

## Chapter 5

# ESTIMATING THE RATES OF ELECTRON CHARGE MIS-IDENTIFICATION

Many physics analyses involve charged leptons in their final states, where leptons typically refer to electrons or muons. Such an electron or a muon leaves a track in the detector, and its charge, or more specifically the sign of the charge, is determined from the curvature — due to the installed magnetic fields — of the track. Because of some factors which will be discussed below, this determination could occasionally be erroneous, leading to what is called charge misidentification.

Electron charge mis-identification is important for analyses that involve same-sign electrons in the final state. Examples of such analyses include measurements of same-sign WW scattering [47], analyses that involve the production of a Higgs in association with a  $t\bar{t}$  pair ( $t\bar{t}H$ ), and supersymmetry search with two same-sign leptons [49]. In general, charge mis-identification rates occur on the order of  $O(1\%)$ , while Standard Model processes that provide opposite-sign dileptons (dominantly  $Z \rightarrow e^+e^-$  bosons) occur approximately  $10^3$  times more commonly than genuine Standard Model sources of same-sign leptons (dominantly  $WZ$  production). As a result, opposite-sign sources of dileptons suffering from charge mis-identification can constitute a large background in these searches, and so it is crucial to estimate the charge mis-identification background precisely.

This chapter describes a method for estimating the rate of charge mis-identification using a likelihood function. Section 5.1 discusses briefly how electron charge mis-identification might arise at ATLAS. Section 5.2 discusses the likelihood method, including the Poisson likelihood used as well as how it is applied to  $Z \rightarrow e^+e^-$  events to measure the charge mis-identification rates. Finally, Section 5.3 provides some conclusions.

It is to be noted that muon charge mis-identification is known to be negligible, except at very high  $p_T$ . Indeed, the magnetic field in the Muon Spectrometer (see Section 3.3.2.3) allows the measurement of the track curvature over a larger radius, thereby reducing the chance the charge could be mis-identified. Analyses such as the supersymmetry search with two same-sign leptons mentioned above have found that muon charge mis-identification is in indeed negligible (cite).

## 5.1 Electron Charge Mis-identification

At ATLAS, the sign of the charge of an electron is determined from its track in the Inner Detector (see Section 3.3.2.1). Indeed, as the electron passes through the Inner Detector, its track is bent by the installed magnetic fields. The direction of the curvature of the track determines the charge of the electron.

Charge mis-identification, where the charge of the electron is identified incorrectly, occurs mainly because of two reasons:

- As the electron passes through the detector and interacts with the materials in the detector, it may radiate photons. These radiated photons may in turn convert to electron-positron pairs. A charge mis-identification occurs when the electron candidate is matched to the wrong track.
- The reconstructed track of the electron appears rather straight, i.e. the curvature of the track is small, at very high momentum or at large pseudorapidity, the latter because the lever arm of the tracker is limited.

## 5.2 The Likelihood Method

Since it is impossible to know with absolute certainty, after the charge of an electron has been measured, that a charge mis-identification has occurred or not, we seek instead to determine the rates of charge mis-identification for an ensemble of electrons. Essentially, we start from a sample of electrons for which the true charges are known. Charge measurements on the sample will result in another sample that consists of electrons with the original charges as well as electrons whose charges have switched. A key step is to write down the probabilistic distribution of charge assuming a rate of charge mis-identification, and then seek to determine this rate from the actual data. This method is called the likelihood method and will be discussed in section.

This chapter uses  $Z \rightarrow e^+e^-$  events because these provide a good source of clean, high-statistics sample of electrons.

