



Linear Models For Regression, Part 2

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Plan for this lecture

- Presentation of the next assignment
- A snippet of theoretical foundation of regression
- Recap of linear models for regression
- Bayesian regression for linear models
- Bayesian sequential learning for regression
- Advanced topic: Full Bayesian approach by computing the predictive distribution by integration over all models.



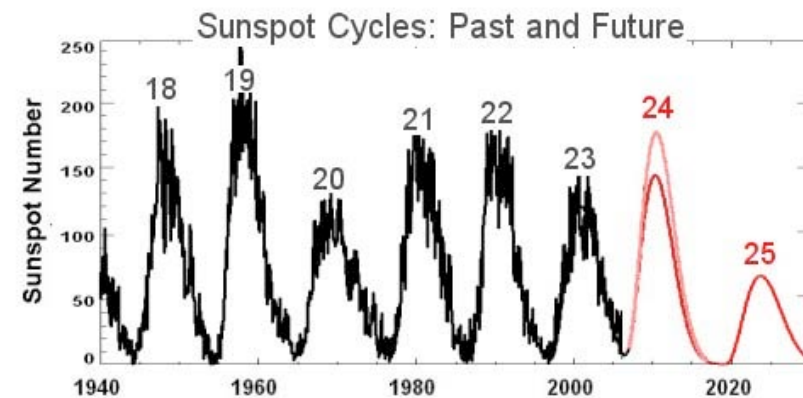
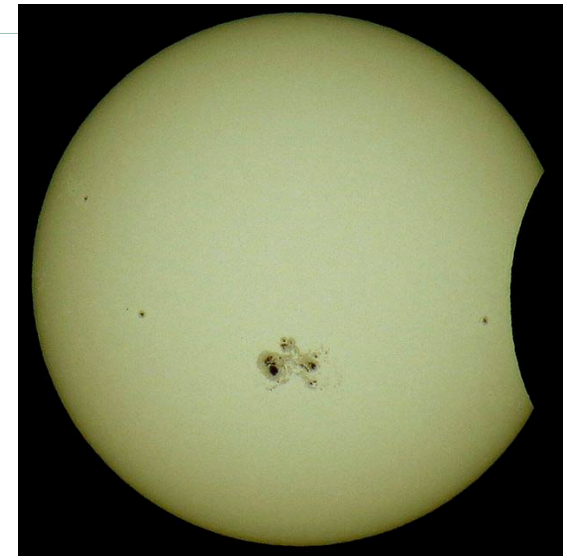
Assignment 2: Basic learning algorithms

- Regression (this lecture):
 - Linear models for regression applied to a real data set in order to predict the number of sunspots from previous years sunspot numbers.
 - A theoretical question about weighted sum of squares



Example: Sunspots (Assignment 2)

- **Input variable:**
 - Number of sunspot in previous years
- **Target variable:**
 - Number of sunspots in following years
- **Your task:**
 - Learn a linear regression model
$$t = \mathbf{y}(\mathbf{x})$$
for predicting sunspot numbers
 - How to do this?
 - We learn today and Tuesday



<http://en.wikipedia.org/wiki/Sunspot>



Assignment 2: Basic learning algorithms

- Regression (this lecture):
 - Linear models for regression applied to a real data set in order to predict the number of sunspots from previous years sunspot numbers.
 - A theoretical question about weighted sum of squares
- Classification (next lecture):
 - Experiment with Linear Discriminant Analysis for classification on the Iris data set.
 - Theoretical question about the Bayes optimal classifier



Recall from last lecture

- The least squares solution is equivalent to maximum likelihood (ML) solution under Gaussian noise model. Both have tendency to overfit the data for $M \geq N$ (poor generalization).

$$\operatorname{argmin}_{\mathbf{w}} \tilde{E}(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N [y(x_n, \mathbf{w}) - t_n]^2 \Leftrightarrow \operatorname{argmax}_{\mathbf{w}} p(\mathbf{T} | \mathbf{X}, \mathbf{w})$$

- Regularized least squares equivalent to maximum a posteriori (MAP) under Gaussian noise model and isotropic Gaussian prior on model parameters. Both behaves well for $M \geq N$ (or at least better than ML solution).

$$\operatorname{argmin}_{\mathbf{w}} \tilde{E}(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N [y(x_n, \mathbf{w}) - t_n]^2 + \frac{\lambda}{2} \|\mathbf{w}\|^2 \Leftrightarrow \operatorname{argmax}_{\mathbf{w}} p(\mathbf{w} | \mathbf{T}, \mathbf{X})$$



Goals and approaches for regression

- Goal: Make prediction of unseen data.
- Approaches to making predictions in regression:
 - $y(x)$ (Regression function)
 - $p(t | x)$ (Predictive distribution)
 - $p(t, x)$ (Joint distribution)
- Today we generalize the regression model and develop the full Bayesian approach to regression.



Decision theoretic interpretation of regression: Minimizing the risk (Recap)

- Assume we have 100 independent data sets
 $S = \{(t_1, \mathbf{x}_1), \dots, (t_N, \mathbf{x}_N)\}_{i=1, \dots, 100}$
- We want to choose a model $y(\mathbf{x})$ that performs well on all these data sets.
- Choose the model that on average (over data sets) gives optimal performance.
- Optimal model? Optimality is defined through the loss function.
- Formally: Minimize the average *loss* $L(t, y(\mathbf{x}))$ (a.k.a. the empirical risk) we incur by modeling data t with the model $y(\mathbf{x})$

$$R_S(y(\mathbf{x})) = \frac{1}{N} \sum_{n=1}^N L(t_n, y(\mathbf{x}_n))$$

- Or minimize the (theoretical) *risk*

$$R_p(y(\mathbf{x})) = E[L] = \iint L(t, y(\mathbf{x})) p(t, \mathbf{x}) d\mathbf{x} dt$$



Decision theoretic interpretation of regression: Minimizing the risk (Recap)

- Common regression loss function: $L(t, y(\mathbf{x})) = (y(\mathbf{x}) - t)^2$

$$R_p(y(\mathbf{x})) = E[L] = \iint (y(\mathbf{x}) - t)^2 p(t, \mathbf{x}) d\mathbf{x} dt$$

- Minimization using calculus of variation (see Appendix D) leads to:

$$y(\mathbf{x}) = \int t p(t | \mathbf{x}) dt = E_t[t | \mathbf{x}]$$

- That is, the optimal solution under squared loss is given by the conditional mean of t given \mathbf{x} with respect to the predictive distribution.
- Or said in another way: The solution is given by the mean of the predictive distribution

