DEPARTMENT OF COMPUTER SCIENCE UNIVERSITY OF COPENHAGEN



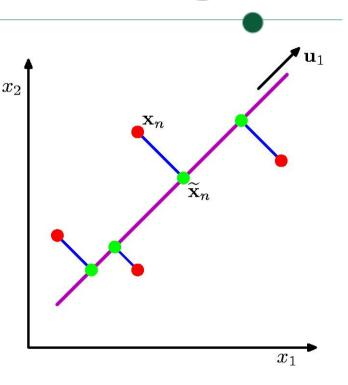
Principal Component Analysis (PCA) Dimensionality and noise reduction

Kim Steenstrup Pedersen

Plan for this lecture



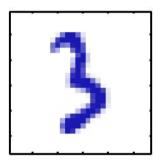
- Continuous latent variable models:
 - Why linear PCA?
- Principal component analysis (PCA):
 - Maximum variance formulation
 - Minimum error formulation
- Applications of PCA:
 - Preprocessing of data for noise reduction
 - Dimensionality reduction
 - Visualization of high dimensional data
- And some computational tricks

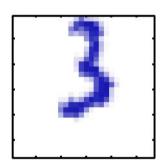


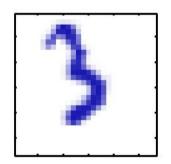


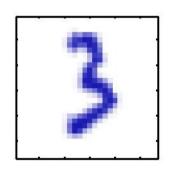


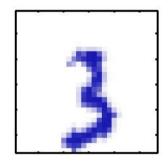
- Often high dimensional data have few degrees of freedom, i.e. a low intrinsic dimensionality.
- Example: Images of hand written digits







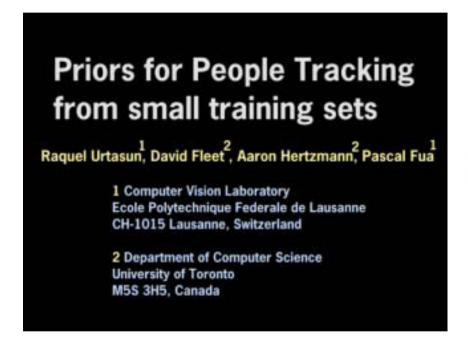




- Dimensionality: 64 x 64 pixels = high dimensionality
- Intrinsic degrees of freedom (< 64 x 64):
 - Easy: Translation (2), rotation (1)
 - Complicated: Degrees of freedom coming from the variability in how to write the digit 3.
 - Not all images represents valid digits the set of digit images is sparsely distributed in the space of images.



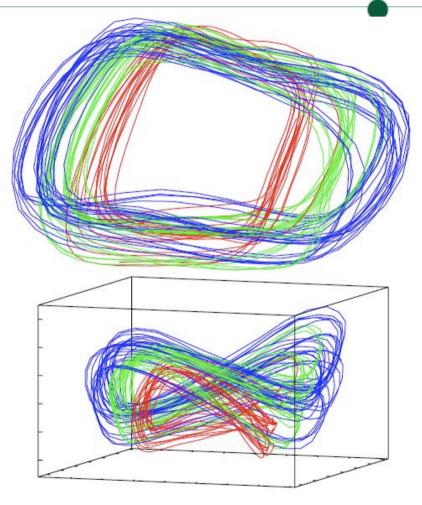




[Urtasun et al, 2005b]

Dimensionality: 15 joints x 3 (3D)

The intrinsic dimensionality of walking is much lower







- We can model the degrees of freedom as latent (hidden) variables z.
- The connection between data representation x and latent variables z is in general some non-linear mapping:

$$\mathbf{x} = \varphi(\mathbf{z}, \varepsilon)$$

Including some noise \mathcal{E} .

• The simplest choice is a linear model with additive noise:

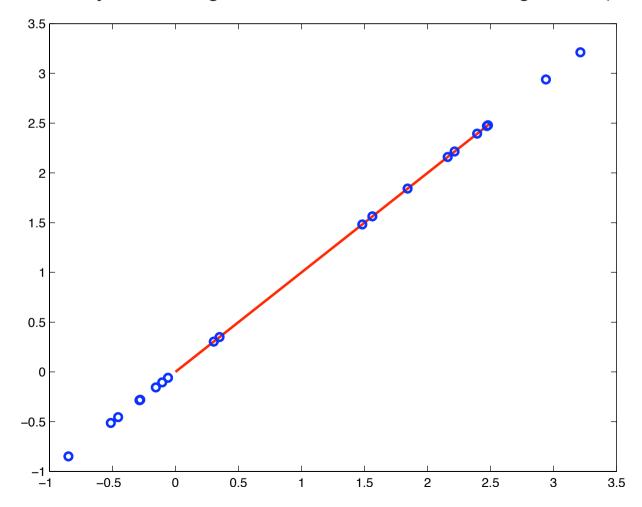
$$\mathbf{x} = \mathbf{A}\mathbf{z} + \mathbf{b} + \boldsymbol{\varepsilon}$$

 Principal component analysis (PCA) is based on a linear model.