### 5.2.1 The Poisson Likelihood

Consider a pair of electrons  $e^+e^-$  (an opposite-sign pair) at truth level. The charges of the electrons are opposite of each other, but because of charge mis-identification there is a chance of this pair being identified as having the same charge (a same-sign pair). Assuming a probability  $p$  of such a chance, then in considering  $n$  pairs  $e^+e^-$ , the probability of seeing exactly  $n_{ss}$  same-sign pairs is given by the binomial distribution

$$P(n_{ss}) = \binom{n}{n_{ss}} p^{n_{ss}} (1-p)^{n-n_{ss}}.$$

Since it is known that the charge mis-identification probability  $p$  is typically small while the number of pairs considered  $n$  is typically very large, the Poisson

1115 distribution may be used to approximate the binomial distribution. Thus, let

$$m_{ss} = np \quad (5.1)$$

1116 denote the expected number of same-sign pairs, it follows that

$$P(n_{ss}) = \frac{m_{ss}^{n_{ss}} e^{-m_{ss}}}{n_{ss}!} \quad (5.2)$$

1117 is the Poisson probability of seeing  $n_{ss}$  same-sign pairs, given the average number  
1118 of same-sign pairs  $m_{ss}$ . This will also be called a likelihood function.

1119 Consider again an opposite-pair  $e^+e^-$  with the probability  $p$  of being identified as  
1120 a same-sign pair. We may speak directly of a probability  $\epsilon$  of an electron in the pair  
1121 having its charge incorrectly identified. Then a same-sign pair results if only one of  
1122 the electrons in the original opposite-sign pair has its charge identified incorrectly,  
1123 which means we may write

$$p = (1 - \epsilon)\epsilon + \epsilon(1 - \epsilon). \quad (5.3)$$

1124 The Poisson likelihood of Equation 5.2 may now be written to depend explicitly  
1125 on  $\epsilon$ :

$$P(n_{ss}|\epsilon) = \frac{m_{ss}^{n_{ss}} e^{-m_{ss}}}{n_{ss}!}, \quad m_{ss} = np = n(1 - \epsilon)\epsilon + \epsilon(1 - \epsilon). \quad (5.4)$$

1126 In an actual measurement of the rates of charge mis-identification, the individual  
1127 rates  $\epsilon$ 's are in general different for the electrons in the pair, if for example the  
1128 dependence of  $\epsilon$  on the transverse momenta  $p_T$  is taken into account. We may  
1129 introduce then a number of bins in  $p_T$  and may speak of a charge mis-identification  
1130 probability associated with a bin  $i$ . Consequently an electron pair is associated with  
1131 a pair of bins  $(i, j)$ , and we have the following quantities:

1132 ○ The probability

$$p_{ij} = (1 - \epsilon_i)\epsilon_j + \epsilon_i(1 - \epsilon_j) \quad (5.5)$$

1133 in place of the probability  $p$  in Equation 5.3. This is the probability an opposite-  
1134 sign pair may be seen as a same-sign pair in the bin pair  $(i, j)$

1135 ○ The number of electron pairs considered,  $n_{ij}$ , in the bin pair  $(i, j)$

1136 ○ The expected number of same-sign pairs

$$m_{ss,ij} = n_{ij}p_{ij} \quad (5.6)$$

1137 in place of the expected number of same-sign pairs in Equation 5.1

1138 ○ The Poisson likelihood

$$P(n_{ss,ij}|\epsilon_i, \epsilon_j) = \frac{m_{ss,ij}^{n_{ss,ij}} e^{-m_{ss,ij}}}{n_{ss,ij}!} \quad (5.7)$$

in place of the Poisson likelihood in Equation 5.4. This will also be denoted simply as  $L_{ij}$

These equations remain the same if instead of one-dimensional bins, two dimensional bins are used. Indeed, suppose the dependency of the charge mis-identification rates on, say  $p_T$  and  $\eta$ , needs to be taken into account. If there are  $a$  bins in  $p_T$  and  $b$  bins in  $\eta$ , the total number of bins  $ab$  could be labelled  $1, 2, \dots, ab$  and treated as an ordered sequence of one-dimensional bins. The actual  $p_T$  and  $\eta$  binning together with the estimation of the charge mis-identification rates will be discussed in the following section.

