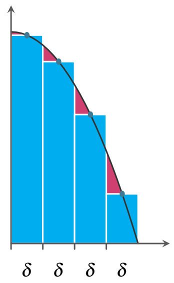
Survey

Fill it out!!
tinyurl.com/cs70-survey

Calculus



Riemann Sum/Integral: $\int_a^b f(x)dx = lim_{\delta \to 0} \sum_i \delta f(a_i)$
"Area is defined as rectangles and add up some thin ones."

Derivative (Rate of change):
$F'(x) = \lim_{h \to 0} \frac{F(x+h)-F(x)}{h}$.
"Rise over run of close together points."

Fundamental Theorem: $F(b) - F(a) = \int_a^b F'(x)dx$.
"Area $(F(\cdot))$ under $f(x)$ grows at x, $F'(x)$, by $f(x)$"
Thus $F'(x) = f(x)$.

CS70: Continuous Probability.

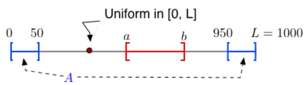
Continuous Probability 1

1. Examples
2. Events
3. Continuous Random Variables

Uniformly at Random in $[0,1]$.

Choose a real number X, uniformly at random in $[0,1]$.

What is the probability that X is exactly equal to $1/3$? Well, ..., 0.



What is the probability that X is exactly equal to 0.6? Again, 0.

In fact, for any $x \in [0,1]$, one has $Pr[X = x] = 0$.

How should we then describe 'choosing uniformly at random in $[0,1]$'?

Here is the way to do it:

$$Pr[X \in [a,b]] = b - a, \forall 0 \le a \le b \le 1.$$

Makes sense: $b - a$ is the fraction of $[0,1]$ that $[a,b]$ covers.

Poll

$$F_X(x) = Pr[X \le x]$$

.

$$f_X(x) = \lim_{\delta \to 0} Pr[X \in (x, x+\delta)]$$

What is true?
(A) $F_X(x) = \int_{-\infty}^{\infty} f_X(y)dy$
(B) $\int_{-\infty}^{\infty} f_X(x) = 1$
(C) $F_X(x) = \int_{-\infty}^{x} f(y)dy$.
(D) $f(x) = F_X'(x)$.
(E) $\int_{-\infty}^{\infty} F_X(x)dx = 1$.
(F) $\int_{-\infty}^{\infty} xf(x)dx = \int_{-\infty}^{\infty} (1 - F(x))dx$.

(A) False. limits wrong. (B) cuz probability distribution.
(C) "sums up probability of rectangles", e.g. calculus.
(D) calculus, fundamental theorem.

(F) is true since $\int_{-\infty}^{\infty} xf(x)dx = \int_{-\infty}^{\infty} F(x)dx = E[X]$.

Next lecture.

Uniformly at Random in $[0,1]$.

Let $[a,b]$ denote the **event** that the point X is in the interval $[a,b]$.

$$Pr[[a,b]] = \frac{\text{length of } [a,b]}{\text{length of } [0,1]} = \frac{b-a}{1} = b - a.$$

Intervals like $[a,b] \subseteq \Omega = [0,1]$ are **events.**
More generally, events in this space are unions of intervals.
Example: the event A - "within 0.2 of 0 or 1" is $A = [0,0.2] \cup [0.8,1]$.
Thus,

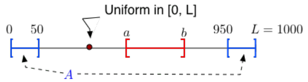
$$Pr[A] = Pr[[0,0.2]] + Pr[[0.8,1]] = 0.4.$$

More generally, if A_n are pairwise disjoint intervals in $[0,1]$, then

$$Pr[\cup_n A_n] := \sum_n Pr[A_n].$$

Many subsets of $[0,1]$ are of this form. Thus, the probability of those sets is well defined. We call such sets events.

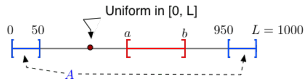
Uniformly at Random in $[0,1]$.



Note: A **radical** change in approach.
Finite prob. space: $\Omega = \{1,2,\ldots,N\}$, with $Pr[\omega] = p_\omega$.
$\implies Pr[A] = \sum_{\omega\in A} p_\omega$ for $A \subset \Omega$.