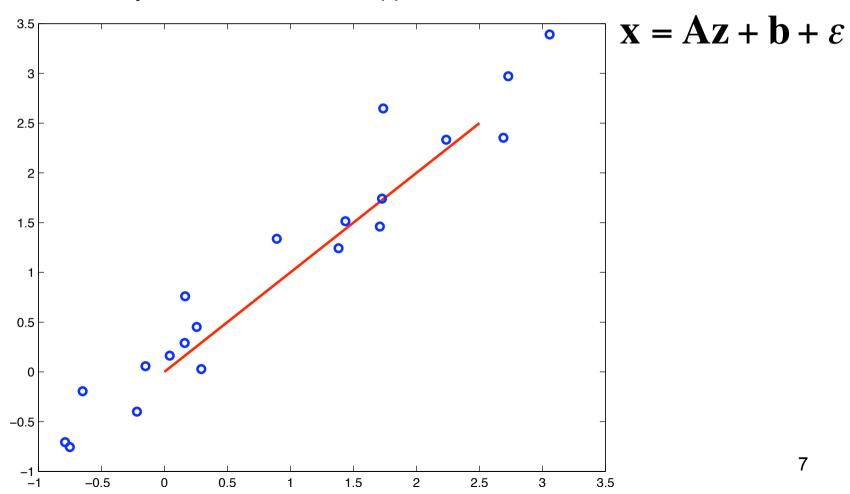
Data is 2D but it only has 1 degree of freedom and lies along a line (linear subspace)





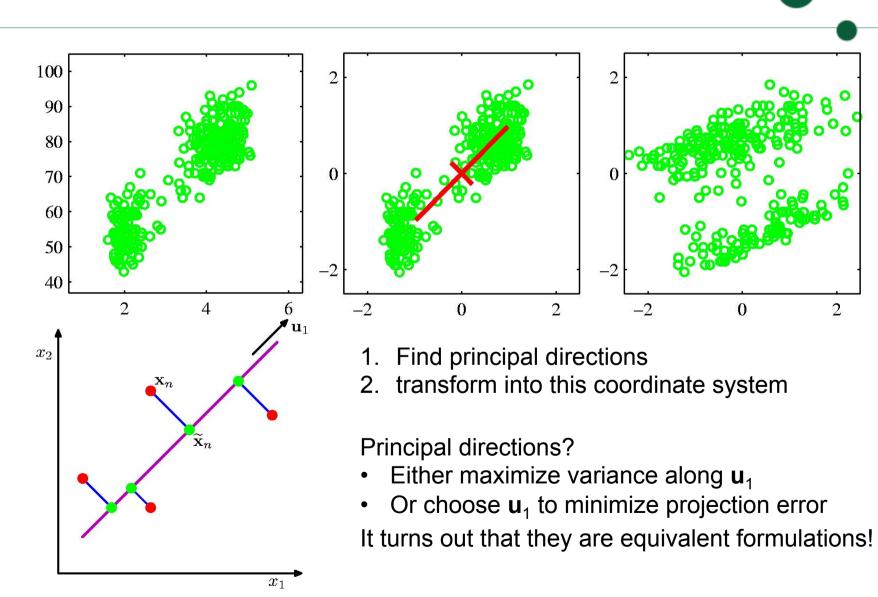


A bit more messy but data can still be approximated with a linear model:



Principal Component Analysis (PCA): The typical steps









Project data $\{x_n\}_{n=1,...,N}$ onto directions $\{u_i\}_{i=1,...,M}$ with $M \ll D$.

We find directions sequentially, u_1 first.

Mean of projected data: $\mathbf{u}_1^T \bar{\mathbf{x}}$ with

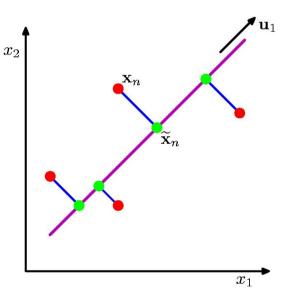
$$\bar{\mathbf{x}} = \frac{1}{N} \sum_{n=1}^{N} \mathbf{x}_n .$$

Variance of projected data:

$$\frac{1}{N} \sum_{n=1}^{N} \left\{ \mathbf{u}_{1}^{T} \left(\mathbf{x}_{n} - \bar{\mathbf{x}} \right) \right\}^{2} = \mathbf{u}_{1}^{T} \mathbf{S} \mathbf{u}_{1} .$$

with the empirical co-variance

$$S = \frac{1}{N} \sum_{n=1}^{N} (x_n - \bar{x}) (x_n - \bar{x})^T$$



PCA: Maximum variance formulation (The first)

Maximize variance

$$\mathbf{u}_{1}^{T}\mathbf{S}\mathbf{u}_{1}$$

with respect to u_1 .

But we need a constraint to avoid $u_1 \to \infty$:

$$\mathbf{u}_1^T \mathbf{S} \mathbf{u}_1 + \lambda_1 (1 - \mathbf{u}_1^T \mathbf{u}_1)$$

Student exercise – take derivative wrt \mathbf{u}_1 and show

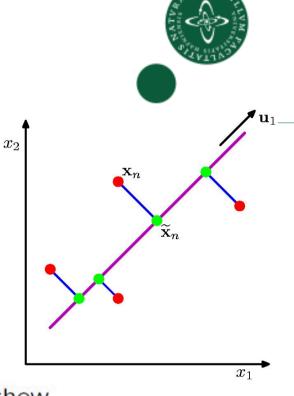
$$Su_1 = \lambda_1 u_1$$
.

