beta_i^{LS} + \lambda/2 & \text{when } \beta_i^{LS} < -\lambda/2 \\ 0 & \text{when } |\beta_i^{LS}| < \lambda/2 \end{cases}$$

Regularized estimators vs OLS



Selection of the tuning parameter for LASSO

- General rule: the reduction of λ_L results in identification of more elements from the true support (true discoveries) but at the same time it produces more falsely identified variables (false discoveries)
- The choice of λ_L is challenging- e.g. crossvalidation typically leads to many false discoveries
- When $X^T X = I$ Lasso selects X_j iff $|\hat{\beta}_j^{LS}| > \lambda$
- Selection $\lambda = \sigma \Phi^{-1}(1 - \alpha/(2p)) \approx \sigma \sqrt{2 \log p}$ corresponds to Bonferroni correction and controls FWER.

Irrepresentability condition

The sign vector of β is defined as

$$S(\beta) = (S(\beta_1), \dots, S(\beta_p)) \in \{-1, 0, 1\}^p,$$

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Let $I := \{i \in \{1, \dots, p\} \mid \beta_i \neq 0\}$, and let $X_I, X_{\bar{I}}$ be matrices whose columns are respectively $(X_i)_{i \in I}$ and $(X_i)_{i \notin I}$.

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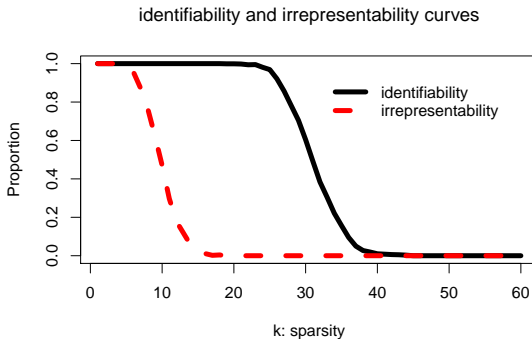
When

$$\|X_{\bar{I}}' X_I (X_I' X_I)^{-1} S(\beta_I)\|_{\infty} > 1$$

then probability of the support recovery by LASSO is smaller than 0.5 (Wainwright, 2009).

Irrepresentability and identifiability curves

$n=100$, $p=300$, elements of X were generated as iid $N(0,1)$



Definition (Identifiability)

Let X be a $n \times p$ matrix. The vector $\beta \in R^p$ is said to be identifiable with respect to the l^1 norm if the following implication holds

$$X\gamma = X\beta \text{ and } \gamma \neq \beta \Rightarrow \|\gamma\|_1 > \|\beta\|_1. \quad (1)$$

Theorem (Tardivel, Bogdan, 2019)

For any $\lambda > 0$ LASSO can separate well the causal and null features if and only if vector β is identifiable with respect to l_1 norm and $\min_{i \in I} |\beta_i|$ is sufficiently large.

Corollary

Appropriately thresholded LASSO can properly identify the sign of sufficiently large β if and only if β is identifiable with respect to l_1 norm.

Conjecture

Adaptive (reweighted) LASSO can properly identify the sign of sufficiently large β if and only if β is identifiable with respect to l_1 norm.

Adaptive LASSO [Zou, *JASA* 2006], [Candès, Wakin and Boyd, *J. Fourier Anal. Appl.* 2008]

$$\beta_{aL} = \operatorname{argmin}_b \left\{ \frac{1}{2} \|y - Xb\|_2^2 + \lambda \sum_{i=1}^p w_i |b|_i \right\}, \quad (2)$$

where $w_i = \frac{1}{\hat{\beta}_i}$, and $\hat{\beta}_i$ is some consistent estimator of β_i .

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Reduces bias and improves model selection properties

$$X_{ij} \sim \mathcal{N}(0, 1/n), \quad z_i \sim \mathcal{N}(0, \sigma^2)$$

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$$\tau^2 = \sigma^2 + \frac{1}{\delta} \mathbb{E} \left(\eta_{\alpha\tau}(\Pi + \tau Z) - \Pi \right)^2,$$

$$\lambda = \left(1 - \frac{1}{\delta} \mathbb{P}(|\Pi + \tau Z| > \alpha\tau) \right) \alpha\tau.$$

Theorem

For any pseudo-Lipschitz function φ , the lasso solution $\hat{\beta}$ with fixed λ obeys

$$\frac{1}{p} \sum_{i=1}^p \varphi(\hat{\beta}_i, \beta_i) \longrightarrow \mathbb{E} \varphi(\eta_{\alpha\tau}(\Pi + \tau Z), \Pi)$$