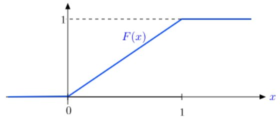
Continuous space: e.g., $\Omega = [0,1]$,
$Pr[\omega]$ is typically 0.
Instead, start with $Pr[A]$ for some events A.
Event A = interval, or union of intervals.

Uniformly at Random in $[0,1]$.



$Pr[X \le x] = x$ for $x \in [0,1]$. Also, $Pr[X \le x] = 0$ for $x < 0$.
$Pr[X \le x] = 1$ for .2$x > 1$.

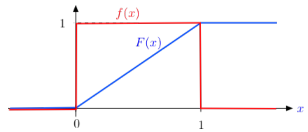
Define $F(x) = Pr[X \le x]$.



Then we have $Pr[X \in (a,b]] = Pr[X \le b] - Pr[X \le a] = F(b) - F(a)$.

Thus, $F(\cdot)$ specifies the probability of all the events!

Uniformly at Random in $[0,1]$.



$Pr[X \in (a,b]] = Pr[X \le b] - Pr[X \le a] = F(b) - F(a)$.

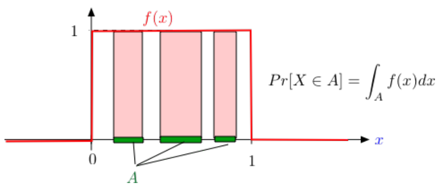
An alternative view is to define $f(x) = \frac{d}{dx}F(x) = 1\{x \in [0,1]\}$. Then

$$F(b) - F(a) = \int_a^b f(x)dx.$$

Thus, the probability of an event is the integral of $f(x)$ over the event:

$$Pr[X \in A] = \int_A f(x)dx.$$

Uniformly at Random in $[0,1]$.



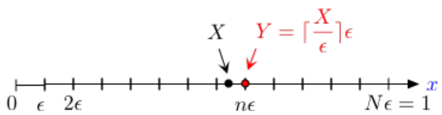
Think of $f(x)$ as describing how
one unit of probability is spread over $[0,1]$: uniformly!

Then $Pr[X \in A]$ is the probability mass over A.

Observe:

- This makes the probability automatically additive.
- We need $f(x) \ge 0$ and $\int_{-\infty}^{\infty} f(x)dx = 1$.

Uniformly at Random in $[0,1]$.



Discrete Approximation: Fix $N \gg 1$ and let $\varepsilon = 1/N$.

Define $Y = n\varepsilon$ if $(n-1)\varepsilon < X \le n\varepsilon$ for $n = 1,\ldots,N$.

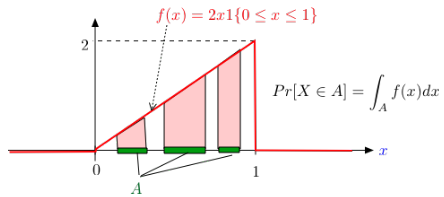
Then $|X - Y| \le \varepsilon$ and Y is discrete: $Y \in \{\varepsilon, 2\varepsilon, \ldots, N\varepsilon\}$.

Also, $Pr[Y = n\varepsilon] = \frac{1}{N}$ for $n = 1,\ldots,N$.

Thus, X is 'almost discrete.'

Calculus view: $Pr[Y = n\varepsilon]$ is area of rectangle in Riemann sum.

Nonuniformly at Random in $[0,1]$.



This figure shows a different choice of $f(x) \ge 0$ with $\int_{-\infty}^{\infty} f(x)dx = 1$.

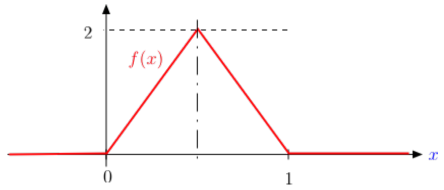
It defines another way of choosing X at random in $[0,1]$.

Note that X is more likely to be closer to 1 than to 0.

One has $Pr[X \le x] = \int_{-\infty}^{x} f(u)du = x^2$ for $x \in [0,1]$.