Multiply by \mathbf{u}_1^T and show

$$\mathbf{u}_1^T \mathbf{S} \mathbf{u}_1 = \lambda_1$$
.

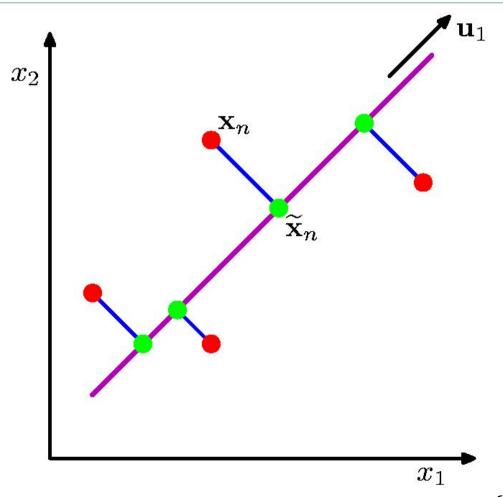
How to interpret these two results?

Recall: The eigenvectors are orthonormal and form a coordinate system



PCA: Minimum error formulation (The second)





Goal: Find the optimal reconstructing orthonormal directions $\left\{\mathbf{u}_{i}\right\}$

PCA: Minimum error formulation (The second)



• Assume complete orthonormal basis given: $\{\mathbf{u}_i\}$, i = 1,...,D

$$\mathbf{u}_{i}^{T}\mathbf{u}_{j} = \delta_{ij} \qquad \delta_{ij} = 1 \text{ if } i = j, \text{ otherwise } \delta_{ij} = 0$$

• In this basis, data \mathbf{x}_n may be represented as $(z_{n1}, z_{n2}, ..., z_{nD})$

$$z_{ni} = \mathbf{x}_n^T \mathbf{u}_i$$
 (projection on PC) And back: $\mathbf{x}_n = \sum_{i=1}^D z_{ni} \mathbf{u}_i$

 Goal: Approximate data with fewer dimensions M < D by an M dimensional linear subspace:

$$\tilde{\mathbf{x}}_{n} = \sum_{i=1}^{M} z_{ni} \mathbf{u}_{i} + \sum_{j=M+1}^{D} b_{j} \mathbf{u}_{j}$$

$$z_{ni} = \mathbf{x}_{n}^{T} \mathbf{u}_{i} , i = 1,...,M$$

$$b_{j} = \overline{\mathbf{x}}^{T} \mathbf{u}_{j} , j = M+1,...,D$$

By minimizing the sum of squares error:

$$J = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \tilde{\mathbf{x}}_n)^2 = \sum_{j=M+1}^{D} \mathbf{u}_j^T \mathbf{S} \mathbf{u}_j \quad \text{subject to constraint } \mathbf{u}_i^T \mathbf{u}_j = \delta_{ij}$$

PCA: Minimum error formulation (The second)



Minimizing the sum of squares error:

$$J = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \tilde{\mathbf{x}}_n)^2 = \sum_{i=M+1}^{D} \mathbf{u}_i^T \mathbf{S} \mathbf{u}_i \quad \text{subject to constraint } \mathbf{u}_i^T \mathbf{u}_j = \delta_{ij}$$

- It can be proven that the general solution is given by the eigenvector equation: $\mathbf{S}\mathbf{u}_i = \lambda_i \mathbf{u}_i$
- Also the corresponding error for this solution is

$$J = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \tilde{\mathbf{x}}_n)^2 = \sum_{i=M+1}^{D} \lambda_i$$

 Hence for fixed M choose the eigenvectors corresponding to the M largest eigenvalues to minimize the sum of squares reconstruction error.

PCA: Summary



- The two formulations are equivalent.
- Solve the eigenvector equation for the data covariance:

$$\mathbf{S}\mathbf{u}_i = \frac{1}{N}\mathbf{X}^T\mathbf{X}\mathbf{u}_i = \lambda_i \mathbf{u}_i$$

Where the data matrix is $\mathbf{X} = (\mathbf{x}_1 - \overline{\mathbf{x}}, \dots, \mathbf{x}_N - \overline{\mathbf{x}})^T \in \mathbb{R}^{N \times D}$