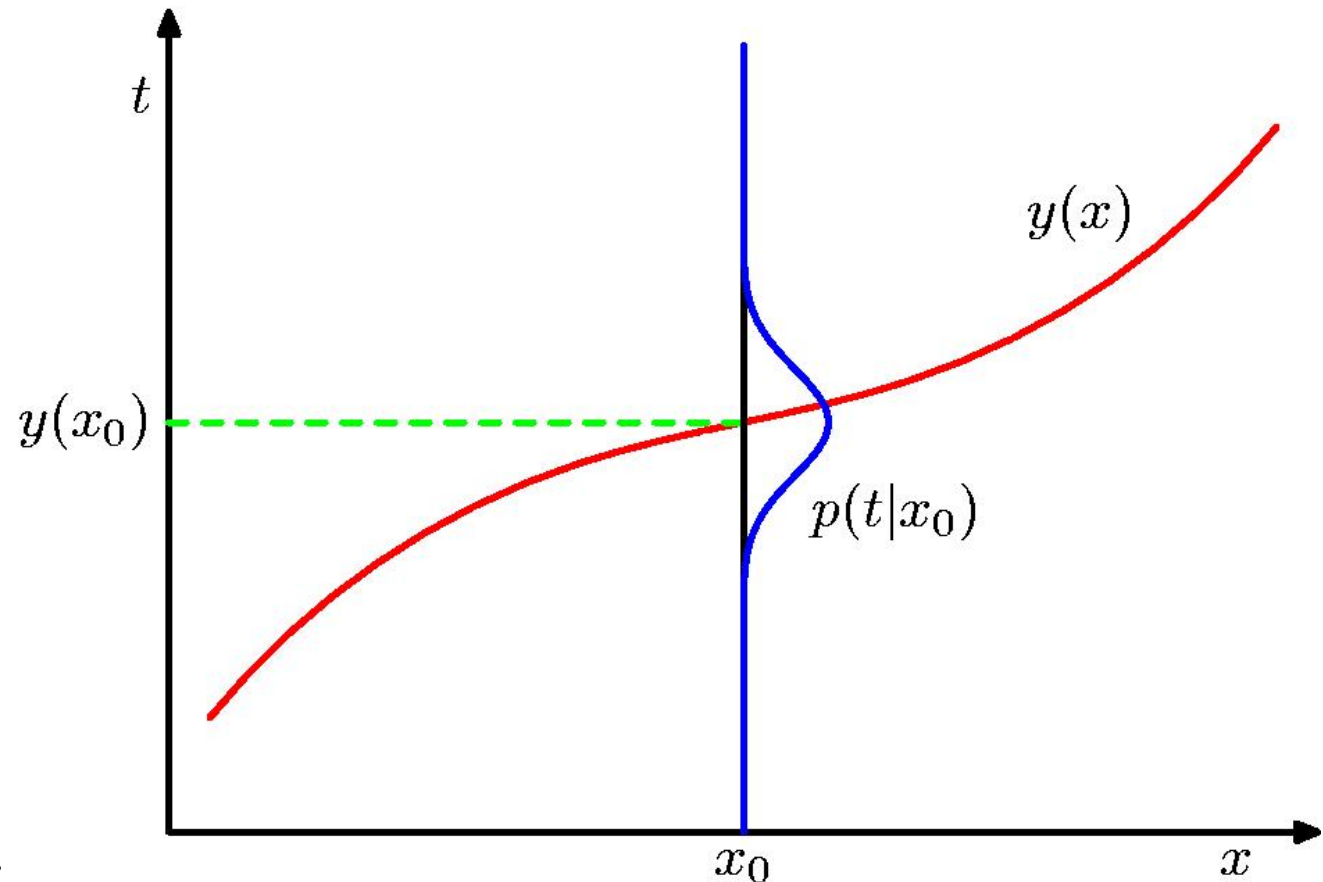


Optimal regression function under quadratic loss

The noise model:

$$t(\mathbf{x}) = y(\mathbf{x}, \mathbf{w}) + \varepsilon$$

$$\varepsilon \sim \mathcal{N}(t | 0, \beta^{-1})$$



Noise model leads to:

$$p(t | \mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t | y(\mathbf{x}, \mathbf{w}), \beta^{-1}) \Rightarrow E_t[t | \mathbf{x}] = \int t p(t | \mathbf{x}, \mathbf{w}, \beta) dt = y(\mathbf{x}, \mathbf{w})$$



Linear basis function models (Recap)

- Training data set: $X = \{x_1, \dots, x_N\}$

$$T = \{t_1, \dots, t_N\}$$

- The $(M-1)$ 'th order polynomial model is linear in the M model parameters:

$$y(x, \mathbf{w}) = w_0 + w_1x + w_2x^2 + \dots + w_{M-1}x^{M-1} = w_0 + \sum_{j=1}^{M-1} w_j x^j$$

- Generalize this model using (non-linear) basis functions:

$$y(x, \mathbf{w}) = w_0 + w_1\phi_1(x) + w_2\phi_2(x) + \dots + w_{M-1}\phi_{M-1}(x) = w_0 + \sum_{j=1}^{M-1} w_j \phi_j(x)$$

- In vector notation using $\phi_0(x) = 1$:

$$y(x, \mathbf{w}) = \sum_{j=0}^{M-1} w_j \phi_j(x) = \mathbf{w}^T \bar{\phi}(x)$$

$$\mathbf{w} = (w_0, \dots, w_{M-1})^T, \quad \bar{\phi}(x) = (\phi_0(x), \dots, \phi_{M-1}(x))^T$$



Examples of basis functions (Recap)

- Simple D -dim. linear model: Assume $\mathbf{x} = (x_1, \dots, x_D)^T$
Basis functions:

$$\phi_j(\mathbf{x}) = x_j \quad , \quad \bar{\phi}(\mathbf{x}) = (1, x_1, \dots, x_D)^T$$

Regression model:

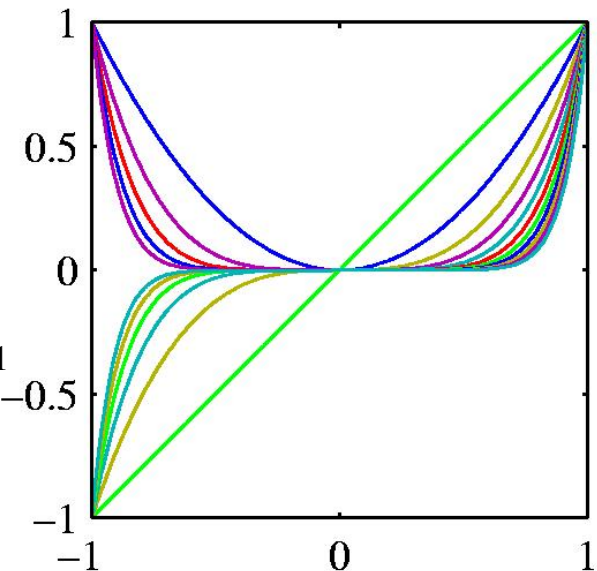
$$y(\mathbf{x}, \mathbf{w}) = \mathbf{w}^T \bar{\phi}(\mathbf{x}) = w_0 + w_1 x_1 + w_2 x_2 + \dots + w_D x_D$$

- Polynomial model (monomial basis):
Basis functions:

$$\phi_j(x) = x^j \quad , \quad \bar{\phi}(x) = (1, x, x^2, \dots, x^{M-1})^T$$

Regression model:

$$y(x, \mathbf{w}) = \mathbf{w}^T \bar{\phi}(x) = w_0 + w_1 x + w_2 x^2 + \dots + w_{M-1} x^{M-1}$$





Examples of basis functions

- Gaussian basis function:

Basis functions:

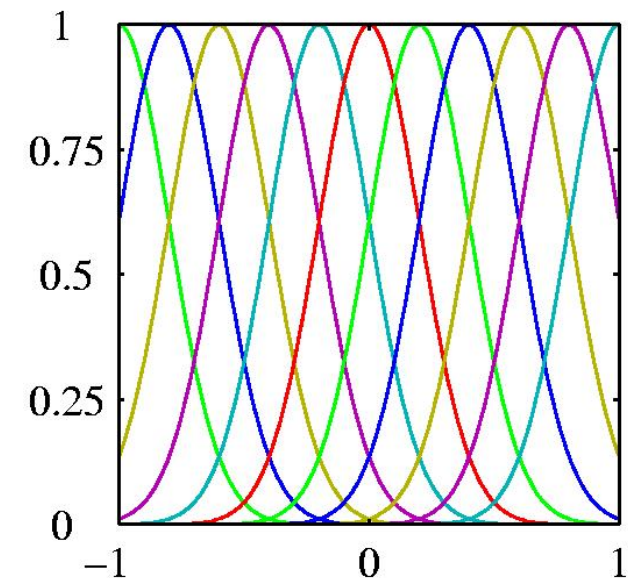
$$\phi_j(x) = \exp\left[-\frac{(x - \mu_j)^2}{2s^2}\right]$$

Regression model:

$$y(x, \mathbf{w}) = w_0 + w_1 \exp\left[-\frac{(x - \mu_1)^2}{2s^2}\right] + \dots + w_{M-1} \exp\left[-\frac{(x - \mu_{M-1})^2}{2s^2}\right]$$

μ_j position of basis function and s scale

- Other basis functions:
 - Sigmoid
 - Fourier
 - Wavelets
 - Splines (piecewise polynomial), ...





