

Exercise 11

Dienstag, 6. Dezember 2022 18:09

(a) The null hypothesis states that there is no γ source: $s_0 = 0$. Under this assumption, what value and what error result for b_0 according to the maximum likelihood method?

$$\ln L = N_{\text{off}} \ln(b) + N_{\text{on}} \ln(\alpha b) - (1+\alpha)b - \ln(N_{\text{off}}!) - \ln(N_{\text{on}}!)$$

$$\frac{d(\ln L)}{db} = \frac{N_{\text{off}}}{b} + \frac{N_{\text{on}}}{b} - (1+\alpha) \stackrel{!}{=} 0$$

$$\Leftrightarrow \frac{N_{\text{off}} + N_{\text{on}}}{b} = 1 + \alpha$$

$$\Rightarrow \hat{b} = \frac{N_{\text{off}} + N_{\text{on}}}{1 + \alpha}$$

$$\frac{d^2(\ln L)}{db^2} = \frac{N_{\text{off}} + N_{\text{on}}}{b^2} = \frac{1 + \alpha}{N_{\text{off}} + N_{\text{on}}}$$

$$\Rightarrow \hat{b}_0 = \frac{N_{\text{off}} + N_{\text{on}}}{1 + \alpha} = \frac{1 + \alpha}{N_{\text{off}} + N_{\text{on}}}$$

b)

$$\lambda = \frac{L(b_0, s_0)}{L(b_1, s_1)}$$

$$b_0 = \frac{N_{\text{off}} + N_{\text{on}}}{1 + \alpha}, \quad s_0 = 0$$

$$b_1 = N_{\text{off}}, \quad s_1 = N_{\text{on}} - \alpha N_{\text{off}}$$

$$L(b, s) = \frac{s^{N_{\text{off}}} (\alpha b + s)^{N_{\text{on}}}}{N_{\text{off}}! N_{\text{on}}!} e^{-(1+\alpha)b - s}$$

$$\begin{aligned} \Rightarrow \lambda &= \frac{\left(\alpha \frac{N_{\text{off}} + N_{\text{on}}}{1 + \alpha}\right)^{N_{\text{on}}} e^{-\frac{(1+\alpha)(N_{\text{off}} + N_{\text{on}})}{1 + \alpha}}}{(N_{\text{on}} - \alpha N_{\text{off}})^{N_{\text{off}}} (\alpha N_{\text{off}} + N_{\text{on}} - \alpha N_{\text{off}})^{N_{\text{on}}} e^{-(1+\alpha)N_{\text{off}} - N_{\text{on}} + \alpha N_{\text{off}}}} \\ &= \left(\frac{\alpha}{N_{\text{on}}} \frac{N_{\text{off}} + N_{\text{on}}}{1 + \alpha}\right)^{N_{\text{on}}} \cdot \left(\frac{1}{N_{\text{on}} - \alpha N_{\text{off}}}\right)^{N_{\text{off}}} \end{aligned}$$

c)

$$D = -2 \ln(\lambda) = -2 N_{\text{on}} \ln\left(\frac{\alpha}{N_{\text{on}}} \frac{N_{\text{off}} + N_{\text{on}}}{1 + \alpha}\right) + 2 N_{\text{off}} \ln(N_{\text{on}} - \alpha N_{\text{off}})$$

Vergleiche: $N(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}}$

u^2 ist χ^2 verteilt, $-2 \ln(\lambda)$ auch

Exercise 12

Donnerstag, 1. Dezember 2022 10:39

a)

Die Zählraten der einzelnen Bin folgen einer Poisson-Verteilung.

