### Lecture #23

CS170 Spring 2021

### Streaming Algorithms Motivation

- · Vast amounts of data "streaming" by, too much to store
  - · Search engine tracking clicks on websites
  - · Router monitoring network traffic
  - · Data arriving from sensors
- . Is there a much (exponentially) smaller data structure that we can quickly update on the fly, and query when needed?
- · Monte Carlo algorithms: fast, probably accurate

# 3 Examples of Streaming Algorithms · Simple: Counting Total Sales · Input: n sales with prices pypz,...pn · Desired output: P= 3 pi · Initialize C=O, Update C=C+pi, Query: return C Memory Reguirement: 110g2 Pl · Morris's Alg. for Approximate Counting · Like above, with pi=1 · Goal: use O(log2 log2n) bits, not llog2n] · Flajolet & Martin (FM) Alg. for Distinct Elements · Count # distinct integers among i,..., im · Goal: use O (log2n) bits, ij ∈ {1,..,n}

Randomized Approximate Counting · Goal: compute estimate ñ of n where P(|ñ-n|>E.n) <5 tor some OLE, SLI, that you can choose, using O(log\_(log\_n)) bits · Chebyshev's Inequality:  $P(|X - E(x)| > \lambda) = P(|X - E(x)|^2 > \lambda^2)$  $\leq \mathbb{E}(|X-\mathbb{E}(X)|^2)/\lambda^2$  by Markov's Ineq. =  $Var(X)/\lambda^2$  definition of Variance - X, Y independent => Var(aX+b1)=a Var(X)+b Var(Y) · Ex: X, ..., Xn independent, identically distributed (i.i.d.) S= L Z X; = ES=E(Xi), Var(S)= L Var(Xi) 3

First Try at Randomized (ounting Initialize: C=0

Update: 
$$c=c+1$$
 with probability  $p$ 
Query: return  $\tilde{n}=c/p$ 

Let  $X_i=1$  w.p.  $p$ ,  $O$  w.p.  $1-p$ , i.i.d.  $\Rightarrow c=\tilde{Z}_i$   $X_i$ 

Thm:  $E(\tilde{n})=\frac{1}{p}E(c)=\frac{1}{p}\tilde{Z}_iE(X_i)=\frac{1}{p}\tilde{Z}_ip=n$ 

Thm:  $Var(\tilde{n})=\frac{1}{p}Var(c)=\frac{1}{p}\tilde{Z}_iVar(X_i)=\frac{n}{p}[p-\tilde{p}]=\frac{n(1-p)}{p}$ 

Do we save any bits?

#bits =  $\Gamma \log_2 c \tilde{I} \approx \lceil \log_2 n p \rceil \approx \log_2 n - \log_2 \frac{1}{p} \Rightarrow \text{Save} \log_2 \frac{1}{p} \text{ bits}$ 

Are we accurate? Not if we save many bits  $(p \ll 1)$ 
 $P(|n-\tilde{n}| > \epsilon n) \leq \frac{Var(\tilde{n})}{(\epsilon n)^2} = \frac{(1-p)}{\epsilon^2 p n}$ 

Morris's Algorithm: (1/3)Initialize: X=0 Update: X=X+1 with probability 2x Query: return ñ=2x-1 · Let Xn = X after n updates ·Thm: E(x) = E(2<sup>xn</sup>-1) = n, or E(2<sup>xn</sup>) = n+1 Proot-induction on n; base case is n=0  $\mathbb{E}(2^{x_{n+1}}) = \mathbb{E}P(x_n = j \text{ and we increment } X) \cdot 2^{j+1}$ +P(Xn=j and we don't) - 2t  $= \sum_{j=0}^{2} P(X_{n}=j) \cdot \frac{1}{2^{j}} \cdot 2^{j+1} + P(X_{n}=j) \cdot (1-\frac{1}{2^{j}}) 2^{j}$  $= \frac{2}{5} P(X_{n} = i) (2 + 2^{i} - 1) = \frac{2}{5} P(X_{n} = i) + \frac{2}{5} P(X_{n} = i) - 2^{i}$   $= 1 + \mathbb{E}(2^{X_{n}}) = 1 + (n+1) = n+2$ 

(2/3)Morris's Algorithm: Initialize: X=0 Update: X=X+1 with probability 2x Query: return ñ=2x-1 · Let Xn = X after n updates · Thm: E(x) = E(2n-1) = n Intuition: we are approximating log\_n insted of n => need log\_ (log\_n) bits exponentially fewer than login needed for n. Could we use even fewer bits to approximate n to within a factor  $1\pm\epsilon^2$ . No: need to distinguish  $[1,(1+\epsilon)^2],[(1+\epsilon)^4]...[(1+\epsilon)^4]...[(1+\epsilon)^4],...[(1+\epsilon)^4],...[(1+\epsilon)^4]...$ Hintervals = log2n(1te) = θ (log2n) = need Ω(log2log2n) bits

