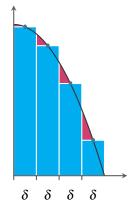
Survey

Fill it out!! https://forms.gle/XL79oruU8BHrQcaeA

Calculus



Riemann Sum/Integral: $\int_a^b f(x)dx = \lim_{\delta \to 0} \sum_i \delta f(a_i)$

Derivative (Rate of change):

$$F'(x) = \lim_{h \to 0} \frac{F(x+h) - F(x)}{h}$$
.

Fundamental Theorem: $F(b) - F(a) = \int_a^b F'(x) dx$.

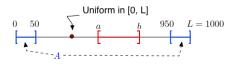
CS70: Continuous Probability.

Continuous Probability 1

- 1. Examples
- Events
- 3. Continuous Random Variables

Choose a real number X, uniformly at random in [0,1].

What is the probability that X is exactly equal to 1/3? Well, ..., 0.



What is the probability that X is exactly equal to 0.6? Again, 0.

In fact, for any $x \in [0,1]$, one has Pr[X = x] = 0.

How should we then describe 'choosing uniformly at random in [0,1]'? Here is the way to do it:

$$Pr[X \in [a,b]] = b - a, \forall 0 \le a \le b \le 1.$$

Makes sense: b - a is the fraction of [0,1] that [a,b] covers.

Let [a,b] denote the **event** that the point X is in the interval [a,b].

$$Pr[[a,b]] = \frac{\text{length of } [a,b]}{\text{length of } [0,1]} = \frac{b-a}{1} = b-a.$$

Intervals like $[a,b] \subseteq \Omega = [0,1]$ are **events.**

More generally, events in this space are unions of intervals.

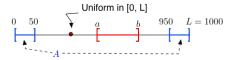
Example: the event A - "within 0.2 of 0 or 1" is $A = [0,0.2] \cup [0.8,1]$. Thus.

$$Pr[A] = Pr[[0,0.2]] + Pr[[0.8,1]] = 0.4.$$

More generally, if A_n are pairwise disjoint intervals in [0,1], then

$$Pr[\cup_n A_n] := \sum_n Pr[A_n].$$

Many subsets of [0,1] are of this form. Thus, the probability of those sets is well defined. We call such sets events.



Note: A radical change in approach.

Finite prob. space: $\Omega = \{1, 2, ..., N\}$, with $Pr[\omega] = p_{\omega}$.

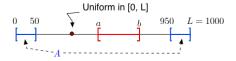
$$\implies Pr[A] = \sum_{\omega \in A} p_{\omega} \text{ for } A \subset \Omega.$$

Continuous space: e.g., $\Omega = [0, 1]$,

 $Pr[\omega]$ is typically 0.

Instead, start with Pr[A] for some events A.

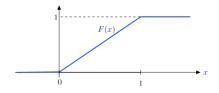
Event A = interval, or union of intervals.



$$Pr[X \le x] = x \text{ for } x \in [0,1]. \text{ Also, } Pr[X \le x] = 0 \text{ for } x < 0.$$

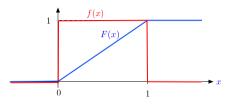
 $Pr[X \le x] = 1 \text{ for } .2x > 1.$

Define $F(x) = Pr[X \le x]$.



Then we have $Pr[X \in (a,b]] = Pr[X \le b] - Pr[X \le a] = F(b) - F(a)$.

Thus, $F(\cdot)$ specifies the probability of all the events!



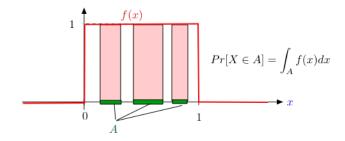
$$Pr[X \in (a,b]] = Pr[X \le b] - Pr[X \le a] = F(b) - F(a).$$

An alternative view is to define $f(x) = \frac{d}{dx}F(x) = 1\{x \in [0,1]\}$. Then

$$F(b) - F(a) = \int_a^b f(x) dx.$$

Thus, the probability of an event is the integral of f(x) over the event:

$$Pr[X \in A] = \int_A f(x) dx.$$

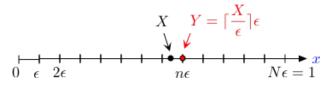


Think of f(x) as describing how one unit of probability is spread over [0,1]: uniformly!