$\hat{\mathcal{S}}$ - set of variables selected by LASSO

$$FDP \equiv \frac{|\hat{\mathcal{S}} \cap \mathcal{H}_0|}{|\hat{\mathcal{S}}|}$$

$$FDR = E(FDP)$$

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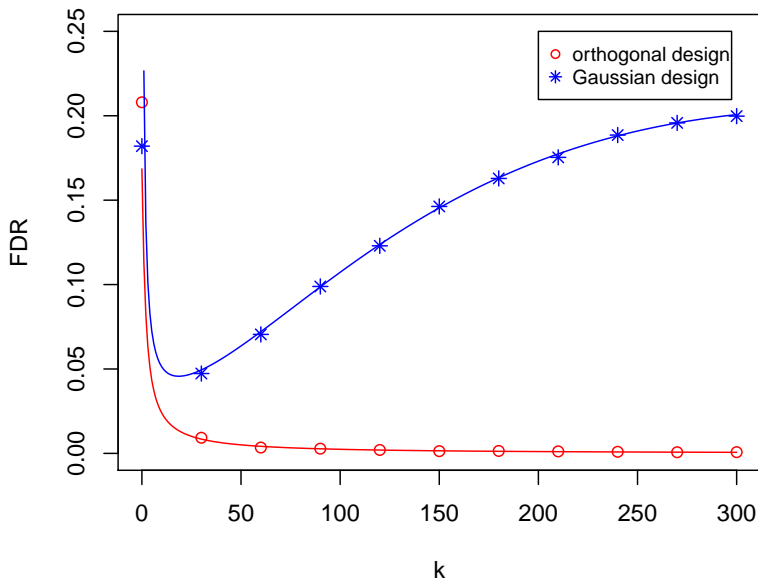
$$FDR = E(FDP)$$

Bogdan, van den Berg, Su and Candés, 2013

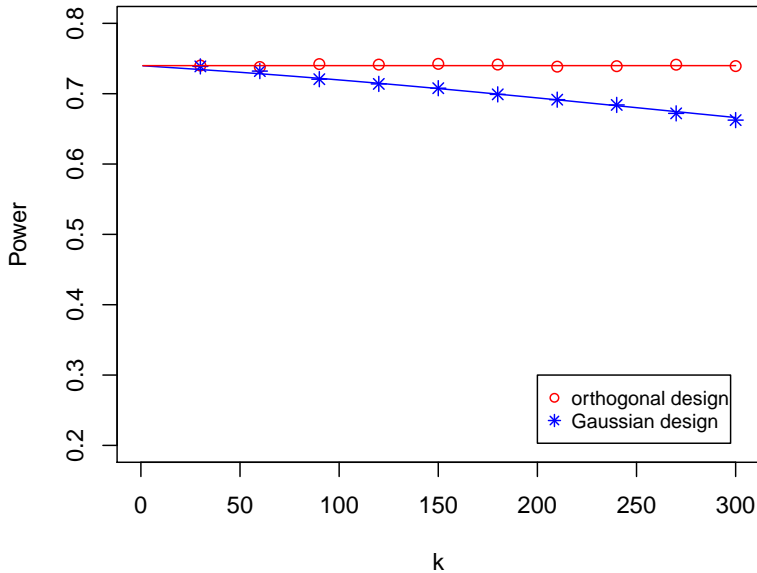
$$FDR \rightarrow \frac{2\mathbb{P}(\Pi = 0)\Phi(-\alpha)}{\mathbb{P}(|\Pi + \tau Z| > \alpha\tau)} ,$$

$$\text{Power} \rightarrow \mathbb{P}(|\Pi + \tau Z| > \alpha\tau | \Pi \neq 0).$$

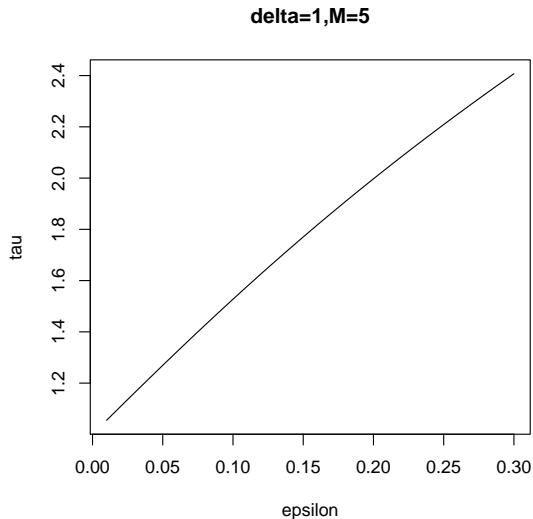
FDR - illustration



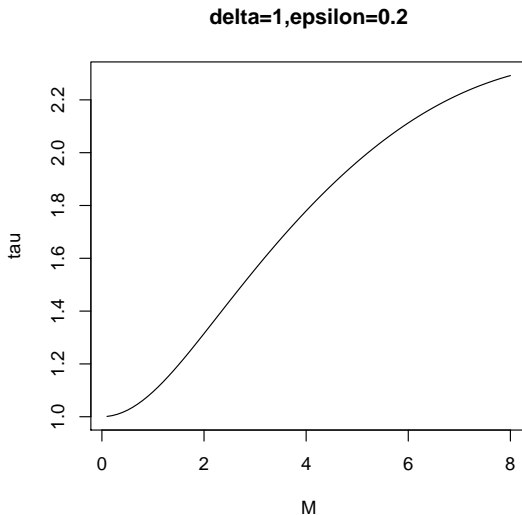
Power - illustration



Magnitude of additional noise (1)