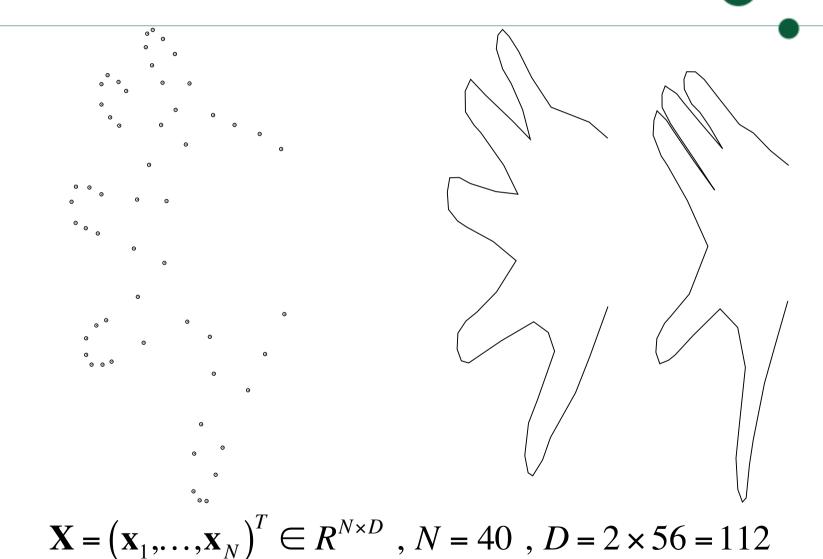
• Eigenvectors form the Principal Components (PC) and an orthogonal coordinate basis.

$$z_{ni} = \mathbf{x}_n^T \mathbf{u}_i$$
 (projection on PC), \mathbf{x}_n in PC space $(z_{n1}, ..., z_{nM})^T$

- Eigenvalues represents the projected data variance along the corresponding PC.
- Now what can this be used for?

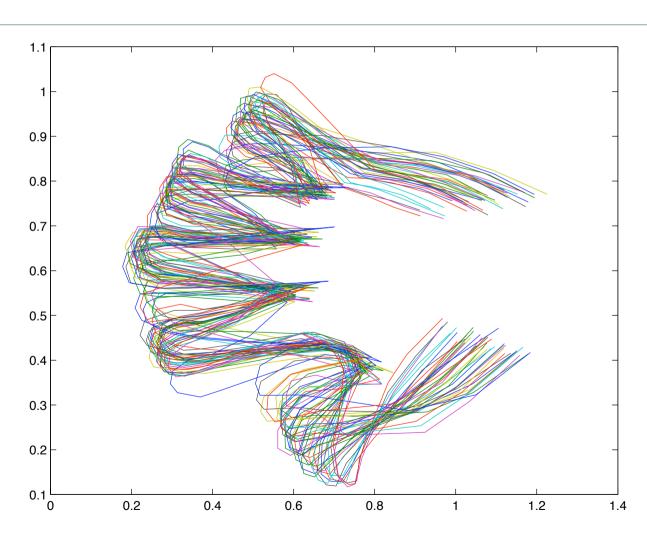












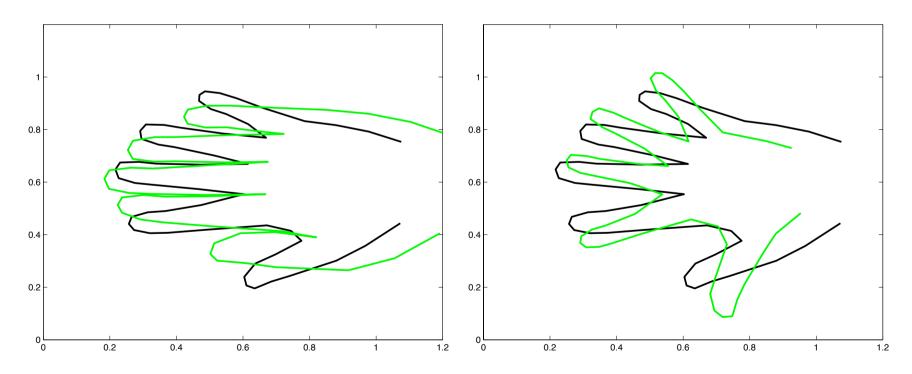


Visualizing the first principle component

Black: Mean data point (mean hand) $\overline{\mathbf{X}}$

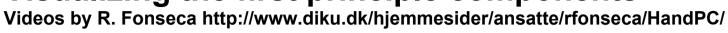
Green: Visualizing +/- 2 standard deviations along first PC $\bar{\mathbf{x}} \pm 2\sqrt{\lambda_1}\mathbf{u}_1$