Then $Pr[X \in A]$ is the probability mass over A.

Observe:

- This makes the probability automatically additive.
- We need $f(x) \ge 0$ and $\int_{-\infty}^{\infty} f(x) dx = 1$.



Discrete Approximation: Fix $N \gg 1$ and let $\varepsilon = 1/N$.

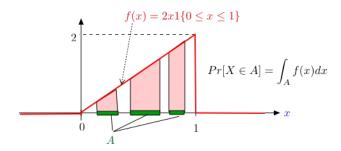
Define $Y = n\varepsilon$ if $(n-1)\varepsilon < X \le n\varepsilon$ for n = 1, ..., N.

Then $|X - Y| \le \varepsilon$ and Y is discrete: $Y \in \{\varepsilon, 2\varepsilon, ..., N\varepsilon\}$.

Also, $Pr[Y = n\varepsilon] = \frac{1}{N}$ for n = 1, ..., N.

Thus, X is 'almost discrete.'

Calculus view: $Pr[Y = n\varepsilon]$ is area of rectangle in Riemann sum.



This figure shows a different choice of $f(x) \ge 0$ with $\int_{-\infty}^{\infty} f(x) dx = 1$.

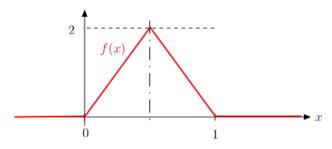
It defines another way of choosing X at random in [0,1].

Note that *X* is more likely to be closer to 1 than to 0.

One has $Pr[X \le x] = \int_{-\infty}^{x} f(u) du = x^2$ for $x \in [0, 1]$.

Also, $Pr[X \in (x, x + \varepsilon)] = \int_{x}^{x+\varepsilon} f(u) du \approx f(x)\varepsilon$.

Another Nonuniform Choice at Random in [0,1].



This figure shows yet a different choice of $f(x) \ge 0$ with $\int_{-\infty}^{\infty} f(x) dx = 1$.

It defines another way of choosing X at random in [0,1].

Note that X is more likely to be closer to 1/2 than to 0 or 1.

For instance, $Pr[X \in [0, 1/3]] = \int_0^{1/3} 4x dx = 2[x^2]_0^{1/3} = \frac{2}{9}$.

Thus, $Pr[X \in [0,1/3]] = Pr[X \in [2/3,1]] = \frac{2}{9}$ and $Pr[X \in [1/3,2/3]] = \frac{5}{9}$.

General Random Choice in R

Let F(x) be a nondecreasing function with $F(-\infty) = 0$ and $F(+\infty) = 1$.

Define X by $Pr[X \in (a,b]] = F(b) - F(a)$ for a < b. Also, for $a_1 < b_1 < a_2 < b_2 < \cdots < b_n$,

$$Pr[X \in (a_1, b_1] \cup (a_2, b_2] \cup (a_n, b_n]]$$

$$= Pr[X \in (a_1, b_1]] + \dots + Pr[X \in (a_n, b_n]]$$

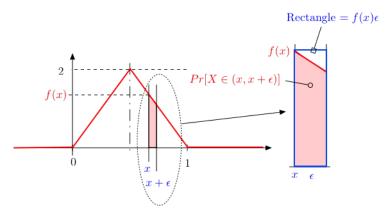
$$= F(b_1) - F(a_1) + \dots + F(b_n) - F(a_n).$$

Let
$$f(x) = \frac{d}{dx}F(x)$$
. Then,
$$Pr[X \in (x, x + \varepsilon]] = F(x + \varepsilon) - F(x) \approx f(x)\varepsilon.$$

F(x) is cumulative distribution function (cdf) of X f(x) is the probability density function (pdf) of X. When F and f correspond RV X: $F_X(x)$ and $f_X(x)$.

$$Pr[X \in (x, x + \varepsilon)]$$

An illustration of $Pr[X \in (x, x + \varepsilon)] \approx f_X(x)\varepsilon$:



Thus, the pdf is the 'local probability by unit length.' It is the 'probability density.'