Die PDF lautet also $\frac{e^{-Np_i} (Np_i)^{n_i}}{n_i!}$ & $\frac{e^{-Mp_i} (Mp_i)^{m_i}}{m_i!}$

b)

Die Likelihood ist das Produkt der beiden PDF:

$$L(p_i) = \frac{e^{-Np_i} (Np_i)^{n_i}}{n_i!} \cdot \frac{e^{-Mp_i} (Mp_i)^{m_i}}{m_i!} = \frac{1}{n_i! m_i!} e^{-p_i(N+M)} \cdot N^{n_i} M^{m_i} (p_i)^{n_i+m_i}$$

Maximiere $F = \ln(L) = -p_i(N+M) + (n_i+m_i) \cdot \ln(p_i) + \text{const}$

$$\frac{\partial F}{\partial p_i} \Big|_{\hat{p}_i} = 0$$

$$\Rightarrow -(N+M) + \frac{n_i+m_i}{\hat{p}_i} = 0 \quad \Leftrightarrow \quad \hat{p}_i = \frac{n_i+m_i}{N+M}$$

c)

$$\chi^2 = \sum_{i=1}^r \frac{(n_i - n_i^{\text{Model}})^2}{n_i^{\text{Model}}} \quad n_i^{\text{Model}} = N\hat{p}_i \quad (\text{analog für } M)$$

$$\Rightarrow \chi^2 = \sum_{i=1}^r \left[\frac{(n_i - N \cdot \frac{n_i+m_i}{N+M})^2}{\frac{n_i+m_i}{N+M} \cdot N} + \frac{(m_i - M \cdot \frac{n_i+m_i}{N+M})^2}{\frac{n_i+m_i}{N+M} \cdot M} \right]$$

d)

- 2r Terme werden aufsummiert
 - (r-1) p_i Werte durch Schätzer \hat{p}_i gegeben
 - (r-1), da $\sum p_i = 1 \Rightarrow$ legt ein p_i fest
 - M, N ebenfalls durch Schätzer $N = \sum_i n_i, M = \sum_i m_i$ gegeben
- $\Rightarrow df = 2r - (r-1) - 2 = r-1$ Freiheitsgrade
- Für einen geringen Stichprobenumfang (<10 Einträge/Bin), folgt die Teststatistik keiner χ^2 -Verteilung, weil die Unsicherheiten in diesem Fall zu groß werden.

e)

• 2 Freiheitsgrade

$$N = 111 + 188 + 333 = 632$$

$$M = 15 + 36 + 30 = 81$$

$$p_i = \frac{n_i+m_i}{N+M}$$

$$\Rightarrow p_1 = \frac{111+15}{632+81} = \frac{126}{713} \approx 0,177$$

$$p_2 = \frac{188+36}{632+81} = \frac{224}{713} \approx 0,314$$

$$p_3 = 1 - p_2 - p_1 = \frac{363}{713} \approx 0,509$$

Habe nicht gesehen, dass man die gar nicht mehr braucht ;)

$$\chi^2 = \frac{1}{NM} \sum_i \frac{(Nm_i - Mn_i)^2}{n_i + m_i}$$

$$= \frac{1}{632 \cdot 81} \left(\frac{(632 \cdot 15 - 81 \cdot 111)^2}{111 + 15} + \frac{(632 \cdot 36 - 81 \cdot 188)^2}{188 + 36} + \frac{(632 \cdot 30 - 81 \cdot 333)^2}{333 + 30} \right)$$

$$\approx 8,43$$

$$= \frac{1}{632.81} \left(\frac{(632.73 - 87.7771)}{111 + 15} + \frac{(632.50 - 87.7881)}{188 + 36} + \frac{(632.30 - 87.5551)}{333 + 30} \right)$$