Likelihood under general linear model

Observations (i.i.d.):

$$\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$$

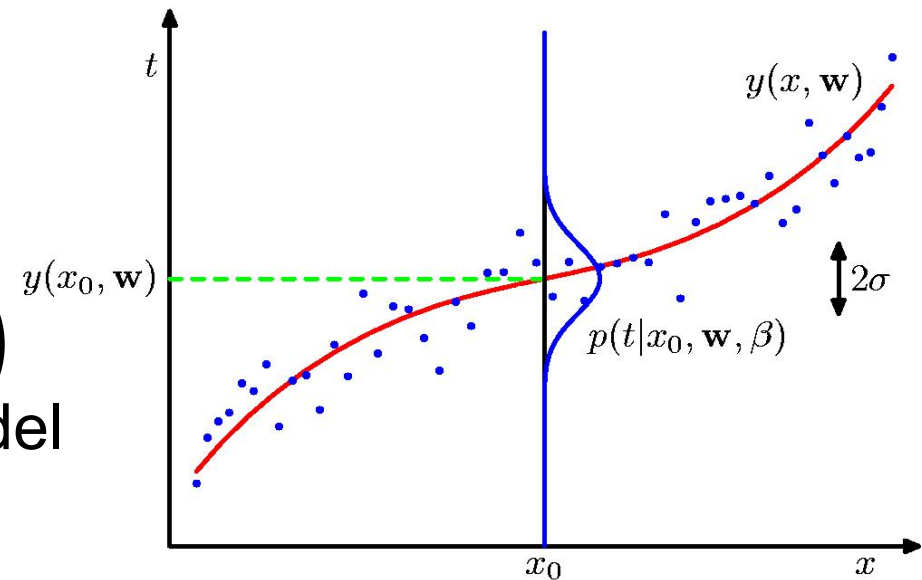
$$\mathbf{T} = (t_1, \dots, t_N)^T$$

$$p(t | \mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t | y(\mathbf{x}, \mathbf{w}), \beta^{-1})$$

Under the Gaussian noise model
we have the likelihood:

$$p(\mathbf{T} | \mathbf{X}, \mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n | \mathbf{w}^T \bar{\phi}(\mathbf{x}_n), \beta^{-1})$$

$$= \left(\frac{\beta}{2\pi} \right)^{N/2} \exp \left[-\frac{\beta}{2} \sum_{n=1}^N (t_n - \mathbf{w}^T \bar{\phi}(\mathbf{x}_n))^2 \right]$$





Maximum Likelihood (ML) solution for the general linear model

Maximize the log-likelihood with respect to \mathbf{w} :

$$\frac{\partial}{\partial w_j} \log p(\mathbf{T} | \mathbf{X}, \mathbf{w}, \beta) = \frac{\partial}{\partial w_j} \left[-\frac{\beta}{2} \sum_{n=1}^N \left(t_n - \mathbf{w}^T \bar{\phi}(\mathbf{x}_n) \right)^2 \right] = 0 \text{ for all } j$$

\Downarrow

$$\mathbf{w} = \left(\left[\sum_{n=1}^N \bar{\phi}(\mathbf{x}_n) \bar{\phi}^T(\mathbf{x}_n) \right]^{-1} \right)^T \left(\sum_{n=1}^N t_n \bar{\phi}^T(\mathbf{x}_n) \right)^T$$

Voila, we get the ML solution – but what an ugly expression!



The design matrix

- Introduce design matrix notation: $\Phi_{nj} = \phi_j(\mathbf{x}_n)$

$$\Phi = \begin{bmatrix} \phi_0(\mathbf{x}_1) & \phi_1(\mathbf{x}_1) & \cdots & \phi_{M-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \phi_1(\mathbf{x}_2) & \cdots & \phi_{M-1}(\mathbf{x}_2) \\ \vdots & \vdots & \ddots & \vdots \\ \phi_0(\mathbf{x}_N) & \phi_1(\mathbf{x}_N) & \cdots & \phi_{M-1}(\mathbf{x}_N) \end{bmatrix} \in \mathbb{R}^{N \times M}$$

- Each row contains the outcome of evaluating the M basis functions in the n 'th data point.



Examples of design matrices

- 1-dim. (M-1)'th order polynomial model:

$$y(x, \mathbf{w}) = w_0 + w_1 x + \cdots + w_{M-1} x^{M-1}$$

$$\Phi = \begin{bmatrix} 1 & x_1 & x_1^2 & \cdots & x_1^{M-1} \\ 1 & x_2 & x_2^2 & \cdots & x_2^{M-1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_N & x_N^2 & \cdots & x_N^{M-1} \end{bmatrix}$$

- 1-dim. linear model: $y(x, \mathbf{w}) = w_0 + w_1 x$

$$\Phi = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_N \end{bmatrix}$$



The design matrix specifies the likelihood

- We can rewrite the likelihood using the design matrix:

$$p(\mathbf{T} | \mathbf{X}, \mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n | \mathbf{w}^T \bar{\phi}(\mathbf{x}_n), \beta^{-1}) = \mathcal{N}(\mathbf{T} | \Phi \mathbf{w}, \beta^{-1} \mathbf{I})$$

- The design matrix and the target vector \mathbf{T} represents the training data in the likelihood.
- With unspecified \mathbf{T} the likelihood is a N-dim. multivariate Gaussian with isotropic covariance.



ML solution for general linear model

- Maximize with respect to parameters \mathbf{w} :

$$\mathbf{w}_{\text{ML}} = \left(\left[\sum_{n=1}^N \bar{\phi}(\mathbf{x}_n) \bar{\phi}^T(\mathbf{x}_n) \right]^{-1} \right)^T \left(\sum_{n=1}^N t_n \bar{\phi}^T(\mathbf{x}_n) \right)^T = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{T}$$

- Maximize with respect to precision β :

$$\frac{\partial}{\partial \beta} \log p(\mathbf{T} | \mathbf{X}, \mathbf{w}_{\text{ML}}, \beta) = \frac{N}{2} \frac{1}{\beta} - \frac{1}{2} \sum_{n=1}^N \left(t_n - \mathbf{w}_{\text{ML}}^T \bar{\phi}(\mathbf{x}_n) \right)^2 = 0 \Rightarrow$$

$$\frac{1}{\beta_{\text{ML}}} = \frac{1}{N} \sum_{n=1}^N \left(t_n - \mathbf{w}_{\text{ML}}^T \bar{\phi}(\mathbf{x}_n) \right)^2$$



Example: Prediction of body fat percentage

- **The problem:**
 - Measuring percentage of body fat accurately is inconvenient/ costly and requires weighing the body underwater and in air and applying the so-called Siri's equation.
 - How accurately can we predict the percentage of body fat from measurements of circumferences of selected body parts?
- **Data set (measurements from $N=252$ men):**
 - Density determined from underwater weighing
 - Percentage body fat from Siri's (1956) equation
 - Age (years)
 - Weight (lbs)
 - Height (inches)
 - Circumferences (cm): Neck, Chest, Abdomen 2, Hip, Thigh, Knee, Ankle, Biceps (extended), Forearm, Wrist



Example: Prediction of body fat percentage

- Consider a 1-dim. linear model of a subset of the percentage body fat data set by selecting column 8 (Abdomen 2) as the x variable:

t [%]	x [cm]
14.7	83.3
17.8	88.2
16.9	90.3
32.6	113.4
5.7	84.5
32.6	108.1
15.2	98.8
25.3	108.8

$$\Phi = \begin{bmatrix} 1 & 83.3 \\ 1 & 88.2 \\ 1 & 90.3 \\ 1 & 113.4 \\ 1 & 84.5 \\ 1 & 108.1 \\ 1 & 98.8 \\ 1 & 108.8 \end{bmatrix}$$

$$\mathbf{T} = \begin{bmatrix} 14.7 \\ 17.8 \\ 16.9 \\ 32.6 \\ 5.7 \\ 32.6 \\ 15.2 \\ 25.3 \end{bmatrix}$$