Discrete Approximation

Fix $\varepsilon \ll 1$ and let $Y = n\varepsilon$ if $X \in (n\varepsilon, (n+1)\varepsilon]$.

Thus,
$$Pr[Y = n\varepsilon] = F_X((n+1)\varepsilon) - F_X(n\varepsilon)$$
.

Note that $|X - Y| \le \varepsilon$ and Y is a discrete random variable.

Also, if
$$f_X(x) = \frac{d}{dx} F_X(x)$$
, then $F_X(x+\varepsilon) - F_X(x) \approx f_X(x)\varepsilon$.

Hence, $Pr[Y = n\varepsilon] \approx f_X(n\varepsilon)\varepsilon$.

Thus, we can think of X of being almost discrete with $Pr[X = n\varepsilon] \approx f_X(n\varepsilon)\varepsilon$.

Example: CDF

Example: hitting random location on gas tank.

Random location on circle.



Random Variable: Y distance from center.

Probability within y of center:

$$Pr[Y \le y] = \frac{\text{area of small circle}}{\text{area of dartboard}}$$

= $\frac{\pi y^2}{\pi} = y^2$.

Hence,

$$F_Y(y) = Pr[Y \le y] = \begin{cases} 0 & \text{for } y < 0 \\ y^2 & \text{for } 0 \le y \le 1 \\ 1 & \text{for } y > 1 \end{cases}$$

Calculation of event with dartboard...

Probability between .5 and .6 of center? Recall CDF.

$$F_Y(y) = Pr[Y \le y] = \begin{cases} 0 & \text{for } y < 0 \\ y^2 & \text{for } 0 \le y \le 1 \\ 1 & \text{for } y > 1 \end{cases}$$

$$Pr[0.5 < Y \le 0.6] = Pr[Y \le 0.6] - Pr[Y \le 0.5]$$

= $F_Y(0.6) - F_Y(0.5)$
= $.36 - .25$
= $.11$

PDF.

Example: "Dart" board.

Recall that

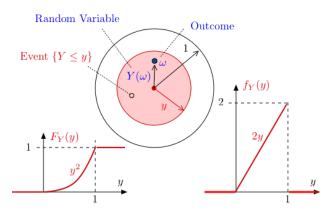
$$F_{Y}(y) = Pr[Y \le y] = \begin{cases} 0 & \text{for } y < 0 \\ y^{2} & \text{for } 0 \le y \le 1 \\ 1 & \text{for } y > 1 \end{cases}$$

$$f_{Y}(y) = F'_{Y}(y) = \begin{cases} 0 & \text{for } y < 0 \\ 2y & \text{for } 0 \le y \le 1 \\ 0 & \text{for } y > 1 \end{cases}$$

The cumulative distribution function (cdf) and probability distribution function (pdf) give full information.

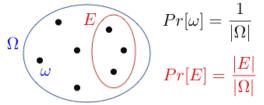
Use whichever is convenient.

