

## Linear Regression: Preamble

The “best” guess about  $Y$ , if we know only the distribution of  $Y$ , is  $E[Y]$ .

If “best” is Mean Squared Error.

More precisely, the value of  $a$  that minimizes  $E[(Y - a)^2]$  is  $a = E[Y]$ .

**Proof:**

Let  $\hat{Y} := Y - E[Y]$ .

Then,  $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$ .

So,  $E[\hat{Y}c] = 0, \forall c$ . Now,

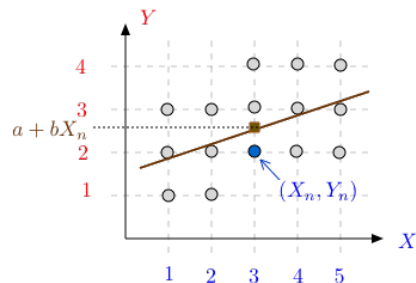
$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a \\ &= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2 \\ &= E[\hat{Y}^2] + 0 + c^2 \geq E[\hat{Y}^2]. \end{aligned}$$

Hence,  $E[(Y - a)^2] \geq E[(Y - E[Y])^2], \forall a$ .  $\square$

## Motivation

Example 2: 15 people.

We look at two attributes:  $(X_n, Y_n)$  of person  $n$ , for  $n = 1, \dots, 15$ :



The line  $Y = a + bX$  is the linear regression.

## Linear Regression: Preamble

Thus, if we want to guess the value of  $Y$ , we choose  $E[Y]$ .

Now assume we make some observation  $X$  related to  $Y$ .

How do we use that observation to improve our guess about  $Y$ ?

The idea is to use a function  $g(X)$  of the observation to estimate  $Y$ .

The simplest function  $g(X)$  is a constant that does not depend of  $X$ .

The next simplest function is linear:  $g(X) = a + bX$ .

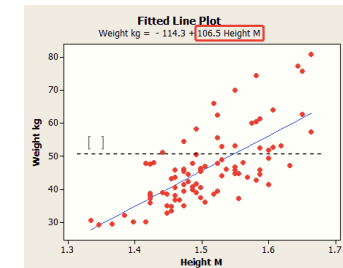
What is the best linear function? That is our next topic.

A bit later, we will consider a general function  $g(X)$ .

## Linear Regression: Motivation

Example 1: 100 people.

Let  $(X_n, Y_n) = (\text{height, weight})$  of person  $n$ , for  $n = 1, \dots, 100$ :



The blue line is  $Y = -114.3 + 106.5X$ . ( $X$  in meters,  $Y$  in kg.)

Best linear fit: [Linear Regression](#).

## LLSE

$LLSE[Y|X]$  - best guess for  $Y$  given  $X$ .

**Theorem**

Consider two RVs  $X, Y$  with a given distribution

$Pr[X = x, Y = y]$ . Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

**Proof 1:**

$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$ .  $E[Y - \hat{Y}] = 0$  by linearity.

Also,  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (next slide)

Combine brown inequalities:  $E[(Y - \hat{Y})(c + dX)] = 0$  for any  $c, d$ .

Since:  $\hat{Y} = \alpha + \beta X$  for some  $\alpha, \beta$ , so  $\exists c, d$  s.t.  $\hat{Y} - \alpha - \beta X = c + dX$ .

Then,  $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$ . Now,

$$\begin{aligned} E[(Y - \alpha - \beta X)^2] &= E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - \alpha - \beta X)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

This shows that  $E[(Y - \hat{Y})^2] \leq E[(Y - \alpha - \beta X)^2]$ , for all  $(\alpha, \beta)$ .

Thus  $\hat{Y}$  is the LLSE.  $\square$

## A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Hence,  $E[Y - \hat{Y}] = 0$ . We want to show that  $E[(Y - \hat{Y})X] = 0$ .

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because  $E[(Y - \hat{Y})E[X]] = 0$ .

Now,

$$\begin{aligned} E[(Y - \hat{Y})(X - E[X])] &= E[(Y - E[Y])(X - E[X])] - \frac{\text{cov}(X, Y)}{\text{var}(X)} E[(X - E[X])(X - E[X])] \\ &= E[(Y - E[Y])(X - E[X])] - \frac{\text{cov}(X, Y)}{\text{var}(X)} \text{var}[X] = 0. \quad \square \end{aligned}$$

(\*) Recall that  $\text{cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$  and  $\text{var}[X] = E[(X - E[X])^2]$ .