All the possible bin pairs  $(i, j)$  need to be used and therefore, assuming statistically-independent rates, we will maximize the likelihood function

$$L = \prod_{i,j} L_{ij}$$

to find the rates  $\epsilon_i$ , the data being  $n_{ij}$ , the numbers of electrons observed in the bin pair  $(i, j)$ , and  $n_{ss,ij}$ , the number of same-sign electron pairs observed in the bin pair  $(i, j)$ .

### 5.2.2 Estimation of the Rates on $Z \rightarrow e^+e^-$ sample

The Poisson likelihood is applied on a  $Z \rightarrow e^+e^-$  data sample to determine the charge mis-identification rates as follows. First, several selections are applied, including

- Logical OR between two single-electron triggers, one with  $E_T > 24$  GeV plus Medium identification, one with  $E_T > 60$  GeV plus Loose identification
- At least two electron candidates with  $|\eta| < 2.47$
- One electron is required to pass the Tight identification requirement, and to have  $E_T > 25$  GeV. The other electron must have  $E_T > 10$  GeV and must satisfy the track quality criteria (the tracks associated with the electron must have at least one hit in the pixel detector and at least seven hits in the pixel and SCT detectors)
- The invariant mass is within  $\pm 15$  GeV of the  $Z$  mass

The invariant mass of the pair of electrons will play an important role in the following discussion. Figure 5.1 [37] shows the invariant mass distribution  $m_{ee}$  for two different  $\eta$  range, each electron having  $0.0 < \eta < 0.8$  and  $2.0 < \eta < 2.47$ ; in both figures the each electron in the pairs is selected to have  $E$  between 25 GeV and 50 GeV. Due to charge mis-identification same-sign electron pairs exist in addition to opposite-sign pairs, and they are plotted alongside opposite-sign pairs. It is seen that same-sign pairs have a broader peak which is also slightly shifted to lower values, consistent with the fact that radiation which causes charge-misidentification also causes energy loss. It is also seen that charge-misidentification is higher at higher  $\eta$ , as have been commented previously.

To continue, an invariant mass interval  $(m_l, m_h)$  is selected, where  $m_l = 15$  GeV is the low mass point and  $m_h = 15$  GeV the high mass point around the  $Z$  mass peak. Then electrons in the events are binned according to their  $\eta$  and  $p_T$  values. Each electron pair then is associated to a pair of bin  $(i, j)$ , and all such bin pairs need to be taken into account. The quantities needed are (see Section 5.2.1):

○  $n_{ij}$ , the number of electrons counted in the bin pair  $(i, j)$

○  $n_{ss,ij}$ , the number of same-sign electron pairs counted in the bin pair  $(i, j)$

There are non-signal events contamination in these quantities and to deal with them we assume that the two sides of the interval  $(m_l, m_h)$  are dominated by background contributions, and adopt following method. To begin, in addition to the original invariant mass interval  $(m_l, m_h)$ , we consider the interval  $(m_l - w_l, m_h + w_h)$  where  $w_l = 15$  GeV and  $w_h = 15$  GeV are some widths. The latter interval is made up of three intervals:

○  $(m_l, m_h)$ . This is the original invariant mass interval.

○  $(m_l - w_l, m_l)$ . This is the interval that lies to the left of the original interval

○  $(m_h, m_h + w_h)$ . This is the interval that lies to the right of the original interval

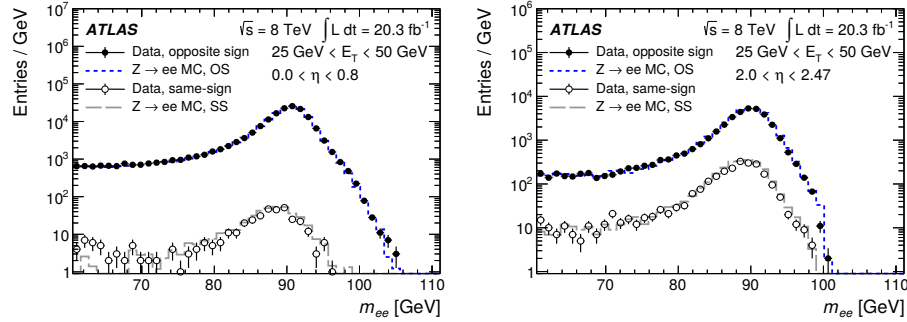


Figure 5.1: Distribution of the invariant mass  $m_{ee}$  for  $E_T$  between 25 and 50 GeV and  $|\eta|$  between 0.0 and 0.8 [37]. Due to charge mis-identification same-sign pairs as well as opposite-sign pairs are seen.