Target



U[a,b]

Uniform Probability Space

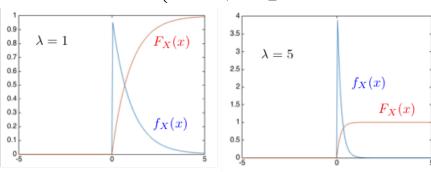


$Expo(\lambda)$

The exponential distribution with parameter $\lambda > 0$ is defined by

$$f_X(x) = \lambda e^{-\lambda x} \mathbf{1}\{x \ge 0\}$$

$$F_X(x) = \left\{ egin{array}{ll} 0, & ext{if } x < 0 \ 1 - e^{-\lambda x}, & ext{if } x \geq 0. \end{array}
ight.$$



Note that $Pr[X > t] = e^{-\lambda t}$ for t > 0.

Continuous Random Variables

Continuous random variable X, specified by

1. $F_X(x) = Pr[X \le x]$ for all x. Cumulative Distribution Function (cdf).

$$Pr[a < X \le b] = F_X(b) - F_X(a)$$

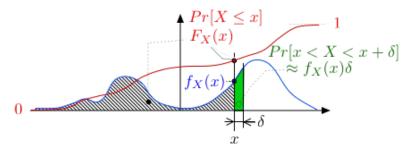
- 1.1 $0 \le F_X(x) \le 1$ for all $x \in \Re$.
- 1.2 $F_X(x) \le F_X(y)$ if $x \le y$.
- 2. Or $f_X(x)$, where $F_X(x) = \int_{-\infty}^x f_X(u) du$ or $f_X(x) = \frac{d(F_X(x))}{dx}$. Probability Density Function (pdf).

$$Pr[a < X \le b] = \int_a^b f_X(x) dx = F_X(b) - F_X(a)$$

- 2.1 $f_X(x) \ge 0$ for all $x \in \Re$.
- 2.2 $\int_{-\infty}^{\infty} f_X(x) dx = 1$.

Recall that $Pr[X \in (x, x + \delta)] \approx f_X(x)\delta$. X "takes" value $n\delta$, for $n \in Z$, with $Pr[X = n\delta] = f_X(n\delta)\delta$

A Picture



The pdf $f_X(x)$ is a nonnegative function that integrates to 1.

The cdf $F_X(x)$ is the integral of f_X .

$$Pr[x < X < x + \delta] \approx f_X(x)\delta$$

 $Pr[X \le x] = F_X(x) = \int_{-\infty}^{x} f_X(u)du$

Multiple Continuous Random Variables

One defines a pair (X, Y) of continuous RVs by specifying $f_{X,Y}(x,y)$ for $x, y \in \Re$ where

$$f_{X,Y}(x,y)dxdy = Pr[X \in (x,x+dx), Y \in (y+dy)].$$

The function $f_{X,Y}(x,y)$ is called the joint pdf of X and Y.

Example: Choose a point (X, Y) uniformly in the set $A \subset \Re^2$. Then

$$f_{X,Y}(x,y) = \frac{1}{|A|} 1\{(x,y) \in A\}$$

where |A| is the area of A.

Interpretation. Think of (X, Y) as being discrete on a grid with mesh size ε and $Pr[X = m\varepsilon, Y = n\varepsilon] = f_{X,Y}(m\varepsilon, n\varepsilon)\varepsilon^2$.

Recall Marginal Distribution:

$$Pr[X = x] = \sum_{v} Pr[X = x, Y = y].$$

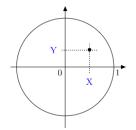
Similarly:

$$f_X(x) = \int f_{X,Y}(x,y) dy$$
.

Sum "goes to" integral.

Example of Continuous (X, Y)

Pick a point (X, Y) uniformly in the unit circle.



Thus,
$$f_{X,Y}(x,y) = \frac{1}{\pi} 1\{x^2 + y^2 \le 1\}.$$

Consequently,

$$Pr[X > 0, Y > 0] = \frac{1}{4}$$

$$Pr[X < 0, Y > 0] = \frac{1}{4}$$

$$Pr[X^2 + Y^2 \le r^2] = \frac{\pi r^2}{\pi} = r^2$$

$$Pr[X > Y] = \frac{1}{2}.$$

Independent Continuous Random Variables

Definition: Continuous RVs X and Y independent if and only if

$$Pr[X \in A, Y \in B] = Pr[X \in A]Pr[Y \in B], \forall A, B.$$

Theorem: Continuous RVs X and Y independent if and only if

$$f_{X,Y}(x,y) = f_X(x)f_Y(y).$$

Note: $f_X(x)$ ($f_Y(y)$) is (marginal) distribution of X (Y).