Also, $Pr[X \in (x, x+\varepsilon)] = \int_x^{x+\varepsilon} f(u)du \approx f(x)\varepsilon$.

Another Nonuniform Choice at Random in $[0,1]$.



This figure shows yet a different choice of $f(x) \geq 0$ with $\int_{-\infty}^{\infty} f(x)dx = 1$.

It defines another way of choosing X at random in $[0,1]$.

Note that X is more likely to be closer to $1/2$ than to 0 or 1.

For instance, $Pr[X \in [0,1/3]] = \int_0^{1/3} 4xdx = 2[x^2]_0^{1/3} = \frac{2}{9}$.

Thus, $Pr[X \in [0,1/3]] = Pr[X \in [2/3,1]] = \frac{2}{9}$ and $Pr[X \in [1/3,2/3]] = \frac{5}{9}$.

General Random Choice in $\Re$

Let $F(x)$ be a nondecreasing function with $F(-\infty) = 0$ and $F(+\infty) = 1$.

Define X by $Pr[X \in (a,b]] = F(b) - F(a)$ for $a < b$. Also, for $a_1 < b_1 < a_2 < b_2 < \cdots < b_n$,

$$\begin{aligned} & Pr[X \in (a_1,b_1] \cup (a_2,b_2] \cup (a_n,b_n]] \\ &= Pr[X \in (a_1,b_1]] + \cdots + Pr[X \in (a_n,b_n]] \\ &= F(b_1) - F(a_1) + \cdots + F(b_n) - F(a_n). \end{aligned}$$

Let $f(x) = \frac{d}{dx}F(x)$. Then,

$$Pr[X \in (x, x+\varepsilon]] = F(x+\varepsilon) - F(x) \approx f(x)\varepsilon.$$

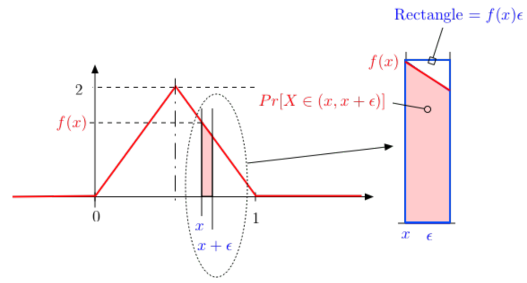
$F(x)$ is cumulative distribution function (cdf) of X

$f(x)$ is the probability density function (pdf) of X.

When F and f correspond RV X: $F_X(x)$ and $f_X(x)$.

$Pr[X \in (x, x+\varepsilon)]$

An illustration of $Pr[X \in (x, x+\varepsilon)] \approx f_X(x)\varepsilon$:



Thus, the pdf is the 'local probability by unit length.'

It is the 'probability density.'

Discrete Approximation

Fix $\varepsilon \ll 1$ and let $Y = n\varepsilon$ if $X \in (n\varepsilon, (n+1)\varepsilon]$.

Thus, $Pr[Y = n\varepsilon] = F_X((n+1)\varepsilon) - F_X(n\varepsilon)$.

Note that $|X - Y| \leq \varepsilon$ and Y is a discrete random variable.

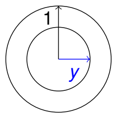
Also, if $f_X(x) = \frac{d}{dx}F_X(x)$, then $F_X(x+\varepsilon) - F_X(x) \approx f_X(x)\varepsilon$.

Hence, $Pr[Y = n\varepsilon] \approx f_X(n\varepsilon)\varepsilon$.

Thus, we can think of X of being almost discrete with $Pr[X = n\varepsilon] \approx f_X(n\varepsilon)\varepsilon$.

Example: CDF, pre-poll

Example: hitting random location on gas tank.
Random location on circle.



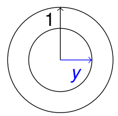
What is probability of being within y of the center, for non-negative $y \leq 1$?

(A) 1.
(B) 0.
(C) $\int_0^y (2\pi y)dy$
(D) y^2.

(D) Next slide.

Example: CDF

Example: hitting random location on gas tank.
Random location on circle.