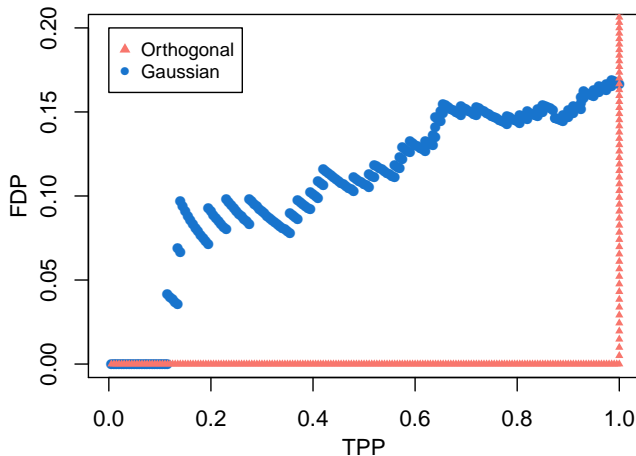


Magnitude of additional noise (2)



False Discoveries along the lasso path

Su, Bogdan and Candes, (2017), $\delta = 1$, $\epsilon = 0.2$



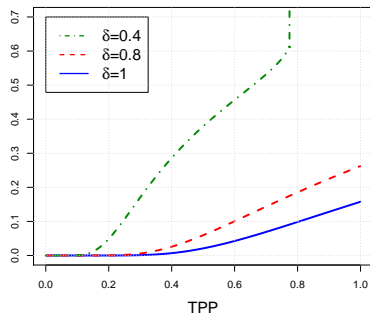
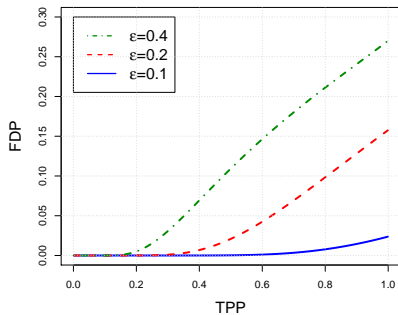
Theorem (Su, Bogdan, Candes, 2017)

Fix $\delta \in (0, \infty)$ and $\epsilon \in (0, 1)$. Then the event

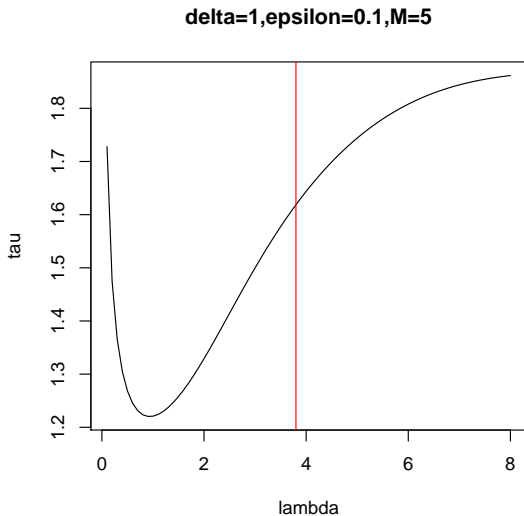
$$\bigcap_{\lambda \geq 0.01} \left\{ FDP(\lambda) \geq q^*(TPP(\lambda)) - 0.001 \right\} \quad (3)$$

holds with probability tending to one.

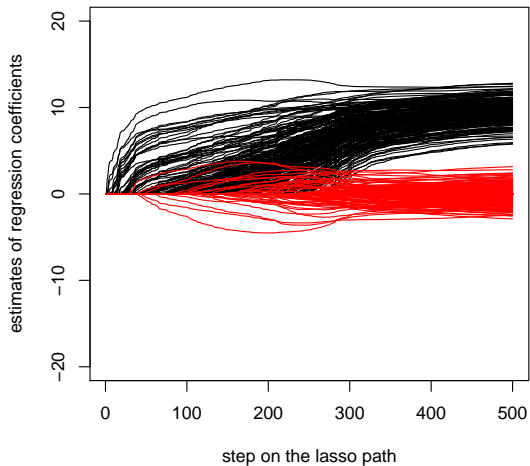
FDR-Power trade-off (2)



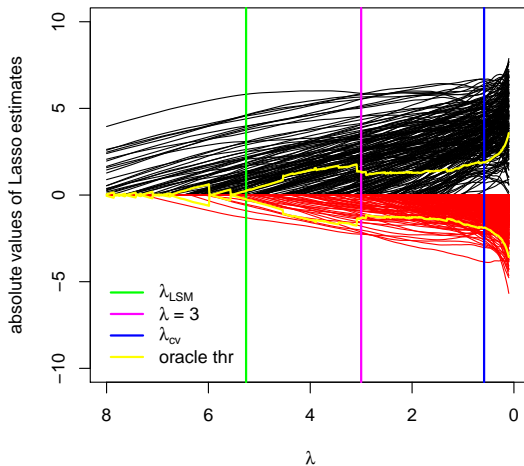
Magnitude of noise



Thresholded LASSO (1)



Thresholded LASSO (2)



Thresholded LASSO (3)