Proof: Intervals: A = [x, x + dx], B = [y, y + dy].

$$Pr[X \in A, Y \in B] = Pr[X \in A] \times Pr[Y \in B]$$

$$\approx f_X(x) \ dx \times f_Y(y) \ dy$$

$$= f_X(x)f_Y(y) \ dxdy.$$

Thus, $f_{X,Y}(x,y) = f_X(x)f_Y(y)$.

Mutual Independence.

Definition: Continuous RVs $X_1, ..., X_n$ are mutually independent if

$$Pr[X_1 \in A_1, \dots, X_n \in A_n] = Pr[X_1 \in A_1] \cdots Pr[X_n \in A_n], \forall A_1, \dots, A_n.$$

Theorem: Continuous RVs $X_1, ..., X_n$ are mutually independent if and only if

$$f_{\mathbf{X}}(x_1,\ldots,x_n)=f_{X_1}(x_1)\cdots f_{X_n}(x_n).$$

Proof: As in the discrete case.

Conditional density.

Conditional Density: $f_{X|Y}(x,y)$.

Conditional Probability:
$$Pr[X \in A | Y \in B] = \frac{Pr[X \in A, Y \in B]}{Pr[Y \in B]}$$

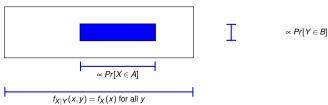
$$Pr[X \in [x, x + dx]|Y \in [y, y + dy]] = \frac{f_{X,Y}(x,y)dxdy}{f_Ydy}$$

$$f_{X|Y}(x,y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{f_{X,Y}(x,y)}{\int_{-\infty}^{+\infty} f_{X,Y}(x,y)dy}$$

Corollary: For independent random variables, $f_{X|Y}(x,y) = f_X(x)$.

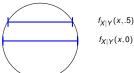
Independent Random Variables?

Uniform on a rectangle? Independent?



Also: $Pr[X \in A, Y \in B] \propto \text{Area of rectangle} \propto Pr[X \in A] \times Pr[Y \in B].$ Independent!

Uniform on a circle? Independent?



 $f_{X|Y}(x,.5)$

Not independent!

Summary

Continuous Probability 1

- 1. pdf: $Pr[X \in (x, x + \delta]] = f_X(x)\delta$.
- 2. CDF: $Pr[X \le x] = F_X(x) = \int_{-\infty}^x f_X(y) dy$.
- 3. U[a,b]: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\}$; $F_X(x) = \frac{x-a}{b-a}$ for $a \le x \le b$.
- 4. $Expo(\lambda)$: $f_X(x) = \lambda \exp\{-\lambda x\} 1\{x \ge 0\}; F_X(x) = 1 - \exp\{-\lambda x\} \text{ for } x \le 0.$
- 5. Target: $f_X(x) = 2x1\{0 \le x \le 1\}$; $F_X(x) = x^2$ for $0 \le x \le 1$.
- 6. Joint pdf: $Pr[X \in (x, x + \delta), Y = (y, y + \delta)) = f_{X,Y}(x, y)\delta^2$.
 - 6.1 Conditional Distribution: $f_{X|Y}(x,y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$.
 - 6.2 Independence: $f_{X|Y}(x,y) = f_X(x)$

Summary

Continuous Probability

- Continuous RVs are essentially the same as discrete RVs
- ▶ Think that $X \approx x$ with probability $f_X(x)\varepsilon$
- Sums become integrals,
- The exponential distribution is magical: memoryless.