Random Variable: Y distance from center.
Probability within y of center:

$$\begin{aligned} Pr[Y \leq y] &= \frac{\text{area of small circle}}{\text{area of dartboard}} \\ &= \frac{\pi y^2}{\pi} = y^2. \end{aligned}$$

Hence,

$$F_Y(y) = Pr[Y \leq y] = \begin{cases} 0 & \text{for } y < 0 \\ y^2 & \text{for } 0 \leq y \leq 1 \\ 1 & \text{for } y > 1 \end{cases}$$

Calculation of event with dartboard..

Probability between .5 and .6 of center?
Recall CDF.

$$F_Y(y) = Pr[Y \leq y] = \begin{cases} 0 & \text{for } y < 0 \\ y^2 & \text{for } 0 \leq y \leq 1 \\ 1 & \text{for } y > 1 \end{cases}$$

$$\begin{aligned} Pr[0.5 < Y \leq 0.6] &= Pr[Y \leq 0.6] - Pr[Y \leq 0.5] \\ &= F_Y(0.6) - F_Y(0.5) \\ &= .36 - .25 \\ &= .11 \end{aligned}$$

PDF.

Example: "Dart" board.
Recall that

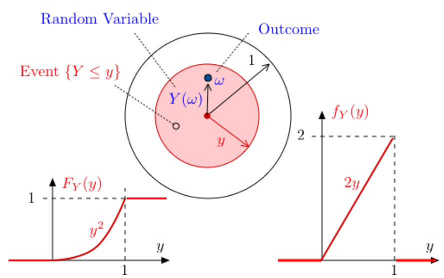
$$F_Y(y) = Pr[Y \leq y] = \begin{cases} 0 & \text{for } y < 0 \\ y^2 & \text{for } 0 \leq y \leq 1 \\ 1 & \text{for } y > 1 \end{cases}$$

$$f_Y(y) = F_Y'(y) = \begin{cases} 0 & \text{for } y < 0 \\ 2y & \text{for } 0 \leq y \leq 1 \\ 0 & \text{for } y > 1 \end{cases}$$

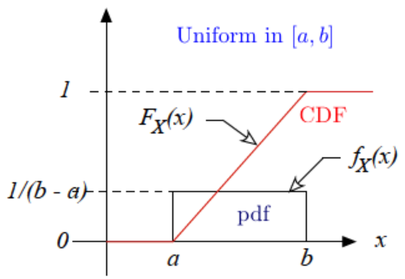
The cumulative distribution function (cdf) and probability distribution function (pdf) give full information.

Use whichever is convenient.

Target



$U[a,b]$



Exponential derivation:Poll.

$Pr[X = i] = (1-p)^{i-1}p$.

Let $p = \lambda/n$. and $Y = X/n$.

What is true?

(A) $X \sim G(p)$
(B) $Pr[X > i] = (1-p)^i$. (C) $Pr[Y > i/n] = (1-\lambda/n)^i$.
(D) $Pr[Y > y] = (1-\lambda/n)^{ny}$.
(E) $\lim_{n\to\infty}(1-\lambda/n)^{ny} = e^{-\lambda y}$.

(A) True by definition.
(B) $Pr[X > i] = (1-p)^i$ at least i coin flips fail.
(C) True, definition of Y
(D) True, $y = i/n$ means $i = ny$.
(E) $(1-\lambda/n)^{ny} = ((1-\lambda/n)^{n/\lambda})^{\lambda y}$
and $\lim_{n\to\infty}(1-\frac{\lambda}{n})^{n/\lambda} = e$.

The limit as $n \to \infty$ of Y has $Pr[Y > y] = e^{-\lambda y}$.

$Pr[Y > y]$ is defined as "Survival function."

$Expo(\lambda)$

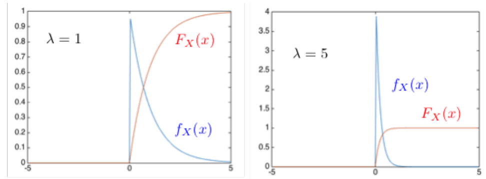
"Limit of geometric."