Conclusion: PC 1 captures finger spread variation and size

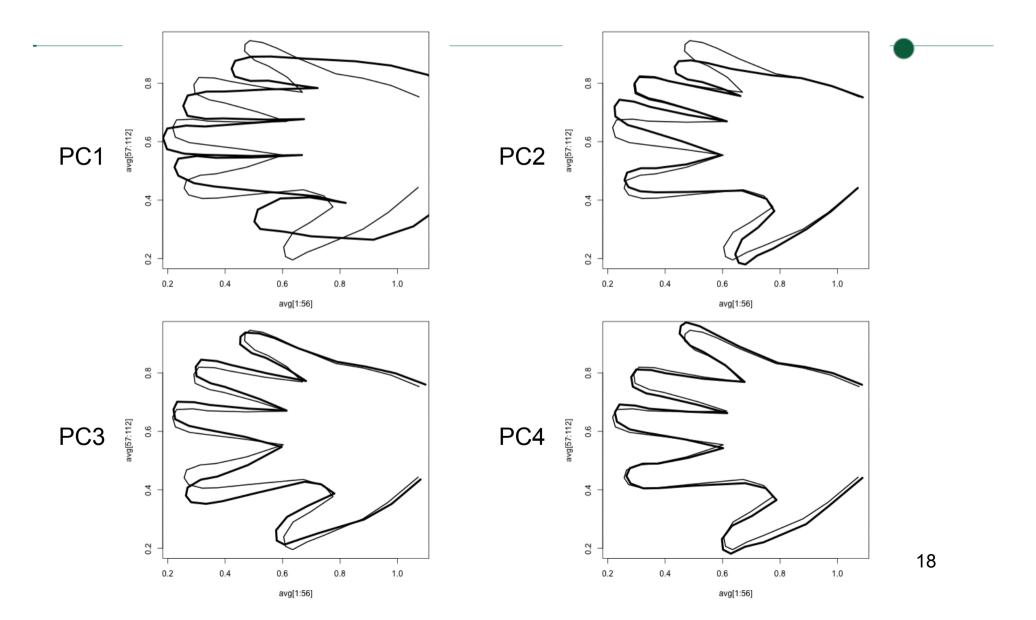


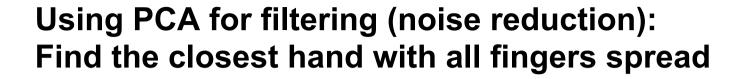
PCA can be used to visualize the largest / important variation in the data set.

Visualizing the first principle components

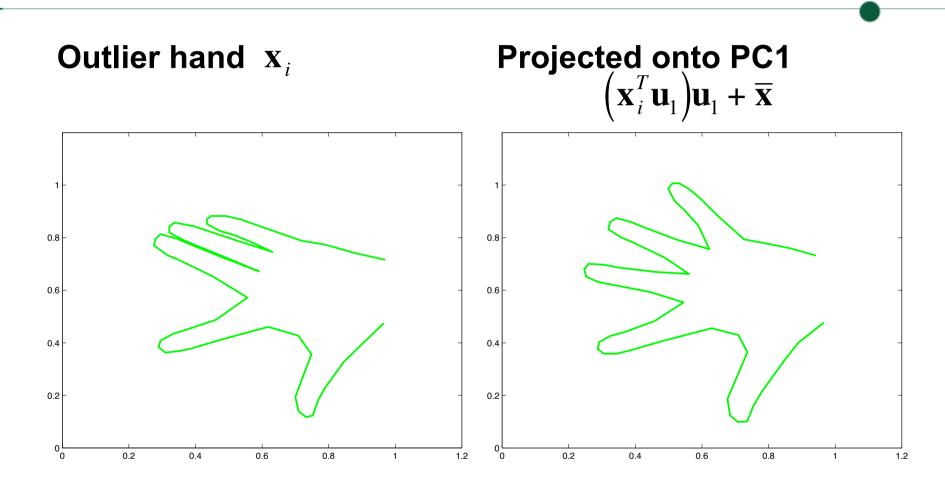












Filtering: We remove irrelevant variation by projection onto principal components. In this example we concentrate on the spread of all fingers.





- Data set (measurements from N=252 men):
 - Density determined from underwater weighing
 - Percentage body fat from Siri's (1956) equation
 - Age (years)
 - Weight (lbs)
 - Height (inches)
 - Circumferences (cm): Neck, Chest, Abdomen 2, Hip, Thigh,
 Knee, Ankle, Biceps (extended), Forearm, Wrist
- Observations x: circumferences, weight, age and height (D=15 dimensional vector)

Target values t: Percentage body fat (scalar)

Data set:
$$\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^T$$
, $\mathbf{T} = (t_1, \dots, t_N)^T$

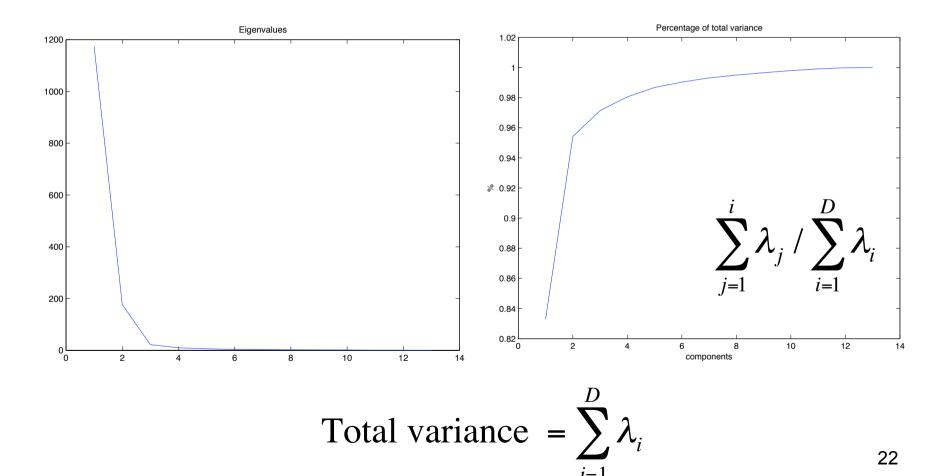




- Picking features by hand, Ex.:
 - Selection 1 : Columns 4 (weight), 7 (chest), 8 (abdomen), and 9 (hip)
 - Selection 2 : Column 8 (abdomen)
- This is based on intuition and picking variables with the largest covariance with the target variable in column 2 (percentage body fat).
- Lets see what we get out of performing PCA on this data set (dimensionality reduction).
 - Do PCA on columns 3 to 15, that is perform PCA on $\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^T \in \mathbb{R}^{N \times D}$, N = 252, D = 13



