## 1. The risk modeling

Given

$$y_n = \beta_n^T \mathcal{Z} + \sqrt{1 - \beta_n^T \beta_n} \epsilon_n$$

where  $\mathcal{Z} \sim \mathcal{N}\left(0, I_{S}\right)$  and  $\epsilon \sim \mathcal{N}\left(0, I_{N}\right)$ , the resulting  $y_{n} \sim \mathcal{N}\left(0, 1\right)$  as shown below

$$\mathbb{E}\left\{y_n\right\} = \sum_{i} \beta_{n,i} \mathbb{E}\left\{\mathcal{Z}_i\right\} + \sqrt{1 - \beta_n^T \beta_n} \mathbb{E}\left\{\epsilon_n\right\} = 0$$

$$var \{y_n\} = \mathbb{E} \left\{ (y_n - \mathbb{E} \{y_n\})^2 \right\}$$

$$= \mathbb{E} \left\{ (\beta_n^T \mathcal{Z} + \sqrt{1 - \beta_n^T \beta_n} \epsilon_n)^2 \right\}$$

$$= \beta_n^T \beta_n \mathbb{E} \left\{ (\mathcal{Z} - 0)^2 \right\} + (1 - \beta_n^T \beta_n) \mathbb{E} \left\{ (\epsilon_n - 0)^2 \right\}$$

$$= \beta_n^T \beta_n var \{\mathcal{Z}\} + (1 - \beta_n^T \beta_n) var \{\epsilon_n\}$$

$$= 1$$

2. Motivation for threshold between different states  $H_{c(n)}^c$ The motivation is to model discrete probability in the credit state matrix with a continuous distribution such as the gaussian. In this cawe, we want to set  $H_{c(n)}^c$  such that

$$p(H_{c(n)}^{c-1} \le y_n \le H_{c(n)}^c) = p_{c(n)}^c$$

therefore we can write

$$p(y_n \le H_{c(n)}^c) = \sum_{\gamma=1}^c p_{c(n)}^{\gamma} \qquad \xrightarrow{y_n \sim \mathcal{N}(0,1)} \qquad \Phi(H_{c(n)}^c) = \sum_{\gamma=1}^c p_{c(n)}^{\gamma}$$

## 3. Confidence interval for monte carlo estimation

$$p(L_N(\mathcal{Z}, \epsilon) \ge l) \in p(L_N(\mathcal{Z}, \epsilon) \ge l) \pm CI$$

Idea is that for the two naive algorithms that are purported to be equivalent, the CI should be approximately the same

### 1. Likelihood Function for Two Level IS

Likelihood for the inner sampling conditioned on Z is given by

$$e^{\theta_x(Z)L+\psi(\theta_x(Z),Z)}$$
 where  $\psi(\theta) = \sum_{k=1}^m \log(1 + p_k(e^{\theta c_k} - 1))$ 

The likelihood function for the outer sampling of Z consists of the following change of distribution

$$Z \sim \mathcal{N}(0, I) \longrightarrow Z \sim \mathcal{N}(\mu, I)$$

where  $\mu$  is the twisting parameter for the outer importance sampling such that the resulting shifted normal distribution resembles the zero variance IS distribution, in other words,

$$\mu = \max_{z} P(L > x | Z = z)e^{\frac{-z^{T}z}{2}}$$

Then the likelihood for the outer IS is then

$$\frac{\mathcal{N}(0,I)}{\mathcal{N}(\mu,I)} = \frac{exp(-\frac{1}{2}z^{T}z)}{exp(-\frac{1}{2}(z-\mu)^{T}(z-\mu))}$$

$$= exp\left(-\frac{1}{2}z^{T}z - \frac{1}{2}z^{T}z + z^{T}\mu - \frac{1}{2}\mu^{T}\mu\right)$$

$$= e^{-\mu^{T}Z + \mu^{T}\mu/2}$$

Therefore, the estimator for probability of tail event is given by

$$\mathbb{1}_{L>x}e^{\theta_x(Z)L+\psi(\theta_x(Z),Z)}e^{-\mu^TZ+\mu^T\mu/2}$$

# Likelihood Expression for Exponential Twisting GL2005

Show

$$\prod_{k=1}^{m} \left( \frac{p_k}{p_{k,\theta}} \right)^{Y_k} \left( \frac{1 - p_k}{1 - p_{k,\theta}} \right)^{1 - Y_k} = e^{-\theta L + \phi(\theta)}$$

where

$$\phi(\theta) = \sum_{k=1}^{m} \log \left( 1 + p_k (e^{\theta c_k} - 1) \right)$$

$$L = \sum_{k=1}^{m} c_k Y_k$$

$$p_{k,\theta} = \frac{p_k e^{\theta c_k}}{1 + p_k (e^{\theta c_k} - 1)}$$

Proof.