From last slide: $S(t) = Pr[X > t] = e^{-\lambda t}$ for $t > 0$.
Note: $f_X(x) = F'(t) = 1 - S(t) = -S'(t)$.

The exponential distribution with parameter $\lambda > 0$ is defined by

$$f_X(x) = \lambda e^{-\lambda x} 1\{x \geq 0\}$$

$$F_X(x) = \begin{cases} 0, & \text{if } x < 0 \\ 1 - e^{-\lambda x}, & \text{if } x \geq 0. \end{cases}$$



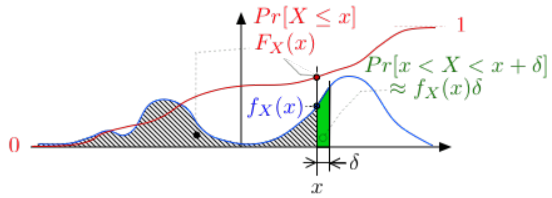
Continuous Random Variables

Continuous random variable X, specified by

1. $F_X(x) = Pr[X \le x]$ for all x.
   Cumulative Distribution Function (cdf).
   $Pr[a < X \le b] = F_X(b) - F_X(a)$
   1.1 $0 \le F_X(x) \le 1$ for all $x \in \Re$.
   1.2 $F_X(x) \le F_X(y)$ if $x \le y$.
2. Or $f_X(x)$, where $F_X(x) = \int_{-\infty}^{x} f_X(u)du$ or $f_X(x) = \frac{d(F_X(x))}{dx}$.
   Probability Density Function (pdf).
   $Pr[a < X \le b] = \int_a^b f_X(x)dx = F_X(b) - F_X(a)$
   2.1 $f_X(x) \ge 0$ for all $x \in \Re$.
   2.2 $\int_{-\infty}^{\infty} f_X(x)dx = 1$.

Recall that $Pr[X \in (x, x+\delta)] \approx f_X(x)\delta$.
X "takes" value $n\delta$, for $n \in Z$, with $Pr[X = n\delta] = f_X(n\delta)\delta$

A Picture



The pdf $f_X(x)$ is a nonnegative function that integrates to 1.

The cdf $F_X(x)$ is the integral of f_X.

$$Pr[x < X < x+\delta] \approx f_X(x)\delta$$
$$Pr[X \le x] = F_x(x) = \int_{-\infty}^{x} f_X(u)du$$

Multiple Continuous Random Variables

One defines a pair (X, Y) of continuous RVs by specifying $f_{X,Y}(x,y)$ for $x, y \in \Re$ where

$f_{X,Y}(x,y)dxdy = Pr[X \in (x, x+dx), Y \in (y+dy)]$.

The function $f_{X,Y}(x,y)$ is called the joint pdf of X and Y.

Example: Choose a point (X, Y) uniformly in the set $A \subset \Re^2$. Then

$f_{X,Y}(x,y) = \frac{1}{|A|} 1\{(x,y) \in A\}$

where $|A|$ is the area of A.

Interpretation. Think of (X, Y) as being discrete on a grid with mesh size ε and $Pr[X = m\varepsilon, Y = n\varepsilon] = f_{X,Y}(m\varepsilon, n\varepsilon)\varepsilon^2$.

Recall Marginal Distribution:

$Pr[X = x] = \sum_y Pr[X = x, Y = y]$.

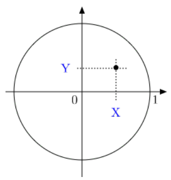
Similarly:

$f_X(x) = \int f_{X,Y}(x,y)dy$.

Sum "goes to" integral.

Example of Continuous (X, Y)

Pick a point (X, Y) uniformly in the unit circle.



Thus, $f_{X,Y}(x,y) = \frac{1}{\pi} 1\{x^2 + y^2 \le 1\}$.