Eigenvalues and total variance

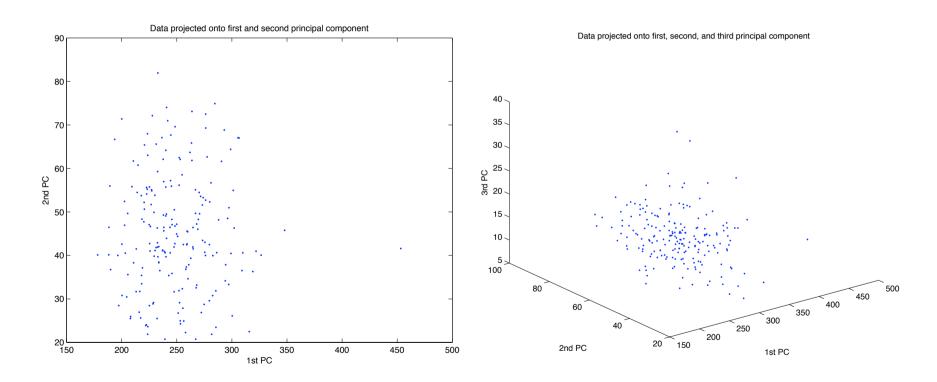




PCA for columns 3 to 15: Projections of data

Data on PC 1 and 2

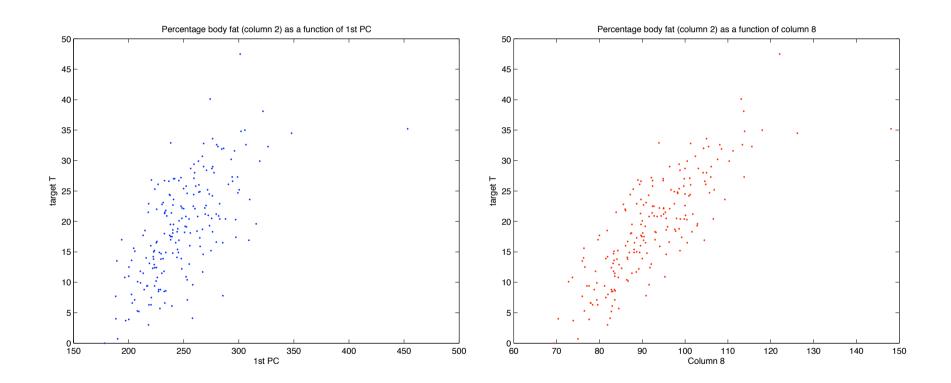
Data on PC 1, 2, and 3



What dimensionality for the latent variable should we choose? 1, 2, or 3?







Regression results revisited: Do we improve our results?

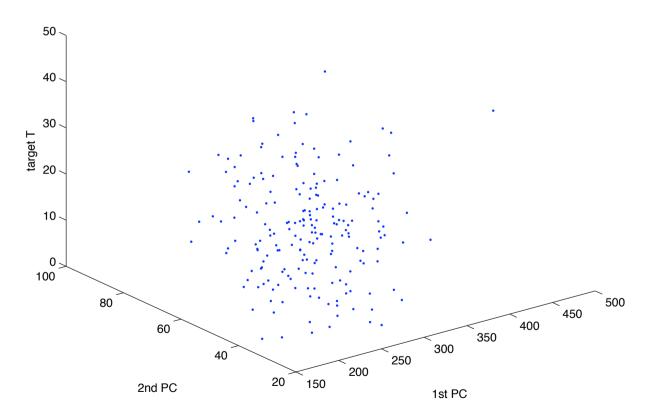


Maximum likelihood RMS error with column 8:
 RMS = 10.9 (% body fat)

- Maximum likelihood RMS error with PC 1: RMS = 9.9 (% body fat)
- (Average results over 50 trials (bootstrapping) to compensate for sensitivity to random partitioning of data into training and test set).
- And maybe we can do better by using both PC 1 and 2?

Target variable plotted against 1st PC and 2nd PC

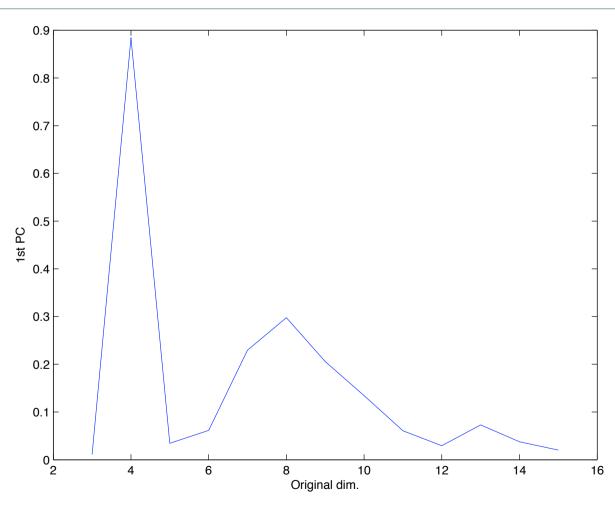
Percentage body fat (column 2) as a function of 1st and 2nd PC



Since the first 2 PCs capture more than 95 % of the total variance a good choice would be these.