Example: Prediction of body fat percentage

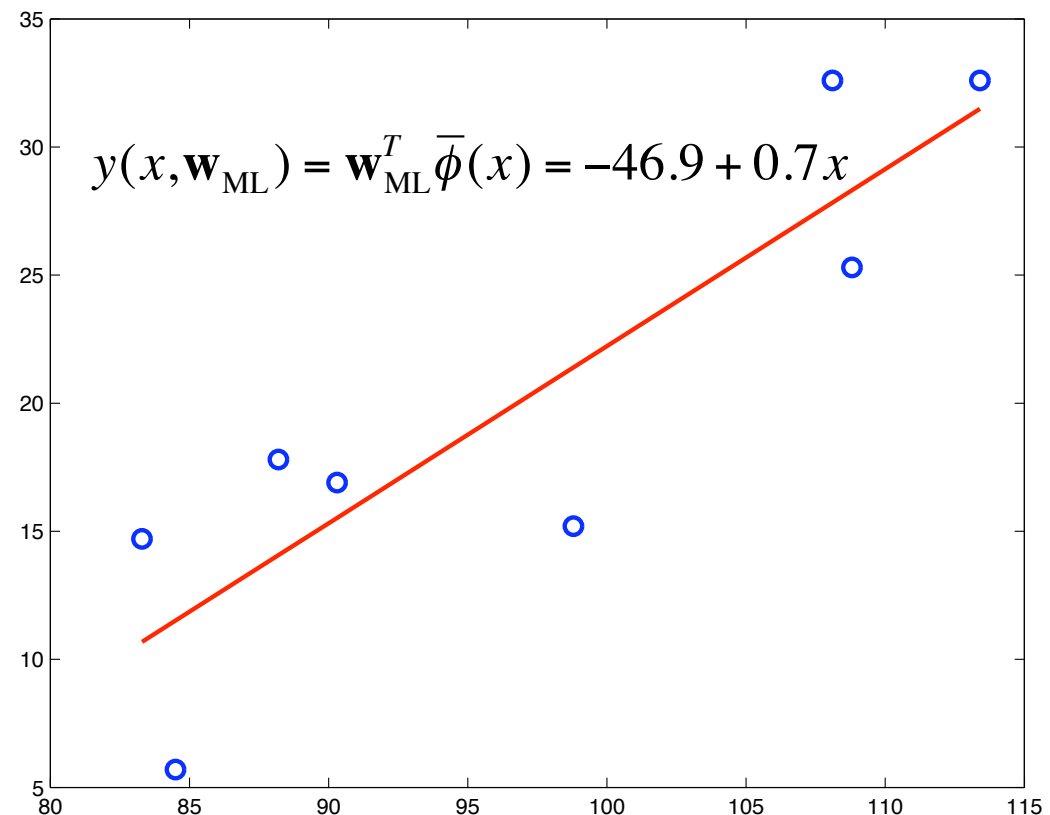
Design matrix

Target vector

$$\Phi = \begin{bmatrix} 1 & 83.3 \\ 1 & 88.2 \\ 1 & 90.3 \\ 1 & 113.4 \\ 1 & 84.5 \\ 1 & 108.1 \\ 1 & 98.8 \\ 1 & 108.8 \end{bmatrix}$$

$$\mathbf{T} = \begin{bmatrix} 14.7 \\ 17.8 \\ 16.9 \\ 32.6 \\ 5.7 \\ 32.6 \\ 15.2 \\ 25.3 \end{bmatrix}$$

$$\mathbf{w}_{\text{ML}} = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{T} = \begin{bmatrix} -46.9 \\ 0.7 \end{bmatrix}$$





Summary: Maximum likelihood (ML) regression

- Learn the model parameters \mathbf{w} on the training set:

$$X = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$$
$$\mathbf{T} = (t_1, \dots, t_N)^T$$
$$\Phi = \begin{bmatrix} \phi_0(\mathbf{x}_1) & \phi_1(\mathbf{x}_1) & \cdots & \phi_{M-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \phi_1(\mathbf{x}_2) & \cdots & \phi_{M-1}(\mathbf{x}_2) \\ \vdots & \vdots & \ddots & \vdots \\ \phi_0(\mathbf{x}_N) & \phi_1(\mathbf{x}_N) & \cdots & \phi_{M-1}(\mathbf{x}_N) \end{bmatrix}$$

$$\mathbf{w}_{\text{ML}} = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{T} \quad \beta_{\text{ML}}^{-1} = \frac{1}{N} \sum_{n=1}^N \left(t_n - \mathbf{w}_{\text{ML}}^T \bar{\phi}(\mathbf{x}_n) \right)^2$$

- Apply the model to new data \mathbf{x} :

$$y(\mathbf{x}, \mathbf{w}_{\text{ML}}) = \mathbf{w}_{\text{ML}}^T \bar{\phi}(\mathbf{x})$$

$$p(t \mid \mathbf{x}, \mathbf{w}_{\text{ML}}, \beta_{\text{ML}}^{-1}) = \mathcal{N}(t \mid y(\mathbf{x}, \mathbf{w}_{\text{ML}}), \beta_{\text{ML}}^{-1})$$



Summary: Maximum likelihood (ML) regression

- Apply the model to the test set

$$\tilde{X} = \{\tilde{\mathbf{x}}_1, \dots, \tilde{\mathbf{x}}_{\tilde{N}}\}$$

$$\tilde{\mathbf{T}} = (\tilde{t}_1, \dots, \tilde{t}_{\tilde{N}})^T$$

- and compute root-mean-square error:

$$\text{RMS} = \sqrt{\frac{1}{\tilde{N}} \sum_{n=1}^{\tilde{N}} (\tilde{t}_n - y(\tilde{\mathbf{x}}_n, \mathbf{w}_{\text{ML}}))^2}$$



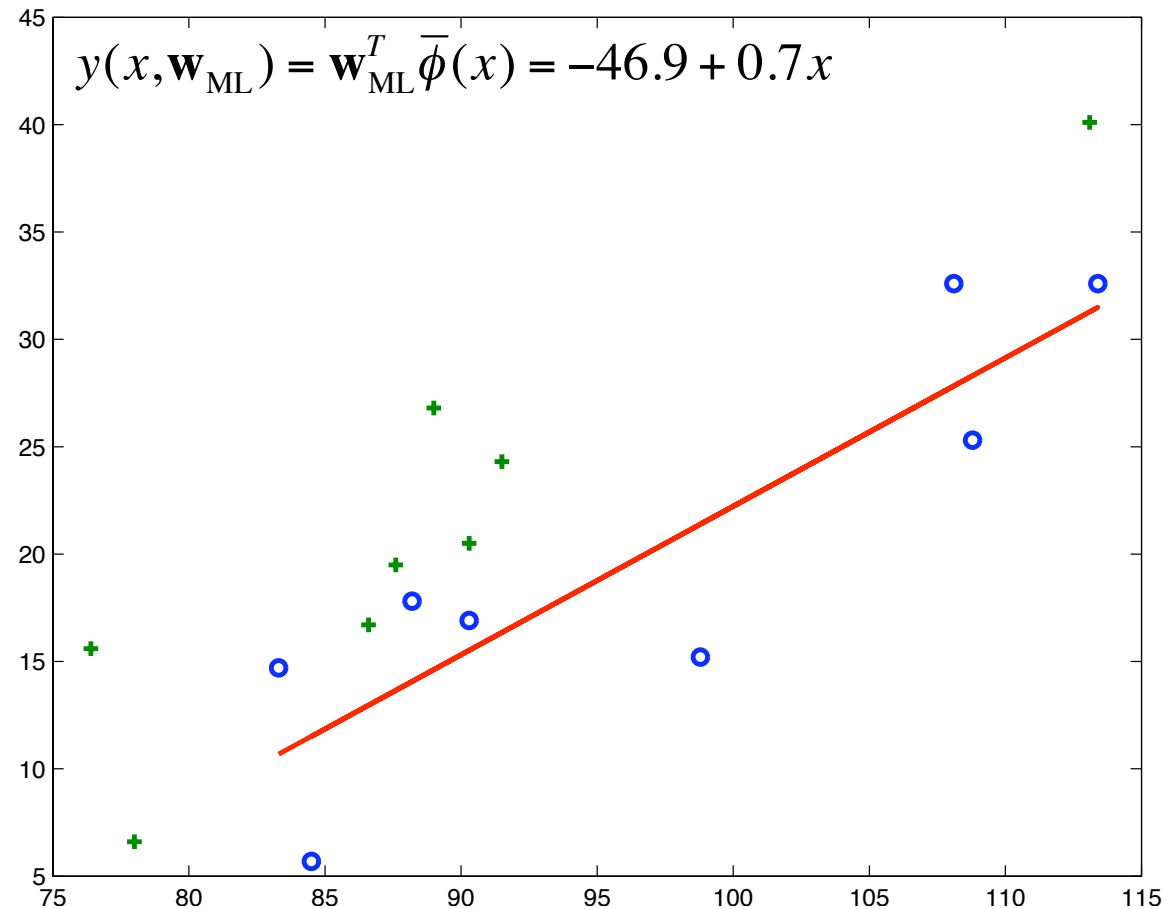
Example: Prediction of body fat percentage