$$\approx 8,43$$

| Degrees of Freedom | Probability of a larger value of χ^2 | | | | | | | | |
|--------------------|---|--------|--------|--------|--------|-------|-------|-------|-------|
| | 0.99 | 0.95 | 0.90 | 0.75 | 0.50 | 0.25 | 0.10 | 0.05 | 0.01 |
| 1 | 0.000 | 0.004 | 0.016 | 0.102 | 0.455 | 1.32 | 2.71 | 3.84 | 6.63 |
| 2 | 0.020 | 0.103 | 0.211 | 0.575 | 1.386 | 2.77 | 4.61 | 5.99 | 9.21 |
| 3 | 0.115 | 0.352 | 0.584 | 1.212 | 2.366 | 4.11 | 6.25 | 7.81 | 11.34 |
| 4 | 0.297 | 0.711 | 1.064 | 1.923 | 3.357 | 5.39 | 7.78 | 9.49 | 13.28 |
| 5 | 0.554 | 1.145 | 1.610 | 2.675 | 4.351 | 6.63 | 9.24 | 11.07 | 15.09 |
| 6 | 0.872 | 1.635 | 2.204 | 3.455 | 5.348 | 7.84 | 10.64 | 12.59 | 16.81 |
| 7 | 1.239 | 2.167 | 2.833 | 4.255 | 6.346 | 9.04 | 12.02 | 14.07 | 18.48 |
| 8 | 1.647 | 2.733 | 3.490 | 5.071 | 7.344 | 10.22 | 13.36 | 15.51 | 20.09 |
| 9 | 2.088 | 3.325 | 4.168 | 5.899 | 8.343 | 11.39 | 14.68 | 16.92 | 21.67 |
| 10 | 2.558 | 3.940 | 4.865 | 6.737 | 9.342 | 12.55 | 15.99 | 18.31 | 23.21 |
| 11 | 3.053 | 4.575 | 5.578 | 7.584 | 10.341 | 13.70 | 17.28 | 19.68 | 24.72 |
| 12 | 3.571 | 5.226 | 6.304 | 8.438 | 11.340 | 14.85 | 18.55 | 21.03 | 26.22 |
| 13 | 4.107 | 5.892 | 7.042 | 9.299 | 12.340 | 15.98 | 19.81 | 22.36 | 27.69 |
| 14 | 4.660 | 6.571 | 7.790 | 10.165 | 13.339 | 17.12 | 21.06 | 23.68 | 29.14 |
| 15 | 5.229 | 7.261 | 8.547 | 11.037 | 14.339 | 18.25 | 22.31 | 25.00 | 30.58 |
| 16 | 5.812 | 7.962 | 9.312 | 11.912 | 15.338 | 19.37 | 23.54 | 26.30 | 32.00 |
| 17 | 6.408 | 8.672 | 10.085 | 12.792 | 16.338 | 20.49 | 24.77 | 27.59 | 33.41 |
| 18 | 7.015 | 9.390 | 10.865 | 13.675 | 17.338 | 21.60 | 25.99 | 28.87 | 34.80 |
| 19 | 7.633 | 10.117 | 11.651 | 14.562 | 18.338 | 22.72 | 27.20 | 30.14 | 36.19 |
| 20 | 8.260 | 10.851 | 12.443 | 15.452 | 19.337 | 23.83 | 28.41 | 31.41 | 37.57 |
| 22 | 9.542 | 12.338 | 14.041 | 17.240 | 21.337 | 26.04 | 30.81 | 33.92 | 40.29 |
| 24 | 10.856 | 13.848 | 15.659 | 19.037 | 23.337 | 28.24 | 33.20 | 36.42 | 42.98 |
| 26 | 12.198 | 15.379 | 17.292 | 20.843 | 25.336 | 30.43 | 35.56 | 38.89 | 45.64 |
| 28 | 13.565 | 16.928 | 18.939 | 22.657 | 27.336 | 32.62 | 37.92 | 41.34 | 48.28 |
| 30 | 14.953 | 18.493 | 20.599 | 24.478 | 29.336 | 34.80 | 40.26 | 43.77 | 50.89 |
| 40 | 22.164 | 26.509 | 29.051 | 33.660 | 39.335 | 45.62 | 51.80 | 55.76 | 63.69 |
| 50 | 27.707 | 34.764 | 37.689 | 42.942 | 49.335 | 56.33 | 63.17 | 67.50 | 76.15 |
| 60 | 37.485 | 43.188 | 46.459 | 52.294 | 59.335 | 66.98 | 74.40 | 79.08 | 88.38 |

(This one gives α .)

H_0 : Werte stammen aus gleicher Verteilung
Berechnet: $\chi^2 \approx 8,43$

$\alpha = 0,1$: $\chi^2 = 4,61 \Rightarrow H_0$ akzeptiert

$\alpha = 0,05$: $\chi^2 = 5,99 \Rightarrow$ " "

$\alpha = 0,01$: $\chi^2 = 9,21 \Rightarrow H_0$ abgelehnt

Typ 2 Fehler: H_0 wird angenommen, obwohl die Hypothese falsch ist.