PCA picks a linear combination of variables 4, 7, 8, 9, and 10 (the rest are fairly small).





- Projection of data onto PC's can visualize the major modes of variation in the dataset.
- PC's versus target values / class labels can help with our the selection of models.
- By inspecting the individual PC vector we can see which features contributes the most to this PC (feature selection).
- PCA with M < D gives a lossy reconstruction of data and can be used for noise removal.





- N<D: The problem is singular with only N eigenvalues!
- The eigenvector equation for the empirical covariance

$$\mathbf{S}\mathbf{u}_i = \frac{1}{N}\mathbf{X}^T\mathbf{X}\mathbf{u}_i = \lambda_i \mathbf{u}_i$$

- Pre-multiply by **X**: $N^{-1}XX^{T}(X\mathbf{u}_{i}) = \lambda_{i}(X\mathbf{u}_{i})$
- Introduce $\mathbf{v}_i = \mathbf{X}\mathbf{u}_i$ and substitute: $N^{-1}\mathbf{X}\mathbf{X}^T\mathbf{v}_i = \lambda_i\mathbf{v}_i$
- Find eigenvectors for this smaller N x N problem.
- Eigenvector for covariance: Pre-multiply by X^T

$$\frac{1}{N} \mathbf{X}^T \mathbf{X} \left(\mathbf{X}^T \mathbf{v}_i \right) = \mathbf{S} \left(\mathbf{X}^T \mathbf{v}_i \right) = \lambda_i \left(\mathbf{X}^T \mathbf{v}_i \right)$$

 $\frac{1}{N} \mathbf{X}^T \mathbf{X} (\mathbf{X}^T \mathbf{v}_i) = \mathbf{S} (\mathbf{X}^T \mathbf{v}_i) = \lambda_i (\mathbf{X}^T \mathbf{v}_i)$ and back to eigenvectors of **S**: $\mathbf{u}_i = \frac{1}{(N\lambda_i)^{1/2}} \mathbf{X}^T \mathbf{v}_i$

Singular Value Decomposition (SVD) : A computational trick



- SVD a simple way of doing PCA: $[\mathbf{U}, \mathbf{K}, \mathbf{V}] = \text{SVD}(\mathbf{X})$ $\mathbf{X} = \mathbf{U}\mathbf{K}\mathbf{V}^T \in R^{N \times D}$
- U and V are N x N and D x D orthogonal matrices:

$$\mathbf{U}^T \mathbf{U} = \mathbf{U} \mathbf{U}^T = \mathbf{I}_N \qquad \qquad \mathbf{V}^T \mathbf{V} = \mathbf{V} \mathbf{V}^T = \mathbf{I}_D$$

- **K** is a *N* x *D* diagonal matrix of sorted singular values (≥0).
- The covariance may be expressed as

$$\mathbf{S} = \frac{1}{N} \mathbf{X}^T \mathbf{X} = \frac{1}{N} \mathbf{V} \mathbf{K}^T \mathbf{U}^T \mathbf{U} \mathbf{K} \mathbf{V}^T = \frac{1}{N} \mathbf{V} \mathbf{K}^T \mathbf{K} \mathbf{V}^T$$

Columns of V are eigenvectors of S (the PCs):

$$\mathbf{S}\mathbf{v}_{i} = \frac{\mathbf{K}_{ii}^{2}}{N}\mathbf{v}_{i} \qquad \qquad \lambda_{i} = \frac{\mathbf{K}_{ii}^{2}}{N}$$

• Projection of data onto *i*'th PC: $\mathbf{X}\mathbf{v}_i = \mathbf{U}\mathbf{K}\mathbf{V}^T\mathbf{v}_i = \mathbf{u}_i\mathbf{K}_{ii}$



PCA in preprocessing: Standardization

 It is common to preprocess data by normalizing the individual variables to have zero mean and unit variance:

$$\tilde{x}_{ni} = \frac{\left(x_{ni} - \bar{x}_i\right)}{\sigma_i}$$

Covariance matrix becomes the correlation matrix

$$\rho_{ij} = \tilde{\mathbf{S}}_{ij} = \frac{1}{N} \sum_{n=1}^{N} \frac{\left(x_{ni} - \bar{x}_{i}\right)}{\sigma_{i}} \frac{\left(x_{nj} - \bar{x}_{j}\right)}{\sigma_{j}}$$

But PCA does more than this!
 We can decorrelate variables as we will see in a minute.

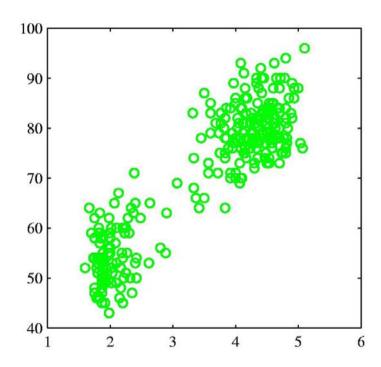




Hydrothermal geyser in Yellowstone National Park,

Wyoming, USA.