Consequently,

$$Pr[X > 0, Y > 0] = \frac{1}{4}$$
$$Pr[X < 0, Y > 0] = \frac{1}{4}$$
$$Pr[X^2 + Y^2 \le r^2] = \frac{\pi r^2}{\pi} = r^2$$
$$Pr[X > Y] = \frac{1}{2}.$$

Independent Continuous Random Variables

Definition: Continuous RVs X and Y independent if and only if

$$Pr[X \in A, Y \in B] = Pr[X \in A]Pr[Y \in B], \forall A, B.$$

Theorem: Continuous RVs X and Y independent if and only if

$$f_{X,Y}(x,y) = f_X(x)f_Y(y).$$

Note: $f_X(x)$ ($f_Y(y)$) is (marginal) distribution of X (Y).

Proof: Intervals: $A = [x, x+dx]$, $B = [y, y+dy]$.

$$\begin{aligned} Pr[X \in A, Y \in B] &= Pr[X \in A] \times Pr[Y \in B] \\ &\approx f_X(x)\ dx \times f_Y(y)\ dy \\ &= f_X(x)f_Y(y)\ dxdy. \end{aligned}$$

Thus, $f_{X,Y}(x,y) = f_X(x)f_Y(y)$. $\square$

Mutual Independence.

Definition: Continuous RVs $X_1, \dots, X_n$ are mutually independent if

$$Pr[X_1 \in A_1, \dots, X_n \in A_n] = Pr[X_1 \in A_1] \cdots Pr[X_n \in A_n], \forall A_1, \dots, A_n.$$

Theorem: Continuous RVs $X_1, \dots, X_n$ are mutually independent if and only if

$$f_{\mathbf{X}}(x_1, \dots, x_n) = f_{X_1}(x_1) \cdots f_{X_n}(x_n).$$

Proof: As in the discrete case.

Conditional density.

Conditional Density: $f_{X|Y}(x,y)$.

Conditional Probability: $Pr[X \in A | Y \in B] = \frac{Pr[X \in A, Y \in B]}{Pr[Y \in B]}$

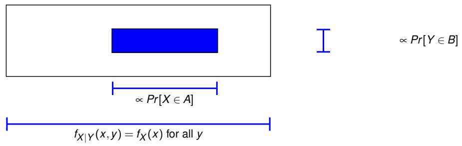
$Pr[X \in [x, x+dx] | Y \in [y, y+dy]] = \frac{f_{X,Y}(x,y)dxdy}{f_Y dy}$

$f_{X|Y}(x,y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{f_{X,Y}(x,y)}{\int_{-\infty}^{+\infty} f_{X,Y}(x,y)dy}$

Corollary: For independent random variables, $f_{X|Y}(x,y) = f_X(x)$.

Independent Random Variables?

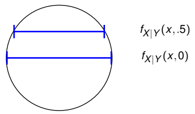
Uniform on a rectangle? Independent?



Also: $Pr[X \in A, Y \in B] \propto$ Area of rectangle $\propto Pr[X \in A] \times Pr[Y \in B]$.

Independent!

Uniform on a circle? Independent?



Not independent!

Summary

Continuous Probability 1

1. pdf: $Pr[X \in (x, x+\delta]] = f_X(x)\delta$.
2. CDF: $Pr[X \le x] = F_X(x) = \int_{-\infty}^{x} f_X(y)dy$.
3. $U[a,b]$: $f_X(x) = \frac{1}{b-a} 1\{a \le x \le b\}; F_X(x) = \frac{x-a}{b-a}$ for $a \le x \le b$.
4. $Expo(\lambda)$:
   $f_X(x) = \lambda \exp\{-\lambda x\} 1\{x \ge 0\}; F_X(x) = 1 - \exp\{-\lambda x\}$ for $x \le 0$.
5. Target: $f_X(x) = 2x 1\{0 \le x \le 1\}; F_X(x) = x^2$ for $0 \le x \le 1$.
6. Joint pdf: $Pr[X \in (x, x+\delta), Y = (y, y+\delta)) = f_{X,Y}(x,y)\delta^2$.
   6.1 Conditional Distribution: $f_{X|Y}(x,y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$.
   6.2 Independence: $f_{X|Y}(x,y) = f_X(x)$

Summary

Continuous Probability

- Continuous RVs are similar to discrete RVs (break into intervals.)
- Think that $X \approx x$ with probability $f_X(x)\varepsilon$
- Sums become integrals, ...