Then, in addition to the quantities  $n_{ij}$  and  $n_{ss,ij}$  in the original central interval  $(m_l, m_h)$ , we will consider the corresponding quantities in the two new intervals, to be denoted  $n_{ij}^l$  and  $n_{ss,ij}^l$  in the left interval and  $n_{ij}^h$  and  $n_{ss,ij}^h$  in the right interval. We assume the left and right intervals are background intervals and subsequently compute the weighted quantities  $b(n_{ij})$ , to mean the background contamination in  $n_{ij}$ , and  $b(n_{ss,ij})$ , to mean background contamination in  $n_{ss,ij}$ :

$$b(n_{ij}) = \frac{w_l \times n_{ij}^l + w_h \times n_{ij}^h}{w_l + w_h}, \quad b(n_{ss,ij}) = \frac{w_l \times n_{ss,ij}^l + w_h \times n_{ss,ij}^h}{w_l + w_h}$$

which will be taken as the backgrounds in  $n_{ij}$  and  $n_{ss,ij}$  in the central interval respectively.

1199 The terms  $n_{ij}$  and  $n_{ss,ij}$  and the background terms  $b(n_{ij})$  and  $b(n_{ss,ij})$  are to be  
 1200 used as follows. According to Equation 5.7 the Poisson likelihood to be fitted is

$$P(n_{ss,ij}|\epsilon_i, \epsilon_j) = \frac{m_{ss,ij}^{n_{ss,ij}} e^{-m_{ss,ij}}}{n_{ss,ij}!}$$

1201 The background terms make a contribution to the expected number of same-sign  
 1202  $m_{ss,ij}$  in the likelihood, modifying it from  $m_{ss,ij} = n_{ij}p_{ij}$  (see Equation 5.6) to

$$m_{ss,ij} = (n_{ij} - b(n_{ij})) \times p_{ij} + b(n_{ss,ij})$$

1203 The first quantity on the right in the equation above is the same-sign contribu-  
 1204 tion from signal events where the background has to be subtracted, and the second  
 1205 quantity is the contribution from background events.

### 1206 5.2.3 Charge Mis-identification Rates and Uncertainties

1207 The rates are obtained upon the maximization of the likelihood function discussed  
 1208 in the previous section. The statistical uncertainties associated with the estimated  
 1209 rates depend on the statistics of the data, and are given by the statistical tool that  
 1210 maximizes the Poisson likelihood.

1211 The following sources of systematic uncertainties are evaluated:

- 1212 ☐ Systematic uncertainty that comes from background subtraction, which is eval-  
 1213 uated by determining the rates with and without background subtraction. The  
 1214 inclusion of this uncertainty ensures a conservative figure of systematic uncer-  
 1215 tainty in the charge mis-identification rates; it has a small impact because the  
 1216 background is small.
- 1217 ☐ The invariant mass interval  $(m_l, m_h)$  may be varied, from 15 GeV around the  
 1218  $Z$  mass to 10 and 20 GeV additionally. In this way an idea of how the selection  
 1219 of an interval may affect the rates may be obtained.
- 1220 ☐ The invariant mass widths  $w_l$  and  $w_h$  may be varied, taking values 20, 25, or  
 1221 30 GeV. This takes into account the uncertainty on the rates due to the choice  
 1222 of a mass width.