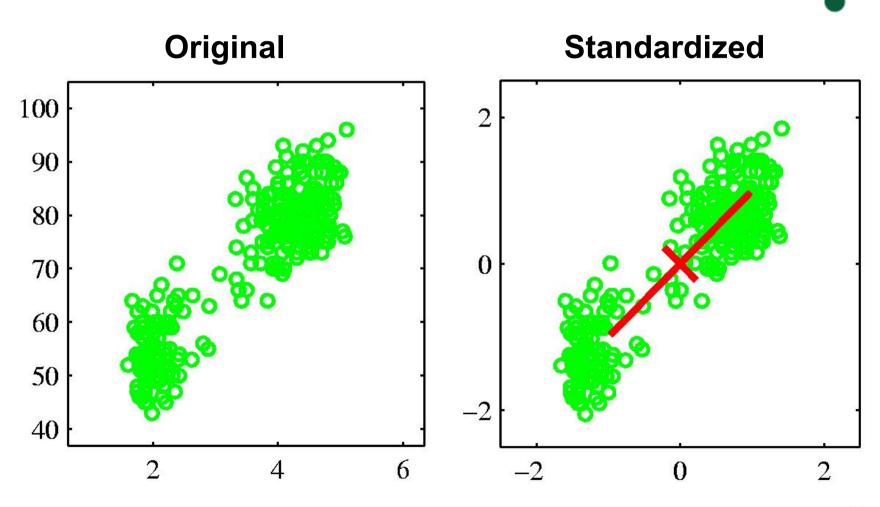




- x-axis duration of eruption in minutes
- y-axis time to next eruption in minutes
- Notice the big difference in magnitude of the two axes.



Example of standardization





PCA in preprocessing: Whitening

• Write the eigenvector equation as SU = UL, where

$$\mathbf{U} = (\mathbf{u}_1, \dots, \mathbf{u}_D) \quad \text{and} \quad \mathbf{L} = \begin{pmatrix} \lambda_1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \lambda_D \end{pmatrix}$$

- Translate, rotate, and scale the data into the coordinate system of the PCs: $\mathbf{y}_n = \mathbf{L}^{-1/2}\mathbf{U}^T(\mathbf{x}_n \overline{\mathbf{x}})$
- In this coordinate system the data is zero mean and have identity covariance

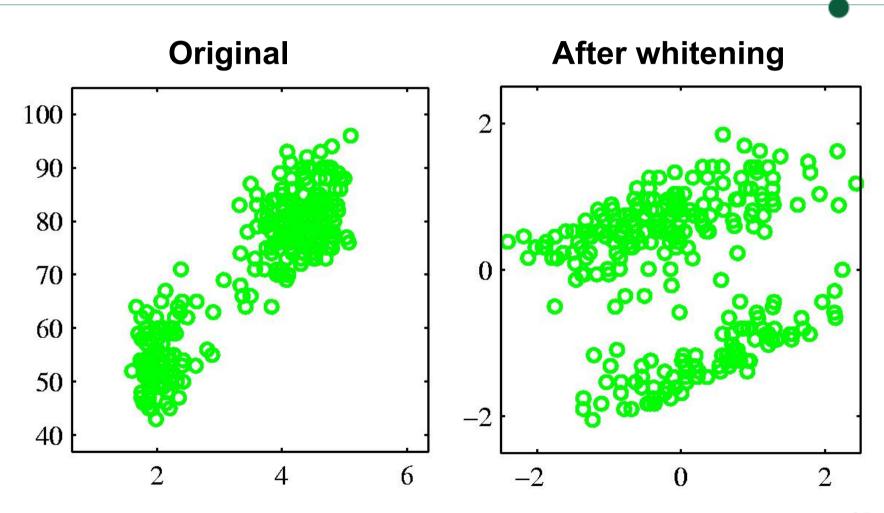
$$\frac{1}{N} \sum_{n=1}^{N} \mathbf{y}_{n} \mathbf{y}_{n}^{T} = \frac{1}{N} \sum_{n=1}^{N} \mathbf{L}^{-1/2} \mathbf{U}^{T} (\mathbf{x}_{n} - \overline{\mathbf{x}}) (\mathbf{x}_{n} - \overline{\mathbf{x}})^{T} \mathbf{U} \mathbf{L}^{-1/2}$$

$$= \mathbf{L}^{-1/2} \mathbf{U}^{T} \mathbf{S} \mathbf{U} \mathbf{L}^{-1/2} = \mathbf{L}^{-1/2} \mathbf{U}^{T} \mathbf{U} \mathbf{L} \mathbf{L}^{-1/2} = \mathbf{L}^{-1/2} \mathbf{L} \mathbf{L}^{-1/2} = \mathbf{I}_{D}$$

This is referred to as whitening the data.



Example of whitening







- Continuous latent variable models
- Linear PCA: Find eigenvectors of data covariance
 - Maximum variance formulation
 - Minimum error formulation
- Applications
 - Preprocessing of data (whitening and filtering)
 - Dimensionality reduction
 - Visualization of high dimensional data

Literature



- Continuous latent variable models: CB page 559 561
- and Principal Component Analysis (PCA):
 CB page 561 570