Root mean square error on the training set:

$$\begin{aligned}\text{RMS} &= \sqrt{\frac{1}{N} \sum_{n=1}^N (t_n - y(\mathbf{x}_n, \mathbf{w}_{\text{ML}}))^2} \\ &= 4.1\end{aligned}$$

Root mean square error on the test set:

$$\begin{aligned}\text{RMS} &= \sqrt{\frac{1}{\tilde{N}} \sum_{n=1}^{\tilde{N}} (\tilde{t}_n - y(\tilde{\mathbf{x}}_n, \mathbf{w}_{\text{ML}}))^2} \\ &= 7.6\end{aligned}$$





Bayesian linear regression (MAP)

- The Gaussian likelihood from before (Assume known noise precision β):

$$p(\mathbf{T} | \mathbf{X}, \mathbf{w}, \beta) = \mathcal{N}(\mathbf{T} | \Phi \mathbf{w}, \beta^{-1} \mathbf{I})$$

- Add a prior to regularize the solution, thereby reducing the risk of *overfitting* to the training.
- Pick the (conjugated) Gaussian prior for the parameters

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w} | \mathbf{m}_0, \mathbf{S}_0)$$

- Posterior:

$$p(\mathbf{w} | \mathbf{T}, \mathbf{X}, \beta) = \frac{p(\mathbf{T} | \mathbf{X}, \mathbf{w}, \beta) p(\mathbf{w})}{p(\mathbf{T})}$$



Bayesian linear regression (MAP)

- The posterior is a Gaussian:

$$\begin{aligned} p(\mathbf{w} | \mathbf{T}, \mathbf{X}, \beta) &= \mathcal{N}(\mathbf{T} | \Phi \mathbf{w}, \beta^{-1} \mathbf{I}) \mathcal{N}(\mathbf{w} | \mathbf{m}_0, \mathbf{S}_0) / p(\mathbf{T}) \\ &= \mathcal{N}(\mathbf{w} | \mathbf{m}_N, \mathbf{S}_N) \end{aligned}$$

- Posterior covariance and mean (by completing the square):

$$p(\mathbf{w} | \mathbf{T}, \mathbf{X}, \beta) = \mathcal{N}(\mathbf{w} | \mathbf{m}_N, \mathbf{S}_N)$$

$$\mathbf{S}_N = \left(\mathbf{S}_0^{-1} + \beta \Phi^T \Phi \right)^{-1} \in \mathbb{R}^{M \times M}$$

$$\mathbf{m}_N = \mathbf{S}_N \left(\mathbf{S}_0^{-1} \mathbf{m}_0 + \beta \Phi^T \mathbf{T} \right) \in \mathbb{R}^M$$

- MAP solution: $\mathbf{w}_{\text{MAP}} = \mathbf{m}_N$



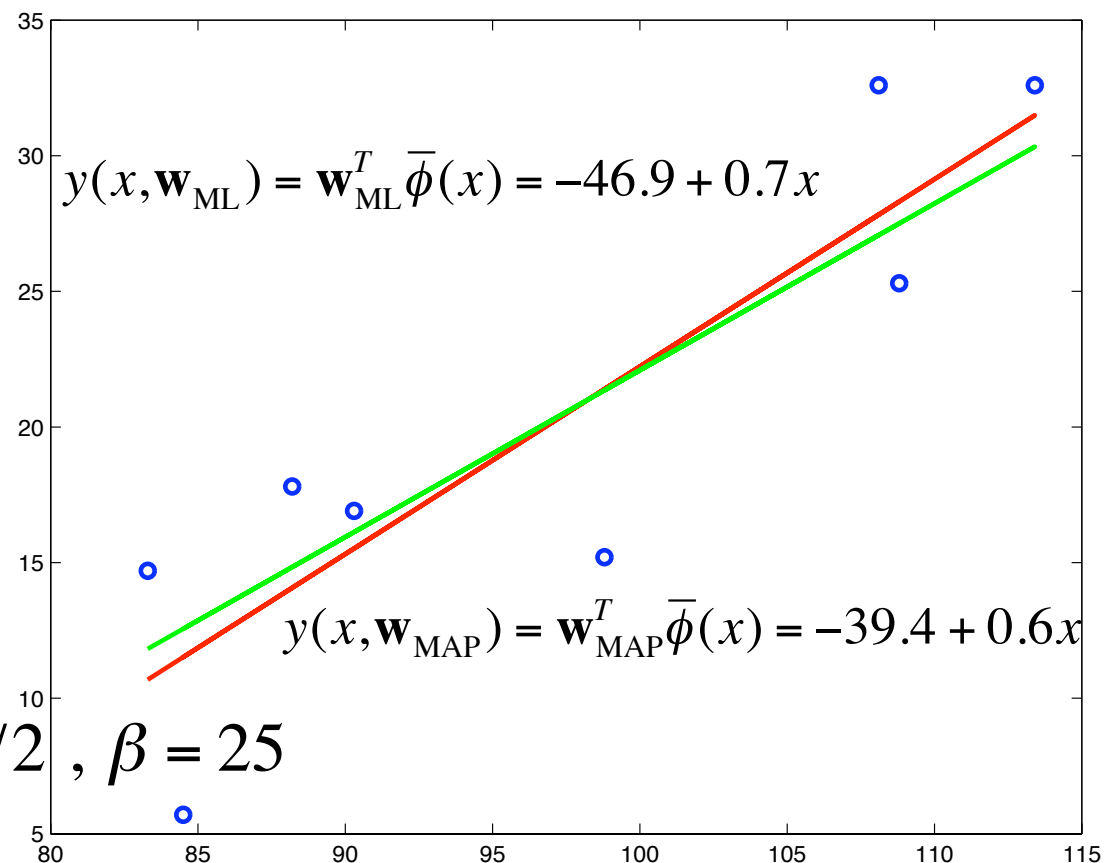
Example: Prediction of body fat percentage

Regression using $S_0 = \alpha^{-1} \mathbf{I}$, $\mathbf{m}_0 = \mathbf{0}$

$$\Phi = \begin{bmatrix} 1 & 83.3 \\ 1 & 88.2 \\ 1 & 90.3 \\ 1 & 113.4 \\ 1 & 84.5 \\ 1 & 108.1 \\ 1 & 98.8 \\ 1 & 108.8 \end{bmatrix} \quad \mathbf{T} = \begin{bmatrix} 14.7 \\ 17.8 \\ 16.9 \\ 32.6 \\ 5.7 \\ 32.6 \\ 15.2 \\ 25.3 \end{bmatrix}$$

$$\mathbf{w}_{\text{ML}} = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{T} = \begin{bmatrix} -46.9 \\ 0.7 \end{bmatrix}$$

$$\mathbf{w}_{\text{MAP}} = \mathbf{m}_N = \begin{bmatrix} -39.4 \\ 0.6 \end{bmatrix}, \quad \alpha = 1/2, \quad \beta = 25$$