1223 The actual rates are estimated for the following three sets of requirements:

- 1224 ☐ Medium: Medium identification requirements
- 1225 ☐ Tight + isolation: Tight identification requirements plus track isolation cut  
 1226  $p_T^{\text{cone } 0.2}/E_T < 0.14$ .
- 1227 ☐ Tight + isolation + impact parameter: Tight identification plus  $E_T^{\text{cone } 0.3}/E_T <$   
 1228  $0.14$  and  $p_T^{\text{cone } 0.2}/E_T < 0.07$  and additionally  $|z_0| \times \sin \theta < 0.5$  mm and  
 1229  $|d_0|/\sigma_{d_0} < 5.0$

Figure 5.2 [37] show the estimated rates in data and simulation. The dashed lines indicate the bins in which the rates are calculated. Total uncertainty, which is computed as the sum in quadrature of statistical and systematic uncertainties, is also showed. In most bins, simulation over-estimates the rates as compared to the data by 5-20% depending on  $\eta$  and electron requirements.

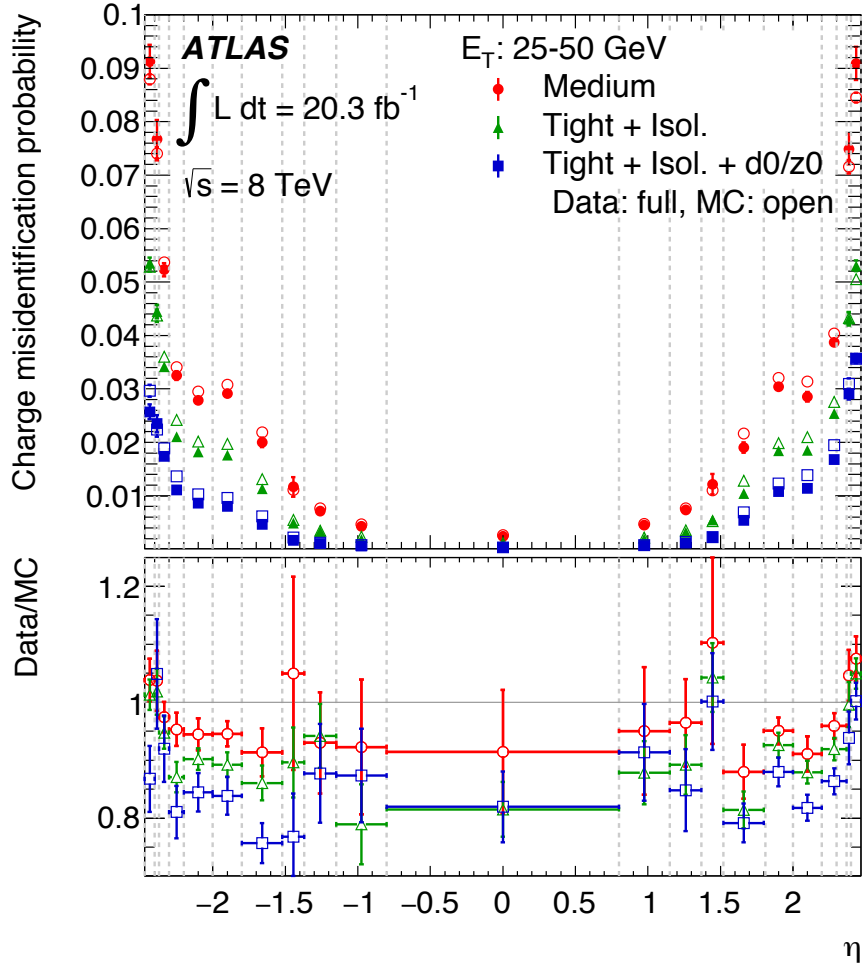


Figure 5.2: Charge mis-identification probabilities in  $\eta$  bins,  $E_T$  between 25 GeV and 50 GeV [37]. Three different sets of selection requirements (Medium, Tight + Isolation, and Tight + Isolation + impact parameter) are shown, along with simulation expectations. Displayed in the lower panel is the data-to-simulation ratios. The uncertainties are the total uncertainties from the sum in quadrature of statistical and systematic uncertainties. The dashed lines indicate the bins in which the rates are calculated.