$$\begin{split} lhs &= \exp\left\{\sum_{k=1}^{m} Y_k \log\left(\frac{p_k}{p_{k,\theta}}\right) + (1-Y_k) \log\left(\frac{1-p_k}{1-p_{k,\theta}}\right)\right\} \\ &= \exp\left\{\sum_{k=1}^{m} Y_k \log\left(1+p_k(e^{\theta c_k}-1)\right) - Y_k \theta c_k + (1-Y_k) \log\left(1+p_k\left(e^{\theta c_k}-1\right)\right)\right\} = rhs \end{split}$$

where

$$1 - p_{k,\theta} = \frac{1 - p_k}{1 + p_k \left( e^{\theta c_k} - 1 \right)}$$

# Glasserman&Li Algorithm

- 1. Outer Level IS for sysmatic risk Z
  - (a) Find shifted parameter  $\mu$  for outer IS for Z
  - (b) Sample  $Z \sim \mathcal{N}(\mu, I)$
- 2. Inner Level IS for each default indicators  $Y_k$ 
  - (a) Calculate conditional default probabilities  $p_k(Z)$  for  $k=1,\cdots,m$

$$p_k(Z) = P(Y_k = 1|Z) = p(X_k > x_k|Z) = P(a_k Z + b_k \epsilon_k > \Phi^{-1}(1 - p_k)|Z)$$

(b) Compute the twisted parameters  $\theta_x(Z)$ 

$$\theta_x(Z) = \begin{cases} \text{solution to } \frac{\partial}{\partial \theta} \psi_m(\theta, Z) = x & \psi'(0) = \mathbb{E}_p \left\{ L | Z \right\} = \sum_{k=1}^m p_k(Z) c_k < x \\ 0 & otherwise \end{cases}$$

(c) Compute default indicators (bernoulli) from twisted conditional default probabilities

$$p_{k,\theta_x(Z)}(Z) = \frac{p_k(Z)e^{\theta_x(Z)c_k}}{1 + p_k(Z)(e^{\theta_x(Z)c_k} - 1)}$$
  $k = 1, \dots, m$ 

- (d) Compute Loss  $L = c_1 Y_1 + \cdots + c_m Y_m$  under twisted distribution
- 3. Return the estimator of tail proabilities

$$\mathbb{1}_{L>x}e^{\theta_x(Z)L+\psi(\theta_x(Z),Z)}e^{-\mu^TZ+\mu^T\mu/2}$$

Therefore,

$$P(L>x) = \mathbb{E}_{Z \sim \mathcal{N}(\mu,I) Y_k \sim p_{k,\theta_x(Z)}} \left\{ \mathbb{1}_{L>x} e^{\theta_x(Z)L + \psi(\theta_x(Z),Z)} e^{-\mu^T Z + \mu^T \mu/2} \right\}$$

# Optimization

Want to solve for

$$\max_{x} \min_{y} f(x, y)$$

Same as

$$\max_{x} f(x, \hat{y}(x)) \quad \text{where} \quad \hat{y}(x) = \arg\min_{y} f(x, y)$$

Simply we write as a function of 1 variable

$$\max_{x} \hat{f}(x) \quad \text{where} \quad \hat{f}(x) = f(x, \hat{y}(x))$$

Want to compute the 1st order and 2nd order derivatives

$$\hat{f}'(x) = \frac{\partial f}{\partial x} f(x, \hat{y}(x)) + \frac{\partial f}{\partial y} f(x, \hat{y}(x)) \hat{y}'(x)$$

Want to compute  $\hat{y}'(x)$  first find critical points

$$\frac{\partial f}{\partial y}f(x,y)|_{y=\hat{y}(x)} = 0$$

Solve for the function

$$f_{yx} = \frac{\partial}{\partial x} \frac{\partial}{\partial y} f(x, y)|_{y = \hat{y}(x)}$$

Solve for  $\hat{y}'(x)$ 

$$f_{yx}(x,\hat{y}(x)) + f_{yy}(x,\hat{y}(x))\hat{y}'(x) = 0 \quad \to \quad \hat{y}'(x) = -\frac{f_{yx}(x,\hat{y}(x))}{f_{yy}(x,\hat{y}(x))}$$

Then compute second derivative, i.e.  $\hat{f}''(x)$