Example: Prediction of body fat percentage

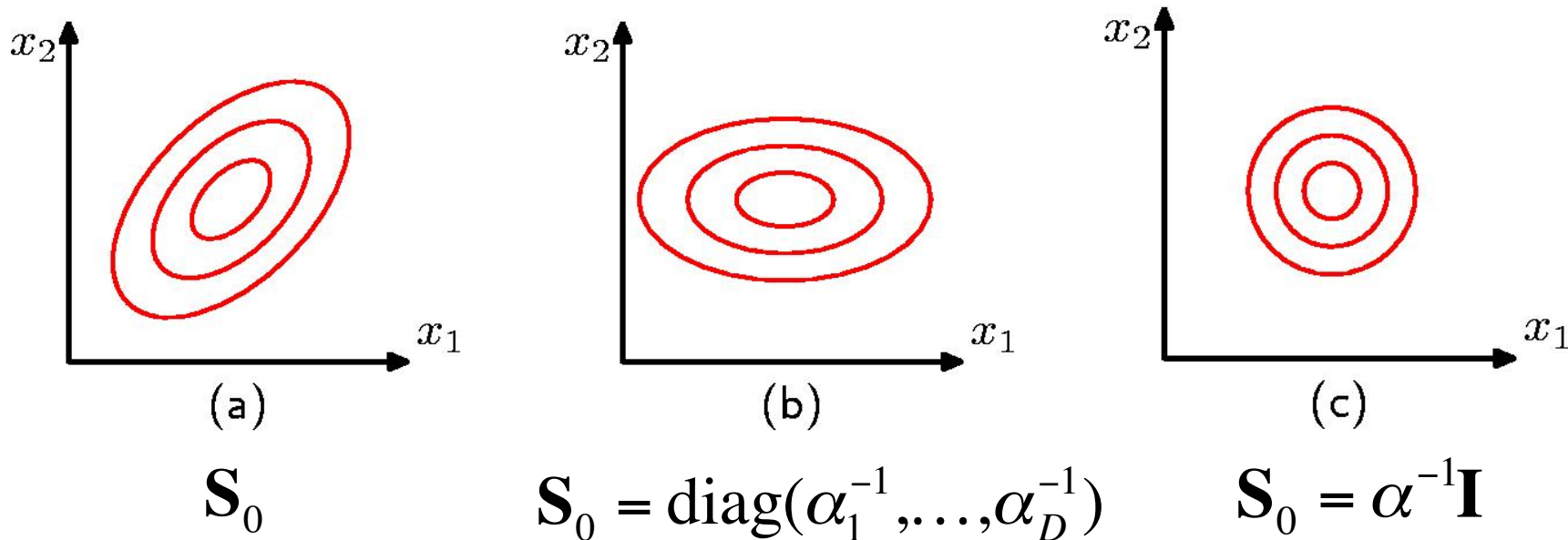
- Root mean square error on the test set:

$$\text{RMS(ML)} = \sqrt{\frac{1}{\tilde{N}} \sum_{n=1}^{\tilde{N}} (\tilde{t}_n - y(\tilde{\mathbf{x}}_n, \mathbf{w}_{\text{ML}}))^2} = 7.6$$

$$\text{RMS(MAP)} = \sqrt{\frac{1}{\tilde{N}} \sum_{n=1}^{\tilde{N}} (\tilde{t}_n - y(\tilde{\mathbf{x}}_n, \mathbf{w}_{\text{MAP}}))^2} = 7.1$$



Simplified Gaussian Prior Models



Choose an appropriate prior, but consider:

- (a) In general the covariance matrix consists of $D(D+1)/2$ free parameters.
- Reduce the amount of parameters to D in the diagonal model (b) and 1 in the isotropic model (c).



Bayesian linear regression: Effect of the prior

- In order to simplify, let's assume an isotropic prior: $\mathbf{S}_0 = \alpha^{-1}\mathbf{I}$
- Uniform prior: $\alpha^{-1} \rightarrow \infty$, $\alpha \rightarrow 0$: $\mathbf{S}_0^{-1} = \alpha\mathbf{I} \rightarrow \underline{\underline{\mathbf{0}}}$
 $\mathbf{S}_N = (\mathbf{S}_0^{-1} + \beta\Phi^T\Phi)^{-1} \rightarrow (\beta\Phi^T\Phi)^{-1}$
 $\mathbf{m}_N = \mathbf{S}_N(\mathbf{S}_0^{-1}\mathbf{m}_0 + \beta\Phi^T\mathbf{T}) \rightarrow$
 $\mathbf{S}_N\beta\Phi^T\mathbf{T} = \beta^{-1}(\Phi^T\Phi)^{-1}\beta\Phi^T\mathbf{T} = (\Phi^T\Phi)^{-1}\Phi^T\mathbf{T} = \mathbf{w}_{\text{ML}}$
- No data: $N = 0$
 $\mathbf{S}_N = \mathbf{S}_0$
 $\mathbf{m}_N = \mathbf{S}_N\mathbf{S}_0^{-1}\mathbf{m}_0 = \mathbf{S}_0\mathbf{S}_0^{-1}\mathbf{m}_0 = \mathbf{m}_0$ (posterior \rightarrow prior)

Bayesian sequential learning for regression

The online learning version



Posterior for $N-1$ acts as prior for parameter at N

$$p(\mathbf{w} \mid \mathbf{T}, \mathbf{X}, \beta) = \frac{p(\mathbf{T} \mid \mathbf{X}, \mathbf{w}, \beta) p(\mathbf{w})}{p(\mathbf{T})} \propto$$

$$\underbrace{p(t_N \mid \mathbf{x}_N, \mathbf{w}, \beta)}_{\text{likelihood for } N} p(\mathbf{w}) \underbrace{\prod_{n=1}^{N-1} p(t_n \mid \mathbf{x}_n, \mathbf{w}, \beta)}_{\text{posterior for } N-1} =$$

$$\prod_{n=1}^N \mathcal{N}(t_n \mid \mathbf{w}^T \bar{\phi}(\mathbf{x}_n), \beta^{-1}) \mathcal{N}(\mathbf{w} \mid \mathbf{m}_0, \mathbf{S}_0) =$$

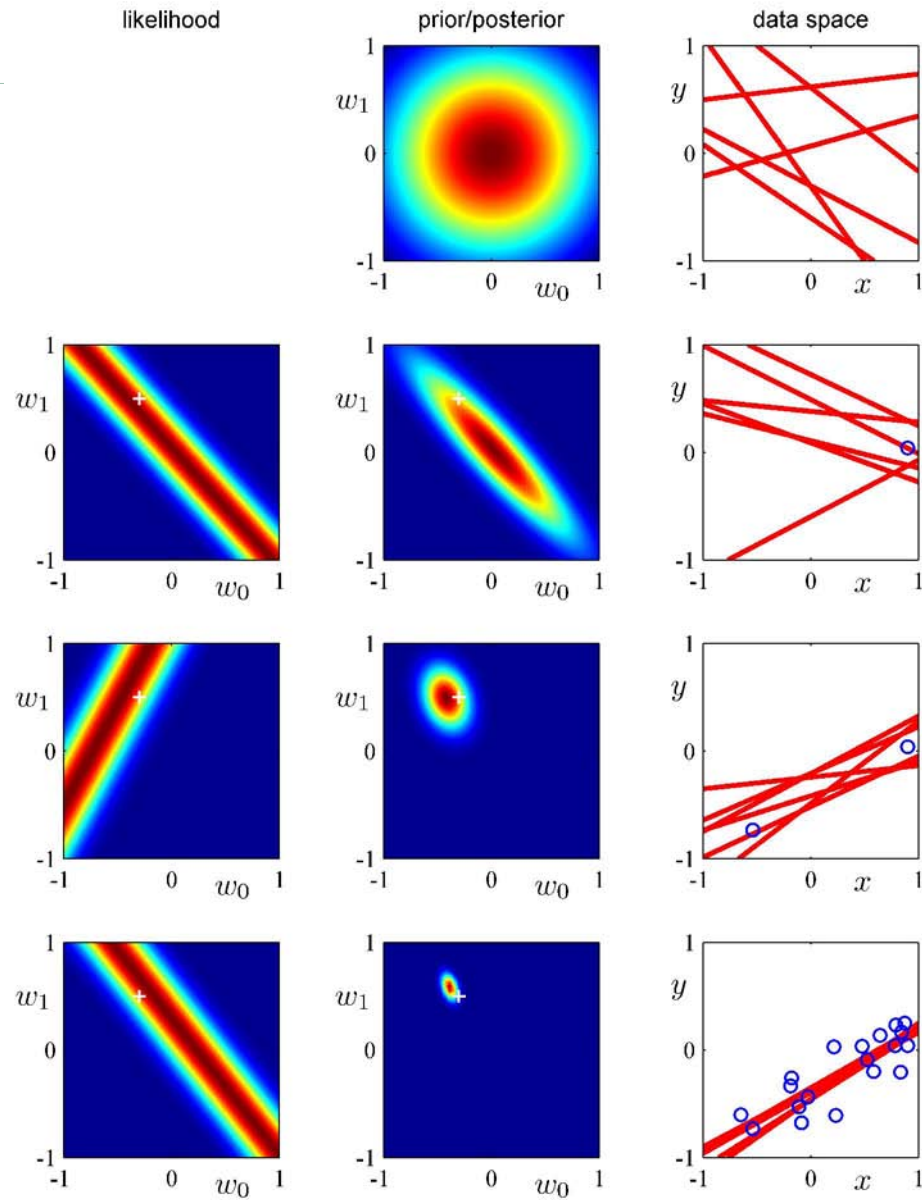
$$\mathcal{N}(t_N \mid \mathbf{w}^T \bar{\phi}(\mathbf{x}_N), \beta^{-1}) \prod_{n=1}^{N-1} \mathcal{N}(t_n \mid \mathbf{w}^T \bar{\phi}(\mathbf{x}_n), \beta^{-1}) \mathcal{N}(\mathbf{w} \mid \mathbf{m}_0, \mathbf{S}_0)$$