#### 5.2.4 Estimating Charge Mis-identification Background from the Charge Mis-identification Rates

In this section we give an example of how the charge mis-identification rates may be used to estimate the charge mis-identification background in analysis with a same-sign lepton pair signature. Suppose a sample of same-sign electron pairs has been

selected in the bin pairs  $(i, j)$  (see Section 5.2). Let there be  $n_{ss,ij}$  of such pairs, and we wish to determine the charge misidentification contribution to this number.

To begin, we have to distinguish between the number of same-sign electron pairs  $n_{ss,ij}$  that has been selected and the number of genuine same-sign electron pairs. The latter is what would be counted if there were no charge misidentification. Denote it by  $\bar{n}_{ss,ij}$ .

A charge misidentification contribution occurs whenever there is an opposite-sign pair of electrons in which one of the electron has its charge mis-identified. The probability for this to happen is, according to Equation 5.5,

$$p_{ij} = (1 - \epsilon_i)\epsilon_j + \epsilon_i(1 - \epsilon_j),$$

where  $\epsilon_i$  and  $\epsilon_j$  are the charge mis-identification rates in the bins. This probability has to be multiplied by the real number of opposite-sign pairs, and not the number of opposite-sign pairs counted in the bin pair, because the latter involves contribution from same-sign pairs  $\bar{n}_{ss,ij}$  as well.

Denote the number of opposite-sign pairs counted in the bin pair  $(i, j)$  by  $n_{os,ij}$ , and the corresponding real quantity by  $\bar{n}_{os,ij}$ . The quantities available are  $n_{ss,ij}$ ,  $n_{os,ij}$ , and the mis-identification rates  $\epsilon_i$  and  $\epsilon_j$ . The unknown are  $\bar{n}_{ss,ij}$  and  $\bar{n}_{os,ij}$ , but they are needed to determine charge mis-identification contribution. Regarding this, the following relation holds

$$n_{os,ij} = \bar{n}_{os,ij} - \bar{n}_{os,ij} \times p_{ij} + \bar{n}_{ss,ij} \times p_{ij},$$

which says that the number of opposite-sign lepton pairs counted in the bin pair  $(i, j)$  is the corresponding real number minus the portion that is identified as same-sign plus the contribution from real same-sign pairs. This may be re-written as

$$n_{os,ij} = \bar{n}_{os,ij} \times (1 - p_{ij}) + \bar{n}_{ss,ij} \times p_{ij}.$$

Similarly we have the following relation

$$n_{ss,ij} = \bar{n}_{ss,ij} \times (1 - p_{ij}) + \bar{n}_{os,ij} \times p_{ij}$$

These two relations form a system of equations from which the unknown  $\bar{n}_{os,ij}$  and  $\bar{n}_{ss,ij}$  may be solved. Then the charge mis-identification contribution to  $n_{ss,ij}$  is simply  $\bar{n}_{os,ij} \times p_{ij}$ .

The method just discussed is to be contrasted with the scale factor method, where scale factors that adjust the charge mis-identification rates in simulations to match the data are provided to different analyses. The former method excludes the need for the use of all systematic uncertainties that are associated with the use of simulation samples.

## 5.3 Conclusions

This chapter described the electron charge mis-identification problem at ATLAS and how the charge mis-identification rates are measured by fitting a Poisson likelihood



1274 function using the  $Z \rightarrow e^+e^-$  data sample ([data set](#)). Three sets of charge mis-  
1275 identification rates are measured and provided to ATLAS analyses, corresponding to  
1276 three different sets of selection requirements (Medium, Tight + Isolation, and Tight  
1277 + Isolation + impact parameter) ([the range of flip rates observed](#)). In general,  
1278 simulation underestimates the charge mis-identification rates as compared to those in  
1279 the data.

1280 It is to be noted that in addition to measuring the charge mis-identification rates,  
1281 a separate effort was started by the physics team at Université de Montréal aiming at  
1282 reducing charge mis-identification. The technique relies on the output of a boosted  
1283 decision tree (BTD) using a simulated sample of single electrons (see Reference [\[37\]](#))  
1284 ([fix reference](#)).