(3/3)Morris's Algorithm: Initialize: X=0 Update: X=X+( with probability 2x Query: return ñ=2x-1 · Let Xn = X after n updates · Thm: E(x) = E(2^n-1) = n \*Thm: Var (ñ) = 1 n2 - 1 n - 1 Proot: E(22xn) = (3/2)n2+(3/2)n+1 by induction  $\mathbb{E}(2^{2\times n+1}) = \underset{j=0}{\overset{\infty}{\geq}} P(x_n = j \text{ and we increment } X) \cdot 2^{2(j+1)}$ + P(Xn=j and we don't) - 22(j)  $= --- = (3/2)(n+1)^2 + (3/2)(n+1) + 1$ · Chebysher: P(|ñ-h|>En) = 1/2 = 1/2 = 00ps 7

Making Morris's Algorithm more accurate

- ·Runs "copies" of Morris, yielding ñ.,..., ñs, return average ñ= 1 3 ñ;
  - · E(ñ) = E(ñi) = n, Var(ñ) = = Var(ñi)
- $-P(\ln-\pi)>_{\epsilon}n)<\frac{1}{5}\frac{1}{2\epsilon^{2}}<\delta$  if  $s>\frac{1}{2\epsilon^{2}\delta}=\Theta(\frac{1}{\epsilon^{2}\delta})$
- · How many bits do we need?
- · Intuition: s copies of Morris > O( 1 loglogzn) bits
- More carefuly: with probability 1-5, need  $O\left(\frac{1}{5^25} \cdot \log_2 \log_2 \frac{n}{55}\right) \text{ bits}$

Flajolet + Muller (FM) Alg. for Distinct Element Counting

- · Given stream i, iz, ..., im, each i; = {1,..,n}
  count t= ## distint elements in stream
- ·Ex if stream is 1,2,7,2,3,7, t=4
- ·Straightforward solutions:
  - 1) Keep array of n bits y, initially all y; =0
    set yix = 1 when ix appears
  - 2) Store whole stream in memory => 0 (m. log, n) bits
- Goal: use o(n) bits to compute  $\tilde{t}$  where  $P(|t-\tilde{t}|>\epsilon t)<\delta$

## I dealized FM Alg.

- · Goal: count t=#distinct elements in i,.., im, ij & ? I... n}
- · Pick random function h: [1, .., n] [0,1]
  - · Each h(i) is i.i.d. random real number uniformly distributed in [0,1]
- · Initialize: X=)

Update: X=min(X, h(i))

Query: return == x-1

Intuition: X = min of t distinct uniform random i.i.d. numbers in [0,1], we expect  $X \sim \frac{1}{t+1}$ , so  $t \sim \frac{1}{x} - 1$  I dealized FM Alg.

· Goal: count t= #distinct elements in i,.., im, ij & Sl.. n}

· X = min of t uniform random i.i.d numbers in [0,1]

· Thm: E(X) = +1

Proof: Analogous to discrete case where
$$\mathbb{E}(X) = \underbrace{Zi}_{i=1} p(i) = \underbrace{Zi}_{i=1} p(j) = \underbrace{ZP(XZi)}_{i=1}$$

$$\mathbb{E}(X) = \int_{0}^{\infty} P(X > \lambda) d\lambda = \int_{0}^{1} P(X > \lambda) d\lambda$$

$$=\int_0^1 \int_{c=1}^{t} (1-\lambda)d\lambda = \int_0^1 (1-\lambda)^t d\lambda = \frac{1}{t+1}$$

Making Idealized FM More Accorate

- · Same as Morris: run s copies, average results
- · Thm: Var (x) = t (++1)2(++2)

Proof: Follows from

$$\mathbb{E}(X^{2}) = \int_{0}^{1} P(X^{2} > \lambda) d\lambda = \int_{0}^{1} P(X > \sqrt{\lambda}) d\lambda$$

$$= \int_{0}^{1} (1 - \sqrt{\lambda})^{t} d\lambda \qquad \text{substitute } u = 1 - \sqrt{\lambda}$$

$$= 2 \int_{0}^{1} u^{t} (1 - u) du = \frac{2}{(t+1)(t+2)}$$

Thm: run  $S = \lceil \frac{1}{2^2 \delta} \rceil$  independent copies  $FM_1, ..., FM_s$  withoutputs  $X_1, ..., X_s$ ,  $Z = \frac{1}{5} \stackrel{>}{\geq} X_i$ , output Z = 1/Z - 1. Then  $P(1 \stackrel{\sim}{\xi} - t \mid 2 O(\varepsilon) t) = \delta$ 

Proof: Apply Chebyshev's Ineg to Z

## Making FM Practical

- · Can't generate uniform random real numbers in practice, need an approximation
- · Generate random integers  $\in \{0,1,...,B\}$ , divide final minimum by B
- · Use pseudorandom number generators!

  to approximate actual random numbers

  · Uses O(logn + log B) bits
- · Recent version: Hyperloglog