Example: Line regression

- Synthetic data set: $f(x, \mathbf{a}) = a_0 + a_1 x$, $a_0 = -0.3$, $a_1 = 0.5$
 $t_n = f(x_n, \mathbf{a}) + \varepsilon$, $\varepsilon \sim \mathcal{N}(\varepsilon | 0, 0.2^2)$, $x_n \sim \mathcal{U}(x | -1, 1)$
- Regression model: $y(x, \mathbf{w}) = w_0 + w_1 x$
- Lets assume the isotropic prior: $\mathbf{S}_0 = \alpha^{-1} \mathbf{I}$, $\mathbf{m}_0 = \mathbf{0}$
 $p(\mathbf{w} | \alpha) = \mathcal{N}(\mathbf{w} | \mathbf{m}_0, \mathbf{S}_0)$
- Then posterior mean and covariance becomes:
$$\mathbf{S}_N = \left(\mathbf{S}_0^{-1} + \beta \Phi^T \Phi \right)^{-1} = \left(\alpha \mathbf{I} + \beta \Phi^T \Phi \right)^{-1}$$
$$\mathbf{m}_N = \mathbf{S}_N \left(\mathbf{S}_0^{-1} \mathbf{m}_0 + \beta \Phi^T \mathbf{T} \right) = \beta \mathbf{S}_N \Phi^T \mathbf{T}$$
- Assume noise precision known $\beta = (1/0.2^2) = 25$ and set prior precision to $\alpha = 2$ and the MAP estimate of parameters is then $\mathbf{w}_{MAP} = \mathbf{m}_N$.

Example: Line regression





Full Bayesian Approach (Advanced): Model independent predictive distribution

- In general, we don't care about the specific choice of parameter, but want to make predictions of new unseen data:

$$p(t | x) \quad (\text{Predictive distribution})$$

- Including the observations (training set), the *model independent predictive distribution* is given by marginalization over all models:

$$p(t | \mathbf{x}, \mathbf{T}, \mathbf{X}, \alpha, \beta) = \int \overbrace{p(t | \mathbf{x}, \mathbf{w}, \beta)}^{\text{Noise model}} \overbrace{p(\mathbf{w} | \mathbf{T}, \mathbf{X}, \alpha, \beta)}^{\text{Posterior}} d\mathbf{w}$$



Gaussian predictive distribution

- Consider the case of Gaussian noise model, prior and posterior:

$$p(t | \mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t | y(\mathbf{x}, \mathbf{w}), \beta^{-1}) \quad (\text{Noise model})$$

$$p(\mathbf{w} | \mathbf{T}, \mathbf{X}, \alpha, \beta) = \mathcal{N}(\mathbf{w} | \mathbf{m}_N, \mathbf{S}_N) \quad (\text{Posterior})$$

- Predictive distribution:

$$\begin{aligned} p(t | \mathbf{x}, \mathbf{T}, \mathbf{X}, \alpha, \beta) &= \int \underbrace{p(t | \mathbf{x}, \mathbf{w}, \beta)}_{\text{Noise model}} \underbrace{p(\mathbf{w} | \mathbf{T}, \mathbf{X}, \alpha, \beta)}_{\text{Posterior}} d\mathbf{w} \\ &= \int \mathcal{N}(t | y(\mathbf{x}, \mathbf{w}), \beta^{-1}) \mathcal{N}(\mathbf{w} | \mathbf{m}_N, \mathbf{S}_N) d\mathbf{w} \end{aligned}$$

$$= \int \mathcal{N}(t, \mathbf{w} | \mathbf{x}, \beta^{-1}, \mathbf{m}_N, \mathbf{S}_N) d\mathbf{w} = \mathcal{N}(t | y(\mathbf{x}, \mathbf{m}_N), \sigma_N^2(\mathbf{x}))$$

$$\sigma_N^2(\mathbf{x}) = \frac{1}{\beta} + \bar{\boldsymbol{\phi}}^T(\mathbf{x}) \mathbf{S}_N \bar{\boldsymbol{\phi}}(\mathbf{x}) \quad (\text{Predictive variance})$$



Example: Sinusoidal data set

- Synthetic sinusoidal data set: $X = (x_1, \dots, x_N)^T$

$$t(x) = \sin(2\pi x) + \chi, \chi \sim N(\chi | 0, 0.3^2) \quad T = (t_1, \dots, t_N)^T$$

- Linear regression with 9 Gaussian basis functions:

$$\phi_j(x) = \exp\left[-\frac{(x - \mu_j)^2}{2s^2}\right]$$

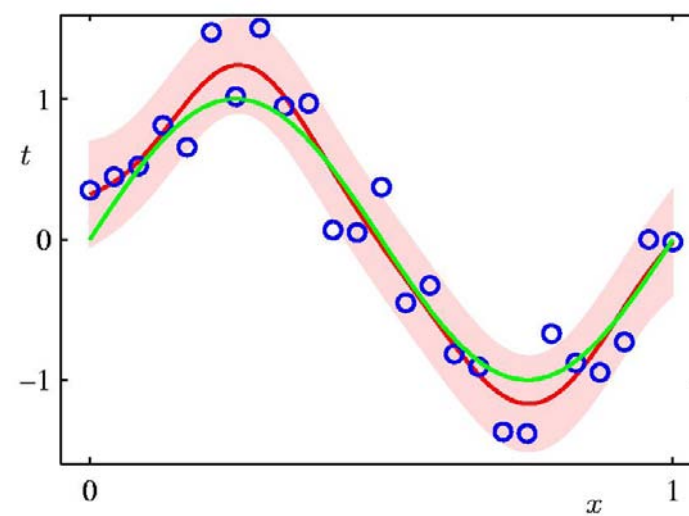
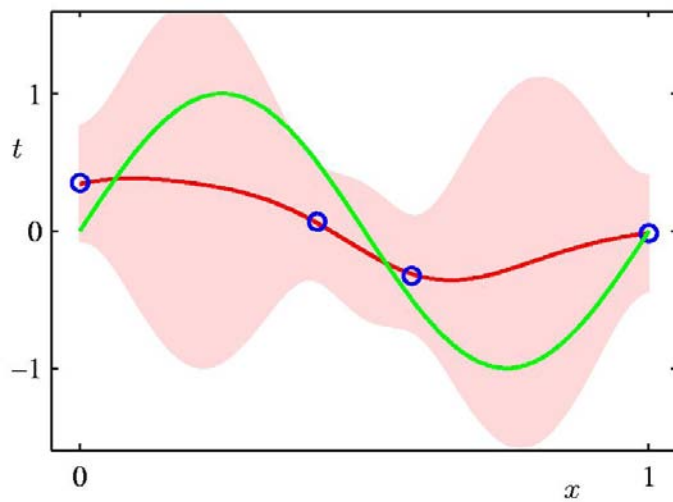
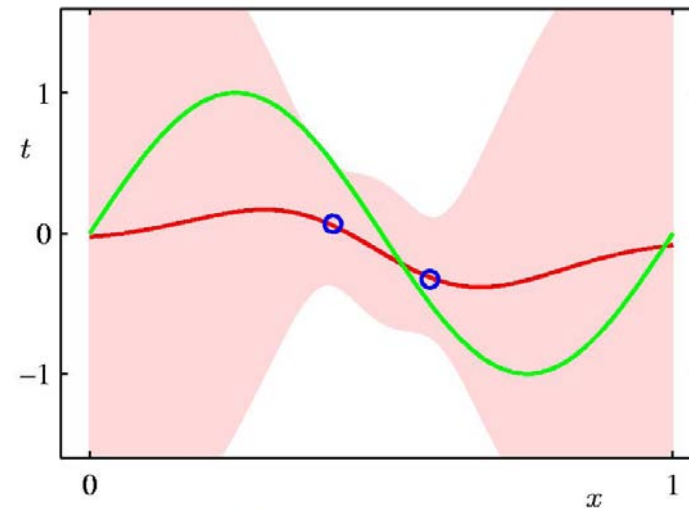
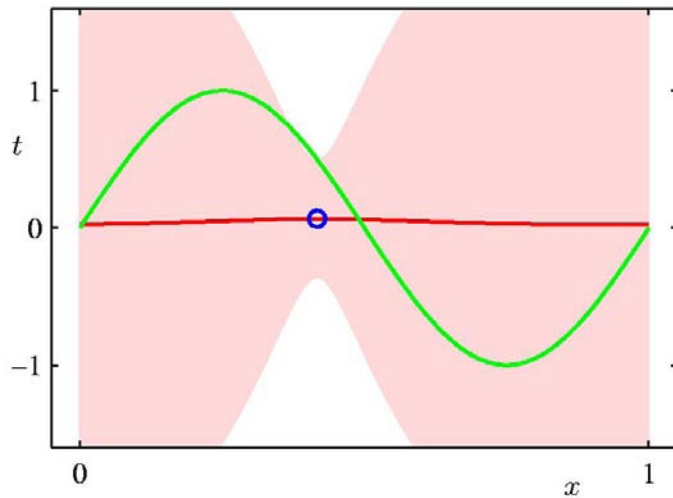
- Lets plot the predictive mean curve:

$$\int t p(t | \mathbf{x}, \mathbf{T}, \mathbf{X}, \alpha, \beta) dt = \int t \mathcal{N}\left(t | y(\mathbf{x}, \mathbf{m}_N), \sigma_N^2(\mathbf{x})\right) dt = y(\mathbf{x}, \mathbf{m}_N)$$

- And the predictive standard deviation curve: $\sigma_N(\mathbf{x})$

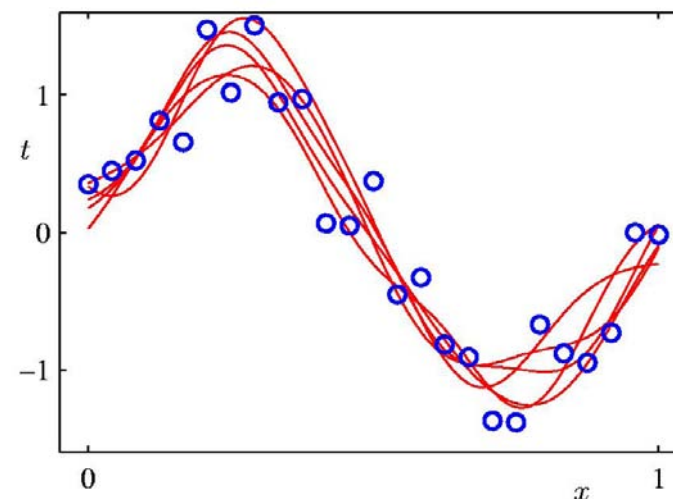
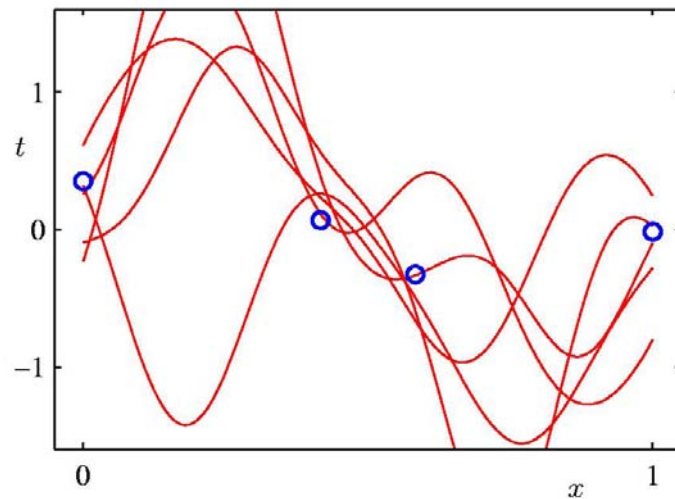
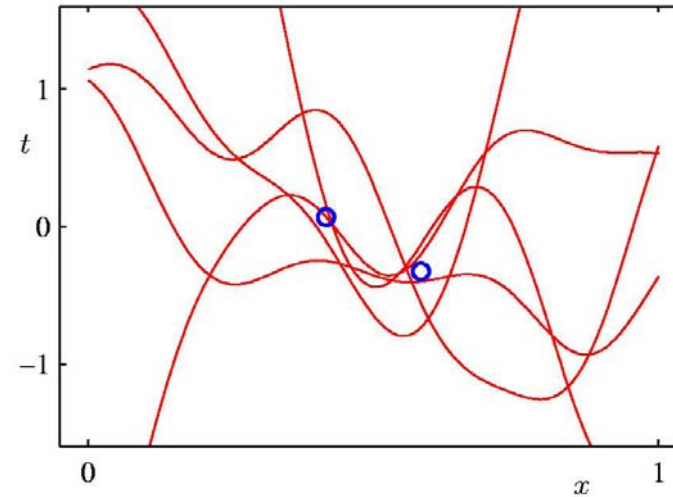
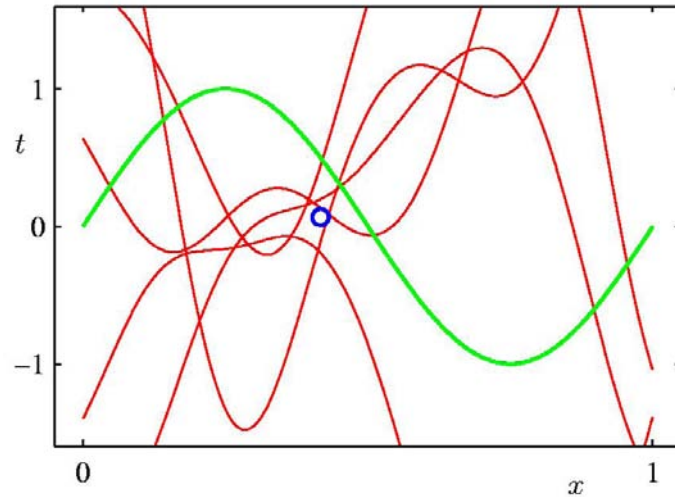


Examples of Predictive Distribution





Samples From Posterior Distribution





Summary

- Linear models for regression
- Bayesian regression for linear models
- Bayesian sequential learning for regression
- Full Bayesian approach – the model independent predictive distribution
- Advanced: It is difficult to simultaneously estimate \mathbf{w} , α , β analytically, however approximations exist (not for this course).

Literature



- Linear models for regression: Sec. 3.1
- Loss function for regression: Sec. 1.5.5
- Bayesian Linear Regression: Sec. 3.3 – 3.3.2
- Limitations on fixed basis functions: Sec. (3.6)